

# hw6

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## Problem 1

We have table shown below:

$n_{11} = x$	$n_{12} = n_1 - x$	$n_1 = n_{1+}$
$n_{21} = y$	$n_{22} = n_2 - y$	$n_2 = n_{2+}$
$n_{+1}$	$n_{+2}$	

The square of Z statistic is:

$$\left( \frac{\hat{p}_1 - \hat{p}_2}{SE_{\hat{p}_1 - \hat{p}_2}} \right)^2 = \frac{(\hat{p}_1 - \hat{p}_2)^2}{\hat{p}(1 - \hat{p})\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}$$

$$\hat{p} = \frac{n_1 \hat{p}_1 + n_2 \hat{p}_2}{n_1 + n_2}$$

The  $\chi^2$  statistic is:

$$\sum \frac{(\text{observed} - \text{expected})^2}{\text{expected}}$$

So we have:

$$\begin{aligned}\chi^2 &= \frac{(n_1 \hat{p}_1 - n_1 \hat{p})^2}{n_1 \hat{p}} + \frac{(n_1(1 - \hat{p}_1) - n_1(1 - \hat{p}))^2}{n_1(1 - \hat{p})} + \frac{(n_2 \hat{p}_2 - n_2 \hat{p})^2}{n_2 \hat{p}} + \frac{(n_2(1 - \hat{p}_2) - n_2(1 - \hat{p}))^2}{n_2(1 - \hat{p})} \\ &= \frac{n_1(\hat{p}_1 - \hat{p})^2}{\hat{p}} + \frac{n_1(\hat{p}_1 - \hat{p})^2}{1 - \hat{p}} + \frac{n_2(\hat{p}_2 - \hat{p})^2}{\hat{p}} + \frac{n_2(\hat{p}_2 - \hat{p})^2}{1 - \hat{p}} \\ &= \frac{n_1(\hat{p}_1 - \hat{p})^2 + n_2(\hat{p}_2 - \hat{p})^2}{\hat{p}(1 - \hat{p})} = \frac{n_1 \hat{p}_1^2 + n_2 \hat{p}_2^2 - 2(n_1 \hat{p}_1 + n_2 \hat{p}_2)\hat{p} + (n_1 + n_2)\hat{p}^2}{\hat{p}(1 - \hat{p})} \\ &= \frac{n_1 \hat{p}_1^2 + n_2 \hat{p}_2^2 - (n_1 + n_2)\hat{p}^2}{\hat{p}(1 - \hat{p})} \\ &= \frac{n_1 \hat{p}_1^2 + n_2 \hat{p}_2^2 - (n_1 \hat{p}_1 + n_2 \hat{p}_2)^2 / (n_1 + n_2)}{\hat{p}(1 - \hat{p})} \\ &= \frac{\frac{n_1 n_2}{n_1 + n_2}(\hat{p}_1 - \hat{p}_2)^2}{\hat{p}(1 - \hat{p})} = \left( \frac{\hat{p}_1 - \hat{p}_2}{SE_{\hat{p}_1 - \hat{p}_2}} \right)^2\end{aligned}$$

## Problem 2

a. Null:  $H_0 : RD = \hat{p}_1 - \hat{p}_2 = 0$

Alternative:  $H_A : \hat{p}_1 - \hat{p}_2 \neq 0$

$$SE_{RD} = \sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}}$$

```
x=2/15
y=4/19
p=60/340
z=(y-x)/sqrt(p*(1-p)*(1/150+1/190))
2*(1-pnorm(z))
```

```
## [1] 0.06375422
```

We could see p-value = 0.06375422 < 0.05, so we reject null and conclude that the two proportions are not equal.

b.

For absolute change:

$H_0 : RD = \hat{p}_1 - \hat{p}_2 = 0$

$H_A : \hat{p}_1 - \hat{p}_2 \neq 0$

CI=[-0.155191935 0.004424391]

```
x=21/152
y=41/192
(x-y)+c(-1,1)*qnorm(0.975)*sqrt(x*(1-x)/152+y*(1-y)/192)
```

```
## [1] -0.155191935 0.004424391
```

For relative risk: we use log-transformation

$$H_0 : \log(\hat{RR}) = \log\left(\frac{\hat{p}_1}{\hat{p}_2}\right) = 0$$

$$H_A : \log(\hat{RR}) = \log\left(\frac{\hat{p}_1}{\hat{p}_2}\right) \neq 0$$

Interval for logRR = [-2.416758 1.503242]

Exponentiate it and we got CI = [0.3871324 1.0361082]

```
x=20/150
y=40/190
se=sqrt((1-x)/(x*150)+(1-y)/(y*190))
log=log(x/y)+c(-1,1)*1.96*se
exp(log)
```

```
## [1] 0.3871324 1.0361082
```

For odds ratio:

CI = [0.3211208 1.0364953]

```
n11=20
n12=130
n21=40
n22=150
n=(n11*n22)/(n12*n21)
se=sqrt(1/n11+1/n12+1/n21+1/n22)
log=log(n)+c(-1,1)*se*qnorm(0.975)
exp(log)
```

```
## [1] 0.3211208 1.0364953
```

c.

