NONLINEAR MARKOVIAN STOCHASTIC APPROXIMATION

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1 Preliminaries

Notations The Euclidean norm is denoted by $\|.\|$. The lowercase letter c and its derivatives c', c_0 , etc. denote universal numerical constants, whose value may change from line to line. As we are primarily interested in dependence of α and k, we adopt the following big-O notation: $\|f\| = \mathcal{O}(h(\alpha, k))$ if it holds that $\|f\| \le s \cdot \|h(\alpha, k)\|$ for some constant s > 0.

We use of the following iteration scheme:

$$\theta_{t+1} = \theta_t + \alpha \left(g(\theta_t, X_{t+1}) + \xi_{t+1}(\theta_t) \right)$$

where $g: \mathbb{R}^d \times \mathcal{X} \to \mathbb{R}^d$ is a deterministic function, $\{\xi_k\}_{k\geq 1}$ are i.i.d zero-mean random fields, and $\alpha > 0$ is a constant stepsize. We shall omit the superscript (α) in θ_k when the dependence on α is clear from the context.

In our settings, $\{X_n, n \in \mathbb{N}\}$ is a *state-dependent* (or controlled) Markov chain, *i.e.*, for any bounded measurable function $g: \mathbb{R}^d \times \mathcal{X} \to \mathbb{R}^d$,

$$\mathbb{E}\left[g\left(\theta_{n},X_{n+1}\right)|\mathcal{F}_{n}\right]=P_{\theta_{n}}g\left(\theta_{n},X_{n}\right)=\int g\left(\theta_{n},x\right)P_{\theta}\left(X_{n},dx\right),$$

where $P_{\theta}: \mathcal{X} \times \mathcal{X} \to \mathbb{R}_+$ is a Markov kernel such that, for each $\theta \in \Theta$, P_{θ} has a unique stationary distribution π_{θ} . Also, \mathscr{F}_n denotes the filtration generated by the random variables $(\theta_0, \{\xi_{m+1}\}_{m \le n-1}, \{X_m\}_{m \le n})$.

1.1 Assumptions

Assumption 1 For each $X \in \mathcal{X}$, the function $g(\theta, X)$ is three times continuously differentiable in θ with uniformly bounded first to third derivatives, i.e., $\sup_{\theta \in \mathbb{R}^d} \|g^{(i)}(\theta, X)\| < \infty$ for $i = 1, 2, 3, X \in \mathcal{X}$. Moreover, there exists a constant $L_1 > 0$ such that $(1) \|g^{(i)}(\theta, X) - g^{(i)}(\theta', X)\| \le L_1$, for all $\theta, \theta' \in \mathbb{R}^d$, i = 0, 1, 2 and $X \in \mathcal{X}$, and $(2) \|g^{(i)}(0, X)\| \le L_1$ for all $X \in \mathcal{X}$ and i = 0, 1, 2.

Assumption 1 implies that $g(\theta, X)$ is L_1 -Lipschitz w.r.t θ uniformly in X. The above assumption immediately implies that the growth of $\|g\|$ and $\|\bar{g}\|$ will be at most linear in θ , i.e., $\|g(\theta, X)\| \le L_1(\|\theta - \theta^*\| + 1)$ and $\|\bar{g}(\theta)\| \le L_1(\|\theta - \theta^*\| + 1)$. Similar implications also hold for the first and second derivatives in the same manner.

Assumption 2 There exists $\mu > 0$ such that $\langle \theta - \theta', \bar{g}(\theta) - \bar{g}(\theta') \rangle \leq -\mu \|\theta - \theta'\|^2, \forall \theta, \theta' \in \mathbb{R}^d$. Consequently, the target equation $\bar{g}(\theta) = 0$ has a unique solution θ^* .

Denote by \mathscr{F}_k the filtration generated by $\{X_{t+1}, \theta_t, \xi_{t+1}\}_{t=0}^{k-1} \cup \{X_{k+1}, \theta_k\}.$

Assumption 3 Let $p \in \mathbb{Z}_+$ be given. The noise sequence $(\xi_k)_{k \geq 1}$ is a collection of i.i.d random fields satisfying the following conditions with $L_{2,p} > 0$:

$$\mathbb{E}\left[\xi_{k+1}(\theta)|\mathcal{F}_k\right] = 0 \quad and \quad \mathbb{E}^{1/(2p)}\left[\|\xi_1(\theta)\|^{2p}\right] \le L_{2,p}\left(\|\theta - \theta^*\| + 1\right), \quad \forall \theta \in \mathbb{R}^d.$$

Define $C(\theta) = \mathbb{E}\left[\xi_1(\theta)^{\otimes 2}\right]$ and assume that $C(\theta)$ is at least twice differentiable. There also exists $M_{\epsilon}, k_{\epsilon} \geq 0$ such that for $\theta \in \mathbb{R}^d$, we have $\max_{i=1,2} \|C^{(i)}(\theta)\| \leq M_{\epsilon}\{1 + \|\theta - \theta^*\|^{k_{\epsilon}}\}$. In the sequel, we set $L := L_1 + L_2$, and without loss of generality, we assume $L \geq 2\mu$ for some technical reasons.

Assumption 4 There exists a Borel measurable function $\hat{g}: \mathbb{R}^d \times \mathcal{X} \to \mathbb{R}^d$ where for each $\theta \in \mathbb{R}^d$, $X \in \mathcal{X}$,

$$\hat{g}(\theta, X) - P_{\theta}\hat{g}(\theta, X) = g(\theta, X) - \bar{g}(\theta)$$
.

 $\begin{aligned} \textbf{Assumption 5} \quad & There \ exists \ L_{PH}^{(0)} < \infty \ and \ L_{PH}^{(1)} < \infty \ such \ that, \ for \ all \ \theta \in \mathbb{R}^d \ and \ X \in \mathcal{X}, \ one \ has \ \|\hat{g}\left(\theta,X\right)\| \leq L_{PH}^{(0)}, \ \|P_{\theta}\hat{g}\left(\theta,X\right)\| \leq L_{PH}^{(0)}. \ Moreover, \ for \ \left(\theta,\theta'\right) \in \mathcal{H}^2, \end{aligned}$

$$\sup_{X \in \mathcal{X}} \|P_{\theta} \, \hat{g} \, (\theta, X) - P_{\theta'} \, \hat{g} \, \left(\theta', X\right)\| \leq L_{PH}^{(1)} \|\theta - \theta'\|.$$

Assumption 6 For any $\theta, \theta' \in \mathbb{R}^d$, we have $\sup_{X \in \mathcal{X}} \|P_{\theta}(X, .) - P_{\theta'}(X, .)\|_{TV} \le L_P \|\theta - \theta'\|$.

Assumption 7 For any $\theta, \theta' \in \mathbb{R}^d$, we have $\sup_{X \in \mathcal{X}} \|g(\theta, X) - g(\theta', X)\| \le L_H \|\theta - \theta'\|$.

Assumption 8 There exists $\rho < 1$, $K_P < \infty$ such that

$$\sup_{\theta \in \mathbb{R}^d, X \in \mathcal{X}} \|P_{\theta}^n(X,.) - \pi_{\theta}(.)\|_{TV} \le \rho^n K_P,$$

Assumption 9 For any $\theta, \theta' \in \mathbb{R}^d$, we have $\sup_{X \in \mathcal{X}} \|\pi_{\theta}(X) - \pi_{\theta'}(X)\| \le L_S \|\theta - \theta'\|$. **Lemma 1** Assume that assumptions 6-8 hold. Then, for any $\theta \in \mathbb{R}^d$ and $X \in \mathcal{X}$,

$$\|\hat{g}(\theta, X)\| \leq \frac{\sigma K_P}{1-\rho},$$

$$\|P_{\theta}\hat{g}(\theta,X)\| \leq \frac{\sigma\rho K_P}{1-\rho}.$$

Moreover, for any $\theta, \theta' \in \mathbb{R}^d$ and $X \in \mathcal{X}$,

$$||P_{\theta}\hat{g}(\theta, X) - P_{\theta'}\hat{g}(\theta', X)|| \le ||\theta - \theta'||,$$

where

$$L_{PH}^{(1)} = \frac{K_P^2 \sigma L_P}{(1-\rho)^2} (2+K_P) + \frac{K_P}{1-\rho} L_H.$$

Proof of this lemma can be found in [2], Lemma 7.

2 Error Bound

2.1 Base Case

We first prove the following lemma because we are going to use that calculation in many different parts of the proof:

Lemma2. Using Assumptions 1, 3, 5, and for sufficiently small α and $t \ge 1$, we have

$$\mathbb{E}[\|\theta_t - \theta_{t-1}\|] \le \alpha L(\mathbb{E}[\|\theta_{t-1} - \theta^*\|] + 1).$$

Proof. We have

$$\begin{split} \mathbb{E}\left[\left\|\theta_{t} - \theta_{t-1}\right\|\right] &\leq \alpha \mathbb{E}\left[\left\|g\left(\theta_{t-1}, X_{t}\right) + \xi_{t}\left(\theta_{t-1}\right)\right\|\right] \\ &\leq \alpha L_{1}\left(\mathbb{E}\left[\left\|\theta_{t-1} - \theta^{*}\right\|\right] + 1\right) + \alpha L_{2}\left(\mathbb{E}\left[\left\|\theta_{t-1} - \theta^{*}\right\|\right] + 1\right) \\ &\leq \alpha L\left(\mathbb{E}\left[\left\|\theta_{t-1} - \theta^{*}\right\|\right] + 1\right) \end{split}$$

where the first line follows from ??, second line from the Lipschitzness condition and the assumption of

$$\mathbb{E}^{1/2}\left[\left\|\xi_{t}\left(\theta_{t-1}\right)\right\|^{2}|\mathcal{F}_{t-1}\right]\leq L_{2}\left(\mathbb{E}\left[\left\|\theta_{t-1}\right\|\right]+1\right),$$

and the third line from ??.

For the base case analysis, we can write:

$$\begin{split} &\mathbb{E}\left[\left\|\boldsymbol{\theta}_{k+1} - \boldsymbol{\theta}^*\right\|^2\right] - \mathbb{E}\left[\left\|\boldsymbol{\theta}_k - \boldsymbol{\theta}^*\right\|^2\right] = \\ &2\alpha\mathbb{E}\left[\left\langle\boldsymbol{\theta}_k - \boldsymbol{\theta}^*, g\left(\boldsymbol{\theta}_k, X_{k+1}\right)\right\rangle\right] + \alpha^2\mathbb{E}\left[\left\|g\left(\boldsymbol{\theta}_k, X_{k+1}\right)\right\|^2\right] + \alpha^2\mathbb{E}\left[\left\|\boldsymbol{\xi}_{k+1}\left(\boldsymbol{\theta}_k\right)\right\|^2\right] = \\ &2\alpha\mathbb{E}\left[\left\langle\boldsymbol{\theta}_k - \boldsymbol{\theta}^*, g\left(\boldsymbol{\theta}_k, X_{k+1}\right) - \bar{g}\left(\boldsymbol{\theta}_k\right)\right\rangle\right] + 2\alpha\mathbb{E}\left[\left\langle\boldsymbol{\theta}_k - \boldsymbol{\theta}^*, \bar{g}\left(\boldsymbol{\theta}_k\right)\right\rangle\right] + \alpha^2\mathbb{E}\left[\left\|g\left(\boldsymbol{\theta}_k, X_{k+1}\right)\right\|\right] + \alpha^2\mathbb{E}\left[\left\|\boldsymbol{\xi}_{k+1}\left(\boldsymbol{\theta}_k\right)\right\|^2\right]. \end{split}$$

It is easy to see that under Strong Monotonicity assumption, we have

$$\left\langle \theta_{k} - \theta^{*}, \bar{g}\left(\theta_{k}\right) \right\rangle = \left\langle \theta_{k} - \theta^{*}, \bar{g}\left(\theta_{k}\right) - \bar{g}\left(\theta^{*}\right) \right\rangle \leq -\mu \|\theta_{k} - \theta^{*}\|^{2}.$$

Additionally, under Assumption 1 and 3, we have the following upper bound

$$\begin{split} &\alpha^{2}\left(\mathbb{E}\left[\left\|g\left(\theta_{k},X_{k+1}\right)\right\|^{2}\right]+\mathbb{E}\left[\left\|\xi_{k+1}\left(\theta_{k}\right)\right\|^{2}\right]\right)\\ &\leq\alpha^{2}\left(L_{1}^{2}\mathbb{E}\left[\left(\left\|\theta_{k}-\theta^{*}\right\|+1\right)^{2}\right]+L_{2}^{2}\mathbb{E}\left[\left(\left\|\theta_{k}-\theta^{*}\right\|+1\right)^{2}\right]\right)\\ &\leq2\alpha^{2}L^{2}\left(\mathbb{E}\left[\left\|\theta_{k}-\theta^{*}\right\|^{2}\right]+1\right). \end{split}$$

Therefore, we have

$$\mathbb{E}\left[\left\|\theta_{k+1} - \theta^*\right\|^2\right] \le \left(1 - 2\alpha\left(-\alpha L^2 + \mu\right)\right)\mathbb{E}\left[\left\|\theta_k - \theta^*\right\|^2\right] + 2\alpha^2 L^2 + 2\alpha\mathbb{E}\left[\left\langle\theta_k - \theta^*, g\left(\theta_k, X_{k+1}\right) - \bar{g}\left(\theta_k\right)\right\rangle\right]$$

Solving this recursion gives us the following inequality:

$$\begin{split} \mathbb{E}\left[\left\|\theta_{k+1}-\theta^*\right\|^2\right] &\leq \left(1-2\alpha\left(-\alpha L^2+\mu\right)\right)^{k+1}\mathbb{E}\left[\left\|\theta_0-\theta^*\right\|^2\right] \\ &+\sum_{t=0}^k\left(1-2\alpha\left(-\alpha L^2+\mu\right)\right)^t 2\alpha^2 L^2 \\ &+\sum_{t=0}^k 2\alpha\left(1-2\alpha\left(-\alpha L^2+\mu\right)\right)^{k-t}\mathbb{E}\left[\left\langle\theta_t-\theta^*,g\left(\theta_t,X_{t+1}\right)-\bar{g}\left(\theta_t\right)\right\rangle\right]. \end{split}$$

For notational simplicity we define $\gamma_t \coloneqq 2\alpha \left(1-2\alpha \left(-\alpha L^2+\mu\right)\right)^{k-t}$ for $0 \le t \le k$.

