

## 3.2 Uniform, Exponential, Gamma Distributions

# Exponential Distribution

Consider a Poisson process with an expected number of occurrences,  $\lambda$ , in a given interval.

Let  $W$  be the waiting time until the first occurrence. Then  $W$  follows an **exponential distribution**.

(also the waiting time between occurrences)

Instead, if we count the number of these occurrences,  $X$ ,  
 $X \sim \text{Pois}(\lambda)$ .

Exponential Distribution:

$$X \sim \text{Exp}(\theta)$$

$$f(x) = \frac{1}{\theta} e^{-x/\theta}, \quad 0 \leq x < \infty$$

$$E[X] = \theta$$

$$\text{Var}[X] = \theta^2$$

Alternatively,  $f(x) = \lambda e^{-x\lambda}$ ,  $0 \leq x < \infty$

# Finding $E[X]$ and $\sigma^2$ for an exponential

$$\begin{aligned} M(t) &= \int_0^\infty e^{tx} \left(\frac{1}{\theta}\right) e^{-x/\theta} dx = \lim_{b \rightarrow \infty} \int_0^b \left(\frac{1}{\theta}\right) e^{-(1-\theta t)x/\theta} dx \\ &= \lim_{b \rightarrow \infty} \left[ -\frac{e^{-(1-\theta t)x/\theta}}{1-\theta t} \right]_0^b = \frac{1}{1-\theta t}, \quad t < \frac{1}{\theta}. \end{aligned}$$

$$M'(t) = \frac{\theta}{(1-\theta t)^2} \qquad M''(t) = \frac{2\theta^2}{(1-\theta t)^3}.$$

$$E[X] = M'(0) = \theta$$

$$E[X^2] = M''(0) = 2\theta^2$$

$$Var[X] = \theta^2$$

# Memoryless Property

The **exponential** and **geometric** distributions are **memoryless**.

## Exponential Distribution

- Radioactive decay
- How long a fly will stay on a table until it takes off?

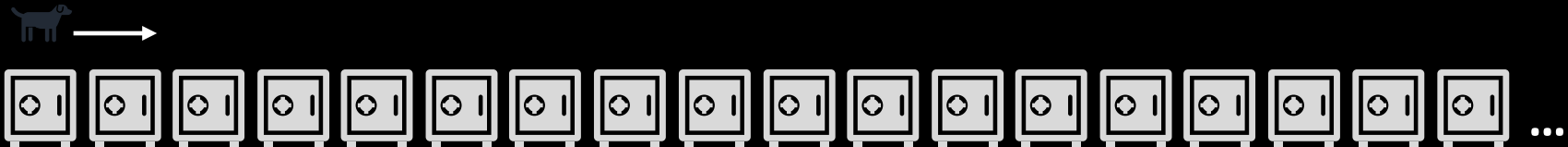
## Geometric Distribution

- How many more times do I need to roll a die until my first success

# Memoryless Property Example

Chloe walks down an infinite hallway of safes.

- Each safe has a different code (1000 possibilities).
  - $P[\text{correct code}] = 1/1000$
- Chloe only tries one code per safe
- Let  $X$  be the number of safes Chloe **still** needs to try before she successfully opens one.  $E[X] = 1000$



# Memoryless Property

Discrete Memorylessness (Geometric):

if  $X$  is the **total** number of trials required for the first success,

$$P[X > m + n \mid X > m] = P[X > n]$$

Continuous Memorylessness (Exponential):

if  $X$  is the **total** time required for the first success,

$$P[X > t + s \mid X > t] = P[X > s]$$

Note: Most random variables/phenomena are not memoryless.

We generally obtain and update information over time.

e.g. (not memoryless):

- A car engine's remaining life (how many miles it has left).
- The amount of time left until class ends.



Notes: memoryless property and Bayes Rule:

$$P[X > m + n \mid X > m] = P[X > n]$$

Suppose  $X \sim \text{Exp}(\theta = 4)$

What is  $P[X > 9]$ ?

What is  $P[X > 9 \mid X > 5]$ ?

