

COVID-19 vaccine-associated mortality in the Southern Hemisphere

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Abstract

Seventeen equatorial and Southern-Hemisphere countries were studied (Argentina, Australia, Bolivia, Brazil, Chile, Colombia, Ecuador, Malaysia, New Zealand, Paraguay, Peru, Philippines, Singapore, South Africa, Suriname, Thailand, Uruguay), which comprise 9.10 % of worldwide population, 10.3 % of worldwide COVID-19 injections (vaccination rate of 1.91 injections per person, all ages), virtually every COVID-19 vaccine type and manufacturer, and span 4 continents.

In the 17 countries, there is no evidence in all-cause mortality (ACM) by time data of any beneficial effect of COVID-19 vaccines. There is no association in time between COVID-19 vaccination and any proportionate reduction in ACM. The opposite occurs.

All 17 countries have transitions to regimes of high ACM, which occur when the COVID-19 vaccines are deployed and administered. Nine of the 17 countries have no detectable excess ACM in the period of approximately one year after a pandemic was declared on 11 March 2020 by the World Health Organization (WHO), until the vaccines are rolled out (Australia, Malaysia, New Zealand, Paraguay, Philippines, Singapore, Suriname, Thailand, Uruguay).

Unprecedented peaks in ACM occur in the summer (January-February) of 2022 in the Southern Hemisphere, and in equatorial-latitude countries, which are synchronous with or immediately preceded by rapid COVID-19-vaccine-booster-dose rollouts (3rd or 4th doses). This phenomenon is present in every case with sufficient mortality data (15 countries). Two of the countries studied have insufficient mortality data in January-February 2022 (Argentina and Suriname).

Detailed mortality and vaccination data for Chile and Peru allow resolution by age and by dose number. It is unlikely that the observed peaks in all-cause mortality in January-February 2022 (and additionally in: July-August 2021, Chile; July-August 2022, Peru), in each of both countries and in each elderly age group, could be due to any cause other than the temporally associated rapid COVID-19-vaccine-booster-dose rollouts. Likewise, it is unlikely that the transitions to regimes of high ACM, coincident with the rollout and sustained administration of COVID-19 vaccines, in all 17 Southern-Hemisphere and equatorial-latitude countries, could be due to any cause other than the vaccines.

Synchronicity between the many peaks in ACM (in 17 countries, on 4 continents, in all elderly age groups, at different times) and associated rapid booster rollouts allows this

firm conclusion regarding causality, and accurate quantification of COVID-19-vaccine toxicity.

The all-ages vaccine-dose fatality rate (vDFR), which is the ratio of inferred vaccine-induced deaths to vaccine doses delivered in a population, is quantified for the January–February 2022 ACM peak to fall in the range 0.02 % (New Zealand) to 0.20 % (Uruguay). In Chile and Peru, the vDFR increases exponentially with age (doubling approximately every 4 years of age), and is largest for the latest booster doses, reaching approximately 5 % in the 90+ years age groups (1 death per 20 injections of dose 4). Comparable results occur for the Northern Hemisphere, as found in previous articles (India, Israel, USA).

We quantify the overall all-ages vDFR for the 17 countries to be (0.126 ± 0.004) %, which would imply 17.0 ± 0.5 million COVID-19 vaccine deaths worldwide, from 13.50 billion injections up to 2 September 2023. This would correspond to a mass iatrogenic event that killed (0.213 ± 0.006) % of the world population (1 death per 470 living persons, in less than 3 years), and did not measurably prevent any deaths.

The overall risk of death induced by injection with the COVID-19 vaccines in actual populations, inferred from excess all-cause mortality and its synchronicity with rollouts, is globally pervasive and much larger than reported in clinical trials, adverse effect monitoring, and cause-of-death statistics from death certificates, by 3 orders of magnitude (1,000-fold greater).

The large age dependence and large values of vDFR quantified in this study of 17 countries on 4 continents, using all the main COVID-19 vaccine types and manufacturers, should induce governments to immediately end the baseless public health policy of prioritizing elderly residents for injection with COVID-19 vaccines, until valid risk-benefit analyses are made.

Table of Contents

Abstract.....	1
Table of Figures	5
Table of Tables	8
1. Introduction.....	9
2. Data.....	11
3. Method to Detect Time Transitions to Regimes of High All-Cause Mortality.....	12
4. Methods to Quantify vDFR from All-Cause Mortality	13
4.1 Historical-trend baseline for a period (or peak) of mortality (Method 1).....	13
4.2 Special case of a single historic integrated point (Method 2)	16
4.3 Application of the methods to the specific countries.....	18
5. Results.....	18
5.1 Transitions to regimes of high ACM	18
5.2 Covid-period excess mortality compared to COVID-19-vaccination-period excess mortality, in the 17 countries of this study	25
5.3 All studied countries all-ages national ACM and vaccine rollouts.....	28
5.4 Chile by age group dose 4 — National ACM and vaccine dose 4 rollouts (90+, 85-89, 80-84, 75-79, 70-74, 65-69, and 60-64 years)	71
5.5 Chile by age group dose 3 — National ACM and vaccine dose 3 rollouts (90+, 85-89, 80-84, 75-79, 70-74, 65-69, and 60-64 years)	80
5.6 Peru by age group dose 3 — National ACM and vaccine dose 3 rollouts (90+, 85-89, 80-84, 75-79, 70-74, 65-69, and 60-64 years)	85
5.7 Peru by age group dose 4 — National ACM and vaccine dose 4 rollouts (90+, 85-89, 80-84, 75-79, 70-74, 65-69, and 60-64 years)	97
5.8 vDFR by age — Age-stratified vaccine dose fatality rates for doses 3 and 4 in Chile and Peru	101
5.9 New Zealand all-ages doses 3 and 4 — National ACM and vaccine dose rollouts	104
6. Discussion	107
6.1 COVID-19 vaccines can cause death	107
6.2 Absence of excess mortality until the COVID-19 vaccines are rolled out.....	109
6.3 The COVID-19 vaccines did not save lives and appear to be lethal toxic agents	110
6.4 Strong evidence for a causal association and vaccine lethal toxicity	113
6.5 Causality in excess mortality is amply demonstrated.....	115
6.6 Assessing other interpretations of the cause of the excess mortality	120

6.7	Implications regarding age-dependence of fatal toxicity of COVID-19 vaccines	123
6.8	Excess ACM by time prior to COVID-19 vaccine rollouts.....	127
7.	Conclusion.....	131
7.1	Causality proven	131
7.2	Actual vaccine mortality much larger than that incorrectly inferred from faulty data	131
7.3	The policy of prioritizing elderly residents for COVID-19 vaccination must be ended.....	133
	References.....	134
	Appendix A: Sources of mortality and vaccination data.....	154
	Appendix B: Examples of all-cause mortality and vaccination data.....	158
	Appendix C: Technical and specific information for applications of the methods to the data	163
	Appendix D: Single-point method quantification of the excess mortality in the Covid period	168
	Appendix E: Single-point method quantification of the excess mortality in the vaccination period	175

Table of Figures

Figure 1: World map showing the 17 countries considered in the present study, in relation to the equator and the tropics — Argentina, Australia, Bolivia, Brazil, Chile, Colombia, Ecuador, Malaysia, New Zealand, Paraguay, Peru, Philippines, Singapore, South Africa, Suriname, Thailand, and Uruguay.	12
Figure 2: Transitions between regimes of mortality — ACM by time (week or month) (blue), vaccine administration by time (week) (orange), and the 1-year backward moving average of the ACM by time (week or month) (red), as per the legend, for the 17 countries in this study. The 11 March 2020 pandemic declaration date is shown by a vertical grey line in each panel. The data sources are specified in Appendix A.	24
Figure 3: (three panels) Covid-period excess ACM versus vaccination-period excess ACM (top), first expanded view (middle), second expanded view (bottom), with 1:1 line as a reference.	28
Figure 4: All-ages ACM by week (2019-2023) or by month (2017-2023) (black), with all-ages vaccine rollouts doses 1 and 2 (blue) and boosters (orange), for the 17 countries	

in this study (as labelled); also showing the 11 March 2020 date and limits of integration for the nominally January–February 2022 peak in ACM (vertical lines). The data sources are specified in Appendix A.....	37
Figure 5: All-ages ACM by week or by month (light blue), 2015–2023, with integrations for the nominally January–February 2022 peak in ACM (dark blue), for the 17 countries (as labelled). The data sources are specified in Appendix A.....	46
Figure 6: All-ages ACM by week (2019–2023) or by month (2017–2023) (black), with all-ages all-doses COVID-19 vaccine rollout (blue), for the 17 countries in this study; also showing the 11 March 2020 date (grey) and limits of integration (dashed, blue) for the vaccination period. The data sources are specified in Appendix A.	55
Figure 7: All-ages ACM by week (ACM/w) or by month (ACM/m) (light blue), 2015–2023, with integrations (dark blue), and trend lines (orange), for the vaccination period, for the 17 countries in this study (as labelled). The data sources are specified in Appendix A.	64
Figure 8: All-ages national excess ACM in the time period of the nominal January–February 2022 ACM peaks (filled circles) or during the vaccination period (open diamonds, Method 2 calculation method) versus the all-ages national total vaccine injections in the same time period, using both by-week (blue) and by-month (orange) ACM data, for all 17 countries (as labelled). Full range (top panel). Expanded view (bottom panel). Pearson correlation coefficient is $r = +0.94$	69
Figure 9: vDFR (vaccination period) versus vDFR (January–February 2022 ACM peak period), with correlation analysis. The vDFR values are expressed as fractions, not %, and the scale is $x1e-3$, so “4” is 0.4 %, and so on. The analysis lines have the same meanings as in Figure 8. Pearson correlation coefficient is $r = +0.74$	70
Figure 10: All-ages vDFR values for the vaccination period (orange, meth. 2) and for the period of the January–February 2022 peak in ACM (blue), by country, in decreasing order, compared to the lists of vaccine manufacturers. Values of all-ages vDFR are from Table 2.	71
Figure 11: Chile (with rollouts), ACM by week (black), 2019–2023, by age group (90+, 85–89 ... 60–64 years), with vaccine rollouts all-doses (blue) and dose-4 (x4 amplified, pink). The data sources are specified in Appendix A.	75
Figure 12: Chile (dose 4), ACM by week (light blue), 2010–2022, by age group (90+, 85–89 ... 60–64 years), with integrations, 14-week dose-4-vaccination-period (dark blue, points) and trend line (orange). The data sources are specified in Appendix A.	79
Figure 13: Chile (dose 3), ACM by week (light blue), 2010–2022, by age group (60+, 90+, 85–89 ... 60–64 years), with integrations, 22-week dose-3-vaccination-period (dark blue, points) and trend line (orange). The data sources are specified in Appendix A....	84

Figure 14: Peru (with rollouts), ACM by week (black), 2019-2022, by age group (60+, 90+, 85-89 ... 60-64, 50-59 ... 20-29, 0-19 years), with vaccine rollouts all-doses (blue) and dose-4 (x4 amplified, pink). The data sources are specified in Appendix A.	92
Figure 15: Peru (dose 3), ACM by week (light blue), 2017-2022, by age group (60+, 90+, 85-89 ... 60-64 years), with integrations, 26-week dose-3-vaccination-period (dark blue, points) and trend line (orange). The data sources are specified in Appendix A....	96
Figure 16: Peru (dose 4), ACM by week (light blue), 2017-2022, by age group (60+, 90+, 85-89 ... 60-64 years), with integrations, 25-week dose-4-vaccination-period (dark blue, points) and trend line (orange). The data sources are specified in Appendix A..	101
Figure 17: vDFR versus age, for Chile and for Peru, for the January-February 2022 peak (Chile dose 4, Peru dose 3), for the July-August 2022 peak (Peru, dose 4), and for the July-August 2021 peak (Chile, dose 3), with exponential fits. Linear scale (top), log scale (bottom).	103
Figure 18: (top) New Zealand (with rollouts), all-ages ACM by week (black), 2019-2023, with all-ages vaccine rollouts dose 1 to dose 4 (colours). (bottom) ACM by week for 80+ years age group (black), and same integrations bounds (dashed vertical lines). The 11 March 2020 pandemic declaration date is shown by a grey vertical line. The data sources are specified in Appendix A.	105
Figure 19: New Zealand (dose 3), all-ages ACM by week (light blue), 2011-2023, with integrations, 28-week dose-3-vaccination-period (dark blue, points) and trend line (orange). The data sources are specified in Appendix A.....	106
Figure 20: New Zealand (dose 4), ACM by week (light blue), 2011-2023, with integrations, 31-week dose-4-vaccination-period (dark blue, points) and trend line (orange). The data sources are specified in Appendix A.....	107
Figure 21: Appendix B - Examples of all-ages all-cause mortality (blue) and all-ages all-doses vaccine-administration (orange) data, for the 17 countries in this study (as labelled). The vertical grey line indicates the 11 March 2020 WHO declaration of a pandemic. The data sources are specified in Appendix A.....	162
Figure 22: Appendix D - Single historic point method (Method 2) for excess ACM in the Covid periods of the 17 countries in this study. Covid period (orange), single-point historic reference period (green), best-line fit (red). The time index is the week or month number from start of 2015. The 11 March 2020 pandemic declaration date is shown by a vertical grey line. The data sources are specified in Appendix A.	174
Figure 23: Appendix E - Single historic point method for excess ACM (Method 2) in the vaccination periods of the 17 countries in this study. Vaccination period (orange), single-point historic reference period (green), best-line fit (red). The time index is the week or month number from start of 2015. The 11 March 2020 pandemic declaration date is shown by a vertical grey line. The data sources are specified in Appendix A.	180

Table of Tables

Table 1: Total excess mortalities in the Covid and vaccination periods, number of injections in the vaccination period, population	26
Table 2: All-ages national vaccine dose fatality rates (vDFR) extracted from the time period of the nominal January–February 2022 ACM peak (“S22-peak”) and from the entire time period of vaccination (“v-period”), for the 17 countries in this study.....	66
Table 3: Extracted doubling ages, increases in age to double risk of death per injection	124
Table 4: Appendix A - Type of data and sources of data.....	154
Table 5: Appendix C - Parameters used to apply the trend-line method (Method 1) to the data	163

1. Introduction

All-cause mortality by time is the most reliable data for detecting and epidemiologically characterizing events causing death, and for gauging the population-level impact of any surge or collapse in deaths from any cause. Such data can be collected by jurisdiction or geographical region, by age group, by sex, and so on; and it is not susceptible to reporting bias or to any bias in attributing causes of death in the mortality itself (Aaby et al., 2020; Bilinski and Emanuel, 2020; Bustos Sierra et al., 2020; Félix-Cardoso et al., 2020; Fouillet et al., 2020; Kontis et al., 2020; Mannucci et al., 2020; Mills et al., 2020; Olson et al., 2020; Piccininni et al., 2020; Rancourt, 2020; Rancourt et al., 2020; Sinnathamby et al., 2020; Tadbiri et al., 2020; Vestergaard et al., 2020; Villani et al., 2020; Achilleos et al., 2021; Al Wahaibi et al., 2021; Anand et al., 2021; Böttcher et al., 2021; Chan et al., 2021; Dahal et al., 2021; Das-Munshi et al., 2021; Deshmukh et al., 2021; Faust et al., 2021; Gallo et al., 2021; Islam, Jdanov, et al., 2021; Islam, Shkolnikov, et al., 2021; Jacobson and Jokela, 2021; Jdanov et al., 2021; Joffe, 2021; Karlinsky and Kobak, 2021; Kobak, 2021; Kontopantelis et al., 2021a, 2021b; Kung et al., 2021a, 2021b; Liu et al., 2021; Locatelli and Rousson, 2021; Miller et al., 2021; Moriarty et al., 2021; Nørgaard et al., 2021; Panagiotou et al., 2021; Pilkington et al., 2021; Polyakova et al., 2021; Rancourt et al., 2021a, 2021b; Rossen et al., 2021; Sanmarchi et al., 2021; Sempé et al., 2021; Soneji et al. 2021; Stein et al., 2021; Stokes et al., 2021; Vila-Corcoles et al., 2021; Wilcox et al., 2021; Woolf et al., 2021; Woolf, Masters and Aron, 2021; Yorifuji et al., 2021; Ackley et al., 2022; Acosta et al., 2022; Engler, 2022; Faust et al., 2022; Ghaznavi et al., 2022; Gobiña et al., 2022; He et al., 2022; Henry et al., 2022; Jha et al., 2022; Johnson and Rancourt, 2022; Juul et al., 2022; Kontis et al., 2022; Kontopantelis et al., 2022; Lee et al., 2022; Leffler et al., 2022; Lewnard et al., 2022; McGrail, 2022; Neil et al., 2022; Neil and Fenton, 2022; Pálinskás and Sándor, 2022; Ramírez-Soto and Ortega-Cáceres, 2022; Rancourt, 2022; Rancourt et al., 2022a, 2022b; Razak et al., 2022; Redert, 2022a, 2022b; Rossen et al., 2022; Safavi-Naini et al., 2022; Schöley et al., 2022; Sy, 2022; Thoma and Declercq, 2022; Wang et al., 2022; Aarstad and Kvistad, 2023; Bilinski et al., 2023; de Boer et al., 2023; de Gier et al., 2023; Demetriou et al., 2023; Donzelli et al., 2023; Haugen, 2023;

Jones and Ponomarenko, 2023; Kuhbandner and Reitzner, 2023; Lytras et al., 2023; Masselot et al., 2023; Matveeva and Shabalina, 2023; Neil and Fenton, 2023; Paglino et al., 2023; Rancourt et al., 2023; Redert, 2023; Schellekens, 2023; Scherb and Hayashi, 2023; Šorli et al., 2023; Woolf et al., 2023).

We have previously reported several cases in which anomalous peaks in all-cause mortality (ACM) are temporally associated with rapid COVID-19 vaccine-dose rollouts and cases in which the start of the COVID-19 vaccination campaign coincides with the start of a new regime of sustained elevated mortality; in India, Australia, Israel, USA, and Canada, including states and provinces (Rancourt, 2022; Rancourt et al., 2022a, 2022b, 2023).

These studies allowed us to make the first quantitative determinations of the vaccine-dose fatality rate (vDFR), which is the ratio of inferred vaccine-induced deaths to vaccine doses administered in a population, based on excess-ACM evaluation on a given time period, compared to the number of vaccine doses administered in the same time period. The all-ages all-doses value of vDFR was typically approximately 0.05 % (1 death per 2,000 injections), with an extreme value of 1 % for the special case of India (Rancourt, 2022). Our work, using extensive data for Australia and Israel, has also shown that vDFR is exponential with age (doubling every 5 years of age), reaching approximately 1 % for 80+ year olds (Rancourt et al., 2023).

The clearest example is that of a relatively sharp ACM peak occurring in January–February 2022 in Australia, which is synchronous with the rapid rollout of Australia's dose 3 of the COVID-19 vaccine; occurring in 5 of 8 of the Australian states and in all of the more-elderly age groups (Rancourt et al., 2022a, 2023). In contrast, often one must contend with the confounding effect of the intrinsic seasonal variation of ACM; however, in this case for Australia, the said January–February 2022 peak occurs at a time in the intrinsic seasonal cycle when one should have a stable (Southern Hemisphere) summer low or summer trough in ACM. There are no previous examples of such a peak in the summer in the historic record of ACM for Australia (Rancourt et al., 2022a).

Few national jurisdictions have the kind of extensive age-stratified mortality and vaccination data available for Australia and Israel. Two other such jurisdictions are Chile and Peru. Here, we show that Chile and Peru, like Australia, has a relatively sharp ACM peak occurring in January–February 2022, which is synchronous with the rapid rollout of Chile's dose 4 and Peru's dose 3 of the COVID-19 vaccine, respectively, occurring for all of the more-elderly age groups.

This shared feature between Chile, Peru and Australia led us to look for more examples of the January–February 2022 ACM-peak phenomenon in the Southern Hemisphere and in equatorial regions. Equatorial countries have no summer and winter seasons and no seasonal variations in their ACM patterns. We found the same phenomenon everywhere that data was available (Australia, Bolivia, Brazil, Chile, Colombia, Ecuador, Malaysia, New Zealand, Paraguay, Peru, Philippines, Singapore, South Africa, Thailand, Uruguay), although incomplete for Bolivia and not as distinctive for New Zealand. Here, we report on those findings.

2. Data

The sources of mortality and vaccine-administration data are given in Appendix A: Sources of mortality and vaccination data.

Appendix B: Examples of all-cause mortality and vaccination data contains examples of the data: all-ages national ACM by time (week or month), from 2015 to 2023, and all-ages all-doses vaccine administration by week, using Y-scales starting from zero, for the 17 countries considered in the present study: Argentina, Australia, Bolivia, Brazil, Chile, Colombia, Ecuador, Malaysia, New Zealand, Paraguay, Peru, Philippines, Singapore, South Africa, Suriname, Thailand, and Uruguay.

Figure 1 shows the said 17 countries considered, in relation to the equator on a world map.

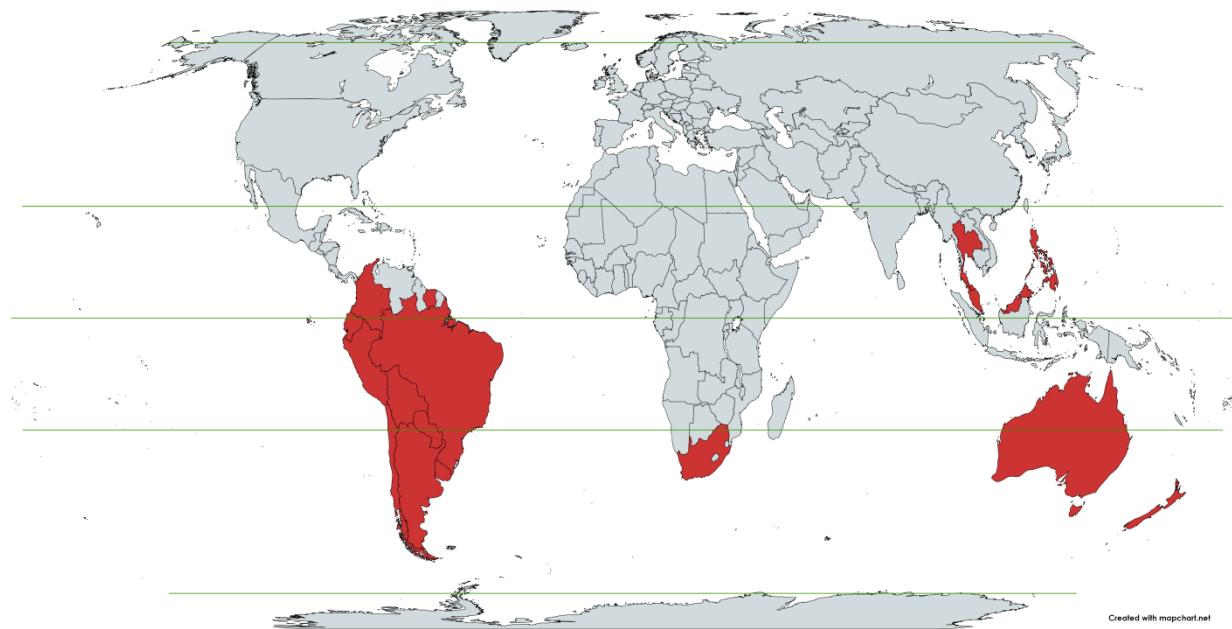


Figure 1: World map showing the 17 countries considered in the present study, in relation to the equator and the tropics — Argentina, Australia, Bolivia, Brazil, Chile, Colombia, Ecuador, Malaysia, New Zealand, Paraguay, Peru, Philippines, Singapore, South Africa, Suriname, Thailand, and Uruguay.

3. Method to Detect Time Transitions to Regimes of High All-Cause Mortality

We implement the following method developed by one of us (JH) for detecting changes in regime in ACM data by time (day, week, month, quarter).

One is interested in detecting transitions in time (as one advances in time from a stable historic period) to regimes of “higher than usual” or “higher than recent” ACM, which may be associated with the declaration of a pandemic or with rollouts of vaccines. Although the trained eye can detect such transitions in the raw ACM by time data itself, it is useful to apply a statistical transformation, which is designed to largely eliminate the

confounding difficulty of seasonal variations in ACM, which occur in non-equatorial countries.

Since the dominant period of the seasonal variations in ACM is 1 year, and since we wish to detect changes moving forward in time, we adopt the following approach. We apply a 1-year backward moving average to the ACM by time data. Each point in time of the 1-year backward moving average is simply the average ACM for the year ending at the said point in time, and we plot this moving average by time. Changes in regime of ACM then appear as breaks (in slope or value) in the moving average by time.

Note that the 1-year backward moving average method produces one significant but easily discerned artifact: Relatively large and sharp peaks in ACM give rise to artificial drops in the moving average at one year ahead of (later than) the said relatively large and sharp peaks in ACM.

4. Methods to Quantify vDFR from All-Cause Mortality

4.1 Historical-trend baseline for a period (or peak) of mortality (Method 1)

Our first method (Method 1) for quantification of vDFR by age group (or all ages) and by vaccine dose number (or all doses) is as follows (Rancourt et al., 2022a, 2023), here improved to adjust for systematic seasonal effects:

- i. Plot the ACM by time (day, week, month) for the age group (or all ages) over a large time scale, including the years prior to the declared pandemic.
- ii. Identify the date (day, week, month) of the start of the vaccine rollout (first dose rollout) for the age group (or all ages).
- iii. Note, for consistency, that the ACM undergoes a step-wise increase to larger values near the date of the start of the vaccine rollout.

- iv. Integrate (add) ACM from the start of the vaccine rollout to the end of available data or end of vaccinations (all doses), whichever comes first. This is the basic integration time window used in the calculation, start to end dates.
- v. Apply this window and this integration over successive and non-overlapping equal-duration periods, moving as far back as the data permits.
- vi. Start each new integration window at the same point in the seasonal cycle as the start of the basic integration window for the vaccine period, even if this introduces gaps between successive integration periods.
- vii. Plot the resulting integration values versus time, and note, for consistency, that the value has an upward jog, well discerned from the historic trend or values, for the vaccination period.
- viii. Extrapolate the historic trend of integrated values into the vaccination period. The difference between the measured and extrapolated (historic trend predicted) integrated values of ACM in the vaccination period is the excess mortality associated with the vaccination period.
- ix. The extrapolation, in practice, is achieved by fitting a straight line to chosen pre-vaccination-period integration points.
- x. If too few points are available for the extrapolation, giving too large an uncertainty in the fitted slope, then impose a slope of zero, which amounts to using an average of recent values. In some cases, even a single point (usually the point for the immediately preceding integration window) can be used.
- xi. The error in the extrapolated value is most often overwhelmingly the dominant source of error in the calculated excess mortality. Estimate the “accuracy error” in the extrapolated value as the mean deviation of the absolute value difference with the fitted line (mean of the absolute values of the residuals) for the chosen points of the fit. This error is a measure of the integration-period variations from all causes over a near region having an assumed linear trend.
- xii. The said “accuracy error” is generally larger than the “precision error” (or statistical error) in the extrapolated value, as it represents the year-to-year

- variability of the integrated ACM in the integration window in the years prior to the Covid or vaccination periods.
- xiii. If there are too few integration windows in the available normal years prior to the peak or region of interest to obtain a good estimate of the historic year-to-year variability, or if the statistical errors in the integrated values are relatively large, then make use of the statistical errors to best estimate the needed uncertainty.
 - xiv. Apply the same integration window (start-to-end dates during vaccination) to count all vaccine doses administered in that time.
 - xv. Depending on particular circumstances in the data, it may be necessary to use different integration bounds (different windows) for the ACM and for the vaccine administration. We saw no need for this, and we did not try to implement or test such an optimization.
 - xvi. Define $vDFR = (\text{vaccination-period excess mortality}) / (\text{vaccine doses administered in the same vaccination period})$. Calculate the uncertainty in $vDFR$ using the estimated error in vaccination-period excess mortality.

The same method is adapted to any region of interest (such as a peak in ACM) of sub-annual duration, by translating the window of integration (of the region of interest) backwards by increments of one year.

The above-described method is robust and ideally adapted to the nature of ACM data. Integrated ACM will generally have a small statistical error.

A large time-wise integration window (e.g., for the entire vaccination period) mostly removes the difficulty arising from intrinsic seasonal variations; and this difficulty is further solved by starting each new integration window at the same point in the seasonal cycle as the start of the basic integration window for the vaccine period (point-vi, above).

The historic trend is analysed without introducing any model assumptions or uncertainties beyond assuming that the near trend can be modelled by a straight line, where justified by the data itself. Such an analysis, for example, takes into account year to year changes in age-group cohort size arising from the age structure of the population. The only assumption is that a locally linear near trend for the unperturbed (ACM-wise unperturbed) population is realistic.

While the above method is designed for cases (jurisdictions) in which there is no evidence in the ACM data for mortality caused by factors other than the vaccine rollouts, such as Covid measures (treatment protocols, societal impositions, isolation and so forth; since no excess mortality occurs in the pre-vaccination period of the Covid period), it can be readily adapted to cases in which mortality in the vaccination period is confounded by additional (Covid period) causal factors that cannot be ruled out.

One approach is simply to adapt the above method to calendar years, irrespective of whether excess mortality occurs prior to the COVID-19 vaccine rollouts. One obtains excess ACM by calendar year, relative to the expected value from the historic trend deduced by linear extrapolation from a chosen range of yearly ACM values for < 2020 (for years prior to 2020, when the 11 March 2020 announcement of a pandemic was made). One then compares the excess ACM for 2020 and for 2021. In many (most) countries, there was essentially no COVID-19 vaccination in 2020, and a rapid rollout essentially started in January 2021.

4.2 Special case of a single historic integrated point (Method 2)

In cases in which it is not possible or practical to obtain more than one integration value for the needed extrapolation (steps v to ix, above), rather than assume a zero slope for the extrapolation (step x, above), the following second method (Method 2) can be applied.

If $Y(-1)$ is the sole historic integrated point, then simply take the needed extrapolated value, $Y(0)$, to be:

$$Y(0) = Y(-1) + m \Delta T W \quad (1)$$

where m is the slope of the best-straight-line fit through the original ACM by time unit (day, week, month...) versus numbered time unit, ΔT is the number of time units between $Y(0)$ and $Y(-1)$ (i.e., between the start of the $Y(0)$ integration window and the start of the $Y(-1)$ integration window), and W is the inclusive width of the integration window in number of time units.

This assumes that the ACM by time varies on a straight line, notwithstanding seasonal variations, on the near segment used to obtain the best-straight-line fit.

The resulting excess mortality for the integration window or period, $xACM(0)$, is then:

$$xACM(0) = ACM(0) - Y(0) \quad (2)$$

where $ACM(0)$ is the integrated ACM in the period of interest.

The statistical error (standard deviation) in $xACM(0)$ is then given by:

$$\text{sig}(xACM(0)) = \sqrt{[ACM(0) + Y(-1) + (\Delta T W \text{ sig}(m))^2]} \quad (3)$$

where $\text{sig}(m)$ is the nominally statistical error in m .

If there is no seasonal variation in ACM, as occurs in equatorial-latitude jurisdictions, then $\text{sig}(m)$ is the actual statistical error in m . With seasonal variations in ACM, $\text{sig}(m)$ extracted from the least squares fitting to a straight line does not have a simple

meaning. In this case, $\text{sig}(m)$ will incorporate uncertainty arising from seasonal variations, and increases with increasing amplitude of the seasonal variation.

4.3 Application of the methods to the specific countries

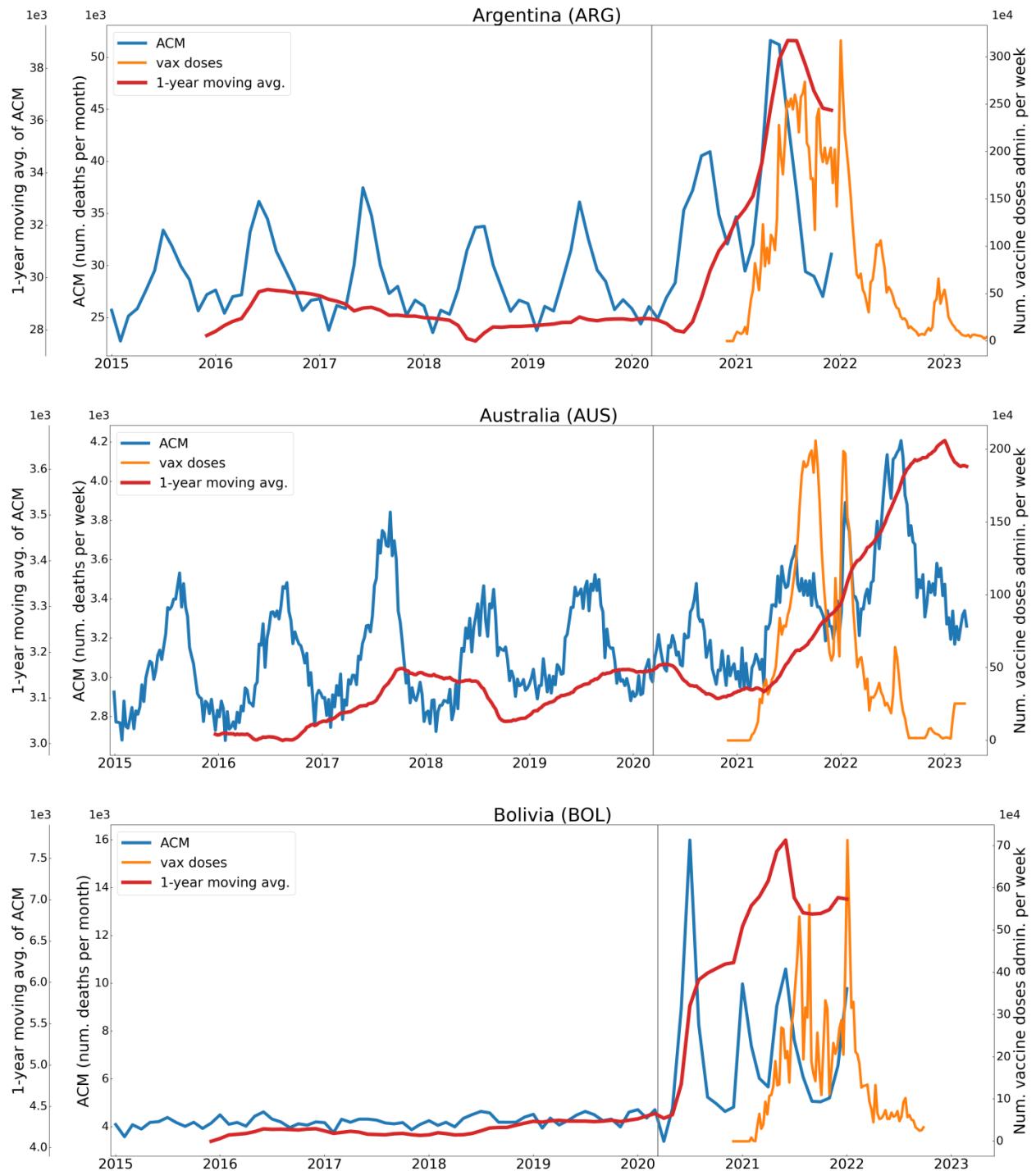
The parameters for applying the methods (Methods 1 and 2) to the data are given in Appendix C: Technical and specific information for applications of the methods to the data.

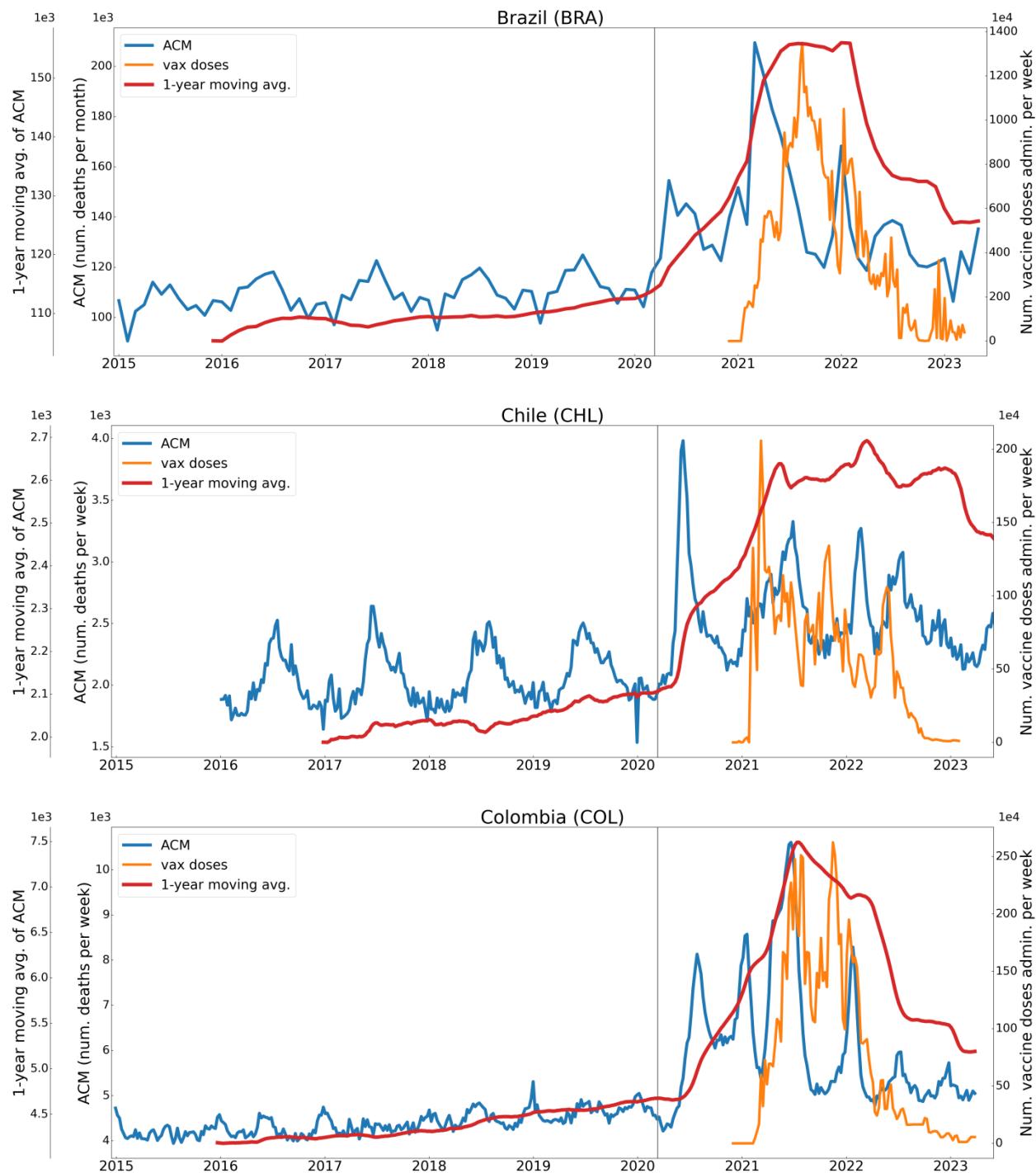
5. Results

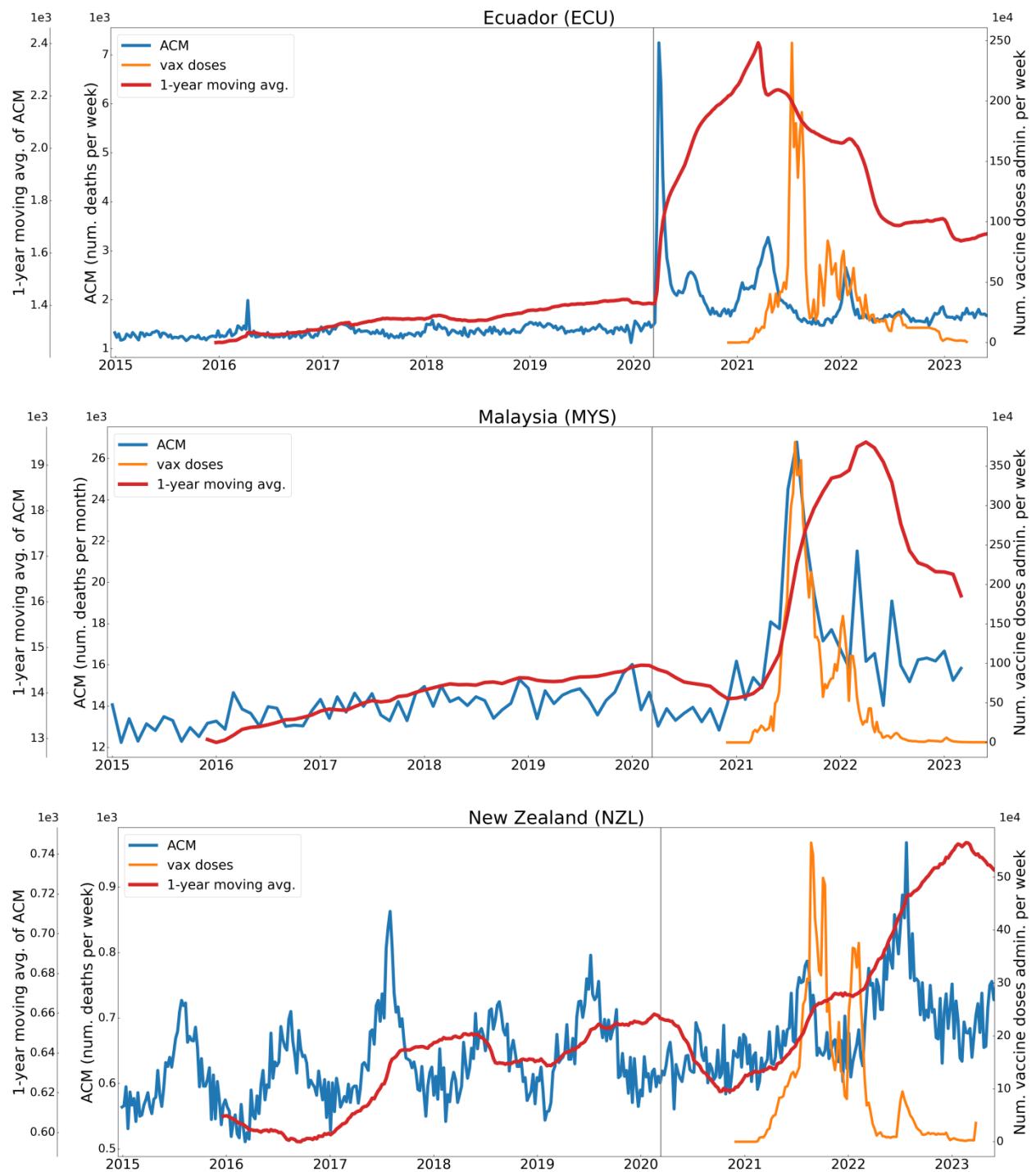
5.1. Transitions to regimes of high ACM

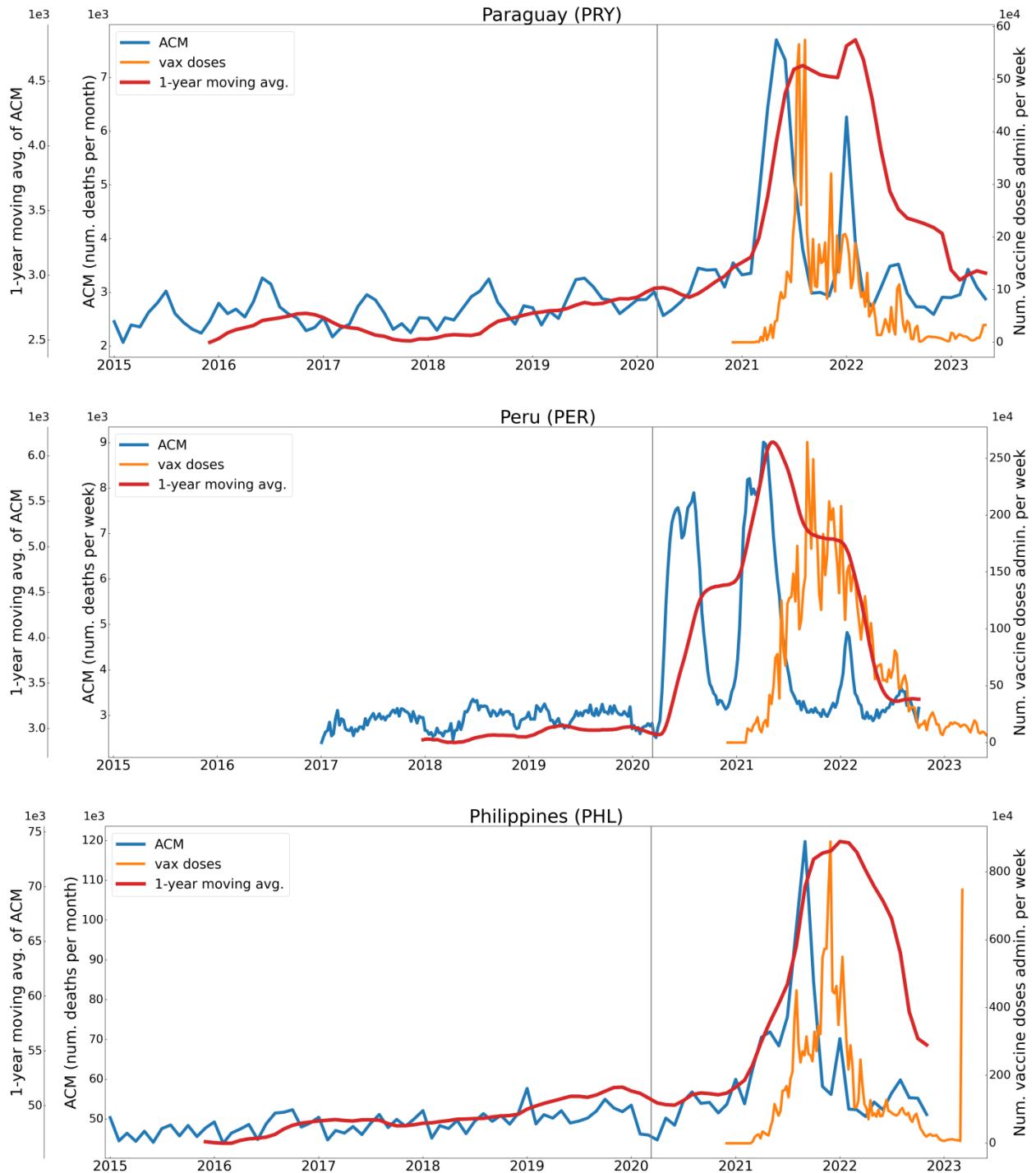
Figure 2 shows the ACM by time (week or month) (blue), vaccine administration by time (week) (orange) (source: OWID, 2023a), and the 1-year backward moving average of the ACM by time (week or month, respectively) (red), for the 17 countries in this study. The 11 March 2020 pandemic declaration date is shown by a vertical grey line in each panel.

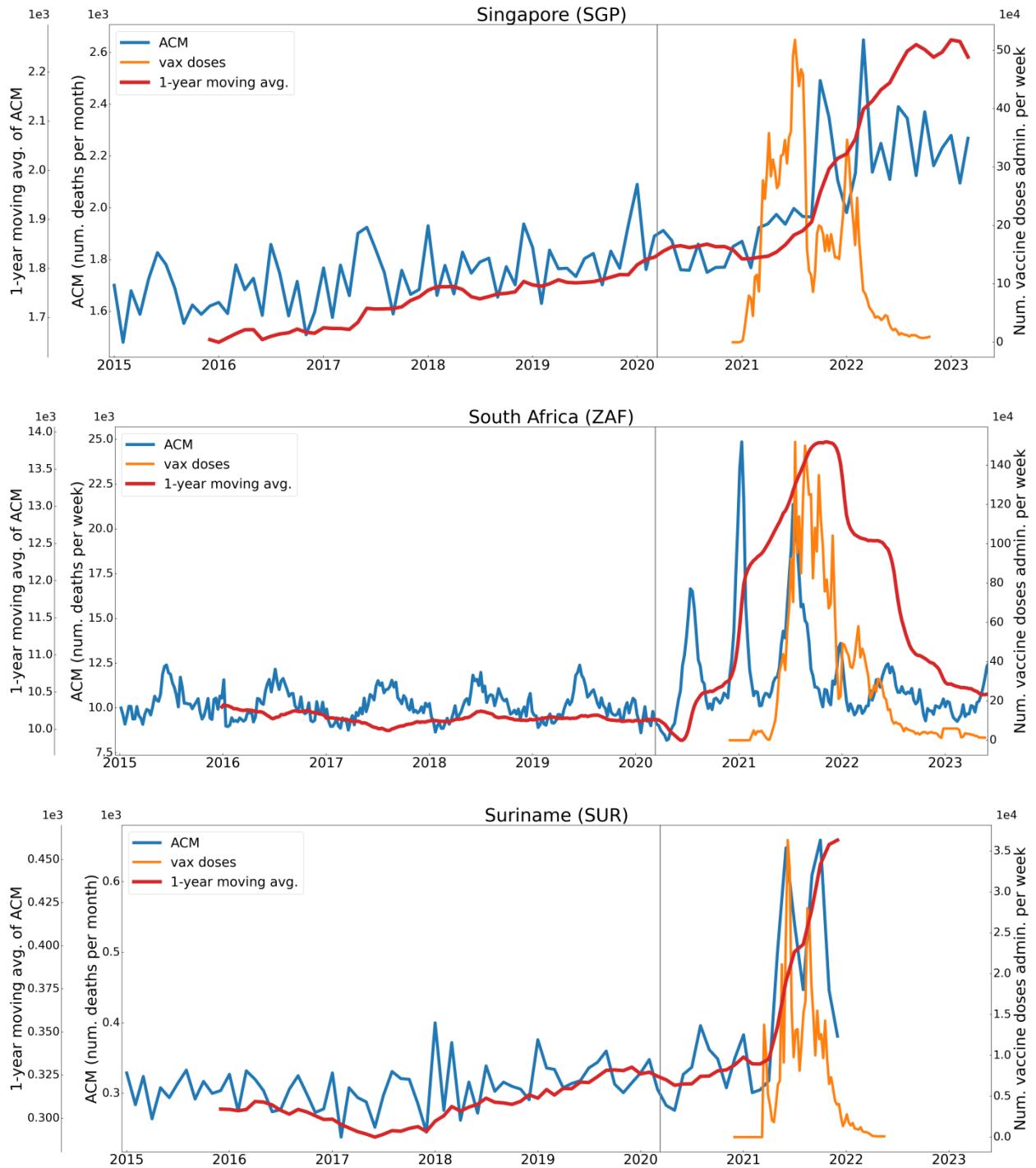
The vaccine administration by week (e.g., orange, Figure 2), for all-ages analyses in the present paper, is obtained from the original cumulative data (OWID, 2023a) by interpolating to obtain all dates, and then summing by week. As a result, where there are sudden jumps in the cumulative data, this can produce a large weekly value as an artifact, such as for the Philippines (Figure 2). Similarly, drops in cumulative values can produce artificial negative weekly values, as seen in a few cases, below.











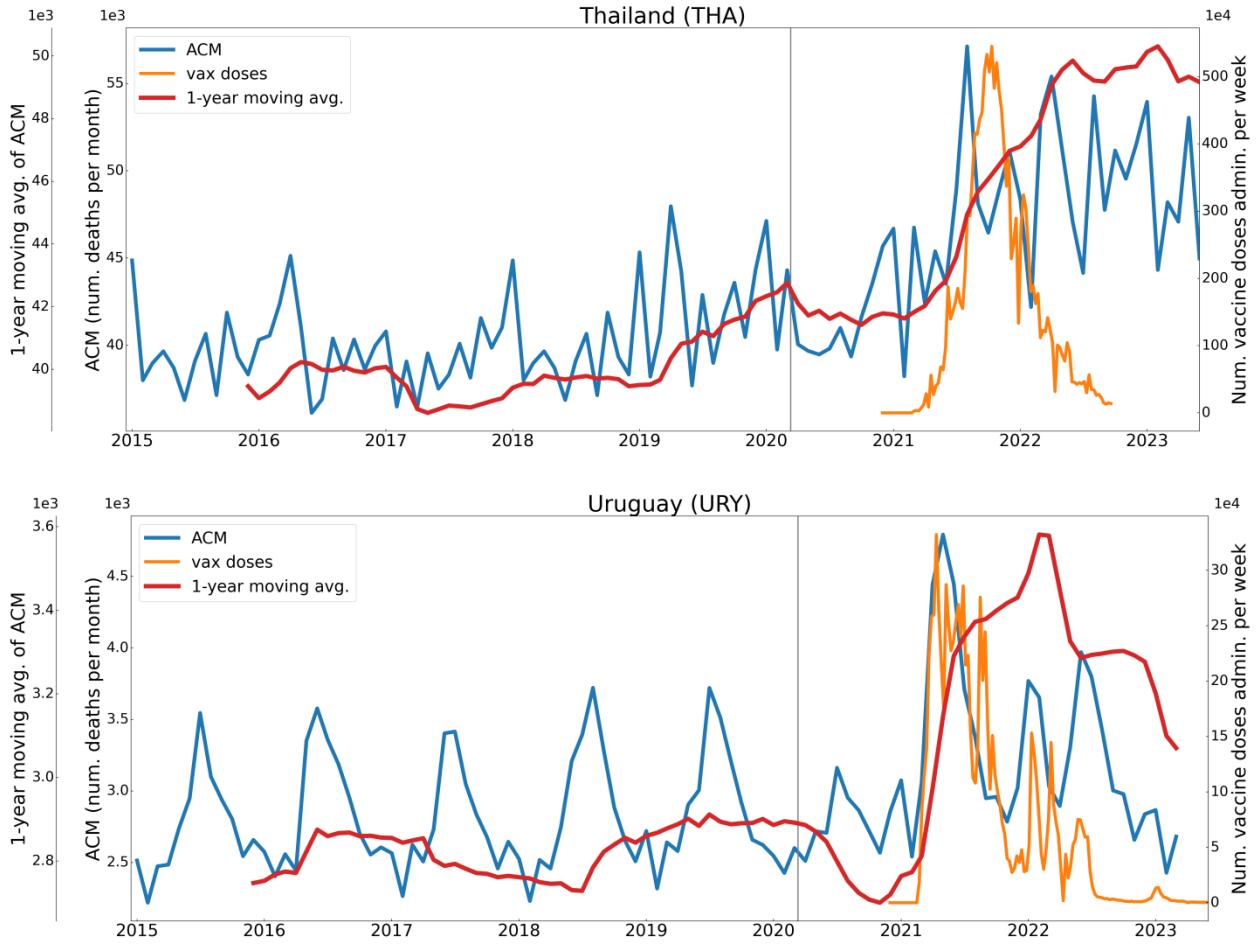


Figure 2: Transitions between regimes of mortality — ACM by time (week or month) (blue), vaccine administration by time (week) (orange), and the 1-year backward moving average of the ACM by time (week or month) (red), as per the legend, for the 17 countries in this study. The 11 March 2020 pandemic declaration date is shown by a vertical grey line in each panel. The data sources are specified in Appendix A.

Here (Figure 2), 9 of the 17 countries have no detectable excess mortality in the year or so between when a pandemic is announced on 11 March 2020 and the starting time of the first vaccine rollout in each country. That is, in 9 of the 17 countries, there is no detectable excess mortality until the vaccines are rolled out (Australia, Malaysia, New Zealand, Paraguay, Philippines, Singapore, Suriname, Thailand, Uruguay).

In the other 8 of the 17 countries, a new regime of higher mortality is initiated after 11 March 2020 and prior to any COVID-19 vaccine administration (Argentina, Bolivia, Brazil, Chile, Colombia, Ecuador, Peru, South Africa).

In all 17 countries, vaccination is associated with a regime of high mortality, and there is no association in time between COVID-19 vaccination and proportionate reduction in ACM.

5.2 Covid-period excess mortality compared to COVID-19-vaccination-period excess mortality, in the 17 countries of this study

We apply the “Special case of a single historic integrated point” method (Method 2, or “single-point method”, Equations 1 to 3, see Methods) to quantify the excess ACM in the entire Covid period (or “Covid period”), from the week or month of 11 March 2020 to the week or month, respectively, of the last (most recent) usable or chosen point in the ACM by time data. This method takes into account the recent decadal linear change in historic ACM in the needed extrapolation for the reference mortality.

The method and calculation of the Covid-period excess ACM are illustrated in Appendix D: Single-point method quantification of the excess mortality in the Covid period, for all 17 countries in this study.

We perform the same calculation for the entire COVID-19-vaccination-period (or “vaccination period”) excess ACM, from the first week or month of COVID-19 vaccination to the same week or month, respectively, which was taken to be the last week or month, respectively, of the Covid period for each country, which is illustrated in Appendix E: Single-point method quantification of the excess mortality in the vaccination period.

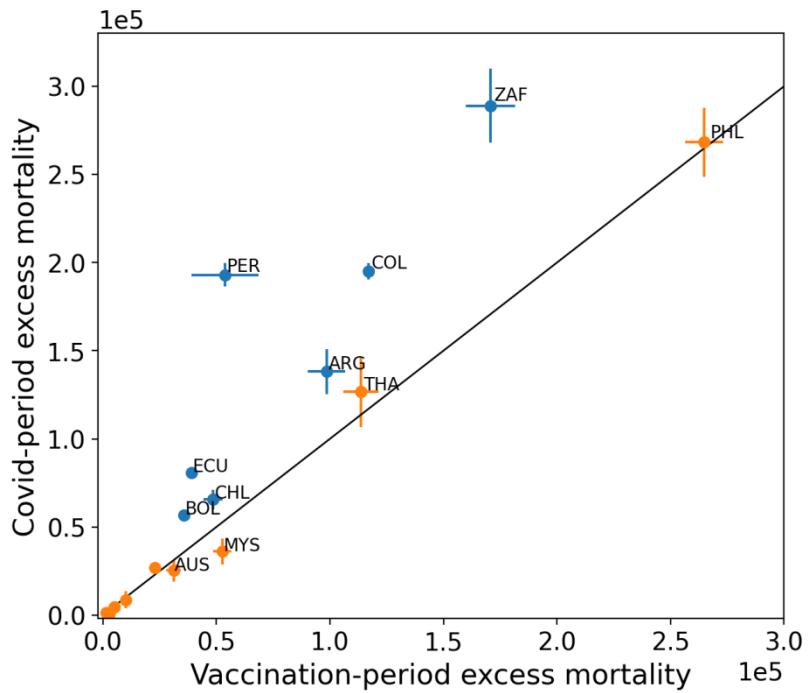
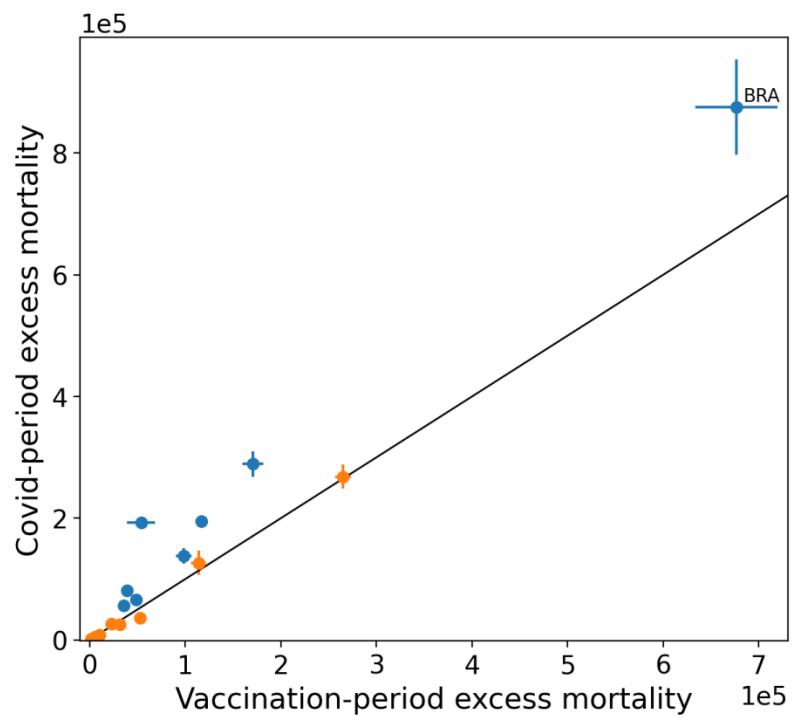
The results of the calculations are given in Table 1, for the excess ACM in the Covid period (column labelled “Covid period ACM”) and for excess ACM in the vaccination period (column labelled “Vaccination period ACM”), and their associated errors (Equation 3, Methods). Total numbers of COVID-19 vaccine injections for the same

vaccination periods are also given (column labelled “Vaccination period injections”), as are projected populations for 2022 (United Nations, 2023). Both the Covid and vaccination periods end on the same date for each country. The totals are also given, with propagated errors. The orange shading corresponds to those countries that have no apparent excess ACM prior to the start of vaccination, pursuant to Figure 2, whereas the blue shading is for countries that exhibit excess ACM prior to the start of vaccination (and after a pandemic is declared on 11 March 2020).

Table 1: Total excess mortalities in the Covid and vaccination periods, number of injections in the vaccination period, population

Country	Country code	Covid excess ACM	error	Vaccination excess ACM	error	Vaccination period injections	Population
Argentina	ARG	138,181	12,674	98,511	8,120	78,405,465	45,510,318
Australia	AUS	25,305	6,186	31,148	3,262	65,492,360	26,177,413
Bolivia	BOL	56,701	944	35,696	610	11,860,582	12,224,110
Brazil	BRA	875,876	78,167	676,668	42,828	486,436,436	215,313,498
Chile	CHL	65,604	5,513	48,659	4,186	62,673,028	19,603,733
Colombia	COL	195,020	4,704	116,994	2,575	90,411,911	51,874,024
Ecuador	ECU	80,835	1,739	39,052	1,290	39,502,549	18,001,000
Malaysia	MYS	36,096	7,349	52,651	4,087	72,552,897	33,938,221
New Zealand	NZL	735	1,547	2,805	855	11,996,667	5,185,288
Paraguay	PRY	26,797	3,344	22,837	1,860	9,621,361	6,780,744
Peru	PER	193,025	6,686	53,810	14,701	84,750,059	34,049,588
Philippines	PHL	268,192	19,719	264,748	8,287	169,068,966	115,559,009
Singapore	SGP	4,559	879	5,158	440	14,727,467	5,975,689
South Africa	ZAF	289,012	20,993	170,667	10,923	38,678,499	59,893,885
Suriname	SUR	1,546	166	1,488	101	490,144	618,040
Thailand	THA	126,698	19,903	113,736	7,740	142,635,014	71,697,030
Uruguay	URY	8,648	4,790	10,201	2,640	8,841,960	3,422,794
TOTAL		2,392,831	87,904	1,744,829	49,285	1,388,145,365	725,824,384

The results of Table 1 are illustrated in Figure 3. Here, the same colour coding of the symbols as for Table 1 is used: orange points for countries that have no apparent excess ACM prior to the start of vaccination, pursuant to Figure 2, and blue points for countries that exhibit excess ACM prior to the start of vaccination (and after a pandemic is declared on 11 March 2020).



