

Normal distribution → Normal r.v is a continuous r.v.

Let us suppose that the normal r.v X has mean μ and SD σ

Here mean μ and SD σ are the parameters of the normal distribution.

The p.d.f fun for the normal variable X is given by

$$f(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, \quad -\infty < x < \infty$$

Mode of Normal Distribution

$$f'(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} \left[-\frac{2(x-\mu)}{2\sigma^2} \right]$$

$$f'(x) = 0 \Rightarrow x = \mu$$

$$f''(x) = \frac{1}{\sqrt{2\pi}\sigma} \left[e^{-\frac{(x-\mu)^2}{2\sigma^2}} \left(-\frac{(x-\mu)}{\sigma^2} \right)^2 + e^{-\frac{(x-\mu)^2}{2\sigma^2}} \left(-\frac{1}{\sigma^2} \right) \right]$$

$$f''(\mu) = \frac{1}{\sqrt{2\pi}\sigma} \left(-\frac{1}{\sigma^2} \right) < 0$$

$\Rightarrow x = \mu$ is the mode of the normal distribution

Median of the normal distribution

If M is the median of the normal distribution, then we have

$$\int_{-\infty}^M f(x) dx = \int_M^{\infty} f(x) dx = \frac{1}{2}$$

$$\text{we have } \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx = 1 \quad \text{--- (1)}$$

$$dz = \frac{dx}{\sigma}$$

Substitute $z = \frac{x-\mu}{\sigma}$ then (1) can be written as

$$\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{z^2}{2}} dz = 2 \int_0^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz = 1 \Rightarrow \int_0^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz = \frac{1}{2} \quad \text{--- (2)}$$

$$\text{we have } \int_M^{\infty} f(x) dx = \frac{1}{2} \Rightarrow \int_M^{\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx = \frac{1}{2} \quad \text{--- (3)}$$

Substitute $z = \frac{x-\mu}{\sigma}$ then (3) can be written as

$$\int_{\frac{M-\mu}{\sigma}}^{\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{z^2}{2}} \sigma dz = \int_{\frac{M-\mu}{\sigma}}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz = \frac{1}{2} \quad \text{--- (4)}$$

$$dz = \frac{dx}{\sigma}$$

$$\begin{matrix} x \rightarrow M & x \rightarrow \infty \\ z \rightarrow \frac{M-\mu}{\sigma} & z \rightarrow 0 \end{matrix}$$

Comparing (2) and (4)

$$\frac{M-\mu}{\sigma} = 0 \Rightarrow \boxed{M = \mu}$$

In Normal distribution

$$\boxed{\text{mean} = \text{mode} = \text{median} = \mu}$$

~~Mean deviation~~

Mean deviation of the normal distribution

$$\text{Mean deviation} = E(|X - \mu|) = \int_{-\infty}^{\infty} |x - \mu| \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx \quad (5)$$

Substitute $z = \frac{x - \mu}{\sigma}$ then (5) can be written as

$$\text{Mean deviation} = \int_{-\infty}^{\infty} |\sigma z| \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz = \frac{2\sigma}{\sqrt{2\pi}} \int_0^{\infty} z e^{-\frac{z^2}{2}} dz$$

$$\Rightarrow \sqrt{\frac{2}{\pi}} \sigma \left[-e^{-\frac{z^2}{2}} \right]_{z=0}^{\infty} = \sqrt{\frac{2}{\pi}} \sigma = 0.7979\sigma \approx 0.8\sigma$$

In normal distribution

$$P(a \leq X \leq b) = \int_a^b f(x) dx = \int_a^b \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

we know

$$\boxed{P(-\infty < X < \infty) = 1}$$

$$\boxed{\int_{-\infty}^{\infty} f(x) dx = 1}$$

Standard Normal Variable

A normal variable X is said to be a standard normal variable if its mean is zero and standard deviation is equal to 1

$$\text{Let } z = \frac{X - \mu}{\sigma}$$

Here mean of z is zero and s.d of z is 1.

Therefore, $z = \frac{X - \mu}{\sigma}$ is a standard normal variable.

Proof

$$\begin{aligned} \text{Var } Z &= E(Z^2) - [E(Z)]^2 \\ &= E(Z^2) - 0 \\ &= E\left[\left(\frac{X - \mu}{\sigma}\right)^2\right] = \frac{1}{\sigma^2} \left[E((X - \mu)^2) \right] \end{aligned}$$

$$\Rightarrow \frac{1}{\sigma^2} E[(X - \mu)^2] = \frac{1}{\sigma^2} E[(X^2 - 2X\mu + \mu^2)]$$

