Superstar Teams

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Motivation: opening the "black box" of firms in macro

- Most production processes involve groups of heterogeneous workers
 - o too complex for any individual to perform all tasks well
 - → firms coordinate groups of workers with specialized skills
 - o individuals also vary in their overall productivity ("talent")

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 - but theoretically & micro oriented
- This paper: tractable model for empirical & quantitative analysis that highlights macro implications: agg. productivity, trends in labor market inequality

 - (1) theory (2) measurement
- (3) application(s)

Modeling firms as team assemblies & core mechanism

- What's a firm: an organized collection of workers with task-specific skills (team)
- **Key idea:** specialization \Rightarrow coworker talent complementarities \Rightarrow sorting
- Implications:
 - $\circ \uparrow$ specialization \Rightarrow talent sorting \uparrow
 - o input/labor market frictions lower output b/c teams are formed inefficiently

- Develop tractable theory of the firm centered on team-production & formation
 - 1 task-based production
 - 2 skill heterogeneity: task-specificity & variable talent
 - **3** team production
 - assembling a team takes time due to search frictions

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 - o theory has implications for agg. productivity, productivity dispersion, job ladders, ...

Relation & contributions to literature

• Firm organization: parsimonious team-model for quantitative applications

Firms: Lucas, 1978; Rosen, 1982; Becker & Murphy, 1992; Hopenhayn & Rogerson, 1993; **Kremer, 1993**; Kremer & Maskin, 1996; **Garicano, 2000**; Klette & Kortum, 2004; **Garicano & Rossi-Hansberg, 2006**; Kohlhepp, 2022; Kuhn et al., 2022; Minni, 2022; Bassi et al., 2023

Task assignment: Costinot & Vogel, 2010; Acemoglu & Restrepo, 2018; Ocampo, 2021; Adenbaum, 2022

Teams: Akcigit et al., 2018; Chade & Eeckhout, 2020; Jarosch et al., 2021; Herkenhoff et al., 2022

- Frictional labor market sorting: endogenize & measure complementarities

 Shimer & Smith, 2000; Cahuc et al., 2006; Eeckhout & Kircher, 2011/2018; Hagedorn et al., 2017; de Melo,
 2018; Herkenhoff et al., 2022; Lindenlaub & Postel-Vinay, 2023
- Wage inequality: structural model of ↑ firm-level inequality due to technological △
 Technology: Katz & Murphy, 1992; Krusell et al., 2000; Autor, Levy & Murnane, 2003; Jones, 2009; Deming, 2017; Acemoglu & Restrepo, 2018; Alon, 2018; Neffke, 2019; Jones, 2021; Atalay et al., 2021
 Firms: Card et al., 2013; Barth et al., 2016; Alvarez et al., 2018; Bloom et al., 2019; Aeppli & Wilmers, 2021; Criscuolo et al. 2021; Hakanson et al., 2021; Sorkin & Wallskog, 2021; Kleinman, 2022

Theory

Overview of environment

- Ex-ante homogeneous firms assemble teams of heterogeneous workers who vary in task specialization (what you're skilled at) & talent (how skilled you are)
 - task-based production
 - hire workers in search-frictional labor market
- Challenge: how to keep this tractable?
- Analyze economy-wide organization of production in 2 steps:
 - 1 team production: how much does a team produce if tasks are optimally allocated?
 - 2 team formation: who is matched to work with whom?

Step 1: production in a single team of given composition

• Firm: 1 team of $n \in \mathbb{Z}_{++}$ workers produces output from unit continuum of tasks \mathcal{T}

$$\ln Y = \int_{\mathcal{T}} \ln q(\tau) d\tau \tag{1}$$

Task aggregation:

$$q(\tau) = \sum_{i=1}^{n} y_i(\tau) \tag{2}$$

• Workers o tasks: i supplies 1 time unit, task-specific productivities $\mathbf{z}_i = \{\mathbf{z}_i(\tau)\}_{\tau \in \mathcal{T}}$

$$y_i(\tau) = z_i(\tau)l_i(\tau) \tag{3}$$

$$1 = \int_{\mathcal{T}} l_i(\tau) d\tau \tag{4}$$

Firm's optimization problem

- Firm solves mini-planner problem: max_{q,{l,}} Y s.t. (1)-(4)
- **Preview:** derive & characterize reduced-form team production function f

$$f(\mathbf{z}_1, ..., \mathbf{z}_n) = \max Y$$

s.t. (1)-(4)

Firm's optimization problem

• Firm solves mini-planner problem: max Y s.t. (1)-(4)

$$\mathcal{L}(\cdot) = \mathbf{Y} + \lambda \left[\underbrace{\left(\int_{\mathcal{T}} \ln q(\tau) d\tau \right) - \ln \mathbf{Y}}_{\text{tasks} \to \text{output}} \right] + \int_{\mathcal{T}} \lambda(\tau) \left(\underbrace{\sum_{i=1}^{n} y_{i}(\tau) - q(\tau)}_{\text{task aggregation}} \right) d\tau$$

$$+ \sum_{i=1}^{n} \lambda_{i}^{L} \underbrace{\left(\int_{\mathcal{T}} \frac{y_{i}(\tau)}{\mathbf{z}_{i}(\tau)} d\tau - 1 \right)}_{\text{time constraint} + \text{task production}} + \text{non-negativity constraints}$$

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FOCs imply

simply
$$\lambda(\tau) = \min_i \left\{ \frac{\lambda_i^L}{\mathbf{z}_i(\tau)} \right\}$$
 shadow cost of τ with $\lambda(\tau) = \min_i \left\{ \frac{\lambda_i^L}{\mathbf{z}_i(\tau)} \right\}$ productivity of t for t

Tractability: leverage insight from trade literature

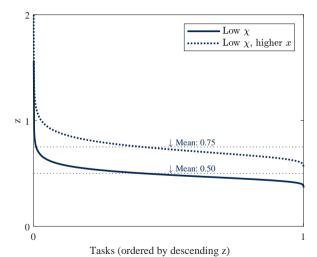
Assumption: Multivariate Fréchet distribution of worker-task skills

$$\Pr[Z_1(\tau) \leq z_1, \dots, Z_n(\tau) \leq z_n] = \exp\left[-\left(\sum_{i=1}^n \left(\left(\frac{z_i}{\iota X_i}\right)^{-\frac{1}{\chi}}\right)^{\frac{1}{\xi}}\right)^{\frac{1}{\xi}}\right]$$

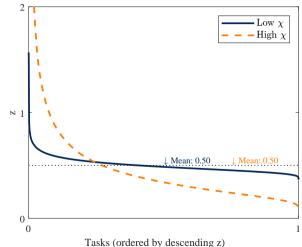
- Payoff #1: low-dim. representation of worker-task skill dist.
 - o x_i : scale term \sim "talent" type
 - χ : within-worker skill dispersion ~ specialization
 - o $\xi \in (0,1]$: correlation in z's across workers \sim are team members specialized in the same or different tasks

Illustration: talent types

 "Talent" type x_i: how skilled is i conditional on whatever task they're specialized in (absolute advantage)



 Parameter *x*: within-worker skill dispersion across tasks \sim specialization



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- Payoff #1: low-dim. representation of worker-task skill dist.
- Payoff #2: max-stability property allows closed-form characterization of key objects

Micro-founded production function



Proposition: Reduced-form production function

Talent types $(x_1,...,x_n)$, χ and ξ are sufficient statistic for team output Y:

$$Y = f(x_1, \cdots, x_n; \chi, \xi)$$

Proof sketch



- solving firm's mini-planner problem implies task-demand and assignment of tasks by comparative advantage with complete division of labor
- \circ Fréchet distribution allows closed-form characterization of distribution of shadow costs $\lambda(au)$ conditional on this assignment
- \circ integrate over tasks and workers, normalizing $\lambda = 1$

Micro-founded production function: characterization

Proposition: Reduced-form production function

$$f(\mathbf{x}_1,\cdots,\mathbf{x}_n;\chi,\xi)=n^{1+\chi\xi}\times\left(\frac{1}{n}\sum_{i=1}^n(\mathbf{x}_i)^{\frac{1}{1+\chi\xi}}\right)^{1+\chi\xi}$$

 \rightarrow vs. no-division-of-labor: $f(x_1,...,x_n)=n imes (rac{1}{n}\sum_i^n x_i)$

1 Output depends on team composition – f is not additively separable

The whole is greater than the sum of its parts...

