

# MATH-517: Assignment 3

GERMI Rémi

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## Theoretical exercise

### Question 1

Let

$$K_i = K \left( \frac{X_i - x}{h} \right), \quad u_i = X_i - x,$$

we can write the problem as a weighted least squares problem.

The minimization problem is:

$$(\hat{\beta}_0(x), \hat{\beta}_1(x)) = \arg \min_{\beta_0, \beta_1 \in \mathbb{R}} \sum_{i=1}^n (Y_i - \beta_0 - \beta_1 u_i)^2 K_i.$$

Let

$$\mathbf{X} = \begin{pmatrix} 1 & u_1 \\ \vdots & \vdots \\ 1 & u_n \end{pmatrix}, \quad \mathbf{W} = \text{diag}(K_1, \dots, K_n), \quad \mathbf{Y} = (Y_1, \dots, Y_n)^\top.$$

Then the weighted least squares solution is

$$\hat{\beta}(x) = (\mathbf{X}^\top \mathbf{W} \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{W} \mathbf{Y}.$$

The fitted value is

$$\hat{m}(x) = \hat{\beta}_0(x) = \mathbf{e}_1^\top (\mathbf{X}^\top \mathbf{W} \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{W} \mathbf{Y},$$

where  $\mathbf{e}_1 = (1, 0)^\top$ . Hence,

$$\hat{m}(x) = \sum_{i=1}^n w_{n,i}(x) Y_i,$$

with weights

$$w_{n,i}(x) = \mathbf{e}_1^\top (\mathbf{X}^\top \mathbf{W} \mathbf{X})^{-1} \begin{pmatrix} 1 \\ u_i \end{pmatrix} K_i.$$

## Question 2

Let

$$S_{n,k}(x) = \frac{1}{nh} \sum_{i=1}^n u_i^k K_i, \quad k = 0, 1, 2.$$

We get

$$\mathbf{X}^\top \mathbf{W} \mathbf{X} = nh \begin{pmatrix} S_{n,0}(x) & S_{n,1}(x) \\ S_{n,1}(x) & S_{n,2}(x) \end{pmatrix} =: nh \mathbf{M}(x),$$

and its inverse is

$$(\mathbf{X}^\top \mathbf{W} \mathbf{X})^{-1} = \frac{1}{nh} \frac{1}{S_{n,0}(x)S_{n,2}(x) - S_{n,1}(x)^2} \begin{pmatrix} S_{n,2}(x) & -S_{n,1}(x) \\ -S_{n,1}(x) & S_{n,0}(x) \end{pmatrix}.$$

Therefore,

$$w_{n,i}(x) = \frac{1}{nh} \frac{S_{n,2}(x) - S_{n,1}(x)(X_i - x)}{S_{n,0}(x)S_{n,2}(x) - S_{n,1}(x)^2} K\left(\frac{X_i - x}{h}\right).$$


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## Question 3

Noting that

$$\mathbf{X}(1,0)^\top = \mathbf{1}_n = \begin{pmatrix} 1 \\ \vdots \\ 1 \end{pmatrix} \quad \text{and} \quad \mathbf{X}^\top \mathbf{W} = \sum_{i=1}^n \begin{pmatrix} 1 \\ u_i \end{pmatrix} K_i,$$

we get

$$\sum_{i=1}^n w_{n,i}(x) = \sum_{i=1}^n \mathbf{e}_1^\top (\mathbf{X}^\top \mathbf{W} \mathbf{X})^{-1} \begin{pmatrix} 1 \\ u_i \end{pmatrix} K_i \tag{1}$$

$$= \mathbf{e}_1^\top (\mathbf{X}^\top \mathbf{W} \mathbf{X})^{-1} \sum_{i=1}^n \begin{pmatrix} 1 \\ u_i \end{pmatrix} K_i \tag{2}$$

$$= \mathbf{e}_1^\top (\mathbf{X}^\top \mathbf{W} \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{W} \mathbf{1} \tag{3}$$

$$= \mathbf{e}_1^\top (\mathbf{X}^\top \mathbf{W} \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{W} \mathbf{X} (1,0)^\top \tag{4}$$

$$= \mathbf{e}_1^\top \mathbf{I}_2 (1,0)^\top = 1. \tag{5}$$

Thus, we obtain:

$$\sum_{i=1}^n w_{n,i}(x) = 1.$$

## Practical exercise

- **Aim of the study :**

The purpose of the study is to estimate and observe the optimal bandwidth  $\hat{h}_{AMISE}$  for a beta distribution and to observe its behavior as a function of the block size  $N$ , the sample size  $n$ , and the parameters of the beta distribution.