by basic definition

$$\Rightarrow \frac{\sigma^2}{\sigma^2} = 1$$

$$\boxed{SD = 1}$$

$$\begin{aligned} E(z) &= \text{mean of } z \\ &= E\left(\frac{X - \mu}{\sigma}\right) \\ &= \frac{1}{\sigma} E(X - \mu) \\ &= \frac{1}{\sigma} [E(X) - \mu] \\ &= \frac{1}{\sigma} [\mu - \mu] = 0 \end{aligned}$$

$$\boxed{\text{Var } X = E[(X - \mu)^2]}$$

To find $P(a \leq X \leq b)$

Let $z_1 = \frac{a-\mu}{\sigma}$ and $z_2 = \frac{b-\mu}{\sigma}$

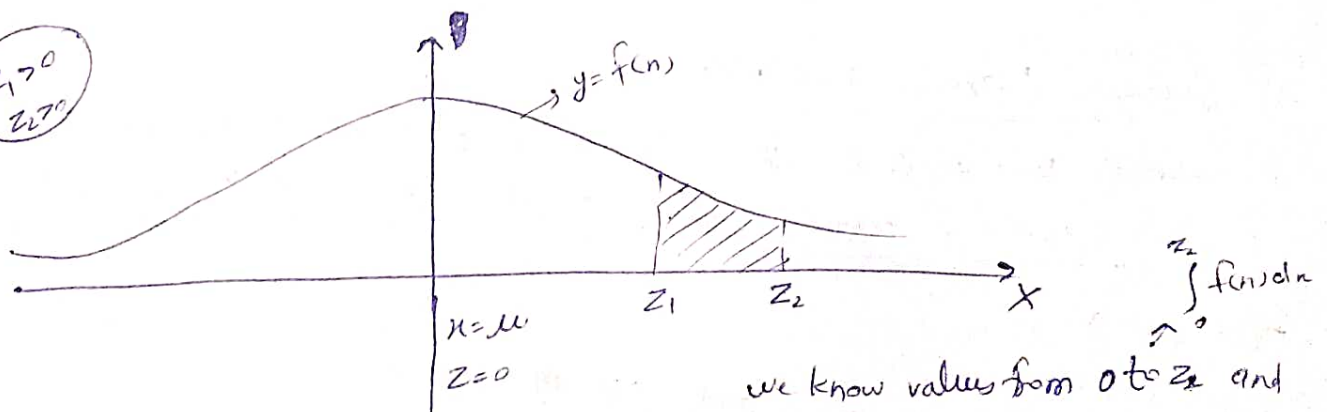
$$P(a \leq X \leq b) = \int_a^b \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx = \int_{z_1}^{z_2} \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz = P(z_1 \leq Z \leq z_2)$$

Here the values of $\int_0^{z_0} \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz$, where $z_0 > 0$ (for various values of z_0) are available in the form of a table.

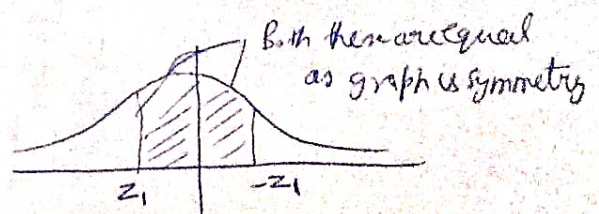
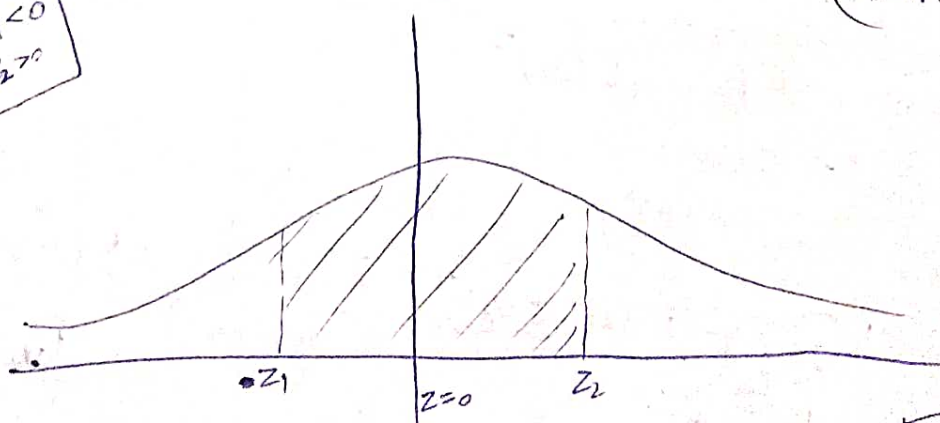
$$P(a \leq X \leq b) = P(z_1 \leq Z \leq z_2)$$

$$= \begin{cases} P(0 \leq Z \leq z_2) - P(0 \leq Z \leq z_1) & , \text{ if } z_1, z_2 > 0 \\ P(0 \leq Z \leq -z_1) + P(0 \leq Z \leq z_2) & , \text{ if } z_1 < 0, z_2 > 0 \\ P(0 \leq Z \leq -z_1) - P(0 \leq Z \leq -z_2) & , \text{ if } z_1, z_2 < 0 \end{cases}$$

