

习题 2.1 3, 4, 8(1), 第一个定理之证明

3.  $X, Y$  相互独立,  $X \sim \text{Exp}(\lambda_1), Y \sim \text{Exp}(\lambda_2)$ 证明 (1)  $\min\{X, Y\} \sim \text{Exp}(\lambda_1 + \lambda_2)$ 

$$P(\min\{X, Y\} < k) = \iint_{\min\{u, v\} < k} \lambda_1 e^{-\lambda_1 u} \lambda_2 e^{-\lambda_2 v} du dv$$

$$= \iint_{\substack{u \leq v \\ u < k}} \lambda_1 \lambda_2 e^{-\lambda_1 u - \lambda_2 v} du dv + \iint_{\substack{v \leq u \\ v < k}} \lambda_1 \lambda_2 e^{-\lambda_1 u - \lambda_2 v} du dv$$

$$= \int_k^{+\infty} \lambda_2 e^{-\lambda_2 v} \left( \int_k^v \lambda_1 e^{-\lambda_1 u} du \right) dv + \int_k^{+\infty} \lambda_1 e^{-\lambda_1 u} \left( \int_k^u \lambda_2 e^{-\lambda_2 v} dv \right) du$$

$$= \int_k^{+\infty} \lambda_2 e^{-\lambda_2 v} [e^{-\lambda_1 k} - e^{-\lambda_1 v}] dv + \int_k^{+\infty} \lambda_1 e^{-\lambda_1 u} [e^{-\lambda_2 k} - e^{-\lambda_2 u}] du$$

$$= e^{-\lambda_1 k} \int_k^{+\infty} \lambda_2 e^{-\lambda_2 v} dv - \lambda_2 \int_k^{+\infty} e^{-(\lambda_1 + \lambda_2)v} dv + e^{-\lambda_2 k} \int_k^{+\infty} \lambda_1 e^{-\lambda_1 u} du - \lambda_1 \int_k^{+\infty} e^{-(\lambda_1 + \lambda_2)u} du$$

$$= e^{-\lambda_1 k} [e^{-\lambda_2 k}] - \lambda_2 \cdot \frac{1}{\lambda_1 + \lambda_2} e^{-(\lambda_1 + \lambda_2)k} + e^{-\lambda_2 k} e^{-\lambda_1 k} - \lambda_1 \cdot \frac{1}{\lambda_1 + \lambda_2} e^{-(\lambda_1 + \lambda_2)k}$$

$$= e^{-(\lambda_1 + \lambda_2)k}, \text{ 因此 } \min\{X, Y\} \sim \text{Exp}(\lambda_1 + \lambda_2)$$

$$(2) P(X < Y) = \iint_{0 \leq u < v} \lambda_1 \lambda_2 e^{-\lambda_1 u - \lambda_2 v} du dv = \int_0^{+\infty} \lambda_2 e^{-\lambda_2 v} \left( \int_0^v \lambda_1 e^{-\lambda_1 u} du \right) dv$$

$$= \int_0^{+\infty} \lambda_2 e^{-\lambda_2 v} (1 - e^{-\lambda_1 v}) dv = \int_0^{+\infty} \lambda_2 e^{-\lambda_2 v} dv - \int_0^{+\infty} \lambda_2 e^{-(\lambda_1 + \lambda_2)v} dv$$

$$= 1 - \frac{\lambda_2}{\lambda_1 + \lambda_2} = \frac{\lambda_1}{\lambda_1 + \lambda_2}$$

4.  $X_1, \dots, X_n$  相互独立,  $P(X_i = k) = (1-p)^{k-1}p, X_i \sim \text{Exp}(\lambda)$ 

$$P(X_1 < X_2) = \sum_{k=1}^{\infty} p(1-p)^{k-1} P(X_1 < X_2 | X_1 = k)$$

$$f_{X_1}(t) = \sum_{k=1}^{\infty} p(1-p)^{k-1} f_{X_1+X_2+\dots+X_k}(t) \quad \text{由于 } X_i \sim \text{Exp}(\lambda) = \Gamma(1, \lambda), \text{ 独立和 } X_1 + \dots + X_k \sim \Gamma(k, \lambda)$$

$$= \sum_{k=1}^{\infty} p(1-p)^{k-1} \frac{\lambda^k t^{k-1} e^{-\lambda t}}{(k-1)!} = \lambda p e^{-\lambda t} \sum_{k=1}^{\infty} \frac{[(1-p)\lambda t]^{k-1}}{(k-1)!} = \lambda p e^{-\lambda t} e^{\lambda t(1-p)} = \lambda p e^{-\lambda p t}$$

