



## Machine learning methods for inflation forecasting in Brazil: New contenders versus classical models<sup>☆</sup>

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### ARTICLE INFO

*JEL classification:*

C14  
C15  
C22  
C53  
C55  
E17  
E31

*Keywords:*

Machine learning  
Big data  
Inflation forecasting

### ABSTRACT

In this paper, we explore machine learning (ML) methods to improve inflation forecasting in Brazil. An extensive out-of-sample forecasting exercise is designed with multiple horizons, a large database of 501 series, and 50 forecasting methods, including new ML techniques proposed here, traditional econometric models and forecast combination methods. We also provide tools to identify the key variables to predict inflation, thus helping to open the ML black box. Despite the evidence of no universal best model, the results indicate that ML methods can, in numerous cases, outperform traditional econometric models in terms of mean-squared error. Moreover, the results indicate the existence of nonlinearities in the inflation dynamics, which are relevant to forecasting inflation. The set of top forecasts often includes forecast combinations, tree-based methods (such as random forest and xgboost), breakeven inflation, and survey-based expectations. Altogether, these findings offer a valuable contribution to macroeconomic forecasting, especially, focused on Brazilian inflation.

### 1. Introduction

Producing reliable inflation forecasts is a constant challenge for policymakers and of greatest importance to economic agents and their investment decisions. Inflation adds uncertainty to investment decisions and shortens the investment horizon, especially in emerging markets, making the construction of accurate forecasts a relevant issue in such economies.

Building accurate forecasts is generally not an easy task because it requires an approach complex enough to incorporate relevant variables but also focused on excluding irrelevant data. In this sense, machine learning (ML) methods, in general, are able to identify

<sup>☆</sup> We are especially grateful for the helpful comments and suggestions from co-editor Nelson Ramírez-Rondán, two anonymous referees, Marcelo Aragão, Jorge Henrique Barbosa, Rafael Cusinato, Angelo Fasolo, Diogo Guillen, Serafín Martínez-Jaramillo, Fabio Kanczuk, Vicente Machado, Marcelo Medeiros, Euler de Mello, André Minella, and José Valentim Vicente. We also benefited from comments by the seminar participants at the 18th ESTE - Time Series and Econometrics Meeting, IV Workshop da Rede de Pesquisa do Banco Central do Brasil, CEMLA XXV Meeting of the Central Bank Researchers Network, Experiências em Data Science no Coaf e no BCB, XXII Encontro Brasileiro de Finanças, and the 44th Meeting of the Brazilian Econometric Society. Gaglianone gratefully acknowledges the financial support from the National Council for Scientific and Technological Development - CNPq (Brazil). The views expressed in the paper are those of the authors and do not necessarily reflect those of the Banco Central do Brasil or the Fundação Getulio Vargas.

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nonlinear patterns in the data, hidden to standard linear models, thus offering an alternative (and compelling) approach to traditional econometric models.<sup>1</sup>

Moreover, despite the low frequency of data in macroeconomics, where the usual variables of interest are collected on an annual, quarterly or monthly basis (leading to much less data accumulation in the time dimension compared, for instance, to an intraday high-frequency database) and the usual split of data into training and test sets (in-sample and out-of-sample, reducing still further the amount of data used for model estimation), there is still a high incentive to use ML methods<sup>2</sup> in applied macroeconomics also because many of these methods can handle datasets with a large number of predictors, in contrast to linear econometric models usually based on a few variables. Furthermore, the continuous improvement in computer technology allows the running of ML algorithms at a much faster speed.

The objective of this paper is to forecast Brazilian inflation based on a large number of macroeconomic and financial variables. Our goal is to assess whether ML approaches can indeed offer improvement in forecast accuracy in applied macroeconomics and contribute to the standard statistical toolkit used in macro forecasting.

To do so, we conduct an extensive *horse-race* (pseudo out-of-sample forecasting exercise) across 50 models (or methods) to forecast inflation in Brazil at multiple horizons (ranging from 1 up to 18 months). The list of competing methods includes several ML methods based on regularization approaches (elastic net, lasso, adaptive lasso, and ridge) or regression trees (random forest, quantile regression forest, and xgboost), as well as traditional econometric approaches (ARMA, VAR, factor models), reduced-form structural models (Phillips curves), survey-based forecasts (Focus), and breakeven inflation (BEI) from financial market data, among many others. Our database covers 501 time series, coming from 167 macroeconomic and financial variables used to build high-dimensional models.

The literature on macroeconomic forecasting using ML methods is relatively new. For instance, see Medeiros et al. (2016) and Garcia et al. (2017) for applications with Brazilian data; Cheng et al. (2019) for aggregating individual survey-based forecasts, using ML tools to improve forecasts of the U.S. inflation; Kohlschein (2021) for an investigation of the drivers of inflation in 20 advanced countries using random forest; and Costa et al. (2021) for oil price point and density forecasting using ML methods.

Our research contributes to this fast-growing literature in five ways. The first original contribution is the proposal of a new *quantile-combination method*, based on Meinshausen's (2006) quantile regression forest model. The idea is to use information of the conditional distribution from a set of estimated conditional quantiles to build an improved conditional mean forecast. The second contribution is to employ a *hybrid machine learning* approach, inspired by the work of Medeiros et al. (2021), to build new ML methods. The goal is to disentangle forecast accuracy due to variable selection from possible nonlinearities in the data-generating process. The third contribution is to use ML not only as a forecast method but also as a *forecast combination* device. The idea is to check whether ML methods can beat traditional forecast combination approaches when combining a given set of point forecasts. The fourth contribution is to provide a simple way to build *fan charts* from ML-based inflation forecasts, where a measure of uncertainty can be attached to the forecasted inflation-path based on past forecast errors. The final contribution is to help opening the ML *black box*<sup>3</sup> by employing a set of auxiliary graphs<sup>4</sup> already available in the literature to further analyze the performance of competing methods far beyond the usual accuracy analysis based on mean-squared forecast error.

