

Superstar Teams

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Motivation: production involves heterogeneous workers

- Most production processes are too complex for any one individual to know how to do all the required tasks well
→ **firms coordinate** groups of **workers** with **specialized skills**

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 - “**superstar firms**” employ the most talented workers
- **Argument:** how production with heterogeneous workers is organized – within and across firms – shapes the firm-level distribution of wages & agg. productivity

Firms as “team assemblies”: theory, measurement, application

- **Firm:** organized collection of ‘complementary’ workers performing tasks (“team”)
- **This paper:**
 - ① develops a tractable **theory** of assignment of workers to tasks & workers into teams
 - ② shows how it can be disciplined using **micro data**
 - ③ **quantifies** implications for labor market inequality & agg. productivity
- **Application:** explain why \uparrow wage inequality is largely a *between*-firm phenomenon

Theory

- **Qs:** how much can one firm produce with a given team of heterogeneous workers?
how do workers sort into different firms?
- Develop a parsimonious **task-based model of team production**
 - many tasks required for production; each worker has limited time and their productivity varies across tasks; firm optimally allocates time to tasks
 - → **analytical microfoundation for coworker complementarities:** strength of complementarities is endogenously increasing in degree of skill specialization
- Integrate team-production into a **dynamic, frictional eqm. matching** environment
 - → **coworker talent complementarities foster firm-level concentration of talent**

Measurement

- **Q:** how can we empirically discipline this model?
- Develop a **theory-guided measurement strategy** designed for standard micro data
 - → **identification:** production complementarities recovered from wage + match data
- Model calibrated to DE panel data endogenously generates **empirically realistic variation in firm-level average wages** – *without* having assumed ex-ante differences

Application: technological change & the “firming up” of inequality

- *Changes* in the organization of production have **macro implications**
- **Data:** wage inequality is increasingly a *between-firm* phenomenon
[Card et al., 2013; Bloom et al., 2019; Criscuolo et al., 2022, ...]
- Model offers **quantitative, structural explanation**
 - ① set of tasks any one worker can perform very well has narrowed: specialization \uparrow
 - ② coworker talent complementarities \sim doubled since 1985
 - ③ individuals of similar talent increasingly work together
 - ④ this explains $\approx 40\%$ of \uparrow between-firm wage inequality share
- Paper: implications for agg. productivity, job ladders, ...

Relation & contributions to literature

- **Firm organization:** **parsimonious task model suitable for quantitative applications**

Firms: Lucas, 1978; Rosen, 1982; Becker & Murphy, 1992; Hopenhayn & Rogerson, 1993; **Kremer, 1993**; Kremer & Maskin, 1996; **Garicano, 2000**; Klette & Kortum, 2004; **Garicano & Rossi-Hansberg, 2006**; Kohlhepp, 2022; Kuhn et al., 2022; Minni, 2022; Bassi et al., 2023

Task assignment: Costinot & Vogel, 2010; **Acemoglu & Restrepo, 2018**; Ocampo, 2021; Adenbaum, 2022

Teams: Akcigit et al., 2018; Chade & Eeckhout, 2020; **Jarosch et al., 2021**; Herkenhoff et al., 2022

- **Frictional labor market sorting:** **endogenize & measure complementarities**

Shimer & Smith, 2000; Cahuc et al., 2006; Eeckhout & Kircher, 2011/2018; Hagedorn et al., 2017; de Melo, 2018; **Herkenhoff et al., 2022**; Lindenlaub & Postel-Vinay, 2023

- **Wage inequality:** **structural model of \uparrow firm-level inequality due to technological Δ**

Technology: Katz & Murphy, 1992; Krusell et al., 2000; Autor, Levy & Murnane, 2003; **Jones, 2009**; Deming, 2017; Acemoglu & Restrepo, 2018; Alon, 2018; Neffke, 2019; Jones, 2021; Atalay et al., 2021

Firms: **Card et al., 2013**; Barth et al., 2016; Alvarez et al., 2018; **Bloom et al., 2019**; Aeppli & Wilmers, 2021; Criscuolo et al. 2021; Hakanson et al., 2021; Sorkin & Wallskog, 2021; Kleinman, 2022

Theory



Overview of environment

- Many **workers** and many **firms**
- **Ex-ante homogeneous firms** assemble – hire workers & **assign tasks** – teams
- **Workers are heterogeneous in productivity**
→ workforce composition is (only) source of ex-post differences across firm
- **How is production organized?**
 - 1 **team production**: how much does a team produce if tasks are optimally allocated?
 - 2 **team formation**: who gets to work with whom?

Step 1: production in a single team of given composition

- Firm: **1 team of** $n \in \mathbb{Z}_{++}$ **workers** produces output from **unit continuum of tasks** \mathcal{T}

$$\ln Y = \int_{\mathcal{T}} \ln q(\tau) d\tau \quad (1)$$

- Task aggregation:**

$$q(\tau) = \sum_{i=1}^n y_i(\tau) \quad (2)$$

- Workers** \rightarrow **tasks:** i supplies 1 time unit, task-specific productivities $\mathbf{z}_i = \{\mathbf{z}_i(\tau)\}_{\tau \in \mathcal{T}}$

$$y_i(\tau) = \mathbf{z}_i(\tau) l_i(\tau) \quad (3)$$

$$1 = \int_{\mathcal{T}} l_i(\tau) d\tau \quad (4)$$

Firm's optimization problem

- **Firm solves mini-planner problem:** $\max_{\mathbf{q}, \{\mathbf{l}_i\}} Y$ s.t. (1)-(4)
- **Preview:** derive & characterize *reduced-form* team production function f

$$f(\mathbf{z}_1, \dots, \mathbf{z}_n) = \max Y$$

s.t. (1)-(4)

Firm's optimization problem

- **Firm solves mini-planner problem:** $\max Y$ s.t. (1)-(4)

$$\begin{aligned}
 \mathcal{L}(\cdot) = & Y + \lambda \left[\underbrace{\left(\int_{\mathcal{T}} \ln q(\tau) d\tau \right)}_{\text{tasks} \rightarrow \text{output}} - \ln Y \right] + \int_{\mathcal{T}} \lambda(\tau) \underbrace{\left(\sum_{i=1}^n y_i(\tau) - q(\tau) \right)}_{\text{task aggregation}} d\tau \\
 & + \sum_{i=1}^n \lambda_i^L \underbrace{\left(\int_{\mathcal{T}} \frac{y_i(\tau)}{z_i(\tau)} d\tau - 1 \right)}_{\text{time constraint + task production}} + \text{non-negativity constraints}
 \end{aligned}$$

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- FOCs** imply

$$\lambda(\tau) = \min_i \left\{ \frac{\lambda_i^L}{z_i(\tau)} \right\}$$

shadow cost of τ \leftarrow

λ_i^L \rightarrow opportunity cost of i 's time

$z_i(\tau)$ \rightarrow productivity of i for τ

Tractability: leverage insight from trade literature

Assumption: Multivariate Fréchet distribution of worker-task productivities

$$\Pr [Z_1(\tau) \leq z_1, \dots, Z_n(\tau) \leq z_n] = \exp \left[- \left(\sum_{i=1}^n \left(\left(\frac{z_i}{x_i} \right)^{-\frac{1}{\chi}} \right)^{\frac{1}{\xi}} \right)^{\xi} \right]$$

- **Payoff #1:** low-dim. representation of worker-task productivity dist.
 - x_i : scale term \sim “**talent**” type
 - χ : *within*-worker productivity dispersion \sim **specialization**
 - $\xi \in (0, 1]$: \sim correlation in z 's *across* workers

Illustration: talent types

- **“Talent” type** $x_i \sim$ absolute advantage

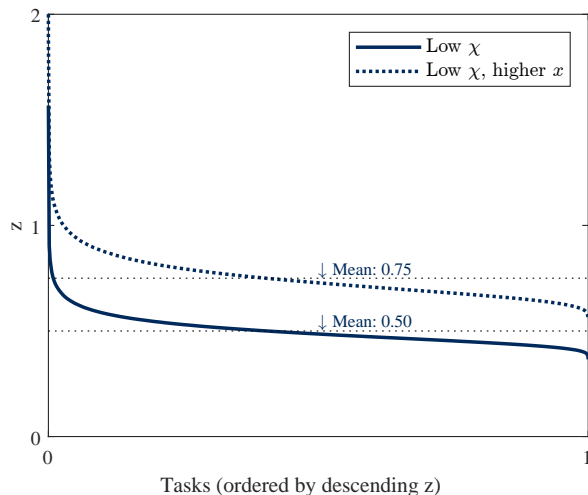
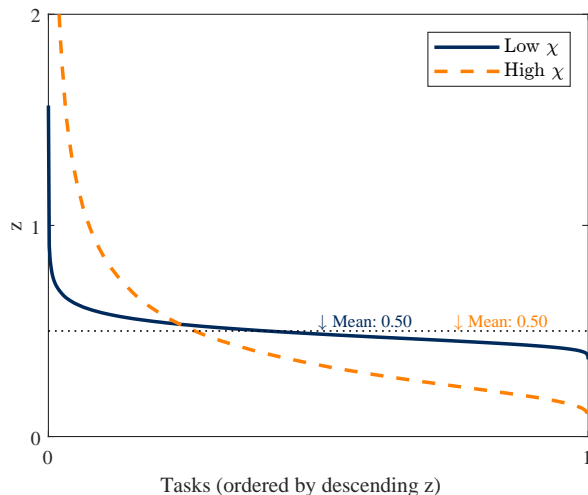


Illustration: specialization parameter

► Illustration: correlation

- Parameter χ : *within-worker* productivity dispersion across tasks \sim **specialization**



Tractability: leverage insight from trade literature

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$$\Pr [Z_1(\tau) \leq z_1, \dots, Z_n(\tau) \leq z_n] = \exp \left[- \left(\sum_{i=1}^n \left(\left(\frac{z_i}{\iota X_i} \right)^{-\frac{1}{\chi}} \right)^{\frac{1}{\xi}} \right)^{\xi} \right]$$

- Payoff #1: low-dim. representation of worker-task productivity dist.
- **Payoff #2:** max-stability property allows closed-form characterization of key objects

Micro-founded production function

[▶ Lemma](#)

Proposition: Reduced-form production function

Talent types (x_1, \dots, x_n) and ξ are sufficient statistic for team output Y given χ :

$$Y = f(x_1, \dots, x_n; \chi, \xi)$$

[▶ Proof sketch](#)

Micro-founded production function: characterization

► Taylor approx.