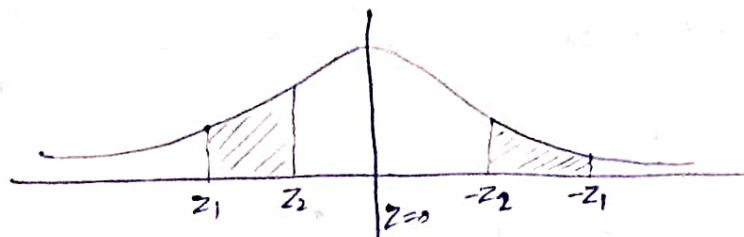
Case $\begin{matrix} z_1 > 0 \\ z_2 > 0 \end{matrix}$



Case $\begin{matrix} z_1 < 0 \\ z_2 > 0 \end{matrix}$



Case $Z < 0$
 $Z_2 < 0$



(25)

Note → A binomial variable X with parameters n and p is referred as a normal variable if n is very large and neither p nor q is small, i.e. if $n \geq 30$ and $p, q > 0.05$, then corresponding binomial variable X is referred as a normal variable.

Q1 Given that random variable X has binomial distribution with $n=50$ and $p=0.25$ find
(i) $P(X > 10)$ (ii) $P(X < 18)$ (iii) $P(X > 21)$ (iv) $P(9 < X < 14)$

Soln So here as we can see $n \geq 30$ and $p, q > 0.05$

as $n=50$ $p=0.25$ $q=0.75$

So X referred as Normal variable

$\text{Mean} = \mu = np$ $SD = \sigma = \sqrt{npq}$

$\mu = 12.5$

$\sigma = 2.6516 \approx 2.65$

(i) $P(X > 10) = P(Z > -0.94)$

$Z = \frac{X - \mu}{\sigma} = \frac{X - 12.5}{2.65}$

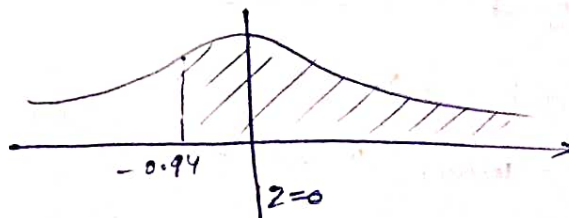
$\Rightarrow P(-0.94 < Z < 0) + P(0 < Z < \infty)$

Here $X < 10$

$\frac{10 - 12.5}{2.65} = \frac{-2.5}{2.65} = -0.94$

$X > 10$
 $n - \mu > 10 - \mu$
 $\frac{n - \mu}{\sigma} > \frac{10 - \mu}{\sigma}$

$Z > \frac{10 - 12.5}{2.65} \Rightarrow Z > -0.94$



$\Rightarrow P(-0.94 < Z < 0) + P(0 < Z < \infty)$

$\Rightarrow P(0 < Z < 0.94) + P(0 < Z < \infty)$

$\Rightarrow 0.3264 + \frac{1}{2} = 0.8264$

$$(ii) P(X < 18) = P\left(\frac{X - \mu}{\sigma} < \frac{18 - \mu}{\sigma}\right)$$

$$= P(Z < 2.08)$$

$$= P(-\infty < Z < 0) + P(0 < Z < 2.08)$$

$$= 0.5 + 0.4812 = 0.9812$$

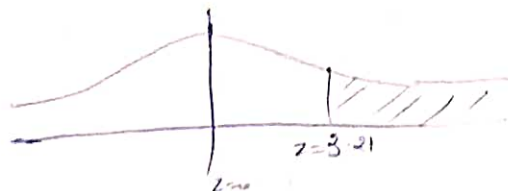
$$Z = \frac{X - \mu}{\sigma}$$

$$= \frac{18 - 12.5}{2.64}$$

$$= 2.08$$

$$= 2.08$$

$$(iii) P(X > 21) = P\left(\frac{X - \mu}{\sigma} > \frac{21 - \mu}{\sigma}\right) = P(Z > 3.21)$$



$$Z = \frac{X - \mu}{\sigma}$$

$$= \frac{21 - 12.5}{2.64}$$

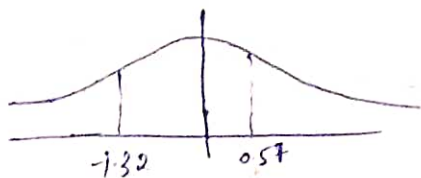
$$= 3.2075$$

$$\Rightarrow P(0 < Z < \infty) - P(0 < Z < 3.21)$$

$$\Rightarrow 0.5 - 0.49994 = 0.0006$$

$$(iv) P(9 < X < 14) = P\left(\frac{9 - \mu}{\sigma} < \frac{X - \mu}{\sigma} < \frac{14 - \mu}{\sigma}\right)$$

$$\Rightarrow P(-1.32 < Z < 0.57)$$



$$\Rightarrow P(0 < Z < 1.32) + P(0 < Z < 0.57)$$

$$\Rightarrow 0.4066 + 0.2157$$

$$\Rightarrow 0.6223$$

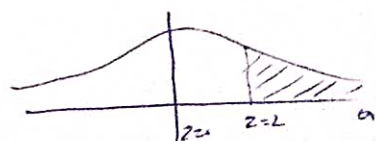
Q2 If X is a normal variable with mean 12 and s.d 4, then

