

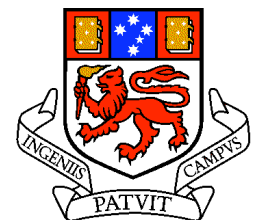
Statistical analysis of genome-wide association (GWAS) data

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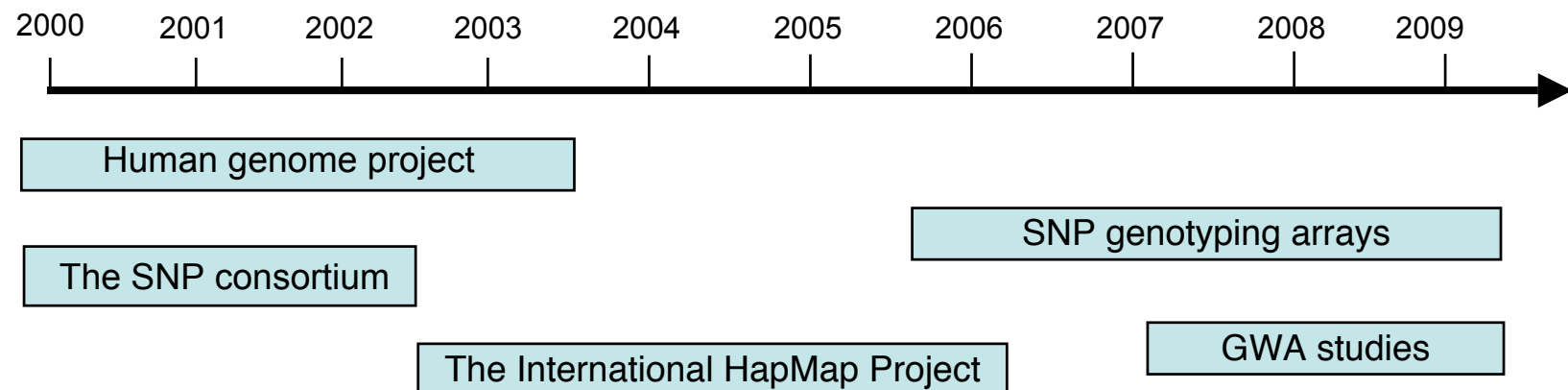


Outline

- Introduction
- Confounding variables and linkage disequilibrium
- Statistical methods to test for association in case-control GWA studies
 - Allele counting chi-square test
 - Logistic regression
- Multiple testing and power
- Example: GWAS for multiple sclerosis (MS)
 - Data cleaning / quality control
 - Results

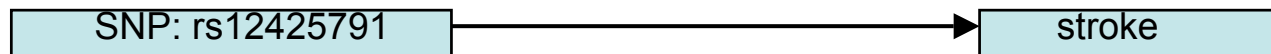
GWA studies have been very successful since 2007

- Prior to the advent of GWA studies, there was **very** little success in identifying genetic risk factors for complex multifactorial diseases
- GWA studies have identified over 200 separate associations with various complex diseases in the past two years
- “Human Genetic Variation” hailed as “Breakthrough of the Year” by Science magazine in 2007



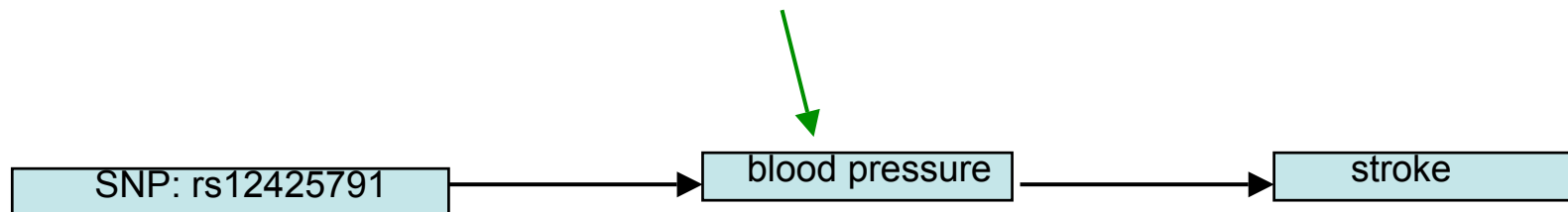
This talk: case-control GWA studies

- Obtain DNA from people with disease of interest (cases) and unaffected controls
- Run each DNA sample on a SNP chip to measure genotypes at 300,000-1,000,000 SNPs in cases and controls
- Identify SNPs where one allele is significantly more common in cases than controls
 - The SNP is *associated* with disease



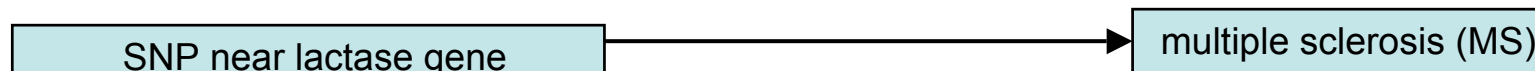
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 - The SNP is *associated* with disease
- Alternative strategy (Peter Visscher's talk): test for association between SNPs and a quantitative trait that underlies the disease (*endophenotype*)



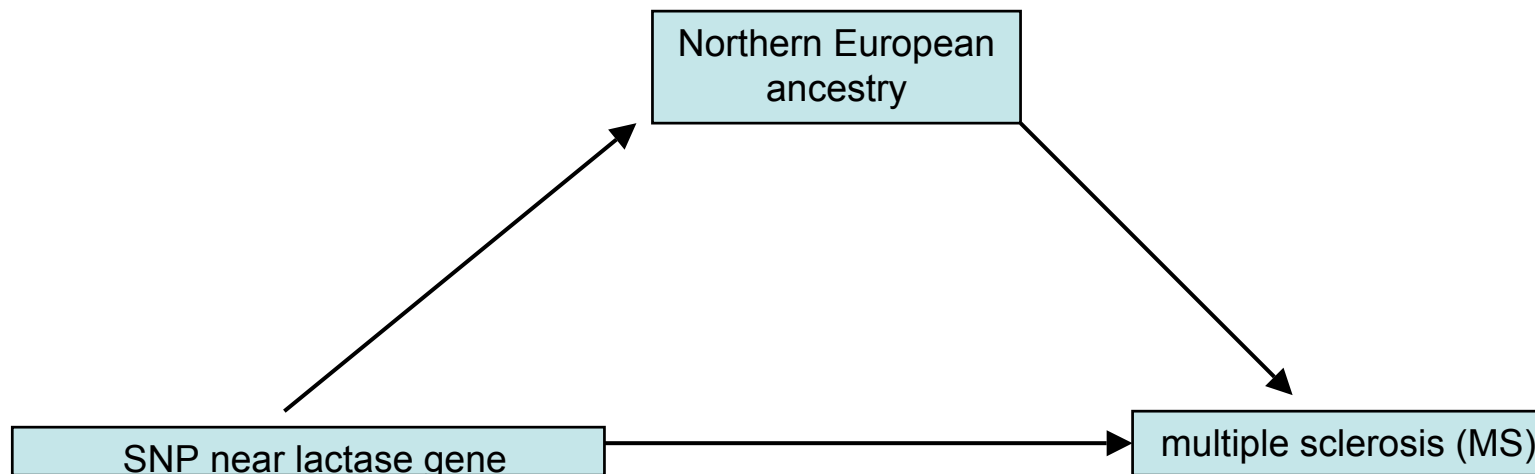
Association does not imply causation

- Suppose that genotypes at a particular SNP are significantly associated with disease
- This may be because the SNP is associated with some other factor (a *confounder*), which is associated with disease but is not in the same causal pathway



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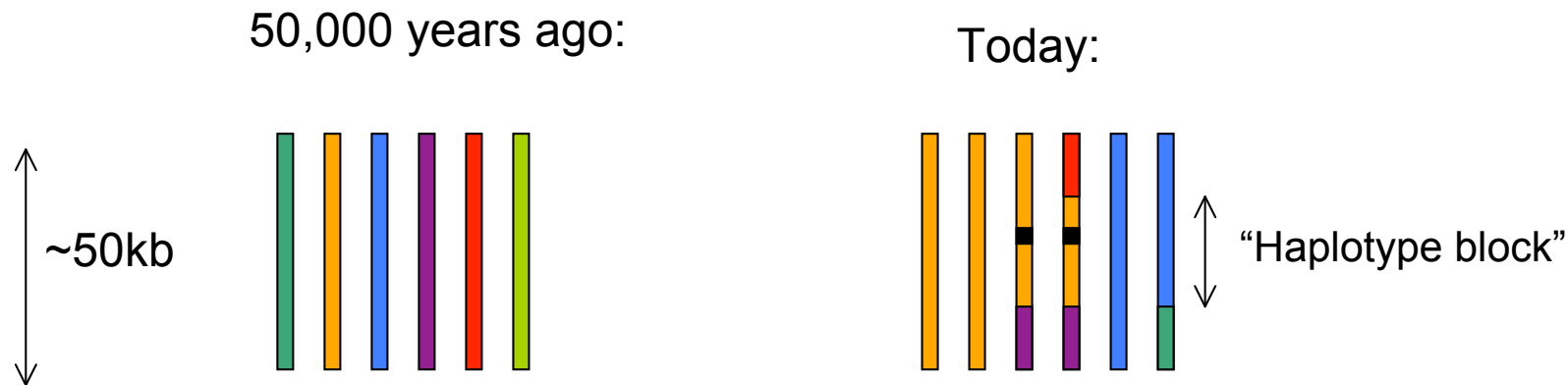


Association does not imply causation

- Suppose that genotypes at a particular SNP are significantly associated with disease
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- Possible confounders of genetic associations:
 - Ethnic ancestry
 - Genotyping batch, genotyping centre
 - DNA quality
- Environmental exposures in the same causal pathway
 - Nicotine receptors --> smoking --> lung cancer
Hung et al, Nature 452: 633 (2008) + other articles in same issue
 - Alcohol dehydrogenase genes --> alcohol consumption --> throat cancer
Hashibe et al, Nature Genetics 40: 707 (2008)

Helpful confounding: linkage disequilibrium

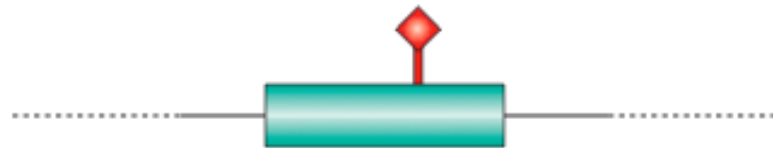
Linkage disequilibrium (LD) is the non-independence of alleles at nearby markers in a population because of a lack of recombinations between the markers



Direct and indirect association testing

Hirschhorn and Daly: Nature Reviews Genetics 6: 95 (2005)