Proposition: Reduced-form production function

$$f(\mathbf{x}_1,\cdots,\mathbf{x}_n;\chi,\xi) = \underbrace{n^{1+\chi\xi}}_{\text{efficiency gains}} \times \left(\frac{1}{n} \sum_{i=1}^{n} (\mathbf{x}_i)^{\frac{1}{1+\chi\xi}}\right)^{1+\chi\xi}$$

- **2** Efficiency gains from teamwork large when workers have skills that are specialized (χ) in different (ξ) tasks
 - o no-division-of-labor: $f(\cdot) = n \times (\frac{1}{n} \sum_{i=1}^{n} x_i)$
 - o intuition: gains from tasks being assigned by comparative advantage

▶ task assignment

...but output is lowered by talent dispersion

Proposition: Reduced-form production function

$$f(x_1, \dots, x_n; \chi, \xi) = \underbrace{n^{1+\chi\xi}}_{\text{efficiency gains}} \times \underbrace{\left(\frac{1}{n} \sum_{i=1}^{n} (x_i)^{\frac{1}{1+\chi\xi}}\right)^{1+\chi\xi}}_{\text{talent complementarity}},$$

- **3** Coworker talent complementarities strong when workers have skills that are specialized (χ) in different (ξ) tasks
 - \circ reduced-form elasticity of complementarity $\gamma := \frac{\partial \ln(f_j/f_i)}{\partial \ln(x_i/x_i)} = \frac{\chi \xi}{1+\chi \xi}$
 - o greater dispersion in team members' talent reduces output, other things equal
 - $\circ \partial f(\cdot)/\partial x_i\partial x_{-i} > 0$

...but output is lowered by talent dispersion

Proposition: Reduced-form production function

$$f(x_1, \dots, x_n; \chi, \xi) = \underbrace{n^{1+\chi\xi}}_{\text{efficiency gains}} \times \underbrace{\left(\frac{1}{n} \sum_{i=1}^{n} (x_i)^{\frac{1}{1+\chi\xi}}\right)^{1+\chi\xi}}_{\text{talent complementarity}},$$

- **3** Coworker talent complementarities are strong when workers have skills that are specialized (χ) in different (ξ) tasks
 - o intuition: greater specialization means higher-x team members can't easily perform more tasks as their skill rapidly diminishes

 ▶ task assignment

Key takeaways

lacktriangledown \uparrow Skill specialization endogenously generates \uparrow coworker talent complementarities

Firm organization meets frictional matching

- So team composition matters for productivity, especially when χ is high
- How is team composition determined?
- Exploit tractability of team production fn. f to integrate team production with dynamic eqm. matching
 - o team composition determined by tradeoff between complementarities and search costs
 - \circ preview: dynamic setting will also be key for identification of f with panel data

Frictional matching into teams: environment

- Random-search + multi-worker firms [Herkenhoff-Lise-Menzio-Phillips, 2024]
 - o teams instead of one-worker firms [e.g., Shimer-Smith, 2000; Hagedorn-Law-Manovskii, 2017]
- **Production technology:** $f(\cdot)$ with $n \in \{0, 1, 2\}$
- **Employment states:** unemp., employed alone, employed with $x' \in \mathcal{X}$
- Nash wage bargaining with continuous renegotiation
- ξ : **match-specific shock** observed by firms & workers before match decision
 - o microfoundation: coworkers' task specialization matters for f, not only x
 - $\circ \xi$ is endogenously determined

Matching - stationary equilibrium



HJ-Bellman equations → values & matching policies

► HJBs

• Flows between/**distribution** over types \times employment states

► KFEs

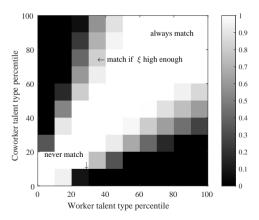
Definition: Stationary equilibrium

A stationary eqm. consists of a value functions & a distribution of agents, s.t.

- 1 the value functions satisfy the HJB equations given the distribution;
- 2 the distribution is stationary given the policy fn's implied by the value fn's.

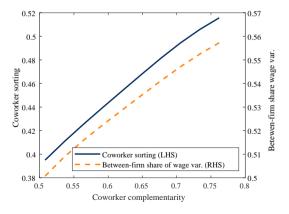
Mechanism: conditional matching probabilities for given χ

• What is the probability that a firm with a type-x worker would hire a type-x' worker conditional on meeting?



Mechanism: more positive assortative matching as $\chi\uparrow$

• What is the correlation between talent x & coworker talent x', ρ_{xx} ?



Key takeaways

- \bigcirc \uparrow Skill specialization endogenously generates \uparrow coworker talent complementarities
- ↑ Talent complementarities lead to ↑ positive assortative matching

Model Meets Data

Taking the model to (micro) data

- Goal: test core mechanisms & analyze quantitative properties of the model
- Steps:
 - matched employer-employee panel data
 - map model objects to data
 - 3 estimate & quantitatively analyze model
 - validate core model mechanisms (brief today)

Primary data: SIEED matched-employer employee panel for W Germany

- 1.5% sample of establishments + entire biographies of associated workers; social security information on employer, daily wage, occupation, demographics
- o initially focus on 2010-2017, later extend to 1985-2017
- \circ selection: f-t employees aged 20-60, \geq 10 observations per establishment-year

What do we need to know to quantitatively analyze the model?

- \circ some parameters e.g. labor market separation and meeting rates can be \sim directly read off micro data or are informed by literature (e.g., discount rate)
- \circ x and x' can be recovered using applied micro methods
- \circ key challenge: χ

Mapping theory to data: worker & coworker types

- **Theory:** wage monotonically \uparrow in x, so can measure using panel dimension
- Implementation: fixed effect (FE) wage regressions
 - AKM model [Abowd et al., 1999]: job switching identifies time-invariant "worker types,"
 controlling for possibility that some employers pay more to all workers
 pre-est. k-means clustering to address limited mobility bias [Bonhomme et al., 2019]
 - AKM misspecified according to structural model, so also use a non-param. ranking algo [Hagedorn et al., 2017] that's consistent → similar worker types
 - \Rightarrow Worker i's talent type \hat{x}_i : decile rank of i's FE within 2d-occupation

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 - \Rightarrow Worker i's talent type \hat{x}_i : decile rank of i's FE within 2d-occupation
- "Representative coworker type" \hat{x}_{-it} : avg. \hat{x} of workers in same estab.-yr.



• Then can compute, e.g., $\hat{\rho}_{xx} = \operatorname{corr}(\hat{x}_i, \hat{x}_{-it})$

Measurement: a useful identification result

- Within-worker skill dispersion χ isn't easily observable
- **Theory:** Proposition 1 ties χ to $\frac{\partial^2 f(\cdot)}{\partial x \partial x'}$, which given x & x' we can identify from w(x|x')

Proposition: Measuring complementarities

Coworker complementarities (CC) in production are proportional to CC in wages:

$$\frac{\partial^2 f(x,x',\xi)}{\partial x \partial x'} \propto \frac{\partial^2 w(x|x',\xi)}{\partial x \partial x'}.$$

$$\Rightarrow \frac{\partial^2 \bar{f}(x,x')}{\partial x \partial x'} \propto \frac{\partial^2 \bar{w}(x|x')}{\partial x \partial x'}.$$
can measure this

▶ Proof sketch

Key takeaways

- \bullet \uparrow Skill specialization endogenously generates \uparrow coworker talent complementarities
- ↑ Talent complementarities lead to ↑ + assortative coworker matching
- **3** Complementarities can be identified from panel micro data on wages and matches

Mapping theory to data: coworker complementarity

- Want to approximate $\frac{\partial^2 \bar{w}(x|x')}{\partial x \partial x'}$
- Estimating polynomial equation:

$$\frac{w_{it}}{\bar{w}_{t}} = \beta_{0} + \beta_{1}\hat{x}_{i} + \beta_{11}\hat{x}_{i}^{2} + \beta_{2}\hat{x}_{-it} + \beta_{22}\hat{x}_{-it}^{2} + \beta_{c}\left(\hat{x}_{i} \times \hat{x}_{-it}\right) \\
+ \psi_{j(it)} + \nu_{o(i)t} + \xi_{s(i)t} + \epsilon_{it}$$

o alternative: non-parametric approximation

• Estimating equation:

coworker complementarity

$$\frac{\mathbf{w}_{it}}{\bar{\mathbf{w}}_{t}} = [\ldots] + \hat{\beta}_{c} \left(\hat{\mathbf{x}}_{i} \times \hat{\mathbf{x}}_{-it}\right) + [\ldots]$$

| | $\hat{eta}_{m{c}}$ | Non-parametric FD method |
|--------------------------|--------------------|--------------------------|
| Coworker complementarity | 0.0058*** | 0.0075 |
| Obs. (1000s) | 4,410 | 4,410 |

Notes. Regressions include FEs for employer; occupation-year; industry-year. Employer-clustered standard errors in parentheses. Observations weighted by the inverse employment share of the respective type and (rounded) coworker type cell. FD: finite differences.

