5. Monte-Carlo methods I

Background

- ► The goal of Monte-Carlo methods is the sampling of a process in order to determine some statistical properties
- ► For instance, we toss a coin 4 times. What is the probability to obtain 3 tail and 1 head?
- ► Mathematics gives us the solution:

$$P(3 \text{ head}) = {4 \choose 3} {1 \choose 2}^3 {1 \choose 2}^3 {1 \choose 2}^1 = {1 \over 4}$$

▶ But we could also do a simulation

A Monte-Carlo computer simulation

```
from random import randint
success=0
attempts=10000
for i in range(attempts):
    if randint(0,1)+randint(0,1)+randint(0,1)+randint(0,1)==3:
        success+=1
print "Number of attempts=", attempts
print "Number of success=", success
                              We get for instance:
Number of attempts= 10000
```

Number of success= 2559

More difficult problems

► For the coin tossing problem, no need for a simulation

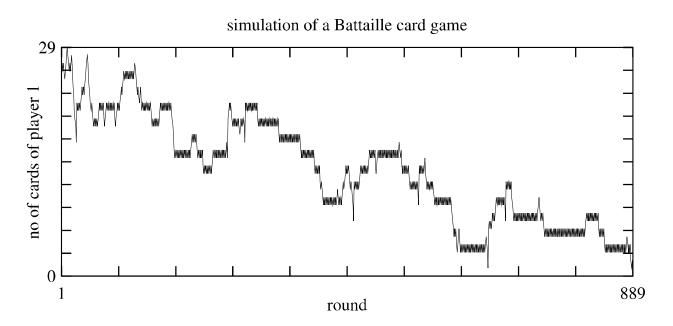
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More difficult problems

- ▶ For the coin tossing problem, no need for a simulation
- ► But we can think of other problems for which probability theory could hardly be applied
- ► For instance: what is the average duration of the card game called "war" (or battle)?

The war card game with 52 cards



Historical note

- ► The method was name in the 1940s by John von Neumann, Stanislaw Ulam and Nicholas Metropolis after the name of the *Monte-Carlo casino*, where Ulam's uncle used to gamble ...and loose his money
- ► The motivation was to find out the probability that a Canfield solitaire will finish successfully.
- ▶ Ulam found it easier to play many Canfield solitaires and estimate the number of successes, rather than trying to apply combinatorics and probability theory.
- ► Then the Monte-Carlo methods was successfully applied to the Manhattan project (nuclear weapon) in the Los Alamos National Laboratory.

[End of module]
Monte-Carlo Methods I

Coming next

Monte-Carlo Methods II

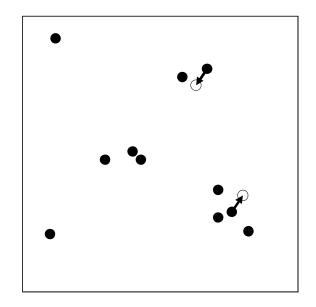
6. Monte-Carlo Methods II

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- \blacktriangleright Let x be a point in this state space. Let us assume that this point moves across the space by jumping randomly to another point x'.
- ▶ The jump from location x to location x' takes place with probability $W_{x\to x'}$. This advanced the system time from t to t+1 (Markov chain)

- We want this process to sample a prescribed probability $\rho(t, x)$. This stochastic process should be at point x at time t with a probability $\rho(t, x)$.
- ▶ How do we choose $W_{x\to x'}$?



$$\rho \propto \exp(-E(x)/k_BT)$$

The probability that our random exportation is at location x at time t is

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- ▶ Let us consider a 1D discrete space: $x \in \mathbf{Z}$.
- ▶ where one can move to the right with probability W_+ , to the left with probability W_- and stay still with probability W_0 .
- ▶ The equation for p(t,x) simplifies to

$$p(t+1,x) = p(t,x-1)W_{+} + p(t,x)W_{0} + p(t,x+1)W_{-}$$

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$$p(t+1,x) = p(t,x-1)W_{+} + p(t,x)W_{0} + p(t,x+1)W_{-}$$

▶ In order to have $p = \rho$, one need $W_+ = W_- = \Delta t D/(\Delta x)^2$ and $W_0 = 1 - 2\Delta t D/(\Delta x)^2 = 1 - W_+ - W_-$, and thus $\Delta t D/(\Delta x)^2 \le 1/2$

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Monte-Carlo simulation of Diffusion

▶ Therefore a random walk is a way to sample a density ρ that obeys the diffusion equation.

Monte-Carlo simulation of Diffusion

- ▶ Therefore a random walk is a way to sample a density ρ that obeys the diffusion equation.
- ▶ With a random walk, it is easy to add obstacles, or aggregation processes, hard to include in the differential equation.