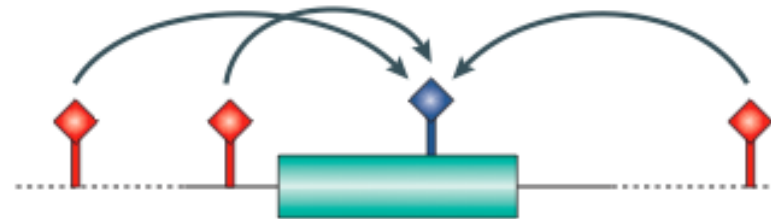
a



Direct association

Functional SNP is genotyped
and an association is found

b

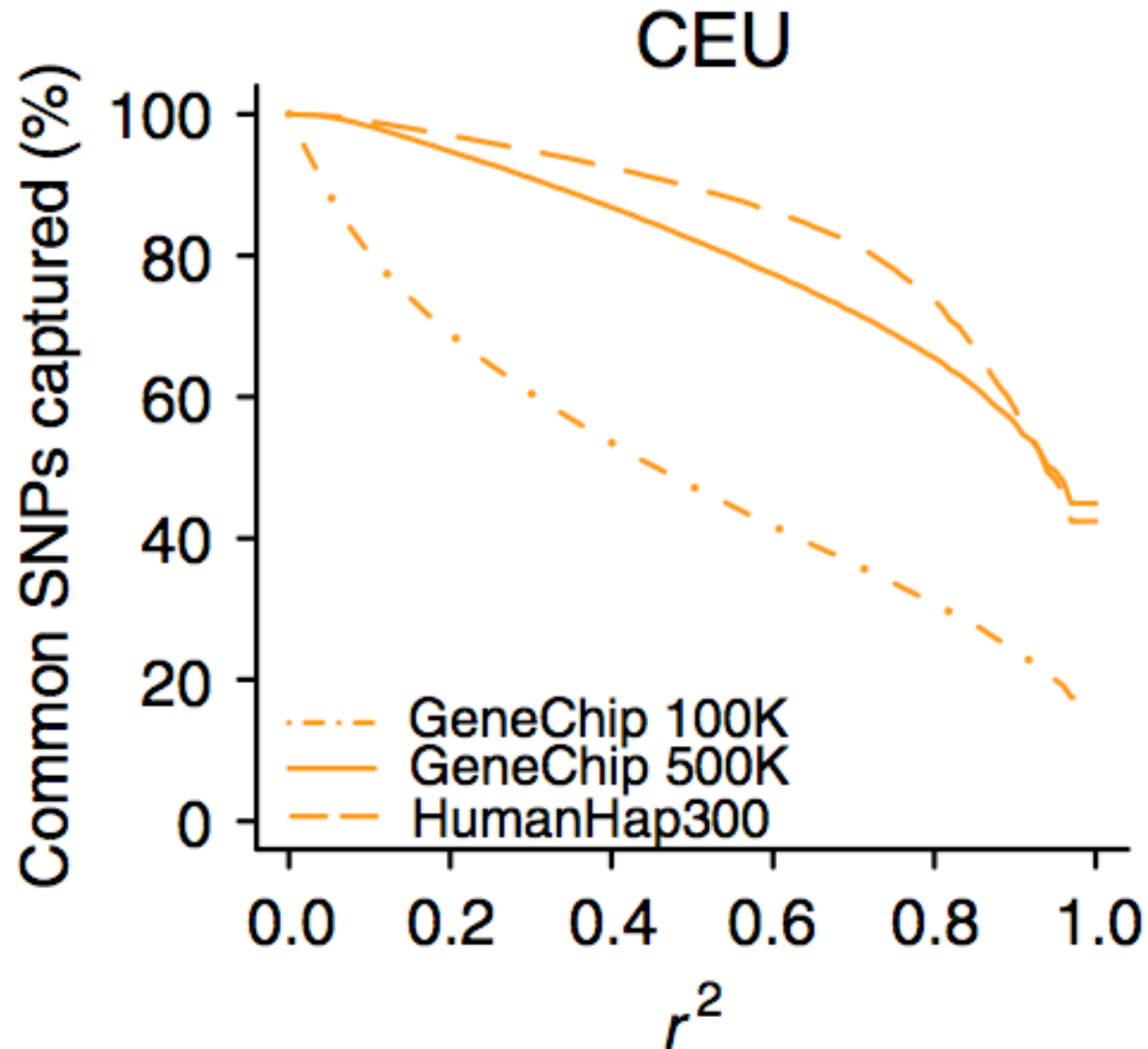


Indirect association

Functional SNP (**blue**) is not
genotyped, but a number of
other SNPs (**red**), in LD with
the functional SNP, are
genotyped, and an
association is found for these
SNPs

LD is helpful, because not all SNPs have to be genotyped

Pe'er et al: Nature Genetics 38: 663 (2006)



Allele counting to test for association between SNP genotype and case / control status

	GG	GT	TT	Total
Cases	r_0	r_1	r_2	R
Controls	s_0	s_1	s_2	S
Total	n_0	n_1	n_2	N

Observed allele counts

	G	T	Total
Cases	$2r_0+r_1$	r_1+2r_2	$2R$
Controls	$2s_0+s_1$	s_1+2s_2	$2S$
Total	$2n_0+n_1$	n_1+2n_2	$2N$

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Total	$2n_0 + n_1$	$n_1 + 2n_2$	$2N$

Expected allele counts

G	T
$2R(2n_0 + n_1)/(2N)$	$2R(n_1 + 2n_2)/(2N)$
$2S(2n_0 + n_1)/(2N)$	$2S(n_1 + 2n_2)/(2N)$

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Cases	r_0	r_1	r_2	R
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Total	n_0	n_1	n_2	N

Observed allele counts			Expected allele counts	
	G	T	G	T
Cases	$2r_0+r_1$	r_1+2r_2	$2R(2n_0+n_1)/(2N)$	$2R(n_1+2n_2)/(2N)$
Controls	$2s_0+s_1$	s_1+2s_2	$2S(2n_0+n_1)/(2N)$	$2S(n_1+2n_2)/(2N)$
Total	$2n_0+n_1$	n_1+2n_2		

Chi-square test for independence of rows and columns (null hypothesis):

$$\sum \frac{(\text{Obs} - \text{Exp})^2}{\text{Exp}} \sim \chi^2 \text{ with 1 df}$$

PLINK **--assoc** option Other options (e.g. dominant/recessive models)
--model

The odds ratio: a measure of effect size

Odds of an event occurring = $\text{Pr}(\text{event occurs}) / \text{Pr}(\text{event doesn't occur})$
= $\text{Pr}(\text{event occurs}) / [1 - \text{Pr}(\text{event occurs})]$

	Allele counts	
	G	T
Cases	<i>a</i>	<i>b</i>
Controls	<i>c</i>	<i>d</i>

Consider all the G alleles in the sample, and pick one at random.
The odds that the G allele occurs in a case: a/c

Consider all the T alleles in the sample, and pick one at random.
The odds that a T allele occurs in a case: b/d

odds ratio = $\frac{\text{odds that G allele occurs in a case}}{\text{odds that T allele occurs in a case}} = \frac{a/c}{b/d} = \frac{a d}{b c}$

Interpretation of the odds ratio

	G	T
Cases	<i>a</i>	<i>b</i>
Controls	<i>c</i>	<i>d</i>

$$\text{odds ratio (OR)} = \frac{\text{odds that G allele occurs in a case}}{\text{odds that T allele occurs in a case}} = \frac{a d}{b c}$$

OR = increase in odds of being a case for each additional G allele

OR = 1: no association between genotype and disease

OR > 1: G allele increases risk of disease

OR < 1: T allele increases risk of disease

If the disease is rare (e.g. ~0.1% for MS), the odds ratio is roughly equal to the *genotype relative risk (GRR)*:

the increase in risk of disease conferred by each additional G allele

e.g. if OR = 1.2,

$$\text{Pr(MS | TT)} = 0.1\%$$

$$\text{Pr(MS | GT)} = 0.12\%$$

$$\text{Pr(MS | GG)} = 0.144\%$$

Logistic regression: more flexible analysis for GWA studies

- Similar to linear regression, used for binary outcomes instead of continuous outcomes
- Let Y_i be the phenotype for individual i
 - $Y_i = 0$ for controls
 - $Y_i = 1$ for cases
- Let X_i be the genotype of individual i at a particular SNP
 - TT $X_i = 0$
 - GT $X_i = 1$
 - GG $X_i = 2$

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- Basic logistic regression model
Let $p_i = E(Y_i | X_i)$, expected value of pheno given geno
Define $\text{logit}(p_i) = \log_e[p_i / (1 - p_i)]$

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$$\text{logit}(p_i) \sim \beta_0 + \beta_1 X_i$$

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Test whether β_1 differs significantly from zero:
roughly equivalent to allele counting chi-square test

Estimate of odds ratio: $\exp(\beta_1)$

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- Let X_i be the genotype of individual i at a particular SNP
TT $X_i = 0$
GT $X_i = 1$
GG $X_i = 2$
- Add extra terms to adjust for potential confounders: e.g. ethnicity, genotyping batch, genotypes at other SNPs
Let $p_i = E(Y_i | X_i, C_i, D_i, \dots)$

$$\text{logit}(p_i) \sim \beta_0 + \beta_1 X_i + \beta_2 C_i + \beta_3 D_i + \dots$$

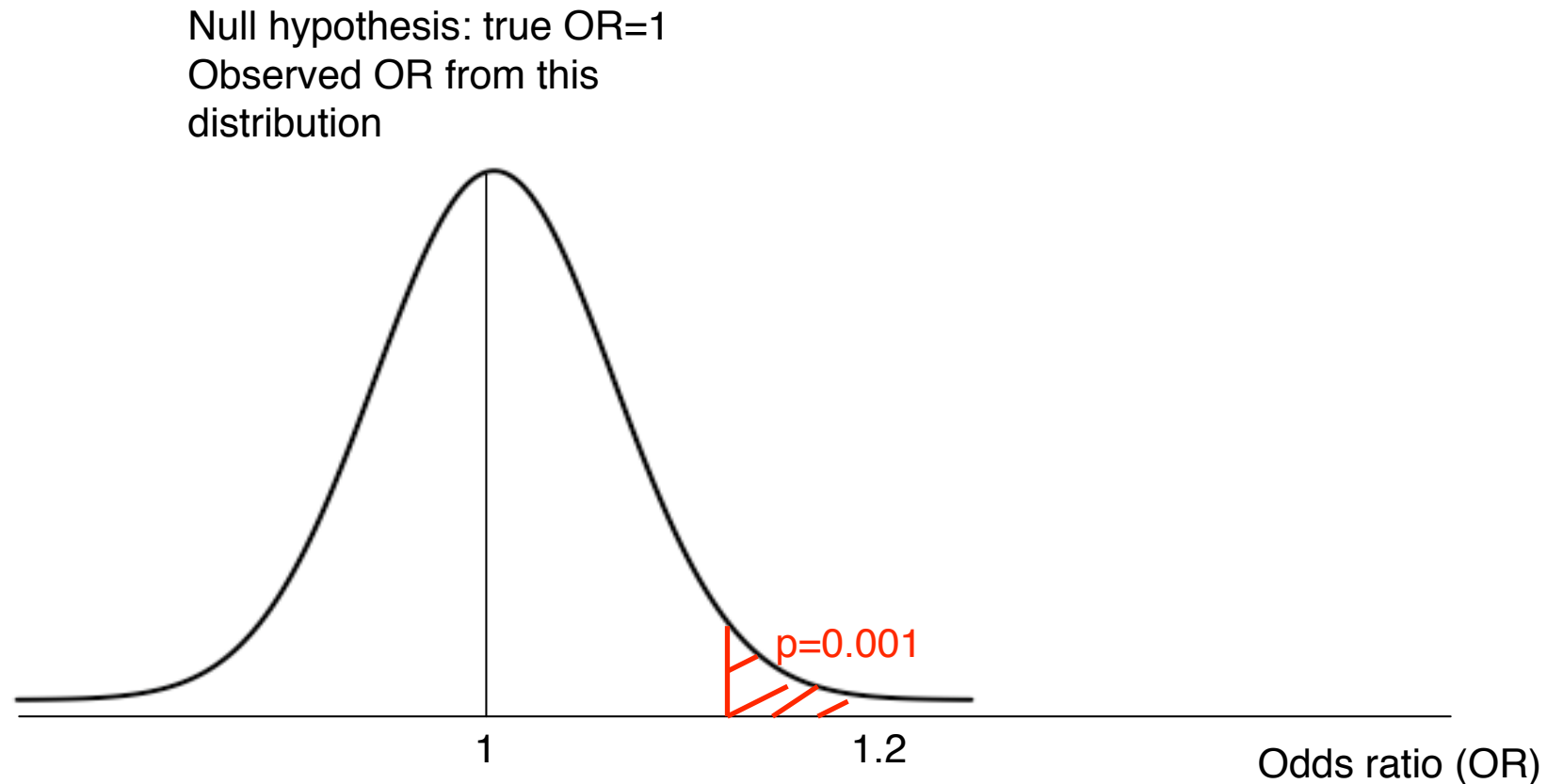
PLINK `--logistic`

Multiple testing

- Suppose you test 500,000 SNPs for association with disease
- Expect around $500,000 \times 0.05 = 25,000$ to have p-value less than 0.05
- More appropriate significance threshold
$$p = 0.05 / 500,000 = 10^{-7}$$
genome-wide significance
- In our MS GWAS we considered SNPs for follow-up if they had p-values less than 0.001
- To detect a smaller p-value need a larger study

