## Homework 9

## Michael Brodskiy

Professor: I. Salama

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(a) Since it is given that  $X, Y \in [0, 1]$ , we find that the range of  $W = \text{Max}(X, Y) \to$  $W \in [0,1]$ . Therefore, we write:

$$S_W = [0, 1]$$

(b) Given that w is the upper bound of x and/or y, we may write:

$$F_W(w) = \int_0^w \int_0^w 4xy \, dx \, dy$$

We evaluate this to get:

$$F_W(w) = 4 \int_0^w y \int_0^w x \, dx \, dy$$
$$F_W(w) = 4 \left[ \frac{y^2}{2} \right] \left[ \frac{x^2}{2} \right] \left[ \frac{y^2}{2} \right]$$

$$F_W(w) = 4 \left[ \frac{y^2}{2} \right] \Big|_0^w \left[ \frac{x^2}{2} \right] \Big|_0^w$$

$$F_W(w) = \begin{cases} w^4, & 0 \le w < 1\\ 1, & w \ge 1\\ 0, & \text{otherwise} \end{cases}$$

From here, we know that:

$$f_W(w) = \frac{d}{dw} [F_W(w)]$$

This gives us:

$$f_W(w) = 4w^3, \quad 0 \le w < 1$$

(a) Given that  $y \ge x$ , we know that the lowest point for W will occur when y = x, or W=0. Similarly, the biggest difference occurs when y=1 and x=0, which produces W = 1. Thus, we write:

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$$S_W = [0, 1]$$

(b) From here, we may write:

$$F_W(w) = \int_0^1 \int_0^{y-w} 6x \, dx \, dy$$

We evaluate to get:

$$F_W(w) = 6 \int_0^1 [x^2/2] \Big|_0^{y-w} dy$$

$$F_W(w) = 3 \int_0^1 (y-w)^2 dy$$

$$F_W(w) = 3 \int_0^1 y^2 - 2yw + w^2 dy$$

$$F_W(w) = y^3 - 3y^2w + 3w^2y \Big|_0^1$$

$$F_W(w) = \begin{cases} 1 - 3w + 3w^2, & 0 \le w \le 1\\ 1, & w > 1\\ 0, & \text{otherwise} \end{cases}$$

We then differentiate to get:

$$f_W(w) = 6w - 3, \quad 0 \le w \le 1$$

3. (a) We can express the PDF of Y using the law of total probability as:

$$f_Y(y) = P[Z = 1]f_x(y) + P[Z = -1]f_x(-y)$$

Substituting our values, since we know that X is normal with  $\mu = 0$  and  $\sigma^2 = 1$ , we get:

$$f_Y(y) = \frac{1}{\sqrt{2\pi}} \left[ pe^{-\frac{y^2}{2}} + (1-p)e^{-\frac{y^2}{2}} \right]$$

Distributing and simplifying gives us:

$$f_Y(y) = \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}}$$

Thus, we see that this returns to simply being a normal distribution.

(b) When we take  $\mu \to 10$ , this gives us:

$$f_Y(y) = \frac{1}{\sqrt{2\pi}} \left[ pe^{-\frac{(y-10)^2}{2}} + (1-p)e^{-\frac{(-y-10)^2}{2}} \right]$$

We simplify to get:

$$f_Y(y) = \frac{1}{\sqrt{2\pi}} \left[ pe^{-\frac{(y-10)^2}{2}} + (1-p)e^{-\frac{(y+10)^2}{2}} \right]$$

5. We begin by finding the moment generating function of X. We can write this as:

$$\phi_x(s) = E[e^{sX}]$$

$$E[e^{sX}] = \sum_{x=0}^{\infty} e^{sx} \left(\frac{e^{-\alpha_1} \alpha_1^x}{x!}\right)$$

We rearrange to get:

$$E[e^{sX}] = e^{-\alpha_1} \sum_{x=0}^{\infty} \left( \frac{(e^s \alpha_1)^x}{x!} \right)$$

We know that the series expansion for an exponential is:

$$e^x \to \sum_{n=0}^{\infty} \frac{x^n}{n!}$$

As such, we get:

$$\phi_x(s) = e^{\alpha_1(e^s - 1)}$$

Similarly, we find the MGF of Y:

$$\phi_y(s) = e^{\alpha_2(e^s - 1)}$$

Since X and Y are independent, the MGF of Z = X + Y will simply be the product of the individual MGFs:

$$\phi_z(s) = e^{(\alpha_1 + \alpha_2)(e^s - 1)}$$

Note that, if we were to apply the same process to find the MGF of Z as we did for X and Y, we would still get the same answer.