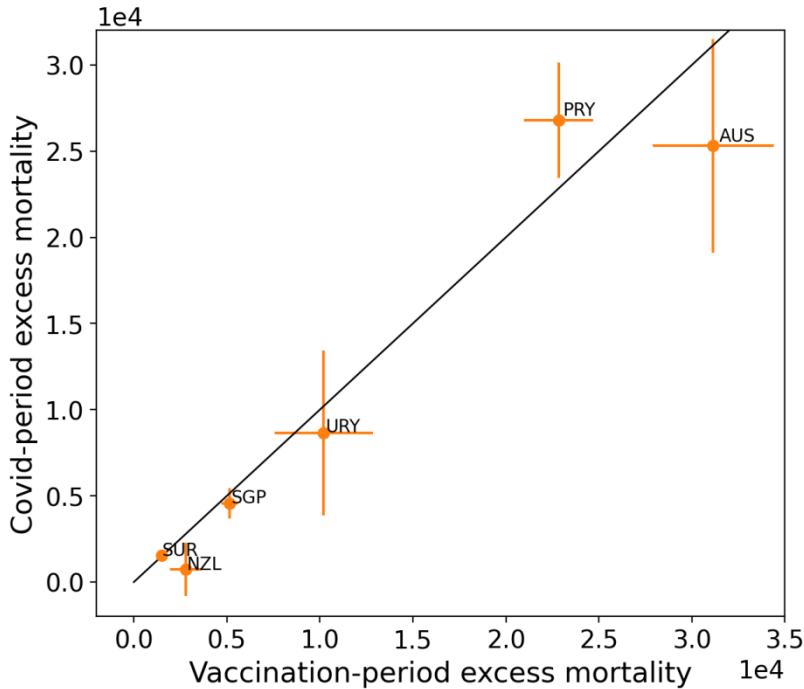
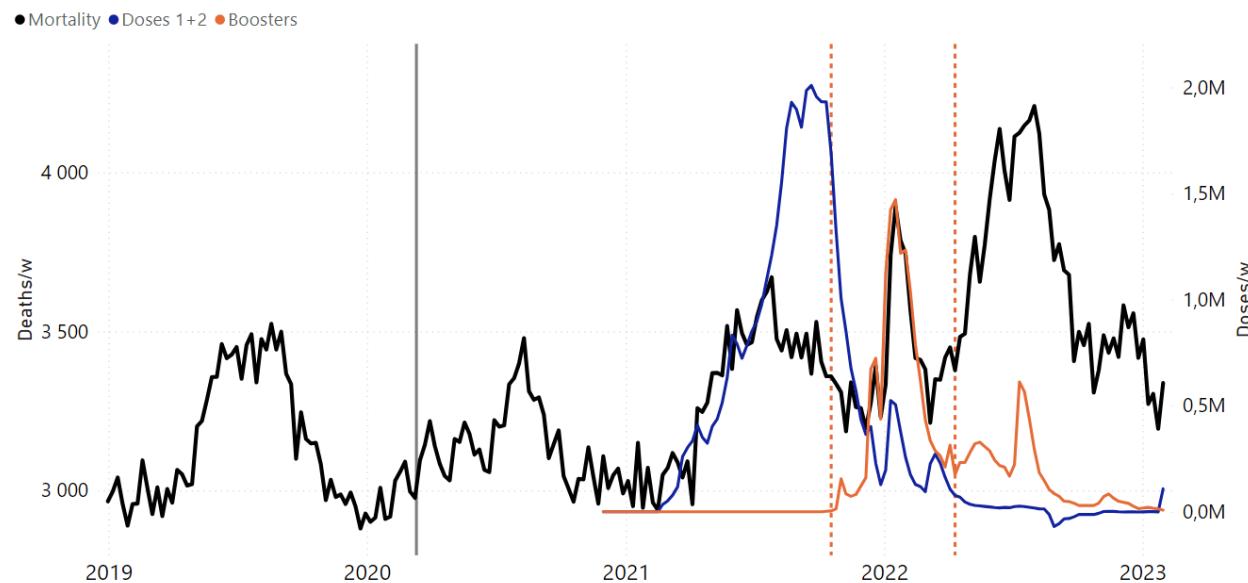


Figure 3: (three panels) Covid-period excess ACM versus vaccination-period excess ACM (top), first expanded view (middle), second expanded view (bottom), with 1:1 line as a reference.

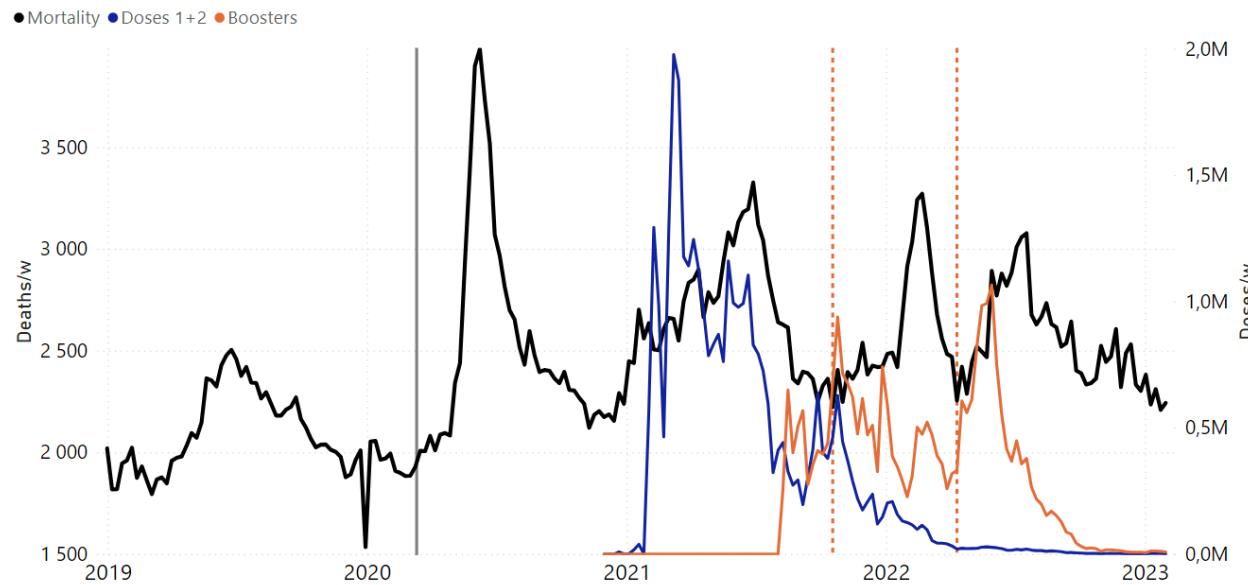
5.3 All studied countries all-ages national ACM and vaccine rollouts

Figure 4 shows the available all-ages national ACM (black) by month (2017-2023) or by week (2019-2023), compared to the COVID-19 vaccine-dose rollouts, for: Australia, Chile, Colombia, Ecuador, New Zealand, Peru, South Africa (by-week ACM data); and Argentina, Bolivia, Brazil, Malaysia, Paraguay, Philippines, Singapore, Suriname, Thailand, and Uruguay (by-month ACM data). Grey vertical lines indicate the 11 March 2020 World Health Organization (WHO) announcement of a pandemic, and orange dashed vertical lines show the bounds of the integration window for the January-February 2022 peak of interest (same bounds as for Figure 5, and Table 2). Vaccine administration for dose 1 and dose 2 by week or by month is in blue; whereas booster administration (doses beyond dose 2) by week or by month is in orange. The vaccine data is from Our World in Data (2023a), where the dose1 and dose 2 data (blue) was obtained by subtracting the booster doses (orange) from the all-doses data.

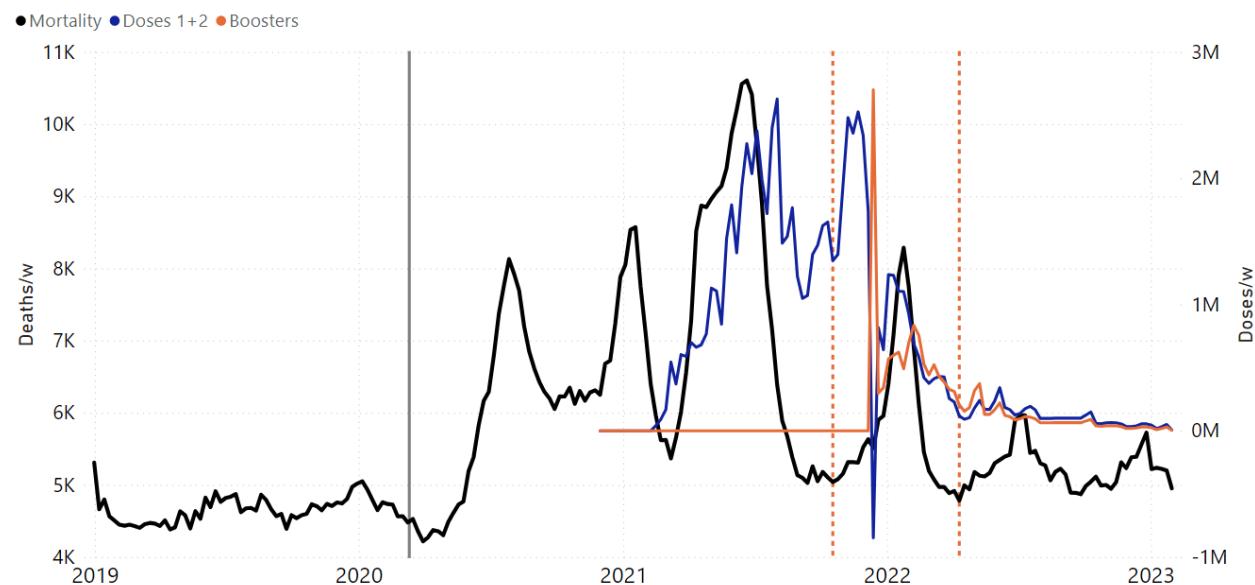
ACM and vaccination by week, Australia, 2019-2023



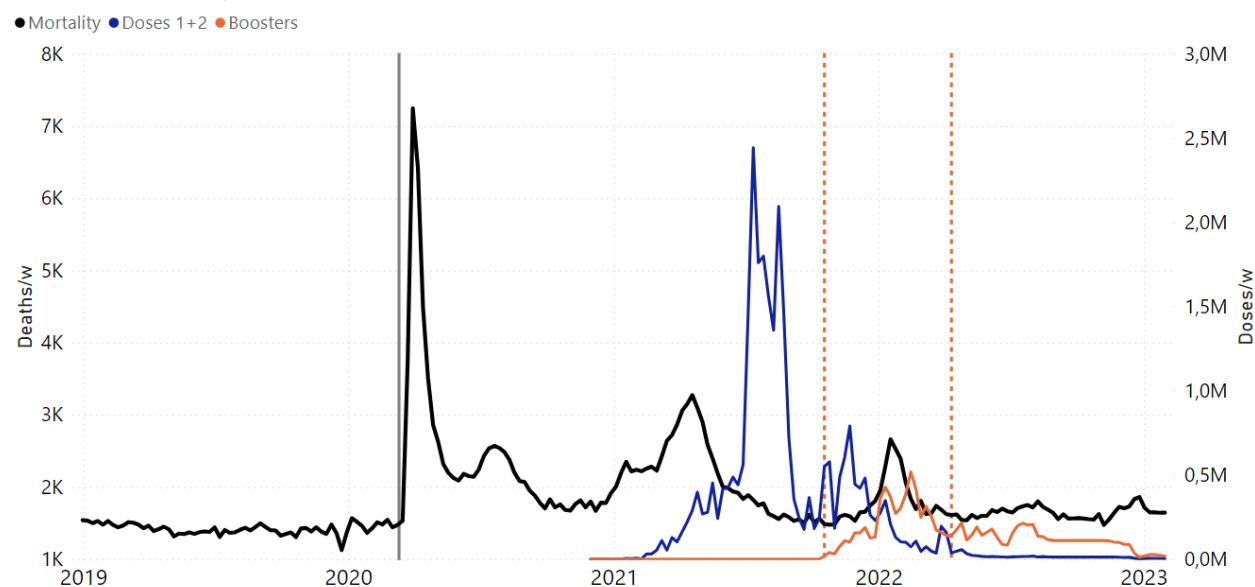
ACM and vaccination by week, Chile, 2019-2023



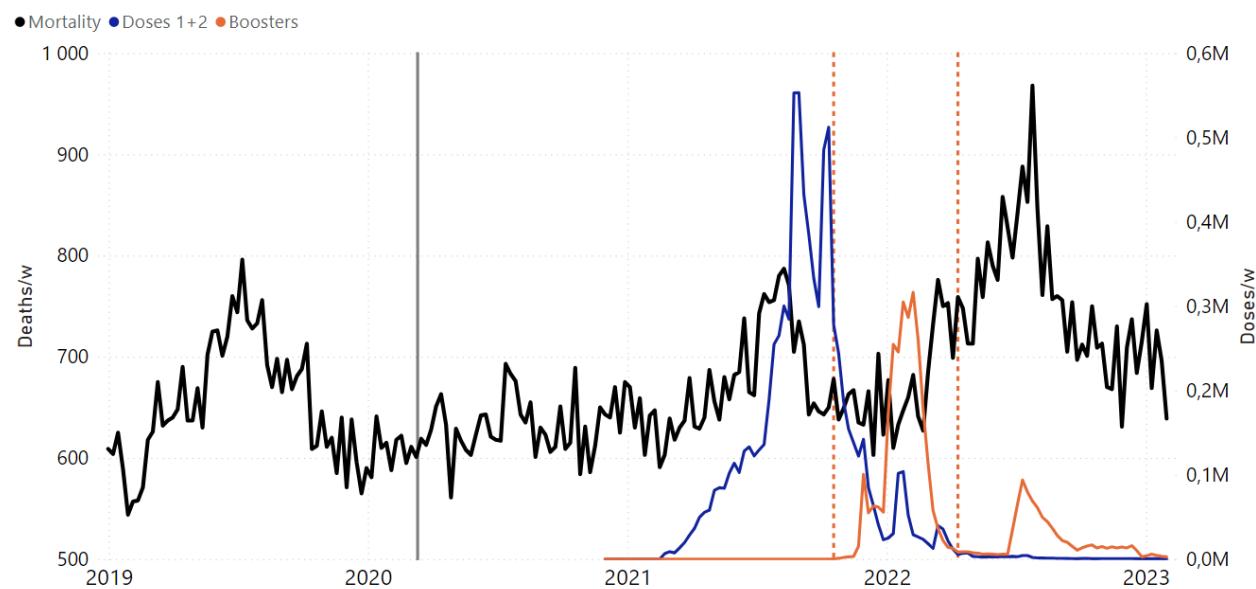
ACM and vaccination by week, Colombia, 2019-2023



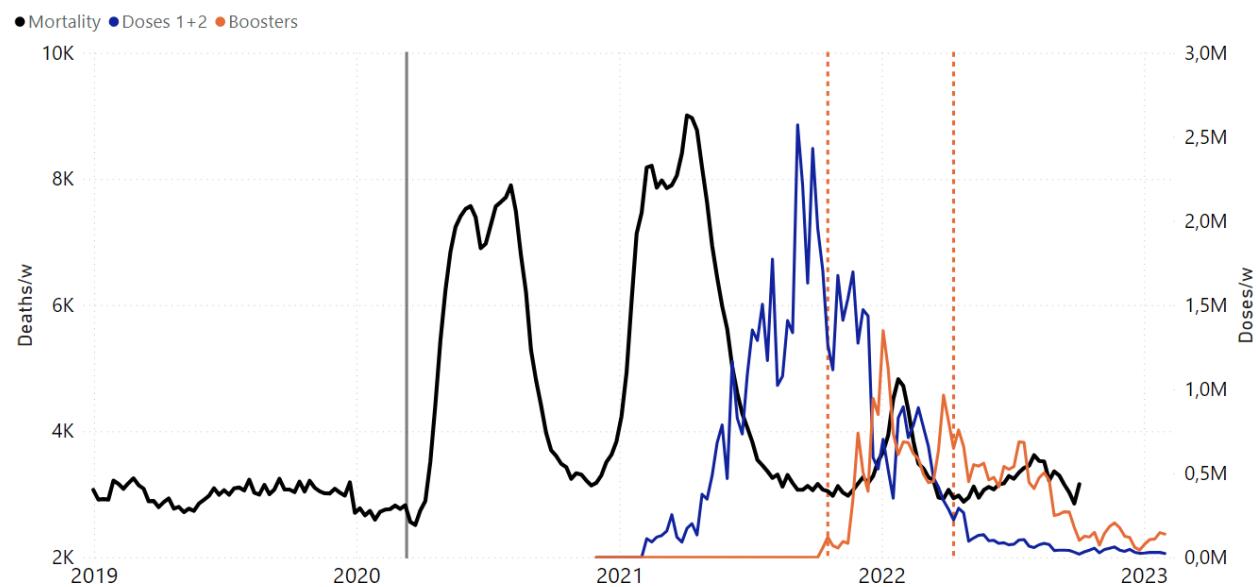
ACM and vaccination by week, Ecuador, 2019-2023



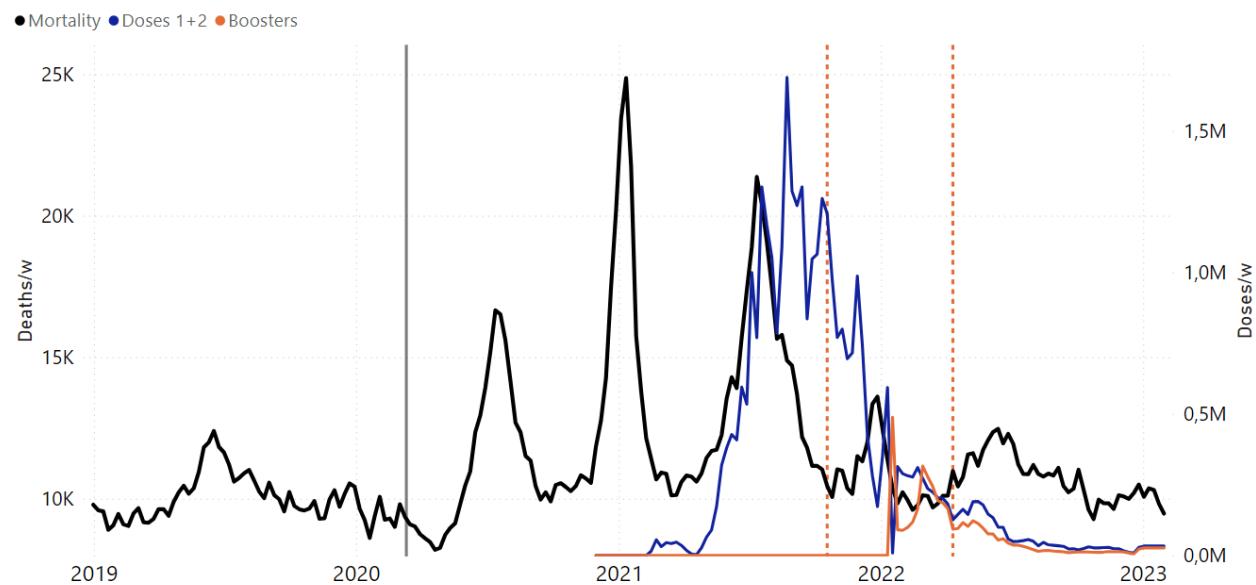
ACM and vaccination by week, New Zealand, 2019-2023



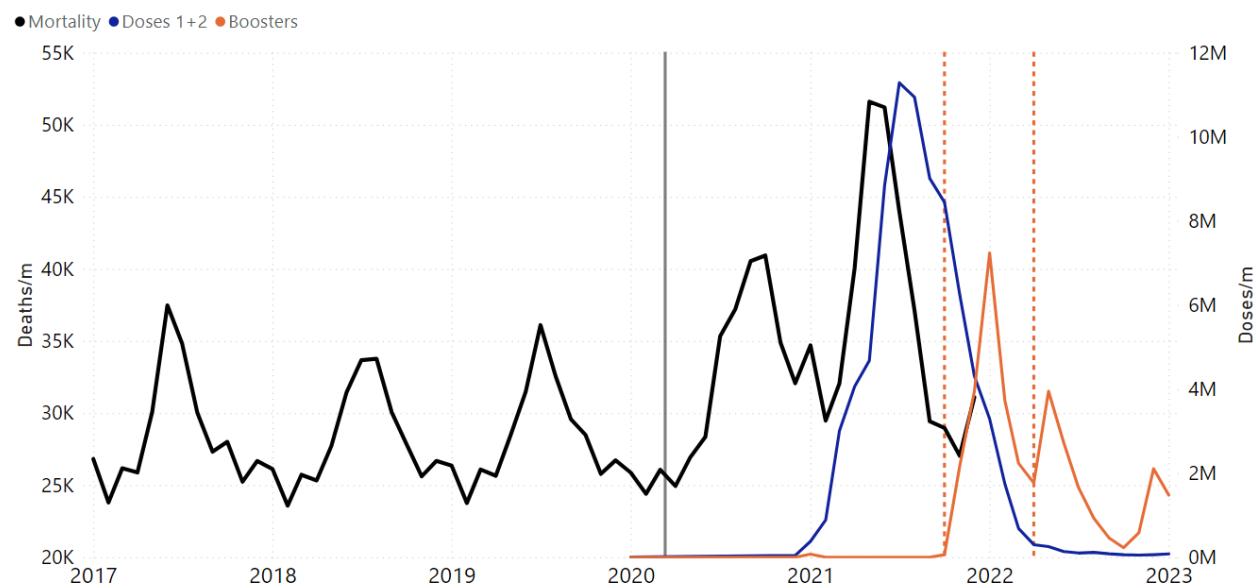
ACM and vaccination by week, Peru, 2019-2023



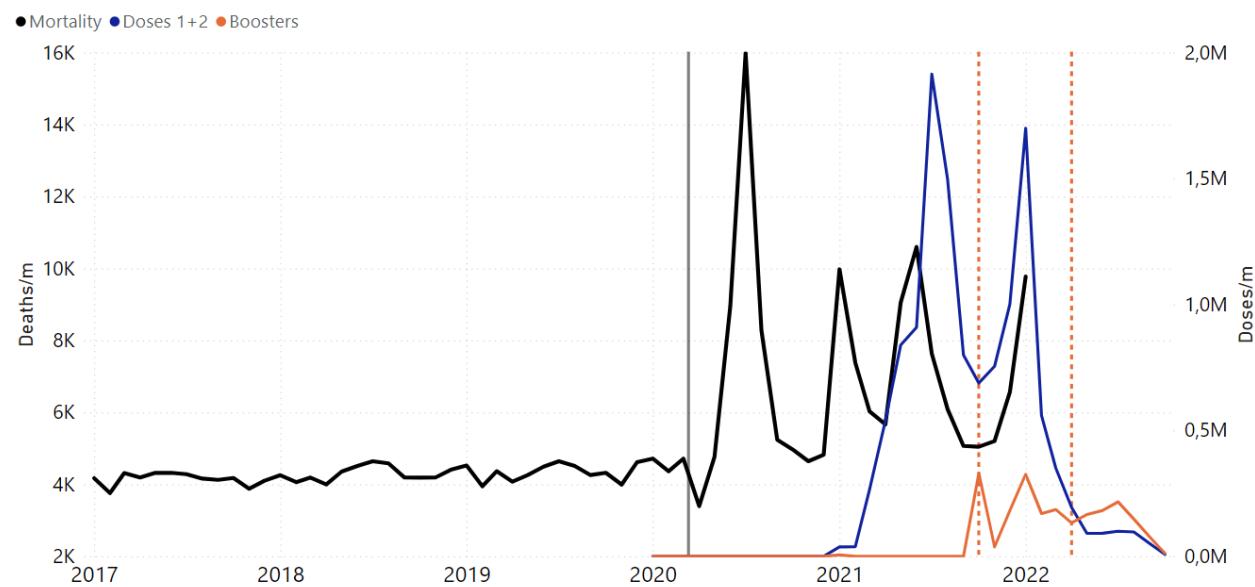
ACM and vaccination by week, South Africa, 2019-2023



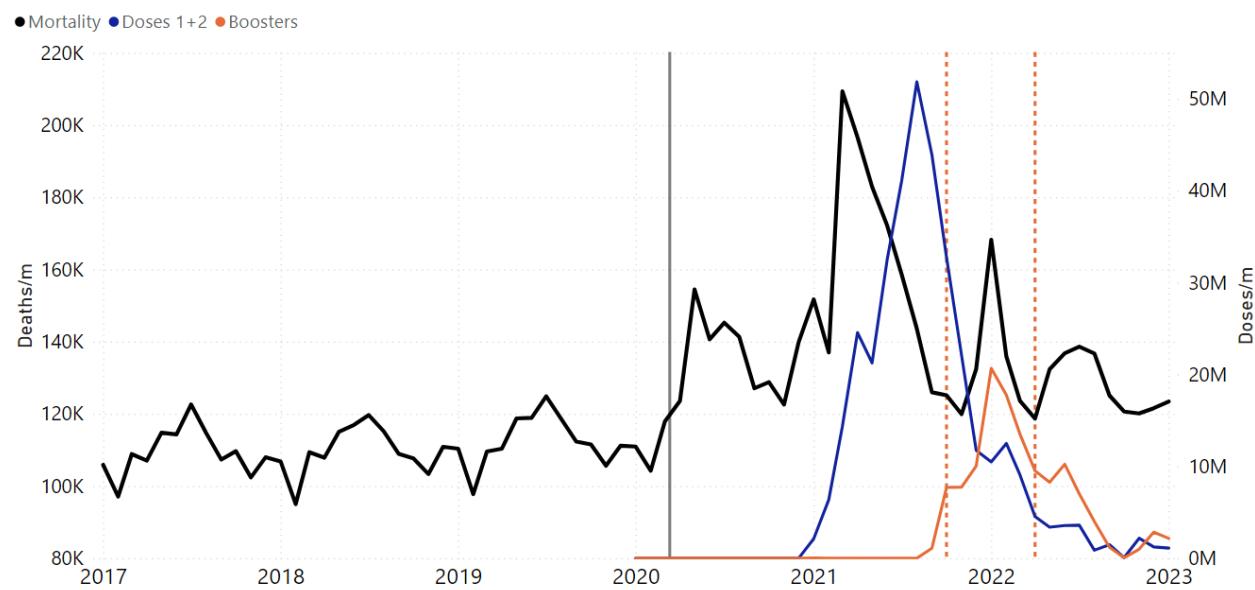
ACM and vaccination by month, Argentina, 2017-2023



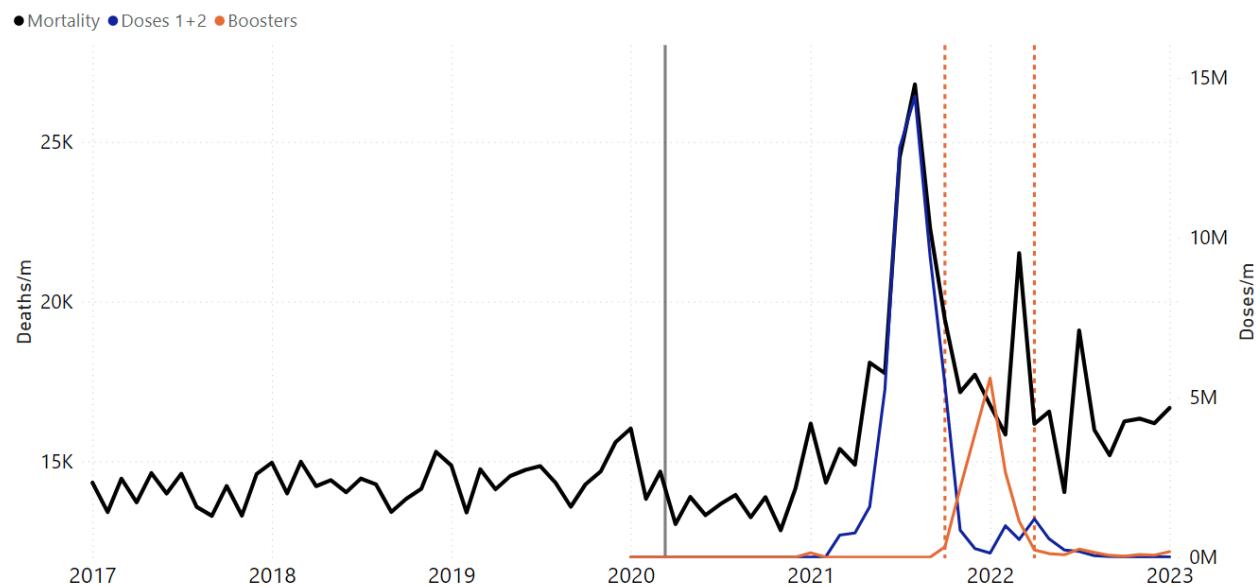
ACM and vaccination by month, Bolivia, 2017-2023



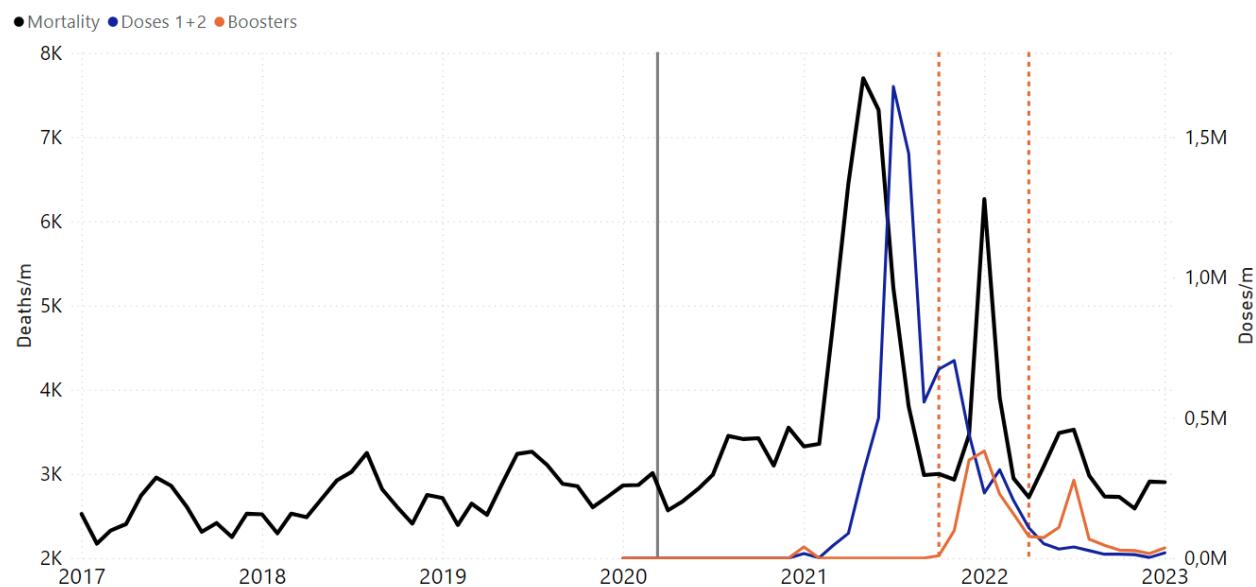
ACM and vaccination by month, Brazil, 2017-2023



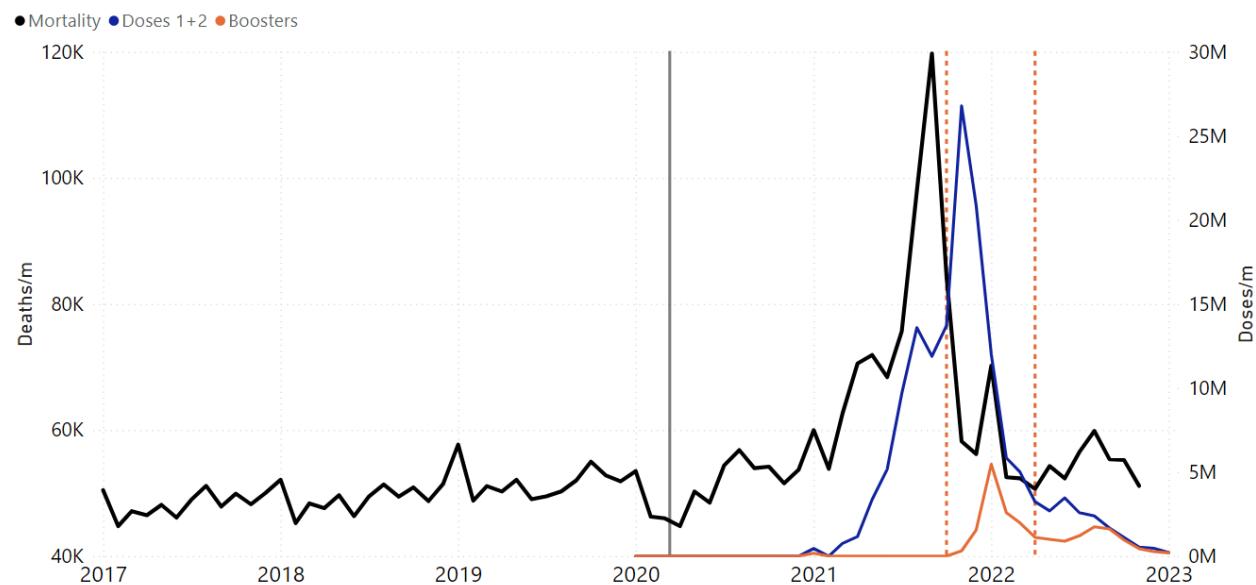
ACM and vaccination by month, Malaysia, 2017-2023



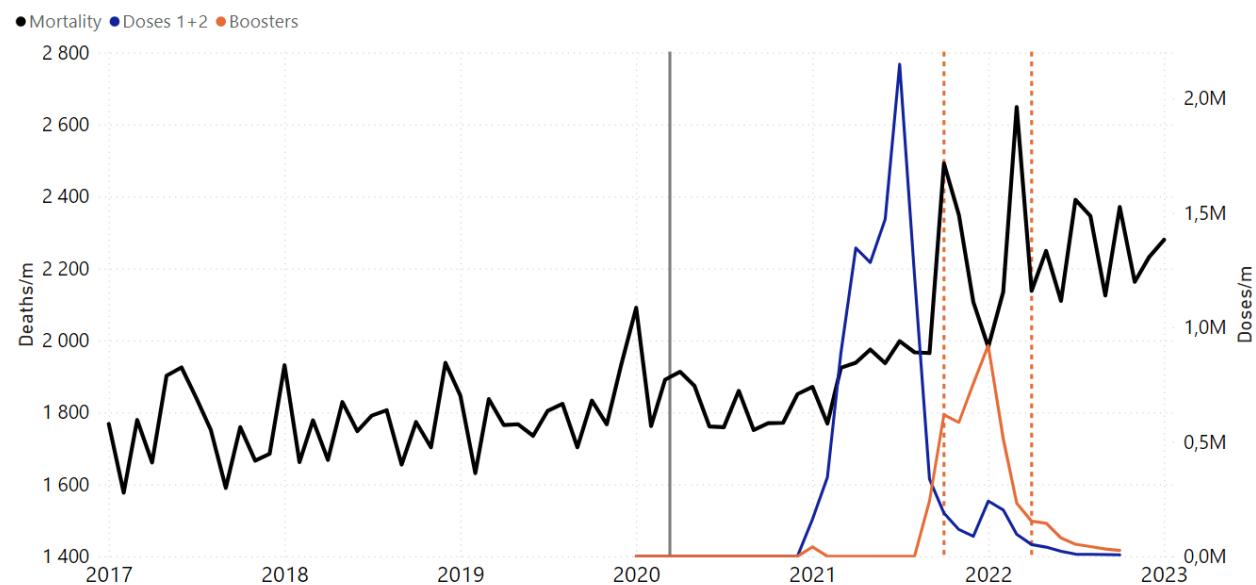
ACM and vaccination by month, Paraguay, 2017-2023



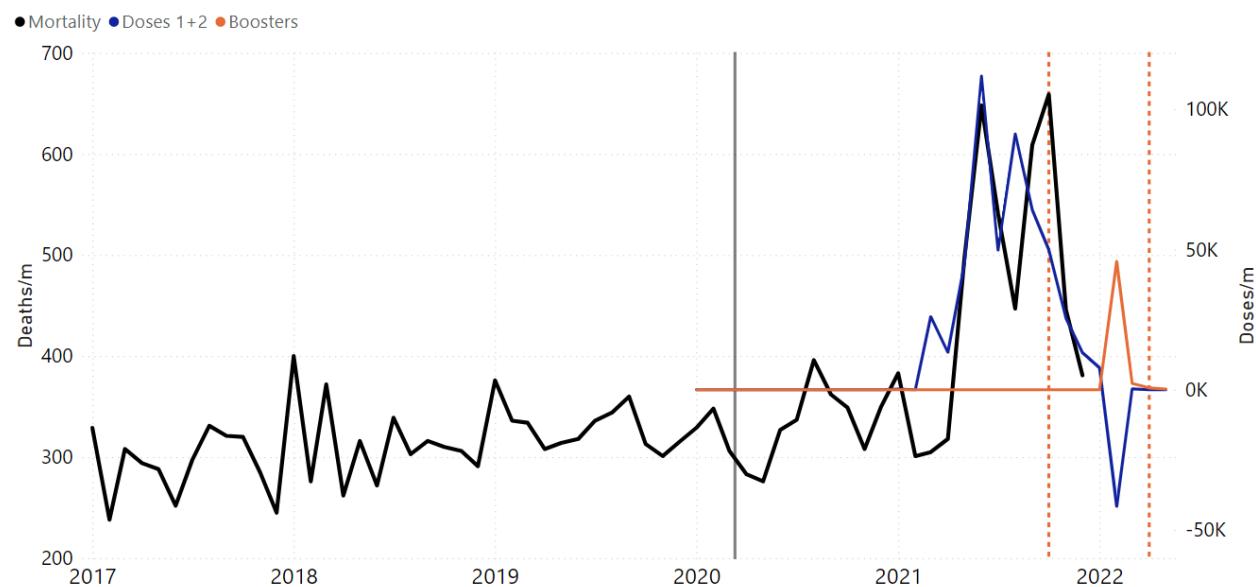
ACM and vaccination by month, Philippines, 2017-2023



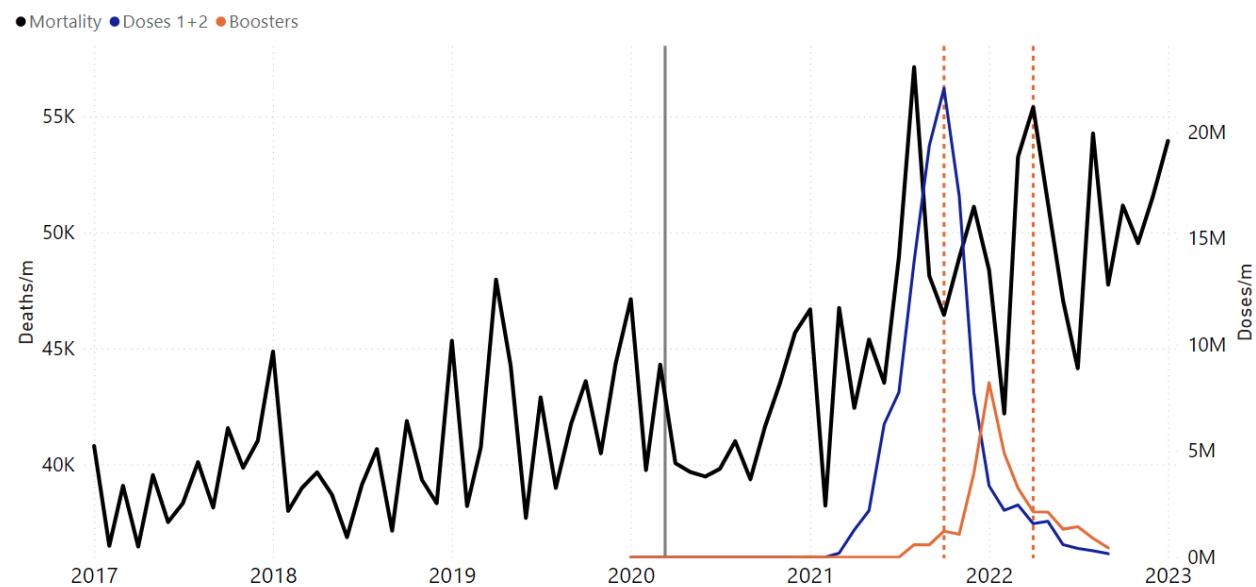
ACM and vaccination by month, Singapore, 2017-2023



ACM and vaccination by month, Suriname, 2017-2023



ACM and vaccination by month, Thailand, 2017-2023



ACM and vaccination by month, Uruguay, 2017-2023

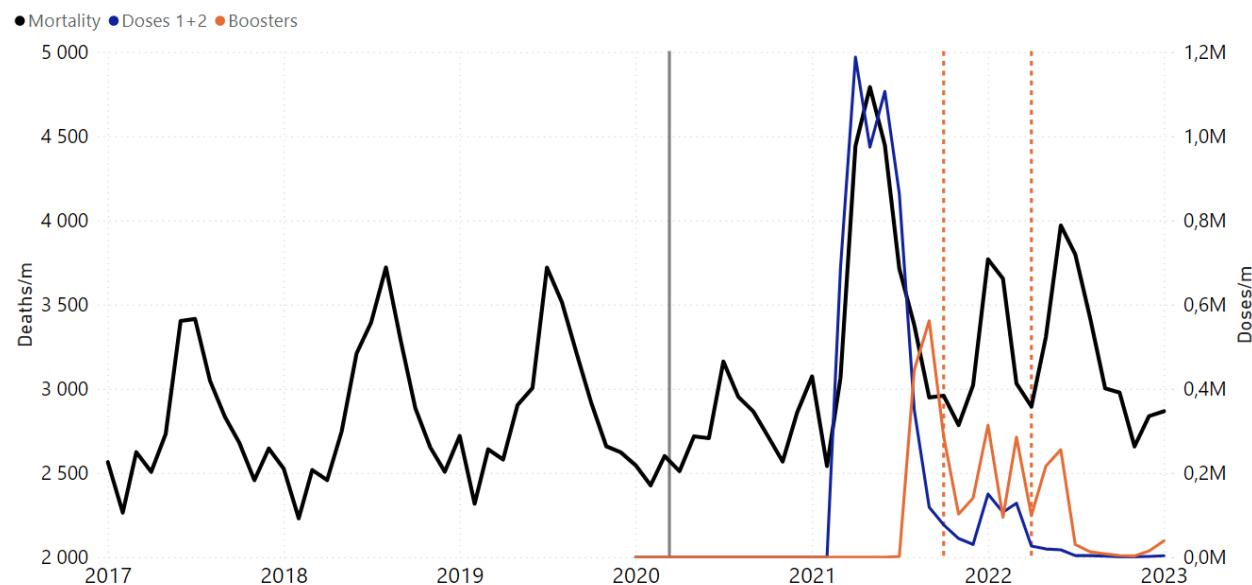


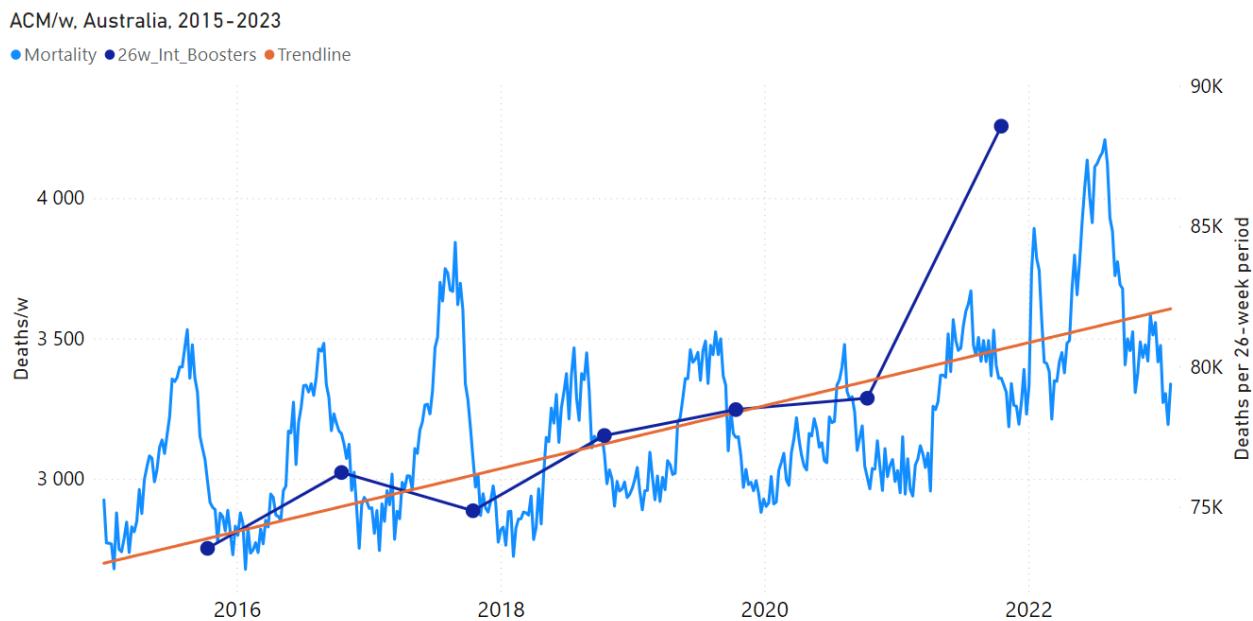
Figure 4: All-ages ACM by week (2019-2023) or by month (2017-2023) (black), with all-ages vaccine rollouts doses 1 and 2 (blue) and boosters (orange), for the 17 countries in this study (as labelled); also showing the 11 March 2020 date and limits of integration for the nominally January-February 2022 peak in ACM (vertical lines). The data sources are specified in Appendix A.

Figure 5 shows the available all-ages national ACM by time (week or month), from 2015 to 2023, the integration points that were used to quantify excess ACM in the January-February 2022 peak of interest, and the linear trend lines that were used to extrapolate to the expected values (“Historical-trend baseline for a period (or peak) of mortality” method, Method 1, see Methods), for: Australia, Chile, Colombia, Ecuador, New Zealand, Peru, South Africa (by-week ACM data); and for Argentina, Bolivia, Brazil, Malaysia, Paraguay, Philippines, Singapore, Suriname, Thailand, and Uruguay (by-month ACM data). The figure has: ACM by week or by month (light blue); integrated ACM by 26-week or 7-month vaccination-period integration window (dark blue, points), the last point being for the actual period of interest itself; extrapolation line used to calculate the excess ACM in the booster period (orange). An integration point is positioned in time at the start of the integration window. See Methods.

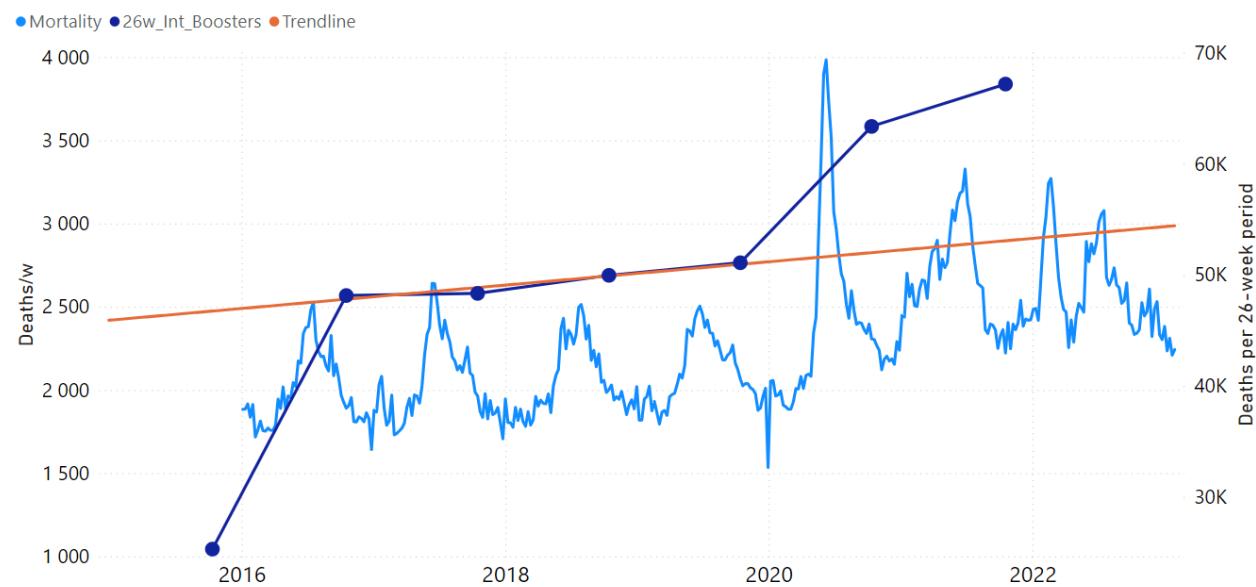
For completeness, in Figure 5 and Figure 4, we include those countries for which the calculation cannot be completed because of insufficient ACM data in the January-February 2022 region of interest. Argentina, Bolivia and Suriname have incomplete

ACM data in the integration region of interest, causing a low integration value as an artifact (Figure 5).

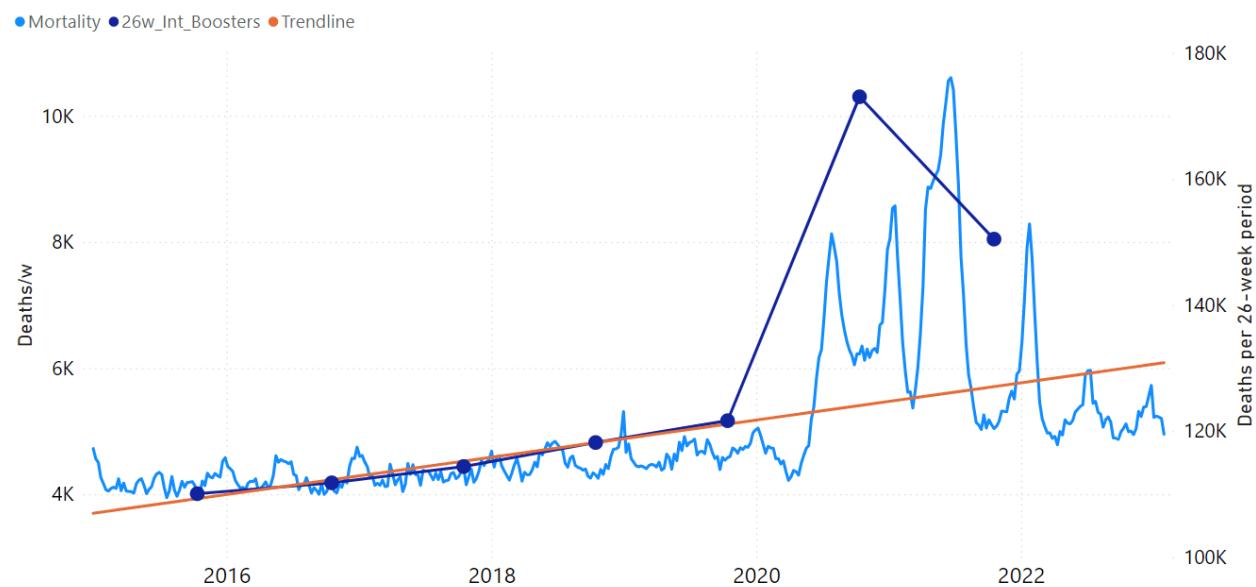
For simplicity and comparison, in Figure 5 and Figure 4 (and in Table 2), we standardized to essentially the same integration window for all countries (see Appendix C): 26-week window with by-week ACM data (2021-week-42 through 2022-week-15, which is 182 days) or 7-month window with by-month ACM data (2021-month-10 through 2022-month-04, which is 212 days).