# Example 1 (textbook)

## Example 3.2-2

Customers arrive in a certain shop according to an approximate Poisson process at a mean rate of 20 per hour. What is the probability that the shopkeeper will have to wait more than 5 minutes for the arrival of the first customer? Let  $X$  denote the waiting time in minutes until the first customer arrives, and note that  $\lambda = 1/3$  is the expected number of arrivals per minute. Thus,

$$\theta = \frac{1}{\lambda} = 3$$

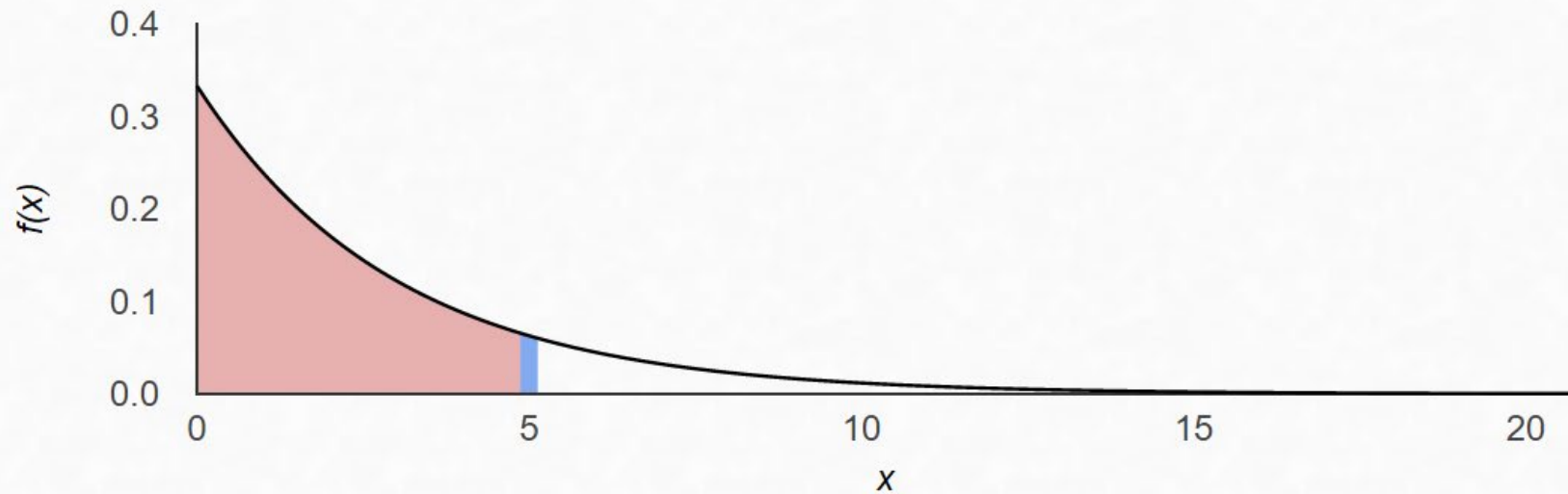
$$f(x) = \frac{1}{3} e^{-(1/3)x}, \quad 0 \leq x < \infty.$$

$$P(X > 5) = \int_5^{\infty} \frac{1}{3} e^{-(1/3)x} dx = e^{-5/3} = 0.1889.$$

Rate  $\lambda = 1/3$

$x = 5$

$P(X < x) = 0.81112$



## Example 2

Suppose the length of time,  $X$ , between occurrences follows an exponential distribution with mean = 5 sec.

- A) Give the pdf and support (sample space) of  $X$ .
- B) What is the probability that it will take more than 10 seconds for the first occurrence (to happen)? (0.135)

# Example 3

Suppose an electronic component has a lifespan which can be modeled as an exponential distribution, with mean = 500 hours.

A) Find the pdf and cdf of this distribution

B) Find  $P[X > x]$   $(e^{-x/500})$

C) If this component has already lasted 200 hours, find the probability that it will last at least 600 hours total.  $(e^{-4/5})$

# Example 4 (Exponential Distribution Derivation)

Given a Poisson Process with rate  $\lambda$ :

Let  $X$  denote the # of occurrences in a time of length  $w$ .

- What is distribution of  $X$ ?

$$X \sim \text{Poisson}(\quad).$$

- What is  $P[X = 0]$ ?

(no occurrences in an interval of length  $w$ )

Let  $W$  represent the *waiting time* until the 1<sup>st</sup> occurrence in this Poisson Process.

$$\begin{aligned} F(w) &= P(W \leq w) = 1 - P(W > w) \\ &= 1 - P(\text{no occurrences in } [0, w]) \\ &= 1 - e^{-\lambda w}, \end{aligned}$$

$$F'(w) = f(w) = \lambda e^{-\lambda w}$$

# The Gamma Distribution

Consider a Poisson process with rate  $\lambda$ :

Let a random variable,  $X$ , denote the waiting time until the  $\alpha$ th occurrence.

$X$  follows a Gamma Distribution.

