# RISK ESTIMATION -QUANTIFYING THE WORLD-

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### Loss functions and risk

Define a function  $\ell: \mathbb{R} \times \mathbb{R} \to \mathbb{R}$  such that smaller values of  $\ell$  indicate better performance

Two important examples:

- $\ell(\hat{f}(X), Y) = (\hat{f}(X) Y)^2$  (regression, square-error)
- $\ell(\hat{f}(X), Y) = \mathbf{1}(\hat{f}(X) \neq Y)$  (classification, 0-1)

These expressions are both random variables.

This leads us to define the prediction or estimation risk of a procedure  $\hat{f}$  to be

$$R(\hat{f}) = \mathbb{E}\ell(\hat{f}(X), Y)$$

<sup>&</sup>lt;sup>1</sup>This is the loss for prediction. Other tasks, such as estimation, may have a different domain.

#### RISK ESTIMATION

The prediction risk can be written

$$R(f) = \mathbb{E}\ell(f(X), Y) \leftrightarrow \text{Bias} + \text{Variance}$$

The overriding theme is that we would like to add a judicious amount of bias to get lower risk

As R isn't known, we need to estimate it

From the Preliminary Materials,  $\hat{R} = \frac{1}{n} \sum_{i=1}^{n} \ell(f(X_i), Y_i)$  isn't very good

(In fact, one tends to not add bias when estimating R with  $\hat{R}$ )

#### RISK ESTIMATION: A GENERAL FORM

The problem is that  $\hat{R}$  is overly optimistic

The average optimism is

$$opt = * \mathbb{E}[R - \hat{R}]$$

Typically,  $\mathrm{opt}$  is positive as  $\hat{R}$  will underestimate the risk

( $\ast$  See Elements of Statistical Learning, Chapter 7 for details for a more precise statement)

## RISK ESTIMATION: A GENERAL FORM

It turns out for a variety of  $\ell$  (such as squared error and 0-1)

$$opt = \frac{2}{n} \sum_{i=1}^{n} Cov(\hat{f}(X_i), Y_i)$$

This is related intimately with degrees of freedom

$$df = \frac{1}{\sigma^2} \sum_{i=1}^n Cov(\hat{f}(X_i), Y_i) = \frac{n}{2\sigma^2} opt$$

$$(\sigma^2=\mathbb{V}Y_i)$$

EXAMPLE: For multiple regression (i.e.  $\hat{f}(X) = \hat{\beta}_{LS}^{\top} X$ ),

$$\mathrm{df} = \mathrm{trace}(\mathbb{X}(\mathbb{X}^{\top}\mathbb{X})^{-1}\mathbb{X}^{\top}) = \mathrm{rank}(\mathbb{X})$$

### A RISK ESTIMATE

Therefore, we get the following expression of risk

$$GIC = \hat{R} + \widehat{opt}$$

(Writing GIC indicates generalized information criterion)

Differing  $\widehat{\mathrm{opt}}$  leads to AIC, BIC, Mallows Cp, and others

opt depends on:

- a variance estimator  $\hat{\sigma}$
- a scaling term

# Various forms of risk estimates

$$\begin{aligned} \text{AIC} &= \hat{R} + 2 \cdot \text{df} \cdot \hat{\sigma}^2 / n \\ \text{AICc}(\hat{\beta}) &= \text{AIC} + \frac{2 \text{df} \cdot \left( \text{df} + 1 \right)}{n - \text{df} - 1} \\ \text{BIC} &= \hat{R} + \log(n) \cdot \text{df} \cdot \hat{\sigma}^2 / n \end{aligned}$$

Including more parameters leads to:

- a smaller  $\hat{R}$
- a larger opt

GOAL: Now, we can use one of the GIC procedures to tell us which model to use

(As long as  $\log n \ge 2$ , BIC picks a smaller model than AIC)

# Cross-validation

# A DIFFERENT APPROACH TO RISK ESTIMATION

Let  $(X_0, Y_0)$  be a test observation, identically distributed as an element in  $\mathcal{D}$ , but also independent of  $\mathcal{D}$ .

$$R(f) = \mathbb{E}\ell(f(X_0), Y_0) \underbrace{=}_{\text{regression}} \mathbb{E}(Y_0 - f(X_0))^2$$

Of course, the quantity  $(Y_0 - f(X_0))^2$  is an unbiased estimator of R(f) and hence we could use it to estimate R(f)

However, we don't have any such new observation

Or do we?

