# MATH-517: Assignment 3

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# Theoretical exercise

Define the quadratic form

$$\mathcal{Q}(\beta) = \sum_{i=1}^n K\left(\frac{X_i - x}{h}\right) (Y_i - \beta_0 - \beta_1 (X_i - x))^2 = (Y - X\beta)^\mathrm{T} W(Y - X\beta),$$

where 
$$W=\mathrm{Diag}\left(W_{1},...,W_{n}\right)=\mathrm{Diag}\left(K\left(\frac{X_{1}-x}{h}\right),...,K\left(\frac{X_{n}-x}{h}\right)\right).$$

Using matrix calculus, we have that  $\frac{\partial \mathcal{Q}(\beta)}{\partial \beta} = 2X^TWX\beta - 2X^TWY$  so that the normal equations are given by  $X^TWX\hat{\beta} = X^TWY$ . Assuming that  $X^TWX$  is invertible, we immediately obtain from the normal equations that  $\hat{\beta} = (X^TWX)^{-1}X^TWY$ .

We will start off by compute some terms,  $X^{T}WX$  and  $X^{T}WY$  separately.

$$\begin{split} X^{\mathrm{T}}WX &= \begin{bmatrix} 1 & \cdots & 1 \\ X_1-x & \cdots & X_n-x \end{bmatrix} \begin{bmatrix} W_1 & & \\ & \ddots & \\ & & W_n \end{bmatrix} \begin{bmatrix} 1 & X_1-x \\ \vdots & \vdots \\ 1 & X_n-x \end{bmatrix} \\ &= \begin{bmatrix} 1 & \cdots & 1 \\ X_1-x & \cdots & X_n-x \end{bmatrix} \begin{bmatrix} W_1 & W_1(X_1-x) \\ \vdots & \vdots \\ W_n & W_n(X_n-x) \end{bmatrix} \end{split}$$

such that

$$X^{\mathrm{T}}WX = \begin{bmatrix} \sum_{i=1}^{n} W_{i} & \sum_{i=1}^{n} W_{i}(X_{i}-x) \\ \sum_{i=1}^{n} W_{i}(X_{i}-x) & \sum_{i=1}^{n} W_{i}(X_{i}-x)^{2} \end{bmatrix}$$

Define  $\tilde{S}_{n,k}(x) := nhS_{n,k}(x)$  for k = 0, 1, 2, where  $S_{n,k}(x)$  are defined as in the README. In view of this definition,  $X^{\mathrm{T}}WX = \begin{bmatrix} \tilde{S}_{n,0} & \tilde{S}_{n,1} \\ \tilde{S}_{n,1} & \tilde{S}_{n,2} \end{bmatrix} = nh\begin{bmatrix} S_{n,0} & S_{n,1} \\ S_{n,1} & S_{n,2} \end{bmatrix}$ .

We now turn to  $X^{\mathrm{T}}WY$ .

$$\begin{split} X^{\mathrm{T}}WY &= \begin{bmatrix} 1 & \cdots & 1 \\ X_1 - x & \cdots & X_n - x \end{bmatrix} \begin{bmatrix} W_1 \\ & \ddots \\ & & W_n \end{bmatrix} \begin{bmatrix} Y_1 \\ \vdots \\ Y_n \end{bmatrix} = \begin{bmatrix} W_1 & \cdots W_n \\ W_1(X_1 - x) & \cdots & W_n(X_n - x) \end{bmatrix} \begin{bmatrix} Y_1 \\ \vdots \\ Y_n \end{bmatrix} \\ &= \begin{bmatrix} \sum_{i=1}^n W_i Y_i \\ \sum_{i=1}^n W_i (X_i - x) Y_i \end{bmatrix} \end{split}$$

Defining  $T_0 = \sum_{i=1}^n W_i Y_i$  and  $T_1 = \sum_{i=1}^n W_i (X_i - x) Y_i$ , we have that  $X^T W Y = \begin{bmatrix} T_0 & T_1 \end{bmatrix}^T$ . We now invert  $X^T W X$ . Since  $X^T W X$  is a 2 × 2 matrix, the inverse is

$$\begin{bmatrix} \tilde{S}_{n,0} & \tilde{S}_{n,1} \\ \tilde{S}_{n,1} & \tilde{S}_{n,2} \end{bmatrix}^{-1} = \frac{1}{D} \begin{bmatrix} \tilde{S}_{n,2} & -\tilde{S}_{n,1} \\ -\tilde{S}_{n,1} & \tilde{S}_{n,2} \end{bmatrix}, \text{ where } D = \tilde{S}_{n,0} \tilde{S}_{n,2} - \tilde{S}_{n,1}^2 \text{ is the determinent.}$$