The second term above is just a geometric series. According to Lemma 12 of [1], this equals to $\frac{\alpha L^2 \left[1-\left(1-2\alpha(-\alpha L^2+\mu)\right)^{k+1}\right]}{-\alpha L^2+\mu}.$

Now, we can upper bound the third summand using below decomposition:

$$\mathbb{E}\left[\left.\sum_{t=0}^{k}\gamma_{t}\left\langle\theta_{t}-\theta^{*},g(\theta_{t},X_{t+1})-\bar{g}(\theta_{t})\right\rangle\right.\right]=A_{1}+A_{2}+A_{3}+A_{4}+A_{5}$$

with

$$\begin{split} A_1 &:= \mathbb{E}\left[\sum_{t=1}^k \gamma_t \left\langle \theta_t - \theta^*, \hat{g}\left(\theta_t, X_{t+1}\right) - P_{\theta_t} \hat{g}\left(\theta_t, X_t\right) \right\rangle\right], \\ A_2 &:= \mathbb{E}\left[\sum_{t=1}^k \gamma_t \left\langle \theta_t - \theta^*, P_{\theta_t} \hat{g}\left(\theta_t, X_t\right) - P_{\theta_{t-1}} \hat{g}\left(\theta_{t-1}, X_t\right) \right\rangle\right], \\ A_3 &:= \mathbb{E}\left[\sum_{t=1}^k \gamma_t \left\langle \theta_t - \theta_{t-1}, P_{\theta_{t-1}} \hat{g}\left(\theta_{t-1}, X_t\right) \right\rangle\right], \\ A_4 &:= \mathbb{E}\left[\sum_{t=1}^k \left(\gamma_t - \gamma_{t-1}\right) \left\langle \theta_{t-1} - \theta^*, P_{\theta_{t-1}} \hat{g}\left(\theta_{t-1} - \theta^*, X_t\right) \right\rangle\right], \\ A_5 &:= \mathbb{E}\left[\gamma_0 \left\langle \theta_0 - \theta^*, \hat{g}\left(\theta_0, X_0\right) \right\rangle\right] + \mathbb{E}\left[\gamma_k \left\langle \theta_k - \theta^*, P_{\theta_k} \hat{g}\left(\theta_k, X_{k+1}\right) \right\rangle\right], \end{split}$$

For A_1 , we note that $\hat{g}(\theta_t, X_{t+1}) - P_{\theta_t} \hat{g}(\theta_t, X_t)$ is a martingale difference sequence [cf. ?] and therefore we have $A_1 = 0$ by taking the total expectation.

For A_2 , applying Cauchy-Schwarz inequality and ??, we have

$$\begin{split} A_{2} &\leq \sum_{t=1}^{k} L_{PH}^{(1)} \gamma_{t} \mathbb{E} \left[\| \theta_{t} - \theta^{*} \| \| \theta_{t} - \theta_{t-1} \| \right] \\ &= \sum_{t=1}^{k} \alpha L_{PH}^{(1)} \gamma_{t} \mathbb{E} \left[\| \theta_{t} - \theta^{*} \| \| g(\theta_{t-1}, X_{t}) + \xi_{t}(\theta_{t-1}) \| \right] \\ &\leq \sum_{t=1}^{k} \alpha L_{PH}^{(1)} \gamma_{t} \mathbb{E} \left[\left(\| \theta_{t} - \theta_{t-1} \| + \| \theta_{t-1} - \theta^{*} \| \right) \left(\| g(\theta_{t-1}, X_{t}) \| + \| \xi_{t}(\theta_{t-1}) \| \right) \right] \\ &\leq \sum_{t=1}^{k} \alpha L_{PH}^{(1)} \gamma_{t} \left(L_{1} \left(\mathbb{E} \left[\| \theta_{t-1} - \theta^{*} \|^{2} \right] + \mathbb{E} \left[\| \theta_{t} - \theta_{t-1} \| \| \theta_{t-1} - \theta^{*} \| \right] + \mathbb{E} \left[\| \theta_{t-1} - \theta^{*} \| \right] + \mathbb{E} \left[\| \theta_{t} - \theta_{t-1} \| \| \xi_{t}(\theta_{t-1}) \| \right] \right) \\ &+ \mathbb{E} \left[\| \theta_{t} - \theta_{t-1} \| \| \xi_{t}(\theta_{t-1}) \| \right] + \mathbb{E} \left[\| \theta_{t-1} - \theta^{*} \| \| \xi_{t}(\theta_{t-1}) \| \right] \right) \end{split}$$

where the second line follows from ?? and the third line follows from the triangle inequality. Now we upper the compound terms in the last line's parentheses:

$$\begin{split} \mathbb{E}\left[\left\|\theta_{t}-\theta_{t-1}\right\|\,\left\|\theta_{t-1}-\theta^{*}\right\|\right] &= \mathbb{E}\left[\mathbb{E}\left[\left\|\theta_{t}-\theta_{t-1}\right\|\,\left\|\theta_{t-1}-\theta^{*}\right\|\right|\mathcal{F}_{t-1}\right]\right] \\ &\leq \mathbb{E}\left[\alpha L\left(\left\|\theta_{t-1}-\theta^{*}\right\|+1\right)\left\|\theta_{t-1}-\theta^{*}\right\|\right] \\ &\leq \frac{\alpha L\left(3\mathbb{E}\left[\left\|\theta_{t-1}-\theta^{*}\right\|^{2}\right]+1\right)}{2} \end{split}$$

where in the second line we used Lemma 2 and in the last line we used $u \le \frac{u^2+1}{2}$.

$$\begin{split} \mathbb{E}\left[\|\theta_{t} - \theta_{t-1}\| \, \|\xi_{t}\left(\theta_{t-1}\right)\| \right] &\leq \mathbb{E}\left[\alpha\left(\|g\left(\theta_{t-1}, X_{t}\right)\| + \|\xi_{t}\left(\theta_{t-1}\right)\| \right) \|\xi_{t}\left(\theta_{t-1}\right)\| \right] \\ &\leq \mathbb{E}\left[\alpha \|\xi_{t}\left(\theta_{t-1}\right)\|^{2} + \alpha L_{1}\left(\|\theta_{t-1} - \theta^{*}\| + 1 \right) \|\xi_{t}\left(\theta_{t-1}\right)\| \right] \\ &\leq L\left(\mathbb{E}\left[\|\theta_{t-1} - \theta^{*}\| \right] + 1 \right) \end{split}$$

where the first inequality follows from ??, second line from Lemma 2 and in the last line we used

boundedness property of the noise and sufficiently small α .

$$\begin{split} \mathbb{E}\left[\left\|\boldsymbol{\theta}_{t-1} - \boldsymbol{\theta}^*\right\| \, \left\|\boldsymbol{\xi}_t\left(\boldsymbol{\theta}_{t-1}\right)\right\|\right] &= \mathbb{E}\left[\mathbb{E}\left[\left\|\boldsymbol{\theta}_{t-1} - \boldsymbol{\theta}^*\right\| \, \left\|\boldsymbol{\xi}_t\left(\boldsymbol{\theta}_{t-1}\right)\right\| \left|\mathcal{F}_{t-1}\right|\right]\right] \\ &\leq \mathbb{E}\left[L_2\|\boldsymbol{\theta}_{t-1} - \boldsymbol{\theta}^*\| \left(\left\|\boldsymbol{\theta}_{t-1} - \boldsymbol{\theta}^*\right\| + 1\right)\right] \\ &\leq \frac{L_2\left(3\|\boldsymbol{\theta}_{t-1} - \boldsymbol{\theta}^*\|^2 + 1\right)}{2} \end{split}$$

where the second line follows from ?? and in the last line we used $u \le \frac{u^2+1}{2}$.

Summing up all these bounds, we can write for A_2 :

$$\begin{split} A_2 &\leq \sum_{t=1}^k \alpha L_{PH}^{(1)} \gamma_t \left(2L \left(\mathbb{E} \left[\| \theta_{t-1} - \theta^* \| \right] + 1 \right) + 2L \left(3\mathbb{E} \left[\| \theta_{t-1} - \theta^* \|^2 \right] + 1 \right) + \mathbb{E} \left[\| \theta_{t-1} - \theta^* \|^2 \right] \right) \\ &\leq \sum_{t=1}^k \alpha L L_{PH}^{(1)} \gamma_t \left(8\mathbb{E} \left[\| \theta_{t-1} - \theta^* \|^2 \right] + 5 \right) \end{split}$$

which in the last line we again used $u \le \frac{u^2+1}{2}$ property.

For A_3 , we obtain

$$\begin{split} A_{3} &\leq \sum_{t=1}^{k} \gamma_{t} \mathbb{E} \left[\| \theta_{t} - \theta_{t-1} \| \| P_{\theta_{t-1}} \hat{g} \left(\theta_{t-1}, X_{t} \right) \| \right] \\ &\leq \sum_{t=1}^{k} L_{PH}^{(0)} \gamma_{t} \mathbb{E} \left[\| g \left(\theta_{t-1}, X_{t} \right) + \xi_{t} (\theta_{t-1}) \| \right] \\ &\leq \sum_{t=1}^{k} L_{PH}^{(0)} \gamma_{t} \left(L_{1} \left(\mathbb{E} \left[\| \theta_{t-1} - \theta^{*} \| \right] + 1 \right) + L_{2} \left(\mathbb{E} \left[\| \theta_{t-1} - \theta^{*} \| \right] + 1 \right) \right) \\ &\leq \sum_{t=1}^{k} \alpha L L_{PH}^{(0)} \gamma_{t} \left(\mathbb{E} \left[\| \theta_{t-1} - \theta^{*} \| \right] + 1 \right) \end{split}$$

where second line follows from ?? and third line follows from ?? .

For A_4 , we have

$$\begin{split} A_4 &\leq \sum_{t=1}^k \left| \gamma_t - \gamma_{t-1} \right| \mathbb{E} \left[\left\| \theta_{t-1} - \theta^* \right\| \left\| P_{\theta_{t-1}} \hat{g} \left(\theta_{t-1}, X_t \right) \right\| \right] \\ &\leq \sum_{t=1}^k L_{PH}^{(0)} \left| \gamma_t - \gamma_{t-1} \right| \mathbb{E} \left[\left\| \theta_{t-1} - \theta^* \right\| \right] \end{split}$$

Finally, for A_5 , we obtain

$$A_5 \le L_{DH}^{(0)} \left(\gamma_0 \mathbb{E} \left[\| \theta_0 - \theta^* \| \right] + \gamma_k \mathbb{E} \left[\| \theta_k - \theta^* \| \right] \right)$$

which follows from Cacuhy-Scwarz inequality and ??.