故  $X \sim \text{Exp}(\lambda p)$ Poisson 分布细分:  $V$  为几何分布, 记录了第一次分给第一类的时刻 (第一次成功的时刻)

8(1) 推论 2.1.8 (马尔可夫性)  $\forall r \geq 1, 0 \leq t_1 < \dots < t_r < t < t+s$ , 单调上升非负整数  $0 \leq k_1 \leq \dots \leq k_r \leq k \leq l$   
 有  $P(X_{t+s}=l | X_t=k, X_{t_1}=k_1, \dots, X_{t_r}=k_r)$   
 $= P(X_{t+s}=l | X_t=k) = P(X_s=l-k)$

证明:  $P(X_{t+s}=l | X_t=k, X_{t_1}=k_1, \dots, X_{t_r}=k_r)$

$$= \frac{P(X_{t+s}=l, X_t=k, X_{t_1}=k_1, \dots, X_{t_r}=k_r)}{P(X_t=k, X_{t_1}=k_1, \dots, X_{t_r}=k_r)}$$

利用独立增量性质

$$\begin{aligned} \text{分子} &= P(X_t=k, X_{t_r}=k_r, \dots, X_{t_1}=k_1) = P(X_t=k, X_{t-t_r}=k-k_r, X_{t-t_r-t_{r-1}}=k-k_r-k_{r-1}, \dots, X_{t-t_1-t_2}=k-k_1-k_2) \\ &= P(X_t=k) P(X_{t_r}=k_r, \dots, X_{t_1}=k_1 | X_t=k) = P(X_t=k) P(X_t-X_{t_r}=k-k_r, X_{t_r}-X_{t_{r-1}}=k_r-k_{r-1}, \dots, \\ &\quad X_{t_r}-X_{t_1}=k_r-k_1) \end{aligned}$$

$$= P(X_t=k) P(X_{t-t_r}=k-k_r) P(X_{t_r-t_{r-1}}=k_r-k_{r-1}) \dots P(X_{t-t_1-t_2}=k-k_1-k_2)$$

同理有: 分子 =  $P(X_{t+s}=l, X_t=k, \dots, X_{t_r}=k_r)$

$$= P(X_{t+s}=l, X_t=k) P(X_{t-t_r}=k-k_r) \dots P(X_{t-t_1-t_2}=k-k_1-k_2)$$

$$\text{因此分数} = P(X_{t+s}=l | X_t=k) = \frac{P(X_{t+s}=l, X_t=k)}{P(X_t=k)} = P(X_{t+s}-X_t=l-k) = P(X_s=l-k)$$

课上第一个定理的证明:

定理: 计数过程  $X = \{X_t | t \geq 0\} \sim PP(\lambda) \Leftrightarrow X$  满足,  $X_0=0$ , 且为独立平稳增量。具有普遍性, 简单性。

证明: ( $\Leftarrow$ ) 即证明  $X_t \sim P(\lambda t)$  泊松分布。记  $P_k(t) = P(X_t=k)$ 。

$$\begin{aligned} P_0(t+h) &= P(X_{t+h}=0) = P(X_t=0, X_{t+h}-X_t=0) \stackrel{\text{独立增量}}{=} P(X_t=0) P(X_{t+h}-X_t=0) \\ &= P_0(t) [1-\lambda h + o(h)] \end{aligned}$$

$$\begin{aligned} P_k(t+h) &= P(X_{t+h}=k) = P(X_t=k, X_{t+h}-X_t=0) + P(X_t=k-1, X_{t+h}-X_t=1) + \dots + P(X_t=0, X_{t+h}-X_t=k) \\ &= P_k(t) (1-\lambda h + o(h)) + P_{k-1}(t) (\lambda h + o(h)) + o(h) [P_{k-2}(t) + \dots + P_0(t)] \end{aligned}$$