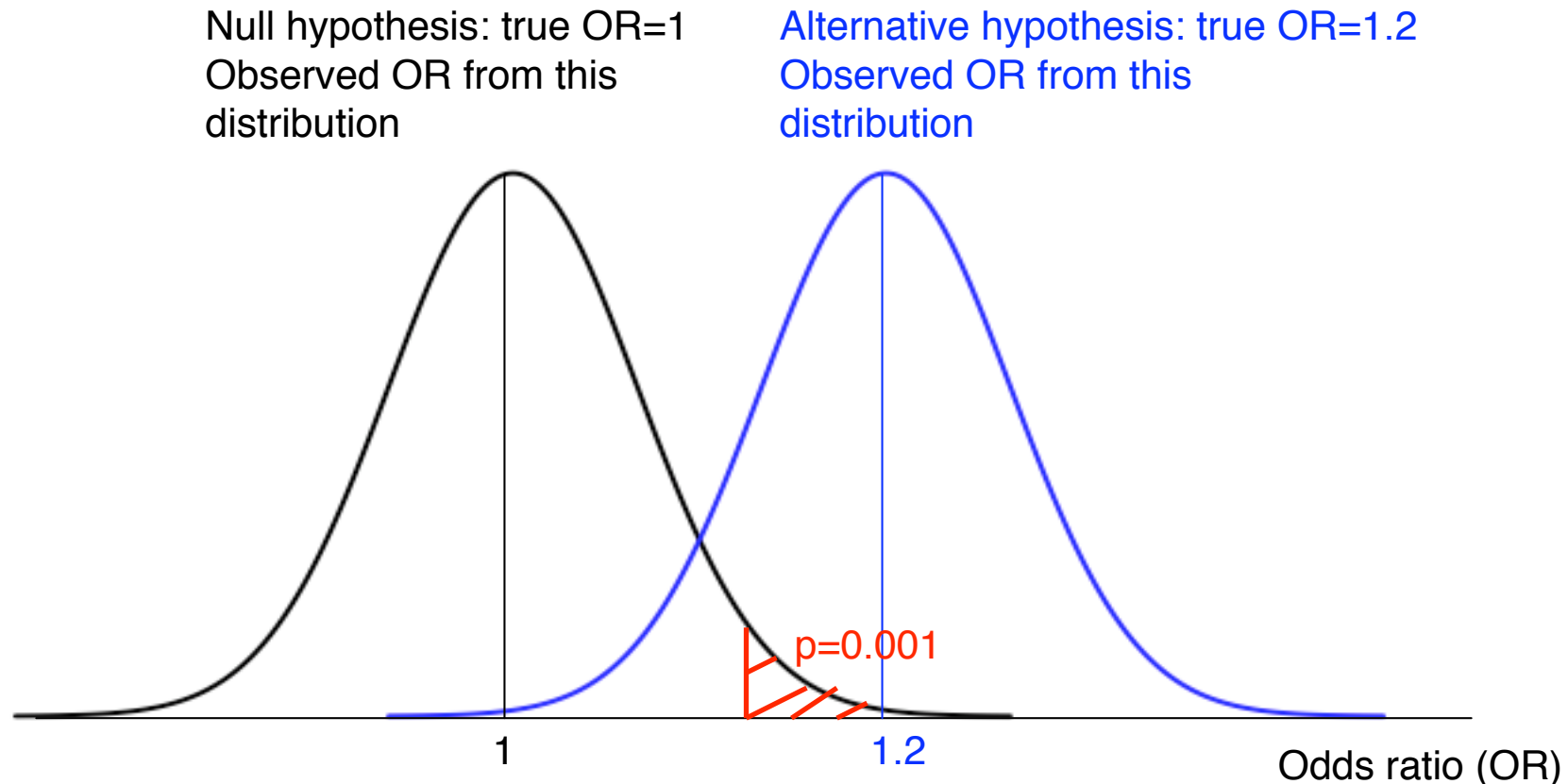
The power to detect an association

- Suppose the G allele of a SNP has frequency 0.2. If each additional G allele increases odds of disease by 1.2, and 1618 cases and 3413 controls are genotyped, what is the *power* (chance) of detecting an association with significance $p < 0.001$?



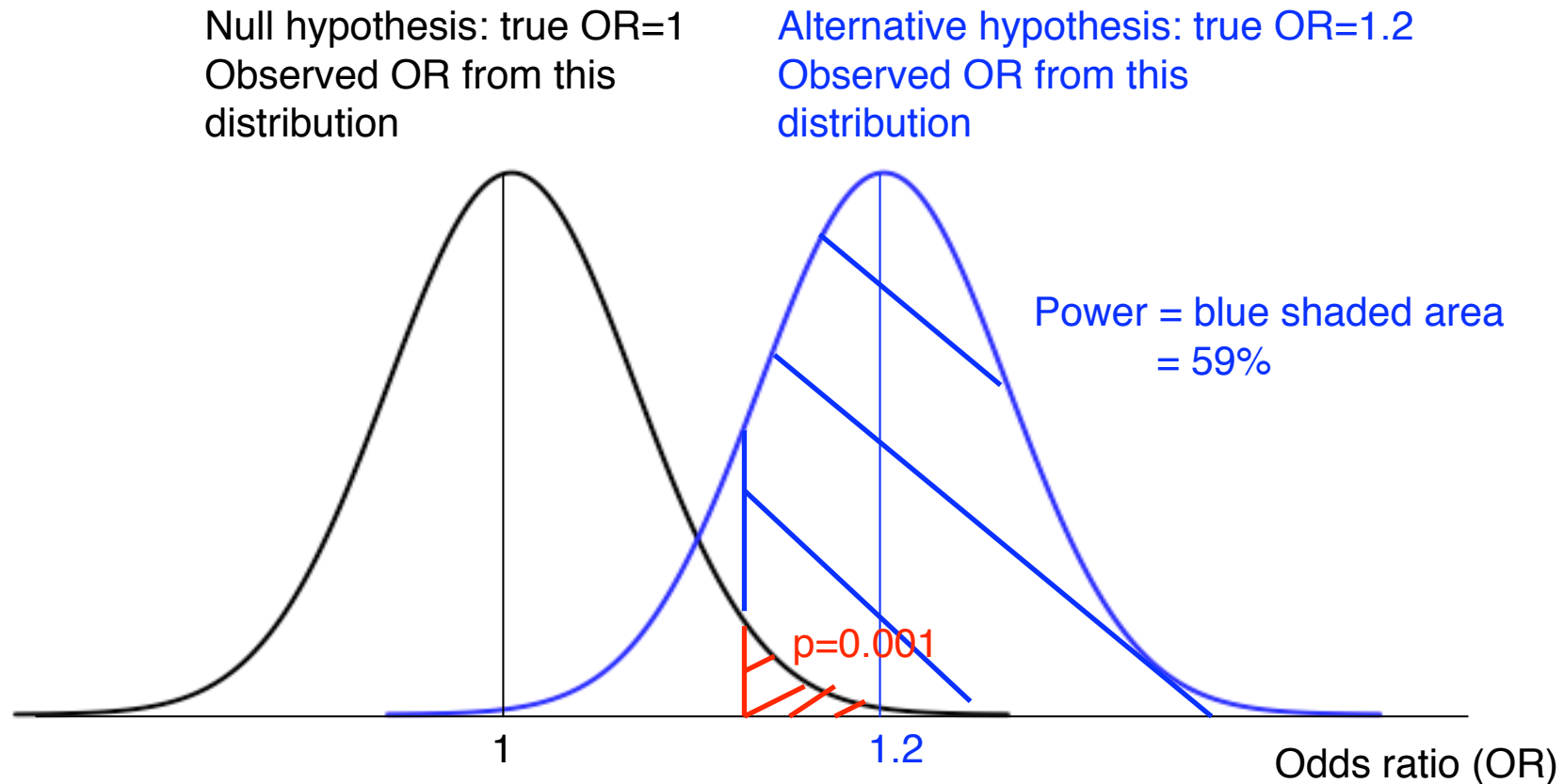
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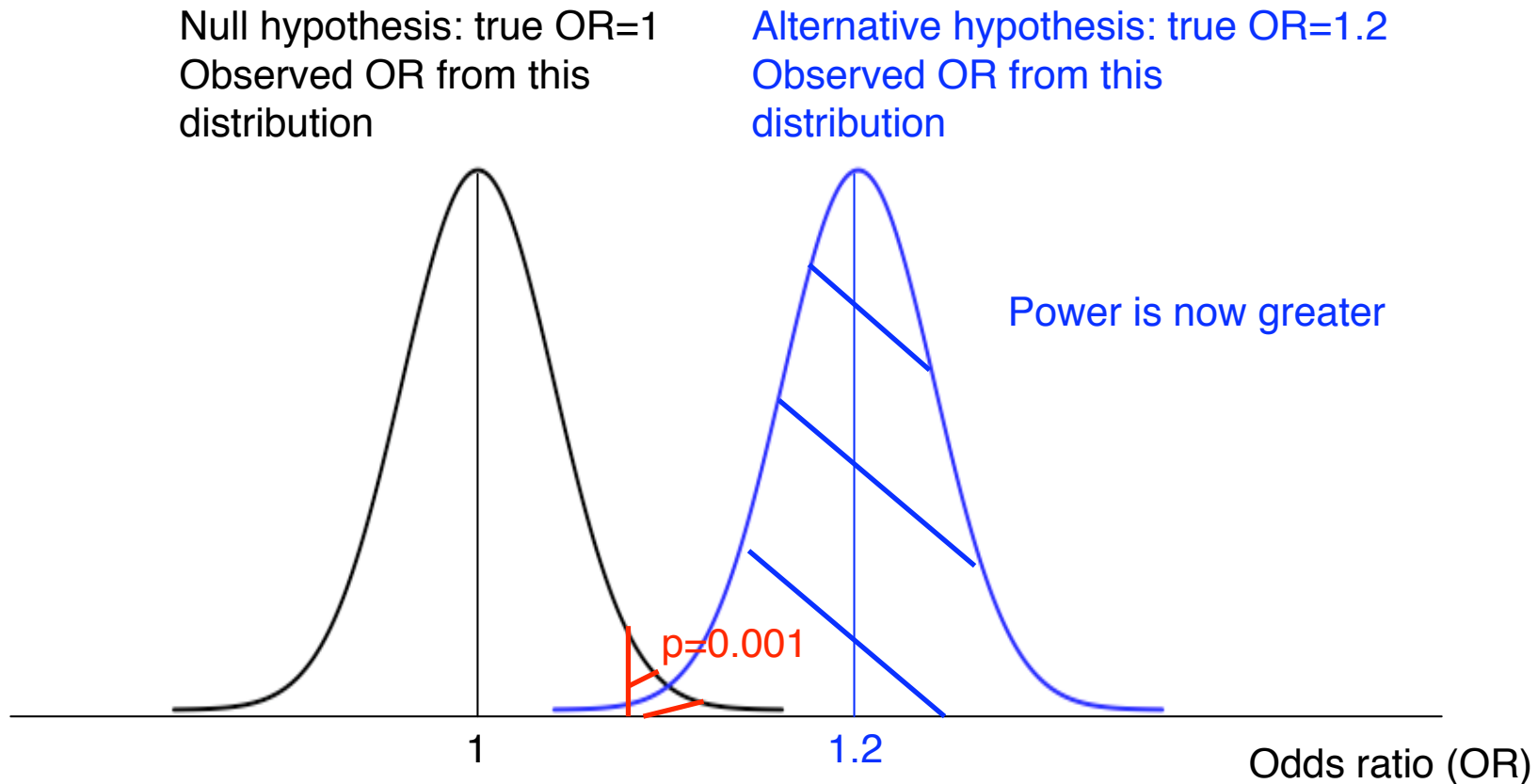