Robustness checks: measuring coworker complementarity



• Types from non-parametric ranking algorithm instead of AKM-based

▶ Jump

Schooling as a non-wage measure of types

▶ Jump

Lagged types

▶ Jump

Small teams

▶ Jump

Movers

▶ Jump

· Non-parametric, finite-differences approximation

► Jump

Excluding managers

▶ lump

Log specification

Quantifying the model: approach



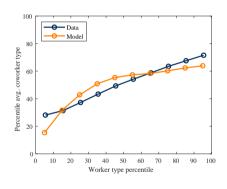
- Calibrate the model to the W German economy (2010-2017)
 - \bullet externally calibrated: discount rate, \bar{n} , bargaining power
 - offline estimation: job separation hazard
 - **3 online estimation** (indirect inference): meeting rate, unemp. flow benefit, production
 - \circ targets: $\hat{eta}_{ extsf{c}}$, total wage variance, avg. wage level, replacement rate, job finding rate
- Production function disciplined by \hat{eta}_c
 - \circ set χ to align regression estimates for this reduced-form moment using model-generated data with empirical estimate



• Macro moments of interest are untargeted: sorting, between-firm wage inequality

- Match coworker sorting patterns
 - $\circ \ \
 ho_{xx} = ext{O.43} \ ext{(vs. O.62 in data)}$
- Match between-firm wage inequality
 - o between-share 0.48 (vs. 0.57 in data)

Extensive validation of core mechanisms





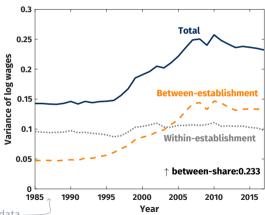
- Non-wage implications of complementarities: direction of EE moves
 - Δ coworker talent positively correlated with own talent
- Cross-sectional analysis: variation across occupations/industries
 - o task-based proxy for $\chi \uparrow \Rightarrow$ coworker talent complementarity \uparrow
 - \circ coworker talent complementarity $\uparrow \Rightarrow$ coworker talent sorting \uparrow
- · Work-in-progress analyses
 - $\circ \xi \uparrow \Rightarrow$ coworker talent complementarity \uparrow
 - ∘ $\chi \uparrow \Rightarrow$ wage dispersion conditional on $(x, x') \uparrow$
 - \circ wage effects induced by separation from a worker \leftrightarrow Jaeger-Heining (2022)





"the variance of firm [wages] explains an increasing share of total inequality in a range of countries"

[Song-Price-Guvenen-Bloom-von Wachter, 2019]

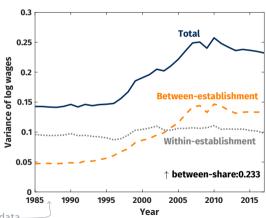


German matched employer-employee data—

Wage inequality has risen – and firms appear to play a key role



Applied question: what is/are the causal driver(s)? implications?



German matched employer-employee data—

Hypothesis

Hypothesis:

- the set of tasks any one worker can perform very well has narrowed: specialization ↑
- 2 coworker talent complementarities ↑
- 3 individuals of similar talent increasingly work together
- 4 this generates greater between-firm wage dispersion

Overview Theory Model Meets Data Firming Up Inequality Other applications Conclusion

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· Approach:

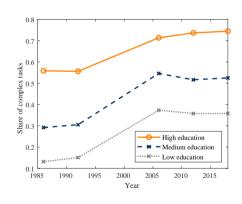
- o evaluate descriptive evidence
- o use structural model for counterfactual exercises quantifying this mechanism

Qualitative evidence suggests specialization (χ) \uparrow

▶ BIBB

► Occ. movements ► Science

- Task complexity ↑: "extensive margin" of \(\chi\)
 - DE longitudinal task survey
 - "complex": cognitive non-routine (e.g., organizing, researching)
- · Literature points in similar direction
 - o Grigsby ('23): transferability ↓ among occ w high/social/manual skills
 - o Jones ('09): burden of knowledge
 - Jovanovic-Rousseau ('08): proliferation of occupation codes

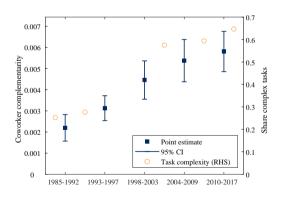


Coworker talent complementarity has strengthened



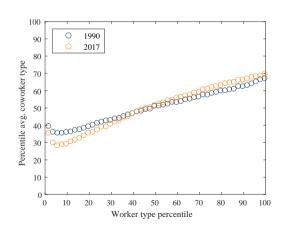
• **Theory:** specialization $(\chi) \uparrow$ is associated with coworker complementarity \uparrow

✓Coworker complementarity has more than doubled between 1985-1992 and 2010-2017



• **Theory:** complementarity ↑ is associated with talent sorting ↑

 Coworker matching has become more positively assortative

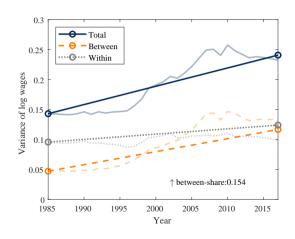


Key takeaways

- ↑ Skill specialization endogenously generates ↑ coworker talent complementarity
- ↑ Coworker complementarity leads to ↑ positive assortative coworker sorting
- 3 Complementarities can be disciplined using panel micro data on wages and matches
- Coworker complementarity has doubled since 1985 & talent sorting has intensified

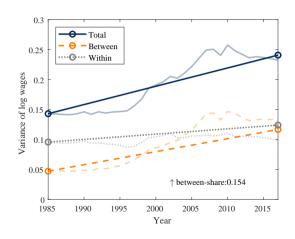
Model matches changes in firm-level wage distribution

- How quantitatively important are these developments in accounting for the firming up of inequality?
- Re-calibrate model: '85-'92
- Model replicates untargeted rise of between-share in data
 - 2/3 of ↑ between-share in data,
 ('85-'92)→('10-'17)



Model matches changes in firm-level wage distribution - why?

- Model replicates untargeted rise of between-share in data
- Reflects several parameters changing
 - $\circ \chi$: XX in '85-'92 (vs. XX in '10-'17)
 - job arrival & separation ↑
 - 0 ...



Complementarity \uparrow explains \approx 25-40% of observed between-share \uparrow

• **Q:** How much of \uparrow between-firm share of wage var. is due to \uparrow complementarities?

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- **Q:** How much of \uparrow between-firm share of wage var. is due to \uparrow complementarities?
- **Counterfactual:** between-firm share in 2010s absent $\chi \uparrow$ since '85-'92

Complementarity \uparrow explains \approx 25-40% of observed between-share \uparrow

- **Q:** How much of \uparrow between-firm share of wage var. is due to \uparrow complementarities?
- **Counterfactual:** between-firm share in 2010s absent $\chi \uparrow$ since '85-'92
- A: $\chi \uparrow$ accounts for 58% of model-predicted $\Delta \leftrightarrow \approx$ 38% of empirical Δ
- Robustness exercises: 25-40%

| | △ model | Implied % Δ model due to Δ parameter |
|-----------------------------------|---------|--|
| Model | 0.154 | - |
| Cf.: fix period-1 complementarity | 0.065 | 58 |

Overview of model robustness checks

- · Declining search frictions
- · Within-industry calibration
- Economy-wide vs. within-occupation analysis
- OJS
- · Increased talent dispersion

- ▶ Jump
- **▶** Jump
- ▶ Jump





Key takeaways

- **1** ↑ Skill specialization endogenously generates ↑ coworker talent complementarity
- ↑ Coworker complementarity leads to ↑ positive assortative coworker sorting
- Complementarities can be disciplined using panel micro data on wages and matches
- Coworker complementarity has doubled since 1985 & talent sorting has intensified
- f G This explains a substantial share of \uparrow between-firm wage inequality share

Other applications

Overview of extensions & other implications

• **Agg. productivity**: complementarities + search frictions ⇒ mismatch costs

▶ Jump

Productivity dispersion: talent sorting ⇒ firm-level productivity gaps



- "Coworker job ladders": workers, especially talented ones, tend to switch toward jobs with better coworkers
- Person-level inequality: segregation by itself need not cause greater person-level inequality, but it may in interaction with, e.g., product market frictions or fairness norms

Conclusion

Conclusion: firms form & organize teams – matters for macro

- Main idea: if individuals have specialized skills, firms assemble teams of 'complementary' coworkers, generating systematic sorting patterns
- This paper:
 - 1 task-based microfoundation for firm-level production fn. with skill complementarities
 - \Rightarrow specialization + team production \rightarrow complementarities
 - 2 measurement combining reduced-form micro evidence with model structure
 - ⇒ between-firm differences w/o assuming ex-ante heterogeneity
 - 3 structural explanation for the "firming up" of inequality
 - ⇒ role of increased complementarities

Thank You!



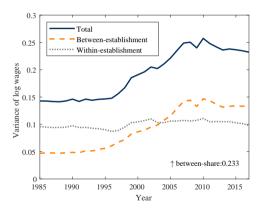
Extra Slides

Fact #1: ↑ between-firm share of wage inequality



- Large empirical literature: "firming up inequality" [e.g., Card et al., 2013; Song et al., 2019]
 - o "superstar firms" [e.g., Autor et al., 2020]
- Fact 1: ↑ wage inequality primarily due to between-component
- Robust pattern





Notes. Model-free statistical decomposition, where the "between" component corresponds to the person-weighted variance of est.-level avg. log wage.

Fact #2: talented workers increasingly collaborate



To what extent do talented workers tend to have talented coworkers?

