MATH-517: Assignment 3

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2025-10-05

Theoretical exercise

Define the quadratic form

$$\mathcal{Q}(\beta) = \sum_{i=1}^n K\left(\frac{X_i - x}{h}\right) (Y_i - \beta_0 - \beta_1 (X_i - x))^2 = (Y - X\beta)^\mathrm{T} W(Y - X\beta),$$

where
$$W=\mathrm{Diag}\left(W_{1},...,W_{n}\right)=\mathrm{Diag}\left(K\left(\frac{X_{1}-x}{h}\right),...,K\left(\frac{X_{n}-x}{h}\right)\right).$$

Using matrix calculus, we have that $\frac{\partial \mathcal{Q}(\beta)}{\partial \beta} = 2X^TWX\beta - 2X^TWY$ so that the normal equations are given by $X^TWX\hat{\beta} = X^TWY$. Assuming that X^TWX is invertible, we immediately obtain from the normal equations that $\hat{\beta} = (X^TWX)^{-1}X^TWY$.

We will start off by compute some terms, $X^{T}WX$ and $X^{T}WY$ separately.

$$\begin{split} X^{\mathrm{T}}WX &= \begin{bmatrix} 1 & \cdots & 1 \\ X_1-x & \cdots & X_n-x \end{bmatrix} \begin{bmatrix} W_1 & & \\ & \ddots & \\ & & W_n \end{bmatrix} \begin{bmatrix} 1 & X_1-x \\ \vdots & \vdots \\ 1 & X_n-x \end{bmatrix} \\ &= \begin{bmatrix} 1 & \cdots & 1 \\ X_1-x & \cdots & X_n-x \end{bmatrix} \begin{bmatrix} W_1 & W_1(X_1-x) \\ \vdots & \vdots \\ W_n & W_n(X_n-x) \end{bmatrix} \end{split}$$

such that

$$X^{\mathrm{T}}WX = \begin{bmatrix} \sum_{i=1}^{n} W_{i} & \sum_{i=1}^{n} W_{i}(X_{i}-x) \\ \sum_{i=1}^{n} W_{i}(X_{i}-x) & \sum_{i=1}^{n} W_{i}(X_{i}-x)^{2} \end{bmatrix}$$

Define $\tilde{S}_{n,k}(x) := nhS_{n,k}(x)$ for k = 0, 1, 2, where $S_{n,k}(x)$ are defined as in the README. In view of this definition, $X^{\mathrm{T}}WX = \begin{bmatrix} \tilde{S}_{n,0} & \tilde{S}_{n,1} \\ \tilde{S}_{n,1} & \tilde{S}_{n,2} \end{bmatrix} = nh\begin{bmatrix} S_{n,0} & S_{n,1} \\ S_{n,1} & S_{n,2} \end{bmatrix}$.

We now turn to $X^{\mathrm{T}}WY$.

$$\begin{split} X^{\mathrm{T}}WY &= \begin{bmatrix} 1 & \cdots & 1 \\ X_1 - x & \cdots & X_n - x \end{bmatrix} \begin{bmatrix} W_1 \\ & \ddots \\ & & W_n \end{bmatrix} \begin{bmatrix} Y_1 \\ \vdots \\ Y_n \end{bmatrix} = \begin{bmatrix} W_1 & \cdots W_n \\ W_1(X_1 - x) & \cdots & W_n(X_n - x) \end{bmatrix} \begin{bmatrix} Y_1 \\ \vdots \\ Y_n \end{bmatrix} \\ &= \begin{bmatrix} \sum_{i=1}^n W_i Y_i \\ \sum_{i=1}^n W_i (X_i - x) Y_i \end{bmatrix} \end{split}$$

Defining $T_0 = \sum_{i=1}^n W_i Y_i$ and $T_1 = \sum_{i=1}^n W_i (X_i - x) Y_i$, we have that $X^T W Y = \begin{bmatrix} T_0 & T_1 \end{bmatrix}^T$. We now invert $X^T W X$. Since $X^T W X$ is a 2 × 2 matrix, the inverse is

$$\begin{bmatrix} \tilde{S}_{n,0} & \tilde{S}_{n,1} \\ \tilde{S}_{n,1} & \tilde{S}_{n,2} \end{bmatrix}^{-1} = \frac{1}{D} \begin{bmatrix} \tilde{S}_{n,2} & -\tilde{S}_{n,1} \\ -\tilde{S}_{n,1} & \tilde{S}_{n,2} \end{bmatrix}, \text{ where } D = \tilde{S}_{n,0} \tilde{S}_{n,2} - \tilde{S}_{n,1}^2 \text{ is the determinent.}$$

