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Part One

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1. Linear Regression

- projection
- orthongonal decomposition
- Gaussian Linear Regression
- prediction (generally of \hat{y})
- different types of errors
- influence
- lack of fit
- \bullet R^2
- Multicollinearity

1.1 Projection in Euclidean Space

Monday August 22

Definition 1.1.1 — Euclidean Space. One way to think of the Euclidean plane is as a set of points satisfying certain relationships, expressible in terms of distance and angle. **Euclidean space** is an abstraction detached from actual physical locations, specific reference frames, measurement instruments, and so on.

Let Euclidian Space be denoted by \mathbb{R}^{P} .

$$\mathbb{R}X \dots X\mathbb{R} = \{(x_1, \dots, x_p) : x_1 \in \mathbb{R} \dots, x_p \in \mathbb{R}^P\}$$

Definition 1.1.2 — **Inner Product.** In linear algebra, an inner product space is a vector space with an additional structure called an inner product. This additional structure associates each pair of vectors in the space with a scalar quantity known as the inner product of the vectors. **Inner products** allow the rigorous introduction of intuitive geometrical notions such as the length of a vector or the angle between two vectors. They also provide the means of defining orthogonality between vectors (zero inner product).

Let $a \in \mathbb{R}^P$, $b \in \mathbb{R}^P$

$$a^T b = \sum_{i=1}^P a_i b_i$$

$$a^T b = \langle a, b \rangle$$

Definition 1.1.3 — **Hilbert Space**. The mathematical concept of a Hilbert space generalizes the notion of Euclidean space. It extends the methods of vector algebra and calculus from the two-dimensional Euclidean plane and three-dimensional space to spaces with any finite or infinite number of dimensions. A Hilbert space is an abstract vector space possessing the structure of an inner product that allows length and angle to be measured. Furthermore, Hilbert spaces are complete: there are enough limits in the space to allow the techniques of calculus to be used.

Hilbert Inner Product Space $\{\mathbb{R}^P, \langle a, b \rangle\}$

General Inner Product

Let $\Sigma \in \mathbb{R}^{P_X P}$ set of all $P_X P$ matrices. Assume Σ is a positive definite matrix.

$$x^T \Sigma x < 0$$
$$\forall x \in \mathbb{R}^P$$

 $x \neq 0$

Then $a^T \Sigma b$ also satisfies the conditions for inner product.

$$a^T \Sigma b = \langle a, b \rangle_{\Sigma}$$

$$a^T b = a^T I b = \langle a, b \rangle_I$$

 $\{\mathbb{R}^P, <, >_{\Sigma}\}$ is a more general inner product space.

Linear Transformation

A matrix, $A \in \mathbb{R}^{PxP}$ can be viewed as linear transformation $T_A : \mathbb{R}^P \to \mathbb{R}^P, x \mapsto Ax$



Bing Li will denote T_A as A.

- \rightarrow means maps to for a domain.
- \mapsto means maps to for a value.
- \Rightarrow means implies.

If $A: \mathbb{R}^P \to \mathbb{R}^P$.

$$ker(A) = \{x \in \mathbb{R}^P, Ax = 0\}$$

 $ran(A) = \{Ax : x \in \mathbb{R}^P\}$

Definition 1.1.4 — Kernel. In linear algebra, the kernel, or sometimes the null space, is the set of all elements v of V for which L(v) = 0, where 0 denotes the zero vector in W.

In coordinate plane, think of a function that crosses the x-axis. The kernel would be all points on x where y = 0.

Definition 1.1.5 — Range. In coordinate plane, how much of the y axis is reached with the function? Now extend this idea to more dimensions.

A linear transformation is **idempotent** if

$$A = A^{2}$$
$$Ax = A(A(x))$$
$$\forall x \in \mathbb{R}^{P}$$

If A were a number it could only be 1 or 0.

Wednesday August 24

Let $T \in \mathbb{R}^{PxP}$ then there exists a unique operator $R \in \mathbb{R}^{PxP}$ such that $\forall x, y \in \mathbb{R}^{P}$,

$$\langle x, Ty \rangle = \langle Rx, y \rangle$$

(general inner product, $a^T \Sigma b$). Aside: What this states is that if you give me any operator in the first you can find one in the second.