The outline of the paper is as follows. In Section 2, we present the methodology comprising machine learning methods and traditional econometric models to predict inflation. Section 3 presents an out-of-sample empirical exercise where competing methods are used to forecast Brazilian inflation, in the same spirit as Medeiros et al. (2016).<sup>5</sup> Section 4 concludes.

## 2. Methodology

### 2.1. Machine learning in a nutshell

ML is a branch of artificial intelligence often described as the art and science of pattern recognition. It is essentially a data-driven approach with mild assumptions about the underlying statistical relationships in the data, and it entails a large variety of methods. It usually consists of two core elements, a learning method and an algorithm, enabling one to automate as many of the modeling choices as possible in a manner that is not subject to the discretion of the forecaster (Hall, 2018).

Most traditional forecasting methods rely on fitting data to a pre-specified relationship between dependent and independent variables, thus assuming a specific functional and stochastic process. In contrast, ML offers a different approach to statistical analysis

<sup>1</sup> Also, there is a big difference in the final objective of such approaches: while in ML it is about forecasting using data (usually, the more data the better), in traditional econometrics it is about explaining the structural relationships between economic variables and testing if they conform to economic theory. In this sense, inflation forecasting is a by-product of such relationships (e.g., the Phillips curve). In turn, ML methods originally do not establish underlying economic relationships and simply try to obtain the best prediction, not to explain causalities (although causal inference has been recently attracting substantial attention in the ML community; see Cui and Athey 2022).

<sup>2</sup> A lot of ML methods are originally designed to deal with large amounts of data (*big data*).

<sup>3</sup> The *black box* term applied to ML techniques has been around for years. It is often employed to criticize the lack of explainability in neural networks.

<sup>4</sup> The set of auxiliary graphs consists of bias-variance decomposition, cumulative squared prediction error, word clouds, and variable (or feature) importance.

<sup>5</sup> Compared to previous papers focused on Brazilian inflation forecasting (e.g., Medeiros et al. 2016, and Garcia et al. 2017), we considered: (i) a broader set of models (proposing new ML methods and forecast combination tools); (ii) a larger database of macroeconomic and financial variables, including new potential predictors; (iii) longer forecast horizons (from one up to 18 months); and (iv) inflation rates measured either by monthly or 12-month accumulated rates.

and forecasting, in particular, because it is to a great extent data-driven, as it makes almost no assumption about the underlying statistical relationships in the data.

According to [Samuel \(1959\)](#), ML is the ability of computers to learn from experience without being explicitly programmed. [Cerulli and Drago \(2021\)](#) point out that ML places itself at the intersection between statistics, computer science, and artificial intelligence. According to these authors: “*The primary objective of ML is turning information into knowledge and value by ‘letting the data speak’*”. [Hansen \(2019\)](#) explains that ML is “*a new and somewhat vague term, but typically is taken to mean procedures which are primarily used for point prediction in settings with unknown structure. Machine learning methods generally allow for large sample sizes, large number of variables, and unknown structural form.*”

In fact, ML encompasses a wide variety of models but often consists of two core elements: a *learning method*, where data is used to determine the best fit for the input variables, and an *algorithm* that captures the relationship between the input and output. In general, ML can be categorized into three types (see [Jung et al. 2018](#)):

(i) *supervised learning*, where the dependent variables are clearly identified, even if the specific relationships in the data are not known (e.g., linear regression, logistic regression);

(ii) *unsupervised learning*, where there is no specific output defined beforehand, and the goal is to recognize data patterns and determine output classification categories (e.g., cluster analysis, principal components); and

(iii) *reinforcement learning*, which iteratively searches for an optimal location for the input variables that yields the highest reward (i.e., maximizing a reward function using no training set. For example: sarsa, Q-learning).

According to [Varian \(2014\)](#), the growing amounts of data and ever complex-growing relationships warrant the usage of ML approaches in economics. In this paper, we build inflation forecasts using different supervised ML algorithms: based either on penalized-regression models (e.g., ridge regression, lasso, adaptive lasso, elastic net) or on tree-based methods (e.g., random forest, quantile regression forest, XGBoost), besides a recurrent neural network model.

The first approach entails regularization techniques that introduce penalties for *overfitting* the data.<sup>6,7</sup> For example, the elastic net model mixes two different kinds of regularization by penalizing both the number of variables in the model and the extent to which any given variable contributes to the model’s forecast. By applying these penalties, the elastic net *learns* which variables are most important, eliminating the need for researchers to make discretionary choices about which variables to include.

The second (tree-based) approach is nonparametric, based on the recursive binary partitioning of the covariate space, which can deal with a very large number of explanatory variables, thus producing highly nonlinear predicted models.

## 2.2. Models (or methods) to forecast inflation

There is a variety of approaches in the literature to model inflation dynamics. According to [Ang et al. \(2007\)](#), economists use four main methods to forecast inflation: time-series models, structural models (e.g., Phillips curve), asset price models (e.g., term-structure of interest rates), and methods that employ survey-based measures (e.g., survey of professional forecasters).

In this paper, inflation forecasts come from 50 forecasting methods listed in [Table 1](#). Besides some traditional econometric approaches to forecast inflation, such as ARMA and VAR models, this paper considers Phillips curves, well-known in the macro literature (e.g., [Stock and Watson, 1999](#)), survey-based inflation expectations, and inflation forecasts embedded in financial market data (breakeven inflation). The set of forecasting methods also includes many nonlinear machine learning methods, based on regularization procedures or regression trees, and several forecast combination techniques.

The list of models, of course, is not an exhaustive list, as models that are more complex can always be included.<sup>8</sup> Nonetheless, the set of inflation forecasting methods listed in [Table 1](#) is a good starting point for comparing the accuracy of traditional econometric approaches with some new competing ML techniques.