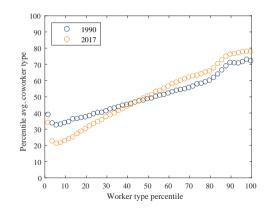
Fact 2: + assortative coworker sorting \(\)

o
$$\rho_{xx} = \operatorname{corr}(\hat{x}_i, \hat{x}_{-it})$$
: 0.43 ('85-'92) \nearrow 0.62 ('10-'17)

· Robust pattern

```
► Table ► Within-occ. nonlinear

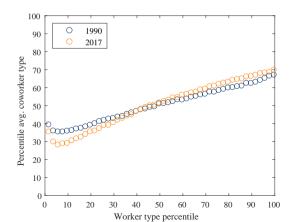
► Hakanson et al. (2021)
```



Evolution of coworker sorting: within-occupation ranking



 The most talented within each occupation – the best engineer, PA, economist, manager, ... – tend to work together, and increasingly so



Fact #3: increased education premium due to workplace effects



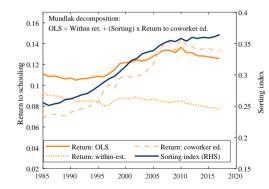
- Fact 3: increase in return to schooling is primarily due to workplace effects
- Mundlak decomposition of year-specific OLS return to schooling:

$$\beta_t^{\text{ols}} = \beta_t^{\text{within}} + \rho_t \times \beta_t^{\text{estab.}}$$

In
$$w_{it} = eta_{ ext{o}} + eta_{ ext{t}}^{ ext{within}} \mathsf{S}_{i} + eta_{ ext{t}}^{ ext{estab}} ar{\mathsf{S}}_{j(i,t),t} + e_{it}$$

where $\bar{S}_{j(i,t),t}$ is avg. years of schooling in establishment j of worker i in year t

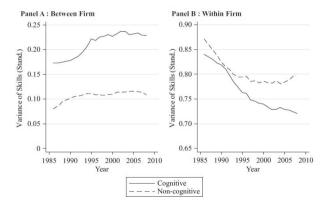
- $\mathbf{1}$ β_t^{within} : within-establishment return
- ${\it 2}$ ${\it \beta}_t^{\rm estab.}$: return to avg. establishment schooling



Notes. Plot of coefficients from year-by-year regressions of log wages.



• Direct measures of cognitive and non-cognitive skills across Swedish firms during 1986–2008, using test data from military enlistment



Workers increasingly tend to perform similar tasks across different jobs

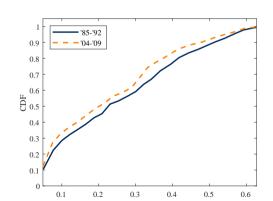


• \(\sqrt{Workers move to jobs with similar tasks, rather than randomly \)

Comparison

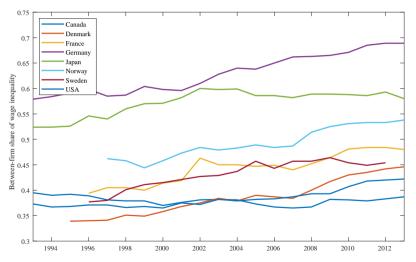
• Q: are workers becoming more likely to perform similar tasks across jobs, over time?

- **Yes:** distribution of moves in ('04-'09) is stochastically dominated by that in ('85-'92)
 - $\circ~$ uncond. average: 0.253 \rightarrow 0.227: 10% decline
- Robust in regression design
 - quantile regressions: ✓at different quantiles



Firming up inequality: cross-country evidence



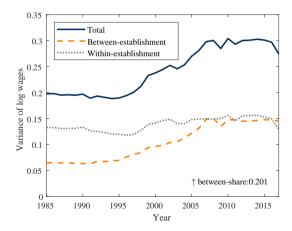


Notes. Data from Tomaskovic-Devey et al. (2020). Measures of earnings differ across countries and, for Germany, between T-D et al. and my study based on the SIEED.





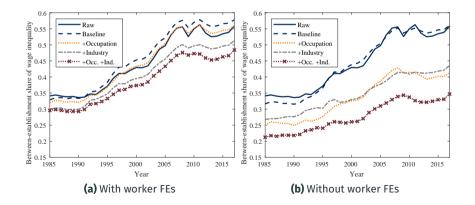
• Instead of considering all employers, restrict attention to "panel establishments"

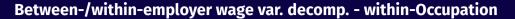


Between-/within-employer wage var. decomp. - alternative w-residuals

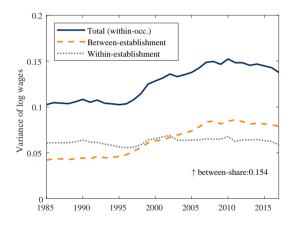


- "With worker FEs": regress $\ln \tilde{w}_{it} = \alpha_i + X'_{it}\hat{\beta} + \epsilon_{it}$, construct $\ln w_{it} = \ln(\tilde{w}_{it} X'_{it}\hat{\beta})$.
- "Without worker FEs": regress $\ln \tilde{w}_{it} = \alpha_0 + X_{it}' \hat{\beta} + \epsilon_{it}$, and consider residuals ϵ_{it}





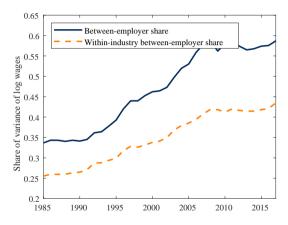




Between-/within-employer wage var. decomp. - within-Industry



► Main ► Other moments

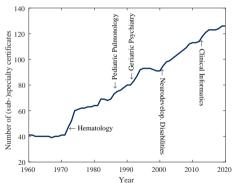


Notes. Based on 'baseline' residualized wages.

Examples: rising specialization

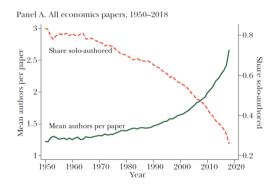


• Deepening medical specialization



Notes. Data from American Board of Medical Specialities. For each year, it shows the number of unique speciality or sub-speciality certificates that have been approved and issued at least once by that year and which are are still beine issued.

• Rise of research teams [Jones, 2021]



X-sectional validation (occ's): tasks \Rightarrow complementarity

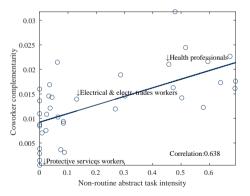


- ↑ Non-routine abstract task intensity
 ⇒ ↑ coworker talent complementarity
- $\hat{\beta}_{c}$ estimated separately for each occupation 0.03 0.025 Coworker complementarity ↓Health professionals[○] 0.02 Electrical & electr 0.005 Non-routine abstract task intensity: proxy for χ Correlation:0.638 Protective services worker 0.1 0.3 0.5 0.6 Non-routine abstract task intensity

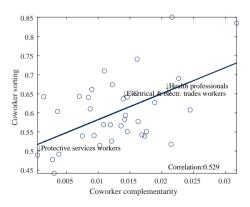
Notes. Quadros de Pessoal microdata. Analysis at ISCO-08-2d level.

X-sectional validation (occ's): tasks \Rightarrow complementarity \Rightarrow sorting

↑ Non-routine abstract task intensity
 ⇒ ↑ coworker talent complementarity



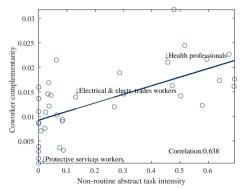
↑ Coworker talent complementarity
 ⇒ ↑ coworker sorting



Occupations: task complexity \Rightarrow complementarity \Rightarrow sorting

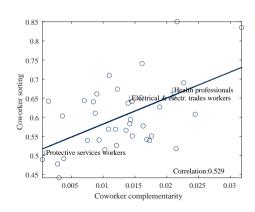


↑ Non-routine abstract task intensity
 ⇒ ↑ coworker wage complementarity



Notes. Quadros de Pessoal microdata. Horizontal axis indicates occupation's reliance on non-routine. abstract (NRA) tasks [Mihaylov and Tidens. 2019].

↑ Coworker wage complementarity
 ⇒ ↑ coworker sorting



Industries: coworker importance \Rightarrow complementarity \Rightarrow sorting



- ↑ Teamwork [Bombardini et al., 2012]
 ⇒ ↑ coworker wage complementarity
- 0.035 Coworker complementarity 0.02 Correlation:0 319 0.5 1.5 O*NET complementarity z-score

Notes. Horizontal axis measures the industry-level weighted mean score of an occupation-level index constructed from O*NET measuring the importance of: teamwork. impact on coworker output. Communication. and contact.

↑ Coworker wage complementarity
 ⇒ ↑ coworker sorting

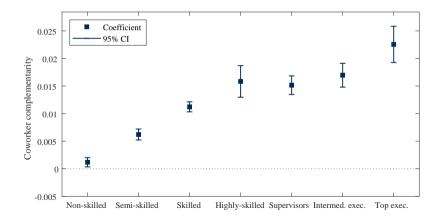


Notes. NACE-4-digit industries.

Hierarchies: complexity \Rightarrow complementarities



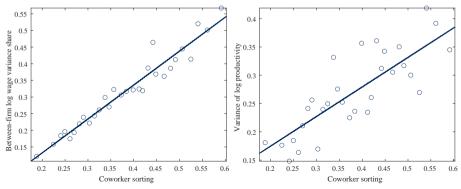
⇒ Coworker wage complementarities are (weakly) ↑ in the layer of a firm's hierarchy



Industries: coworker sorting ⇒ between-firm inequality



⇒ Measures of between-firm inequality in productivity and pay are increasing in the degree of coworker sorting at the industry-level.

