

Daniel Lupercio - STAT 724 HW 1

```
In [2]: # %load ../standard_import.txt
import pandas as pd
import numpy as np
import matplotlib.pyplot as plt
from mpl_toolkits.mplot3d import axes3d
import seaborn as sns

from sklearn.preprocessing import scale
import sklearn.linear_model as skl_lm
from sklearn.metrics import mean_squared_error, r2_score
import statsmodels.api as sm
import statsmodels.formula.api as smf

%matplotlib inline
plt.style.use('seaborn-white')
```

Chapter 3, Exercise 10

```
In [3]: import os
UP_DIR = '/Users/daniel421/Desktop/STAT_724/ISLR_data'
csv_file = os.path.join(UP_DIR, 'Carseats.csv')
car_seats = pd.read_csv(csv_file)
car_seats.head()
```

```
Out[3]:
```

	Unnamed: 0	Sales	CompPrice	Income	Advertising	Population	Price	ShelveLoc	Age	Education	Urban	US
0	1	9.50	138	73	11	276	120	Bad	42	17	Yes	Yes
1	2	11.22	111	48	16	260	83	Good	65	10	Yes	Yes
2	3	10.06	113	35	10	269	80	Medium	59	12	Yes	Yes
3	4	7.40	117	100	4	466	97	Medium	55	14	Yes	Yes
4	5	4.15	141	64	3	340	128	Bad	38	13	Yes	No

(a) Fit a multiple regression model to predict *Sales* using *Price*, *Urban*, and *US*.

```
In [4]: model_fit = smf.ols("Sales ~ Price + Urban + US", car_seats).fit()
```

(b) Provide an interpretation of each coefficient in the model. Be careful—some of the variables in the model are qualitative!

```
In [5]: model_fit.summary().tables[1]
```

```
Out[5]:
```

	coef	std err	t	P> t	[0.025	0.975]
Intercept	13.0435	0.651	20.036	0.000	11.764	14.323
Urban[T.Yes]	-0.0219	0.272	-0.081	0.936	-0.556	0.512
US[T.Yes]	1.2006	0.259	4.635	0.000	0.691	1.710
Price	-0.0545	0.005	-10.389	0.000	-0.065	-0.044

For fixed values of *Urban* and *US*, a 1-unit increase in **Price** results in a change of *Sales* of -0.0545 units (54 sales).

For fixed values of *Price* and *Urban*, the effect of the store being located in the **US** is a change of *Sales* of 1.2006 units (1,200 sales).

For fixed values of *Price* and *US*, the effect of the store being located in an **Urban** location is a change of *Sales* of -.0219 units (decrease of 22 sales).

(c) Write out the model in equation form, being careful to handle the qualitative variables properly.

$$\hat{y} = 13.0435 - 0.0219 * Urban + 1.2006 * US - 0.0545 * Price$$

- *Urban* = 1 for a store in an urban location, 0 elsewhere
- *US* = 1 for a store in the US, 0 elsewhere

(d) For which of the predictors can you reject the null hypothesis $H_0 : \beta_j = 0$?

$$H_0 : \beta_{2,3} = 0 \text{ \& } H_A : \beta_{2,3} \neq 0$$

Based on the p-values of *US* and *Price*, we can reject $H_0 : \beta_{2,3} = 0$

(e) On the basis of your response to the previous question, fit a smaller model that only uses the predictors for which there is evidence of association with the outcome.

```
In [6]: model_fit2 = smf.ols("Sales ~ Price + US", car_seats).fit()
```

```
In [7]: model_fit2.summary().tables[0]
```

```
Out[7]:
```

OLS Regression Results			
Dep. Variable:	Sales	R-squared:	0.239
Model:	OLS	Adj. R-squared:	0.235
Method:	Least Squares	F-statistic:	62.43
Date:	Sat, 25 Sep 2021	Prob (F-statistic):	2.66e-24
Time:	13:54:09	Log-Likelihood:	-927.66
No. Observations:	400	AIC:	1861.
Df Residuals:	397	BIC:	1873.
Df Model:	2		
Covariance Type:	nonrobust		

(f) How well do the models in (a) and (e) fit the data?

