

Lecture 08: Finite Sample PCA I

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Advanced Statistics

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Motivation

- To do PCA of data vectors $\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_N \in \mathbb{R}^d$, coming from some distribution $f_{\mathbf{X}}(\mathbf{x})$, we needed to have the covariance matrix Σ .
- But what happens in cases when we do not know the covariance matrix Σ . Can we still do PCA then as well?
- Specifically, assume we have data vectors $\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_N$, coming from some unknown distribution $f_{\mathbf{X}}(\mathbf{x})$ and unknown covariance matrix Σ .
- How to do PCA then? And what will be the error we make compared to the case when we knew the covariance matrix?

Finite Sample PCA

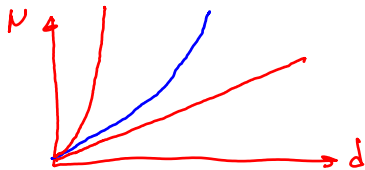
- Let us have a sample of N i.i.d. data vectors $\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_N \in \mathbb{R}^d$, coming from some unknown distribution $f_{\mathbf{X}}(\mathbf{x})$ and unknown covariance matrix Σ .
- Let us estimate the unknown covariance matrix Σ as $\hat{\Sigma}_N$ using the N samples as

$$\hat{\Sigma}_N = \frac{1}{N} \sum_{i=1}^N \mathbf{X}_i \mathbf{X}_i^T, \quad (1)$$

- Then, let's do PCA, using the estimated covariance matrix $\hat{\Sigma}_N$, on the data vectors $\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_N$.
- Hopefully, the results using the estimated covariance matrix $\hat{\Sigma}_N$ will be close to doing the PCA, using the actual covariance matrix Σ , on the data vectors $\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_N$.

Finite Sample PCA

- An immediate question in when using the estimated covariance matrix $\hat{\Sigma}_N$ to do PCA is the following.
- How large should the sample size N be in order for the two PCA results to be close?
- Since each of the samples $\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_N$ is of dimension d , Σ is a $d \times d$ matrix.
- Then, for sure, the accuracy of the PCA using the estimated covariance matrix $\hat{\Sigma}_N$ will depend on how N depends on d .
- How will N depend on d ? Maybe for a good accuracy we need $N = O(d)$, or maybe $N = O(d^2)$ or maybe $N = O(e^d)$



$$d = 10^6$$

$$N = \exp(10^6) \sim$$

$$N = 10^{12}$$

$$N = 10^6$$

Finite Sample PCA

- We know that the PCA result using the covariance matrix Σ is a function of its eigenvalues, $\lambda_i(\Sigma)$, and eigenvectors, $\mathbf{v}_i(\Sigma)$.
- Hence, if $\lambda_i(\hat{\Sigma}_N)$ and $\mathbf{v}_i(\hat{\Sigma}_N)$ are the eigenvalues and eigenvectors of the estimated covariance matrix, $\hat{\Sigma}_N$, then for the two PCA results to be close, the following has to hold

$$\begin{aligned}\lambda_i(\Sigma) &\approx \lambda_i(\hat{\Sigma}_N) \\ \text{AND} \\ \mathbf{v}_i(\Sigma) &\approx \mathbf{v}_i(\hat{\Sigma}_N),\end{aligned}\tag{2}$$

i.e., the eigenvalues and eigenvectors of Σ and $\hat{\Sigma}_N$ are very close to each other.

- How large should the sample size N be in order for the eigenvalues and eigenvectors of Σ and $\hat{\Sigma}_N$ to be very close to each other?

Covariance Estimation Problem

Step 1:

- Let us measure how far Σ and $\hat{\Sigma}_N$ are from each other using some measurement/metric.
- For matrices, so far, we learned that we can use the Frobenius norm or the operator norm.
- Let's try the operator norm:

$$\|\hat{\Sigma}_N - \Sigma\|_{\text{op}} = \sigma_1(\hat{\Sigma}_N - \Sigma), \quad (3)$$

where $\sigma_1(\hat{\Sigma}_N - \Sigma)$ is the largest singular value of the matrix $\hat{\Sigma}_N - \Sigma$

- Now, we have

$$\begin{aligned} \sigma_1(\hat{\Sigma}_N - \Sigma) &= \left| \lambda_1(\hat{\Sigma}_N - \Sigma) \right| = \max_{\mathbf{v}, \|\mathbf{v}\|_2=1} \left| \mathbf{v}^T (\hat{\Sigma}_N - \Sigma) \mathbf{v} \right| \\ &= \max_{\mathbf{v}, \|\mathbf{v}\|_2=1} \left| \mathbf{v}^T \hat{\Sigma}_N \mathbf{v} - \mathbf{v}^T \Sigma \mathbf{v} \right| \end{aligned} \quad (4)$$