Combining the above terms gives us:

$$\begin{split} \mathbb{E}\left[\sum_{t=1}^{k}\gamma_{t}\left\langle\theta_{t}-\theta^{*},g\left(\theta_{t},X_{t+1}-\bar{g}\left(\theta_{t}\right)\right)\right\rangle\right] \leq \sum_{t=0}^{k-1}\alpha LL_{PH}^{(1)}\gamma_{t+1}\left(5+8\mathbb{E}\left[\left\|\theta_{t}-\theta^{*}\right\|^{2}\right]\right) + \sum_{t=0}^{k-1}\alpha LL_{PH}^{(0)}\gamma_{t+1}\left(\mathbb{E}\left[\left\|\theta_{t-1}-\theta^{*}\right\|\right]+1\right) + \\ \sum_{t=0}^{k-1}L_{PH}^{(0)}|\gamma_{t}-\gamma_{t+1}|\,\mathbb{E}\left[\left\|\theta_{t}-\theta^{*}\right\|\right] + L_{PH}^{(0)}\left(\gamma_{0}\mathbb{E}\left[\left\|\theta_{0}-\theta^{*}\right\|\right]+\gamma_{k}\mathbb{E}\left[\left\|\theta_{k}-\theta^{*}\right\|\right]\right) \end{split}$$

now it should be noticed that as long as the α satisfies $\alpha \leq \frac{\mu}{L^2}$, we have $\gamma_t \leq \gamma_{t+1}$. Thus, we can simplify

the above upper bound and write it this way:

$$\begin{split} \mathbb{E}\left[\left.\sum_{t=0}^{k}\gamma_{t}\left\langle\theta_{t}-\theta^{*},g\left(\theta_{t},X_{t+1}-\bar{g}\left(\theta_{t}\right)\right)\right\rangle\right] &\leq \sum_{t=0}^{k-1}\alpha LL_{PH}^{(1)}\gamma_{t+1}\left(5+8\mathbb{E}\left[\left\|\theta_{t}-\theta^{*}\right\|^{2}\right]\right)+\\ &\qquad \qquad \sum_{t=0}^{k-1}L_{PH}^{(0)}\left(\left(\alpha L+1\right)\gamma_{t+1}-\gamma_{t}\right)\mathbb{E}\left[\left\|\theta_{t}-\theta^{*}\right\|\right]+\\ &\qquad \qquad \sum_{t=0}^{k-1}\alpha LL_{PH}^{(0)}\gamma_{t+1}+L_{PH}^{(0)}\left(\gamma_{0}\mathbb{E}\left[\left\|\theta_{0}-\theta^{*}\right\|\right]+\gamma_{k}\mathbb{E}\left[\left\|\theta_{k}-\theta^{*}\right\|\right]\right) \end{split}$$

Hence, using the derived upper bounds from the above terms, we have:

$$\begin{split} \mathbb{E}\left[\left\|\theta_{k+1} - \theta^*\right\|^2\right] &\leq \sum_{t=0}^{k-1} \alpha L L_{PH}^{(1)} \gamma_{t+1} \left(5 + 8\mathbb{E}\left[\left\|\theta_t - \theta^*\right\|^2\right]\right) + \sum_{t=0}^{k-1} L_{PH}^{(0)} \left((\alpha L + 1) \gamma_{t+1} - \gamma_t\right) \mathbb{E}\left[\left\|\theta_t - \theta^*\right\|\right] + \\ & \left(1 - 2\alpha \left(-\alpha L^2 + \mu\right)\right) \gamma_0 \mathbb{E}\left[\left\|\theta_0 - \theta^*\right\|^2\right] + L_{PH}^{(0)} \gamma_0 \mathbb{E}\left[\left\|\theta_0 - \theta^*\right\|\right] + L_{PH}^{(0)} \gamma_k \mathbb{E}\left[\left\|\theta_k - \theta^*\right\|\right] + \\ & 2\alpha^2 L L_{PH}^{(0)} \left[\frac{1 - \left(1 - 2\alpha \left(-\alpha L^2 + \mu\right)\right)^k}{1 - \left(1 - 2\alpha \left(-\alpha L^2 + \mu\right)\right)}\right] + \frac{\alpha L^2 \left[1 - \left(1 - 2\alpha \left(-\alpha L^2 + \mu\right)\right)^{k+1}\right]}{-\alpha L^2 + \mu} \end{split}$$

for further notation simplicity we define $c_{1,t} := 2\alpha^2 L L_{PH}^{(0)} \left[\frac{1 - (1 - 2\alpha(-\alpha L^2 + \mu))^t}{1 - (1 - 2\alpha(-\alpha L^2 + \mu))} \right] + \frac{\alpha L^2 \left[1 - (1 - 2\alpha(-\alpha L^2 + \mu))^{t+1} \right]}{-\alpha L^2 + \mu}$ for $0 \le t \le k$. Now to write down this upper bound in a way in which it only depends on $\|\theta_0 - \theta^*\|$ related terms and constants, we can write:

$$\begin{split} \mathbb{E} \left[\| \theta_{k+1} - \theta^* \|^2 \right] &\leq \sum_{t=0}^{k-1} \left[8\alpha L L_{PH}^{(1)} \gamma_{t+1} \mathbb{E} \left[\| \theta_t - \theta^* \|^2 \right] + 5\alpha L L_{PH}^{(1)} \gamma_{t+1} \right] + \sum_{t=0}^{k-1} L_{PH}^{(0)} \left((\alpha L + 1) \gamma_{t+1} - \gamma_t \right) \mathbb{E} \left[\| \theta_t - \theta^* \| \right] + \\ & \left(1 - 2\alpha \left(-\alpha L^2 + \mu \right) \right) \gamma_0 \mathbb{E} \left[\| \theta_0 - \theta^* \|^2 \right] + L_{PH}^{(0)} \gamma_0 \mathbb{E} \left[\| \theta_0 - \theta^* \| \right] + L_{PH}^{(0)} \gamma_k \mathbb{E} \left[\| \theta_k - \theta^* \| \right] + c_{1,k} \\ &= \sum_{t=0}^{k-1} 8\alpha L L_{PH}^{(1)} \gamma_{t+1} \mathbb{E} \left[\| \theta_t - \theta^* \|^2 \right] + 5\alpha L L_{PH}^{(1)} \sum_{t=0}^{k-1} \gamma_{t+1} + \sum_{t=0}^{k-1} L_{PH}^{(0)} \left((\alpha L_1) \gamma_{t+1} - \gamma_t \right) \mathbb{E} \left[\| \theta_t - \theta^* \| \right] + \\ & \left(1 - 2\alpha \left(-\alpha L^2 + \mu \right) \right) \gamma_0 \mathbb{E} \left[\| \theta_0 - \theta^* \|^2 \right] + L_{PH}^{(0)} \gamma_0 \mathbb{E} \left[\| \theta_0 - \theta^* \| \right] + L_{PH}^{(0)} \gamma_k \mathbb{E} \left[\| \theta_k - \theta^* \| \right] + \\ & \left(1 - 2\alpha \left(-\alpha L^2 + \mu \right) \right) \gamma_0 \mathbb{E} \left[\| \theta_0 - \theta^* \|^2 \right] + L_{PH}^{(0)} \gamma_0 \mathbb{E} \left[\| \theta_0 - \theta^* \| \right] + L_{PH}^{(0)} \gamma_k \mathbb{E} \left[\| \theta_k - \theta^* \| \right] + c_{1,k} + \\ & \frac{10\alpha^2 L L_{PH}^{(1)} \left[1 - \left(1 - 2\alpha \left(-\alpha L^2 + \mu \right) \right)^k \right]}{\left[1 - \left(1 - 2\alpha \left(-\alpha L^2 + \mu \right) \right) \right]} \end{split}$$

where the last equality follows from the definition of γ_t s. Similarly we define $c_{2,t} \coloneqq \frac{10\alpha^2 L L_{PH}^{(1)} \left[1 - (1 - 2\alpha(-\alpha L^2 + \mu))^t\right]}{\left[1 - (1 - 2\alpha(-\alpha L^2 + \mu))\right]}$ for $0 \le t \le k$. So we can write it as

$$\begin{split} \mathbb{E}\left[\left\|\theta_{k+1} - \theta^*\right\|^2\right] &\leq \sum_{t=0}^{k-1} 8\alpha L L_{PH}^{(1)} \gamma_{t+1} \mathbb{E}\left[\left\|\theta_t - \theta^*\right\|^2\right] + \sum_{t=0}^{k-1} L_{PH}^{(0)} \left((\alpha L + 1)\gamma_{t+1} - \gamma_t\right) \mathbb{E}\left[\left\|\theta_t - \theta^*\right\|\right] + \\ & \left(1 - 2\alpha \left(-\alpha L^2 + \mu\right)\right) \gamma_0 \mathbb{E}\left[\left\|\theta_0 - \theta^*\right\|^2\right] + L_{PH}^{(0)} \gamma_0 \mathbb{E}\left[\left\|\theta_0 - \theta^*\right\|\right] + L_{PH}^{(0)} \gamma_k \mathbb{E}\left[\left\|\theta_k - \theta^*\right\|\right] + c_{1,k} + c_{2,k} \right] \end{split}$$

Now for the second term on RHS, we note that since $L \ge 2\mu$,

$$(\alpha L + 1)\gamma_{t+1} - \gamma_t \le 2\alpha L\gamma_{t+1}, \quad \mathbb{E}[\|\theta_t - \theta^*\|] \le \sqrt{\mathbb{E}[\|\theta_t - \theta^*\|^2]},$$

and consequently

$$\begin{split} &\frac{1}{(1-2\alpha(-\alpha L^2+\mu))^k} \sum_{t=0}^{k-1} L_{PH}^{(0)} \big((\alpha L+1) \gamma_{t+1} + \gamma_t \big) \mathbb{E} \big[\|\theta_t - \theta^*\| \big] \\ & \leq 4 L_{PH}^{(0)} L \alpha^2 \sum_{t=0}^{k-1} \frac{1}{(1-2\alpha(-\alpha L^2+\mu))^{t+1}} \sqrt{\mathbb{E} \big[\|\theta_t - \theta^*\|^2 \big]} \\ & \leq 4 L_{PH}^{(0)} L \alpha^2 \Big(\sum_{t=0}^{k-1} \frac{1}{(1-2\alpha(-\alpha L^2+\mu))^{t+1}} \Big)^{1/2} \Big(\sum_{t=1}^{k-1} \frac{1}{(1-2\alpha(-\alpha L^2+\mu))^{t+1}} \mathbb{E} \big[\|\theta_t - \theta^*\|^2 \big] \Big)^{1/2} \\ & \leq 4 L_{PH}^{(0)} L \alpha^2 \cdot \sum_{t=0}^{k-1} \frac{1}{(1-2\alpha(-\alpha L^2+\mu))^{t+1}} \mathbb{E} \big[\|\theta_t - \theta^*\|^2 \big] + \frac{1}{-\alpha L^2 + \mu} \cdot \frac{2\alpha L L_{PH}^{(0)}}{(1-2\alpha(-\alpha L^2+\mu))^k}. \end{split}$$

We also note that

$$\frac{\gamma_k}{(1 - 2\alpha(-\alpha L^2 + \mu))^k} \mathbb{E}[\|\theta_k - \theta^*\|] \le \alpha \frac{\mathbb{E}[\|\theta_k - \theta^*\|^2]}{(1 - 2\alpha(-\alpha L^2 + \mu))^k} + \frac{\alpha}{(1 - 2\alpha(-\alpha L^2 + \mu))^k}.$$

similarly

$$\frac{\gamma_0}{(1-2\alpha(-\alpha L^2+\mu))^k} \mathbb{E}\big[\|\theta_0-\theta^*\|\big] \le \alpha \mathbb{E}\big[\|\theta_0-\theta^*\|^2\big] + \alpha.$$

and we also define for $0 \le t \le k$

$$c_{3,t} \coloneqq \frac{1}{-\alpha L^2 + \mu} \frac{2\alpha L L_{PH}^{(0)}}{\left(1 - 2\alpha \left(-\alpha L^2 + \mu\right)\right)^t} + \frac{\alpha L_{PH}^{(0)}}{\left(1 - 2\alpha \left(-\alpha L^2 + \mu\right)\right)^t} + \alpha L_{PH}^{(0)}$$

to wrap up all the remainder terms.

