

Estimation - Simple Linear Regression

 A simple linear regression is the special case of an OLS regression model with a single predictor variable.

$$Y = \beta_0 + \beta_1 X + \epsilon \tag{1}$$

• For the ith observation we will denote the regression model by

$$Y_i = \beta_0 + \beta_1 X_i + \epsilon_i. \tag{2}$$

• For the random sample Y_1, Y_2, \ldots, Y_n we can estimate the parameters β_0 and β_1 by minimizing the sum of the squared errors,

$$\operatorname{Min} \sum_{i=1}^{n} \epsilon_i^2 \tag{3}$$

which is equivalent to minimizing

$$\min \sum_{i=1}^{n} (Y_i - \beta_0 - \beta_1 X_i)^2.$$
 (4)

Estimators and Estimates for Simple Linear Regression

ullet The estimators for eta_0 and eta_1 can be computed analytically and are given by

$$\widehat{\beta}_1 = \frac{\sum (Y_i - \bar{Y})(X_i - \bar{X})}{(X_i - \bar{X})^2} = \frac{\operatorname{Cov}(Y, X)}{\operatorname{Var}(X)}$$
(5)

and

$$\hat{\beta}_0 = \bar{Y} - \hat{\beta}_1 \bar{X} \tag{6}$$

- The regression line always goes through the *centroid* (\bar{X}, \bar{Y}) .
- We refer to the formulas for $\hat{\beta}_0$ and $\hat{\beta}_1$ as the *estimators* and the values that these formulas can take for a given random sample as the *estimates*. In statistics we put *hats* on all estimators and estimates.
- Given $\hat{\beta}_0$ and $\hat{\beta}_1$ the *predicted value* or *fitted value* is given by

$$\widehat{Y} = \widehat{\beta}_0 + \widehat{\beta}_1 X. \tag{7}$$

Estimation - The General Case

- We seldom build regression models with a single predictor variable. Typically we have multiple predictor variables denoted by X_1, X_2, \ldots, X_k , and hence the standard regression case is sometimes referred to as *multiple regression* in introductory regression texts.
- We can still think about the estimation of β_0 , β_1 , β_2 , ..., β_k in the same manner as the simple linear regression case

$$\operatorname{Min} \sum_{i=1}^{n} (Y_i - \beta_0 - \beta_1 X_{1i} - \beta_2 X_{2i} - \dots - \beta_k X_{ki})^2, \tag{8}$$

but the computations will be performed as matrix computations.

General Estimation - Matrix Notation

Before we set up the matrix formulation for the OLS model, let's begin by defining some matrix notation.

- The error vector $\epsilon = [\epsilon_1 \cdots \epsilon_n]^T$.
- The response vector $\mathbf{Y} = [Y_1 \cdots Y_n]^T$.
- The design matrix or predictor matrix $X = [1 X_1 X_2 \cdots X_k]$.
- The parameter vector $\beta = [\beta_0 \ \beta_1 \ \beta_2 \cdots \beta_k]^T$.
- All vectors are column vectors, and the superscript T denotes the vector or matrix transpose.

General Estimation - Matrix Computations

ullet We minimize the sum of the squared error by minimizing $S(m{eta}) = m{\epsilon}^T m{\epsilon}$ which can be re-expressed as

$$S(\beta) = (\mathbf{Y} - \mathbf{X}\beta)^{T} (\mathbf{Y} - \mathbf{X}\beta). \tag{9}$$

• Taking the matrix derivative of $S(\beta)$, we get

$$S_{\beta}(\hat{\beta}) = -2\mathbf{X}^T \mathbf{Y} + 2\mathbf{X}^T \mathbf{X} \hat{\beta}. \tag{10}$$

 Setting the matrix derivative to zero, we can write the expression for the least squares normal equations

$$\mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} = \mathbf{X}^T \mathbf{Y},\tag{11}$$

which yield the estimator

$$\hat{\beta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{Y}. \tag{12}$$

• The estimator form $\hat{\beta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{Y}$ assumes that the inverse matrix $(\mathbf{X}^T \mathbf{X})^{-1}$ exists and can be computed. In practice your statistical software will directly solve the normal equations using a QR factorization.

