

$$\begin{aligned}
 1. \text{ Proof: } \hat{\beta}_1 &= \frac{SS_{xy}}{SS_{xx}} \\
 &= \frac{\sum_{i=1}^n y_i (x_i - \bar{x})}{SS_{xx}}, \quad SS_{xx} = \sum_{i=1}^n (x_i - \bar{x})^2 \\
 &= \frac{\sum_{i=1}^n \frac{(x_i - \bar{x})}{S_{xx}} \cdot y_i}{\downarrow C_i} \\
 &= \sum_{i=1}^n C_i \cdot y_i, \quad C_i = \frac{x_i - \bar{x}}{SS_{xx}}
 \end{aligned}$$

$$SS_{xy} \equiv S_{xy}, \quad SS_{xx} \equiv S_{xx}$$

$$\begin{aligned}
 \hat{\beta}_0 &= \bar{y} - \hat{\beta}_1 \bar{x} \\
 &= \frac{\sum_{i=1}^n y_i}{n} - \bar{x} \cdot \sum_{i=1}^n C_i y_i \\
 &= \sum_{i=1}^n \underbrace{\left( \frac{1}{n} - C_i \bar{x} \right)}_{d_i} y_i \\
 &= \sum_{i=1}^n d_i y_i
 \end{aligned}$$

$$2. \quad \text{Var}(\hat{\beta}_1) = \text{Var}\left(\sum_{i=1}^n C_i y_i\right)$$

Because  $y_i$ 's are uncorrelated, and  $C_i$ s are non-random.

$$\begin{aligned} \text{Var}(aX+bY) \\ = a^2 \text{Var}(X) + b^2 \text{Var}(Y) \end{aligned}$$

$$\begin{aligned}
 &= \sum_{i=1}^n C_i^2 \text{Var}(y_i) \\
 &= \sum_{i=1}^n \frac{(x_i - \bar{x})^2}{S_{xx}^2} \cdot \sigma^2 \\
 &= \frac{\sigma^2 \left( \sum_{i=1}^n (x_i - \bar{x})^2 \right)}{S_{xx}^2} \quad \text{--- } S_{xx} \\
 &= \frac{\sigma^2}{S_{xx}}
 \end{aligned}$$

$$\bullet \quad \sum_{i=1}^n e_i = 0$$

$$\begin{aligned}
 \text{Proof: } \sum_{i=1}^n e_i &= \sum_{i=1}^n (y_i - \hat{y}_i), \quad \hat{y}_i = \hat{\beta}_0 + \hat{\beta}_1 x_i \\
 &= \sum_{i=1}^n (y_i - (\hat{\beta}_0 + \hat{\beta}_1 x_i)) \\
 &= \sum_{i=1}^n \underbrace{(y_i - (\bar{y} - \hat{\beta}_1 \bar{x}))}_{\text{---}} - \hat{\beta}_1 \underbrace{\sum_{i=1}^n (x_i - \bar{x})}_{\text{---}} \\
 &= \sum_{i=1}^n (y_i - \bar{y}) - \hat{\beta}_1 \sum_{i=1}^n (x_i - \bar{x}) \\
 &= \underbrace{\sum_{i=1}^n y_i - n \cdot \bar{y}}_0 - \hat{\beta}_1 \underbrace{\left( n \bar{x} - \sum_{i=1}^n x_i \right)}_0 = 0
 \end{aligned}$$

$$\sum_{i=1}^n y_i = n \cdot \bar{y} = n \left( \frac{\sum_{i=1}^n y_i}{n} \right)$$

$$\bullet \quad \sum_{i=1}^n y_i = \sum_{i=1}^n \hat{y}_i$$

$$\begin{aligned}
 \text{Proof: } \sum_{i=1}^n e_i &= \sum_{i=1}^n (y_i - \hat{y}_i) = 0 \\
 \text{so } \sum_{i=1}^n y_i - \sum_{i=1}^n \hat{y}_i &= 0 \\
 \Rightarrow \sum_{i=1}^n y_i &= \sum_{i=1}^n \hat{y}_i
 \end{aligned}$$

$$\bullet \quad \sum_{i=1}^n x_i e_i = 0$$

$$\begin{aligned}
 \text{Proof: } \sum_{i=1}^n x_i (y_i - \hat{y}_i) &= \sum_{i=1}^n x_i (y_i - (\hat{\beta}_0 + \hat{\beta}_1 x_i)), \quad \hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x} \\
 &= \sum_{i=1}^n x_i (y_i - (\bar{y} - \hat{\beta}_1 \bar{x}) - \hat{\beta}_1 x_i) \\
 &= \sum_{i=1}^n x_i (y_i - \bar{y}) + \hat{\beta}_1 \sum_{i=1}^n x_i (\bar{x} - x_i) \\
 &= S_{xy} + \hat{\beta}_1 (-S_{xx}) \\
 &= S_{xy} + \frac{S_{xy}}{S_{xx}} (-S_{xx}) \\
 &= S_{xy} - S_{xy} \\
 &= 0
 \end{aligned}$$

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$$\bullet \quad \sum_{i=1}^n \hat{y}_i e_i = 0$$

$$\begin{aligned}
 \text{Proof: } \sum_{i=1}^n \hat{y}_i e_i &= \sum_{i=1}^n (\hat{\beta}_0 + \hat{\beta}_1 x_i) e_i \\
 &= \hat{\beta}_0 \underbrace{\sum_{i=1}^n e_i}_0 + \hat{\beta}_1 \underbrace{\sum_{i=1}^n x_i e_i}_0 \\
 &= 0
 \end{aligned}$$

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• The Least-square regression line always passes through the centroid  $(\bar{x}, \bar{y})$

Proof: fitted regression line

$$\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x$$

plug-in  $\bar{x}$  to the RHS.

$$\hat{\beta}_0 + \hat{\beta}_1 \bar{x} = \bar{y} - \hat{\beta}_1 \bar{x} + \hat{\beta}_1 \bar{x} = \bar{y}$$

$\Rightarrow (\bar{x}, \bar{y})$  is on the line

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