Bayesian credible intervals:

```
x1 = 20; n1 = 150
x2 = 40; n2 = 190
alpha1 = 1; beta1 = 1
alpha2 = 1; beta2 = 1
p1 = rbeta(1000,x1+alpha1, n1-x1+beta1)
p2 = rbeta(1000,x2+alpha2, n2-x2+beta2)
```

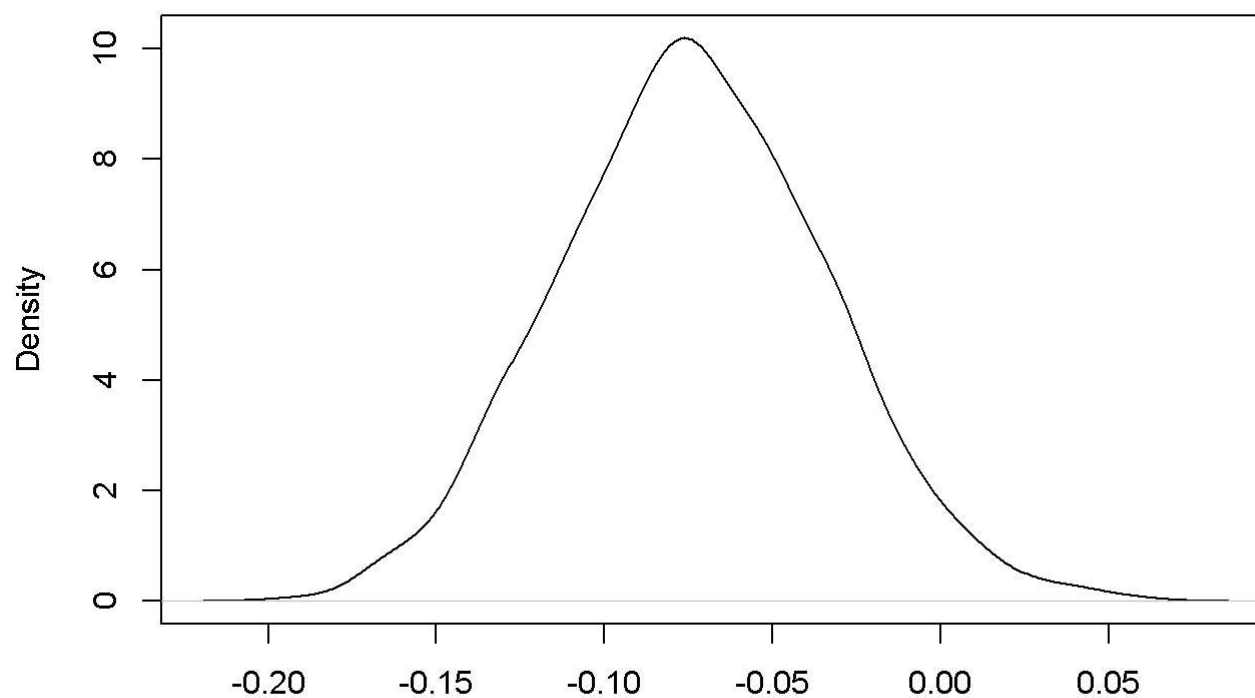
For RD:

```
rd = p1-p2
quantile(rd,c(.025,.975))
```

```
##          2.5%      97.5%  
## -0.149046911  0.004687509
```

```
plot(density(rd))
```