Substituting back and rearranging with also defining $c'_{2,t} := \frac{c_{2,t}}{(1-2\alpha(-\alpha L^2+\mu))^t}$ and $c'_{1,t} := \frac{c_{1,t}}{(1-2\alpha(-\alpha L^2+\mu))^t}$, yields

$$\frac{\mathbb{E}\left[\|\theta_{k+1} - \theta^*\|^2\right]}{(1 - 2\alpha(-\alpha L^2 + \mu))^k} \leq \frac{\alpha}{\left(1 - 2\alpha\left(-\alpha L^2 + \mu\right)\right)^k} \mathbb{E}\left[\|\theta_k - \theta^*\|^2\right] + \sum_{t=0}^{k-1} \frac{\alpha\left(8LL_{PH}^{(1)} + 4\alpha\left(1 - 2\alpha\left(-\alpha L^2 + \mu\right)\right)^{-1}LL_{PH}^{(0)}\right)}{\left(1 - 2\alpha\left(-\alpha L^2 + \mu\right)\right)^t} \mathbb{E}\left[\|\theta_t - \theta^*\|^2\right] + \alpha\mathbb{E}\left[\|\theta_0 - \theta^*\|^2\right] + c_{1,k}' + c_{2,k}' + c_{3,k}.$$

for sufficiently small α s, we have

$$\alpha \left(\frac{3}{2} L L_{PH}^{(1)} + 2\alpha \left(1 - 2\alpha \left(-\alpha L^2 + \mu \right) \right)^{-1} L L_{PH}^{(0)} \right) \le 4\alpha \left((\alpha L + 1) L_{PH}^{(0)} + 2L L_{PH}^{(1)} \right)$$

using above simplification we can rewrite our upper bound as

$$\frac{\mathbb{E} \left[\left\| \theta_{k+1} - \theta^* \right\|^2 \right]}{(1 - 2\alpha (-\alpha L^2 + \mu))^k} \leq 4\alpha \left((\alpha L + 1) \, L_{PH}^{(0)} + 2L L_{PH}^{(1)} \right) \sum_{t=0}^k \frac{\mathbb{E} \left[\left\| \theta_t - \theta^* \right\|^2 \right]}{\left(1 - 2\alpha \left(-\alpha L^2 + \mu \right) \right)^t} + \alpha \mathbb{E} \left[\left\| \theta_0 - \theta^* \right\|^2 \right] + c_{1,k}' + c_{2,k}' + c_{3,k}'.$$

For solving the above recursion, we first define $S_t \coloneqq 4\alpha \left((\alpha L + 1) \, L_{PH}^{(0)} + 2L L_{PH}^{(1)} \right) \sum_{l=0}^t \frac{\mathbb{E}[\|\theta_l - \theta^*\|^2]}{(1 - 2\alpha (-\alpha L^2 + \mu))^l}$ for

 $0 \le t \le k$. Also we use $C_t \coloneqq c'_{1,t} + c'_{2,t} + c_{3,t}$ and for $0 \le t \le k$, defining constant terms. Now we can write

$$\frac{\mathbb{E}\left[\left\|\theta_{t+1} - \theta^*\right\|^2\right]}{\left(1 - 2\alpha\left(-\alpha L^2 + \mu\right)\right)^t} \leq S_t + \alpha \mathbb{E}\left[\left\|\theta_0 - \theta^*\right\|^2\right] + C_t.$$

using this expansion, we should first notice that

$$\begin{split} \frac{S_{t}}{S_{t-1}} &= \frac{S_{t-1} + 4\alpha \left((\alpha L + 1) L_{PH}^{(0)} + 2L L_{PH}^{(1)} \right) \frac{\mathbb{E}[\|\theta_{t} - \theta^{*}\|^{2}]}{(1 - 2\alpha (-\alpha L^{2} + \mu))^{t}}}{S_{t-1}} \\ &= 1 + 4\alpha \left((\alpha L + 1) L_{PH}^{(0)} + 2L L_{PH}^{(1)} \right) \frac{S_{t-1} + \alpha \mathbb{E}\left[\|\theta_{0} - \theta^{*}\|^{2} \right] + C_{t-1}}{S_{t-1}} \\ &\leq 1 + 8\alpha \left((\alpha L + 1) L_{PH}^{(0)} + 2L L_{PH}^{(1)} \right). \end{split}$$

Now, since we have $S_0 = 4\alpha \left((\alpha L + 1) L_{PH}^{(0)} + 2L L_{PH}^{(1)} \right) \mathbb{E} \left[\|\theta_0 - \theta^*\|^2 \right]$, thus

$$S_t \leq 4\alpha \left((\alpha L + 1) L_{PH}^{(0)} + 2L L_{PH}^{(1)} \right) \left[1 + 8\alpha \left((\alpha L + 1) L_{PH}^{(0)} + 2L L_{PH}^{(1)} \right) \right]^t \mathbb{E} \left[\|\theta_0 - \theta^*\|^2 \right].$$

Substituting this upper bound into previous equations we get

$$\begin{split} & \mathbb{E}\left[\left\| \theta_{k+1} - \theta^* \right\|^2 \right] \leq \\ & \left(1 - 2\alpha \left(-\alpha L^2 + \mu \right) \right)^k \left[4\alpha \left((\alpha L + 1) \, L_{PH}^{(0)} + 2L L_{PH}^{(1)} \right) \left(1 + 8\alpha \left((\alpha L + 1) \, L_{PH}^{(0)} + 2L L_{PH}^{(1)} \right) \right)^k \mathbb{E}\left[\left\| \theta_0 - \theta^* \right\|^2 \right] + \alpha \mathbb{E}\left[\left\| \theta_0 - \theta^* \right\|^2 \right] + C_k \right] \end{split}$$

Choosing α sufficiently small, we can simplify the above inequality and write it as follows

$$\mathbb{E}\left[\left\|\boldsymbol{\theta}_{k+1} - \boldsymbol{\theta}^*\right\|^2\right] \leq \tilde{c}_1 \cdot \left(1 - 2\alpha\mu\right)^{k+1} \mathbb{E}\left[\left\|\boldsymbol{\theta}_0 - \boldsymbol{\theta}^*\right\|^2\right] + \tilde{c}_2 \cdot 2\alpha L\left(L_{PH}^{(0)} + \alpha L_{PH}^{(1)}\right)$$

where \tilde{c}_1 and \tilde{c}_2 are $\mathcal{O}(1)$ constants.

2.2 General Case

Similar to the previous case, we first prove a useful lemma:

Lemma3. Using Assumptions 1, 3, 5, and $m, n \ge 1$, we have

$$\mathbb{E}\left[\left\|\theta_{t}-\theta_{t-1}\right\|^{m}\left\|\theta_{t-1}-\theta^{*}\right\|^{n}\right]\leq 2^{2m+n-2}\alpha^{m}L^{m}\left(\mathbb{E}\left[\left\|\theta_{t-1}-\theta^{*}\right\|^{m+n}\right]+1\right).$$

Also when m = 0, this upper bound can be written as

$$\mathbb{E}\left[\|\theta_{t-1} - \theta^*\|^n\right] \le 2^{n-1} \left(\mathbb{E}\left[\|\theta_{t-1} - \theta^*\|^n\right] + 1\right). \tag{1}$$

Proof. We have

$$\begin{split} \mathbb{E} \big[\| \theta_{t} - \theta_{t-1} \|^{m} \| \theta_{t-1} - \theta^{*} \|^{n} \big] &\leq \alpha^{m} \mathbb{E} \big[\| g \left(\theta_{t-1}, X_{t} \right) + \xi_{t} \left(\theta_{t-1} \right) \|^{m} \| \theta_{t-1} - \theta^{*} \|^{n} \big] \\ &\leq 2^{m-1} \alpha^{m} \mathbb{E} \big[\big(\| g \left(\theta_{t-1}, X_{t} \right) \|^{m} + \| \xi_{t} \left(\theta_{t-1} \right) \|^{m} \big) \| \theta_{t-1} - \theta^{*} \|^{n} \big] \\ &= 2^{m-1} \alpha^{m} \mathbb{E} \big[\mathbb{E} \big[\big(\| g \left(\theta_{t-1}, X_{t} \right) \|^{m} + \| \xi_{t} \left(\theta_{t-1} \right) \|^{m} \big) \| \mathscr{F}_{t-1} \big] \| \theta_{t-1} - \theta^{*} \|^{n} \big] \\ &\leq 2^{m-1} \alpha^{m} \mathbb{E} \big[\big(\| \theta_{t-1} - \theta^{*} \| + 1 \big)^{m} \big] + L_{2}^{m} \mathbb{E} \big[\big(\| \theta_{t-1} - \theta^{*} \| + 1 \big)^{m} \big] \big) \| \theta_{t-1} - \theta^{*} \|^{n} \big] \\ &\leq 2^{m-1} \alpha^{m} L^{m} \mathbb{E} \left[\big(\| \theta_{t-1} - \theta^{*} \| + 1 \big)^{m+n} \right] \\ &\leq 2^{2m+n-2} \alpha^{m} L^{m} \big(\mathbb{E} \big[\| \theta_{t-1} - \theta^{*} \|^{m+n} \big] + 1 \big) \end{split}$$

where the second and the last line follows from the mclaurin's inequality. Also the fourth line follows from ??. Proof for the other case is a trivial consequence of these calculations.

Now to do the proof in this case, we assume that the moment bound in [??] has been proven for $k \le n-1$, we now proceed to show that the desired moment convergence holds for n with $2 \le n \le p$.

We start with the following decomposition of $\|\theta_{k+1} - \theta^*\|^{2n}$

$$\begin{split} \|\boldsymbol{\theta}_{k+1} - \boldsymbol{\theta}^*\|^{2n} &= \left(\|\boldsymbol{\theta}_k - \boldsymbol{\theta}^*\|^2 + 2\alpha \left\langle \boldsymbol{\theta}_k - \boldsymbol{\theta}^*, g\left(\boldsymbol{\theta}_k, X_{k+1}\right) + \boldsymbol{\xi}_{k+1}\left(\boldsymbol{\theta}_k\right)\right\rangle + \alpha^2 \|g\left(\boldsymbol{\theta}_x, X_{k+1}\right) + \boldsymbol{\xi}_{k+1}\left(\boldsymbol{\theta}_k\right)\|^2\right)^n \\ &= \sum_{\substack{i,j,l\\i+j+l=n}} \binom{n}{i,j,l} \|\boldsymbol{\theta}_k - \boldsymbol{\theta}^*\|^{2i} \left(2\alpha \left\langle \boldsymbol{\theta}_k - \boldsymbol{\theta}^*, g\left(\boldsymbol{\theta}_k, X_{k+1}\right) + \boldsymbol{\xi}_{k+1}\left(\boldsymbol{\theta}_k\right)\right\rangle\right)^j \left(\alpha \|g\left(\boldsymbol{\theta}_k, X_{k+1}\right) + \boldsymbol{\xi}_{k+1}\left(\boldsymbol{\theta}_k\right)\|\right)^{2l} \end{split}$$

We note the following cases.

- 1. i = n, j = l = 0. In this case, the summand is simply $\|\theta_k \theta^*\|^{2i}$.
- 2. When i = n 1, j = 1 and l = 0. In this case, the summand is of order α , i.e.,

$$\alpha \cdot 2n \langle \theta_k - \theta^*, g(\theta_k, X_{k+1}) + \xi_{k+1}(\theta_k) \rangle^j \|\theta_k - \theta^*\|^{2(n-1)}$$
.

We can further decompose it as

$$2n\alpha\left\langle \theta_{k}-\theta^{*},g\left(\theta_{k},X_{k+1}\right)+\xi_{k+1}\left(\theta_{k}\right)\right\rangle \left\Vert \theta_{k}-\theta^{*}\right\Vert ^{2(n-1)}\\ =\underbrace{2n\alpha\left\langle \theta_{k}-\theta^{*},g\left(\theta_{k},X_{k+1}\right)-\bar{g}\left(\theta_{k}\right)+\xi_{k+1}\left(\theta_{k}\right)\right\rangle \left\Vert \theta_{k}-\theta^{*}\right\Vert ^{2(n-1)}}_{T_{1}}+\underbrace{2n\alpha\left\langle \theta_{k}-\theta^{*},\bar{g}\left(\theta_{k}\right)\right\rangle \left\Vert \theta_{k}-\theta^{*}\right\Vert ^{2(n-1)}}_{T_{2}}.$$

Note that, when (X_k) is i.i.d or from a martingale noise, we have

$$\mathbb{E}\left[T_1|\theta_k\right] = 0$$

However, when (X_k) is Markovian, the above inequality does not hold and T_1 requires careful analysis.

Nonetheless, under the strong monotonicity assumption, we have

$$T_2 \leq -2n\alpha\mu\|\theta_k - \theta^*\|^{2n}$$
.

3. For the remaining terms, we see that they are of higher orders of α . Therefore, when α is selected sufficiently small, these terms do not raise concern.