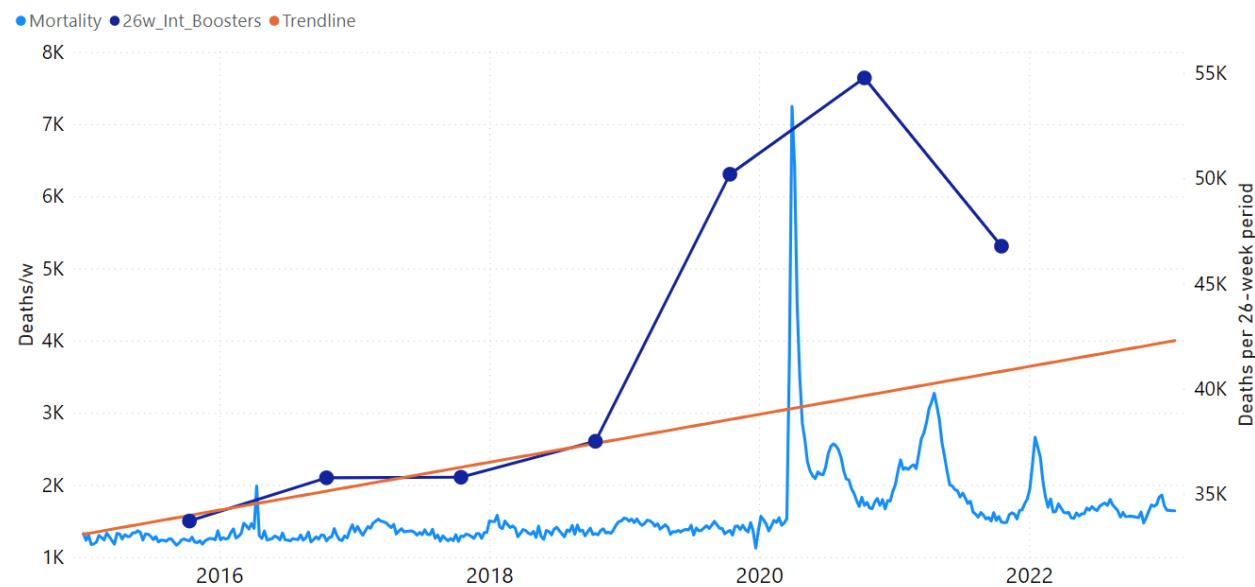
ACM/w, Chile, 2015-2023



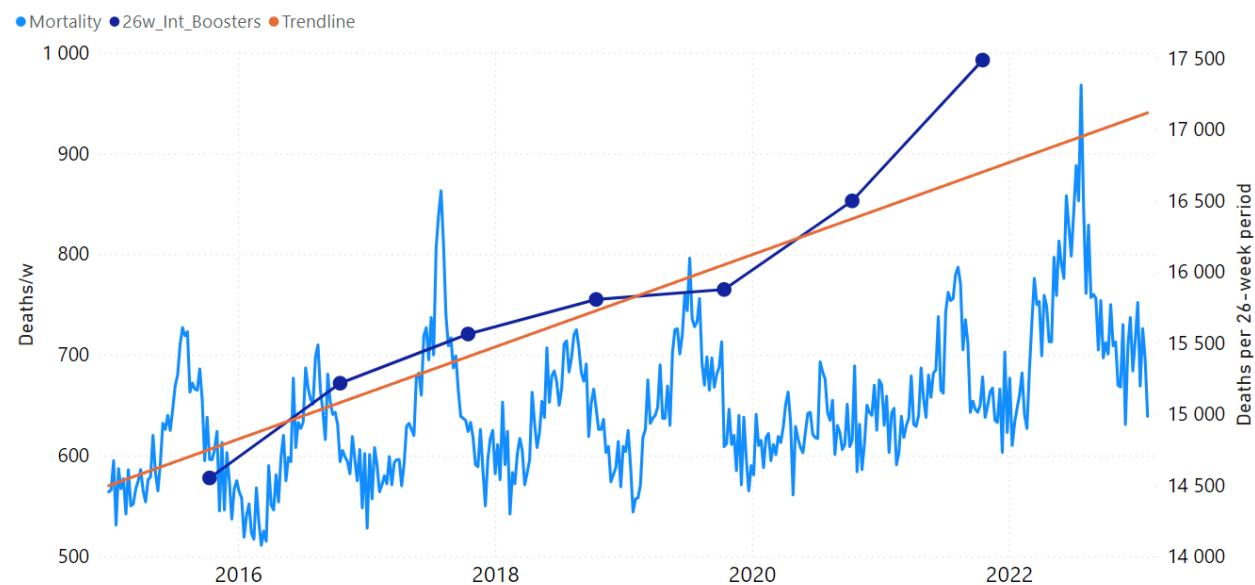
ACM/w, Colombia, 2015-2023



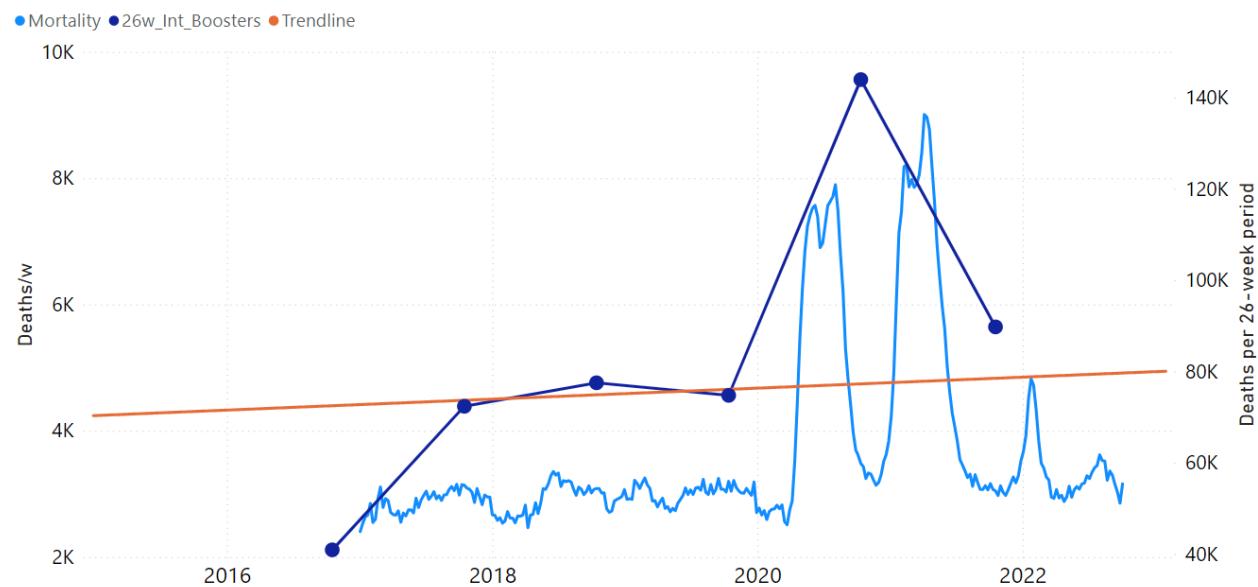
ACM/w, Ecuador, 2015-2023



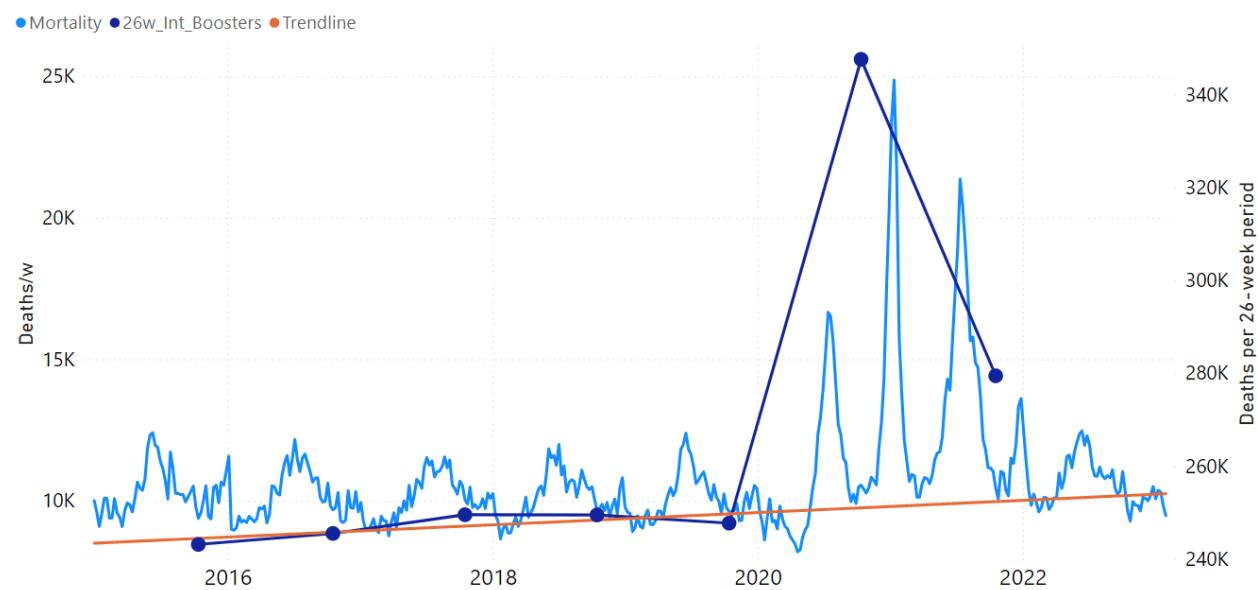
ACM/w, New Zealand, 2015-2023



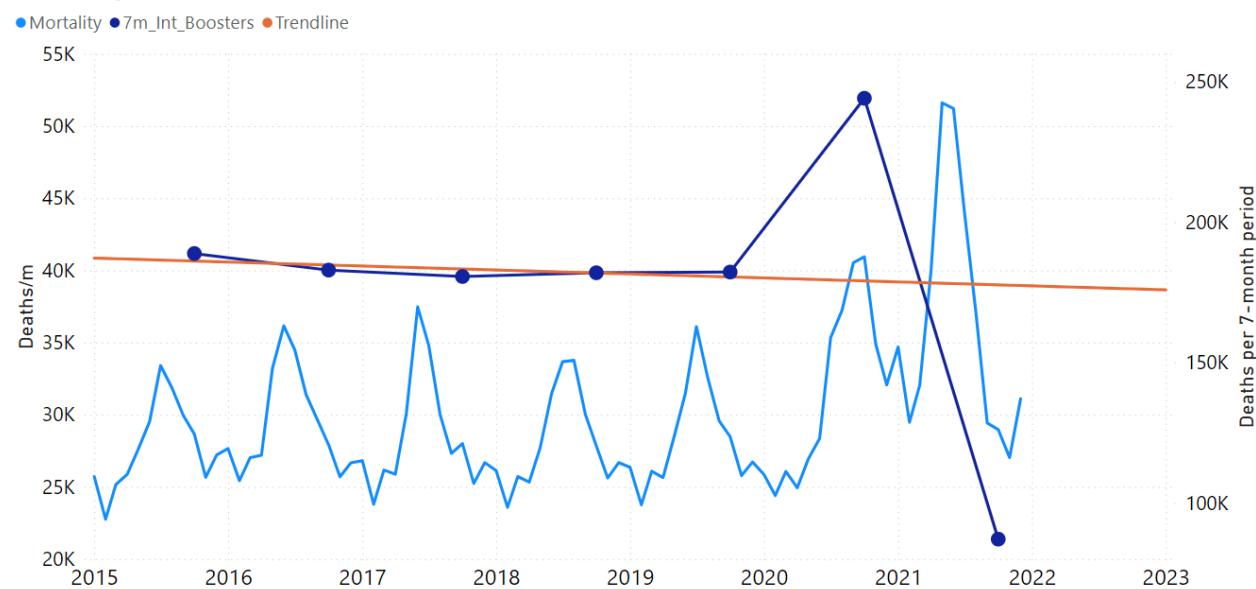
ACM/w, Peru, 2015-2023



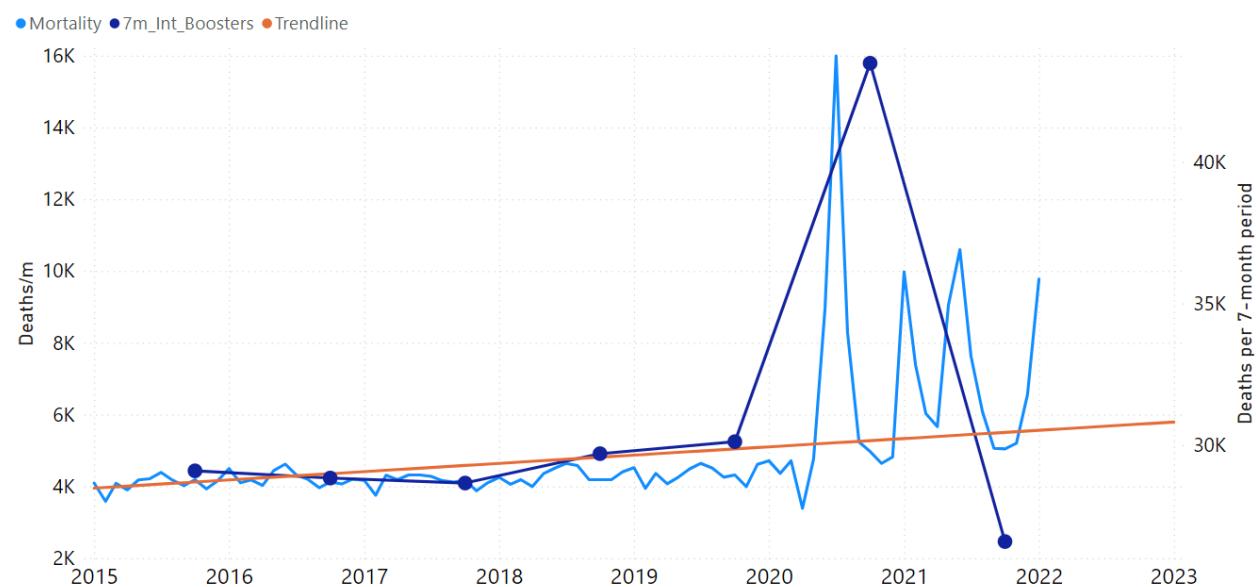
ACM/w, South Africa, 2015-2023



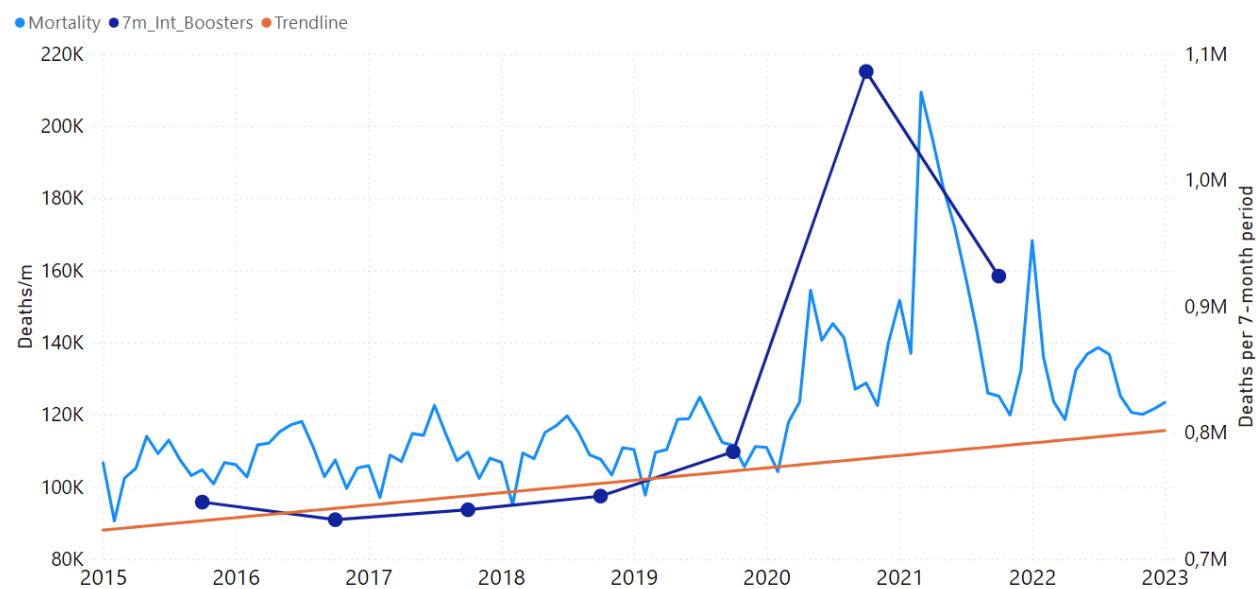
ACM/m, Argentina, 2015-2022



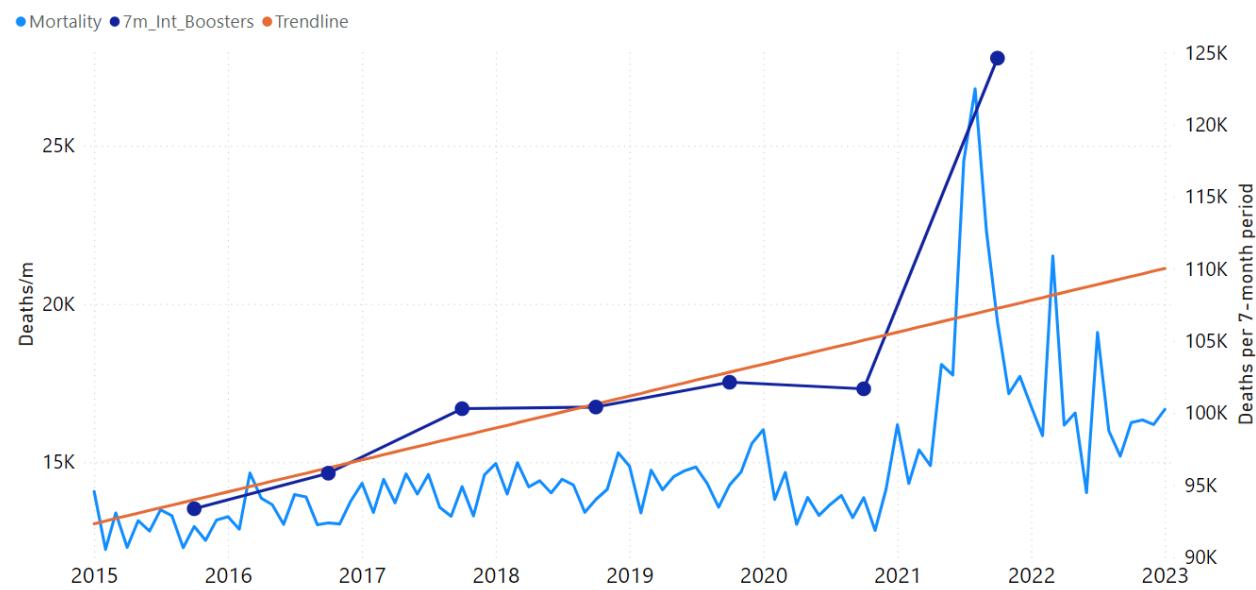
ACM/m, Bolivia, 2015-2022



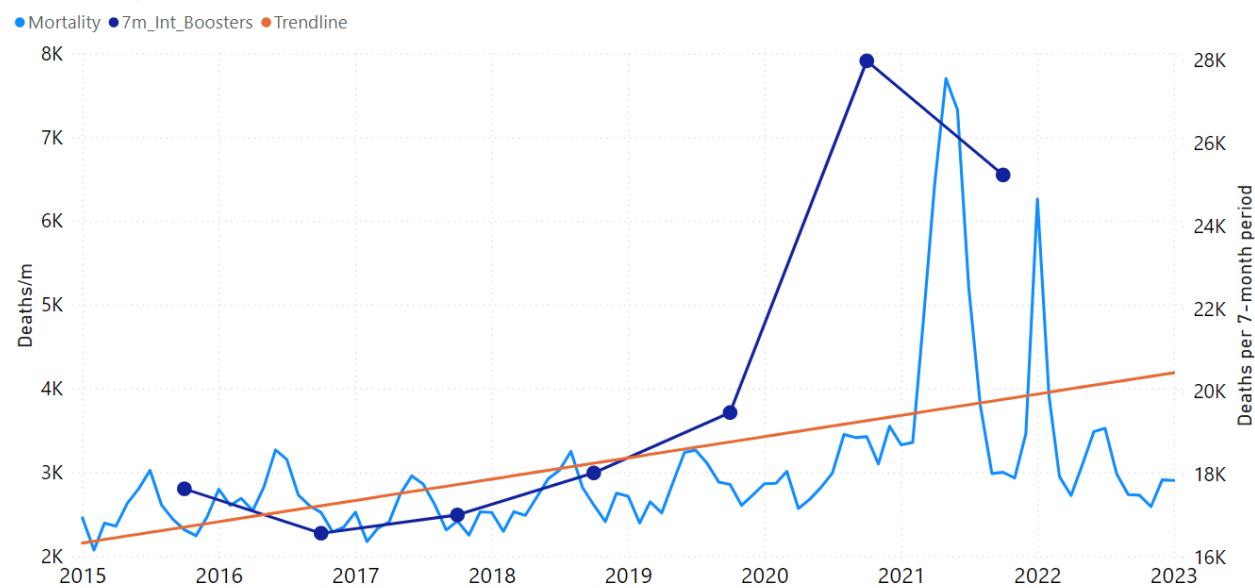
ACM/m, Brazil, 2015-2023



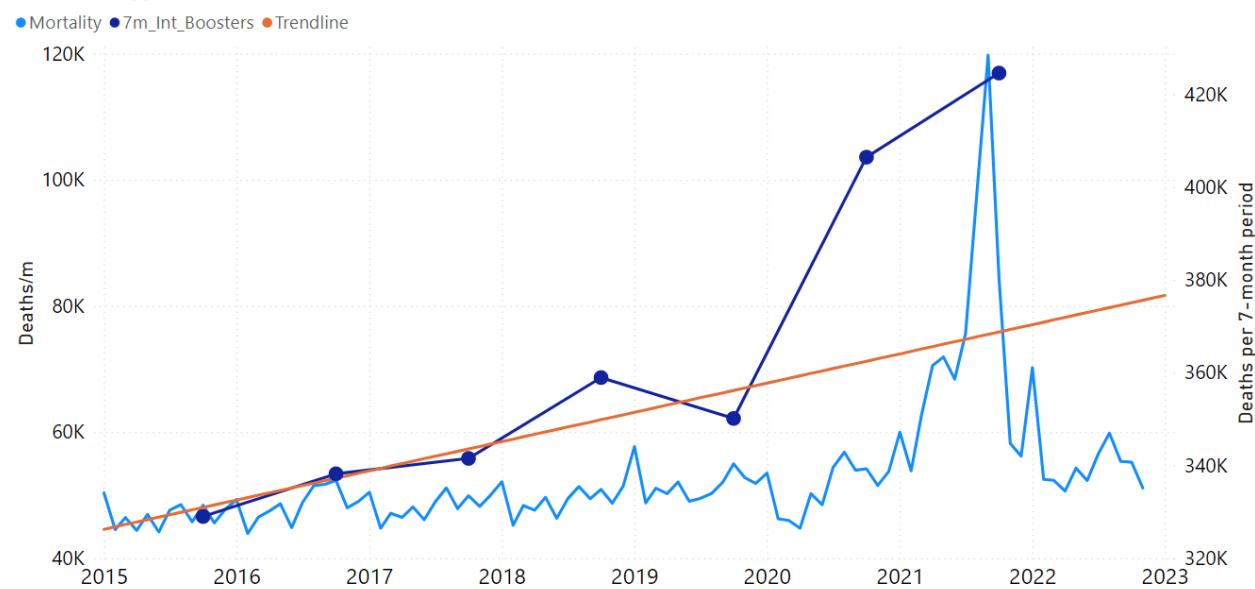
ACM/m, Malaysia, 2015-2023



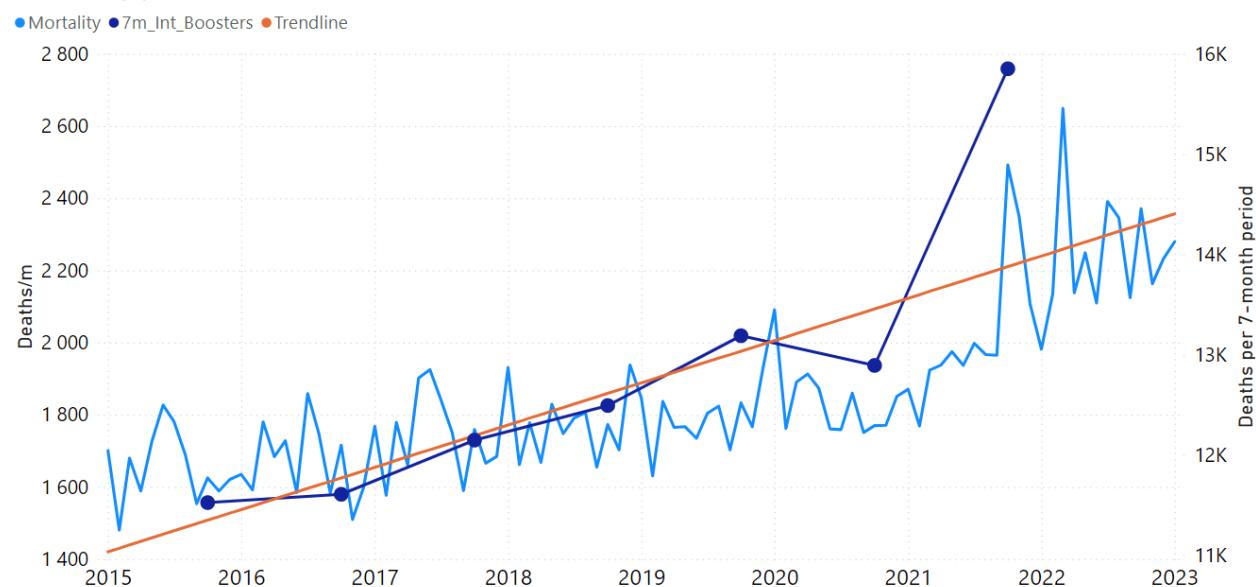
ACM/m, Paraguay, 2015-2023



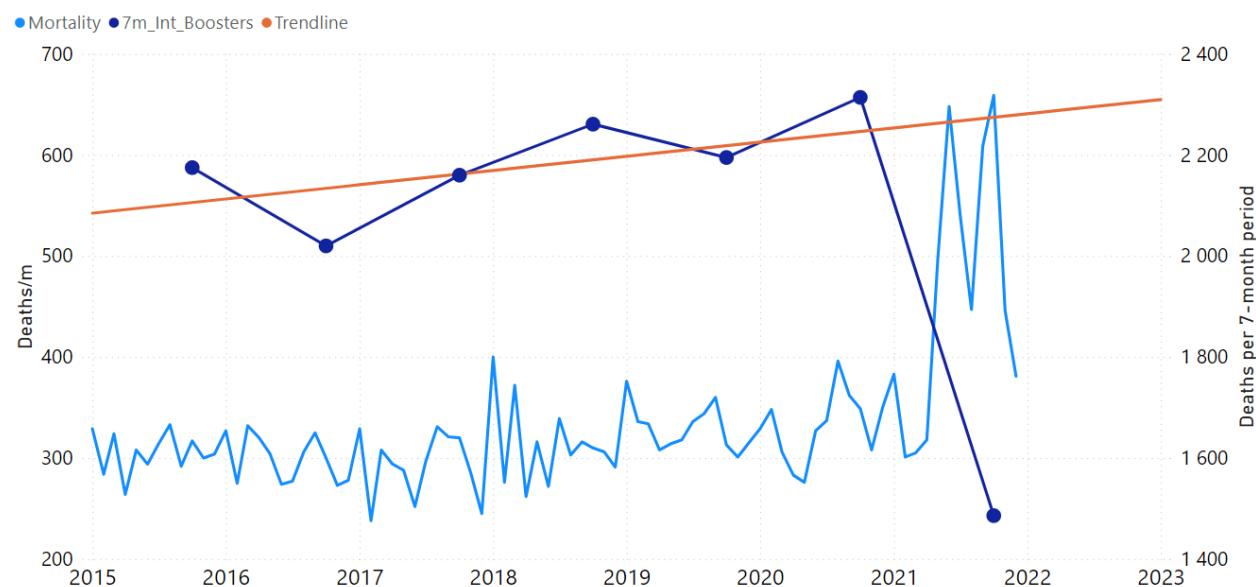
ACM/m, Philippines, 2015-2023



ACM/m, Singapore, 2015-2023



ACM/m, Suriname, 2015-2022



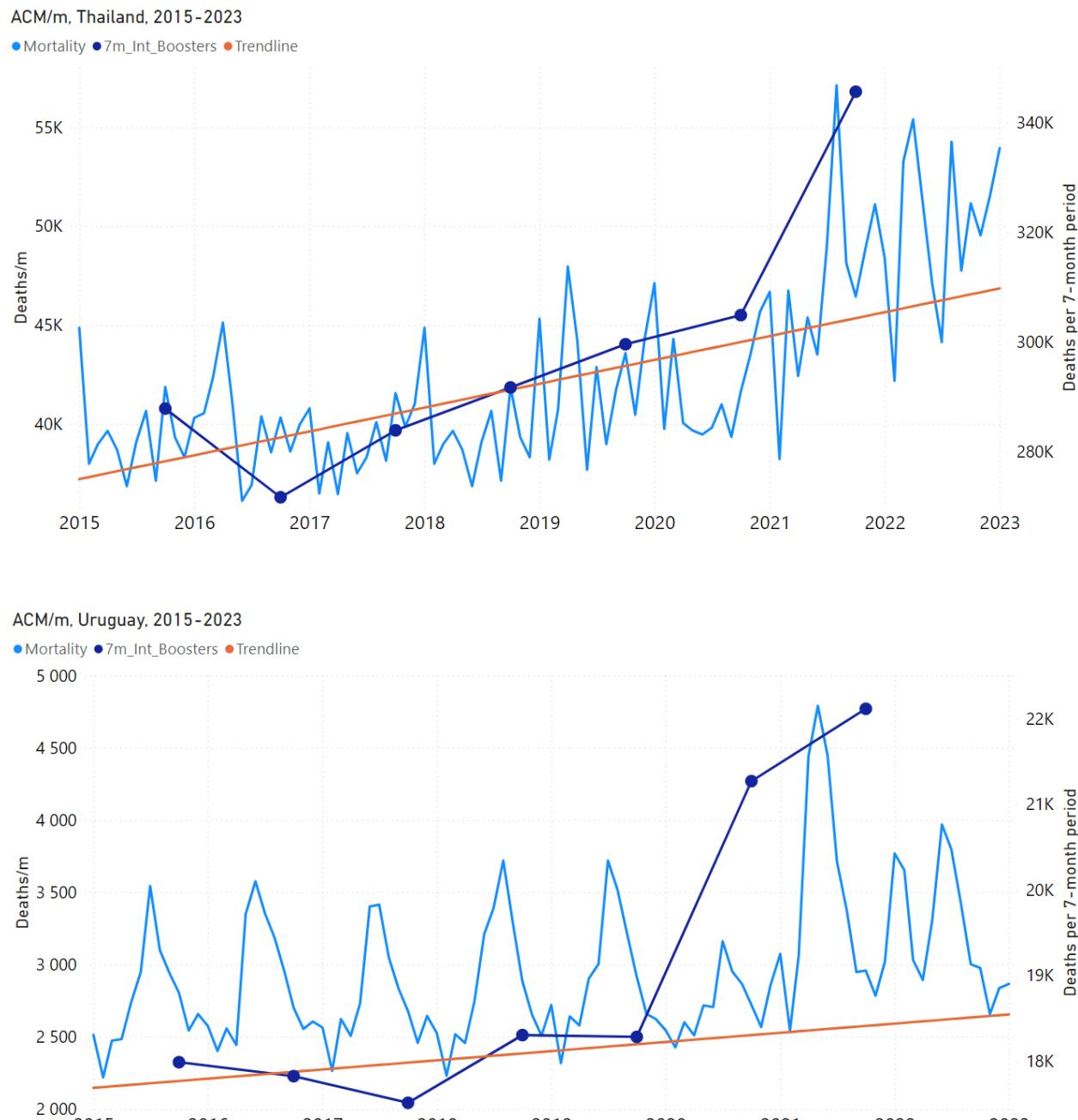


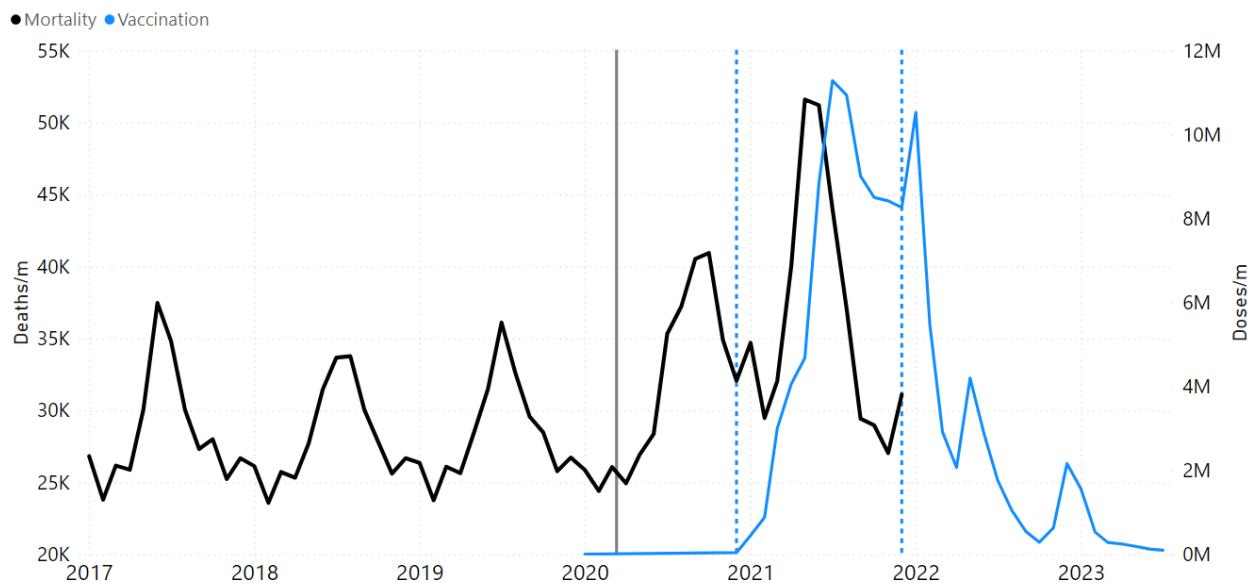
Figure 5: All-ages ACM by week or by month (light blue), 2015-2023, with integrations for the nominally January-February 2022 peak in ACM (dark blue), for the 17 countries (as labelled). The data sources are specified in Appendix A.

Figure 6 shows the same kind of graph as depicted in Figure 4, except that the entire available COVID-19 vaccine period is highlighted, instead of just the booster-rollout region related to the nominally January-February 2022 peak in ACM, for all 17 countries

in this study (Argentina, Australia, Bolivia, Brazil, Chile, Colombia, Ecuador, Malaysia, New Zealand, Paraguay, Peru, Philippines, Singapore, South Africa, Suriname, Thailand, Uruguay).

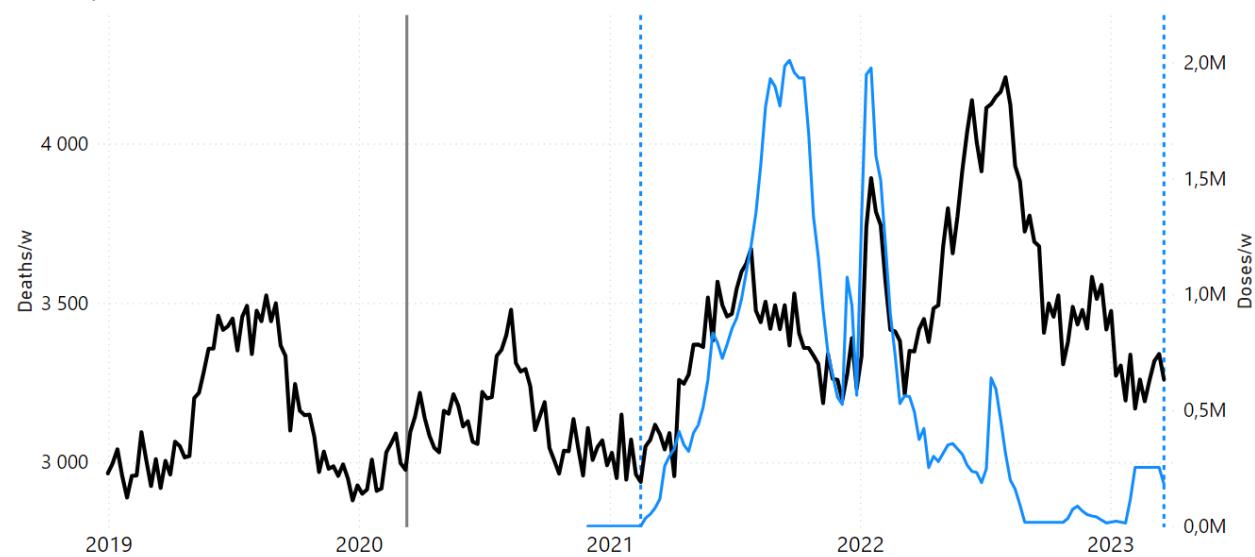
Here (Figure 6 and Figure 7), the vaccination period is defined to start at the first week or month of COVID-19 vaccine administration, and to end at effectively the latest time of vaccine administration or at the latest date of available ACM data. See Appendix C: Technical and specific information for applications of the methods to the data for the specified ranges used. The idea is to capture the all-ages all-doses long-time impact of vaccination in each country, with a corresponding calculated value of vDFR, and to compare it with the value obtained from the targeted focus on the booster-associated nominally January-February 2022 peak in ACM.

ACM and vaccination by month, Argentina, 2017-2023



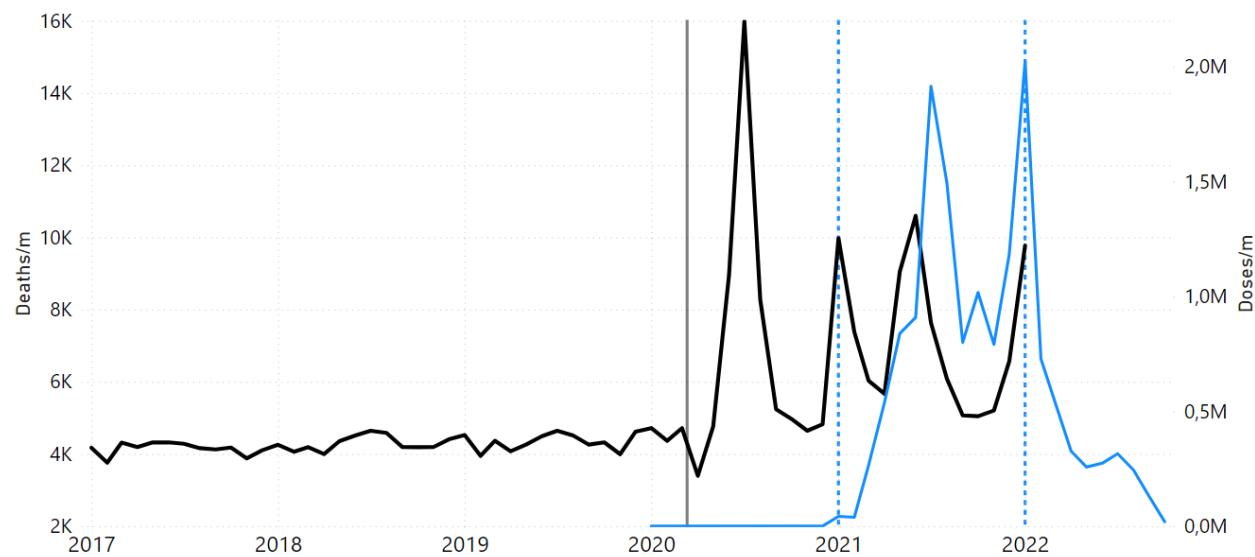
ACM and vaccination by week, Australia, 2019-2023

● Mortality ● Vaccination

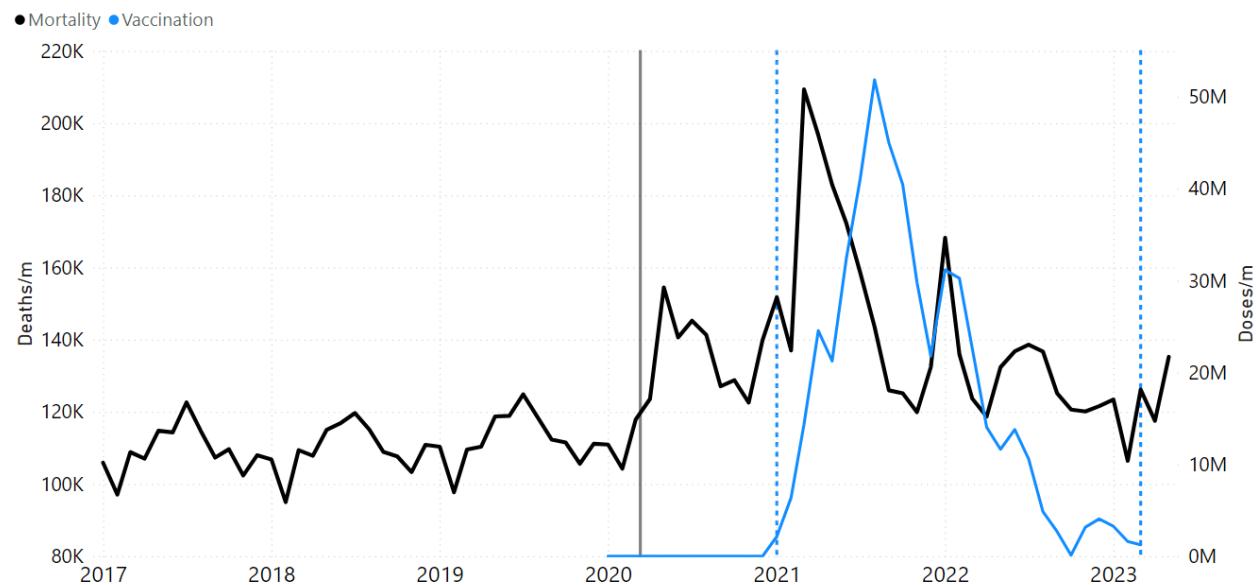


ACM and vaccination by month, Bolivia, 2017-2022

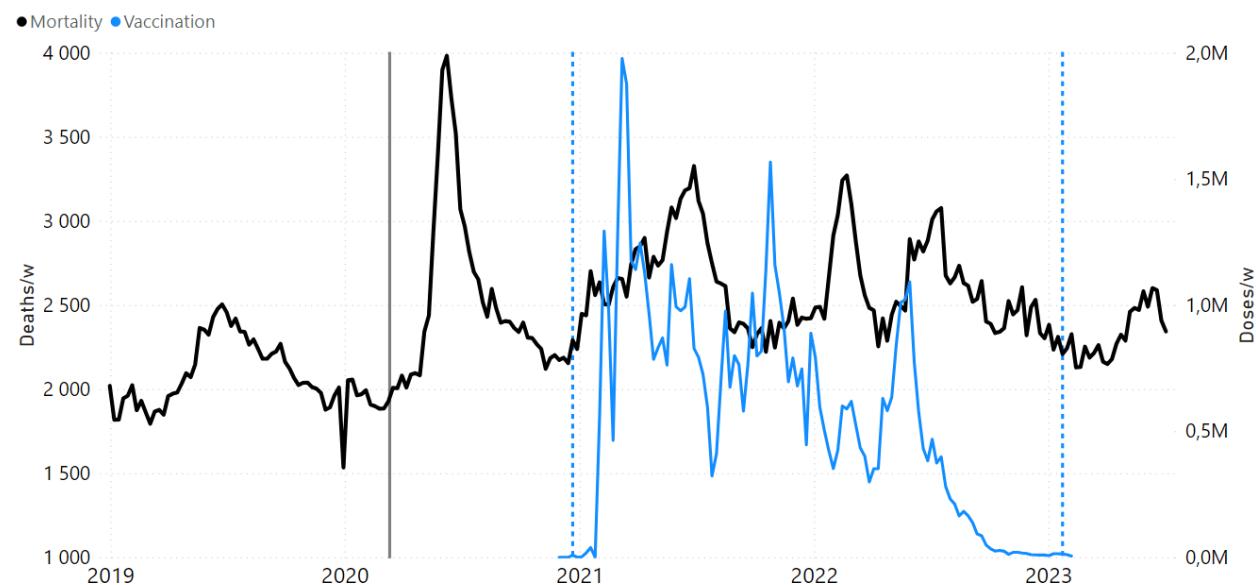
● Mortality ● Vaccination



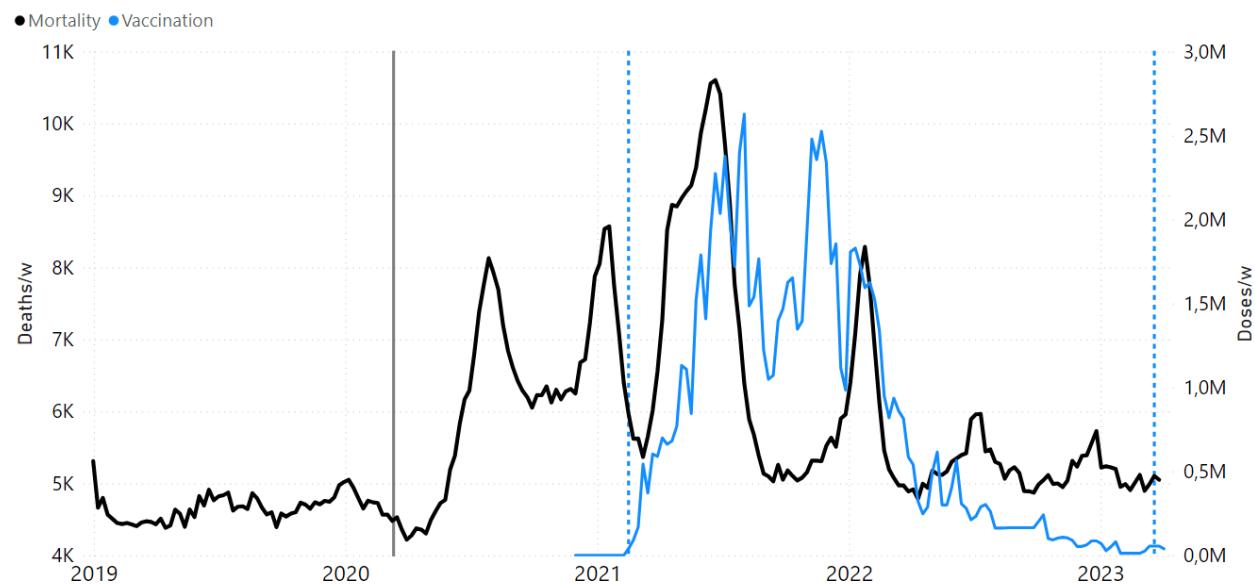
ACM and vaccination by month, Brazil, 2017-2023



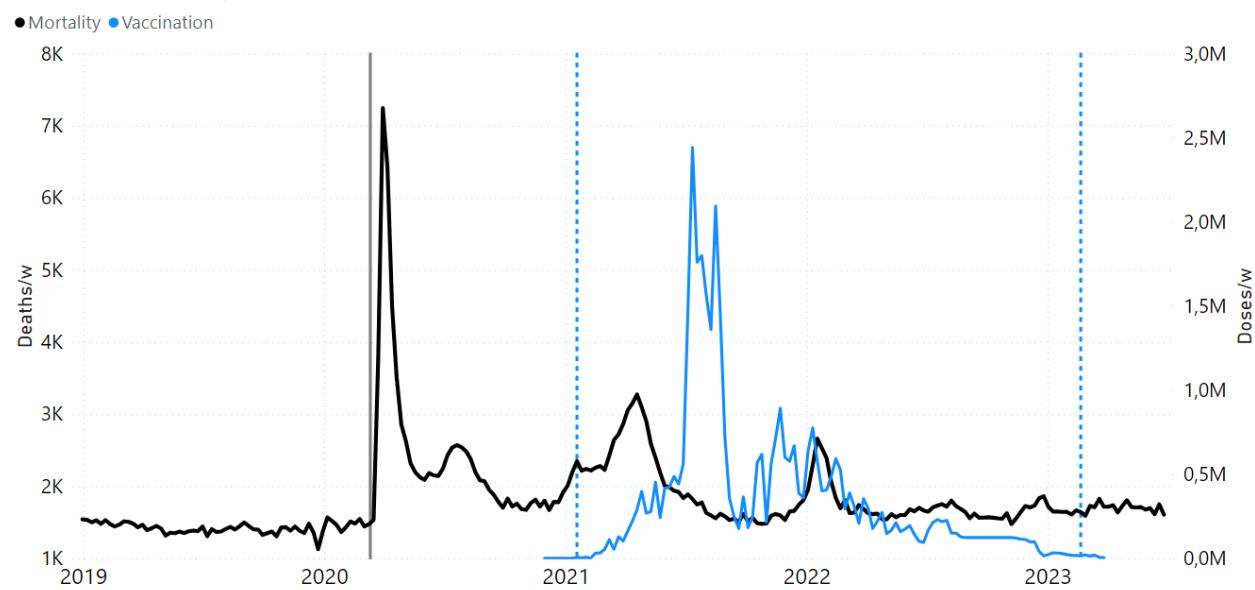
ACM and vaccination by week, Chile, 2019-2023



ACM and vaccination by week, Colombia, 2019–2023

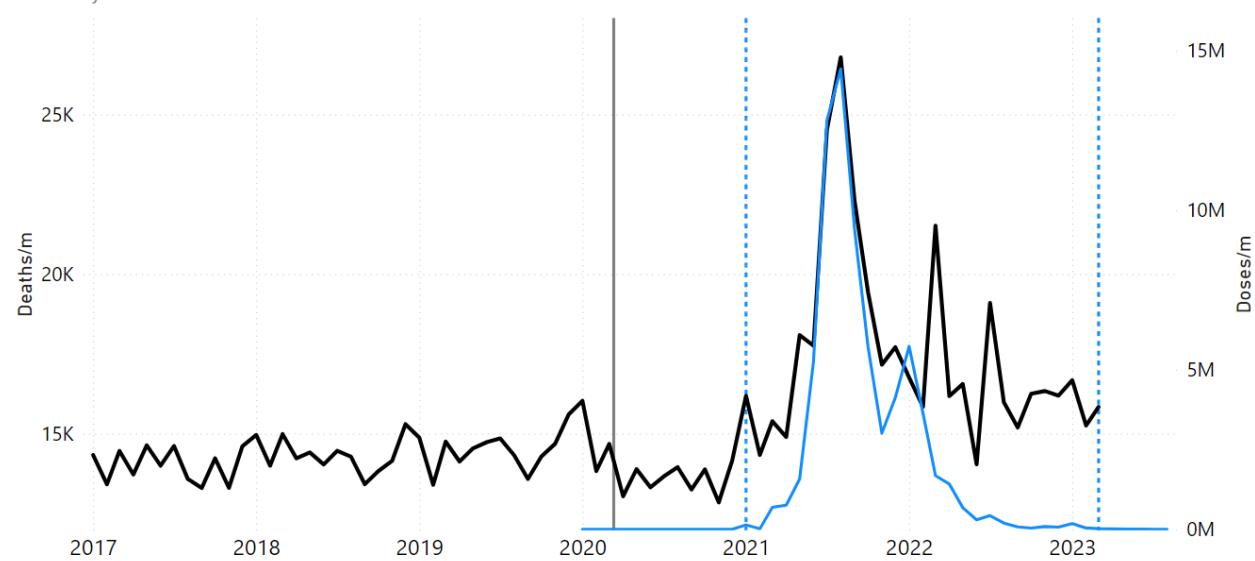


ACM and vaccination by week, Ecuador, 2019–2023



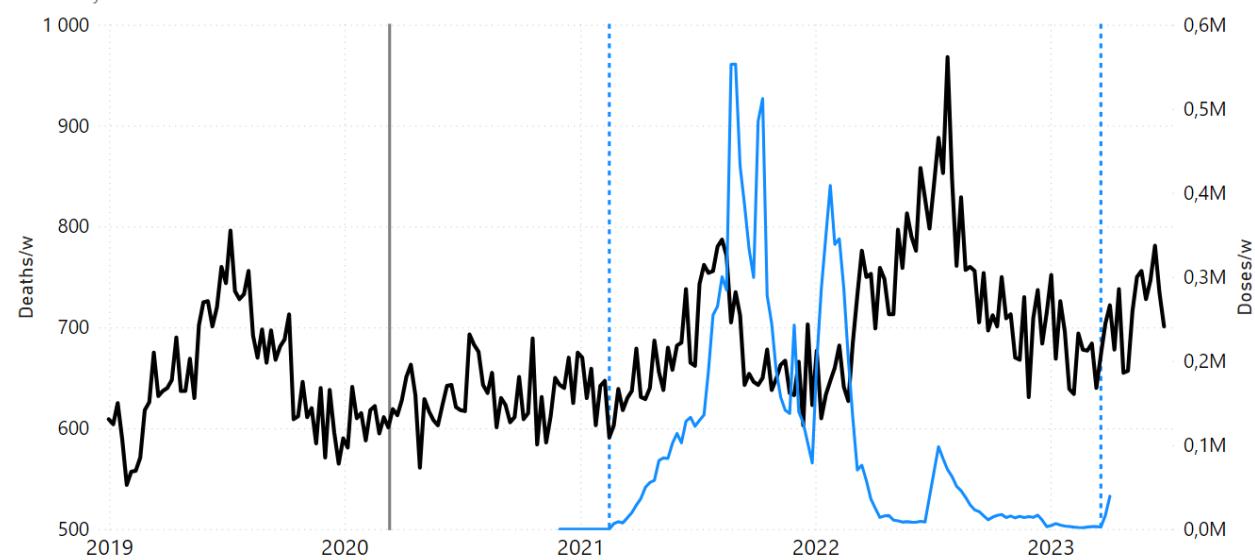
ACM and vaccination by month, Malaysia, 2017-2023

● Mortality ● Vaccination



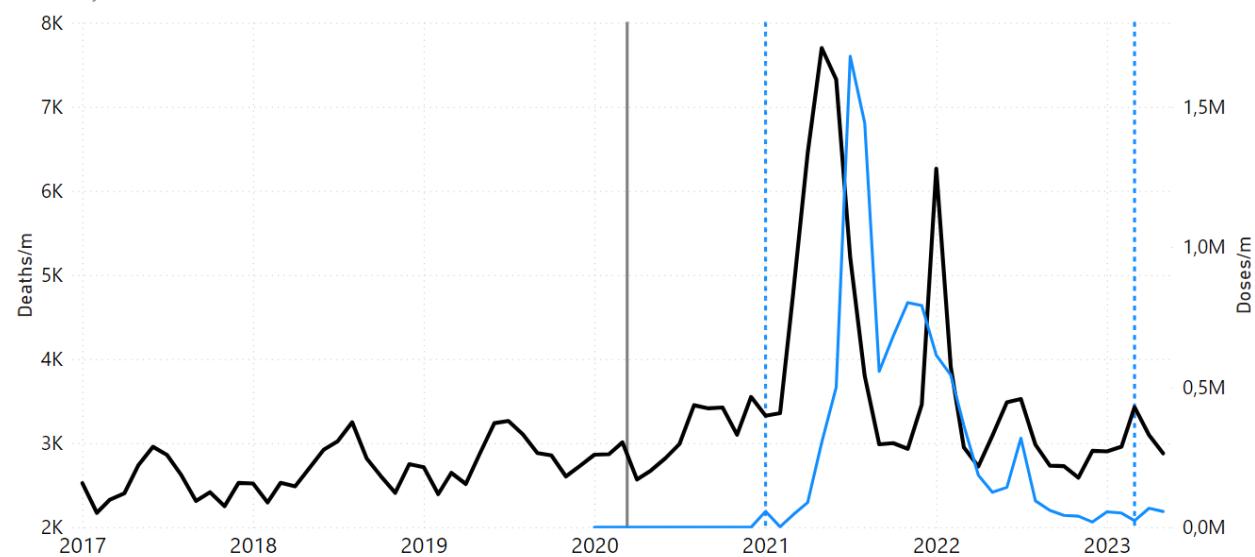
ACM and vaccination by week, New Zealand, 2019-2023

● Mortality ● Vaccination



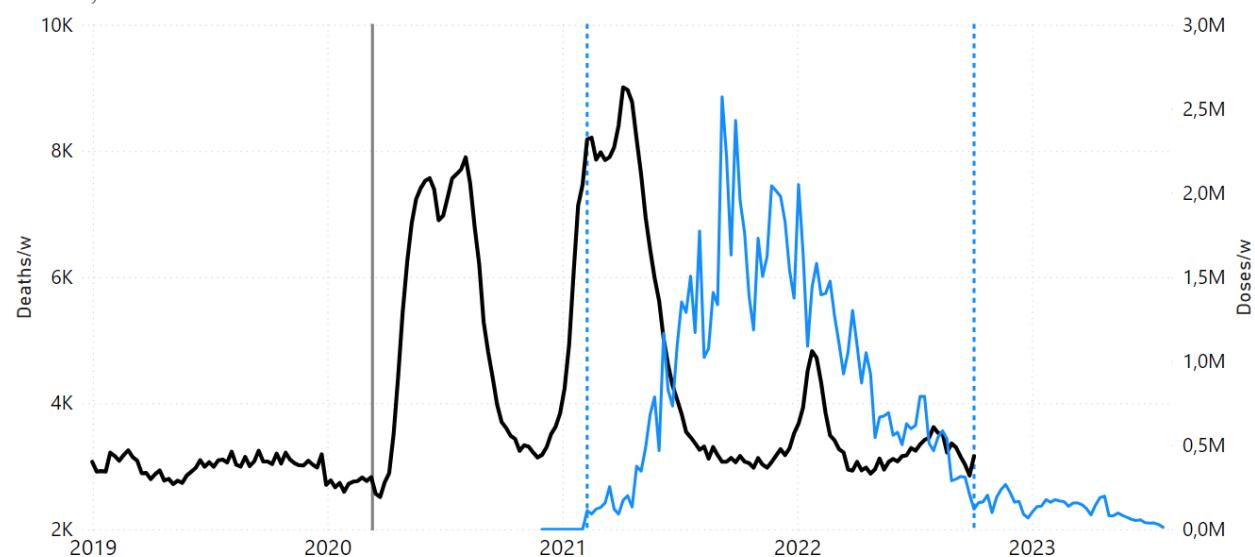
ACM and vaccination by month, Paraguay, 2017-2023

● Mortality ● Vaccination

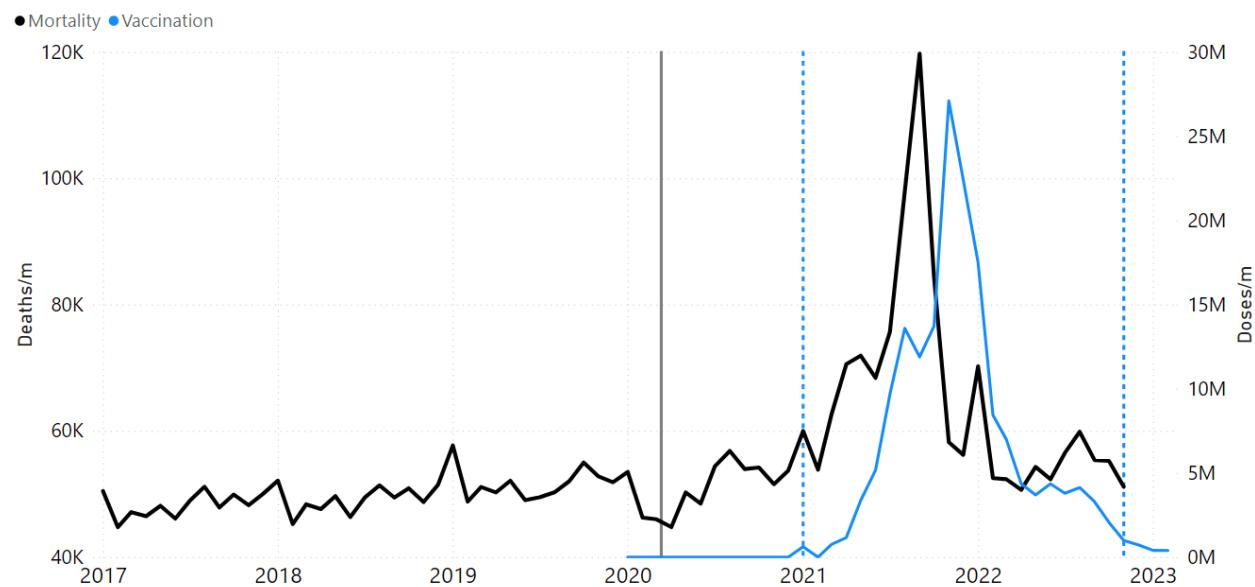


ACM and vaccination by week, Peru, 2019-2023

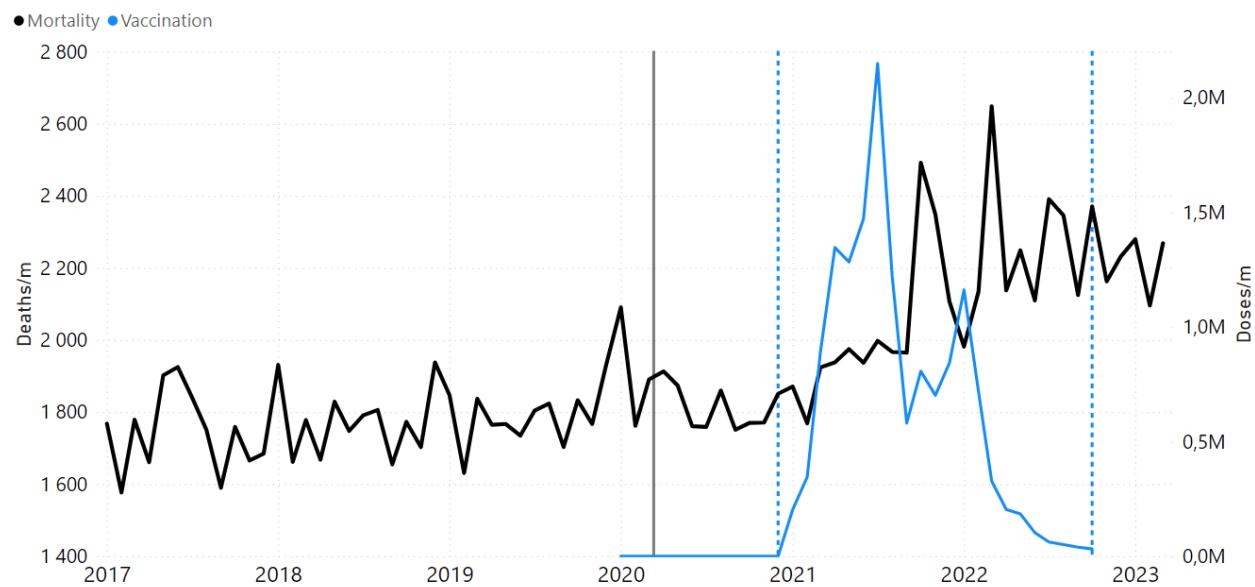
● Mortality ● Vaccination



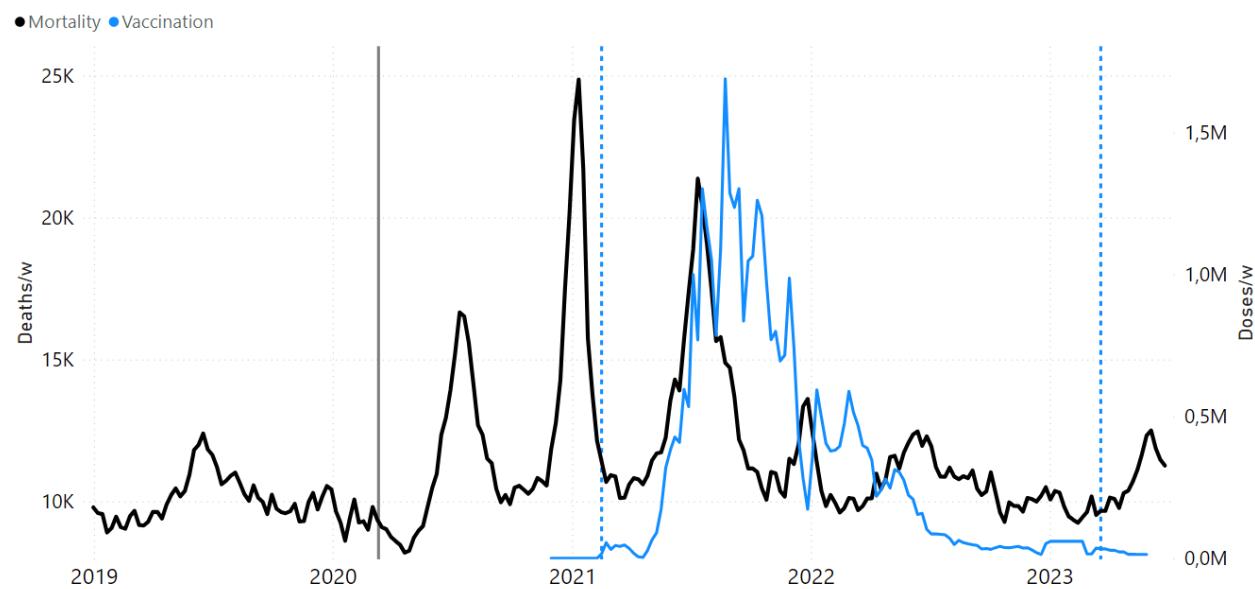
ACM and vaccination by month, Philippines, 2017-2023



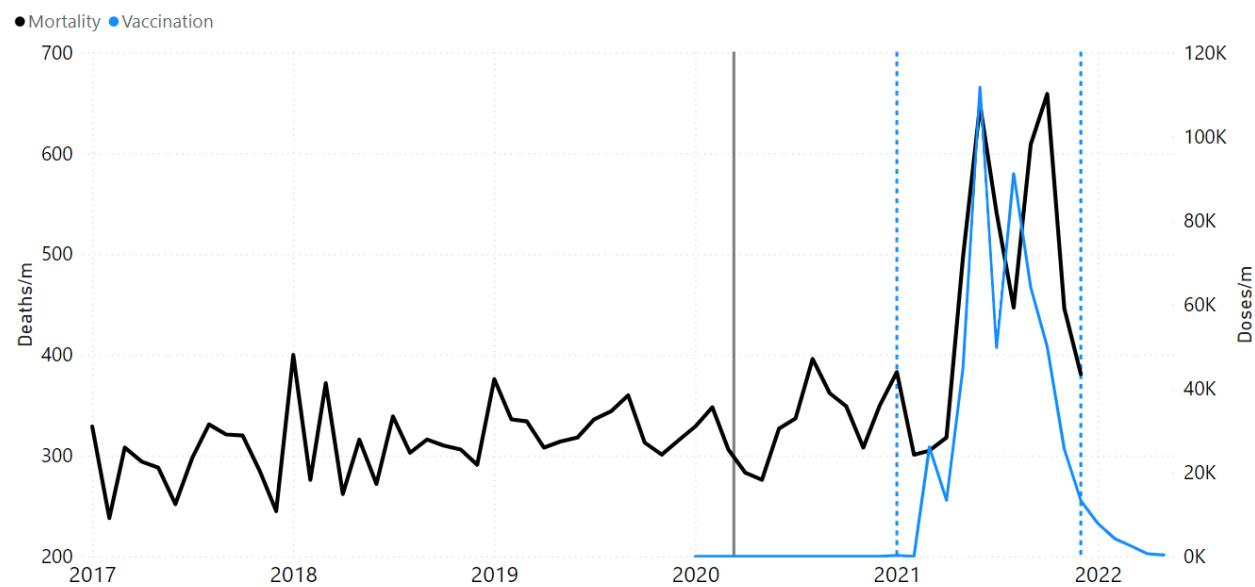
ACM and vaccination by month, Singapore, 2017-2023



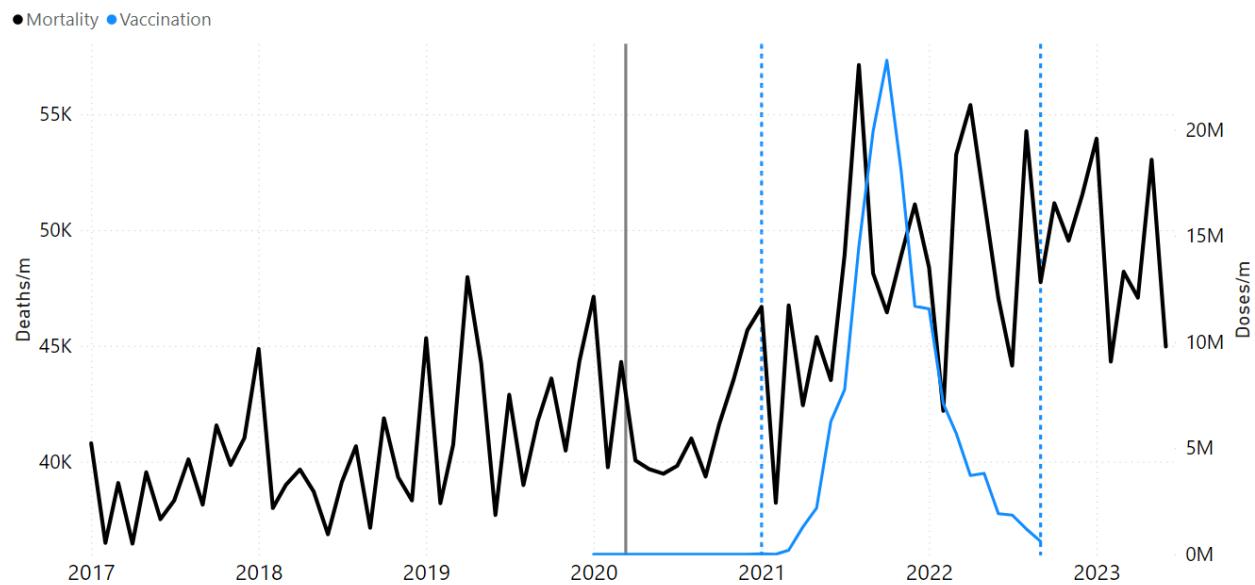
ACM and vaccination by week, South Africa, 2019-2023



ACM and vaccination by month, Suriname, 2017-2022



ACM and vaccination by month, Thailand, 2017-2023



ACM and vaccination by month, Uruguay, 2017-2023

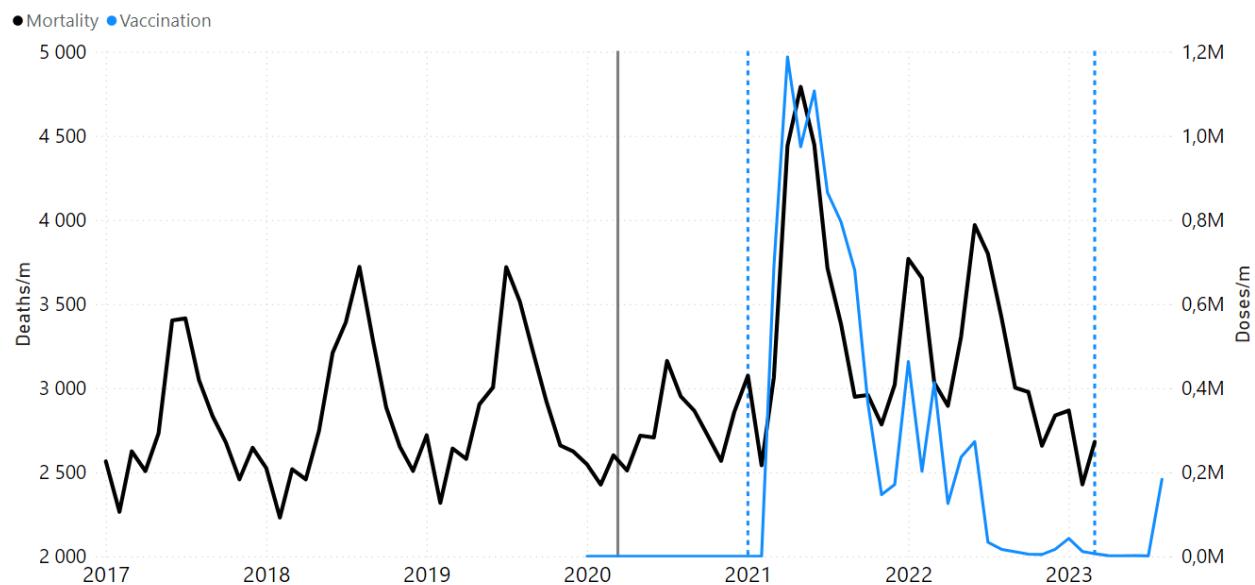
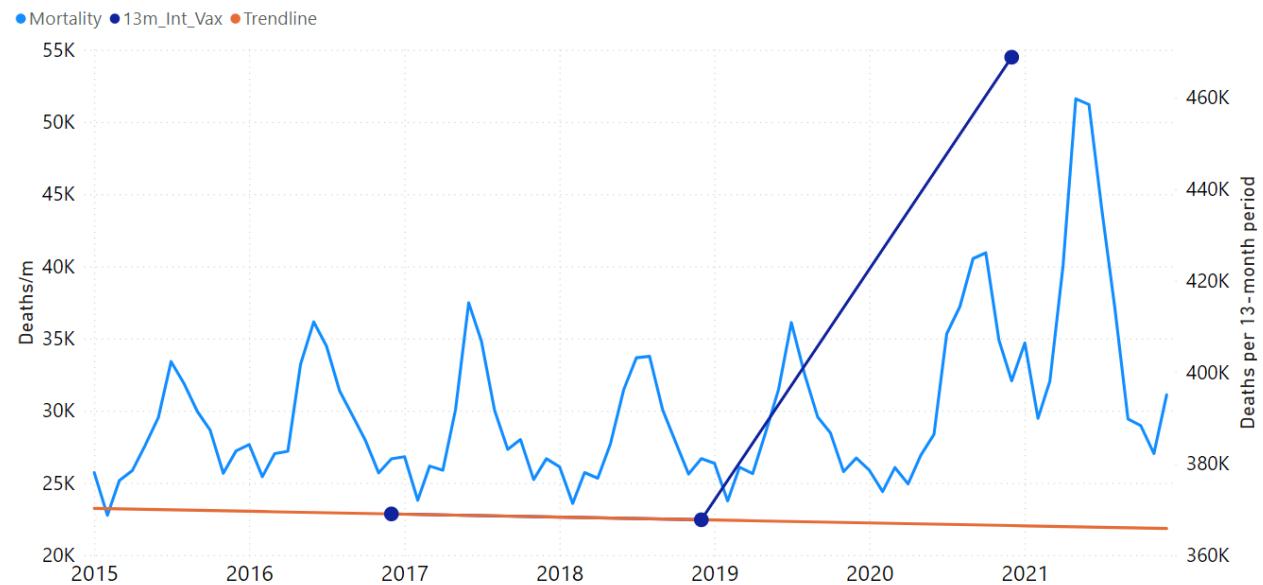


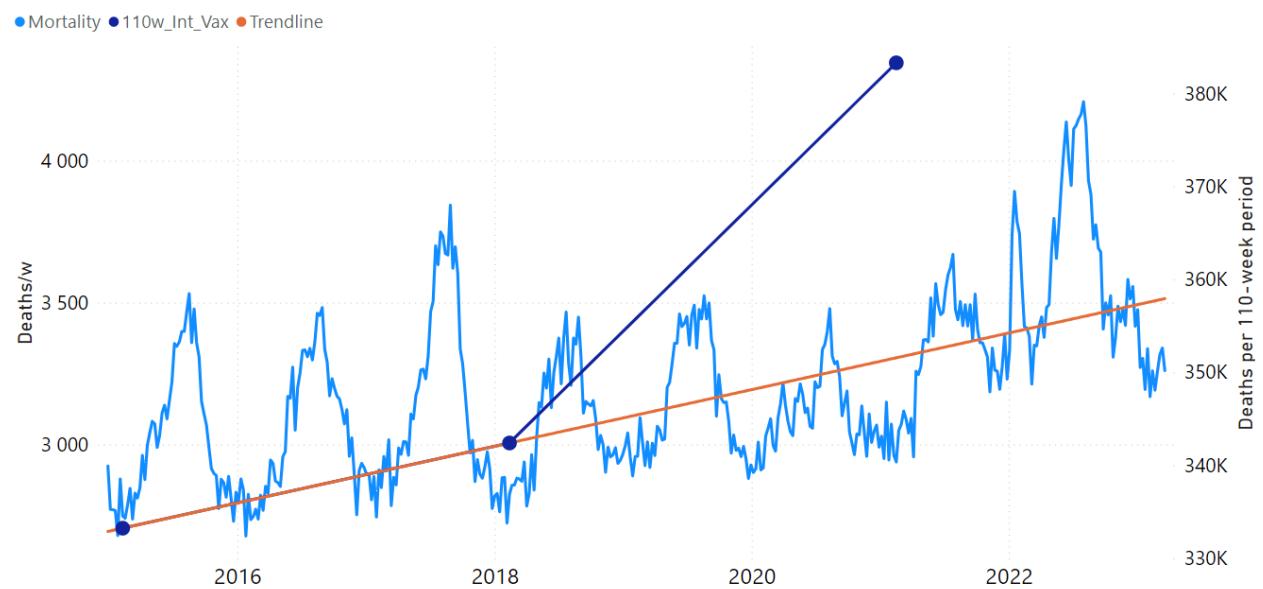
Figure 6: All-ages ACM by week (2019-2023) or by month (2017-2023) (black), with all-ages all-doses COVID-19 vaccine rollout (blue), for the 17 countries in this study; also showing the 11 March 2020 date (grey) and limits of integration (dashed, blue) for the vaccination period. The data sources are specified in Appendix A.

Figure 7 shows the integrations corresponding to the vaccination periods illustrated in Figure 6.

