

# Assignment 2

April 27, 2022

## 1 Exercise 1

It was finished on the website.

## 2 Exercise 2

1. The initial entropy of "Appealing" is

$$Entropy(t_0) = -\frac{5}{10} \log_2 \frac{5}{10} = \frac{1}{2} \quad (1)$$

2. "Taste" is the root of the decision tree. Then the information gain is

$$\begin{aligned} InfoGain &= Entropy(t_0) - \sum_{k=1}^K \frac{n_k}{n} Entropy(t_k) \\ &= \frac{1}{2} - \left( \frac{0}{3} \log_2 \frac{0}{3} + \frac{3}{3} \log_2 \frac{3}{3} \right) - \left( \frac{2}{4} \log_2 \frac{2}{4} + \frac{2}{4} \log_2 \frac{2}{4} \right) - \left( \frac{3}{3} \log_2 \frac{3}{3} + \frac{0}{3} \log_2 \frac{0}{3} \right) \\ &= \frac{1}{2} - 0 + 1 - 0 \\ &= 1.5 \end{aligned} \quad (2)$$

3. The full decision tree without pruning is as below:

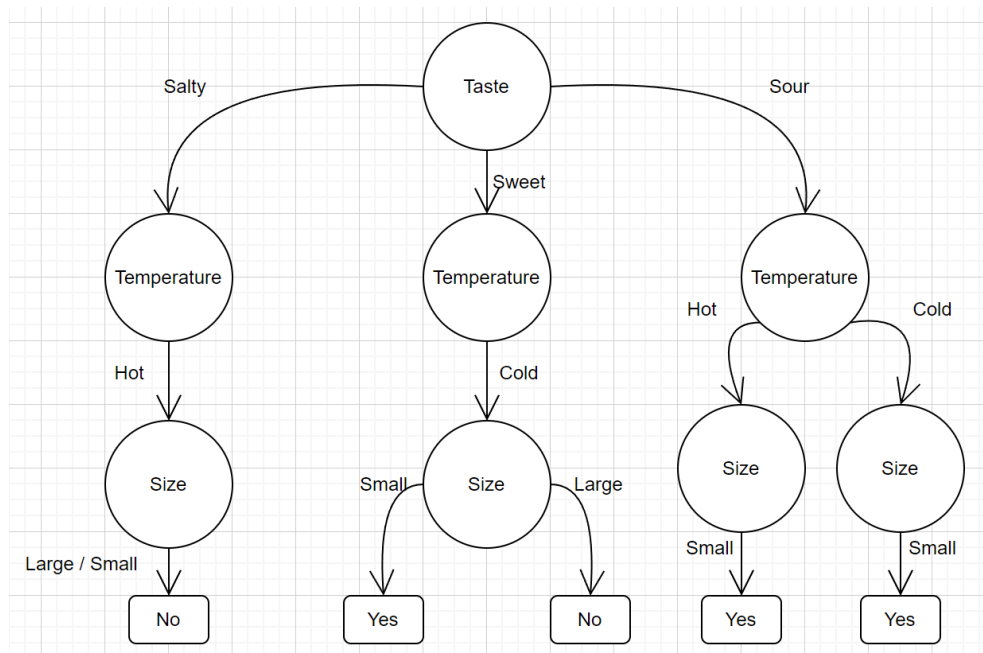


Figure 1: Full decision tree without pruning

### 3 Exercise 3

1. Since

$$f_{\theta} = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{1}{2\sigma^2}(x - \mu)^2\right) \quad (3)$$

then we have

$$\begin{aligned} \ln L(\theta) &= \ln \prod_{i=1}^n f_{\theta}(x_i) = \sum_{i=1}^n \ln f_{\theta}(x_i) \\ &= \sum_{i=1}^n \left(-\frac{1}{2} \ln(2\pi\sigma^2) - \frac{1}{2\sigma^2}(x_i - \mu)^2\right) \\ &= -\frac{n}{2} \ln(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2 \end{aligned} \quad (4)$$

So

$$\frac{\partial \ln L(\theta)}{\partial \mu} = \frac{\mu}{\sigma^2} \sum_{i=1}^n (x_i - \mu) = \frac{\mu}{\sigma^2} \left(\sum_{i=1}^n x_i - n\mu\right) \quad (5)$$

$$\frac{\partial \ln L(\theta)}{\partial \sigma} = -\frac{n}{\sigma} + \frac{1}{\sigma^3} \sum_{i=1}^n (x_i - \mu)^2 \quad (6)$$

Let equation (3) and (4) equal to 0 and combine them, then

$$\hat{\mu} = \frac{1}{n} \sum_{i=1}^n x_i, \quad \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2$$

2.

$$E\hat{\mu} = E\left(\frac{1}{n} \sum_{i=1}^n x_i\right) = \frac{1}{n} \sum_{i=1}^n E(x_i) = \frac{1}{n} \sum_{i=1}^n \mu = \frac{1}{n} n\mu = \mu \quad (7)$$