Proposition: Reduced-form production function

$$f(x_1, \dots, x_n; \chi, \xi) = \underbrace{n^{1+\chi\xi}}_{\text{efficiency gains}} \times \left(\frac{1}{n} \sum_{i=1}^n (x_i)^{\frac{1}{1+\chi\xi}} \right)^{1+\chi\xi}$$

vs. no-division-of-labor: $f(x_1, \dots, x_n) = n \times \left(\frac{1}{n} \sum_{i=1}^n x_i \right)$

- 1 Efficiency gains from teamwork are large when workers have skills that are *specialized* (χ) in *different* (ξ) tasks

Micro-founded production function: characterization

► Quality & task mismatch

► Example

Proposition: Reduced-form production function

$$f(x_1, \dots, x_n; \chi, \xi) = \underbrace{n^{1+\chi\xi}}_{\text{efficiency gains}} \times \underbrace{\left(\frac{1}{n} \sum_{i=1}^n (x_i)^{\frac{1}{1+\chi\xi}} \right)^{1+\chi\xi}}_{\text{complementarity}},$$

2 Coworker talent complementarities are important when workers have skills that are specialized (χ) in different (ξ) tasks

- elasticity of complementarity $\gamma := \frac{\partial \ln(f_j/f_i)}{\partial \ln(x_i/x_j)} = \frac{\chi\xi}{1+\chi\xi}$
- output more 'vulnerable' to the lowest-x member(s) \leftrightarrow talented workers gain disproportionately more from more talented colleagues

Intuition: features of optimal organization

► Graphic: task assignment

► Extension to communication frictions

- **What is the intuition for these properties?**
- Solution of firm's mini-planner problem implies:
 - ❶ **Complete division of labor**, with tasks assigned by **comparative advantage**
 - i 's task set $\mathcal{T}_i = \left\{ \tau \in \mathcal{T} : \frac{z_i(\tau)}{\lambda_i^L} \geq \max_{k \neq i} \frac{z_k(\tau)}{\lambda_k^L} \right\}$
 - **classic source of efficiency gains**
 - ❷ i 's share of tasks **↑ in i 's talent**, **↓ in coworkers' talent**
 - i 's task share $\pi_i = (x_i^{\frac{1}{1+\chi\xi}}) \left(\sum_{k=1}^n (x_k)^{\frac{1}{1+\chi\xi}} \right)^{-1}$

Intuition: comparative statics for task shares

- Suppose that $x_i > x_j$. Then
 - 1 i performs a strictly larger share of tasks than j for $\chi < \infty$



Intuition: comparative statics for task shares

- Suppose that $x_i > x_j$. Then
 - ① i performs a strictly larger share of tasks than j for $\chi < \infty$
 - ② the difference in task shares is decreasing in χ



⇒ **Greater skill specialization implies a larger share of tasks is performed by relatively less talented team members** – more talented coworkers can't easily compensate

Firm organization meets frictional matching

- **Recap:** optimal assignment of workers to tasks
→ team composition matters for productivity, especially when χ is high
- **Questions:**
 - 1 how is team composition determined in a multi-worker, multi-firm economy?
 - 2 how can we empirically pin down χ ?
- **To address Qs: integrate team production with (dynamic) frictional eqm. matching**
 - 1 mechanism: tradeoff determining assignment of workers to firms
 - 2 measurement device: identification of $f(\cdot)$ off of labor mkt. histories +wages

Frictional matching into teams: environment

- **Random-search + multi-worker firms** *[Herkenhoff-Lise-Menzio-Phillips, 2024]*
 - why search frictions: important for mechanism + measurement
- **Production technology:** $f(\cdot)$ with $n \in \{0, 1, 2\}$
- **Employment states:** unemp., employed alone, employed with $x' \in \mathcal{X}$
- **Nash wage bargaining** with continuous renegotiation
- ξ : **match-specific shock**
 - ξ can be observed and contracted on by firms & workers
 - ξ unobserved in standard micro data sets

Matching – stationary equilibrium

► Details

- HJ-Bellman equations → **values & matching policies**
- Flows between/**distribution** over types \times employment states

► HJBs

► KFEs

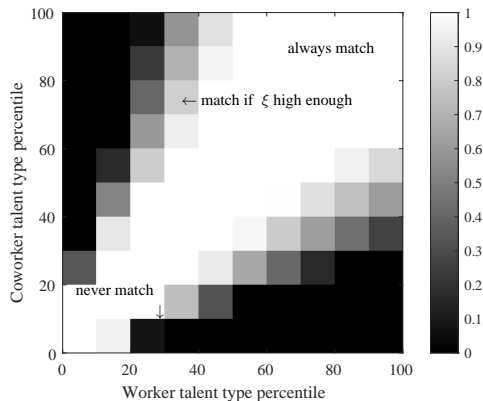
Definition: Stationary equilibrium

A stationary eqm. consists of a value functions & a distribution of agents, s.t.

- 1 the value functions satisfy the HJB equations given the distribution;
- 2 the distribution is stationary given the policy fn's implied by the value fn's.

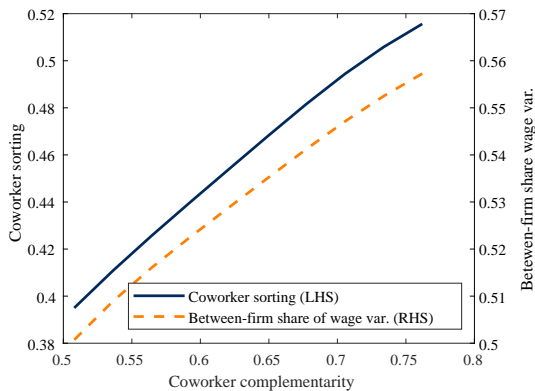
Mechanism: conditional matching probabilities for given χ

- What is the probability that a firm with a type- x worker would hire a type- x' worker conditional on meeting?



Mechanism: more positive assortative matching as $\chi \uparrow$

- What is the correlation between talent x & coworker talent x' (coworker sorting)?



Measurement: a useful identification result

- **Strategy:** recover χ from production, rather than inferring from sorting
 \Rightarrow How to quantify $\frac{\partial^2 f(\cdot)}{\partial x \partial x'}$?

► Discussion

Proposition: Measuring complementarities

Coworker complementarities (CC) in production are proportional to CC in wages:

$$\frac{\partial^2 f(x, x', \xi)}{\partial x \partial x'} \propto \frac{\partial^2 w(x|x', \xi)}{\partial x \partial x'}.$$

$$\Rightarrow \frac{\partial^2 \bar{f}(x, x')}{\partial x \partial x'} \approx \boxed{\frac{\partial^2 \bar{w}(x|x')}{\partial x \partial x'}} \quad \leftarrow \text{can measure this}$$

► Proof sketch

Model Meets Data

Taking the model to (micro) data

- ➊ Matched employer-employee panel data
- ➋ Mapping model objects to data
- ➌ Model calibration: quantitative analysis
- ➍ Validation (*brief today, more in paper*)

Taking the model to (micro) data: panel data for Germany

► SIEED

- **Primary data:** **SIEED matched-employer employee panel for W Germany** ► Processing
 - 1.5% sample of establishments + entire biographies of associated workers; social security information on employer, daily wage, occupation, demographics
 - initially focus on 2010-2017, later extend to 1985-2017
 - selection: full-time employees aged 20-60 and meeting min. earnings requirements, ≥ 10 observations per establishment-year
- Supplement with:
 - Portuguese matched employer-employee panel + balance sheet data
 - German repeated cross-sections of individual-level survey data on tasks

Mapping theory to data: worker & coworker types

► Implementation

- **Worker talent types:** measure using standard methods (wages monotonically \uparrow in x)
 - Baseline: 2-way fixed effect (FE) wage regressions [*Abowd et al., 1999* – AKM]
 - intuition: job switching identifies time-invariant “worker types,” controlling for possibility that some employers pay more to all workers
 - pre-est. k-means clustering \rightarrow address limited mobility bias [*Bonhomme et al., 2019*]
 - robustness: non-param. ranking algo instead of AKM [*Hagedorn et al., 2017*]
- \Rightarrow Worker “type” \hat{x}_i : decile rank of worker FE

Mapping theory to data: worker & coworker types

[► Implementation](#)

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- **“Representative coworker type”** \hat{X}_{-it} : average \hat{x}_i of coworkers in same establishment-year

[► Discussion](#)

Mapping theory to data: coworker complementarity

- Recall structural wage equation:

$$\bar{w}(x|x') = \omega \bar{f}(x, x') + g(x) - h(x')$$

$$\Rightarrow \frac{\partial^2 \bar{w}(x|x')}{\partial x \partial x'} \propto \frac{\partial^2 \bar{f}(x, x')}{\partial x \partial x'}$$

- Estimating polynomial equation:

$$\frac{w_{it}}{\bar{w}_t} = \beta_0 + \beta_1 \hat{x}_i + \beta_{11} \hat{x}_i^2 + \beta_2 \hat{x}_{-it} + \beta_{22} \hat{x}_{-it}^2 + \beta_c (\hat{x}_i \times \hat{x}_{-it})$$

$$+ \psi_{j(it)} + \nu_{o(i)t} + \xi_{s(i)t} + \epsilon_{it}$$

- alternative: non-parametrically approximate $\frac{\partial^2 w(x|x')}{\partial x \partial x'}$

Evidence on coworker complementarity (2010-2017)

► Robustness

► B-o-E calc. γ

► Peer effects

- **Estimating equation:**

coworker complementarity

$$\frac{w_{it}}{\bar{w}_t} = [\dots] + \beta_c (\hat{x}_i \times \hat{x}_{-it}) + \psi_{j(it)} + \nu_{o(i)t} + \xi_{s(i)t} + \epsilon_{it}$$

	$\hat{\beta}_c$	Non-parametric FD method
Coworker complementarity	0.0091*** (0.00035)	0.0097
Obs. (1000s)	4,410	4,410

Notes. Regressions include FEs for employer; occupation-year; industry-year. Employer-clustered standard errors in parentheses. Observations weighted by the inverse employment share of the respective type and (rounded) coworker type cell. FD: finite differences.

Robustness checks: measuring coworker complementarity

► Main

- Types from non-parametric ranking algorithm instead of AKM-based
- Schooling as a non-wage measure of types
- Lagged types
- Small teams
- Movers
- Non-parametric, finite-differences approximation
- Excluding managers
- Log specification

► Jump

► Jump

► Jump

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Quantifying the model: approach

► Parameter values

► Identification validation

► Within-industry

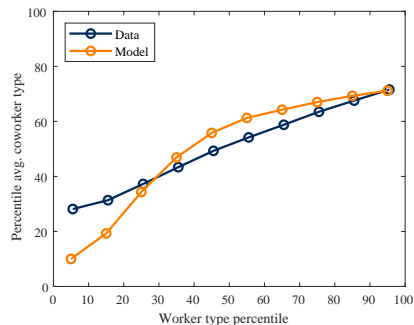
- **Calibrate** the model to the W German economy (2010-2017)
 - 1 externally calibrated: discount rate, team-benefit, bargaining power
 - 2 offline estimation: job separation hazard
 - 3 **online estimation** (indirect inference): meeting rate, unemp. flow benefit, production
 - targets: $\hat{\beta}_c$, total wage variance, avg. wage level, replacement rate, job finding rate
- **Production complementarity informed by $\hat{\beta}_c$**
 - set χ to align regression estimates for this reduced-form moment using model-generated data with empirical estimate
- **Macro moments of interest are untargeted**: sorting, between-firm wage inequality

► Monte-Carlo

Calibration results & model validation

► Between-share adjustment method

- ✓ **Match coworker sorting patterns**
 - $\rho_{xx} = 0.57$ (vs. 0.62 in data)
- ✓ **Match between-firm wage inequality**
 - between-share 0.56 (vs. 0.57 in data)
- Extensive **validation** exercises
 - non-wage implications of complementarities (e.g., direction of EE moves)
 - cross-sectional analysis: variation across occupations/industries