Our main goal here is to forecast the inflation rate  $y_{t+h}$  at period  $t+h$  using the information set available at period  $t$ . In this sense, inflation is modeled as a function of a set of predictors  $\tilde{x}_t$ , measured at time  $t$ , as follows:

$$y_{t+h} = \Psi_h(\tilde{x}_t) + \varepsilon_{t+h}, \quad (1)$$

where  $\Psi_h(\cdot)$  is a possibly nonlinear mapping of a set of predictors (a single model or an ensemble of different specifications),  $\varepsilon_{t+h}$  is the forecasting error, and predictors  $\tilde{x}_t$  may include weakly exogenous predictors, lagged values of inflation, and a number of factors computed from a large number of potential covariates ([Garcia et al., 2017](#)). Here, we consider  $\tilde{x}_t = \{\mathbf{1}_t, x_t, x_{t-1}, \dots, x_{t-s}\}'$ , where  $\mathbf{1}_t$  is a constant term,  $x_t = \{x_{1,t}, \dots, x_{n,t}\}$  is a set of  $n$  predictors, and  $s$  is the maximum lag adopted for the set of variables  $x_t$  when forming the vector of variables  $\tilde{x}_t$ .

<sup>6</sup> In statistics, *overfitting* denotes the production of an analysis, which is assumed to be valid for the entire population (for instance, an estimated input-output relationship), which corresponds too closely to a particular set of data, but may fail to fit additional data or reliably forecast future observations. In other words, when the model learns the training sample too well and shows low prediction capability out-of-sample, it is overfitted.

<sup>7</sup> According to [Hall \(2018\)](#), ML algorithms usually deliver a model complex enough to avoid underfitting the data but not too complex to overfit it.

<sup>8</sup> For instance, the set of models could include simpler reduced-form models such as Bayesian VARs ([Andres-Escayola et al., 2021](#)) or Factor Augmented VARs ([Figueiredo and Guillén, 2013](#)), the Unobserved Component Stochastic Volatility (UCSV) approach of [Stock and Watson \(2007\)](#), small- or medium-scale semi-structural macroeconomic models ([Brasil, 2021a,b](#)), Dynamic Stochastic General Equilibrium (DSGE) structural models ([Azevedo Costa, 2016](#); [Gonçalves et al., 2016](#); [Sin and Gaglianone, 2006](#)), among others, which nowadays are commonly used in central banks to forecast inflation.

**Table 1**  
Models (or methods) selected to forecast inflation.

|    |   |    |  |
|----|---|----|--|
| 1  | Random Walk                             | 26 | Hybrid Random Forest - Adalasso            |
| 2  | Random Walk (Atkeson-Ohanian)           | 27 | Hybrid Random Forest - XGBoost             |
| 3  | ARMA                                    | 28 | Inflation Expectations (Breakeven)         |
| 4  | VAR                                     | 29 | Inflation Expectations (Focus Survey)      |
| 5  | Phillips Curve (Backward)               | 30 | Combination 1 (Mean)                       |
| 6  | Phillips Curve (Hybrid)                 | 31 | Combination 1 (Median)                     |
| 7  | Factor Model 1                          | 32 | Combination 1 (Granger-Ramanathan)         |
| 8  | Factor Model 2                          | 33 | Combination 1 (Constrained Least Squares)  |
| 9  | Factor Model 3                          | 34 | Combination 1 (Complete Subset Regression) |
| 10 | Factor Model 4                          | 35 | Combination 1 (Adalasso)                   |
| 11 | Elastic Net                             | 36 | Combination 1 (Random Forest)              |
| 12 | Lasso                                   | 37 | Combination 2 (Mean)                       |
| 13 | Adaptive Lasso (Adalasso)               | 38 | Combination 2 (Median)                     |
| 14 | Ridge Regression                        | 39 | Combination 2 (Granger-Ramanathan)         |
| 15 | Random Forest                           | 40 | Combination 2 (Constrained Least Squares)  |
| 16 | Quantile Regression Forest              | 41 | Combination 2 (Complete Subset Regression) |
| 17 | XGBoost                                 | 42 | Combination 2 (Adalasso)                   |
| 18 | Recurrent Neural Network (RNN)          | 43 | Combination 2 (Random Forest)              |
| 19 | Disaggregated Inflation (ARMA)          | 44 | Combination 3 (Mean)                       |
| 20 | Disaggregated Inflation (Adalasso)      | 45 | Combination 3 (Median)                     |
| 21 | Disaggregated Inflation (Random Forest) | 46 | Combination 3 (Granger-Ramanathan)         |
| 22 | Hybrid Adalasso - OLS                   | 47 | Combination 3 (Constrained Least Squares)  |
| 23 | Hybrid Adalasso - Random Forest         | 48 | Combination 3 (Complete Subset Regression) |
| 24 | Hybrid Adalasso - XGBoost               | 49 | Combination 3 (Adalasso)                   |
| 25 | Hybrid Random Forest - OLS              | 50 | Combination 3 (Random Forest)              |

Notes: Combination 1 is based on models 1–27. Combinations 2 and 3 are based on the superior models of the *model confidence set* by Hansen et al. (2011), considering models 1–27 or 1–29, respectively.

To build our forecasting exercise, we split the sample into three consecutive time subperiods, where time is indexed by  $t = 1, 2, \dots, T_1, \dots, T_2, \dots, T$ . The first subperiod ( $t = 1, \dots, T_1$ ), usually called the *estimation sample*, is used for model estimation and forecast inflation  $y_t$  in the subsequent periods.<sup>9</sup> In the second subperiod ( $t = T_1 + 1, \dots, T_2$ ), realizations of  $y_t$  are confronted with forecasts produced in the estimation sample, and forecast combination weights are estimated, if that is the case. The first and second subperiods, together, are labeled as the *training set*. The final subperiod, also known as the *test set*, is where a genuine out-of-sample forecast is entertained, comprising the last  $P$  observations of the sample ( $t = T_2 + 1, \dots, T$ ). Thus, we have  $P = T - T_2$  observations to compare forecasts and compute accuracy measures.