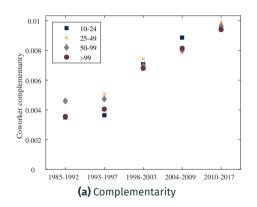


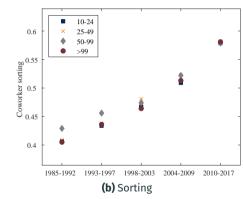
(a) Between-firm share of wage dispersion

(b) Productivity dispersion

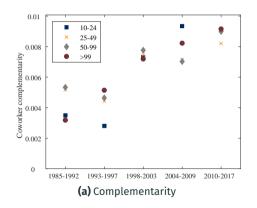
Coworker complementarity & sorting by team size

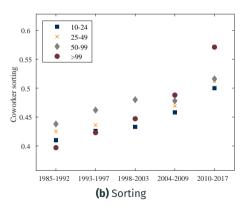




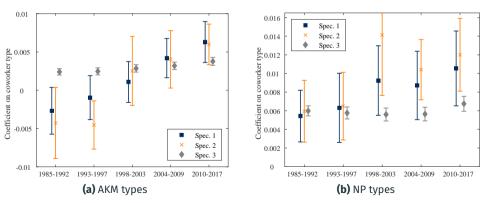


Coworker complementarity & sorting by team size - panel estab. only





$$\ln w_{it} = \beta_0 + \beta_1 \hat{x}_i + \frac{\beta_2}{2} \hat{x}_{-it} + \psi_{j(it)} + \nu_{o(i)t} + \xi_{s(i)t} + \epsilon_{it}$$



Notes. Specifications vary by ranking method – within-economy (spec. 1) vs. within-occupation (spec. 2/spec.3) and coworker group definition – establishment-year (spec. 1/spec.2) vs. establishment-occupation-year (spec.3).

Sorting & complementarity based on non-parametric ranking algorithm



• Instead of ranking workers based on AKM worker FEs, use non-param. ranking algo [Hagedorn et al., 2017]

| | Sorting | | Complementarities | | |
|-----------|-----------------|------|-------------------|---------|--|
| Period | Spec. 1 Spec. 2 | | Spec. 1 | Spec. 2 | |
| 1985-1992 | 0.47 | 0.38 | 0.001 | 0.000 | |
| 1993-1997 | 0.56 | 0.46 | 0.002 | 0.001 | |
| 1998-2003 | 0.60 | 0.48 | 0.004 | 0.002 | |
| 2004-2009 | 0.65 | 0.50 | 0.005 | 0.002 | |
| 2010-2017 | 0.68 | 0.51 | 0.005 | 0.004 | |

Notes. This table indicates, under the column "Sorting" the correlation between a worker's estimated type and that of their average coworker, separately for five sample periods. The column "Complementarities" indicates the point estimate of the regression coefficient \(\triangle _C\). Under "Specification 1" workers are ranked economy wide, while under "Specification 2" they are ranked within two-digit occupations. Worker rankings are based on the non-parametric method.

Coworker complementarity: excluding managers



- Concern regarding complementarity estimates: driven by managers?
 - o only managers benefit from team quality, e.g. via larger span of control
 - o the only coworkers that matter are managers

| Period | Baseline | Exclude as recipients | Exclude entirely |
|-----------|-----------|-----------------------|------------------|
| 1985-1992 | 0.0036*** | 0.0036*** | 0.0038*** |
| 1993-1997 | 0.0042*** | 0.0041*** | 0.0043*** |
| 1998-2003 | 0.0070*** | 0.0074*** | 0.0076*** |
| 2004-2009 | 0.0082*** | 0.0084*** | 0.0092*** |
| 2010-2017 | 0.0091*** | 0.0097*** | 0.0093*** |

Notes. Managed are defined based on KldB-1988-3d, as in Jarosch et al. (2023).

- Consider sub-samples of job movers, job movers with contiguous employment spells ($t \rightarrow t+1$), and job movers with non-contiguous E spells ($t \rightarrow t+s$, s>1)
- Caveat: annual panel given data size, no direct observation of U/N spells in SIEED

| Period | Baseline | All movers | Contig. E spells | Non-contig. E spells |
|-------------------------|-----------|------------|------------------|----------------------|
| 1985-1992 | 0.0043*** | 0.0043*** | 0.0045*** | 0.0039*** |
| 1993-1997 | 0.0049*** | 0.0052*** | 0.0052*** | 0.0051*** |
| 1998-2003 | 0.0078*** | 0.0085*** | 0.0083*** | 0.0082*** |
| 2004-2009 | 0.0090*** | 0.0107*** | 0.0104*** | 0.0102*** |
| 2010-2017 | 0.0088*** | 0.0103*** | 0.0101*** | 0.0090*** |
| Obs. in '10-'17 (1000s) | 4,410 | 538 | 355 | 375 |

Notes. Unweighted observations. Regressions include FEs for employer; occupation-year; industry-year. Employer-clustered standard errors in parentheses.

Coworker complementarity: finite-differences approximation



- Regression approach imposes strong functional form assumptions on approximated empirical wage function $\hat{w}(x|x')$
 - ofc, mirrored inside structural model when calibrating
- Alternative: construct non-parametric $\hat{w}(x|x')$, then use finite-difference methods to compute the cross-partial derivative (but w/o FE controls)

| Period | Regression | Non-parametric FD method | | |
|-----------|------------|--------------------------|--|--|
| 1985-1992 | 0.0036 | 0.0073 | | |
| 1993-1997 | 0.0042 | 0.0074 | | |
| 1998-2003 | 0.0070 | 0.0081 | | |
| 2004-2009 | 0.0082 | 0.0120 | | |
| 2010-2017 | 0.0091 | 0.0098 | | |

- Concern with both regression approach and non-parametric FD approach: mechanical relationship between wages ("LHS") and (within-period time-invariant) worker types, which are estimated from wages themselves ("RHS")
- Robustness check #1: years of schooling as type measure



- Robustness check #2: assign to each individual i in periods $p \in \{2,3,4,5\}$ the FE estimated for i in period p-1; re-compute worker deciles and average coworker types, \hat{x}_i^{p-1} and $\hat{x}_{-it}^{p-1} = (|S_{-it}|)^{-1} \sum_{k \in \mathcal{S}} \hat{x}_k^{p-1}$; re-estimate wage regression
- Results (see paper): magnitude of estimated $\hat{\beta}_c$ around 50% smaller when using lagged types, but evolution over time similar to baseline

Complementarity estimates using years of schooling



| | '85-'92 | '93-'97 | '98-'03 | '04-'09 | '10-'17 |
|----------------|-----------|-----------|-----------|-----------|-----------|
| Interaction | o.oo63*** | o.oo6o*** | 0.0099*** | 0.0112*** | 0.0129*** |
| | (o.ooo8) | (o.ooo7) | (0.0008) | (0.0007) | (0.0009) |
| Obs. (1000s) | 3,613 | 2,508 | 2,694 | 3,836 | 4,376 |
| R ² | 0.5033 | 0.5451 | 0.5746 | 0.6330 | 0.6425 |

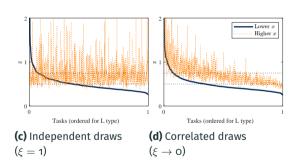
Notes. Dependent variable is the wage level over the year-specific average wage. Independent variables are a constant, years of schooling, coworker years of schooling, and the interaction between those two terms. All regressions include industry-year, occupation-year and employer fixed effects. Employer-clustered standard errors in parentheses. Observations are unweighted. The sample is unchanged from the main text, except that 96,517 observations with missing years of schooling are dropped. Observation count rounded to 1000ss.

| | | Baseline | | | | Within-in | dustry a | ıvg. |
|---------------|--------------------|---|------------|-------------------|--------------------|---|------------|-------------------|
| Sample Period | $\sigma_{\rm W}^2$ | $\sigma_{\bar{\mathrm{W}}}^{\mathrm{2}}/\sigma_{\mathrm{W}}^{\mathrm{2}}$ | $ ho_{XX}$ | $\hat{eta}_{f c}$ | $\sigma_{\rm W}^2$ | $\sigma_{\bar{\mathrm{W}}}^{\mathrm{2}}/\sigma_{\mathrm{W}}^{\mathrm{2}}$ | $ ho_{xx}$ | $\hat{eta}_{f c}$ |
| 1 | 0.143 | 0.337 | 0.427 | 0.0036 | 0.125 | 0.249 | 0.333 | 0.00283 |
| 2 | 0.148 | 0.391 | 0.458 | 0.0042 | 0.125 | 0.288 | 0.351 | 0.00342 |
| 3 | 0.191 | 0.456 | 0.495 | 0.0070 | 0.150 | 0.324 | 0.369 | 0.00585 |
| 4 | 0.234 | 0.547 | 0.547 | 0.0082 | 0.168 | 0.388 | 0.405 | 0.00738 |
| 5 | 0.241 | 0.568 | 0.617 | 0.0091 | 0.171 | 0.412 | 0.464 | 0.00823 |

Notes. Within-industry avg. is person-year weighted average across OECD STAN-A38 (2-digit) industries.