Effect of increasing sample size

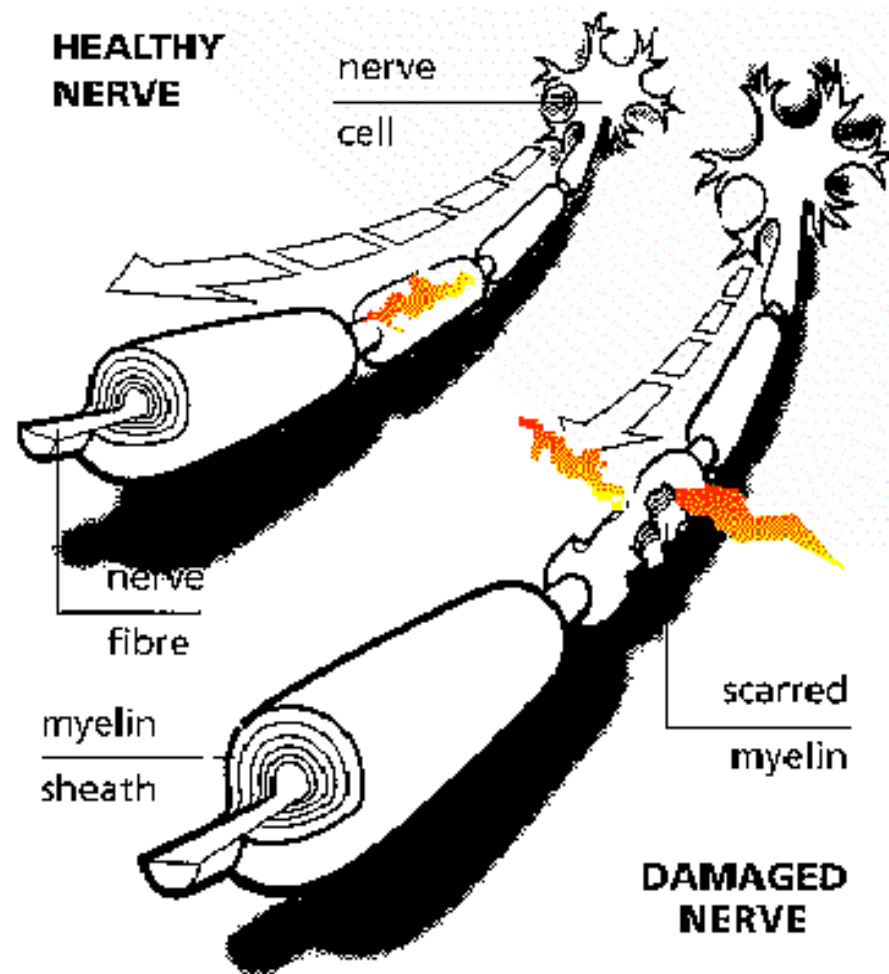
Observed OR tends to be closer to true OR (narrower distributions)

⇒ Null and **alternative** distributions become more separate

⇒ Power increases



Multiple sclerosis - degradation of myelin sheath around nerve fibres

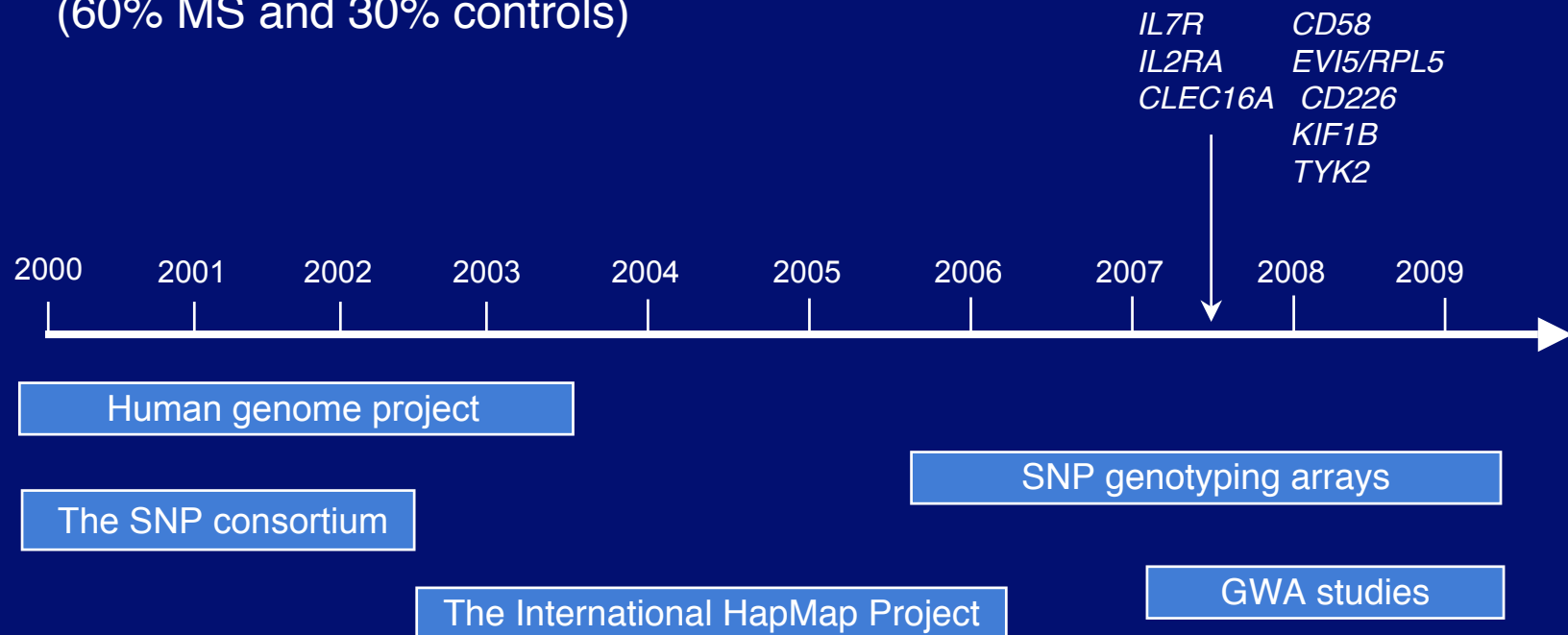


Multiple sclerosis

- neurodegenerative disease
- autoimmune attack on myelin sheaths around nerve cells
- more females affected than males (3:1)
- average age-at-onset ~30
- ~16,000 people with MS in Australia (\$2 billion p.a.)
- no cure

Risk factors

- Epstein-Barr virus
- Exposure to infant siblings (Ponsonby et al, JAMA, 2005)
- Latitude gradient, childhood sun exposure (van der Mei et al, Lancet, 2003)
- Only genetic risk factor known before 2007 (first GWAS):
HLA-DRB1*1501 discovered in 1972
(60% MS and 30% controls)



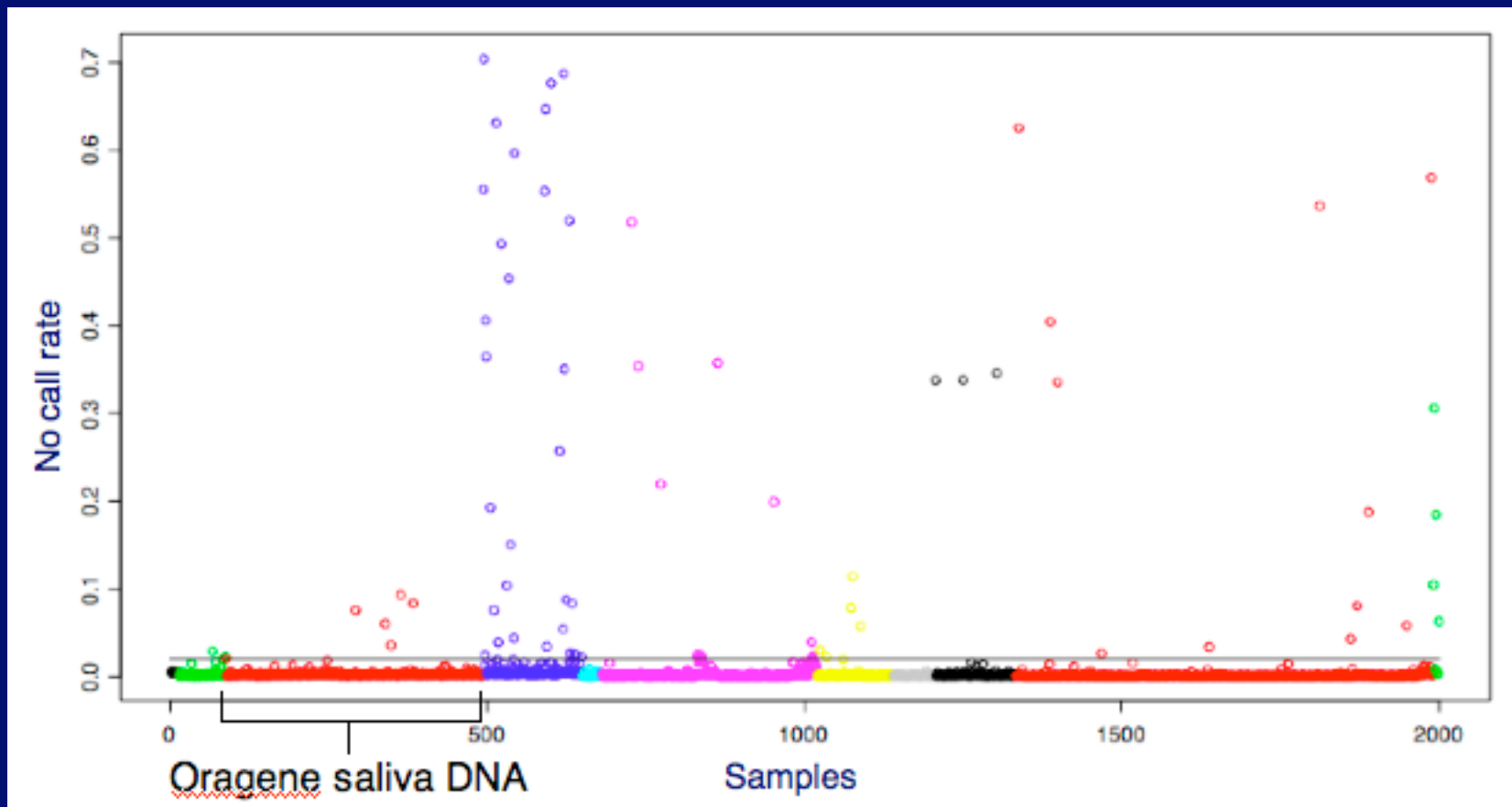
Australian and New Zealand MS GWAS

- Assemble collection of DNA samples (all states + NZ)
- Genotype 1952 MS cases from around Australia and New Zealand with Illumina 370CNV BeadChips
(Patrick Danoy, Matt Brown, Diamantina Institute, UQ)
- Analyse GWAS data
 - Quality control (Devindri Perera, Menzies)
 - Impute genotypes at millions of other SNPs
(Sharon Browning, Univ of Auckland)
 - Compare case genotypes with >3500 controls from the UK and US
(publicly available data)
- Replication genotyping
(Justin Rubio's lab, Howard Florey Institute, Univ of Melbourne)