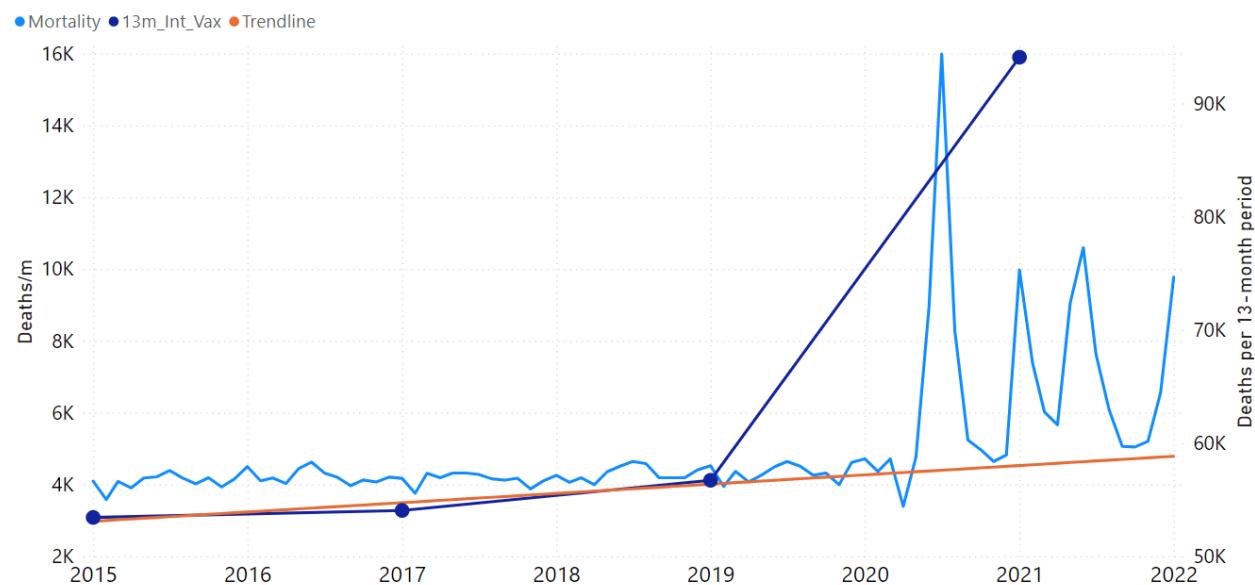
ACM/m, Argentina, 2015-2021



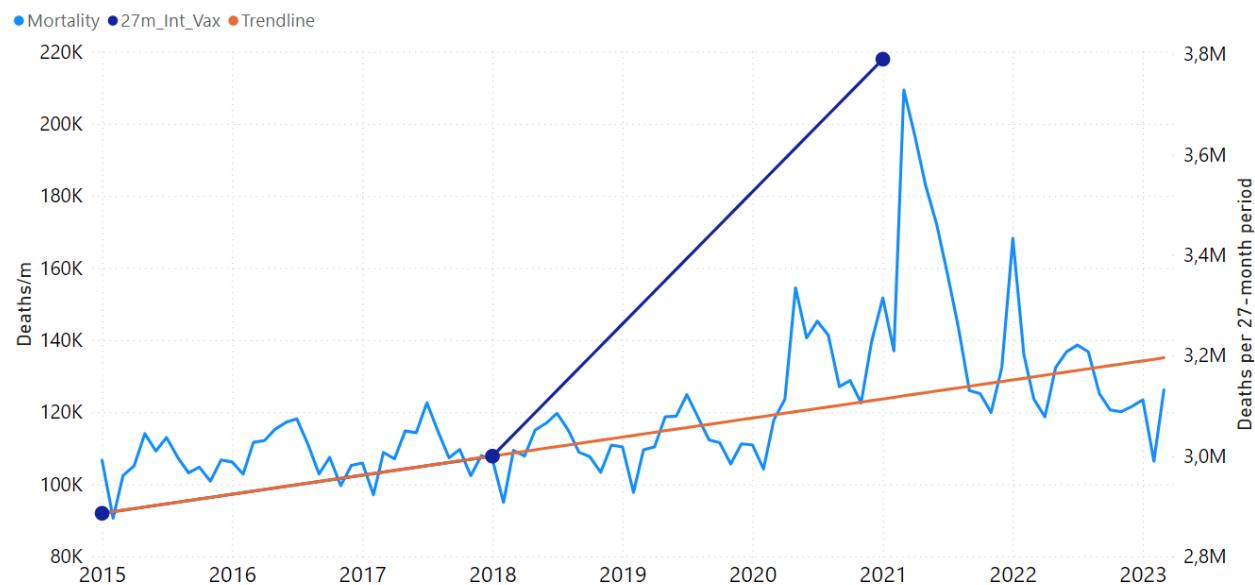
ACM/w, Australia, 2015-2023



ACM/m, Bolivia, 2015-2022



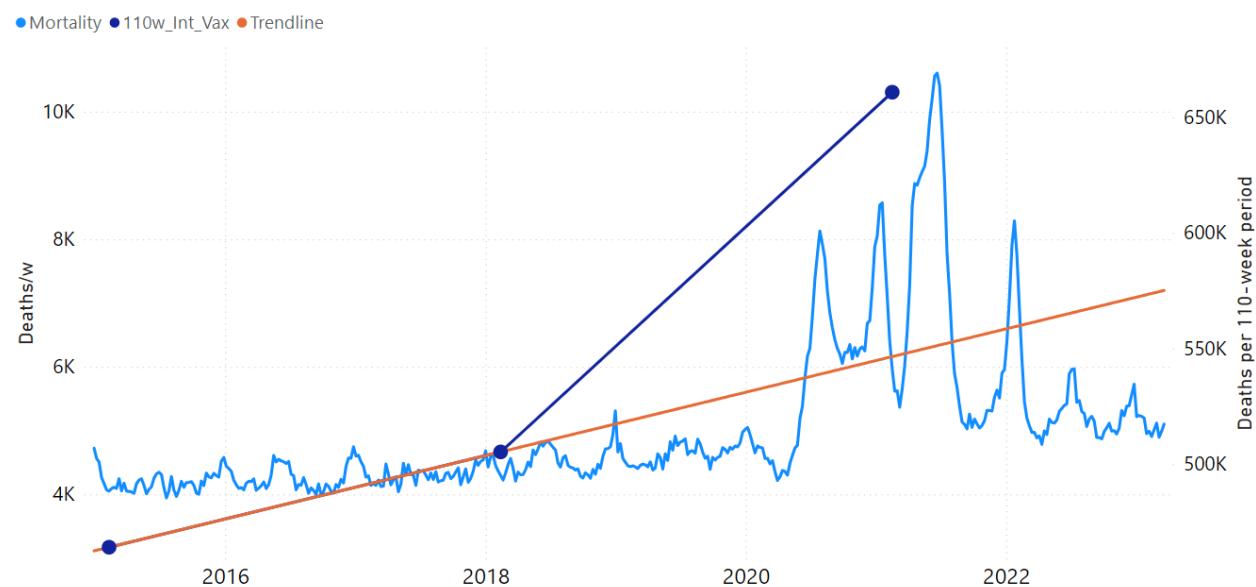
ACM/m, Brazil, 2015-2023



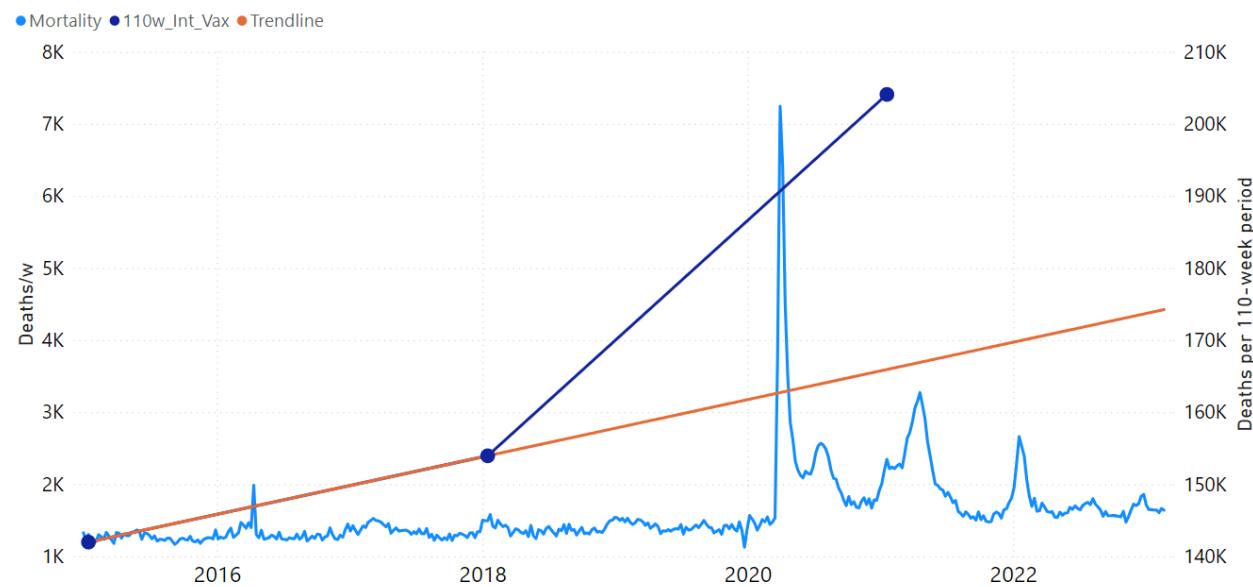
ACM/w, Chile, 2016-2023



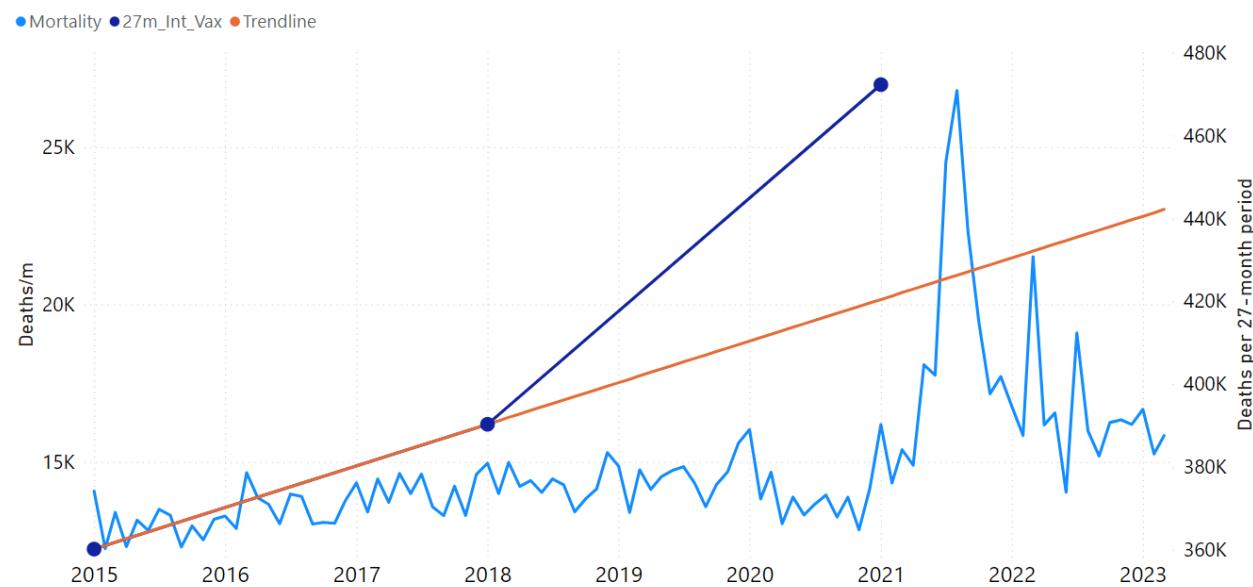
ACM/w, Colombia, 2015-2023



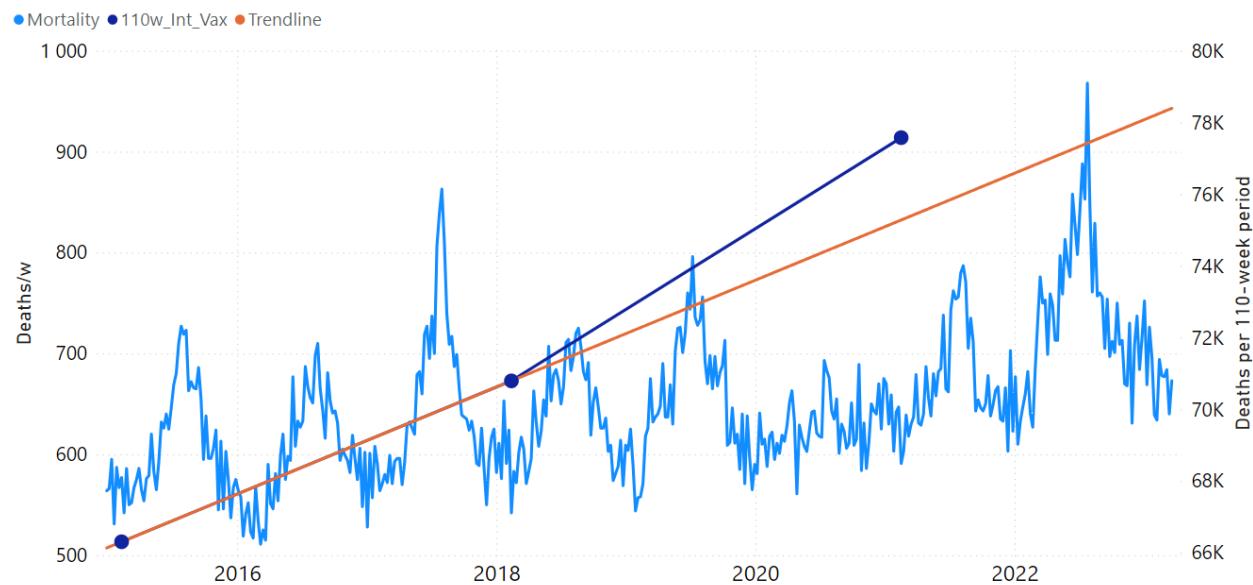
ACM/w, Ecuador, 2015-2023



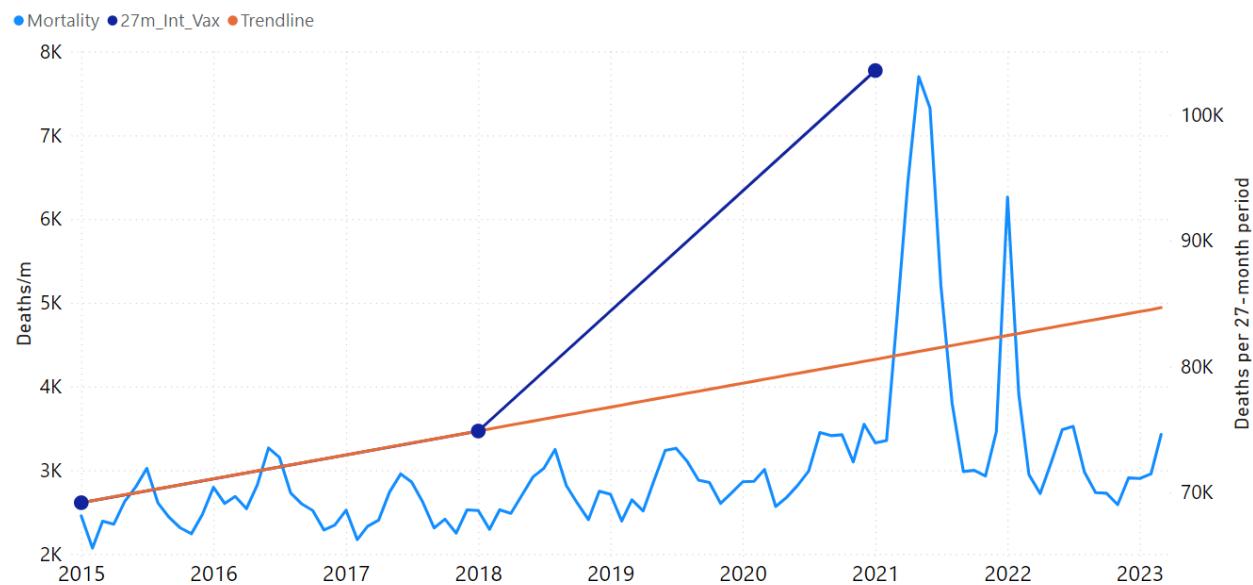
ACM/m, Malaysia, 2015-2023



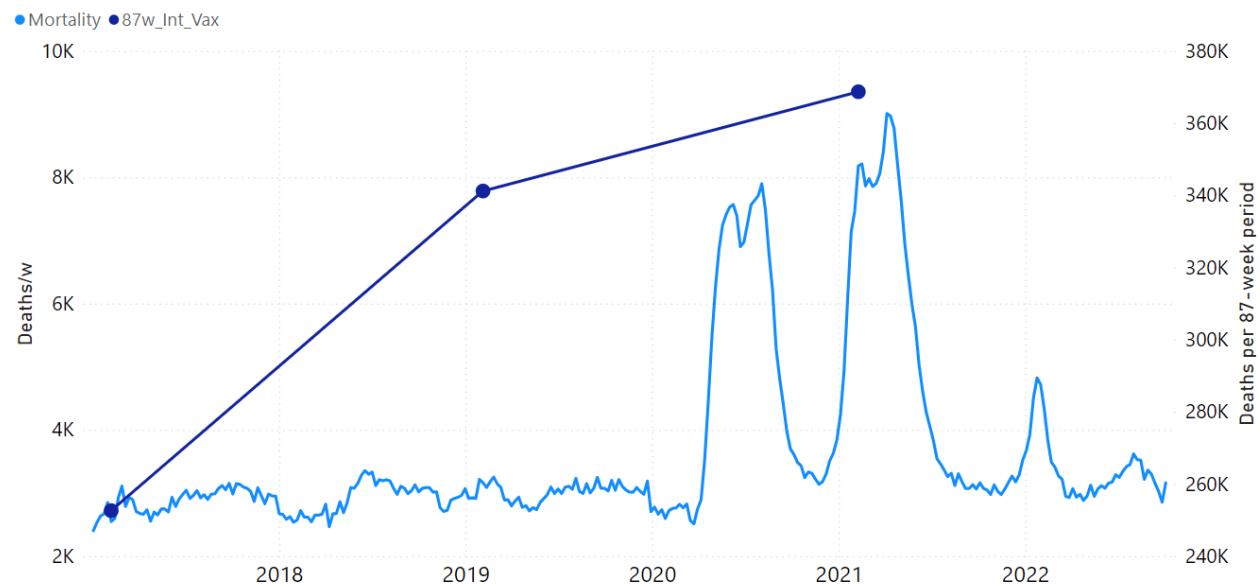
ACM/w, New Zealand, 2015-2023



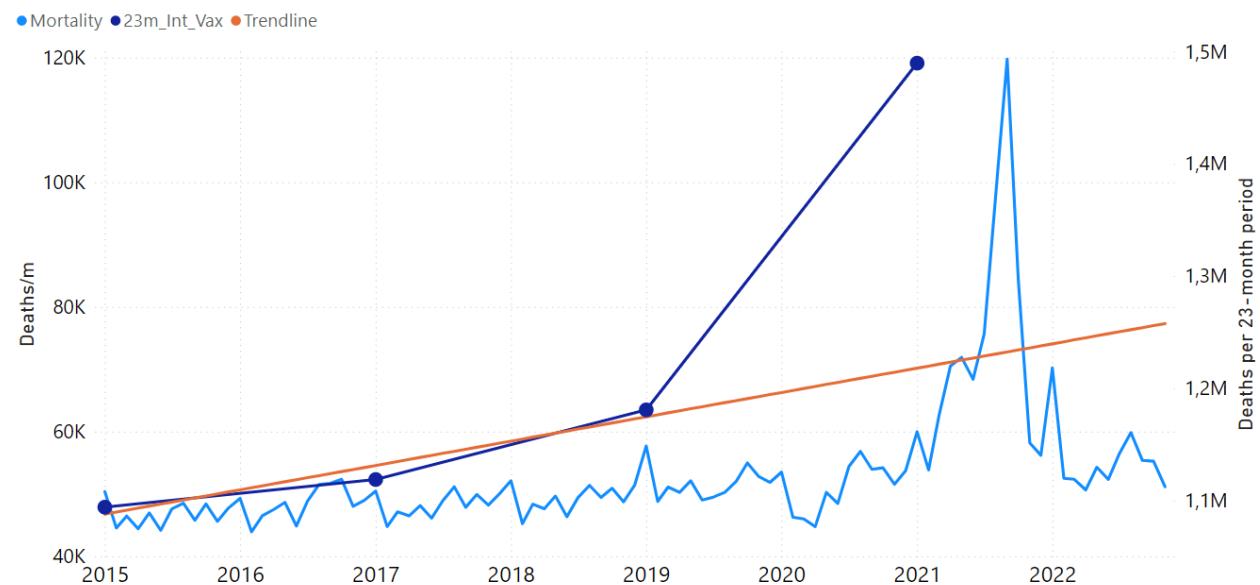
ACM/m, Paraguay, 2015-2023



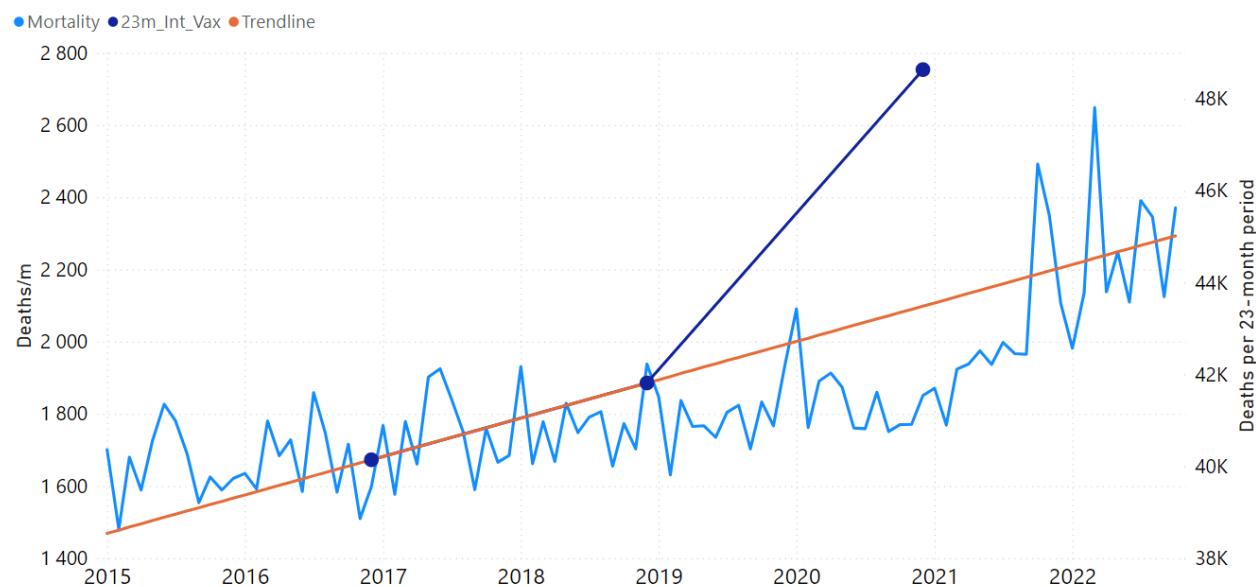
ACM/w, Peru, 2017-2022



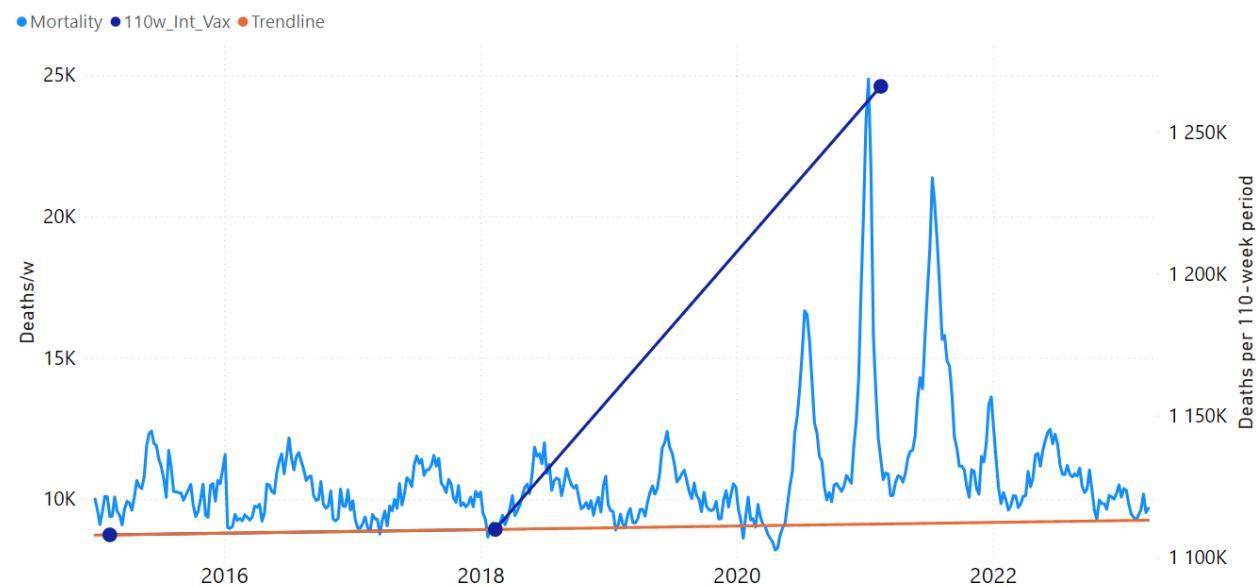
ACM/m, Philippines, 2015-2022



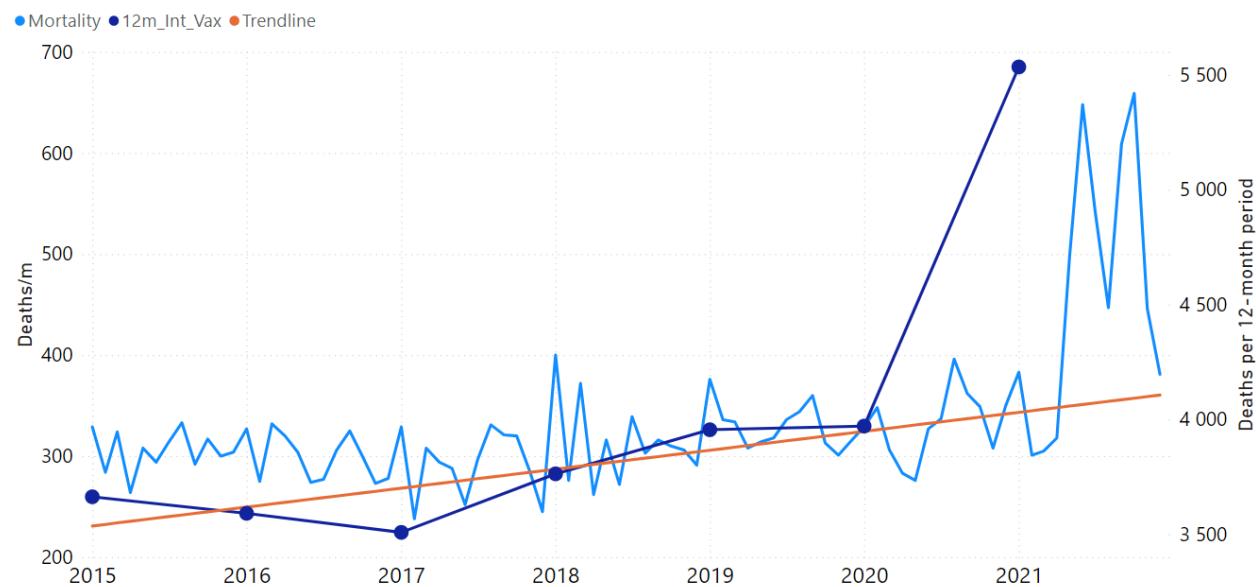
ACM/m, Singapore, 2015-2022



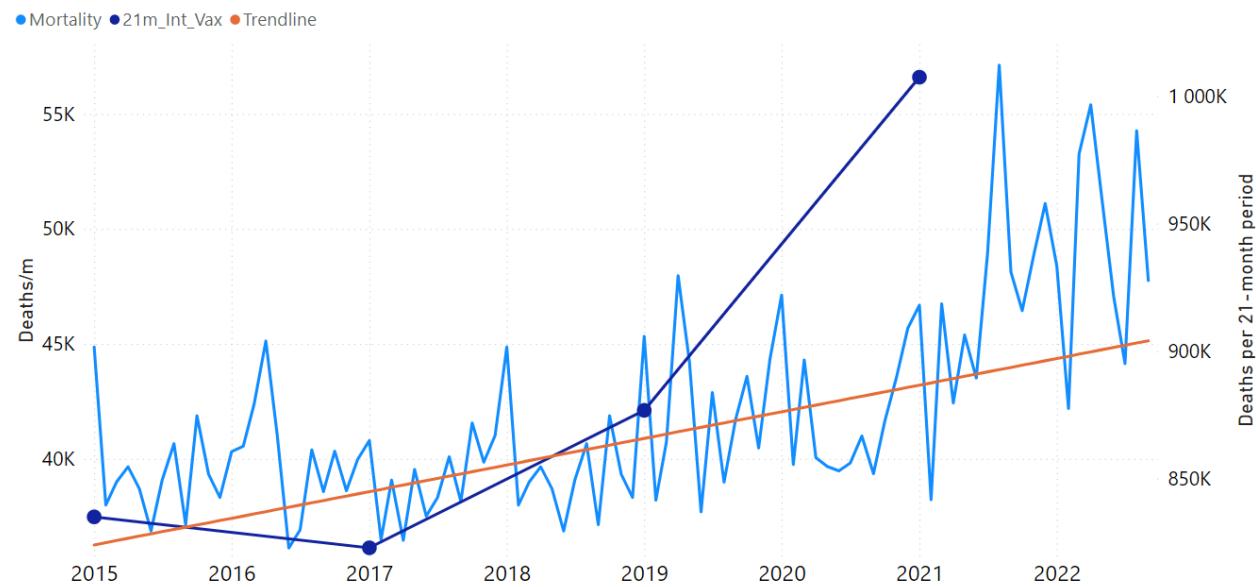
ACM/w, South Africa, 2015-2023



ACM/m, Suriname, 2015-2021



ACM/m, Thailand, 2015-2022



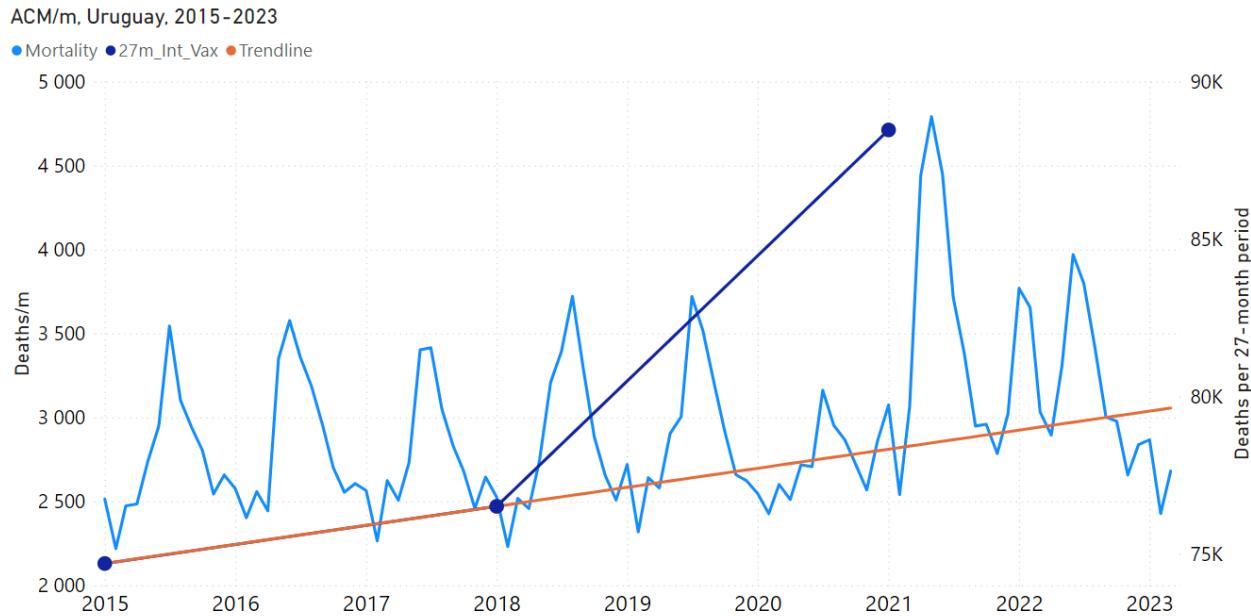


Figure 7: All-ages ACM by week (ACM/w) or by month (ACM/m) (light blue), 2015-2023, with integrations (dark blue), and trend lines (orange), for the vaccination period, for the 17 countries in this study (as labelled). The data sources are specified in Appendix A.

We also applied our alternative method (Method 2: Equations 1 to 3; “Special case of a single historic integrated point” method) to obtain the excess mortality in the vaccination periods of all 17 countries, using all the same integration limits as used in applying the first method (Method 1), above. These calculations are illustrated in Appendix E: Single-point method quantification of the excess mortality in the vaccination period.

Table 2 compares calculated all-ages values of vDFR, for the 17 countries in the study, for two different time periods of interest, and using two different quantification methods:

- (1) The first (left to right) column of numbers provides the calculated all-ages national vDFR values using the “Historical-trend baseline for a period (or peak) of mortality” method (Method 1) and the time period of the nominal January–February 2022 ACM peak (as illustrated in Figure 5 and Figure 4), for 14 of the 17 countries in the study, which have ACM data covering the said peak. Three countries have incomplete ACM data in the time period of interest (Argentina, Bolivia, Suriname). The time periods of integration, for both ACM and total

vaccine administration, are: 2021-week-42 through 2022-week-15 (26-week = 182-day integration, with by-week ACM data) and 2021-month-10 through 2022-month-04 (7-month = 212-day integration, with by-month ACM data).

- (2) The second column of numbers provides the error in the first-column numbers, which is from the mean deviation magnitude from the straight-line points used for the historic trend (Method 1, see Methods). It is therefore an “accuracy” error, arising from a variability due to often unavoidable seasonal uncertainty (see Methods).
- (3) The third column of numbers provides the calculated all-ages national vDFR values using the “Historical-trend baseline for a period (or peak) of mortality” method (Method 1) and the all-doses vaccination period (as illustrated in Figure 6 and Figure 7), for 15 of the 17 countries in the study, which have enough usable historic ACM data for the method to be applied. Meaningful “accuracy” errors in most cases cannot be evaluated for the numbers in this column. Two countries do not have enough ACM data (Chile, Peru), for more than one needed point of historic ACM integration, for a reliable vDFR value to be calculated, with the dataset used.
- (4) The fourth column of numbers provides the calculated all-ages national vDFR values using the “Special case of a single historic integrated point” method (Method 2) and the all-doses vaccination period (as illustrated in Appendix E), for all 17 countries in the study.
- (5) The fifth column of numbers provides the error in the fourth-column numbers, which is calculated using Equation 3 of the “Special case of a single historic integrated point” method (Method 2), with the slope (m) values given in Figure 23. As such, this error represents an estimate of the “accuracy” error, which is larger than a statistical error, since the statistically calculated error in m is affected by the seasonal spread in ACM-by-time-unit values.

Table 2: All-ages national vaccine dose fatality rates (vDFR) extracted from the time period of the nominal January–February 2022 ACM peak (“S22-peak”) and from the entire time period of vaccination (“v-period”), for the 17 countries in this study.

Country	Code	vDFR (%) (S22-peak)	error (%) (S22-peak)	vDFR (%) (v-period, Method 1)	vDFR (%) (v-period, Method 2)	error (%) (v-period, Method 2)
Argentina	ARG	—	—	0.13	0.13	0.01
Australia	AUS	0.032	0.003	0.049	0.048	0.005
Bolivia	BOL	—	—	0.30	0.301	0.005
Brazil	BRA	0.071	0.006	0.14	0.139	0.009
Chile	CHL	0.081	0.002	—	0.078	0.007
Colombia	COL	0.063	0.001	0.13	0.129	0.003
Ecuador	ECU	0.047	0.003	0.097	0.099	0.003
Malaysia	MYS	0.068	0.003	0.072	0.073	0.006
New Zealand	NZL	0.017	0.003	0.019	0.023	0.007
Paraguay	PRY	0.14	0.02	0.24	0.24	0.02
Peru	PER	0.030	0.005	—	0.06	0.02
Philippines	PHL	0.056	0.004	0.16	0.157	0.005
Singapore	SGP	0.041	0.003	0.035	0.035	0.003
South Africa	ZAF	0.19	0.01	0.40	0.44	0.03
Suriname	SUR	—	—	0.31	0.30	0.02
Thailand	THA	0.051	0.007	0.085	0.080	0.005
Uruguay	URY	0.20	0.01	0.11	0.12	0.03

Figure 8 shows all-ages national excess ACM in the time period of the nominal January–February 2022 ACM peaks (filled circles) or in the time periods of vaccination (open diamonds), as obtained from the analyses illustrated in Figure 5 or Figure 23, respectively, versus the all-ages all-doses national total vaccine injections in each of the same time periods as the time windows of ACM integration, as obtained from the analyses illustrated in Figure 4 or Figure 6, respectively, for the 17 countries of the study, subject to data availability. The full range is shown in the top panel, and an expanded view near the origin is shown in the bottom panel.

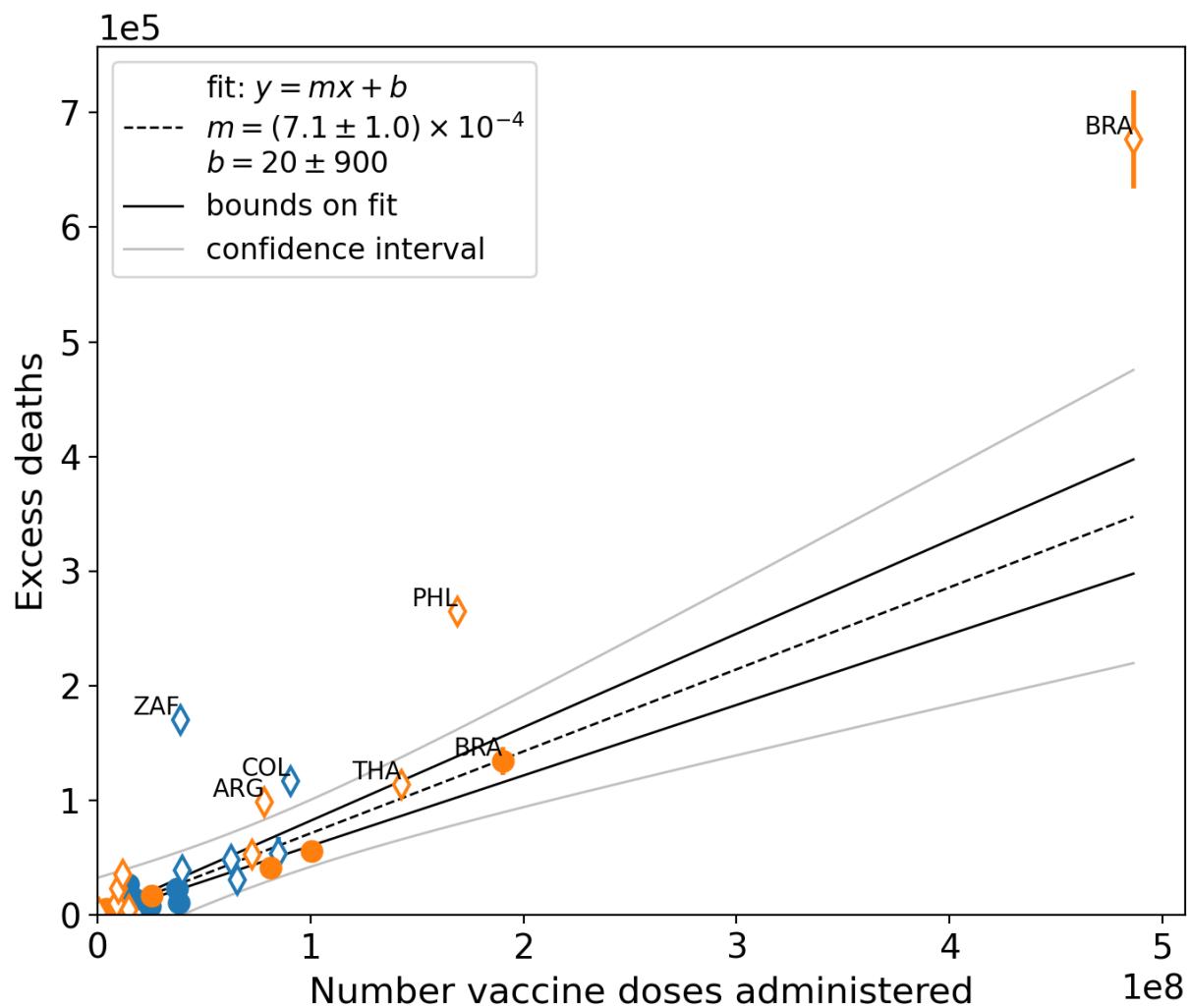
The results from by-week ACM data (blue symbols) and by-month ACM data (orange symbols) are comparable, even though they use different integration-window widths, because they are excess mortalities in anomalous peaks rather than total mortalities.

The overall (all points in Figure 8) Pearson correlation coefficient is $r = +0.94$.

As per its legend, Figure 8 also shows:

- (1) The best straight-line fit (least squares fit) using the Y-errors in the points, assumed to be proportional to the true error (dashed black line)
- (2) Upper and lower bounds on the best straight-line fit, using largest intercept and largest slope or smallest intercept and smallest slope, respectively, within the error bounds on intercept and slope (solid black lines)
- (3) The standard confidence-interval limits, not using Y-errors, and therefore assuming equal and normal Y-errors for all points (curved grey lines)

The best-line fit parameters are as specified in the figure: Y-intercept = $+20 \pm 900$ deaths, slope = $(7.1 \pm 1.0) \times 10^{-4}$ deaths per injection ($r = +0.94$). For comparison, if only the data for the time period of the nominal January–February 2022 ACM peak is used (filled circles in Figure 8), one obtains: Y-intercept = $-500 \pm 1,000$ deaths, slope = $(6.4 \pm 0.9) \times 10^{-4}$ deaths per injection ($r = +0.97$).



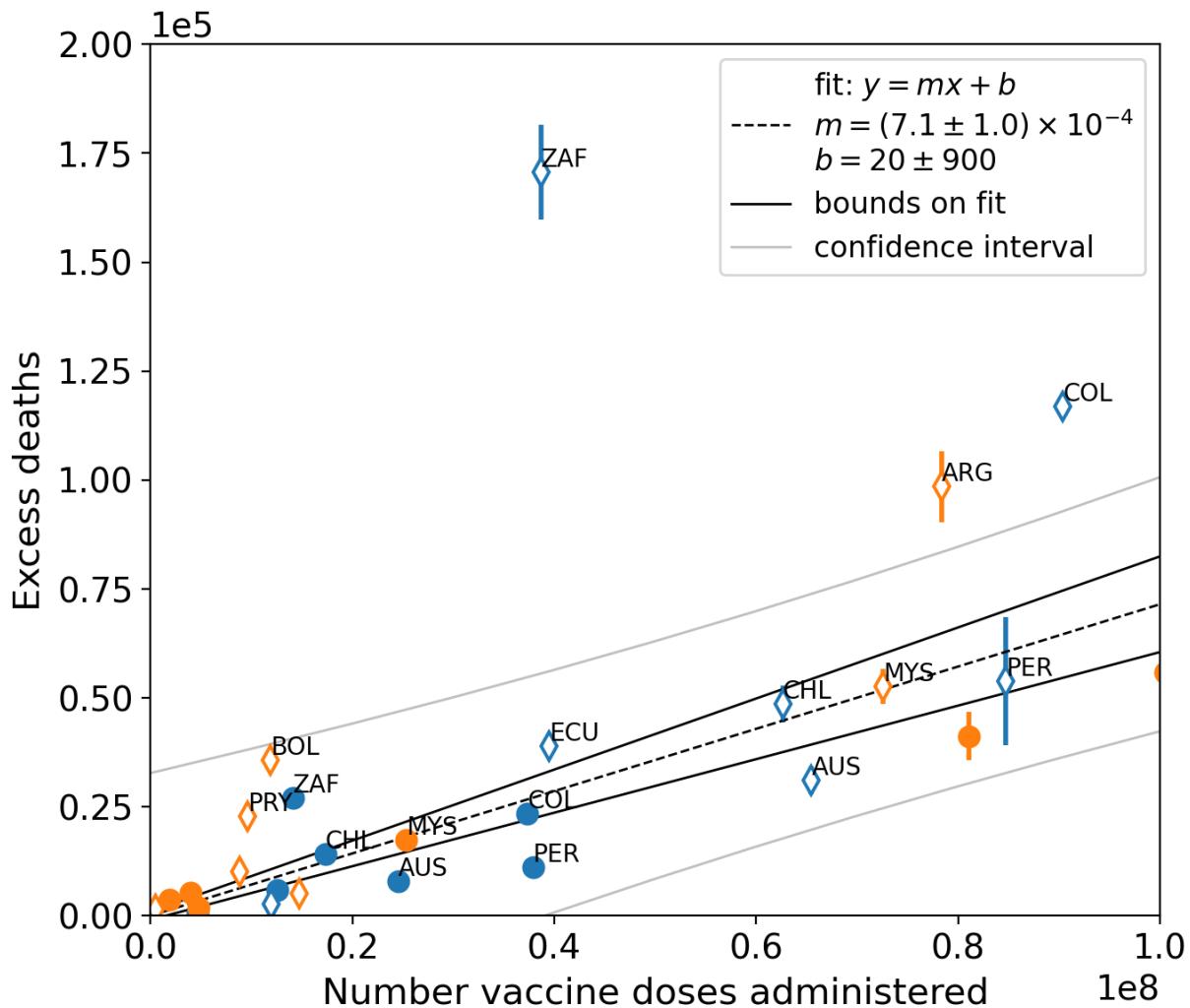


Figure 8: All-ages national excess ACM in the time period of the nominal January–February 2022 ACM peaks (filled circles) or during the vaccination period (open diamonds, Method 2 calculation method) versus the all-ages national total vaccine injections in the same time period, using both by-week (blue) and by-month (orange) ACM data, for all 17 countries (as labelled). Full range (top panel). Expanded view (bottom panel). Pearson correlation coefficient is $r = +0.94$.

In Figure 8, the points for the vaccination period (open diamonds) have generally larger Y-values than the points for the time period of the nominal January–February 2022 ACM peak (filled circles), relative to the overall trend line. This would correspond to generally larger values of country-specific all-ages vDFR, since $vDFR = Y\text{-value}/X\text{-value}$ for each point. This systematic difference is analyzed in Figure 9.

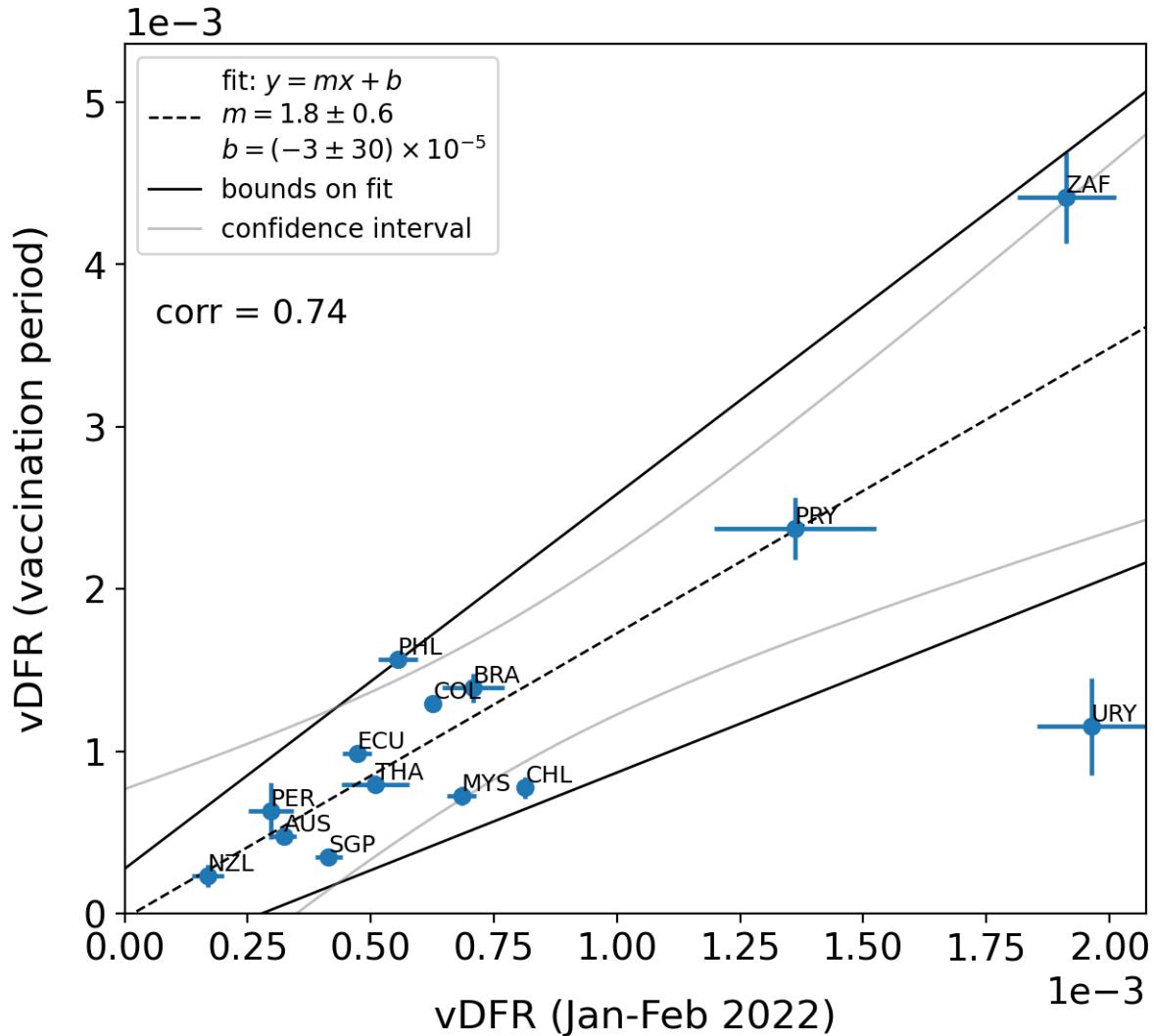


Figure 9: vDFR (vaccination period) versus vDFR (January-February 2022 ACM peak period), with correlation analysis. The vDFR values are expressed as fractions, not %, and the scale is $\times 10^{-3}$, so “4” is 0.4 %, and so on. The analysis lines have the same meanings as in Figure 8. Pearson correlation coefficient is $r = +0.74$.

As per its legend, in Figure 9 the slope is 1.8 ± 0.6 and the intercept is zero, within error (-3 ± 30). The Pearson correlation coefficient is $+0.74$.

South Africa has the largest value of all-ages vDFR (0.44 %), of the 17 study countries, followed by Suriname (0.30 %), Bolivia (0.301 %), Paraguay (0.24 %), and Philippines (0.157 %). This is illustrated in Figure 10, for all the countries and both integration-time periods, compared to the lists of vaccine manufacturers by country (OWID, 2023b).

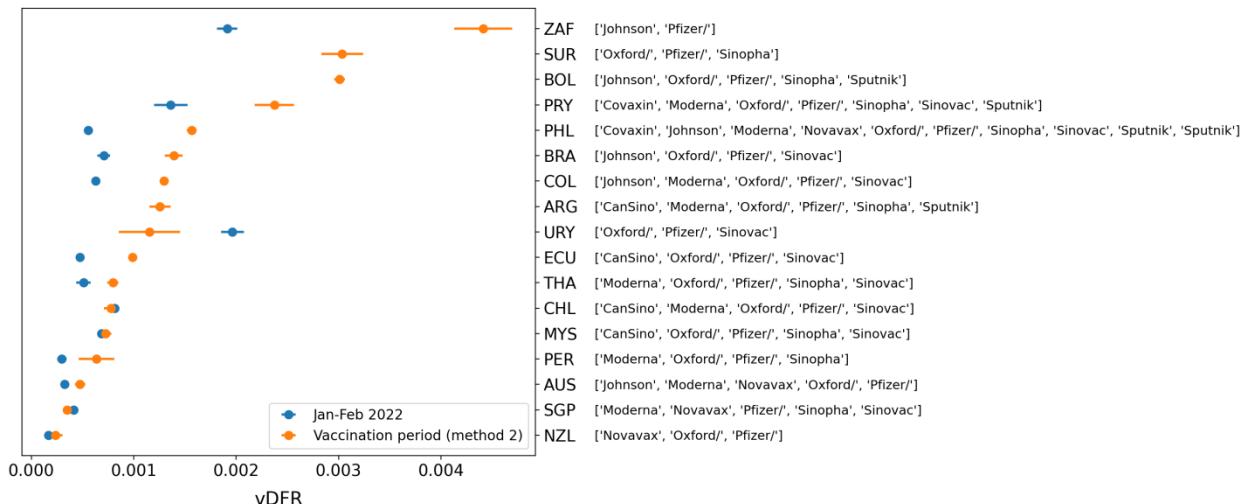
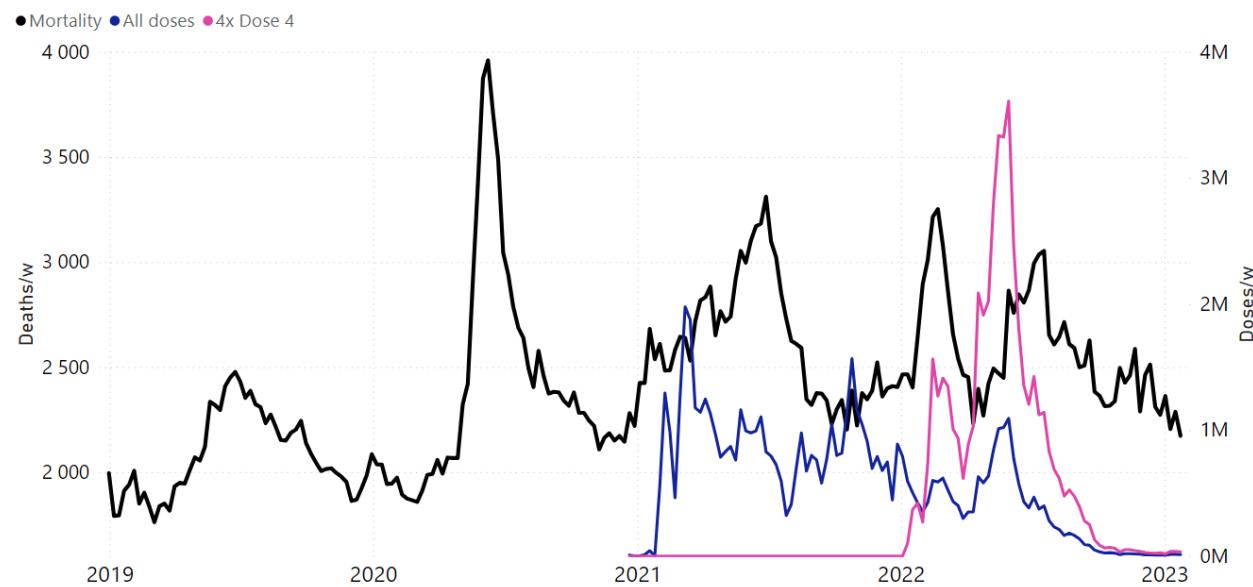


Figure 10: All-ages vDFR values for the vaccination period (orange, meth. 2) and for the period of the January–February 2022 peak in ACM (blue), by country, in decreasing order, compared to the lists of vaccine manufacturers. Values of all-ages vDFR are from Table 2.

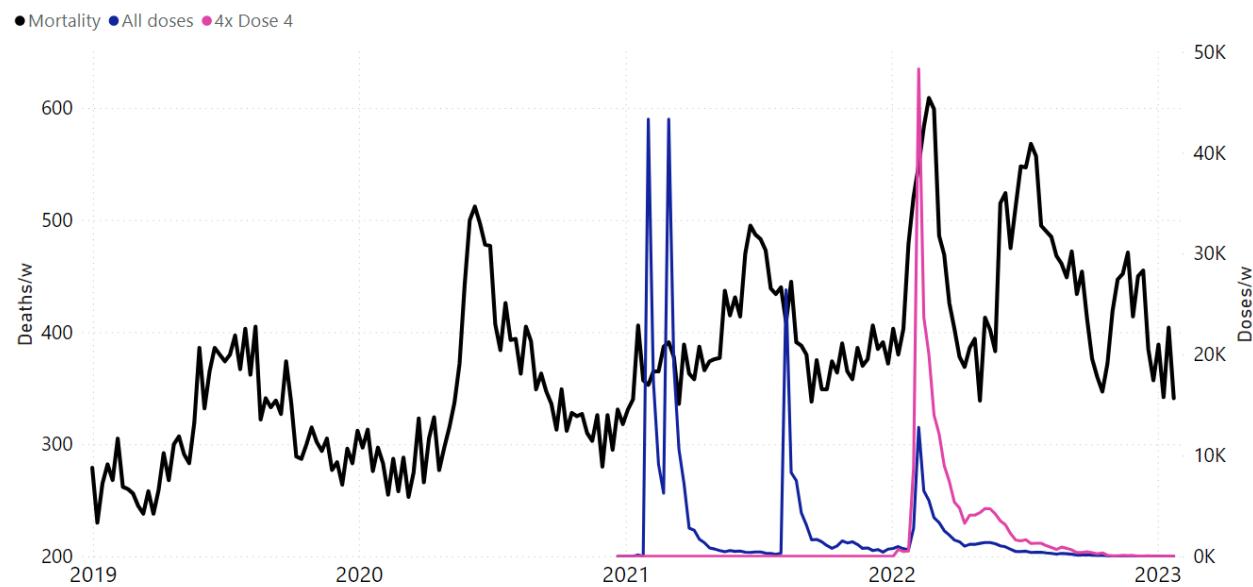
5.4 Chile by age group dose 4 — National ACM and vaccine dose 4 rollouts (90+, 85-89, 80-84, 75-79, 70-74, 65-69, and 60-64 years)

Figure 11 shows the national ACM by week for Chile (black), from 2019 to 2023, and all-doses (blue) and dose-4 (x4 amplification) (pink) COVID-19 vaccine rollouts, for all ages and for each of the age groups analyzed (90+ years through 60-64 years), as labelled. For many of the age groups, the four different doses are visually resolved as distinct peaks in the all-doses data, which is confirmed by the dose-specific data (not shown).