For the regularization approaches considered in this paper (e.g., elastic net), the mapping  $\Psi_h(\cdot)$  is linear, such that:

$$y_{t+h} = \tilde{x}'_t \beta_h + \varepsilon_{t+h}, \quad (2)$$

where  $\beta_h$  is a vector of unknown parameters. The inflation forecast from the linear ML approach,  $f_{y_{T_2+h}}^{ML}$ , using a sample of  $t = 1, \dots, T_2$  observations, is given by:

$$f_{y_{T_2+h}}^{ML} = \tilde{x}'_{T_2} \widehat{\beta}_h, \quad \text{for } h = 1, \dots, H. \quad (3)$$

To evaluate forecast  $f_{y_t}^{ML}$ , we compute the respective mean-squared error as follows:  $MSE_h = \frac{1}{(P-h+1)} \sum_{t=T_2+h}^T (y_t - f_{y_t}^{ML})^2$ . Note that we adopt the *direct forecast* approach, where the inflation  $h$  periods ahead ( $y_{T_2+h}$ ) is modeled as a function of a set of predictors  $\tilde{x}'_t$  measured at time  $T_2$ . In other words, for each horizon  $h$ , we estimate a different vector of unknown parameters  $\beta_h$  (in contrast to the iterated multistep approach; see Marcellino et al. 2006). Thus, we avoid the necessity of estimating a model for the time-evolution of  $\tilde{x}_t$ .

### 2.2.1. Elastic net, lasso, adaptive lasso, and ridge regression

**Elastic Net:** This is a regularization method proposed by Zou and Hastie (2005)<sup>10</sup> which simultaneously performs automatic variable selection and continuous shrinkage and can select groups of correlated variables. The elastic net encourages a grouping effect where highly correlated regressors tend to be jointly included (or excluded) from the model, and it can be particularly useful when the number of predictors  $k$  is high compared to the number of observations  $T$ . For a nonnegative shrinkage parameter  $\lambda$ , and a combination parameter  $\alpha$  strictly between 0 and 1, the elastic net solves the following problem:

$$\hat{\beta} = \arg \min_{\{\beta_1, \dots, \beta_k\}} \left( \frac{1}{T} \sum_{t=1}^T \left( y_t - \sum_{j=1}^k x'_{j,t} \beta_j \right)^2 + \lambda P_\alpha(\beta) \right), \quad (4)$$

<sup>9</sup> As will be discussed in Section 3.1, we use a recursive estimation scheme (i.e., increasing sample size).

<sup>10</sup> According to the authors: “It is like a stretchable fishing net that retains ‘all the big fish’.”

where

$$P_\alpha(\beta) = \sum_{j=1}^k \alpha |\beta_j| + \frac{(1-\alpha)}{2} \beta_j^2, \quad (5)$$

and  $\beta$  is the  $k \times 1$  vector of parameters,  $y_t$  is the dependent variable, and  $\{x_{1,t}, \dots, x_{k,t}\}$  is the  $k \times 1$  vector of regressors. The tuning parameter  $\lambda$  controls the overall strength of the penalty term  $P_\alpha(\beta)$ , which interpolates between the  $l_1$ -norm of  $\beta$  and the squared  $l_2$ -norm of  $\beta$ . Note that by setting  $\lambda = 0$ , the elastic net becomes the ordinary least squares (OLS) regression. Note also that the objective function has no-closed form solution, but it is convex and can be minimized using any convex optimization method such as gradient or coordinate descent.<sup>11</sup>

**Lasso:** The least absolute shrinkage and selection operator (*lasso*) was proposed by Tibshirani (1996). The core idea is to shrink to zero the irrelevant coefficients. The lasso is a penalized least squares method imposing an  $l_1$ -penalty on the regression coefficients, which allows lasso to perform continuous shrinkage and automatic variable selection simultaneously. It is also a particular case of the elastic net estimator (4), considering  $\alpha = 1$  in penalty term (5).

According to Cheng et al. (2019), lasso is “*the most intensively studied statistical method in the past 15 years*.” Indeed, it has shown success in many practical situations, as it can handle more variables than observations. Nonetheless, it has limitations and might even become an inappropriate method for selecting variables in some cases. Zou and Hastie (2005) list a few examples: (i) when the number of predictors  $k$  is greater than the number of observations  $T$ , lasso selects at most  $T$  variables before it saturates, due to the nature of the convex optimization problem; (ii) in the case of a *grouping effect*<sup>12</sup> lasso tends to select only one variable from the group; (iii) when  $T > k$  and with high correlated predictors, the ridge regression tends to perform better than lasso.

**Adaptive Lasso:** Zou (2006) shows that lasso is inconsistent for variable selection under certain circumstances and proposes the adaptive lasso (or *adalasso*), where adaptive weights are used for penalizing different coefficients in the  $l_1$ -penalty. According to the author, the adaptive lasso enjoys oracle properties (i.e., it performs as well as if the true underlying model were known) and does not select useless variables that may damage the accuracy of a forecast. The core idea behind the model is to use previously known information to select the variables more efficiently. In practice, it is a two-step estimation that first generates different weights  $w_j$  for each candidate variable  $x_{j,t}$ , which are used in the second step (lasso estimation) as additional information. The adalasso estimator is defined as follows:

$$\hat{\beta} = \arg \min_{\{\beta_1, \dots, \beta_k\}} \left( \frac{1}{T} \sum_{t=1}^T \left( y_t - \sum_{j=1}^k x'_{j,t} \beta_j \right)^2 + \lambda \sum_{j=1}^k w_j |\beta_j| \right), \quad (6)$$

where  $w_j = |\hat{\beta}_j^*|^{-\tau}$  represents the weights;  $\hat{\beta}_j^*$  is a parameter estimated in the first-step, and  $\tau > 0$  is an additional tuning parameter that determines how much one wants to emphasize the difference in the weights. In general,  $\tau$  is set to unity and  $\hat{\beta}_j^*$  is the respective lasso coefficient estimated in the first step.<sup>13</sup>

**Ridge Regression:** In contrast to lasso, the ridge regression (Hoerl and Kennard, 1970) minimizes the squared sum of the residuals subject to a bound on the  $l_2$ -norm of the parameters. It is a particular case of the elastic net estimator (4), considering  $\alpha = 0$  in penalty term (5). Since ridge is a continuous shrinkage method, in some cases it can achieve better out-of-sample performance through a *bias-variance* trade-off (i.e., using regularization to balance the forecast errors due to bias and variance). In particular, ridge is good at improving the OLS counterpart when multicollinearity is present. However, ridge cannot be used for variable selection (and to produce a parsimonious model) because it retains all regressors in the model; that is, it only shrinks the coefficients close (but never equal) to zero.