Therefore, to prove the desired moment bound, we spend the remaining section analyzing T_1 . Immediately, we note that

$$\begin{split} \mathbb{E}\left[T_{1}\right] &= \mathbb{E}\left[2n\alpha\left\langle\theta_{k} - \theta^{*}, g\left(\theta_{k}, X_{k+1}\right) - \bar{g}\left(\theta_{k}\right) + \mathbb{E}\left[\xi_{k+1}\left(\theta_{k}\right) \left|\theta_{k}\right|\right\rangle \left\|\theta_{k} - \theta^{*}\right\|^{2(n-1)}\right] \\ &= 2n\alpha\mathbb{E}\left[\underbrace{\left\langle\theta_{k} - \theta^{*}, g\left(\theta_{k}, X_{k+1}\right) - \bar{g}\left(\theta_{k}\right)\right\rangle \left\|\theta_{k} - \theta^{*}\right\|^{2(n-1)}}_{T_{1}'}\right]. \end{split}$$

Subsequently, we focus on analyzing T_1' ; but before that, we write the general recursion of the error bound. First, we define $T_{1,t}' := \langle \theta_t - \theta^*, g(\theta_t, X_{t+1}) - \bar{g}(\theta_t) \| \theta_t - \theta^* \|^{2(n-1)} \rangle$ to make T_1' dependent on the iteration index. Now, following the above decomposition and taking the expectations, we have:

$$\mathbb{E}\left[\left\|\boldsymbol{\theta}_{k+1} - \boldsymbol{\theta}^*\right\|^{2n}\right] \leq \mathbb{E}\left[\left\|\boldsymbol{\theta}_{k} - \boldsymbol{\theta}^*\right\|^{2n}\right] + 2n\alpha\mathbb{E}\left[T_{1,k}'\right] - 2n\alpha\mu\mathbb{E}\left[\left\|\boldsymbol{\theta}_{k} - \boldsymbol{\theta}^*\right\|^{2n}\right] + o\left(\alpha\right) = \left(1 - 2n\alpha\mu\right)\mathbb{E}\left[\left\|\boldsymbol{\theta}_{k} - \boldsymbol{\theta}^*\right\|^{2n}\right] + 2n\alpha\mathbb{E}\left[T_{1,k}'\right] + o\left(\alpha\right)$$

similarly to the previous case we define $\gamma_t := 2n\alpha (1 - 2n\alpha\mu)^{k-t}$ for $0 \le t \le k$. Solving the above recursion will give us

$$\mathbb{E}\left[\left\|\theta_{k+1} - \theta^*\right\|^{2n}\right] \le \sum_{t=0}^{k} \gamma_t \mathbb{E}\left[T'_{1,t}\right] + \gamma_0 \mathbb{E}\left[\left\|\theta_0 - \theta^*\right\|^{2n}\right] + o\left(\alpha\right)$$

We have to upper bound the first term in the RHS above. For this purpose, we use a similar decomposition to our base case analysis:

$$\sum_{t=0}^{k} \gamma_{t} \mathbb{E}\left[T_{1,t}'\right] = \mathbb{E}\left[\sum_{t=0}^{k} \gamma_{t} \left\langle \theta_{t} - \theta^{*}, g\left(\theta_{t}, X_{t+1}\right) - \bar{g}\left(\theta_{t}\right) \right\rangle \|\theta_{t} - \theta^{*}\|^{2(n-1)}\right] = A_{1} + A_{2} + A_{3} + A_{4} + A_{5}$$

with

$$\begin{split} A_{1} &:= \mathbb{E}\left[\sum_{t=1}^{k} \gamma_{t} \left\langle \theta_{t} - \theta^{*}, \hat{g}\left(\theta_{t}, X_{t+1}\right) - P_{\theta_{t}} \hat{g}\left(\theta_{t}, X_{t}\right) \right\rangle \|\theta_{t} - \theta^{*}\|^{2(n-1)}\right], \\ A_{2} &:= \mathbb{E}\left[\sum_{t=1}^{k} \gamma_{t} \left\langle \theta_{t} - \theta^{*}, P_{\theta_{t}} \hat{g}\left(\theta_{t}, X_{t}\right) - P_{\theta_{t-1}} \hat{g}\left(\theta_{t-1}, X_{t}\right) \right\rangle \|\theta_{t} - \theta^{*}\|^{2(n-1)}\right], \\ A_{3} &:= \mathbb{E}\left[\sum_{t=1}^{k} \gamma_{t} \left\langle \theta_{t} - \theta_{t-1}, P_{\theta_{t-1}} \hat{g}\left(\theta_{t-1} X_{t}\right) \right\rangle \|\theta_{t} - \theta^{*}\|^{2(n-1)}\right], \\ A_{4} &:= \mathbb{E}\left[\sum_{t=1}^{k} \left(\gamma_{t} - \gamma_{t-1}\right) \left\langle \theta_{t-1} - \theta^{*}, P_{\theta_{t-1}} \hat{g}\left(\theta_{t-1} - \theta^{*}, X_{t}\right) \right\rangle \|\theta_{t} - \theta^{*}\|^{2(n-1)}\right], \\ A_{5} &:= \mathbb{E}\left[\sum_{t=1}^{k} \gamma_{t-1} \left\langle \theta_{t-1} - \theta^{*}, P_{\theta_{t-1}} \hat{g}\left(\theta_{t-1} - \theta^{*}, X_{t}\right) \right\rangle \left(\|\theta_{t} - \theta^{*}\|^{2(n-1)} - \|\theta_{t-1} - \theta^{*}\|^{2(n-1)}\right)\right], \\ A_{6} &:= \mathbb{E}\left[\gamma_{0} \left\langle \theta_{0} - \theta^{*}, \hat{g}\left(\theta_{0}, X_{0}\right) \right\rangle \|\theta_{0} - \theta^{*}\|^{2(n-1)}\right] + \mathbb{E}\left[\gamma_{k} \left\langle \theta_{k} - \theta^{*}, P_{\theta_{k}} \hat{g}\left(\theta_{k}, X_{k+1}\right) \right\rangle \|\theta_{k} - \theta^{*}\|^{2(n-1)}\right]. \end{split}$$

For A_1 , we note that $\hat{g}(\theta_t, X_{t+1}) - P_{\theta_t} \hat{g}(\theta_t, X_t)$ is a martingale difference sequence [cf. ?] and therefore we have $A_1 = 0$ by taking the total expectation.

For A_2 , applying Cauchy-Schwarz inequality and ??, we have

$$\begin{split} A_2 &\leq \sum_{t=1}^k L_{PH}^{(1)} \gamma_t \mathbb{E} \left[\| \theta_t - \theta_{t-1} \| \| \theta_t - \theta^* \|^{2n-1} \right] \\ &\leq \sum_{t=1}^k L_{PH}^{(1)} \gamma_t \mathbb{E} \left[\| \theta_t - \theta_{t-1} \| \left(\| \theta_t - \theta_{t-1} \| + \| \theta_{t-1} - \theta^* \| \right)^{2n-1} \right] \\ &\leq \sum_{t=1}^k 2^{2n-2} L_{PH}^{(1)} \gamma_t \mathbb{E} \left[\| \theta_t - \theta_{t-1} \|^{2n} + \| \theta_t - \theta_{t-1} \| \| \theta_{t-1} - \theta^* \|^{2n-1} \right] \\ &\leq \sum_{t=1}^k 2^{2n-2} L_{PH}^{(1)} \gamma_t \left(2^{4n} \alpha^{2n} L^{2n} \left(\mathbb{E} \left[\| \theta_{t-1} - \theta^* \|^{2n} \right] + 1 \right) + 2^{2n-1} \alpha L \left(\mathbb{E} \left[\| \theta_{t-1} - \theta^* \|^{2n} \right] + 1 \right) \right) \\ &\leq \sum_{t=1}^k 2^{4n-2} \alpha L L_{PH}^{(1)} \gamma_t \left(\mathbb{E} \left[\| \theta_{t-1} - \theta^* \|^{2n} \right] + 1 \right) \end{split}$$

where the second line follows from triangle inequality, third line from mclaurin's inequality, fourth line from Lemma 3, and the last line for $\alpha < \frac{1}{4L}$.

For A_3 , we obtain

$$\begin{split} A_{3} &\leq \sum_{t=1}^{k} \gamma_{t} \mathbb{E} \left[\| \theta_{t} - \theta_{t-1} \| \| P_{\theta_{t-1}} \hat{g} \left(\theta_{t-1}, X_{t} \right) \| \| \theta_{t} - \theta^{*} \|^{2(n-1)} \right] \\ &\leq \sum_{t=1}^{k} L_{PH}^{(0)} \gamma_{t} \mathbb{E} \left[\| \theta_{t} - \theta_{t-1} \| \left(\| \theta_{t} - \theta_{t-1} \| + \| \theta_{t-1} - \theta^{*} \| \right)^{2(n-1)} \right] \\ &\leq \sum_{t=1}^{k} 2^{2n-3} L_{PH}^{(0)} \gamma_{t} \mathbb{E} \left[\| \theta_{t} - \theta_{t-1} \|^{2n-1} + \| \theta_{t} - \theta_{t-1} \| \| \theta_{t-1} - \theta^{*} \|^{2(n-1)} \right] \\ &\leq \sum_{t=1}^{k} 2^{2n-3} L_{PH}^{(0)} \gamma_{t} \left(2^{4n-1} \alpha^{2n-1} L^{2n-1} \left(\mathbb{E} \left[\| \theta_{t-1} - \theta^{*} \|^{2n-1} \right] + 1 \right) + 2^{2n-2} \alpha L \left(\mathbb{E} \left[\| \theta_{t-1} - \theta^{*} \|^{2n-1} \right] + 1 \right) \right) \\ &\leq \sum_{t=1}^{k} 2^{4n-4} \alpha L L_{PH}^{(0)} \gamma_{t} \left(\mathbb{E} \left[\| \theta_{t-1} - \theta^{*} \|^{2n-1} \right] + 1 \right) \end{split}$$

where second line follows from triangle inequality, third line from mclaurin's inequality, fourth line from Lemma 3, and the last line for $\alpha < \frac{1}{4L}$.

For A_4 , we have

$$\begin{split} A_{4} &\leq \sum_{t=1}^{k} |\gamma_{t} - \gamma_{t-1}| \, \mathbb{E} \left[\|\theta_{t-1} - \theta^{*}\| \|P_{\theta_{t-1}} \hat{g} \left(\theta_{t-1}, X_{t}\right) \| \, \|\theta_{t} - \theta^{*}\|^{2(n-1)} \right] \\ &\leq \sum_{t=1}^{k} L_{PH}^{(0)} |\gamma_{t} - \gamma_{t-1}| \, \mathbb{E} \left[\|\theta_{t-1} - \theta^{*}\| \, \|\theta_{t} - \theta^{*}\|^{2(n-1)} \right] \\ &\leq \sum_{t=1}^{k} L_{PH}^{(0)} |\gamma_{t} - \gamma_{t-1}| \mathbb{E} \left[\|\theta_{t-1} - \theta^{*}\| \, \left(\|\theta_{t} - \theta_{t-1}\| + \|\theta_{t-1} - \theta^{*}\| \right)^{2(n-1)} \right] \\ &\leq \sum_{t=1}^{k} 2^{2n-3} L_{PH}^{(0)} |\gamma_{t} - \gamma_{t-1}| \mathbb{E} \left[\|\theta_{t} - \theta_{t-1}\|^{2(n-1)} \, \|\theta_{t-1} - \theta^{*}\| + \|\theta_{t-1} - \theta^{*}\|^{2n-1} \right] \\ &\leq \sum_{t=1}^{k} 2^{2n-3} L_{PH}^{(0)} |\gamma_{t} - \gamma_{t-1}| \left(2^{4n-5} \alpha^{2n-2} L^{2n-2} \left(\mathbb{E} \left[\|\theta_{t-1} - \theta^{*}\|^{2n-1} \right] + 1 \right) + 2^{2n-2} \left(\mathbb{E} \left[\|\theta_{t-1} - \theta^{*}\|^{2n-1} \right] + 1 \right) \right) \\ &\leq \sum_{t=1}^{k} 2^{4n-4} L_{PH}^{(0)} |\gamma_{t} - \gamma_{t-1}| \left(\mathbb{E} \left[\|\theta_{t-1} - \theta^{*}\|^{2n-1} \right] + 1 \right) \end{split}$$

where the third line follows from triangle inequality, fourth line from mclaurin's inequality, fifth line from Lemma 3, and the last line for $\alpha < \frac{1}{4L}$.