Quality control - MS samples (PLINK)

- Start with 1952 samples
- Exclusions
 - Samples with >2% of SNPs not called 70 `--mind`

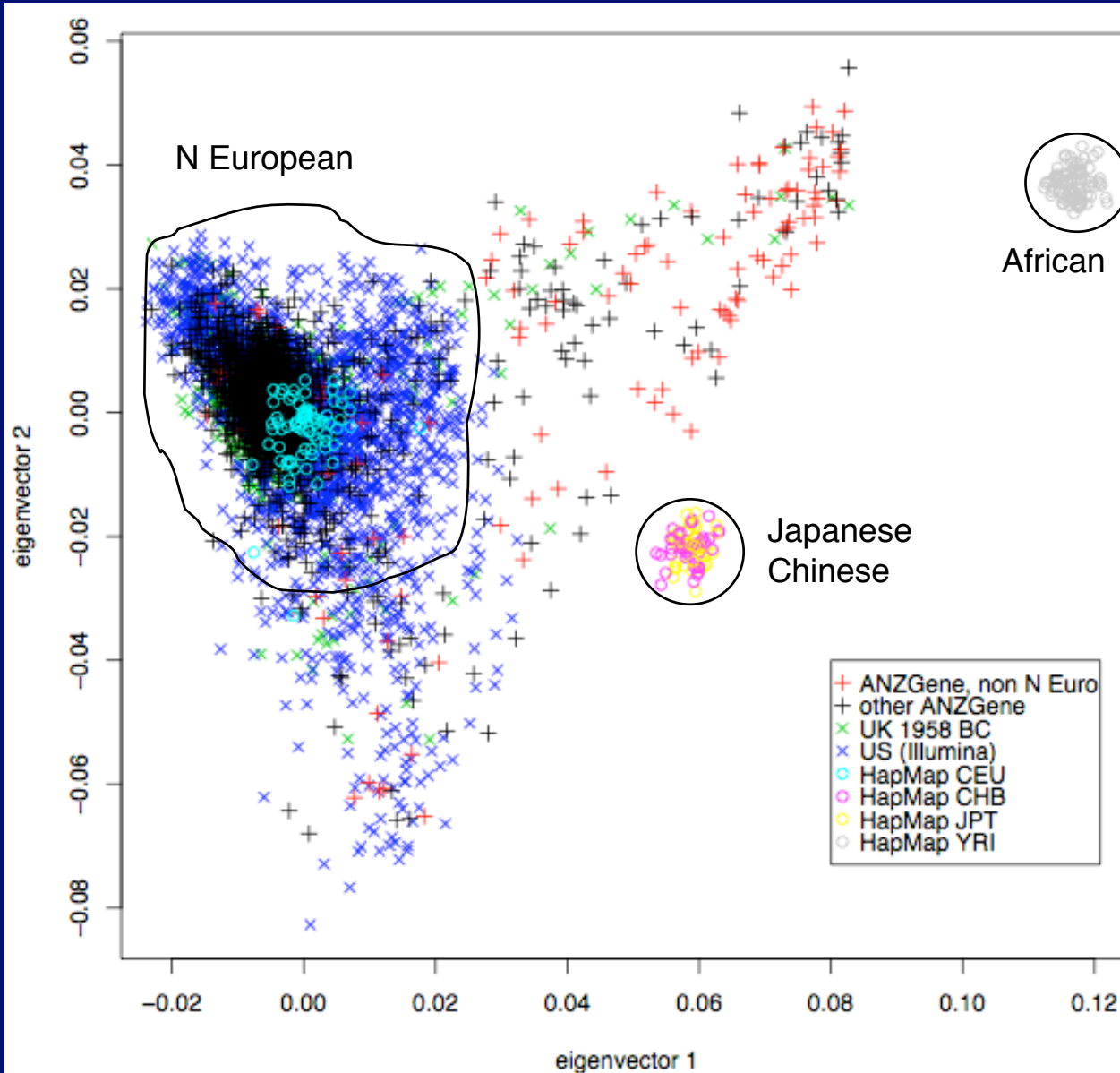
Genotype call rate



Quality control - MS samples (PLINK)

- Start with 1952 samples
- Exclusions
 - Samples with >2% of SNPs not called 70 `--mind`
 - Suspect batch of samples 128
 - Uncertain phenotype 10
 - Duplicates / relatives 88 `--genome`
 - Ancestry outliers 35

Quality control - ethnicity



- Principal components analysis: EIGENSTRAT Price et al (2006). Nat Genet 38: 904

- Use an independent set of ~77,000 SNPs
--indep-pairwise
- 178 outliers removed:
 - 35 MS
 - 143 controls

Quality control - MS samples

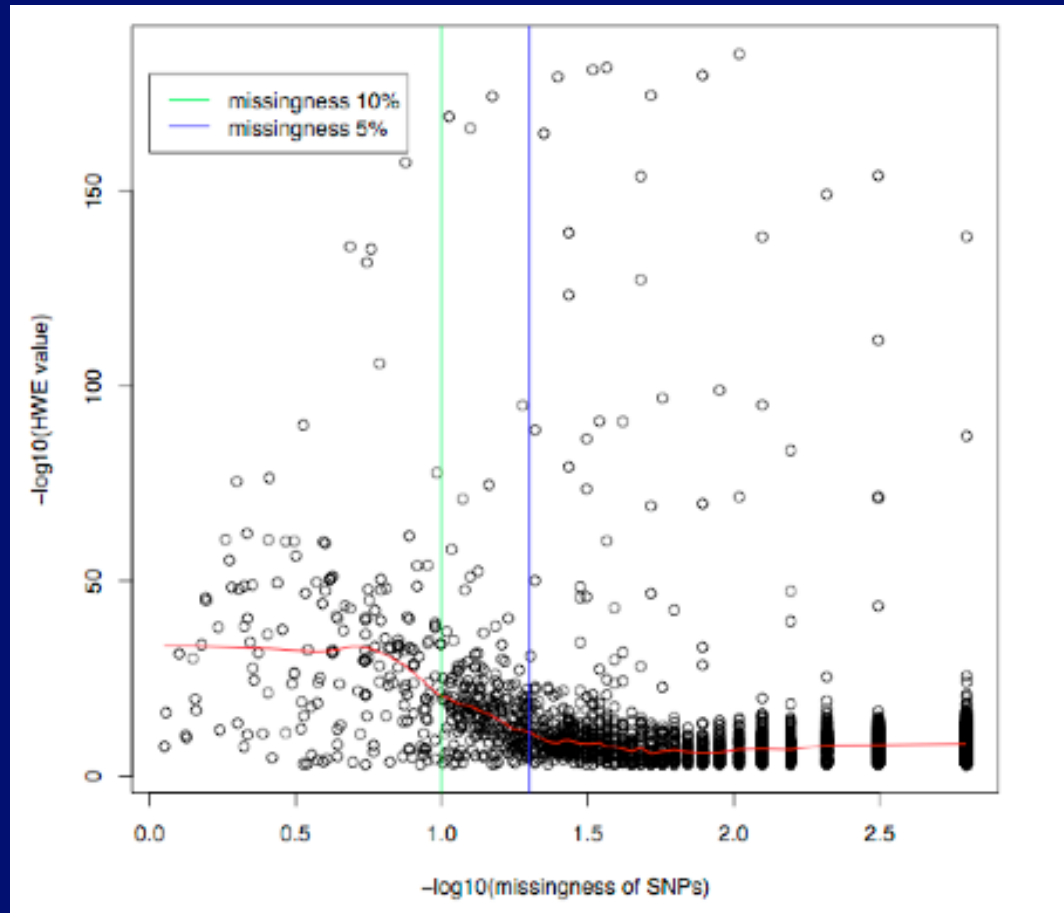
- Start with 1952 samples
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 - Samples with >2% of SNPs not called 70 `--mind`
 - Suspect batch of samples 128
 - Uncertain phenotype 10
 - Duplicates / relatives 88 `--genome`
 - Ancestry outliers 35
 - Sex discrepancies 3 `--check-sex`
- Leaves 1618 samples

Quality control - SNPs

- Start with 310,504 SNPs in both case and control datasets
- Exclude SNPs
 - Not called in >5% of samples `--geno`
 - In Hardy-Weinberg disequilibrium `--hwe`
 - Where one allele has frequency < 1% `--maf`
- Leaves 302,098 SNPs

Choice of 5% no-call threshold

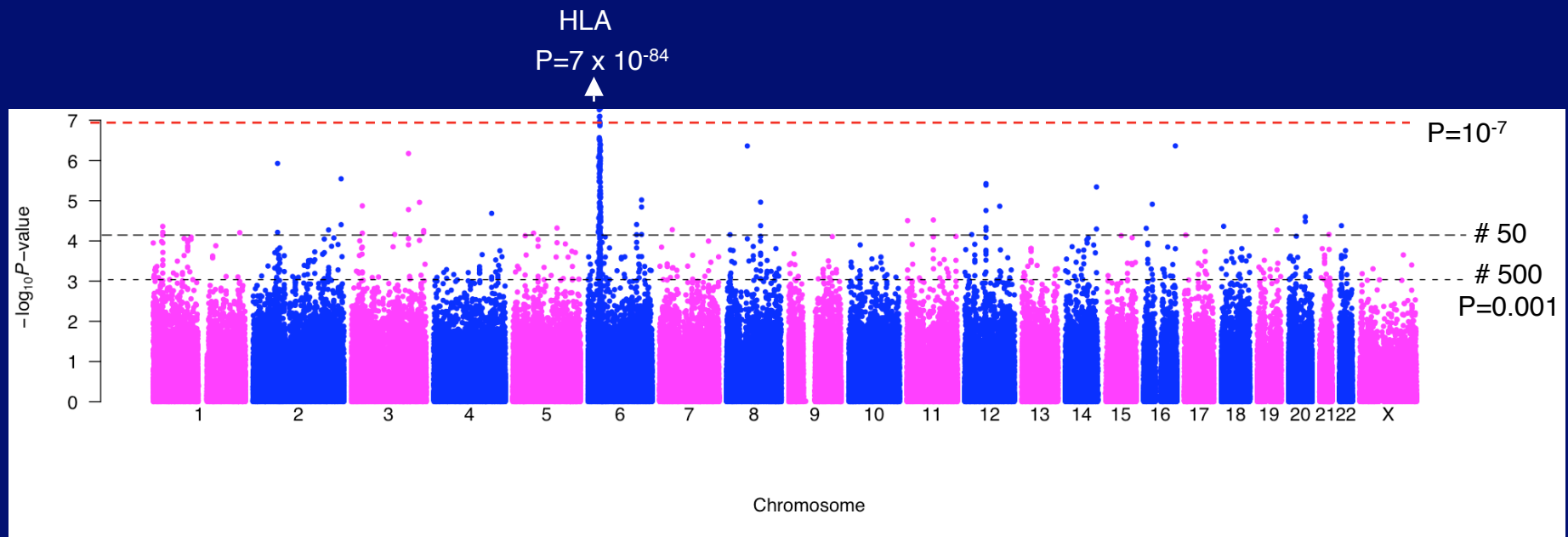
- We originally planned to use a 10% threshold, but lots of SNPs with no call rate 5-10% showed deviations from Hardy-Weinberg equilibrium



