# IMA205 TP3 Theoretical Questions

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March 18, 2025

## OLS

#### Questions

1. Demonstrate that OLS is the estimator with the smallest variance: compute  $\mathbf{E}[\tilde{\beta}]$  and  $\mathrm{Var}(\tilde{\beta}) = \mathbf{E}[(\tilde{\beta} - \mathbf{E}[\tilde{\beta}])(\tilde{\beta} - \mathbf{E}[\tilde{\beta}])^{\top}]$  and show when and why  $\mathrm{Var}(\beta^*) < \mathrm{Var}(\tilde{\beta})$ . Which assumption of OLS do we need to use?

#### Answers

Let the linear model be

$$y = X\beta + \varepsilon$$
,  $E[\varepsilon] = 0$ ,  $Var(\varepsilon) = \sigma^2 I$ .

A linear estimator is

$$\tilde{\beta} = C y$$
.

For unbiasedness we require

$$E[\tilde{\beta}] = C X \beta = \beta \implies C X = I.$$

Its variance is

$$\operatorname{Var}(\tilde{\beta}) = \sigma^2 C C^T$$
.

The OLS estimator is

$$\beta^* = (X^T X)^{-1} X^T y,$$

with variance

$$Var(\beta^*) = \sigma^2 (X^T X)^{-1}.$$

Any other unbiased estimator can be written as

$$C = (X^T X)^{-1} X^T + D$$
, with  $DX = 0$ .

Then one may show that

$$\operatorname{Var}(\tilde{\beta}) - \operatorname{Var}(\beta^*) = \sigma^2 D D^T \succeq 0.$$

Thus, the OLS estimator has the smallest variance among all linear unbiased estimators.

# Ridge regression

### Questions

- 2. Show that the estimator of ridge regression is biased (that is  $\mathbf{E}[\beta_{\text{ridge}}^*] \neq \beta$ ).
- 3. Recall that the SVD decomposition is  $\mathbf{x_c} = UDV^T$ . Write down by hand the solution  $\beta_{\text{ridge}}^*$  using the SVD decomposition. When is it useful using this decomposition? Hint: do you need to invert a matrix?
- 4. Remember that  $Var(\beta_{OLS}^*) = \sigma^2(\mathbf{x}^T\mathbf{x})^{-1}$ . Show that  $Var(\beta_{OLS}^*) \ge Var(\beta_{ridge}^*)$ .
- 5. When  $\lambda$  increases what happens to the bias and to the variance? Hint: Compute MSE =  $\mathbf{E}[(y_0 x_0^T \beta_{ridge}^*)^2]$  at the test point  $(x_0, y_0)$  with  $y_0 = x_0^T \beta + \epsilon_0$  being the true model and  $x_0^T \beta_{ridge}^*$  the ridge estimate.
- 6. Show that  $\beta_{\text{ridge}}^* = \frac{\beta_{\text{oLS}}^*}{1+\lambda}$  when  $\mathbf{x_c}^T \mathbf{x_c} = I_d$ .

### **Answers:**

• 2)

The ridge estimator is

$$\beta_{\text{ridge}}^* = (X_c^T X_c + \lambda I)^{-1} X_c^T y_c.$$

Taking expectations,

$$E[\beta_{\text{ridge}}^*] = (X_c^T X_c + \lambda I)^{-1} X_c^T X_c \beta.$$

Since  $(X_c^T X_c + \lambda I)^{-1} X_c^T X_c \neq I$  for  $\lambda > 0$ , the estimator is biased.

• 3)

Let  $X_c = UDV^T$  so that

$$X_c^T X_c = V D^T D V^T.$$

Then,

$$\beta_{\text{ridge}}^* = V(D^T D + \lambda I)^{-1} D^T U^T y_c.$$

This form is useful when  $X_c^T X_c$  is singular or ill–conditioned, as the addition of  $\lambda I$  guarantees invertibility.

•  $\frac{4)}{\text{For OLS}}$ 

$$Var(\beta_{\text{OLS}}^*) = \sigma^2 (X_c^T X_c)^{-1}.$$

For ridge regression, one finds

$$\operatorname{Var}(\beta_{\operatorname{ridge}}^*) = \sigma^2 (X_c^T X_c + \lambda I)^{-1} X_c^T X_c (X_c^T X_c + \lambda I)^{-1},$$

which is smaller (in the positive semidefinite sense).

the overall mean squared error at an appropriate choice of  $\lambda$ .

• 5)
As  $\lambda$  increases, the bias increases (since  $E[\beta_{\text{ridge}}^*]$  shrinks toward zero), while the variance decreases, lowering

•  $\underline{\mathbf{6}}$ )
If  $X_c^T X_c = I$  then

$$\beta_{\text{ridge}}^* = (I + \lambda I)^{-1} X_c^T y_c = \frac{1}{1+\lambda} \beta_{\text{OLS}}^*.$$

# Elastic Net

#### Questions

7. Compute by hand the solution of Eq.2 supposing that  $\mathbf{x_c}^T \mathbf{x_c} = I_d$  and show that the solution is:

$$(\beta_{\text{ElNet}}^*)_j = \frac{(\beta_{\text{OLS}}^*)_j \pm \frac{\lambda_1}{2}}{1 + \lambda_2}$$

# Answers

The Elastic Net estimator is defined as

$$\beta_{\text{ElNet}}^* = \arg\min_{\beta} \left\{ (y_c - X_c \beta)^T (y_c - X_c \beta) + \lambda_2 \|\beta\|_2^2 + \lambda_1 \|\beta\|_1 \right\}.$$

Define

$$\lambda = \lambda_1 + \lambda_2$$
 and  $\alpha = \frac{\lambda_2}{\lambda_1 + \lambda_2}$ .

Then the problem can be written as

$$\min_{\beta} \left\{ (y_c - X_c \beta)^T (y_c - X_c \beta) + \lambda \left[ \alpha \|\beta\|_2^2 + (1 - \alpha) \|\beta\|_1 \right] \right\}.$$

Assume  $X_c^T X_c = I$  so that the OLS solution is  $\beta_{\text{OLS}}^* = X_c^T y_c$ . Then, the minimization decouples for each coordinate j:

$$\min_{\beta_j} \left\{ (\beta_{\text{OLS},j} - \beta_j)^2 + \lambda_2 \, \beta_j^2 + \lambda_1 \, |\beta_j| \right\}.$$

Collecting the quadratic terms gives

$$(1 + \lambda_2)\beta_j^2 - 2\beta_{\text{OLS},j}\beta_j + \lambda_1 |\beta_j|.$$

The minimizer is obtained via soft-thresholding:

$$\beta_{\text{ElNet},j} = \text{sgn}(\beta_{\text{OLS},j}) \max \left\{ |\beta_{\text{OLS},j}| - \frac{\lambda_1}{2(1+\lambda_2)}, 0 \right\}.$$

This is the closed–form solution in the special case of orthonormal predictors.