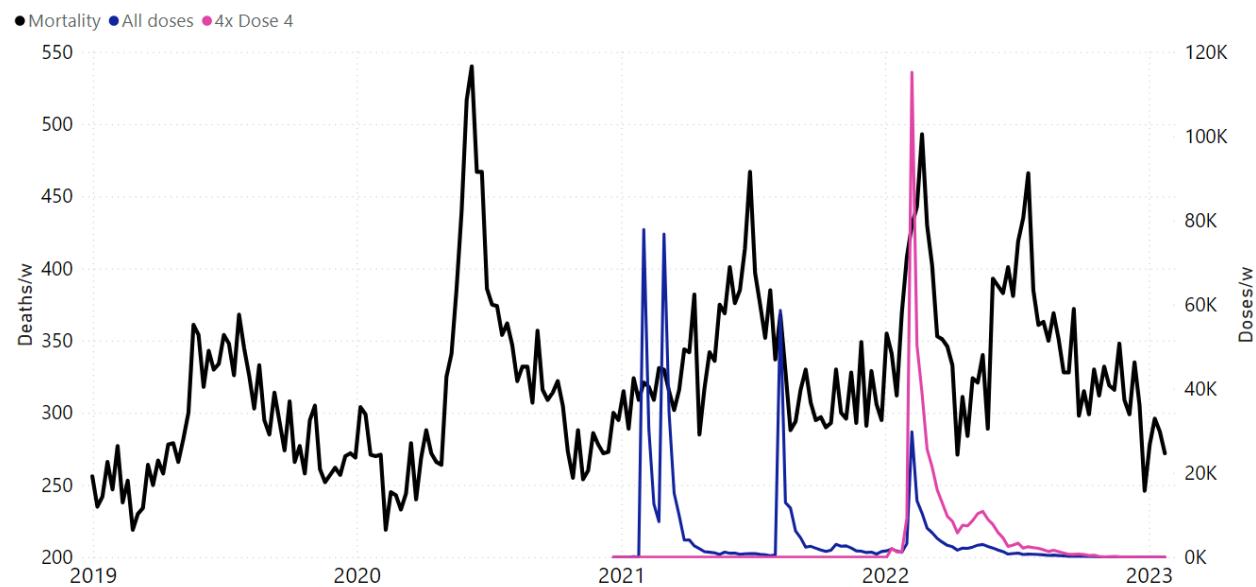
ACM and vaccination by week, Chile, all ages, 2019–2022



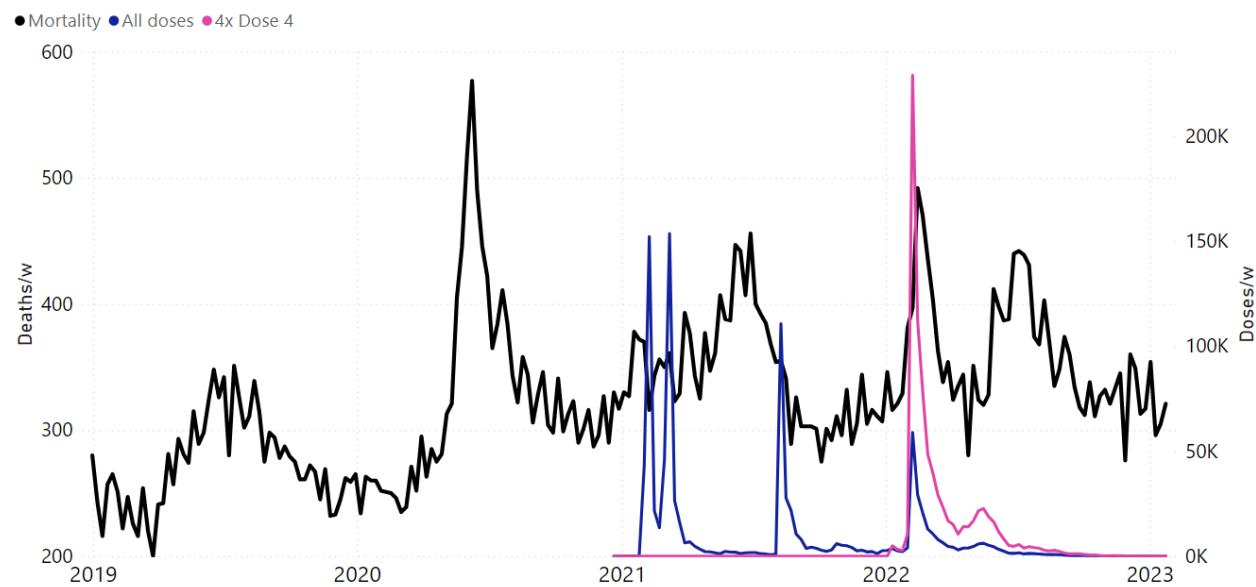
ACM and vaccination by week, Chile, 90+, 2019–2022



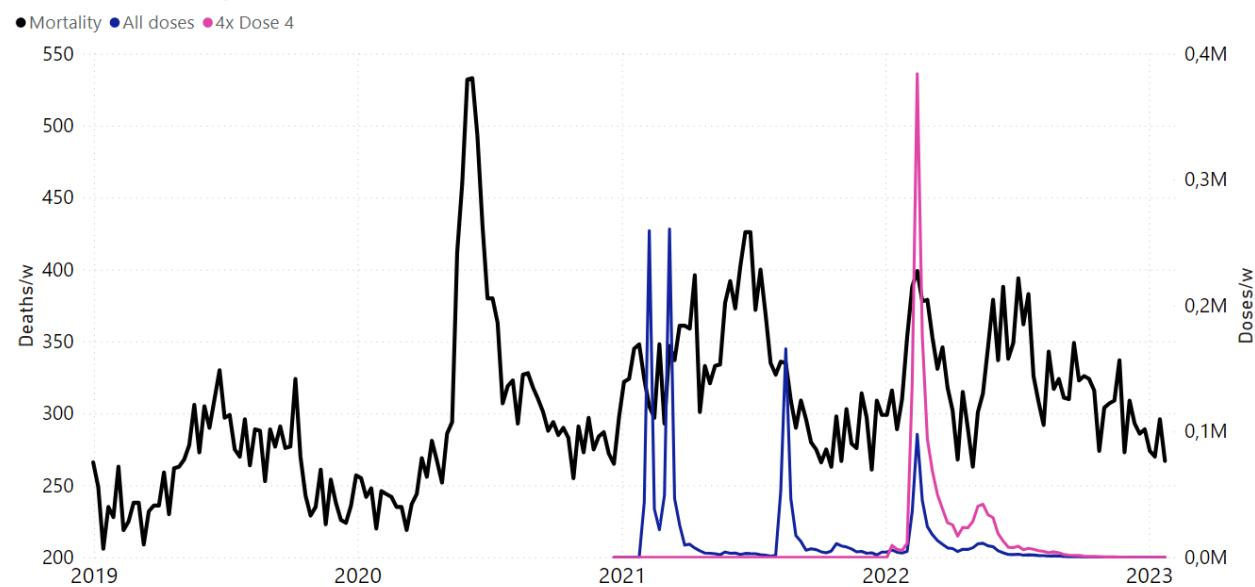
ACM and vaccination by week, Chile, 85–89, 2019–2022



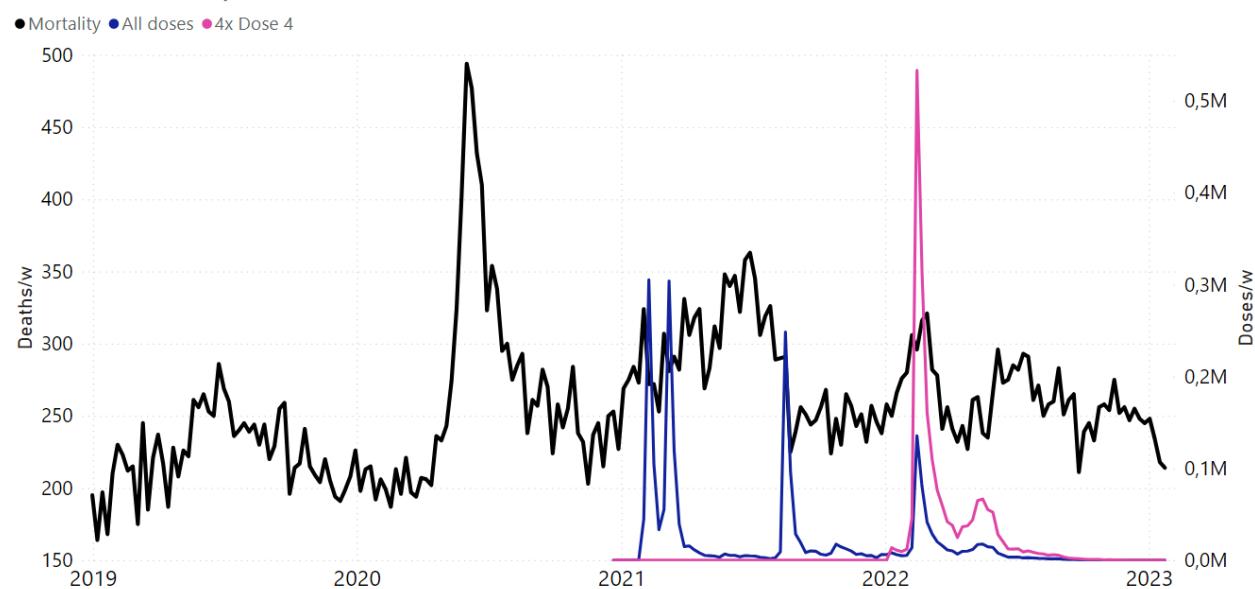
ACM and vaccination by week, Chile, 80–84, 2019–2022



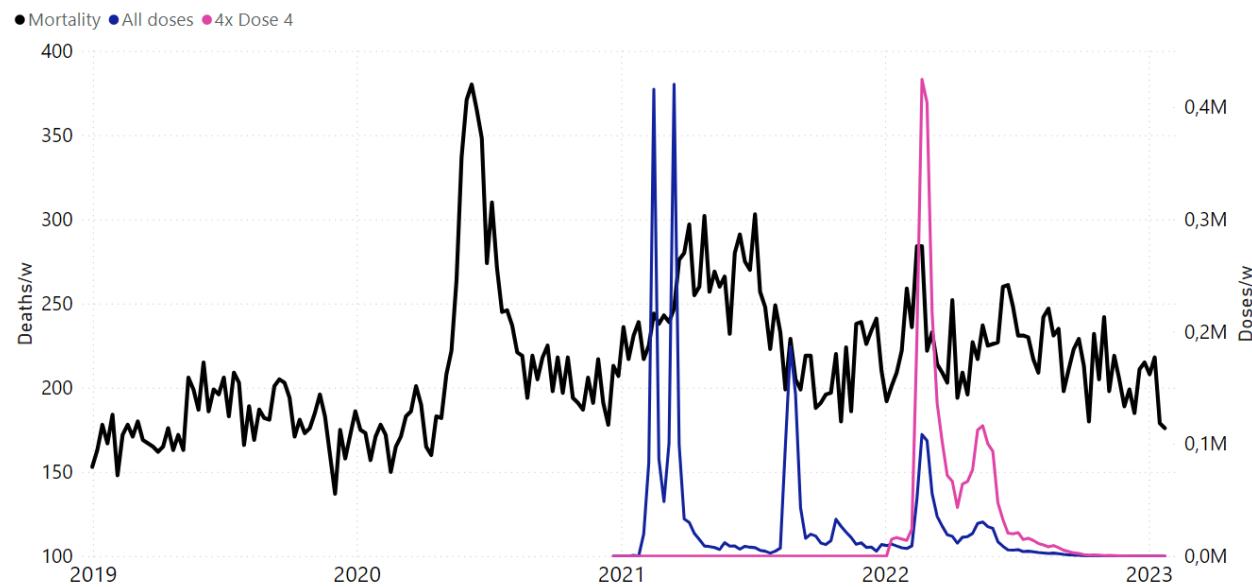
ACM and vaccination by week, Chile, 75-79, 2019-2022



ACM and vaccination by week, Chile, 70-74, 2019-2022



ACM and vaccination by week, Chile, 65-69, 2019-2022



ACM and vaccination by week, Chile, 60-64, 2019-2022

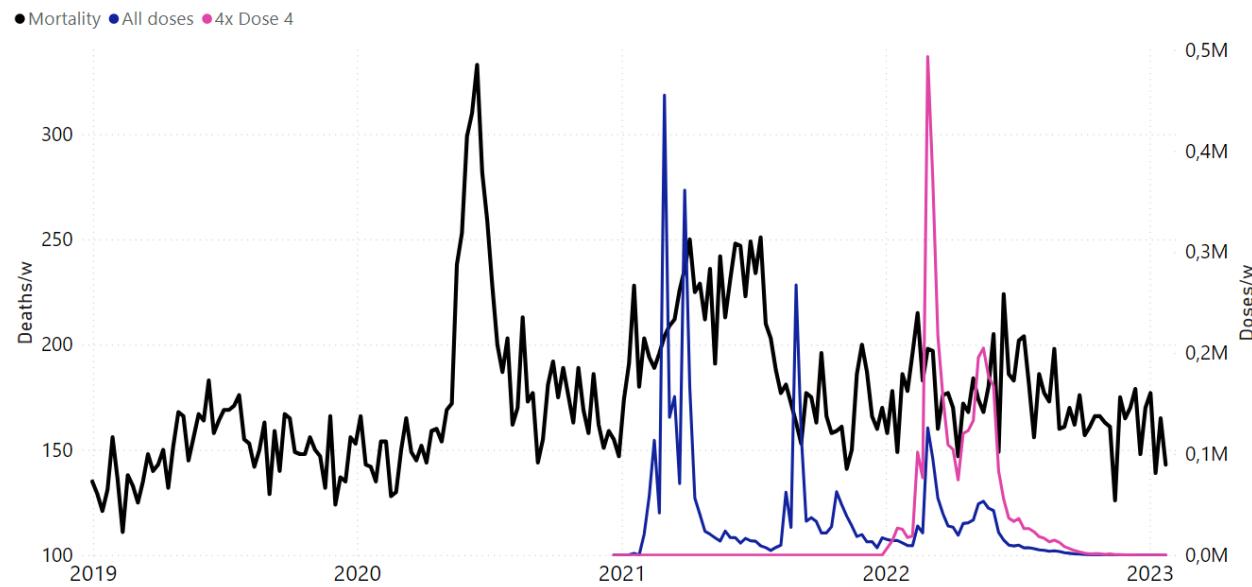
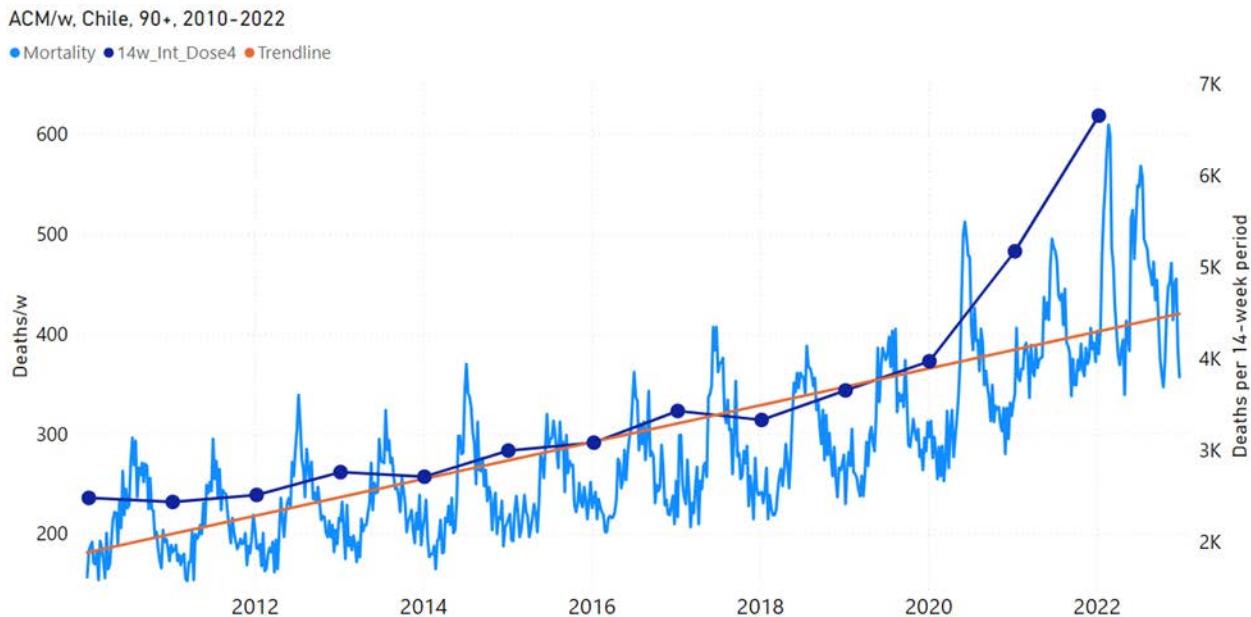


Figure 11: Chile (with rollouts), ACM by week (black), 2019-2023, by age group (90+, 85-89 ... 60-64 years), with vaccine rollouts all-doses (blue) and dose-4 (x4 amplified, pink). The data sources are specified in Appendix A.

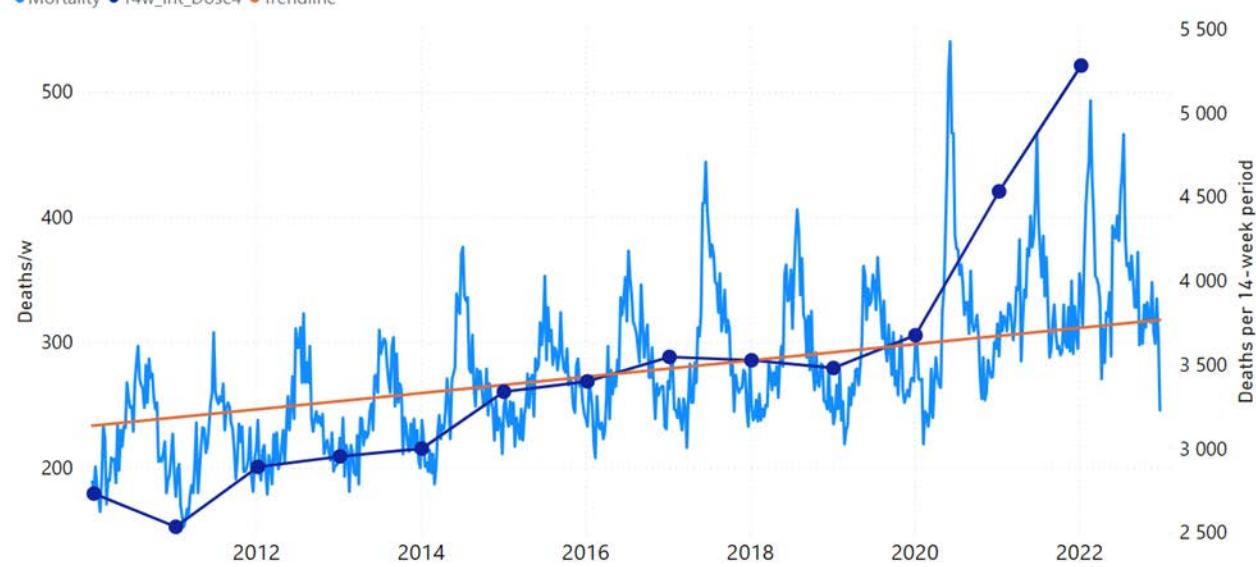
Figure 12 shows the national ACM by week for Chile (light blue), from 2010 to 2023, and the 14-week integration points (dark blue) and trend line (orange) that were used to quantify excess ACM in the January-February 2022 peak of interest ("Historical-trend

baseline for a period (or peak) of mortality" method, Method 1, see Methods), corresponding to the vaccine dose 4 rollout, for each of the age groups (90+ years through 60-64 years, as indicated). Here the integration windows (14 weeks) are selected for the age-group data, and are therefore different than the integration window used for all-ages (Figure 5 and Figure 4) (see Appendix C). The last integration point (in 2022) is for the vaccine dose 4 period itself, and each integration point is positioned in time at the start of the integration window. See Methods for details.



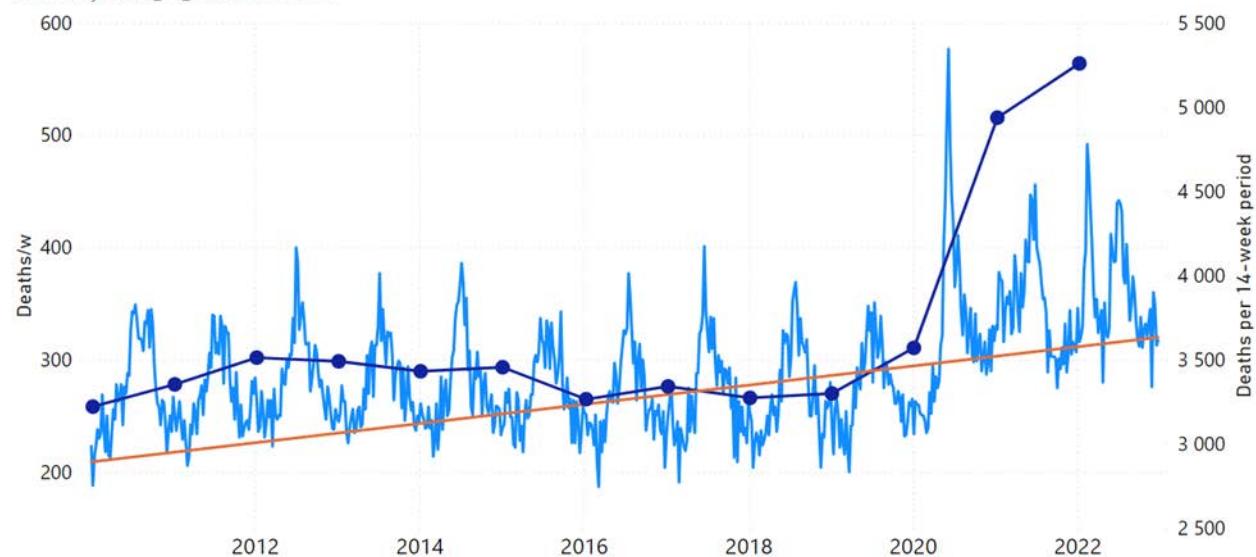
ACM/w, Chile, 85-89, 2010-2022

● Mortality ● 14w_Int_Dose4 ● Trendline



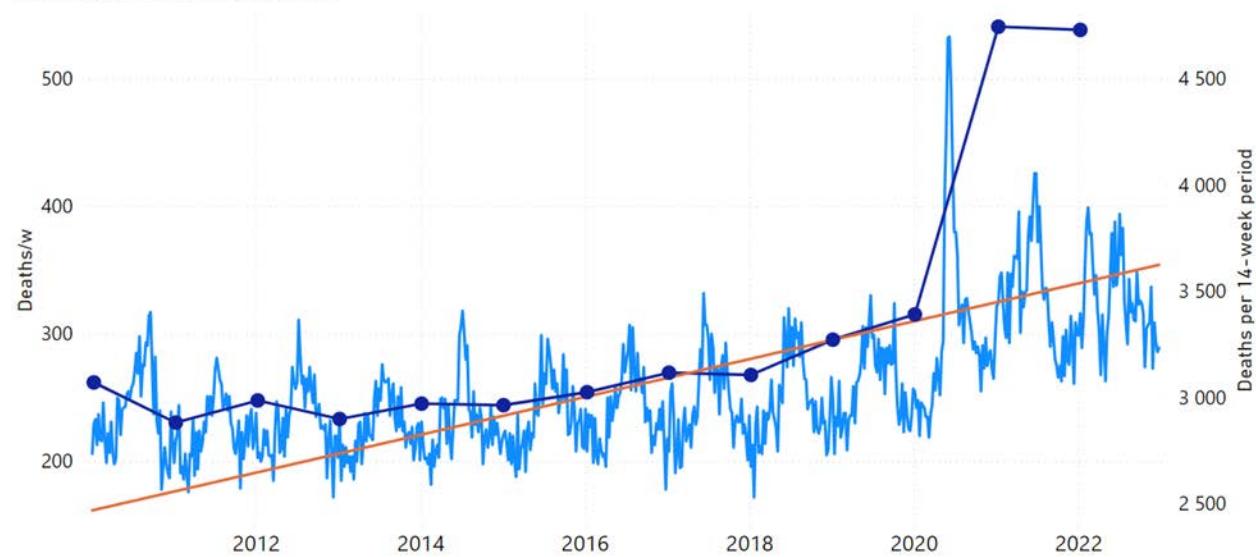
ACM/w, Chile, 80-84, 2010-2022

● Mortality ● 14w_Int_Dose4 ● Trendline



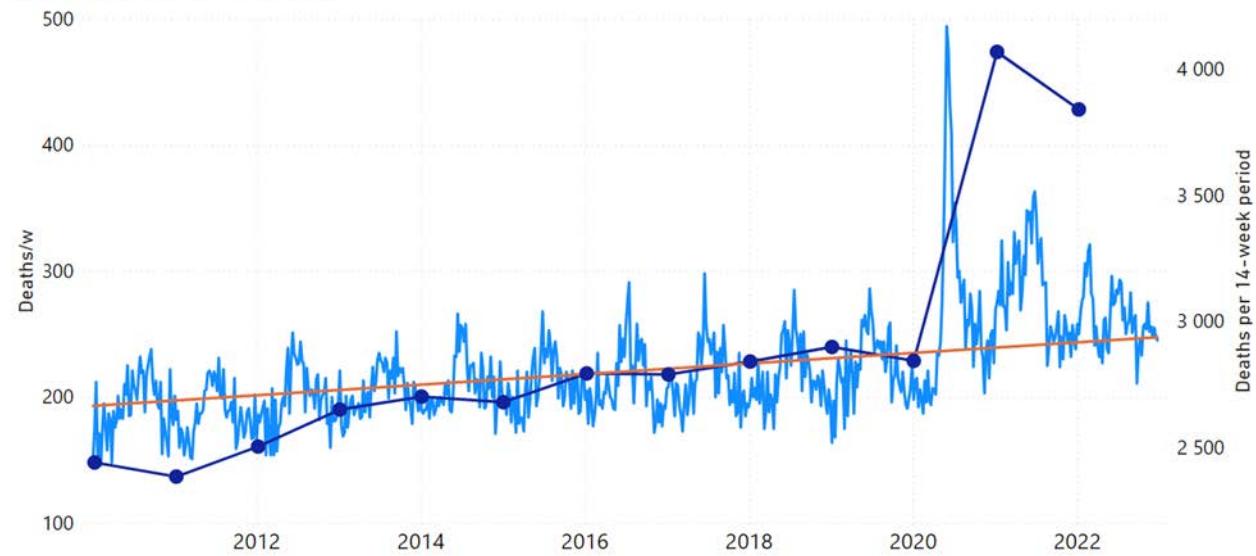
ACM/w, Chile, 75-79, 2010-2022

● Mortality ● 14w_Int_Dose4 ● Trendline



ACM/w, Chile, 70-74, 2010-2022

● Mortality ● 14w_Int_Dose4 ● Trendline



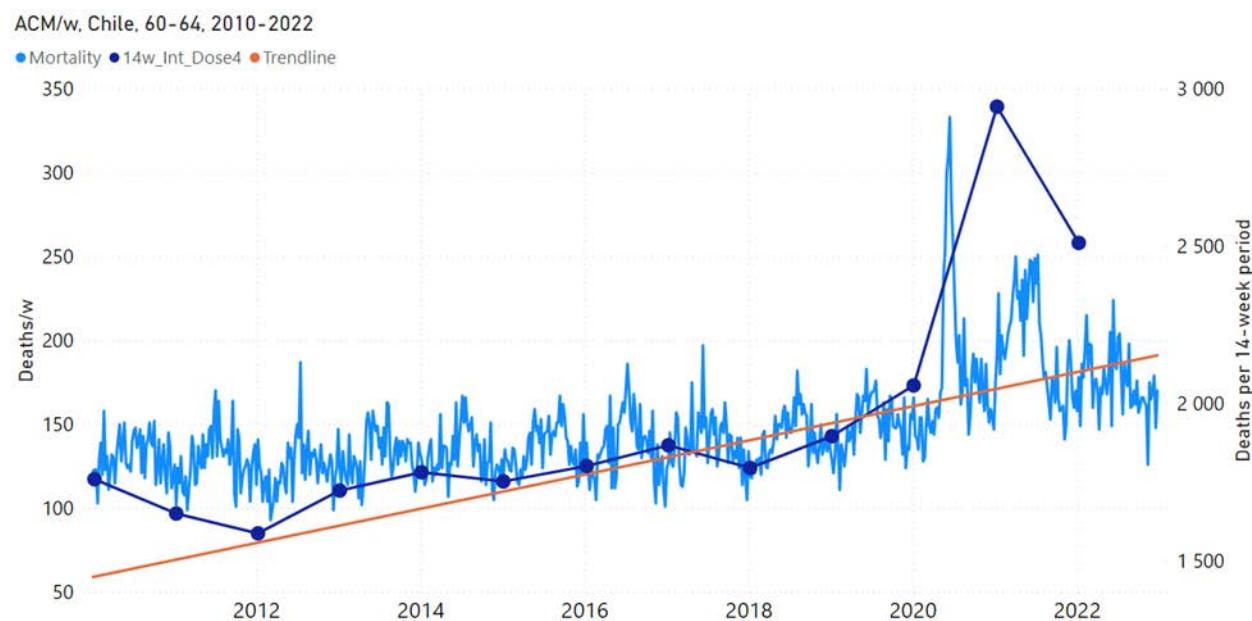
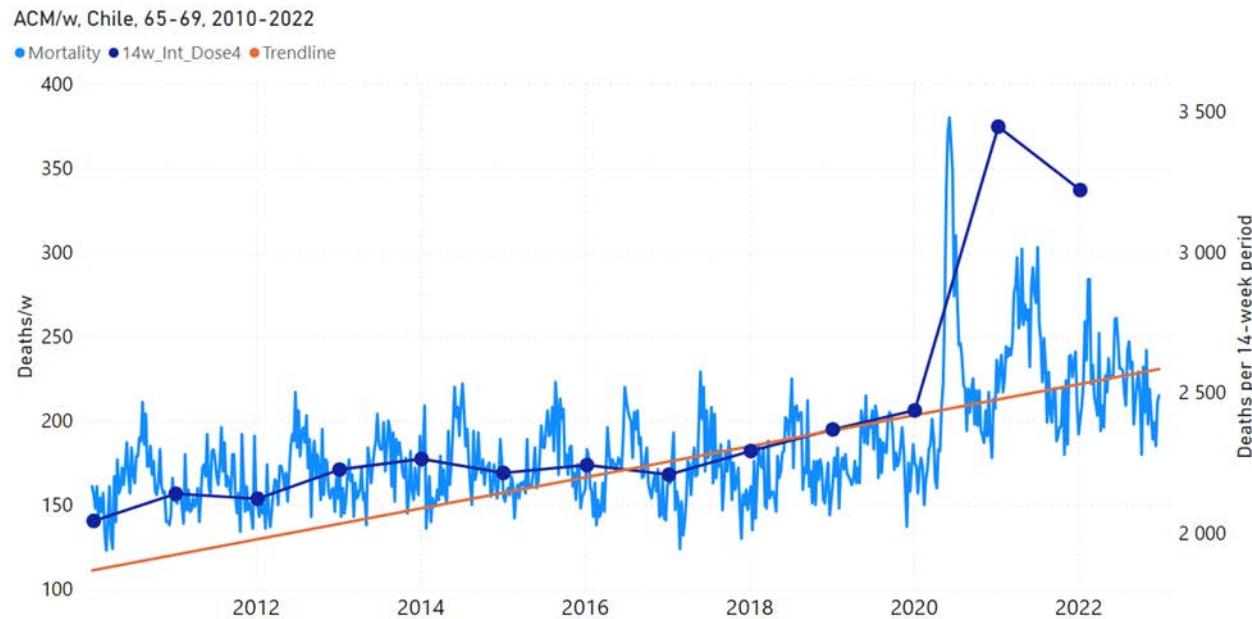


Figure 12: Chile (dose 4), ACM by week (light blue), 2010-2022, by age group (90+, 85-89 ... 60-64 years), with integrations, 14-week dose-4-vaccination-period (dark blue, points) and trend line (orange). The data sources are specified in Appendix A.

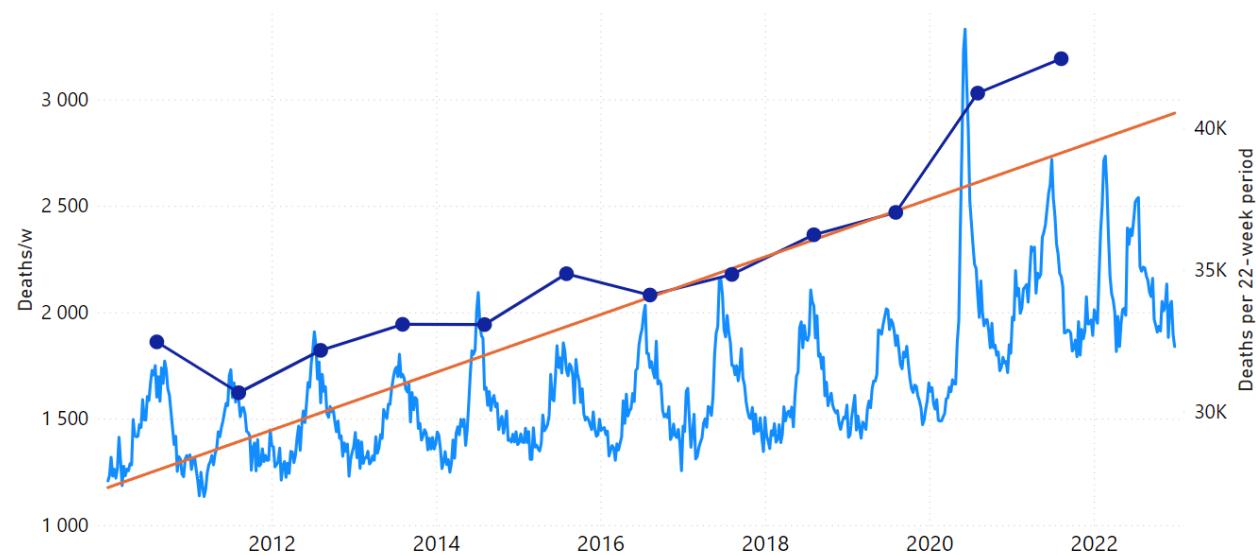
5.5 Chile by age group dose 3 — National ACM and vaccine dose 3 rollouts (90+, 85-89, 80-84, 75-79, 70-74, 65-69, and 60-64 years)

Since the dose-3 rollout is well defined and distinct from the other rollouts, in specific age groups in Chile (Figure 11), and since there is excess mortality over the time period of the dose-3 rollout, we additionally quantified for dose 3 in Chile, even though there is conflation with a seasonal peak in ACM and no sharp and distinct isolated associated peak in ACM, although 90+ years does show a well-formed shoulder (Figure 11) (see Methods, and see below).

Figure 13 shows the national ACM by week for Chile (light blue), from 2010 to 2023, and the 22-week-integration points (dark blue) and trend line (orange) that were used to quantify excess ACM in the nominal July-August 2021 region of interest (“Historical-trend baseline for a period (or peak) of mortality” method, Method 1, see Methods), corresponding to the vaccine dose 3 rollout, for each of the age groups (60+ years, and 90+ years through 60-64 years, as indicated). Here, as in Figure 12 and Figure 11, the integration windows (22 weeks) are selected for the age-group data (Appendix C). The last integration point (in 2021) is for the vaccine dose 3 period itself, and each integration point is positioned in time at the start of the integration window. See Methods for details.

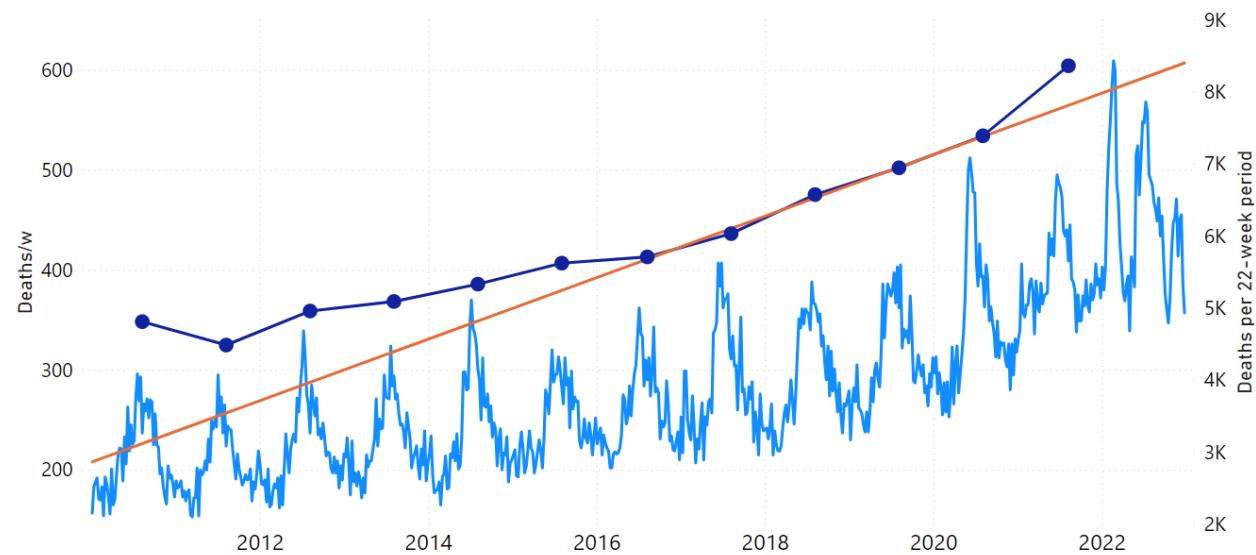
ACM/w, Chile, 60+, 2010-2022

● Mortality ● 22w_Int_Dose3 ● Trendline



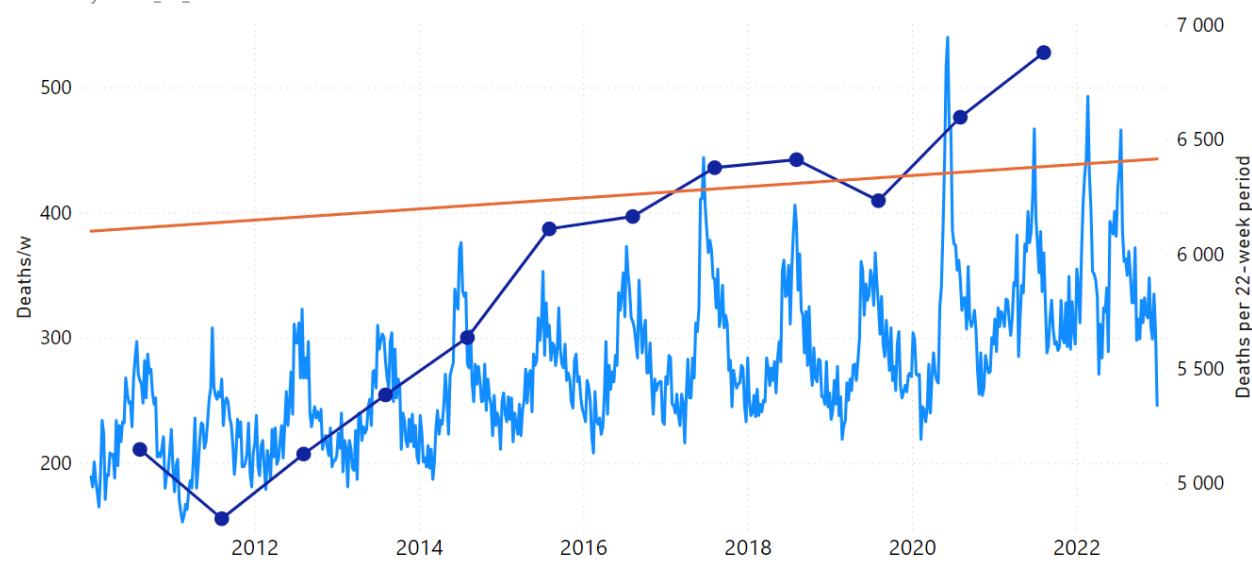
ACM/w, Chile, 90+, 2010-2022

● Mortality ● 22w_Int_Dose3 ● Trendline



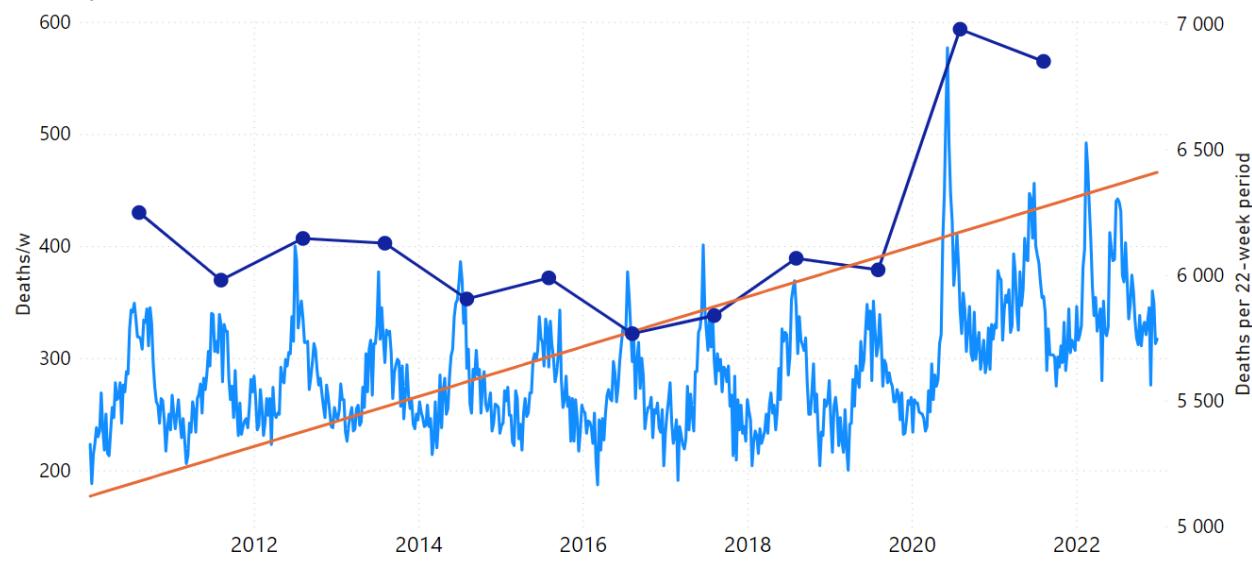
ACM/w, Chile, 85-89, 2010-2022

● Mortality ● 22w_Int_Dose3 ● Trendline

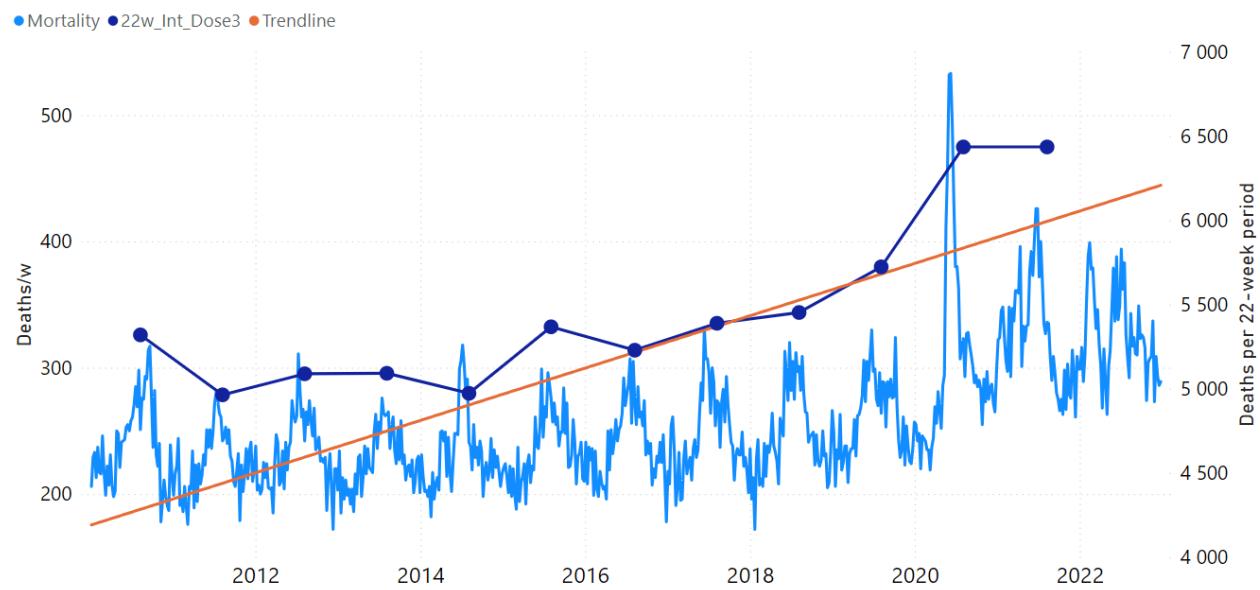


ACM/w, Chile, 80-84, 2010-2022

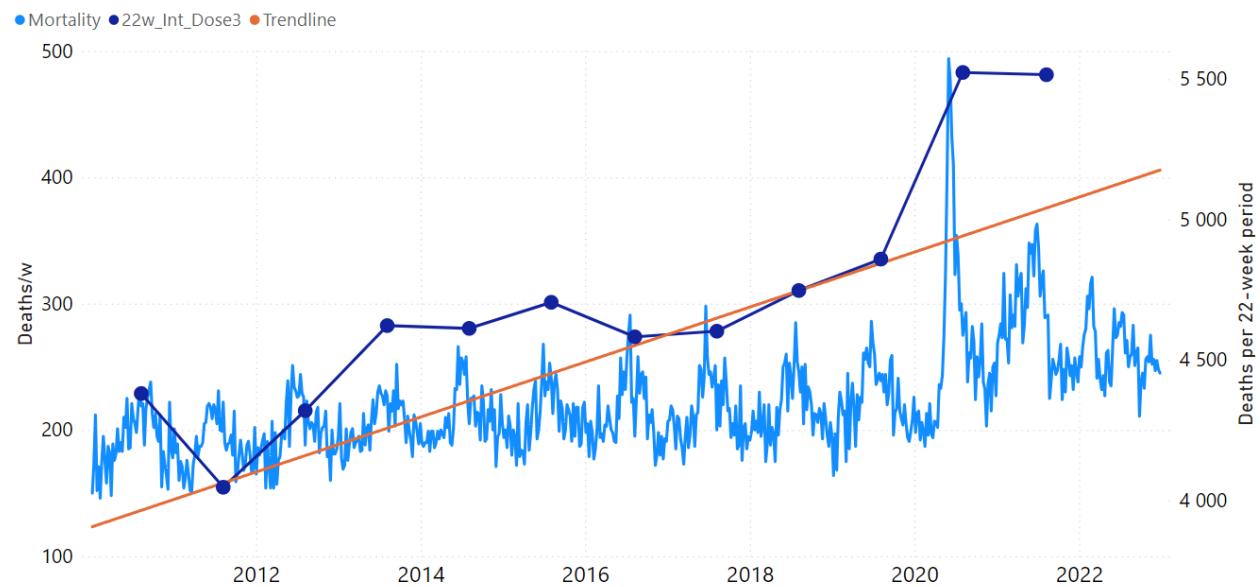
● Mortality ● 22w_Int_Dose3 ● Trendline



ACM/w, Chile, 75-79, 2010-2022



ACM/w, Chile, 70-74, 2010-2022



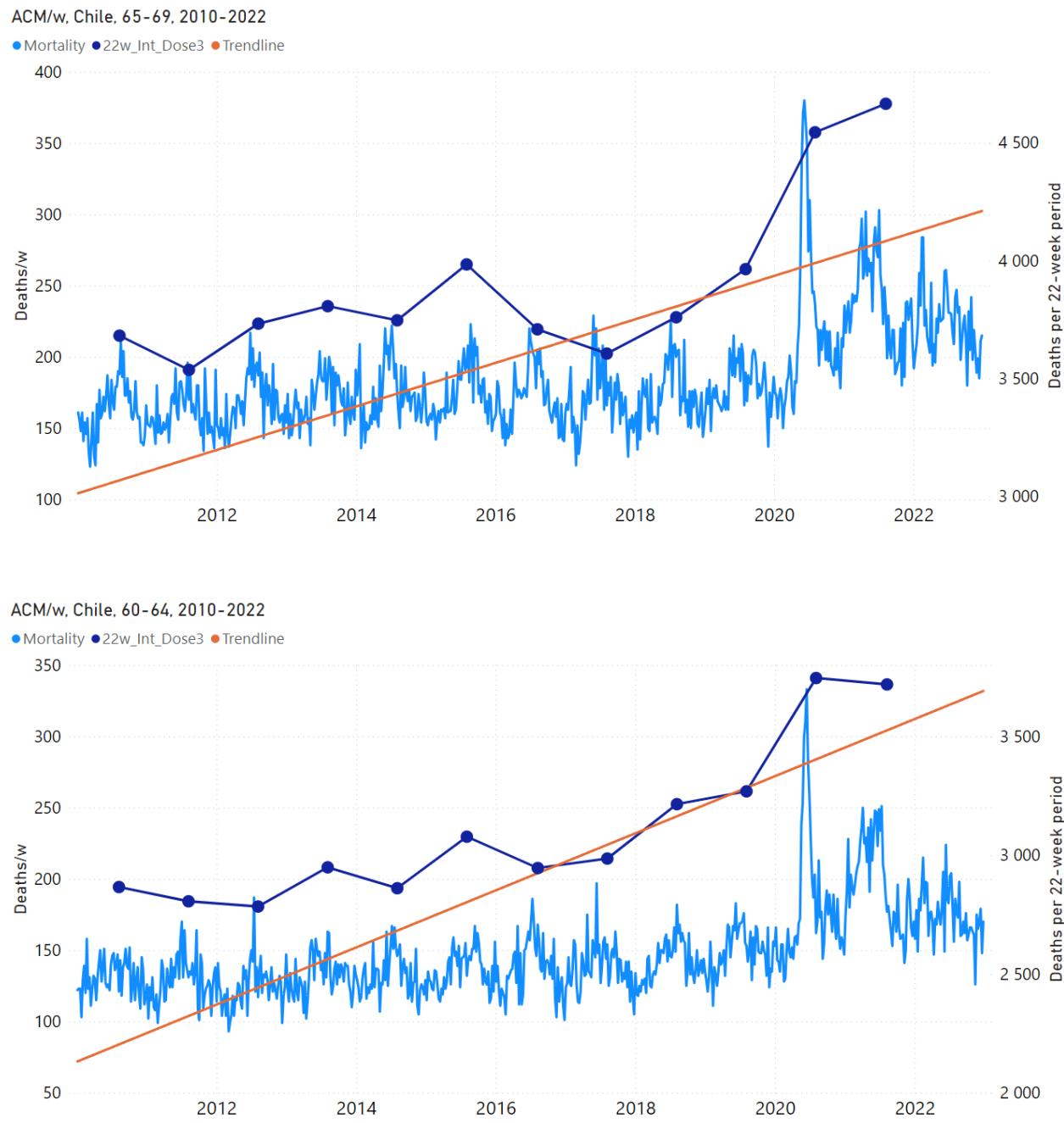
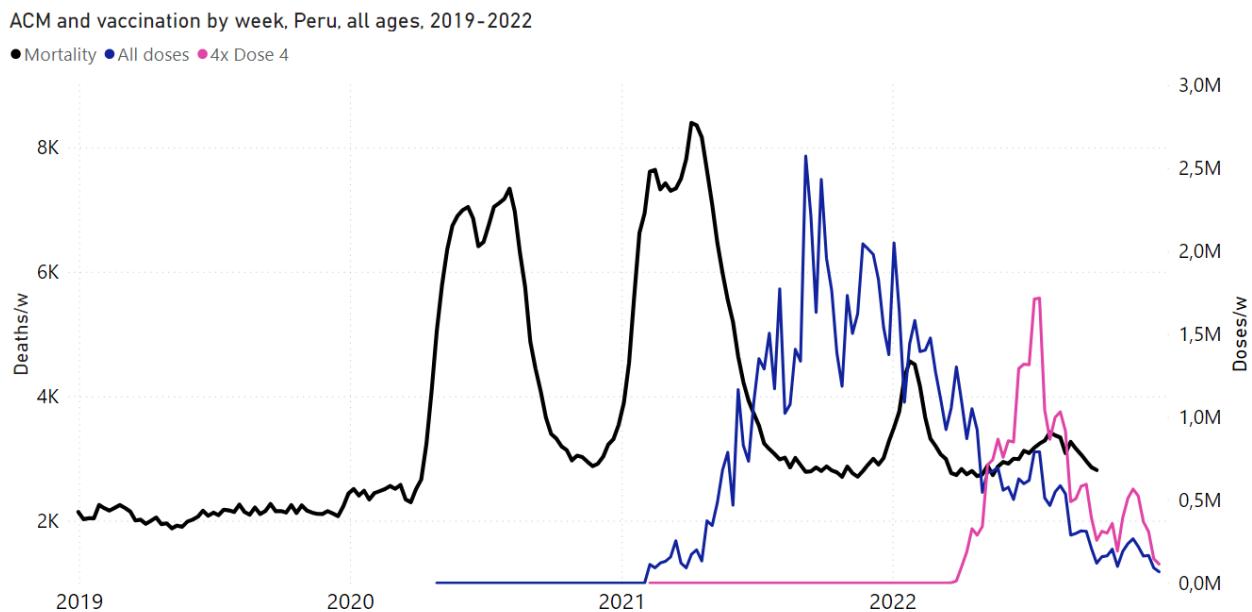


Figure 13: Chile (dose 3), ACM by week (light blue), 2010-2022, by age group (60+, 90+, 85-89 ... 60-64 years), with integrations, 22-week dose-3-vaccination-period (dark blue, points) and trend line (orange). The data sources are specified in Appendix A.

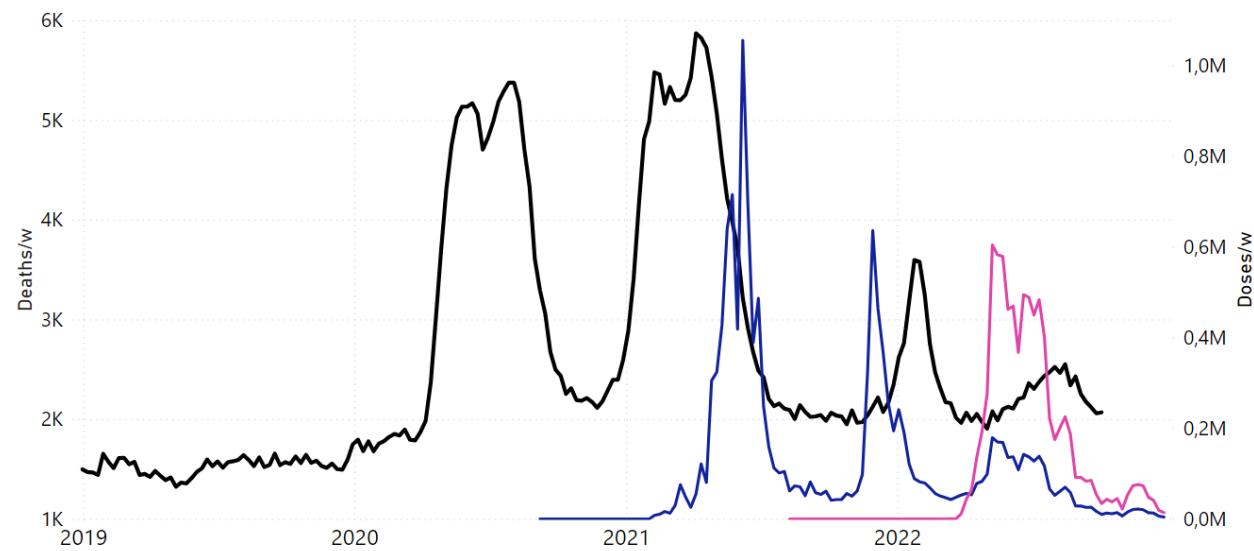
5.6 Peru by age group dose 3 — National ACM and vaccine dose 3 rollouts (90+, 85-89, 80-84, 75-79, 70-74, 65-69, and 60-64 years)

Figure 14 shows the national ACM by week for Peru (black), from 2019 to 2022, and dose-specific (all doses, blue; and dose 4, pink) COVID-19 vaccine rollouts, for all-ages and for each of the age groups examined (60+, 90+, 85-89, 80-84, 75-79, 70-74, 65-69, 60-64, 50-59, 40-49, 30-39, 20-29, and 0-19 years), as labelled. All-doses vaccine administration by week is in blue, dose-4 vaccine administration by week (x4 amplification) is in pink. For many of the age groups, the four different doses are visually resolved as distinct peaks in the all-doses data, which is confirmed by the dose-specific data (not shown).



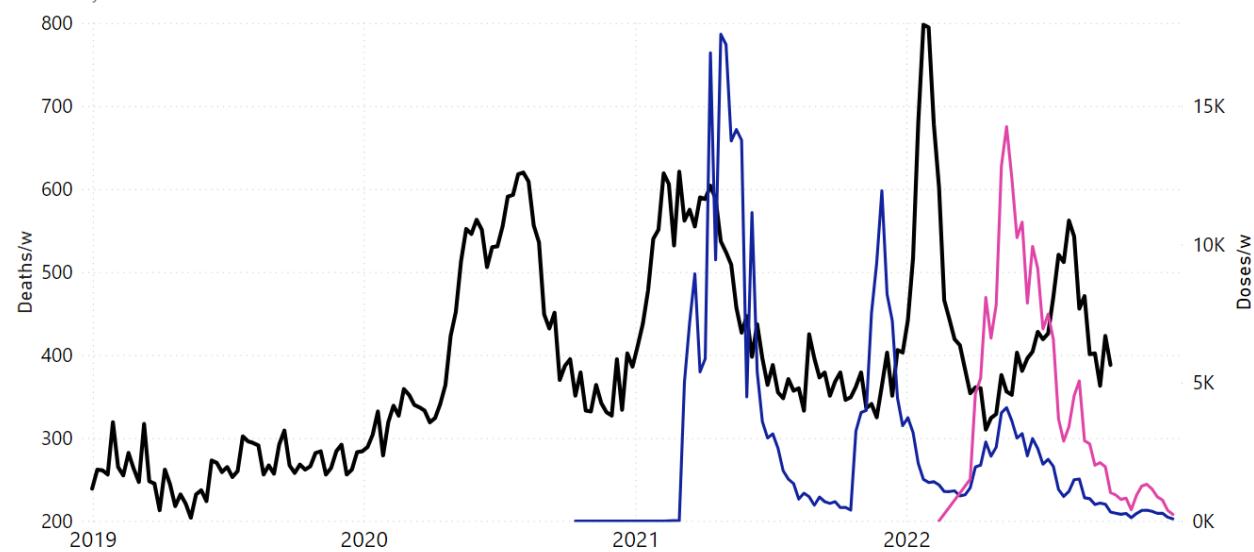
ACM and vaccination by week, Peru, 60+, 2019-2022

● Mortality ● All doses ● 4x Dose 4

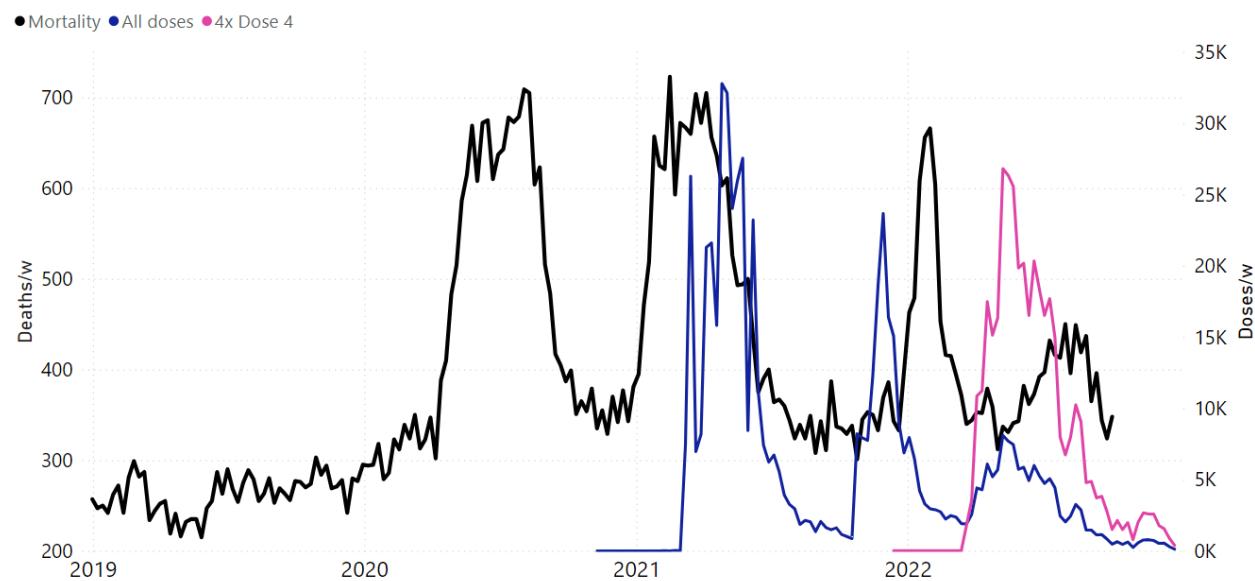


ACM and vaccination by week, Peru, 90+, 2019-2022

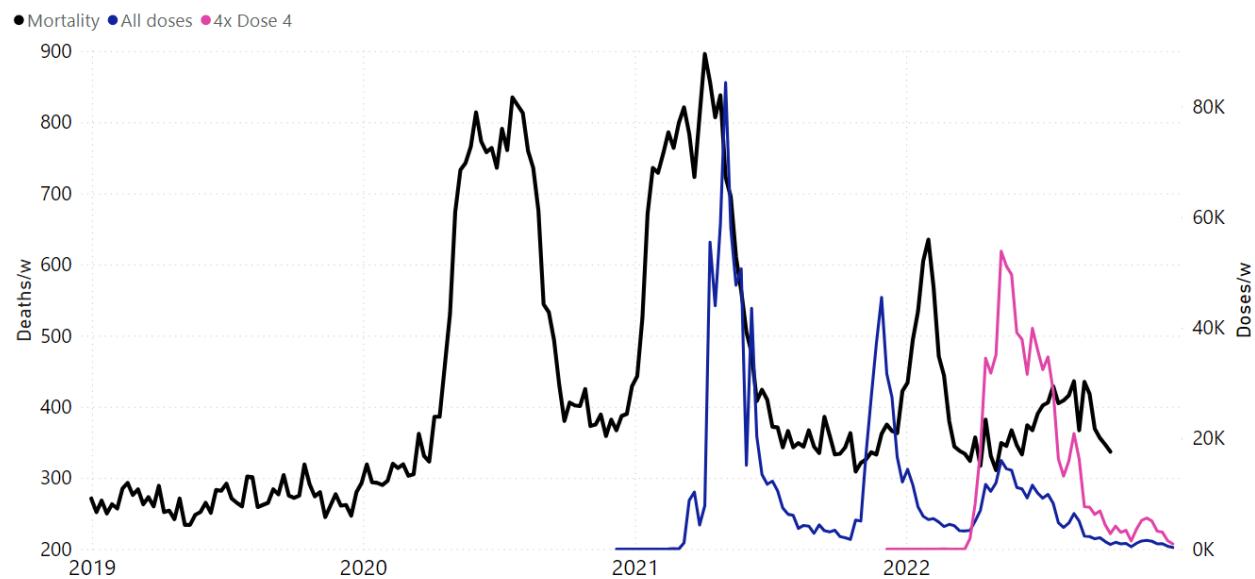
● Mortality ● All doses ● 4x Dose 4



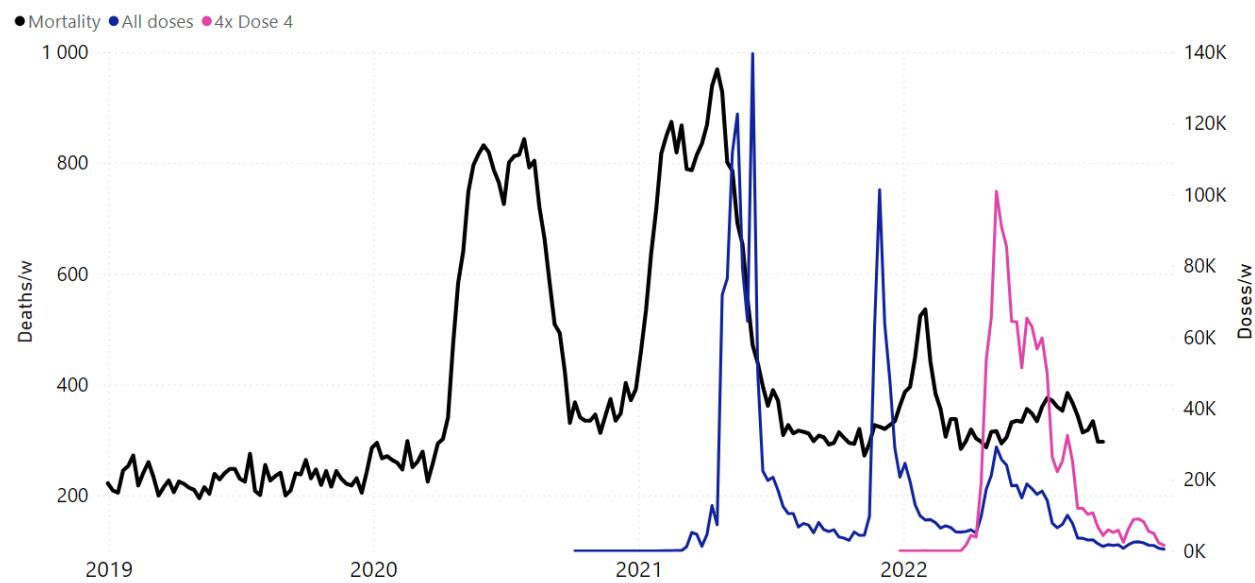
ACM and vaccination by week, Peru, 85-89, 2019-2022



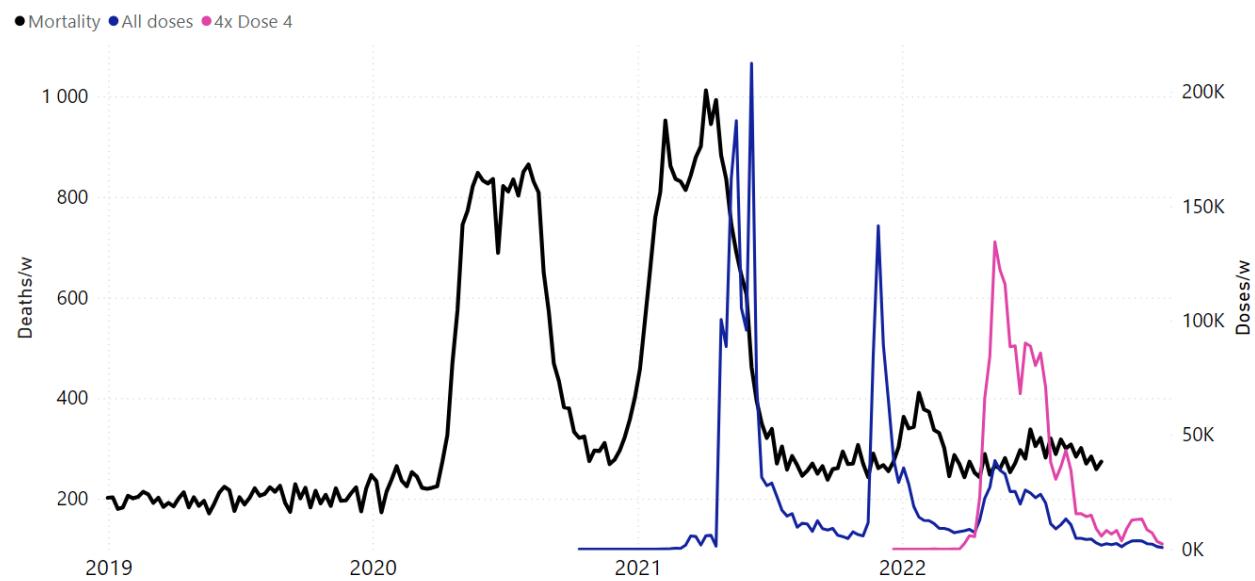
ACM and vaccination by week, Peru, 80-84, 2019-2022



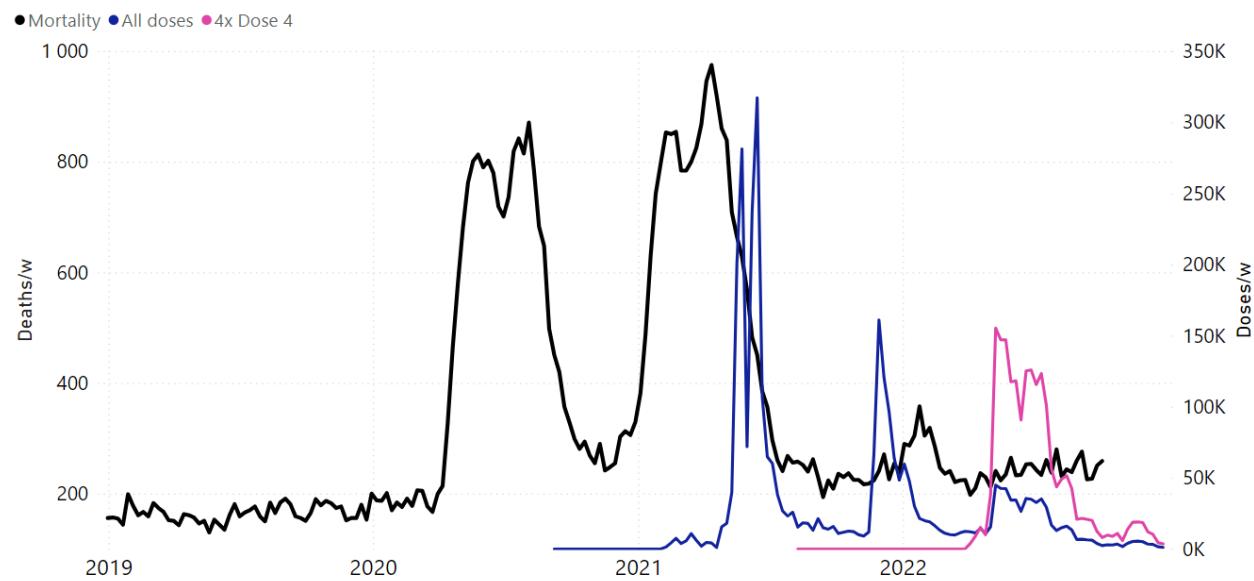
ACM and vaccination by week, Peru, 75-79, 2019-2022



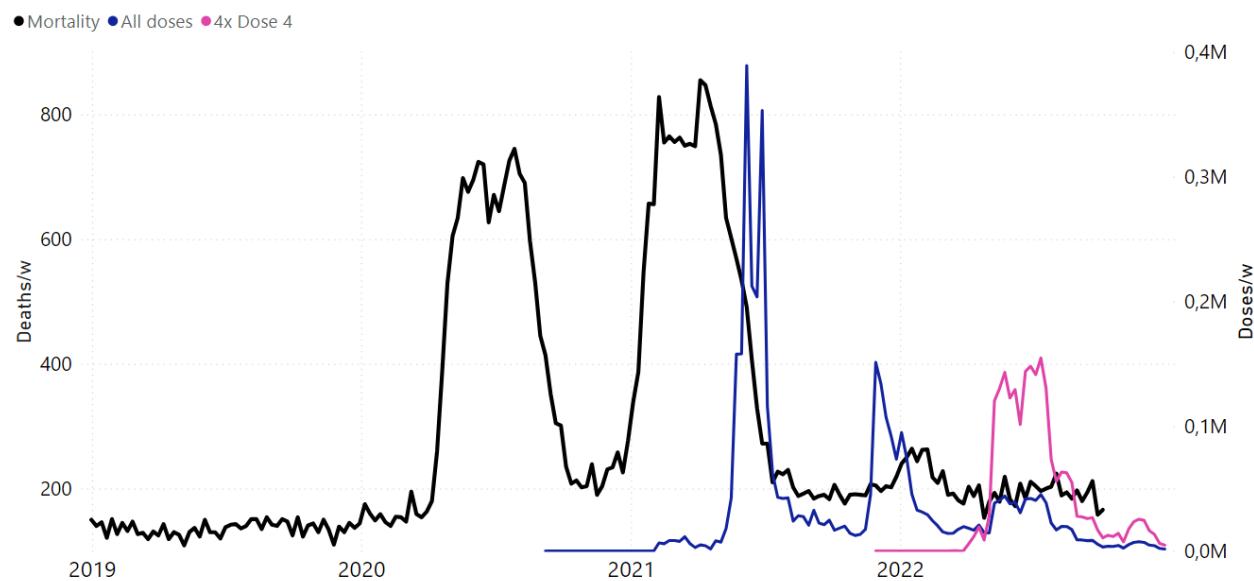
ACM and vaccination by week, Peru, 70-74, 2019-2022