(a) find the probability of (i) $X \geq 20$ (ii) $0 \leq X \leq 12$

(b) find μ_1 , if $P(X > \mu_1) = 0.24$

(c) find μ_0 and μ_1 when $P(\mu_0 < X < \mu_1) = 0.5$ and $P(X > \mu_1) = 0.25$

$$\text{Sol}^n \quad P(X \geq 20) = P\left(\frac{X - \mu}{\sigma} \geq \frac{20 - \mu}{\sigma}\right) = P(Z \geq 2)$$



$$\Rightarrow P(0 < Z < \infty) - P(0 < Z < 2)$$

$$\Rightarrow 0.5 - 0.4772$$

$$\Rightarrow$$

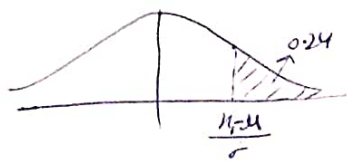
$$\frac{20 - \mu}{\sigma} = \frac{20 - 12}{4} = \frac{8}{4} = 2$$

$$(iv) P(0 \leq X \leq 12) = P\left(\frac{0-\mu}{\sigma} \leq \frac{X-\mu}{\sigma} \leq \frac{12-\mu}{\sigma}\right) = P(-3 \leq Z \leq 0)$$

$$\Rightarrow P(0 \leq Z \leq 3) = 0.4987$$

$$(b) P(X > \mu_1) = 0.24$$

$$P\left(\frac{X-\mu}{\sigma} > \frac{\mu_1-\mu}{\sigma}\right) = 0.24 \Rightarrow P\left(Z > \frac{\mu_1-12}{4}\right) = 0.24$$



$$P(0 < Z < \infty) - P\left(0 < Z < \frac{\mu_1-12}{4}\right) = 0.24$$

$$0.5 - 0.24 = P\left(0 < Z < \frac{\mu_1-12}{4}\right)$$

$$0.26 = P\left(0 < Z < \frac{\mu_1-12}{4}\right)$$

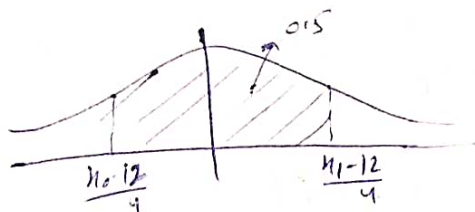
By using table values

$$\frac{\mu_1-12}{4} = 0.705$$

$$\mu_1 = 0.704 \times 4 + 12$$

$$(iii) P(\mu_0 < X < \mu_1) = 0.5$$

$$P\left(\frac{\mu_0-12}{4} < Z < \frac{\mu_1-12}{4}\right) = 0.5$$



$$P(X > \mu_1) = 0.25$$

$$P\left(\frac{X-\mu}{\sigma} > \frac{\mu_1-\mu}{\sigma}\right) = 0.25$$

$$P\left(Z > \frac{\mu_1-12}{4}\right) = 0.25$$

$$P(0 < Z < \infty) - P\left(Z > \frac{\mu_1-12}{4}\right) = 0.25$$

$$0.5 - 0.25 = P\left(Z > \frac{\mu_1-12}{4}\right)$$

$$P\left(Z > \frac{\mu_1-12}{4}\right) = 0.25 \quad \text{--- (2)}$$

$$P\left(\frac{\mu_0-12}{4} < Z < 0\right) + P\left(0 < Z < \frac{\mu_1-12}{4}\right) = 0.5$$

$$P\left(\frac{\mu_0-12}{4} < Z < 0\right) + 0.25 = 0.5$$

$$P\left(\frac{\mu_0-12}{4} < Z < 0\right) = 0.25 \quad \text{--- (1)}$$

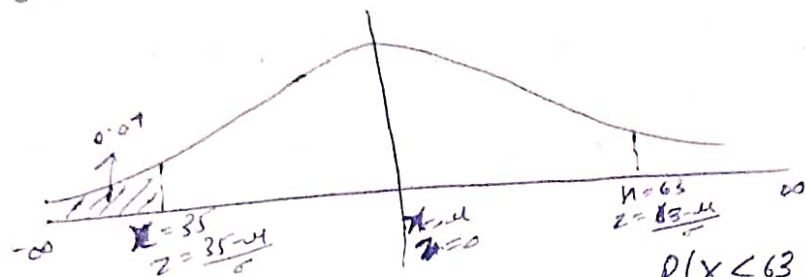
from (1) and (2)