**density.default(x = rd)**



For

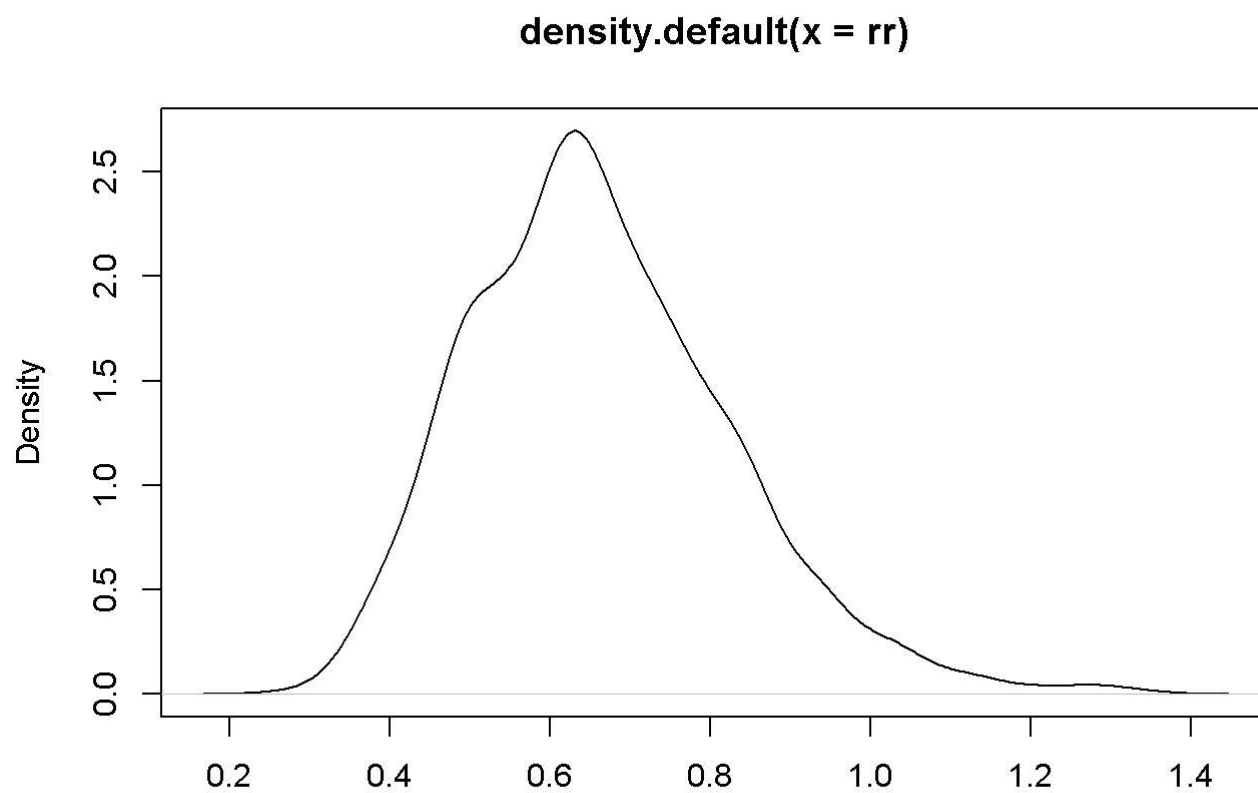
N = 1000 Bandwidth = 0.00894

RR:

```
rr = p1/p2  
quantile(rr,c(0.025,0.975))
```

```
##          2.5%      97.5%  
## 0.3955785  1.0290853
```

```
plot(density(rr))
```



For

N = 1000 Bandwidth = 0.03563

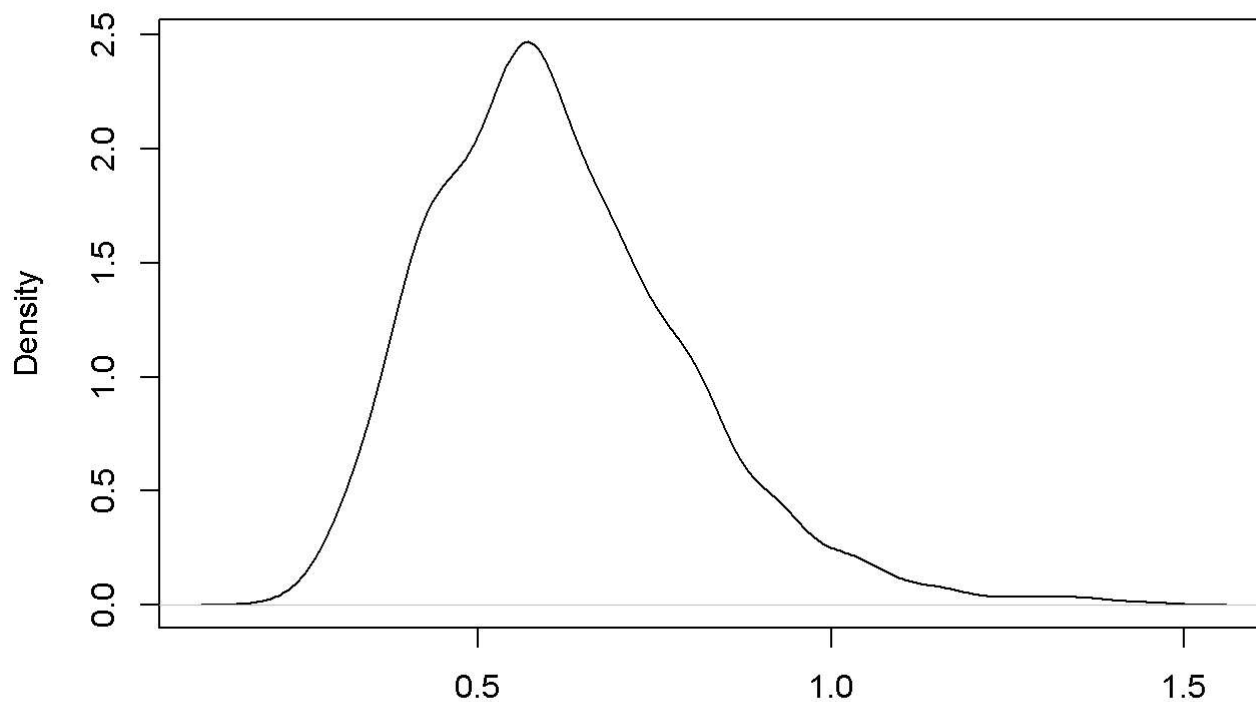
OR:

```
or = p1/(1-p1)/(p2/(1-p2))  
quantile(or,c(0.025,0.975))
```

```
##      2.5%      97.5%  
## 0.3338704 1.0348896
```

```
plot(density(or))
```