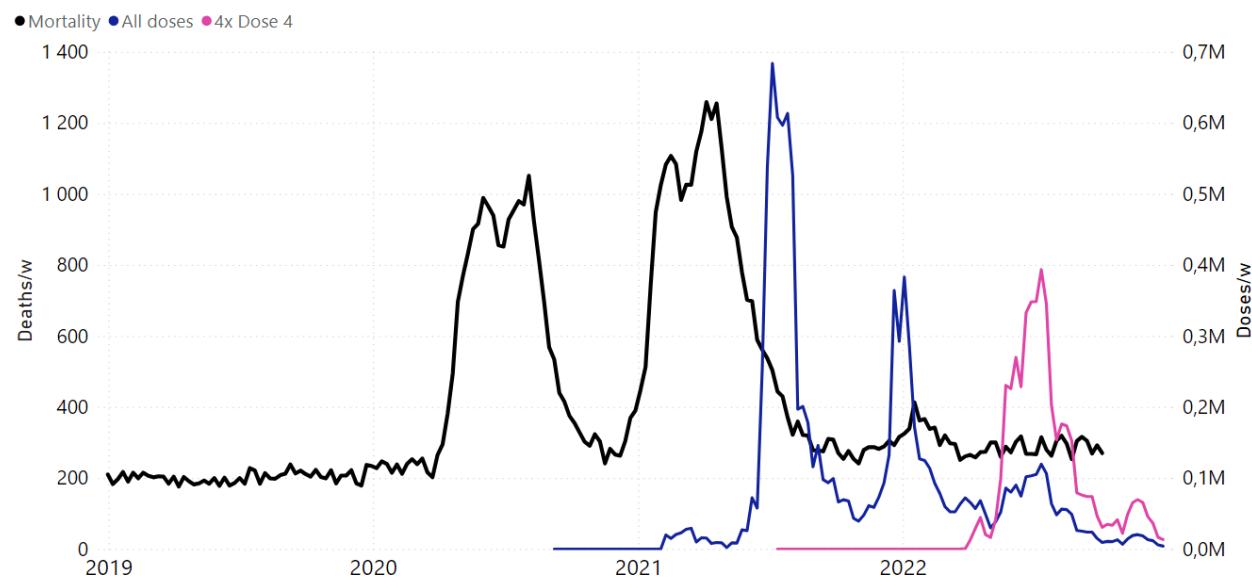
ACM and vaccination by week, Peru, 65-69, 2019-2022



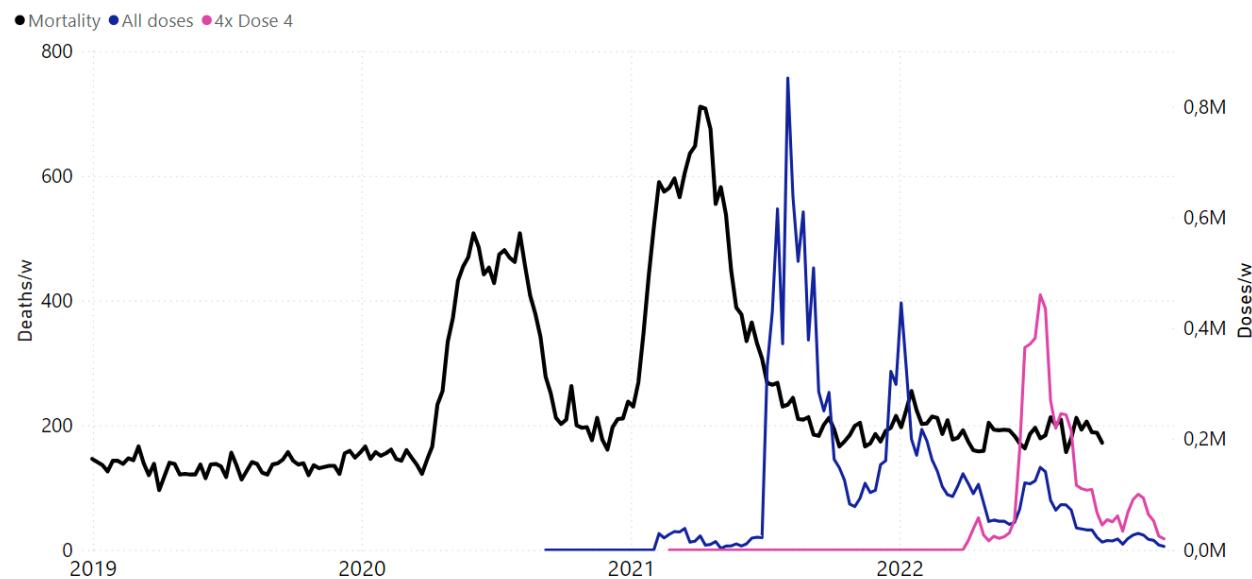
ACM and vaccination by week, Peru, 60-64, 2019-2022



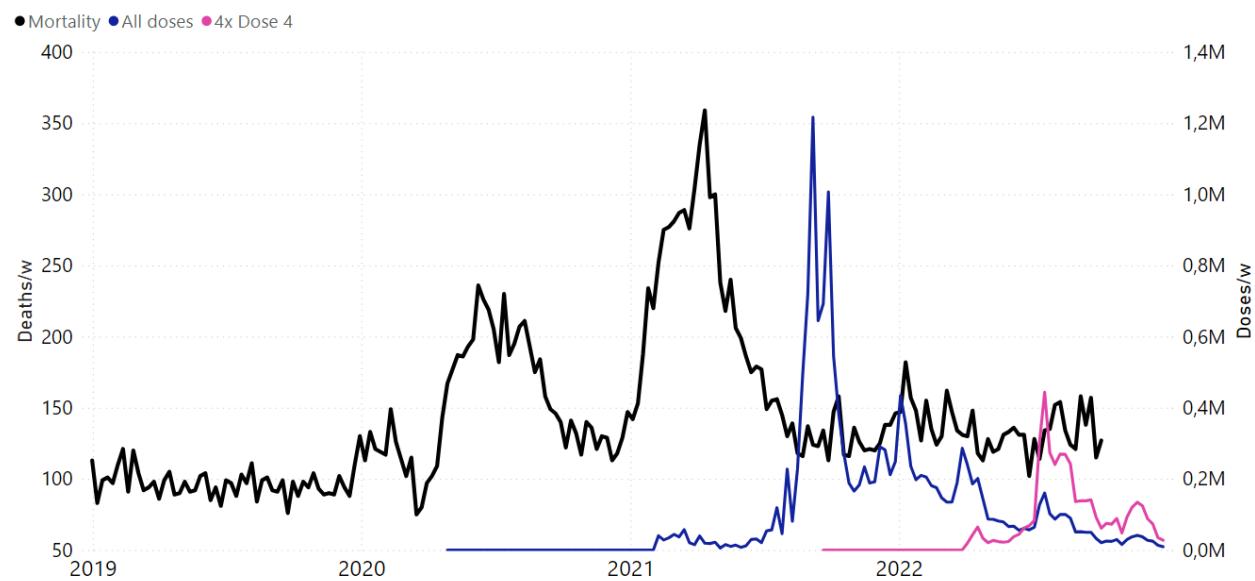
ACM and vaccination by week, Peru, 50-59, 2019-2022



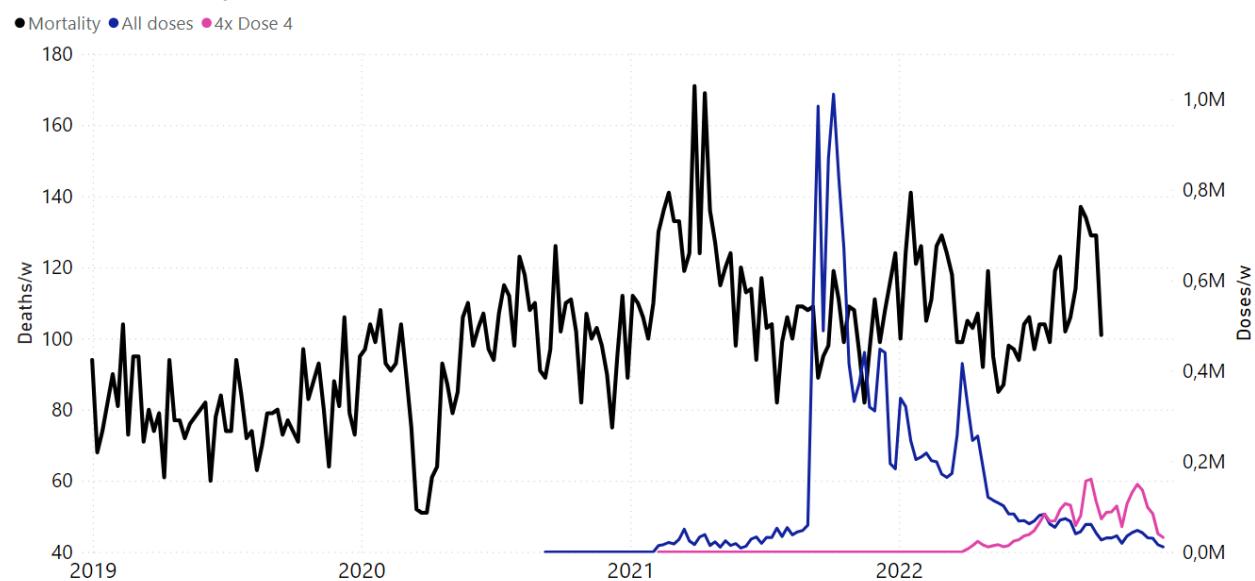
ACM and vaccination by week, Peru, 40-49, 2019-2022



ACM and vaccination by week, Peru, 30-39, 2019-2022



ACM and vaccination by week, Peru, 20-29, 2019-2022



ACM and vaccination by week, Peru, 0-19, 2019-2022

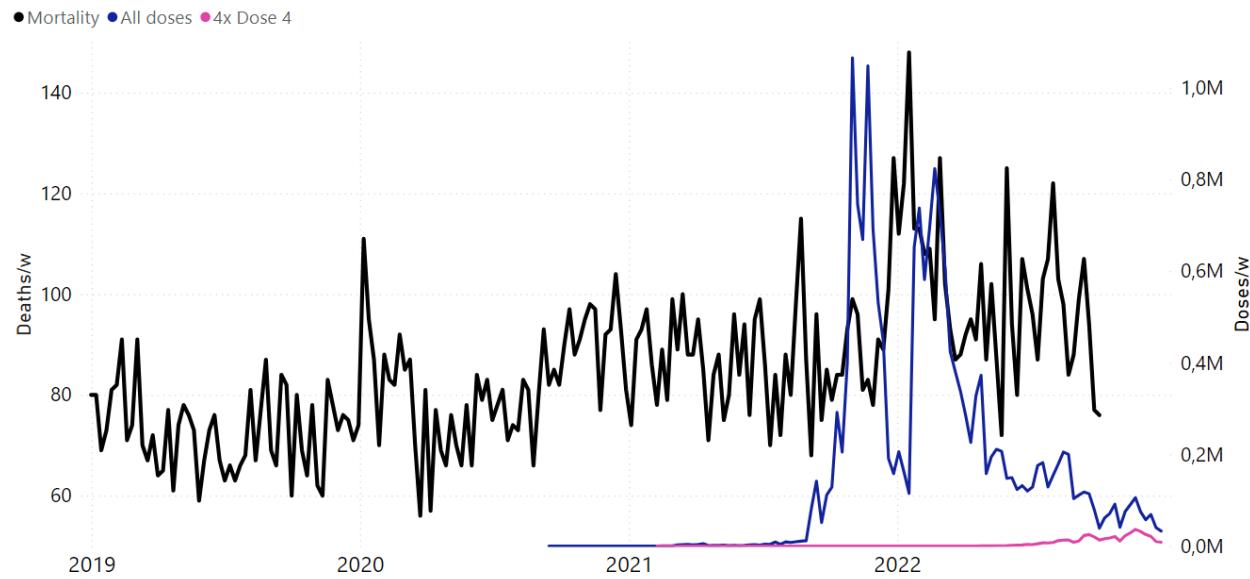
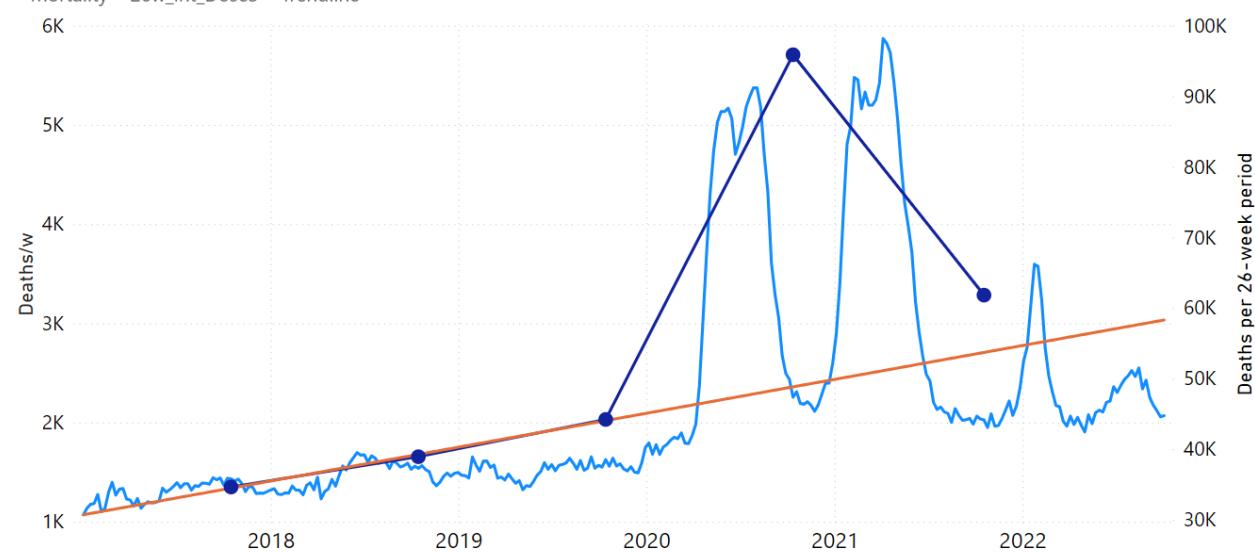


Figure 14: Peru (with rollouts), ACM by week (black), 2019-2022, by age group (60+, 90+, 85-89 ... 60-64, 50-59 ... 20-29, 0-19 years), with vaccine rollouts all-doses (blue) and dose-4 (x4 amplified, pink). The data sources are specified in Appendix A.

Figure 15 shows the national ACM by week for Peru (light blue), from 2017 to 2022, and the 26-week integration points (dark blue) and trend line (orange) that were used to quantify excess ACM in the January-February 2022 peak of interest (“Historical-trend baseline for a period (or peak) of mortality” method, Method 1, see Methods), associated with Peru’s dose-3 rollout, for the 60+ year age group, and for each of the more specific age groups analyzed (90+, 85-89, 80-84, 75-79, 70-74, 65-69, and 60-64 years). The last point (positioned in late 2021) is for the dose-3 vaccination period itself, and each integration point is positioned in time at the start of the integration window. See Methods for details.

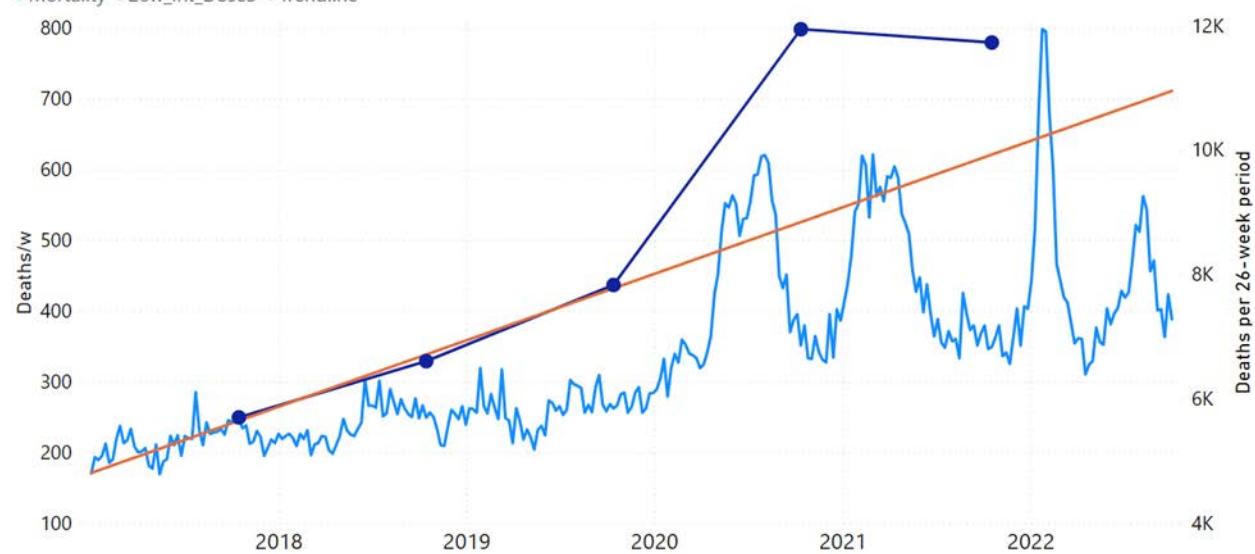
ACM/w, Peru, 60+, 2017-2022

● Mortality ● 26w_Int_Dose3 ● Trendline

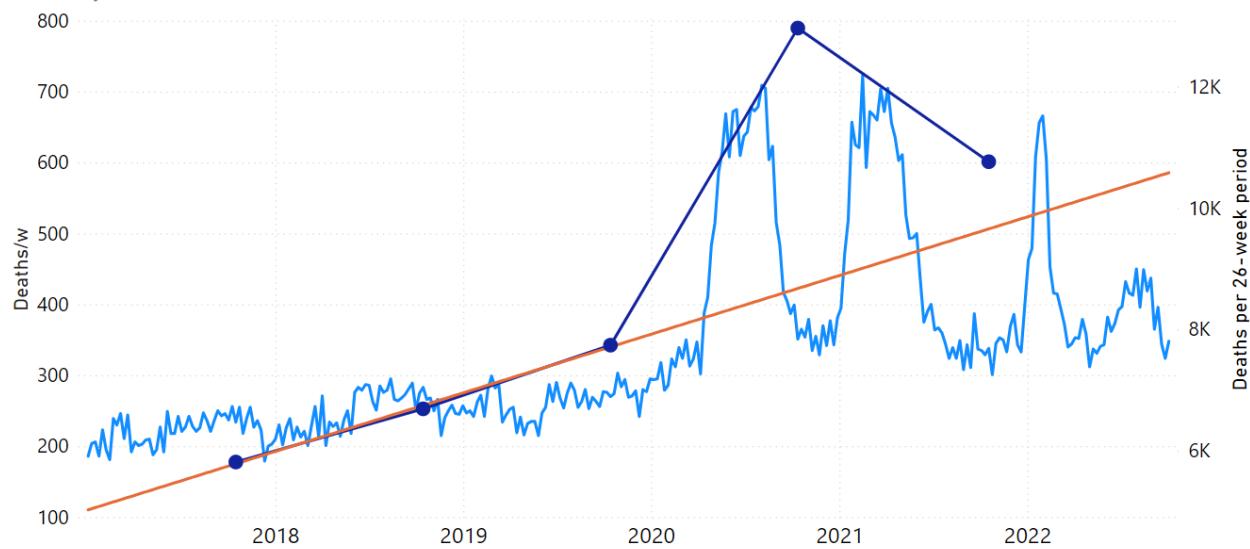


ACM/w, Peru, 90+, 2017-2022

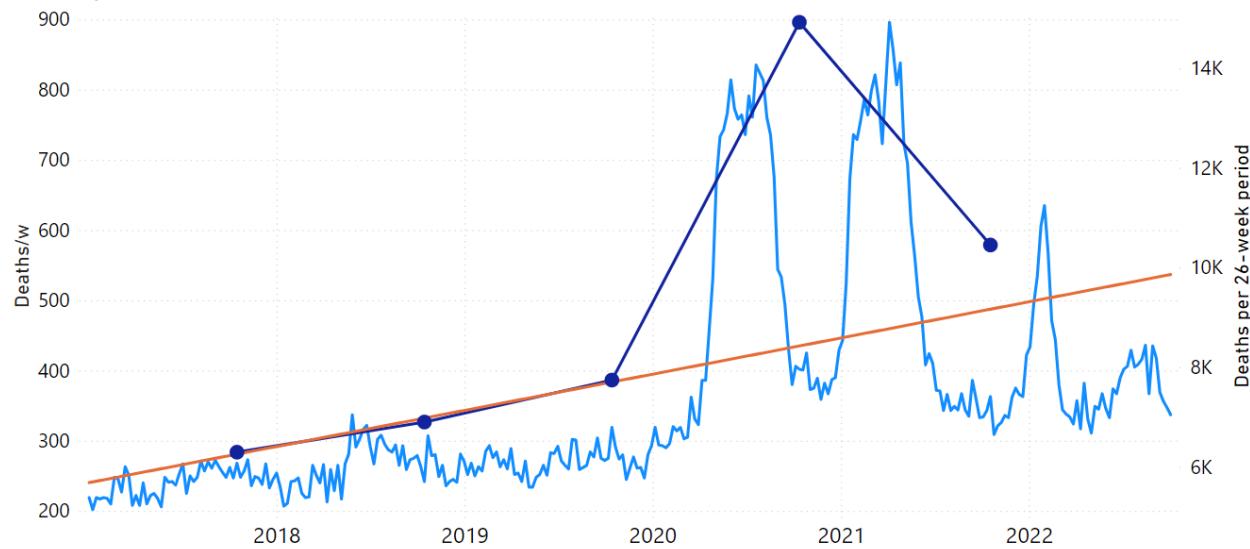
● Mortality ● 26w_Int_Dose3 ● Trendline



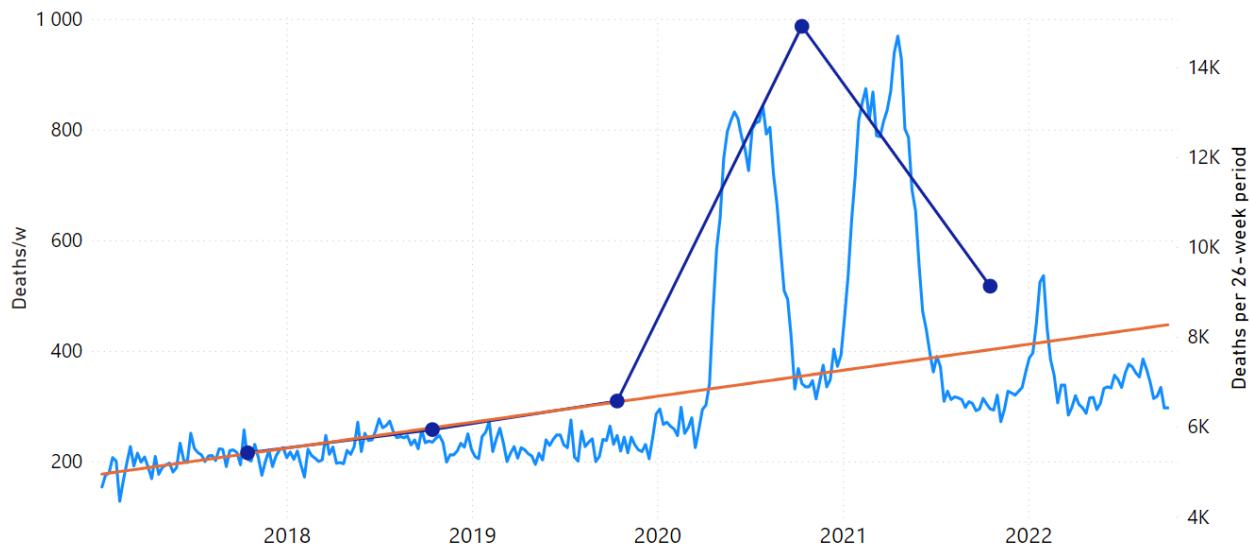
ACM/w, Peru, 85-89, 2017-2022
 ● Mortality ● 26w_Int_Dose3 ● Trendline



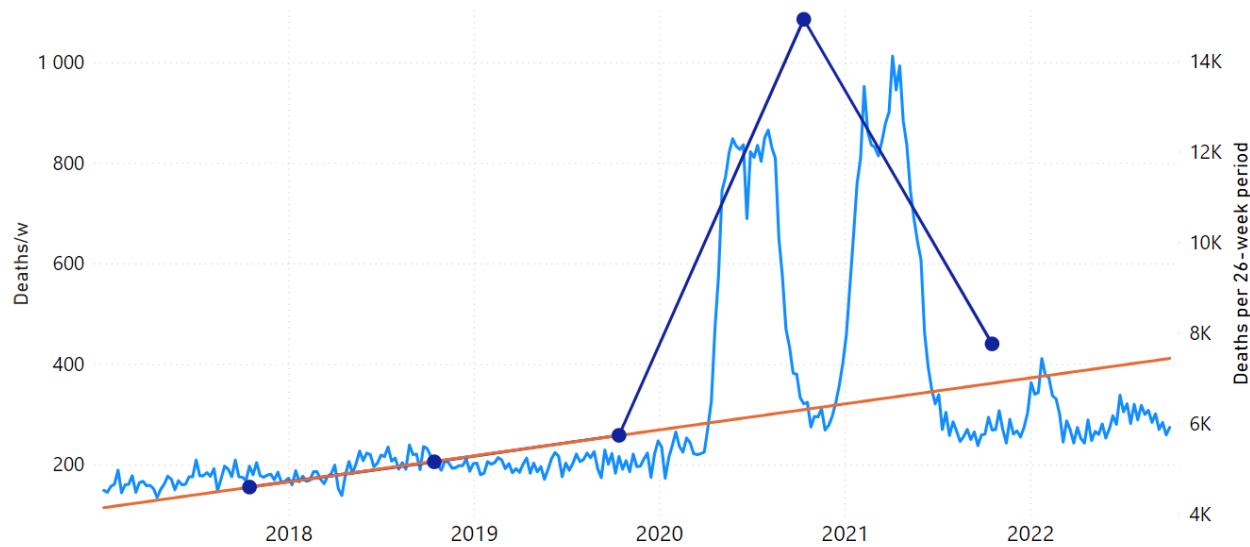
ACM/w, Peru, 80-84, 2017-2022
 ● Mortality ● 26w_Int_Dose3 ● Trendline



ACM/w, Peru, 75-79, 2017-2022
 ● Mortality ● 26w_Int_Dose3 ● Trendline



ACM/w, Peru, 70-74, 2017-2022
 ● Mortality ● 26w_Int_Dose3 ● Trendline



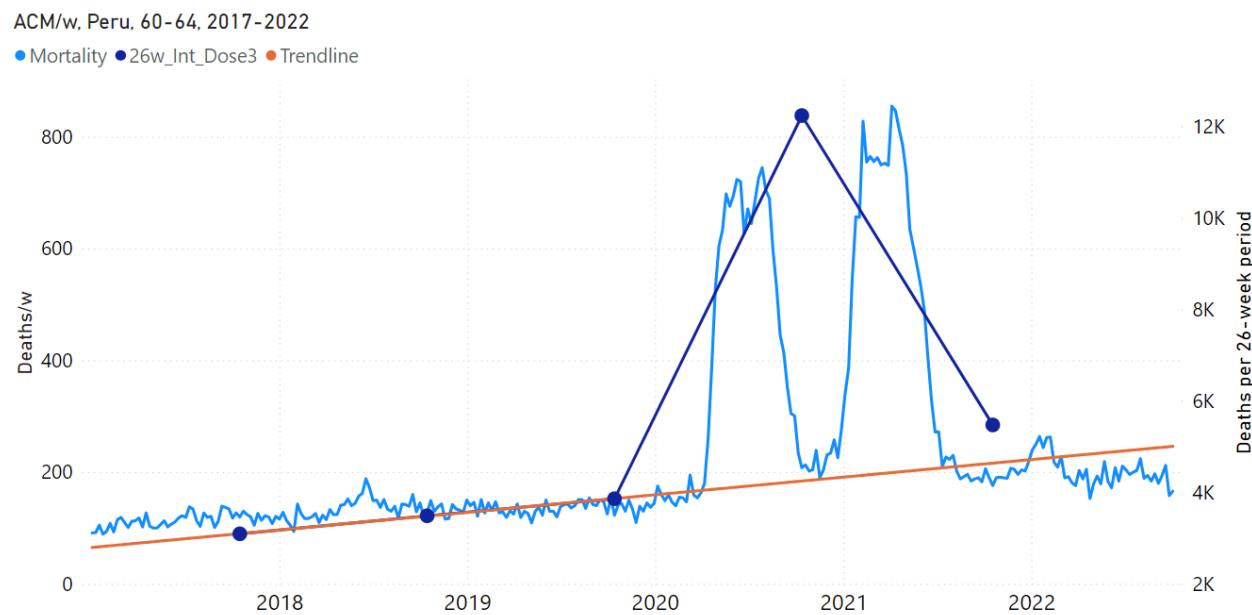
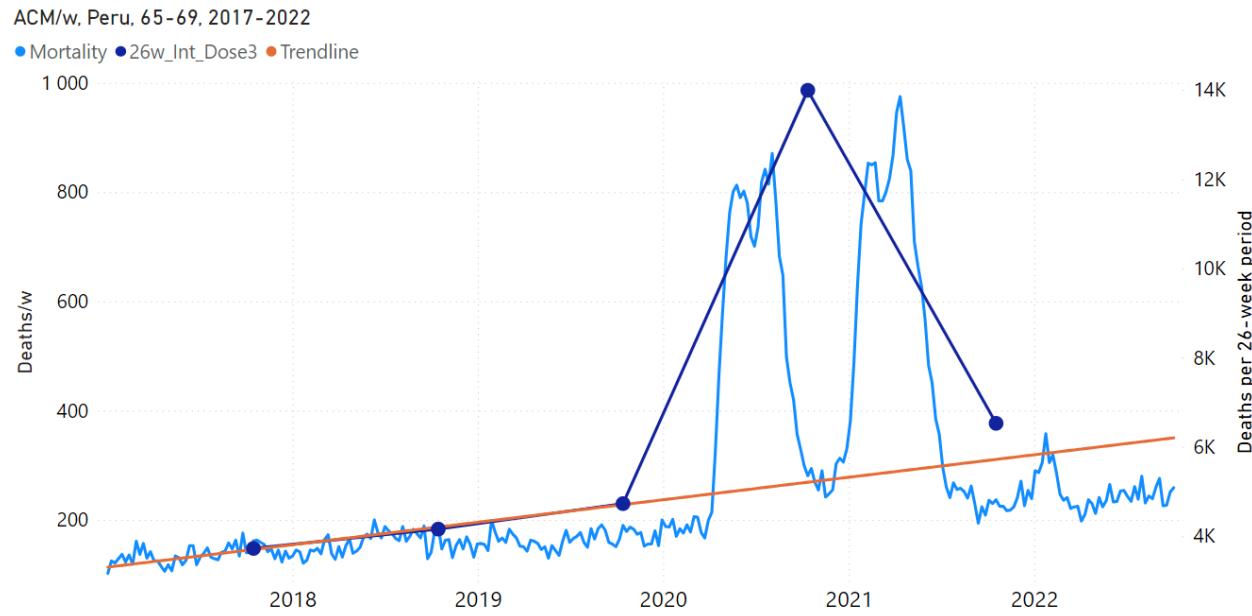
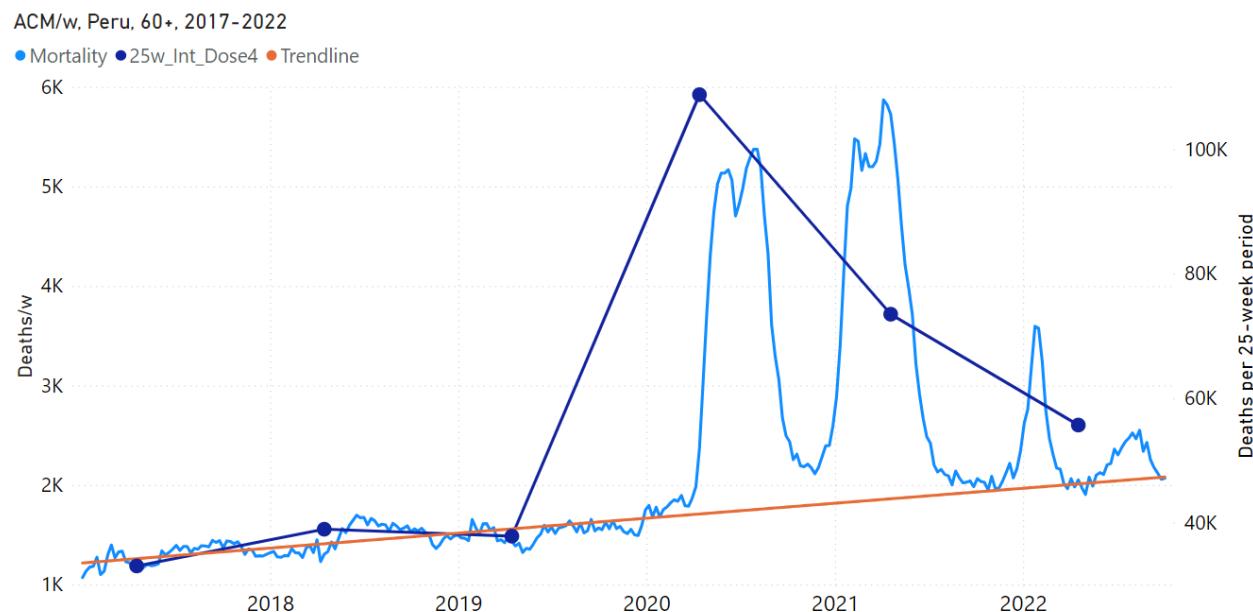


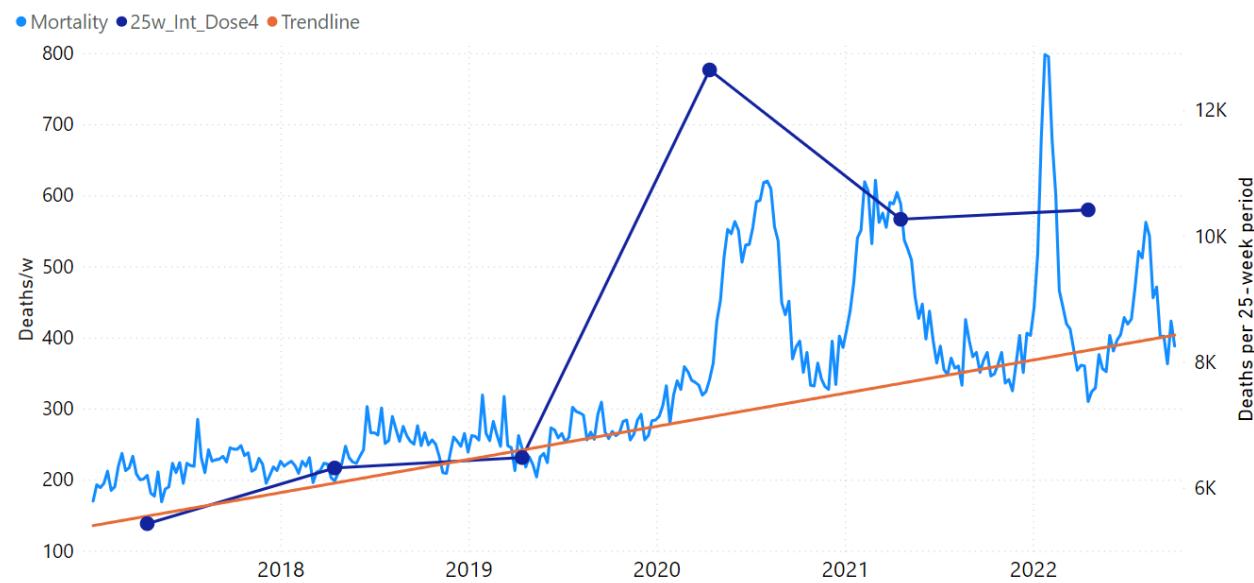
Figure 15: Peru (dose 3), ACM by week (light blue), 2017-2022, by age group (60+, 90+, 85-89 ... 60-64 years), with integrations, 26-week dose-3-vaccination-period (dark blue, points) and trend line (orange). The data sources are specified in Appendix A.

5.7 Peru by age group dose 4 — National ACM and vaccine dose 4 rollouts (90+, 85-89, 80-84, 75-79, 70-74, 65-69, and 60-64 years)

Similarly to Figure 15, Figure 16 shows the national ACM by week for Peru (light blue), from 2017 to 2022, and the 25-week integration points (dark blue) and trend line (orange) that were used to quantify excess ACM in a peak (nominally, July-August 2022) (“Historical-trend baseline for a period (or peak) of mortality” method, Method 1, see Methods) associated with Peru’s dose-4 rollout, for the 60+ year age group, and for each of the more specific age groups analyzed (90+, 85-89, 80-84, 75-79, 70-74, 65-69, and 60-64 years). The last point (positioned in early 2022) is for the dose-4 vaccination period itself, and each integration point is positioned in time at the start of the integration window. See Methods for details.



ACM/w, Peru, 90+, 2017-2022

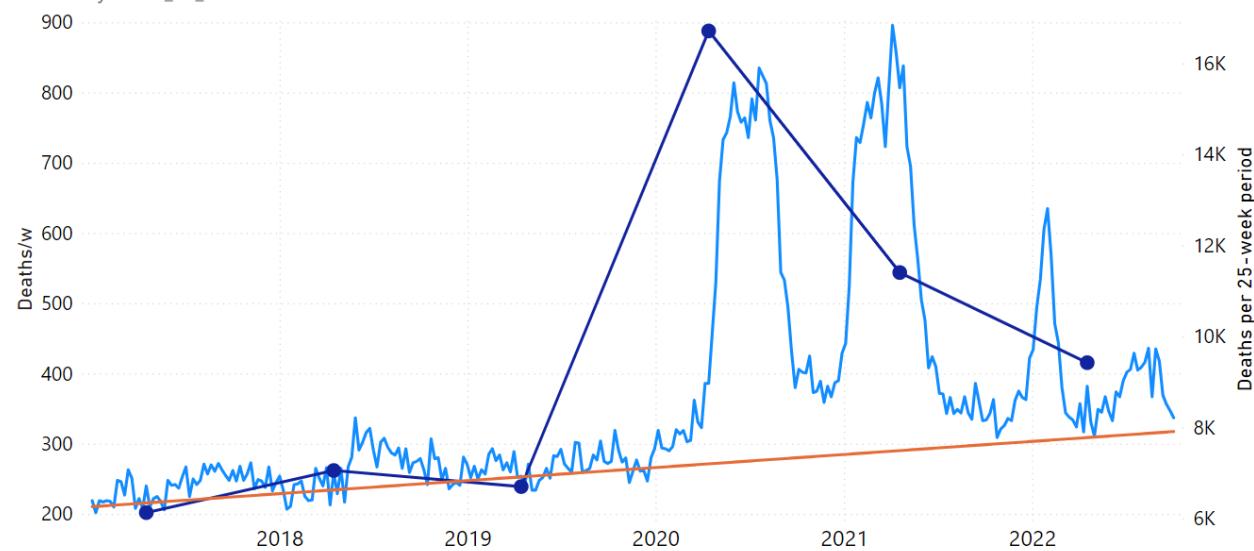


ACM/w, Peru, 85-89, 2017-2022



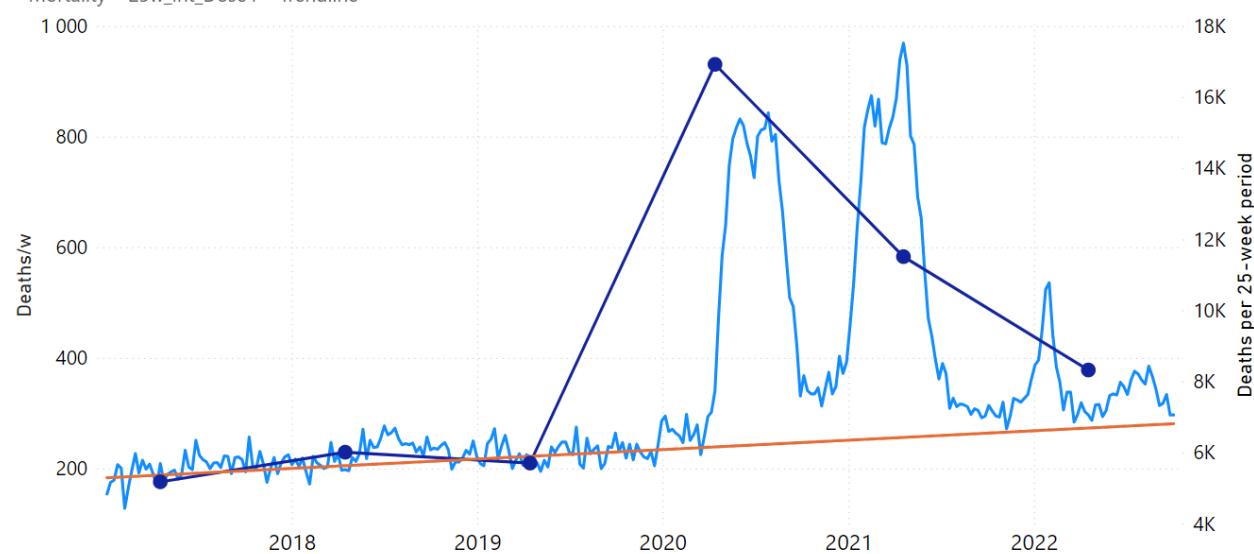
ACM/w, Peru, 80-84, 2017-2022

● Mortality ● 25w_Int_Dose4 ● Trendline

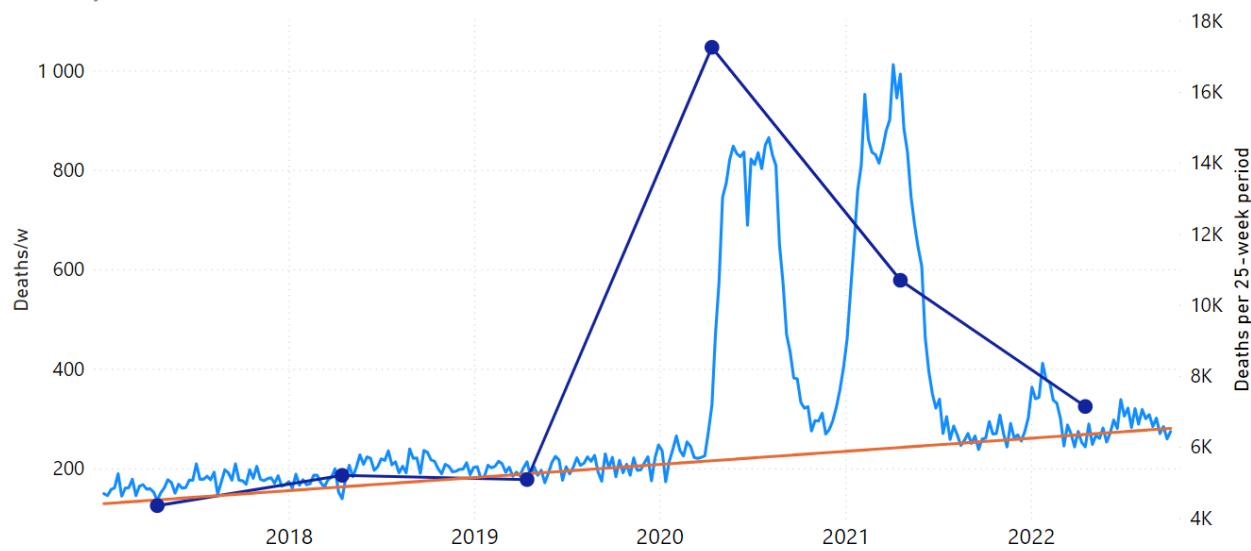


ACM/w, Peru, 75-79, 2017-2022

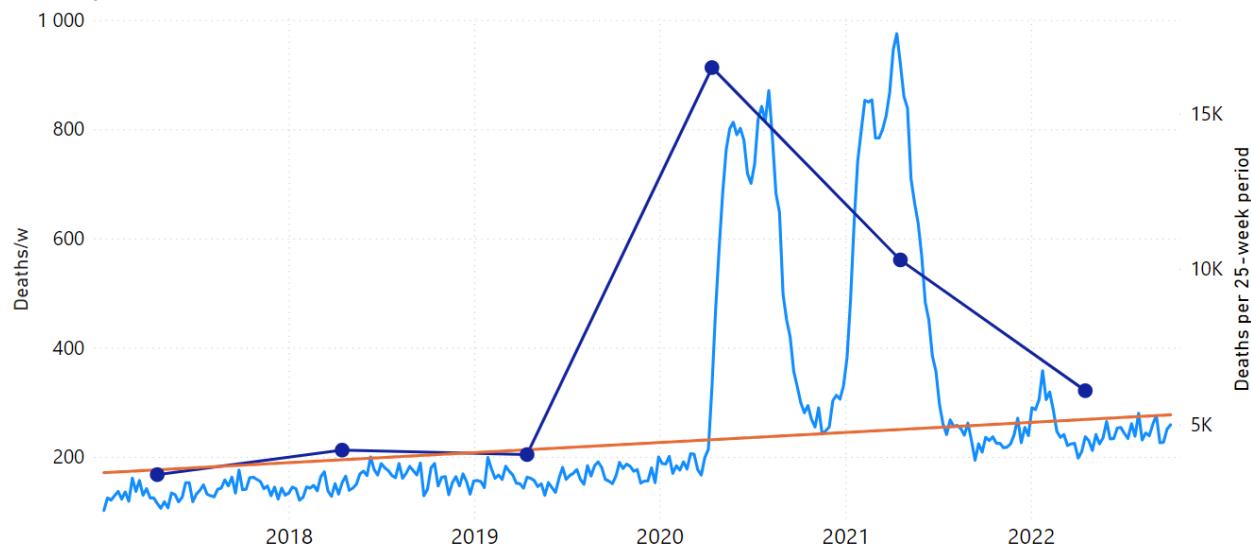
● Mortality ● 25w_Int_Dose4 ● Trendline



ACM/w, Peru, 70-74, 2017-2022
 ● Mortality ● 25w_Int_Dose4 ● Trendline



ACM/w, Peru, 65-69, 2017-2022
 ● Mortality ● 25w_Int_Dose4 ● Trendline



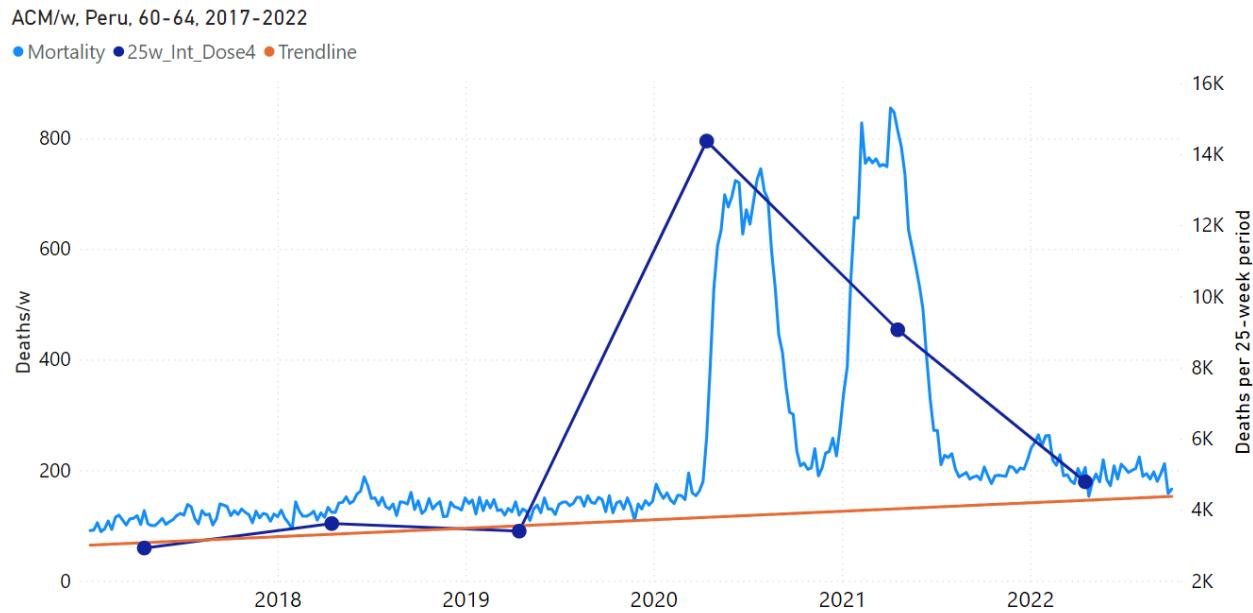
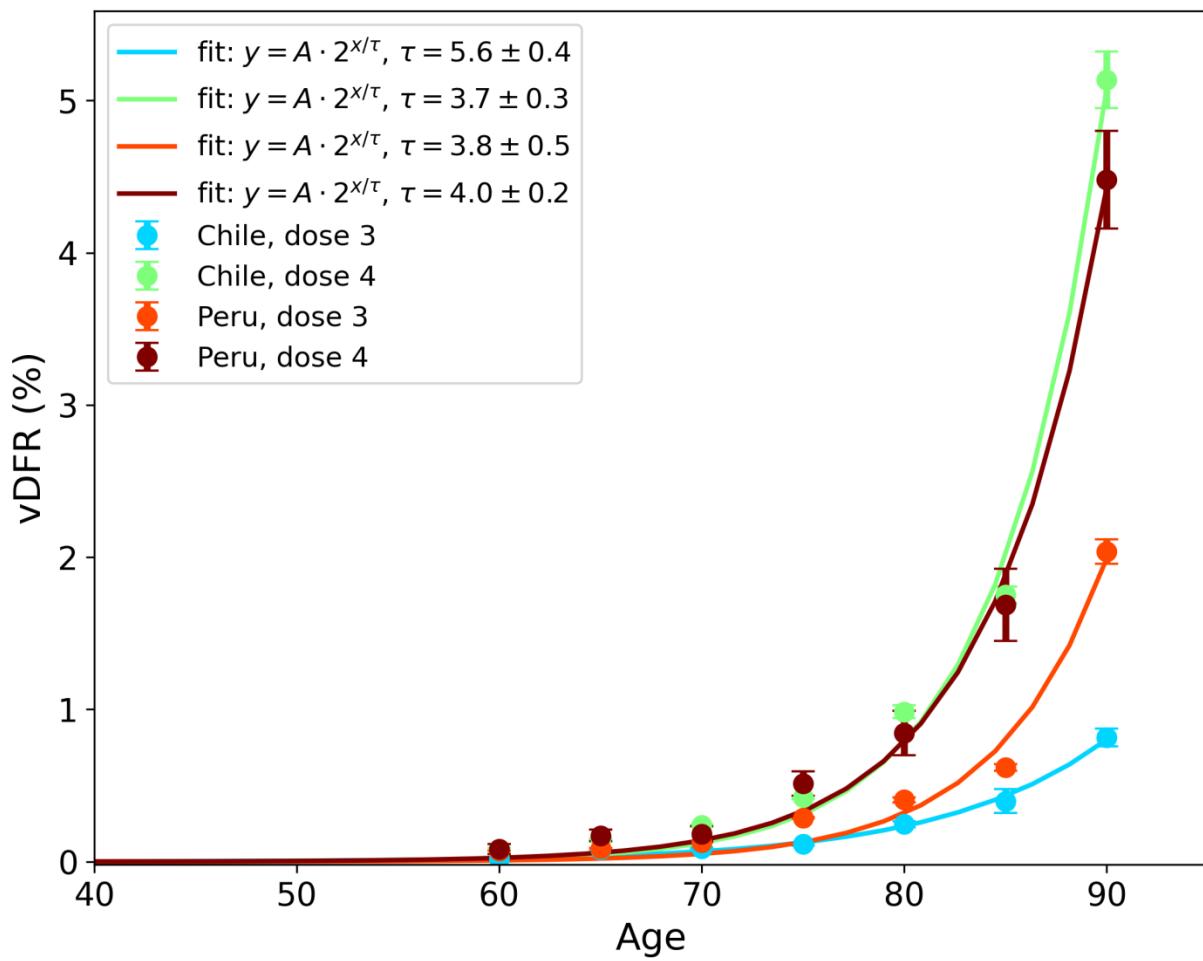


Figure 16: Peru (dose 4), ACM by week (light blue), 2017-2022, by age group (60+, 90+, 85-89 ... 60-64 years), with integrations, 25-week dose-4-vaccination-period (dark blue, points) and trend line (orange). The data sources are specified in Appendix A.

5.8 vDFR by age — Age-stratified vaccine dose fatality rates for doses 3 and 4 in Chile and Peru

Figure 17 (top panel) shows national vDFR (linear scale) versus age (“Historical-trend baseline for a period (or peak) of mortality”, Method 1, see Methods), for Chile and for Peru, for the January-February 2022 peak of interest (Chile dose 4; Peru dose 3), for the peak (nominally, July-August 2022) associated with Peru’s dose-4 rollout, and for the peak (nominally, July-August 2021) associated with Chile’s dose 3 rollout, with exponential fits. Figure 17 (bottom panel) shows the same data using a logarithmic scale for vDFR. Each point is positioned in time at the starting age of the age group.

The exponential fits were performed without using the Y-specific errors (i.e., errors in vDFR) shown. Therefore, the resulting error on the doubling-time fit parameter T2 (τ , *tau*, in the figure legends) is based on the unweighted residual Y-values.



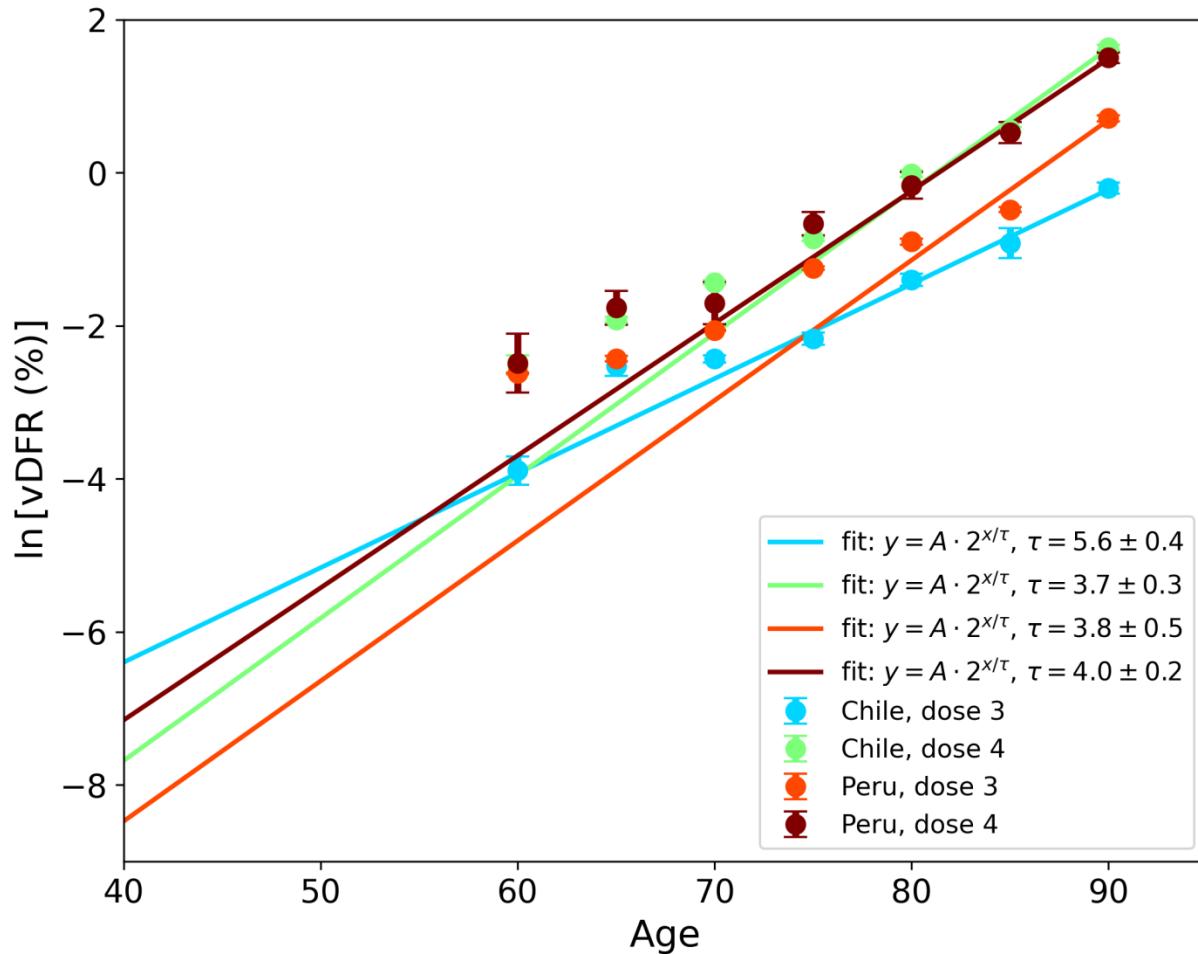


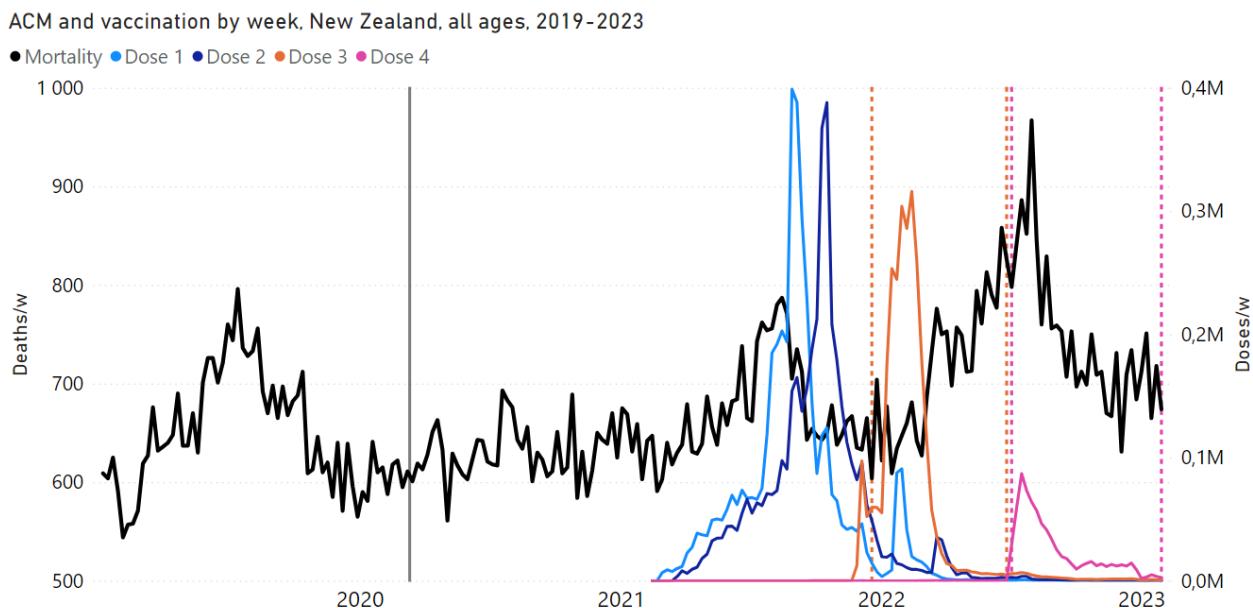
Figure 17: vDFR versus age, for Chile and for Peru, for the January–February 2022 peak (Chile dose 4, Peru dose 3), for the July–August 2022 peak (Peru, dose 4), and for the July–August 2021 peak (Chile, dose 3), with exponential fits. Linear scale (top), log scale (bottom).

In Figure 17, the age (X-axis value, in years) assigned to a given age group is the starting age of the window of ages for the age group. This particular choice makes little difference because translating the x values by any constant number, for example, does not affect the doubling time obtained by fitting an exponential function, and only slightly affects the y intercept at $x = 0$ (the prefactor in the exponential).

5.9 New Zealand all-ages doses 3 and 4 — National ACM and vaccine dose rollouts

New Zealand is the case, among our 15 countries having sufficient ACM data, in which a nominally January-February 2022 peak in ACM is the least prominent. Nonetheless, the peak is measurably present for New Zealand, especially in the 80+ years age group.

Figure 18 shows the all-ages national ACM by week for New Zealand (black), from 2019 to 2023, compared to the COVID-19 vaccine-dose rollouts. Vertical lines indicate the 11 March 2020 World Health Organization (WHO) announcement of a pandemic (grey), and the bounds of the integration windows for the regions of interest corresponding to the rollouts of doses 3 and 4 (orange and pink, respectively). The bottom panel shows the ACM data for the 80+ years age group, and the same vertical lines, as references.



ACM and vaccination periods by week, New Zealand, 80+, 2019-2023

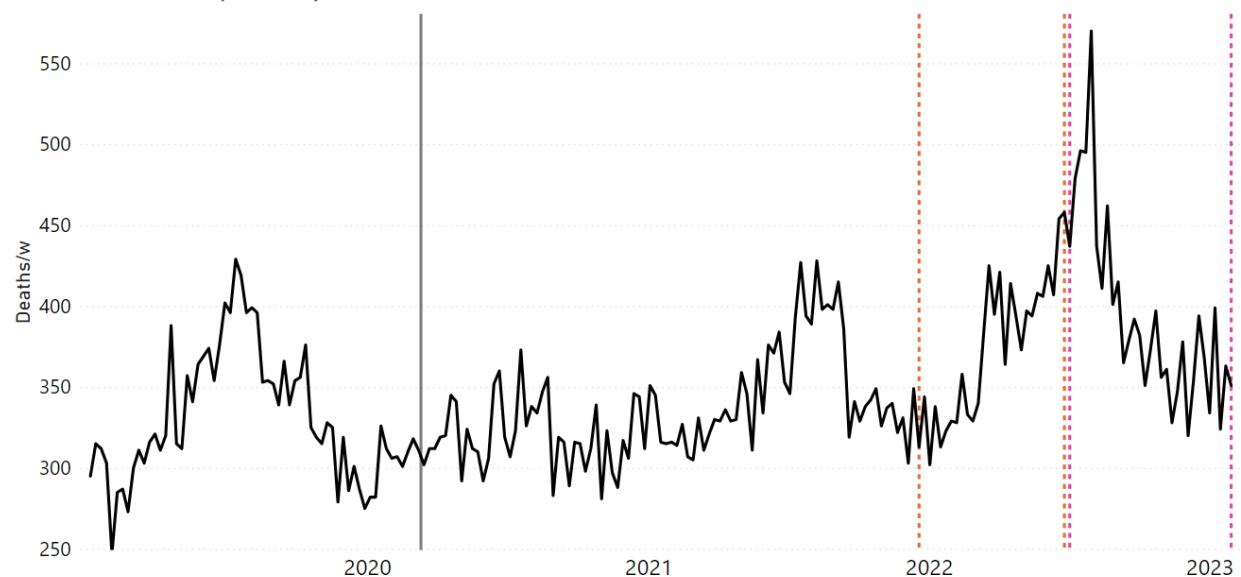


Figure 18: (top) New Zealand (with rollouts), all-ages ACM by week (black), 2019-2023, with all-ages vaccine rollouts dose 1 to dose 4 (colours). (bottom) ACM by week for 80+ years age group (black), and same integrations bounds (dashed vertical lines). The 11 March 2020 pandemic declaration date is shown by a grey vertical line. The data sources are specified in Appendix A.

Figure 19 has the all-ages national ACM by week for New Zealand (light blue), from 2011 to 2023, and the 28-week-integration points that were used to quantify excess ACM in the January-February 2022 peak of interest (“Historical-trend baseline for a period (or peak) of mortality” method, Method 1, see Methods) (dark blue points and orange trend line), associated with New Zealand’s dose-3 rollout (28-week period, orange dashed vertical lines bounds shown in Figure 18).