$$-\left(\frac{\mu_0-12}{4}\right) = +\left(\frac{\mu_1-12}{4}\right) = 0.675$$

Q If X is a normal variable with mean = 30 and S.D 5 find prob that $|X-30| \geq 5$

$$\begin{aligned} \text{Soln } P(|X-30| \geq 5) &= 1 - P(|X-30| \leq 5) \Rightarrow 1 - P(5 \leq X-30 \leq 5) \\ &\Rightarrow 1 - P(25 \leq X \leq 35) \\ &\Rightarrow 1 - P\left(\frac{25-30}{5} \leq \frac{X-30}{5} \leq \frac{35-30}{5}\right) \\ &\Rightarrow 1 - P(-1 \leq Z \leq 1) \\ &\Rightarrow 1 - [2P(0 \leq Z \leq 1)] \\ &\Rightarrow 1 - 2 \times 0.3413 = 0.3174 \end{aligned}$$

Q In a distribution exactly normal, 7% of the items are under 35 and 89% of the items are under 63. What are the mean and SD of distribution?



$$P(X < 35) = 7\% = 0.07$$

$$P(X < 63) = 89\% = 0.89$$

Since 0.07 is less than 0.5 so 0.07 must be on left side

$P(-\infty < X < 35) = 0.07$ and we know $P(-\infty < X < \mu) = 0.5$ so it must be on left side

By both of these we can say $35 < \mu$

$P(-\infty < X < 63) = 0.89$ and we know $P(-\infty < X < \mu) = 0.5$

By both of these we can say that $63 > \mu$

$$P(-\infty < X < 63) \Rightarrow P(-\infty < X < \mu) + P(\mu < X < 63) = 0.89$$

$$\Rightarrow P(-\infty < Z < \frac{\mu - \mu}{\sigma}) + P(\frac{\mu - \mu}{\sigma} < Z < \frac{63 - \mu}{\sigma}) = 0.89$$

$$\Rightarrow P(-\infty < Z < 0) + P(0 < Z < \frac{63 - \mu}{\sigma}) = 0.89$$

$$\Rightarrow 0.5 + P(0 < Z < \frac{63 - \mu}{\sigma}) = 0.89$$

$$\Rightarrow P(0 < Z < \frac{63 - \mu}{\sigma}) = 0.89 - 0.5 = 0.39 \quad \text{--- (1)}$$

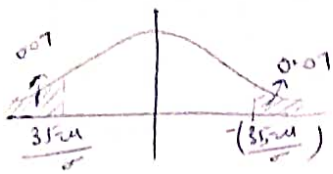
~~$$P(-\infty < X < 35) = P(-\infty < \frac{X - \mu}{\sigma} < \frac{35 - \mu}{\sigma}) = 0.07$$~~

~~$$\Rightarrow P(-\infty < Z < \frac{35 - \mu}{\sigma}) = 0.07$$~~

~~$$\Rightarrow P(\frac{35 - \mu}{\sigma} < Z < \infty) = 0.07$$~~

~~$$\Rightarrow P(0 < Z < \frac{35 - \mu}{\sigma}) = 0.07$$~~

~~$$\Rightarrow P(0 < Z < \frac{35 - \mu}{\sigma}) = 0.07$$~~



~~P(0.07)~~

$$P(-\infty < Z < \frac{35-\mu}{\sigma}) = 0.07$$

$$P(Z \geq -(\frac{35-\mu}{\sigma})) = 0.07$$

$$P(0 < Z < \infty) - P(0 < Z < -(\frac{35-\mu}{\sigma})) = 0.07$$

$$0.5 - 0.07 = P(0 < Z < -(\frac{35-\mu}{\sigma}))$$

$$0.43 = P(0 < Z < -(\frac{35-\mu}{\sigma}))$$

By using
table values

$$\boxed{-\left(\frac{35-\mu}{\sigma}\right) = 1.475} - A$$

from (1)

$$P(0 < Z < \frac{63-\mu}{\sigma}) = 0.39$$

By using
table values

$$\boxed{\frac{63-\mu}{\sigma} = 1.225} - B$$

Solving A and B we get μ and σ

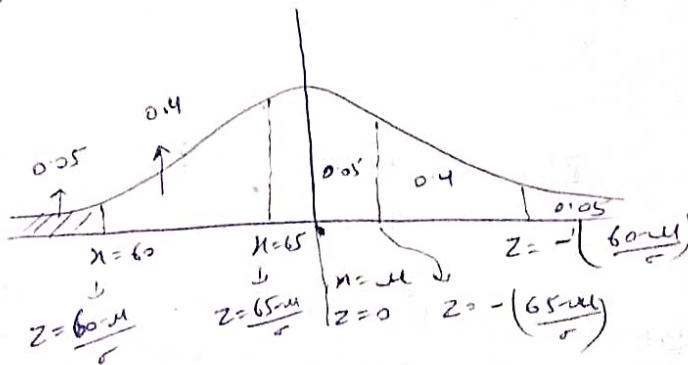
Q5 Of a large group of people 5% are under 60 inches in height and the height of the 40% of the people lie between 60 and 65 inches. Find the mean and Sd of the distribution.

