

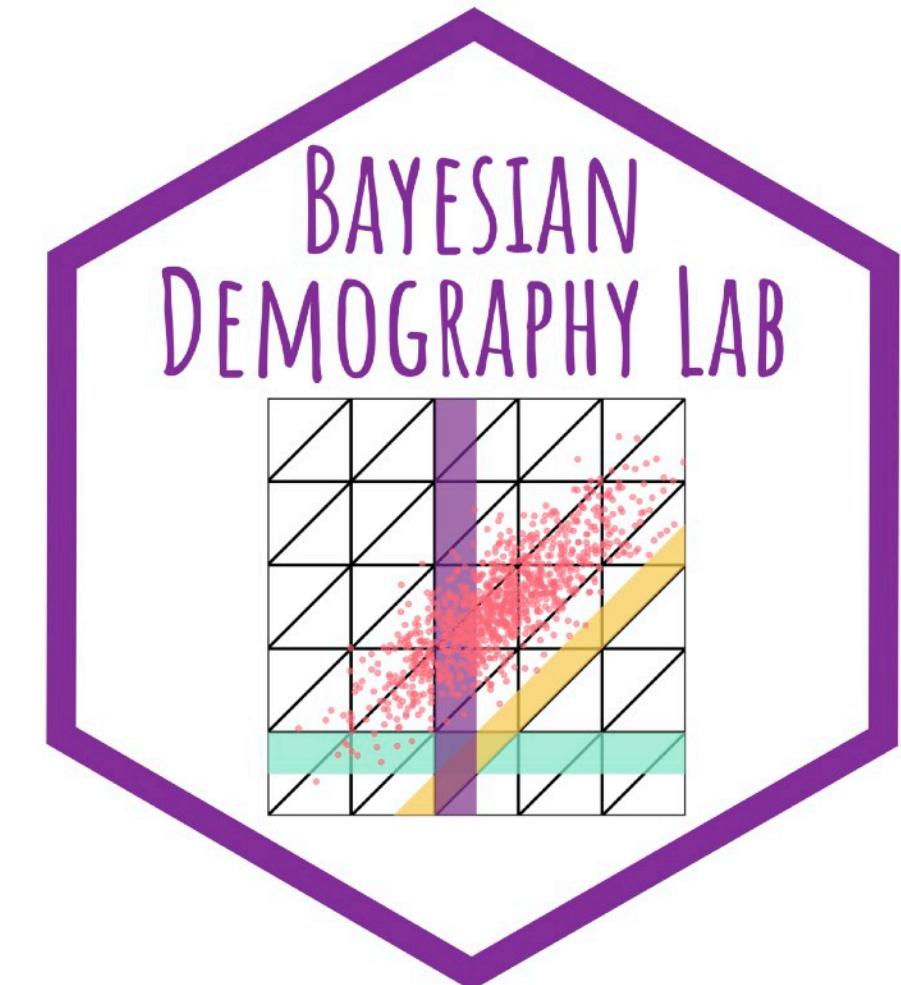
# Measuring and understanding inequalities in parental loss

New statistical methods and estimates

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University of Toronto

# Introduction

- Demographer, with a background in statistics
- Interested in demographic questions in contexts where populations/outcomes are hard to measure
- A large part of my work centres on the development of Bayesian methods for demographic estimation



Grew up in Hobart, Tasmania, Australia



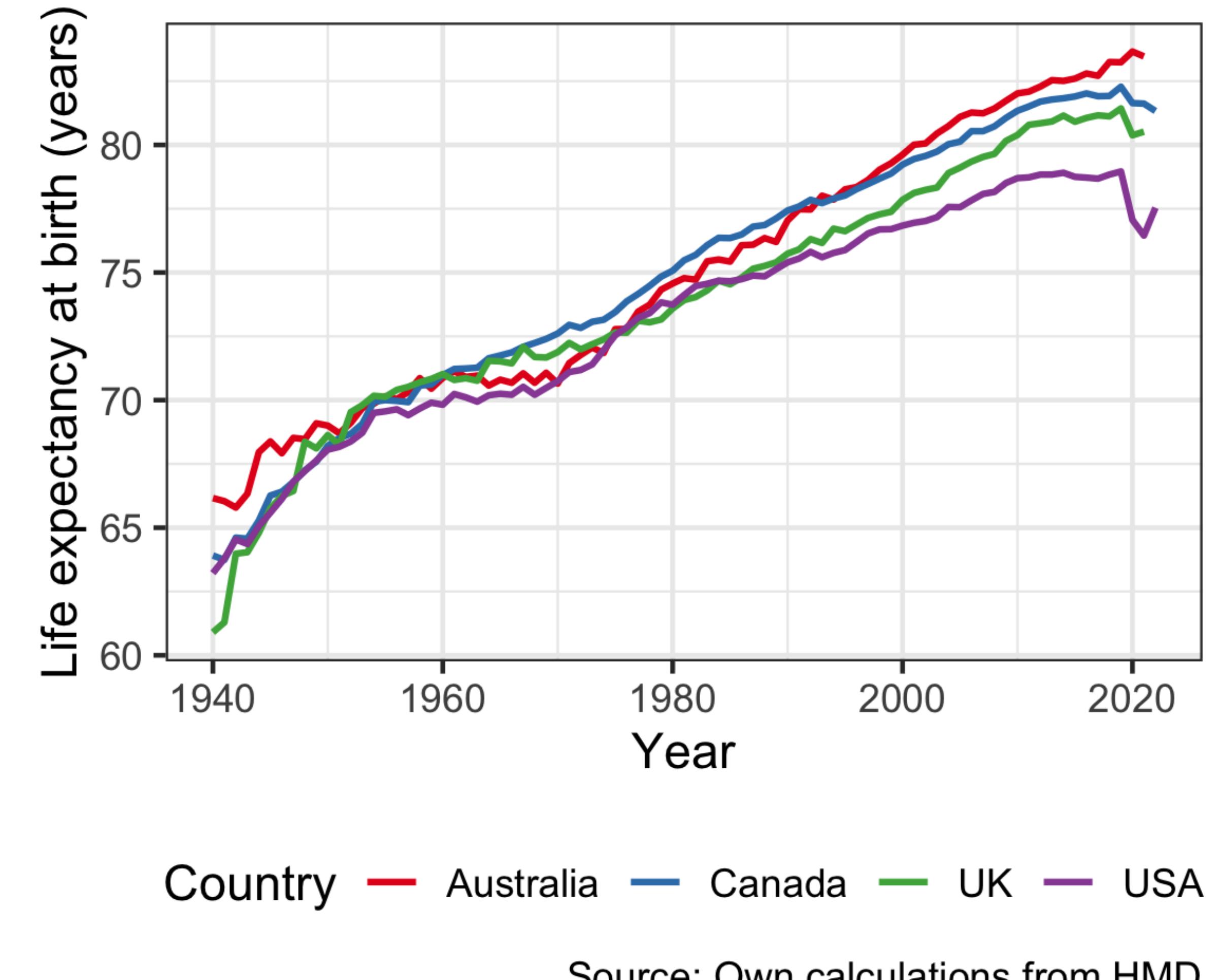
Masters/PhD at UC Berkeley



Associate Prof at University of Toronto

# Motivation

- Life expectancy in the United States has been stagnating (Abrams et al, 2023)
- The rise of young- and middle-age adult mortality is a well-documented contributor (Polizzi and Dowd, 2024)
- Disparities in mortality by race/ethnicity persist (Dwyer-Lindgren et al, 2022, Hendi 2024)



Source: Own calculations from HMD

# Beyond death counts: changing the viewpoint

- Focus on **parental loss**
- An increase in adult mortality implies an increase in children who have lost a parent at a young age
- Taking key demographic quantities and changing the viewpoint
- Analyzing demographic rates from a different perspective uncovers new information about social phenomena (e.g. Bumpass and Lu 2000; Alburez-Gutierrez 2022)

# The importance of studying parental loss

- The premature death of a parent is negatively associated with a range of outcomes (educational, health, economic) (Niederkrotenthaler et al. 2012; Pham et al. 2018; Patterson et al. 2022)
- Differing levels of parental death in different societies has implications for evolution of family structure, support of extended kin (Umberson 1995; Daw et al. 2016; Jiang et al. 2023)
- Macro implications of support required from healthcare and foster care systems (Roehrkasse 2021)
- Substantial racial/ethnic disparities in mortality translate into vastly different exposures to parental death, contributing to cumulative disadvantage (Umberson 2017, Umberson et al. 2017)

# How do we estimate parental loss?

- Quantifying the number of children who have lost a parent is not necessarily straightforward
- Not recorded on death certificates, linked data is restricted
- Parental loss captured in some surveys, but sample sizes are small, some groups are underrepresented, and cause of death is usually not known

# How do we estimate parental loss?

- Estimation is possible using vital rates (births and deaths) and demographic projection models
- But these have strong assumptions, and do not allow for the combination of data sources, or uncertainty quantification
- (Bayesian) statistical methods provide a framework to extend classical demographic methods to overcome some limitations

# This talk

- Two projects on parental loss in the United States context
  - Youth who have lost a parent due to drugs and firearms
  - Time lived without parents
  - Highlight differences by race/ethnicity
- Methodological contributions:
  - Extend existing demographic methods to estimate parental loss by cause
  - Develop a Bayesian statistical framework to combine multiple data sources on parental loss

Youth experiencing parental  
deaths due to drug poisonings  
and firearm violence

# Motivation

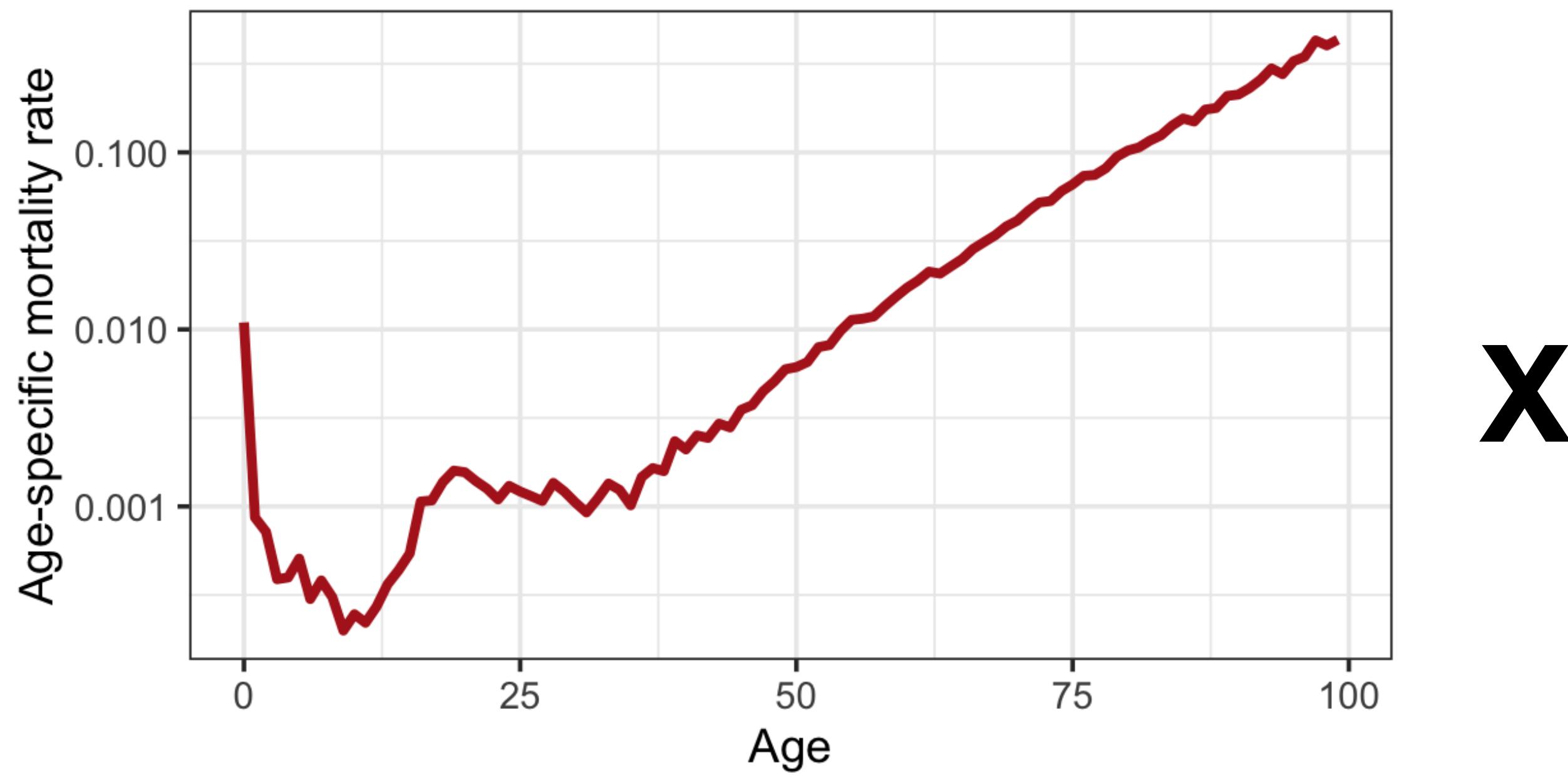
- The United States is experiencing dual and overlapping public health crises: drug poisonings and firearm deaths
- Since 1999, over 1 million people have died due to drug overdoses, and more than 750,000 by firearms (NCHS 2024)
- The majority of these deaths occur to those aged 15-54, a period when many have young children (NCHS 2024)
- Drug and firearms deaths have also increased for youths (<18 years); drug poisonings now the third leading cause of death (Goldstick et al 2022, Woolf et al 2023)
- Dual burden for youth: increased mortality, and losing a parent

There are no data available on parental loss by cause, so we obtain estimates using a demographic projection model

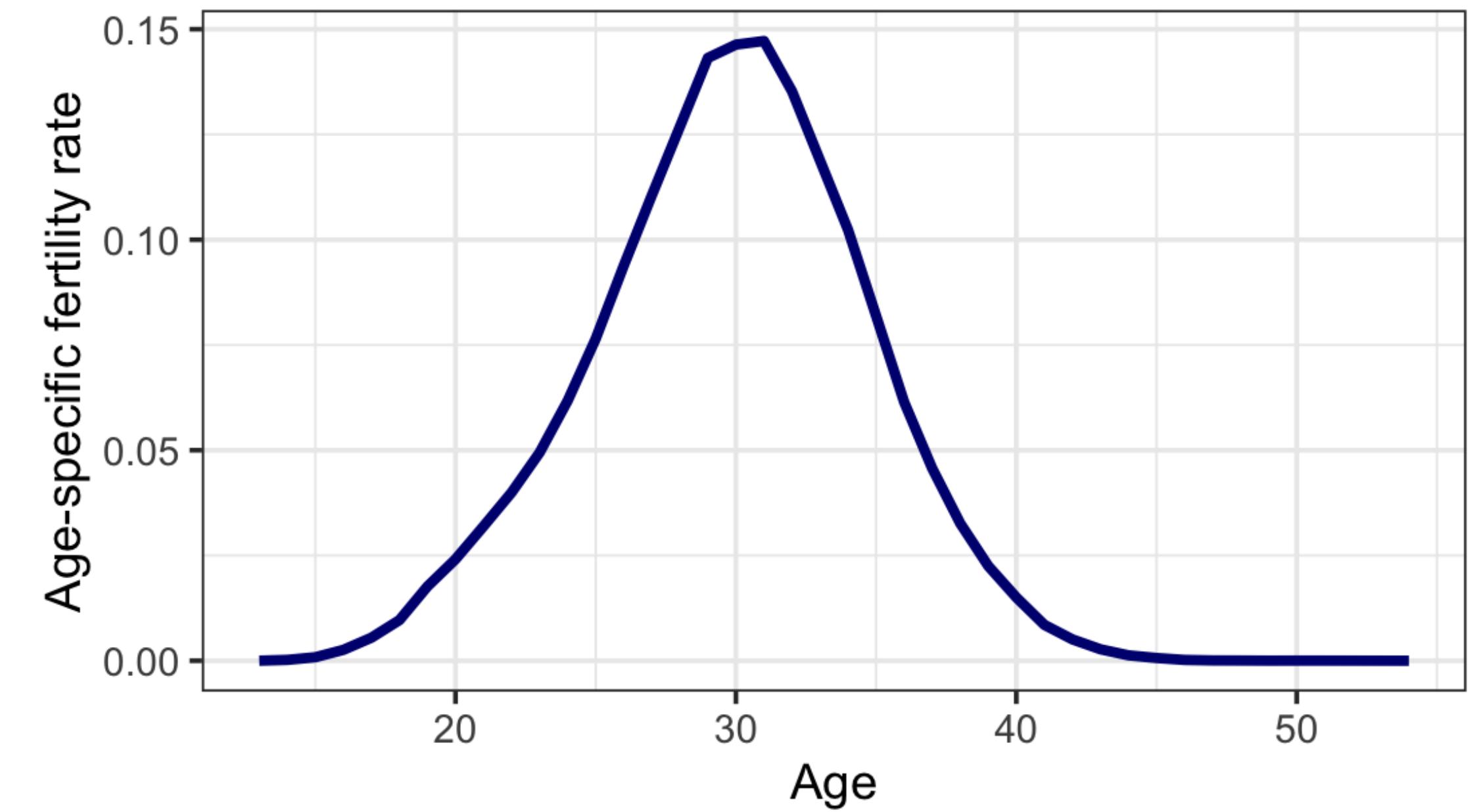
# Estimating parental loss with demographic rates

# Estimating parental loss with demographic rates

Intuition: If you know the mortality rates by age and the fertility rates by age in a population, then you can estimate expected number of mothers / fathers still alive

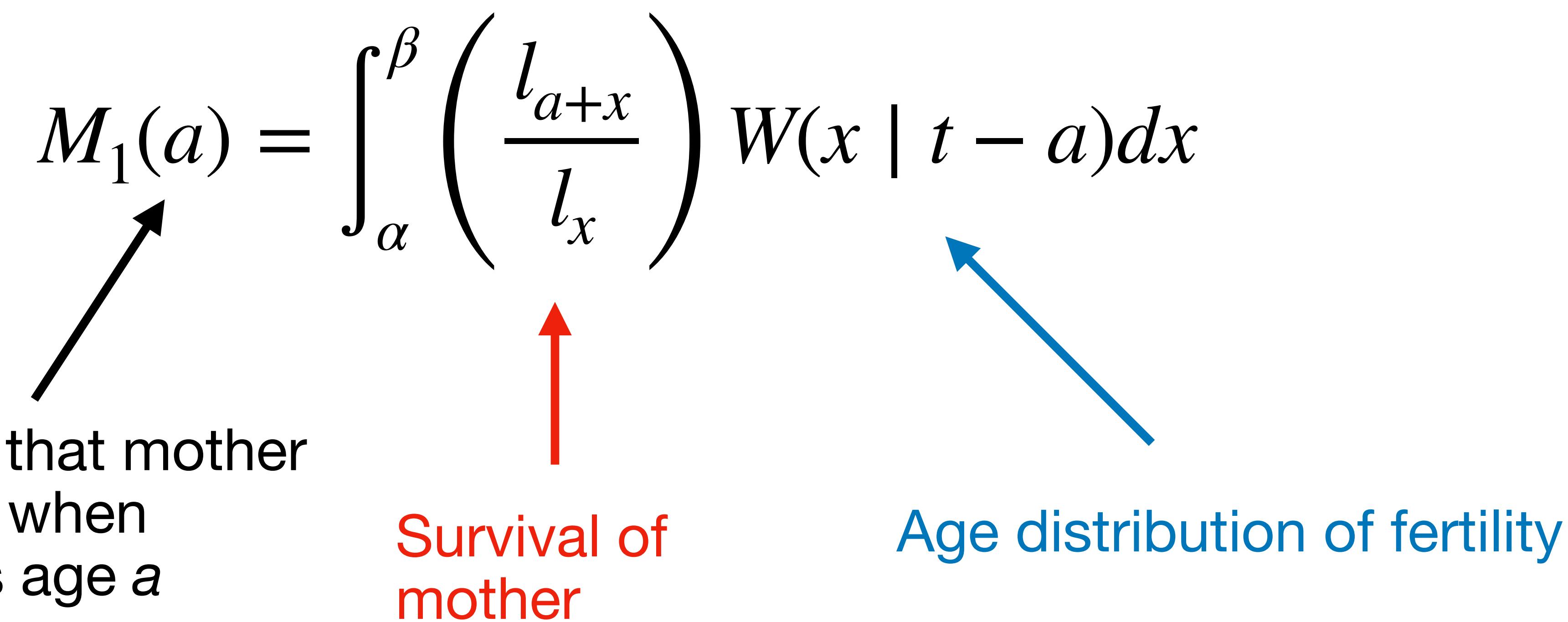


X



# Estimating parental loss with demographic rates

Goodman, Keyfitz, and Pullum (1974) laid out the mathematical foundations for quantifying surviving kin using demographic rates. E.g.

$$M_1(a) = \int_{\alpha}^{\beta} \left( \frac{l_{a+x}}{l_x} \right) W(x \mid t - a) dx$$


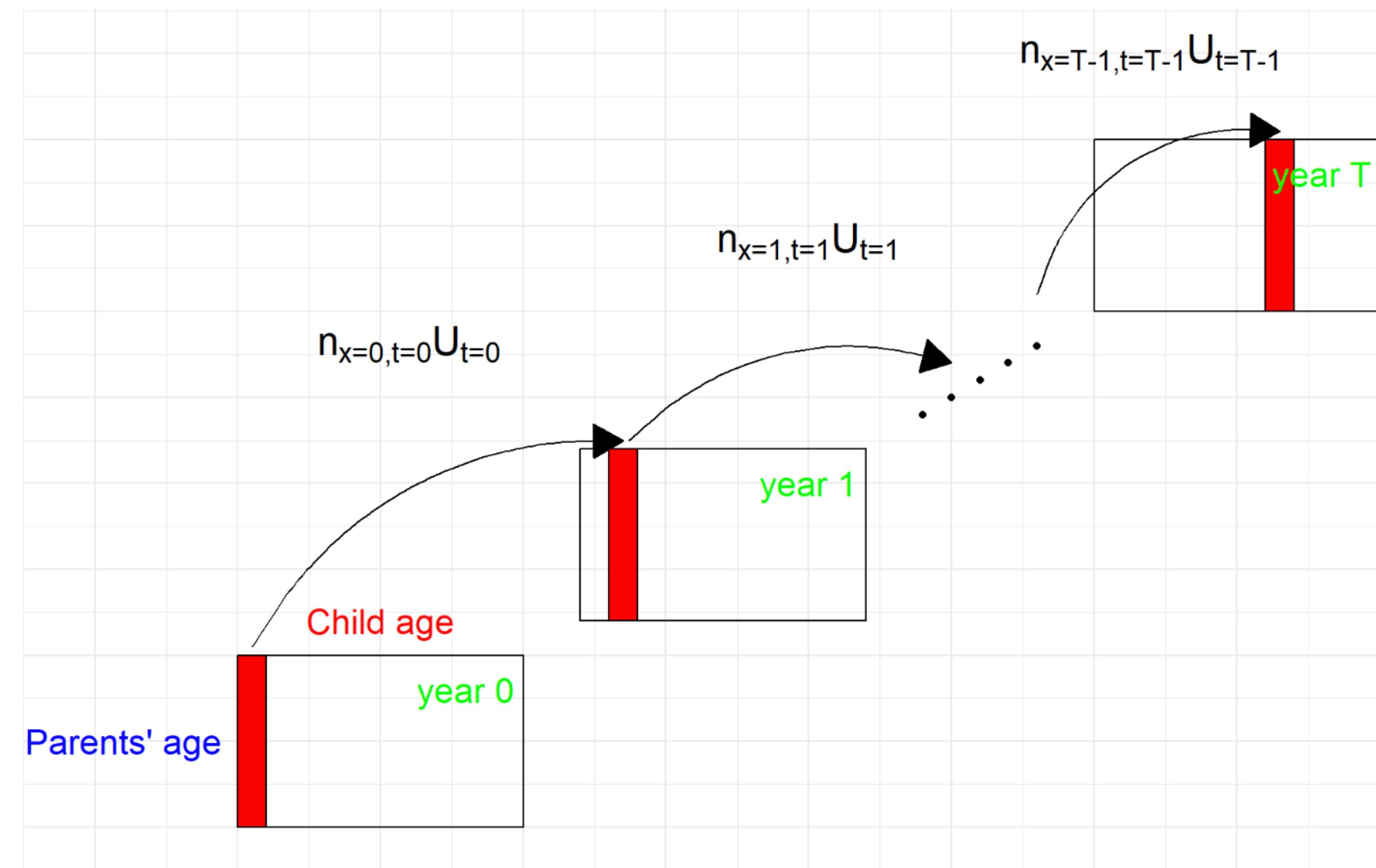
Probability that mother is still alive when daughter is age  $a$

Survival of mother

Age distribution of fertility

# Enter the Matrix

- Caswell (2019+): takes the Goodman, Keyfitz, Pullum equations, and turns them into a matrix projection model
- Extends the idea of the cohort component projection model
- Treat the children and parents as two different, but connected, populations
- We extended to allow us to track deaths of parents by cause (drugs, firearms, all other causes)



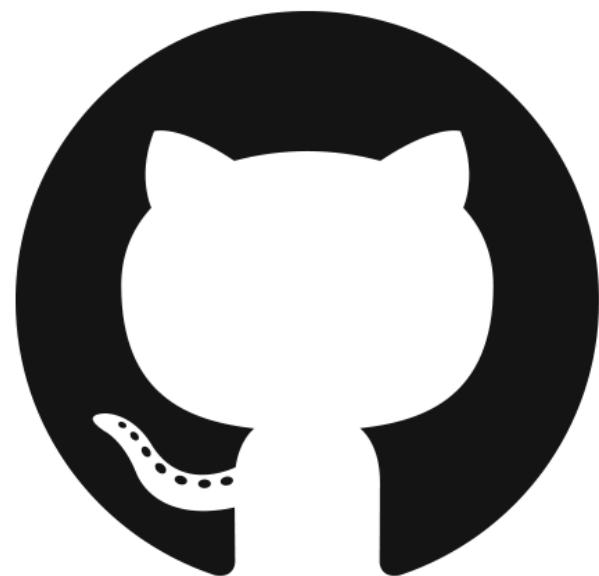
# Data

**Sources:** National Center for Health Statistics (NCHS) and US Census Bureau

- Multiple causes of death (ICD10 codes) from 1999 (firearm-related, drug-related deaths, all other causes of death combined)
- Calculated cause-specific deaths for all NCHS/CDC bridged-race groups by one-year age, and sex
  - Non-Hispanic Black, non-Hispanic White, Hispanic, and Total
- Age- and race/ethnicity-specific female fertility rates
- Population counts by age, sex, and race/ethnicity

# Reproducible materials

- Paper was published in *JAMA* earlier this year: doi:10.1001/jama.2024.8391
- All the data and code to produce the results, plots, and supplementary analyses is here: [https://github.com/benjisamschlu/parental\\_deaths](https://github.com/benjisamschlu/parental_deaths)



## Notes on Reproducibility

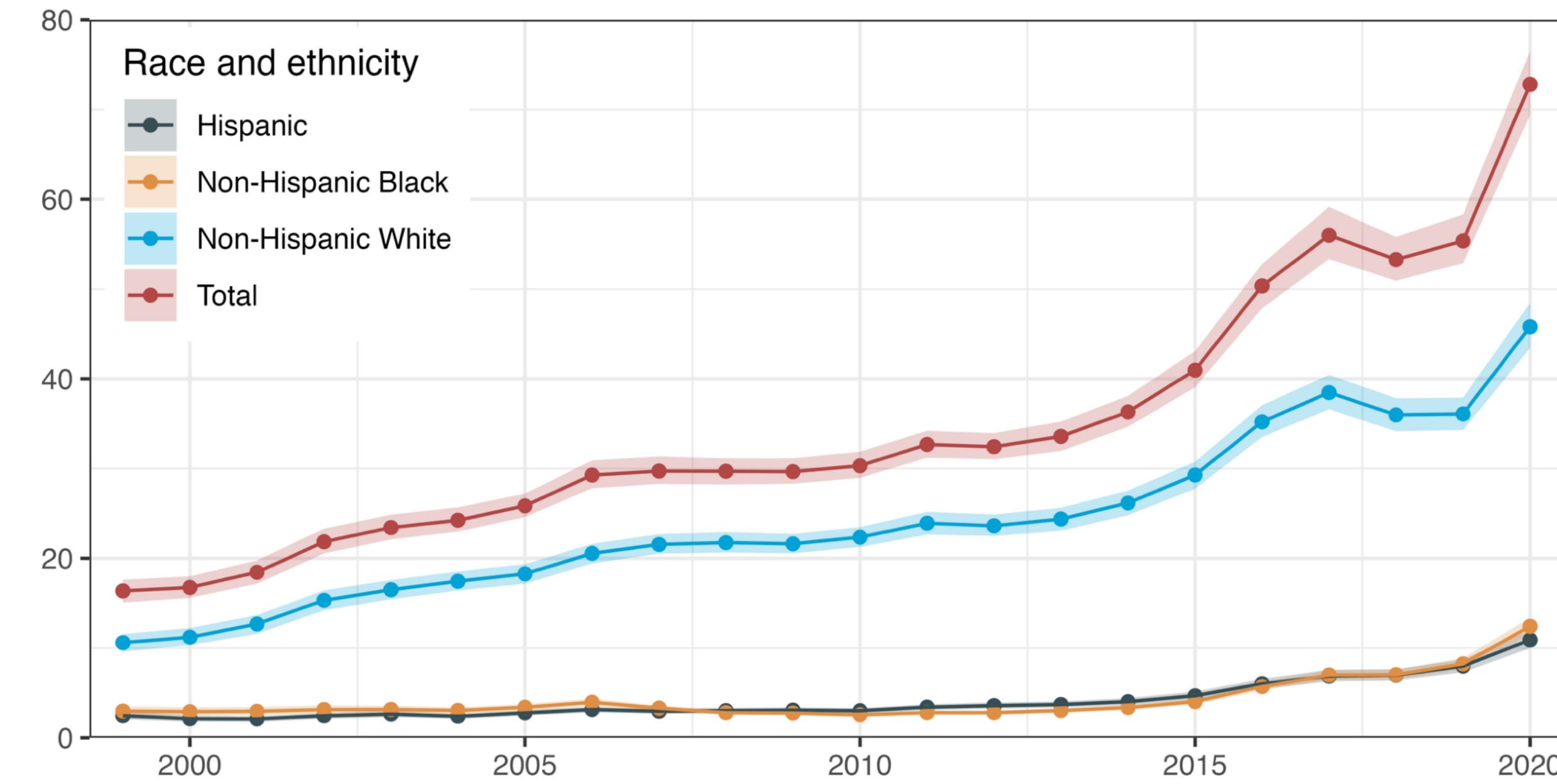
We use publicly available data and provide code that will download and/or munge these data as necessary. For full reproducibility, we also provide [all of our simulation files](#), which will allow for exact reproduction of our results. These files are stores on [the Harvard Dataverse](#); however, due to limitations on file size, we split each of the (>5GB) simulation files into 10 smaller chunks and provide code to recombine them (`./codes/00_combine_sim_files.R`). If you wish to fully reproduce our results, use the `00` code file to download the files noting that they take up approximately 55GB of space and require ~16GB of RAM (each file) to recombine.