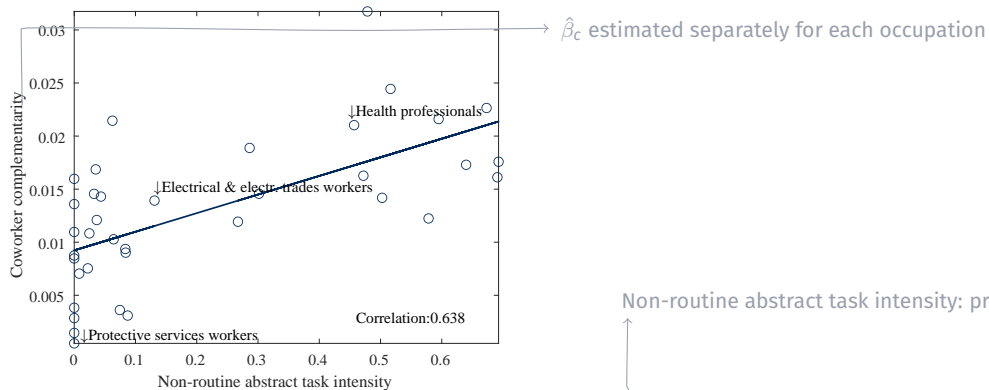


► Details

X-sectional validation (occ's): tasks \Rightarrow complementarity

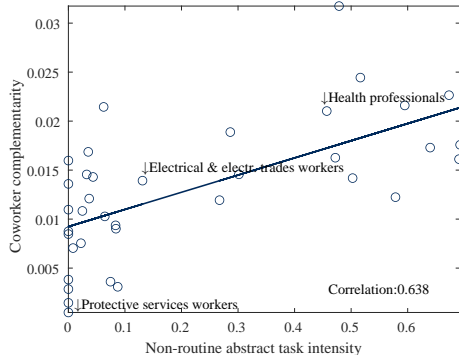
[▶ More validation](#)

- \uparrow **Non-routine abstract task intensity**
 $\Rightarrow \uparrow$ **coworker talent complementarity**

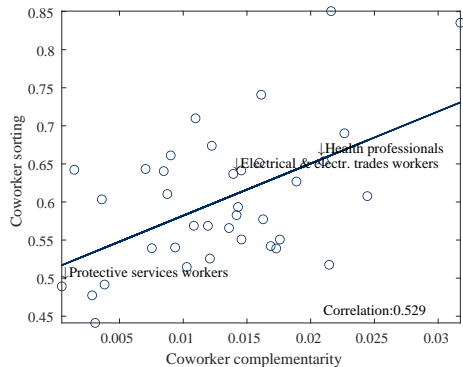


X-sectional validation (occ's): tasks \Rightarrow complementarity \Rightarrow sorting

- \uparrow Non-routine abstract task intensity
 $\Rightarrow \uparrow$ coworker talent complementarity



- \uparrow **Coworker talent complementarity**
 $\Rightarrow \uparrow$ **coworker sorting**



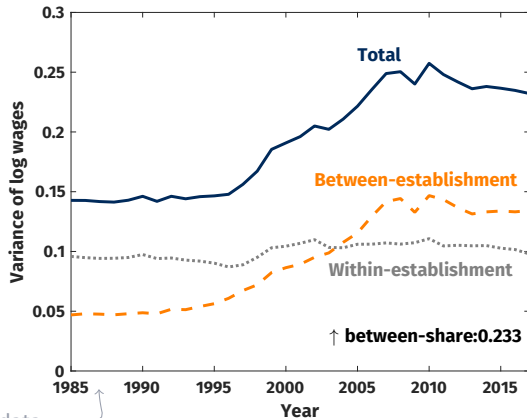
Firming Up Inequality

The rise in wage inequality is mostly a between-firm phenomenon

[Details](#)

“the variance of firm [wages] explains an increasing share of total inequality in a range of countries”

[Song-Price-Guvenen-Bloom-von Wachter, 2019]

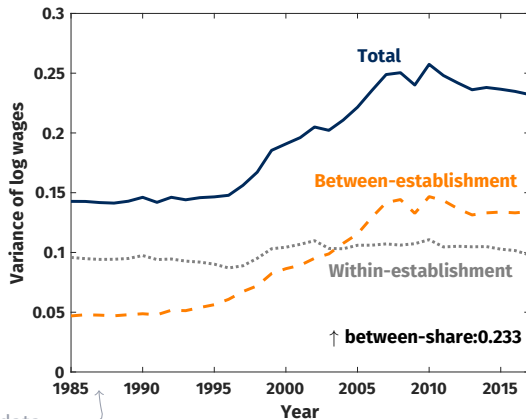


German matched employer-employee data

Wage inequality has risen – and firms appear to play a key role

[Details](#)

Applied question: what is/are the causal driver(s)? implications?



German matched employer-employee data

Hypothesis

- **Hypothesis:**

- ① the set of tasks any one worker can perform very well has narrowed: specialization ↑
- ② coworker talent complementarities ↑
- ③ individuals of similar talent increasingly work together
- ④ this generates greater between-firm wage dispersion

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- **Approach:**

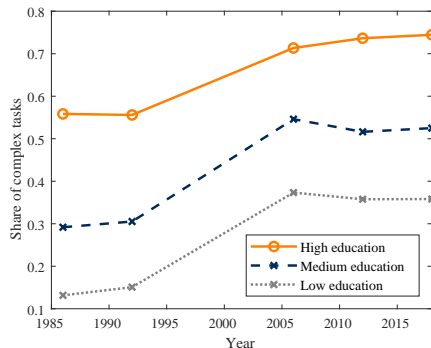
- evaluate descriptive evidence
- use structural model for counterfactual exercises quantifying this mechanism

Qualitative evidence suggests specialization (χ) \uparrow

► Occ. movements

► Science

- **Task complexity \uparrow :**
“extensive margin” of χ
 - DE longitudinal task survey [► BIBB](#)
 - “complex”: cognitive non-routine (e.g., organizing, researching)



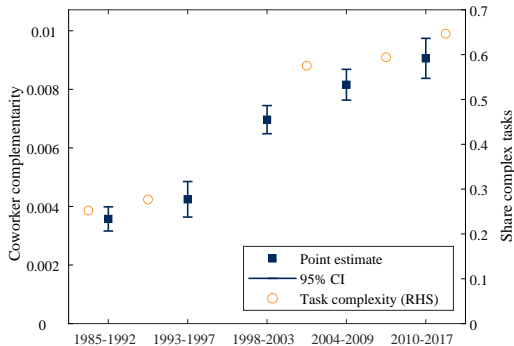
Coworker talent complementarity has strengthened

► Schooling

► Peer effect trends

- **Theory:** specialization (χ) \uparrow is associated with coworker complementarity \uparrow

✓ **Coworker complementarity has more than doubled between 1985-1992 and 2010-2017**

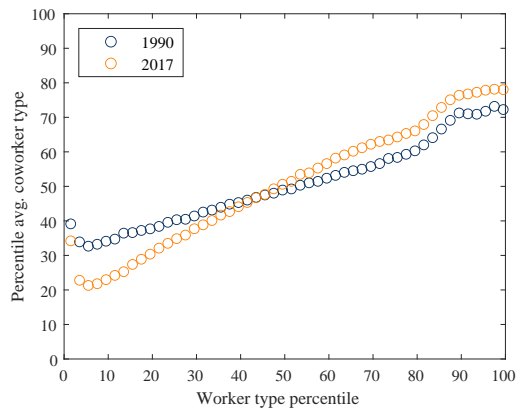


Talent sorting has intensified

[Details](#)

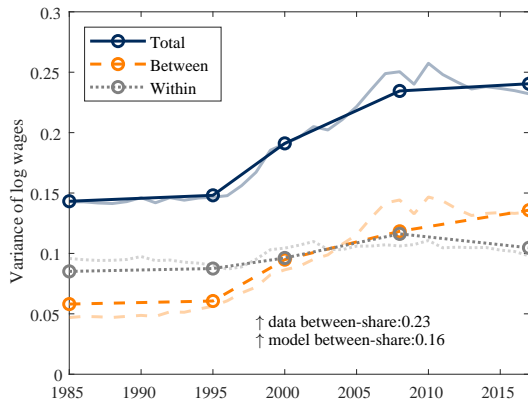
- **Theory:** complementarity \uparrow is associated with talent sorting \uparrow

✓ Coworker matching has become more positively assortative



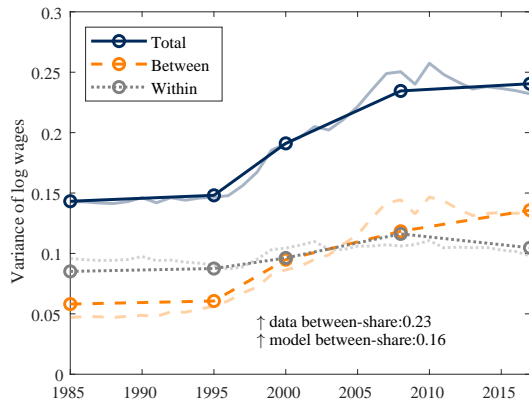
Model matches *changes* in firm-level wage distribution

- How quantitatively important are these developments in accounting for the firming up of inequality?
- **Re-calibrate model:** 4 earlier periods
- ✓ **Model replicates untargeted rise of between-share in data**
 - 68% of \uparrow between-share in data, ('85-'92) \rightarrow ('10-'17)



Model matches changes in firm-level wage distribution – why?

- ✓ Model replicates untargeted rise of between-share in data
- **Reflects several parameters changing**
 - elasticity of compl.: 0.43 in '85-'92 (vs. 0.84 in '10-'17)
 - job arrival & separation ↑
 - ...



Complementarity \uparrow explains $\approx 40\%$ of observed between-share \uparrow

- **Q:** How much of \uparrow between-firm share of wage var. is due to \uparrow complementarities?

Complementarity \uparrow explains $\approx 40\%$ of observed between-share \uparrow

- **Q:** How much of \uparrow between-firm share of wage var. is due to \uparrow complementarities?
- **Counterfactual:** between-firm share in 2010s absent $\chi \uparrow$ since '85-'92

Complementarity \uparrow explains $\approx 40\%$ of observed between-share \uparrow

- **Q:** How much of \uparrow between-firm share of wage var. is due to \uparrow complementarities?
- **Counterfactual:** between-firm share in 2010s absent $\chi \uparrow$ since '85-'92
- **A:** $\chi \uparrow$ **accounts for 59%** of model-predicted $\Delta \leftrightarrow \approx 40\%$ of empirical Δ

	Δ model	Implied % Δ model due to Δ parameter
Model baseline	0.16	-
Cf.: fix period-1 complementarity	0.065	59

Overview of model robustness checks

- Declining search frictions
- Within-industry calibration
- Outsourcing & within-occupation analysis
- OJS
- Increased talent dispersion

► Jump

► Jump

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Productivity & more

Overview of extensions & other implications

- **Aggregate productivity**
- Productivity dispersion
- “Coworker job ladders”
- Person-level inequality
- For fun: generative AI

▶ Jump

▶ Jump

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Conclusion

Conclusion: firms form & organize teams – matters for macro

- **Key idea:** if individuals have specialized skills, firms assemble teams of ‘complementary’ coworkers, generating systematic sorting patterns
- **This paper:**
 - 1 **task-based microfoundation for coworker complementarities**
⇒ specialization + team production → complementarities
 - 2 **measurement** combining reduced-form micro evidence with model structure
⇒ between-firm differences w/o assuming ex-ante heterogeneity
 - 3 structural explanation for the **“firming up” of inequality**
⇒ role of **increased complementarities**

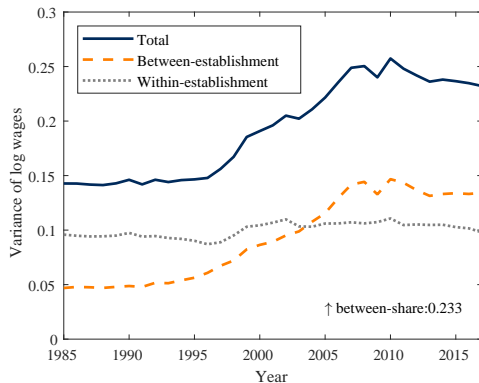
Thank You!