Recall that we are interested in $\hat{\beta}_0$ and in view of the above reasoning,

$$\hat{\beta} = \frac{1}{D} \begin{bmatrix} \tilde{S}_{n,2} & -\tilde{S}_{n,1} \\ -\tilde{S}_{n,1} & \tilde{S}_{n,0} \end{bmatrix} \begin{bmatrix} T_0 \\ T_1 \end{bmatrix} \text{ such that } \hat{\beta}_0 = \frac{1}{D} \left(\tilde{S}_{n,2} T_0 - \tilde{S}_{n,1} T_1 \right) = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left(\tilde{S}_{n,2} - \tilde{S}_{n$$

Recall that $\tilde{S}_{n,k}=nhS_{n,k}$. It follows that $D=(nh)^2\left(S_{n,0}S_{n,2}-S_{n,1}^2\right)$. It immediately follows that

$$\hat{\beta}_0 = \sum_{i=1}^n \frac{nh}{(nh)^2} \frac{W_i \left(S_{n,2} - S_{n,1}(X_i - x)\right)}{S_{n,0} S_{n,2} - S_{n,1}^2} Y_i = \sum_{i=1}^n \frac{1}{nh} \frac{W_i (S_{n,2} - S_{n,1}(X_i - x)}{S_{n,0} S_{n,2} - S_{n,1}^2} Y_i = \sum_{i=1}^n w_{n,i} Y_i,$$

where

$$w_{n,i} = w_{n,i}(x) = \frac{1}{nh} \frac{K\left(\frac{X_i - x}{h}\right) \left(S_{n,2}(x) - S_{n,1}(x)(X_i - x)\right)}{S_{n,0}(x)S_{n,2}(x) - S_{n,1}(x)^2}.$$

With this result, we answer both question (1) and (2) at the same time.

It remains to show that $\sum_{i=1}^{n} w_{n,i}(x) = 1$. We proceed as follows:

$$\sum_{i=1}^n w_{n,i}(x) = \frac{1}{nh} \sum_{i=1}^n \frac{W_i \left(S_{n,2}(x) - (X_i - x) S_{n,1}(x) \right)}{S_{n,0}(x) S_{n,2}(x) - S_{n,1}(x)^2}$$

$$=\frac{1}{nh}\frac{S_{n,2}(x)}{S_{n,0}(x)S_{n,2}(x)-S_{n,1}(x)^2}\sum_{i=1}^nW_i-\frac{1}{nh}\frac{S_{n,1}(x)}{S_{n,0}(x)S_{n,2}(x)-S_{n,1}(x)^2}\sum_{i=1}^nW_i(X_i-x)$$

In view of the definitions of $S_{n,0}(x)$ and $S_{n,1}(x)$, we obtain

$$\sum_{i=1}^n w_{n,i}(x) = \frac{S_{n,2}(x)S_{n,0}(x)}{S_{n,0}(x)S_{n,2}(x) - S_{n,1}(x)^2} - \frac{S_{n,1}(x)^2}{S_{n,0}(x)S_{n,2}(x) - S_{n,1}(x)^2} = 1,$$

which concludes the third (3) part of the exercise.

This concludes the theoretical exercise.

Practical exercise

Unless stated otherwise, your answer to the practical part should include the following elements:

- Description of the aim of the simulation study.
- Description of the different quantities that intervene in your simulation study. Explain how these quantities (fixed or random) are defined and the reasoning behind the choices you made.
- Description of your findings using appropriate graphics and/or tables (with a caption!) that are well commented in the text.

The code should not appear in the PDF report, unless there is a specific reason to include it.

Structure and details:

Aim of the simulation study:

The goal is to observe the impact of various parameters on the estimation of the IMSE-optimal bandwith h_{IMSE} , σ^2 and θ_{22} . In particular how it varies with respect to

- various values (combinations) of the Beta function, $\alpha, \beta > 0$ (see below)
- various values (combinations) of sample size n and block size / number of blocks N (see below)

Fixed quantities:

• ADD IF NECESSARY

Code architecture:

The <u>technical functions</u> are located in a file functions.py with usual alias fct. There, we have the following functions:

- generate_sample takes inputs alpha, beta, n_samples, sigma_2 which respectively represent α , β , the number of samples n and the variance of the error ϵ , σ^2 . This function returns two one-dimensional arrays, covariate and response representing X and Y respectively.
- estimate_parameters takes inputs covariate, response, bandwith, p, reg. They represent respectively X, Y, a default bandwith needed to perform an initial KDE and p which represents the order. Note that p=1 by default and reg is set to 1e-5. It computes the estimated $\hat{\beta}$ by solving the normal equation. To efficiently compute the solution, we use Cholesky decomposition and np.linalg.solve instead of np.linalg.inv, which is significantly more stable and efficient in this situation. For even more stability, we add a small regularization (reg) before using np.linalg.solve. Throughout the calls of the functions, the reg parameter is never changed and remains set to its default value of 1e-5. The returns is a function beta_est that, given a array of covariates, computes the fitted values.
- estimate_sigma_theta takes inputs covariate, repsonse, bandwith, N_blocks which respectively represent X,Y, a default bandwith needed to perform the initial KDE and the number of blocks to compute estimates of $\hat{\theta}_{22}$ and $\hat{\sigma}^2$. Note that in the README, it is usually mentioned the size of the blocks instead of the number. We decided to implement in terms of the number of blocks. This choice is due to (my) reasoning logic. The computation of the size, given the number of blocks is given by size = sample_size / N_blocks. To compute blocks, we compute the size, and then split the data in blocks [B1 | ... | Bk | Bextra] Bextra is the remaining number of samples. This function returns an estimate of θ_{22} and σ^2 , following the procedure recommended in the README.
- est_h_IMSE_support takes inputs sigma_2_hat, theta_22, size, covariate as inputs which respectively represent $\hat{\sigma}^2$, $\hat{\theta}_{22}$ and the sample size n. It returns h_{IMSE} as given in the README. Note that size is not a necessary argument as size=len(covariate). This choice was made for clarity. The covariate array is a necessary argument as the support of the realizations is given by support=max(covariate)-min(covariate) (note

that this is always a nonnegative quantity and a positive quantity with probability one). This function returns an estimate of h_{IMSE} .

- compute_mallow_C_p takes inputs covariate, response, fixed_bandwith, N_blocks (as previously seen). It computes $C_p(N)$ as suggested in the README and returns its value.
- optimal_mallow takes inputs covariate, repsonse, fixed_bandwith, max_N_blocks. The first two inputs are as seen previously and the third is the same as bandwith seen previously. The last argument represents the maximal number of blocks on which the function computes the C_p statistics (for m "model" selection). The function returns the number of blocks that minimized the Mallow C_p statistic. The function calls compute_mallow_C_p for each block size.
- h_IMSE_Cp_optimized takes inputs covariate, reponse, number_of_samples, default_bandwith, max_number_of_blocks, which respectively represent X, Y, n, the fixed bandwith as before (previously called bandwith or fixed_bandwith and the maximum number of blocks to test (previously called max_N_blocks. The function calls optimal_mallow and re-estimates σ^2 and θ_{22} , given the Mallow-optimal number of blocks (the return of optimal_mallow). This function returns an estimation of σ^2 and θ_{22} with Mallow-optimal selection of the number of blocks.
- simulate takes inputs alpha, beta, number_of_samples, error_variance, default_bandwith, number_of_blocks, which respectively represent $\alpha,\beta,n,\sigma^2,$ a default bandwith (as previously seen) and a number of blocks. It first calls generate_sample with $\alpha,\beta,n,\sigma^2,$ which returns covariate, response. With that return, it calls estimate_sigma_theta with the default bandwith and the number of blocks. Finally, it calls est_h_IMSE_support with the computed quantities. This function returns the estimations of h_{IMSE},σ^2 and $\theta_{22}.$ We mostly use this function to ease the production of plots.

The <u>plotting functions</u> are accessible through a file called **plot_results.py**. There, one can use any of the functions below:

- plot_simple_fit takes inputs covariate, response, fixed_bandwith, sigma_2, representing X, Y, the fixed bandwith as previously seen and σ^2 (for the plot title). It returns a plot showing visually how the choice of bandwith impacts the fits. To produce this plot simply, refer to the subsection below "How to simply plot results". An example of call would be:
 - py plot_results.py simple_plot 0.1 0.2 10000 0.2 1
- plot_alpha_beta_impact takes inputs alphas, betas, error_variance, default_bandwith, number_of_blocks, which respectively represent a list of alphas $[\alpha_1,...,\alpha_{n_\alpha}]$, as list of betas $[\beta_1,...,\beta_{n_\beta}]$, the default bandwith as seen previously and error_variance represents σ^2 . This function produces three heat maps of the values of estimated

 h_{IMSE} , σ^2 and θ_{22} for different values of α_i , β_j . To produce this plot simply, refer to the subsection below "**How to simply plot results**". An example of call would be:

- py plot_results.py alpha_beta_impact 100 1.0 0.1 5
- plot_sample_blocks_impact takes inputs number_of_samples_range, number_of_blocks, alpha, beta, error_variance, default_bandwith, which respectively represent a list $[n_1,...,n_{k_n}]$ of values for the number of samples n, a list $[N_1,...,N_{k_N}]$ of values for the number of blocks N, fixed values of α , β and σ^2 . This function produces three heat maps of the values of estimated h_{IMSE} , σ^2 and θ_{22} for different values of n_i and N_j . To produce this plot simply, refer to the subsection below "How to simply plot results". An example of call would be:
 - py plot_results.py sample_blocks_impact 0.1 0.2 1.0 0.1
- plot_sample_size_impact_mallow takes inputs alpha, beta, max_number_of_samples, step, error_variance, default_bandwith, max_number_of_blocks, as seen previously. max_number_of_blocks is set to 10 by default. The additional new parameter represents the number of steps between two different sample sizes. The list of sample sizes to test are then given by n_range = np.arange(start=default_start, stop=max_number_of_samples, step=step). This function produces three (sub) plots representing how the estimations of h_{IMSE} , σ^2 and θ_{22} evolve when computed on n samples with Mallow-optimal number of blocks. To produce this plot simply, refer to the subsection below "How to simply plot results". An example of call would be:
 - py plot_results.py sample_size_mallow_evolution 0.1 0.2 2000 10 1.0
 0.1

How to simply plot results:

To call them, use the console with one of the following keywords key∈ { simple_plot, alpha_beta_impact, sample_blocks_impact, plot_sample_size_impact_mallow } and call as follows: py plot_results.py key parameter1 parameter2 parameter3 ...

where parameter1 parameter2 ... are the functions parameters, detailed below:

- when key=simple_plot: The parameters are in the following order alpha, beta, number_of_samples, fixed_bandwith, variance, representing respectively α, β, n , the fixed bandwith (as previously mentioned) and σ^2 .
- when key=alpha_beta_impact: The parameters are in the following order number_of_samples, error_variance, default_bandwith, number_of_blocks, representing respectively the number of samples n to test, σ^2 , the default bandwith (as previously mentioned) and the number of blocks.
- when key=sample_blocks_impact: The parameters are in the following order alpha, beta, error_variance, default_bandwith, as previously seen.

• when key=plot_sample_size_impact_mallow: The parameters are in the following order alpha, beta, max_number_of_samples, step, error_variance, default_bandwith.

When a call is successful, the console will display TASK: <followed by the actual task it is performing > and will print TASK completed when the plot has been produced.

Findings and reports:

We start off with a simple visualization of the setup for fixed α, β, n . The plot below showcases how different bandwith choices visually impact the fit. In orange is the true m(x). In red is a (relatively good fitting) line. We observe that as the bandwith goes closer to zero, the fits become better. However, when to close to zero, the fitted function appears to be to sensible and does not provide a robust fit to the truth.

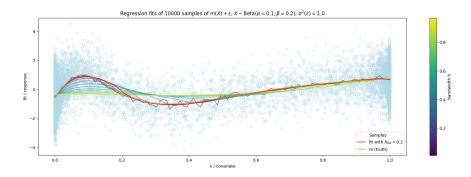


Figure 1: Visual plot of bandwith impact, call of py plot_results.py simple_plot 0.1 0.2 10000 0.2 1

Here below is another plot, after calling py plot_results.py simple_plot 10 10 10000 0.07 1

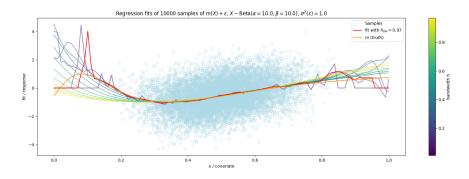


Figure 2: Visual plot of bandwith impact, call of py plot_results.py simple_plot 10 10 100000 0.07 1

NB: Note that to produce these variations, one must directly change export name of the .png files.

• The choice of N should possibly depend on α , β (and m) as the concentrations of samples will vary in [0, 1]. Given m and depending on N, we may not be able to capture the volatility of θ_{22} accurately.