故令  $h \rightarrow 0$  有:  $P'_k(t) = -\lambda P_k(t) + \lambda P_{k-1}(t)$ , 也值条件  $P_k(0) = 0 \quad (\forall k)$

$$\Rightarrow P_k(t) = P(X_t=k) = \frac{(\lambda t)^k}{k!} e^{-\lambda t}$$

( $\Rightarrow$ )  $X$  为  $PP(\lambda)$ , 当然满足右边条件。

$$X \sim P(\lambda) \Rightarrow \frac{\lambda^x e^{-\lambda}}{x!}$$

$$X \sim \text{Exp}(\lambda) \Rightarrow \lambda e^{-\lambda t}$$

Date

No.

Week 10 ② 作业 2.1节 5, 7, 8(2); 定理 2.1.12(2) 叙述

5. (1) 合并后的车流  $\{Z_t\} \sim PP(\lambda_1 + \lambda_2)$

$$Z_1 - Z_0 \stackrel{d}{=} Z_1 \sim P(\lambda_1 + \lambda_2)$$

$$\text{故 } P(Z_1 = 3) = \frac{(\lambda_1 + \lambda_2)^3}{3!} e^{-(\lambda_1 + \lambda_2)}$$

(2) 等待第一辆脚踏车的时刻  $S_1^{(X)}$ , 在  $[0, S_1^{(X)})$  时间内, 无车数目  $Y_{S_1^{(X)}} \sim P(S_1^{(X)} \lambda_2)$

$$\text{而 } S_1^{(X)} \stackrel{d}{=} S_1 \sim \text{Exp}(\lambda)$$

$$\text{故 } P(Y_{S_1^{(X)}} = 3) = \int_0^\infty P(Y_{S_1^{(X)}} = 3 | S_1^{(X)} = t) P(S_1^{(X)} = t) dt$$

$$= \int_0^\infty \frac{(\lambda_2 t)^3 e^{-\lambda_2 t}}{3!} \cdot \lambda_1 e^{-\lambda_1 t} dt = \frac{\lambda_1 \lambda_2^3}{6} \int_0^\infty t^3 e^{-(\lambda_1 + \lambda_2)t} dt$$

$$= \frac{\lambda_1 \lambda_2^3}{6} \cdot \frac{6}{(\lambda_1 + \lambda_2)^4} = \frac{\lambda_1 \lambda_2^3}{(\lambda_1 + \lambda_2)^3}$$

7. (1)  $p = \frac{800 - 300}{2000 - 300} = \frac{5}{17}$ , 故读利(未注)为  $\{X_t\} \sim PP(\frac{5}{17}\lambda)$

$$X_{30} \sim P(\frac{150}{17}\lambda), P(X_{30} = k) = \frac{1}{k!} \cdot (\frac{150\lambda}{17})^k \cdot e^{-\frac{150}{17}\lambda}$$

$$(2) E(e^{aY}) = E(e^{a(\delta_1 + \dots + \delta_X)}) = E(E(e^{a(\delta_1 + \dots + \delta_X)} | X))$$

$$= E(E(e^{a\delta_1} | X)) = E\left(\left(\frac{e^{2a} - e^a}{a}\right)^X\right) \quad \text{母函数 } E(\delta^X) = e^{\lambda(S-1)}$$

$$= e^{\frac{150}{17}\lambda \left(\frac{e^{2a} - e^a}{a} - 1\right)}$$

8(2) 命题 2.1.10 在  $\{X_T = k\}$  条件下,  $(S_1, \dots, S_k)$  的条件密度为  $\frac{k!}{T^k} 1_{\{0 < S_1 < \dots < S_k < T\}}$