- **Behavior of  $h$  when  $N$  growth:**

From the  $h$  vs  $N$  plot (for Beta(2,2)): We observe that  $\hat{h}_{AMISE}$  decreases as  $N$  increases. We can explain this by the fact that increasing  $N$  divides the sample into smaller blocks. The local polynomial fits then become more localized and can better capture curvature in  $m(x)$ .

This leads to a larger estimated  $\hat{\theta}_{22}$  (the integrated squared second derivative), and possibly to a smaller estimated residual variance  $\hat{\sigma}^2$ .

And from the definition of an increase in  $\hat{\theta}_{22}$  typically implies a decrease in  $\hat{h}$ . Empirically, we observe exactly this behavior in the simulation.

- **Should  $N$  depend on  $n$ ? Why?**

Yes, the block size  $N$  (number of blocks) should be chosen in relation to the sample size  $n$ . In fact, if  $N$  is too large relative to  $n$ , blocks contain very few points, making degree-4 polynomial fits unreliable.

Variance of the estimates increases and the denominator  $(n - 5N)$  in the computation of  $\hat{\sigma}^2$  can even become small or negative. If  $N$  is too small, the fits are overly global and may underestimate curvature, leading to a smaller  $\hat{\theta}_{22}$  and therefore an overly large  $\hat{h}$ . Hence we use the  $C_p$ -like rule for try to get the best  $N$ .

$N$  should scale with  $n$  — larger data-sets allow more blocks, but each block must remain large enough for stable estimation.

- **What happens when we change the parameter of the Beta law?**

When the number of observations varies a lot between different regions in the support of  $X$  e.g between Beta(5,1) or Beta(1,5) most observations concentrate near one end of the support, leaving few in the tails.

As a result: Blocks in sparse regions have few points  $\rightarrow$  unstable polynomial fits.

Estimates of  $\hat{\theta}_{22}$  and  $\hat{\sigma}^2$  become noisier and  $\hat{h}$  values fluctuate more, reflecting the heterogeneity in data density. The  $h$  vs Beta plot illustrates this.

For instance, the U-shaped Beta(2, 2) distribution produced larger  $\hat{h}$  values, since most observations lie near the boundaries and curvature estimates are less stable.