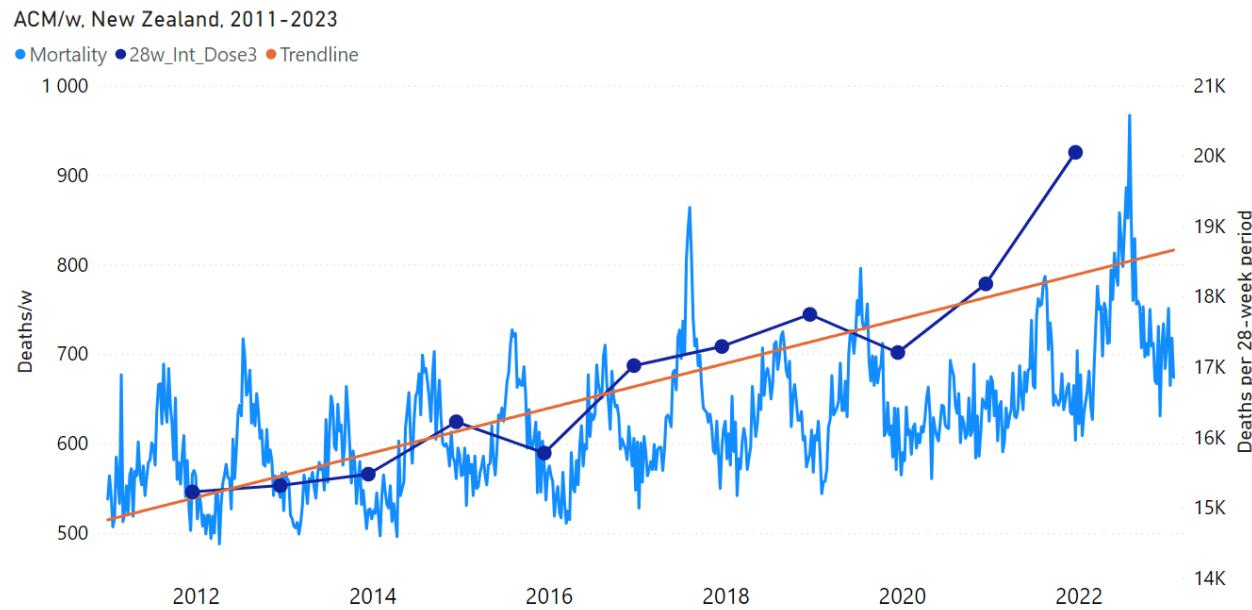


Figure 19: New Zealand (dose 3), all-ages ACM by week (light blue), 2011-2023, with integrations, 28-week dose-3-vaccination-period (dark blue, points) and trend line (orange). The data sources are specified in Appendix A.

Figure 20 shows the all-ages national ACM by week for New Zealand (light blue), from 2011 to 2023, and the 31-week-integration points that were used to quantify excess ACM in a July-August 2022 peak (“Historical-trend baseline for a period (or peak) of mortality” method, Method 1, see Methods) (dark blue points and orange trend line), associated with New Zealand’s dose-4 rollout (31-week period, pink dashed vertical lines bounds shown in Figure 18).

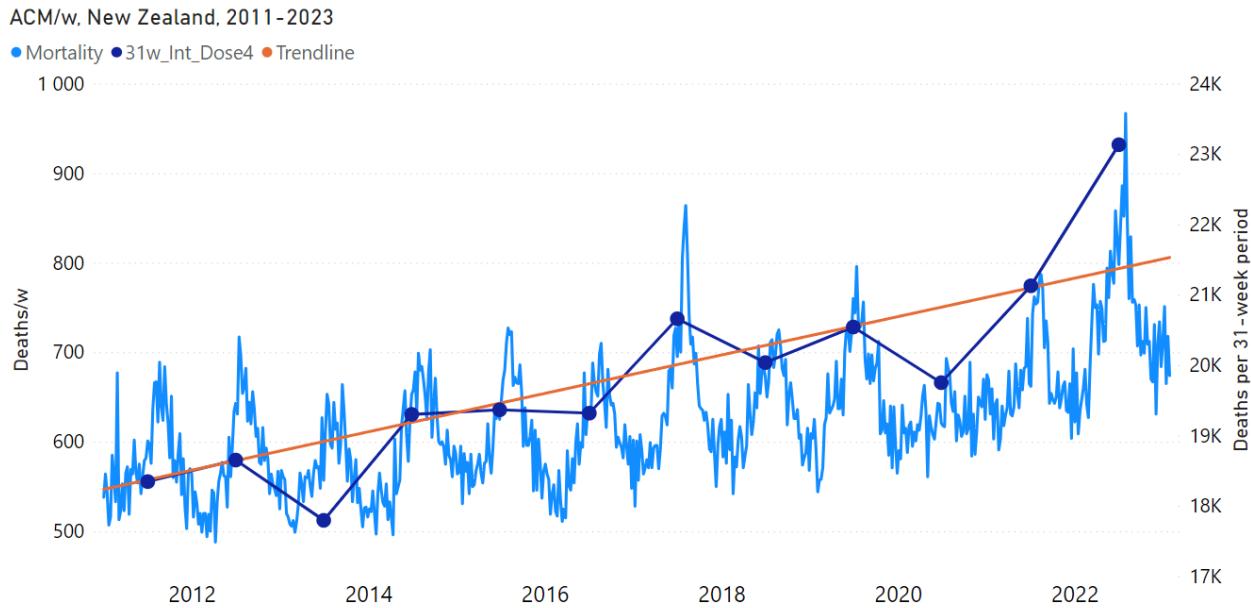


Figure 20: New Zealand (dose 4), ACM by week (light blue), 2011-2023, with integrations, 31-week dose-4-vaccination-period (dark blue, points) and trend line (orange). The data sources are specified in Appendix A.

The corresponding dose-specific all-ages vDFR values for New Zealand are: (0.05 ± 0.01) % (nominally January-February 2022, dose 3) and (0.21 ± 0.03) % (nominally July-August 2022, dose 4).

6. Discussion

6.1 COVID-19 vaccines can cause death

It is important to discern the question of whether a COVID-19 vaccine injection can cause the death of the patient and the question of whether excess ACM (population level, by definition) is causally associated with the COVID-19 vaccine rollouts.

Even if there is clinical and pathological proof that injections can cause the deaths of individual subjects, this does not demonstrate a causal relation between a rapid vaccine rollout and a temporally associated peak in excess ACM. With necessarily limited numbers of documented cases of individual deaths, it only proves that the said causal relation is possible. A formal consideration of causality in excess mortality must nonetheless be made, which is done in sections below.

In this context, it is well established that COVID-19 vaccine injections have caused and are likely to cause the deaths of individuals, as shown by:

- many detailed autopsy studies (e.g., Choi et al., 2021; Edler et al., 2021; Schneider et al., 2021; Sessa et al., 2021; Gill et al., 2022; Mörz, 2022; Murata et al., 2022; Suzuki et al., 2022; Takahashi et al., 2022; Tan et al., 2022; Yeo et al., 2022; Yoshimura et al., 2022; Chaganti et al., 2023; de Boer, Crawford, Parsons, 2023; Esposito et al., 2023; Hulscher et al., 2023; Jeon et al., 2023; Manu, 2023; Nushida et al., 2023; Onishi et al., 2023; Schwab et al., 2023),
- adverse effect monitoring (Rose and McCullough, 2021; Hickey and Rancourt, 2022),
- studies of vaccine-induced pathologies (e.g., Goldman et al., 2021; Kuvandik et al., 2021; Turni and Lefringhausen, 2022; Edmonds et al., 2023; Wong et al., 2023),
- an established causal link to vaccine-induced pathology, by histopathology and immunohistochemical staining of skin biopsy specimens (Sano et al., 2023),
- secondary analysis of serious adverse events reported in placebo-controlled, industry phase III randomized clinical trials (e.g., Fraiman et al., 2022),
- more than 1,250 peer-reviewed publications about COVID-19 vaccine adverse effects (React 19, 2022), and
- the known vaccine injury compensation programs of states worldwide, which include death resulting from the COVID-19 vaccines (Mungwira et al. 2020; Wood et al., 2020; Crum et al., 2021; Kamin-Friedman and Davidovitch, 2021), where Japan, Canada and the UK have granted compensation for COVID-19

vaccine induced deaths (*The Japan Times*, 26 July 2022; Corbett, 6 September 2022; Wise, 2022).

These data and circumstances support but do not in themselves prove the conclusion of a likely causal link, which is made in several population-level studies, including:

- a recent survey study (Skidmore, 2023),
- our prior quantitative evaluations of vaccine dose fatality rate (vDFR) from all-cause mortality (ACM) data in several countries (Rancourt, 2022; Rancourt et al., 2022a, 2022b, 2023), and
- the present study, which will conclude that causality in excess mortality is demonstrated.

Finally, there are detailed explanations based on principles of immunology as to the cellular, molecular and systems mechanisms for severe harm from COVID-19 vaccines, and as to why, for example, repeated doses can engender non-linearly-more-severe reactions (Palmer et al., 2023).

6.2 Absence of excess mortality until the COVID-19 vaccines are rolled out

Previously, we have demonstrated that in several countries and states there is no detectable or virtually no excess ACM during the declared pandemic, compared to the recent historic record, until the COVID-19 vaccines are rolled out, followed by large excess mortality during and after vaccine rollout:

1. **India** (Rancourt, 2022)
2. **Australia, and each of its eight states** (Rancourt et al., 2022a)
3. **Israel** (relatively less excess mortality) (Rancourt et al., 2023)

4. Canada (relatively less excess mortality) (Rancourt et al., 2021b, 2022c; Rancourt, 2023)

Likewise, in the present paper, we show that 9 of the 17 countries studied have no detectable excess ACM until the vaccines are rolled out: Australia, Malaysia, New Zealand, Paraguay, Philippines, Singapore, Suriname, Thailand, and Uruguay.

This is conclusively shown in Figure 2, and corresponding quantifications are given in Table 1 and Figure 3.

This means that, in these 9 countries (Australia, Malaysia, New Zealand, Paraguay, Philippines, Singapore, Suriname, Thailand, Uruguay), for approximately one year after the 11 March 2020 WHO declaration of a pandemic, there were no net extra deaths that could be attributed to a pandemic or to pandemic-response medical or government measures. These countries had no extra deaths until they were subjected to rapid mass COVID-19 vaccine administration (Figure 2).

This is part of the strong evidence for a causal link, described below.

6.3 The COVID-19 vaccines did not save lives and appear to be lethal toxic agents

Since excess mortality occurred solely after initiating (and during) vaccine rollouts in 9 countries (Australia, Malaysia, New Zealand, Paraguay, Philippines, Singapore, Suriname, Thailand, Uruguay; Figure 2), the vaccines did not reduce serious illness (as claimed by manufacturers) enough to reduce any risk of death.

In the 17 countries of the present study, there is no evidence in ACM by time data of any beneficial effect of COVID-19 vaccines. There is no association in time between COVID-19 vaccination and any proportionate reduction in ACM.

On the contrary, vaccine administration exhibits known epidemiological characteristics of sudden population-wide exposure to a lethal toxic substance, in all of the 17 countries, and the age groups studied:

- i. A transition to a regime of higher ACM on initiation of prolonged challenge with the lethal toxic agent (Figure 2, Figure 6, Figure 7)
- ii. Specific peaks of ACM temporally associated with (synchronous with or immediately following) all sharp peaks in mass challenge with (population-wide administration of) a lethal toxic agent (booster rollouts) (Figure 5, Figure 4, Figure 12, Figure 11, Figure 13, Figure 15, Figure 16, Figure 14, Figure 18, Figure 19, Figure 20)
- iii. Proportionality between the number of population-wide individual assaults with the presumed lethal toxic agent (individual injections) and the temporally associated excess mortality (Figure 8, Figure 9)
- iv. An exponentially or near-exponentially increasing lethal toxicity (vDFR) with age of the individuals subjected to the toxic agent (Figure 17)
- v. Consistent values of the calculated lethal toxicity of the toxic agent (values of vDFR) across many populations and for different episodes of sudden rollouts (all doses, and booster rollouts) (Table 2, Figure 8, Figure 9, Figure 10)
- vi. Apparent dependence of fatal toxicity on the toxicological nature of the administered agent (vDFR dependence on vaccine type/manufacturer)

Regarding point-iv, increasing and exponentially increasing lethal toxicity with age is known from animal and human studies of poisoning and overdoses (e.g., Wiberg et al., 1970; Rogers and Heard, 2007; Chen et al., 2009; Phua et al., 2009; Shively et al., 2017).

Regarding point-v, the values of all-ages vDFR, in the 17 countries studied, are consistent within one order of magnitude (Table 2, Figure 9, Figure 10). In Figure 9, the generally larger values in all-ages vDFR for the vaccination period compared to the

January–February 2022 ACM peak period may be due to such factors as: different populations being preferentially vaccinated, fragile individuals being eliminated prior to the peak period, differences in vaccine type used, booster policy or societal pressure wherein fragile individuals are avoided in the peak period, a larger contribution from non-COVID-19-vaccine all-ages deaths (i.e., from medical and government measures) in the earlier stages of COVID-19 vaccination, and so forth. The USA Vaccine Adverse Events Reporting System (VAERS) data also shows smaller all-ages fatal vaccine toxicity per injection for boosters relative to pre-booster injections (Hickey and Rancourt, 2022; their Table 1). Other factors that could affect the systematic difference between booster and pre-booster doses include: smaller quantities of active ingredients used in boosters compared to pre-booster doses, and changing immune response on being subjected to repeated challenges (Palmer et al., 2023).

Regarding point-vi (in relation to points iv and v), country-to-country differences in values of all-ages vDFR will be dependent on the age structure of the elderly population (> 60 years) and on health status of the elderly. Despite these large confounding effects, there is some evidence of the possible effect of COVID-19 vaccination type or manufacturer:

- South Africa has the largest value of all-ages vDFR (Table 2, Figure 9, Figure 10) and a large fraction of its vaccines were Johnson & Johnson: 24.6 %, compared to 2.1 % for the European Union, which is the largest fraction of use of this vaccine in the database (OWID, 2023c).
- In the USA's VAERS, the Janssen (Johnson & Johnson) vaccine is 4 times more fatally toxic (per injection) than the Pfizer vaccine, for both 18-64 years and 65+ years age groups (Hickey and Rancourt, 2022; their Table 1).
- Unlike the mRNA lipid-nanoparticle vaccines, the Johnson & Johnson vaccine is a viral vector vaccine based on a human adenovirus modified to contain a gene for making the spike protein of SARS-CoV-2. The vaccine has a history of being paused and discontinued.

- Paraguay used the Indian vaccine Covaxin, whereas India had an all-ages vDFR of approximately 1 % (Rancourt, 2022).
- Uruguay had the Chinese vaccine Sinovac as the dominant fraction of its vaccines: 55 %, although only Pfizer boosters were later used (OWID, 2023c).
- Suriname relied largely on the Indian vaccine Covishield; on China's Covilo, and on Moderna's Spikevax.
- Bolivia used Sinovac, Covilo, Covishield, Janssen, Sputnik V, and Pfizer's Comirnaty.

6.4 Strong evidence for a causal association and vaccine lethal toxicity

First: Some of the strongest evidence for a causal association and vaccine lethal toxicity is the remarkable temporal associations between rapid first-doses and booster rollouts and immediate peaks in ACM, in specific elderly age groups, including peaks of mortality occurring at times in seasonal cycles when peaks virtually never occur. This was previously graphically illustrated for Israel (Rancourt et al., 2023; their Appendix 2: Figure A2 F1). In the present paper, there are multiple examples of this type of evidence, for Chile (Figure 12, Figure 11, Figure 13) and for Peru (Figure 15, Figure 16, Figure 14), in every case where age-stratified mortality and age-stratified (and dose-specific) vaccination data are available.

These findings are conclusive. The associations are numerous and systematic, and there are no counter examples. We have found no evidence in our extensive research on ACM that COVID-19 vaccines had any beneficial effect. If vaccines prevented transmission, infection or serious illness, then there should be decreases in mortality following vaccine rollouts, not increases, as in every observed elderly age group subjected to rapid booster rollouts. And, mortality would not increase solely when vaccines are rolled out, where no excess mortality occurs prior to vaccine rollouts, as we have documented here, in 9 countries across 3 continents.

Second: Another strong line of evidence for a causal association and vaccine lethal toxicity is present in data that is not stratified by age. Such mortality and vaccination data is partially confounded by the lack of age stratification. This is mostly because of spread and overlap of vaccine administration to different age groups. Nonetheless, there are several clear examples in non-age-stratified data of temporal associations between peaks in all-ages ACM and peaks in all-ages vaccination rollouts. This was previously demonstrated for Australia and each of 5 of its 8 states (Rancourt et al., 2022a), regarding the January-February 2022 features studied again here (Figure 5, Figure 4, Figure 8, Figure 9, Figure 10).

In the present paper, every country with sufficient mortality data (15 countries: Australia, Bolivia, Brazil, Chile, Colombia, Ecuador, Malaysia, New Zealand, Paraguay, Peru, Philippines, Singapore, South Africa, Thailand, Uruguay) exhibits an unprecedented and relatively sharp peak or surge in all-ages ACM during or near January-February 2022 (mid-summer in the Southern Hemisphere), which is synchronous with or immediately preceded by a rapid rollout (all-ages) of a COVID-19 vaccine booster (dose 3 or dose 4, depending on the country), while non-booster doses are also administered at the same time: Figure 5, Figure 4, Figure 11 (top panel), Figure 14 (top two panels), Figure 18, Figure 19, and Figure 21 (Appendix B: Examples of all-cause mortality and vaccination data).

Third: Regarding evidence in support of causality and toxicity, we must include the striking examples described above (“Absence of excess mortality until the COVID-19 vaccines are rolled out” section) where no detectable excess mortality occurs until the COVID-19 vaccines are rolled out, in many countries, on several continents.

Fourth: And we must include the reproducible and consistent nature of the phenomenon (“The COVID-19 vaccines did not save lives and appear to be lethal toxic agents” section).

We should also include prior studies of the same phenomenon, in different jurisdictions, which find compatible results with the present study. In addition to India (Rancourt, 2022), Australia (Rancourt et al., 2022a, 2023), Canada (Rancourt et al., 2021b, 2022c; Rancourt, 2023), and Israel (Rancourt et al., 2023), the USA has a seasonally anomalous peak in ACM occurring in late-summer or fall-2021, which is prominent in the 25-64 years age group in 21 states of the USA, most notably including Alabama, Mississippi, Georgia, Florida and Louisiana. This peak in ACM is synchronous with a rapid surge in vaccine administration (seen as a step-wise increase in cumulative all-dose vaccine administration), in the period corresponding to the so-called “vaccine equity” campaigns in the USA (Rancourt et al., 2022b; their Figures 10C, 10D, and 11A-F). For the entire USA and all-ages, the said late-summer or fall-2021 peak in ACM corresponds to an excess mortality of approximately 160 K deaths, during a period of more than 2 months when approximately 60 M injections (all-doses) were administered. This corresponds to a national all-ages vDFR value of approximately 0.3 % for that period, which would be larger for the states and ages contributing most to the excess mortality (high poverty, most elderly).

6.5 Causality in excess mortality is amply demonstrated

The above-described heads of evidence in support of a causal relation between COVID-19 vaccine administration and temporally associated excess ACM can be summarized as follows:

- i. There is no evidence in ACM-by-time data of any beneficial effect of COVID-19 vaccines. There is no association in time between COVID-19 vaccination and any proportionate reduction in ACM.
- ii. On the contrary, there occurs an onset or increase of a large excess ACM on rolling out the COVID-19 vaccines, in every country and state or province, studied to date, on virtually all continents, including for initial rollouts occurring at significantly different times (by several months).

- iii. The said onset manifests itself as a new sustained regime of high excess ACM — punctuated by additional peaks in ACM associated with specific rapid rollouts (boosters) — which surges at the start of vaccination, irrespective of whether or not there was excess mortality prior to vaccination.
- iv. In many jurisdictions (including each state in Australia), there is no detectable excess ACM until the vaccines are rolled out, when new regimes of high excess ACM are initiated.
- v. In all studied cases in which available ACM and vaccine rollout data is resolved by age (Israel, Chile, Peru), there is synchronicity between every booster rollout and a sharp peak or feature in ACM, in every elderly age group (> 60 years ages), which is some 30 separate events of synchronicity, at different times in seasonal patterns, in both hemispheres.
- vi. With ACM and vaccination data that is not resolved by age, there are nonetheless several examples of synchronicity between a burst of vaccination and an anomalous peak in ACM, especially the nominally January-February 2022 peak in ACM, in all 15 countries having sufficient ACM data in the present study, and in 5 of the 8 states of Australia (Rancourt et al., 2022a).
- vii. There is observed proportionality (not mere correlation) between number of COVID-19 vaccine injections (all ages) and temporally associated all-ages excess ACM, whether one uses specific peaks in ACM or the nominally entire vaccination period. This is shown in Figure 8.
- viii. The corresponding all-ages toxicities (vDFR, excess-ACM/number-vaccine-injections) for specific ACM peaks and for the vaccination period are proportional to each other, irrespective of the country-to-country differences in their values. This is shown in Figure 9.
- ix. The values of all-ages vDFR vary from country to country, as one would expect from national differences in population health status and its domestic heterogeneity, and possibly from national differences in vaccine type used, but always lie in the range of approximately 0.02 % to 0.40 % (Figure 8, Figure 9, Figure 10, and Table 2), attaining 1 % in the exceptional case of India (Rancourt, 2022).

- x. The inferred toxicity (vDFR) increases significantly with age (Figure 17), which is a known characteristic of dose-dependent severe outcomes and fatality from poisons (Wiberg et al., 1970; Rogers and Heard, 2007; Chen et al., 2009; Phua et al., 2009; Shively et al., 2017).
- xi. The COVID-19 vaccines are known to be intrinsically dangerous and have caused deaths in individuals of all ages (see: “COVID-19 vaccines can cause death” section, and references therein).
- xii. There is evidence of the possible effect of COVID-19 vaccination type or manufacturer (see: “The COVID-19 vaccines did not save lives and appear to be lethal toxic agents” section).

As such, the robust criteria described by Ioannidis (2016) for proving causality are amply satisfied:

- **Experiment:** The same phenomenon is independently observed in distinct jurisdictions, for distinct age groups, and at different times, which constitutes ample verification in independent real-world large-scale experiments.
- **Temporality:** The many step-wise increases and anomalous peaks in ACM are synchronous with vaccine rollouts; including in jurisdictions in which excess mortality did not occur until vaccination was implemented after approximately one year into the declared pandemic.
- **Consistency:** The phenomenon is qualitatively the same and of comparable magnitude each time it is observed.

There can be little doubt that the mass COVID-19 vaccination campaigns caused the temporally associated excess mortality in the 17 countries of the present study, and in other countries studied to date.

Regarding quantification (calculation of all-ages vDFR), a main uncertainty arises when the nominally entire vaccination period is used. In this case, when there is excess ACM prior to the start of COVID-19 vaccination, due to causes other than COVID-19

vaccination (i.e., medical and government measures), one should consider whether these causes continue, to some degree, into the COVID-19 vaccination period.

There is some evidence for this. For example, three countries having among the largest vaccination-period all-ages vDFR to peak-period all-ages vDFR ratios (Peru, Ecuador, Colombia; Figure 9) also have very large pre-vaccination all-ages excess ACM features (Figure 2, Figure 6). Three countries having the smallest vaccination-period all-ages vDFR to peak-period all-ages vDFR ratios (Uruguay, Singapore, Malaysia; Figure 9) have no pre-vaccination all-ages excess ACM (Figure 2, Figure 6). However, of the five countries having the largest values of vaccination-period all-ages vDFR (South Africa, Suriname, Bolivia, Paraguay, Philippines; Figure 10, Table 2), three have no all-ages excess ACM in the pre-vaccination period (Suriname, Paraguay, Philippines; Figure 2, Figure 6). Overall, there is no consistent and systematic variation between all-ages vDFR values and pre-vaccination period all-ages excess ACM, and there is essentially the same proportionality (with zero intercept) between all-ages excess ACM and injections administered for both integration periods examined (Figure 8). Importantly, we must also keep in mind, with all-ages evaluations, that pre-COVID-19-vaccination-period interventions causing excess ACM (e.g., government measures) can have large contributions in the younger age groups, whereas vaccine deaths (for boosters in particular; doses 3 and 4) are essentially confined to 60+ year-olds (Figure 11, Figure 14; and see Rancourt et al., 2021a, 2022b, 2023). Furthermore, COVID-19 vaccination is intended to alleviate or eliminate the need for measures other than COVID-19 vaccination, and to reduce mortality, which should reduce non-COVID-19-vaccination measures in the vaccination period. And, finally, it is possible that some of the countries had increased seasonal-flu-vaccination campaigns at the end of 2020.

When an ACM-peak integration region is used (in quantifying an all-ages vDFR), an uncertainty might arise from a non-optimized region of integration for counting the injections that are temporally associated with the peak in ACM, which can lead to some undercounting or overcounting. Generally, it appears that all such uncertainties are small compared to the large country-to-country variability of all-ages vDFR values, which is reported below.

Therefore, the total number of deaths inferred in the present study of 17 equatorial and Southern-Hemisphere countries to have been caused by the COVID-19 vaccine injections is $1,745,000 \pm 49,000$, for a total of 1,388,145,365 injections over the same nominally entire vaccination periods of the 17 countries, compared to the overall Covid-period excess mortality of $2,393,000 \pm 88,000$ (Table 1).

As such, the overall, vaccination-period value of all-ages vDFR for the 17 countries in the study is: $vDFR = (1,745,000 \pm 49,000 \text{ deaths}) / (1,388,145,365 \text{ injections}) = (0.1257 \pm 0.0035) \%$. The mean country-to-country variability, relative to this overall value of all-ages vDFR is 0.079 % (Table 1, see also Table 2). Therefore, the all-ages vDFR results from the 17 countries can be summarized as:

$$\text{overall all-ages vDFR} = (0.126 \pm 0.004) \% \quad (4a)$$

$$\text{country-to-country mean variability in all-ages vDFR} = 0.079 \% \quad (4b)$$

Previously, we estimated that a representative global value would be 0.1 %, and that this would represent approximately 13 million deaths from the COVID-19 vaccines, from 13.25 billion injections up to 24 January 2023 (Rancourt et al., 2023). This can be updated as follows: $(0.1257 \pm 0.0035) \% \times 13.50 \text{ billion injections}$ (2 September 2023, OWID, 2023a) = $(16.97 \pm 0.47) \text{ million COVID-19 vaccine deaths worldwide, to date}$. This current estimate is based on: 10.3 % of worldwide COVID-19 injections, 9.10 % of worldwide population, and a vaccination rate of 1.91 injections per person (all ages), in 17 countries.

This implies that 17.0 ± 0.5 million COVID-19 vaccine deaths would have occurred globally, up to 2 September 2023. In other words, analysis of ACM by time in the 17 countries studied shows that the global COVID-19 vaccination campaign was in effect a mass iatrogenic event that killed $(0.213 \pm 0.006) \%$ of the world population, and did not measurably prevent any deaths.

Keeping in mind the caveats mentioned above, Table 1 also suggests that, overall, in the 17 countries studied,

$$(1,744,829 \pm 49,285) / (2,392,831 \pm 87,904) = (73 \pm 3) \%$$

of excess deaths of the entire nominal Covid period (from 11 March 2020 to the end, usually in 2023, of available data, see Appendix C) were early deaths causally induced by the COVID-19 vaccine injections, keeping in mind that in 9 of the 17 countries this fraction was 100 % (“Absence of excess mortality until the COVID-19 vaccines are rolled out” section). This occurred in circumstances in which there is no evidence in any ACM-by-time data that the vaccination rollouts prevented any deaths whatsoever.

6.6 Assessing other interpretations of the cause of the excess mortality

Competing interpretations of the cause or causes of the excess mortality in the nominally January-February 2022 ACM peaks occurring in equatorial-latitude and Southern-Hemisphere countries are stated (in italics) and discounted as follows:

- I. *The peaks are seasonal peaks or are confounded by seasonal variations in ACM.* Unlikely. The peaks are relatively narrow and occur during the Southern-Hemisphere summer period, normally a trough in mortality. There is no seasonal component of increased mortality corresponding to increased mortality in January-February in the Southern Hemisphere; and equatorial countries do not exhibit seasonal patterns of mortality.
- II. *The peaks are due to intense summer heat waves.* Unlikely. Heat-wave mortality peaks, when they occur, are much narrower and never occur synchronously in a global hemisphere. There were no such climatic events recorded in January-February 2022 in the countries included in this study.
- III. *The peaks are due to one or several earthquakes.* Unlikely. Earthquake peaks are initially very narrow, from people crushed in falling buildings, and they occur locally, near the earthquake epicenter.

- IV. *The peaks are due to simultaneous and short-lived outbreaks of war, famine or some deadly disease in January-February 2022, across equatorial latitudes and the Southern Hemisphere.* Unlikely. No such events occurred in the countries studied. (Sudden emergence of deadly COVID-19 variants is addressed in point-VII, and below.)
- V. *The peaks are due to aggressive Covid measures, treatments or responses (other than vaccine booster administration) applied in January-February 2022.* Unlikely. Covid measures, treatments or responses vary widely from jurisdiction to jurisdiction, as do the demographics of the populations at highest risk (e.g., Johnson and Rancourt, 2022). No measures, treatments or responses were uniformly and synchronously applied in January-February 2022 in the equatorial regions and the Southern Hemisphere.
- VI. *The peaks are due to underlying population health conditions and/or comorbidities of the deceased.* Unlikely. The fatal consequences of underlying health conditions and the presence of chronic conditions among those with the shortest life expectancies cannot suddenly arise by accident, at the population level, to cause a sharp peak in mortality. Rather, these underlying circumstances will participate in causing country-to-country variability in fatal response to any rapidly introduced toxic agent. Furthermore, there is no evidence in the ACM data of a so-called dry tinder effect: Those who died in the mortality peaks would not have died in a short time that followed, and the excess mortalities are large net excesses on the timescale of several years.
- VII. *The peaks are due to the emergence of one or more variant(s) of SARS-CoV-2 causing synchronous mortality peaks in January-February 2022, across the equatorial regions and the Southern Hemisphere.* Unlikely. Epidemiological theory of a contact-spreading viral respiratory disease predicts a wide range of delay (months, years) between seeding of a new variant and measurable exponential growth of mortality (or peak of new infections), depending sensitively on characteristics of the society (e.g., Parham and Michael, 2011; Hasegawa and Nemoto, 2016; Ma et al., 2022).

On the question of comorbidities, it is known that newly imposed conditions or challenges against frail or susceptible individuals can induce premature death. Premature death is death. The higher the frailty, the more likely and immediate the fatal response. In such circumstances, the said newly imposed conditions or challenges unambiguously caused the death. This would be true of a violent traffic accident and it is equally true for challenges with a toxic substance or with newly imposed circumstances causing significant psychological stress. Regarding logically attributing the cause of death, it is irrelevant, for example, that in all cases the heart ultimately stopped functioning and breathing stopped.

Regarding the theory of emergence of one or more variant(s) of SARS-CoV-2, this emergence would have to cause simultaneous peaks and surges of mortality in 17 countries across 4 continents (Figure 1, Figure 2, Figure 4, Figure 11, Figure 14, Figure 18), which is a statistically impossible occurrence if we accept the theories of spontaneous viral mutations and contact spreading of viral respiratory diseases; and all the resulting peaks of mortality would have the remarkable coincidence of occurring precisely when vaccine boosters were rolled out.

In advancing interpretations of excess mortality based on changing variant predominances, one should be wary that variant predominance evaluation based on data from genomic surveillance networks is a new methodology in its tentative infancy, essentially introduced during the latest (COVID-19) pandemic declared by the WHO, which may be driven by institutional and industry interests, rather than rigorous efforts to determine analytic reliability. At this stage, the public health establishment relies in part on surveys of experts to gauge robustness of the method, and even these selected experts opine confidence scores in the broad range from “very low” to “very high”, depending on methodological details and intended application (Worp et al., 2023).

A main limitation of genomic surveillance, which has not been quantified or resolved, is that the genomic measurements, assumed to be valid and intended to represent a given region at a given time, are taken from small non-randomized cohorts. Randomization is

essential because people go into clinics and hospitals (where samples for genomic measurement are volunteered) precisely because they have particular concerns or conditions. In the words of the many authors of a recent overview of genomic surveillance: “The findings of this report are subject to at least four limitations. First, estimates might be biased by nonrandom sampling of specimens or differential timing of reporting...” (Lambrou et al., 2022). Also, the surveillance itself targets so-called “consensus variants”, which are arbitrarily chosen (Lambrou et al., 2022).

In addition, the viral characteristics that are assigned by authors to a particular variant believed to be predominant are generally assigned by circular reasoning to achieve the desired interpretation of epidemiological data, rather than obtained in controlled clinical trials. This was the case in a recent attempt to explain an excess mortality event in India: Dhar et al. (2021) postulated that the April-July 2021 “second wave” event in Delhi (India) was due to the Delta variant, which would have quickly swept Delhi to become predominant because it would have had higher transmissibility and larger immune escape than concomitantly circulating variants. However, Dhar et al. estimated the needed characteristics of Delta by fitting a model to the epidemiological data and to the variant predominance (estimated by genomic measurements from small non-randomized cohorts). The inferred characteristics of Delta were obtained by fitting the said characteristics to the epidemiological data, rather than being independently determined via any clinical, animal-model or *in vitro* study.

6.7 Implications regarding age-dependence of fatal toxicity of COVID-19 vaccines

The detailed age and dose-discriminated data for Chile and Peru is remarkable and allows unprecedented certainty in assigning temporal associations between sharp peaks in ACM and sharp peaks in specific booster rollouts to specific age groups. It also allows unprecedented certainty in quantifying the age dependence of the fatal toxicity from COVID-19 vaccines.

The risk of death per injection (vDFR) increases exponentially with age, for 60+ years ages, doubling approximately every 4 years in age, and attains values of approximately 5 % for 90+ year-olds subjected to dose 4 (2nd booster) in both Chile and Peru (Figure 17), for which age and dose-discriminated data is available. vDFR = 5 % corresponds to 1 death per 20 injections of dose 4 for 90+ year-olds in Chile and Peru (Figure 11, Figure 14, Figure 17).

The exponential increases of vDFR with age found for Chile and Peru in the present study (Figure 17) are essentially the same as those previously measured for Israel and Australia (Rancourt et al., 2023; their Figure 3), and the extracted doubling ages (T2) (increases in age to double the risk of death per injection, vDFR) are consistent between all values determined to date:

Table 3: Extracted doubling ages, increases in age to double risk of death per injection

Fitted function: $vDFR = A 2^{Age/T2}$			
Country	Dose No.	Study	T2 (\pm) (years)
Chile	3	This study	5.6(0.4)
Chile	4	This study	3.7(0.3)
Peru	3	This study	3.8(0.5)
Peru	4	This study	4.0(0.2)
Australia	all doses	Rancourt et al., 2023	5.4(0.3)
Israel	all doses	Rancourt et al., 2023	4.9(0.5)

In Table 3, the quoted errors are 1-sigma errors. For Australia and Israel, the dose numbers were not discriminated, and the calculation is based on the nominally entire vaccination periods, rather than specific ACM peaks (Rancourt et al., 2023).

This doubling age (T₂) of approximately 4-5 years for risk of dying per injection of the COVID-19 vaccines (Table 3) is approximately half of the doubling age of 10 years for risk per year (or per any sufficiently small fixed unit of time) of dying of all causes in a modern human population, and of the main old-age diseases cancer, pneumonia and heart disease (Strehler and Mildvan, 1960).

This difference between doubling age for risk of dying per injection and risk per year of dying of all causes means that the age effect is greater on risk of dying from being challenged by injection with the COVID-19 vaccines than on the risk of dying (per fixed unit time) of old age in general. A doubling age may be a characteristic of the type of challenge. For example, risk of death from challenge by violent automobile accident would be relatively age-independent, corresponding to a large T₂. Likewise, if the risk of death from the challenge is tied to a physiological system that degrades more rapidly with age, for advanced ages, than overall death-causing aging itself, then T₂ would be smaller than its value for risk (per fixed unit time) of dying of all causes. Therefore, doubling age (T₂) provides information about the mechanism of death from challenge by COVID-19 vaccine injection, for elderly subjects.

More importantly and concretely, our results establish large and age-dependent values of vDFR in elderly people, as large as 5 % (1 death per 20 injections) measured for 90+ year-olds in Chile and Peru (Figure 17), 1 % (1 death per 100 injections) (nominally all ages) in India (Rancourt, 2022), 0.55 % (1 death per 180 injections) for 80+ year-olds in Israel (Rancourt et al., 2023; their Table 2), and 0.93 % (1 death per 110 injections) for 85+ year-olds in Australia (Rancourt et al., 2023; their Table 1).

This turns the public health policy of prioritizing elderly people for injection with COVID-19 vaccines on its head, in actual circumstances in which there were no reliable controlled clinical trials focussing on representative elderly subjects.

The public health notion that elderly and vulnerable individuals must be prioritized for COVID-19 vaccination assumes:

- i. a constant age-independent vDFR
- ii. a small value of vDFR (optimistically estimated from managed trials, funded by the pharmaceutical industry)

Our research shows that both assumptions (i and ii) are false, and far from reality in the field, on the scale of nations.

The said public health notion has always been baseless since it was not anchored in any sufficient evaluation of age-stratified risk of fatality from the COVID-19 injections (e.g., Veronese et al., 2021; Abbatecola et al., 2022; Gao et al., 2022), and is now proven to be incorrect.

Prioritizing elderly people for COVID-19 vaccination, in the absence of relevant data, was reckless. Norway may be the only jurisdiction that immediately and publicly recognized a problem and changed its policy regarding vaccinating the most elderly and frail (*Reuters*, 18 January 2021; *Fortune*, 15 January 2021).

The same criticism may be valid for other treatments supposedly intended to protect the elderly, such as seasonal influenza vaccines, or any medication having intrinsic toxicity and not sufficiently (and independently) tested on elderly cohorts having age-related chronic conditions.

Some will be tempted to compare our results (Figure 17, Table 3) with published age-stratified COVID-19 infection fatality rates (IFR) (e.g., COVID-19 Forecasting Team, 2022; Pezzullo et al., 2023).

While in principle this is a correct approach of risk-benefit analysis, we believe that the IFR studies are not reliable, for the following reasons:

- i. The deaths in the numerator of IFR are “COVID-19 deaths”, and this cause of death assignation is susceptible to bias and is highly uncertain (Rancourt,

- 2020; Rancourt et al., 2020, 2021a, 2021b, 2022c; and the present study, esp. Figure 2).
- ii. The number of infections, in the denominator of IFR, is reliant on molecular antibody tests, which are not specific and have not been sufficiently validated (e.g., Rancourt, 2021).
 - iii. If the IFR evaluations were valid, then it would be virtually impossible for jurisdictions like Australia, Canada, India, Israel, Malaysia, New Zealand, Paraguay, Philippines, Singapore, Suriname, Thailand, and Uruguay to have no detectable or relatively very little excess ACM in the pre-vaccination period of the declared pandemic.
 - iv. We do not detect any excess ACM that can be attributed to COVID-19 in the jurisdictions that we have studied in detail (USA and all its states; Canada and its provinces; France and its departments and regions; Australia and its states) (Rancourt, 2020, 2022; Rancourt et al., 2020, 2021a, 2021b, 2022a).

At this stage, we would urge the WHO and governments to pause the misguided and baseless public health policy of prioritizing elderly people for injection with COVID-19 vaccines, and with influenza vaccines, until valid risk-benefit analyses are made and publicly reported.

6.8 Excess ACM by time prior to COVID-19 vaccine rollouts

None of the 17 countries studied have anomalous excess mortality peaks prior to the WHO's 11 March 2020 declaration of a pandemic (Figure 2, Figure 12, Figure 15, Figure 20, Figure 21, Figure 22), as is generally true for all countries studied previously.

Eight of the 17 countries studied have significant excess ACM peaks in the approximately one year period between 11 March 2020 and the first COVID-19-vaccine

injections (hereafter, “pre-vaccination period”) (Figure 2, Figure 6, Figure 21): Argentina, Bolivia, Brazil, Chile, Colombia, Ecuador, Peru, and South Africa.

Two of the 8 countries with significant excess ACM peaks in the pre-vaccination period (Argentina and Brazil) simply have larger seasonal peaks in mid-2020, and no other anomalous excess ACM features in the pre-vaccination period (Figure 6).

The other 6 of the 8 countries with significant excess ACM peaks in the pre-vaccination period (Bolivia, Chile, Colombia, Ecuador, Peru, South Africa) have very large, unprecedented, and relatively sharp peaks in ACM in the pre-vaccination period (Figure 6):

- Ecuador has a stunning peak that surges literally immediately after the 11 March 2020 declaration of a pandemic (Figure 6), which is reminiscent of the synchronous anomalous mortality events in New York (USA), Lombardy (Italy), Madrid (Spain), Stockholm (Sweden) and other such hot spots where this anomaly occurred. These hot-spot peaks have been interpreted to be caused by sudden changes in medical and institutional protocols and government responses, tied to the declaration of a pandemic, and that they cannot be due to a globally spreading viral respiratory disease (Rancourt, 2020, 2023; Rancourt et al., 2020, 2021a, 2021b).
- Bolivia, Chile, Peru and South Africa have similarly very large, unprecedented, and relatively sharp peaks in ACM, as seen for Ecuador, which occur not immediately but soon after the 11 March 2020 declaration (Figure 6). We would interpret these peaks in the same way as for Ecuador, as being caused by sudden changes in medical and institutional protocols and government responses, following the 11 March 2020 declaration of a pandemic, and that they are not due to a globally spreading viral respiratory disease.

- Peru has a stunning second sharp peak in ACM that immediately follows its first such peak, creating a double-peak structure in the pre-vaccination period (Figure 6). Ecuador also has a (relatively smaller) second peak in ACM, which follows the first peak in the pre-vaccination period (Figure 6).
- In addition to the above-described features, Colombia and South Africa have large sharp peaks in ACM, which immediately precede the start of vaccination (Figure 6). Similarly, Peru, Ecuador and Bolivia have ACM increases, which immediately precede and lead into the vaccination period (Figure 6). It is possible that such peaks and increases are due to augmented influenza vaccination programs at that time, or sudden changes imposing aggressive government measures.

Also, Australia, Malaysia, New Zealand and Uruguay appear to have somewhat smaller ACM in the approximately-one-year-duration pre-vaccination period, compared to the recent historic trend (Figure 2, Figure 6, Figure 20, Figure 21). These depressed mortalities in the pre-vaccination period are near the limit of statistical significance compared to historic behaviours of ACM by time. If they are real, and if they are due to institutional and government measures implemented after 11 March 2020, then some questions arise:

- i. Why did this occur in only 4 of the 17 countries studied, and the opposite in 8 of the countries?
- ii. Why did these countries implement coerced COVID-19 vaccination, after one year of such positive observations?

Overall, the patterns and magnitudes of ACM by time vary tremendously, from country to country, in the pre-vaccination period (Figure 6):

- from complete absence of excess mortality (9 of 17 countries),
- to significantly but moderately larger mid-2020 seasonal peaks (Brazil, Argentina),

- to variable numbers (1, 2 or almost 3, counting a pre-vaccination rise in ACM) of very large peaks in ACM, occurring at different times from country to country (Bolivia, Chile, Colombia, Ecuador, Peru, South Africa).

In addition, a country such as Bolivia, which has an extraordinarily large peak of excess ACM in the pre-vaccination period, can share a major land border with a country such as Paraguay, which has no detectable excess ACM in the pre-vaccination period.

We attribute this variability in the pre-vaccination period to large country-to-country differences in aggressive medical and government measures, convoluted with underlying population structures of fragile individuals, similarly to the large state-to-state differences in the USA (Rancourt et al., 2021a; Johnson and Rancourt, 2022).

By comparison, the patterns and magnitudes of ACM by time are relatively consistent, from country to country, in the vaccination period (Figure 6):

- all 17 countries have significant excess ACM in the vaccination period,
- virtually all 17 countries have large peaks of excess ACM in early-2021 when the COVID-19 vaccines are first rolled out,
- followed by a peak in excess ACM in early-2022 (the nominally January-February 2022 peak) when boosters are rolled out.

For Singapore, both vaccination-period peaks are later, but each is immediately preceded by a maximum in vaccine-injection administration (Figure 6). For Suriname, the first vaccination-period peak in ACM actually has a two-peak structure, but each peak in the pair is preceded by a maximum in vaccine-injection administration (Figure 6).

We attribute this relative constancy from country to country of excess ACM by time of the vaccination period as being due to the large uniformity in timing of COVID-19 vaccine and booster rollouts across the 17 countries; whereas medical and government measures were generally more disparate in the pre-vaccination period.

7. Conclusion

7.1 Causality proven

The 17 countries studied (Argentina, Australia, Bolivia, Brazil, Chile, Colombia, Ecuador, Malaysia, New Zealand, Paraguay, Peru, Philippines, Singapore, South Africa, Suriname, Thailand, Uruguay) comprise 9.10 % of worldwide population, 10.3 % of worldwide COVID-19 injections (vaccination rate of 1.91 injections per person, all ages), virtually every COVID-19 vaccine type and manufacturer, and span 4 continents.

The scientific tests for causality are amply satisfied, as extensively demonstrated in these sections of the present paper:

- COVID-19 vaccines can cause death
- Absence of excess mortality until the COVID-19 vaccines are rolled out
- The COVID-19 vaccines did not save lives and appear to be lethal toxic agents
- Strong evidence for a causal association and vaccine lethal toxicity
- Causality in excess mortality is amply demonstrated
- Assessing other interpretations of the cause of the excess mortality
- Implications regarding age-dependence of fatal toxicity of COVID-19 vaccines

Also, there are no known facts that disprove the inferred and quantitative causal relation between the observed excess ACM peaks and the temporally associated COVID-19 vaccine and booster rollouts.

7.2 Actual vaccine mortality much larger than that incorrectly inferred from faulty data

There can be little doubt that the peaks in excess ACM are caused by the COVID-19 vaccinations, with a mean all-ages fatal toxicity by injection of vDFR = $(0.126 \pm$

0.004) %, or approximately 1 death per 800 injections, which is reasonably expected to be globally representative.

This is a staggering number, compared to what is generally believed about traditional vaccines, which is approximately one serious adverse effect per million (Malhotra, 2023). It is three orders of magnitude (one thousand times) larger.

In contrast to this large number from ACM data, the effective all-ages vDFR values inferred from other sources — the small number of autopsy confirmations, adverse effect monitoring of deaths, the small number of national vaccine injury compensation payments for deaths, pharmaceutical-industry clinical trial reports, and death-certificates and corresponding government-reported cause-of-death statistics — for COVID-19 vaccines are significantly smaller, again by orders of magnitude.

Therefore, we conclude with a high degree of certainty that adverse-effect monitoring, clinical trial reports, and death-certificate statistics greatly underestimate the fatal toxicity of the injections.

The large gap between incorrectly inferred and actual population-level vDFR values probably occurs because of systemic avoidance of admitting the injections as a possible cause of death in frail and vulnerable patients. For example, the cause(s) of death will be attributed to particular organ or system failures, without reference to the vaccine, even if the injection was recent and the patient was previously stable.

The measured all-ages vDFR of $(0.126 \pm 0.004) \%$ implies that 17.0 ± 0.5 million COVID-19 vaccine deaths would have occurred globally, up to 2 September 2023. It appears that the global COVID-19 vaccination campaign was in effect a mass iatrogenic event that killed $(0.213 \pm 0.006) \%$ of the world population (1 death per 470 living persons, in less than 3 years), and did not measurably prevent any deaths.

7.3 The policy of prioritizing elderly residents for COVID-19 vaccination must be ended

The detailed age and dose-discriminated data for Chile and Peru allows unprecedented certainty in quantifying the age dependence of the fatal toxicity from COVID-19 vaccines. The risk of death per injection (vDFR) increases exponentially with age, for 60+ years ages, doubling approximately every 4 years in age, and attains values of approximately 5 % for 90+ year-olds subjected to dose 4 (2nd booster). vDFR = 5 % corresponds to 1 death per 20 injections of dose 4 for 90+ year-olds.

These vDFR values are consistent with all prior evaluations based on ACM:

- 1 % (1 death per 100 injections) (nominally all ages) in India (Rancourt, 2022)
- 0.55 % (1 death per 180 injections) for 80+ year-olds in Israel (Rancourt et al., 2023; their Table 2)
- 0.93 % (1 death per 110 injections) for 85+ year-olds in Australia (Rancourt et al., 2023; their Table 1)

The population-level age-dependence of vDFR (doubling every 4 years in age) and its large magnitude should induce governments to immediately implement an absolute pause on the baseless public health policy of prioritizing elderly people for injection with COVID-19 vaccines, until valid risk-benefit analyses are made and publicly reported.

The same may be true of seasonal influenza vaccines.

(see appendixes after references)

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Appendix A: Sources of mortality and vaccination data

Table 4 describes the data and the sources of the data used in this work: All-cause mortality (ACM) and COVID-19 vaccine administration, for the 17 countries.

Table 4: Appendix A - Type of data and sources of data

Country	Data	Period	Time unit	Filters	Source
Australia	ACM	2015-2023	Week	-	WMD, 2023
Chile	ACM	2016-2023	Week	-	WMD, 2023
Colombia	ACM	2015-2023	Week	-	WMD, 2023
Ecuador	ACM	2015-2023	Week	-	WMD, 2023
New Zealand	ACM	2015-2023	Week	-	WMD, 2023
Peru	ACM	2017-2022	Week	-	WMD, 2023
South Africa	ACM	2015-2023	Week	-	WMD, 2023
Argentina	ACM	2015-2021	Month	-	WMD, 2023
Bolivia	ACM	2015-2022	Month	-	WMD, 2023
Brazil	ACM	2015-2023	Month	-	WMD, 2023
Malaysia	ACM	2015-2023	Month	-	WMD, 2023
Paraguay	ACM	2015-2023	Month	-	WMD, 2023
Philippines	ACM	2015-2022	Month	-	WMD, 2023
Singapore	ACM	2015-2023	Month	-	WMD, 2023
Suriname	ACM	2015-2021	Month	-	WMD, 2023
Thailand	ACM	2015-2023	Month	-	WMD, 2023

Uruguay	ACM	2015-2023	Month	-	WMD, 2023
Australia	Vaccination	2021-2023	Day	-	OWID, 2023
Chile	Vaccination	2020-2023	Day	-	OWID, 2023
Colombia	Vaccination	2021-2023	Day	-	OWID, 2023
Ecuador	Vaccination	2021-2023	Day	-	OWID, 2023
New Zealand	Vaccination	2021-2023	Day	-	OWID, 2023
Peru	Vaccination	2021-2023	Day	-	OWID, 2023
South Africa	Vaccination	2021-2023	Day	-	OWID, 2023
Argentina	Vaccination	2020-2023	Day	-	OWID, 2023
Bolivia	Vaccination	2021-2022	Day	-	OWID, 2023
Brazil	Vaccination	2021-2023	Day	-	OWID, 2023
Malaysia	Vaccination	2021-2023	Day	-	OWID, 2023
Paraguay	Vaccination	2021-2023	Day	-	OWID, 2023
Philippines	Vaccination	2021-2023	Day	-	OWID, 2023
Singapore	Vaccination	2020-2022	Day	-	OWID, 2023
Suriname	Vaccination	2021-2022	Day	-	OWID, 2023
Thailand	Vaccination	2021-2022	Day	-	OWID, 2023
Uruguay	Vaccination	2021-2023	Day	-	OWID, 2023
Chile	ACM	1990-2023	Day	By age ¹	DEIS, 2023
Chile	Vaccination	2020-2023	Day	By age ¹ , by dose ²	MinCiencia, 2023
Peru	ACM	2017-2023	Day	By age ¹	MINSA, 2023a
Peru	Vaccination	2021-2023	Day	By age ¹ , by dose ³	MINSA, 2023b
New Zealand	ACM	2011-2023	Week	By age group ⁴	Stats NZ, 2023a
New Zealand	Vaccination	2021-2023	Day	By dose ⁵	Stats NZ, 2023b ⁶

Notes for Table 4:

¹ Single year of age available. The data for the yearly ages are summed to create data for the specific age groups used in this study.

² Five categories: Dose 1, Dose 2, Dose 3, Dose 4, Unique dose

³ Six doses: Dose 1, Dose 2, Dose 3, Dose 4, Dose 5, Dose 6. There was no data for a dose 7; and we did not use dose 8 and dose 9 data, which consisted of small numbers.

⁴ Four age groups: 0-29, 30-59, 60-79, 80+

⁵ Five categories: Dose 1, Dose 2, Dose 3 (referred to as “First Boosters” in the source data), Dose 4 (referred to as “Second Boosters” in the source data), Third primary dose

⁶ The vaccination data was originally downloaded on 24 February 2023 (Health NZ, 2023), which we verified to be essentially exactly the same as that more recently downloaded (14 September 2023, Stats NZ, 2023b).

We use a sum of days to generate ACM by week (ACM/w), similarly for vaccination data.

The vaccination data from OWID (2023) is of two types:

- Vaccine doses, cumulative or new per day
- Vaccine booster doses, cumulative

In Figure 2, in Figure 4, in Figure 6 and in Figure 21, ACM data is from WMD (2023) and vaccination data is from OWID (2023).

In Figure 5, in Figure 7, in Figure 22 and in Figure 23, ACM data is from WMD (2023).

In Figure 11, ACM data is from DEIS (2023) and vaccination data is from MinCiencia (2023).

In Figure 12 and in Figure 13, ACM data is from DEIS (2023).

In Figure 14, ACM data is from MINSA (2023a) and vaccination data is from MINSA (2023b).

In Figure 15 and in Figure 16, ACM data is from MINSA (2023a).

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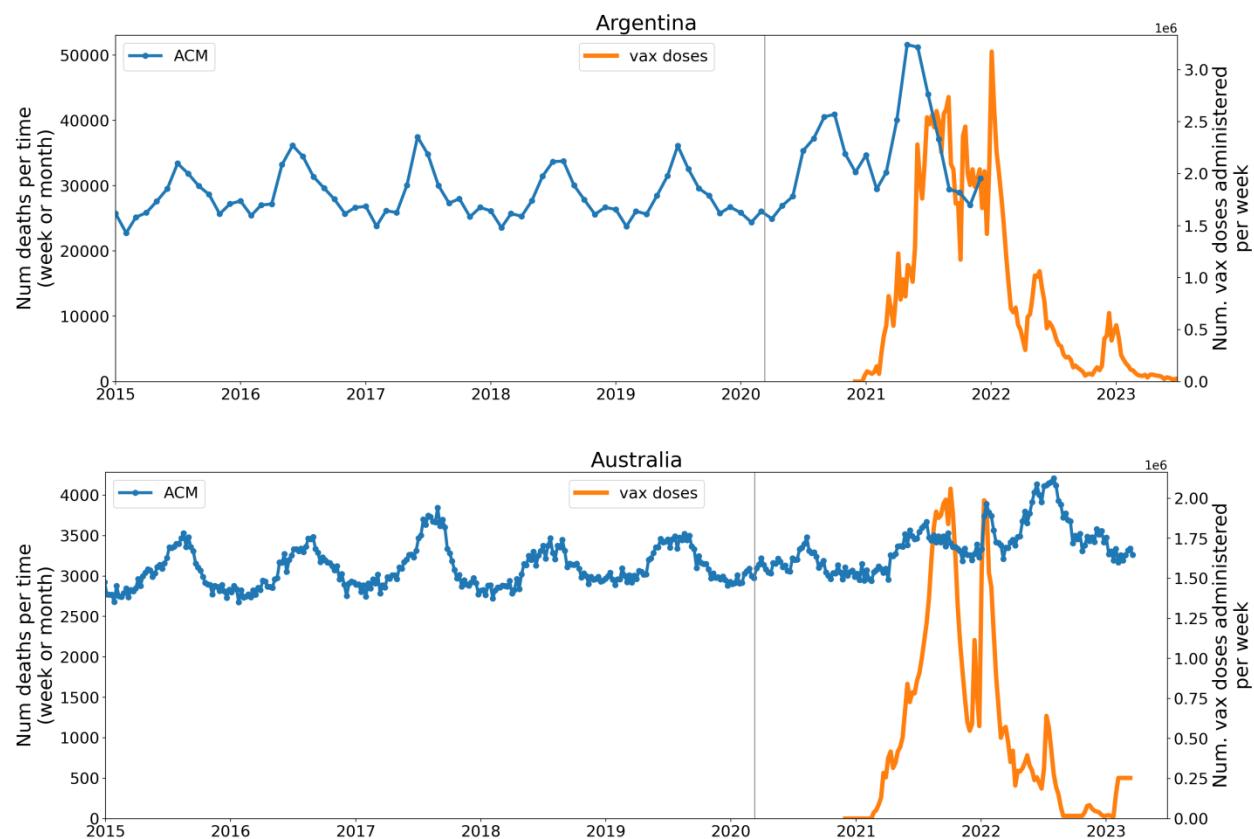
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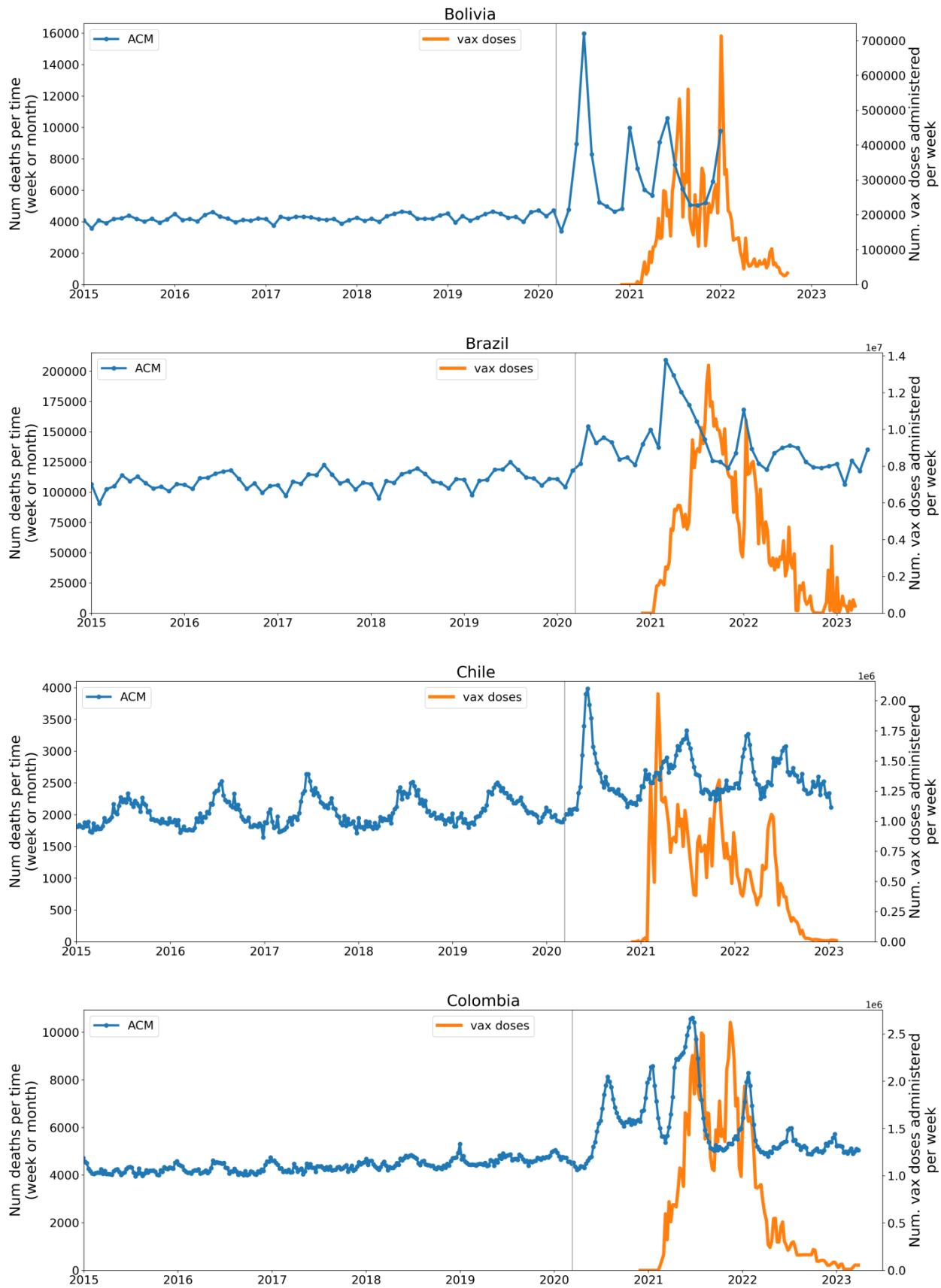
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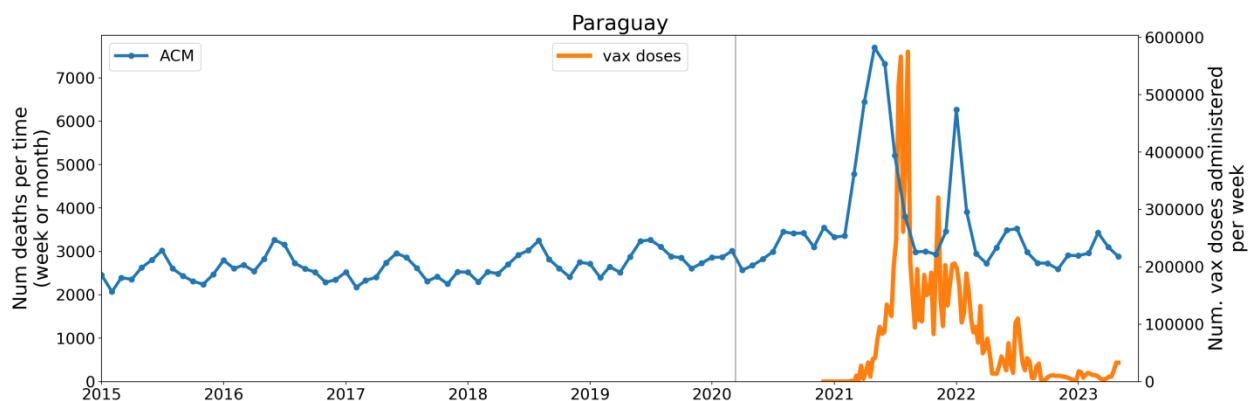
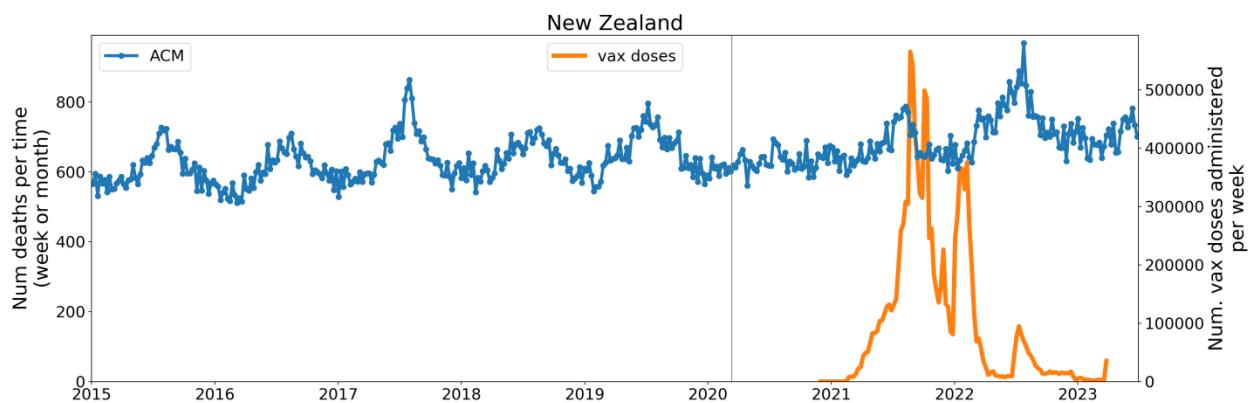
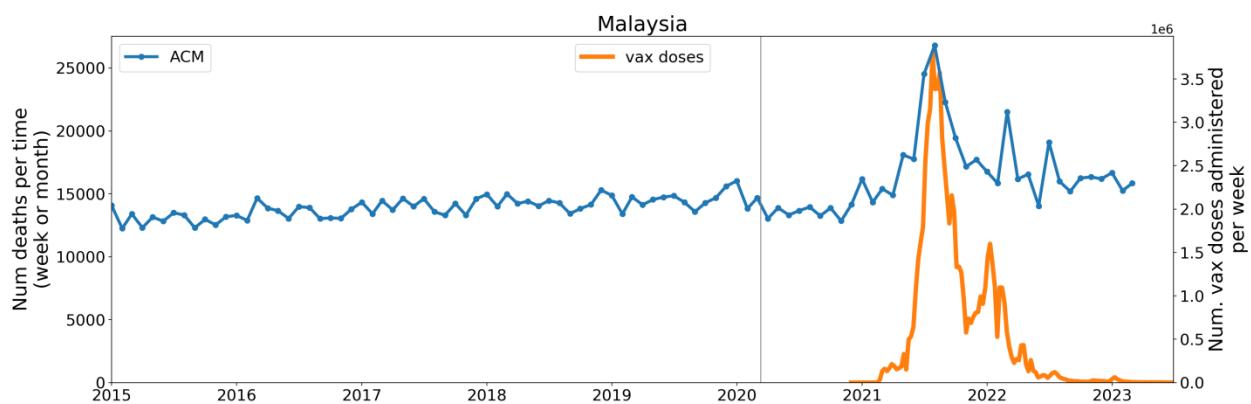
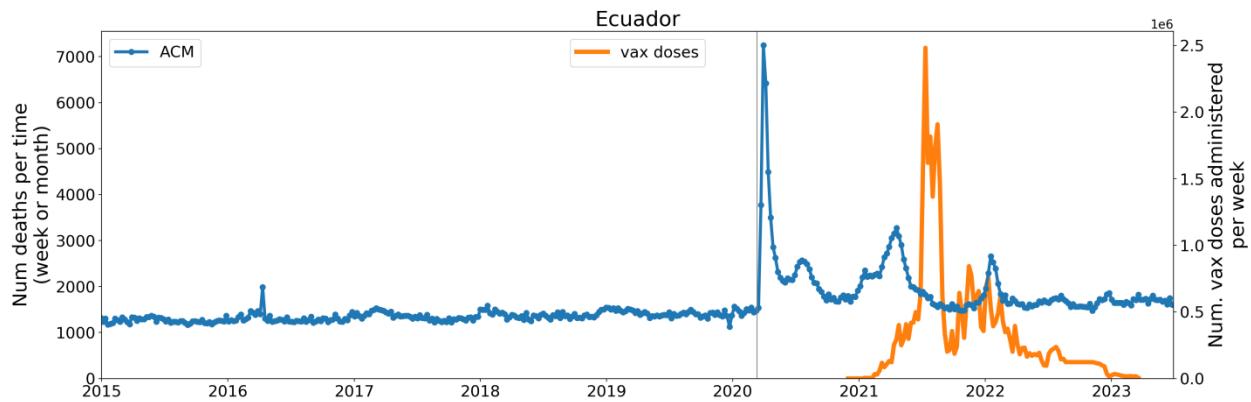
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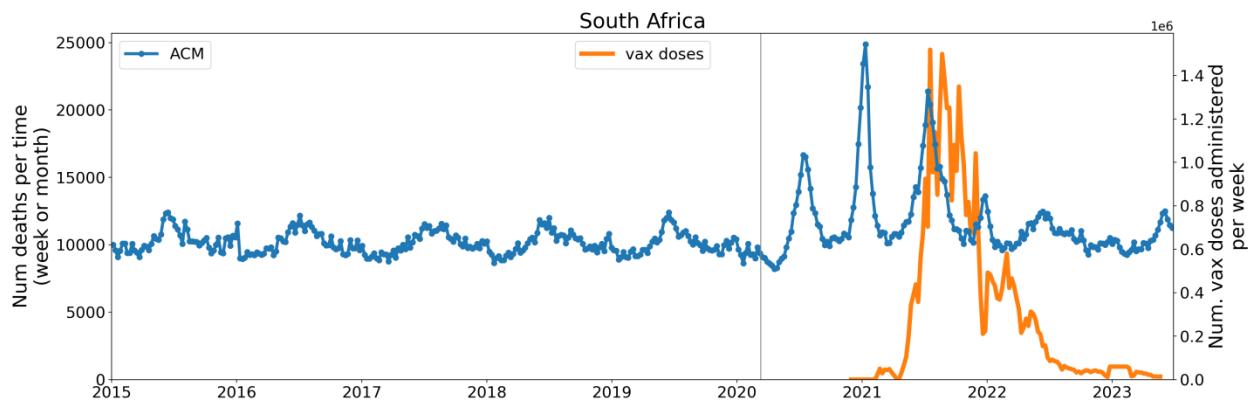
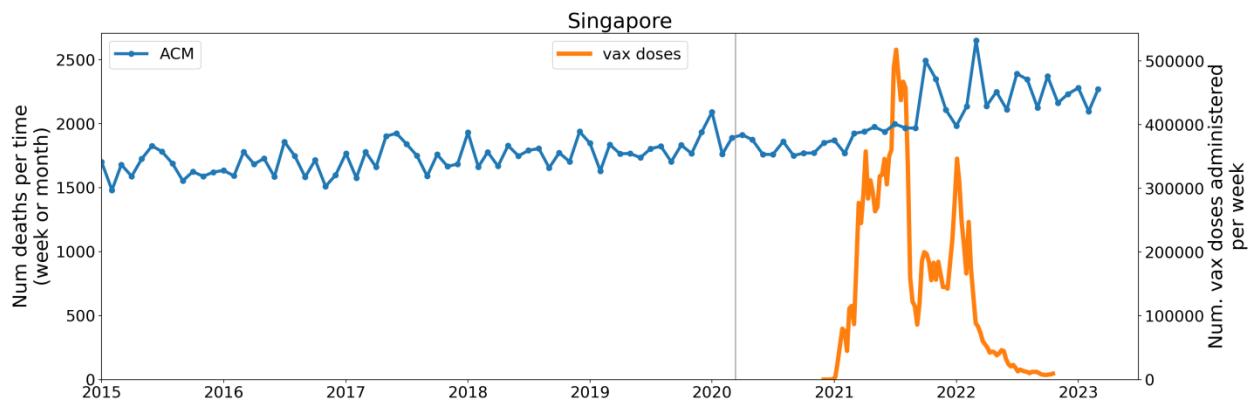
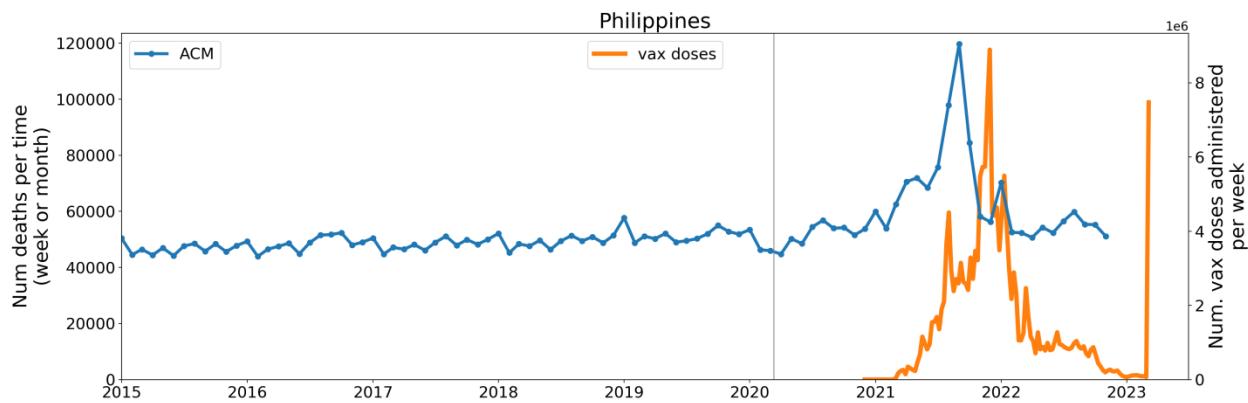
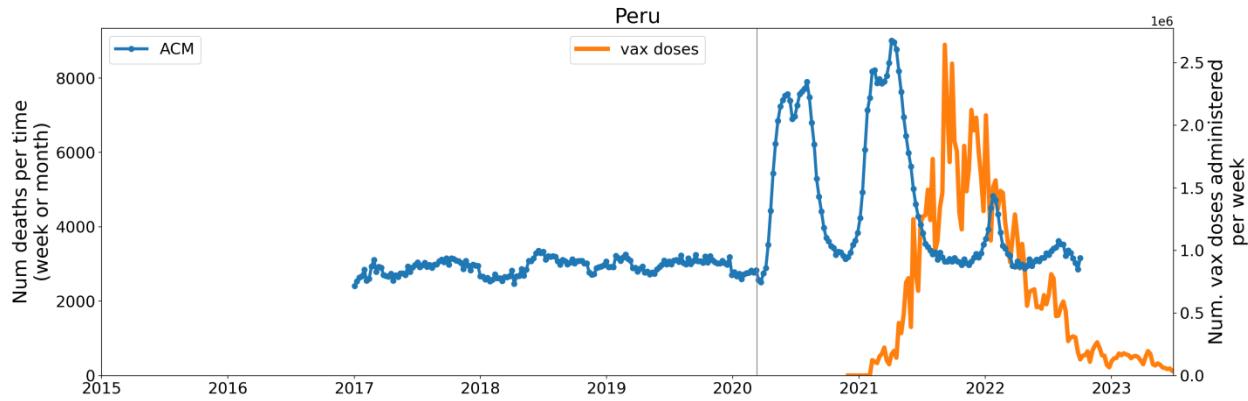
Appendix B: Examples of all-cause mortality and vaccination data

The following panels show ACM by time (week or month, as seen on the graphs) (blue, left scale), and vaccine administration (all dose numbers) by week (orange, right scale), 2015-2023, for the 17 countries in this study, as labelled (Argentina, Australia, Bolivia, Brazil, Chile, Colombia, Ecuador, Malaysia, New Zealand, Paraguay, Peru, Philippines, Singapore, South Africa, Suriname, Thailand, and Uruguay). The thin vertical grey line, in each panel, indicates the date (11 March 2020) at which a pandemic was declared by the World Health Organization (WHO).









