

ME613 - Análise de Regressão

Parte 10

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Região de Confiança

Propriedades

$$\hat{\boldsymbol{\beta}}_{p \times 1} \sim \mathcal{N}_{p}(\boldsymbol{\beta}, (\mathbf{X}^{T}\mathbf{X})^{-1}\sigma^{2})$$

$$\frac{1}{\sigma}(\mathbf{X}^{T}\mathbf{X})^{1/2}(\boldsymbol{\beta} - \hat{\boldsymbol{\beta}}) \sim \mathcal{N}_{p}(\mathbf{0}, \mathbf{I})$$

$$\frac{1}{\sigma^{2}}(\boldsymbol{\beta} - \hat{\boldsymbol{\beta}})^{T}(\mathbf{X}^{T}\mathbf{X})(\boldsymbol{\beta} - \hat{\boldsymbol{\beta}}) \sim \chi^{2}(p)$$

$$(n - p)\frac{s^{2}}{\sigma^{2}} \sim \chi^{2}(n - p)$$

Portanto:

$$\frac{1}{ps^2} (\boldsymbol{\beta} - \hat{\boldsymbol{\beta}})^T \mathbf{X}^T \mathbf{X} (\boldsymbol{\beta} - \hat{\boldsymbol{\beta}}) \sim F(p, n - p)$$



Intervalo de Confiança para eta_k

Um intervalo de $100(1-\alpha)\%$ de confiança para β_k é dado por:

$$IC(\beta_k, 1 - \alpha) = \left[\hat{\beta}_k - t_{n-p,\alpha/2} \sqrt{\widehat{Var}(\hat{\beta}_k)}; \right]$$
$$\hat{\beta}_k + t_{n-p,\alpha/2} \sqrt{\widehat{Var}(\hat{\beta}_k)}$$



Estudo sobre diversidade das espécies em Galápagos.

Conjunto de dados: 30 ilhas, 7 variáveis.

Species: the number of plant species found on the island

Endemics: the number of endemic species

Area: the area of the island (km²)

Elevation: the highest elevation of the island (m)

Nearest: the distance from the nearest island (km)

Scruz: the distance from Santa Cruz island (km)

Adjacent: the area of the adjacent island (square km)





IC 95% para $\beta_{Adjacent}$:

```
confint(lmod)[6,]
```

```
## 2.5 % 97.5 %
## -0.11133622 -0.03827344
```



Região de Confiança para $oldsymbol{eta}$

Região de Confiança de $100 \times (1 - \alpha)\%$:

$$\frac{1}{ps^2} (\boldsymbol{\beta} - \hat{\boldsymbol{\beta}})^T \mathbf{X}^T \mathbf{X} (\boldsymbol{\beta} - \hat{\boldsymbol{\beta}}) \le F(p, n - p; 1 - \alpha)$$

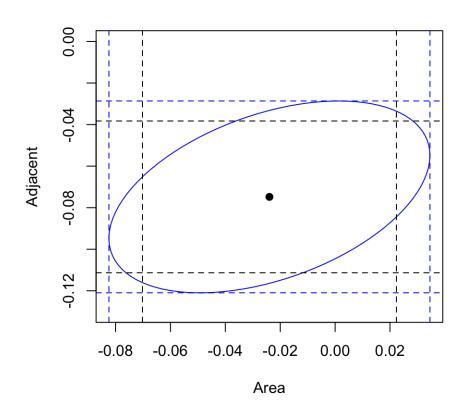


RC 95% para $\beta_{Adjacent}$ e β_{Area}

```
library(ellipse)
aa <- ellipse(lmod,which=c(2,6),level=0.95)
plot(aa,type="1",ylim=c(-0.13,0),col="blue")
points(coef(lmod)[2], coef(lmod)[6], pch=19)
abline(v=confint(lmod)[2,],lty=2)
abline(h=confint(lmod)[6,],lty=2)
abline(h=c(max(aa[,2]),min(aa[,2])),lty=2,col="blue")
abline(v=c(max(aa[,1])),min(aa[,1])),lty=2,col="blue")</pre>
```