## density.default(x = or)



N = 1000 Bandwidth = 0.03926

## Problem 3

95% CI =

$$\left[ \log\left(\frac{x}{n}\right) - 1.96 * \sqrt{\frac{1}{x} - \frac{1}{n}}, \log\left(\frac{x}{n}\right) + 1.96 * \sqrt{\frac{1}{x} - \frac{1}{n}} \right]$$

## problem 4

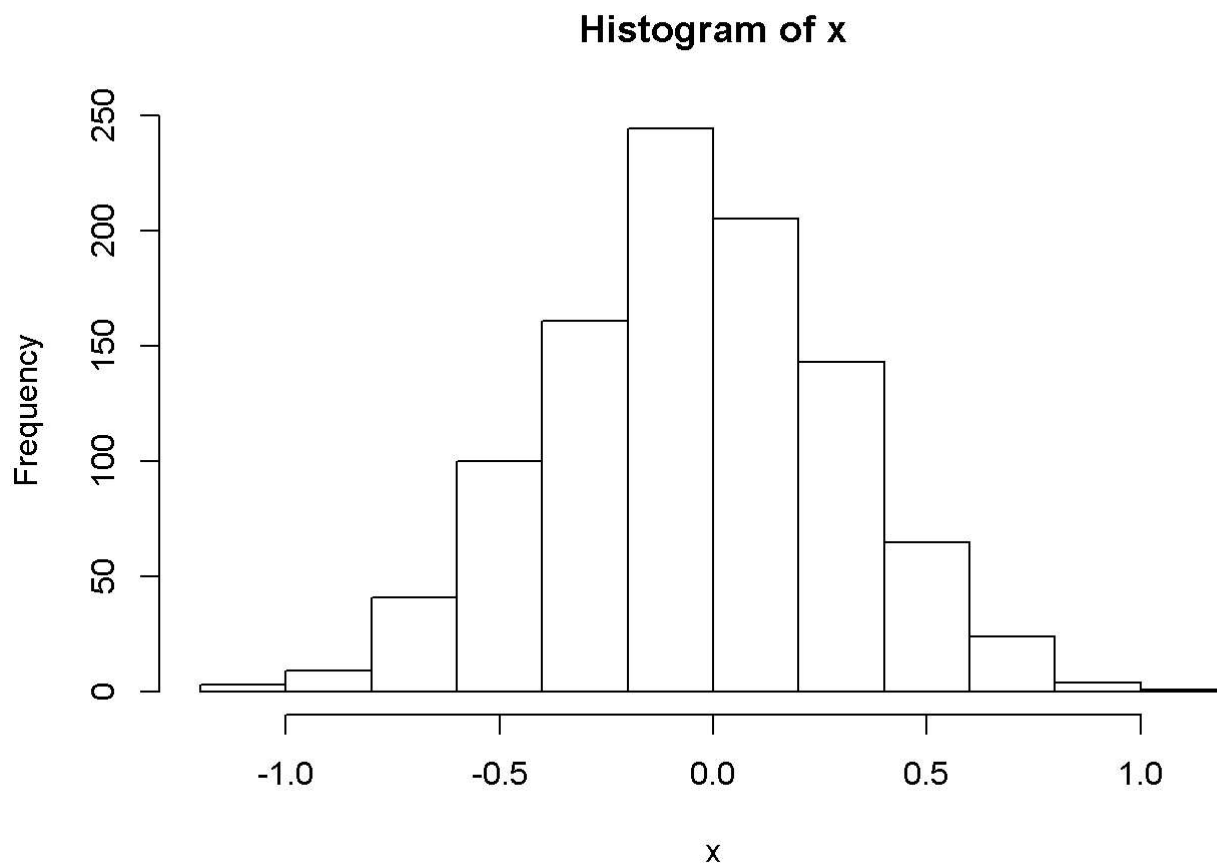
a.  $\frac{\hat{p} - p}{\hat{SE}_{\hat{p}}} \rightarrow N(0, 1)$

$$\frac{\sqrt{\hat{p} - \sqrt{p}}}{\frac{1}{2\sqrt{p}} \hat{SE}_{\hat{p}}} \rightarrow N(0, 1)$$