Illustration: specialization parameter Peack

 ξ: controls correlation of coworkers' task-specific productivities





- **1** Derive $G(p) := \Pr{\{\tilde{\lambda}(\tau) \leq p\}}$ given $G_i(p) := \Pr{\{\lambda_i(\tau) \leq p\}}$, using FOC and max-stability property
- ② Use G(p) + standard CES shadow price index to solve (int. by sub.) for

$$\lambda = \left(\int_{\mathcal{T}} \tilde{\lambda}(\tau)^{1-\eta} d\tau \right)^{\frac{1}{1-\eta}}.$$
 (5)

- 3 Use G(p) and $G_i(p)$ to derive probability that i produces some task τ , which by LLN (continuum assumption!) is equal to share of tasks produced, π_i
- **Q** Relate λ_i^L to value of all tasks produced by the worker, $\lambda_i^L = \pi_i \lambda Y$
- **6** Normalize $\lambda = 1$, then algebra yields $Y = f(x_1, ..., x_n; \chi)$

Lemma

Lemma: Lemma

Implied task share and shadow-cost index equal

$$\pi_{i} = \frac{\left(x_{i}/\lambda_{i}^{L}\right)^{\frac{1}{\chi\xi}}}{\sum_{k=1}^{n}\left(x_{i}/\lambda_{i}^{L}\right)^{\frac{1}{\chi\xi}}} \quad x_{i} \lambda = \left(\sum_{i=1}^{n}\left(\frac{x_{i}}{\lambda_{i}^{L}}\right)^{\frac{1}{\chi\xi}}\right)^{-\chi\xi}$$



- · What is the intuition for these properties?
- Solution of firm's mini-planner problem implies:
 - **1** Complete division of labor, with tasks assigned by comparative advantage

$$\circ \ \textit{i's task set} \ \mathcal{T}_{\textit{i}} = \left\{ \tau \in \mathcal{T} : \frac{\mathsf{z}_{\textit{i}}(\tau)}{\lambda_{\textit{i}}^{\textit{L}}} \geq \mathsf{max}_{k \neq \textit{i}} \, \frac{\mathsf{z}_{\textit{k}}(\tau)}{\lambda_{\textit{k}}^{\textit{L}}} \right\}$$

- o classic source of efficiency gains
- **2** i's share of tasks \uparrow in i's talent, \downarrow in coworkers' talent

$$\circ$$
 i's task share $\pi_i = (x_i^{\frac{1}{1+\chi\xi}})(\sum_{k=1}^n (x_k)^{\frac{1}{1+\chi\xi}})^{-1}$



- Suppose that $x_i > x_j$. Then
 - **1** i performs a strictly larger share of tasks than j for $\chi < \infty$



Intuition: comparative statics for task shares

- Suppose that $x_i > x_j$. Then
 - **1** i performs a strictly larger share of tasks than j for $\chi < \infty$
 - $oldsymbol{2}$ the difference in task shares is decreasing in χ



⇒ Greater skill specialization implies a larger share of tasks is performed by relatively less talented team members – more talented coworkers can't easily compensate

Surplus max. determines which worker types a firm w/ worker x hires \bigcirc Main



• Joint value of firm with worker x, $\Omega(x)$, satisfies:

$$\rho\Omega(\mathbf{x}) = f(\mathbf{x}) + \delta\big[-\Omega(\mathbf{x}) + V_u(\mathbf{x}) + V_{f,o}\big] + (1 - \omega)\lambda_{v,u} \int \frac{d_u(\tilde{\mathbf{x}}')}{u} \max\{S(\tilde{\mathbf{x}}'|\mathbf{x}), o\}d\tilde{\mathbf{x}}'$$

- o $V_{\mu}(x)$: value for unemp. worker; $V_{f,o}$: value for vacant firm; S(x): surplus from zero-worker firm hiring x
- o $d_u(x)$: density of unemployed workers of type x; $u = \int d_u(x) dx$
- \circ ω : worker bargaining wgt; δ : sep. rate; $\lambda_{V,U}$: rate of vacancy meeting unmatched worker

Surplus max. determines which worker types a firm w/ worker x hires Main



• Joint value of firm with worker x, $\Omega(x)$, satisfies:

$$\rho\Omega(\mathbf{x}) = f(\mathbf{x}) + \delta\big[-\Omega(\mathbf{x}) + V_u(\mathbf{x}) + V_{f.o}\big] + (1 - \omega)\lambda_{v.u} \int \frac{d_u(\tilde{\mathbf{x}}')}{u} \max\{S(\tilde{\mathbf{x}}'|\mathbf{x}), o\}d\tilde{\mathbf{x}}'$$

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- o $d_u(x)$: density of unemployed workers of type x; $u = \int d_u(x) dx$
- \circ ω : worker bargaining wgt; δ : sep. rate; $\lambda_{V,U}$: rate of vacancy meeting unmatched worker
- Surplus S(x|x') reflects complementarities and hiring decisions reflect surplus

$$S(x|x')(\rho + 2\delta) = f(x,x') - \rho(V_u(x) + V_u(x') + V_{f.o}) + \delta S(x) - (\rho + \delta)S(x')$$
$$h(x|x') = \mathbf{1}\{S(x|x') > 0\}$$

HJB: unmatched



· Unmatched firm:

$$\rho V_{f.o} = (1 - \omega) \lambda_{v.u} \int \frac{d_u(x)}{u} S(x)^+ dx$$
 (6)

· Unmatched worker:

$$\rho V_u(x) = b(x) + \lambda_u \omega \left[\int \frac{d_{f.0}}{v} S(x)^+ + \int \frac{d_{m.1}(\tilde{x}')}{v} S(x|\tilde{x}')^+ d\tilde{x}' \right]$$
 (7)

HJB: surpluses

• Surplus of coalition of firm with worker x

$$(\rho + \delta)S(x) = f(x) - \rho(V_u(x) + V_{f.o}) + \lambda_{v.u}(1 - \omega) \int \frac{d_u(\tilde{x}')}{u} S(\tilde{x}'|x)^+ d\tilde{x}'$$
 (8)

• Surplus from adding x to x'

$$S(x|x')(\rho+2\delta) = f(x,x') - \rho(V_u(x) + V_u(x') + V_{f,o}) + \delta S(x) - (\rho+\delta)S(x')$$
 (9)

KFE: unemployed

$$\delta\bigg(d_{m.1}(x) + \int d_{m.2}(x,\tilde{x}')d\tilde{x}'\bigg) = d_u(x)\lambda_u\bigg(\int \frac{d_{f.o}}{v}h(x,\tilde{y}) + \int \frac{d_{m.2}(\tilde{x}')}{v}h(x|\tilde{x}')d\tilde{x}'\bigg). \tag{10}$$

KFE: one-worker matches

$$d_{m.1}(x)\bigg(\delta + \lambda_{v.u} \int \frac{d_u(\tilde{x}')}{u} h(\tilde{x}'|x) d\tilde{x}'\bigg) = d_u(x) \lambda_u \frac{d_{f.o}}{v} h(x) + \delta \int d_{m.2}(x, \tilde{x}') d\tilde{x}'. \tag{11}$$

KFE: two-worker matches

$$2\delta d_{m.2}(x,x') = d_u(x)\lambda_u \frac{d_{m.1}(x')}{v}h(x|x') + d_u(x')\lambda_u \frac{d_{m.1}(x)}{v}h(x'|x). \tag{12}$$

Measurement: a useful identification result



- **Q:** How to quantify $\frac{\partial^2 f(x,x')}{\partial x \partial x'}$?
- Proposition: production complementarities are proportional to wage compl.
- **Proof sketch:** wage level for worker x with coworker x'

▶ Wage eq.

$$w(x|x',\xi) = \omega f(x,x',\xi) + g(x) - h(x')$$

where $g:[0,1]\to\mathbb{R}$ and $h:[0,1]\to\mathbb{R}$ are strictly increasing

⇒ outside options are separable: affect level of wage but not the cross-partial

Mapping model to data: coworker types



• Defining $S_{-it} = \{k : j(kt) = j(it), k \neq i\}$ as the set of *i*'s coworkers in year *t*, compute the average type of *i*'s coworkers in year *t* as $\hat{x}_{-it} = \frac{1}{|S_{-it}|} \sum_{k \in S_{-it}} \hat{x}_k$.