More general case: Master equation

The probability to find the random exploration at location x at time t is p(t,x) given by

$$p(t+1,x) = \sum_{x'} p(t,x')W_{x'\to x}$$

$$= \sum_{x'\neq x} p(t,x')W_{x'\to x} + p(t,x)W_{x\to x}$$

$$= \sum_{x'\neq x} p(t,x')W_{x'\to x} + p(t,x)(1 - \sum_{x'\neq x} W_{x\to x'})$$

$$= p(t,x) + \sum_{x'\neq x} [p(t,x')W_{x'\to x} - p(t,x)W_{x\to x'}]$$

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Detailed balance

In a steady state, the condition $p(x) = \rho(x)$ requires that

$$\sum_{x' \neq x} \left[\rho(x') W_{x' \to x} - \rho(x) W_{x \to x'} \right] = 0$$

We can then choose $W_{x\to x'}$ according to the **detailed balance** condition

$$\rho(x')W_{x'\to x} - \rho(x)W_{x\to x'} = 0$$

Metropolis Rule

Let us consider a physical system at equilibrium whose probability to be in state x is given by the Maxwell-Boltzmann distribution

$$\rho(x) = \Gamma \exp(-E(x)/kT)$$

We can sample this distribution with a stochastic process by choosing $W_{x\to x'}$ according to the **Metropolis rule**:

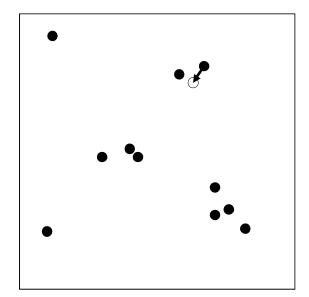
$$W_{x \to x'} = \begin{cases} 1 & \text{si } E' < E \\ \exp[-(E' - E)/kT] & \text{si } E' > E \end{cases}$$

The Metropolis Rule in practice

- ▶ In a gas, one selects one particle at random.
- \blacktriangleright One moves it by an amount Δx .
- \blacktriangleright One computes the energy E' of the gas with this new position.
- ► One accepts this change if

$$rand(0,1) < min(1, exp[-(E'-E)/kT])$$

▶ By sampling ρ with $W_{x\to x'}$, one can compute average physical properties, such as for instance the pressure in the gas.



The Metropolis obeys the detailed balance

Let us assume that E' > E. Detailed balance is obeyed because

$$\rho(x)W_{x\to x'} = \Gamma \exp(-E/kT) \exp[-(E'-E)/kT]$$

$$= \Gamma \exp(-E'/kT)$$

$$= \rho(x') \times 1$$

$$= \rho(x')W_{x'\to x}$$

And similarly if $E' \leq E$

Glauber Rule

This is an alternative to the Metropolis rule. $W_{x\to x'}$ is given by

$$W_{x \to x'} = \frac{\rho(x')}{\rho(x) + \rho(x')}$$

which also clearly obeys detailed balance With $\rho = \Gamma \exp(-E(x)/kT)$, one obtains

$$W_{x\to x'} = \frac{\exp(-E'/kT)}{\exp(-E/kT) + \exp(-E'/kT)}$$

End of module

Monte-Carlo Methods II

Coming next

Monte-Carlo Methods III

7. Monte-Carlo Methods III

Kinetic / Dynamic Monte-Carlo

Let us consider the chemical equations

$$A \stackrel{k_1}{\rightarrow} B \qquad B \stackrel{k_2}{\rightarrow} A$$

They can be written as an ordinary equation

$$\frac{d}{dt} \begin{pmatrix} A \\ B \end{pmatrix} = \begin{pmatrix} -k_1 & k_2 \\ k_1 & -k_2 \end{pmatrix} \begin{pmatrix} A \\ B \end{pmatrix}$$

Analytical solution

$$A(t) = \frac{k_2}{k_1 + k_2} (A_0 + B_0) + \frac{A_0 k_1 - B_0 k_2}{k_1 + k_2} e^{-(k_1 + k_2)t}$$

$$B(t) = \frac{k_1}{k_1 + k_2} (A_0 + B_0) - \frac{A_0 k_1 - B_0 k_2}{k_1 + k_2} e^{-(k_1 + k_2)t}$$

where A_0 and B_0 are the initial concentration of A and B. When $t \to \infty$,

$$A \to A_{\infty} = \frac{k_2}{k_1 + k_2} (A_0 + B_0)$$
 $B \to B_{\infty} = \frac{k_1}{k_1 + k_2} (A_0 + B_0)$

Monte-Carlo Simulation

1 One defines a time step Δt , small enough so that $k_1 \Delta t$ et $k_2 \Delta t$ are smaller than 1. They are the **probabilities** that, during Δt , one A particle get transformed into one B particle, or conversely.

Monte-Carlo Simulation

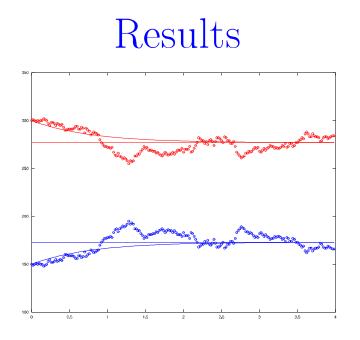
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 - 5 One repeats (2)-(4) until $t = t_{max}$



 $\Delta t = 0.02$ and $k_1 = 0.5$, $k_2 = 0.8$.

The Monte-Carlo simulation fluctuate around analytic solution. We should average over several runs

Let r_i be the rate at which the possible events occur in the system. $i = 1, \ldots n$.

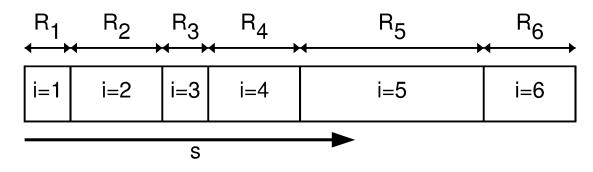
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- ▶ Note that here Δt is calculated according to a decreasing exponential distribution. It gives the average time of occurrence of the next event.
- ▶ Only one event takes place during the time interval Δt .

End of module

Monte-Carlo method III

End of Week 1

Thank you for your attention!