$$\begin{aligned} E\left(\frac{n}{n-1} \hat{\sigma}^2\right) &= E\left(\frac{n}{n-1} \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2\right) = \frac{1}{n-1} E\left(\sum_{i=1}^n (x_i - \bar{x})^2\right) = \frac{1}{n-1} E\left(\sum_{i=1}^n (x_i - \mu + \mu - \bar{x})^2\right) \\ &= \frac{1}{n-1} E\left(\sum_{i=1}^n ((x_i - \mu)^2 - 2(x_i - \mu)(\bar{x} - \mu) + (\bar{x} - \mu)^2)\right) \\ &= \frac{1}{n-1} E\left(\sum_{i=1}^n (x_i - \mu)^2 - 2 \sum_{i=1}^n (x_i - \mu)(\bar{x} - \mu) + n(\bar{x} - \mu)^2\right) \\ &= \frac{1}{n-1} E\left(\sum_{i=1}^n (x_i - \mu)^2 - 2n(\bar{x} - \mu)^2 + n(\bar{x} - \mu)^2\right) \\ &= \frac{1}{n-1} E\left(\sum_{i=1}^n (x_i - \mu)^2 - n(\bar{x} - \mu)^2\right) \\ &= \frac{1}{n-1} \left(\sum_{i=1}^n E(x_i - \mu)^2 - nE(\bar{x} - \mu)^2\right) \\ &= \frac{1}{n-1} (n \text{Var}(x) - n \text{Var}(\bar{x})) \\ &= \frac{n}{n-1} \left(\sigma^2 - \frac{\sigma^2}{n}\right) = \sigma^2 \end{aligned} \quad (8)$$

### 4 Exercise 4

For prior probability, we have

$$P(Y = k) = p_k^{I(y=k)} (1 - p_k)^{I(y \neq k)} \quad (9)$$

$$\begin{aligned}
L(P) &= \arg \max_{p_k} \sum_{y \in Y} \ln(p_k^{I(y=k)} (1-p_k)^{I(y \neq k)}) \\
&= \arg \max_{p_k} (\sum_{y=k} \ln p_k + \sum_{y \neq k} \ln(1-p_k))
\end{aligned} \tag{10}$$

$$\begin{aligned}
&= \arg \max_{p_k} (n_1 \ln p_k + n_2 \ln(1-p_k)) \quad (n_1 = \sum_{i=1}^n I(y=k), n_2 = \sum_{i=1}^n I(y \neq k)) \\
\frac{\partial L}{\partial p_k} &= \frac{\partial (n_1 \ln p_k + n_2 \ln(1-p_k))}{\partial p_k} \\
&= \frac{n_1}{p_k} - \frac{n_2}{1-p_k}
\end{aligned} \tag{11}$$

Let equation (11) be zero, then we can get

$$\hat{p}_k = \frac{n_1}{n_1 + n_2} = \frac{\sum_{i=1}^n I(y=k)}{n} \tag{12}$$

For conditional probability, we have

$$P(X=s|Y=k) = \frac{P(X=s, Y=k)}{P(Y=k)} \tag{13}$$

Since we have found the value of  $P(Y=k)$ , we need to find the MLE for  $P(X=s, Y=k)$ .

$$\begin{aligned}
L(P) &= \arg \max_p \sum_{x \in X, y \in Y} \ln(p^{I(x=s, y=k)} (1-p)^{1-I(x=s, y=k)}) \\
&= \arg \max_p (n_1 \ln p + n_2 \ln(1-p)) \quad (n_1 = \sum_{i=1}^n I(x=s, y=k), n_2 = \sum_{i=1}^n (1-I(x=s, y=k)))
\end{aligned} \tag{14}$$

$$\begin{aligned}
\frac{\partial L}{\partial p} &= \frac{\partial (n_1 \ln p + n_2 \ln(1-p))}{\partial p} \\
&= \frac{n_1}{p} - \frac{n_2}{1-p}
\end{aligned} \tag{15}$$

Let equation (15) be zero, then we can get  $\hat{p} = \frac{n_1}{n_1 + n_2} = \frac{\sum_{i=1}^n I(x=s, y=k)}{n}$ . To sum up, the MLE of  $p_{sk}$  is

$$\hat{p}_{sk} = \frac{\hat{p}}{\hat{p}_k} = \frac{\sum_{i=1}^n I(x=s, y=k)}{\sum_{i=1}^n I(y=k)} \tag{16}$$

## 5 Exercise 5

$$\begin{aligned}
E_{S \sim P^n} \mathcal{E}(f^{1NN}) &= E_{S \sim P^n} E_{(\mathbf{X}, Y) \sim P} 1_{Y \neq f^{1NN}} \\
&= E_{S \sim P^n} E_{(\mathbf{X}, Y) \sim P} 1_{Y \neq y_{\pi_S(\mathbf{X})}} \\
&= E_{S \sim P^n} (1 - \eta(\mathbf{X})) p_{\mathbf{X}}(\mathbf{X}) \\
&= E_{S \sim P^n} \left[ \int_P \eta(\mathbf{X}) (1 - \eta(\mathbf{X}') + \eta(\mathbf{X}')(1 - \eta(\mathbf{X})) dx \right]
\end{aligned} \tag{17}$$

$$2\mathcal{E}(f^*) + c E_{S \sim P^n} E_{\mathbf{X} \sim p_{\mathbf{X}}} \|\mathbf{X} - \mathbf{X}_{\pi_S(\mathbf{X})}\| \geq E_{S \sim P^n} \left[ 2 \int_P |\eta(\mathbf{X}) - \eta(\mathbf{X}_{\pi_S(\mathbf{X})})| dx + \int_P |\eta(\mathbf{X}) - \eta(\mathbf{X}')| dx \right] \tag{18}$$

So we need to prove that

$$a(1-b) + b(1-a) \leq 2|a - a_\pi| + |a - b| \tag{19}$$

where  $a, b \in [0, 1]$ .

Let  $f(b) = 2|a - a_\pi| + |a - b| + 2ab - a - b$ , we have  $\min_b f(b) = \min f(0), f(1), f(a) = \min 2|a - a_\pi|, 2|a - a_\pi|, 2(a^2 - a + |a - a_\pi|)$ . Therefore, it was proved.