So standard error for  $\sqrt{\hat{p}}$  is  $\sqrt{\frac{1-\hat{p}}{4n}}$

b. Simulation:

```
x=rep(0,length=1000)
for (i in 1:1000){
x[i]=(sqrt(mean(rbinom(200,1,0.5)))-sqrt(0.5))/((1-sqrt(mean(rbinom(200,1,0.5))))/4)
}
hist(x)
```



## Problem 5

For  $p_1 = .1; p_2 = .1; n_1 = 100; n_2 = 100$

```
p1 = .1; p2 = .1; n1 = 100; n2 = 100
p = log(p1*(1-p2)/((p2)*(1-p1)))
x=rbinom(1000,n1,p1)
y=rbinom(1000,n2,p2)
z=cbind(x,y)
OR = apply(z, 1, function(x) x[1] * (n2 - x[2]) / (x[2] * (n1 - x[1])) )
SELOGOR = apply(z, 1, function(x) sqrt(1 / x[1] + 1 / (n1 - x[1]) + 1 / x[2] + 1 / (n2 - x[2]) ) )
INTERVAL = cbind(log(OR)-1.96*SELOGOR,log(OR)+1.96*SELOGOR)
s=0
for (i in 1:1000) {
  if (p>=INTERVAL[i,1] & p<=INTERVAL[i,2]){
    s= s+1
  }
}
s/1000
```

```
## [1] 0.943
```

For  $p_1 = .1; p_2 = .5; n_1 = 100; n_2 = 100$

```

p1 = .1; p2 = .5; n1 = 100; n2 = 100
p = log(p1*(1-p2)/((p2)*(1-p1)))
x=rbinom(1000,n1,p1)
y=rbinom(1000,n2,p2)
z=cbind(x,y)
OR = apply(z, 1, function(x) x[1] * (n2 - x[2]) / (x[2] * (n1 - x[1])) )
SELOGOR = apply(z, 1, function(x) sqrt(1 / x[1] + 1 / (n1 - x[1]) + 1 / x[2] + 1 / (n2 - x[2]) ) )
INTERVAL = cbind(log(OR)-1.96*SELOGOR,log(OR)+1.96*SELOGOR)
s=0
for (i in 1:1000) {
  if (p>=INTERVAL[i,1] & p<=INTERVAL[i,2]){
    s= s+1
  }
}
s/1000

```

```
## [1] 0.964
```

For  $p1 = .1; p2 = .9; n1 = 100; n2 = 100$

```

p1 = .1; p2 = .9; n1 = 100; n2 = 100
p = log(p1*(1-p2)/((p2)*(1-p1)))
x=rbinom(1000,n1,p1)
y=rbinom(1000,n2,p2)
z=cbind(x,y)
OR = apply(z, 1, function(x) x[1] * (n2 - x[2]) / (x[2] * (n1 - x[1])) )
SELOGOR = apply(z, 1, function(x) sqrt(1 / x[1] + 1 / (n1 - x[1]) + 1 / x[2] + 1 / (n2 - x[2]) ) )
INTERVAL = cbind(log(OR)-1.96*SELOGOR,log(OR)+1.96*SELOGOR)
s=0
for (i in 1:1000) {
  if (p>=INTERVAL[i,1] & p<=INTERVAL[i,2]){
    s= s+1
  }
}
s/1000

```

```
## [1] 0.963
```

For  $p1 = .5; p2 = .5; n1 = 100; n2 = 100$



```

p1 = .5; p2 = .5; n1 = 100; n2 = 100
p = log(p1*(1-p2)/((p2)*(1-p1)))
x=rbinom(1000,n1,p1)
y=rbinom(1000,n2,p2)
z=cbind(x,y)
OR = apply(z, 1, function(x) x[1] * (n2 - x[2]) / (x[2] * (n1 - x[1])) )
SELOGOR = apply(z, 1, function(x) sqrt(1 / x[1] + 1 / (n1 - x[1]) + 1 / x[2] + 1 / (n2 - x[2]) ) )
INTERVAL = cbind(log(OR)-1.96*SELOGOR,log(OR)+1.96*SELOGOR)
s=0
for (i in 1:1000) {
  if (p>=INTERVAL[i,1] & p<=INTERVAL[i,2]){
    s= s+1
  }
}
s/1000