# The Gamma Function, $\Gamma$

$$\Gamma(t) = \int_0^{\infty} y^{t-1} e^{-y} dy, \quad 0 < t.$$

← This is the definition of the gamma function

$$\begin{aligned} \Gamma(t) &= \left[ -y^{t-1} e^{-y} \right]_0^{\infty} + \int_0^{\infty} (t-1) y^{t-2} e^{-y} dy \\ &= (t-1) \int_0^{\infty} y^{t-2} e^{-y} dy = (t-1) \Gamma(t-1). \end{aligned}$$

$$\Gamma(n) = (n-1) \Gamma(n-1) = (n-1)(n-2) \cdots (2)(1) \Gamma(1).$$

$$\Gamma(1) = \int_0^{\infty} e^{-y} dy = 1.$$

When  $n$  is an integer

$$\Gamma(n) = (n-1)!$$



# Gamma Distribution

$X \sim \text{Gamma}(\alpha, \theta)$

$$f(x) = \frac{1}{\Gamma(\alpha)\theta^\alpha} x^{\alpha-1} e^{-x/\theta}, \quad 0 \leq x < \infty$$

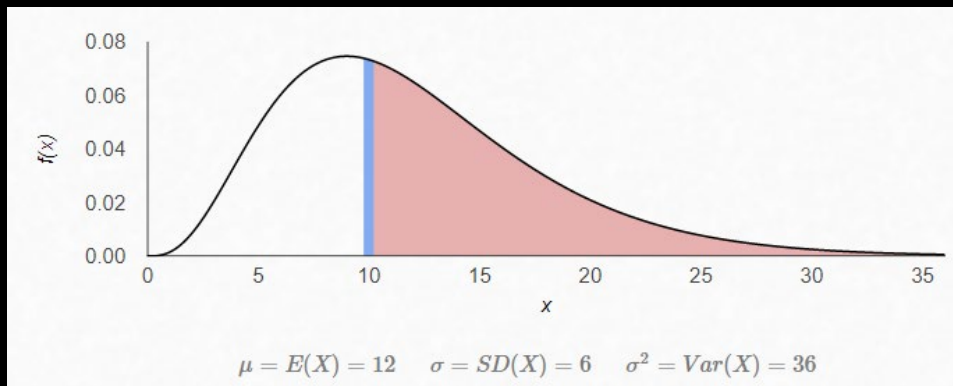
$$E[X] = \alpha\theta$$

$$\text{Var}[X] = \alpha\theta^2$$

# Gamma Example

Customers arrive in a shop according to a Poisson process with a mean rate of 20 per hour. What is the probability that the shopkeeper will have to wait more than 10 minutes for the arrival of the 4<sup>th</sup> customer?

$$\int_{10}^{\infty} \frac{1}{\Gamma(4)3^4} x^{4-1} e^{-x/3} dx = 0.57$$



What is the probability that the shopkeeper will have to wait between than 5 and 20 minutes for the arrival of the 3<sup>rd</sup> customer?

