

IE 7374: Machine Learning

Answer to Question 1

a)

$$E(x) = \sum_{x=1}^4 xp(x) = 1 \cdot \frac{1}{10} + 2 \cdot \frac{2}{10} + 3 \cdot \frac{3}{10} + 4 \cdot \frac{4}{10} = 3$$

$$Var(x) = E(x^2) - [E(x)]^2 = 10 - 9 = 1, \text{ where } E(x^2) = \sum_{x=1}^4 x^2 p(x) = 10$$

b)

$$p(x|y) = \frac{p(y|x)p(x)}{p(y)}$$

$$1) y = 0, p(y) = \sum_{x=1}^4 p(y = 0|x)p(x) = \frac{1}{4} \cdot \frac{2}{10} + \frac{1}{3} \cdot \frac{3}{10} + \frac{3}{8} \cdot \frac{4}{10} = \frac{3}{10}$$

- $x = 1, p(y = 0|x = 1) = 0, p(x = 1) = \frac{1}{10}, p(x = 1|y = 0) = 0$
- $x = 2, p(y = 0|x = 2) = \frac{1}{4}, p(x = 1) = \frac{2}{10}, p(x = 2|y = 0) = \frac{1}{4} \cdot \frac{2}{10} / \frac{3}{10} = \frac{1}{6}$
- $x = 3, p(y = 0|x = 3) = \frac{1}{3}, p(x = 3) = \frac{3}{10}, p(x = 3|y = 0) = \frac{1}{3}$
- $x = 4, p(y = 0|x = 4) = \frac{3}{8}, p(x = 4) = \frac{4}{10}, p(x = 4|y = 0) = \frac{1}{2}$

Similarly, $y = 1$

- $p(x = 1|y = 1) = \frac{1}{7}$
- $p(x = 2|y = 1) = \frac{3}{14}$
- $p(x = 3|y = 1) = \frac{2}{7}$
- $p(x = 4|y = 1) = \frac{5}{14}$

c)

$$E[X|Y = 1] = \sum_{x=1}^4 p(x|y = 1)p(x) = \frac{1}{7} \cdot 1 + \frac{3}{14} \cdot 2 + \frac{2}{7} \cdot 3 + \frac{5}{14} \cdot 4 = \frac{20}{7}$$

$$\begin{aligned} \theta_{MAP} &= \operatorname{argmax} \prod_{i=1}^N P(y_i|\theta^T x_i, \sigma^2) P(\theta) \\ &= \operatorname{argmax} \prod_{i=1}^N \left[\frac{1}{\sqrt{2\pi\sigma^2}} \exp \frac{-(y_i - \theta^T x_i)^2}{2\sigma^2} \cdot \frac{\lambda}{2} \exp^{-\lambda|\theta|} \right] \\ \Rightarrow \operatorname{argmax} \sum_{i=1}^N \left[\frac{-(y_i - \theta^T x_i)^2}{2\sigma^2} - \lambda|\theta| \right] \end{aligned}$$

Answer to Question 2

$$\begin{aligned}
Cov(x, y) &= E[(x - E[x])(y - E[y])] \\
&= E[xy - xE[y] - yE[x] + E[x]E[y]] \\
&= E[xy] - E[xE[y]] - E[yE[x]] + E[x]E[y] \\
&= E[x][y] - E[y]E[x] - E[x]E[y] + E[x]E[y] \\
&= 0
\end{aligned}$$

The prove holds trues only when $E[xy] = E[x]E[y]$, which means variables x, y are independent.

Answer to Question 3

a)
 $E[f[X]] = \sum_{x_i \in \{a, b, c\}} p(f(x_i))p(x_i) = 10 \cdot 0.1 + 5 \cdot 0.2 + \frac{10}{7} \cdot 0.7 = 3$

b)
 $E[1/p(X)] = \sum_{x_i \in \{a, b, c\}} \frac{1}{p(x_i)} p(x_i) = 3$

Answer to Question 4

a)
 $X = \{X_1, X_2, \dots, X_i, \dots, X_n\}$, $X_i \sim \mathcal{N}(\mu, \sigma^2)$, $\theta = (\mu, \sigma^2)$
For MLE,

$$\begin{aligned}
P = P(X|\theta) &= \prod_{i=1}^n P(X_i|\theta) \\
&\xrightarrow{\text{Apply log}} \sum_{i=1}^n \log P(X_i|\theta) \\
&= \sum_{i=1}^n \log \left(\frac{1}{\sqrt{2\pi\sigma^2}} \cdot \exp \frac{-(X_i - \mu)^2}{2\sigma^2} \right) \\
&= \sum_{i=1}^n \left[\log \frac{1}{\sqrt{2\pi}} - \log \sigma - \frac{(X_i - \mu)^2}{2\sigma^2} \right]
\end{aligned}$$

Find the partial deviates of μ and σ^2 ,

$$\frac{\partial P}{\partial \mu} = \sum_{i=1}^n \frac{X_i - \mu}{\sigma^2}$$

$$\text{let } \frac{\partial P}{\partial \sigma} = \sum_{i=1}^n \left[\frac{1}{\sigma} + \frac{1}{2} (X_i - \mu)^2 \cdot (-2) \sigma^{-3} \right]$$

let $\frac{\partial P}{\partial \mu} = 0$, $\frac{\partial P}{\partial \sigma} = 0$,

$$\hat{\mu} = \frac{1}{n} \sum_{i=1}^n X_i$$

$$\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \hat{\mu})^2$$

b)