#### AN INTUITIVE IDEA

Let's set aside one observation and predict it

For example: Set aside  $(X_1, Y_1)$  and fit  $\hat{f}^{(1)}$  on  $(X_2, Y_2), \dots, (X_n, Y_n)$ 

(The notation  $\hat{f}^{(1)}$  just symbolizes leaving out the first observation before fitting  $\hat{f}$ )

$$R_1(\hat{f}^{(1)}) = (Y_1 - \hat{f}^{(1)}(X_1))^2$$

As the left off data point is independent of the data points used for estimation,

$$\mathbb{E}R_1(\hat{f}^{(1)}) \approx R(\hat{f})$$

#### LEAVE-ONE-OUT CROSS-VALIDATION

Cycling over all observations and taking the average produces leave-one-out cross-validation

$$CV_n(\hat{f}) = \frac{1}{n} \sum_{i=1}^n R_i(\hat{f}^{(i)}) = \frac{1}{n} \sum_{i=1}^n (Y_i - \hat{f}^{(i)}(X_i))^2.$$

# More General Cross-Validation Schemes

Let  $\mathcal{N} = \{1, \dots, n\}$  be the index set for  $\mathcal{D}$ 

• K-FOLD: Fix  $V = \{v_1, \dots, v_K\}$  such that  $v_j \cap v_k = \emptyset$  and  $\bigcup_j v_j = \mathcal{N}$ 

$$CV_{K}(\hat{f}) = \frac{1}{K} \sum_{v \in V} \frac{1}{|v|} \sum_{i \in v} (Y_{i} - \hat{f}^{(v)}(X_{i}))^{2}$$

- BOOTSTRAP: Instead of partitioning, we could make K bootstrap draws and average
- ullet FACTORIAL: We could make all subsets of  ${\mathcal N}$  and average

# More general cross-validation schemes: A comparison

- $\mathrm{CV}_K$  gets more computationally demanding as  $K \to n$
- The bias of  $\mathrm{CV}_K$  goes down, but the variance increases as  $K \to n$
- The factorial version isn't commonly used (Very computationally demanding)

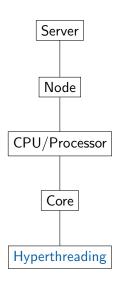
#### SUMMARY TIME

- CV Prediction risk consistent. Generally selects a model larger than necessary
- AIC Provides optimal risk estimation (and is asymptotically equivalent to CV). Inconsistent for model selection
- BIC Consistent for model selection consistent, sub optimal for risk estimation

Aside: There exist impossibility theorems stating that risk estimation procedures good at prediction are bad at model selection (and vice-versa)

# Parallelism

#### DISTRIBUTED COMPUTING HIERARCHY



#### EXAMPLE: A server might have

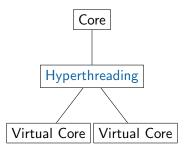
- 64 nodes
- 2 processors per node
- 16 cores per processor
- hyper threading

The goal is to somehow allocate a job so that these resources are used efficiently

Jobs are composed of threads, which are specific computations

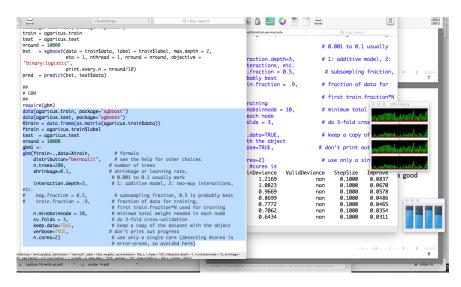
#### HYPERTHREADING

Developed by Intel, Hypertheading allows for each core to pretend to be two cores



This works by trading off computation and read-time for each core

# GBM: FIGURES



# PARALLELISM FOR CV IN R

Go to crossValidationParallel.R