Extra Slides

Fact #1: ↑ between-firm share of wage inequality

[▶ Intro](#)

- Large empirical literature: “firming up inequality” [e.g., Card et al., 2013; Song et al., 2019]
 - “superstar firms” [e.g., Autor et al., 2020]
- **Fact 1: ↑ wage inequality primarily due to between-component**
- Robust pattern

[▶ Cross-country](#)
[▶ Panel est.](#)
[▶ Wage resid. alternatives](#)
[▶ Within-occ](#)
[▶ Within-ind](#)


Notes. Model-free statistical decomposition, where the “between” component corresponds to the person-weighted variance of est.-level avg. log wage.

Fact #2: talented workers increasingly collaborate

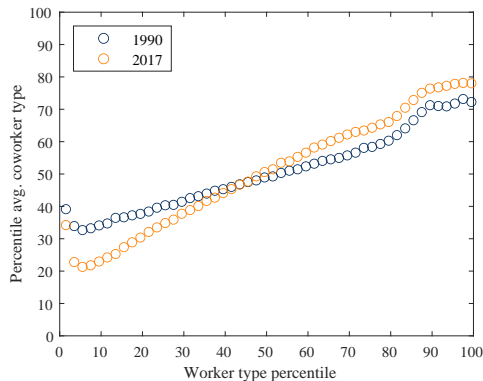
[Intro](#)[Main](#)[Var. decomp.](#)[Fact #3](#)

- To what extent do “talented workers” tend to have “talented coworkers”?

- Fact 2: + assortative coworker sorting** ↑

- $\rho_{xx} = \text{corr}(\hat{x}_i, \hat{x}_{-it})$: 0.43 ('85-'92) ↗ 0.62 ('10-'17)

- Robust pattern

[Table](#)[Within-occ. nonlinear](#)[Hakanson et al. \(2021\)](#)

Fact #3: increased education premium due to workplace effects

[Main](#)

- **Fact 3: increase in return to schooling is primarily due to workplace effects**

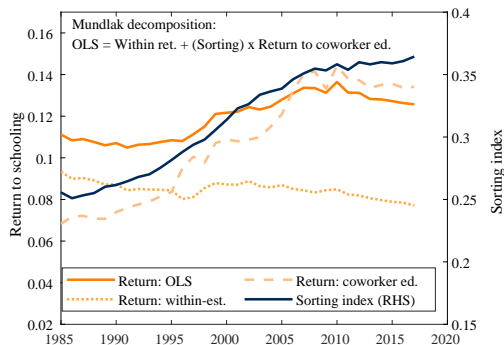
- Mundlak decomposition of year-specific OLS return to schooling:

$$\beta_t^{\text{ols}} = \beta_t^{\text{within}} + \rho_t \times \beta_t^{\text{estab.}}$$

$$\ln w_{it} = \beta_0 + \beta_t^{\text{within}} S_i + \beta_t^{\text{estab.}} \bar{S}_{j(i,t),t} + e_{it}$$

where $\bar{S}_{j(i,t),t}$ is avg. years of schooling in establishment j of worker i in year t

- 1 β_t^{within} : within-establishment return
- 2 $\beta_t^{\text{estab.}}$: return to avg. establishment schooling

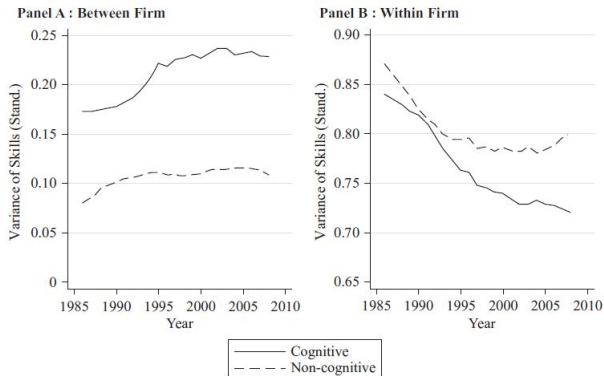


Notes. Plot of coefficients from year-by-year regressions of log wages.

Evidence from the literature: Hakanson et al. (2021)

[▶ Main](#)

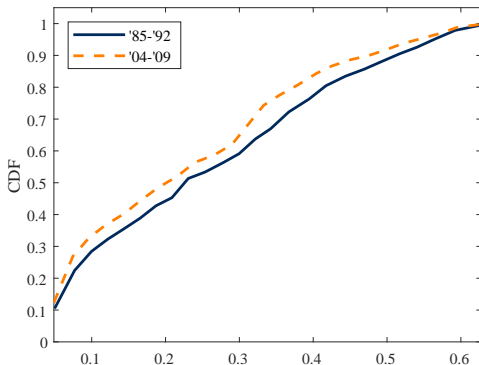
- *Direct* measures of cognitive and non-cognitive skills across Swedish firms during 1986–2008, using test data from military enlistment



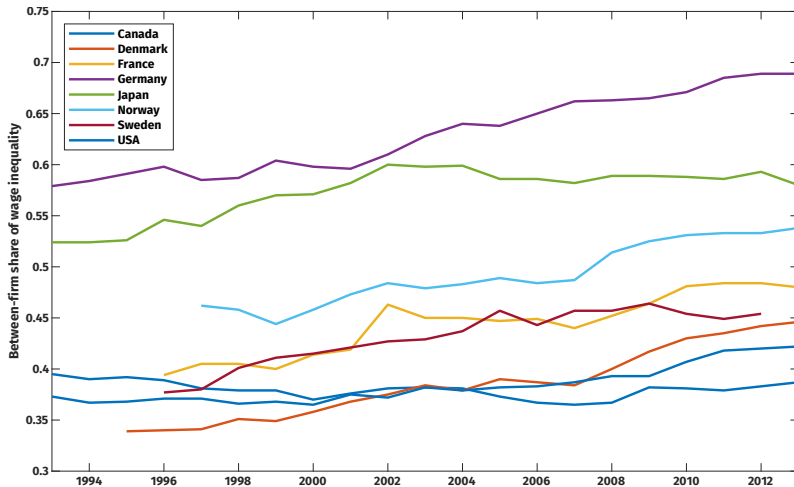
Workers increasingly tend to perform similar tasks across different jobs

[▶ Back](#)[▶ Comparison](#)

- ✓ Workers move to jobs with similar tasks, rather than randomly
- **Q:** are workers becoming *more* likely to perform similar tasks across jobs, over time?
- **Yes:** distribution of moves in ('04-'09) is stochastically dominated by that in ('85-'92)
 - uncond. average: 0.253 → 0.227: 10% decline
- Robust in regression design
 - quantile regressions: ✓ at different quantiles



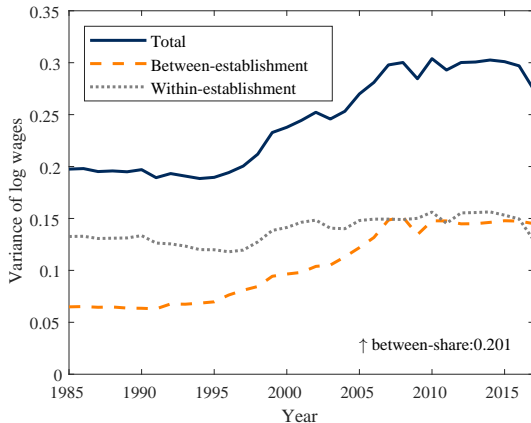
Firming up inequality: cross-country evidence

[▶ Main](#)

Notes. Data from Tomaskovic-Devey et al. (2020). Measures of earnings differ across countries and, for Germany, between T-D et al. and my study based on the SIEED.

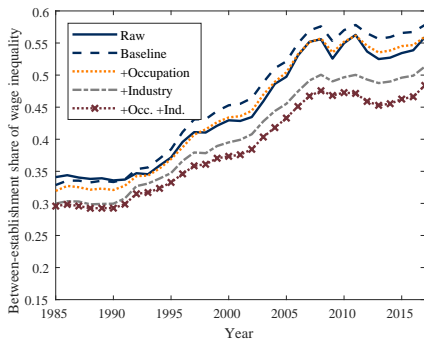
Between-/within-employer wage var decomp. - panel establishments

- Instead of considering *all* employers, restrict attention to “panel establishments”

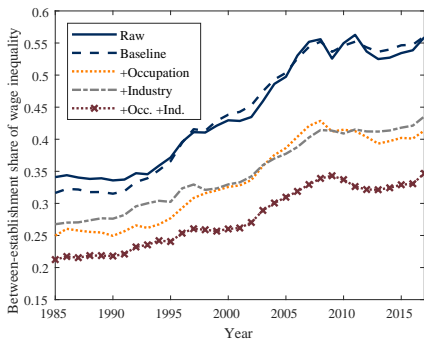


Between-/within-employer wage var. decomp. - alternative w-residuals

- “With worker FEs”: regress $\ln \tilde{w}_{it} = \alpha_i + X'_{it}\hat{\beta} + \epsilon_{it}$, construct $\ln w_{it} = \ln(\tilde{w}_{it} - X'_{it}\hat{\beta})$.
- “Without worker FEs”: regress $\ln \tilde{w}_{it} = \alpha_0 + X'_{it}\hat{\beta} + \epsilon_{it}$, and consider residuals ϵ_{it}

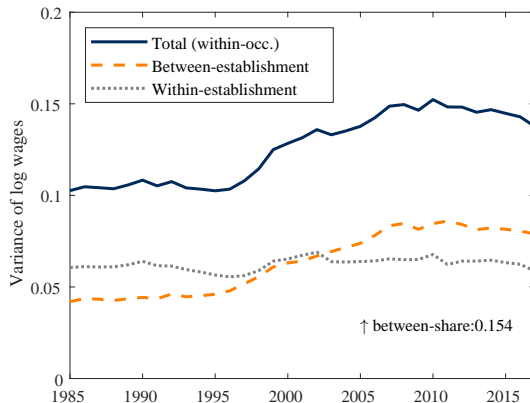


(a) With worker FEs

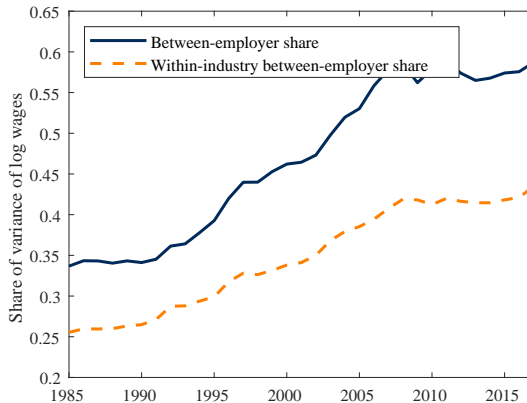


(b) Without worker FEs

Between-/within-employer wage var. decomp. - within-Occupation

[▶ Main](#)

Between-/within-employer wage var. decomp. - within-Industry

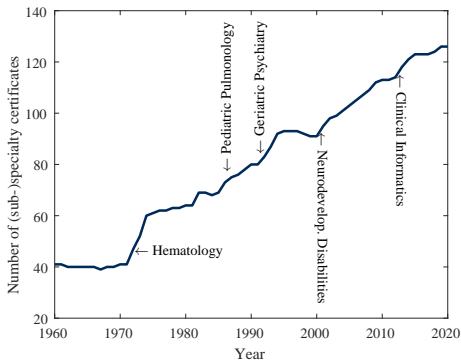
[▶ Main](#)[▶ Other moments](#)

Notes. Based on 'baseline' residualized wages.

