

1 Variational Bayes - Normal Mixture Model, Fixed π

We have *iid* observations y_i generated by a two level Normal Mixture Model with means μ_1 and μ_2 and known variance 1, so

$$p(y_i|\mu_1, \mu_2, k_i) = \mathcal{N}(\mu_1, 1)^{k_i} \mathcal{N}(\mu_2, 1)^{1-k_i}.$$

where the latent variable $k_i = 1$ if y_i is drawn from $\mathcal{N}(\mu_1, 1)$ and $k_i = 0$ otherwise.

Further, \mathbf{k} is modelled as *iid* Bernoulli with known parameter π , so

$$p(k_i|\pi) = \pi^{k_i} (1 - \pi)^{1-k_i}.$$

Introducing the priors $p(\pi) \sim U(0, 1)$ and $p(\mu_1, \mu_2) \propto 1$, the joint distribution becomes

$$\begin{aligned} p(y, k, \mu_1, \mu_2|\pi) &= \prod_{i=1}^n p(y_i|k_i, \mu_1, \mu_2, \pi) p(k_i|\pi) p(\pi) p(\mu_1 \mu_2) \\ &\propto \prod_{i=1}^n \left(\frac{1}{\sqrt{2\pi}} \exp \left\{ \frac{-(y_i - \mu_1)^2}{2} \right\} \right)^{k_i} \left(\frac{1}{\sqrt{2\pi}} \exp \left\{ \frac{-(y_i - \mu_2)^2}{2} \right\} \right)^{1-k_i} \\ &\quad \times \pi^{k_i} (1 - \pi)^{1-k_i} \\ \ln(p(y, k, \mu_1, \mu_2|\pi)) &= \sum_{i=1}^N \left[\ln \left(\exp \left\{ \frac{-(y_i - \mu_1)^2}{2} \right\}^{k_i} \right) \right] + \sum_{i=1}^N \left[\ln \left(\exp \left\{ \frac{-(y_i - \mu_2)^2}{2} \right\}^{1-k_i} \right) \right] \\ &\quad + \sum_{i=1}^N k_i \ln(\pi) + \sum_{i=1}^N (1 - k_i) \ln(1 - \pi) \\ &= \sum_{i=1}^N \left[k_i \frac{-(y_i - \mu_1)^2}{2} \right] + \sum_{i=1}^N \left[(1 - k_i) \frac{-(y_i - \mu_2)^2}{2} \right] \\ &\quad + \sum_{i=1}^N k_i \ln(\pi) + \sum_{i=1}^N (1 - k_i) \ln(1 - \pi) + c. \end{aligned} \tag{1}$$

We can take the variational approximation factorisation $q(k_{1:n}, \mu_1, \mu_2) = \prod_{i=1}^n q(k_i) q(\mu_1) q(\mu_2)$, which implies independence of k_i, k_j for $i \neq j$:

It can be shown that the factorisable distribution that minmises the KL Divergence between $q(\theta)$ and $p(\theta|y)$ satisfies

$$q_i \propto \exp(\mathbb{E}_{q_{j \neq i}}(\ln(p(y, x, \theta)))) \quad (2)$$

for all q_i , where y is the observed data, x is a latent variable and θ is a vector of unknown parameters.

Substituting (1) into (2) and ignoring all terms that do not depend on μ_1 yields

$$\begin{aligned} \ln(q(\mu_1)) &= \mathbb{E}_{k_{1:n}} \sum_{i=1}^n -k_i \frac{(y_i - \mu_1)^2}{2} + c \\ &= -\frac{1}{2} \left(\sum_{i=1}^n \mathbb{E}(k_i) (y_i - \mu_1)^2 \right) + c \\ &= -\frac{1}{2} \left(\sum_{i=1}^n \mathbb{E}(k_i) ((y_i - \tilde{y}_1) + (\tilde{y}_1 - \mu_1))^2 \right) + c \\ &= -\frac{1}{2} \left(\sum_{i=1}^n \mathbb{E}(k_i) ((y_i - \tilde{y}_1)^2 + (\tilde{y}_1 - \mu_1)^2 - 2(y_i - \tilde{y}_1)(\tilde{y}_1 - \mu_1)) \right) + c. \end{aligned}$$

Where

$$\tilde{y}_1 = \frac{\sum_{i=1}^n \mathbb{E}(k_i) y_i}{\sum_{i=1}^n \mathbb{E}(k_i)}.$$