Recall that we are interested in  $\hat{\beta}_0$  and in view of the above reasoning,

$$\hat{\beta} = \frac{1}{D} \begin{bmatrix} \tilde{S}_{n,2} & -\tilde{S}_{n,1} \\ -\tilde{S}_{n,1} & \tilde{S}_{n,0} \end{bmatrix} \begin{bmatrix} T_0 \\ T_1 \end{bmatrix} \text{ such that } \hat{\beta}_0 = \frac{1}{D} \left( \tilde{S}_{n,2} T_0 - \tilde{S}_{n,1} T_1 \right) = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n,1} (X_i - x) \right)}{D} Y_i = \sum_{i=1}^n \frac{W_i \left( \tilde{S}_{n,2} - \tilde{S}_{n$$

Recall that  $\tilde{S}_{n,k}=nhS_{n,k}$ . It follows that  $D=(nh)^2\left(S_{n,0}S_{n,2}-S_{n,1}^2\right)$ . It immediately follows that

$$\hat{\beta}_0 = \sum_{i=1}^n \frac{nh}{(nh)^2} \frac{W_i \left(S_{n,2} - S_{n,1}(X_i - x)\right)}{S_{n,0} S_{n,2} - S_{n,1}^2} Y_i = \sum_{i=1}^n \frac{1}{nh} \frac{W_i (S_{n,2} - S_{n,1}(X_i - x)}{S_{n,0} S_{n,2} - S_{n,1}^2} Y_i = \sum_{i=1}^n w_{n,i} Y_i,$$

where

$$w_{n,i} = w_{n,i}(x) = \frac{1}{nh} \frac{K\left(\frac{X_i - x}{h}\right) \left(S_{n,2}(x) - S_{n,1}(x)(X_i - x)\right)}{S_{n,0}(x)S_{n,2}(x) - S_{n,1}(x)^2}.$$

With this result, we answer both question (1) and (2) at the same time.

It remains to show that  $\sum_{i=1}^{n} w_{n,i}(x) = 1$ . We proceed as follows:

$$\sum_{i=1}^n w_{n,i}(x) = \frac{1}{nh} \sum_{i=1}^n \frac{W_i \left( S_{n,2}(x) - (X_i - x) S_{n,1}(x) \right)}{S_{n,0}(x) S_{n,2}(x) - S_{n,1}(x)^2}$$

$$=\frac{1}{nh}\frac{S_{n,2}(x)}{S_{n,0}(x)S_{n,2}(x)-S_{n,1}(x)^2}\sum_{i=1}^nW_i-\frac{1}{nh}\frac{S_{n,1}(x)}{S_{n,0}(x)S_{n,2}(x)-S_{n,1}(x)^2}\sum_{i=1}^nW_i(X_i-x)$$

In view of the definitions of  $S_{n,0}(x)$  and  $S_{n,1}(x)$ , we obtain

$$\sum_{i=1}^n w_{n,i}(x) = \frac{S_{n,2}(x)S_{n,0}(x)}{S_{n,0}(x)S_{n,2}(x) - S_{n,1}(x)^2} - \frac{S_{n,1}(x)^2}{S_{n,0}(x)S_{n,2}(x) - S_{n,1}(x)^2} = 1,$$

which concludes the third (3) part of the exercise.

This concludes the theoretical exercise.

#### **Practical exercise**

Unless stated otherwise, your answer to the practical part should include the following elements:

- Description of the aim of the simulation study.
- Description of the different quantities that intervene in your simulation study. Explain how these quantities (fixed or random) are defined and the reasoning behind the choices you made.
- Description of your findings using appropriate graphics and/or tables (with a caption!) that are well commented in the text.

The code should not appear in the PDF report, unless there is a specific reason to include it.