Here below are heatmaps of the results of estimations of h_{IMSE} , σ^2 , θ_{22} against various values of α , β . The x-axis represents β and the y-axis represents α . Here the alphas and betas are chosen in np.arange(0.1, 10, step = 0.5). We observe what seems to be a boundary effect when either α or β is close to zero on the estimation of h_{IMSE} and θ_{22} in particular. For h_{IMSE} , when α or β (or both) are close to zero, the values of h_{IMSE} are of larger magnitude in comparison to the other cases. The inverted observation can be made for θ_{22} 's estimation where the values are lower when α , β are close to zero in comparison to the other case.

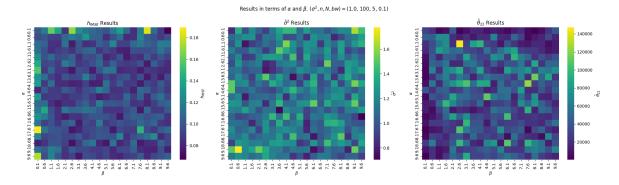


Figure 3: Impact of the choices of alpha and beta on estimations, call of py plot_results.py alpha_beta_impact 100 1.0 0.1 5

To better emphasize on this boundary effect, we recomputed the heatmaps for α, β taking values in np.arange(0.01, 5, step = 0.5). There, the boundary effect on the estimation of h_{IMSE} and θ_{22} are even more noticeable.

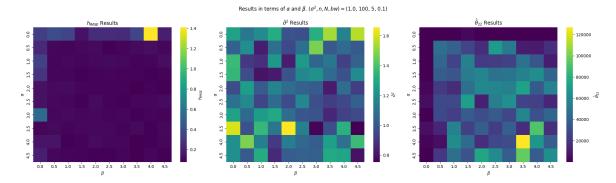


Figure 4: Impact of the choices of alpha and betas, close to zero

Note that in both cases, the estimations of σ^2 remain relatively stable between 0.8 and 1.2 approximately.

To complete the observations around the impacts of α, β on the estimations, we perform a last run for α, β taking values in np.arange(0.01, 1, step = 0.01). There, we observe a fade from the left and upper boundaries to the center for θ_{22} , and a strong boundary effect on h_{IMSE} , as already expected. We additionally observe that the estimations of h_{IMSE} remain very stable for values of α, β far enough from zero (both at the same time) and that the estimations of σ^2 are reasonably stable on average.

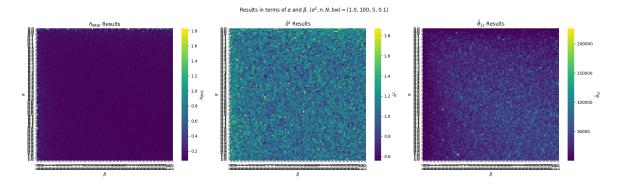


Figure 5: Impact of the choices of alpha and beta with increased granularity and closer to zero

In conclusion,

- The estimation of h_{IMSE} , θ_{22} are more stable for values of α, β far enough from zero (boundary effect). The estimation of h_{IMSE} has very few extreme values (so low variance) for values of α, β large enough.
- The estimation of σ^2 appears to not be affected a lot by changes of α, β .

• initial estimations are that h_{IMSE} is around 0.2 on average, σ^2 is, as expected, around 1.0 and θ_{22} appears to be around 50000 on average.

NB: Note that to produce these variations, one must directly comment / uncomment parts of the code as well as change export name of the .png files. The default computation is made for values of α, β in np.arange(0.1, 10, step = 0.5).

We now take a look at how the number of blocks (or similarly the size of the blocks) and the sample size impact the estimations.

As already mentioned, we think of number of blocks instead of size of blocks. As the number of blocks decrease, the size increases and vice-versa. First notice that we obtain on-average inconsistent estimations of h_{IMSE} and σ^2 when the sample size is low and especially when n s small and N is large. For θ_{22} on the other hand, the computations appear to be inconsistent in comparison to the average when the number of blocks is small. It becomes more stable as the number of blocks increase. This is an observation that one can also make for h_{IMSE} and σ^2 .

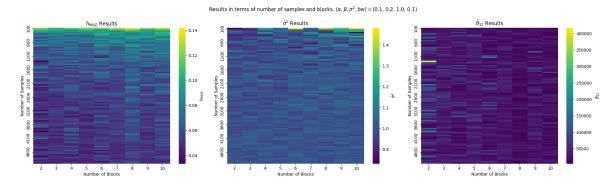


Figure 6: Impact of sample size and number of blocks on estimations (color map)

In conclusion, for this choice of $(\alpha, \beta) = (0.1, 0.2)$,

- As expected, the sample size positively increases the on-average consistency of the estimations.
- The size of the blocks decreases (and the number of blocks increases) the estimations appear to become stable when n (the sample size) is large enough. However, this effect does not seem to be very pronounced as soon as the sample size is large enough.