$R^2 = 0.239$ and $\bar{R}^2 = 0.234$ for the (a) model
 $R^2 = 0.239$ and $\bar{R}^2 = 0.235$ for the (e) model

Both models can explain approximately 23.9% of the variance in *Sales*. However, the \bar{R}^2 for model (e) has a slight increase. This can be attributed in part, to the removal of the *Urban* variable. Although we have very limiting information, it would be best to use the model in (e).

(g) Using the model from (e), obtain 95% confidence intervals for the coefficient(s).

```
In [8]: model_fit2.conf_int(alpha=0.05, cols=None)
```

```
Out[8]:
```

	0	1
Intercept	11.79032	14.271265
US[T.Yes]	0.69152	1.707766
Price	-0.06476	-0.044195

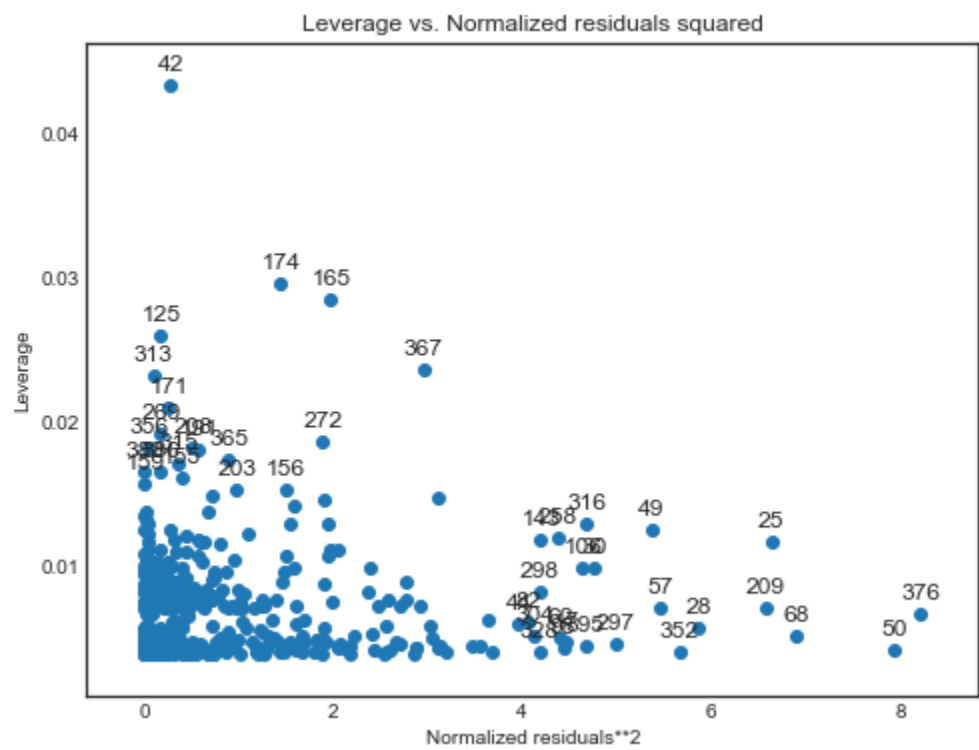
We can say that there is a 95% probability that, on average, the true parameter for *Price* (β_2) falls within (-0.0648, -0.0442).

We can say that there is a 95% probability that, on average, the true parameter for *US* (β_1) falls within (0.692, 1.708).

(h) Is there evidence of outliers or high leverage observations in the model from (e)?

```
In [9]: from statsmodels.graphics.regressionplots import plot_leverage_resid2

fig, ax = plt.subplots(figsize=(8, 6))
fig = plot_leverage_resid2(model_fit2, ax=ax)
```