R is called the **adjoint operator** of T. Written as T^* , that is,

$$\langle x, Ty \rangle = \langle T^*x, y \rangle$$

Derived Facts

$$< x, Ty > = < T^*, y >$$

= $< y, T^*x >$
= $< (T^*)^*y, x >$
= $< x, (T^*)^*y >$

(by the definition)
(inner products the order doesn't matter)
(Use the definition again)
(swap order)

So,
$$T = (T^*)^*$$
.

It is easy to see in our case

$$\langle x, Ty \rangle_{\Sigma} = x^{T} \Sigma Ty$$

$$= x^{T} \Sigma T \Sigma^{-1} \Sigma y$$

$$= (\Sigma^{-1} T^{T} \Sigma x)^{T} \Sigma y$$

$$= \langle \Sigma^{-1} T^{T} \Sigma x, y \rangle_{\Sigma}$$

So, $T^* = \Sigma^{-1}T^T\Sigma$ when $\Sigma = I_P$ (identity) and $T^* = T^T$.

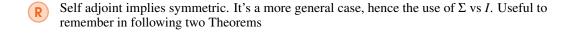
Derived Facts

An operator is **self adjoint** if its adjoint is itself. (i.e. if $T = T^*$ or $\langle x, Ty \rangle = \langle Tx, y \rangle$). In the case of $<,>_{\Sigma}$,

$$T = \Sigma^{-1}T^T\Sigma$$

if

$$\Sigma = I_P$$
, $T = T^T$



Theorem 1.1.1 If $A \in \mathbb{R}^{PxP}$ is symmetric, then there exists **eigenvalue-eigenvector pairs**. $(\lambda_1, v_1), \dots (\lambda_P, v_P)$ such that $v_1 \perp \dots \perp v_P$. Orthoginal basis (ONB) such that

$$A = \sum_{i=1}^{P} \lambda_i v_i v_i^T \text{(spectral decomposition)}$$

More generally, if A is a linear operator in \mathcal{H} (finite dimential inner product such as $(\mathbb{R}^P,<,>_{\Sigma})$). its eigen pair (linear operator now) (λ,ν) is defined by

$$\begin{cases} Av = \lambda v \\ < v, v >= 1 \end{cases}$$

Definition 1.1.6 — Orthogonal Basis. In the following, $(\mathbb{R}^P, <, >_{\Sigma}) = \mathcal{H}$ (H for Hilbert) ONB is defined by:

- 1. $v_i \perp v_j, \langle v_i, v_j \rangle = 0$ 2. $||v_i|| = 1$ 3. $span\{v_1, \dots, v_P\} = \mathcal{H}$

Theorem 1.1.2 Suppose $A: \mathcal{H} \to \mathcal{H}$ is a self adjoint linear operator. Then A has eigen pairs: $(\lambda_1, \nu_1, \dots, (\lambda_P, \nu_P))$ where $\{\nu_1, \dots, \nu_P\}$ is ONB of \mathbb{R} such that

$$A = \sum_{i=1}^{P} \lambda_i v_i v_i^T \Sigma$$

Proof. (λ, v) is eigen pair of A, which means

$$Av = \lambda v$$

$$< v, v > = 1$$

$$v^T \Sigma v = 1$$

Let $u = \sum_{i=1}^{n} v_i$.

Aside: $\Sigma^{\alpha} = \Sigma \lambda_i^{\alpha} v_i v_i^T$

Let
$$v = \Sigma^{-\frac{1}{2}}u$$
.

$$A\Sigma^{-\frac{1}{2}}u = \lambda\Sigma^{-\frac{1}{2}}u$$
$$\Sigma^{-\frac{1}{2}}u = \lambda u$$