· Coworker group:

- o alternative: same establishment-occupation-year cell
- but CC arise precisely when workers are differentiated in their task-specific productivities

Averaging step:

- o equally-weighted averaging ignores non-linearity in coworker aggregation
- paper: show using non-linear averaging method that baseline results in bias, but it's minor in magnitude
- **Firm size variation:** averaging ensures that a single move will induce a smaller change in the *average* coworker quality in a large team than in a small one

Mapping model to data: identification strategy for χ



- **Literature:** complementarities primarily between workers and firms usually inferred indirectly from sorting patterns
 - o exception: Hagedorn-Law-Manovskii (2017)
- This paper: directly measure coworker complementarity in the data, recover χ structurally given $\gamma=\frac{\chi}{\chi+1}$
- Paper does $\it not$ use microfoundation itself to measure χ , respectively γ
- Experiment: fit a (truncated) Fréchet distribution to Grigsby's (2023) non-parametric estimates of the multi-dimensional skill dist. estimated from CPS data
 - $\circ \ \ {
 m recover} \ \gamma = {
 m o.84} \ {
 m for 2006} \ {
 m but} \ {\it very noisy estimates}$



• Ongoing work: use the extended microfoundation to identify χ



Direct estimation of χ : proof of concept

- Grigsby (2023): only paper that provides a cardinal measure of skill task-specificity
- Evidence on time trends are qualitatively consistent with "specialization hypothesis": cross-type average of within-type variance across specific skills grew by nearly 50% b/w 1980s and 2000s & skill transferability has declined amongst high-skill occupations
- His operationalization of worker types and tasks does not directly map onto my model (no identifying assumption; coarse occupational skills; US vs DE data)
- **Proof of concept:** but suppose we just take those data, extract moments capturing average within-worker cross-task efficiency dispersion, fit a (truncated) Fréchet, recover $\gamma = \frac{\chi}{1+\chi}$
 - \Rightarrow / γ similar to structural estimation result based on evidence from wage CC

Semi-structural back-of-envelope calculation for γ



- Structurally recover $\gamma \frac{\chi}{\chi+1}$ by estimating $\frac{\partial^2 w(x|x')}{\partial x \partial x'}$ in the data, which was shown to be proportional to $\frac{\partial^2 f(x,x')}{\partial x \partial x'}$
- But how is $\frac{\partial^2 f(\mathbf{x}, \mathbf{x}')}{\partial \mathbf{x} \partial \mathbf{x}'}$ related to γ ?
- Definitionally, $\gamma = (ff_{ij})/(f_if_j)$ for any $i \neq j$
- Can we avoid full structural model? \Rightarrow If have measures not only of f_{ij} but also output f and marginal products f_i
- Suppose, for any x and x', we use wages to back out marginal products –
 competitive wage determination rather bargaining! and recover output from sum
 of wages divided by labor share
- Find $\gamma \approx$ 0.79 very close to structural estimate!

Overview of validation exercises: direction EE transitions & cross-section



- 2 additional types of validation exercise:
 - ✓EE transitions reallocate workers to more + assortative matches

▶ Details

- o do model-implied relationships also hold in cross-section?
 - 1 $\chi \uparrow \Rightarrow$ coworker complementarity \uparrow
 - 2 coworker complementarity $\uparrow \Rightarrow$ + assortative matching \uparrow

can test predictions because
we have measures of complementarity!

- Implementation of cross-sectional exercises: rich Portuguese micro data
 - o universe of private-sector actors, employer-employee data & income statements
- · Cross-sectional exercises:
 - ✓ Hierarchies
 - ✓ Industries
 - ✓ Occupations



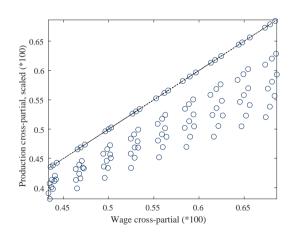






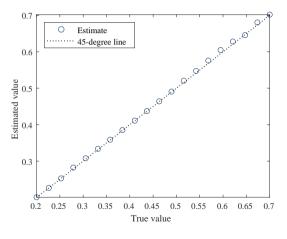


- Solve model for many combinations of χ , λ_e and b
- Compare FD approx of $f_{xx'}(x, x')$, scaled by ω , and $w_{xx'}(x|x')$
- Main parameter driving wedge: λ_e



Monte Carlo study





• Production complementarities imply coworker sorting matters for agg productivity

$$\circ f(\mathbf{X}_1,\cdots,\mathbf{X}_n) = n^{1+\chi\xi} \times \left(\frac{1}{n} \sum_{i=1}^n (\mathbf{X}_i)^{\frac{1}{1+\chi\xi}}\right)^{1+\chi\xi}$$

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- Search frictions induce misallocation \sim coworker **mismatch**

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- Quantify mismatch costs: compare eqm outcome vs to productivity under pure PAM
 ⇒ 2010s gap: 2.05%, similar for earlier periods

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- Search frictions induce misallocation \sim coworker mismatch
- Quantify mismatch costs: compare eqm outcome vs to productivity under pure PAM
 ⇒ 2010s gap: 2.05%, similar for earlier periods
- Trends: ↑ talent sorting limited ↑ in mismatch costs given χ ↑
 ⇒ no-reallocation counterfactual: productivity gap 4.65%

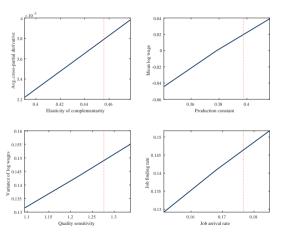
Parameterization, including estimation results (2010s)



| Parameter | Description | Targeted moment | Value | m | ĥ |
|-------------|--------------------------|-------------------|-------|--------|--------|
| χ | Specialization | \hat{eta}_{c} | 0.67 | 0.0058 | 0.0058 |
| a_{o} | Production, constant | Avg. wage (norm.) | 0.29 | 1 | 1 |
| a_1 | Production, scale | Var. log wage | 1.71 | 0.241 | 0.241 |
| b_1 | Replacement rate, scale | Replacement rate | 0.60 | 0.63 | 0.63 |
| δ | Separation hazard | Job loss rate | 0.008 | 0.008 | 0.008 |
| λ_u | Meeting hazard | Job finding rate | 0.22 | 0.162 | 0.162 |
| ρ | Discount rate | External | 0.008 | | |
| ω | Worker bargaining weight | External | 0.50 | | |
| п | Effective team size | External | 25 | | |

Identification validation exercise 1





Notes. This figure plots the targeted moment against the relevant parameter, holding constant all other parameters.

Identification validation exercise 2



Notes. This figure plots the distance function $\mathcal{G}(\psi_i, \psi_{-i}^*)$ when varying a given parameter ψ_i around the estimated value ψ_i^* . The remaining parameters are allowed to adjust to minimize \mathcal{G} .

What about shifts in the talent distribution?



- **Q**: could \uparrow economy-wide x-dispersion, σ_x , also explain \uparrow firm-level inequality?
 - $\circ~$ Kremer-Maskin (1996): as the dispersion of skills (talent) \uparrow in the economy, the relative dispersion of talent within firms \downarrow
- **Obs. #1:** Hakanson et al. (2021): direct skill measures (from military enlistment tests) point to Flynn effect but not not to increased dispersion (in cognitive test scores)
- Obs. #2: model exercise
 - o method: instead of rank interpretation, i.e. $x \sim U$, we separately parameterize σ_x by assuming a $\mathcal N$ distribution
 - finding: $\sigma_x \uparrow \Rightarrow \rho_{xx} \uparrow$, $\sigma_w \uparrow$, and $\sigma_{\bar{w}}/\sigma_w \uparrow$ but no measured increase in $\hat{\beta}_c$
- **Conclusion:** empirically unclear whether $\sigma_x \uparrow$, and if so, this would not explain the observed increase in coworker complementarity, i.e. latter is a distinct channel

Training policies in a team production context (w-i-p)



- "Training policies" ~ non-parametric perturbations of the talent distribution
 - o left-tail intervention: give everyone in 1st decile productivity of those in 2nd decile
 - o right-tail intervention: give everyone in 9th decile productivity of those in 10th decile
- Team production: effect of "training policies" partially via coworker spillovers!
- Relative effectiveness of left-tail vs right-tail intervention:
 - 1 the stronger are coworker complementarities, the relatively greater are the realized productivity gains from a left-tail intervention, b/c low-x tend to be weak links
 - but raising the productivity of coworkers of workers with high productive potential generates greater gains – and with sorting, those coworkers are themselves high-x
- Equilibrium: relative effectiveness of left-/right-tail training depends on both forces
 - tentative quantitative finding: right-tail intervention boosts average productivity by more but left-tail training also lowers inequality

The effect of declining search frictions



- ↓ search frictions could also explain ↑ coworker sorting
- Job arrival & separation rates estimated to ↑ from p1 to p2
- Counterfactual analysis: explains 6% of model-implied ↑ in between-employer share of wage variance

| | △ model | Implied % Δ model due to Δ parameter |
|--------------------------------------|---------|--|
| Model 1: baseline | 0.159 | - |
| Cf. a: fix period-1 complementarity | 0.065 | 59 |
| Cf. b: fix period-1 search frictions | 0.150 | 6 |

Outsourcing & within-occupation ranking analysis



- · Concern: confounding shifts in labor boundary of firm, e.g. outsourcing
- · Address this concern in multiple steps:
 - empirically rank workers within occupation ("good engineer vs. mediocre engineer")
 - ② empirically re-estimate coworker sorting & complementarity (lower but similar ↑)
 - 3 re-estimate model for both periods & re-do counterfactual exercises
- Result: qualitatively & quantitatively similar findings

| | △ model | Implied % Δ model due to Δ parameter | |
|------------------------------|---------|--|--|
| Model 2: within-occ. ranking | 0.198 | - | |
| Cf. a: fix period-1 comp. | 0.076 | 61.47 | |