Examples: rising specialization

[▶ Main](#)

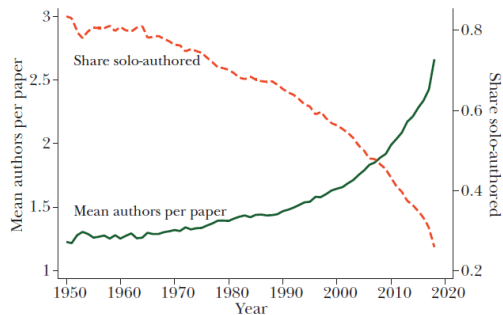
• Deepening medical specialization



Notes. Data from American Board of Medical Specialities. For each year, it shows the number of unique specialty or sub-specialty certificates that have been approved and issued at least once by that year and which are still being issued.

• Rise of research teams [Jones, 2021]

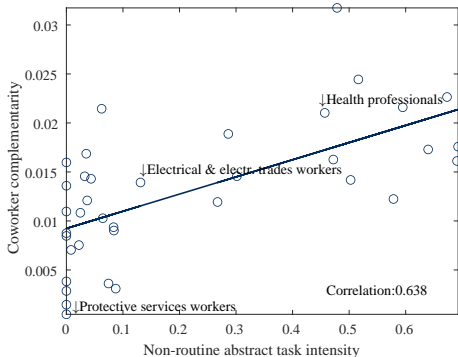
Panel A. All economics papers, 1950–2018



Occupations: task complexity \Rightarrow complementarity \Rightarrow sorting

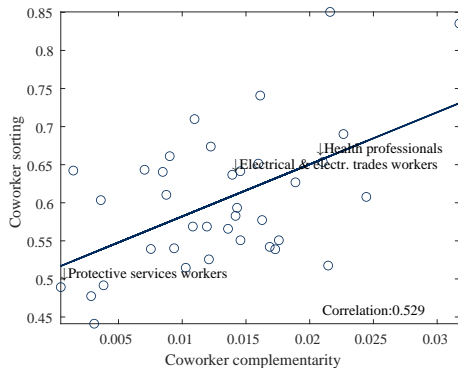
[Main](#)

- \uparrow **Non-routine abstract task intensity**
 $\Rightarrow \uparrow$ **coworker wage complementarity**



Notes. Quadros de Pessoal microdata. Horizontal axis indicates occupation's reliance on non-routine, abstract (NRA) tasks [Mihaylov and Tidens, 2019].

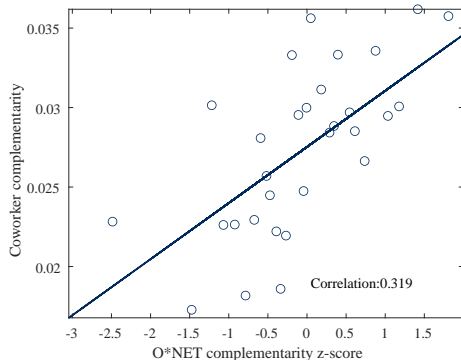
- \uparrow **Coworker wage complementarity**
 $\Rightarrow \uparrow$ **coworker sorting**



Industries: coworker importance \Rightarrow complementarity \Rightarrow sorting

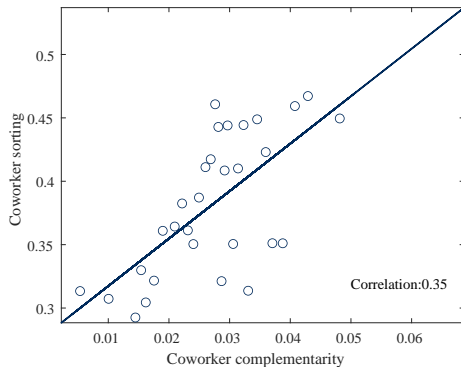
[▶ Main](#)

- \uparrow **Teamwork** [Bombardini et al., 2012]
 $\Rightarrow \uparrow$ **coworker wage complementarity**



Notes. Horizontal axis measures the industry-level weighted mean score of an occupation-level index constructed from O*NET measuring the importance of: teamwork, impact on coworker output, communication, and contact.

- \uparrow **Coworker wage complementarity**
 $\Rightarrow \uparrow$ **coworker sorting**

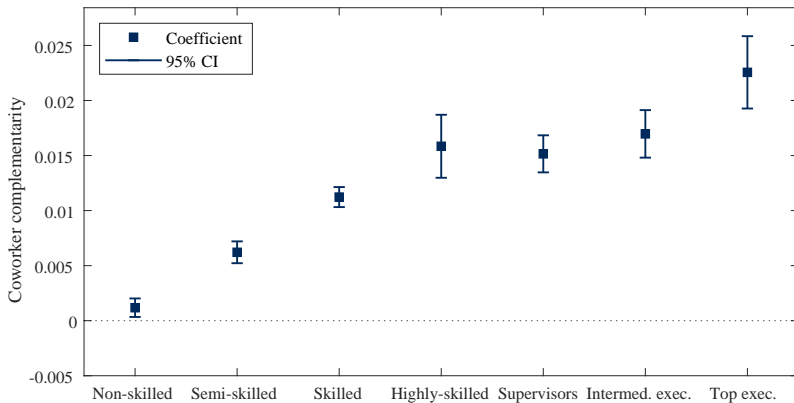


Notes. NACE-4-digit industries.

Hierarchies: complexity \Rightarrow complementarities

[▶ Main](#)[▶ Occupations](#)

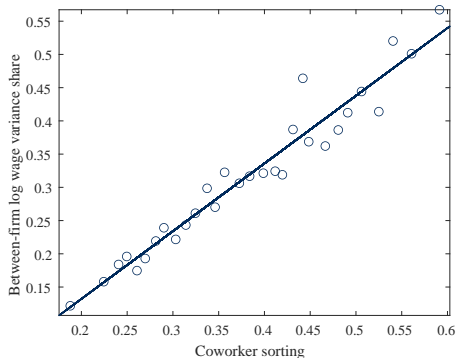
\Rightarrow Coworker wage complementarities are (weakly) \uparrow in the layer of a firm's hierarchy



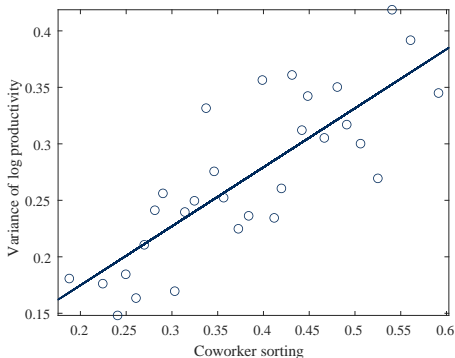
Industries: coworker sorting \Rightarrow between-firm inequality

[▶ Main](#)

\Rightarrow Measures of between-firm inequality in productivity and pay are increasing in the degree of coworker sorting at the industry-level.



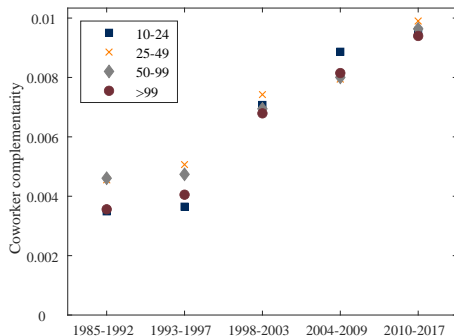
(a) Between-firm share of wage dispersion



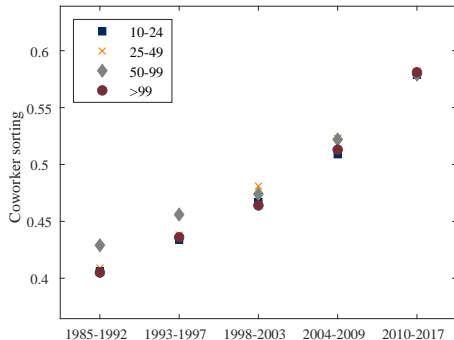
(b) Productivity dispersion

Coworker complementarity & sorting by team size

► Robustness

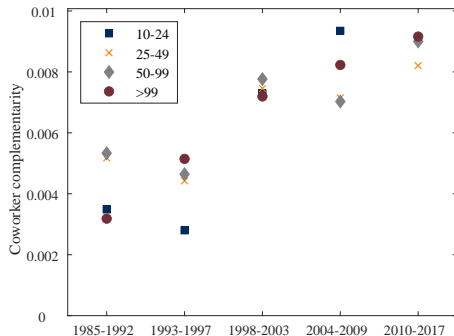


(a) Complementarity

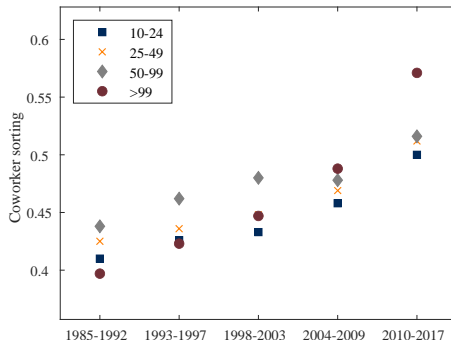


(b) Sorting

Coworker complementarity & sorting by team size – panel estab. only



(a) Complementarity

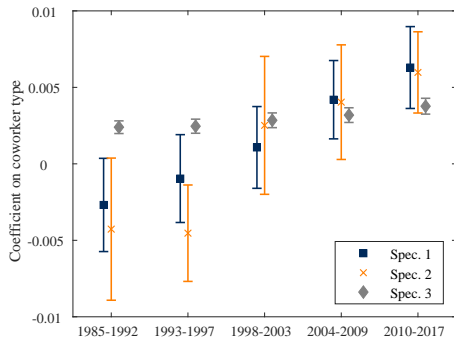


(b) Sorting

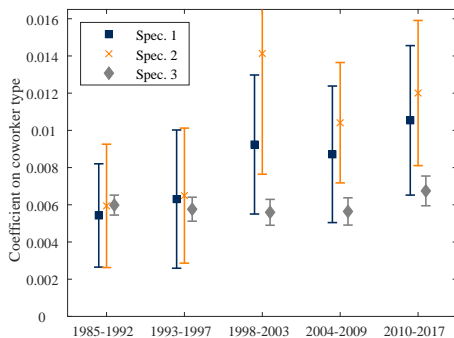
Coworker effects: log wage regression

[▶ Back: cross-section](#)
[▶ Back: time series](#)

$$\ln w_{it} = \beta_0 + \beta_1 \hat{x}_i + \beta_2 \hat{x}_{-it} + \psi_{j(it)} + \nu_{o(i)t} + \xi_{s(i)t} + \epsilon_{it}$$



(a) AKM types



(b) NP types

Notes. Specifications vary by ranking method – within-economy (spec. 1) vs. within-occupation (spec. 2/spec.3) and coworker group definition – establishment-year (spec. 1/spec.2) vs. establishment-occupation-year (spec.3).