#### Structure and details:

## Aim of the simulation study:

The goal is to observe the impact of various parameters on the estimation of the IMSE-optimal bandwith  $h_{IMSE}$ ,  $\sigma^2$  and  $\theta_{22}$ . In particular how it varies with respect to

- various values (combinations) of the Beta function,  $\alpha, \beta > 0$  (see below)
- various values (combinations) of sample size n and block size / number of blocks N (see below)

## Fixed quantities:

# • ADD IF NECESSARY

## Code architecture:

The <u>technical functions</u> are located in a file functions.py with usual alias fct. There, we have the following functions:

- generate\_sample takes inputs alpha, beta, n\_samples, sigma\_2 which respectively represent  $\alpha$ ,  $\beta$ , the number of samples n and the variance of the error  $\epsilon$ ,  $\sigma^2$ . This function returns two one-dimensional arrays, covariate and response representing X and Y respectively.
- estimate\_parameters takes inputs covariate, response, bandwith, p, reg. They represent respectively X, Y, a default bandwith needed to perform an initial KDE and p which represents the order. Note that p=1 by default and reg is set to 1e-5. It computes the estimated  $\hat{\beta}$  by solving the normal equation. To efficiently compute the solution, we use Cholesky decomposition and np.linalg.solve instead of np.linalg.inv, which is significantly more stable and efficient in this situation. For even more stability, we add a small regularization (reg) before using np.linalg.solve. Throughout the calls of the functions, the reg parameter is never changed and remains set to its default value of 1e-5. The returns is a function beta\_est that, given a array of covariates, computes the fitted values.
- estimate\_sigma\_theta takes inputs covariate, repsonse, bandwith, N\_blocks which respectively represent X,Y, a default bandwith needed to perform the initial KDE and the number of blocks to compute estimates of  $\hat{\theta}_{22}$  and  $\hat{\sigma}^2$ . Note that in the README, it is usually mentioned the size of the blocks instead of the number. We decided to implement in terms of the number of blocks. This choice is due to (my) reasoning logic. The computation of the size, given the number of blocks is given by size = sample\_size / N\_blocks. To compute blocks, we compute the size, and then split the data in blocks [B1 | ... | Bk | Bextra] Bextra is the remaining number of samples. This function returns an estimate of  $\theta_{22}$  and  $\sigma^2$ , following the procedure recommended in the README.
- est\_h\_IMSE\_support takes inputs sigma\_2\_hat, theta\_22, size, covariate as inputs which respectively represent  $\hat{\sigma}^2$ ,  $\hat{\theta}_{22}$  and the sample size n. It returns  $h_{IMSE}$  as given in the README. Note that size is not a necessary argument as size=len(covariate). This choice was made for clarity. The covariate array is a necessary argument as the support of the realizations is given by support=max(covariate)-min(covariate) (note

that this is always a nonnegative quantity and a positive quantity with probability one). This function returns an estimate of  $h_{IMSE}$ .

- compute\_mallow\_C\_p takes inputs covariate, response, fixed\_bandwith, N\_blocks (as previously seen). It computes  $C_p(N)$  as suggested in the README and returns its value.
- optimal\_mallow takes inputs covariate, repsonse, fixed\_bandwith, max\_N\_blocks. The first two inputs are as seen previously and the third is the same as bandwith seen previously. The last argument represents the maximal number of blocks on which the function computes the  $C_p$  statistics (for m "model" selection). The function returns the number of blocks that minimized the Mallow  $C_p$  statistic. The function calls compute\_mallow\_C\_p for each block size.
- h\_IMSE\_Cp\_optimized takes inputs covariate, reponse, number\_of\_samples, default\_bandwith, max\_number\_of\_blocks, which respectively represent X, Y, n, the fixed bandwith as before (previously called bandwith or fixed\_bandwith and the maximum number of blocks to test (previously called max\_N\_blocks. The function calls optimal\_mallow and re-estimates  $\sigma^2$  and  $\theta_{22}$ , given the Mallow-optimal number of blocks (the return of optimal\_mallow). This function returns an estimation of  $\sigma^2$  and  $\theta_{22}$  with Mallow-optimal selection of the number of blocks.
- simulate takes inputs alpha, beta, number\_of\_samples, error\_variance, default\_bandwith, number\_of\_blocks, which respectively represent  $\alpha,\beta,n,\sigma^2,$  a default bandwith (as previously seen) and a number of blocks. It first calls generate\_sample with  $\alpha,\beta,n,\sigma^2,$  which returns covariate, response. With that return, it calls estimate\_sigma\_theta with the default bandwith and the number of blocks. Finally, it calls est\_h\_IMSE\_support with the computed quantities. This function returns the estimations of  $h_{IMSE},\sigma^2$  and  $\theta_{22}.$  We mostly use this function to ease the production of plots.