$E[\hat{\mu}] = \mu$, unbiased; $E[\hat{\sigma}^2] = \frac{n-1}{n}\sigma^2$, biased

Prove:

To verify if biased, We test the expectation of respective predicting parameters in sample data using MLEs with these real parameters under normal distribution having overall data records. If these values are equal, then biased, otherwise, unbiased.

$$E[\hat{\mu}] = E\left[\frac{1}{n} \sum_{i=1}^n X_i\right] = \frac{1}{n} \sum_{i=1}^n E[X_i] = \mu$$

$$\begin{aligned} E[\hat{\sigma}^2] &= E\left[\frac{1}{n} \sum_{i=1}^n (X_i - \hat{\mu})^2\right] \\ &= E\left[\frac{1}{n} \sum_{i=1}^n [(X_i - \mu) - (\hat{\mu} - \mu)]^2\right] \\ &= E\left[\frac{1}{n} \sum_{i=1}^n (X_i - \mu)^2 - \frac{2}{n} \sum_{i=1}^n (X_i - \mu)(\hat{\mu} - \mu) + \frac{1}{n} \sum_{i=1}^n (\hat{\mu} - \mu)^2\right] \\ &= E\left[\frac{1}{n} \sum_{i=1}^n (X_i - \mu)^2 - 2(\hat{\mu} - \mu)^2 + (\hat{\mu} - \mu)^2\right] \\ &= E\left[\frac{1}{n} \sum_{i=1}^n (X_i - \mu)^2 - (\hat{\mu} - \mu)^2\right] \\ &= \frac{1}{n} \sum_{i=1}^n E(X_i - \mu)^2 - E[(\hat{\mu} - \mu)^2] \\ &= \sigma^2 - \text{Var}[\hat{\mu}] \\ &= \sigma^2 - \frac{1}{n}\sigma^2 \\ &= \frac{n-1}{n}\sigma^2 \end{aligned}$$

An unbiased for σ^2 can be obtained, $\sigma_{ub} = \frac{n}{n-1}\hat{\sigma}^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \hat{\mu})^2$

Answer to Question 5

Suppose variable X and Y are continuous one,

$$\begin{aligned} E[XY] &= \iint_{x,y} xy \cdot p(xy) dx dy \\ &= \iint_{x,y} xy \cdot p(x)p(y) dx dy \\ &= \int_x xp(x) dx \cdot \int_y yp(y) dy \\ &= E[x]E[y] \end{aligned}$$

$p(xy) = p(x)p(y)$ holds if variables x, y are independent.

Answer to Question 6

a) ii

$$P(H|e1, e2) = \frac{P(e1, e2|H)P(H)}{P(e1, e2)}$$

b) i & ii

$$\begin{aligned} P(H|e1, e2) &= \frac{P(e1, e2|H)P(H)}{P(e1, e2)} \\ &= \frac{P(e1|H)P(e2|H)P(H)}{P(e1, e2)} \end{aligned}$$

Answer to Question 7

$$\begin{aligned} Var[X + Y] &= E[(X + Y - E[X + Y])^2] \\ &= E[(X + Y)^2 - 2(X + Y)E[X + Y] + (E[X + Y])^2] \\ &= E[X^2] + 2E[XY] + E[Y^2] - 2E[X + Y]E[X + Y] \\ &\quad + (E[X])^2 + (E[Y])^2 + 2E[X]E[Y] \\ &= E[X^2] + 2E[XY] + E[Y^2] - 2[(E[X])^2 + 2E[X]E[Y] + (E[Y])^2] \\ &\quad + (E[X])^2 + (E[Y])^2 + 2E[X]E[Y] \\ &= E[X^2] - (E[X])^2 + E[Y^2] - (E[Y])^2 + 2(E[XY] - E[X]E[Y]) \\ &= Var[X] + Var[Y] + 2Cov[X, Y] \end{aligned}$$

If variables X, Y are independent, $Cov[X, Y] = 0$, according the insights from Question 5. Then, $Var[X + Y] = Var[X] + Var[Y]$

Answer to Question 8

We can easily tell that $p(x|\theta_0)$ is an increasing function with respect to θ_0 . This can be obtain by calculating the derivative of joint probability function, $\frac{\partial \log p(x|\theta_0)}{\partial \theta_0} = n > 0$. It shows $\log p(x|\theta_0)$ is an increasing function, and $\log(x)$ is also an increasing function. Then, $p(x|\theta_0)$ increase, too. As $\theta_0 \leq x$, the maximum like-hood estimate of θ_0 is $\min(x_1, x_2, \dots, x_n)$.

Answer to Question 9

a)

To prove a matrix A is a positive-semi definite, which is to show $x^T A x \geq 0$.

$$x^T (A^T A) x = (Ax)^T A x = \|Ax\|^2 \geq 0$$

b)

If A is full rank, $\text{Rank}(A) = n$ & $m \geq n$, then $A^T A$ is positive definite.