Statistical Inference with the t-Test

- In OLS regression the statistical inference for the individual regression coefficients can be performed by using a t-test.
- When performing a t-test there are three primary components: (1) Stating the null and alternate hypotheses, (2) Computing the value of the test statistic, and (3) Deriving a statistical conclusion based on a desired significance level.
- ullet Step 1: The *null* and *alternate* hypotheses for eta_i are given by

$$H_0: \beta_i = 0 \text{ versus } H_1: \beta_i \neq 0.$$
 (13)

Statistical Inference with the t-Test - Continued

• Step 2: The t statistic for β_i is computed by

$$t_i = \frac{\hat{\beta}_i}{\mathsf{SE}(\hat{\beta}_i)} \tag{14}$$

and has a degrees of freedom equal to the sample size minus the number of model parameters, i.e. df = n - dim(Model). For example if you had a regression model with two predictor variables and an intercept estimated on a sample of size 50, then the t statistic would have 47 degrees of freedom.

• Step 3: Reject H_0 or Fail to Reject H_0 based on the value of your t statistic and your significance level. This decision can be made by using the p-value of your t statistic or by using the critical value for your significance level.

Confidence Intervals for Parameter Estimates

An alternative to performing a formal hypothesis test is to use a confidence interval for your parameter estimate. There is a duality between confidence intervals and formal hypothesis testing for regression parameters.

ullet The confidence interval for \widehat{eta}_i is given by

$$\hat{\beta}_i \pm t(df, \alpha/2) * SE(\hat{\beta}_i),$$
 (15)

where $t(df, \alpha/2)$ is a t-value from a theoretical t-distribution, not a t-statistic value.

• If the confidence interval does not contain zero, then this is equivalent to rejecting the null hypothesis $H_0: \beta_i = 0$.

Statistical Intervals for Predicted Values

The phrase *predicted value* is used in statistics to refer to the in-sample *fitted values* from the estimated model or to refer to the out-of-sample *forecasted values*. The dual use of this phrase can be confusing. A better habit is use the phrases *in-sample fitted values* and the *out-of-sample predicted values* to clearly reference these different values.

- Given $\hat{\beta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{Y}$ the vector of fitted values can be computed by $\hat{\mathbf{Y}} = \mathbf{X}\hat{\beta} = \mathbf{H}\mathbf{Y}$, where $\mathbf{H} = \mathbf{X}(\mathbf{X}^T\mathbf{X})^{-1}\mathbf{X}^T$. The matrix \mathbf{H} is called the hat matrix since it puts the hat on \mathbf{Y} .
- The point estimate \hat{Y}_0 at the point \mathbf{x}_0 can be computed by $\hat{Y}_0 = \mathbf{x}_0^T \hat{\boldsymbol{\beta}}$.

Statistical Intervals for Predicted Values - Continued

ullet The confidence interval for an in-sample point \mathbf{x}_0 on the estimated regression function is given by

$$\mathbf{x}_0^T \widehat{\boldsymbol{\beta}} \pm \widehat{\boldsymbol{\sigma}} \sqrt{\mathbf{x}_0^T (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{x}_0}.$$
 (16)

• The prediction interval for the point estimate \widehat{Y}_0 for an out-of-sample \mathbf{x}_0 is given by

$$\mathbf{x}_0^T \hat{\boldsymbol{\beta}} \pm \hat{\boldsymbol{\sigma}} \sqrt{1 + \mathbf{x}_0^T (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{x}_0}.$$
 (17)

• Note that the out-of-sample prediction interval is always wider than the in-sample confidence interval.

Further Notation and Details

In order to compute the t statistic you need the standard error for the parameter estimate. Most statistical software packages should provide this estimate and compute this t statistic for you. However, it is always a good idea to know from where this number comes. Here are the details needed to compute the standard error for $\hat{\beta}_i$.

ullet The estimated parameter vector $\widehat{oldsymbol{eta}}$ has the covariance matrix given by

$$Cov(\hat{\beta}) = \hat{\sigma}^2 \mathbf{X}^T \mathbf{X},\tag{18}$$

where

$$\hat{\sigma}^2 = \frac{\mathsf{SSE}}{n - k - 1}.\tag{19}$$

ullet The variance of \widehat{eta}_i is the ith diagonal element of the covariance matrix

$$Var(\hat{\beta}_i) = \hat{\sigma}^2(\mathbf{X}^T \mathbf{X})_{ii}.$$
 (20)