The <u>plotting functions</u> are accessible through a file called **plot\_results.py**. There, one can use any of the functions below:

- plot\_simple\_fit takes inputs covariate, response, fixed\_bandwith, sigma\_2, representing X, Y, the fixed bandwith as previously seen and  $\sigma^2$  (for the plot title). It returns a plot showing visually how the choice of bandwith impacts the fits. To produce this plot simply, refer to the subsection below "How to simply plot results". An example of call would be:
  - py plot\_results.py simple\_plot 0.1 0.2 10000 0.2 1
- plot\_alpha\_beta\_impact takes inputs alphas, betas, error\_variance, default\_bandwith, number\_of\_blocks, which respectively represent a list of alphas  $[\alpha_1,...,\alpha_{n_\alpha}]$ , as list of betas  $[\beta_1,...,\beta_{n_\beta}]$ , the default bandwith as seen previously and error\_variance represents  $\sigma^2$ . This function produces three heat maps of the values of estimated

 $h_{IMSE}$ ,  $\sigma^2$  and  $\theta_{22}$  for different values of  $\alpha_i$ ,  $\beta_j$ . To produce this plot simply, refer to the subsection below "**How to simply plot results**". An example of call would be:

- py plot\_results.py alpha\_beta\_impact 100 1.0 0.1 5
- plot\_sample\_blocks\_impact takes inputs number\_of\_samples\_range, number\_of\_blocks, alpha, beta, error\_variance, default\_bandwith, which respectively represent a list  $[n_1,...,n_{k_n}]$  of values for the number of samples n, a list  $[N_1,...,N_{k_N}]$  of values for the number of blocks N, fixed values of  $\alpha$ ,  $\beta$  and  $\sigma^2$ . This function produces three heat maps of the values of estimated  $h_{IMSE}$ ,  $\sigma^2$  and  $\theta_{22}$  for different values of  $n_i$  and  $N_j$ . To produce this plot simply, refer to the subsection below "How to simply plot results". An example of call would be:
  - py plot\_results.py sample\_blocks\_impact 0.1 0.2 1.0 0.1
- plot\_sample\_size\_impact\_mallow takes inputs alpha, beta, max\_number\_of\_samples, step, error\_variance, default\_bandwith, max\_number\_of\_blocks, as seen previously. max\_number\_of\_blocks is set to 10 by default. The additional new parameter represents the number of steps between two different sample sizes. The list of sample sizes to test are then given by n\_range = np.arange(start=default\_start, stop=max\_number\_of\_samples, step=step). This function produces three (sub) plots representing how the estimations of  $h_{IMSE}$ ,  $\sigma^2$  and  $\theta_{22}$  evolve when computed on n samples with Mallow-optimal number of blocks. To produce this plot simply, refer to the subsection below "How to simply plot results". An example of call would be:
  - py plot\_results.py sample\_size\_mallow\_evolution 0.1 0.2 2000 10 1.0
    0.1

# How to simply plot results:

To call them, use the console with one of the following keywords key∈ { simple\_plot, alpha\_beta\_impact, sample\_blocks\_impact, plot\_sample\_size\_impact\_mallow } and call as follows: py plot\_results.py key parameter1 parameter2 parameter3 ...

where parameter1 parameter2 ... are the functions parameters, detailed below:

- when key=simple\_plot: The parameters are in the following order alpha, beta, number\_of\_samples, fixed\_bandwith, variance, representing respectively  $\alpha, \beta, n$ , the fixed bandwith (as previously mentioned) and  $\sigma^2$ .
- when key=alpha\_beta\_impact: The parameters are in the following order number\_of\_samples, error\_variance, default\_bandwith, number\_of\_blocks, representing respectively the number of samples n to test,  $\sigma^2$ , the default bandwith (as previously mentioned) and the number of blocks.
- when key=sample\_blocks\_impact: The parameters are in the following order alpha, beta, error\_variance, default\_bandwith, as previously seen.

• when key=plot\_sample\_size\_impact\_mallow: The parameters are in the following order alpha, beta, max\_number\_of\_samples, step, error\_variance, default\_bandwith.

When a call is successful, the console will display TASK: <followed by the actual task it is performing > and will print TASK completed when the plot has been produced.