Consider the case where $(\alpha, \beta) = (10, 10)$ below.

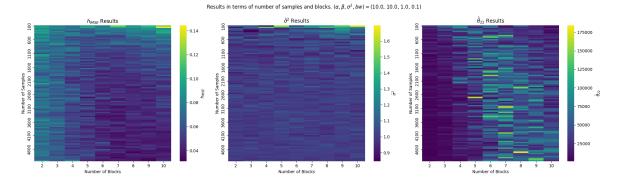


Figure 7: Impact of \$n\$ and \$N\$ when \$\alpha,\beta=10\$

We observe that

- Similarly to before, as n increases, the estimations become more stable in comparison to the average.
- On the other hand, the choice of N appears to impact the stability significantly (especially for θ_{22} and h_{IMSE}) in comparison to the previous case where $(\alpha, \beta) = (0.1, 0.2)$. N will change significantly how the estimates are captured when the density of samples move (this is controlled by α, β . We can deduce that N should (or at least could) depend on α, β .

NB: Note that to produce these variations, one must directly change export name of the .png files.

Let us now take a look at how the estimations behave against the sample size when the number of blocks (or similarly the size) is selected using Mallow's C_p selection criterion. This is a reasonable way to compare the impact of sample sizes alone.

We observe, as expected and previously observed, that small sample sized (even with optimal number of blocks) yield unreasonable estimations in comparison to the average. Moreover, the estimation of h_{IMSE} appears to average around the value of 0.07 after the elbow. The estimation of σ^2 also becomes stable and converges to, it appears, 1.0 (as expected). Furthermore, as already observed previously, the estimation of θ_{22} is relatively unstable.

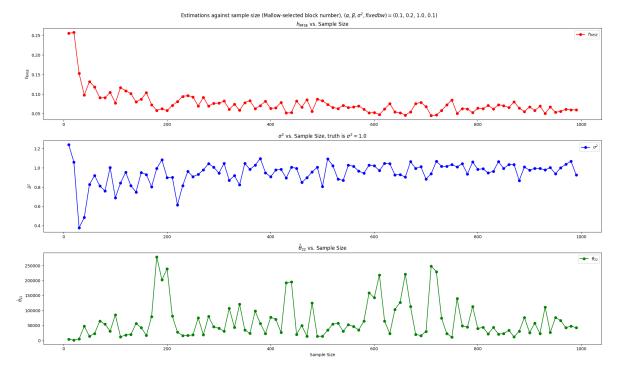


Figure 8: Estimations against sample size with Mallow-selected block number / size

For confirmation purposes, here is a version going up to 2000 samples:

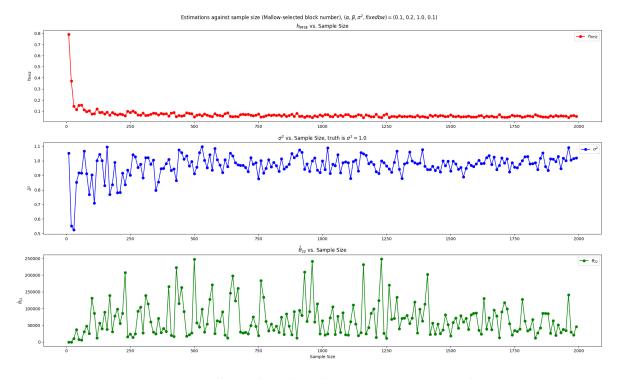


Figure 9: Mallow selected evolution up to 2000 samples

Here below is the call of py plot_results.py simple_plot $0.1\ 0.2\ 10000\ 0.07\ 1$ with the bandwith of 0.7 mentioned below.

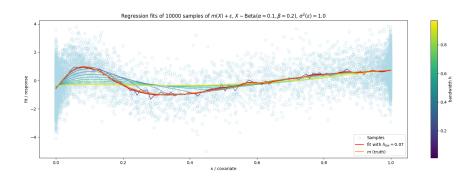


Figure 10: Plot with bandwith of 0.7

 $\it NB: Note that to produce these variations, one must directly change export name of the .pmg files.$