Distance Between True and Sampled Covariance Matrices

- The first element in (4) is

$$\begin{aligned} \mathbf{v}^T \hat{\Sigma}_N \mathbf{v} &= \frac{1}{N} \sum_{i=1}^N \mathbf{v}^T \mathbf{X}_i \mathbf{X}_i^T \mathbf{v} = \frac{1}{N} \sum_{i=1}^N Y_i Y_i = \frac{1}{N} \sum_{i=1}^N Y_i^2 \\ &= \frac{1}{N} \sum_{i=1}^N \langle \mathbf{v}, \mathbf{X}_i \rangle^2, \end{aligned} \quad (5)$$

where $Y_i = \mathbf{v}^T \mathbf{X}_i = \mathbf{X}_i^T \mathbf{v} = \langle \mathbf{v}, \mathbf{X}_i \rangle$.

- The second element in (4) is

$$\begin{aligned} \mathbf{v}^T \Sigma \mathbf{v} &= \mathbf{v}^T E[\mathbf{X} \mathbf{X}^T] \mathbf{v} = E[\mathbf{v}^T \mathbf{X} \mathbf{X}^T \mathbf{v}] = E[Y Y] = E[Y^2] \\ &= E[\langle \mathbf{v}, \mathbf{X} \rangle^2] \end{aligned} \quad (6)$$

where $Y = \mathbf{v}^T \mathbf{X} = \mathbf{X}^T \mathbf{v} = \langle \mathbf{v}, \mathbf{X} \rangle$.

Distance Between True and Sampled Covariance Matrices

- Inserting (5) and (6) into (4), and then (4) into (3), we obtain

$$\|\hat{\Sigma}_N - \Sigma\|_{\text{op}} = \max_{\mathbf{v}, \|\mathbf{v}\|_2=1} \left| \frac{1}{N} \sum_{i=1}^N \langle \mathbf{v}, \mathbf{X}_i \rangle^2 - E[\langle \mathbf{v}, \mathbf{X} \rangle^2] \right|, \quad (7)$$

which is a random variable (RV).

- Let us define $Z(\mathbf{v})$ as

$$Z(\mathbf{v}) = \frac{1}{N} \sum_{i=1}^N \langle \mathbf{v}, \mathbf{X}_i \rangle^2 - E[\langle \mathbf{v}, \mathbf{X} \rangle^2], \text{ s.t. } \|\mathbf{v}\|_2 = 1. \quad (8)$$

Then

$$\|\hat{\Sigma}_N - \Sigma\|_{\text{op}} = \max_{\mathbf{v}, \|\mathbf{v}\|_2=1} |Z(\mathbf{v})| \quad (9)$$

Covariance Estimation Theorem

- Thm (Covariance Estimation Theorem): Let $\mathbf{X}_i \in \mathbb{R}^d$ be N sample vectors independently and identically drawn from the Gaussian distribution $\mathcal{N}(0, \mathbf{\Sigma})$. Let each element of \mathbf{X}_i be with variance σ^2 . Let $\delta \geq 0$. Then, the following holds

$$\Pr \left\{ \|\hat{\mathbf{\Sigma}}_N - \mathbf{\Sigma}\|_{\text{op}} \leq \delta \right\} \geq 1 - 2 \exp \left(-d \sqrt{\frac{2}{\epsilon}} \right) \exp \left(-N \frac{(1-\epsilon)\delta}{2\sigma^2} \min \left\{ \frac{(1-\epsilon)\delta}{2\sigma^2}, 1 \right\} \right) \quad (10)$$

- What this Thm says is that the distance between $\hat{\mathbf{\Sigma}}_N$ and $\mathbf{\Sigma}$ can be made smaller than δ , with probability that goes to one as N increases, as long as N satisfies

$$N \frac{(1-\epsilon)\delta}{2\sigma^2} \min \left\{ \frac{(1-\epsilon)\delta}{2\sigma^2}, 1 \right\} > d \sqrt{\frac{2}{\epsilon}} \quad (11)$$

Covariance Estimation Theorem

- What the Covariance Estimation Theorem tells us is that $N = o(d)$.
- In other words, the number of vector samples N we need to accurately estimate the covariance matrix Σ using

$$\hat{\Sigma}_N = \frac{1}{N} \sum_{i=1}^N \mathbf{X}_i \mathbf{X}_i^T \quad (12)$$

depends linearly on the vectors \mathbf{X}_i length d .

- This is very good news since there is no curse-of-dimensionality in this problem.
- Imagine if it was $N = o(e^d)$. Then, it would have been impossible to estimate $\hat{\Sigma}_N$ using (12) since for $d = 100$ we would need at least e^{100} samples, which is more than the number of atoms in the observable universe.