Robustness: model with OJS - brief overview



- **Baseline:** no job-to-job transitions but is main story robust when workers can switch to better job after accepting job out of unemployment?
 - two opposing effects from increased complementarities
- **Extension:** employed workers also meet vacancies at Poisson rate λ_e
- Main findings:
 - o better fit to empirical sorting patterns in cross-section



- contribution of ↑ complementarities to ↑ firm-level wage inequality slightly smaller, more attributed to ↑ labor market transitions
 - conservative estimates (endogenous search effort, forward-looking wage specification)
- Opens door to thinking about coworker complementarities and **job ladders**



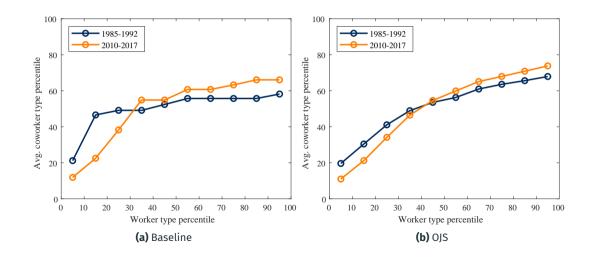
Extension: model with OJS



- Baseline model abstracted from OJS
 - o transparent trade-off, connection to analytical results
- Consider extension to OJS: employed worker meet vacancies at Poisson rate λ_e
 - wages both off and on the job are continuously renegotiated under Nash bargaining, with unemployment serving as the outside option [cf. di Addario et al., 2021]
 - $\circ~$ re-estimate, with empirical labor market flows disciplining λ_e
- Qualitative question: is coworker sorting outcome robust, even if workers can switch to better job after accepting job out of unemployment?
- Analyses:
 - coworker sorting patterns & changes
 - additional model validation: direction of EE flows in model & data

Model-implied coworker sorting patterns: without and with OJS







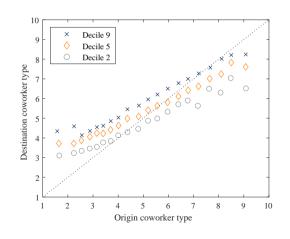
- Theoretical prediction: EE transitions move workers in surplus-maximizing direction $\Rightarrow \Delta \hat{x}_{-it} = \hat{x}_{-i,t} \hat{x}_{-i,t-1}$ should be *positively* correlated with \hat{x}_i
 - o $h_{2.1}(x, x''|x') = 1$ worker x in a two-worker firm with coworker x'' would move to an employer that currently has one employee of type x' if S(x|x') S(x|x'') > 0
- **Empirical analysis**: use SIEED *spell* data to create worker-originMonth-destinationMonth-originJob-destinationJob panel, with information on characteristics of origin and destination job
 - o subsample period 2008-2013 (huge panel at monthly frequency)
 - o count as "EE" if employer change between two adjacent months
- **Regression analysis:** regress $\Delta \hat{x}_{-it}$, scaled by std. σ_{Δ} of coworker quality changes, on *own* type and *origin* coworker type

$$\frac{\Delta \hat{\mathbf{x}}_{-it}}{\sigma_{\mathbf{A}}} = \beta_{\mathbf{O}} + \frac{\beta_{\mathbf{1}}}{\beta_{\mathbf{1}}} \hat{\mathbf{x}}_{i} + \beta_{\mathbf{2}} \hat{\mathbf{x}}_{-i,t-1} + \epsilon_{it}$$

Empirical coworker sorting changes due to EE moves



- EE transitions push toward greater coworker sorting: for any given origin, higher x-workers move to workplaces with better coworkers than lower-x workers do
- But in data EE transitions "move up" low types more than theory predicts
- "Coworker job ladder" with both absolute and type-specific dimension?
- **Next:** change in the job ladder [e.g., Haltiwanger-Spetzler, 2021]



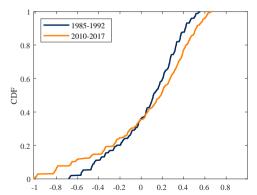
Evidence that EE increasingly reallocate toward PAM: in data & model

| | Data | | Model | |
|-------------------------|---------------------------------|--------------------------------|----------|----------|
| Change in coworker type | '85-'92 | '10-'17 | Period-1 | Period-2 |
| Own type | 0.0883 *** (0.000799) | 0.118 *** (0.000918) | 0.214 | 0.270 |
| Controls | Year FEs, Origin | Year FEs, Origin | Origin | Origin |
| N | 196,098 | 282,718 | ∞ | ∞ |
| adj. R² | 0.284 | 0.204 | | |

Table 1: Change in coworker type due to EE moves positively related to own type – increasingly so

Notes. For the data columns, individual-level clustered standard errors are given in parentheses. Model counterparts are computed simulation-free in population. Dependent variable is scaled throughout by the standard deviation of the change in coworker type.

• Firm dynamics literature: increased firm-level productivity dispersion [Autor et al., 2020; de Ridder, 2024]; correlated with wage & talent dispersion [Berlingieri et al., 2017; Sorkin-Wallskoa, 2020]

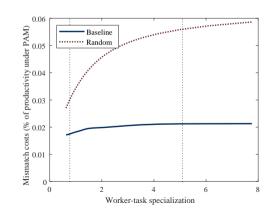






"The benefits of the division of labor are limited by the functioning of the labor market"

- Microfoundation: $\uparrow \chi \Rightarrow \uparrow$ efficiency benefit from teamwork but also \uparrow mismatch costs
- Q: how does the gap to potential vary depending on labor market structure?
- A: under random sorting, productivity gap due to misallocation \uparrow more sharply as $\chi\uparrow$
- Outside model: severe labor mkt frictions (e.g., dev'ing countries [Donovan et al., 2023]) may inhibit specialization [cf. Atencio et al.,



Implications for overall inequality?



- Coworker complementarities do not necessarily \(\) variance of person-level wages
 - (un-)surprising? Varriance decomposition perspective vs. common intuition [Kremer, 1993]
 - o (i) **reallocation effect**, (ii) valuation effect, (iii) outside option effect
- Several mechanisms though through which \uparrow sorting could \uparrow person-level inequality
 - 1 regulation or norms that lead to within-firm wage compression [Akerlof-Yellen, 1990]
 - 2 coworker learning [Jarosch et al., 2021; HLMP, 2023]
 - 3 increasing returns to labor quality [Kremer, 1993]

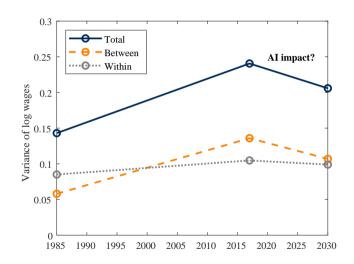
Implications for AI impact on labor markets (1): overview



- Literature moving toward richer way of thinking about tech. change. Here: shifts in across-worker productivity differences and in specialization/interdependence
- Example: AI early/conjectural evidence on impact of LLMs etc
 - o absolute adv. less important ("leveller") [e.g., Brynjolfsson et al., 2023]
 - o comparative adv. less important (e.g., can interpret medical imaging w/o radiologist)
- Illustrative model counterfactual: relative to the 2010s
 - everyone's productivity ↑ by equivalent to 20% of lowest type's productivity so in proportional terms, lower-x benefit more
 - 2 coworker complementarity ↓ by 20%

Implications for AI impact on labor markets (2): illustrative exercise

- Model prediction: AI could lead to reversal of historical trends $\sigma_{\rm W}^2\downarrow$, $\sigma_{\rm W}^2\downarrow$, $\rho_{\rm XX}\downarrow$ along side a productivity boom
 - o driven by ↓ firm-level wage inequality



Implications for AI impact on labor markets (3): conjectures



- Literature moving toward richer way of thinking about tech. change. Here: shifts in across-worker productivity differences and in specialization/interdependence
- Example: AI early/conjectural evidence on impact of LLMs etc
 - o absolute adv. less important? ("leveller") [e.g., Brynjolfsson et al., 2023]
 - o comparative adv. less important? (e.g., can interpret medical imaging w/o radiologist)
- Illustrative model counterfactual: AI could lead to reversal of historical trends...
- ...and perhaps also...
 - $oldsymbol{1}$...**flatter organizations**, with managerial span of control \uparrow
 - ② ...↓ barriers to entry for self-employment/start-ups
 - 3 ...easier job transitions & shorter training durations

Within-industry calibration: overview



- Baseline: calibration and evaluation based on economy-wide moments
- But the model does not incorporate between-industry differences in, e.g., production technology
- Alternative considered here: target $\hat{\beta}_c$ and σ_w^2 computed as within-industry average, evaluate against within-industry trends
 - keep other targets (e.g., job separation) as before

Within-industry calibration: model fit & counterfactual

 Counterfactual: χ↑ explains 83% of model-implied↑in between-share