## 2.2.2. Random forest

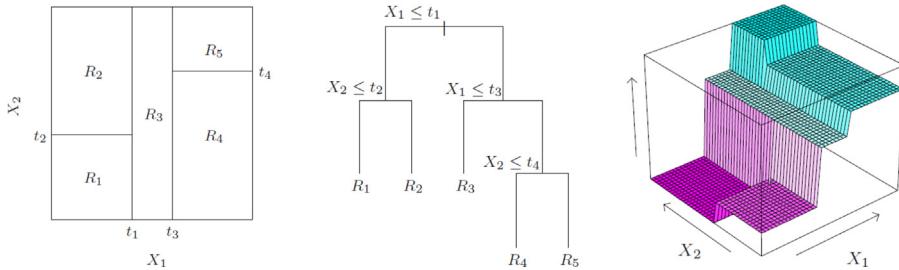
A random forest (RF) is a collection of decision trees introduced as a ML tool in Breiman (2001). It is a very popular and powerful method used in high-dimensional regression or classification. The main idea is to reduce the forecast variance by using bootstrap aggregation (*bagging*) of randomly constructed trees. A *regression tree* is a nonparametric model based on the recursive binary partitioning of the covariate space  $X$ . According to Garcia et al. (2017), the model is usually displayed in a graph, which has the format of a binary decision tree with parent nodes (or split nodes) and terminal nodes (called leaves; which represent different partitions of  $X$ ).<sup>14</sup> Fig. 1 shows an example of a regression tree with two covariates.

<sup>11</sup> Although we defined the elastic net by using  $(\lambda, \alpha)$ , this is not the only choice as the tuning parameters; see Zou and Hastie (2005). Also,  $(\lambda, \alpha)$  can be chosen using cross-validation (CV) or information criteria. In general, CV is not appropriate for time series data because temporal dependency imposes correlation in the time dimension, while CV usually assumes that data are independent and identically distributed. This way, CV splits the dataset randomly, losing the chronological order of observations, which is troublesome in forecasting, because one would be using the future to predict the past. To avoid such a problem, we select the tuning parameters for lasso, adalasso, and elastic net using the Bayesian Information Criterion (see Medeiros et al. 2016).

<sup>12</sup> The grouping effect occurs if the regression coefficients of a group of highly correlated variables tend to be equal (up to a change of sign if negatively correlated).

<sup>13</sup> In our empirical exercise, we adopt  $\tau = 0.3$ , following Medeiros et al. (2016). Alternative adalasso models with  $\tau = 1$  or 2, in general, generate less accurate forecasts in terms of MSE.

<sup>14</sup> According to the authors, the partitions are often defined by a set of hyperplanes, each of which is orthogonal to the axis of a given predictor variable (also called the split variable).



**Fig. 1.** Example of a recursive binary splitting in a regression tree. Notes: The graph on the left shows the partition of a two-dimensional covariate space. The graph in the center displays the corresponding tree, and the graph on the right shows the prediction surface. Source: [Hastie et al. \(2009\)](#).

Note that we first split the covariate space into two regions ( $X_1 \leq t_1$  and  $X_1 > t_1$ )<sup>15</sup> and model the dependent variable by the mean of  $Y$  in each region. The selected variable ( $X_1$ ) and the corresponding split-point ( $t_1$ ) are chosen in order to achieve the best fit. Then, one (or both) of these regions is (are) split into two more regions, and this process is continued until a stopping rule is applied.

In the example shown in Fig. 1, the regression tree model predicts  $Y$  with the constant  $c_m$  in region  $R_m$ ,  $m = 1, \dots, 5$ , as follows:

$$\mathbb{E}_{\text{regression tree}}(Y | (X_1, X_2)) = \sum_{m=1}^5 c_m \mathbf{1}_{\{(X_1, X_2) \in R_m\}}. \quad (7)$$

In practice, one major problem with regression trees is their high prediction variance. Usually, a small change in the data leads to a very different sequences of splits. The main reason for such instability is the hierarchical nature of the algorithm (the effect of a big error in the top split is propagated down to all of the splits below it).<sup>16</sup>

To overcome this issue, one can employ the *bagging* (bootstrap aggregation) method, which consists of fitting the same tree several times to bootstrap-sampled versions of the training data and, then, average the result. This approach often improves model performance because it decreases the forecast variance without increasing the bias too much.<sup>17</sup>

RF uses a modified-bagging method (*random subspace projection*) that selects a random *subset* of covariates at each candidate split. The reason for doing this is the correlation of trees in the ordinary bootstrap: if few covariates are strong predictors for the dependent variable, such covariates will be selected in many of the bootstrapped trees, causing them to be correlated. According to [Hansen \(2019\)](#), the modification adopted in RF aims at *decorrelating* the bootstrap trees by introducing extra randomness.<sup>18</sup>

### 2.2.3. Quantile regression forest

RF approximates the conditional mean of  $Y$  by constructing a weighted average over the sample observations of  $Y$ . Nonetheless, random forests can also provide information about the full conditional distribution of the response variable, not only about the conditional mean. This information can be used, for instance, to build prediction intervals and account for outliers in the data. In this way, conditional quantiles can be inferred with quantile regression forest (QRF), a generalization of RF proposed by [Meinshausen \(2006\)](#).<sup>19</sup>