We use [renv](#) for package management but below we also post the relevant session information to ensure full reproducibility.

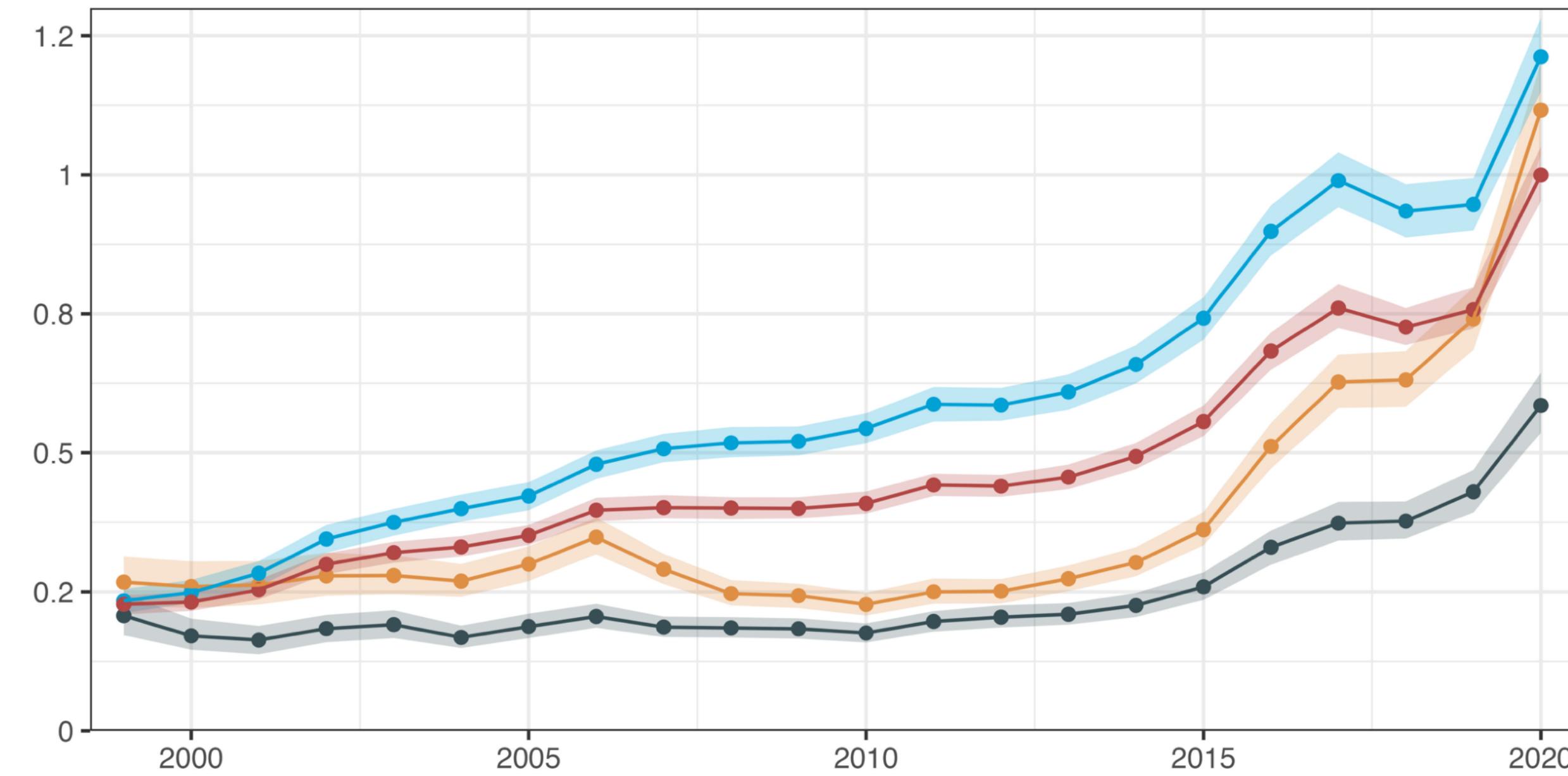
# Results

# Parental loss due to drugs

Number of youth <18 years impacted (95% CI), thousands

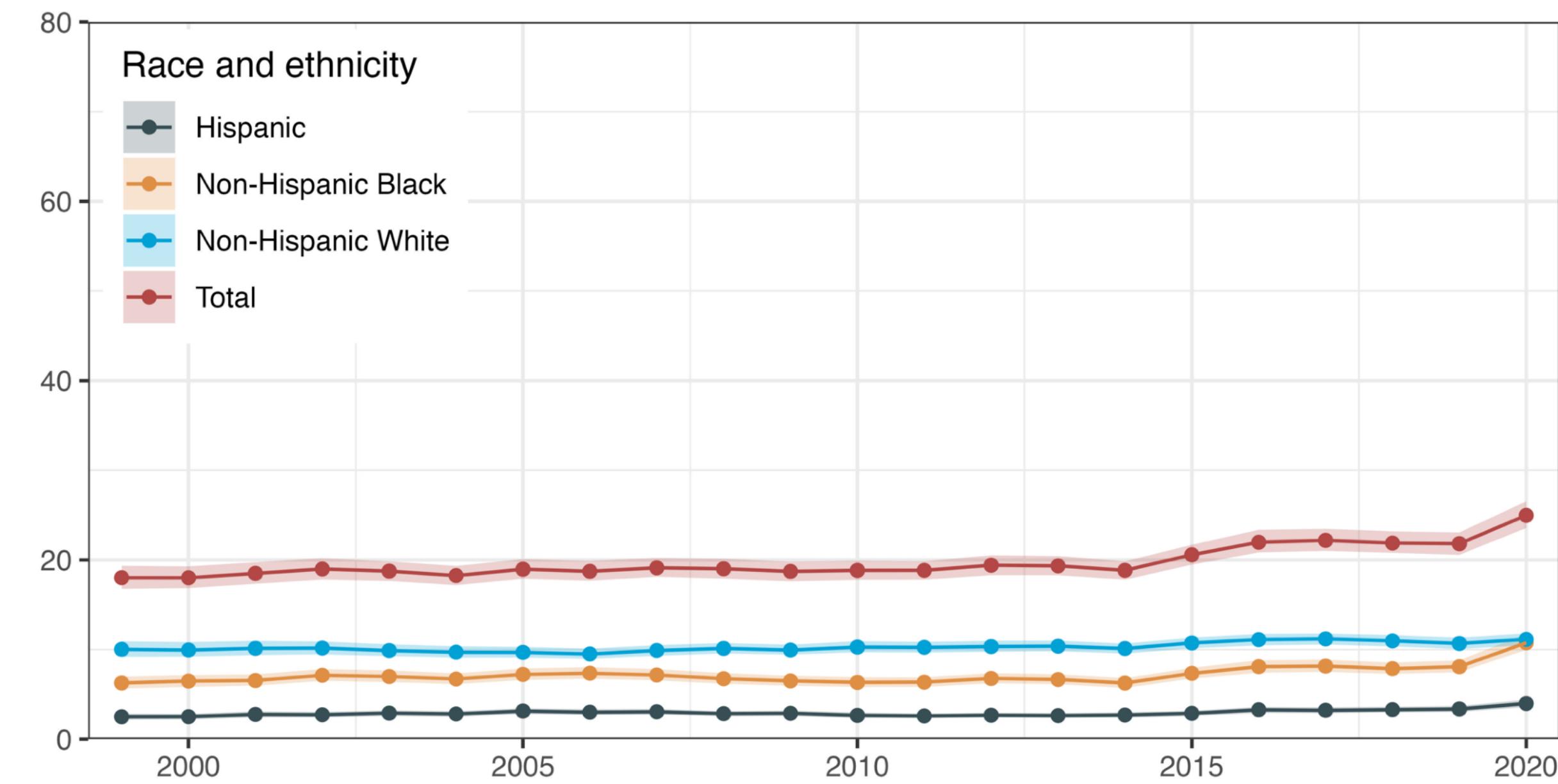


Youth impacted per 1,000 population (95% CI)

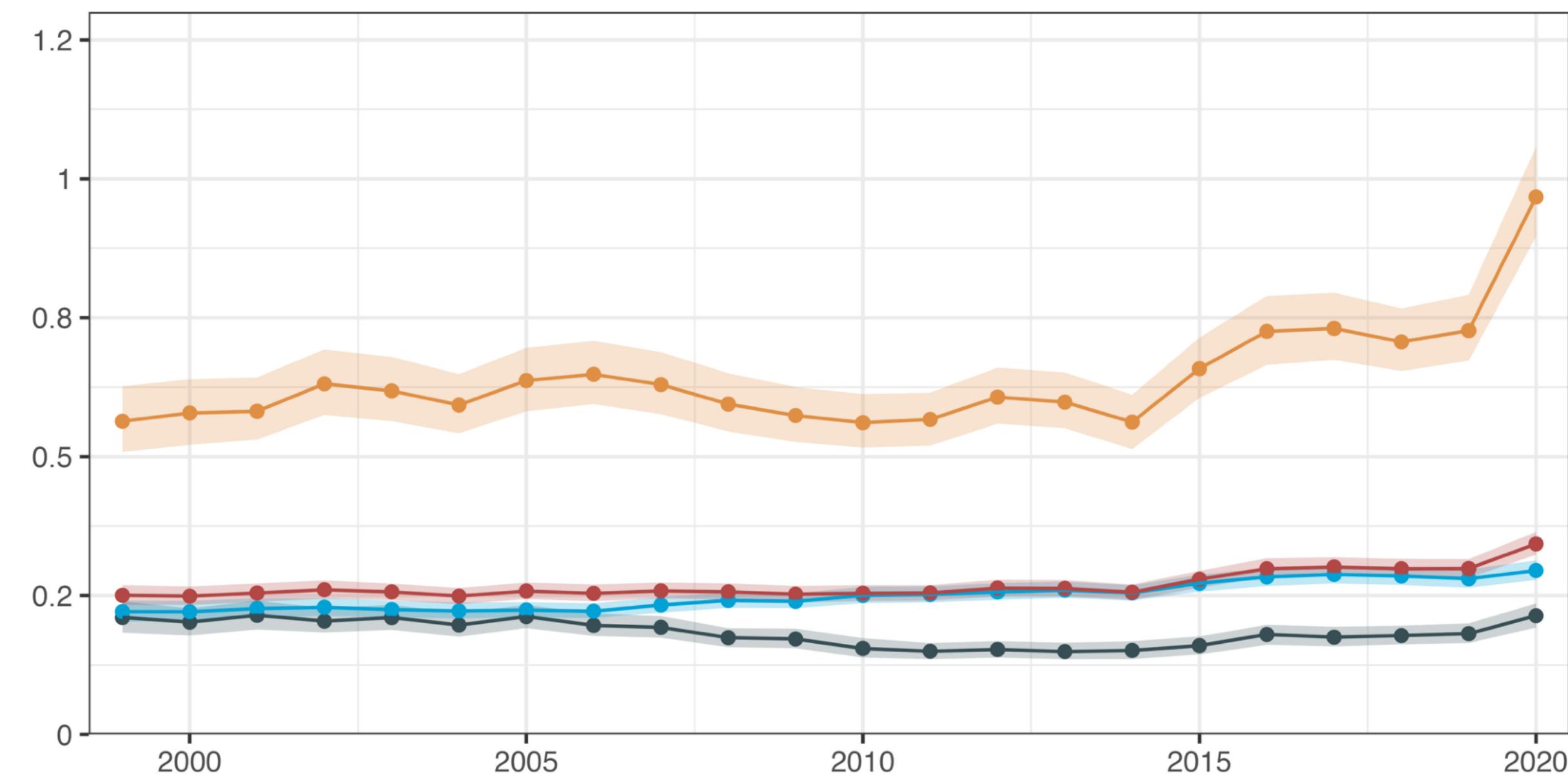


# Due to firearms

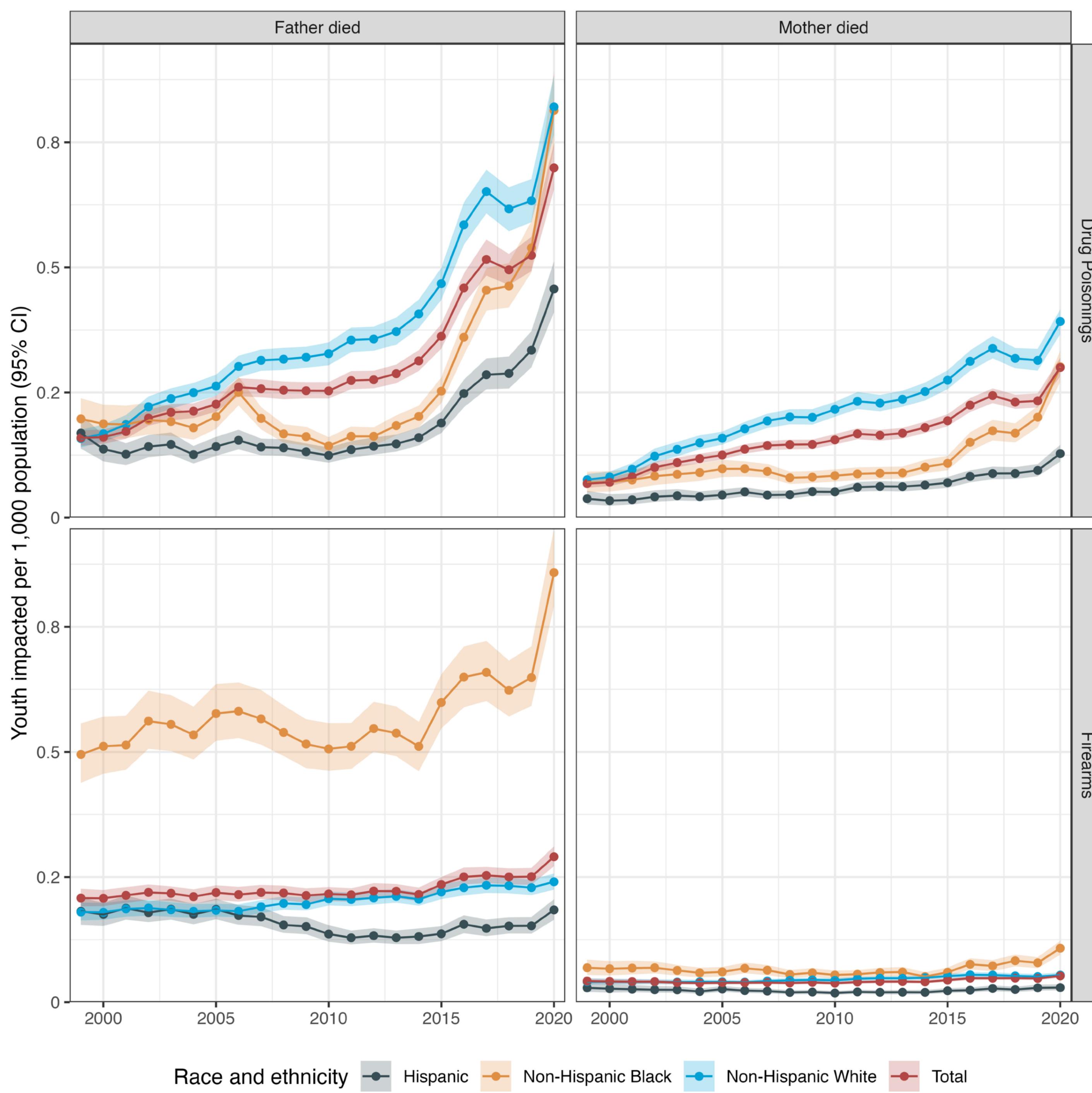
Number of youth <18 years impacted (95% CI), thousands



Youth impacted per 1,000 population (95% CI)



# By parent



# Takeaways

- An estimated more than 1 million youth lost a parent due to drugs or firearms over the period 1999-2020
- In 2020, drugs and firearms accounted for almost a quarter of all parental deaths (double the 1999 level)
- Black youth experienced a disproportionate burden of parental deaths, primarily due to firearm deaths among fathers
- A whole generation of youth who have lost a parent to traumatic causes
- Large racial differences of early parental loss, cumulative disadvantage and weathering over the life course
- Shifting policy to think about the indirect impact on families (US Surgeon General 2024)

# Methodological assumptions and limitations

These are model-based estimates!

- Definition of a parent, assumption of two parents
- Homogeneity of fertility
- Race classification, no multi-racial groups
- Uncertainty is most likely underreported

We did validation and sensitivity tests, but potential for future methodological work

# Estimating years of life lived without parents

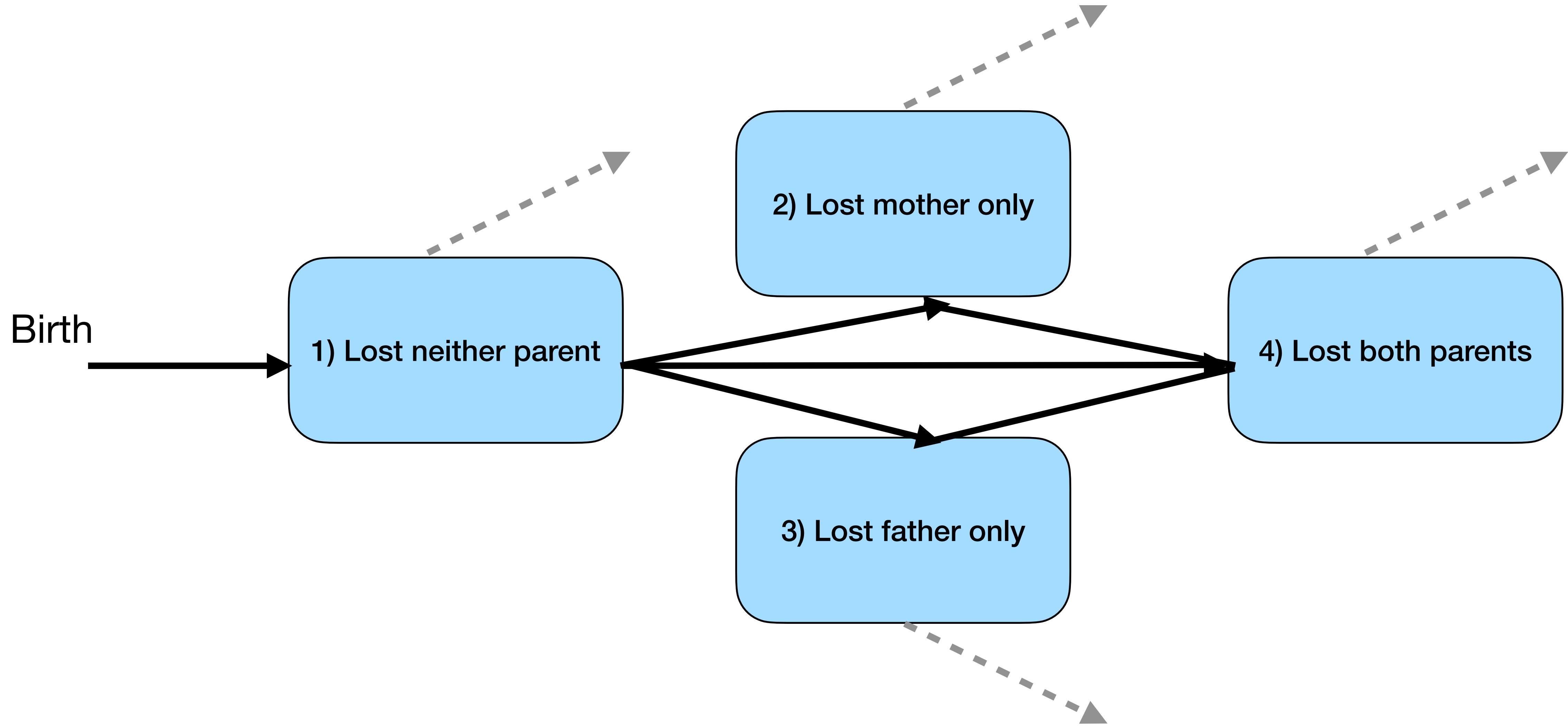
# Motivation

- Shifting focus from magnitude to timing and **length of exposure to parental loss**
- How much time can we expect to have with our parents?
- A function of mortality/fertility of parents and also mortality of ego
- Presence of parents at key points in life course (Cox et al 2016, Liu et al 2022)
- Shared lifetimes, shared resources, multigenerational mobility and capital (Song and Mare 2019)
- Uncertainty about lifespan of family members affects planning (van Raalte 2011, Grote and Pfrombeck 2020)

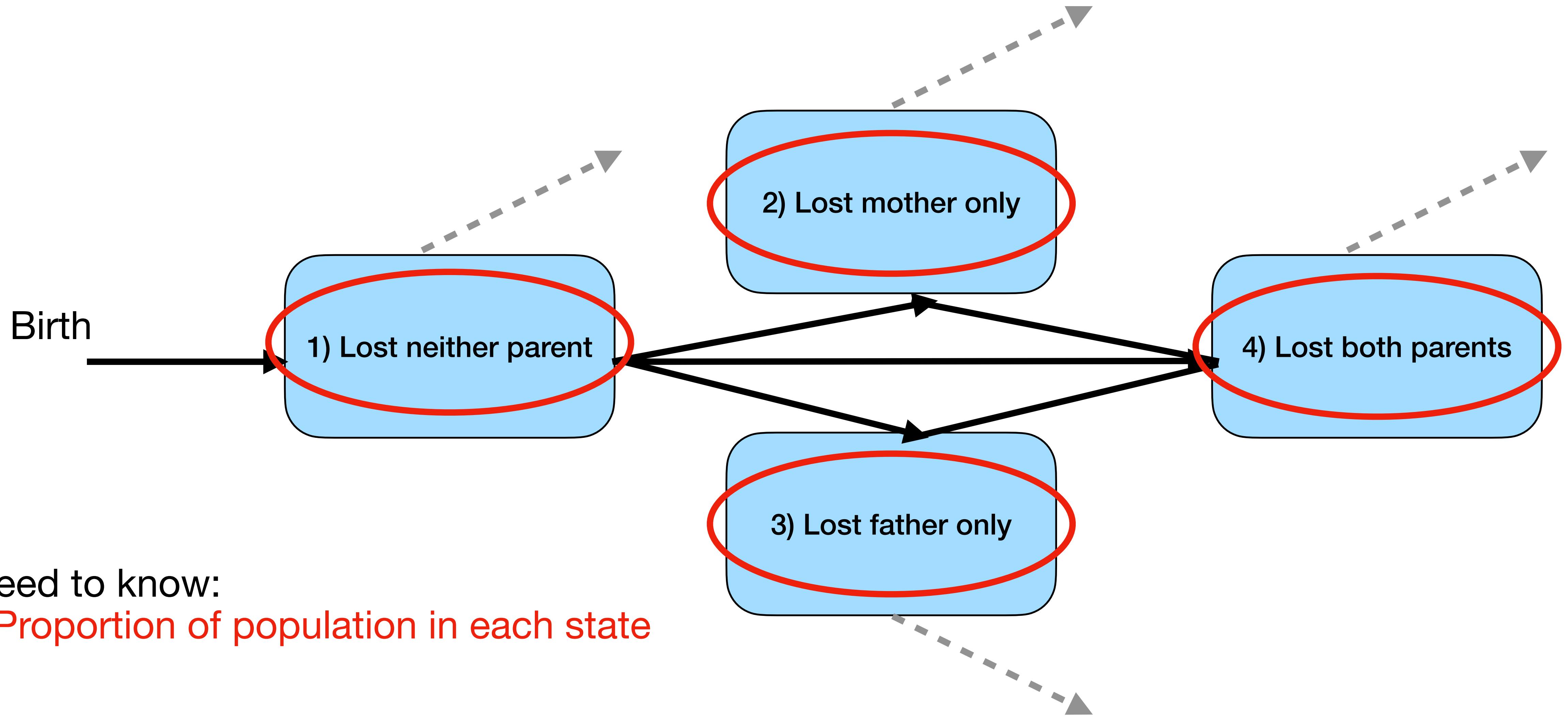
# Estimating time lived without parents

- We use a multistate life table approach
- Multistate lifetables track the mortality experience for a cohort of people over age, partitioning them into different ‘states’
- Here, our states relate to parental loss: 1) lost neither, 2) lost mother only, 3) lost father only, 4), lost both
- Key quantity of interest: life expectancy above age  $x$  in each state  $j$ ,  $e_x(j)$
- For example:  $e_0(1)$  is an estimate of the number of years a baby could expect to live with both of their parents alive

# States of parental loss



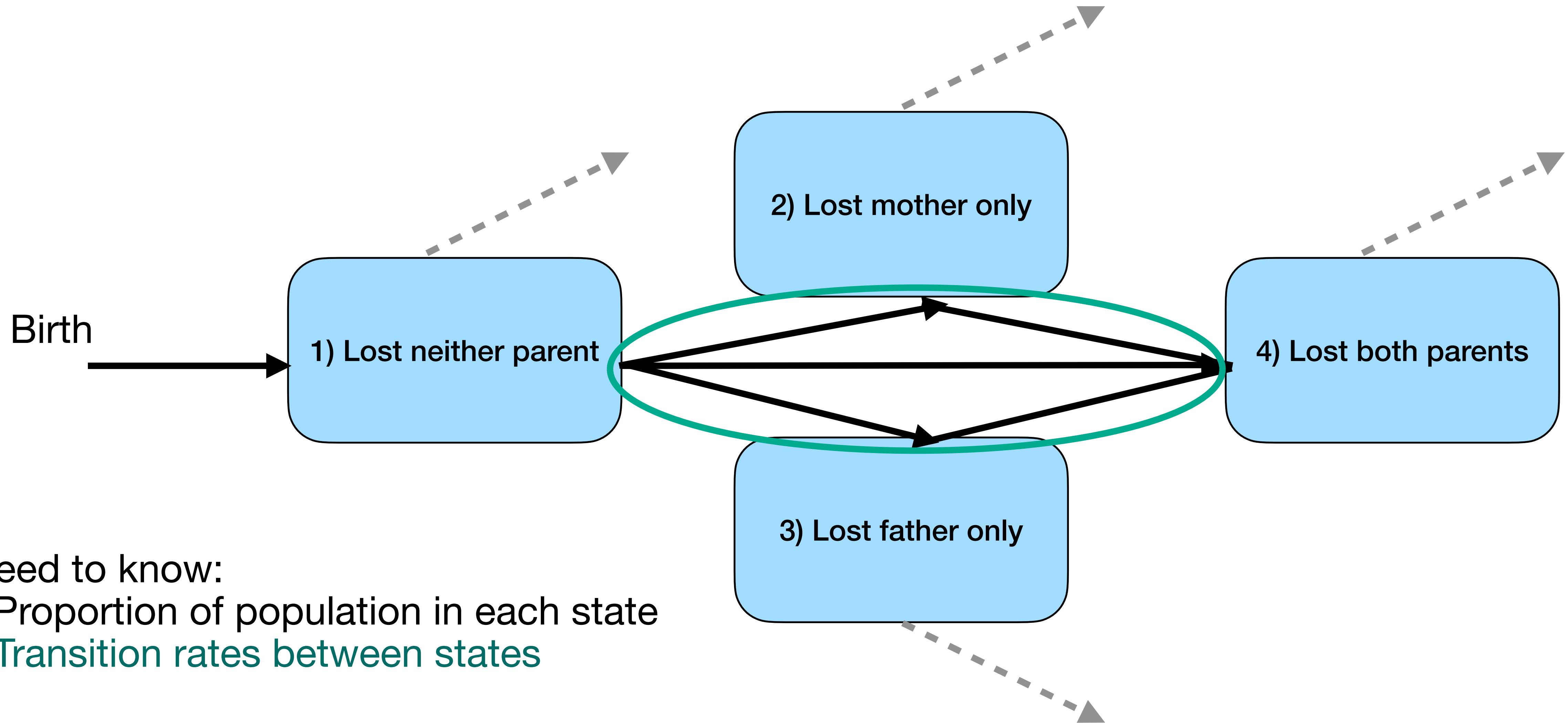
# States of parental loss



Need to know:

- Proportion of population in each state

# States of parental loss



Need to know:

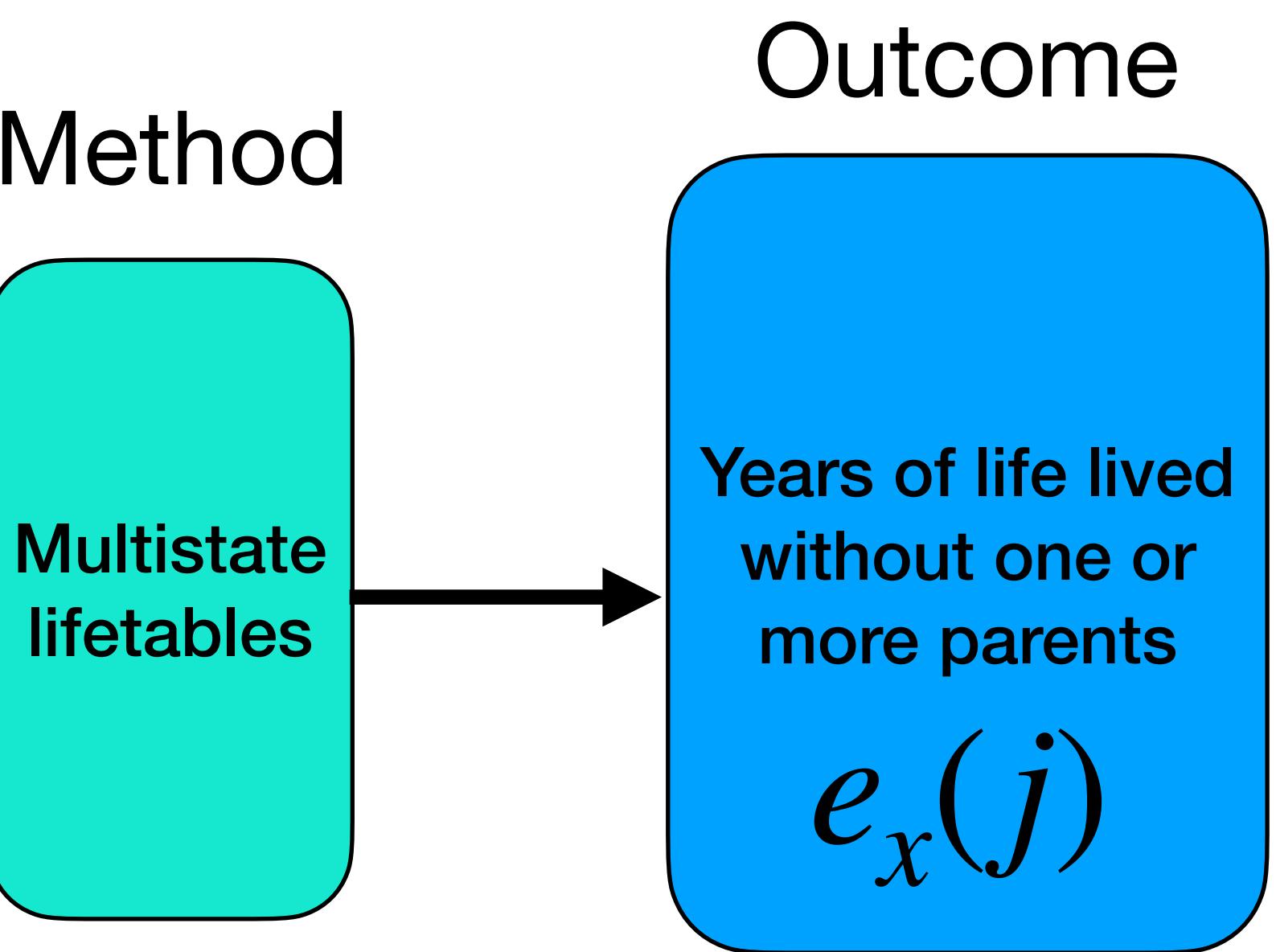
- Proportion of population in each state
- **Transition rates between states**

# What info do we need?

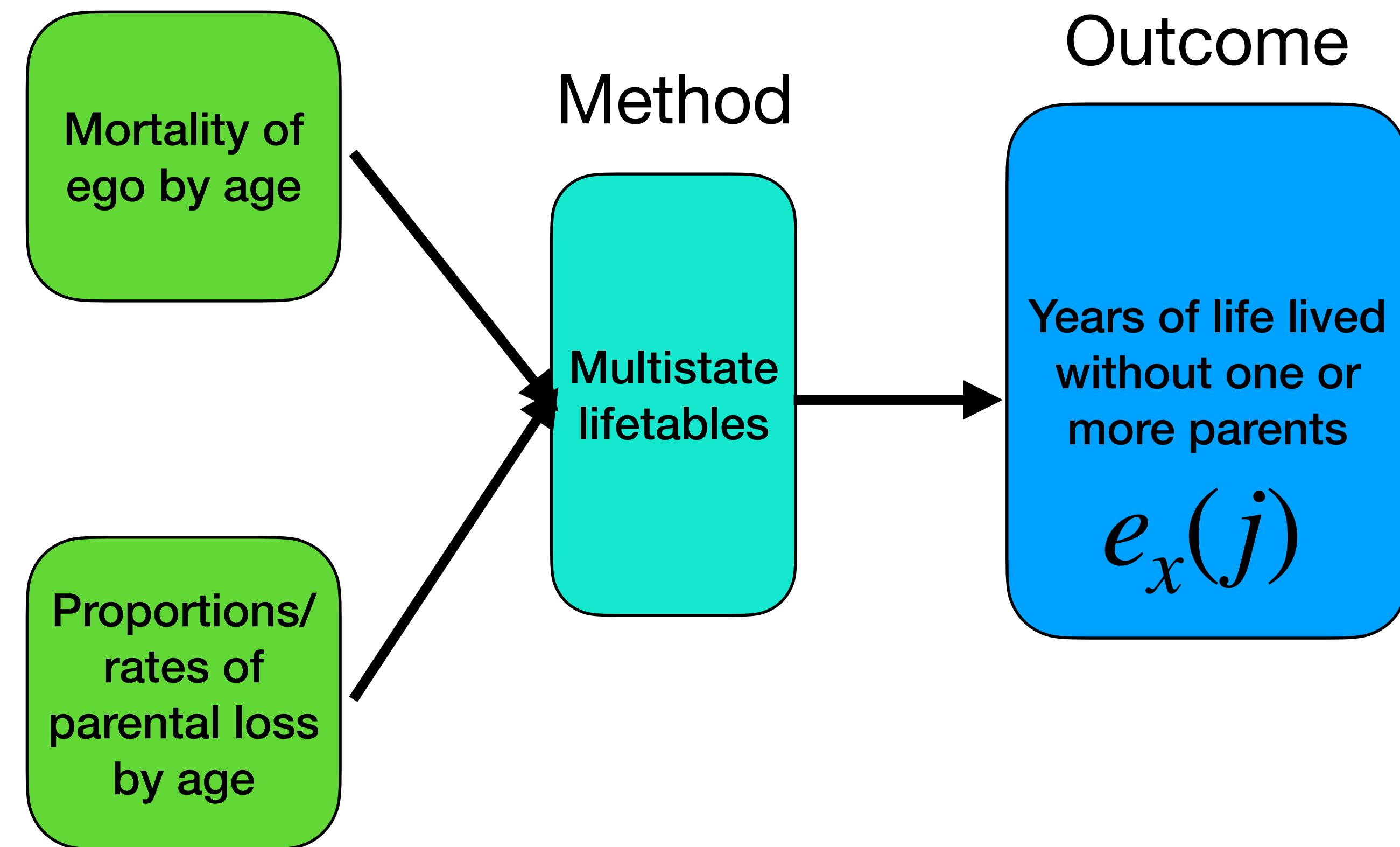
Outcome

Years of life lived  
without one or  
more parents

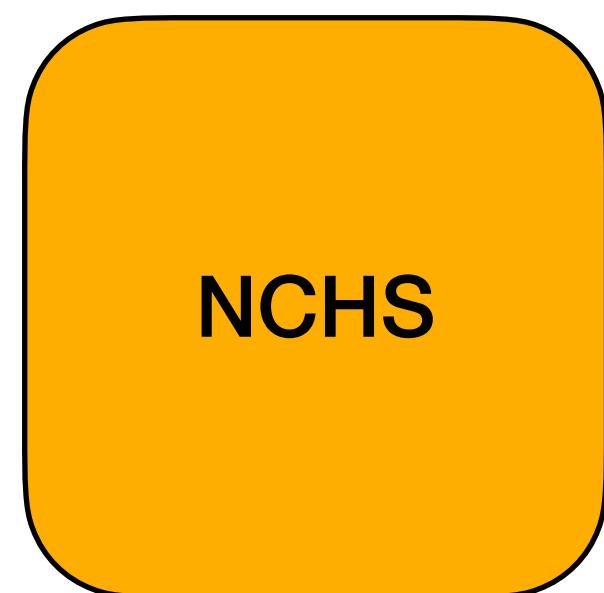
$e_x(j)$



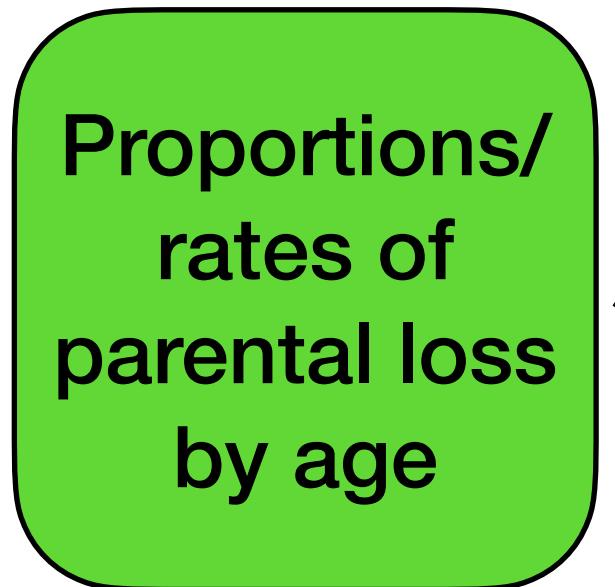
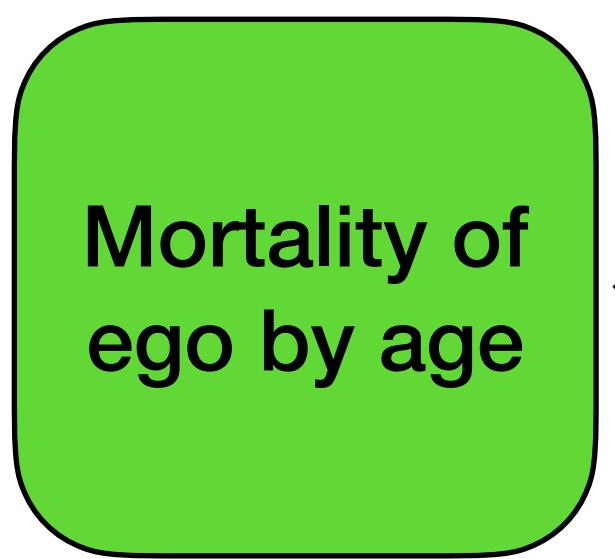
## Inputs required



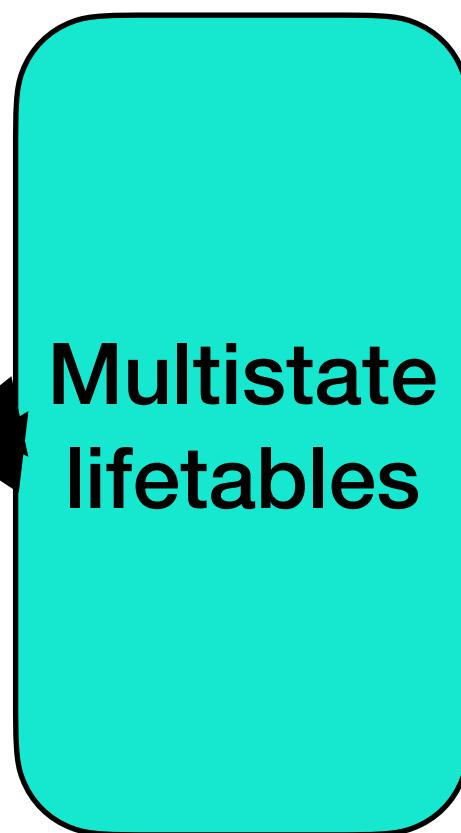
## Data Sources



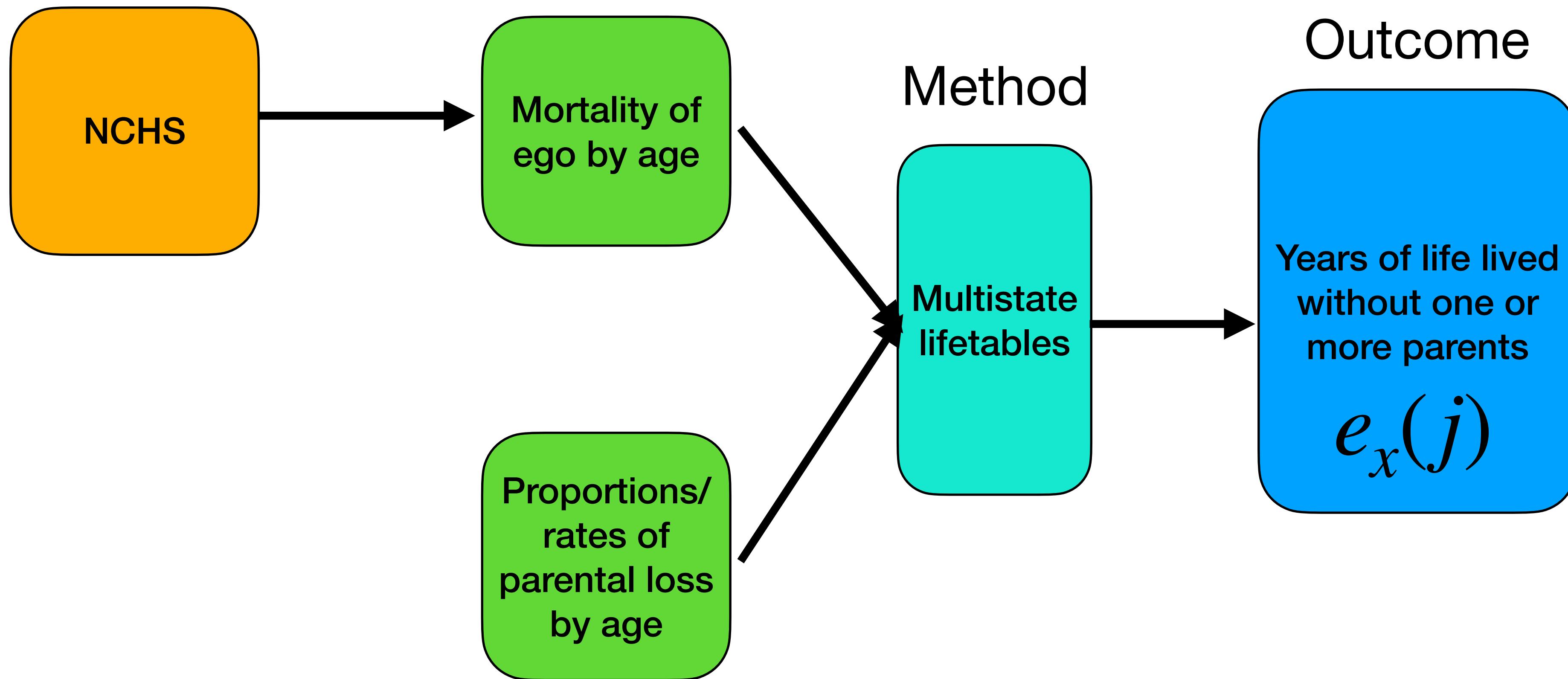
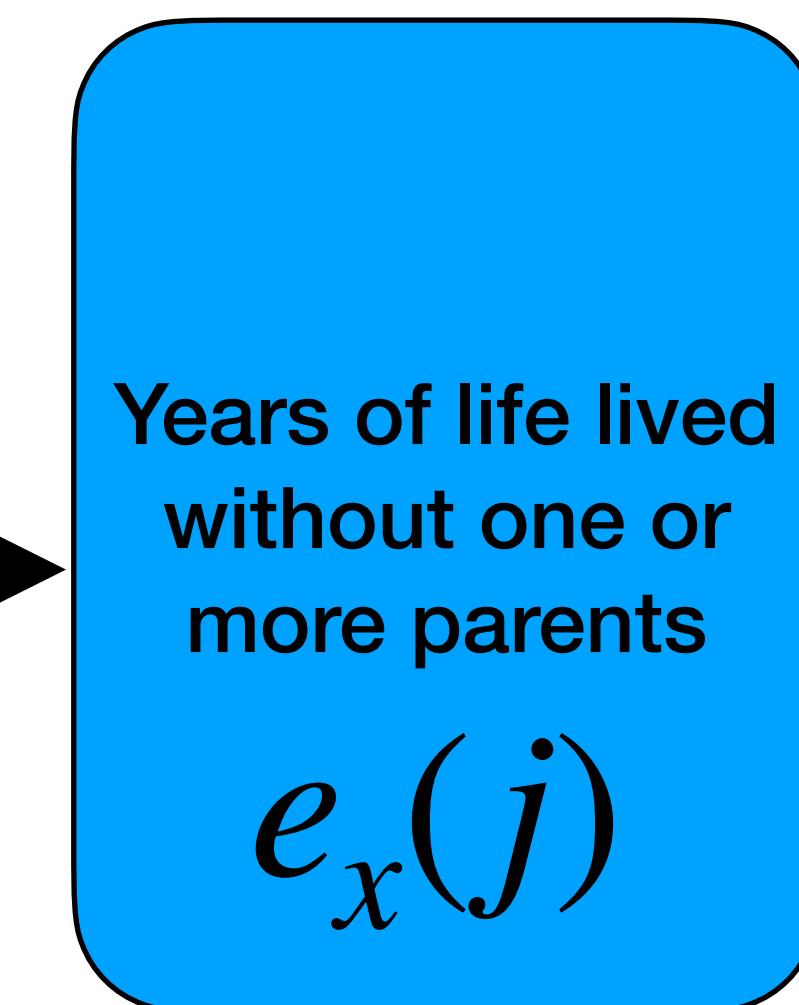
## Inputs required



## Method



## Outcome



## Data Sources

NCHS

Survey data

## Inputs required

Mortality of  
ego by age

Proportions/  
rates of  
parental loss  
by age

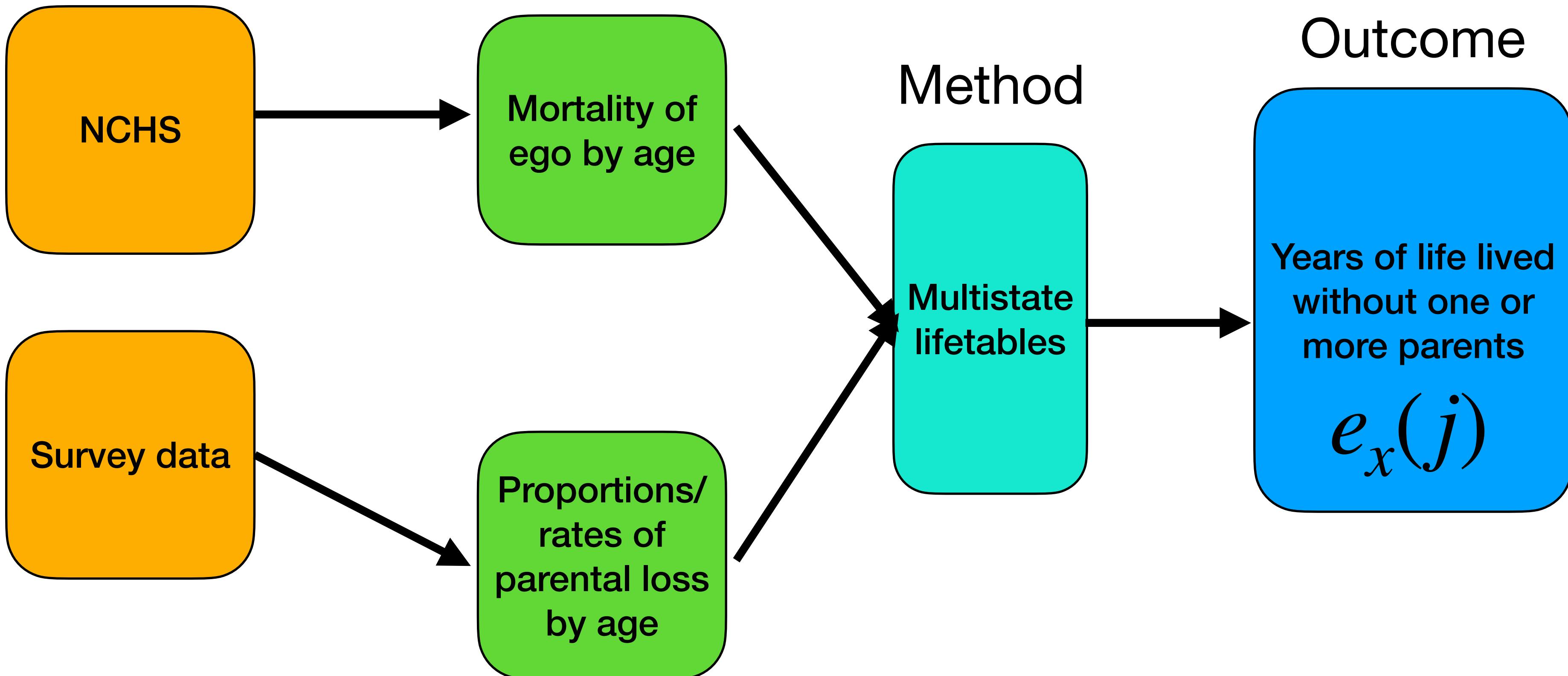
## Method

Multistate  
lifetables

## Outcome

Years of life lived  
without one or  
more parents

$e_x(j)$



# Survey data on parental loss

- Survey of Income and Program Participation (2021)
- Run by US Census Bureau. Nationally representative longitudinal survey, on income, employment, household composition, and includes the family context of individuals
- Survey asked whether parents were alive or deceased as well as the participant's age at the time of losing each parent
- Sample size: 43,400 respondents (30,700 NHW; 8,300 Hispanic; 4,400 NHB)
- But many (10-15%) age-specific rates are missing