So, (λ, ν) is an eigen pair of A in $(\mathbb{R}, <, >_{\Sigma}) \Leftrightarrow (\lambda, u)$ '...' of $\Sigma^{\frac{1}{2}} A \Sigma^{-\frac{1}{2}}$ in $(\mathbb{R}, <, >_{I})$. Note that, A is self adjoint in $(\mathbb{R}, <, >_{\Sigma})$. So, $A = \Sigma^{-1} A^{T} \Sigma$

$$\begin{split} \Sigma^{\frac{1}{2}} A \Sigma^{-\frac{1}{2}} &= \Sigma^{\frac{1}{2}} A^T \Sigma \Sigma^{-\frac{1}{2}} \\ &= \Sigma^{-\frac{1}{2}} A \Sigma^{\frac{1}{2}} \\ &= (\Sigma^{\frac{1}{2}} A \Sigma^{-\frac{1}{2}})^T \end{split}$$

Note: $\Sigma^{\frac{1}{2}}A\Sigma^{-\frac{1}{2}}$ is symmetric!! So by Theorem 1.1, $\Sigma^{\frac{1}{2}}A\Sigma^{-\frac{1}{2}} = \sum \lambda_i v_i v_i^T$ where (λ_i, v_i) eigenpairs of $\Sigma^{\frac{1}{2}}A\Sigma^{-\frac{1}{2}}$.

That means $(\lambda_i, \Sigma^{\frac{1}{2}} v_i)$ are eigen pairs of A.

So,
$$\Sigma^{\frac{1}{2}}A\Sigma^{-\frac{1}{2}} = \sum_{i=1}^{P} \Sigma^{\frac{1}{2}} u_i u_i^T \Sigma^{\frac{1}{2}} \Rightarrow A = \sum_{i=1}^{P} \lambda u_i u_i^T \Sigma$$

Definition 1.1.7 — Projection. If P is an operator in $(\mathbb{R}^P, <, >)$ then P is called a **projection** if it is both idempotent $(P = P^2)$ and self adjoint $(P = P^*)$.

Preposition 1.1 If *A* is a linear operator then $ker(A) = ran(A^*)^{\perp}$

Proof. Take
$$x \in ker(A) (\Rightarrow Ax = 0)$$
.
 $\forall y \in ran(A^*), x \perp y$
 $\Rightarrow x \perp y \forall y = A^*z, z \in \mathbb{R}^P$
Hence,

$$\langle x, y \rangle = \langle x, A^*z \rangle$$

$$= \langle Ax, z \rangle$$

$$= \langle 0, z \rangle$$

$$= 0$$

$$\Rightarrow x \perp y$$

$$\Rightarrow x \in ran(A^*)^{\perp}$$

Or vice versa.

Friday August 26

ightharpoonup \perp means orthogonal complement.

$$\mathscr{S}^{\perp} = \{ v \in \mathbb{R}^P, v \perp \mathscr{S} \}$$

$$v \perp w \forall w \in \mathscr{S}$$

$$< v, w > = 0 \forall w \in \mathcal{S}$$

= $\{v \in \mathbb{R}^P, < v, w > = 0 \forall w \in \mathcal{S}\}$

Recall,
$$ker(A) = ran(A^*)^{\perp}$$

So, if A is slef adjoint then this is true and ran(A) is also span(A) which is the subspace spanned all columns of A.

Theorem 1.1.3 If P is a projection, then

- 1. $Pv = v, \forall v \in ran(P)$
- 2. Pv = 0, $\forall v \perp ran(P)$
- 3. If Q is another projections such that the ran(Q) = ran(P) then Q = P. (The range determines the operator, because it is what decomposes the operator.)

Asside: P acts like one on some spaces, and zero on orthogonal space.

Proof. 1. Let
$$v \in ran(P)$$
. Since $P^2 = P$ (idempotent) then
$$P^2v = Pv$$

$$\Rightarrow P^2v - PV = 0$$

$$\Rightarrow P(Pv - v) = 0$$

$$\Rightarrow Pv - v \in ker(P)$$

$$\Rightarrow Pv - v \perp ran(P)$$

$$\Rightarrow < Pv - v, Pv - v >= 0$$

$$\Rightarrow ||Pv - v|| = 0$$

$$\Rightarrow Pv - v = 0$$

$$\Rightarrow Pv = v$$
2. If
$$v \perp ran(P)$$

$$\Rightarrow v \in ker(P)$$

$$\Rightarrow Pv = 0$$
3. If Q is another operator with $ran(Q) = ran(P) = \mathscr{S}$ then $\forall v \in \mathscr{S}$

$$Qv = v = Pf(\forall v \perp \mathscr{S})$$

$$Qv = 0 = Pv$$

$$Qv = Pv \forall, v \in \mathscr{S}$$

$$Q = P$$

Theorem 1.1.4 Suppose \mathscr{S} is a subspace of \mathbb{R}^P , R v_1, \ldots, v_m is a basis of \mathscr{S} .