```

```
## [1] 0.954
```

For  $p1 = .5; p2 = .9; n1 = 100; n2 = 100$

```

p1 = .5; p2 = .9; n1 = 100; n2 = 100
p = log(p1*(1-p2)/((p2)*(1-p1)))
x=rbinom(1000,n1,p1)
y=rbinom(1000,n2,p2)
z=cbind(x,y)
OR = apply(z, 1, function(x) x[1] * (n2 - x[2]) / (x[2] * (n1 - x[1])) )
SELOGOR = apply(z, 1, function(x) sqrt(1 / x[1] + 1 / (n1 - x[1]) + 1 / x[2] + 1 / (n2 - x[2]) ) )
INTERVAL = cbind(log(OR)-1.96*SELOGOR,log(OR)+1.96*SELOGOR)
s=0
for (i in 1:1000) {
  if (p>=INTERVAL[i,1] & p<=INTERVAL[i,2]){
    s= s+1
  }
}
s/1000

```

```
## [1] 0.958
```

For  $p1 = .9; p2 = .9; n1 = 100; n2 = 100$

```

p1 = .9; p2 = .9; n1 = 100; n2 = 100
p = log(p1*(1-p2)/((p2)*(1-p1)))
x=rbinom(1000,n1,p1)
y=rbinom(1000,n2,p2)
z=cbind(x,y)
OR = apply(z, 1, function(x) x[1] * (n2 - x[2]) / (x[2] * (n1 - x[1])) )
SELOGOR = apply(z, 1, function(x) sqrt(1 / x[1] + 1 / (n1 - x[1]) + 1 / x[2] + 1 / (n2 - x[2]) ) )
INTERVAL = cbind(log(OR)-1.96*SELOGOR,log(OR)+1.96*SELOGOR)
s=0
for (i in 1:1000) {
  if (p>=INTERVAL[i,1] & p<=INTERVAL[i,2]){
    s= s+1
  }
}
s/1000

```

```
## [1] 0.958
```

Summary: There's no big difference in intervals' performance

## Problem 6

For  $p1 = .1; p2 = .1; n1 = 100; n2 = 100$

```

p1 = .1; p2 = .1; n1 = 100; n2 = 100
p = log(p1/p2)
x=rbinom(1000,n1,p1)
y=rbinom(1000,n2,p2)
z=cbind(x,y)
RR = apply(z, 1,
function(x) (x[1] / n1) / (x[2] / n2))
SELOGRR = apply(z, 1,
function(x) {
  phat1 <- x[1] / n1
  phat2 <- x[2] / n2
  sqrt((1 - phat1) / phat1 / n1 + (1 - phat2) / phat2 / n2)
})
INTERVAL = cbind(log(OR)-1.96*SELOGOR,log(OR)+1.96*SELOGOR)
s=0
for (i in 1:1000) {
  if (p>= INTERVAL[i,1] & p<=INTERVAL[i,2]){
    s= s+1
  }
}
s/1000

```

```
## [1] 0.958
```

For  $p1 = .1; p2 = .5; n1 = 100; n2 = 100$

```

p1 = .1; p2 = .5; n1 = 100; n2 = 100
p = log(p1/p2)
x=rbinom(1000,n1,p1)
y=rbinom(1000,n2,p2)
z=cbind(x,y)
RR = apply(z, 1,
function(x) (x[1] / n1) / (x[2] / n2))
SELOGRR = apply(z, 1,
function(x) {
  phat1 <- x[1] / n1
  phat2 <- x[2] / n2
  sqrt((1 - phat1) / phat1 / n1 + (1 - phat2) / phat2 / n2)
})
INTERVAL = cbind(log(RR)-1.96*SELOGRR,log(RR)+1.96*SELOGRR)
s=0
for (i in 1:1000) {
  if (p>= INTERVAL[i,1] & p<=INTERVAL[i,2]){
    s= s+1
  }
}
s/1000

```

```
## [1] 0.953
```

For  $p1 = .1; p2 = .9; n1 = 100; n2 = 100$

```

p1 = .1; p2 = .9; n1 = 100; n2 = 100
p = log(p1/p2)
x=rbinom(1000,n1,p1)
y=rbinom(1000,n2,p2)
z=cbind(x,y)
RR = apply(z, 1,
function(x) (x[1] / n1) / (x[2] / n2))
SELOGRR = apply(z, 1,
function(x) {
  phat1 <- x[1] / n1
  phat2 <- x[2] / n2
  sqrt((1 - phat1) / phat1 / n1 + (1 - phat2) / phat2 / n2)
})
INTERVAL = cbind(log(RR)-1.96*SELOGRR,log(RR)+1.96*SELOGRR)
s=0
for (i in 1:1000) {
  if (p>= INTERVAL[i,1] & p<=INTERVAL[i,2]){
    s= s+1
  }
}
s/1000