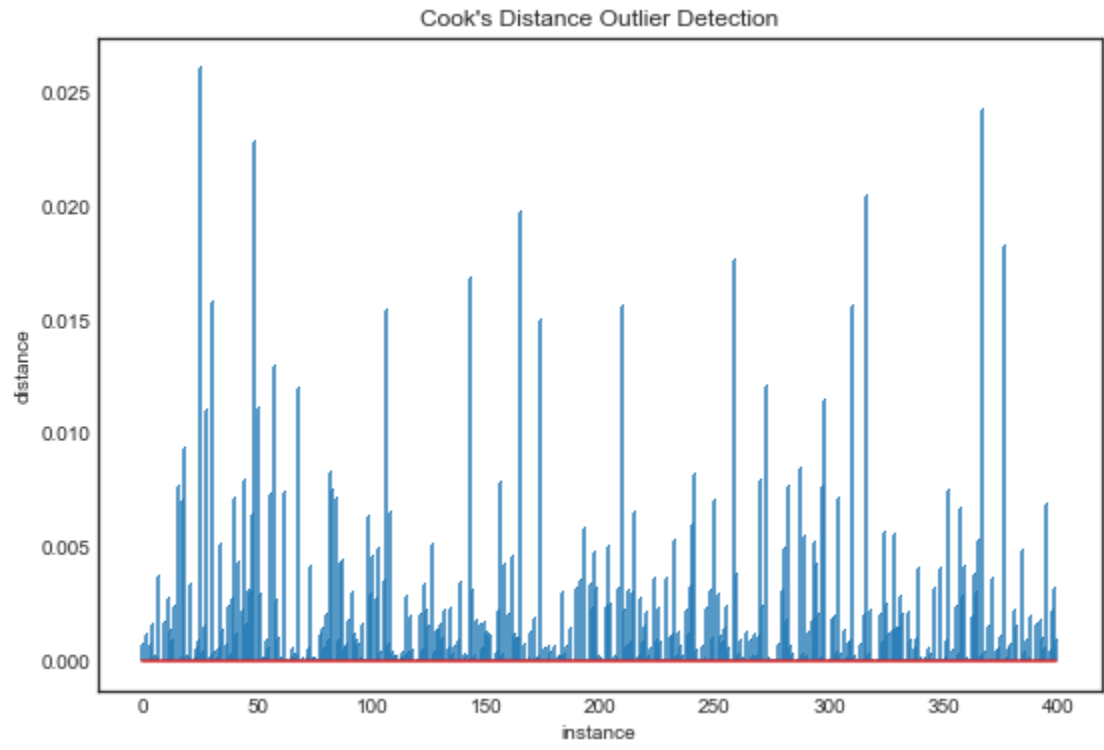
```
In [10]: from statsmodels.stats.outliers_influence import OLSInfluence as influence
# Compute the influence to get Cook's distance
inf = influence(model_fit2)

# cooks_distance is an attribute of influence, here C, not sure about P (p-value maybe?)
C, P = inf.cooks_distance
```

```
In [11]: def plot_cooks_distance(c):
_, ax = plt.subplots(figsize=(9,6))
ax.stem(c, markerfmt=" ")
ax.set_xlabel("instance")
ax.set_ylabel("distance")
ax.set_title("Cook's Distance Outlier Detection")
return ax

plot_cooks_distance(C)
```

```
Out[11]: <AxesSubplot:title={'center': 'Cook's Distance Outlier Detection'}, xlabel='instance', ylabel='distance'>
```



There are instances of high leverage observations, however, there is no indication of strong outliers. This is based off of low cook's distances.

Chapter 3 Exercises

5)

$$\hat{y}_i = x_i \hat{\beta}$$

where

$$\hat{\beta} = \left(\sum_{i=1}^n x_i y_i \right) / \left(\sum_{i=1}^n x_i^2 \right) \quad (3.38)$$

show that we can write

$$\hat{y}_i = \sum_{i'=1}^n a_{i'} y_{i'}$$

From (3.38)

$$\hat{y}_i = \frac{x_i \left(\sum_{i=1}^n x_i y_i \right)}{\sum_{i'=1}^n x_i'^2}$$

Let $d_i = \frac{x_i}{\sum_{i'=1}^n x_i'^2}$, as d_i will be a constant for some $i \in \mathbb{Z}$.

Thus,

$$\hat{y}_i = d_i \sum_{i=1}^n x_i y_i$$

$$\hat{y}_i = \sum_{i=1}^n (d_i x_i) y_i$$

algebra laws

Let $a_{i'} = d_i x_i$

$$\hat{y}_i = \sum_{i=1}^n a_{i'} y_i$$

The fitted values are linear combinations of weighted response values. With $a_{i'}$ being the weight of each y_i , $i \in \mathbb{Z}$