

# Progeny Array Statistical Genetics Methods

Vince Buffalo

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## 1 Parentage Inference Methods

Parentage inference methods are inspired by those from [Meagher and Thompson \(1986\)](#), which are reviewed in [Marshall, Slate, Kruuk, and Pemberton \(1998\)](#). The major difference in our model is the incorporation of an error model that allows for different error rates for heterozygous and homozygous genotypes. We'll step through Meagher and Thompson's original model before deriving the model with error.

### 1.1 Original Model (without Genotyping Error)

[Meagher and Thompson](#)'s basic model is to evaluate the relatedness of three individuals,  $M$ ,  $F$ , and  $O$ , using a likelihood calculated from the joint probability of these individuals' genotypes. Here,  $M$ ,  $F$ , and  $O$  represent alleged mother, alleged father, and offspring, respectively. For each loci, we calculate the probability of these individuals' genotypes  $g_m$ ,  $g_f$ , and  $g_o$  (respectively) under two pedigree models: that  $M$  and  $F$  are the parents of  $O$ , and that  $M$ ,  $F$ , and  $O$  are all unrelated individuals. Following Meagher and Thompson's notation, these two relationships are denoted  $QQ$  and  $UU$ . Another possible relationship between these individuals is that neither  $M$  or  $F$  is a parent of  $O$ ; we denote this as  $QU$  (we'll skip including this possibility for now). Thus, for a single locus the joint probability of genotypes  $g_m$ ,  $g_m$ , and  $g_o$  given that individuals  $M$  and  $F$  are  $O$ 's parents is:

$$P(g_m, g_f, g_o|QQ) = T(g_o|g_m, g_f)P(g_m)P(g_f)$$

(from [Meagher and Thompson 1986](#)). Here,  $T(g_o|g_m, g_f)$  is the Mendelian transmission matrix — essentially a conditional probability matrix.  $P(g_m)$  and  $P(g_f)$  are the probability of the observed genotypes for mother and father. Assuming random mating, these are just the Hardy-Weinberg genotype probabilities. For any genotype with alternate allele frequency  $p$ , the vector  $\mathbf{h}$  encodes the genotype probabilities under Hardy-Weinberg:

$$\mathbf{h} = \begin{bmatrix} (1-p)^2 \\ 2p(1-p) \\ p^2 \end{bmatrix}$$

digraph structure gmp [texlbl="g'\_m"] gfp [texlbl="g'\_f"] gm [texlbl="g\_m"] gf [texlbl="g\_f"] go  
[texlbl="g\_o"] gop [texlbl="g'\_o"]  
gmp -i gm [dir=back, color="blue"] gfp -i gf [dir=back, color="blue"] gm -i go [color="orange"]  
gf -i go [color="orange"] go -i gop [color="blue"] rank=same; gmp; gfp rank=same; gm; gf

Figure 1: A directed graph model with hidden genotype variables  $g_m$ ,  $g_f$ , and  $g_o$ , and observed genotypes  $g'_m$ ,  $g'_f$ , and  $g'_o$ .

## 1.2 Genotype Likelihoods with Error

At this point, this model assumes  $g_m$ ,  $g_f$ , and  $g_o$  are known with certainty. In reality, the true genotypes  $g_m$ ,  $g_f$ , and  $g_o$  are hidden and only the genotypes with error  $g'_m$ ,  $g'_f$ , and  $g'_o$  are observed. Graphically, this is depicted in [Figure 2](#).

To model genotyping errors, we derive an expression for  $P(g'_m, g'_f, g'_o)$  (given  $QQ$ ) that marginalizes over the hidden genotype variables  $g_m$ ,  $g_f$ , and  $g_o$ . So,

$$P(g'_m, g'_f, g'_o) = \sum_{g_m, g_f, g_o \in \Omega} P(g'_m, g'_f, g'_o | g_m, g_f, g_o) P(g_m, g_f, g_o)$$

The advantage of this approach is that we know through Mendelian transmission and Hardy-Weinberg that the component  $P(g_m, g_f, g_o) = T(g_o | g_m, g_f) P(g_m) P(g_f)$ , so,

$$P(g'_m, g'_f, g'_o) = \sum_{g_m, g_f, g_o \in \Omega} P(g'_m, g'_f, g'_o | g_m, g_f, g_o) T(g_o | g_m, g_f) P(g_m) P(g_f)$$