## Data Sources

NCHS

Survey data

## Inputs required

Mortality of  
ego by age

Proportions/  
rates of  
parental loss  
by age

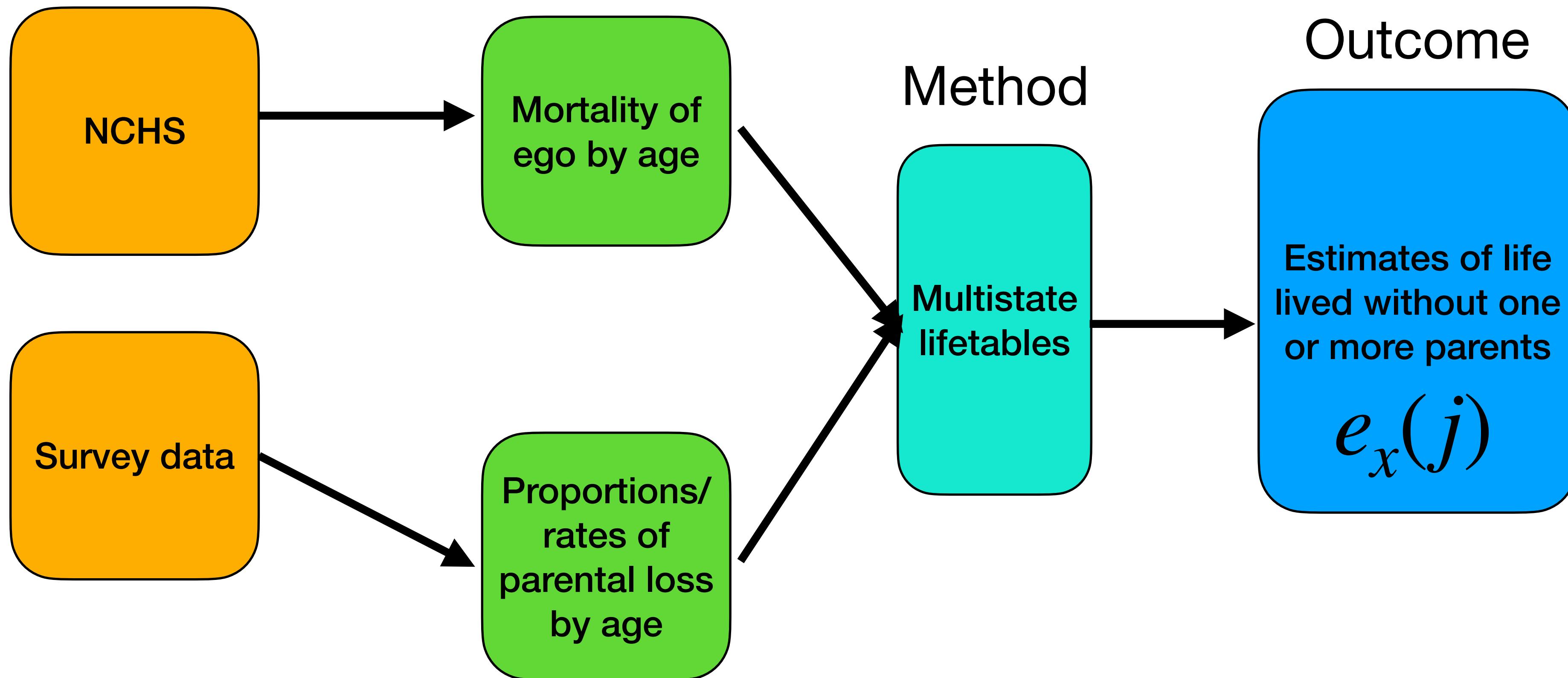
## Method

Multistate  
lifetables

## Outcome

Estimates of life  
lived without one  
or more parents

$e_x(j)$



## Data Sources

NCHS

Survey data

Projection  
model  
estimates

## Inputs required

Mortality of  
ego by age

Proportions/  
rates of  
parental loss  
by age

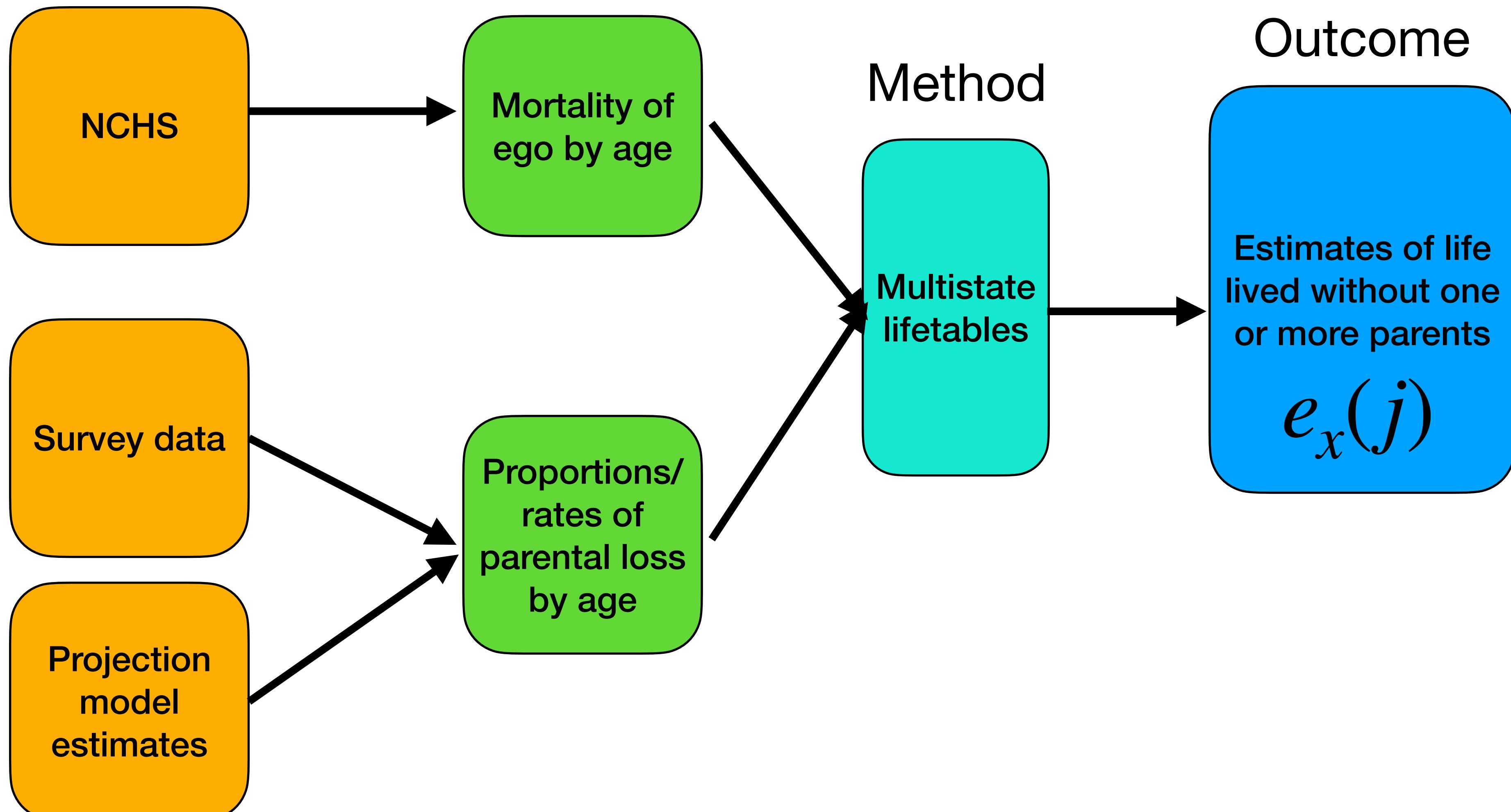
## Method

Multistate  
lifetables

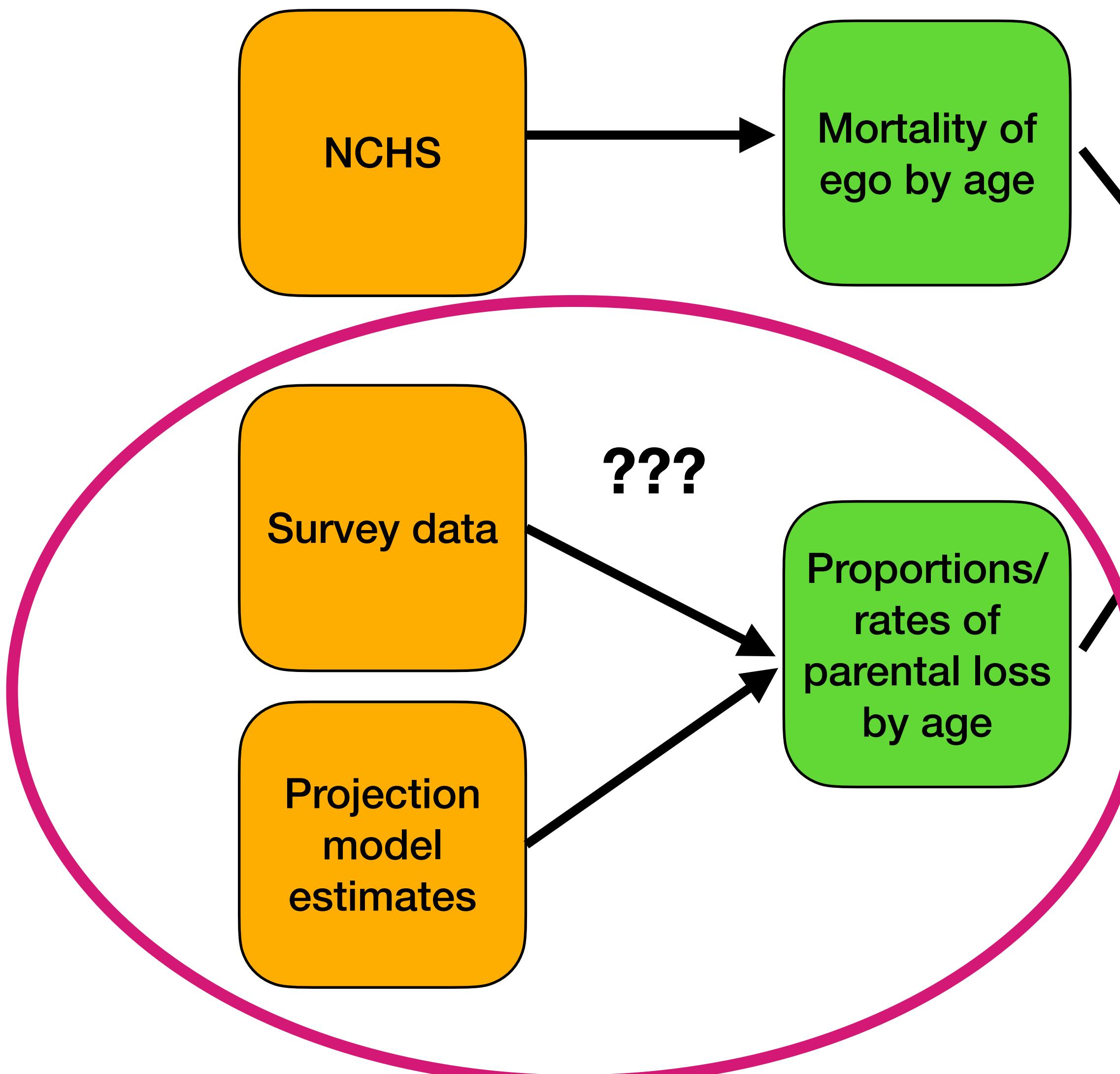
## Outcome

Estimates of life  
lived without one  
or more parents

$e_x(j)$



## Data Sources

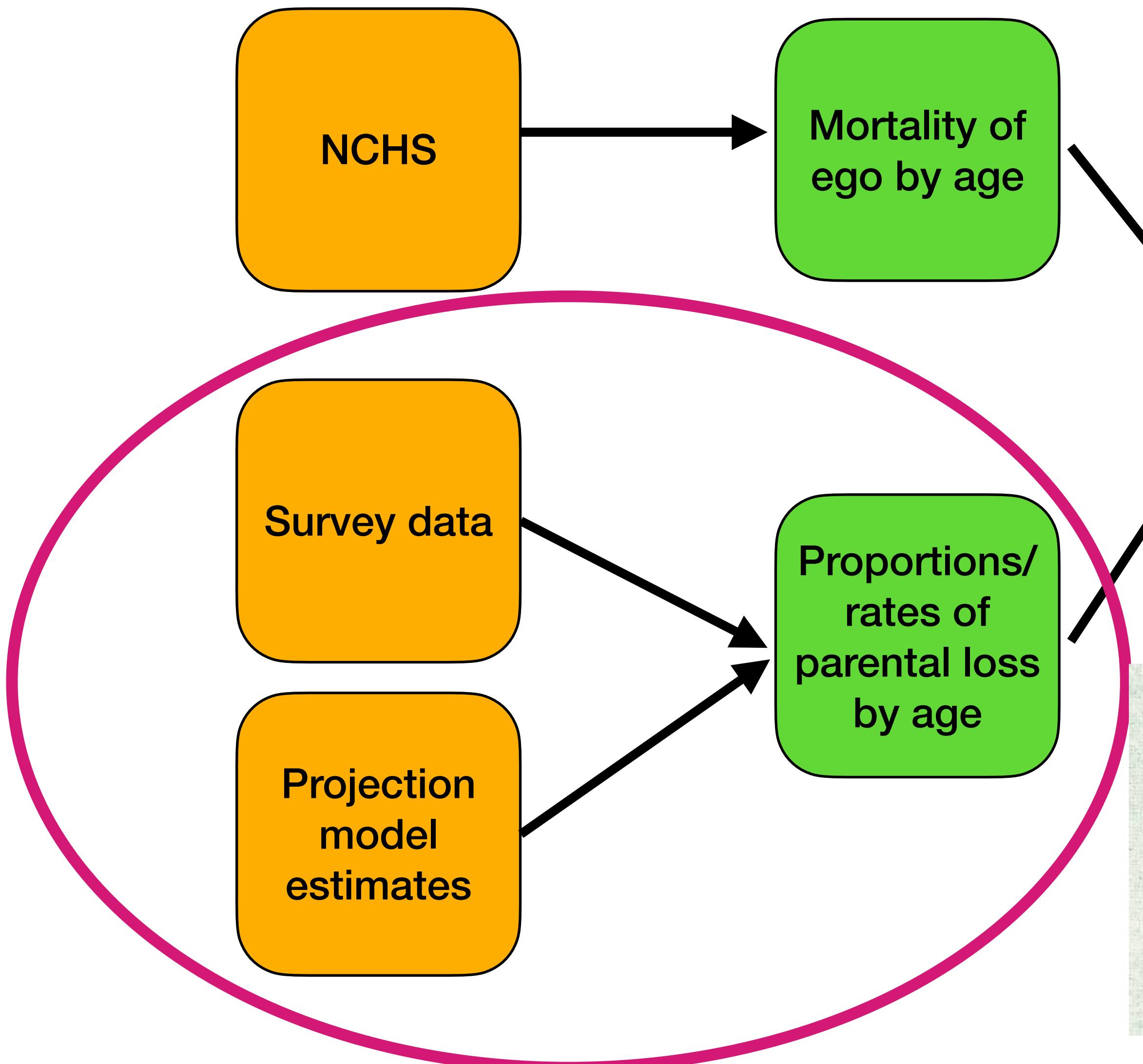


## Inputs required

## Method

## Outcome

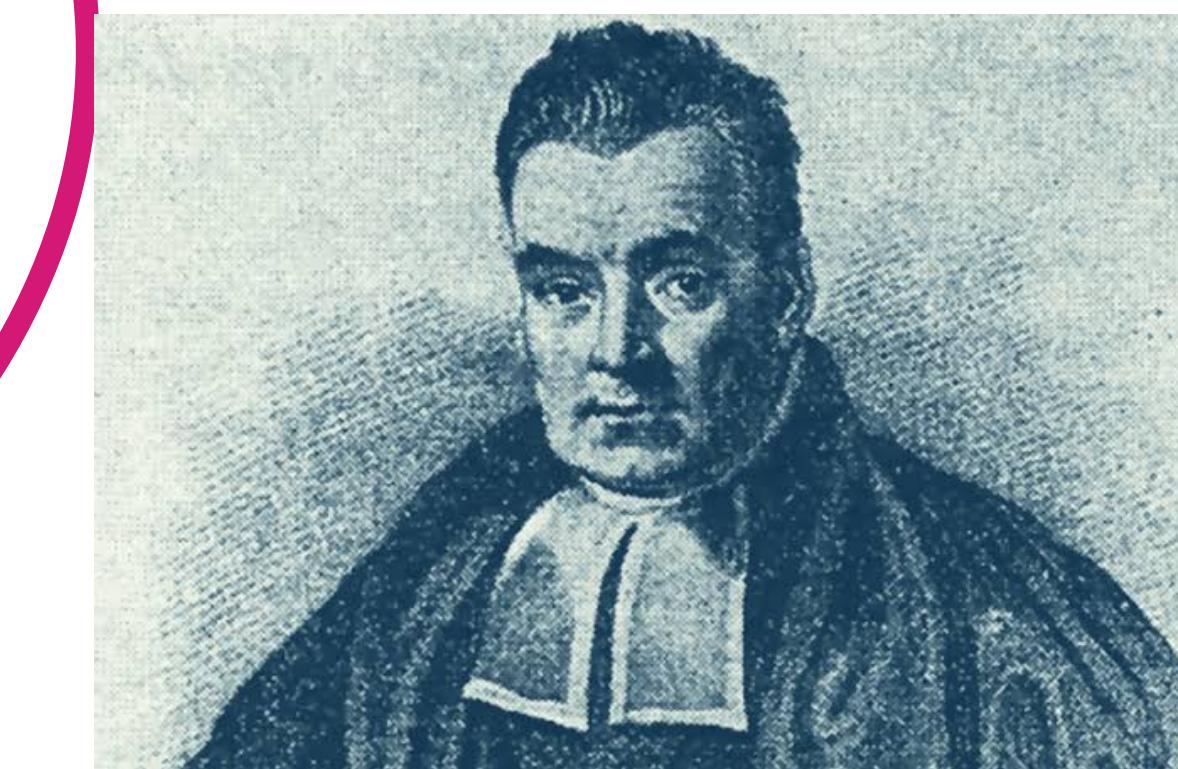
## Data Sources



## Inputs required

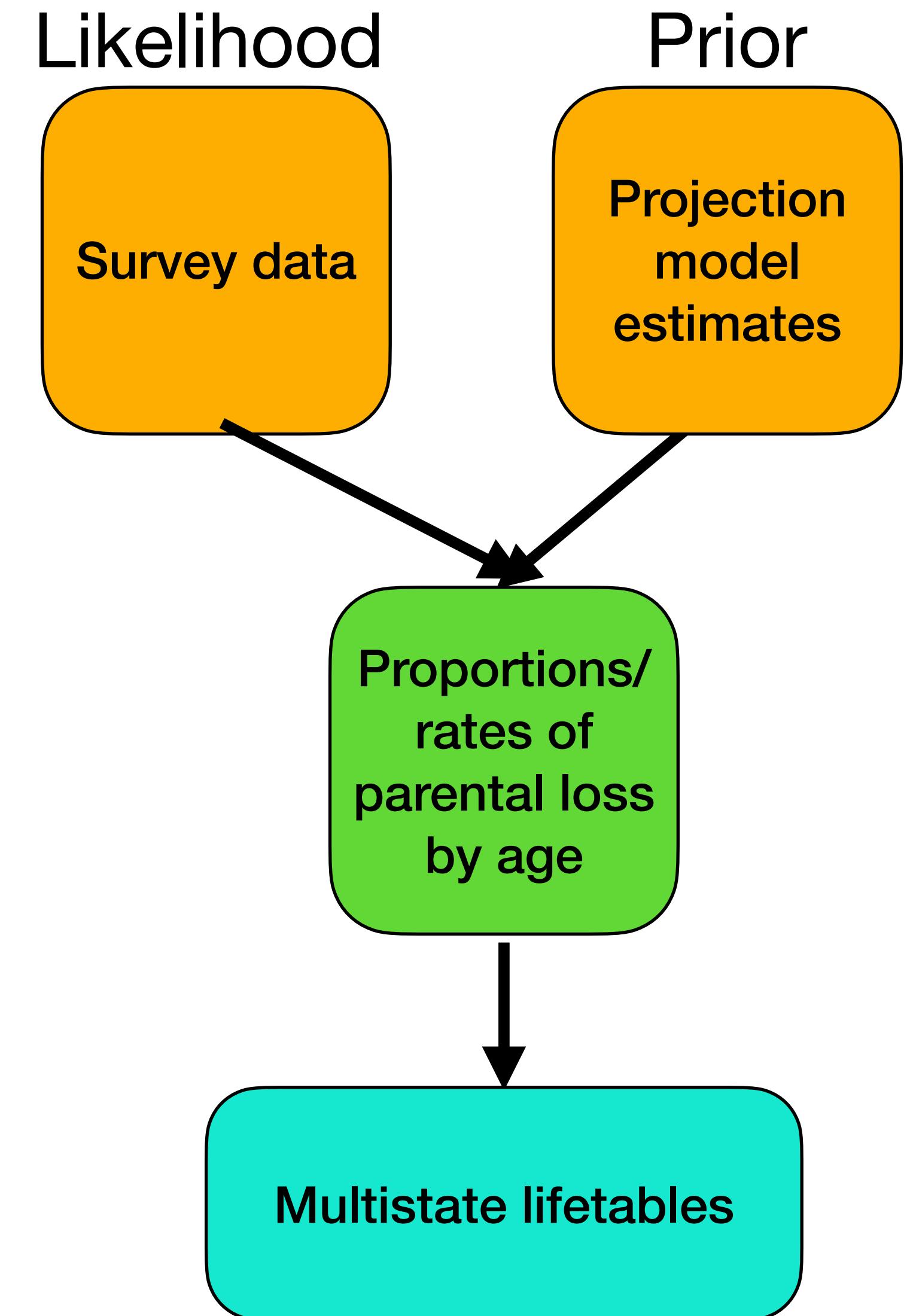
## Method

## Outcome



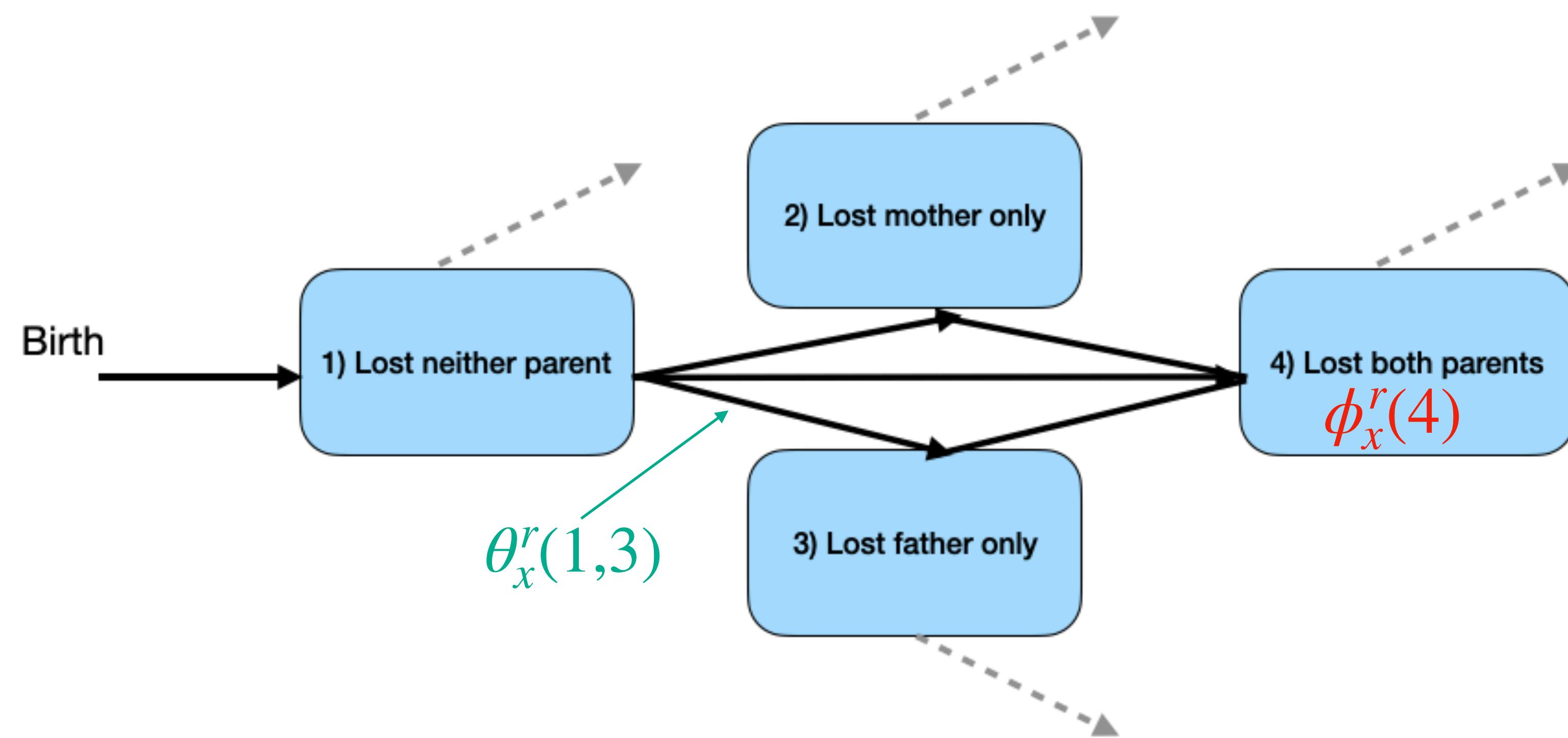
# Bayesian estimation of the parental loss proportions

- In this case, we have some idea of the proportions in each parental loss state from the demographic projection model, even before we ‘see’ survey data (**prior**)
- Survey data is then put into the model through the **likelihood**
- The resulting **posterior** distribution of the proportions is informed by both sources
- These estimates (and their uncertainty) are then used to generate multistate lifetables (with uncertainty)



# Model overview (transition rates)

- Denote  $\theta_x^r(u, v)$  to be the **latent transition rate** at age  $x$ , race  $r$ , from state  $u$  to state  $v$
- Also denote  $\phi_x^r(l)$  as the proportion of the population in state  $l$



# Transition rates: likelihood

- The **observed transition rates** from the SIPP survey are denoted  $m_x^r(u, v)$  and assumed to follow a Normal likelihood

$$m_x^r(u, v) | \theta_x^r(u, v) \sim N\left(\theta_x^r(u, v), \left(SE(m_x^r(u, v))\right)^2\right)$$

- Where  $SE(m_x^r(u, v))$  are the sampling standard errors (calculated taking sampling design into account)

# Transition rates: prior

- From the projection model we only know about the marginal transition rates. Denote these  $\hat{m}_x^r(\cdot, v)$  (for  $v = 2, 3$ )
- Can express the latent marginal transitions as a function of  $\theta_x^r$  and  $\phi_x^r$ . e.g. for losing a mother:

$$\theta_x^r(\cdot, 2) = \frac{(\theta_x^r(1,2) + \theta_x^r(2,4)) \cdot \phi_x^r(1) + \theta_x^r(3,4) \cdot \phi_x^r(3)}{\phi_x^r(1) + \phi_x^r(3)}$$

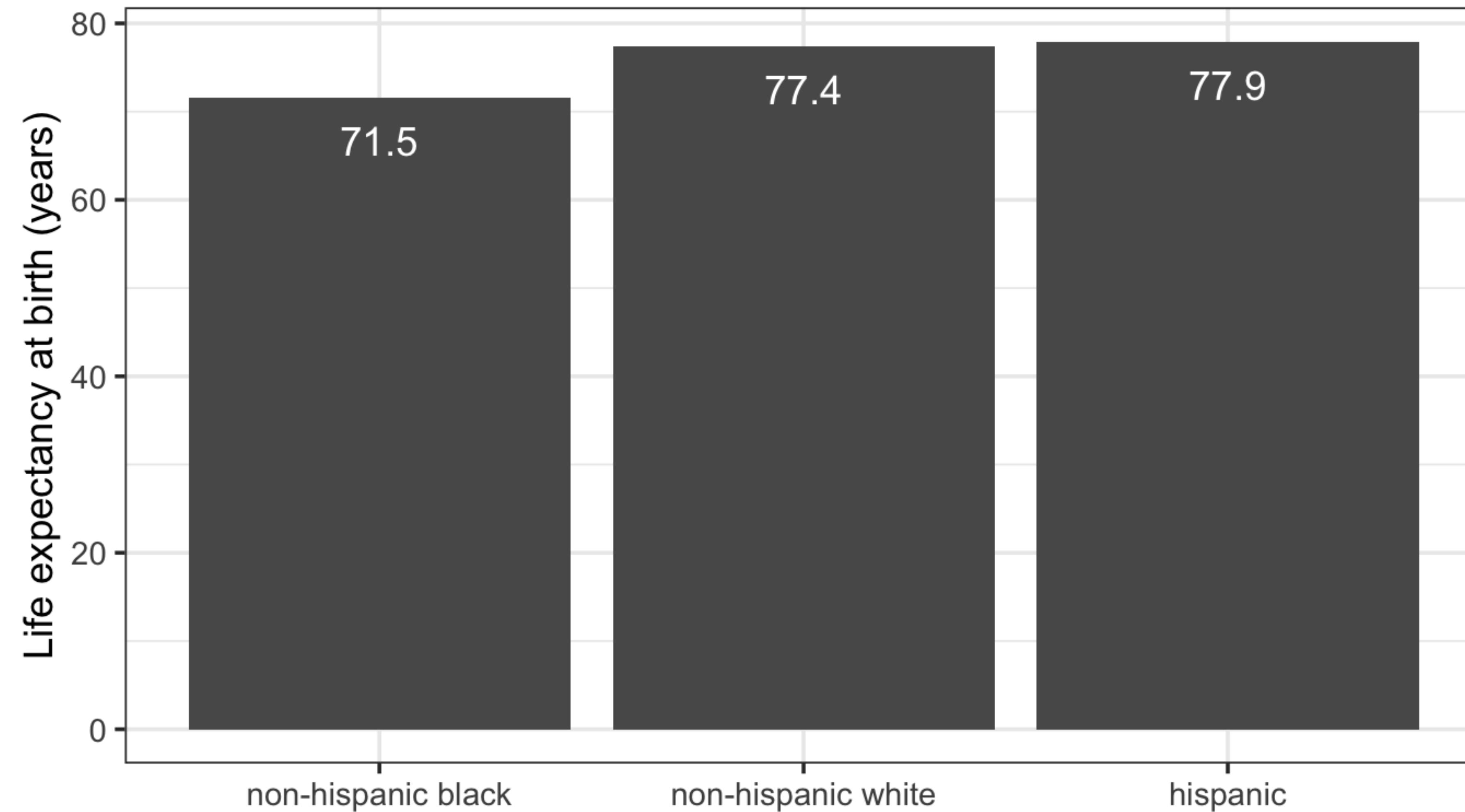
- The latent marginal rates then have prior distributions that are informed by the estimates from the projection model:

$$\theta_x^r(\cdot, v) | \sigma_v^2 \sim N(\hat{m}_x^r(\cdot, v), \sigma_v^2)$$

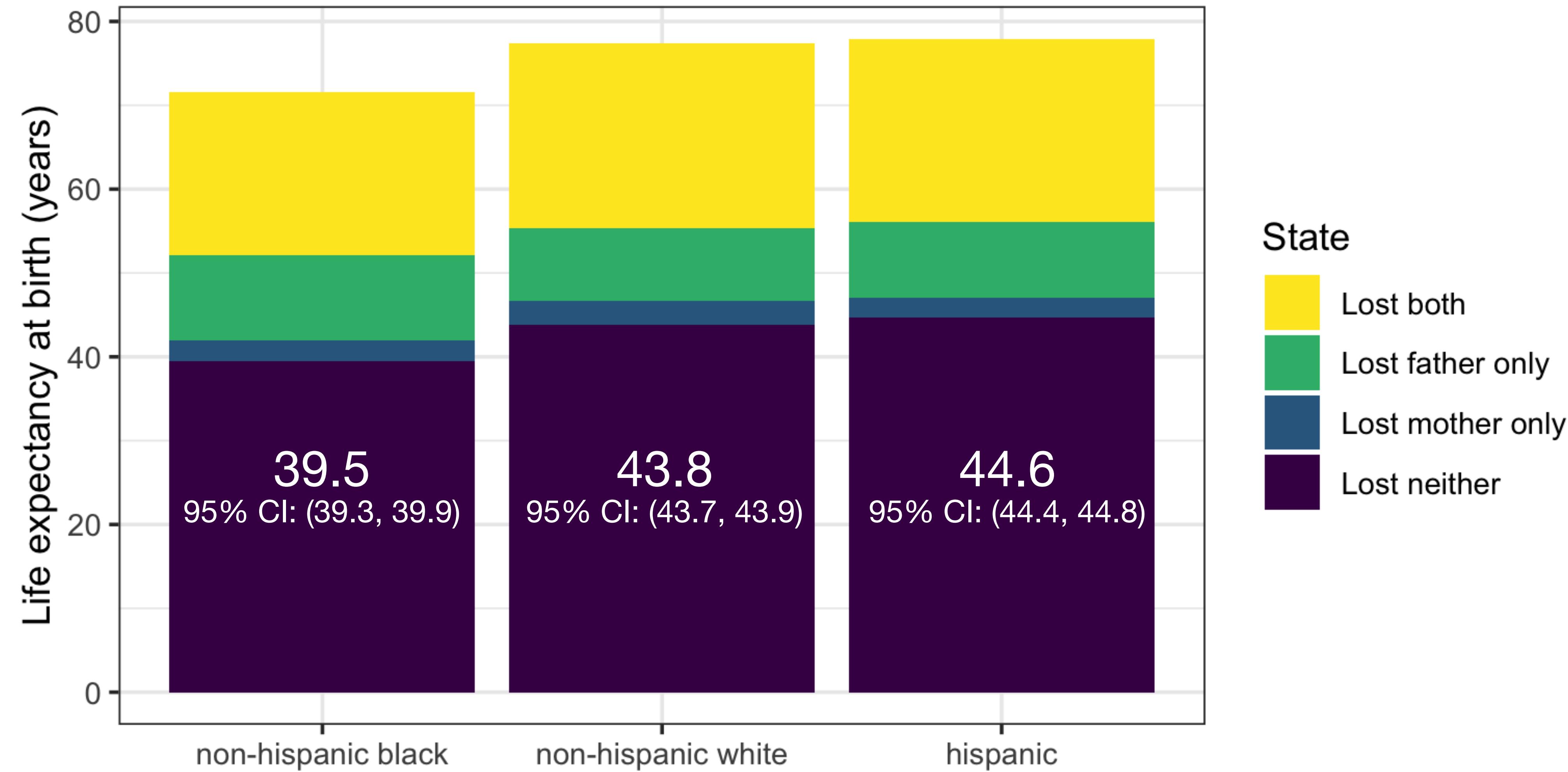
- The resulting posterior estimates for  $\theta_x^r$  are thus informed by both survey data and the projection model