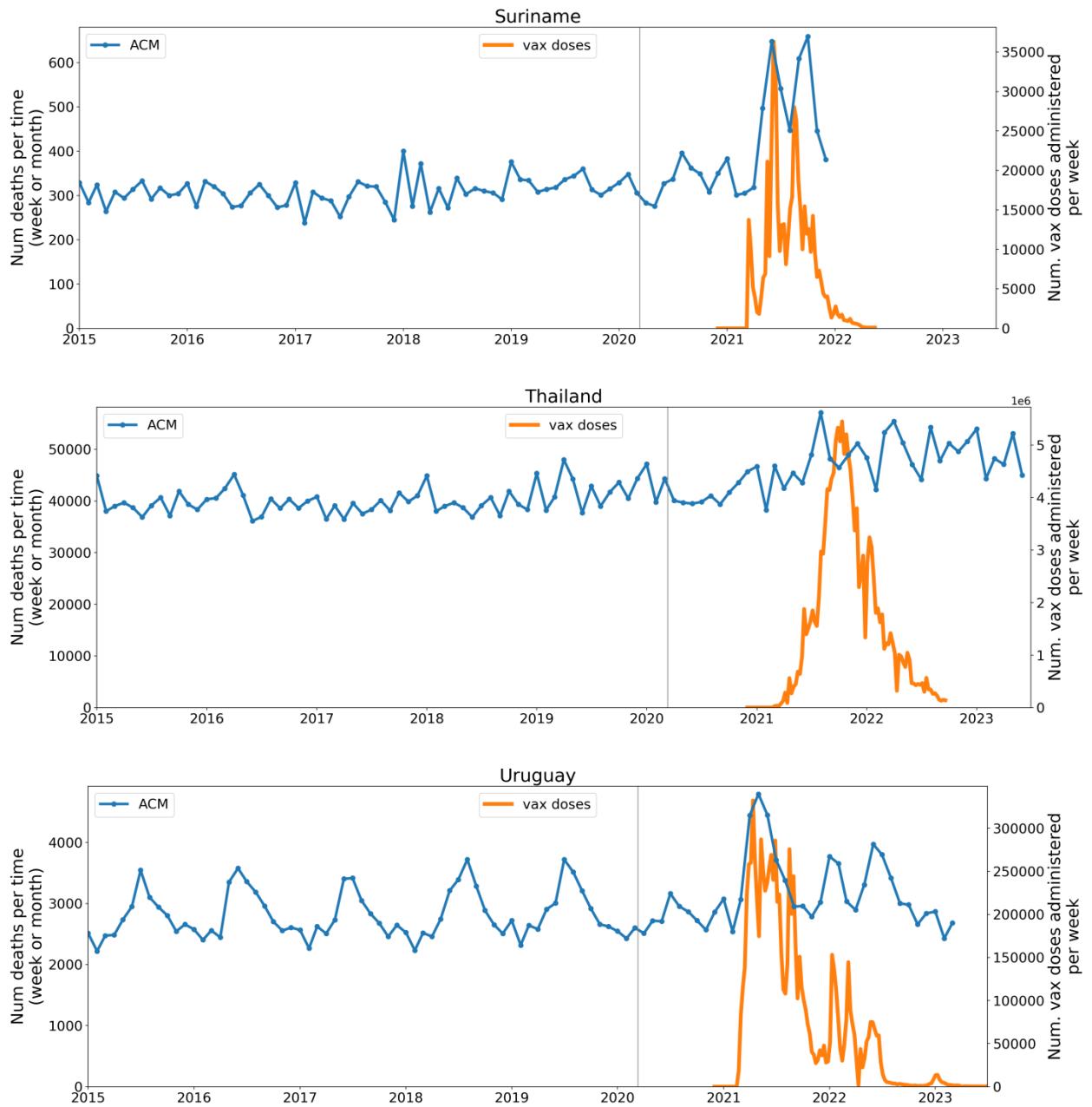


Figure 21: Appendix B - Examples of all-ages all-cause mortality (blue) and all-ages all-doses vaccine-administration (orange) data, for the 17 countries in this study (as labelled). The vertical grey line indicates the 11 March 2020 WHO declaration of a pandemic. The data sources are specified in Appendix A.

Appendix C: Technical and specific information for applications of the methods to the data

Table 5 provides the parameters used in applying the method described in “Historical-trend baseline for a period (or peak) of mortality” (Method 1, see Methods) to the data. In this method, one calculates the trend-lines fitted (straight lines fitted by least squares minimization) to ACM integrated over the periods of equal duration as the period of interest (the vaccination period or a specific-rollout period). The said trend lines are used to calculate the baseline integrated mortality in the periods of interest, in order to obtain the excess ACM of the periods of interest. All vaccine dose counting is on the same time interval as the integration interval for ACM in the period of interest.

The same integration windows (start date to end date, by week or by month, inclusive) have been used for Method 2, for the cases where Method 2 has also been applied.

Table 5: Appendix C - Parameters used to apply the trend-line method (Method 1) to the data

Country	Period of interest	Age group	Weeks / Months*	Duration	Number of integration points used ⁺	From ⁺
Australia	Vaccination	All ages	[2021-W07 – 2023-W12]	110 weeks	2	-1
Chile	Vaccination	All ages	[2020-W52 – 2023-W04]	110 weeks	N/A	N/A
Colombia	Vaccination	All ages	[2021-W07 – 2023-W12]	110 weeks	2	-1
Ecuador	Vaccination	All ages	[2021-W03 – 2023-W08]	110 weeks	2	-1
New Zealand	Vaccination	All ages	[2021-W07 – 2023-W12]	110 weeks	2	-1
Peru	Vaccination	All ages	[2021-W06 – 2022-W40]	87 weeks	N/A	N/A

South Africa	Vaccination	All ages	[2021-W07 – 2023-W12]	110 weeks	2	-1
Argentina	Vaccination	All ages	[2020-M12 – 2021-M12]	13 months	2	-1
Bolivia	Vaccination	All ages	[2021-M01 – 2022-M01]	13 months	3	-1
Brazil	Vaccination	All ages	[2021-M01 – 2023-M03]	27 months	2	-1
Malaysia	Vaccination	All ages	[2021-M01 – 2023-M03]	27 months	2	-1
Paraguay	Vaccination	All ages	[2021-M01 – 2023-M03]	27 months	2	-1
Philippines	Vaccination	All ages	[2021-M01 – 2022-M11]	23 months	3	-1
Singapore	Vaccination	All ages	[2020-M12 – 2022-M10]	23 months	2	-1
Suriname	Vaccination	All ages	[2021-M01 – 2021-M12]	12 months	6	-1
Thailand	Vaccination	All ages	[2021-M01 – 2022-M09]	21 months	3	-1
Uruguay	Vaccination	All ages	[2021-M01 – 2023-M03]	27 months	2	-1
Australia	Booster [†]	All ages	[2021-W42 – 2022-W15]	26 weeks	5	-2
Chile	Booster [†]	All ages	[2021-W42 – 2022-W15]	26 weeks	4	-2
Colombia	Booster [†]	All ages	[2021-W42 – 2022-W15]	26 weeks	5	-2
Ecuador	Booster [†]	All ages	[2021-W42 – 2022-W15]	26 weeks	4	-3
New Zealand	Booster [†]	All ages	[2021-W42 – 2022-W15]	26 weeks	5	-2
Peru	Booster [†]	All ages	[2021-W42 – 2022-W15]	26 weeks	3	-2
South Africa	Booster [†]	All ages	[2021-W42 – 2022-W15]	26 weeks	5	-2
Argentina	Booster [†]	All ages	[2021-M10 – 2022-M04]	7 months	5	-2
Bolivia	Booster [†]	All ages	[2021-M10 – 2022-M04]	7 months	5	-2
Brazil	Booster [†]	All ages	[2021-M10 – 2022-M04]	7 months	5	-2
Malaysia	Booster [†]	All ages	[2021-M10 – 2022-M04]	7 months	5	-2
Paraguay	Booster [†]	All ages	[2021-M10 – 2022-M04]	7 months	5	-2
Philippines	Booster [†]	All ages	[2021-M10 – 2022-M04]	7 months	5	-2
Singapore	Booster [†]	All ages	[2021-M10 – 2022-M04]	7 months	5	-2

Suriname	Booster [†]	All ages	[2021-M10 – 2022-M04]	7 months	5	-2
Thailand	Booster [†]	All ages	[2021-M10 – 2022-M04]	7 months	5	-2
Uruguay	Booster [†]	All ages	[2021-M10 – 2022-M04]	7 months	5	-2
Chile	Dose 3 rollout	90+	[2021-W32 – 2022-W01]	22 weeks	4	-2
Chile	Dose 3 rollout	85-89	[2021-W32 – 2022-W01]	22 weeks	4	-2
Chile	Dose 3 rollout	80-84	[2021-W32 – 2022-W01]	22 weeks	4	-2
Chile	Dose 3 rollout	75-79	[2021-W32 – 2022-W01]	22 weeks	4	-2
Chile	Dose 3 rollout	70-74	[2021-W32 – 2022-W01]	22 weeks	4	-2
Chile	Dose 3 rollout	65-69	[2021-W32 – 2022-W01]	22 weeks	4	-2
Chile	Dose 3 rollout	60-64	[2021-W32 – 2022-W01]	22 weeks	4	-2
Chile	Dose 4 rollout	90+	[2022-W02 – 2022-W15]	14 weeks	5	-2
Chile	Dose 4 rollout	85-89	[2022-W02 – 2022-W15]	14 weeks	5	-2
Chile	Dose 4 rollout	80-84	[2022-W02 – 2022-W15]	14 weeks	5	-2
Chile	Dose 4 rollout	75-79	[2022-W02 – 2022-W15]	14 weeks	5	-2
Chile	Dose 4 rollout	70-74	[2022-W02 – 2022-W15]	14 weeks	5	-2
Chile	Dose 4 rollout	65-69	[2022-W02 – 2022-W15]	14 weeks	5	-2
Chile	Dose 4 rollout	60-64	[2022-W02 – 2022-W15]	14 weeks	5	-2
Peru	Dose 3 rollout	90+	[2021-W42 – 2022-W15]	26 weeks	3	-2
Peru	Dose 3 rollout	85-89	[2021-W42 – 2022-W15]	26 weeks	3	-2
Peru	Dose 3 rollout	80-84	[2021-W42 – 2022-W15]	26 weeks	3	-2
Peru	Dose 3 rollout	75-79	[2021-W42 – 2022-W15]	26 weeks	3	-2
Peru	Dose 3 rollout	70-74	[2021-W42 – 2022-W15]	26 weeks	3	-2
Peru	Dose 3 rollout	65-69	[2021-W42 – 2022-W15]	26 weeks	3	-2
Peru	Dose 3 rollout	60-64	[2021-W42 – 2022-W15]	26 weeks	3	-2
Peru	Dose 4 rollout	90+	[2022-W16 – 2022-W40]	25 weeks	3	-3

Peru	Dose 4 rollout	85-89	[2022-W16 – 2022-W40]	25 weeks	3	-3
Peru	Dose 4 rollout	80-84	[2022-W16 – 2022-W40]	25 weeks	3	-3
Peru	Dose 4 rollout	75-79	[2022-W16 – 2022-W40]	25 weeks	3	-3
Peru	Dose 4 rollout	70-74	[2022-W16 – 2022-W40]	25 weeks	3	-3
Peru	Dose 4 rollout	65-69	[2022-W16 – 2022-W40]	25 weeks	3	-3
Peru	Dose 4 rollout	60-64	[2022-W16 – 2022-W40]	25 weeks	3	-3
New Zealand	Dose 3 rollout	All ages	[2021-W50 – 2022-W25]	28 weeks	6	-2
New Zealand	Dose 4 rollout	All ages	[2022-W26 – 2023-W04]	31 weeks	6	-3

[†] The “Booster” period of interest refers to the time period framing the nominal January–February 2022 peak in ACM, in the data for all ages.

* Format is the following: YEAR-WEEK or YEAR-MONTH. For example:

- [2021-W07 – 2023-W12] means that the period of interest is from the week 7 of 2021 to the week 12 of 2023, inclusive.
- [2021-M01 – 2023-M03] means that the period of interest is from the month 1 (January) of 2021 to the month 3 (March) of 2023, inclusive.

The definition of the week number, throughout the paper, is as given by ABS (2022): “The week number is based on the ISO (International Organization for Standardisation) week date system. In this system, weeks are defined as seven-day periods which start on a Monday. Week 1 of any given year is the week which starts on the Monday closest to 1 January, and for which the majority of its days fall in January (i.e. four days or more). Week 1 therefore always contains the 4th of January and always contains the first Thursday of the year. Using the ISO structure, some years (e.g. 2015 and 2020) contain 53 weeks.”

[‡] From the integration point indicated in the table, included, moving backwards. The last integration point (the integration point of interest) is 0. For example:

- 2 integration points used from the -1 means that 2 integration points were used to fit the trend line: the integration point (-1), directly preceding the last one (0, the one of interest), and one extra integration point (-2) prior to that.

- 6 integration points used from the -3 means that 6 integration points were used to fit the trend line: from the -3 through the -8 integration point, included, where 0 is the integration point of interest (the most recent one).

This applies to the method using several integration points from the historic period (Method 1). Chile (all ages) and Peru (all ages) data were insufficient to apply Method 1.

References of Appendix C

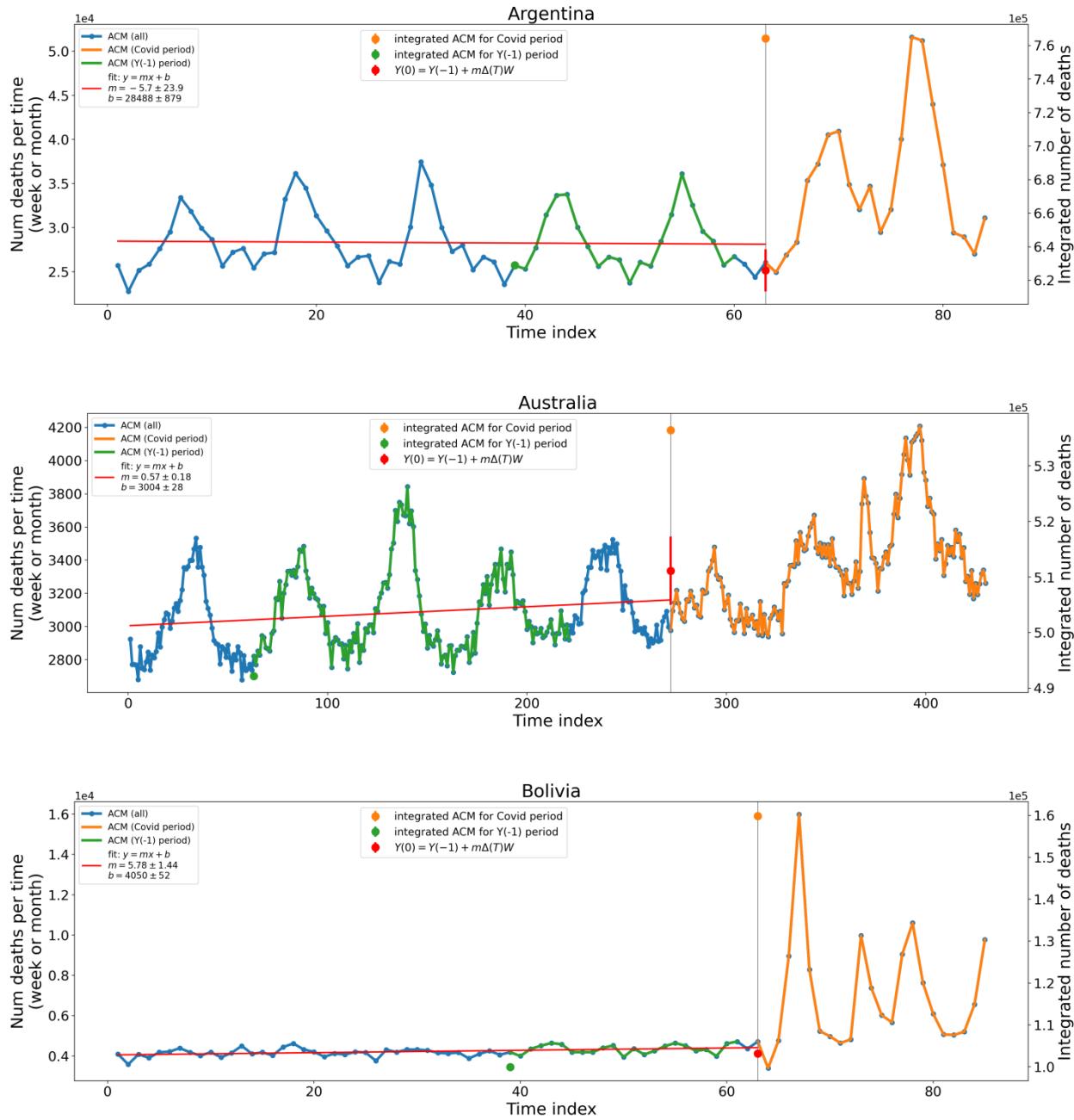
ABS (2022): Australian Bureau of Statistics /// Provisional Mortality Statistics /// files "Provisional Mortality Statistics, Weekly Dashboard, Jan - Sep 2022" and "Deaths by week of occurrence, 2015-21" /// accessed 23 January

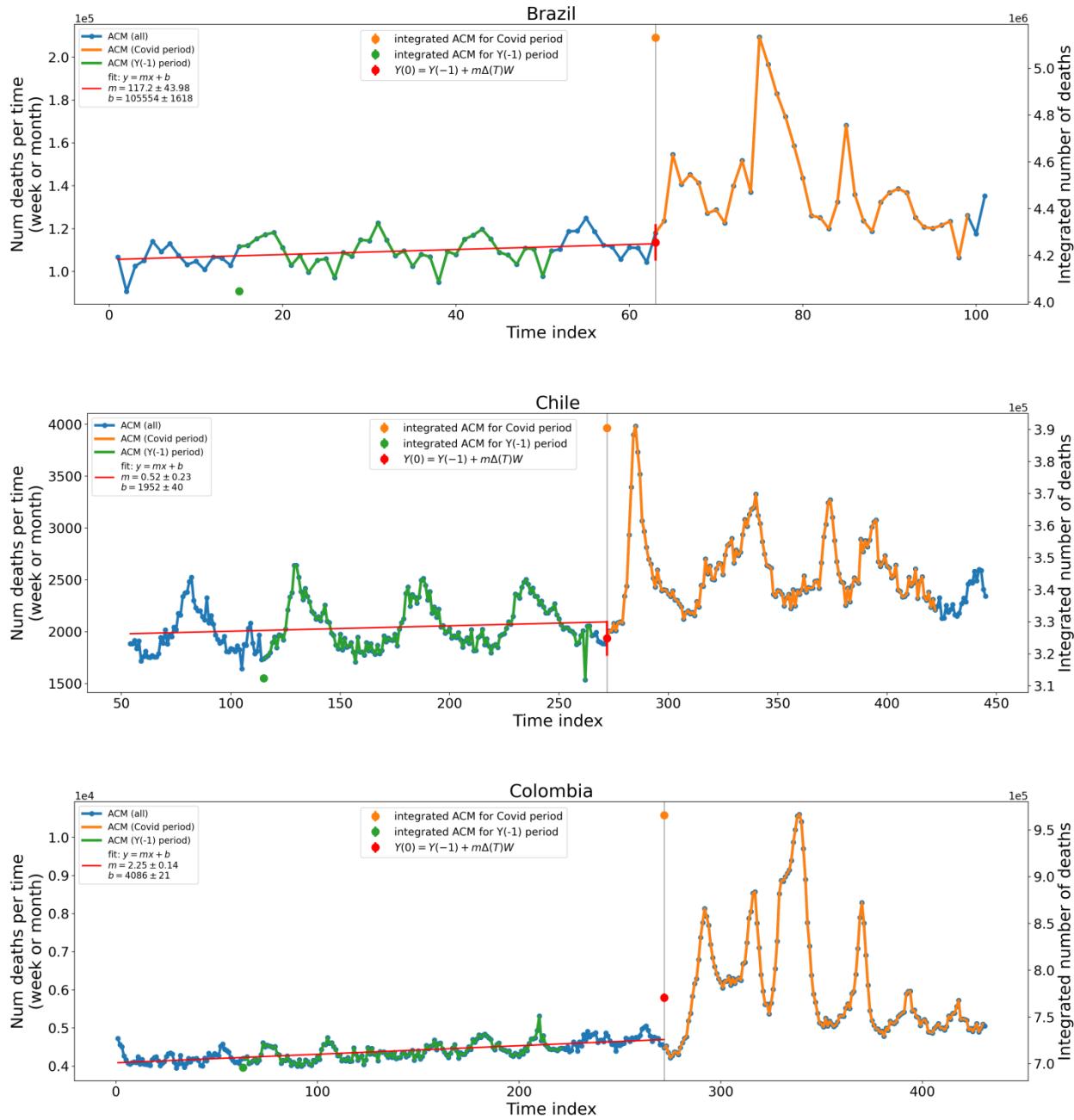
2023 <https://www.abs.gov.au/statistics/health/causes-death/provisional-mortality-statistics/latest-release>

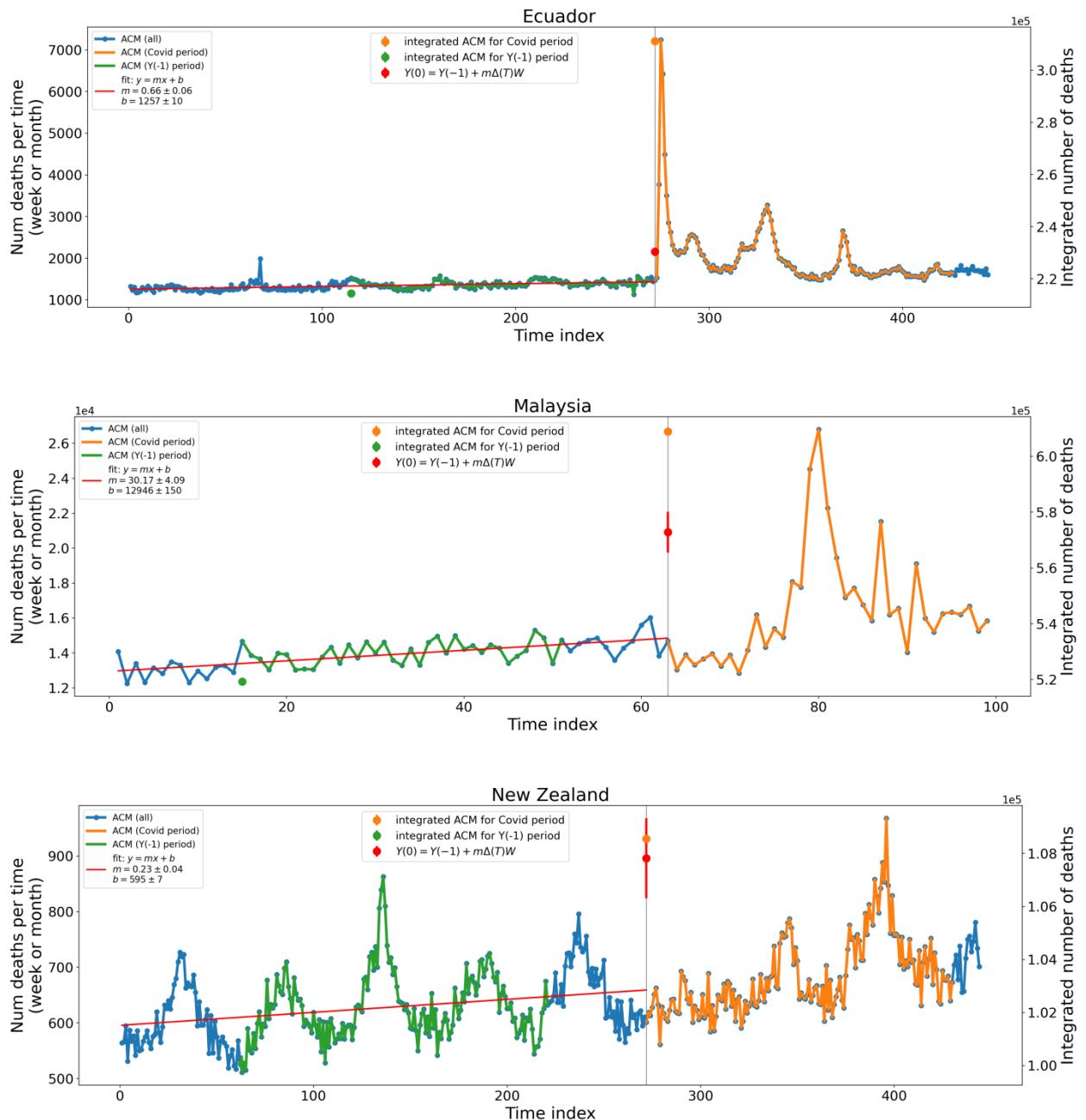
Appendix D: Single-point method quantification of the excess mortality in the Covid period

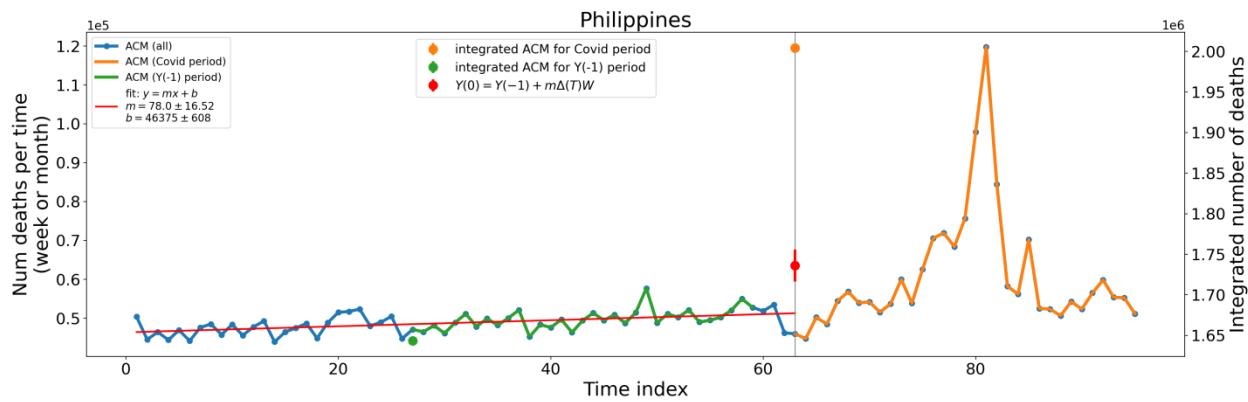
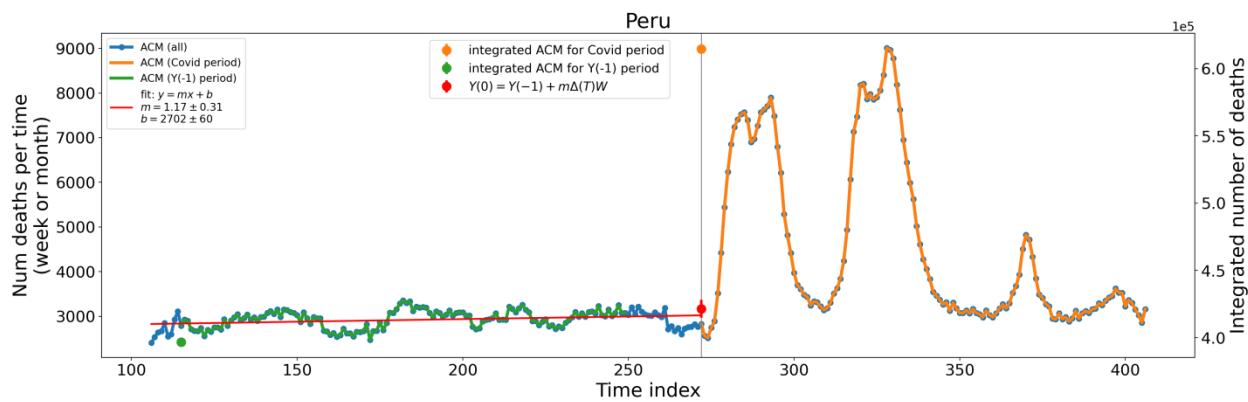
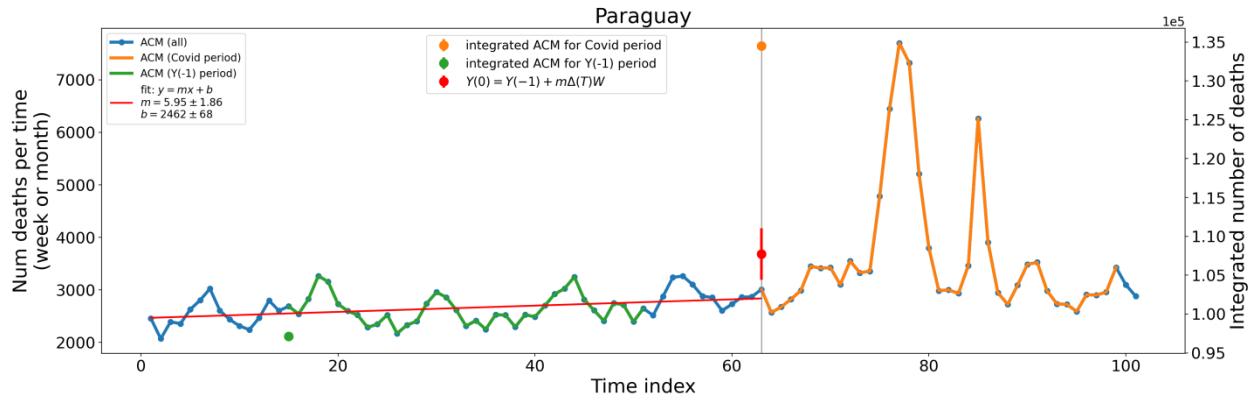
The following Figure 22 illustrates the application of our “Special case of a single historic integrated point” method (Method 2) to quantify the excess mortality in the nominally entire Covid period, from 11 March 2020 when a pandemic was declared to a practical end date taken to be the same as the end date of the nominal vaccination period (which depends on mortality and vaccination data availability), for the 17 countries in this study. The bounds of integration for the periods for each country are specified in Appendix C: Technical and specific information for applications of the methods to the data.

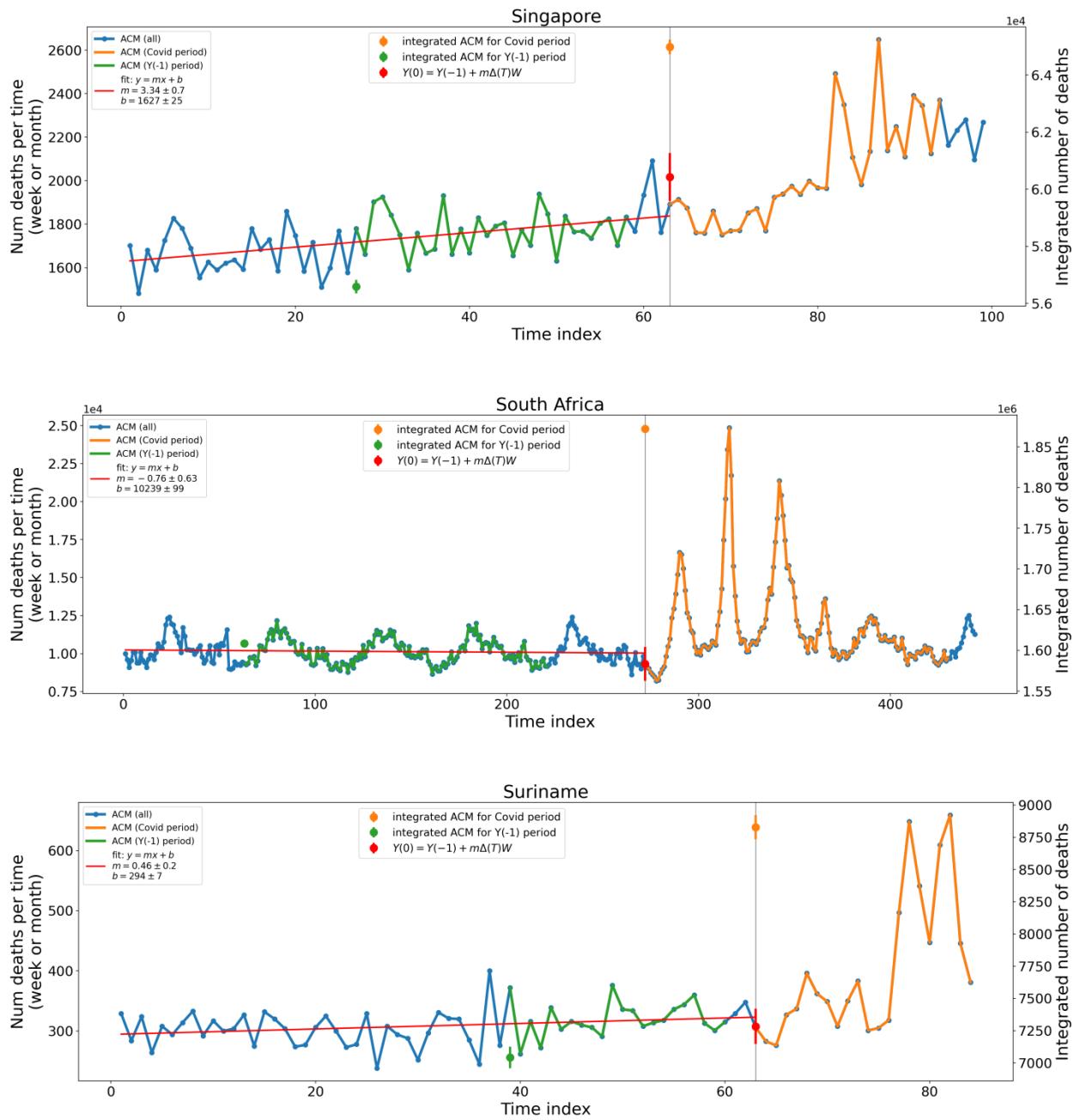
Here, as per the legend, the Covid period is in orange (both its ACM by time and its integration point), the single historic reference period used is in green (both its ACM by time and its integration point), the best-line fit (least squares fit) of the historic ACM by time is in red, and the extrapolated predicted value of the integration for the Covid period is also in red, with its error bar. The weeks or months are counted consecutively in sequence using a time index, where week number 1 is week-1 of 2015, and month number 1 is month-1 (January) of 2015. Week-1 for 2015 contains the dates 29 December 2014 through 4 January 2015. The thin vertical grey line, in each panel, indicates the date (11 March 2020) at which a pandemic was declared by the World Health Organization (WHO).











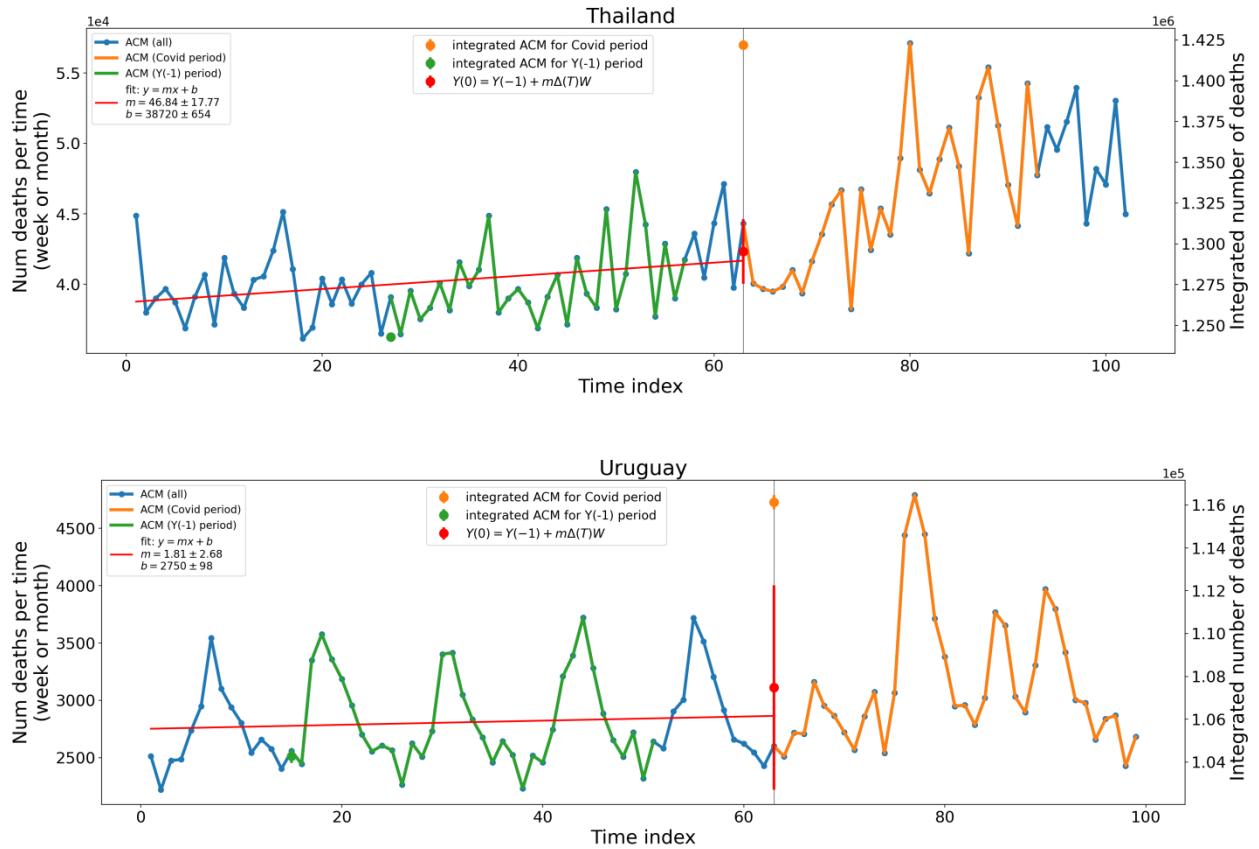
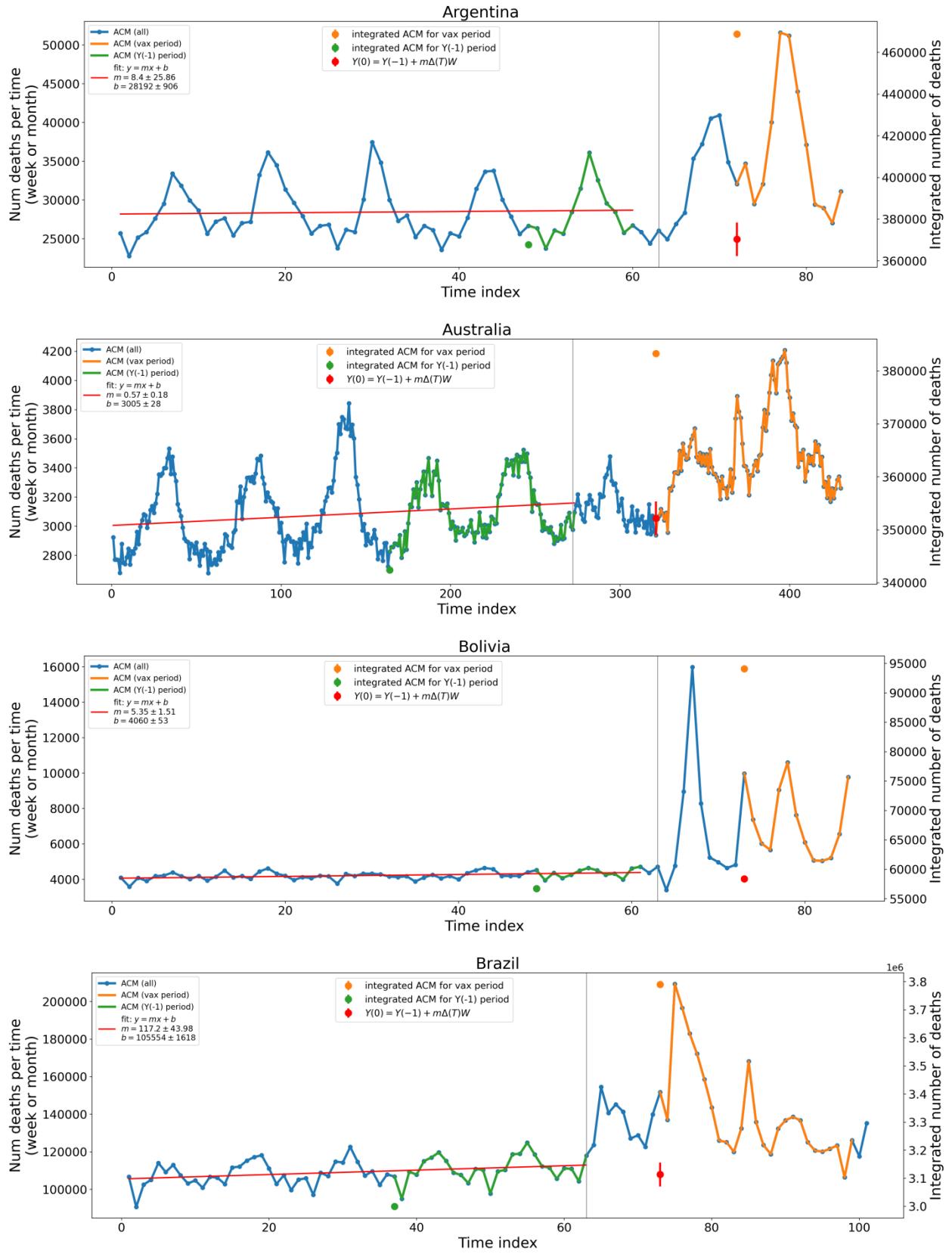


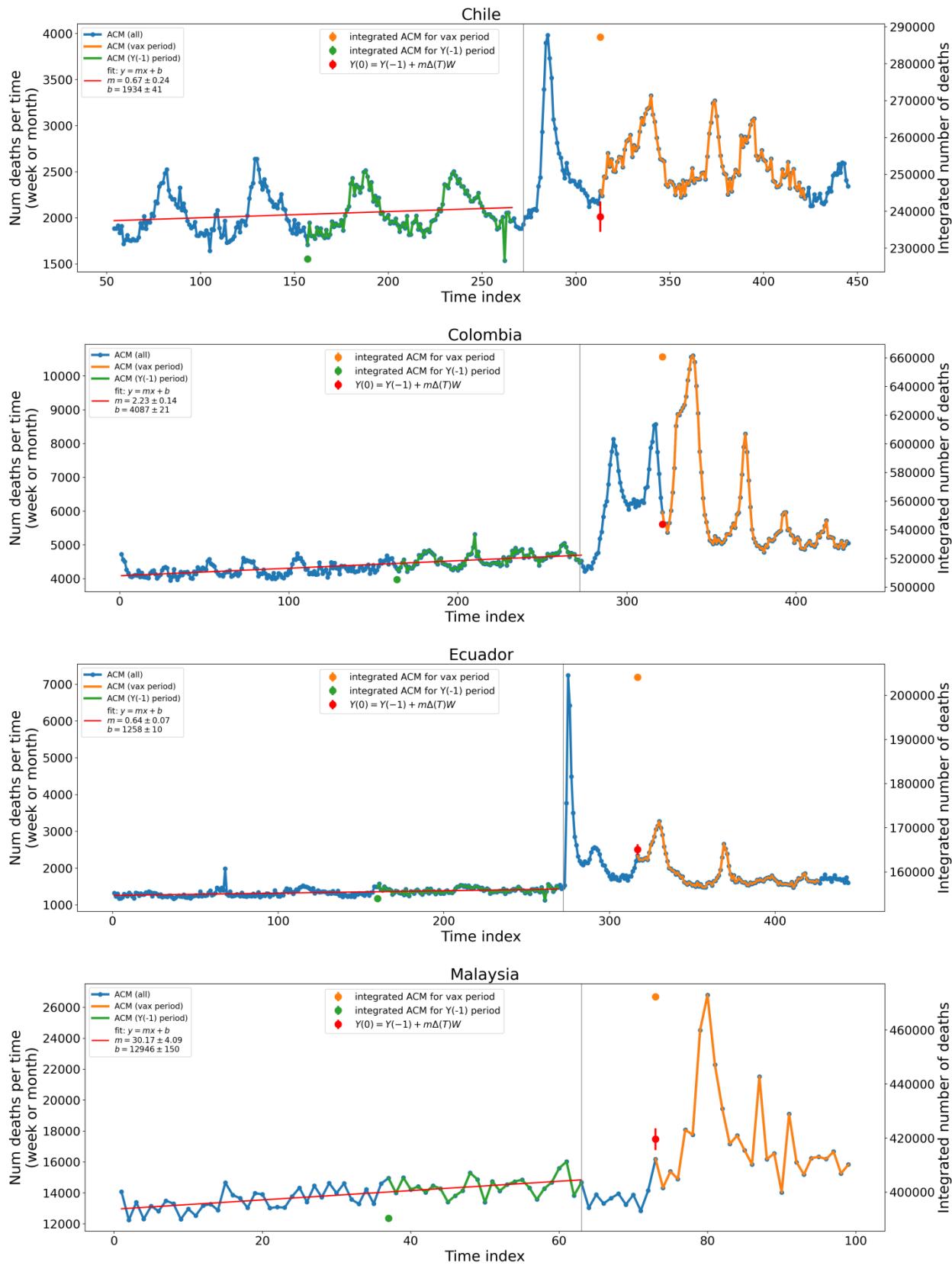
Figure 22: Appendix D - Single historic point method (Method 2) for excess ACM in the Covid periods of the 17 countries in this study. Covid period (orange), single-point historic reference period (green), best-line fit (red). The time index is the week or month number from start of 2015. The 11 March 2020 pandemic declaration date is shown by a vertical grey line. The data sources are specified in Appendix A.

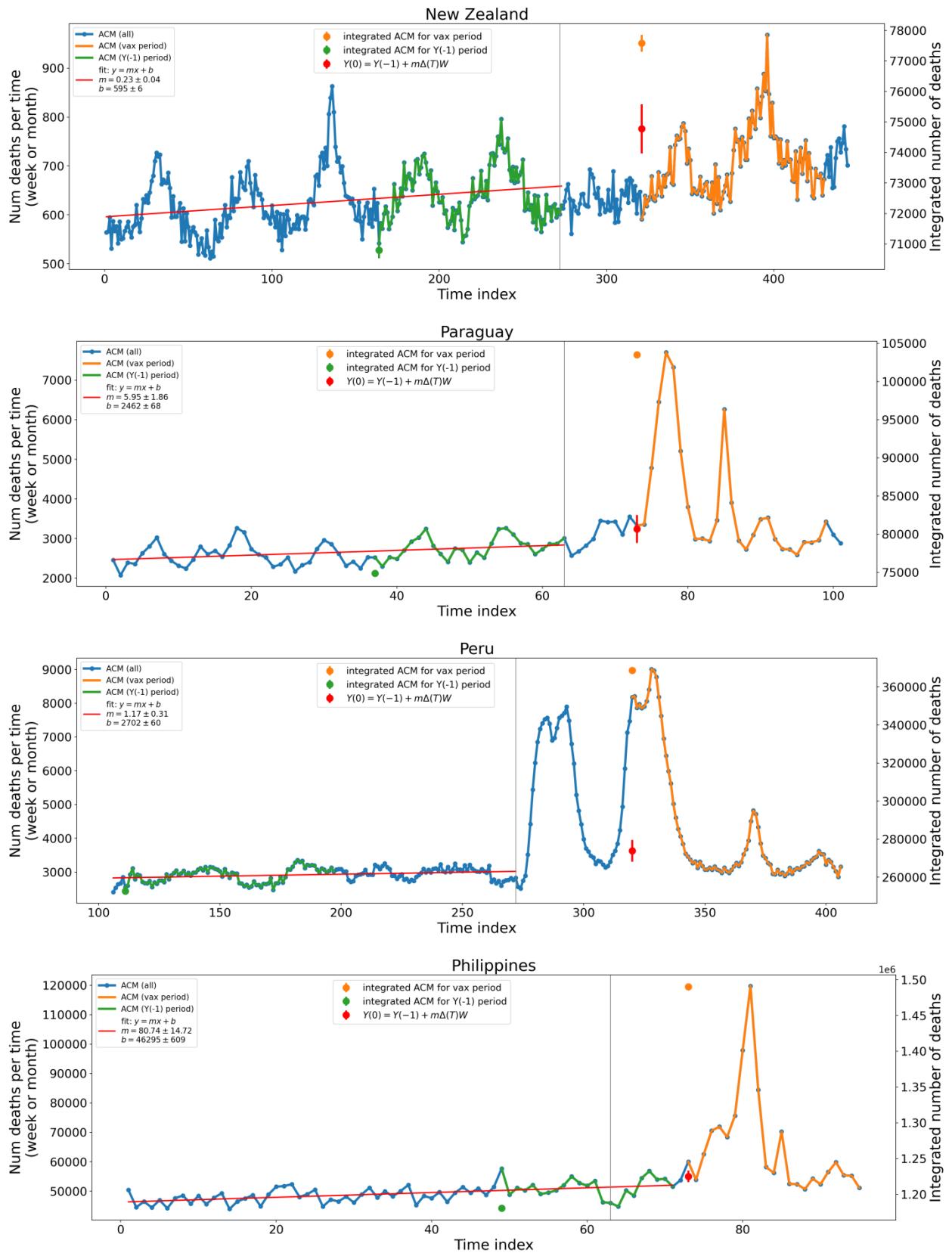
Appendix E: Single-point method quantification of the excess mortality in the vaccination period

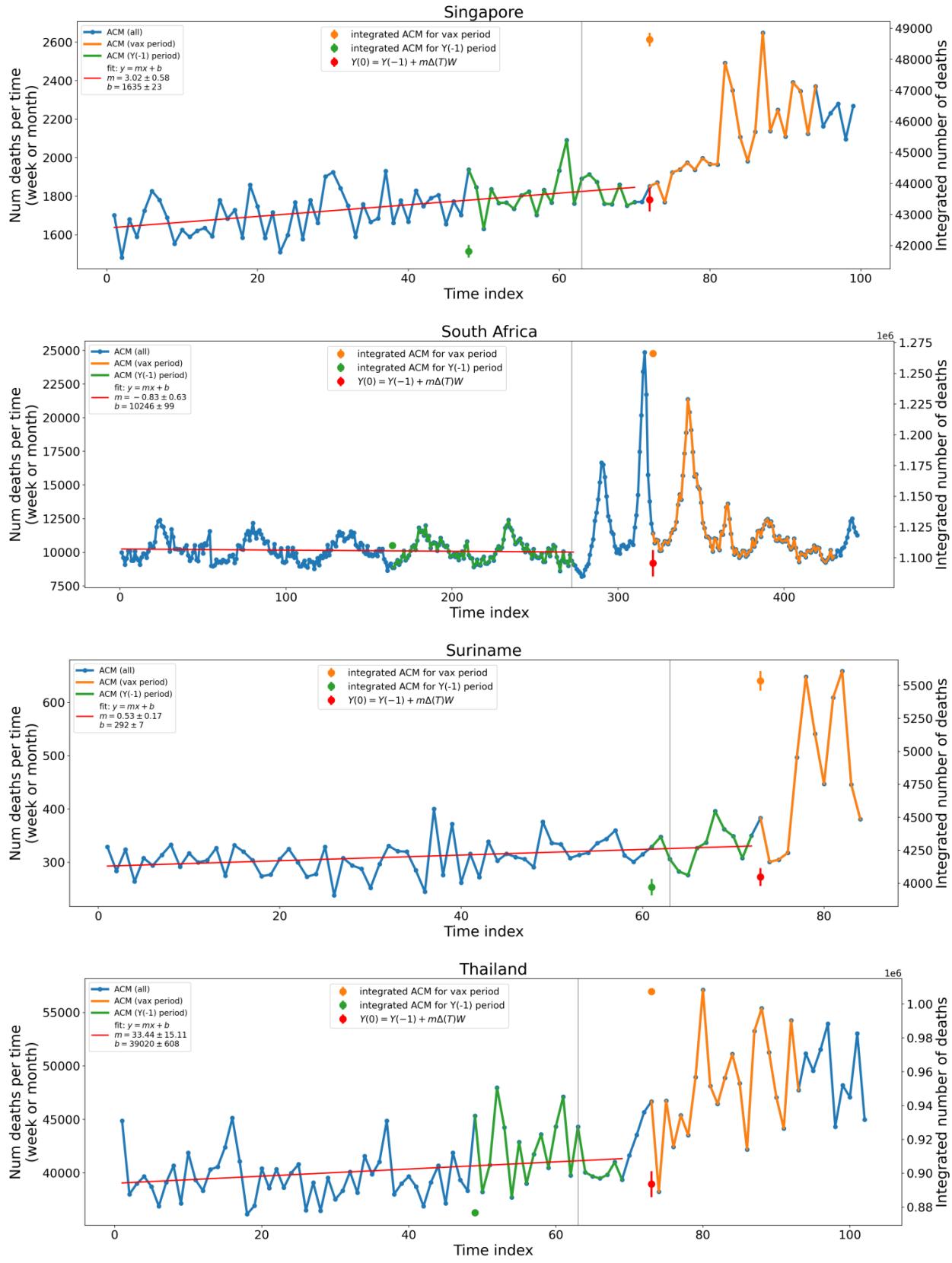
The following Figure 23 illustrates the application of our “Special case of a single historic integrated point” method (Method 2) to the same limits of integration, for the vaccination period, as for Method 1 (historic trend method, shown in Figure 6 and Figure 7), for the 17 countries in this study. The limits of integration are specified in Appendix C.

Here, as per the legend, the vaccination period is in orange (both its ACM by time and its integration point), the single historic reference period used is in green (both its ACM by time and its integration point), the best-line fit (least squares fit) of the historic ACM by time is in red, and the extrapolated predicted value of the integration for the vaccination period is also in red, with its error bar. The weeks or months are counted in sequence using a time index, where week or month number 1 is the week or month containing 1 January 2015. Week-1 for 2015 contains the dates 29 December 2014 through 4 January 2015. The thin vertical grey line, in each panel, indicates the date (11 March 2020) at which a pandemic was declared by the World Health Organization (WHO).









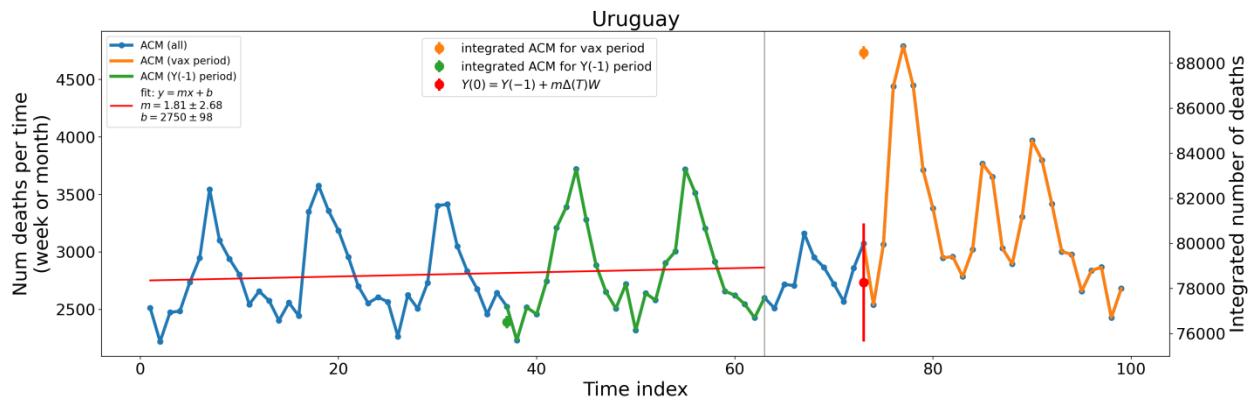


Figure 23: Appendix E - Single historic point method for excess ACM (Method 2) in the vaccination periods of the 17 countries in this study. Vaccination period (orange), single-point historic reference period (green), best-line fit (red). The time index is the week or month number from start of 2015. The 11 March 2020 pandemic declaration date is shown by a vertical grey line. The data sources are specified in Appendix A.