Sorting & complementarity based on non-parametric ranking algorithm

- Instead of ranking workers based on AKM worker FEs, use non-param. ranking algo
[Hagedorn et al., 2017]

Period	Sorting		Complementarities	
	Spec. 1	Spec. 2	Spec. 1	Spec. 2
1985-1992	0.47	0.38	0.001	0.000
1993-1997	0.56	0.46	0.002	0.001
1998-2003	0.60	0.48	0.004	0.002
2004-2009	0.65	0.50	0.005	0.002
2010-2017	0.68	0.51	0.005	0.004

Notes. This table indicates, under the column "Sorting" the correlation between a worker's estimated type and that of their average coworker, separately for five sample periods. The column "Complementarities" indicates the point estimate of the regression coefficient β_C . Under "Specification 1" workers are ranked economy wide, while under "Specification 2" they are ranked within two-digit occupations. Worker rankings are based on the non-parametric method.

Coworker complementarity: excluding managers

[▶ Robustness overview](#)

- **Concern** regarding complementarity estimates: driven by managers?
 - only managers benefit from team quality, e.g. via larger span of control
 - the only coworkers that matter are managers

Period	Baseline	Exclude as recipients	Exclude entirely
1985-1992	0.0036***	0.0036***	0.0038***
1993-1997	0.0042***	0.0041***	0.0043***
1998-2003	0.0070***	0.0074***	0.0076***
2004-2009	0.0082***	0.0084***	0.0092***
2010-2017	0.0091***	0.0097***	0.0093***

Notes. Managed are defined based on KldB-1988-3d, as in Jarosch et al. (2023).

Coworker complementarity: movers

[▶ Robustness overview](#)

- Consider sub-samples of job movers, job movers with contiguous employment spells ($t \rightarrow t + 1$), and job movers with non-contiguous E spells ($t \rightarrow t + s$, $s > 1$)
- Caveat: annual panel given data size, no direct observation of U/N spells in SIEED

Period	Baseline	All movers	Contig. E spells	Non-contig. E spells
1985-1992	0.0043***	0.0043***	0.0045***	0.0039***
1993-1997	0.0049***	0.0052***	0.0052***	0.0051***
1998-2003	0.0078***	0.0085***	0.0083***	0.0082***
2004-2009	0.0090***	0.0107***	0.0104***	0.0102***
2010-2017	0.0088***	0.0103***	0.0101***	0.0090***
Obs. in '10-'17 (1000s)	4,410	538	355	375

Notes. Unweighted observations. Regressions include FEs for employer; occupation-year; industry-year. Employer-clustered standard errors in parentheses.

Coworker complementarity: finite-differences approximation

[▶ Robustness overview](#)

- Regression approach imposes strong functional form assumptions on approximated empirical wage function $\hat{w}(x|x')$
 - ofc, mirrored inside structural model when calibrating
- Alternative: construct non-parametric $\hat{w}(x|x')$, then use finite-difference methods to compute the cross-partial derivative (but w/o FE controls)

Period	Regression	Non-parametric FD method
1985-1992	0.0036	0.0073
1993-1997	0.0042	0.0074
1998-2003	0.0070	0.0081
2004-2009	0.0082	0.0120
2010-2017	0.0091	0.0098

Coworker complementarity: lagged types

[▶ Robustness overview](#)

- Concern with both regression approach and non-parametric FD approach: mechanical relationship between wages (“LHS”) and (within-period time-invariant) worker types, which are estimated from wages themselves (“RHS”)
- Robustness check #1: years of schooling as type measure [▶ Jump](#)
- Robustness check #2: assign to each individual i in periods $p \in \{2, 3, 4, 5\}$ the FE estimated for i in period $p - 1$; re-compute worker deciles and average coworker types, \hat{x}_i^{p-1} and $\hat{x}_{-it}^{p-1} = (|S_{-it}|)^{-1} \sum_{k \in S} \hat{x}_k^{p-1}$; re-estimate wage regression
- Results (see paper): magnitude of estimated $\hat{\beta}_c$ around 50% smaller when using lagged types, but evolution over time similar to baseline

Complementarity estimates using years of schooling

[▶ Robustness overview](#)

	'85-'92	'93-'97	'98-'03	'04-'09	'10-'17
Interaction	0.0063*** (0.0008)	0.0060*** (0.0007)	0.0099*** (0.0008)	0.0112*** (0.0007)	0.0129*** (0.0009)
Obs. (1000s)	3,613	2,508	2,694	3,836	4,376
R^2	0.5033	0.5451	0.5746	0.6330	0.6425

Notes. Dependent variable is the wage level over the year-specific average wage. Independent variables are a constant, years of schooling, coworker years of schooling, and the interaction between those two terms. All regressions include industry-year, occupation-year and employer fixed effects. Employer-clustered standard errors in parentheses. Observations are unweighted. The sample is unchanged from the main text, except that 96,517 observations with missing years of schooling are dropped. Observation count rounded to 1000s.

Within-industry empirical analysis

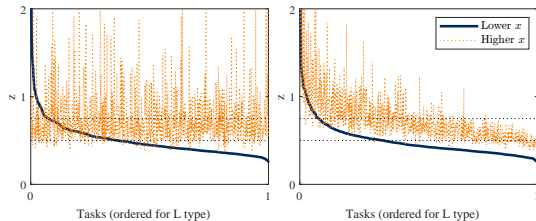
[► Overview: robustness](#)
[► Within-industry calibration](#)

Sample Period	Baseline				Within-industry avg.			
	σ_w^2	σ_w^2/σ_w^2	ρ_{xx}	$\hat{\beta}_c$	σ_w^2	σ_w^2/σ_w^2	ρ_{xx}	$\hat{\beta}_c$
1	0.143	0.337	0.427	0.0036	0.125	0.249	0.333	0.00283
2	0.148	0.391	0.458	0.0042	0.125	0.288	0.351	0.00342
3	0.191	0.456	0.495	0.0070	0.150	0.324	0.369	0.00585
4	0.234	0.547	0.547	0.0082	0.168	0.388	0.405	0.00738
5	0.241	0.568	0.617	0.0091	0.171	0.412	0.464	0.00823

Notes. Within-industry avg. is person-year weighted average across OECD STAN-A38 (2-digit) industries.

Illustration: specialization parameter [▶ Back](#)

- ξ : controls correlation of coworkers' task-specific productivities



(c) Independent draws
($\xi = 1$)

(d) Correlated draws
($\xi \rightarrow 0$)

Proof sketch

- ① Derive $G(p) := \Pr\{\tilde{\lambda}(\tau) \leq p\}$ given $G_i(p) := \Pr\{\lambda_i(\tau) \leq p\}$, using FOC and max-stability property
- ② Use $G(p)$ + standard CES shadow price index to solve (int. by sub.) for

$$\lambda = \left(\int_{\mathcal{T}} \tilde{\lambda}(\tau)^{1-\eta} d\tau \right)^{\frac{1}{1-\eta}}. \quad (5)$$

- ③ Use $G(p)$ and $G_i(p)$ to derive probability that i produces some task τ , which by LLN (continuum assumption!) is equal to share of tasks produced, π_i
- ④ Relate λ_i^L to value of all tasks produced by the worker, $\lambda_i^L = \pi_i \lambda Y$
- ⑤ Normalize $\lambda = 1$, then algebra yields $Y = f(x_1, \dots, x_n; \chi)$

Lemma

Lemma: Lemma

Implied task share and shadow-cost index equal

$$\pi_i = \frac{(x_i/\lambda_i^L)^{\frac{1}{\chi\xi}}}{\sum_{k=1}^n (x_k/\lambda_k^L)^{\frac{1}{\chi\xi}}} \quad x; \lambda = \left(\sum_{i=1}^n \left(\frac{x_i}{\lambda_i^L} \right)^{\frac{1}{\chi\xi}} \right)^{-\chi\xi}$$

Measurement: a useful identification result

[▶ Main](#)
[▶ Non-separable case: scatterplot](#)

- **Q:** How to quantify $\frac{\partial^2 f(x, x')}{\partial x \partial x'}$?
- **Proposition:** production complementarities are proportional to wage compl.
- **Proof sketch:** wage level for worker x with coworker x'

[▶ Wage eq.](#)

$$w(x|x', \xi) = \omega f(x, x', \xi) + g(x) - h(x')$$

where $g : [0, 1] \rightarrow \mathbb{R}$ and $h : [0, 1] \rightarrow \mathbb{R}$ are strictly increasing

\Rightarrow *outside options are separable: affect level of wage but not the cross-partial*

Mapping model to data: coworker types

- Defining $S_{-it} = \{k : j(kt) = j(it), k \neq i\}$ as the set of i 's coworkers in year t , compute the average type of i 's coworkers in year t as $\hat{x}_{-it} = \frac{1}{|S_{-it}|} \sum_{k \in S_{-it}} \hat{x}_k$.
- **Coworker group:**
 - alternative: same establishment-occupation-year cell
 - but CC arise precisely when workers are *differentiated* in their task-specific productivities
- **Averaging step:**
 - equally-weighted averaging ignores non-linearity in coworker aggregation
 - paper: show using non-linear averaging method that baseline results in bias, but it's minor in magnitude
- **Firm size variation:** averaging ensures that a single move will induce a smaller change in the *average* coworker quality in a large team than in a small one

Mapping model to data: identification strategy for χ

[▶ Main](#)

- **Literature:** complementarities – primarily between workers and firms – usually inferred indirectly from sorting patterns
 - exception: Hagedorn-Law-Manovskii (2017)
- **This paper:** directly measure coworker complementarity in the data, recover χ structurally given $\gamma = \frac{\chi}{\chi+1}$
- Paper does *not* use microfoundation itself to measure χ , respectively γ
- Experiment: fit a (truncated) Fréchet distribution to Grigsby's (2023) non-parametric estimates of the multi-dimensional skill dist. estimated from CPS data
 - recover $\gamma = 0.84$ for 2006 but *very* noisy estimates
- **Ongoing work:** use the extended microfoundation to identify χ

[▶ Jump](#)[▶ Jump](#)

Direct estimation of χ : proof of concept

- Grigsby (2023): only paper that provides a *cardinal* measure of skill task-specificity
- Evidence on time trends are qualitatively consistent with “specialization hypothesis”: cross-type average of within-type variance across specific skills grew by nearly 50% b/w 1980s and 2000s & skill transferability has declined amongst high-skill occupations
- His operationalization of worker types and tasks does *not* directly map onto my model (no identifying assumption; coarse occupational skills; US vs DE data)
- **Proof of concept:** but *suppose* we just take those data, extract moments capturing average within-worker cross-task efficiency dispersion, fit a (truncated) Fréchet, recover $\gamma = \frac{\chi}{1+\chi}$
 $\Rightarrow \checkmark \gamma$ **similar to structural estimation result based on evidence from wage CC**

Semi-structural back-of-envelope calculation for γ

- Structurally recover $\gamma \frac{x}{x+1}$ by estimating $\frac{\partial^2 w(x|x')}{\partial x \partial x'}$ in the data, which was shown to be proportional to $\frac{\partial^2 f(x, x')}{\partial x \partial x'}$
- But how is $\frac{\partial^2 f(x, x')}{\partial x \partial x'}$ related to γ ?
- Definitionally, $\gamma = (f f_{ij}) / (f_i f_j)$ for any $i \neq j$
- Can we avoid full structural model? \Rightarrow If have measures not only of f_{ij} but also output f and marginal products f_i
- Suppose, for any x and x' , we use wages to back out marginal products – competitive wage determination rather bargaining! – and recover output from sum of wages divided by labor share
- Find $\gamma \approx 0.79$ – very close to structural estimate!