- Closer look at SNPs with call rates between 5% and 10% suggested that they were unreliable

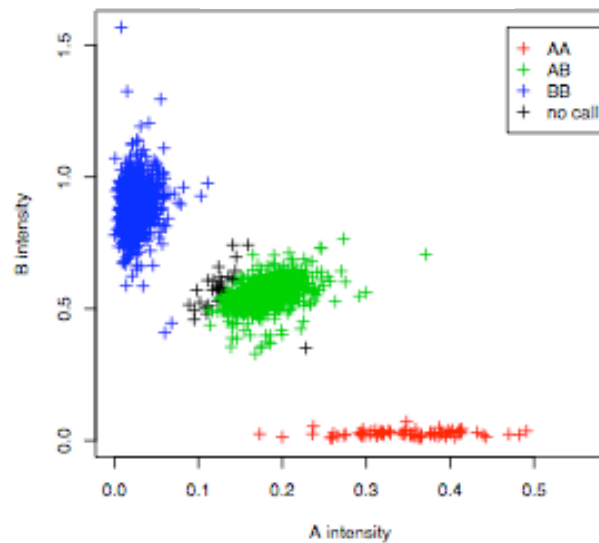
GWAS - results

Total sample = 1618 MS cases + 3413 controls

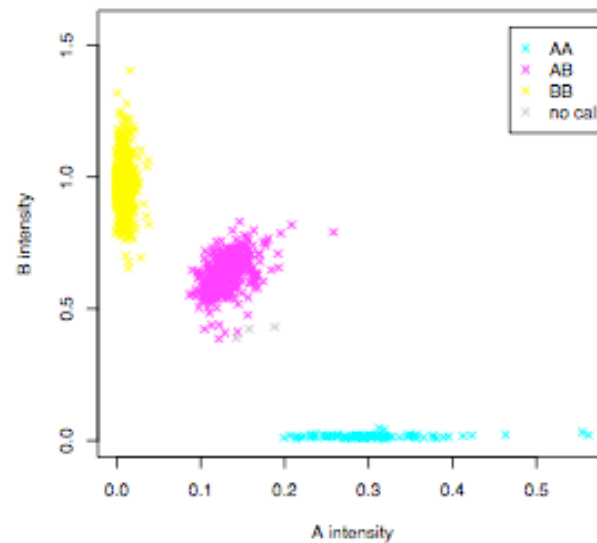


Extra QC for associated SNPs: cluster plots

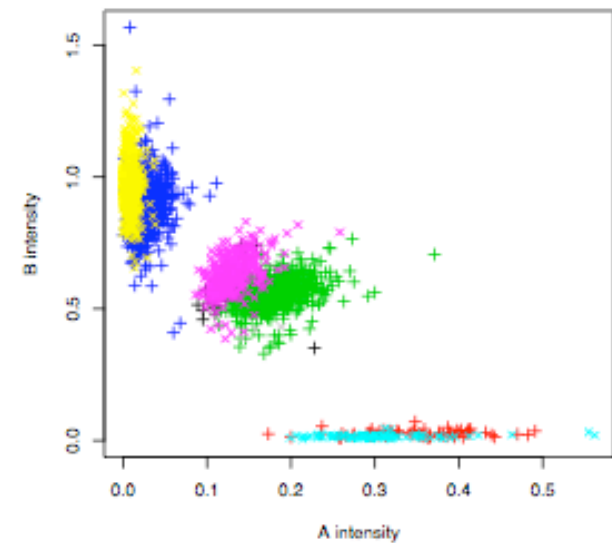
UK controls



ANZ cases



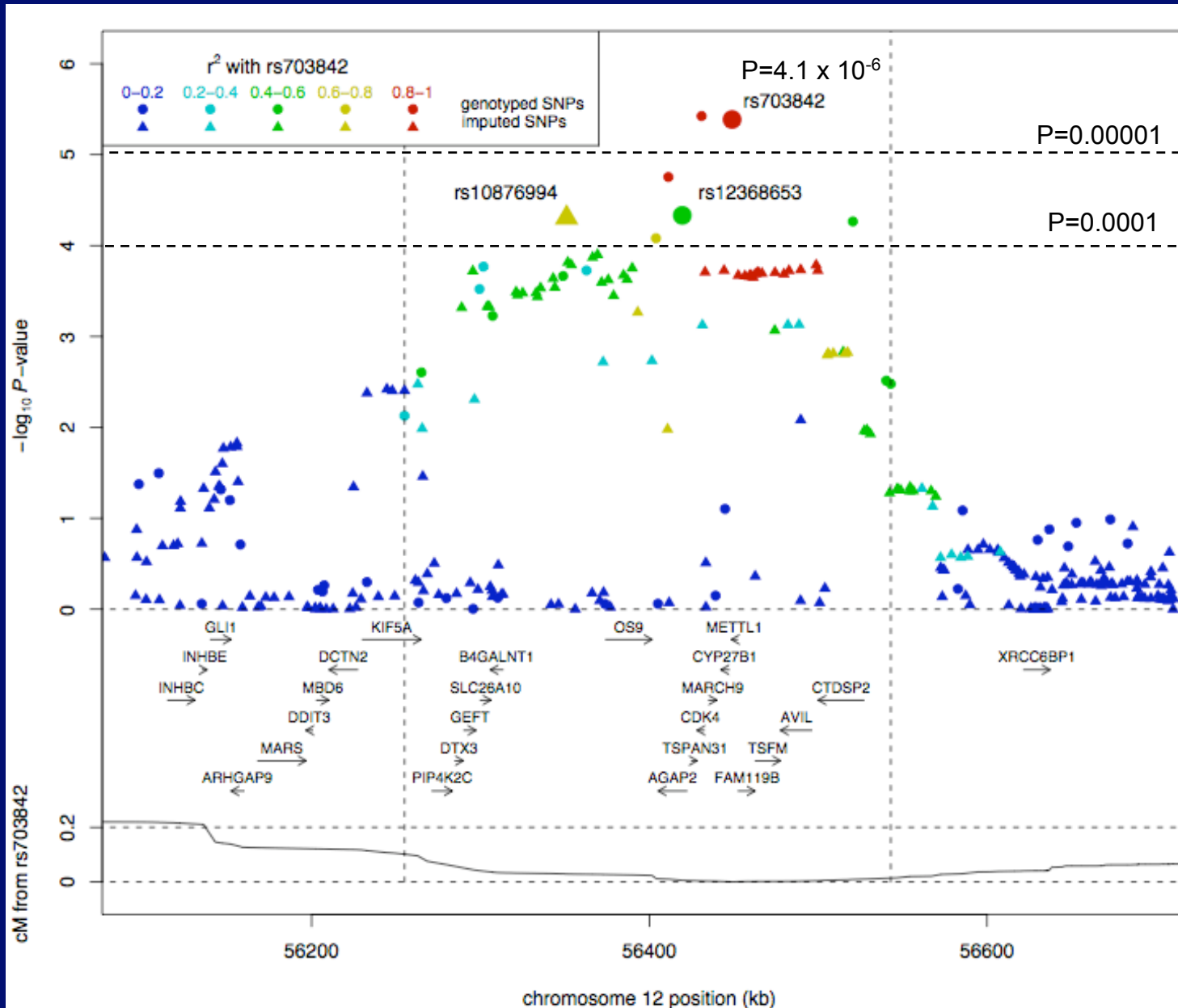
both



The replication phase

- Selected 100 SNPs for replication genotyping
- 2,256 ANZ MS cases + 2,310 ANZ controls
- Two chromosome regions on chr 12 and chr 20 showed (almost) genome-wide significant ($p < 5 \times 10^{-7}$) association with MS after combining GWAS and replication data
- SNPs in 13/53 other regions with replication p-values < 0.1 : more than expected by chance ($p=0.002$)