证明:  $P(S_1 = s_1, \dots, S_k = s_k | X_T = k)$

$$= P(\delta_1 = s_1, \delta_2 = s_2 - s_1, \dots, \delta_{k-1} = s_{k-1} - s_{k-2}, \delta_k = s_k - s_{k-1}, \delta_{k+1} > T - s_k) / P(X_T = k)$$

$$[\text{由独立性}] = \int \left[ \int \left( \int_{T-s_k}^\infty \lambda e^{-\lambda t_{k+1}} dt_{k+1} \right) dt_k \right] \dots$$

$$\begin{aligned} \text{分子} &= P(\delta_1 = s_1) P(\delta_2 = s_2 - s_1) \dots P(\delta_k = s_k - s_{k-1}) \int_{T-s_k}^\infty \lambda e^{-\lambda t} dt \cdot 1_{\{0 < s_1 < \dots < s_k < T\}} \\ &= \lambda e^{-\lambda s_1} \cdot \lambda e^{-\lambda(s_2 - s_1)} \dots \lambda e^{-\lambda(s_k - s_{k-1})} \cdot \int_{T-s_k}^\infty \lambda e^{-\lambda t} dt \cdot 1_{\{0 < s_1 < \dots < s_k < T\}} \\ &= \lambda^{k+1} e^{-\lambda s_k} \cdot 1_{\{0 < s_1 < \dots < s_k < T\}} \end{aligned}$$

$$\text{分母 } P(X_T = k) = \frac{(\lambda T)^k e^{-\lambda T}}{k!}$$

$$\begin{aligned} \text{分子: } P(S_1 = s_1, \dots, S_k = s_k, X_T = k) &= P(\delta_1 = s_1, \delta_2 = s_2 - s_1, \dots, \delta_{k-1} = s_{k-1} - s_{k-2}, \delta_k = s_k - s_{k-1}, \delta_{k+1} > T - s_k) \\ &= \lambda^k e^{-\lambda s_k} \cdot \int_{T-s_k}^\infty \lambda e^{-\lambda t} dt = \lambda^k e^{-\lambda s_k} e^{-\lambda(T-s_k)} \cdot 1_{\{0 < s_1 < \dots < s_k < T\}} = \lambda^k e^{-\lambda T} \cdot 1_{\{0 < s_1 < \dots < s_k < T\}} \end{aligned}$$

$$\text{分母: } \lambda^k e^{-\lambda T} \cdot 1_{\{0 < s_1 < \dots < s_k < T\}} \cdot \frac{k!}{T^k \lambda^k} e^{\lambda T} = \frac{k!}{T^k} \cdot 1_{\{0 < s_1 < \dots < s_k < T\}} \quad \square$$



## 定理 2.12 的续

$\{X_t\} \sim PP(\lambda)$ ,  $\{Y_t\} \sim PP(\mu)$ , 则  $\{X_t + Y_t\} \sim PP(\lambda + \mu)$

$X_{S_1}, X_{S_2} - X_{S_1}, \dots, X_{S_n} - X_{S_{n-1}}$  独立,  $P(\lambda)(S_k - S_{k-1})$

$$\Rightarrow X_{S_1} + Y_{S_1}, (X_{S_2} + Y_{S_2}) - (X_{S_1} + Y_{S_1}), \dots$$

, ..., 也独立,  $\sim P(\lambda + \mu)(S_k - S_{k-1})$

$$\Rightarrow \{Z_t\} \sim PP(\lambda + \mu)$$

## Week 11 作业

### 1. Kolmogorov 后返方程

$$S = \{1, 2\}, Q = \begin{pmatrix} -\lambda & \lambda \\ \mu & -\mu \end{pmatrix}$$

$$\begin{bmatrix} p'_{11}(t) & p'_{12}(t) \\ p'_{21}(t) & p'_{22}(t) \end{bmatrix} = \begin{bmatrix} -\lambda & \lambda \\ \mu & -\mu \end{bmatrix} \begin{bmatrix} p_{11}(t) & p_{12}(t) \\ p_{21}(t) & p_{22}(t) \end{bmatrix}$$

$$\text{有 } \begin{pmatrix} p'_{11}(t) \\ p'_{21}(t) \end{pmatrix} = \begin{pmatrix} -\lambda & \lambda \\ \mu & -\mu \end{pmatrix} \begin{pmatrix} p_{11}(t) \\ p_{21}(t) \end{pmatrix}$$