RC 95% para $eta_{Adjacent}$ e eta_{Area}





Teste de Hipótese Linear

Teste de Hipótese Linear

Teste de hipótese linear:

$$H_0$$
: $\mathbf{R}_{r \times p} \boldsymbol{\beta}_{p \times 1} = \mathbf{q}_{r \times 1}$

$$H_1: \mathbf{R}_{r \times p} \boldsymbol{\beta}_{p \times 1} \neq \mathbf{q}_{r \times 1}$$

Para testar, começamos pensando no vetor de discrepância com relação à H_0 :

$$\mathbf{R}_{r \times p} \hat{\boldsymbol{\beta}}_{p \times 1} - \mathbf{q}_{r \times 1} = \mathbf{m}_{r \times 1}$$

queremos medir quão longe ${f m}$ está de ${f 0}$.



Teste de Wald para Hipótese Linear

Precisamos então conhecer a distribuição de ${\bf m}$, sob H_0 :

$$E(\mathbf{m}) = E(\mathbf{R}\hat{\boldsymbol{\beta}} - \mathbf{q}) = \mathbf{R}E(\hat{\boldsymbol{\beta}}) - \mathbf{q} = \mathbf{R}\boldsymbol{\beta} - \mathbf{q} \stackrel{H_0}{=} \mathbf{0}$$
$$Var(\mathbf{m}) = Var(\mathbf{R}\hat{\boldsymbol{\beta}} - \mathbf{q}) = \mathbf{R}Var(\hat{\boldsymbol{\beta}})\mathbf{R}^T = \sigma^2\mathbf{R}(\mathbf{X}^T\mathbf{X})^{-1}\mathbf{R}^T$$

Estatística do teste:

$$W = \mathbf{m}^{T} [Var(\mathbf{m})]^{-1} \mathbf{m}$$

$$= (\mathbf{R}\hat{\boldsymbol{\beta}} - \mathbf{q})^{T} [\sigma^{2} \mathbf{R} (\mathbf{X}^{T} \mathbf{X})^{-1} \mathbf{R}^{T}]^{-1} (\mathbf{R}\hat{\boldsymbol{\beta}} - \mathbf{q})$$

$$\stackrel{H_{0}}{\sim} \chi^{2}(r)$$



Teste de Wald para Hipótese Linear

Problema: não conhecemos σ^2 . Temos que utilizar um estimador para σ^2 : s^2 .

Sabemos que
$$(n-p)\frac{s^2}{\sigma^2} \sim \chi^2(n-p)$$
.

A estatística do teste é:

$$F = \frac{W}{r} \frac{\sigma^{2}}{s^{2}}$$

$$= \frac{(\mathbf{R}\hat{\boldsymbol{\beta}} - \mathbf{q})^{T} [\mathbf{R}(\mathbf{X}^{T}\mathbf{X})^{-1}\mathbf{R}^{T}]^{-1} (\mathbf{R}\hat{\boldsymbol{\beta}} - \mathbf{q})}{r\sigma^{2}} \frac{\sigma^{2}}{s^{2}}$$

$$= \frac{(\mathbf{R}\hat{\boldsymbol{\beta}} - \mathbf{q})^{T} [s^{2}\mathbf{R}(\mathbf{X}^{T}\mathbf{X})^{-1}\mathbf{R}^{T}]^{-1} (\mathbf{R}\hat{\boldsymbol{\beta}} - \mathbf{q})}{r}$$

$$\stackrel{H_{0}}{\sim} F_{r,n-p}$$



· Para H_0 : $\beta_j = 0$, definimos

$$\mathbf{R} = [0 \quad 0 \quad \dots \quad 1 \quad 0 \quad \dots \quad 0] e \mathbf{q} = 0.$$