# Uniform Distribution

$$X \sim Unif(a, b)$$

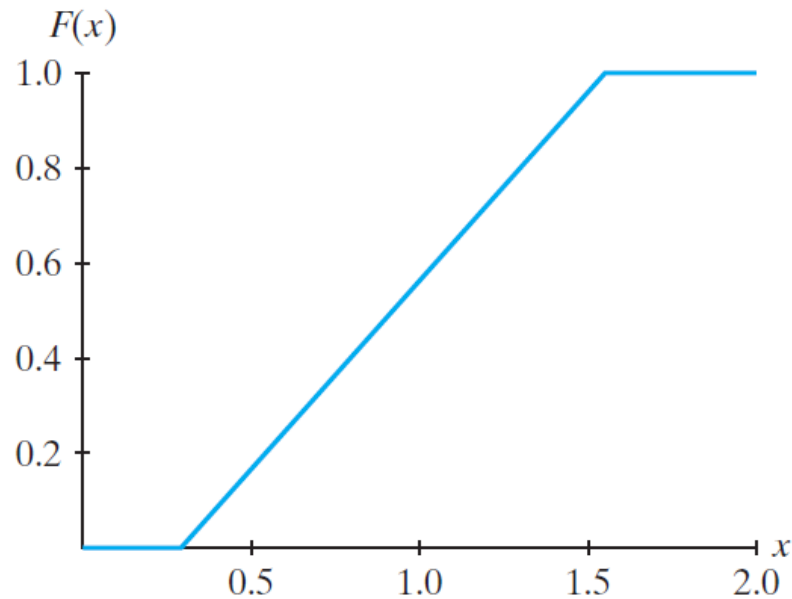
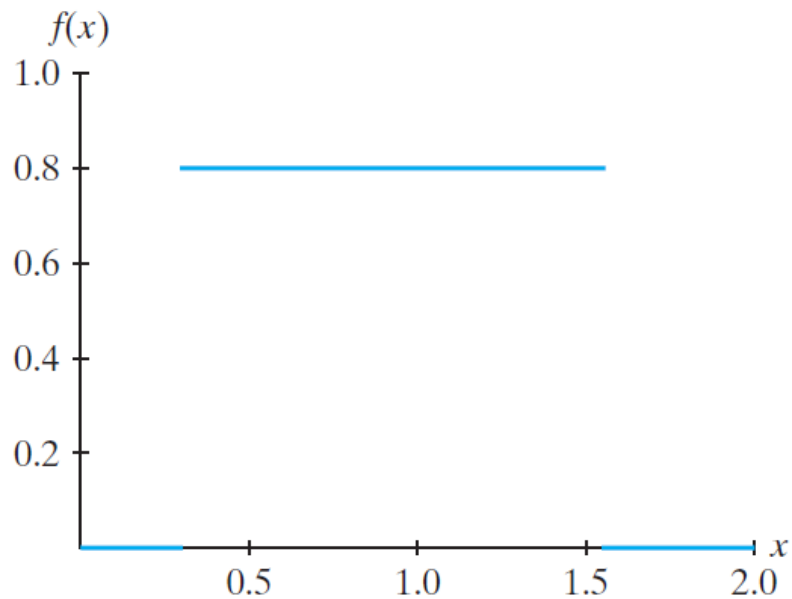
A random variable,  $X$ , has a **uniform distribution** if its pdf is equal to a constant on its support.

$$f(x) = \frac{1}{b-a}, \quad a \leq x \leq b.$$

$$E[X] = \frac{a+b}{2}$$

$$Var[X] = \frac{(b-a)^2}{12}$$

$$f(x) = \frac{1}{b-a}, \quad a \leq x \leq b$$



**Figure 3.1-1** Uniform pdf and cdf

# Uniform Example

Suppose that you arrive at the intersection of Green and Wright, it takes between 0 and 30 seconds for the walk sign to come on. Assume that the time it takes,  $X$ , follows a uniform distribution.

What is the probability that it takes between 10 and 20 seconds for the walk sign to come on?  $(1/3)$

What are  $E[X]$  and  $Var[X]$ ?  $(15), \left(\frac{30^2}{12}\right)$

# Transformation Theorem

**Theorem** Let  $U \sim \text{Uniform}(0,1)$  and  $F$  be a CDF which is strictly increasing. Also, consider a random variable  $X$  defined as

$$X = F^{-1}(U).$$

Then,

$$X \sim F \quad (\text{The CDF of } X \text{ is } F)$$

Proof:

$$\begin{aligned} P(X \leq x) &= P(F^{-1}(U) \leq x) \\ &= P(U \leq F(x)) \quad (\text{increasing function}) \\ &= F(x) \end{aligned}$$

# Transformation Theorem Example 1

Goal: Generate a sample from an exponential

Random Variable with parameter  $\theta = 1$

**Theorem** Let  $U \sim \text{Uniform}(0,1)$  and  $F$  be a CDF which is strictly increasing. Also, consider a random variable  $X$  defined as

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Then,

$$X \sim F \quad (\text{The CDF of } X \text{ is } F)$$

Proof:

$$\begin{aligned} P(X \leq x) &= P(F^{-1}(U) \leq x) \\ &= P(U \leq F(x)) \quad (\text{increasing function}) \\ &= F(x) \end{aligned}$$

$U = \text{runif}(1; \text{min} = 0; \text{max} = 1);$

$X = -\log(1-U)$  OR  $X = -\log(U)$

$$F(x) = 1 - e^{-x} \quad x > 0$$

$$U \sim \text{Uniform}(0,1)$$

$$X = F^{-1}(U)$$

$$= -\ln(1 - U)$$

$$X \sim F$$



# Notes: Transformation Example 2

Use the inverse transform method to simulate  $n$  random samples from  $f(x) = 3x^2$ ,  $0 < x < 1$ .

$$F(x) = x^3, 0 < x < 1$$

$$F^{-1}(u) = u^{1/3}$$

```
n = 10000 #make up a number
u = runif(n)
my_sim = u^(1/3)
windows();hist(my_sim)
x = seq(from=0, to=1, by = 0.01)
lines(x, 3x^2)
```