

$$Q1 (a) M(t) = E(e^{tx}) = \sum_{x=0}^{\infty} e^{tx} \frac{\lambda^x e^{-\lambda}}{x!}$$

$$= e^{-\lambda} \sum_{x=0}^{\infty} \frac{(\lambda e^t)^x}{x!}$$

$$= e^{-\lambda} e^{\lambda(e^t - 1)}$$

$$M'(t) = \lambda e^t e^{\lambda(e^t - 1)} \text{ So } EX_1 = M'(0) = \lambda$$

$$\lambda = \frac{1}{n} \sum_{i=1}^n X_i = \bar{X} \rightarrow \hat{\lambda}_{\text{mom}} = \bar{X}$$

$$E[\hat{\lambda}_{\text{mom}}] = E(\bar{X}) = \lambda_0$$

$$(b) \text{Bias}(\hat{\lambda}_{\text{mom}}) = 0$$

$$\text{Var}(\hat{\lambda}_{\text{mom}}) = \text{Var}\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{\lambda_0}{n}$$

$$\text{MSE}(\hat{\lambda}_{\text{mom}}) = \text{Bias}^2 + \text{Var} = \frac{\lambda_0}{n}$$

$$(c) P(X_1 > 5) = 1 - P(X_1 \leq 5)$$

$$= 1 - \sum_{k=0}^5 \frac{\lambda_0^k}{k!} e^{-\lambda_0}$$

$$(d) L(\lambda; X_1 \dots X_n) = \prod_{i=1}^n \frac{\lambda^{x_i} e^{-\lambda}}{x_i!}$$

$$\log L(\lambda; X_1 \dots X_n) = \sum_{i=1}^n (x_i \log \lambda - \lambda - \log(x_i!))$$

$$\frac{d}{d\lambda} \log L(\lambda) = \sum_{i=1}^n \left( \frac{x_i}{\lambda} - 1 \right) = 0, \quad \frac{d^2}{d\lambda^2} \log L(\lambda) = \sum_{i=1}^n \left( -\frac{x_i}{\lambda^2} \right) < 0$$

$$\text{So } \hat{\lambda}_{\text{MLE}} = \frac{1}{n} \sum_{i=1}^n X_i = \bar{X}$$

$$(e) \text{ let } \theta = \log \lambda_0 \rightarrow \lambda_0 = e^\theta$$

$$\log L(\theta) = \sum_{i=1}^n (x_i \theta - e^\theta - \log(x_i!))$$

$$\frac{d}{d\theta} \log L(\theta) = \sum_{i=1}^n (x_i - e^\theta), \quad \frac{d^2}{d\theta^2} \log L(\theta) = \sum_{i=1}^n (-e^\theta) < 0$$

$$\hat{\theta}_{\text{MLE}} = \log\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \log(\bar{X})$$



(+)  $\hat{\lambda}_{MLE} = \bar{X}$  apply the law of the large number,  $\forall \lambda_0, E(X_i) = \lambda_0$

$\bar{X}_n \xrightarrow{P} \lambda_0$  as  $n \rightarrow \infty$  Since  $\hat{\lambda}_{MLE} = \bar{X}_n$ ,

$\bar{X}_n$  converges in probability to  $\lambda$

So,  $\hat{\lambda}_{MLE} = \bar{X}_n \xrightarrow{P} \lambda_0$  as  $n \rightarrow \infty$

So  $\hat{\lambda}_{MLE}$  for parameter  $\lambda_0$  of Poisson distribution is consistent



Q2.

Date

Page

$$(a) \textcircled{1} E(X) = \lambda \quad \bar{X} = \frac{1}{n} \sum_{i=1}^n X_i \quad E(X) = \bar{X} \rightarrow \lambda = \bar{X} \quad \lambda_{\text{mom}} = \frac{1}{\bar{X}}$$

$$\textcircled{2} E(\hat{\lambda}) = E(\bar{X}) \quad E(X_i) = \frac{a_0}{b_0}$$

$$E(\bar{X}) = E\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n} \sum_{i=1}^n E(X_i) = \frac{a_0}{b_0}$$

$$\rightarrow E(\hat{\lambda}) = \frac{a_0}{b_0}$$

$$(b) L(\lambda) = \prod_{i=1}^n \lambda e^{-\lambda X_i} = \lambda^n e^{-\lambda \sum_{i=1}^n X_i}$$

$$\log L(\lambda) = n \log(\lambda) - \lambda \sum_{i=1}^n X_i$$

$$\frac{d}{d\lambda} \log L(\lambda) = \frac{n}{\lambda} - \sum_{i=1}^n X_i = 0$$

$$\rightarrow \hat{\lambda}_{MLE} = \frac{n}{\sum_{i=1}^n X_i} = \frac{1}{\bar{X}} \quad \text{where } \frac{d^2}{d\lambda^2} \log L(\lambda) = -\frac{n}{\lambda^2} < 0$$

$$\text{So } \hat{\lambda}_{MLE} = \frac{1}{\bar{X}}$$

$$(c) E(\hat{\lambda}) = \frac{b_0 a_0}{\Gamma(a_0)} X^{a_0-1} \exp(-b_0 X)$$

$\lambda = \frac{1}{\bar{X}}$  will converge to  $\frac{b_0}{a_0}$  which is not equal to  $\lambda_0$  when  $a_0 \neq b_0$ , which means inconsistent. But if  $a_0 = b_0$ ,  $\hat{\lambda}$  is consistent.



# Running Sample Means for 10 Repeated Experiments