• Para H_0 : $\beta_k = \beta_j$, definimos

$$\mathbf{R} = [0 \quad 0 \quad 1 \quad \dots \quad -1 \quad 0 \quad \dots \quad 0] e \mathbf{q} = 0.$$

· Para H_0 : $\beta_1 + \beta_2 + \beta_3 = 1$, definimos

$$\mathbf{R} = [0 \ 1 \ 1 \ 1 \ 0 \ \dots \ 0] e \mathbf{q} = 1.$$

• Para H_0 : $\beta_0 = 0$, $\beta_1 = 0$ e $\beta_2 = 0$, definimos

$$\mathbf{R} = \begin{pmatrix} 1 & 0 & 0 & 0 & \dots & 0 \\ 0 & 1 & 0 & 0 & \dots & 0 \\ 0 & 0 & 1 & 0 & \dots & 0 \end{pmatrix} \qquad \mathbf{q} = \mathbf{0}$$

· Para H_0 : $\beta_1 + \beta_2 = 1$, $\beta_3 + \beta_5 = 0$ e $\beta_4 + \beta_5 = 0$, definimos

$$\mathbf{R} = \begin{pmatrix} 0 & 1 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 & 1 \\ 0 & 0 & 0 & 0 & 1 & 1 \end{pmatrix} \qquad \mathbf{q} = \begin{pmatrix} 1 \\ 0 \\ 0 \end{pmatrix}$$



```
library(systemfit)
data( "Kmenta" )
modelo <- lm(consump ~ price + income, data=Kmenta)
summary(modelo)$coef</pre>
##

Estimate Std. Error t value Pr(>|t|)
```

Estimate Std. Error t value
$$Pr(>|t|)$$
 ## (Intercept) 99.8954229 7.51936214 13.285093 2.090605e-10 ## price -0.3162988 0.09067741 -3.488177 2.815290e-03 ## income 0.3346356 0.04542183 7.367285 1.099860e-06

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \varepsilon$$

 X_1 : price

 X_2 : income



```
H_0: \beta_1 = 0
```

```
R <- matrix( c(0,1,0),ncol=length(coef(modelo)),byrow=TRUE)
linearHypothesis(modelo,R)</pre>
```

```
## Linear hypothesis test
##
## Hypothesis:
## price = 0
##
## Model 1: restricted model
## Model 2: consump ~ price + income
##
              RSS Df Sum of Sq F Pr(>F)
    Res.Df
## 1
       18 108.660
## 2
        17 63.332 1 45.328 12.167 0.002815 **
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```



```
H_0: \beta_1 = 2
```

```
R <- matrix( c(0,1,0),ncol=length(coef(modelo)),byrow=TRUE)
q=2
linearHypothesis(modelo,R,q)</pre>
```

```
## Linear hypothesis test
##
## Hypothesis:
## price = 2
##
## Model 1: restricted model
## Model 2: consump ~ price + income
##
              RSS Df Sum of Sq F Pr(>F)
   Res.Df
## 1
        18 2494.21
## 2
        17 63.33 1 2430.9 652.52 5.311e-15 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```



```
H_0: \beta_1 = 2 e \beta_2 = 1.
```

```
R \leftarrow matrix(c(0,1,0,
              0,0,1),ncol=length(coef(modelo)),byrow=TRUE)
q=matrix(c(2,1),ncol=1)
linearHypothesis(modelo,R,q)
## Linear hypothesis test
## Hypothesis:
## price = 2
\#\# income = 1
##
## Model 1: restricted model
## Model 2: consump ~ price + income
##
  Res.Df
           RSS Df Sum of Sq F Pr(>F)
       19 7146.8
## 1
## 2
       17 63.3 2 7083.4 950.7 < 2.2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```



Agradecimento

Slides criados por Samara F Kiihl / IMECC / UNICAMP



Leituras

- The Matrix Cookbook
- Faraway Linear Models with R: Seção 3.5.
- · Draper & Smith Applied Regression Analysis: Seção 9.1.