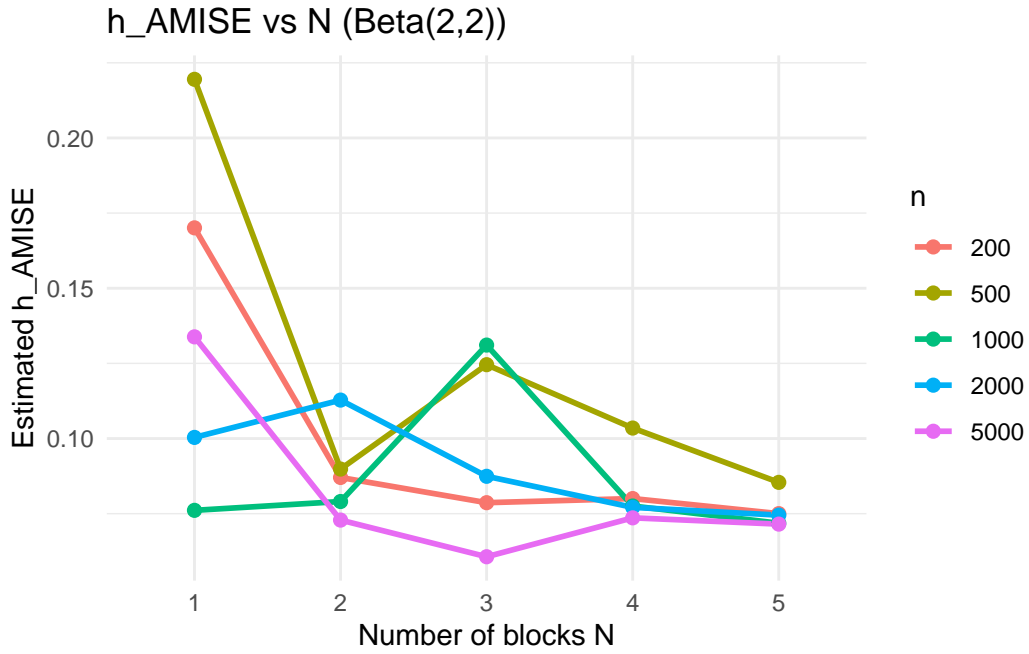
- **4. Effect of sample size  $n$**

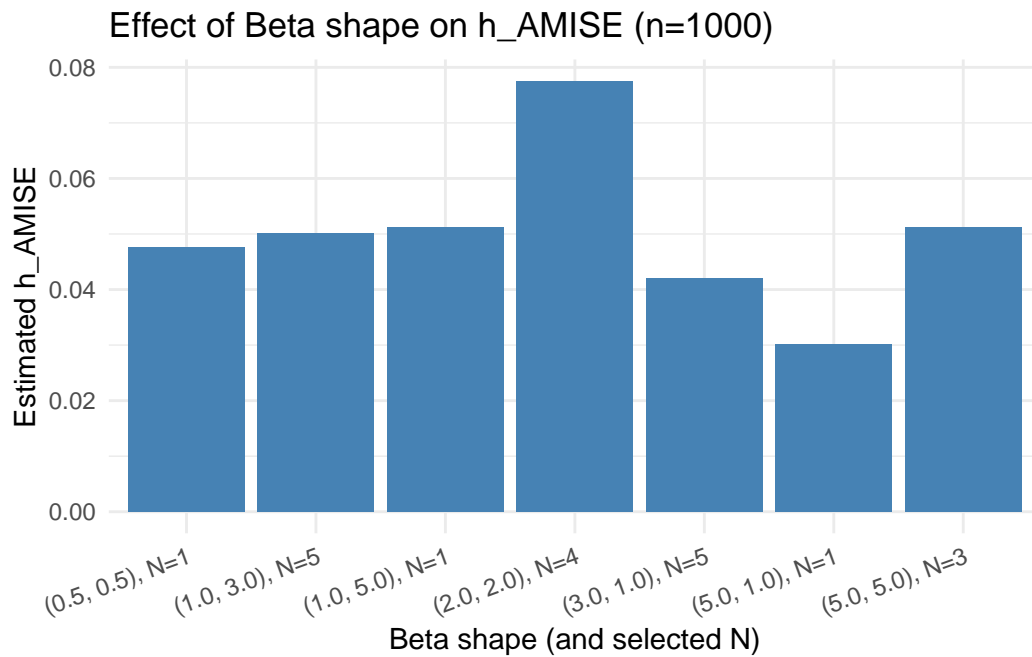
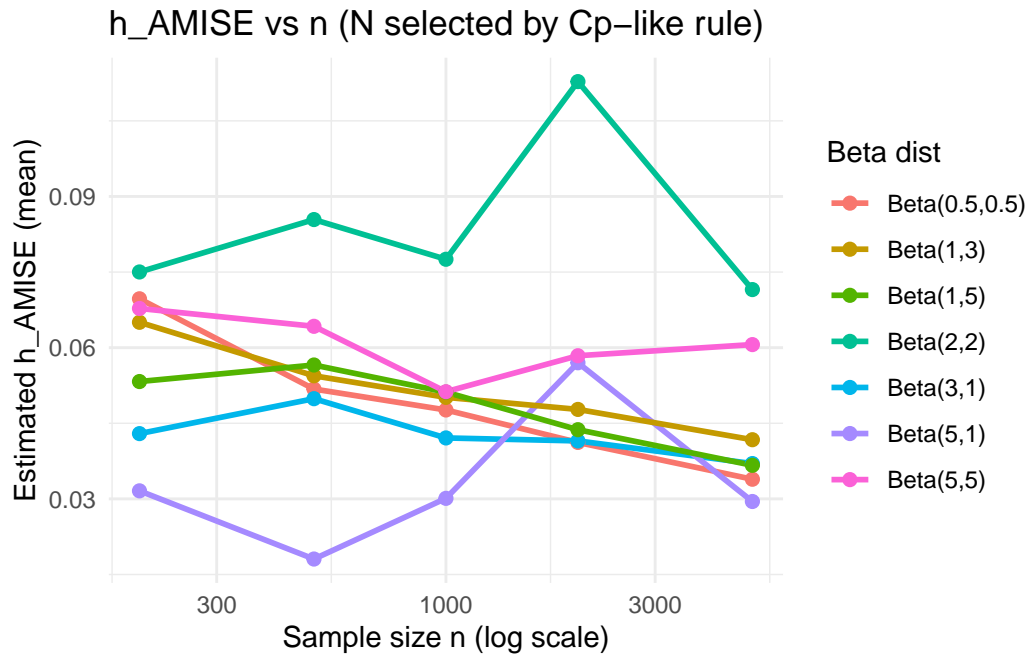
From the  $h$  vs  $n$  plot (using  $C_p$ -selected  $N$ ):

$\hat{h}_{AMISE}$  decreases as  $n$  increases, consistent with the theoretical scaling:

$$h_{AMISE} \propto n^{-1/5}.$$

However, the exact slope and smoothness depend on the Beta shape and the selected  $N$ . When  $N$  is optimally reselected for each  $n$  (as done via the  $C_p$  rule), the dependence of  $\hat{h}$  on  $n$  becomes smooth and closely follows the expected theoretical trend.





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[1] "Simulation finished. Files produced:"
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[1] "plot_h_vs_beta.png"    "plot_h_vs_n.png"      "plot_h_vs_Nblock.png"
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[1] "Saved tables: simulation_results_summary.csv, selected_N_by_Cp.csv"
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