Now for A_5 , we have to first note that, using mean-value theorem and with $a \in [0,1]$, we'll get

$$\begin{split} \|\theta_t - \theta^*\|^{2(n-1)} - \|\theta_{t-1} - \theta^*\|^{2(n-1)} &= \|\theta_t - \theta_{t-1}\| \cdot 2(n-1) \|a(\theta_t - \theta^*) + (1-a)(\theta_{t-1} - \theta^*)\|^{2n-3} \\ &\leq \|\theta_t - \theta_{t-1}\| \cdot 2(n-1) \|a(\theta_t - \theta_{t-1}) + \theta_{t-1} - \theta^*\|^{2n-3} \\ &\leq 2^{2n-3}(n-1) \|\theta_t - \theta_{t-1}\| \left(\|\theta_t - \theta_{t-1}\|^{2n-3} + \|\theta_{t-1} - \theta^*\|^{2n-3} \right) \end{split}$$

where the last line follows using the mclaurin's inequality. Plugging in the above upper bound to A_5 gives us

$$\begin{split} A_5 &\leq \sum_{t=1}^k 2^{2n-3} (n-1) \gamma_{t-1} \mathbb{E} \left[\left\langle \theta_{t-1} - \theta^*, P_{\theta_{t-1}} \hat{g} \left(\theta_{t-1} - \theta^*, X_t \right) \right\rangle \|\theta_t - \theta_{t-1}\| \left(\|\theta_t - \theta_{t-1}\|^{2n-3} + \|\theta_{t-1} - \theta^*\|^{2n-3} \right) \right] \\ &\leq \sum_{t=1}^k 2^{2n-3} (n-1) \gamma_{t-1} \mathbb{E} \left[\|\theta_{t-1} - \theta^*\| \|P_{\theta_{t-1}} \hat{g} \left(\theta_{t-1}, X_t \right) \| \|\theta_t - \theta_{t-1}\| \left(\|\theta_t - \theta_{t-1}\|^{2n-3} + \|\theta_{t-1} - \theta^*\|^{2n-3} \right) \right] \\ &\leq \sum_{t=1}^k 2^{2n-3} (n-1) L_{PH}^{(0)} \gamma_{t-1} \mathbb{E} \left[\|\theta_{t-1} - \theta^*\| \|\theta_t - \theta_{t-1}\| \left(\|\theta_t - \theta_{t-1}\|^{2n-3} + \|\theta_{t-1} - \theta^*\|^{2n-3} \right) \right] \\ &\leq \sum_{t=1}^k 2^{2n-3} (n-1) L_{PH}^{(0)} \gamma_{t-1} \mathbb{E} \left[\|\theta_t - \theta_{t-1}\|^{2n-2} \|\theta_{t-1} - \theta^*\| + \|\theta_t - \theta_{t-1}\| \|\theta_{t-1} - \theta^*\|^{2n-2} \right] \\ &\leq \sum_{t=1}^k 2^{2n-3} (n-1) L_{PH}^{(0)} \gamma_{t-1} \left(2^{4n-5} \alpha^{2n-2} L^{2n-2} \left(\mathbb{E} \left[\|\theta_{t-1} - \theta^*\|^{2n-1} \right] + 1 \right) + 2^{2n-2} \alpha L \left(\mathbb{E} \left[\|\theta_{t-1} - \theta^*\|^{2n-1} \right] + 1 \right) \right) \\ &\leq \sum_{t=1}^k 2^{4n-4} (n-1) \alpha L L_{PH}^{(0)} \gamma_{t-1} \left(\mathbb{E} \left[\|\theta_{t-1} - \theta^*\|^{2n-1} \right] + 1 \right) \end{split}$$

where the fifth line follows from Lemma 3 and the last line follows for $\alpha < \frac{1}{4L}$. Finally, for A_6 , we obtain

$$A_6 \leq L_{PH}^{(0)} \left(\gamma_0 \| \theta_0 - \theta^* \|^{2n-1} + \gamma_k \| \theta_k - \theta^* \|^{2n-1} \right)$$

which follows from Cauchy-Schwarz inequality and ??.

Combining the above terms gives us:

$$\begin{split} \sum_{t=0}^{k} \gamma_{t} \mathbb{E}\left[T_{1,t}'\right] &= \mathbb{E}\left[\sum_{t=0}^{k} \gamma_{t} \left\langle \theta_{t} - \theta^{*}, g\left(\theta_{t}, X_{t+1}\right) - \bar{g}\left(\theta_{t}\right) \right\rangle \|\theta_{t} - \theta^{*}\|^{2(n-1)}\right] \\ &\leq \sum_{t=0}^{k-1} 2^{4n-2} \alpha L L_{PH}^{(1)} \gamma_{t+1} \left(\mathbb{E}\left[\|\theta_{t} - \theta^{*}\|^{2n}\right] + 1\right) + \sum_{t=0}^{k-1} 2^{4n-4} \alpha L L_{PH}^{(0)} \gamma_{t+1} \left(\mathbb{E}\left[\|\theta_{t} - \theta^{*}\|^{2n-1}\right] + 1\right) \\ &+ \sum_{t=0}^{k-1} 2^{4n-4} L_{PH}^{(0)} |\gamma_{t+1} - \gamma_{t}| \left(\mathbb{E}\left[\|\theta_{t} - \theta^{*}\|^{2n-1}\right] + 1\right) + \sum_{t=0}^{k-1} 2^{4n-4} \left(n-1\right) \alpha L L_{PH}^{(0)} \gamma_{t+1} \left(\mathbb{E}\left[\|\theta_{t} - \theta^{*}\|^{2n-1}\right] + 1\right) \\ &+ L_{PH}^{(0)} \left(\gamma_{0} \|\theta_{0} - \theta^{*}\|^{2n-1} + \gamma_{k} \|\theta_{k} - \theta^{*}\|^{2n-1}\right) \\ &= \sum_{t=0}^{k-1} 2^{4n-2} \alpha L L_{PH}^{(1)} \gamma_{t+1} \left(\mathbb{E}\left[\|\theta_{t} - \theta^{*}\|^{2n}\right] + 1\right) + \sum_{t=0}^{k-1} 2^{4n-4} L_{PH}^{(0)} \left(\left(n\alpha L + 1\right) \gamma_{t+1} - \gamma_{t}\right) \left(\mathbb{E}\left[\|\theta_{t} - \theta^{*}\|^{2n-1}\right] + 1\right) \\ &+ L_{PH}^{(0)} \left(\gamma_{0} \|\theta_{0} - \theta^{*}\|^{2n-1} + \gamma_{k} \|\theta_{k} - \theta^{*}\|^{2n-1}\right) \end{split}$$

where in the last equality we used the fact that for $\alpha \leq \frac{\mu}{L^2}$, $\gamma_t \leq \gamma_{t+1}$.

Now, if we make use of the inequality $2|x|^3 \le x^2 + x^4$ and consolidating the terms, for sufficiently small

 α , we have

$$\begin{split} \mathbb{E}\left[\|\theta_{k+1} - \theta^*\|^{2n}\right] &\leq \sum_{t=0}^{k-1} 2^{4n-5} \left[8\alpha L L_{PH}^{(1)} \gamma_{t+1} + L_{PH}^{(0)} \left((n\alpha L + 1) \, \gamma_{t+1} - \gamma_{t}\right)\right] \left(\mathbb{E}\left[\|\theta_{t} - \theta^*\|^{2n}\right] + 1\right) \\ &+ \sum_{t=0}^{k-1} 2^{4n-5} L_{PH}^{(0)} \left((n\alpha L + 1) \, \gamma_{t+1} - \gamma_{t}\right) \left(\mathbb{E}\left[\|\theta_{t} - \theta^*\|^{2(n-1)}\right] + 1\right) + \frac{L_{PH}^{(0)}}{2} \left(\gamma_{0} \|\theta_{0} - \theta^*\|^{2(n-1)} + \gamma_{k} \|\theta_{k} - \theta^*\|^{2(n-1)}\right) \\ &+ \left(\frac{L_{PH}^{(0)}}{2} + 1\right) \gamma_{0} \mathbb{E}\left[\|\theta_{0} - \theta^*\|^{2n}\right] + \frac{L_{PH}^{(0)}}{2} \gamma_{k} \|\theta_{k} - \theta^*\|^{2n} + o\left(\alpha\right) \\ &\leq \sum_{t=0}^{k-1} 2^{4n-4} \alpha L \gamma_{t+1} \left[4L_{PH}^{(1)} + nL_{PH}^{(0)}\right] \mathbb{E}\left[\|\theta_{t} - \theta^*\|^{2n}\right] + \sum_{t=0}^{k-1} 2^{4n-4} n\alpha L L_{PH}^{(0)} \mathbb{E}\left[\|\theta_{t} - \theta^*\|^{2(n-1)}\right] \\ &+ \frac{L_{PH}^{(0)}}{2} \left(\gamma_{0} \|\theta_{0} - \theta^*\|^{2(n-1)} + \gamma_{k} \|\theta_{k} - \theta^*\|^{2(n-1)}\right) + \left(\frac{L_{PH}^{(0)}}{2} + 1\right) \gamma_{0} \mathbb{E}\left[\|\theta_{0} - \theta^*\|^{2n}\right] + \frac{L_{PH}^{(0)}}{2} \gamma_{k} \mathbb{E}\left[\|\theta_{k} - \theta^*\|^{2n}\right] + \tilde{c}_{1,k} + o\left(\frac{L_{PH}^{(0)}}{2} + 1\right) \gamma_{0} \mathbb{E}\left[\|\theta_{0} - \theta^*\|^{2n}\right] + \frac{L_{PH}^{(0)}}{2} \gamma_{k} \mathbb{E}\left[\|\theta_{k} - \theta^*\|^{2n}\right] + \tilde{c}_{1,k} + o\left(\frac{L_{PH}^{(0)}}{2} + 1\right) \gamma_{0} \mathbb{E}\left[\|\theta_{0} - \theta^*\|^{2n}\right] + \frac{L_{PH}^{(0)}}{2} \gamma_{k} \mathbb{E}\left[\|\theta_{k} - \theta^*\|^{2n}\right] + \tilde{c}_{1,k} + o\left(\frac{L_{PH}^{(0)}}{2} + 1\right) \gamma_{0} \mathbb{E}\left[\|\theta_{0} - \theta^*\|^{2n}\right] + \frac{L_{PH}^{(0)}}{2} \gamma_{k} \mathbb{E}\left[\|\theta_{k} - \theta^*\|^{2n}\right] + \tilde{c}_{1,k} + o\left(\frac{L_{PH}^{(0)}}{2} + 1\right) \gamma_{0} \mathbb{E}\left[\|\theta_{0} - \theta^*\|^{2n}\right] + \frac{L_{PH}^{(0)}}{2} \gamma_{k} \mathbb{E}\left[\|\theta_{k} - \theta^*\|^{2n}\right] + \tilde{c}_{1,k} + o\left(\frac{L_{PH}^{(0)}}{2} + 1\right) \gamma_{0} \mathbb{E}\left[\|\theta_{0} - \theta^*\|^{2n}\right] + \frac{L_{PH}^{(0)}}{2} \gamma_{k} \mathbb{E}\left[\|\theta_{k} - \theta^*\|^{2n}\right] + \tilde{c}_{1,k} + o\left(\frac{L_{PH}^{(0)}}{2} + 1\right) \gamma_{0} \mathbb{E}\left[\|\theta_{0} - \theta^*\|^{2n}\right] + \frac{L_{PH}^{(0)}}{2} \gamma_{k} \mathbb{E}\left[\|\theta_{k} - \theta^*\|^{2n}\right] + \tilde{c}_{1,k} + o\left(\frac{L_{PH}^{(0)}}{2} + 1\right) \gamma_{0} \mathbb{E}\left[\|\theta_{0} - \theta^*\|^{2n}\right] + \tilde{c}_{1,k} + o\left(\frac{L_{PH}^{(0)}}{2} + 1\right) \gamma_{0} \mathbb{E}\left[\|\theta_{0} - \theta^*\|^{2n}\right] + \tilde{c}_{1,k} + o\left(\frac{L_{PH}^{(0)}}{2} + 1\right) \gamma_{0} \mathbb{E}\left[\|\theta_{0} - \theta^*\|^{2n}\right] + \tilde{c}_{1,k} + o\left(\frac{L_{PH}^{(0)}}{2} + 1\right) \gamma_{0} \mathbb{E}\left[\|\theta_{0} - \theta^*\|^{2n}\right] + \tilde{c}_{1,k} + o\left(\frac{L_{PH}^{(0)}}{2} + 1\right) \gamma_{0} \mathbb{E}\left[\|\theta_{0} - \theta^*\|^{2n}\right] + \tilde{c}_{1,k} + o\left$$

where the last inequality follows for $(n\alpha L+1)\gamma_{t+1}-\gamma_t \leq 2n\alpha L\gamma_{t+1}$ and also with

$$\tilde{c}_{1,t} := \sum_{l=0}^{t-1} 2^{4n-4} n\alpha L \gamma_{t+1} \left[4L_{PH}^{(1)} + nL_{PH}^{(0)} \right] + \sum_{l=0}^{t-1} 2^{4n-4} n\alpha L L_{PH}^{(0)}$$

being defined for $1 \le t \le k-1$. It can be seen that for sufficiently small α , $\tilde{c}_{1,t}$ would be of order $4^n n\alpha L\left(nL_{PH}^{(0)} + 4\alpha L_{PH}^{(1)}\right)$. Now from induction hypothesis we know that

$$\mathbb{E}\left[\left\|\theta_{t} - \theta^{*}\right\|^{2(n-1)}\right] \leq \tilde{C}_{n-1,1} \cdot \left(1 - 2n\alpha\mu\right)^{t} \mathbb{E}\left[\left\|\theta_{0} - \theta^{*}\right\|^{2n}\right] + \tilde{C}_{n-1,2} \cdot 4^{n-1}(n-1)\alpha L\left((n-1)L_{PH}^{(0)} + 2\alpha L_{PH}^{(1)}\right)\right]$$