```

```
## [1] 0.951
```

For  $p1 = .5; p2 = .5; n1 = 100; n2 = 100$

```

p1 = .5; p2 = .5; n1 = 100; n2 = 100
p = log(p1/p2)
x=rbinom(1000,n1,p1)
y=rbinom(1000,n2,p2)
z=cbind(x,y)
RR = apply(z, 1,
function(x) (x[1] / n1) / (x[2] / n2))
SELOGRR = apply(z, 1,
function(x) {
  phat1 <- x[1] / n1
  phat2 <- x[2] / n2
  sqrt((1 - phat1) / phat1 / n1 + (1 - phat2) / phat2 / n2)
})
INTERVAL = cbind(log(RR)-1.96*SELOGRR,log(RR)+1.96*SELOGRR)
s=0
for (i in 1:1000) {
  if (p>= INTERVAL[i,1] & p<=INTERVAL[i,2]){
    s= s+1
  }
}
s/1000

```

```
## [1] 0.945
```

For  $p1 = .5; p2 = .9; n1 = 100; n2 = 100$

```

p1 = .5; p2 = .9; n1 = 100; n2 = 100
p = log(p1/p2)
x=rbinom(1000,n1,p1)
y=rbinom(1000,n2,p2)
z=cbind(x,y)
RR = apply(z, 1,
function(x) (x[1] / n1) / (x[2] / n2))
SELOGRR = apply(z, 1,
function(x) {
  phat1 <- x[1] / n1
  phat2 <- x[2] / n2
  sqrt((1 - phat1) / phat1 / n1 + (1 - phat2) / phat2 / n2)
})
INTERVAL = cbind(log(RR)-1.96*SELOGRR,log(RR)+1.96*SELOGRR)
s=0
for (i in 1:1000) {
  if (p>= INTERVAL[i,1] & p<=INTERVAL[i,2]){
    s= s+1
  }
}
s/1000

```

```
## [1] 0.957
```

For  $p1 = .9; p2 = .9; n1 = 100; n2 = 100$

```

p1 = .9; p2 = .9; n1 = 100; n2 = 100
p = log(p1/p2)
x=rbinom(1000,n1,p1)
y=rbinom(1000,n2,p2)
z=cbind(x,y)
RR = apply(z, 1,
function(x) (x[1] / n1) / (x[2] / n2))
SELOGRR = apply(z, 1,
function(x) {
  phat1 <- x[1] / n1
  phat2 <- x[2] / n2
  sqrt((1 - phat1) / phat1 / n1 + (1 - phat2) / phat2 / n2)
})
INTERVAL = cbind(log(RR)-1.96*SELOGRR,log(RR)+1.96*SELOGRR)
s=0
for (i in 1:1000) {
  if (p>= INTERVAL[i,1] & p<=INTERVAL[i,2]){
    s= s+1
  }
}
s/1000

```

```
## [1] 0.958
```

Summary: There's no big difference in intervals' performance

## Problem 7

a. The 95% CI of the odds ratio is [2.055483 10.624095]

```

n11=243
n12=259-243
n21=39
n22=51-39
OR = n11*n22/(n21*n12)
SE = sqrt(1/n11+1/n12+1/n21+1/n22)
INTERVAL = exp(c(log(OR)-1.96*SE,log(OR)+1.96*SE))
OR

```

```
## [1] 4.673077
```

```
INTERVAL
```

```
## [1] 2.055483 10.624095
```

b. The 95% CI of the relative risk is [1.050307 1.433203]

```

n1=259
n2=51
p1=243/259
p2=39/51
RR=p1/p2
SE = sqrt((1-p1)/(p1*n1)+(1-p2)/(p2*n2))
INTERVAL = exp(c(log(RR)-1.96*SE,log(RR)+1.96*SE))
RR

```

```
## [1] 1.226908
```

```
INTERVAL
```

```
## [1] 1.050307 1.433203
```

c. The 95% CI of the risk difference is [0.05346365 0.29357246]

```

n1=259
n2=51
p1=243/259
p2=39/51
RD=p1 - p2
SE = sqrt(p1*(1-p1)/n1+p2*(1-p2)/n2)
INTERVAL=c(RD-1.96*SE,RD+1.96*SE)
RD

```

```
## [1] 0.1735181
```

```
INTERVAL
```

```
## [1] 0.05346365 0.29357246
```

d.  $H_0 : p1 = p2$

$H_A : p1 \neq p2$

p-value = 7.767783e-05 < 0.05, so we reject the null and conclude that  $p1 \neq p2$

```

O11=243;E11=282/310*259
O21=39;E21=282/310*51
O12=16;E12=28/310*259
O22=12;E22=28/310*51
TS = sum((O11-E11)^2/E11,(O12-E12)^2/E12,(O21-E21)^2/E21,(O22-E22)^2/E22)
pchisq(TS,1,lower.tail = FALSE)