Distance Between Finite Sample PCA and True PCA

- So far we only proved that we can make $\|\hat{\Sigma}_N - \Sigma\|_{\text{op}}$ vary small by increasing N , given that condition (11) holds.
- But PCA does not depend on $\|\hat{\Sigma}_N - \Sigma\|_{\text{op}}$ explicitly, and instead depends on the distances between the eigenvalues and eigenvectors of Σ and $\hat{\Sigma}_N$.
- In fact, for the PCA results using $\hat{\Sigma}_N$ to be very close to the PCA results using Σ , we needed the following to hold for all dimensions i over which we do the PCA (what happens in the other dimensions is irrelevant)

$$\lambda_i(\Sigma) \approx \lambda_i(\hat{\Sigma}_N) \quad \text{AND} \quad \mathbf{v}_i(\Sigma) \approx \mathbf{v}_i(\hat{\Sigma}_N), \quad (18)$$

i.e., the eigenvalues and eigenvectors of Σ and $\hat{\Sigma}_N$ are very close to each other in all dimensions i over which we do the PCA.

- In the following, we will show that by making $\|\hat{\Sigma}_N - \Sigma\|_{\text{op}}$ vary small, we will implicitly achieve (18).

Distance Between Eigenvalues

- Thm (Weyl's Inequality): $\forall d \times d$ symmetric matrices \mathbf{A} and \mathbf{B} , the following holds

$$\max_i |\lambda_i(\mathbf{A}) - \lambda_i(\mathbf{B})| \leq \|\mathbf{A} - \mathbf{B}\|_{\text{op}} \quad (19)$$

- We will not prove it!
- What this theorem says is that the largest distance between the eigenvalues of any two symmetric matrices \mathbf{A} and \mathbf{B} is upper bounded by the operator norm of $\mathbf{A} - \mathbf{B}$
- Hence, if we make $\|\hat{\Sigma}_N - \Sigma\|_{\text{op}}$ vary small, we implicitly make the distance between the eigenvalues of Σ and $\hat{\Sigma}_N$ also very small.
i.e., the following holds

$$\lambda_i(\Sigma) \approx \lambda_i(\hat{\Sigma}_N) \quad (20)$$

- How about the eigenvectors. Now, they are more tricky, but we can still do it as shown in the following.

Distance Between Eigenvectors

- Thm (Davis-Kahan Inequality): For any symmetric matrices Σ and $\hat{\Sigma}_N$ with eigenvectors $\mathbf{v}_1(\Sigma), \mathbf{v}_2(\Sigma), \dots, \mathbf{v}_k(\Sigma)$ and $\mathbf{v}_1(\hat{\Sigma}_N), \mathbf{v}_2(\hat{\Sigma}_N), \dots, \mathbf{v}_k(\hat{\Sigma}_N)$, respectively, let P_Σ and $P_{\hat{\Sigma}_N}$ be defined as

$$P_\Sigma = \sum_{i=1}^k \mathbf{v}_i(\Sigma) \mathbf{v}_i^T(\Sigma) \quad (21)$$

$$P_{\hat{\Sigma}_N} = \sum_{i=1}^k \mathbf{v}_i(\hat{\Sigma}_N) \mathbf{v}_i^T(\hat{\Sigma}_N). \quad (22)$$

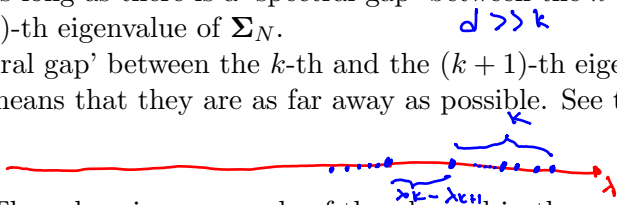
Then, the following holds

$$\|P_{\hat{\Sigma}_N} - P_\Sigma\|_{\text{op}} \leq \frac{\|\hat{\Sigma}_N - \Sigma\|_{\text{op}}}{\lambda_k(\Sigma) - \lambda_{k+1}(\Sigma)} \quad (23)$$

- We will not prove it!

Distance Between Eigenvectors

- What the Davis-Kahan Inequality Thm says is that by making $\|\hat{\Sigma}_N - \Sigma\|_{\text{op}}$ vary small, we implicitly make the first k eigenvectors of $\hat{\Sigma}_N$ to be very close to the first k eigenvectors of Σ_N , as long as there is a 'spectral gap' between the k -th and the $(k+1)$ -th eigenvalue of Σ_N .
- 'Spectral gap' between the k -th and the $(k+1)$ -th eigenvalue of Σ_N , means that they are as far away as possible. See this Figure:



- This Thm also gives us a rule of thumb used in the current PCA algos: Do PCA dimension reduction up to the point when we observe a spectral gap. Anything beyond the spectral gap consider as noise and throw it away.
- In the next lecture, we will arrive at the same concept, noise that needs to be thrown away, via a different method which will reveal us even more insights into random matrices.