Thus, if we plug in this upper bound into the previous inequality, for $\alpha < \frac{1}{4^n}$ we will have

$$\begin{split} \mathbb{E} \left[\| \boldsymbol{\theta}_{k+1} - \boldsymbol{\theta}^* \|^{2n} \right] &\leq \sum_{t=0}^{k-1} 2^{4n-4} \alpha L \gamma_{t+1} \left[4L_{PH}^{(1)} + nL_{PH}^{(0)} \right] \mathbb{E} \left[\| \boldsymbol{\theta}_t - \boldsymbol{\theta}^* \|^{2n} \right] \\ &+ \frac{L_{PH}^{(0)}}{2} \gamma_k \mathbb{E} \left[\| \boldsymbol{\theta}_k - \boldsymbol{\theta}^* \|^{2n} \right] + \tilde{c}_{2,k} \mathbb{E} \left[\| \boldsymbol{\theta}_0 - \boldsymbol{\theta}^* \|^{2n} \right] + \tilde{c}_{1,k} + \tilde{c}_{3,k} + o\left(\alpha\right) \end{split}$$

in which $\tilde{c}_{3,k}$ would be of order $4^n n\alpha L\left(nL_{PH}^{(0)}+4\alpha L_{PH}^{(1)}\right)$ and $\tilde{c}_{2,k}$ is an $\mathcal{O}(1)$ constant dependent to k.

To solve the above recursion, we define $\tilde{S}_t := \frac{L_{PH}^{(0)}}{2} \gamma_t \mathbb{E}\left[\|\theta_t - \theta^*\|^{2n}\right] + \sum_{l=0}^{t-1} 2^{4n-4} \alpha L \gamma_{l+1} \left[4L_{PH}^{(1)} + nL_{PH}^{(0)}\right] \mathbb{E}\left[\|\theta_l - \theta^*\|^{2n}\right]$ for $1 \le t \le k$. Using this we can write

$$\mathbb{E} \left[\|\theta_{t+1} - \theta^*\|^{2n} \right] \leq \tilde{S}_t + \tilde{c}_{2,t} \mathbb{E} \left[\|\theta_0 - \theta^*\|^{2n} \right] + \tilde{c}_{1,t} + \tilde{c}_{3,t} + o(\alpha).$$

Now notice that we have (recall that $\gamma_t = 2n\alpha (1 - 2n\alpha\mu)^{k-t}$.)

$$\begin{split} \frac{\tilde{S}_{t}}{\tilde{S}_{t-1}} &\leq \frac{\tilde{S}_{t-1} + \left(2^{4n-4}\alpha L\gamma_{t} \left[4L_{PH}^{(1)} + nL_{PH}^{(0)}\right] - \frac{L_{PH}^{(0)}}{2}\gamma_{t-1}\right) \mathbb{E}\left[\|\theta_{t-1} - \theta^{*}\|^{2n}\right] + \frac{L_{PH}^{(0)}}{2}\gamma_{t}\mathbb{E}\left[\|\theta_{t} - \theta^{*}\|^{2n}\right]}{\tilde{S}_{t-1}} \\ &= \frac{\tilde{S}_{t-1} + \left(2^{4n-4}\alpha L\gamma_{t} \left[4L_{PH}^{(1)} + nL_{PH}^{(0)}\right] - \frac{L_{PH}^{(0)}}{2}\gamma_{t-1}\right) \mathbb{E}\left[\|\theta_{t-1} - \theta^{*}\|^{2n}\right] + \frac{L_{PH}^{(0)}}{2}\gamma_{t} \left[\tilde{S}_{t-1} + \tilde{c}_{2,t-1}\mathbb{E}\left[\|\theta_{0} - \theta^{*}\|^{2n}\right] + \tilde{c}_{1,t} + \tilde{c}_{3,t} + o\left(\alpha\right)\right]}{\tilde{S}_{t-1}} \end{split}$$

it is not very hard to see that the if we take α small enough above fraction could get as close as possible

to 1. This means that we can write our final error bound as the following

$$\mathbb{E}\left[\left\|\boldsymbol{\theta}_{k+1} - \boldsymbol{\theta}^*\right\|^{2n}\right] \leq \tilde{C}_{n,1} \cdot \left(1 - 2n\alpha\mu\right)^{k+1} \mathbb{E}\left[\left\|\boldsymbol{\theta}_{0} - \boldsymbol{\theta}^*\right\|^{2n}\right] + \tilde{C}_{n,2} \cdot 4^{n} n\alpha L\left(nL_{PH}^{(0)} + 4\alpha L_{PH}^{(1)}\right)$$

in which $\tilde{C}_{n,1}$ and $\tilde{C}_{n,2}$ are $\mathcal{O}(1)$ constants and also α has been taken small enough.

3 Bias Characterization

Before proceeding to the proof, we introduce the following shorthands and notations:

First, we recall that as a consequence of Assumption 2, the target equation $\bar{g}(\theta) = \mathbb{E}_{X \sim \pi_{\theta}} \left[g(\theta, X) \right] = 0$, has a unique solution which we denote by θ^* . Moreover, we define X^* to be the samples being drawn from the stationary distribution of the Markov chain related to θ^* ; *i.e.*, π_{θ^*} . Note that, from stationarity we know this is equal to the case that $\{X_n^*, n \in \mathbb{N}\}$ is a Markov chain initialized same as the other $\{X_n\}$ chain and evolving w.r.t. the P_{θ^*} . In this case we also take $\mathscr{F}_n = \{\theta_0, \{\xi_{m+1}\}_{m \leq n-1}, \{X_m\}_{m \leq n}, \{X_m^*\}_{m \leq n}\}$. For notational simplicity, we omit indexes related to the *optimal Chain(i.e., \{X_n^*\})*.

Additionally, For $X \in \mathcal{X}$,

$$z_i(X) := \mathbb{E}\left[\left(\theta_{\infty} - \theta^*\right)^{\otimes i} | X^* = X\right],$$

$$\delta_i(X) := z_i(X) - \pi_{\theta^*} z_i.$$

We also denote

$$\begin{split} & \bar{g}\left(\theta\right) \coloneqq \mathbb{E}_{X \sim \pi_{\theta^*}} \left[g\left(\theta, X\right) \right], \quad \bar{g}_2\left(\theta\right) \coloneqq \mathbb{E}_{X \sim \pi_{\theta^*}} \left[\left(g\left(\theta, X\right) \right)^{\otimes 2} \right] \\ & \bar{g}^{(1)}\left(\theta\right) \coloneqq \mathbb{E}_{X \sim \pi_{\theta^*}} \left[g'\left(\theta, X\right) \right], \quad \bar{g}_2^{(1)} \coloneqq \mathbb{E}_{X \sim \pi_{\theta^*}} \left[\left(g'\left(\theta, X\right) \right)^{\otimes 2} \right], \quad \bar{g}^{(2)}\left(\theta\right) \coloneqq \mathbb{E}_{X \sim \pi_{\theta^*}} \left[g''\left(\theta, X\right) \right]. \end{split}$$

$$\begin{split} \mathbb{E}\left[g\left(\theta_{\infty},X_{\infty+1}\right)\right] &= \mathbb{E}\left[g\left(\theta^{*},X_{\infty+1}\right) - g\left(\theta^{*},X^{*}\right)\right] + \mathbb{E}\left[g\left(\theta^{*},X^{*}\right)\right] \\ &+ \mathbb{E}\left[\left(g'\left(\theta^{*},X_{\infty+1}\right) - g'\left(\theta^{*},X^{*}\right)\right)\left(\theta_{\infty} - \theta^{*}\right)\right] + \mathbb{E}\left[g'\left(\theta^{*},X^{*}\right)\left(\theta_{\infty} - \theta^{*}\right)\right] \\ &+ \frac{1}{2}\mathbb{E}\left[\left(g''\left(\theta^{*},X_{\infty+1}\right) - g''\left(\theta^{*},X^{*}\right)\right)\left(\theta_{\infty} - \theta^{*}\right)^{\otimes 2}\right] + \frac{1}{2}\mathbb{E}\left[g''\left(\theta^{*},X^{*}\right)\left(\theta_{\infty} - \theta^{*}\right)^{\otimes 2}\right] \\ &+ \mathbb{E}\left[R_{3}\left(\theta_{\infty},X_{\infty+1}\right)\right] \end{split}$$

Firstly, note that from the definition we have $\mathbb{E}[g(\theta^*, X^*)] = 0$. Also,

$$\begin{split} \mathbb{E}\left[g\left(\theta^*, X_{\infty+1}\right) - g\left(\theta^*, X^*\right)\right] &= \mathbb{E}\left[\mathbb{E}\left[g\left(\theta^*, X_{\infty+1}\right) | \theta_{\infty}, X_{\infty}\right] - g\left(\theta^*, X^*\right)\right] \\ &= \mathbb{E}\left[P_{\theta_{\infty}}g\left(\theta^*, X_{\infty}\right) - g\left(\theta^*, X^*\right)\right] \\ &= \mathbb{E}\left[P_{\theta_{\infty}}g\left(\theta^*, X_{\infty}\right) - P_{\theta^*}g\left(\theta^*, X_{\infty}\right) + P_{\theta^*}g\left(\theta^*, X_{\infty}\right) - P_{\theta^*}g\left(\theta^*, X^*\right)\right] \\ &= \mathbb{E}\left[\left(P_{\theta_{\infty}} - P_{\theta^*}\right)g\left(\theta^*, X_{\infty}\right) + P_{\theta^*}\left(g\left(\theta^*, X_{\infty}\right) - g\left(\theta^*, X^*\right)\right)\right] \end{split}$$

where the third equality follows from the fact that X^* comes from the stationary distribution of θ^* . Now notice that we have from Assumption 1 that $\mathbb{E}[g(\theta^*, X_{\infty})] \leq L_1$. Additionally,

$$\mathbb{E}\left[\|P_{\theta_{\infty}} - P_{\theta^*}\|\right] \leq L_P \mathbb{E}\left[\|\theta_{\infty} - \theta^*\|\right]$$

from Assumption 7. For the remaining term we also have

$$\mathbb{E}\left[P_{\theta^*}\left(g\left(\theta^*, X_{\infty}\right) - g\left(\theta^*, X^*\right)\right)\right] = \mathbb{E}\left[P_{\theta^*}\int_{\mathbb{R}} g\left(\theta^*, x\right) \pi_{\theta_{\infty}}(x) - g\left(\theta^*, x\right) \pi_{\theta^*}(x) dx\right]$$

$$\leq \mathbb{E}\left[P_{\theta^*}L_S \|\theta_{\infty} - \theta^*\| \int_{\mathbb{R}} g\left(\theta^*, x\right) dx\right]$$

$$\leq L_1 L_S \mathbb{E}\left[\|\theta_{\infty} - \theta^*\|\right]$$

where the second line follows from Assumption 8. Summing up all these would give us

$$\mathbb{E}\left[g\left(\theta^*, X_{\infty+1}\right) - g\left(\theta^*, X^*\right)\right] \le L_1\left(L_P + L_S\right) \mathbb{E}\left[\|\theta_\infty - \theta^*\|\right],$$

which is of $\mathcal{O}(\alpha^{1/2})$.