The idea here is to provide a nonparametric way of estimating conditional quantiles for a high-dimensional set of predictor variables. According to the author, the QRF algorithm is shown to be consistent and competitive in terms of predictive power. First, recall that the conditional quantile of  $Y$ , given  $X$ , at quantile level  $\tau$ , is defined by

$$Q_\tau(Y | X) = \inf\{y : F(y | X) \geq \tau\} \quad \text{or, equivalently,} \quad (8)$$

$$F(y | X) = \Pr(Y \leq y | X) = \mathbb{E}(I_{\{Y \leq y\}} | X), \quad (9)$$

where  $F(y | X)$  is the conditional cumulative distribution function (CDF) and  $I_{\{Y \leq y\}}$  is an indicator function. Note that the probability of  $Y$  being smaller than  $Q_\tau(\cdot)$  is equal to  $\tau$ . Next, we approximate the CDF by the weighted average of  $I_{\{Y_i \leq y\}}$  over  $n$  observations, as follows:

$$\hat{F}(y | X) = \sum_{i=1}^n w_i(x) I_{\{Y_i \leq y\}}, \quad (10)$$

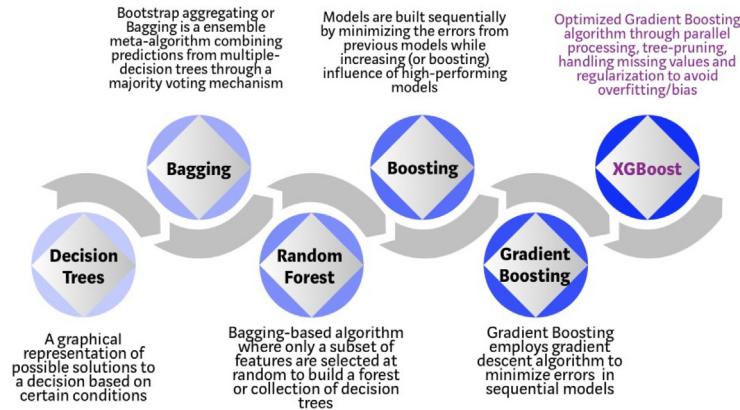
<sup>15</sup> Rather than splitting each node into just two groups, one might consider multiple splits at each stage. However, according to [Hastie et al. \(2009, p.311\)](#), this is not a good strategy because multiple splits fragment the data too quickly, leaving insufficient data at the next level down.

<sup>16</sup> According to [Hastie et al. \(2009\)](#), regression trees tend to learn irregular data patterns and overfit training sets, producing low bias and high variance. To reduce variance, RF averages multiple trees, trained on different parts of the training set. This often generates a small bias, but generally improves forecast accuracy.

<sup>17</sup> Training many trees on a single training set would yield strongly correlated trees, whereas bootstrap sampling helps decorrelating the trees by showing them different training sets.

<sup>18</sup> Thus, forecast variance can be reduced in two ways: (i) in each node, the variable being split is selected from a random subset of variables (instead of the full set); and (ii) each tree is learned on a bootstrapped subsample. See [Hastie et al. \(2009\)](#) for further details.

<sup>19</sup> The main difference between QRF and RF is that for each node, RF keeps only the mean of the observations that fall into this node (and neglects all other information). In contrast, QRF keeps the value of all observations in this node and assesses the conditional distribution based on this full information.



**Fig. 2.** Main algorithms used in decision trees. Notes: Boosting is an *ensemble* technique (i.e., it makes an average of the predictions of a group of models) that constructs models sequentially, and each subsequent model corrects the errors of the previous one, whereas *bagging* constructs models independently. Source: [Morde and Setty \(2019\)](#).

using the weights  $w_i(x)$ . This way, estimates of the conditional quantiles  $\hat{Q}_\tau(\cdot)$  can be obtained by plugging  $\hat{F}(\cdot)$ , instead of  $F(\cdot)$  into (8). Now, we go one-step further, by relating the conditional quantiles with the conditional mean of  $Y$ . This could be accomplished by integrating the conditional quantile function of  $Y$  over the entire domain  $\tau \in [0, 1]$  (see [Koenker 2005](#), p.302). The conditional mean  $\mathbb{E}(Y | X)$  can thus be approximated<sup>20</sup> by a sum of estimated conditional quantiles, as follows:<sup>21</sup>

$$\mathbb{E}(Y | X) = \int_0^1 Q_\tau(Y | X) d\tau = \lim_{P \rightarrow \infty} \left( \sum_{p=1}^P \hat{Q}_{\tau_p}(Y | X) \Delta \tau_p \right). \quad (11)$$

The idea is to aggregate information from different conditional quantiles in order to achieve an improved conditional mean. The approximation of the conditional mean through a combination of conditional quantiles is not a novel approach in the literature. Indeed, it has a long tradition in statistics (see [Judge et al. 1988](#)) and has been previously applied in the forecasting literature. Nonetheless, our original contribution is to propose a new quantile-combination approach, based on QRF, to build conditional mean forecasts through [equations \(8\), \(10\)](#) and (11). The approach proposed here follows the spirit of the averaging scheme applied to quantiles conditional on predictors selected by *lasso*, as proposed by [Lima and Meng \(2017\)](#), and of [Jiang et al. \(2020\)](#), which show that aggregating information over different quantiles can produce superior forecasts for the prediction of stock returns.

The advantage of such approaches relies on the fact that quantiles are robust to outliers (in our case, extreme unanticipated inflationary shocks), which potentially improves forecast accuracy and likely impacts the performance of standard models, usually designed to account for average responses. In sum, we propose the following three-step algorithm:

1. Choose a finite (equidistant) grid of quantile levels. For example:  $\Gamma \equiv [0.05, 0.10, \dots, 0.95]$ ;
2. For each  $\tau \in \Gamma$ , period  $t$ , forecast horizon  $h$ , and information set  $\mathcal{F}_t$ , estimate the conditional quantile  $\hat{Q}_\tau(y_{t+h} | \mathcal{F}_t)$  using the QRF method of [Meinshausen \(2006\)](#); and
3. Compute the average of  $\hat{Q}_\tau(y_{t+h} | \mathcal{F}_t)$  across all  $\tau \in \Gamma$ , and consider it as proxy for  $\mathbb{E}(y_{t+h} | \mathcal{F}_t)$ , i.e., as our (QRF-based) inflation forecast.

#### 2.2.4. Extreme gradient boosting

Extreme Gradient Boosting (or XGBoost) is a decision tree-based ensemble algorithm that uses a gradient boosting setup, as proposed by [Chen and Guestrin \(2016\)](#). It improves upon the previous gradient boosting setups through systems optimization and algorithmic enhancements. According to [Morde and Setty \(2019\)](#), the XGBoost algorithm has the best combination of prediction performance and processing time compared to other algorithms. [Fig. 2](#) shows a brief comparison of the most common decision tree algorithms.