Chromosome 12 association: the downside of LD



rs703842

GWAS

$P = 4.1 \times 10^{-6}$

replication

$P = 1.4 \times 10^{-6}$

GWAS + rep

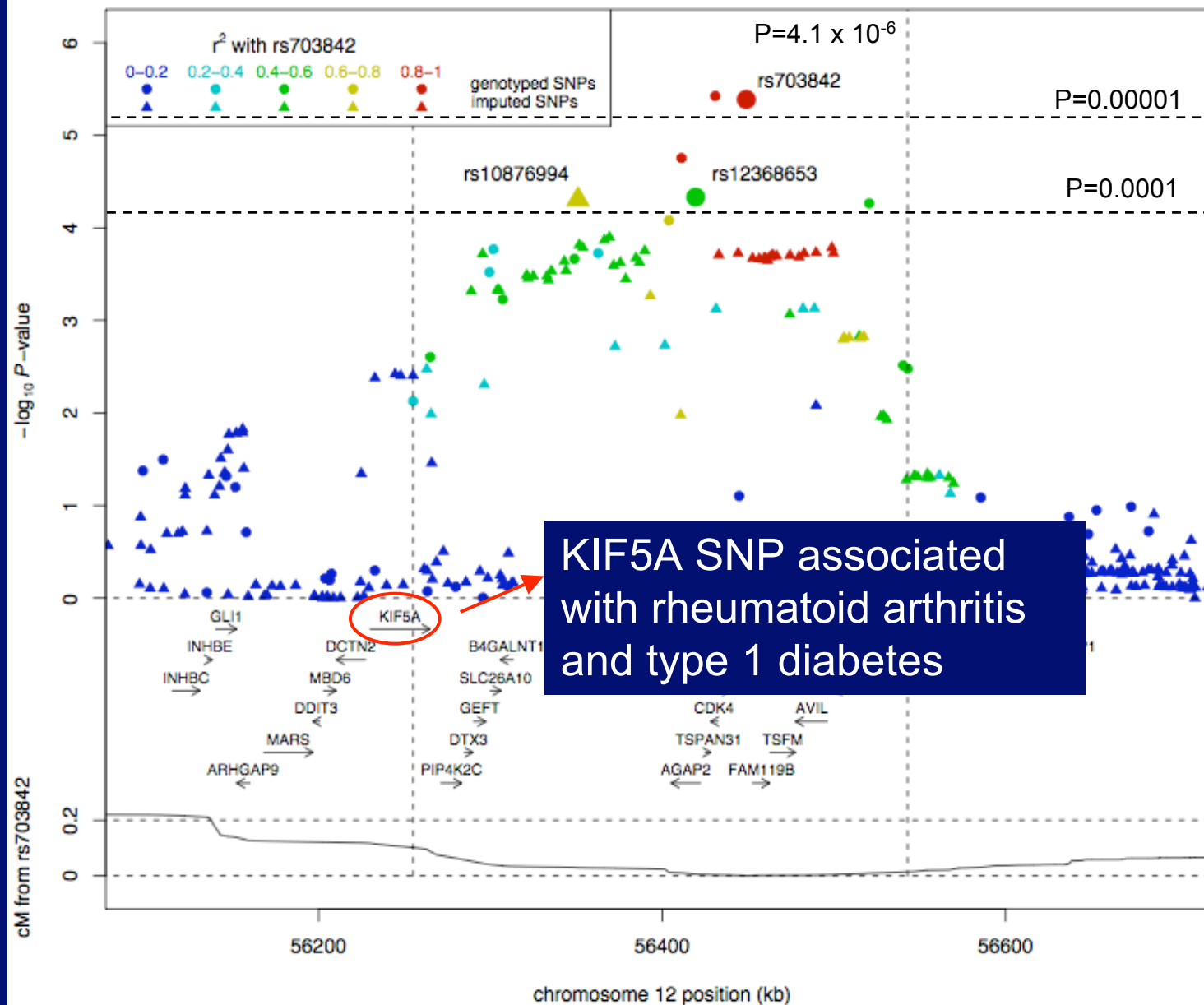
$P = 5.4 \times 10^{-11}$

Allele

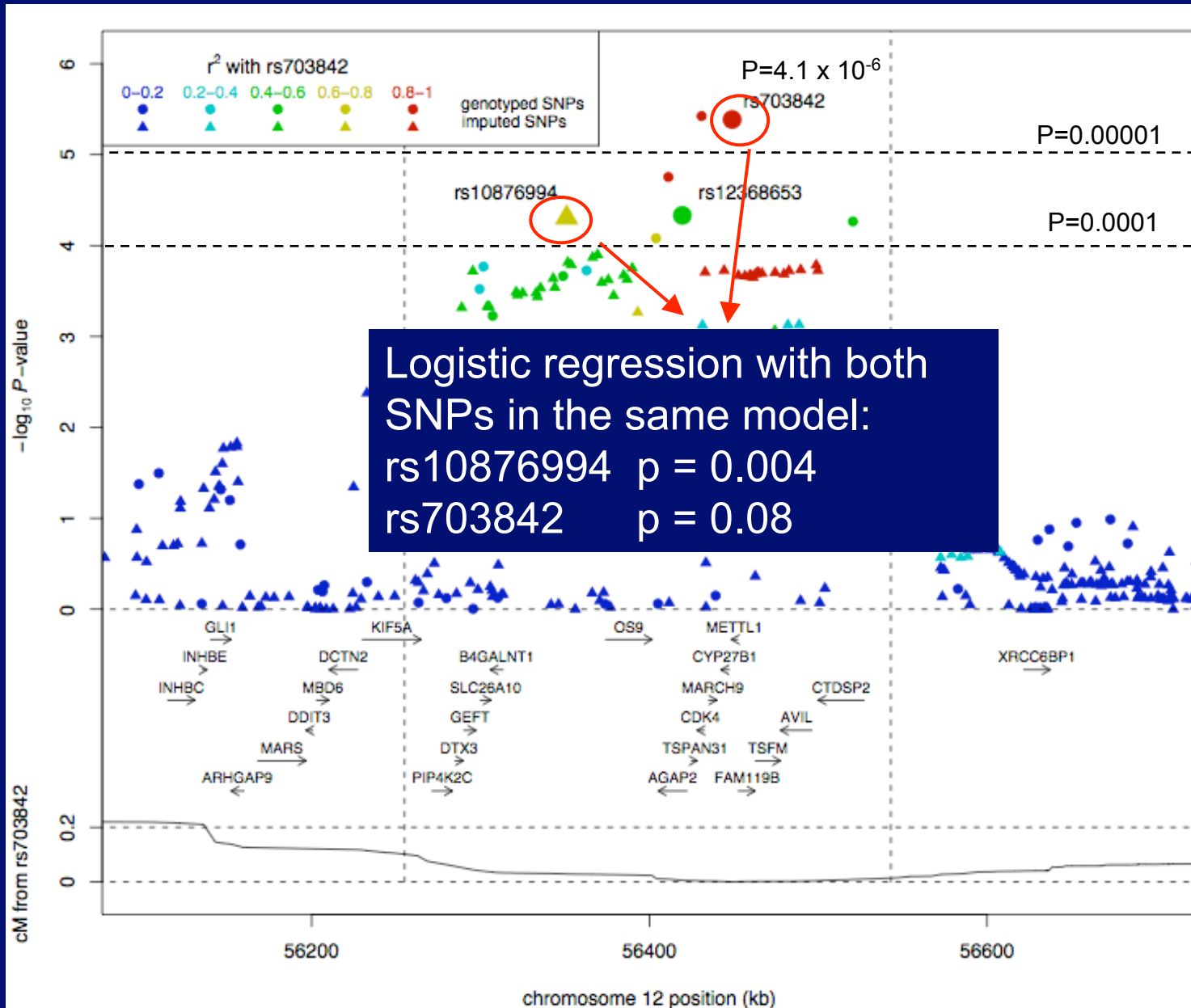
frequency 0.33

Odds ratio 0.81
(protective)

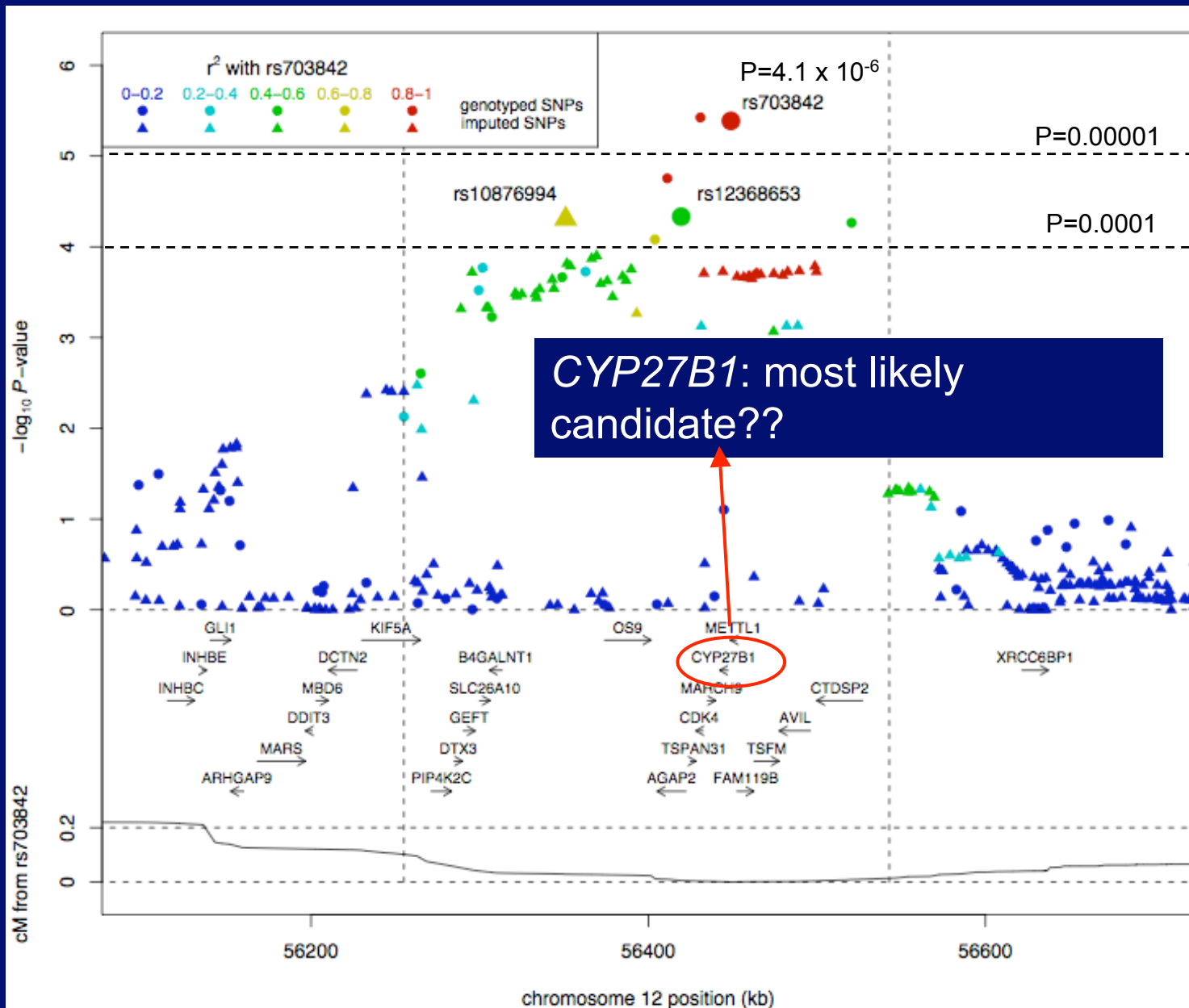
Chromosome 12 association: the downside of LD



Chromosome 12 association: the downside of LD

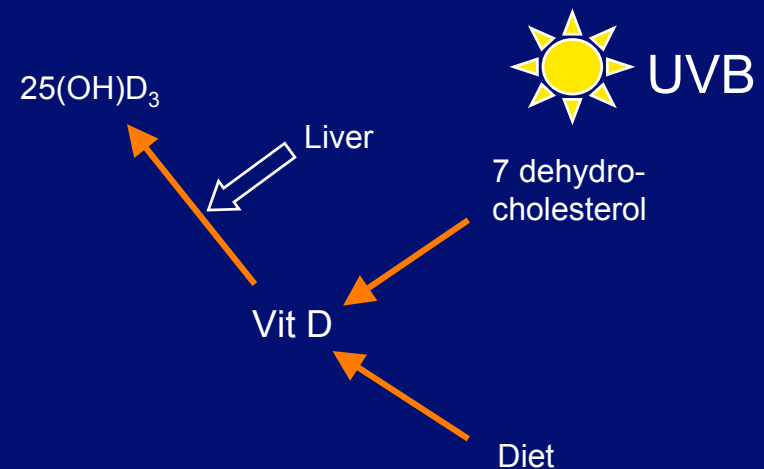
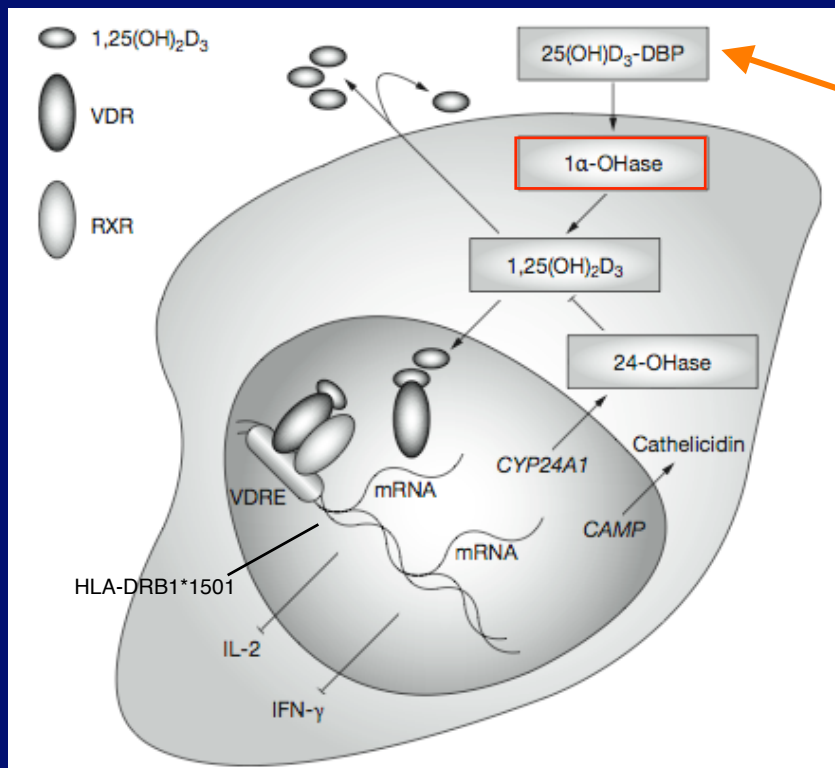