# Results

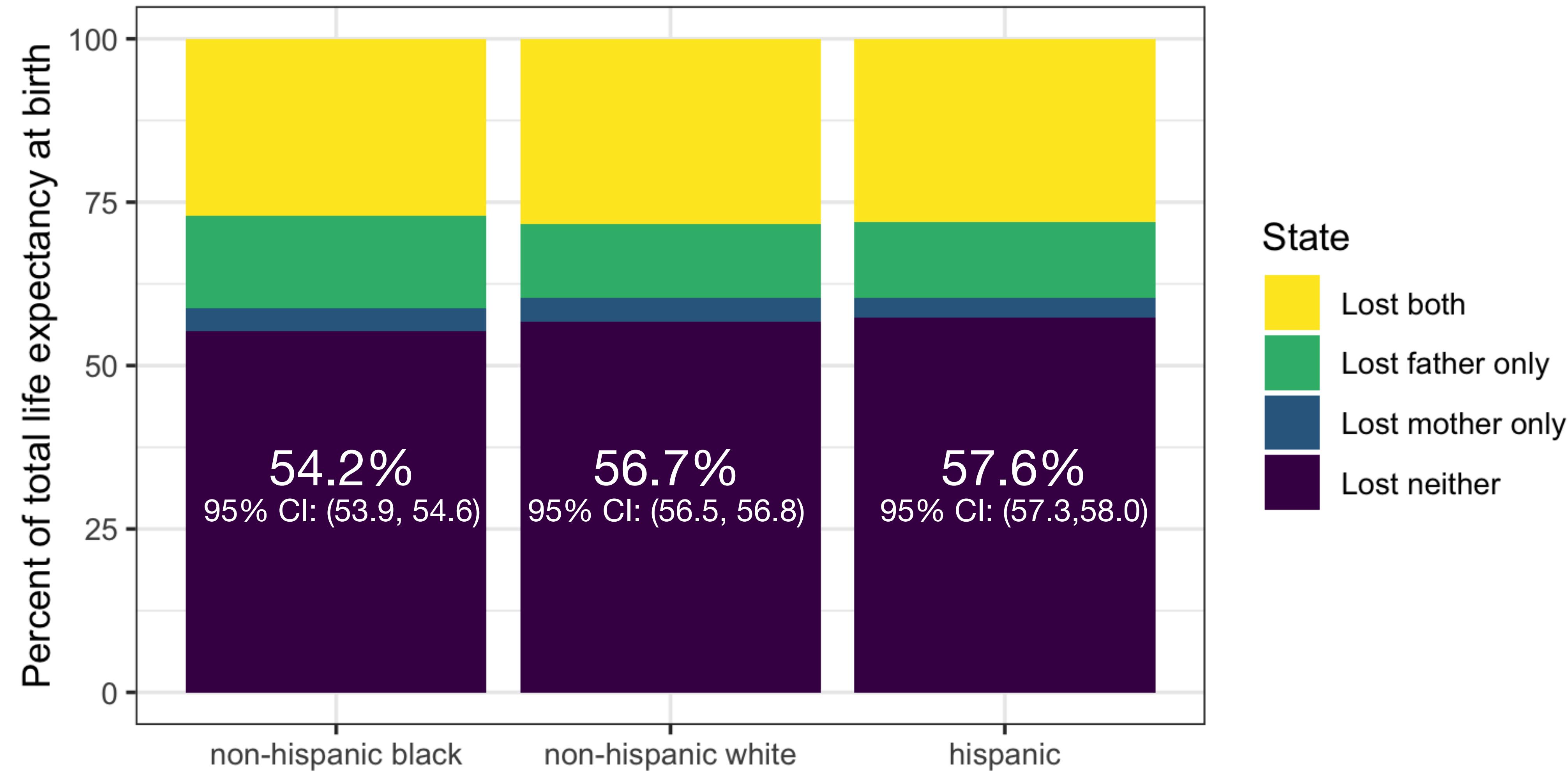
# Life expectancy at birth



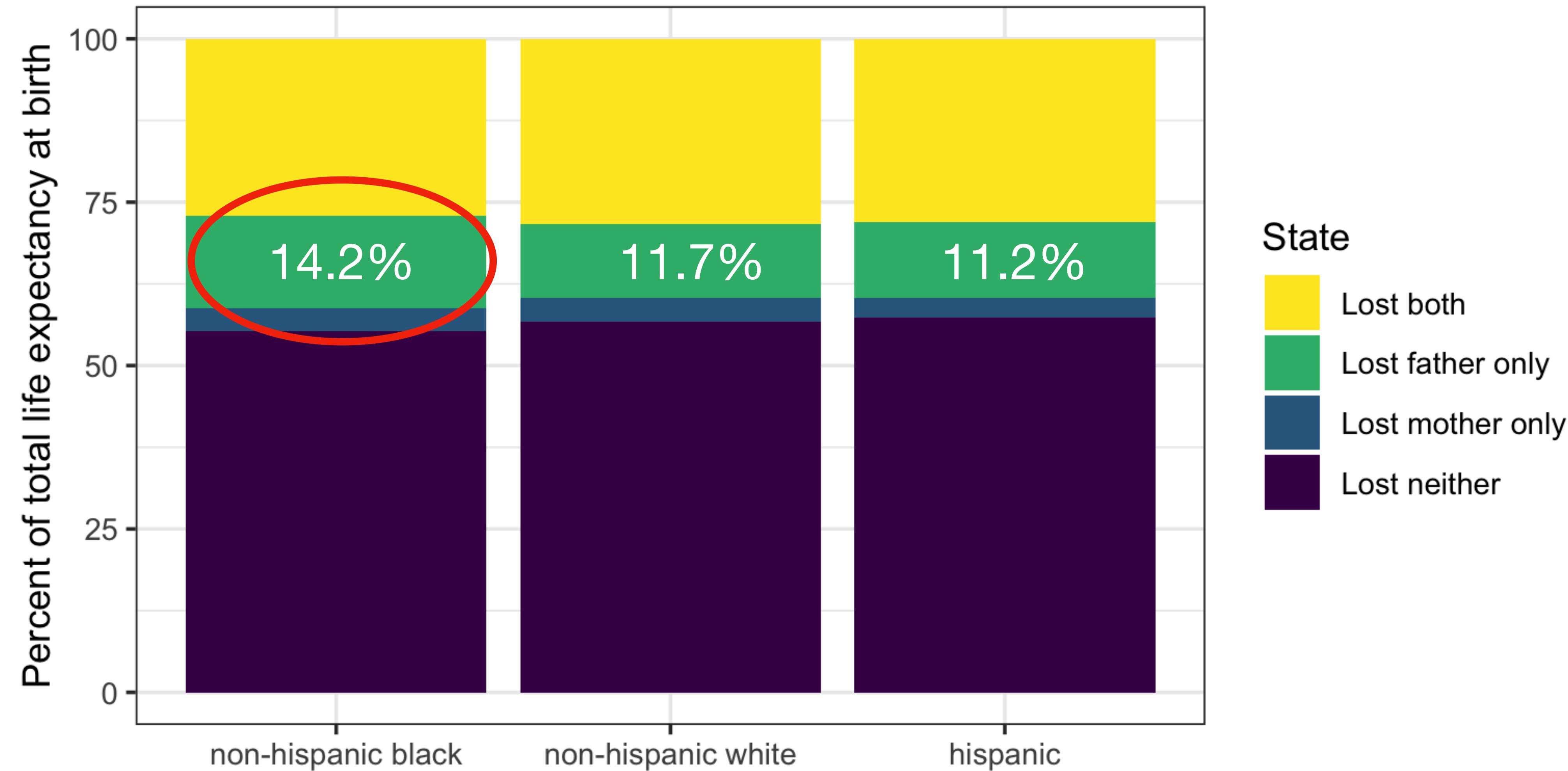
# By parental loss state



# Percent of total life expectancy

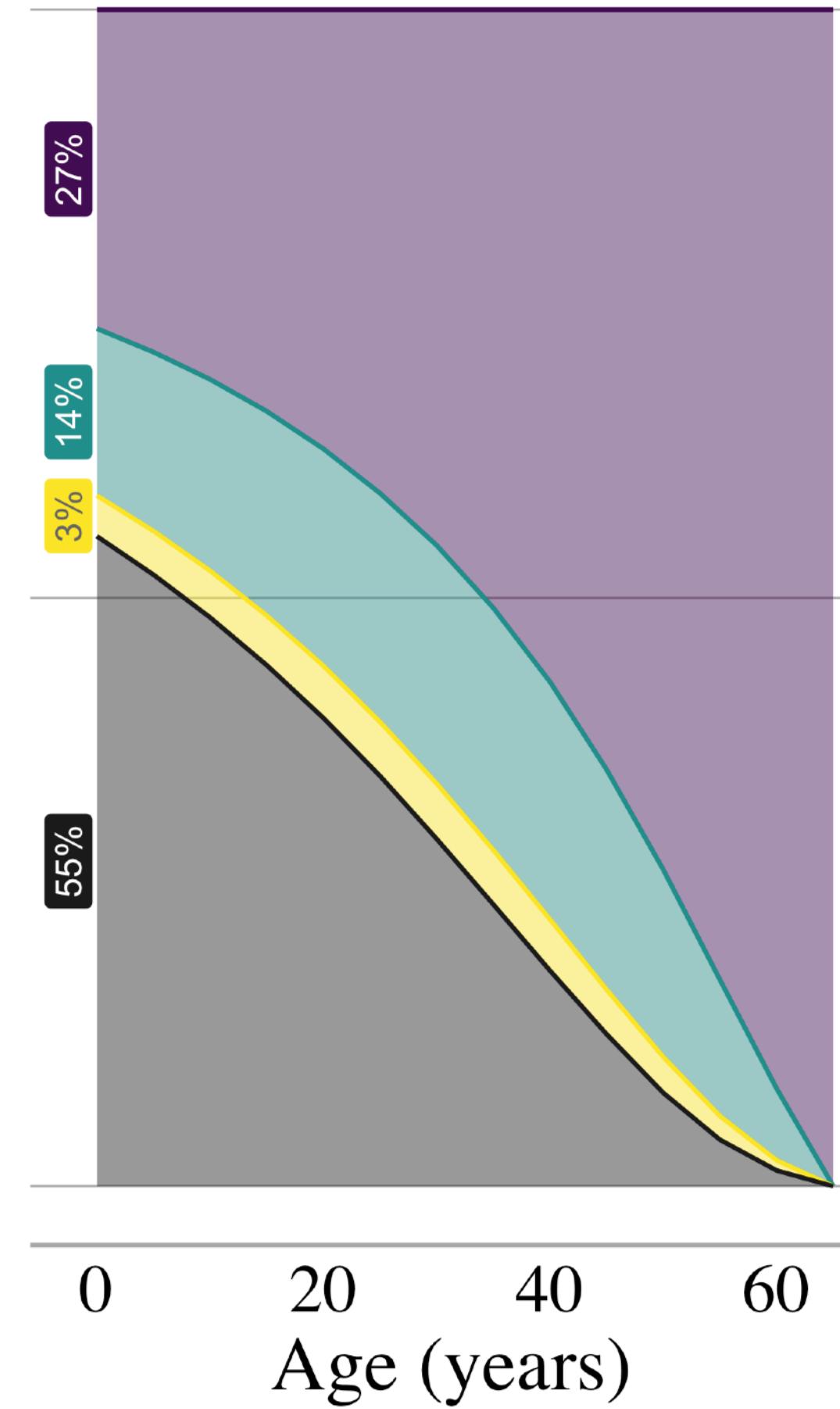


# Percent of total life expectancy

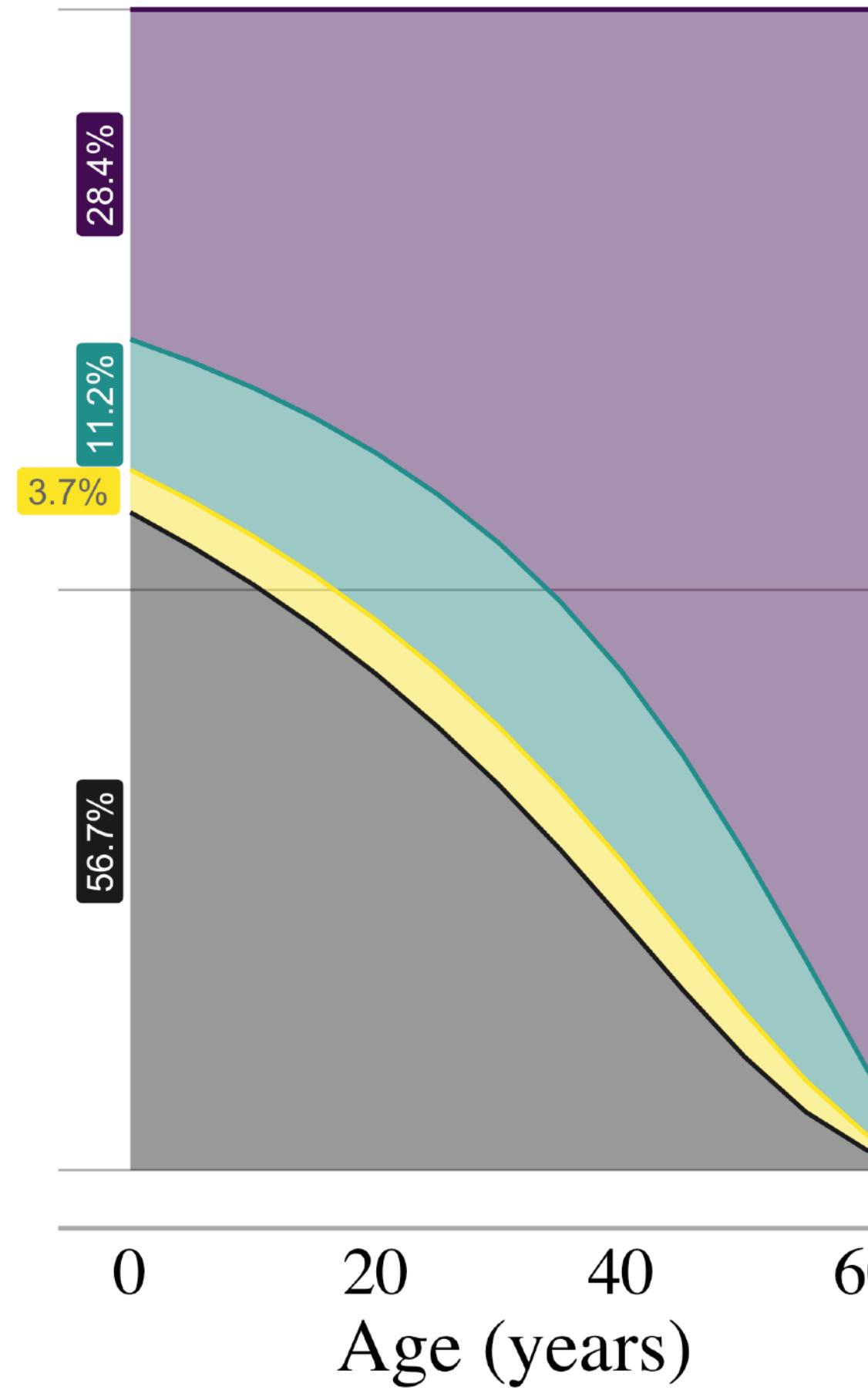


# Over the life course

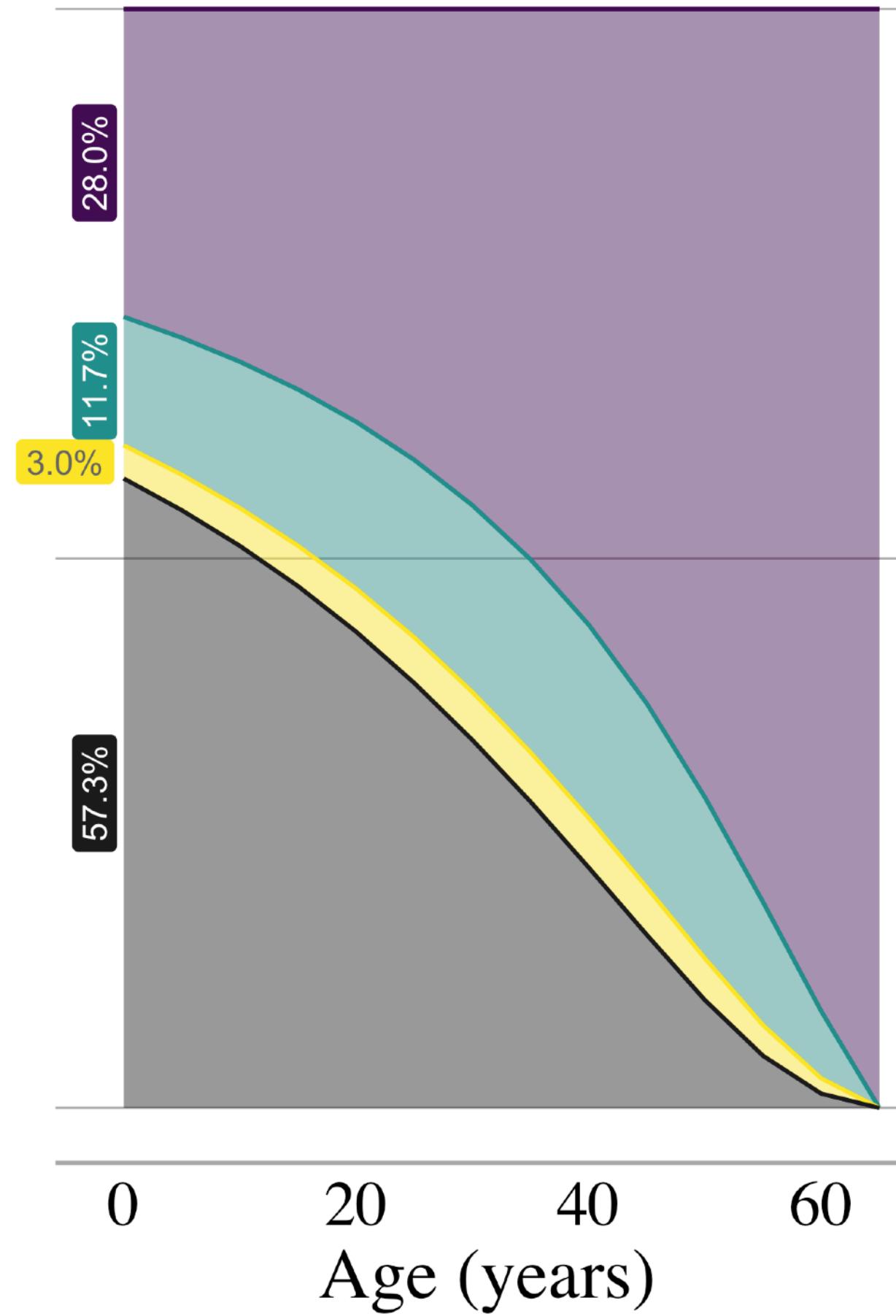
Proportion  
of expected life remaining



a) Non-Hispanic black



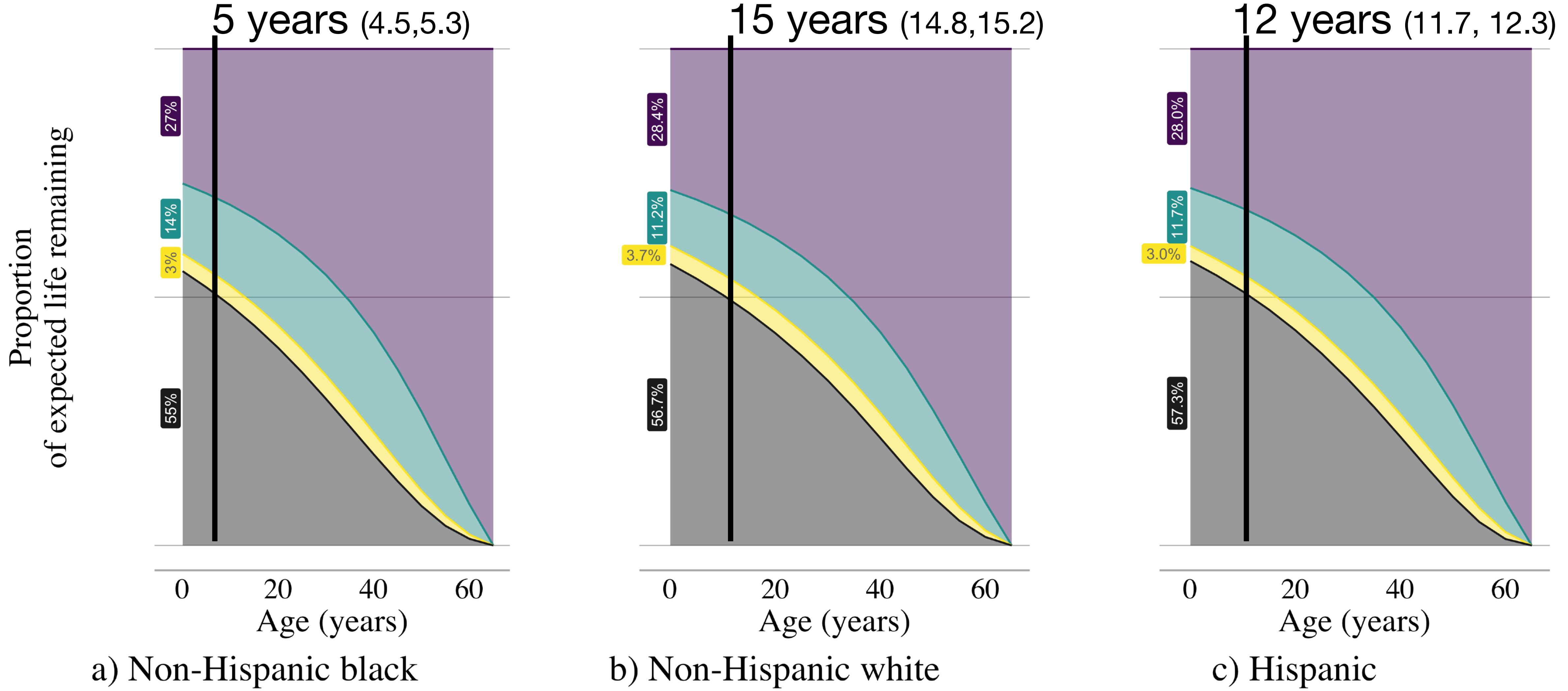
b) Non-Hispanic white



c) Hispanic

- Lost both
- Lost father only
- Lost mother only
- Lost none

# Age at which majority of life is without both parents



- Lost both
- Lost father only
- Lost mother only
- Lost none

# Takeaways

- At birth, can expect to live 42-46% of life without at least one parent
- Black children have less time with their parents but also are expected to lose a parent at a younger age
- Biggest difference across racial/ethnic lines is loss of father: Black children can expect to live 2 years longer without their father alive, even though overall life expectancy is ~6 years shorter
- Implications for life course outcomes and transitions
- Different consequences based on timing in life course: early loss versus sandwich generation

# Summary and future work

# Summary

- Demographic processes link the individual to the aggregate, and back again
- Transitions in demographic rates have implications for the availability of different types of kin, which impacts individual level outcomes
- Patterns of parental loss in the United States are not constant
- Increase in premature parental loss due to traumatic causes
- Loss of a parent is racially patterned in both magnitude and timing

**Estimates here help to quantify and reframe, but it's just one piece**

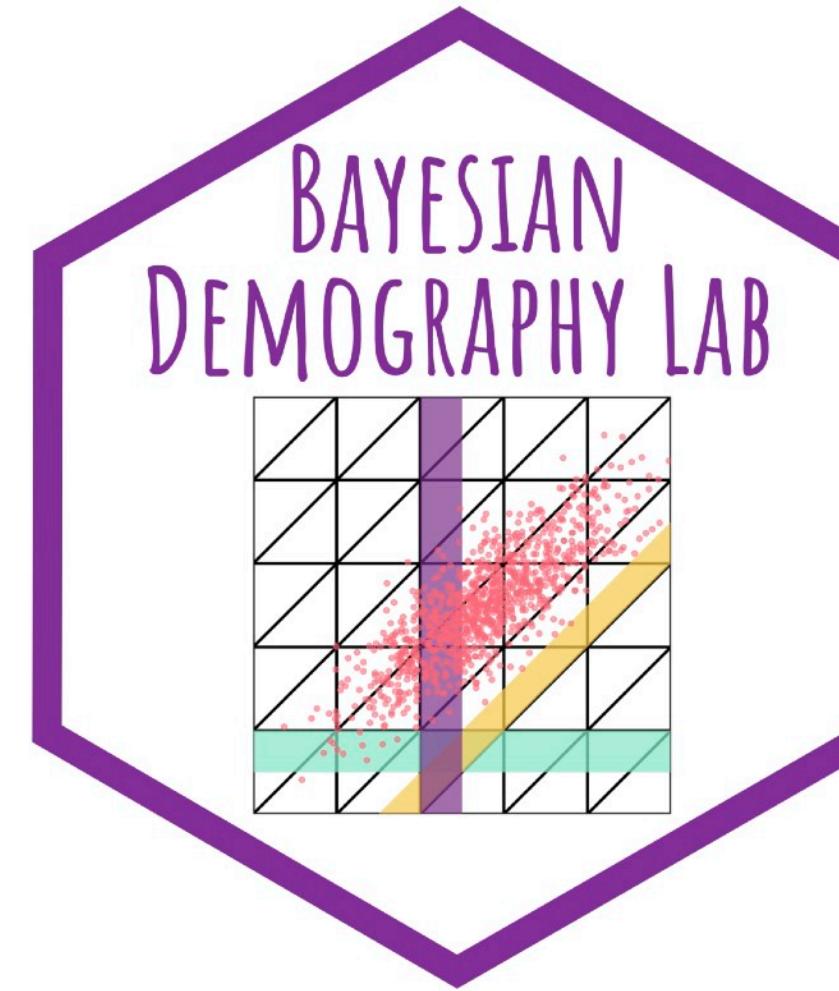
# Future work

- Extending methods to investigate heterogeneity in fertility and mortality
- Compare / integrate concepts of non-biological parents, one parent households, same-sex parents
- Estimation and understanding in low- and middle-income countries
- Extension of statistical models to more complex kin structures

# Team



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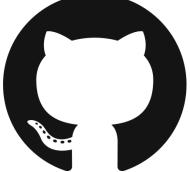
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# Thanks!

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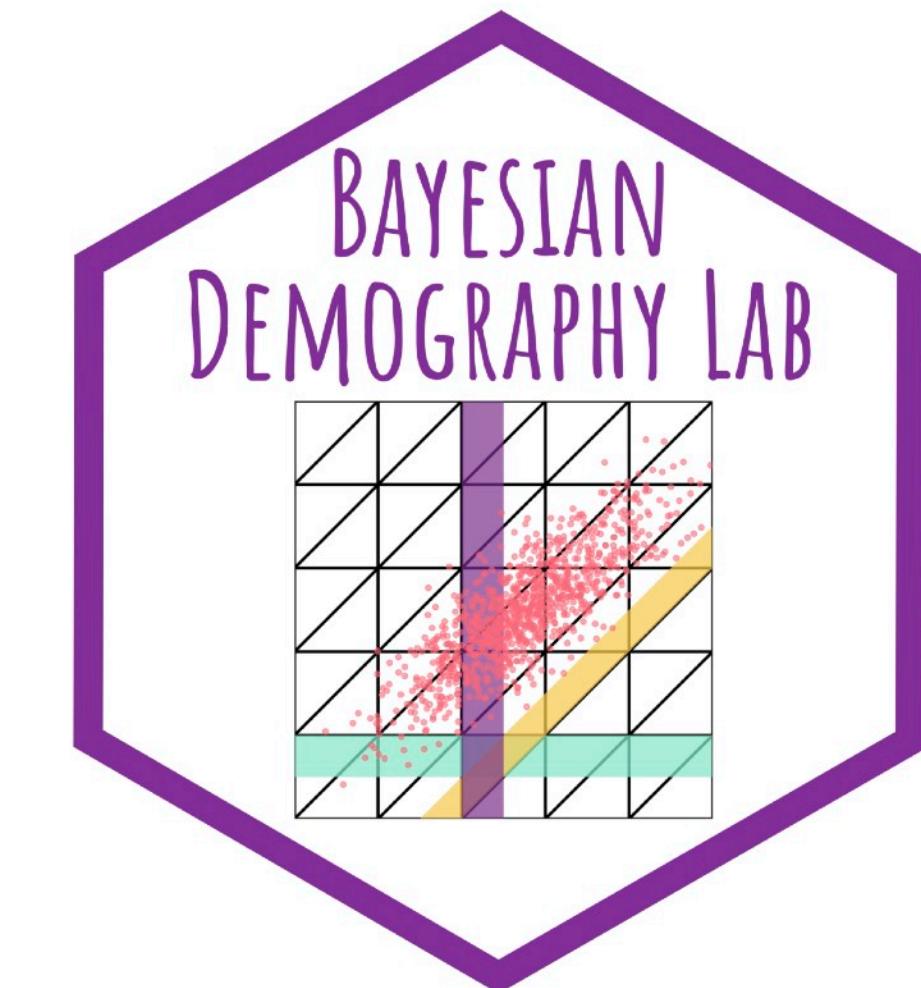
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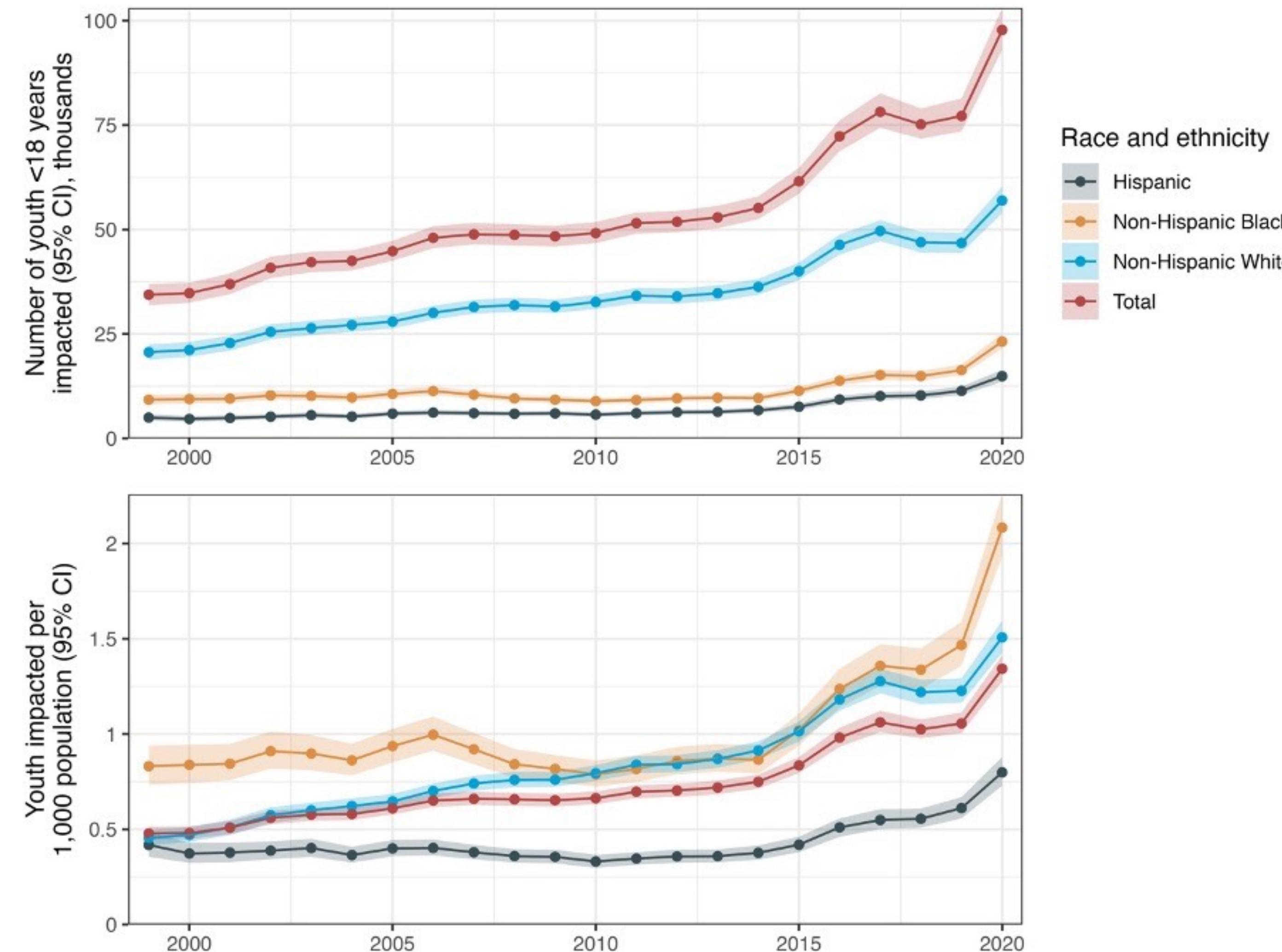
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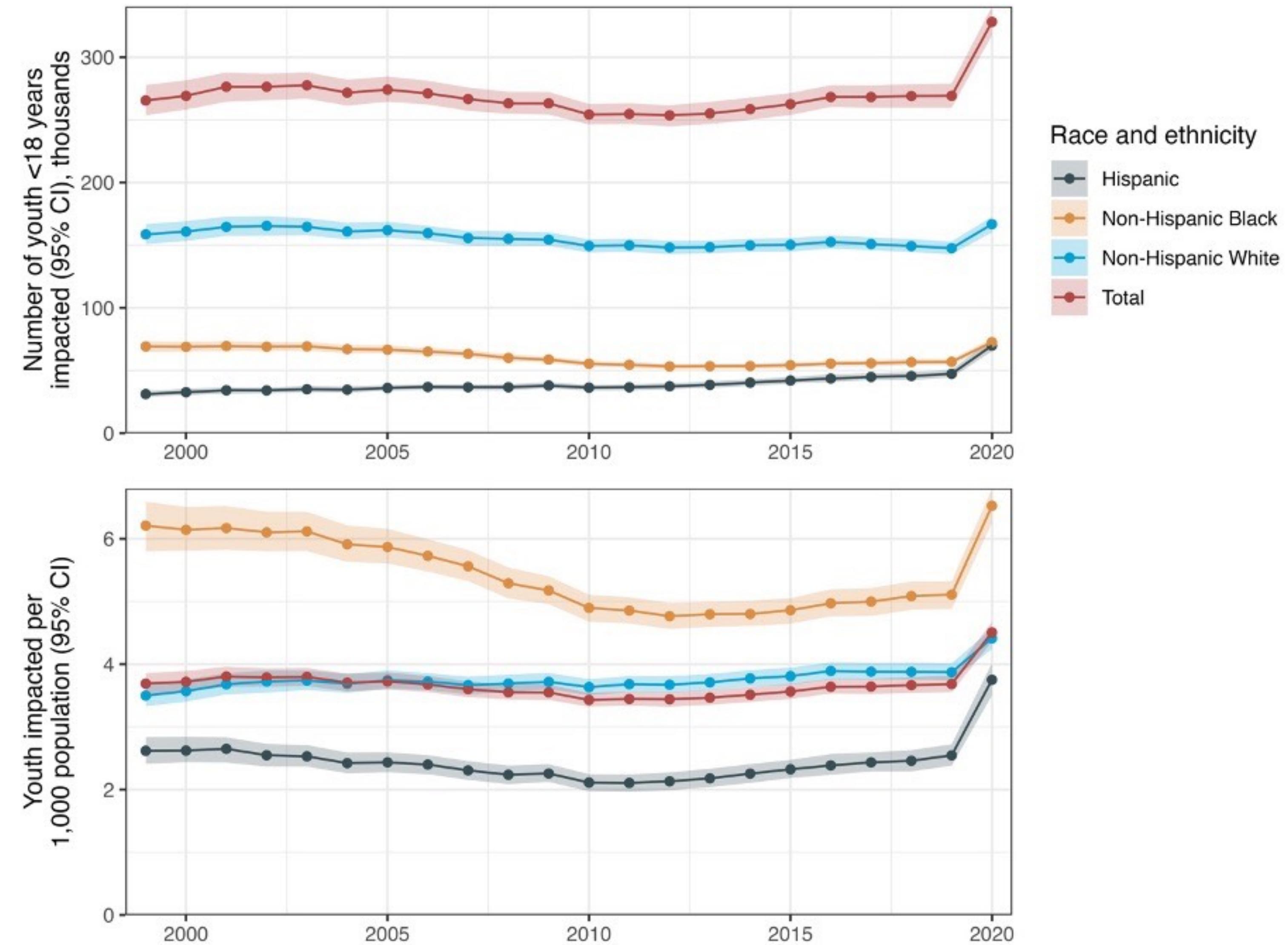
# Extra