Soln $X \rightarrow$ height of the people

$$P(X < 60) = 5\% = 0.05$$

$$P(60 < X < 65) = \frac{40}{100} = 0.4$$

$$P(0 \leq X < 60) = 0.05$$



$$P(0 < Z < -(\frac{65-\mu}{\sigma})) = 0.05$$

$$P(0 < Z < -(\frac{60-\mu}{\sigma})) = 0.45$$

from table we get values for $-(\frac{65-\mu}{\sigma})$ and $-(\frac{60-\mu}{\sigma})$

Q6

Soln Estimation I \rightarrow N.V with $\mu = 325$, $\sigma = 60$

Estimation II \rightarrow N.V with $\mu = 300$, $\sigma = 50$

$X =$ sales in the quarter

(i) $P(X > 350 \text{ w.r.t Estimation-I})$

$$Z = \frac{X - \mu}{\sigma}$$

Here we use μ and σ of Estimation (I)

$$P\left(Z > \frac{350 - \mu}{\sigma}\right) = P\left(Z > \frac{350 - 325}{60}\right) = P\left(Z > \frac{25}{60}\right) = P\left(Z > 0.41\right)$$

$$\Rightarrow P(0 < Z < \infty) - P(0 < Z < 0.41)$$

$$\Rightarrow 0.5 - 0.1591 =$$

(ii) $P(X > 350 \text{ w.r.t Estimation-II})$

Here we'll use μ and σ of Estimation II

$$P\left(Z > \frac{350 - \mu}{\sigma}\right) = P\left(Z > \frac{350 - 300}{50}\right) = P(Z > 1)$$

$$\Rightarrow P(0 < Z < \infty) - P(0 < Z < 1)$$

$$\Rightarrow 0.5 - 0.2420 =$$

(iii) $P(\text{Estimation-I} / X > 350) = \frac{P(X > 350 \text{ w.r.t E-I}) P(E-I)}{P(X > 350)}$

using Bayes's Theorem

$$P(E-I) = P(E-II) = 0.5$$

$$P(X > 350) = P\left(\frac{X > 350}{E-I}\right) P(E-I) + P\left(\frac{X > 350}{E-II}\right) P(E-II)$$

Total probability

(iv) $P(E-II / X > 350) = \frac{P(X > 350 \text{ w.r.t E-II}) P(E-II)}{P(X > 350)}$

Q Show that for Normal distribution the quartile deviation, mean deviation and standard deviation are approximately in the ratio 10:12:15

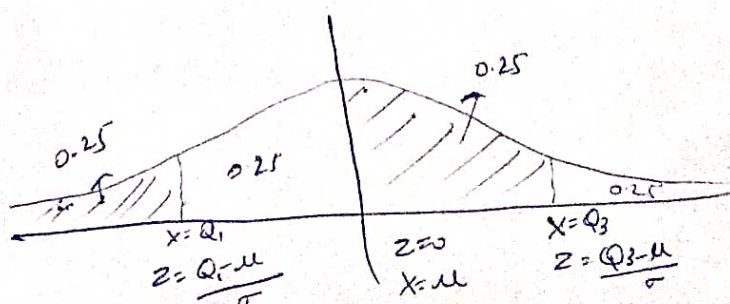
Soln Q.D : M.D : S.D = 10 : 12 : 15

$$M.D = \frac{4}{5} S.D$$

In the derivation we have seen this

$$Q.D = \frac{Q_3 - Q_1}{2}$$

where $Q_1 =$ 1st quartile
 $Q_3 =$ 3rd quartile



$$-\left(\frac{\phi_1 - \mu}{\sigma}\right) = \left(\frac{\phi_3 - \mu}{\sigma}\right)$$

$$P(0 < Z < \frac{\phi_3 - \mu}{\sigma}) = 0.25$$

from table

$$\frac{\phi_3 - \mu}{\sigma} = 0.675$$

$$\phi_3 = \mu + 0.675\sigma$$

$$-\left(\frac{\phi_1 - \mu}{\sigma}\right) = 0.675$$

$$-\phi_1 + \mu = 0.675\sigma$$

$$\phi_1 = \mu - 0.675\sigma$$

$$\phi.D = \frac{\phi_3 - \phi_1}{2} = \frac{2 \times 0.675\sigma}{2}$$

$$\phi.D = 0.675\sigma$$

$$\phi.D : m.D : s.D$$

$$0.675\sigma : \frac{4}{5}\sigma : \sigma = 0.675 : \frac{4}{5} : 1$$

multiply by 15

$$10.125 : 12 : 15 \approx 10 : 12 : 15$$

Q8 Fit the Normal distribution to the following data:

Class interval	60-65	65-70	70-75	75-80	80-85	85-90	90-95	95-100
Frequency	3	21	150	335	326	135	26	4

these are observed frequencies

Solⁿ

Fitting of data \Rightarrow finding of theoretical frequencies

Theoretical frequency = (probability of X) \times total frequency of X

for Normal distribution \Rightarrow parameters are μ (mean), σ (S.D)

class interval

C.I	freq (fi)	mid-value of C.I
60-65	3	62.5
65-70	21	67.5
70-75	150	72.5
75-80	335	77.5
80-85	326	82.5
85-90	135	87.5
90-95	26	92.5
95-100	4	97.5
	1000	

Total frequency Σf_i

$$\mu = \text{mean} = \frac{\Sigma f_i x_i}{\Sigma f_i}$$

=

$$= 79.945$$

$$S.D = \sqrt{\text{var}}$$

$$\text{var} = \sigma^2 = \frac{1}{\Sigma f_i} \Sigma f_i x_i^2 - (\text{mean})^2$$

$$= 29.65$$

$$\sigma = s.D = \sqrt{\text{var}} = 5.445$$

$$P(60 \leq X < 65) = P\left(\frac{60-\mu}{\sigma} \leq Z < \frac{65-\mu}{\sigma}\right)$$

and so on

$$\text{frequency}(X \text{ lies between } 60 \text{ and } 65) = P(60 < X < 65) \times 1000 =$$

for every class interval we have to do these steps.

Exponential distribution:- A r.v. X is said to follow exponential distribution with parameter $\theta > 0$ if its p.d.f is given by

$$f(x) = \begin{cases} \theta e^{-\theta x} & x \geq 0 \\ 0 & \text{otherwise} \end{cases}$$

Characteristics of Exponential distribution i.e. mean and variance
let's find these

M.G.F for exponential distribution about origin

$$M_0(t) = E(e^{tn}) = \int_0^{\infty} e^{tn} f(n) dn = \int_0^{\infty} e^{tx} \theta e^{-\theta x} dx = \int_0^{\infty} \theta e^{(t-\theta)x} dx$$

$$\Rightarrow \left[\frac{\theta e^{(t-\theta)x}}{t-\theta} \right]_{x=0}^{\infty} = \frac{\theta}{\theta-t} = \frac{1}{1-\frac{t}{\theta}} = 1 + \frac{t}{\theta} + \frac{t^2}{\theta^2} + \frac{t^3}{\theta^3} + \dots$$

$$\text{mean} = \frac{d}{dt} M_0(t) \Big|_{t=0}$$

Here we take
 $t=0$
 $e^{-\infty} = \frac{1}{e^{\infty}} = 0$

$$\Downarrow \left(1 - \frac{t}{\theta}\right)^{-1}$$

Mean and Variance of exponential distribution

$$\text{Mean} = E(X) = \frac{d}{dt} M_0(t) \Big|_{t=0} = \frac{1}{\theta}$$

$$E(X^2) = \frac{d^2}{dt^2} M_0(t) \Big|_{t=0} = \frac{2}{\theta^2}$$

$$\text{Variance} = E(X^2) - [E(X)]^2 = \frac{2}{\theta^2} - \frac{1}{\theta^2} = \frac{1}{\theta^2}$$

Important result on exponential distribution (24)

Show that the exponential distribution "lacks memory" i.e. if X follows an exponential distribution then for every constant $a \geq 0$, it has $P(Y \leq x / X \geq a) = P(X \leq x)$ for all x , where $Y = X - a$.

Proof The pdf for the exponential distribution with parameter θ is given by $f(x) = \theta e^{-\theta x}$, $\theta > 0$, $x \geq 0$

$$\text{we have } \boxed{P(Y \leq x / X \geq a)} = \frac{P(Y \leq x \cap X \geq a)}{P(X \geq a)} \quad \text{--- (1)}$$

$$\Rightarrow \boxed{P(Y \leq x \cap X \geq a)} = P(X - a \leq x \cap X \geq a) \text{ since } (Y = X - a)$$

$$\Rightarrow P(X \leq x + a \cap X \geq a) = P(a \leq X \leq a + x)$$

$$\Rightarrow \int_a^{a+x} \theta e^{-\theta n} dn = \left[\frac{\theta e^{-\theta n}}{-\theta} \right]_a^{a+x} = -[e^{-\theta(a+x)} - e^{-\theta a}]$$