Now,  $P(g'_m, g'_f, g'_o | g_m, g_f, g_o)$  seems complex, but each of these observed genotypes conditioned on the real genotype is independent of the others. For example,  $g'_o \perp\!\!\!\perp g'_m \mid g_m$ , meaning that knowing  $g'_m$  doesn't provide any information about the observed offspring's genotype  $g'_o$  given that we know  $g_m$ . This allows us to say  $P(g'_m, g'_f, g'_o | g_m, g_f, g_o) = P(g'_m | g_m) P(g'_f | g_f) P(g'_o | g_o)$ . So:

$$P(g'_m, g'_f, g'_o) = \sum_{g_m, g_f, g_o \in \Omega} P(g'_m | g_m) P(g'_f | g_f) P(g'_o | g_o) T(g_o | g_m, g_f) P(g_m) P(g_f)$$

Every component of this model is now tractable. Our error model enters through the stochastic transition matrix  $\mathbf{E}$  for any true, latent genotype  $g$  to an observed genotype  $g'$ :

$$\mathbf{E} = P(g' | g) = \begin{bmatrix} 1 - e & e/2 & e/2 \\ e/2 & 1 - \epsilon & \epsilon/2 \\ e/2 & e/2 & 1 - e \end{bmatrix}$$

Where  $e$  is the homozygous error rate and  $\epsilon$  is the heterozygous error rate. This error model is similar to models in which the probability of error is distributed uniformly over the two erroneous genotypes ([Sobel, Papp, and Lange, 2002](#); [Lincoln and Lander, 1992](#)).

Extending this to all loci is trivial; we assume independence across loci, so for all loci we sum the log probability. Allowing locus  $l$ 's joint probability to be written as  $P(g'_{m,l}, g'_{f,l}, g'_{o,l})$ , the probability of  $M$  and  $F$  being  $O$ 's parents is:

$$P(G'_m, G'_f, G'_o | QQ) = \sum_{l \in L} \log P(g'_{m,l}, g'_{f,l}, g'_{o,l})$$

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digraph parentimputenode[fixedsize = shape]gmp[texlbl = "g'm" ] gm [texlbl="gm" ] go1
[texlbl="go,1" ] go2 [texlbl="go,2" ] go3 [texlbl="go,3" ] go4 [texlbl="go,4" ] go5 [texlbl="go,5" ] gop1
[texlbl="g'o,1" ] gop2 [texlbl="g'o,2" ] gop3 [texlbl="g'o,3" ] gop4 [texlbl="g'o,4" ] gop5 [texlbl="g'o,5" ]
gmp -i gm [dir=back, color="blue" ] gm -i go1 [color="orange" ] gm -i go2 [color="orange" ] gm
-i go3 [color="orange" ] gm -i go4 [color="orange" ] gm -i go5 [color="orange" ] go1 -i gop1
[color="blue" ] go2 -i gop2 [color="blue" ] go3 -i gop3 [color="blue" ] go4 -i gop4 [color="blue" ]
go5 -i gop5 [color="blue" ]

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Figure 2: A half-sib (with shared mother) with five offspring.

### 1.3 Parental Genotype Imputation

$$\begin{aligned}
P(g_m, g'_{o,\cdot}) &= P(g_m) \prod_{i=1}^n P(g'_{o,i} | g_m) \\
P(g_m, g'_{o,\cdot}) &= P(g_m) \prod_{i=1}^n \sum_{g_o \in \Omega} P(g'_{o,i}, g_{o,i} | g_m) \\
P(g_m, g'_{o,\cdot}) &= P(g_m) \prod_{i=1}^n \sum_{g_o \in \Omega} P(g'_{o,i} | g_{o,i}) P(g_{o,i} | g_m)
\end{aligned} \tag{1}$$

The form of this model is similar to a naive Bayesian model (Koller and Friedman 2009, p. 50), with an added layer of uncertainty from true offspring genotypes to observed offspring genotypes.

### 1.4 Parental Haplotype Inference

Adopting the design of (Browning and Browning, 2009), we build a duo leveled HMM ( $H^2$ ) for GBS data with high error rate. Our hidden states are the alleles on three haplotypes:

## References

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