Overview of validation exercises: direction EE transitions & cross-section

[▶ Main](#)

- 2 additional types of validation exercise:

- ✓ **EE transitions** reallocate workers to more + assortative matches
- do model-implied relationships also hold in **cross-section**?

[▶ Details](#)

① $\chi \uparrow \Rightarrow$ coworker complementarity \uparrow

② coworker complementarity $\uparrow \Rightarrow$ + assortative matching \uparrow

can test predictions *because*
we have measures of comple-
mentarity!

- Implementation of cross-sectional exercises: rich Portuguese micro data

- universe of private-sector actors, employer-employee data & income statements

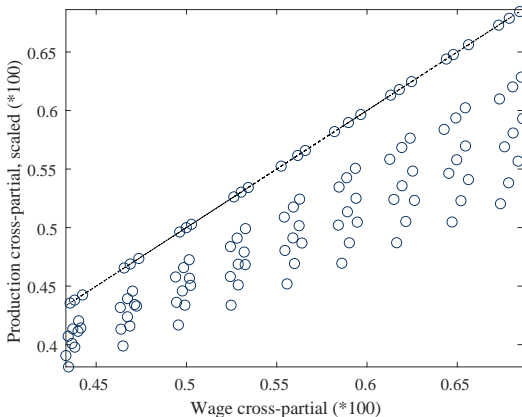
- Cross-sectional exercises:

- ✓ **Hierarchies**
- ✓ **Industries**
- ✓ **Occupations**

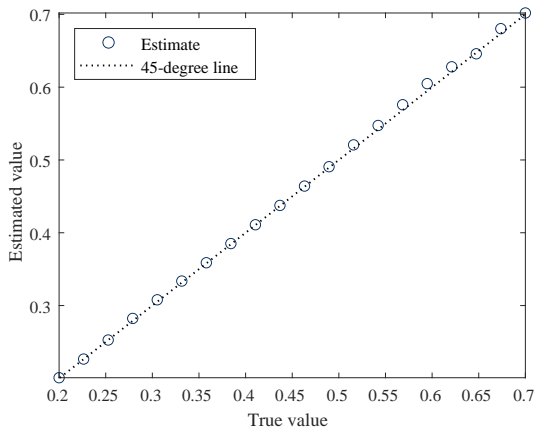
[▶ Details](#)[▶ Details](#)[▶ Details](#)

Wage and production cross-partials beyond the benchmark

- Solve model for many combinations of χ , λ_e and b
- Compare FD approx of $f_{xx'}(x, x')$, scaled by ω , and $w_{xx'}(x|x')$
- Main parameter driving wedge: λ_e



Monte Carlo study

[▶ Main](#)

Implications for aggregate productivity

[► Effects of \$\chi\$ ↑: random vs eqm](#)

- **Production complementarities imply coworker sorting matters for agg productivity**

$$\circ f(x_1, \dots, x_n) = n^{1+\chi\xi} \times \left(\frac{1}{n} \sum_{i=1}^n (x_i)^{\frac{1}{1+\chi\xi}} \right)^{1+\chi\xi}$$

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 \Rightarrow 2010s gap: 2.05%, similar for earlier periods

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- **Search** frictions induce misallocation \sim coworker **mismatch**
- **Quantify** mismatch costs: compare eqm outcome vs to productivity under pure PAM
 \Rightarrow 2010s gap: 2.05%, similar for earlier periods
- **Trends:** \uparrow talent sorting limited \uparrow in mismatch costs given $\chi \uparrow$
 \Rightarrow no-reallocation counterfactual: productivity gap 4.65%

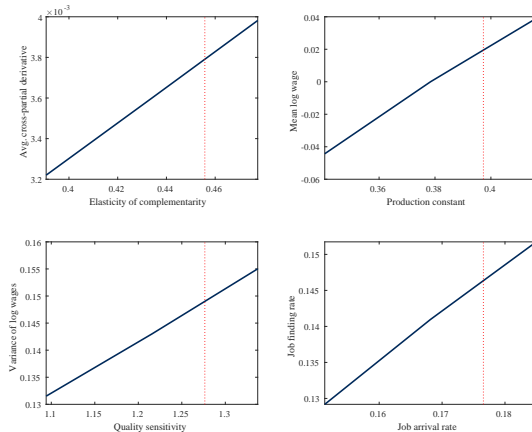
Parameterization, including estimation results (2010s)

Parameter	Description	Targeted moment	Value	m	\hat{m}
γ	Elasticity of complementarity	$\hat{\beta}_c$	0.837	0.0091	0.0091
a_0	Production, constant	Avg. wage (norm.)	0.239	1	1
a_1	Production, scale	Var. log wage	1.557	0.241	0.241
b_1	Replacement rate, scale	Replacement rate	0.664	0.63	0.63
δ	Separation hazard	Job loss rate	0.008	0.008	0.008
λ_u	Meeting hazard	Job finding rate	0.230	0.162	0.162
ρ	Discount rate	External	0.008		
ω	Worker bargaining weight	External	0.50		
a_2	Production, team advantage	External	1.10		

Estimated parameters and targeted moments

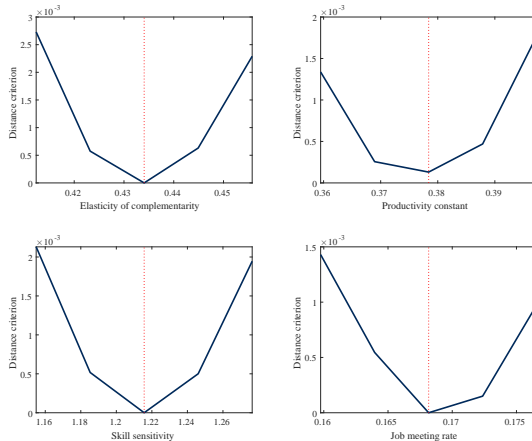
Parameter	Targeted moment	2010-2017			1985-1992		
		Value	m	\hat{m}	Value	m	\hat{m}
γ	$\hat{\beta}_c$	0.837	0.0091	0.0091	0.434	0.0035	0.0035
a_0	Avg. wage (norm.)	0.239	1	1	0.378	1	1
a_1	Var. log wage	1.557	0.241	0.241	1.216	0.143	0.143
b_1	Replacement rate	0.664	0.63	0.63	0.740	0.69	0.69
λ_u	Job finding rate	0.230	0.162	0.162	0.168	0.141	0.141
δ	Job loss rate	0.008	0.008	0.008	0.007	0.007	0.007

Identification validation exercise 1

[▶ Main](#)

Notes. This figure plots the targeted moment against the relevant parameter, holding constant all other parameters.

Identification validation exercise 2



Notes. This figure plots the distance function $\mathcal{G}(\psi_i, \psi_i^*)$ when varying a given parameter ψ_i around the estimated value ψ_i^* . The remaining parameters are allowed to adjust to minimize \mathcal{G} .

What about shifts in the talent distribution?

[▶ Robustness overview](#)

- **Q:** could \uparrow economy-wide x -dispersion, σ_x , also explain \uparrow firm-level inequality?
 - Kremer-Maskin (1996): as the dispersion of skills (talent) \uparrow in the economy, the relative dispersion of talent within firms \downarrow
- **Obs. #1:** Hakanson et al. (2021): direct skill measures (from military enlistment tests) point to Flynn effect but not to increased dispersion (in cognitive test scores)
- **Obs. #2:** model exercise
 - method: instead of rank interpretation, i.e. $x \sim U$, we separately parameterize σ_x by assuming a \mathcal{N} distribution
 - finding: $\sigma_x \uparrow \Rightarrow \rho_{xx} \uparrow, \sigma_w \uparrow$, and $\sigma_{\bar{w}}/\sigma_w \uparrow$ – *but* no measured increase in $\hat{\beta}_c$
- **Conclusion:** empirically unclear whether $\sigma_x \uparrow$, and if so, this would not explain the observed increase in coworker complementarity, i.e. latter is a distinct channel

Training policies in a team production context (w-i-p)

[► Overview](#)

- “Training policies” \sim non-parametric perturbations of the talent distribution
 - left-tail intervention: give everyone in 1st decile productivity of those in 2nd decile
 - right-tail intervention: give everyone in 9th decile productivity of those in 10th decile
- Team production: effect of “training policies” partially via coworker spillovers!
- Relative effectiveness of left-tail vs right-tail intervention:
 - 1 the stronger are coworker complementarities, the relatively greater are the realized productivity gains from a left-tail intervention, b/c low- x tend to be weak links
 - 2 but raising the productivity of coworkers of workers with high productive potential generates greater gains – and with sorting, those coworkers are themselves high- x
- Equilibrium: relative effectiveness of left-/right-tail training depends on both forces
 - tentative quantitative finding: right-tail intervention boosts average productivity by more *but* left-tail training also lowers inequality

The effect of declining search frictions

- ↓ search frictions could also explain ↑ coworker sorting
- Job arrival & separation rates estimated to ↑ from p1 to p2
- **Counterfactual analysis:** explains 6% of model-implied ↑ in between-employer share of wage variance

	Δ model	Implied % Δ model due to Δ parameter
Model 1: baseline	0.159	-
Cf. a: fix period-1 complementarity	0.065	59
Cf. b: fix period-1 search frictions	0.150	6

Outsourcing & within-occupation ranking analysis

- **Concern:** confounding shifts in labor boundary of firm, e.g. outsourcing
- **Address this concern in multiple steps:**
 - ① empirically rank workers *within* occupation (“good engineer vs. mediocre engineer”)
 - ② empirically re-estimate coworker sorting & complementarity (lower but similar \uparrow)
 - ③ re-estimate model for both periods & re-do counterfactual exercises
- **Result:** qualitatively & quantitatively similar findings

	Δ model	Implied % Δ model due to Δ parameter
Model 2: within-occ. ranking	0.198	-
Cf. a: fix period-1 comp.	0.076	61.47

Robustness: model with OJS - brief overview

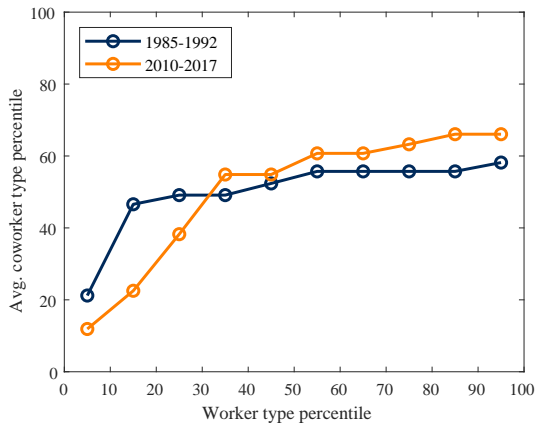
[▶ Main](#)