$$\Rightarrow e^{-\theta a} - e^{-\theta(a+x)} = e^{-\theta a} [1 - e^{-\theta x}]$$

$$\boxed{P(X \geq a)} = \int_a^{\infty} \theta e^{-\theta n} dn = \left[\frac{\theta e^{-\theta n}}{-\theta} \right]_a^{\infty} = -[e^{-\theta n}]_a^{\infty} = -[e^{-\infty} - e^{-\theta a}] \Rightarrow \boxed{e^{-\theta a}}$$

$$\# \text{ from (1) } P(Y \leq x / X \geq a) = \frac{e^{-\theta a} [1 - e^{-\theta x}]}{e^{-\theta a}} = 1 - e^{-\theta x} \quad \text{--- (A)}$$

$$\Rightarrow P(X \leq x) = \int_0^x f(n) dn = \int_0^x \theta e^{-\theta n} dn = \left[\frac{\theta e^{-\theta n}}{-\theta} \right]_0^x = -[e^{-\theta n}]_0^x \Rightarrow 1 - e^{-\theta x} \quad \text{--- (B)}$$

from (A) and (B)

$$P(Y \leq x / X \geq a) = P(X \leq x)$$

Hypergeometric Distribution → when the population is finite and sampling is done without replacement, so that the events are stochastically dependent (although random), we obtain hypergeometric distribution

Note → Here population means number of items in sample space
 # Consider an urn with N balls. M of which are white and $N-M$ are non white. Suppose that we draw a sample of n balls from the urn at random without replacement. then the probability of getting k white balls out of n ($k \leq n$) drawn balls is given by

$$\frac{\binom{M}{k} \binom{N-M}{n-k}}{\binom{N}{n}} \quad \text{where } \binom{M}{k} = {}^M C_k$$

Definition → A s.v X is said to follow hypergeometric distribution if it assume only non-negative values and its p.d.f is given by

$$p(x) = P(X=x) = h(x, N, M, n) = \begin{cases} \frac{\binom{M}{x} \binom{N-M}{n-x}}{\binom{N}{n}} & \text{for } x=0, 1, 2, \dots, n \\ 0 & \text{otherwise} \end{cases}$$

Hypergeometric distribution approximated as a binomial distribution:

Hypergeometric distribution tends to a binomial distribution if

$$N \rightarrow \infty \quad \text{and} \quad \frac{M}{N} \rightarrow p \quad \left(\text{mean } \frac{M}{N} < 1 \right)$$

Problems A taxi cab has 12 Maruti Swift cars and 8 Tata Vista cars. If 5 of these cars in the shop are in repair then find the prob that

(i)

Solⁿ (i) $\frac{{}^{12}C_3 \times {}^8C_2}{{}^{20}C_5}$

(ii) $P(X=3) + P(X=4) + P(X=5)$
 $\frac{{}^{12}C_3 \times {}^8C_2}{{}^{20}C_5} + \frac{{}^{12}C_4 \times {}^8C_1}{{}^{20}C_5} + \frac{{}^{12}C_5 \times {}^8C_0}{{}^{20}C_5}$

(iii) $P(\text{All are maruti}) + P(\text{all are Tata Vista})$
 $\frac{{}^{12}C_5 \times {}^8C_0}{{}^{20}C_5} + \frac{{}^{12}C_0 \times {}^8C_5}{{}^{20}C_5}$

Q2 200 students of first year MCA probability that $R \geq 3$.

Solⁿ Here we are discussing about batch.

So in a batch either a student will be student or not

So it is the case of Binomial distribution

$$p = \frac{M}{N} = \frac{40}{200} = \frac{1}{5} = 0.2$$

$$h = 10$$

$$p = 0.2$$

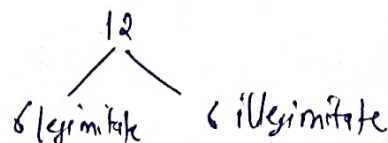
M = student students
 because we are talking
 about students
 every time

$p(R)$ = prob that there are R student in a batch

$$p(R) = {}^n C_R p^R q^{n-R} = {}^{10} C_R (0.2)^R (0.8)^{10-R}$$

$$P(R \geq 3) = 1 - P(R < 3) = 1 - [{}^{10} C_0 (0.2)^0 (0.8)^{10} + {}^{10} C_1 (0.2)^1 (0.8)^9 + {}^{10} C_2 (0.2)^2 (0.8)^8]$$

Q3 Find the prob that the income tax official will catch 3 income tax returns with the illegitimate deductions, if the official randomly affects 5 returns among from 12 returns of which 6 returns contains illegitimate deductions



$$\text{Req Prob} = \frac{{}^6 C_3 \times {}^6 C_2}{{}^{12} C_5}$$