For the first and second derivative cases, conditioning on θ_{∞} would give us a similar setting and thus we will have

$$\begin{split} \mathbb{E}\left[\left(g'\left(\theta^{*},X_{\infty+1}\right)-g'\left(\theta^{*},X^{*}\right)\right)\left(\theta_{\infty}-\theta^{*}\right)\right] &= \mathbb{E}\left[\mathbb{E}\left[\left(\mathbb{E}\left[g'\left(\theta^{*},X_{\infty+1}\right)|X_{\infty},\theta_{\infty}\right]-g'\left(\theta^{*},X^{*}\right)\right)\left(\theta_{\infty}-\theta^{*}\right)|\theta_{\infty}\right]\right] \\ &= \mathbb{E}\left[\mathbb{E}\left[\left(P_{\theta_{\infty}}g'\left(\theta^{*},X_{\infty}\right)-g'\left(\theta^{*},X^{*}\right)\right)\left(\theta_{\infty}-\theta^{*}\right)|\theta_{\infty}\right]\right] \\ &= \mathbb{E}\left[\mathbb{E}\left[\left(P_{\theta_{\infty}}g'\left(\theta^{*},X_{\infty}\right)-P_{\theta^{*}}g'\left(\theta^{*},X_{\infty}\right)+P_{\theta^{*}}g'\left(\theta^{*},X_{\infty}\right)-P_{\theta^{*}}g'\left(\theta^{*},X^{*}\right)\right)\left(\theta_{\infty}-\theta^{*}\right)|\theta_{\infty}\right] \\ &= \mathbb{E}\left[\mathbb{E}\left[\left(\left(P_{\theta_{\infty}}-P_{\theta^{*}}\right)g'\left(\theta^{*},X_{\infty}\right)+P_{\theta^{*}}\left(g'\left(\theta^{*},X_{\infty}\right)-g'\left(\theta^{*},X^{*}\right)\right)\right)\left(\theta_{\infty}-\theta^{*}\right)|\theta_{\infty}\right]\right] \\ &\leq \mathbb{E}\left[\mathbb{E}\left[L_{1}\left(L_{P}+L_{S}\right)\|\theta_{\infty}-\theta^{*}\|\left(\theta_{\infty}-\theta^{*}\right)|\theta_{\infty}\right]\right] \\ &\leq L_{1}\left(L_{P}+L_{S}\right)\mathbb{E}\left[\|\theta_{\infty}-\theta^{*}\|\left(\theta_{\infty}-\theta^{*}\right)\right] \end{split}$$

which is of $\mathcal{O}(\alpha)$ in the end. In the same manner for the second derivative we can calculate this

$$\begin{split} \mathbb{E}\left[\left(g''\left(\theta^{*},X_{\infty+1}\right)-g''\left(\theta^{*},X^{*}\right)\right)\left(\theta_{\infty}-\theta^{*}\right)^{\otimes 2}\right] &= \mathbb{E}\left[\mathbb{E}\left[\left(\mathbb{E}\left[g''\left(\theta^{*},X_{\infty+1}\right)|X_{\infty},\theta_{\infty}\right]-g''\left(\theta^{*},X^{*}\right)\right)\left(\theta_{\infty}-\theta^{*}\right)^{\otimes 2}|\theta_{\infty}\right]\right] \\ &= \mathbb{E}\left[\mathbb{E}\left[\left(P_{\theta_{\infty}}g''\left(\theta^{*},X_{\infty}\right)-g''\left(\theta^{*},X^{*}\right)\right)\left(\theta_{\infty}-\theta^{*}\right)^{\otimes 2}|\theta_{\infty}\right]\right] \\ &= \mathbb{E}\left[\mathbb{E}\left[\left(P_{\theta_{\infty}}g''\left(\theta^{*},X_{\infty}\right)-P_{\theta^{*}}g''\left(\theta^{*},X_{\infty}\right)+P_{\theta^{*}}g''\left(\theta^{*},X_{\infty}\right)-P_{\theta^{*}}g''\left(\theta^{*},X^{*}\right)\right)\left(\theta_{\infty}-\theta^{*}\right)^{\otimes 2}|\theta_{\infty}\right]\right] \\ &= \mathbb{E}\left[\mathbb{E}\left[\left(\left(P_{\theta_{\infty}}-P_{\theta^{*}}\right)g''\left(\theta^{*},X_{\infty}\right)+P_{\theta^{*}}\left(g''\left(\theta^{*},X_{\infty}\right)-g''\left(\theta^{*},X^{*}\right)\right)\right)\left(\theta_{\infty}-\theta^{*}\right)^{\otimes 2}|\theta_{\infty}\right]\right] \\ &\leq \mathbb{E}\left[\mathbb{E}\left[L_{1}\left(L_{P}+L_{S}\right)\|\theta_{\infty}-\theta^{*}\|\left(\theta_{\infty}-\theta^{*}\right)^{\otimes 2}|\theta_{\infty}\right]\right] \\ &\leq L_{1}\left(L_{P}+L_{S}\right)\mathbb{E}\left[\|\theta_{\infty}-\theta^{*}\|\left(\theta_{\infty}-\theta^{*}\right)^{\otimes 2}\right] \end{split}$$

in this case the difference will be of $\mathcal{O}(\alpha^{3/2})$.

4 Experiments

Problem Setting

We consider a stochastic approximation (SA) scheme where the control parameter $\theta_t \in \mathbb{R}^d$ is updated via the rule:

$$\theta_{t+1} = \theta_t + \alpha \left(g(\theta_t, X_{t+1}) + \xi_{t+1} \right),\,$$

where:

• $\alpha > 0$ is a constant step size,

- $g: \mathbb{R}^d \times \mathbb{R} \to \mathbb{R}^d$ is a Lipschitz and strongly monotone function,
- ξ_t is a zero-mean noise sequence with bounded moments,
- $\{X_t\}$ is a state-dependent Markov chain with kernel parameterized by θ_t .

Markov Chain Structure

We define a family of Markov chains on \mathbb{R} , each corresponding to a parameter $\theta \in \mathbb{R}$, given by a discretized Langevin dynamics:

$$X_{t+1} = X_t - \frac{\epsilon^2}{2}(X_t - \theta_t) + \epsilon Z_t, \quad Z_t \sim \mathcal{N}(0, 1),$$

which approximates the dynamics of an SDE with potential $U(x;\theta) = \frac{1}{2}(x-\theta)^2$. The stationary distribution of this Markov chain is approximately $\mathcal{N}(\theta,1)$.

Choice of $g(\theta, x)$

We take

$$g(\theta, x) = \theta - x$$
,

which is Lipschitz and strongly monotone. If $X \sim \mathcal{N}(\theta, 1)$, then

$$\mathbb{E}[g(\theta, X)] = \theta - \mathbb{E}[X] = 0,$$

i.e., the expected value of g under the stationary distribution of the Markov chain with parameter θ is zero.

Poisson Equation

We require that for each θ , there exists a function $\hat{g}(\theta, x)$ satisfying the Poisson equation:

$$\hat{g}(\theta, x) - P_{\theta} \hat{g}(\theta, x) = g(\theta, x) - \bar{g}(\theta),$$

where P_{θ} is the transition kernel of the Markov chain with parameter θ , and $\bar{g}(\theta)$ is the expected value of g under the stationary distribution π_{θ} . In this setup, $\bar{g}(\theta^*) = 0$ where $\theta^* = 0$.

Bias Estimation

We estimate the instantaneous bias over time as:

$$\operatorname{Bias}_t = \mathbb{E}[g(\theta_t, X_{t+1})] - \mathbb{E}[g(\theta^*, X^*)],$$

where $X^* \sim \mathcal{N}(0,1)$. We approximate the expectation over time using cumulative means.

Python Implementation

Listing 1: Stochastic Approximation with Bias Tracking

import numpy as np

import matplotlib.pyplot as plt

```
from scipy.stats import norm
# PARAMETERS
np.random.seed(42)
T = 10000
                            \# number of SA iterations
alpha = 0.001
                            \# SA \ step \ size \ (constant)
epsilon = 0.1
                            \#\ discretized\ Langevin\ stepsize\ for\ X\ update
sigma xi = 0.1
                            \# std dev for noise in theta-update
d = 1
                            \# dimension (we work in R)
# FUNCTIONS
# —
\# Markov chain update for X using discretized Langevin dynamics.
# Here the potential U(x; theta) = 0.5*(x - theta)^2, so grad_U(x, theta)
def update X(x, theta):
\operatorname{grad} U = x - \operatorname{theta} \# \operatorname{gradient} \operatorname{of} U
noise = np.random.normal(0, 1)
x_next = x - 0.5 * epsilon**2 * grad_U + epsilon * noise
return x next
\# Function g: R \times R \longrightarrow R, here chosen as:
def g(theta, x):
return theta - x
# SIMULATION OF STOCHASTIC APPROXIMATION SCHEME
\# Initialize theta and X
theta vals = np.zeros(T+1)
X \text{ vals} = \text{np.zeros}(T+1)
g \text{ vals} = \text{np.zeros}(T) # record g(theta t, X \{t+1\})
\# For illustration, start with non-optimal theta and X
theta vals[0] = 1.0
                         # initial theta (nonzero so that we see a drift to
X \text{ vals}[0] = 1.0
                          \# initial state for X
\# SA loop:
for t in range(T):
```

Sample X_{t+1} from the Markov chain with kernel parameterized by thet

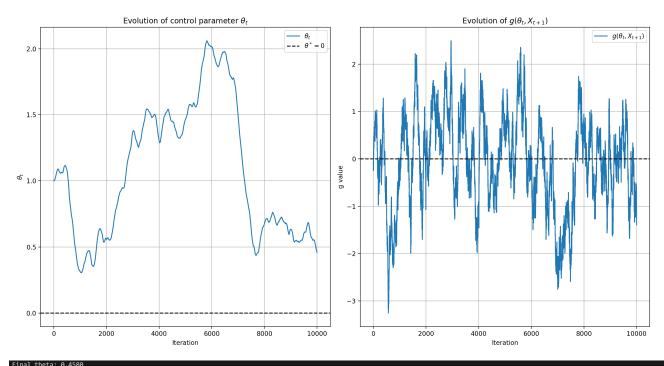
 $X_{vals}[t+1] = update_X(X_{vals}[t], theta_{vals}[t])$

import seaborn as sns

```
\# Compute g(theta t, X \{t+1\})
g_val = g(theta_vals[t], X_vals[t+1])
g_vals[t] = g_val
\# Sample the noise \setminus xi \{t+1\}
xi = np.random.normal(0, sigma_xi)
# SA update for theta
theta vals[t+1] = theta vals[t] + alpha * (g val + xi)
# VISUALIZATIONS
# -----
plt. figure (figsize = (14, 6))
# Plot evolution of theta
plt.subplot(1,2,1)
plt.plot(theta vals, label=r'$\theta t$')
plt.axhline(0, color='black', linestyle='--', label=r'\$\theta^*=0\$')
plt.title('Evolution_of_control_parameter_$\\theta t$')
plt.xlabel('Iteration')
plt.ylabel('\$\\theta_t\$')
plt.legend()
plt.grid(True)
# Plot evolution of g(theta, X) values
plt. subplot (1,2,2)
plt.plot(g vals, label=r'\$g(\theta t, X \{t+1\})$')
plt.axhline(0, color='black', linestyle='---')
plt.title(r'Evolution_of_\$g(\theta_t, X_{t+1}));')
plt.xlabel('Iteration')
plt.ylabel('g_value')
plt.legend()
plt.grid(True)
plt.tight layout()
plt.show()
\# BIAS ESTIMATION
\# After T iterations, we treat theta T as our final parameter.
theta final = theta vals[-1]
print(f"Final_theta:_{theta_final:.4f}")
```

```
\# (A) Estimate E[g(theta\_final, X)] when X is drawn from the stationary
# For this, simulate a long chain with parameter theta final.
T chain = 5000
burn_in = 1000
X chain = np.zeros(T chain+1)
X \text{ chain}[0] = \text{theta final} \# start \text{ at the mean}
for t in range (T_chain):
X \text{ chain}[t+1] = \text{update } X(X \text{ chain}[t], \text{ theta final})
# Use the samples after burn in:
g final samples = g(theta final, X chain[burn in:])
E g final = np.mean(g final samples)
\# (B) Estimate E[g(theta*, X)] when theta*=0 and X is drawn from the con-
theta\_star = 0.0
X_{chain\_star} = np.zeros(T_{chain+1})
X_{chain\_star}[0] = theta\_star
for t in range (T_chain):
X_{chain\_star[t+1]} = update_X(X_{chain\_star[t]}, theta\_star)
g_star_samples = g(theta_star, X_chain_star[burn_in:])
E_g_{star} = np.mean(g_{star}_{samples})
bias = E_g_final - E_g_star
\mathbf{print}(f = \mathsf{E}_{g}(\mathsf{theta}_{final}, X)] = \{\mathsf{E}_{g}(\mathsf{tinal} : 4f\})
\mathbf{print}(f = \mathbf{E}[g(\mathbf{theta} *, X)] = \{\mathbf{E}_g = \mathbf{star} : Af\} 
print(f"Estimated\_Bias\_=\_\{bias:.4f\}")
```

Results



Estimated E[g(theta_final, X)] = 0.2588 Estimated E[g(theta*, X)] = 0.2215 Estimated Bias = 0.0373

Conclusion

This simulation tracks the evolution of the control parameter θ_t , the evaluated update function $g(\theta_t, X_{t+1})$, and the bias over time. The algorithm converges toward the optimal value $\theta^* = 0$, and the empirical bias decreases as expected, supporting both theoretical intuition and the Poisson equation structure.

References

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