Thus, XGBoost is a bagging-based algorithm with a key difference wherein only a subset of features is selected at random. Compared to RF, XGBoost is normally used to train gradient-boosted decision trees and other gradient boosted models, whereas RF uses the same model representation and inference (as gradient-boosted decision trees), but a different training algorithm. Moreover, XGBoost supports *missing values*, as branch directions for missing values are learned during training.

In practice, it requires the right configuration of the algorithm for a dataset by tuning the *hyperparameters*. Most of them are related to the bias-variance trade-off. When one allows the model to become more complicated (e.g., more depth), the model has a

<sup>20</sup> By applying the second fundamental theorem of calculus (or the Newton-Leibniz axiom) to the sum of quantiles, the Riemann integral is obtained in the limit  $P \rightarrow \infty$  (see [Apostol 1967](#)) and the partitions  $\Delta \tau_p = \frac{1}{P+1}$  get finer (i.e.,  $\Delta \tau_p \rightarrow 0$  as long as  $P \rightarrow \infty$ ).

<sup>21</sup> We rely on the fact that the conditional quantiles are consistently estimated using the QRF approach.

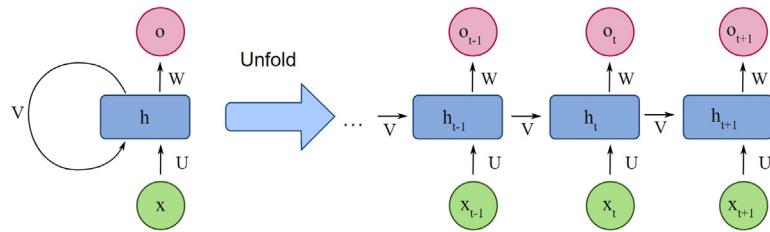


Fig. 3. Basic representation of a RNN. Source: [https://en.wikipedia.org/wiki/Recurrent\\_neural\\_network](https://en.wikipedia.org/wiki/Recurrent_neural_network).

better ability to fit the training data (in-sample), resulting in a less biased model. However, such a complicated model requires more data. The best model should trade the model complexity with its predictive power carefully. See [Chen and Guestrin \(2016\)](#) for further details.<sup>22</sup>

#### 2.2.5. Recurrent neural network

A recurrent neural network (RNN) is a class of artificial neural networks commonly used for time series data ([Elman, 1990](#)). RNNs are highly *nonlinear* models that use training data to learn and represent complex dynamic relationships between variables. They are distinguished by their internal state (memory), as they make use of sequential information to capture long-term temporal dependencies between input variables and the output (dependent variable). While traditional deep neural networks assume that inputs and outputs are independent of each other, the RNN output depends on the prior elements within the sequence and, thus, can exhibit dynamic temporal behavior.<sup>23</sup> See [Dupond \(2019\)](#) for further details.

[Fig. 3](#) shows a basic scheme of the RNN, where  $x_t$  is the input vector of predictors,  $o_t$  is the output vector (dependent variable),  $h_t$  is the hidden layer (a set of neurons) and  $W$  and  $U$  are model parameter matrices. The left graph illustrates the recurrent connections as the arc labeled “V”. The graph on the right unfolds the RNN across time, pointing out that RNN is a class of neural network that exhibits temporal dynamic behavior.

The RNN architecture is essentially driven by the number of *hidden layers*  $h_t$  that control the overall model complexity<sup>24</sup> and the *activation function* (or basis function) that determines whether a given neuron in the network should be activated. These nonlinear functions typically convert the output of a given neuron to a value between 0 and 1 (or -1 and 1). Here, we adopt the logistic (or sigmoid) function, given by  $g(x) = 1/(1 + e^{-x})$ , as the activation function.<sup>25</sup>

#### 2.2.6. Hybrid machine learning

The use of hybrid models in ML to forecast macroeconomic variables is relatively recent (e.g., [Smyl 2020](#)). Here, the idea is to mix *methodologies* to produce a hybrid inflation forecasting method, which allows investigating whether the forecasting accuracy of a given ML model is primarily due to variable selection or to potential nonlinearities in the inflation data-generating process. Inspired by the hybrid method of [Medeiros et al. \(2021\)](#), we propose the following three-step approach:

1. Pre-selection of variables using the *target predictor* procedure of [Bai and Ng \(2008\)](#). First, we separately run individual OLS regressions (i.e., we run the dependent variable onto the intercept and a given individual predictor). Then, each predictor is retained in the set of selected variables only if it is statistically significant at a 5% level, based on its respective OLS regression;
2. Choice of the most relevant variables (e.g., Top20), according to either the adalasso or the RF (impurity corrected) approaches; and
3. Use of the set of top variables as input (e.g., regressors) in the following models: OLS, adalasso, RF, or XGBoost in order to forecast the inflation rate.

In this way, we build six hybrid models, labelled as Ada-OLS, Ada-RF, Ada-XGB, RF-OLS, RF-Ada or RF-XGB, where *OLS* denotes Ordinary Least Squares, *Ada* means adalasso, *RF* indicates random forest, and *XGB* refers to XGBoost. The first nickname stands for the variable selection method, whereas the second one denotes the estimation/forecasting approach. For example, the hybrid model Ada-OLS employs the adalasso to choose the top predictors, which are then used as regressors in the OLS model to forecast inflation.

Although the set of nonlinear approaches could be further enlarged, we believe that the six hybrid models designed here can help us check the importance of variable selection, as part of many ML methodologies, and the role that nonlinearities play when

<sup>22</sup> For more detail, also see <https://towardsdatascience.com/a-guide-to-xgboost-hyperparameters-87980c7f44a9> and <https://xgboost.readthedocs.io/en/latest/index.html>.

<sup>23</sup> RNNs can have additional storage states, which may incorporate feedback loops. Such extra states are referred to in the literature as gated states (or gated memory), and are part of long short-term memory networks (LSTMs) and gated recurrent units (GRUs). Here, we employ the standard RNN setup to save computational time and avoid unnecessary model complexity (overfitting).