- **Baseline:** no job-to-job transitions – but is main story robust when workers can switch to better job after accepting job out of unemployment?
 - two opposing effects from increased complementarities
- **Extension:** employed workers also meet vacancies at Poisson rate λ_e
- **Main findings:**
 - better fit to empirical sorting patterns in cross-section [▶ Comparison](#)
 - contribution of \uparrow complementarities to \uparrow firm-level wage inequality slightly smaller, more attributed to \uparrow labor market transitions
 - conservative estimates (endogenous search effort, forward-looking wage specification)
- Opens door to thinking about coworker complementarities and **job ladders** [▶ Jump](#)

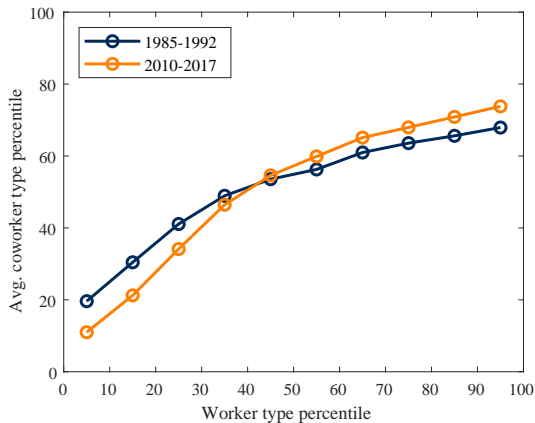
Extension: model with OJS

- Baseline model abstracted from OJS
 - transparent trade-off, connection to analytical results
- **Consider extension to OJS:** employed worker meet vacancies at Poisson rate λ_e
 - wages both off and on the job are continuously renegotiated under Nash bargaining, with unemployment serving as the outside option *[cf. di Addario et al., 2021]*
 - re-estimate, with empirical labor market flows disciplining λ_e
- Qualitative question: is coworker sorting outcome robust, even if workers can switch to better job after accepting job out of unemployment?
- **Analyses:**
 - 1 coworker sorting patterns & changes
 - 2 additional model validation: direction of EE flows in model & data

Model-implied coworker sorting patterns: without and with OJS

[▶ Main](#)

(a) Baseline



(b) OJS

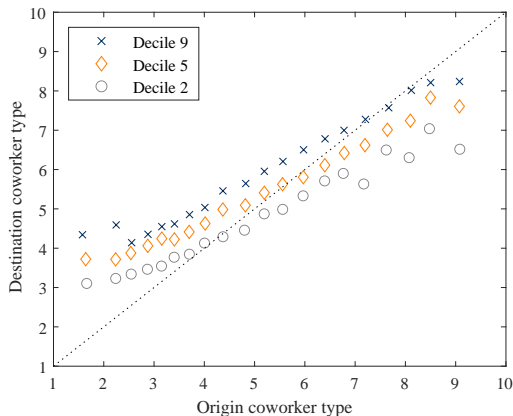
EE transitions in theory and data

- **Theoretical prediction:** EE transitions move workers in surplus-maximizing direction
 $\Rightarrow \Delta \hat{x}_{-it} = \hat{x}_{-i,t} - \hat{x}_{-i,t-1}$ should be *positively* correlated with \hat{x}_i
 - $h_{2.1}(x, x'' | x') = 1$ – worker x in a two-worker firm with coworker x'' would move to an employer that currently has one employee of type x' – if $S(x|x') - S(x|x'') > 0$
- **Empirical analysis:** use SIEED *spell* data to create worker-originMonth-destinationMonth-originJob-destinationJob panel, with information on characteristics of origin and destination job
 - subsample period 2008-2013 (huge panel at monthly frequency)
 - count as “EE” if employer change between two adjacent months
- **Regression analysis:** regress $\Delta \hat{x}_{-it}$, scaled by std. σ_{Δ} of coworker quality changes, on *own* type and *origin* coworker type

$$\frac{\Delta \hat{x}_{-it}}{\sigma_{\Delta}} = \beta_0 + \beta_1 \hat{x}_i + \beta_2 \hat{x}_{-i,t-1} + \epsilon_{it}$$

Empirical coworker sorting changes due to EE moves

- **EE transitions push toward greater coworker sorting:** for any given origin, higher x-workers move to workplaces with better coworkers than lower-x workers do
- *But* in data EE transitions “move up” low types more than theory predicts
- “**Coworker job ladder**” with both absolute and type-specific dimension?
- **Next:** change in the job ladder [e.g., Haltiwanger-Spetzler, 2021]



Evidence that EE *increasingly* reallocate toward PAM: in data & model

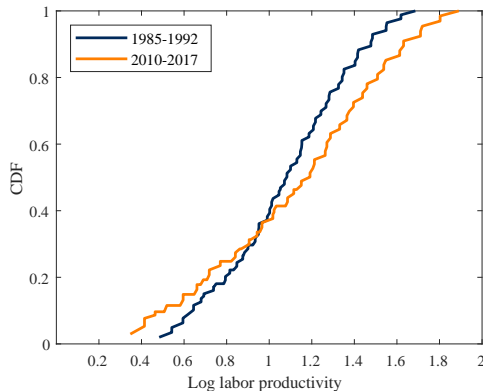
	Data		Model	
<i>Change in coworker type</i>	'85-'92	'10-'17	Period-1	Period-2
Own type	0.0883 ^{***} (0.000799)	0.118 ^{***} (0.000918)	0.214	0.270
Controls	Year FEs, Origin	Year FEs, Origin	Origin	Origin
<i>N</i>	196,098	282,718	∞	∞
adj. <i>R</i> ²	0.284	0.204		

Table 1: Change in coworker type due to EE moves positively related to own type – increasingly so

Notes. For the data columns, individual-level clustered standard errors are given in parentheses. Model counterparts are computed simulation-free in population. Dependent variable is scaled throughout by the standard deviation of the change in coworker type.

Productivity dispersion

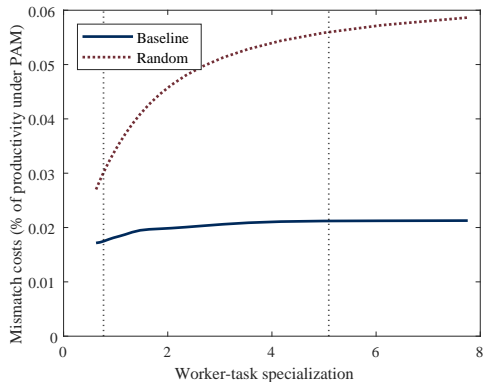
- Firm dynamics literature: increased firm-level productivity dispersion [Autor et al., 2020; de Ridder, 2023]



Productivity costs of complementarity & labor market functioning

“The benefits of the division of labor are limited by the functioning of the labor market”

- Microfoundation: $\uparrow \chi \Rightarrow \uparrow$ efficiency *benefit* from teamwork *but* also \uparrow mismatch costs
- **Q:** how does the gap to potential vary depending on labor market structure?
- **A:** under random sorting, productivity gap due to misallocation \uparrow more sharply as $\chi \uparrow$
- Outside model: severe labor mkt frictions (e.g., dev'ing countries [Donovan et al., 2023]) may inhibit specialization [cf. Atencio et al., 2023; Bassi et al., 2023]



Implications for overall inequality?

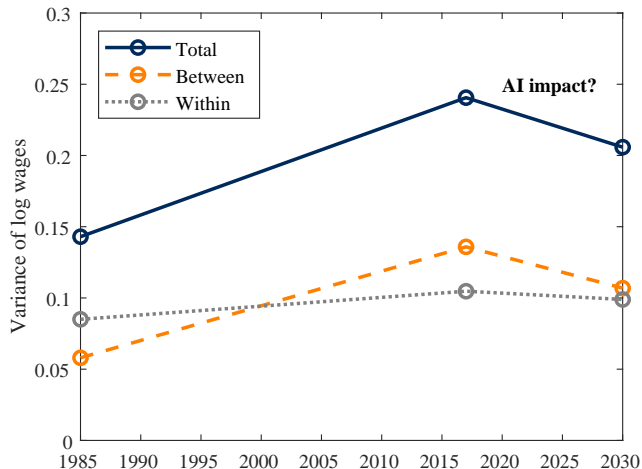
- **Coworker complementarities do not necessarily \uparrow variance of person-level wages**
 - (un-)surprising? Variance decomposition perspective vs. common intuition [Kremer, 1993]
 - (i) **reallocation effect**, (ii) valuation effect, (iii) outside option effect
- Several mechanisms through which \uparrow sorting could \uparrow person-level inequality
 - 1 regulation or norms that lead to within-firm wage compression [Akerlof-Yellen, 1990]
 - 2 coworker learning [Jarosch et al., 2021; HLMP, 2023]
 - 3 increasing returns to labor quality [Kremer, 1993]

Implications for AI impact on labor markets (1): overview

- Literature moving toward richer way of thinking about tech. change. Here: shifts in across-worker productivity differences *and* in specialization/interdependence
- Example: **AI** – early/conjectural evidence on impact of LLMs etc
 - *absolute* adv. less important (“leveller”) [e.g., Brynjolfsson et al., 2023]
 - *comparative* adv. less important (e.g., can interpret medical imaging w/o radiologist)
- **Illustrative model counterfactual:** relative to the 2010s
 - ① everyone's productivity ↑ by equivalent to 20% of lowest type's productivity – so in proportional terms, lower-x benefit more
 - ② coworker complementarity ↓ by 20%

Implications for AI impact on labor markets (2): illustrative exercise

- Model prediction: AI could lead to reversal of historical trends – $\sigma_W^2 \downarrow$, $\sigma_{\bar{W}}^2 \downarrow$, $\rho_{XX} \downarrow$ – along side a productivity boom
 - driven by \downarrow firm-level wage inequality



Implications for AI impact on labor markets (3): conjectures

[▶ Main](#)

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- Example: AI – early/conjectural evidence on impact of LLMs etc
 - *absolute* adv. less important? (“leveller”) [e.g., Brynjolfsson et al., 2023]
 - *comparative* adv. less important? (e.g., can interpret medical imaging w/o radiologist)
- Illustrative model counterfactual: **AI could lead to reversal of historical trends...**
- **...and perhaps also...**
 - ① **...flatter organizations**, with managerial span of control ↑
 - ② **...↓ barriers to entry** for self-employment/start-ups
 - ③ **...easier job transitions** & shorter training durations

Within-industry calibration: overview

[▶ Main](#)[▶ Data moments](#)

- Baseline: calibration and evaluation based on economy-wide moments
- But the model does not incorporate between-industry differences in, e.g., production technology
- Alternative considered here: target $\hat{\beta}_c$ and σ_w^2 computed as within-industry average, evaluate against within-industry trends
 - keep other targets (e.g., job separation) as before

Within-industry calibration: model fit & counterfactual

- Counterfactual: $\chi \uparrow$ explains 83% of model-implied \uparrow in between-share