$$\uparrow$$

$$z_1 = 0, v_1 = \begin{pmatrix} 1 \\ 1 \end{pmatrix}$$

$$z_2 = -(\lambda + \mu), v_2 = \begin{pmatrix} \lambda \\ -\mu \end{pmatrix}$$

$$\Rightarrow c_1 \begin{pmatrix} 1 \\ 1 \end{pmatrix} + c_2 e^{-(\lambda + \mu)t} \begin{pmatrix} \lambda \\ -\mu \end{pmatrix} \text{ 且 } p_{11}(0) = 1, p_{21}(0) = 0$$

$$\text{有 } p_{11}(t) = \frac{\mu}{\lambda + \mu} + \frac{\lambda}{\lambda + \mu} e^{-(\lambda + \mu)t}$$

$$p_{21}(t) = \frac{\mu}{\lambda + \mu} - \frac{\mu}{\lambda + \mu} e^{-(\lambda + \mu)t}$$

同理, 有

$$p_{12}(t) = \frac{\lambda}{\lambda + \mu} - \frac{\lambda}{\lambda + \mu} e^{-(\lambda + \mu)t}$$

$$p_{22}(t) = \frac{\lambda}{\lambda + \mu} + \frac{\mu}{\lambda + \mu} e^{-(\lambda + \mu)t}$$

2. 设  $X = \{X_t\}$  有限状态 CTMC, 平稳转移阵标准, 则  $Q$  保守且  $P'(t) = P(t)Q$ ,  $P'(t) = QP(t)$

$$\text{证明: } q_{ij} = p'_{ij}(0^+) = \lim_{t \rightarrow 0^+} \frac{p_{ij}(t) - 1}{t} = \lim_{t \rightarrow 0^+} \frac{-\sum_{k \neq i} p_{ik}(t)}{t} \cdot (-\sum_{j \neq i} q_{ij})$$

$$\text{故 } q_i - q_{ji} = \sum_{j \neq i} q_{ij} < +\infty, Q \text{ 保守}$$

$$p'_{ij}(t) = \lim_{\Delta t \rightarrow 0^+} \frac{p_{ij}(t + \Delta t) - p_{ij}(t)}{\Delta t} = \lim_{\Delta t \rightarrow 0^+} \frac{\sum_{k \neq i} p_{ik}(t) (p_{kj}(\Delta t) - \delta_{kj})}{\Delta t}$$

$$= (P(t)Q)_{ij}$$

$$p'_{ij}(t) = \lim_{\Delta t \rightarrow 0^+} \frac{p_{ij}(t - \Delta t) - p_{ij}(t)}{-\Delta t} = \lim_{\Delta t \rightarrow 0^+} \frac{\sum_{k \neq i} p_{ik}(t - \Delta t) (\delta_{ij} - p_{kj}(\Delta t))}{-\Delta t}$$

$$\text{由半群性质, } P(t - \Delta t)P(\Delta t) = P(t)$$

$$\Rightarrow P(t - \Delta t)P(\Delta t) = \lim_{\Delta t \rightarrow 0^+} P(t - \Delta t) \cdot I \Rightarrow \lim_{\Delta t \rightarrow 0^+} P(t - \Delta t) = P(t)$$

$$\therefore p'_{ij}(t) = \sum_{k \neq i} p_{ik}(t) q_{kj} = (P(t)Q)_{ij} \Rightarrow P'(t) = P(t)Q$$

$$\text{即 } p'_{ij}(t) = \lim_{\Delta t \rightarrow 0^+} \frac{p_{ij}(t+\Delta t) - p_{ij}(t)}{\Delta t} = \lim_{\Delta t \rightarrow 0^+} \frac{\sum_{k \in S} (p_{ik}(\Delta t) - \delta_{ij}) p_{kj}(t)}{\Delta t} = \sum_{k \in S} I_{ik} p_{kj}(t)$$

$$p'_{ij}(t^-) = \sum_{k \in S} I_{ik} p_{kj}(t)$$

$$\text{故 } p'(t) = -Qp(t)$$