Let
$$v = (v_1, \ldots, v + m) \in \mathbb{R}^{xM}$$
.

Then,

1. $A = v(v^T \Sigma v)^{-1} v^T \Sigma$ is a projection.

2. $ran(A) = \mathcal{S}$

Proof. 1. idempotent.
$$A^{2} = v(v^{T}\Sigma v)^{-1}v^{t}\Sigma v(v^{T}\Sigma v)^{-1}v^{T}\Sigma$$

$$= v(v^{T}\Sigma v)^{-1}v^{T}\Sigma$$

$$= A$$

2. Self adjoint.

Let
$$x, y \in \mathbb{R}^P$$

 $\langle x, Ay \rangle = x^T \sum v (v^T \sum v)^{-1} v^{\Sigma} y$
 $= (v(v^T \sum v)^{-1} v^T \sum x)^T \sum y$
 $= \langle Ax, y \rangle$

3. $ran(A) = \mathcal{S}$?

Let $x \in \mathbb{R}^P$.

$$Ax = v(v^T \Sigma v)^{-1} v^T \Sigma x \in span(v) = \mathscr{S}$$

So let $x \in \mathcal{S}$,

$$x \in ran(v)$$

$$x = vy$$

for some $y \in \mathbb{R}^P$

$$= v(v^T \Sigma v)^{-1} v^T \Sigma v y$$

 $\in ran(A)$

So, $\mathscr{S} \subseteq ran(A)$ and then $\mathscr{S} = ran(A)$.

We write *A* as $P_{\mathscr{S}}(\Sigma)$ (orthogonal projection on to \mathscr{S} with respect to Σ - product).

In the following, let $I : \mathbb{R}^P \to \mathbb{R}^P$ be the identity mapping. $(x \mapsto x)$ Let \mathscr{S} be a subspace in \mathbb{R}^P .

Let
$$Q_{\mathcal{S}}(\Sigma) = I - P_{\mathcal{S}}(\Sigma)$$

Proprosition 1.2
$$Q_{\mathscr{S}}(\Sigma) = P_{\mathscr{S}^{\perp}}(\Sigma)$$

Proof. Show $Q_{\mathcal{S}}(\Sigma)$ is projection.

1. Idempotent

$$\begin{aligned} Q_{\mathscr{S}}^{2}(\Sigma) &= Q_{\mathscr{S}}(\Sigma)Q_{\mathscr{S}}(\Sigma) \\ &= (I - P_{\mathscr{S}}(\Sigma))(I - P_{\mathscr{S}}(\Sigma)) \\ &= I - P_{\mathscr{S}}(\Sigma) - P_{\mathscr{S}}(\Sigma) + P_{\mathscr{S}}P_{\mathscr{S}} \\ &= Q_{\mathscr{S}}(\Sigma) \end{aligned}$$

2. Self-adjoint

$$x, y \in \mathbb{R}^P$$

3. Range

$$ran(Q_{\mathscr{S}}(\Sigma)) = \mathscr{S}^{\perp}$$
. Take $x \perp \mathscr{S} = ran(P_{\mathscr{S}}(\Sigma))^{\perp} = ker(P_{\mathscr{S}}(\Sigma))$.

$$\Rightarrow P_{\mathscr{S}}(\Sigma) = 0$$

$$\Rightarrow Q_{\mathscr{S}}(\Sigma)x = x - P_{\mathscr{S}}(\Sigma)x = x$$

$$X \in ran(Q_{\mathscr{S}}(\Sigma))$$

$$\Rightarrow \mathscr{S}^{\perp} \subseteq ran(Q_{\mathscr{S}}(\Sigma))$$
Take $x \in ran(Q_{\mathscr{S}}(\Sigma))$, $\forall y \in \mathscr{S} = ran(P_{\mathscr{S}}(\Sigma))$

$$y = P_{\mathscr{S}}(\Sigma)z \text{ for some } z \in \mathbb{R}^{P}$$

$$< x, y > = < x, P_{\mathscr{S}}(\Sigma)z > = < P_{\mathscr{S}}(\Sigma)x, z > = 0$$

$$\Rightarrow x \in \mathscr{S}^{\perp}$$

$$\Rightarrow ran(Q_{\mathscr{S}}(\Sigma)) = \mathscr{S}^{\perp}$$

1.2 Cochran's Theorem

This section will be about the distribution of the squared norm of a projection of a Gaussian random vector.