# Project 1

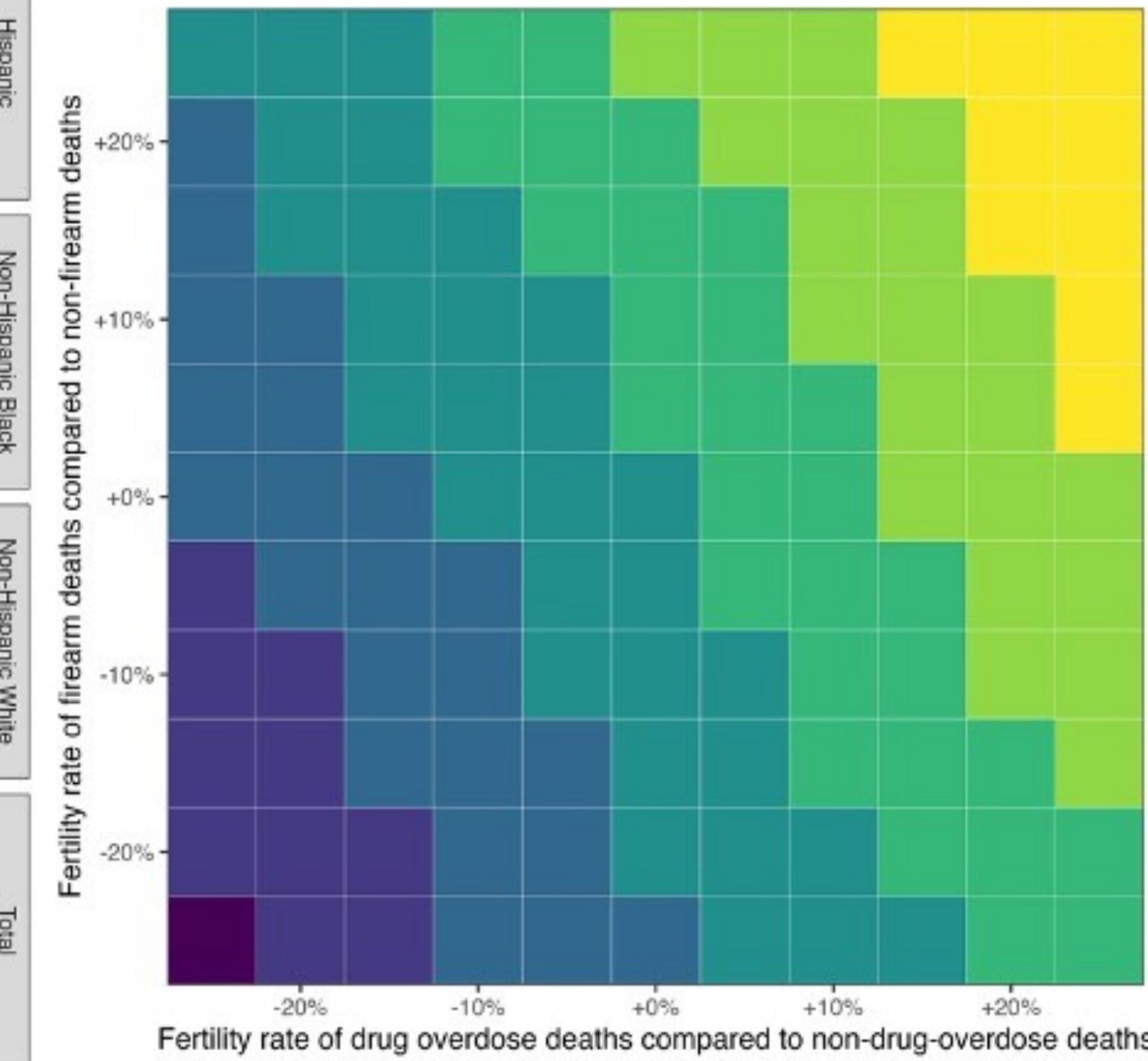
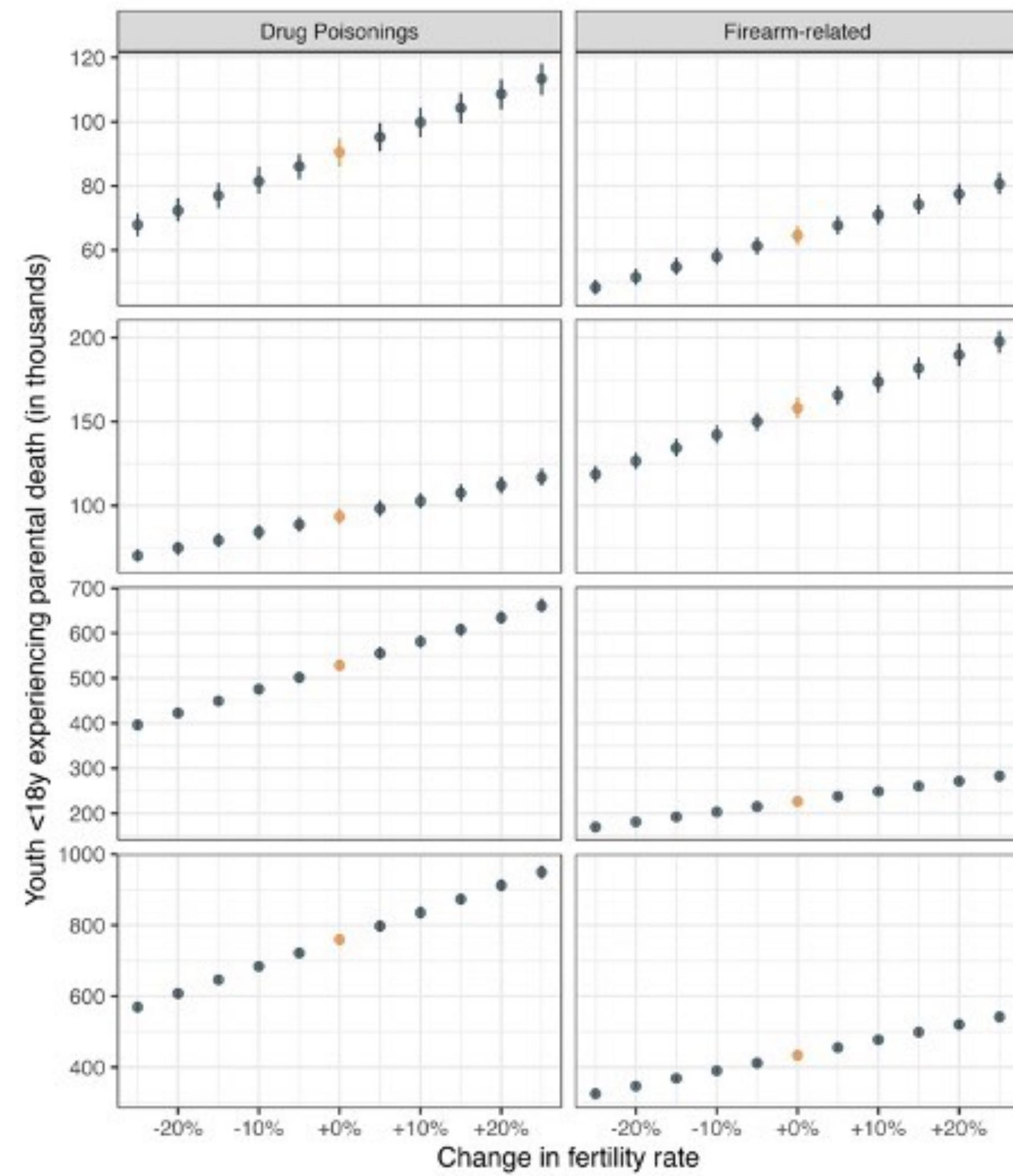
# Both causes combined



# All causes



# Sensitivity



◆ Sensitivity analysis    ■ Primary analysis

Youth <18y experiencing  
parental death (in thousands)

|     |      |      |      |      |      |
|-----|------|------|------|------|------|
| 900 | 1000 | 1100 | 1200 | 1300 | 1400 |
|-----|------|------|------|------|------|

# Expected number of living children

$$\begin{pmatrix} a^f \\ a^m \end{pmatrix}(x+1, t+1) = \begin{pmatrix} U_t^f & 0 \\ 0 & U_t^m \end{pmatrix} \begin{pmatrix} a^f \\ a^m \end{pmatrix}(x, t) + \begin{pmatrix} \bar{\alpha}F_t^f & \bar{\alpha}F_t^m \\ \alpha F_t^f & \alpha F_t^m \end{pmatrix} \begin{pmatrix} \phi^f \\ \phi^m \end{pmatrix}(x, t)$$

# Quantities of interest

First, we only consider the first 18 entries of  $a^f(x, t)$  and  $a^m(x, t)$  reflecting the expected number of living children less than 18 years old of an individual aged  $x$  in year  $t$ , and sum these entries over both child sexes. Define this quantity to  $a_{<18}(x, t)$ . In any given year, we can estimate the number of children losing a parent as

$$\text{Number of children } < 18 \text{ yo experiencing parental death} = \sum_{x=0}^{\omega} a_{<18}(x, t) \times D(x, t) \quad (4)$$

where  $D(x, t)$  reflects the death counts of individuals aged  $x$ , in year  $t$ , from a given cause. This term can be computed by sex, giving the number of children who lost a mother (father). The probability for youth aged less than 18 years old experiencing a parental death by cause in year  $t$ , can be expressed as the ratio

$$p^g(t) = \frac{\text{Number of children } < 18 \text{ yo experiencing parental death}}{\text{Number of children } < 18 \text{ yo}} = \frac{\sum_{x=0}^{\omega} a_{<18}(x, t) \times D(x, t)}{\sum_{x=0}^{\omega} a_{<18}(x, t) \times N(x, t)}$$

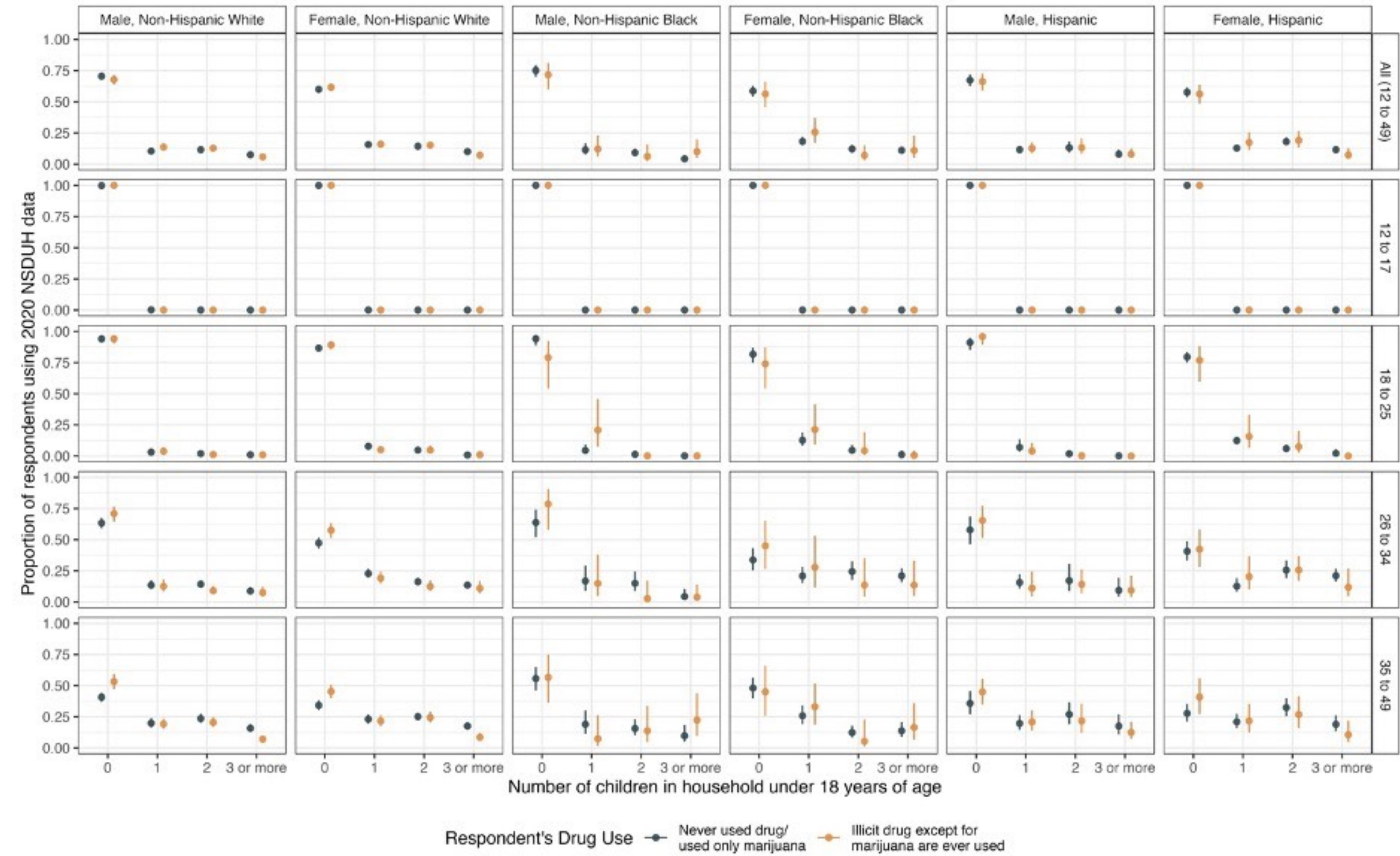
# Male fertility

- Model ratio of race/ethnicity fertility to national fertility over time
- Male race/ethnicity specific fertility is:

Ratio x male national fertility x race-specific female national fertility x (male/female national fertility)

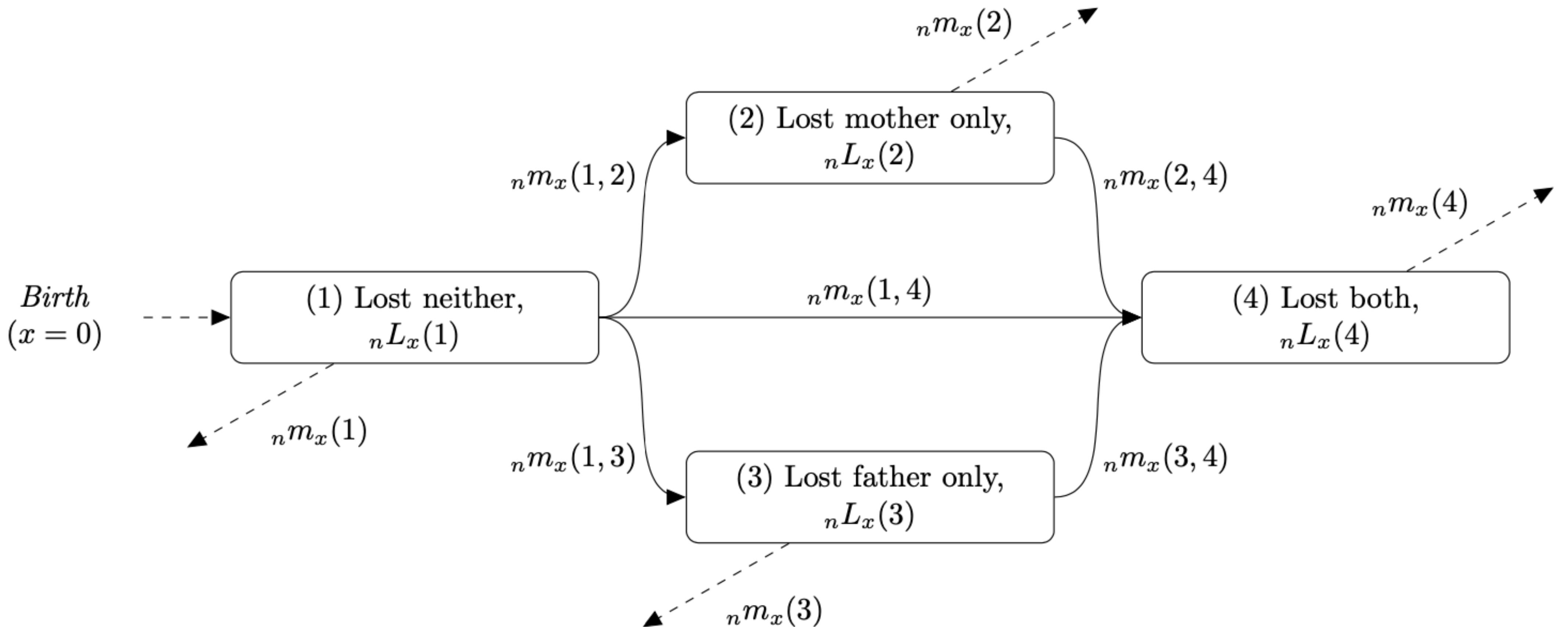
$$\hat{f}_{r,a,t} = (\widehat{\alpha_{a,r}} + \hat{\beta}_{a,r}t) \times F_{a,t} \times TFR_{r,,t} \times \frac{TFR_{M,t}}{TFR_{F,t}}$$

# NSDUH



# Project 2

# States of parental loss



# Multistate lifetable for parental loss

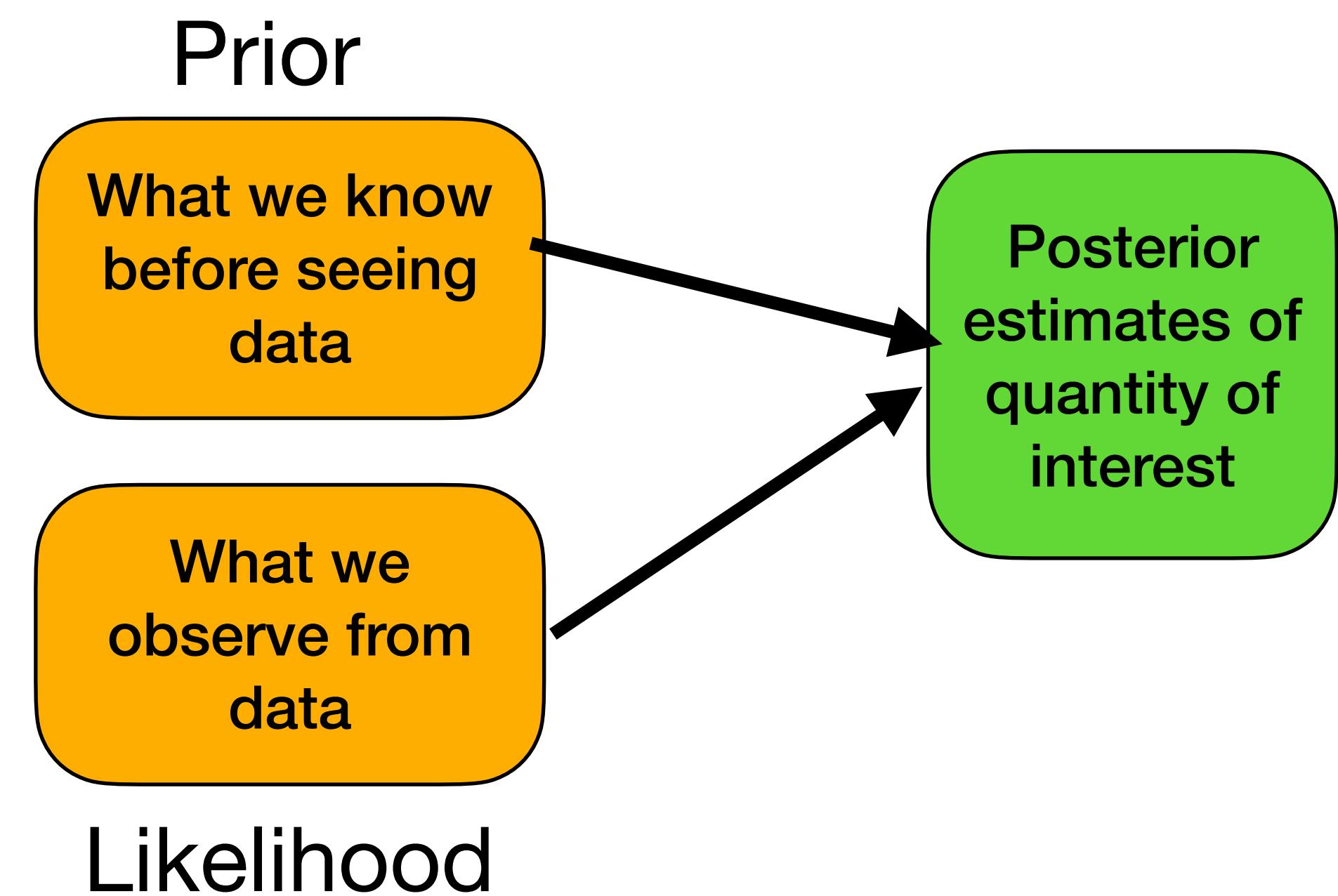
- Interested in four different states over age (lost none, lost mother, lost father, lost both)
- Partition person-years lived in each age group into person-years in each state
- Calculate the life expectancy lived in each state  $i$  for each age  $x$

$$e_x(i) = \frac{\sum_{k=0}^{\infty} L_{(x+k \cdot n)}(i)}{\ell_x} \quad \text{for } i = 1, 2, 3, 4$$

- E.g.  $e_0(1)$  is the expected number of years lived with both parents at age 0.