# Findings and reports:

We start off with a simple visualization of the setup for fixed  $\alpha, \beta, n$ . The plot below showcases how different bandwith choices visually impact the fit. In orange is the true m(x). In red is a (relatively good fitting) line. We observe that as the bandwith goes closer to zero, the fits become better. However, when to close to zero, the fitted function appears to be to sensible and does not provide a robust fit to the truth.

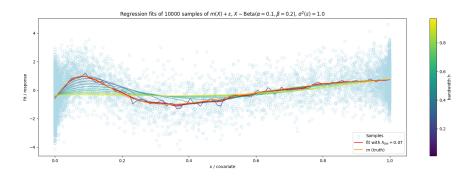


Figure 1: Visual plot of bandwith impact, call of py plot\_results.py simple\_plot 0.1 0.2 10000 0.2 1

This guides us with the belief that an optimal bandwith may have to be chosen close to zero, but not too close either (at least for these parameter choices).

Here below are heatmaps of the results of estimations of  $h_{IMSE}$ ,  $\sigma^2$ ,  $\theta_{22}$  against various values of  $\alpha$ ,  $\beta$ . The x-axis represents  $\beta$  and the y-axis represents  $\alpha$ . Here the alphas and betas are chosen in np.arange(0.1, 10, step = 0.5). We observe what seems to be a boundary effect when either  $\alpha$  or  $\beta$  is close to zero on the estimation of  $h_{IMSE}$  and  $\theta_{22}$  in particular. For  $h_{IMSE}$ , when  $\alpha$  or  $\beta$  (or both) are close to zero, the values of  $h_{IMSE}$  are of larger magnitude in comparison to the other cases. The inverted observation can be made for  $\theta_{22}$ 's estimation where the values are lower when  $\alpha$ ,  $\beta$  are close to zero in comparison to the other case.

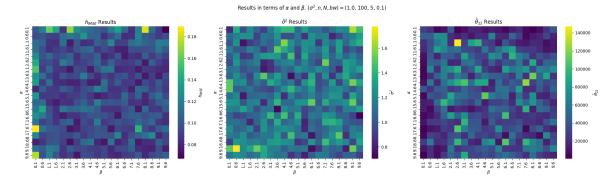


Figure 2: Impact of the choices of alpha and beta on estimations, call of py plot\_results.py alpha\_beta\_impact 100 1.0 0.1 5

To better emphasize on this boundary effect, we recomputed the heatmaps for  $\alpha, \beta$  taking values in np.arange(0.01, 5, step = 0.5). There, the boundary effect on the estimation of  $h_{IMSE}$  and  $\theta_{22}$  are even more noticeable.

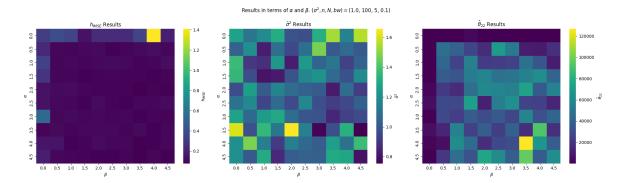


Figure 3: Impact of the choices of alpha and betas, close to zero

Note that in both cases, the estimations of  $\sigma^2$  remain relatively stable between 0.8 and 1.2 approximately.

To complete the observations around the impacts of  $\alpha, \beta$  on the estimations, we perform a last run for  $\alpha, \beta$  taking values in np.arange(0.01, 1, step = 0.01). There, we observe a fade from the left and upper boundaries to the center for  $\theta_{22}$ , and a strong boundary effect on  $h_{IMSE}$ , as already expected. We additionally observe that the estimations of  $h_{IMSE}$  remain very stable for values of  $\alpha, \beta$  far enough from zero (both at the same time) and that the estimations of  $\sigma^2$  are reasonably stable on average.

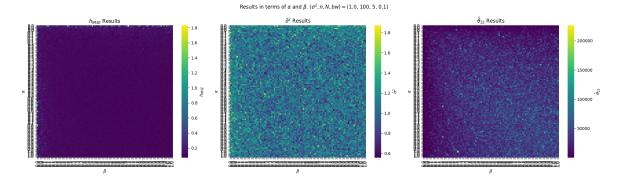


Figure 4: Impact of the choices of alpha and beta with increased granularity and closer to zero

# In conclusion,

- The estimation of  $h_{IMSE}$ ,  $\theta_{22}$  are more stable for values of  $\alpha, \beta$  far enough from zero (boundary effect). The estimation of  $h_{IMSE}$  has very few extreme values (so low variance) for values of  $\alpha, \beta$  large enough.
- The estimation of  $\sigma^2$  appears to not be affected a lot by changes of  $\alpha, \beta$ .
- initial estimations are that  $h_{IMSE}$  is around 0.2 on average,  $\sigma^2$  is, as expected, around 1.0 and  $\theta_{22}$  appears to be around 50000 on average.