7. (a) Using the moment generating function, we know that:

$$E[Y] = \phi_y(0)'$$

As such, we differentiate to get:

$$\phi_y'(s) = \frac{2}{(1-s)^3}$$

We evaluate at s = 0, which gives us:

$$E[Y] = 2$$

Similarly, we may write:

$$E[Y^2] = \phi_u(0)''$$

We differentiate a second time to get:

$$\phi_y''(s) = \frac{6}{(1-s)^4}$$

Evaluating at s = 0, we find:

$$E[Y^2] = 6$$

(b) Similar to (a), we find:

$$E[V] = \phi_w(0)'$$

$$E[V] = \frac{3}{(1-0)^4}$$

$$E[V] = 3$$

This gives us:

$$E[W] = E[Y] + E[V]$$
$$E[W] = 2 + 3$$
$$E[W] = 5$$

We then find:

$$E[V^{2}] = \phi_{w}(0)''$$

$$E[V^{2}] = \frac{12}{(1-0)^{5}}$$

$$E[V^{2}] = 12$$

From here, we may write:

$$Var[W] = Var[Y] + Var[V]$$

This gives us:

$$Var[W] = (6 - 2^{2}) + (12 - 3^{2})$$
$$Var[W] = 5$$

(c) Given that Y and V are independent, we may take their product to find:

$$\phi_w(s) = \frac{1}{(1-s)^5}$$

We may observe that, accordingly,  $W = \text{Erlang}(\lambda = 1, k = 5)$ , which makes sense, given that Y = Erlang(1, 2) and V = Erlang(1, 3). As such, we get:

$$f_W(w) = \frac{w^4 e^{-w}}{4!}, \quad w \ge 0$$

(d) We may re-express this as:

$$W_1 = \frac{1}{2}(Y+V) \to W_1 = \frac{1}{2}W$$

Accordingly, we get:

$$E[W_1] = \frac{1}{2}E[W]$$

$$E[W_1] = 2.5$$

And then:

$$Var[W_1] = \frac{1}{4} Var[W]$$

$$\boxed{\mathrm{Var}[W_1] = 1.25}$$

And finally:

$$f_{W_1}(w_1) = \frac{2(2w_1)^4 e^{-2w_1}}{4!}, \quad w_1 \ge 0$$

(a) Knowing our results for a standard normal distribution, may write the moment generating function as:

$$\phi_x(s) = E[e^{sX}]$$

$$\phi_x(s) = e^{\frac{s^2}{2}}$$

$$\phi_x(s) = e^{\frac{s^2}{2}}$$

(b) Given that each  $X_n$  is independent of the others, we can multiply the MGFs ktimes to write:

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$$E[e^{sV}|K=k] = e^{\frac{ks^2}{2}}$$

From here, we use the given probabilities to write:

$$\phi_v(s) = \sum_{k=1}^{\infty} P_K(k) e^{\frac{ks^2}{2}}$$

This gives us:

$$\phi_v(s) = \frac{.8e^{\frac{s^2}{2}}}{1 - .2e^{\frac{s^2}{2}}}$$

Before evaluating the expected value with the MGF, we may find the value to be:

$$E[V] = E[E[V|K]]$$

From above, we know:

$$E[V|K] = kE[X]$$

And since each X is normally distributed, E[X] = 0. Therefore, we may conclude:

$$\boxed{E[V] = 0}$$

We confirm using the MGF:

$$\phi_v(s)' = \frac{.8se^{\frac{s^2}{2}}}{\left(1 - .2e^{\frac{s^2}{2}}\right)^2}$$

At s = 0 this gives:

$$\phi_v(0)' = 0$$

And therefore, we confirm:

$$E[V] = \phi_v'(0) = 0$$

9. (a) We may observe that the given distribution is binomial, such that we may write:

$$K_{50} = \binom{50}{k} .7^k .3^{1-k}$$

Accordingly, we may find:

$$E[K_{50}] = np$$

$$E[K_{50}] = 50(.7)$$

$$E[K_{50}] = 35 \text{ video packets}$$

(b) Again, given this is a binomial distribution, we write:

$$\sigma_{50} = \sqrt{np(1-p)}$$
 
$$\sigma_{50} = \sqrt{35(.3)}$$
 
$$\sigma_{50} = 3.2404 \text{ video packets}$$

(c) We begin by standardizing to a normal variable, Z:

$$Z = \frac{X - 35}{3.2404}$$

We take the limits as  $30 \le X \le 40$ , such that:

$$\frac{-5}{3.2404} \le Z \le \frac{5}{3.2404}$$
$$-1.543 \le Z \le 1.543$$

We can find this probability by first finding:

$$P[Z \ge -1.543] = .9386$$

We then subtract the upper bound:

$$P[30 \le X \le 40] = .9386 - P[Z \ge 1.543]$$
  
 $P[30 \le X \le 40] = .9386 - .061415$   
 $P[30 \le X \le 40] = .8772$ 

(d) We begin by using continuity correction to write:

$$P[a \le K_n \le b] \approx P\left[\frac{a-\mu}{\sigma} \le Z \le \frac{b-\mu}{\sigma}\right]$$

This gives us:

$$P[30 \le K_n \le 50] \approx P[29.5 \le K_n \le 40.5]$$
  
 $P[30 \le K_n \le 50] \approx P[-1.6973 \le Z \le 1.6973]$ 

We can use the inverse normal operation to find:

$$P[30 \le K_n \le 50] \approx P[Z \le 1.6973] - P[Z \le -1.6973]$$
  
 $P[30 \le K_n \le 50] \approx .955 - .045$   
 $P[30 \le K_n \le 50] \approx .91$ 

11. (a) We know that the expectation value for a single test is given by the inverse of its probability, or:

$$E[X_n] = \frac{1}{p}$$

Given that we sum together a series consisting of n of these geometric distributions, we can find:

$$E[K_n] = E[X_1] + E[X_2] + \dots + E[X_n] = \frac{n}{p}$$

This gives us:

$$E[K_n] = \frac{900}{.9}$$

$$E[K_n] = 1000 \text{ tests}$$

(b) Similarly, the variance of  $K_n$  may be written as:

$$Var[K_n] = nVar[X_n]$$

We may find:

$$\operatorname{Var}[X_n] = \frac{1-p}{p^2}$$

$$\operatorname{Var}[X_n] = \frac{.1}{.9^2}$$

$$\operatorname{Var}[X_n] = .1235 \text{ tests}^2$$

We then multiply by n to find:

$$Var[K_n] = 111.11 \text{ tests}^2$$

(c) Using the CLT for at least 1000 tests, we write:

$$Z = \frac{1000 - \mu}{\sigma} = \frac{1000 - 1000}{\sqrt{111.11}}$$
$$Z = 0$$

From here, we apply a normal distribution to write:

$$P[Z \ge 0] = .5$$

## 12. We can begin by finding:

$$E[V] = 20 - 10E[W^3]$$

We can find this by using:

$$E[W^{3}] = \int_{-1}^{1} \frac{w^{3}}{2} dw$$

$$E[W^{3}] = \frac{w^{4}}{8} \Big|_{-1}^{1}$$

$$E[W^{3}] = 0$$

This gives us:

$$E[V] = 20$$

Similarly, we write:

$$Var[V] = 100 Var[W^{3}]$$

$$Var[V] = 100 \int_{-1}^{1} (w^{3})^{2} \cdot \frac{1}{2} dw$$

$$Var[V] = 100 \int_{-1}^{1} \frac{w^{6}}{2} dw$$

$$Var[V] = 100 \left[ \frac{w^{7}}{14} \right]_{-1}^{1}$$

$$Var[V] = 100/7 \approx 14.285$$

As such, we may write expressions for X as:

$$E[X] = 30 \cdot E[V]$$

$$\sigma_X = \sqrt{30 \cdot \text{Var}[V]}$$

These give us:

$$E[X] = 600 \text{ Mb}$$
  
 $\sigma_X = 20.73 \text{ Mb}$ 

We then want to find:

$$P\left[\frac{S_{30}}{6} \ge 95\right] \to P\left[Z \ge \frac{X - \mu}{\sigma}\right]$$

$$P\left[S_{30} \ge 570\right] \to P\left[Z \ge \frac{X - \mu}{\sigma}\right]$$

We apply the central limit theorem to get:

$$Z = \frac{570 - 600}{20.73}$$
$$Z = -1.4472$$

From here, we find:

$$P[S_{30} \ge 570] = 1 - P[Z < -1.4472]$$

$$P[X \ge 95] \approx .9261$$