Preposition 1.3 If A is idempotent, then its eigenvalues are either 0 or 1.

Proof. λ is eigenvalue of A.

$$\Rightarrow Av = \lambda v(||v|| = 1)$$

$$\lambda = Av = A^2v = \lambda Av = \lambda^2$$

So, λ is 0 or 1.

Monday August 29

Lemma 1.1 Suppose $V \sim N(0, \sigma^2 I_P)$.

P is projection with I_P - inner product. Then $V^T P V \sim \sigma^2 \chi_S^2$ where df = rank(P).

Proof. P is symmetric, and it has spectral decomposisition,

$$ARA^{T}$$

where the A's are orthogonal and R is diagonal with diagonal entries 0 or 1.

Then,

$$A^T V \sim N_P(0, A^T(\sigma^2 I_P)A) = N_P(0, \sigma I_P)$$

Let,

$$Z = RA^T V$$

then,

$$Z \sim N_P(0, \sigma^2 R^2) = N_P(0, \sigma^2 R)$$

That means among the components of Z, some are distributied as N(0, 1) and the rest are zero and they are independant. So,

$$Z^T Z \sim \chi_S^2 = V^T P V$$

Corollary 1.1 Suppose $X \sim N(0, \Sigma)$. Consider the Hilbert space $(\mathbb{R}^P, <, >_{\Sigma^{-1}})$.

$$\langle a,b\rangle_{\Sigma^{-1}}=a^T\Sigma^{-1}b$$

Let $\mathscr S$ be a subspace of $\mathbb R^P$ and $P_{\mathscr S}(\sigma^{-1})$ be the projection onto $\mathscr S$ with respect to $<,>_{\Sigma}^{-1}$ (special case of Fisher information inner product)

Then,

$$||P_{\mathscr{S}}(\Sigma^{-1})x||_{\Sigma^{-1}}^2 \sim \chi_r^2$$

where $r = dim(\mathcal{S})$.

Proof. Let V be a basis matrix of \mathscr{S} (i.e. the col of V form basis in \mathscr{S}).

$$\begin{aligned} ||P_{\mathscr{S}}(\Sigma^{-1})x||_{\Sigma^{-1}}^{2} &= < P_{\mathscr{S}}(\Sigma^{-1})x, P_{\mathscr{S}}(\Sigma^{-1})x > \\ &= x^{T} P_{\mathscr{S}}(\Sigma^{-1}) \Sigma^{-1} P_{\mathscr{S}}(\Sigma^{-1})x \\ &= x^{T} (V(V^{T}\Sigma^{-1}v)^{-1}v^{T}\Sigma^{-1})^{T} \Sigma^{-1} (V(V^{T}\Sigma^{-1}v)^{-1}v^{T}\Sigma^{-1})^{T}x \\ &= x^{T} \Sigma^{-1} V(V^{T}\Sigma^{-1}v)^{-1}v^{T}\Sigma^{-1}V(V^{T}\Sigma^{-1}v)^{-1}v^{T}\Sigma^{-1})^{T}x \\ &= (\Sigma^{-\frac{1}{2}}x)^{T} || \end{aligned}$$

But,

$$\Sigma^{-\frac{1}{2}}x \sim N(0, I_P)$$

$$\Sigma^{-\frac{1}{2}}V(V^T\Sigma^{-1}v)^{-1}v^T\Sigma^{-\frac{1}{2}}$$

is a projection with repect to I_P -inner producted (idempotent, self adjoint, YES). By Lemme 1.1, $\sim \chi_r^2$.