```

```
## [1] 7.767783e-05
```

Or we could use following codes:

```
dat = matrix(c(243,16,39,12),2)
chisq.test(dat, correct = FALSE)
```

```
## Warning in chisq.test(dat, correct = FALSE): Chi-squared approximation may
## be incorrect
```

```
##
## Pearson's Chi-squared test
##
## data:  dat
## X-squared = 15.614, df = 1, p-value = 7.768e-05
```

## Problem 8

95% CI of the odds ratio is [3.649598 10.925327]

```
n11=50;n12=61
n21=27;n22=208
OR = n11*n22/(n12*n21)
SE = sqrt(1/n11+1/n12+1/n21+1/n22)
exp(c(log(OR)-1.96*SE,log(OR)+1.96*SE))
```

```
## [1]  3.649598 10.925327
```

## Problem 9

$$H_0 : p_1 = p_2$$

$$H_A : p_1 \neq p_2$$

p-value = 2.777e-08 < 0.05, so we reject the null meaning we fail to prove the treatment is effective

```
dat = matrix(c(50,10,20,40),2)
chisq.test(dat, correct = FALSE)
```

```
##
## Pearson's Chi-squared test
##
## data:  dat
## X-squared = 30.857, df = 1, p-value = 2.777e-08
```

## Problem 10

95% CI =

$$\left[ \log\left(\frac{\hat{p}}{1-\hat{p}}\right) - 1.96 \frac{1}{\hat{p}(1-\hat{p})} \sqrt{\hat{p}(1-\hat{p})/n}, \log\left(\frac{\hat{p}}{1-\hat{p}}\right) + 1.96 \frac{1}{\hat{p}(1-\hat{p})} \sqrt{\hat{p}(1-\hat{p})/n} \right]$$

## Problem 11

95% CI =

$$\left[ \log\left(\frac{\hat{p}_1}{1 - \hat{p}_1}\right) - \log\left(\frac{\hat{p}_2}{1 - \hat{p}_2}\right) - \frac{1.96}{\left(\frac{1}{\hat{p}_1(1-\hat{p}_1)} - \frac{1}{\hat{p}_2(1-\hat{p}_2)}\right)} \left(\sqrt{\hat{p}_1(1 - \hat{p}_1)/n} + \sqrt{\hat{p}_2(1 - \hat{p}_2)/n}\right), \right. \\ \left. \log\left(\frac{\hat{p}_1}{1 - \hat{p}_1}\right) - \log\left(\frac{\hat{p}_2}{1 - \hat{p}_2}\right) + \frac{1.96}{\left(\frac{1}{\hat{p}_1(1-\hat{p}_1)} - \frac{1}{\hat{p}_2(1-\hat{p}_2)}\right)} \left(\sqrt{\hat{p}_1(1 - \hat{p}_1)/n} + \sqrt{\hat{p}_2(1 - \hat{p}_2)/n}\right) \right]$$

## Problem 12

a.  $H_0 : p_1 = p_2$  $H_A : p_1 > p_2$ 

Using score test:

```
p1=30/40
p2=10/40
p=40/80
TS = (p1-p2)/sqrt(p*(1-p)*(2/40))
```

TS = 4.472136 > qnorm(0.95) = 1.644854. So we reject the null, meaning Drug A has a higher probability of side effect than Drug B.

b. For odds ratio:

95% CI = [3.270965 24.763337]

```
n11 = 30; n12 = 10; n21 = 10; n22 = 30
OR = n11*n22/(n12*n21)
SE = sqrt(1/n11+1/n12+1/n21+1/n22)
INTERVAL = exp(c(log(OR)-1.96*SE, log(OR)+1.96*SE))
INTERVAL
```

## [1] 3.270965 24.763337

For relative risk:

95% CI = [1.703711 5.282585]

```
p1=30/40
p2=10/40
RR = p1/p2
SE = sqrt((1-p1)/(p1*40)+(1-p2)/(p2*40))
INTERVAL = exp(c(log(RR)-1.96*SE, log(RR)+1.96*SE))
INTERVAL
```

## [1] 1.703711 5.282585

For risk difference:

95% CI = [0.3102238 0.6897762]



```
p1=30/40
p2=10/40
RD=p1-p2
SE = sqrt(p1*(1-p1)/40+p2*(1-p2)/40)
INTERVAL = c(RD-1.96*SE,RD+1.96*SE)
INTERVAL
```

```
## [1] 0.3102238 0.6897762
```

## Problem 13

Using score test:

$$H_0 : p = 0.5$$

$$H_A : p \neq 0.5$$

```
pbinom(7,10,0.5,lower.tail = FALSE)
```

```
## [1] 0.0546875
```

p-value = 0.0546875 > 0.05 , so we fail to reject the hypothesis that the coin is fair.