# Using Bayesian methods to combine data sources

- Estimates of quantity of interest are a combination of the **prior** knowledge and evidence we see from data, encoded in the **likelihood**
- The **posterior** combines both of these types of evidence to give estimates of the quantity of interest
- Bayesian inference is very useful in a wide range of demographic settings:
  - Dealing with multiple data sources
  - Different types of uncertainty
  - Hierarchical structures

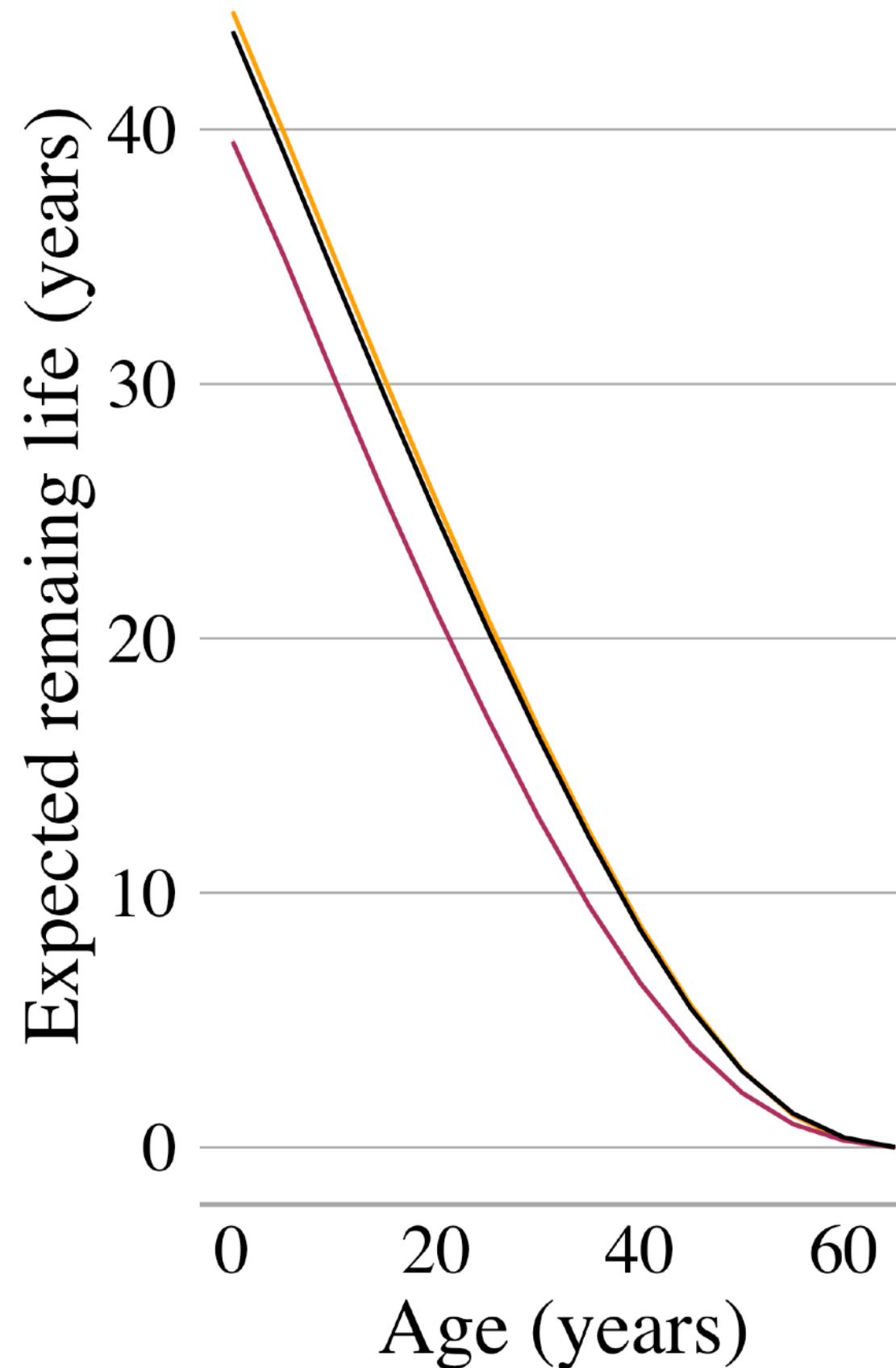


# Computation

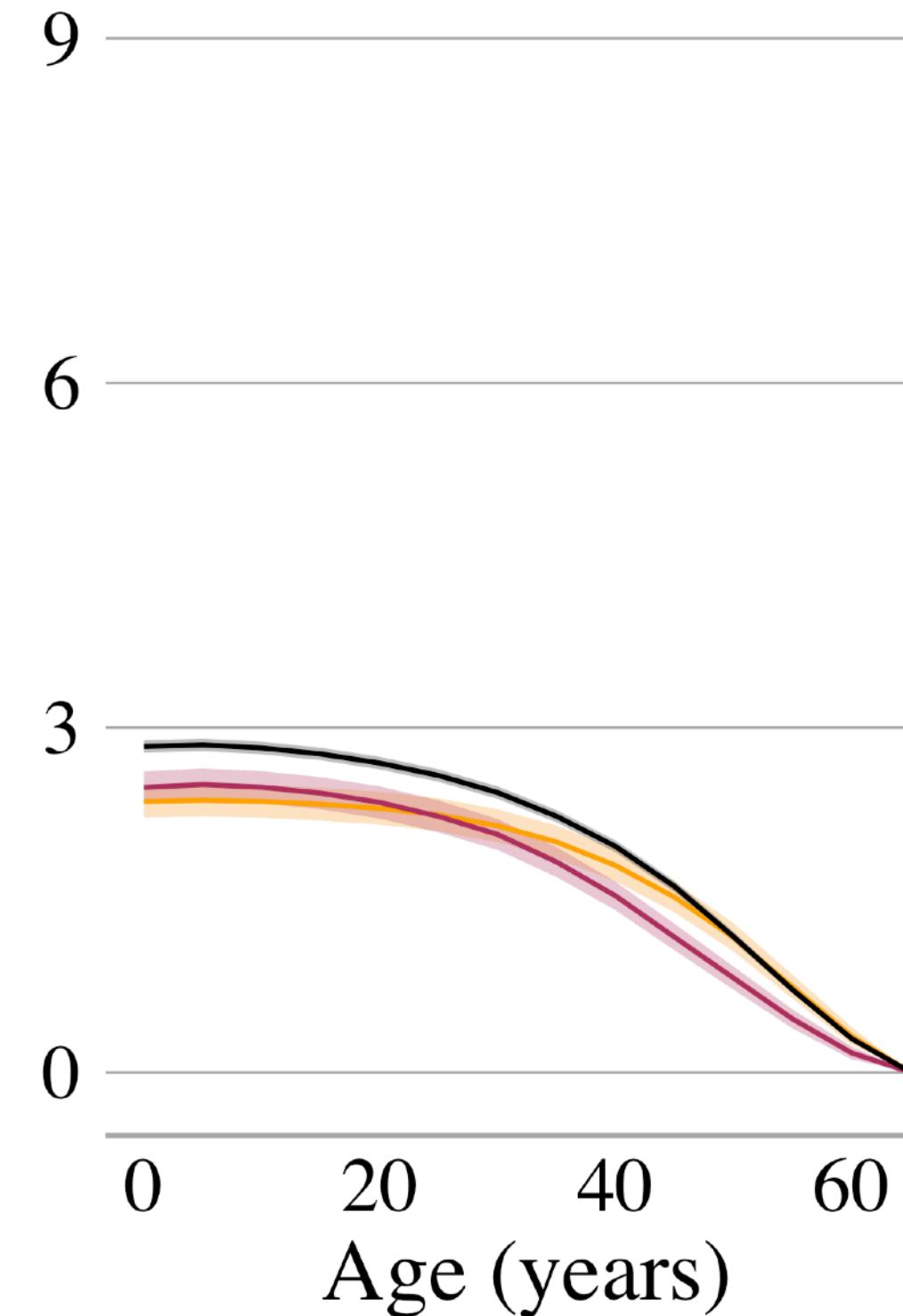
- Samples from the posterior distribution for the transition rates (and proportions) are obtained using a Markov Chain Monte Carlo (MCMC) algorithm (using Stan in R)
- Posterior samples for rates and proportions are then used to calculate posterior samples of multistate lifetables by parental mortality status
- 95% credible intervals are calculated based on the 2.5th and 97.5th percentiles



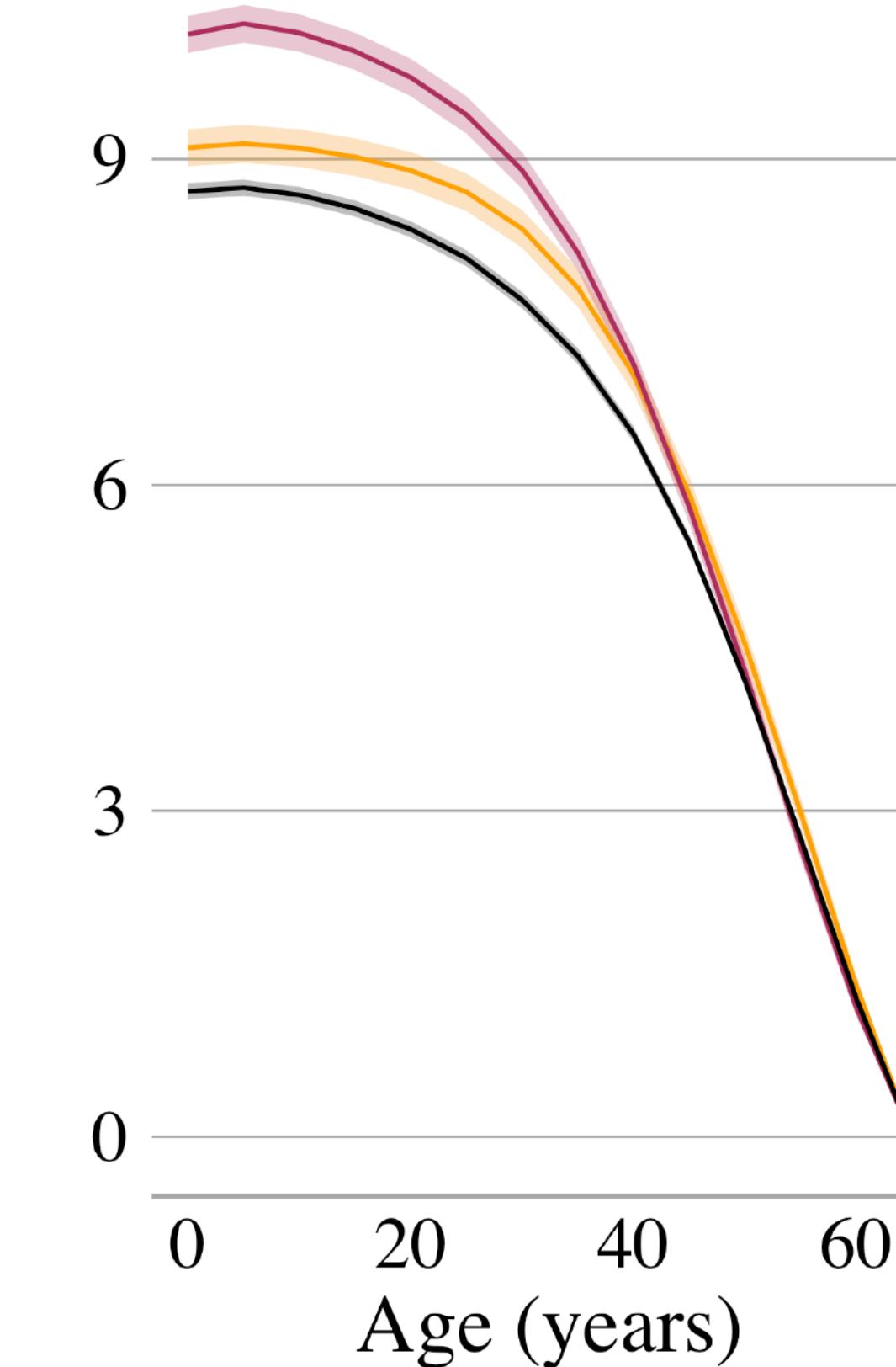
# Remaining life expectancy



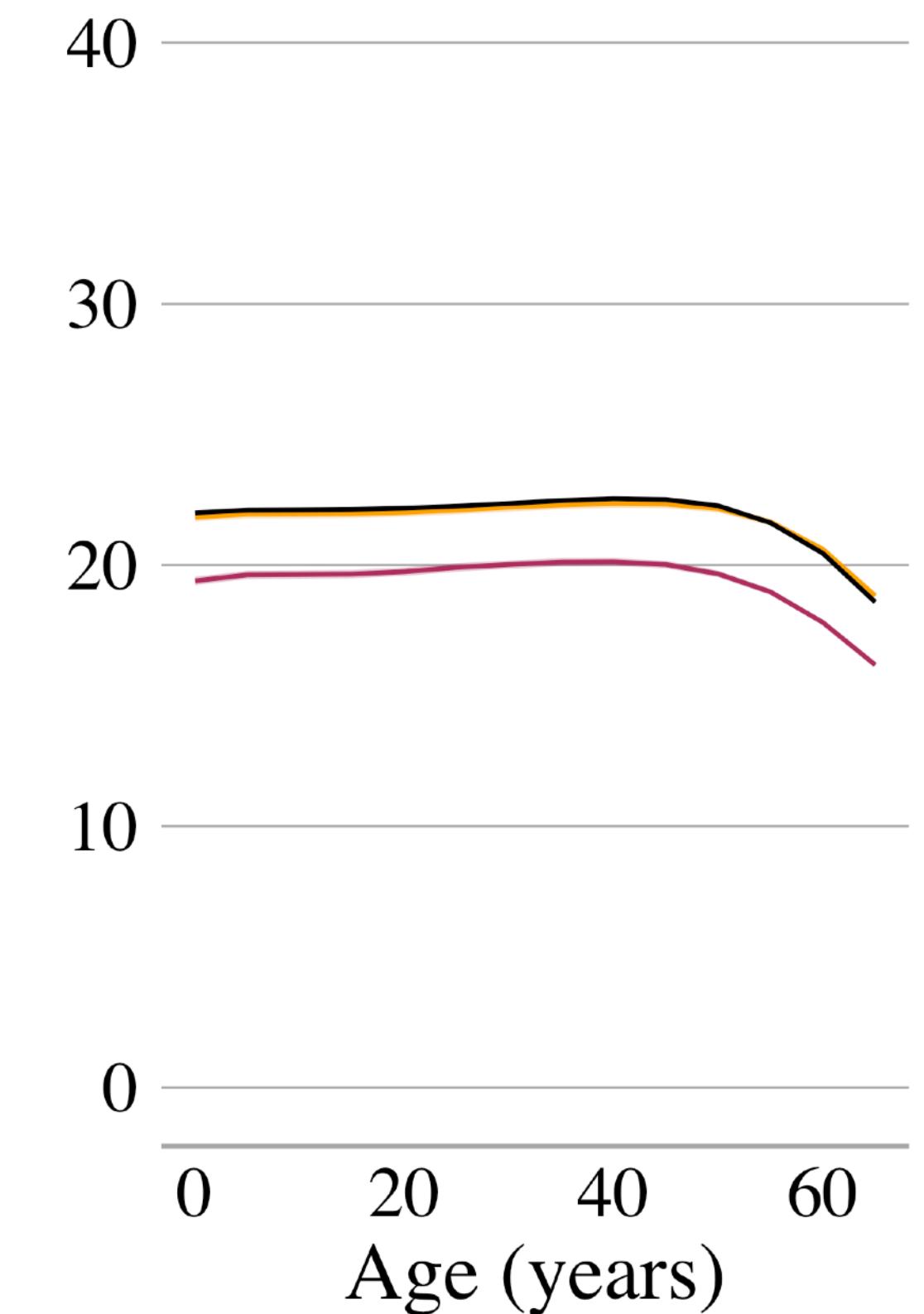
a) Lost none



b) Lost mother only



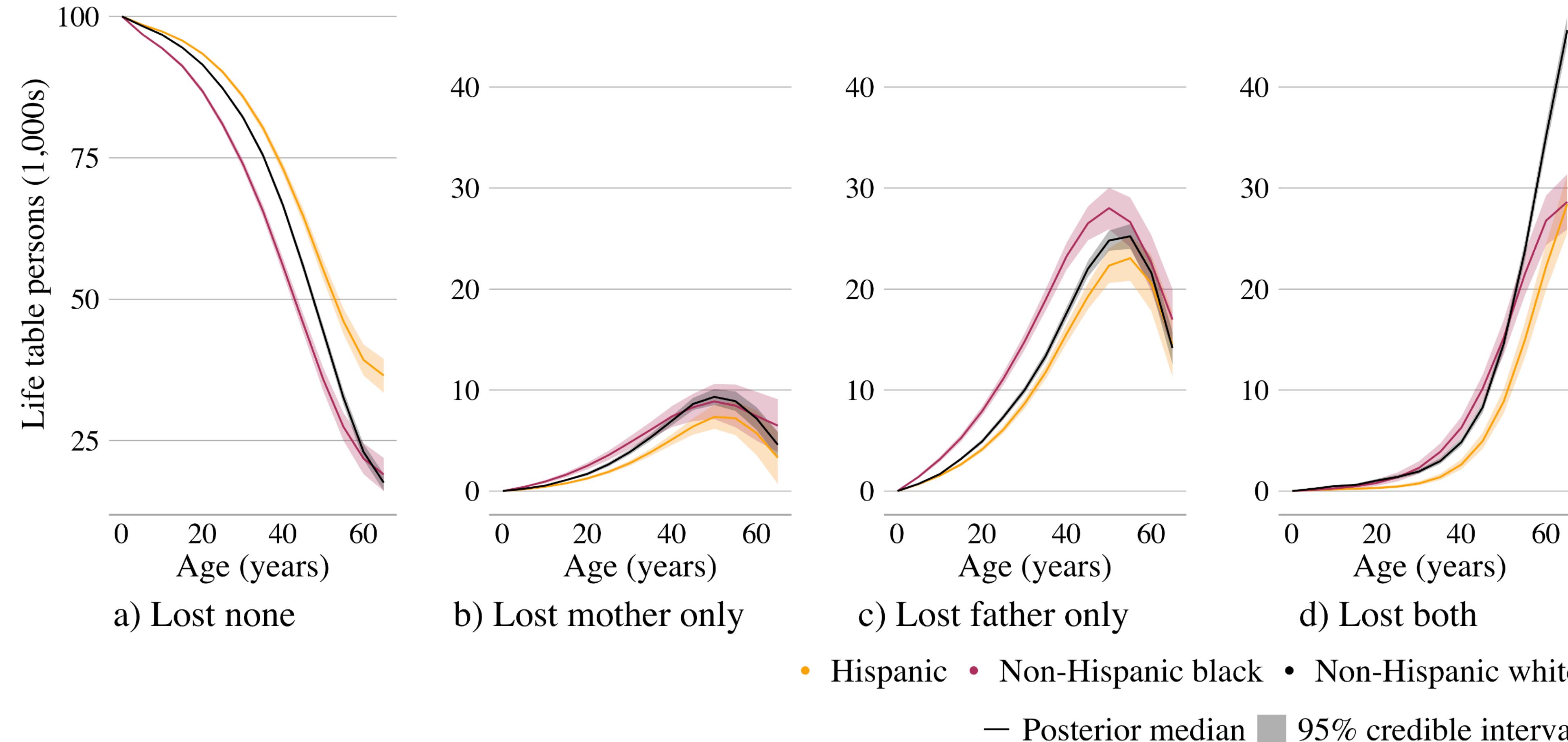
c) Lost father only

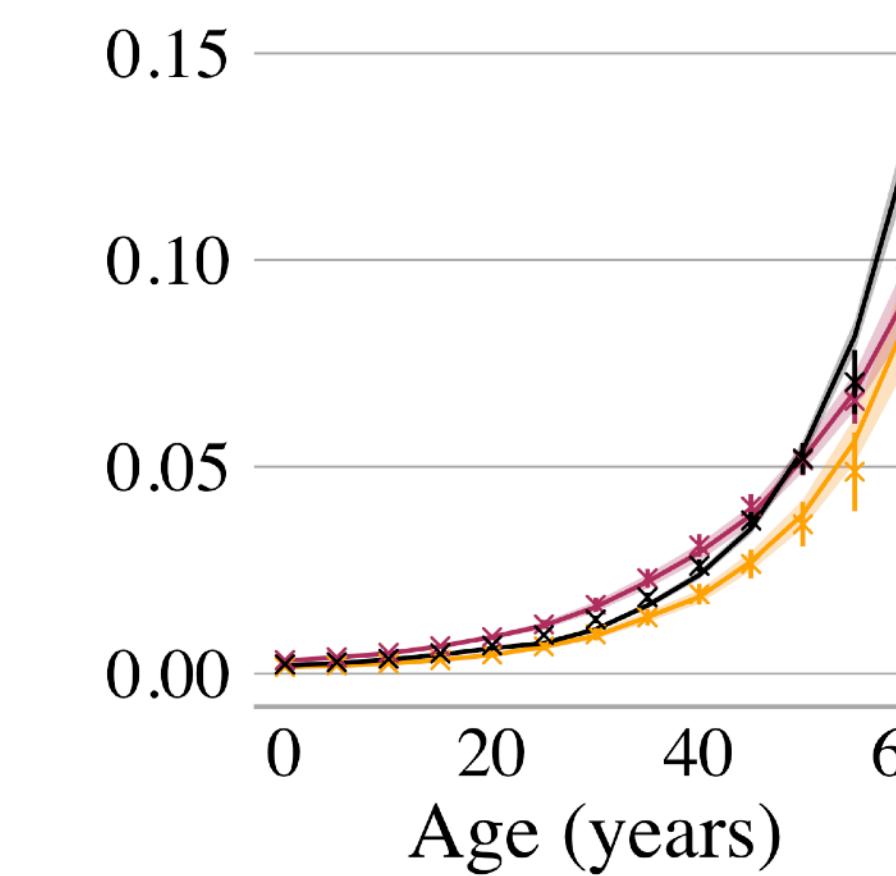
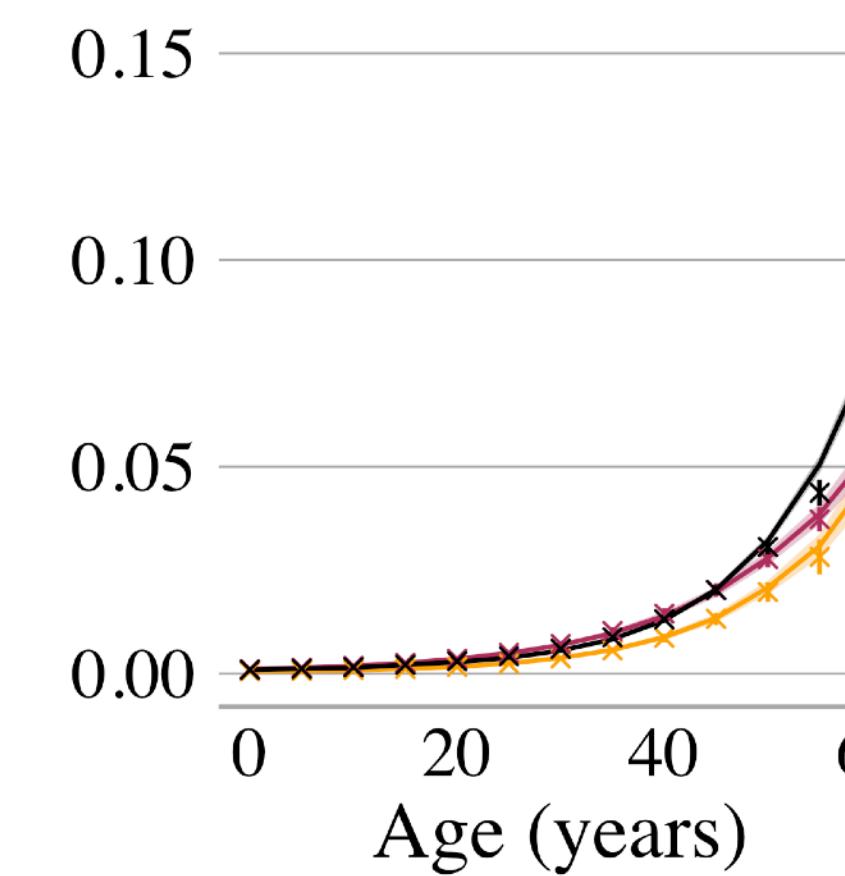
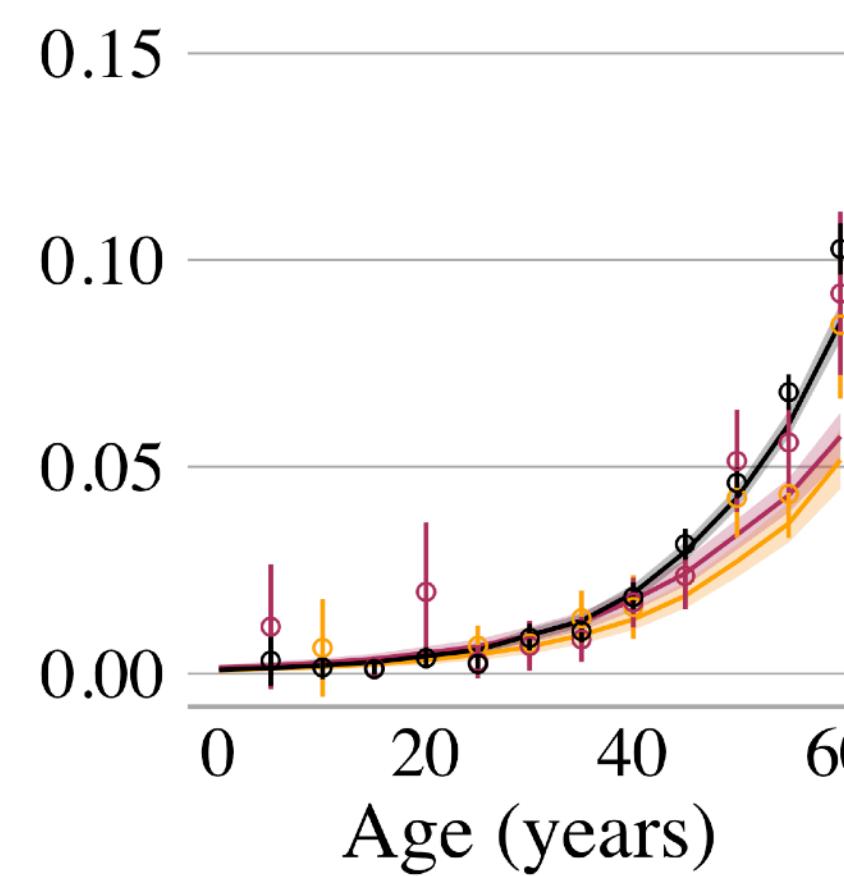
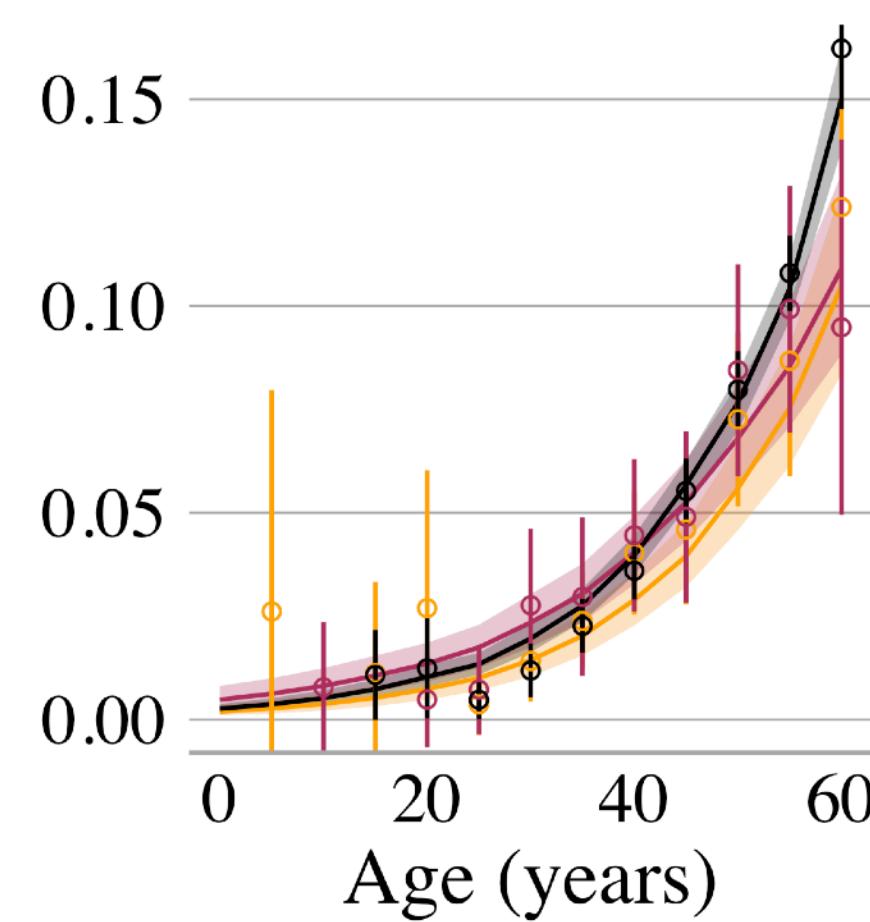
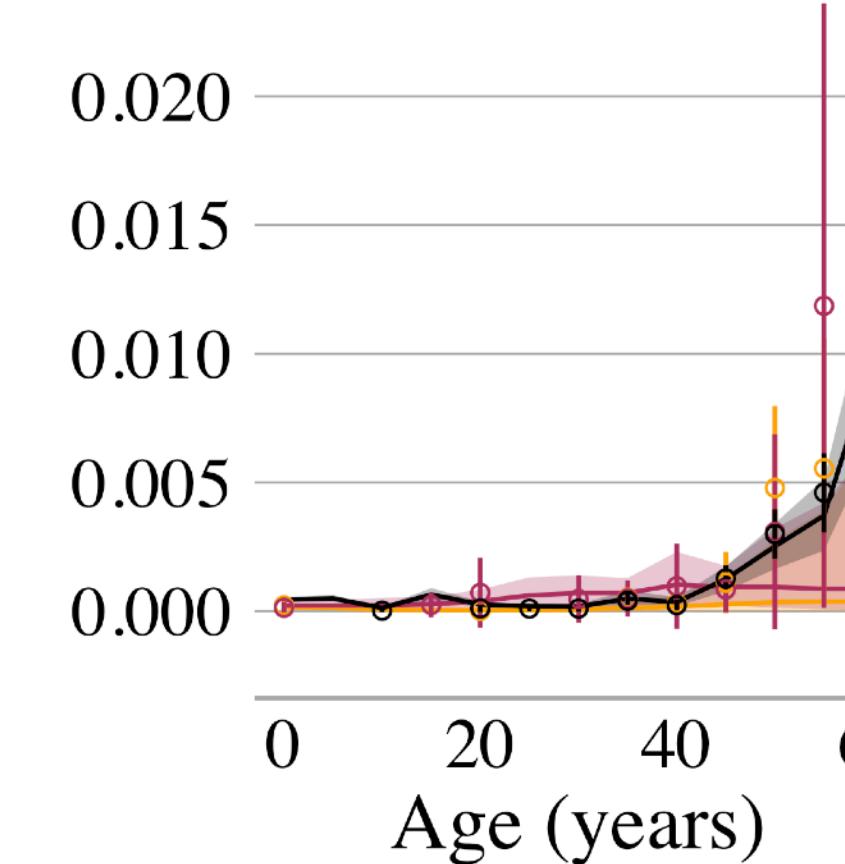
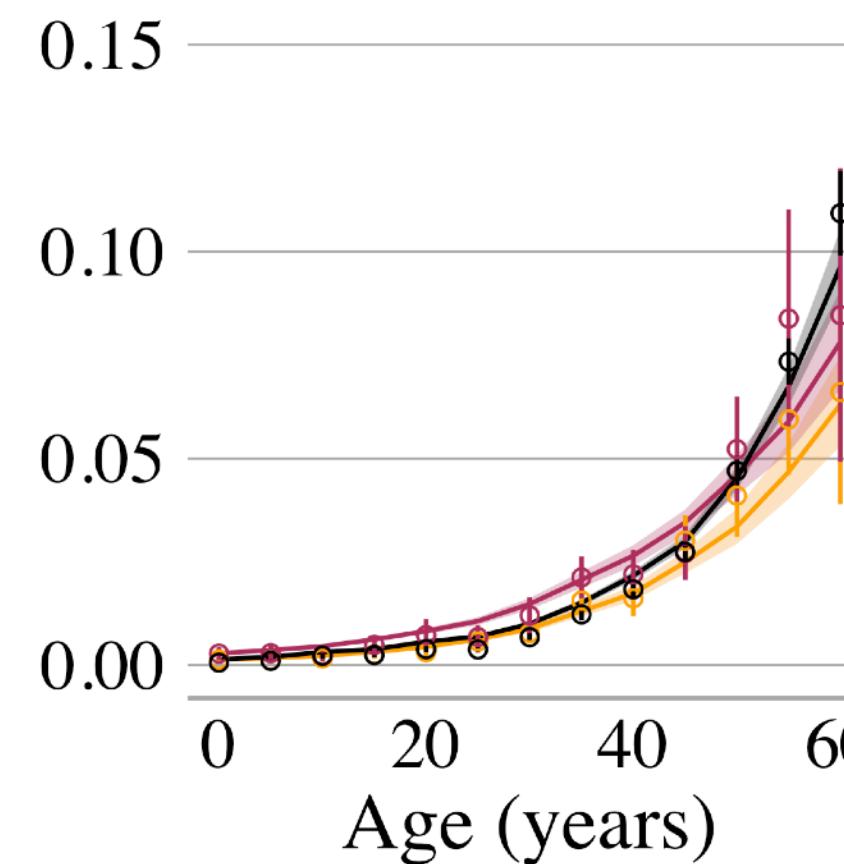
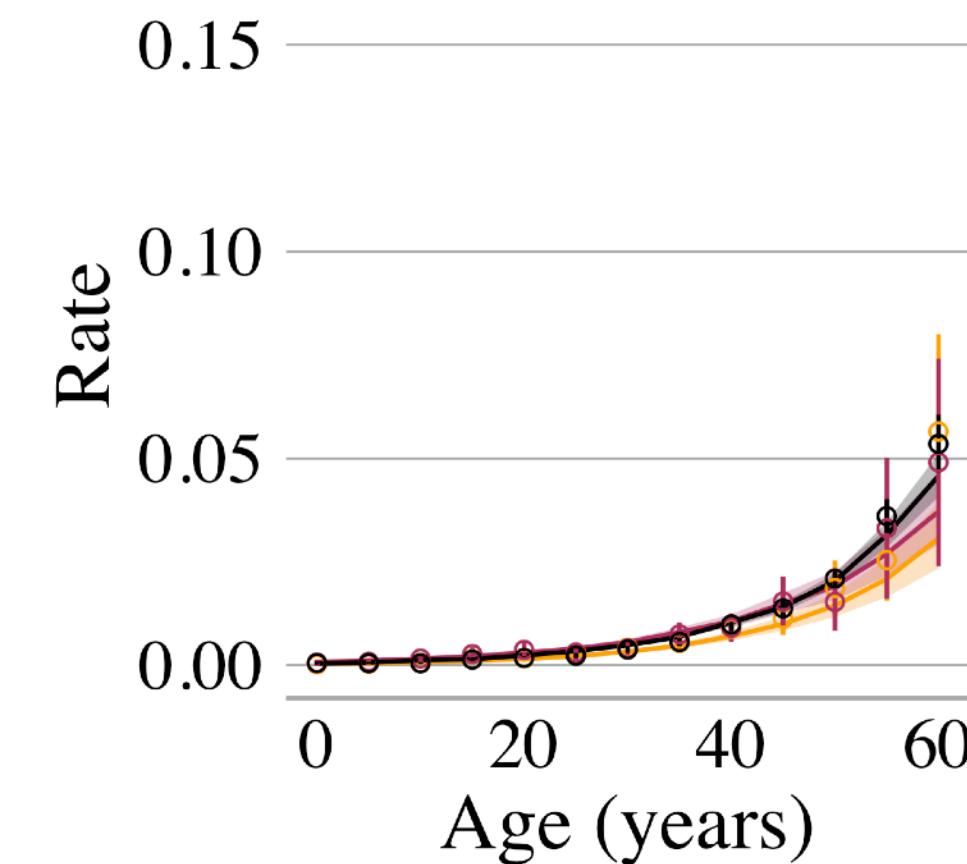


d) Lost both

- Hispanic
- Non-Hispanic black
- Non-Hispanic white
- Posterior median
- 95% credible interval

