

Statistical Foundations of Data Science

Bayes Theorem and The Normal Distribution

University of the Witwatersrand

2025

Review Question

- Compute $\binom{10}{4}$

Lesson Plan

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- Many things are actually normal. Because of the CLT.

Normal distribution - Example 1

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- Goliath's Z-score is $\frac{X-\mu}{\sigma} = \frac{220-175}{15} = 3$. We use a table to see 0.9987 or 99.87 percent of the population is shorter.

Normal distribution - Example 2

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$$\begin{aligned} Z &= \frac{X - \mu}{\sigma} \\ 3 &= \frac{75 - 50}{\sigma} \\ \sigma &= 8.333 \end{aligned}$$

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$$Z = \frac{X - \mu}{\sigma}$$

$$4 = \frac{83 - \mu}{2}$$

$$\mu = 75$$

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- Nearly normal in the sense that the distribution of the means approximates a normal better and better as n gets larger.
- This also assumes that the distribution we're sampling from has a mean and variance. Most distributions do.

Central Limit Theorem

- More particularly if X_i have mean μ and variance σ^2 \bar{X} is distributed $N(\mu, \frac{\sigma^2}{n})$. Same mean but standard deviation is divided by \sqrt{n} . For large n the sample mean is therefore pretty accurate.

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- There are variants of the CLT with relaxed assumptions! These include weak dependence within the samples.

Formal Statement of the Central Limit Theorem

Let X_1, X_2, \dots, X_n be a sequence of i.i.d. random variables with:

- Mean $\mathbb{E}[X_i] = \mu$,
- Variance $\text{Var}(X_i) = \sigma^2$, assuming $0 < \sigma^2 < \infty$.

Define the standardized sum:

$$Z_n = \frac{\sum_{i=1}^n X_i - n\mu}{\sigma\sqrt{n}}$$

Then as $n \rightarrow \infty$,

$$Z_n \rightarrow \mathcal{N}(0, 1).$$

Proof Outline

- We use characteristic functions: $\varphi_X(t) = \mathbb{E}[e^{itX}]$.
- The characteristic function is given by:

$$\varphi_X(t) = \mathbb{E}[e^{itX}].$$

- The characteristic function can be thought of as a clothesline to hang moments from. To see this, we expand e^{itX} into its Taylor series:

$$e^{itX} = \sum_{k=0}^{\infty} \frac{(itX)^k}{k!}.$$

- Taking expectation term by term:

$$\varphi_X(t) = \sum_{k=0}^{\infty} \frac{(it)^k}{k!} \mathbb{E}[X^k].$$

Proof Outline

- Comparing with the power series of e^{itX} , we see:

$$\mathbb{E}[X^k] = \left. \frac{d^k}{dt^k} \varphi_X(t) \right|_{t=0}.$$

- Thus, the characteristic function generates moments through differentiation.
- The characteristic function of a sum is given by:

$$\varphi_{S_n}(t) = \mathbb{E}[e^{itS_n}].$$

- To see this notice $S_n = X_1 + X_2 + \dots + X_n$, we expand:

$$\varphi_{S_n}(t) = \mathbb{E}[e^{it(X_1+X_2+\dots+X_n)}].$$

- Using the property of exponentials:

$$e^{it(X_1+X_2+\dots+X_n)} = e^{itX_1} e^{itX_2} \dots e^{itX_n}.$$

- By independence, expectation distributes:

$$\mathbb{E}[e^{itX_1} e^{itX_2} \dots e^{itX_n}] = \mathbb{E}[e^{itX_1}] \mathbb{E}[e^{itX_2}] \dots \mathbb{E}[e^{itX_n}].$$

$$\varphi_{S_n}(t) = (\varphi_X(t))^n.$$

Expand $\varphi_X(t)$ around $t = 0$:

$$\varphi_X(t) = 1 + it\mu - \frac{t^2\sigma^2}{2} + o(t^2).$$

Proof outline

$$\varphi_{S_n}(t) = \left(1 + it\mu - \frac{t^2\sigma^2}{2} + o(t^2)\right)^n.$$

Using $(1 + x)^n \approx e^{nx}$ for small x :

$$\varphi_{S_n}(t) \approx e^{n(it\mu - \frac{t^2\sigma^2}{2})}.$$

Proof outline

Define $Z_n = \frac{S_n - n\mu}{\sigma\sqrt{n}}$. The characteristic function of Z_n is:

$$\varphi_{Z_n}(t) = e^{-\frac{t^2}{2}}.$$

This is exactly the characteristic function of $\mathcal{N}(0, 1)$.

Conclusion

- By Lévy's continuity theorem, $Z_n \xrightarrow{d} \mathcal{N}(0, 1)$.
- This completes the proof of the Central Limit Theorem.

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- Linear combinations of all variables are still normal
- All conditional distributions are normal.

Conditional Probability

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- What if I say that they are a male aged between 25 and 30?
- The point is that some knowledge of a dataset tells us something about the population as a whole.

Classic Example

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- Example. What if you're in 2017 and someone time travelled back with a test? Or you're on the International Space Station?
- It turns out that your prior probability matters!

Classic Example



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- Two kinds of people get back positive covid tests. Those who have covid and for whom the test works. These make up $0.1 \times 0.95 = 0.095$ of the population. The other kind are those who don't have covid but who the test failed for. These make up $0.9 \times 0.05 = 0.045$ of the population.

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- So your actual probability of having the disease given the test is $\frac{0.095}{0.095+0.045} = 0.67857142857$.
- These numbers can be tweaked a lot. Again the extreme example is testing for a non-existent disease.

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- Well the possible primes are 2, 3 and 5. So we know we rolled one of those. Hence $\frac{1}{3}$
- Generally getting some information puts us in a subset of the sample space.

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- The exception is when $\mathbb{P}(A) = 0$, this happens with continuous models. There are entirely natural and good and rigorous ways to handle this, for the most part we'll just notice it's an issue.
- When we get to the regression part of this course we'll do a lot of this. For example we might consider the conditional distribution of someone's weight given their height, while the probability of the height being exactly $1.812312382904382382239283m$ is zero.

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- Find $\mathbb{P}(XY > 0.5 | X = 0.8)$
- $\mathbb{P}(0.8Y > 0.5) = \mathbb{P}(Y > 0.5/0.8) = 0.375$
- Here the probability that $X = 0.8$ is zero but you know what to do.