It is then easy to derive Cocharan's Theorem. (see proof in Homework 1)

Theorem 1.2.1 Let $X \sim N(0, \Sigma)$ and $\mathcal{H} = \{\mathbb{R}^P, <, >_{\Sigma^{-1}}\}$. Let $\mathcal{S}_1, dots, \mathcal{S}_k$ be linear subspaces of \mathbb{R}^P such that $\mathcal{S}_i \perp \mathcal{S}_j$ in $<, >_{\Sigma^{-1}}$

Let $r_i = dim(\mathcal{S}_i)$.

Let
$$w_i = ||P_{\mathcal{S}_i}(\Sigma^{-1})X||_{\Sigma^{-1}}^2$$

Then,

- 1. $W_i \sim \chi_{r_i}^2$
- 2. $W_1 \perp \!\!\! \perp, \dots, \perp \!\!\! \perp W_k$ where $\perp \!\!\! \perp$ indicates independence.

1.3 Gaussian Linear Regresson Model

$$Y = \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix}$$

$$X = \begin{pmatrix} x_{11} & \dots & x_{1P} \\ \vdots & \ddots & \vdots \\ x_{n1} & \dots & x_{np} \end{pmatrix} \in \mathbb{R}^{nxp}$$

Consider the linear model,

$$y = X\beta + \varepsilon$$

$$\varepsilon \sim N(0, \sigma^2 I_n)$$

where X has full comlumn rank $(n \ge p)$.

Here X is treated as fixed

• Maximum Likelihood Estimator

$$E(y) = X\beta \in \mathbb{R}^n$$

$$Var(y) = \sigma^2 I_n$$

$$y \sim N_p(X\beta, \sigma^2 I_n)$$

Multivariate normal density

$$y \sim N(\mu, \Sigma)$$

$$f_Y(y) = \frac{1}{(2\pi)^{\frac{n}{2}} [det(\Sigma)]^{\frac{1}{2}}} \exp(-\frac{1}{2}(y-\mu)^T \Sigma^{-1}(y-\mu))$$

In our case,

$$\Sigma = \sigma I_n$$

$$det(\Sigma) = det(\sigma^2 I_n) = \sigma^2 det(I_n) = \sigma^{2n}$$

50,

$$f_Y(y) = \frac{1}{(2\pi)^{\frac{n}{2}} \sqrt{\sigma^{2n}}} \exp(-\frac{1}{2\sigma^2} ||y - \mu||^2)$$

$$\log(f_{y}(\eta)) = \frac{n}{2}\log(\sigma^{2}) - \frac{1}{2\sigma^{2}}||y - \mu||^{2} = \ell(\beta, \sigma^{2}, y)$$

$$\frac{\partial}{\partial \beta} = \dots = -\frac{1}{2\sigma^2} 2X^T (y - X\beta) = 0$$

$$\hat{\boldsymbol{\beta}} = (\boldsymbol{X}^T \boldsymbol{x})^{-1} \boldsymbol{X}^T \boldsymbol{Y} \in \mathbb{R}^P$$

$$\frac{\partial}{\partial \sigma^2} l(\beta, \sigma^2, y) = \dots = -\frac{n}{2\sigma^2} + \frac{1}{2\sigma^4} ||y - X\beta||^2 = 0$$

$$\hat{\sigma^2} = \frac{1}{n} ||y - X\hat{\beta}||^2$$

In summary, the MLE for (β, σ^2) in Gaussian Linear Model are

$$\hat{\boldsymbol{\beta}} = (\boldsymbol{X}^T \boldsymbol{x})^{-1} \boldsymbol{X}^T \boldsymbol{Y}$$

$$\hat{\sigma^2} = \frac{1}{n} ||y - X\hat{\beta}||^2$$

Note that

$$X\hat{\beta} = X(X^TX)^{-1}X^Ty = \hat{y}$$

So,
$$\hat{y} = P_{span(x)}(I_P) = P_X$$
.
Now,

$$\hat{\sigma}^2 = \frac{1}{n} ||(I_n - P_X)y||^2 = \frac{1}{n} ||Q_X y||^2$$

where $(I_n - P_X)$ is projection on to $span(X)^{\perp}$.

It turns out that $(X^T y, y^T y)$ is complete, sufficient statistic for this Gaussian linear model.