Note that

$$\sum_{i=1}^n \mathbb{E}(k_i) (y_i - \tilde{y}_1) = \sum_{i=1}^n \mathbb{E}(k_i) \left(y_i - \frac{\sum_{i=1}^n \mathbb{E}(k_i) y_i}{\sum_{i=1}^n \mathbb{E}(k_i)} \right) = 0,$$

hence

$$\ln(q(\mu_1)) = -\frac{\sum_{i=1}^n \mathbb{E}(k_i) (\tilde{y}_1 - \mu_1)^2}{2} + c.$$

Recognizing the kernel of a Gaussian distribution, we can see that $q(\mu_1) \sim \mathcal{N}(\bar{\mu}_1 = \tilde{y}_1, \lambda_1 = (\sum_{i=1}^n \mathbb{E}(k_i))^{-1})$. Similarly, $q(\mu_2) \sim \mathcal{N}(\bar{\mu}_2 = \tilde{y}_2, \lambda_2 = \sum_{i=1}^n \mathbb{E}(1 - k_i)^{-1})$ with

$$\tilde{y}_2 = \frac{\sum_{i=1}^n \mathbb{E}(1 - k_i) y_i}{\sum_{i=1}^n \mathbb{E}(1 - k_i)}.$$

Through independence, all $q(k_i)$ have the same form,

$$\begin{aligned}
\ln(q(k_i)) &= \mathbb{E}_{\mu_1, \mu_2} \left[k_i \frac{-(y_i - \mu_1)^2}{2} + (1 - k_i) \frac{-(y_i - \mu_2)^2}{2} + k_i \ln(\pi) + (1 - k_i) \ln(1 - \pi) + c \right] \\
&= k_i \frac{\mathbb{E}_{\mu_1} - (y_i - \mu_1)^2}{2} + (1 - k_i) \frac{\mathbb{E}_{\mu_2} - (y_i - \mu_2)^2}{2} + k_i \ln(\pi) + (1 - k_i) \ln(1 - \pi) + c \\
&= k_i \frac{2 \ln(\pi) - ((y_i - \bar{\mu}_1)^2 + \lambda_1)}{2} + (1 - k_i) \frac{2 \ln(1 - \pi) - ((y_i - \bar{\mu}_2)^2 + \lambda_2)}{2} + c \\
q(k_i) &\propto \exp \left\{ \frac{2 \ln(\pi) - ((y_i - \bar{\mu}_1)^2 + \lambda_1)}{2} \right\}^{k_i} \exp \left\{ \frac{2 \ln(1 - \pi) - ((y_i - \bar{\mu}_2)^2 + \lambda_2)}{2} \right\}^{1-k_i}
\end{aligned}$$

We expanded the quadratic term and substituted in $(y_i^2 - \bar{\mu}_j)^2 + \lambda_j$ for $\mathbb{E}(y_i - \mu_j)^2$.

Each k_i has a Bernoulli distribution with parameters $p_i = \exp \left\{ \frac{2 \ln(\pi) - ((y_i - \bar{\mu}_1)^2 + \lambda_1)}{2} \right\}$, and $q_i = \exp \left\{ \frac{2 \ln(1 - \pi) - ((y_i - \bar{\mu}_2)^2 + \lambda_2)}{2} \right\}$.

This gives us the update rules for the Variational Bayes iterations:

$$\begin{aligned}
\bar{\mu}_1 &= \frac{\sum_{i=1}^n y_i p_i / (p_i + q_i)}{\sum_{i=1}^n p_i / (p_i + q_i)} \\
\lambda_1 &= \left(\sum_{i=1}^n \frac{p_i}{p_i + q_i} \right)^{-1} \\
\bar{\mu}_2 &= \frac{\sum_{i=1}^n y_i q_i / (p_i + q_i)}{\sum_{i=1}^n q_i / (p_i + q_i)} \\
\lambda_2 &= \left(\sum_{i=1}^n \frac{q_i}{p_i + q_i} \right)^{-1} \\
p_i &= \exp \left\{ \frac{2 \ln(\pi) - ((y_i - \bar{\mu}_1)^2 + \lambda_1)}{2} \right\} \\
q_i &= \exp \left\{ \frac{2 \ln(1 - \pi) - ((y_i - \bar{\mu}_2)^2 + \lambda_2)}{2} \right\}
\end{aligned}$$