Chromosome 12 association: the downside of LD



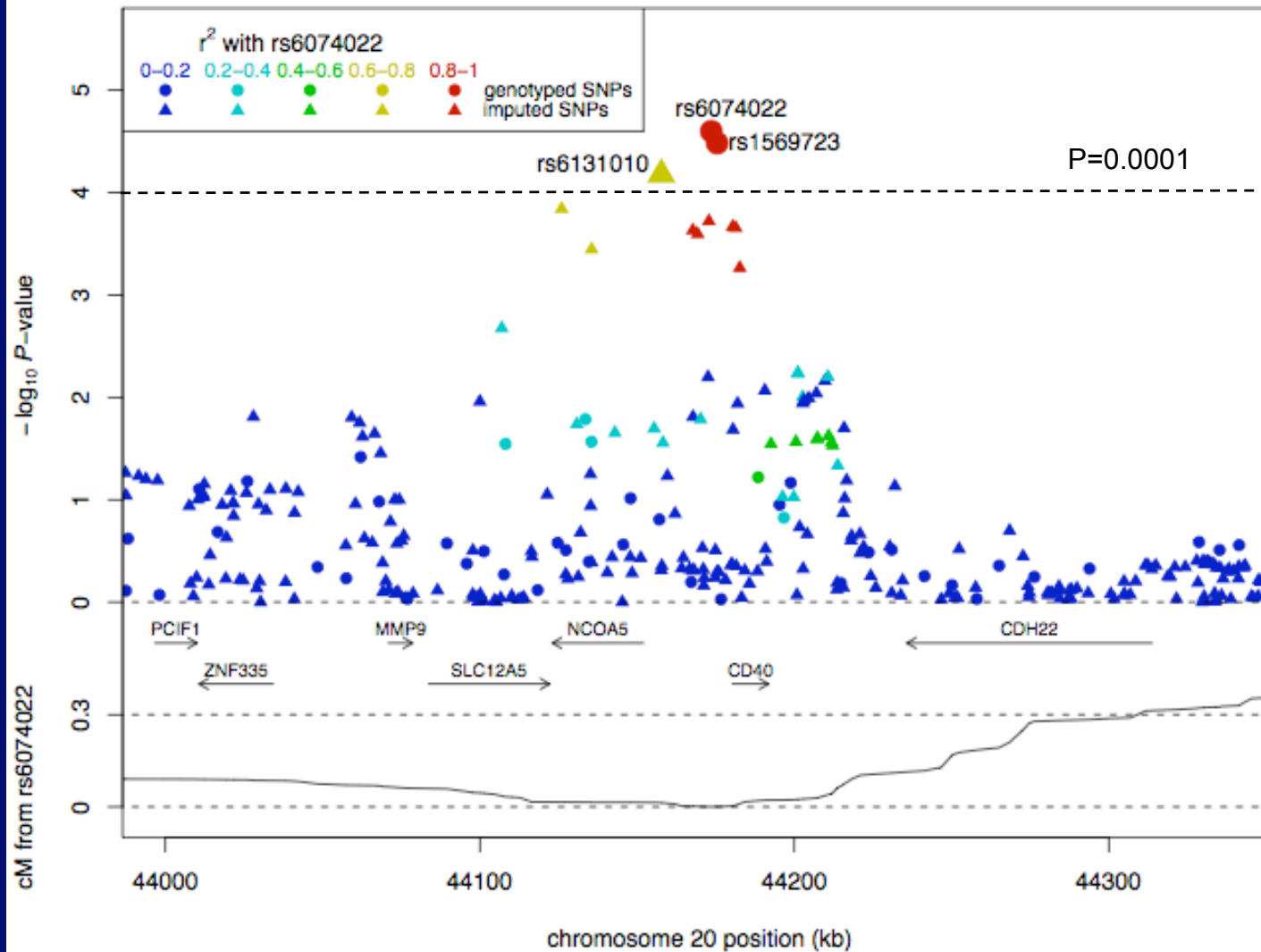
CYP27B1

- Cytochrome p450 gene family (drug metabolizing)
- Encodes 25-hydroxyvitamin D-1 alpha hydroxylase (1α -OHase)
- Converts $25(\text{OH})\text{D}_3$ to bioactive $1,25(\text{OH})_2\text{D}_3$
- $1,25(\text{OH})_2\text{D}_3$ regulates calcium metabolism and the immune system via vitamin D receptor (VDR)



Adorini and Penna (2008)
Nat Clin Prac Rheum 4: 404-12

The chromosome 20 association



rs6074022

GWAS

$P = 2.5 \times 10^{-5}$

replication

$P = 4.6 \times 10^{-4}$

GWAS + rep

$P = 1.3 \times 10^{-7}$

Allele

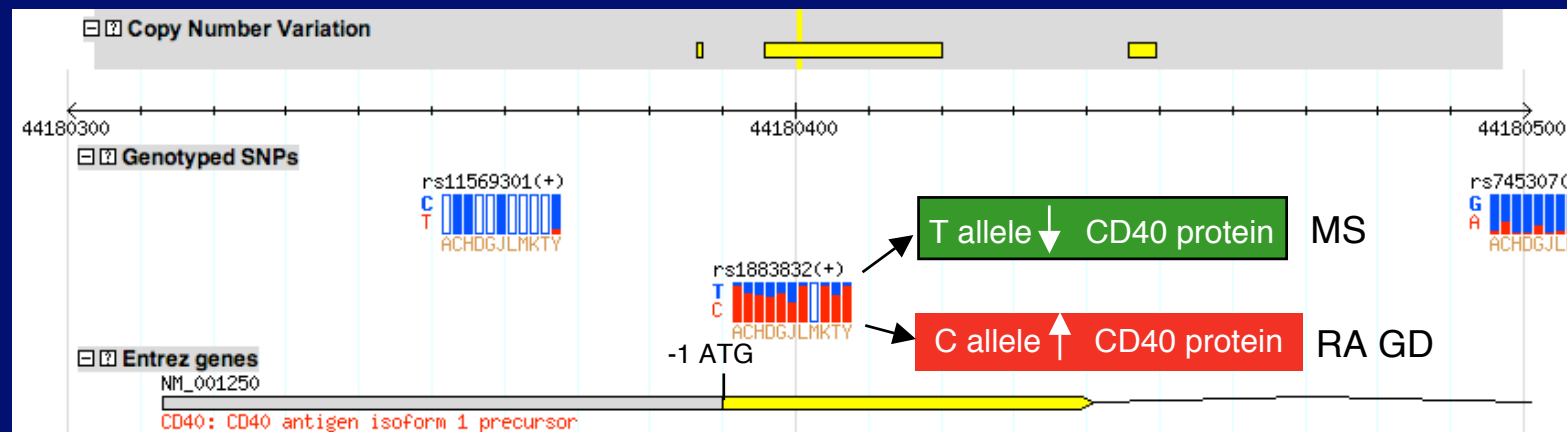
frequency 0.25

Odds ratio 1.20

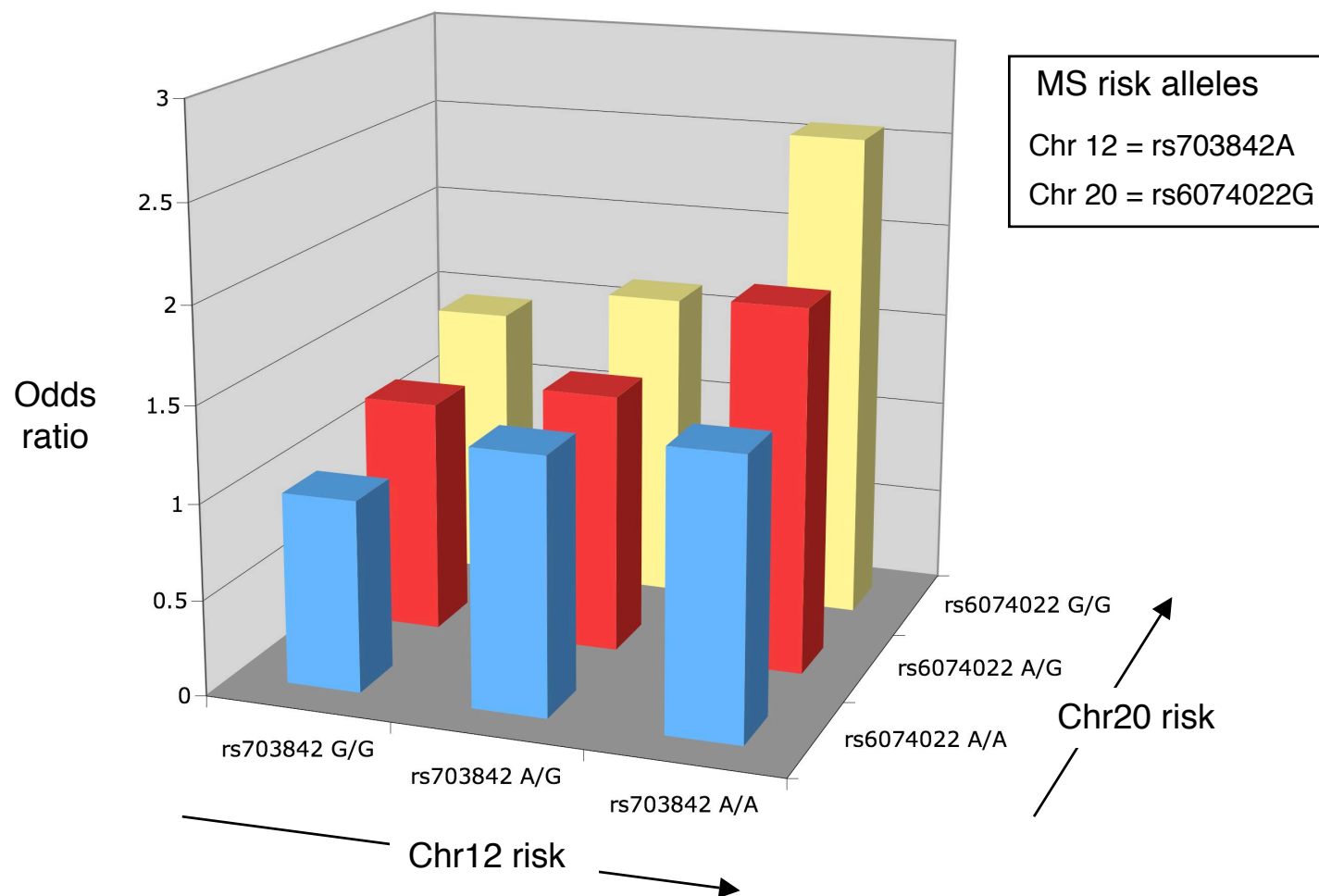
(increased risk)

CD40

- Member of TNF receptor superfamily: regulates many cell- and antibody-mediated immune responses
- SNPs in CD40 are associated with risk of rheumatoid arthritis and Graves' disease
- Functional SNP rs1883832C>T, 1 base pair upstream of the ATG translation initiation codon
- Allelic heterogeneity



Another use of logistic regression: test for gene-gene interaction



Modest evidence that each risk allele has a bigger effect in the presence of the other risk allele ($p = 0.03$)

Summary

- Case-control GWA studies have been very successful in the past couple of years
- Linkage disequilibrium means that most, but not all, common human genetic variation is captured by genotyping a few hundred thousand SNPs
- Small effect sizes (e.g. OR 1.2) mean that GWA studies need to be large, with thousands of cases and controls --> big collaborations
- Methods of statistical analysis are fairly straightforward, but care is required to clean data
- The ultimate test of any association: replication in an independent population

Acknowledgments - MS GWAS

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Pablo Moscato
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Judith Greer

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Glynnis Clarke
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Marilyn Merriman
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Nat Genet 41: 824

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