Wednesday August 31

Recall,

$$\hat{\beta} = (X^T x)^{-1} X^T Y$$

$$\hat{\sigma^2} = \frac{1}{n} ||y - X \hat{\beta}||^2$$

$$Q_x = I_n - P_x$$

$$P_X + X (X^T X)^{-1} X^T$$

Several properties,

$$E(\hat{\boldsymbol{\beta}}) = B \text{ (unbiased)}$$

$$Var(\hat{\boldsymbol{\beta}}) = (X^T X)^{-1} X^T \sigma^2 I_n X (X^T X)^{-1} = \sigma^2 (X^T X)^{-1}$$

$$\hat{\boldsymbol{\beta}} \sim N(\boldsymbol{\beta}, \sigma^2 (X^T X)^{-1})$$

Because P_x has rank p and Q_x has rank (n-p), then

$$||Q_x y||^2 \sim \chi^2_{(n-p)}$$

Let's find an unbiased estimator for σ^2

$$E(\hat{\sigma^2}) = E(\frac{1}{n}||Q_x y||^2)$$
$$= \frac{n-p}{n}\sigma^2$$
$$E(\frac{n}{n-p}\hat{\sigma^2}) = \sigma^2$$

Moreover, $\hat{\beta}$ has one-to-one transformation with

$$(X^T X)^{-1} X^t y \leftrightarrow X (X^T X)^{-1} X^t y = P_{xy}$$

$$Cov(P_{Xy}, Q_{Xy}) = P_X \sigma^2 I_n Q_X$$

= $\sigma^2 P_X Q_X$
= 0

$$P_{Xy} \perp \!\!\!\perp Q_{Xy}$$
 (due to normality)

$$\hat{\beta} \leftrightarrow P_{Xy}$$

 $\hat{\sigma}^2$ is a funciton of Q_{Xy} , so $\hat{\beta} \perp \!\!\! \perp \hat{\sigma}^2$

In your homework, $\hat{\beta}$, $\hat{sigma}^2 \leftrightarrow$ complete sufficient.

 $\hat{\beta}$, $\tilde{sigma^2}$ is UMVUE (Lehmann-Sheffe).

Theorem 1.3.1 — Gaussian Regression Model. Under this model:

1.
$$\hat{\beta}$$
, $\tilde{\sigma}^2$ UMVUE for β , σ^2
2. $\hat{\beta} \sim N(\beta, \sigma^2(X^TX)^{-1})$
3. $(n-p)\tilde{\sigma}^2 \sim \sigma^2 \chi^2_{(n-p)}$

2.
$$\hat{\boldsymbol{\beta}} \sim N(\boldsymbol{\beta}, \sigma^2(X^TX)^{-1})$$

3.
$$(n-p)\tilde{\sigma}^2 \sim \sigma^2 \chi^2_{(n-p)}$$

4.
$$\hat{\beta} \perp \tilde{\sigma}^2$$

Statistical Inference for β , σ^2

Suppose we want to test

$$H_0: \beta_1 = \beta_{i0}$$

Let $M = (X^T X)^{-1}$.

Then,

$$\hat{\beta} \sim N(\beta_i 0, \sigma^2 M_{ii})$$

where, $M_{ii} \leftarrow (i, i)^{th}$ entry of M

Also,
$$\frac{(n-p)\tilde{\sigma^2}}{\sigma^2} \sim \chi^2_{(n-p)}$$

$$\hat{eta}$$
 \perp $\tilde{\sigma^2}$

$$\frac{\frac{\hat{\beta_i} - \beta_{i0}}{\sqrt{\sigma^2 M_{ii}}} \sim N(0,1)}{\sqrt{\frac{(n-p)\sigma^2 \tilde{/}\sigma^2}{n-p}}} \sim t_{(n-p)}$$

$$T = \frac{\hat{\beta}_i - \beta_{i0}}{\sigma \tilde{\sqrt{M}_{ii}}} \sim t_{(n-p)} = (*)$$