# Survivors by age

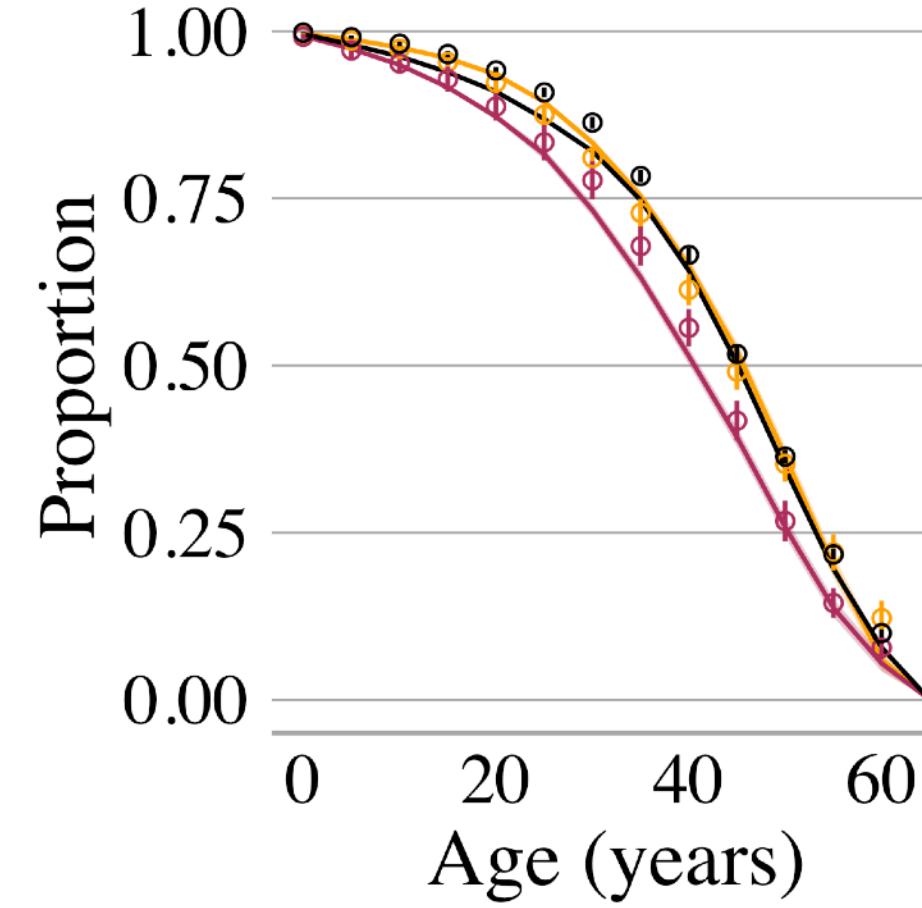




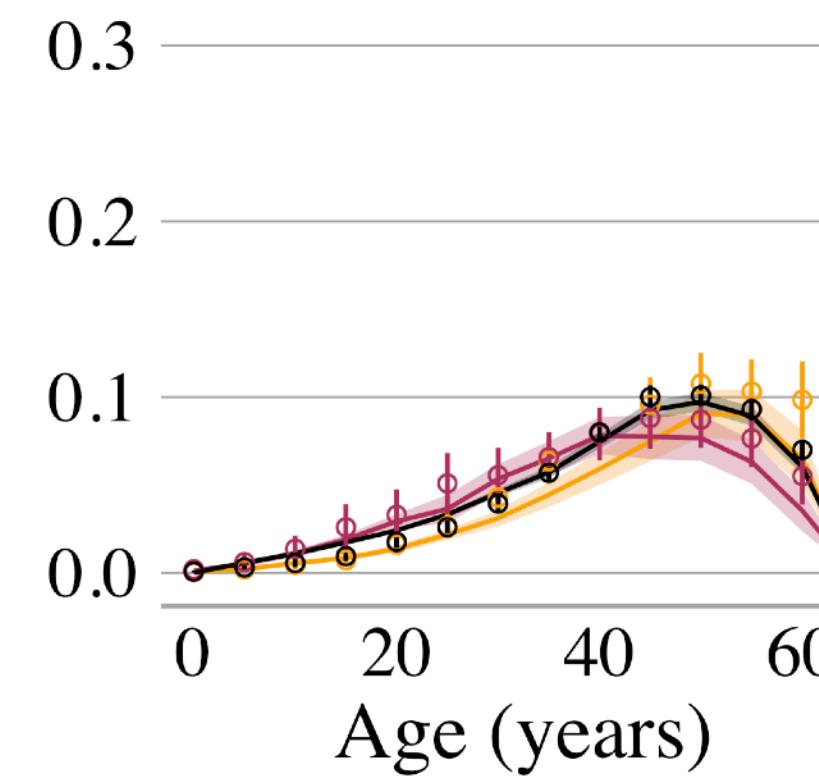
◦ SIPP estimates × Projection matrix estimates

• Hispanic • Non-Hispanic black • Non-Hispanic white

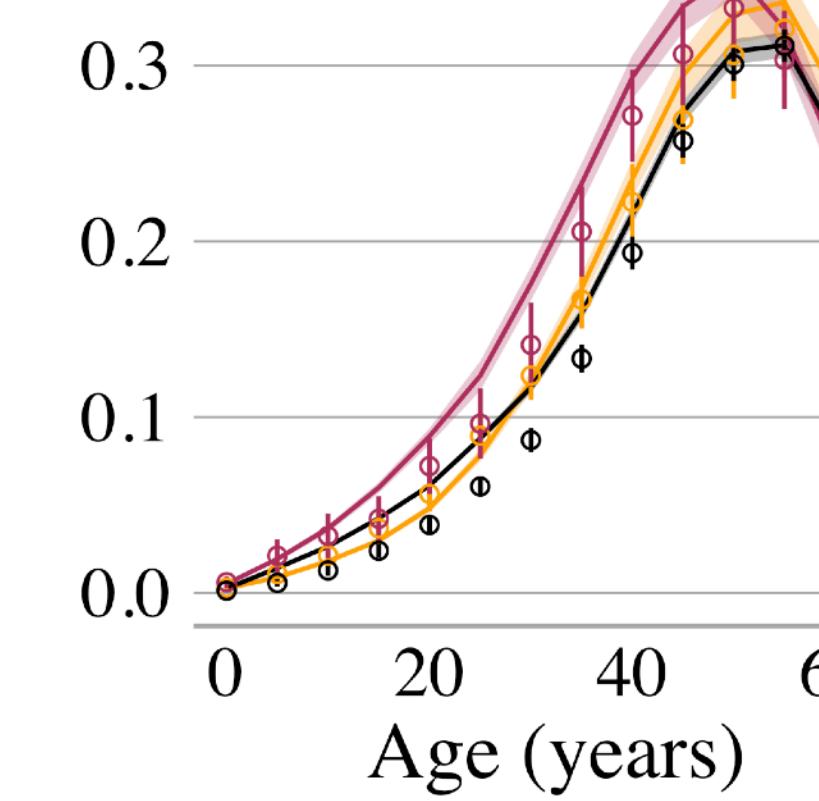
— Posterior median ■ 95% credible interval



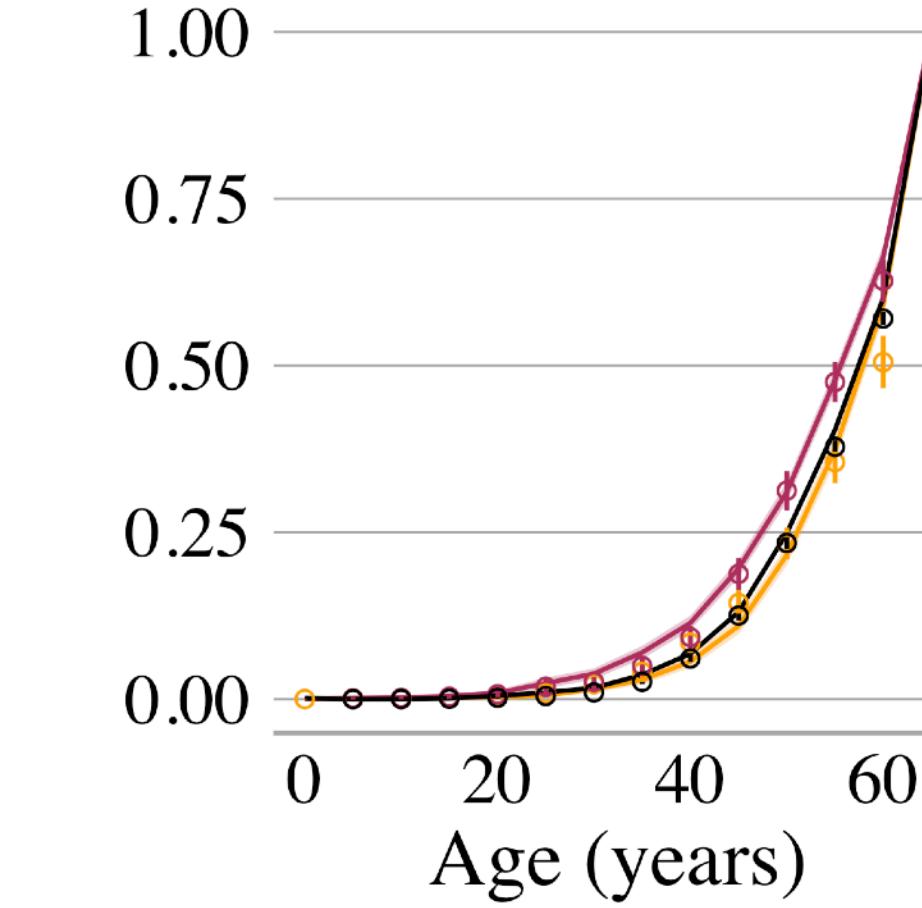
a) Lost none



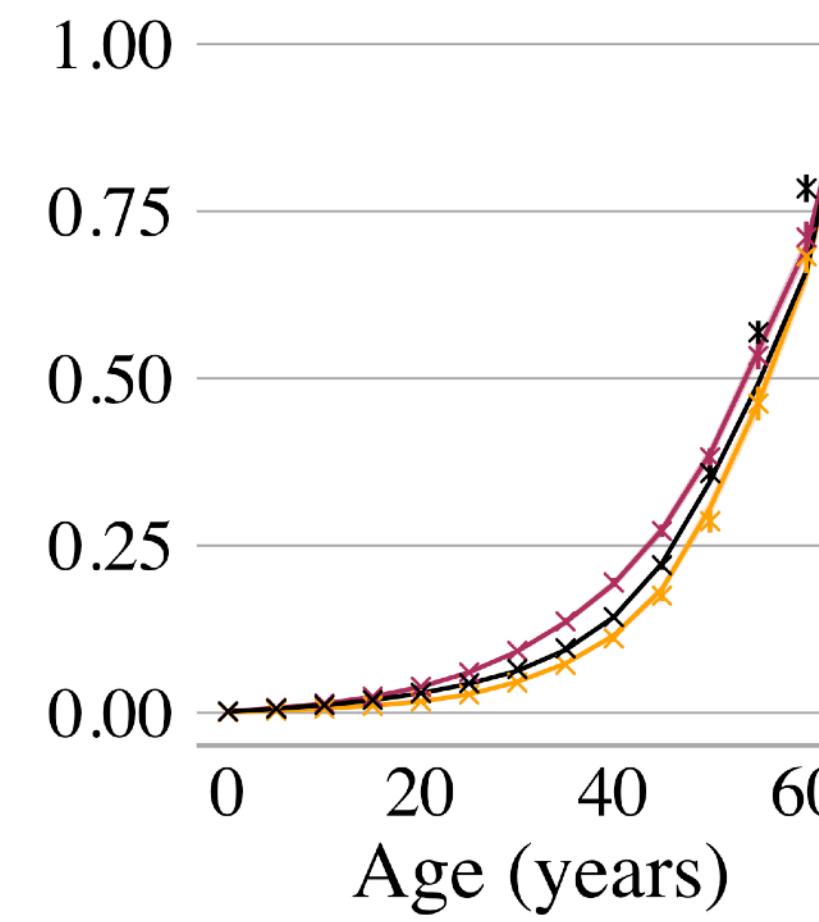
b) Lost mother only



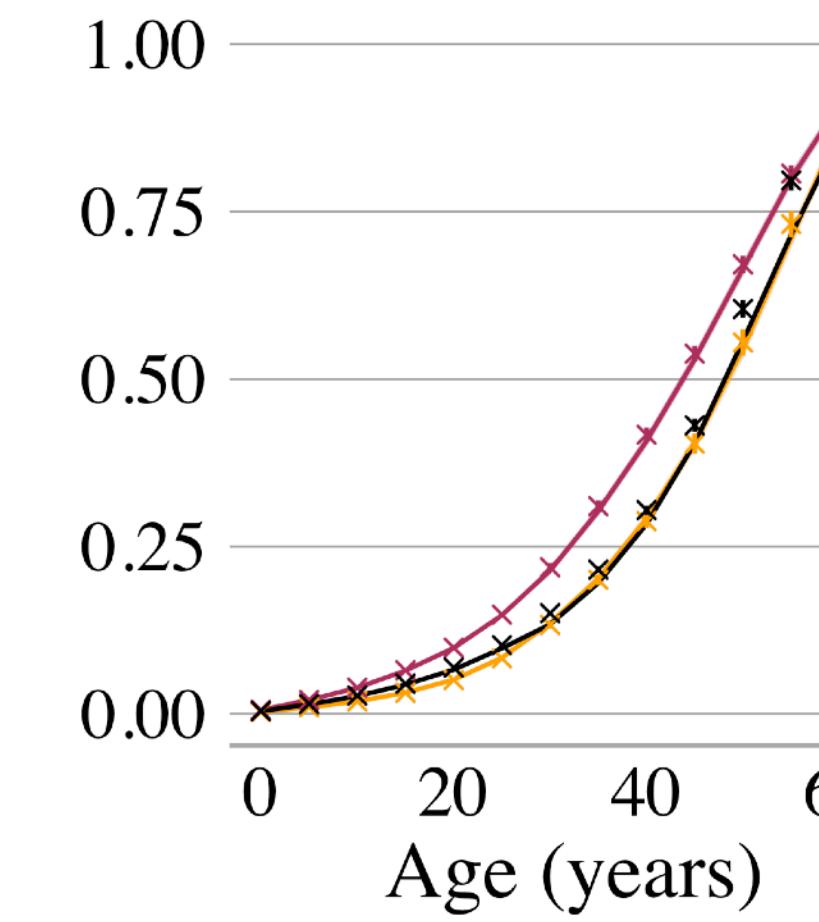
c) Lost father only



d) Lost both



e) Lost mother or both

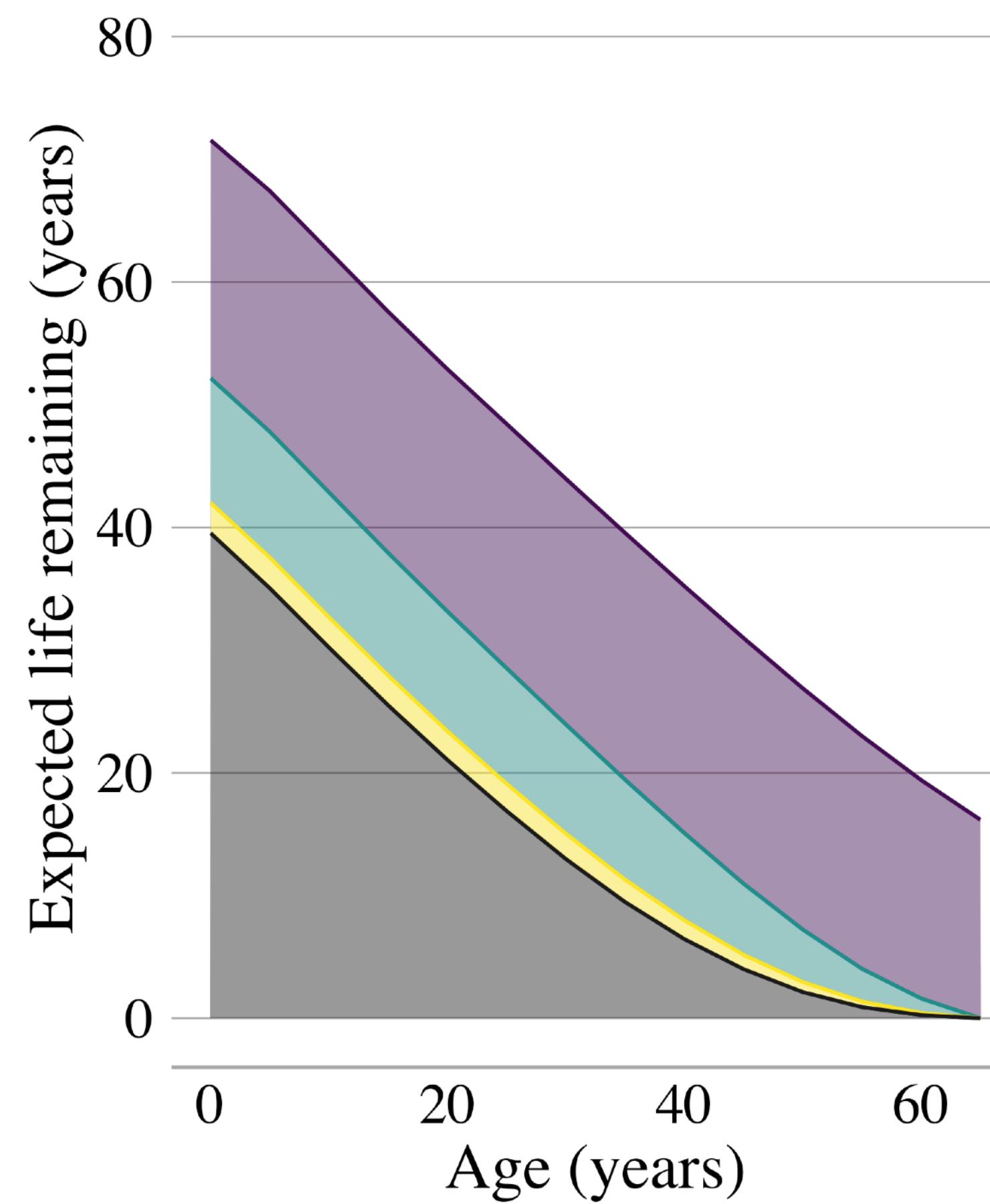


f) Lost father or both

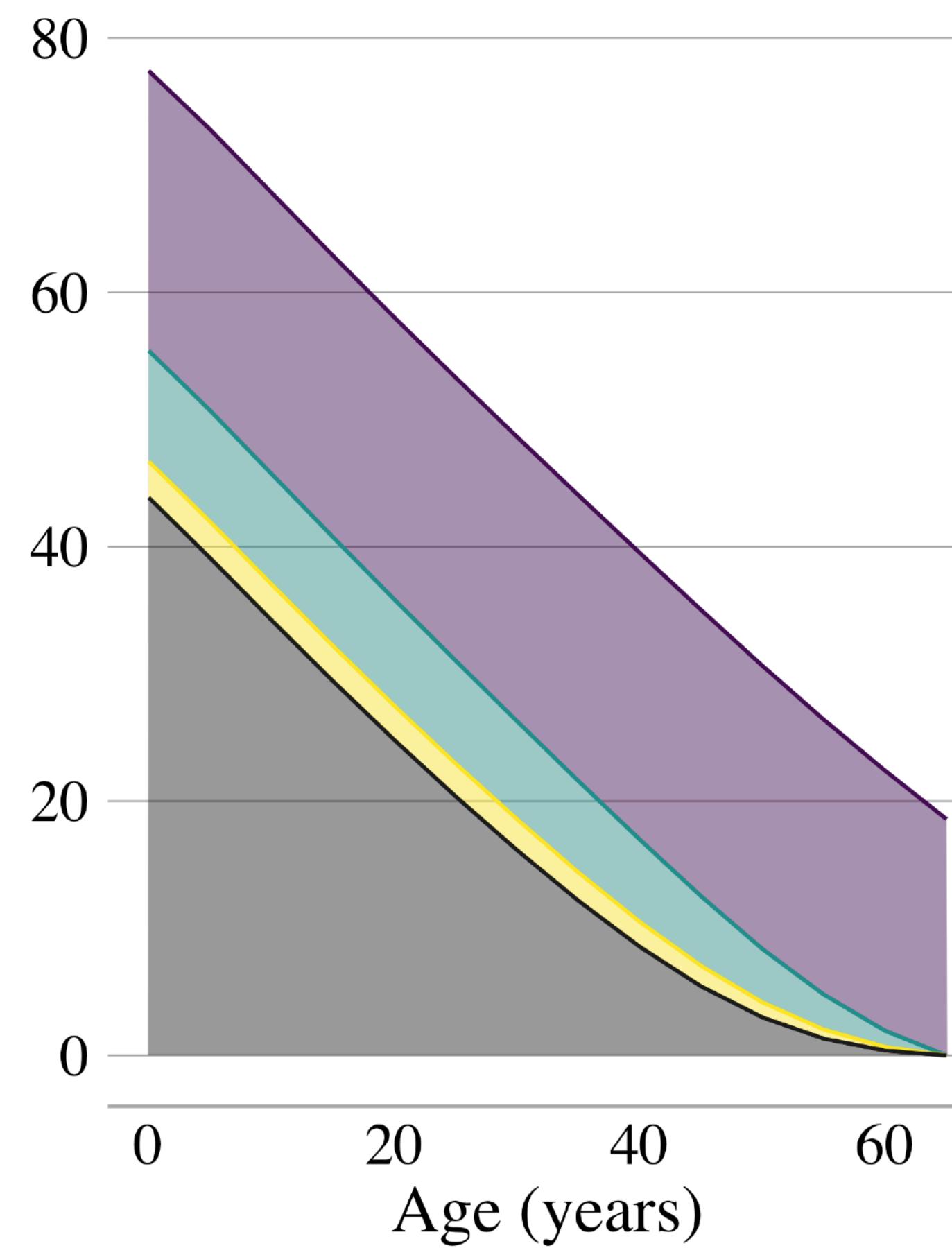
◦ SIPP estimates × Projection matrix estimates

• Hispanic • Non-Hispanic black • Non-Hispanic white

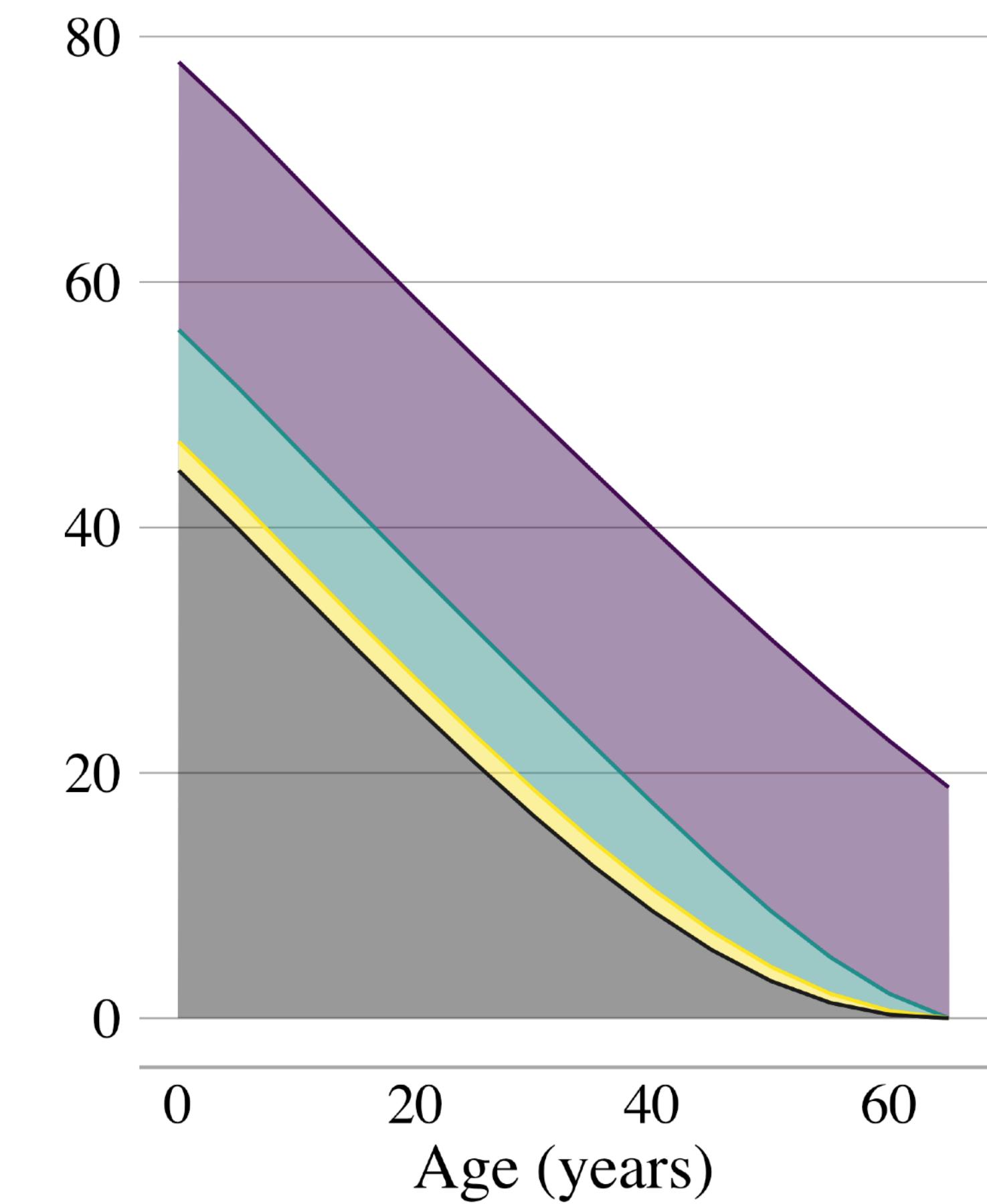
— Posterior median ■ 95% credible interval



a) Non-Hispanic black



b) Non-Hispanic white



c) Hispanic

- Lost both • Lost father only • Lost mother only • Lost none