NB: Note that to produce these variations, one must directly comment / uncomment parts of the code as well as change export name of the .png files. The default computation is made for values of  $\alpha, \beta$  in np.arange(0.1, 10, step = 0.5).

We now take a look at how the number of blocks (or similarly the size of the blocks) and the sample size impact the estimations.

As already mentioned, we think of number of blocks instead of size of blocks. As the number of blocks decrease, the size increases and vice-versa. First notice that we obtain on-average inconsistent estimations of  $h_{IMSE}$  and  $\sigma^2$  when the sample size is low and especially when n s small and N is large. For  $\theta_{22}$  on the other hand, the computations appear to be inconsistent in comparison to the average when the number of blocks is small. It becomes more stable as the number of blocks increase. This is an observation that one can also make for  $h_{IMSE}$  and  $\sigma^2$ .

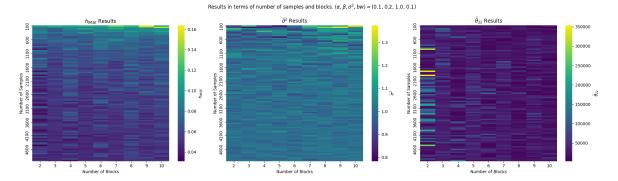


Figure 5: Impact of sample size and number of blocks on estimations (color map)

# In conclusion,

- As expected, the sample size positively increases the on-average consistency of the estimations.
- The size of the blocks decreases (and the number of blocks increases) the estimations appear to become stable when n (the sample size) is large enough.

Let us now take a look at how the estimations behave against the sample size when the number of blocks (or similarly the size) is selected using Mallow's  $C_p$  selection criterion. This is a reasonable way to compare the impact of sample sizes alone.

We observe, as expected and previously observed, that small sample sized (even with optimal number of blocks) yield unreasonable estimations in comparison to the average. Moreover, the estimation of  $h_{IMSE}$  appears to average around the value of 0.07 after the elbow. The estimation of  $\sigma^2$  also becomes stable and converges to, it appears, 1.0 (as expected). Furthermore, as already observed previously, the estimation of  $\theta_{22}$  is relatively unstable.

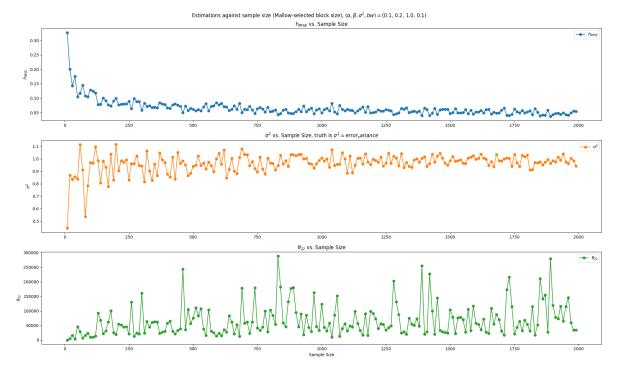


Figure 6: Estimations against sample size with Mallow-selected block number / size

Here below is the call of py plot\_results.py simple\_plot  $0.1\ 0.2\ 10000\ 0.07\ 1$  with the bandwith of 0.7 mentioned below.

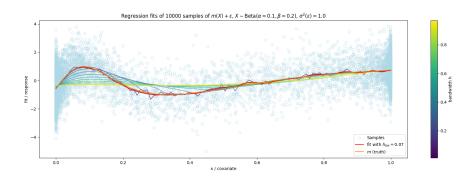


Figure 7: Plot with bandwith of 0.7

 $\it NB: Note that to produce these variations, one must directly change export name of the .prg files.$ 

Answers to suggested questions to address:

- 1. How does  $h_{IMSE}$  evolve as N grows and why could this be:
- 2. Should N depend on n and why:
- 3. What happens when the number of observations varies a lot between different regions in the support of X and link to  $\text{Beta}(\alpha, \beta)$ :