Reject H_0 if

$$|\frac{\hat{eta}_i - eta_{i0}}{\sigma \sqrt{\tilde{M}_{ii}}}| > t_{\frac{lpha}{2}(n-p)}$$

$$X \sim N(\mu, 1) \ y \sim \chi_r^2 \ X \perp \!\!\! \perp y$$

 $\frac{X}{\sqrt{\frac{y}{r}}} \sim t_n(\mu)$

Power at β_{i1}

$$\hat{\beta}_i \sim N(\beta_{i1}, \sigma^2 M_{i1})$$

So,

$$rac{\hat{eta}_i - eta_{i0}}{ ilde{\sigma}\sqrt{M_{ii}}} \sim t_{(n-p)} (rac{eta_{i1} - eta_{i0}}{\sigma\sqrt{M_{ii}}})$$

(alternative distrabution of T)

By this (*),

$$P(\ \in (-t_{\frac{\alpha}{2}(n-p)}, t_{\frac{\alpha}{2}(n-p)}))$$

Convert this to put β_{i0} in between $(1 - \alpha)100$ percent C.I. for $\beta_{i.}$.

$$(\hat{eta_1} - t_{rac{n}{2}(n-p)}\hat{f \sigma}\sqrt{M_{ii}},\hat{eta_1} + t_{rac{n}{2}(n-p)}\hat{f \sigma}\sqrt{M_{ii}})$$

1.5 **Delete One Prediciton**

Very useful in variable selection, cross validation, diagnostics.

Prediction: $\hat{y} = X\hat{\beta} = P_x y$

But this has a drawback as it favors overfitting. Projectioning onto larger spaces will always decrease the norm, $||Q_Xy||^2$. (This can decrease errors which would cause you to think it's better, even though it's not.)

To prevent overfitting, try to be objective, withhold y_i when predicting y_i (inverse of a matrix, rank 1 perpendicular)

Theorem 1.5.1 Suppose $A \in \mathbb{R}^{P \times P}$ is a symmetric, nonsingular matrix. and $v \in \mathbb{R}^{P}$. Then,

$$(A \pm vv^T)^{-1} = A^{-1} \pm \frac{A^{-1}vv^tA^{-1}}{1 \pm v^TA^{-1}v}$$

Use what is left to compute $\hat{\beta}_{-i}$.

$$\hat{\beta_{-i}} = (X_{-1}^T X_{-i})^{-1} X_{-i}^T y_{-i}$$

This can be expanded in simple sum, so that you don't have to do n regressions.

$$()^{T} + ()^{-1}$$

$$= A^{-1} +$$

$$= ()^{-1} +$$

$$=$$

$$= (P_{x})_{ii}$$

$$= P_{i}$$

Delete-one regression.

$$X\hat{\beta}_{-i} = \hat{y}i - \frac{P_i}{1-P_i}(y_i - \hat{y}_i)$$



- General linear models
- Scheffe's simulteaneous confidence
- Singular decomposition
- Non Gaussian error



- Orthogonal design
- Additive 2 way ANOVA
- simultaneous intervals
- nonadditive
- decomposition of sum of squares
- Latin square
- nested design



$$\bullet \ \ \bar{X}_{\dot{i}} - \bar{X}_{\dot{i}}$$



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• deviance <-> sum of squares





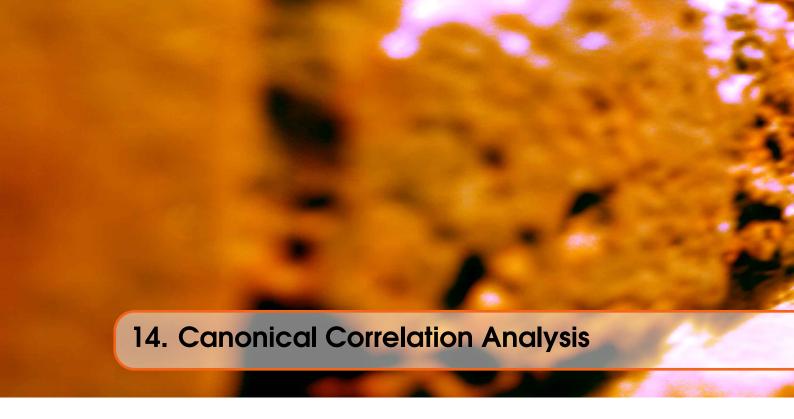




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