<sup>24</sup> The so-called *fully* recurrent neural networks connect the outputs to the inputs of all neurons. This is the most general topology, since all others can be represented by setting some connection weights to zero to simulate the lack of connections between neurons. By considering more than one layer, RNNs are also part of a machine learning field called *deep learning*.

<sup>25</sup> Neural networks have been found to be very successful in complex settings, with a large number of features. However, they often require a substantial amount of data in order to work well in practice.

modeling inflation dynamics. We also bring to ML a flavor of model interpretation by first revealing the top features and then using this superior dataset as input in a given ML forecasting method.

### 2.2.7. Traditional inflation forecasting

Next, we analyze more traditional forecasting methods that are often used by economic agents when producing inflation forecasts.

**Random walk:** The standard random walk (RW) model assumes that the  $h$ -period inflation change is an unforecastable martingale difference sequence (MDS), i.e.,  $\mathbb{E}(y_{t+h} - y_t | \mathcal{F}_t) = 0$ . The out-of-sample inflation forecast, for all  $h = 1, \dots, H$ , is then  $f_{y_{T_2+h}}^{rw} = y_{T_2}$ .

**RW-AO:** This is the variant of the RW model considered by Atkeson and Ohanian (2001), which takes the average inflation over the previous 12 months<sup>26</sup> as the forecast for  $y_{T_2+h}$ , as follows:  $f_{y_{T_2+h}}^{rw-ao} = \frac{1}{12} \sum_{j=0}^{11} y_{T_2-j}$ .

**ARMA:** One of the most common statistical models used in time-series forecasting is the autoregressive moving average (ARMA) model, which assumes that future observations are essentially driven by recent observations. Inflation, which often exhibits persistent behavior, is largely consistent with this assumption.<sup>27</sup>

**VAR:** The vector autoregression (VAR) model is also a traditional method based on a backward-looking approach. Here we use one lag<sup>28</sup> and the following endogenous variables: market price inflation, administered price inflation, 1-year real interest rate, log-difference of nominal exchange rate (BRL/USD), and output gap (proxied by the HP-filtered IBC-BR series).<sup>29</sup> The choice of variables recognizes different time dynamics of the two main components of inflation in Brazil<sup>30</sup> and incorporates the pass-through of imported inflation to domestic inflation. The forecast of headline inflation is built by aggregating the  $h$ -step ahead forecasts of the two inflation components using the respective weights.<sup>31</sup>

**Inflation expectations (Focus):** The Focus survey of professional forecasts is a panel database put together by the Central Bank of Brazil (*Banco Central do Brasil – BCB*) that has collected daily information since 1999, after the implementation of the inflation-targeting regime in Brazil. The survey covers more than 100 professional forecasters (e.g., banks, asset management firms, consulting firms, relevant nonfinancial institutions), which are followed throughout time. The survey is constantly used by market agents, specialized media and the BCB itself to monitor inflation expectations. The forecasts are supplied over various horizons and for a large array of macroeconomic series (see Gaglianone et al. 2022). Here, we consider the median of individual inflation forecasts (IPCA) across survey participants.

**Inflation expectations (BEI):** Another key source of inflation expectations in Brazil is financial data. For instance, using inflation-linked treasury bond market data<sup>32</sup> and the parametric model of Svensson (1994), one can extract the so-called *breakeven inflation* (BEI), which is available on a daily basis and focused on actual financial market agents' decisions (see Val and Araujo 2019). However, the usage of such data usually embodies risk premium issues and maturities with different market liquidity. Since there is no consensus in the finance literature on how to properly compute risk premium, we do not extract it from the BEI series. Besides, it is often neglected for short horizons, although it is usually relevant for longer horizons.<sup>33</sup>

**PC-backward:** The Phillips curve model (PC) has a long tradition in forecasting inflation (Stock and Watson, 1999). We consider here a backward-looking version of the curve, only including past inflation (inertia), imported inflation (pass-through channel)<sup>34</sup> and output gap (the traditional monetary policy channel via aggregate demand).<sup>35</sup> Following the VAR approach, we also disaggregate inflation into two components (market price inflation and administered price inflation), which are modeled separately. First, we estimate a PC for inflation of market prices. Then, we estimate an ARMA( $p, q$ ) model for the administered price inflation. Finally, the headline inflation forecast is built by aggregating the forecasts of the two inflation components using the respective weights.

**PC-hybrid:** This approach considers a hybrid (New Keynesian) version of the PC, which includes backward- and forward-looking terms, imported inflation and output gap (see Arruda et al. 2011, and Gaglianone et al. 2017). The extra term is expected inflation,

<sup>26</sup> According to Atkeson and Ohanian (2001, p.10), “economists have not produced a version of the Phillips curve that makes more accurate inflation forecasts than those from a naive model that presumes inflation over the next four quarters will be equal to inflation over the last four quarters.”

<sup>27</sup> The best ARMA( $p, q$ ) model is recursively selected using the Schwarz information criterion.

<sup>28</sup> According to the Schwarz information criterion and diagnostic testing.

<sup>29</sup> See variables 2, 3, 27, 51, and 120 in Appendix A.

<sup>30</sup> The administered price inflation is in some way regulated by a public agency or set by contracts (often including backward indexation clauses) rather than by the interaction between domestic demand and supply conditions. According to Minella et al. (2003), the dynamics of such prices differ from the market prices in three ways: (i) dependence on international prices in the case of refined petroleum products; (ii) greater pass-through from the exchange rate; and (iii) stronger backward-looking behavior.

<sup>31</sup> For details on IPCA weights, see <https://www.bcb.gov.br/content/ri/inflationreport/201912/ri201912b7i.pdf>

<sup>32</sup> We use data from NTN-B, which is an acronym for *Nota do Tesouro Nacional*, type B, similar to Treasury Inflation-Protected Securities (TIPS) in the U.S.

<sup>33</sup> We use end-of-month data when considering inflation expectations at a monthly frequency. For example, when the goal is to forecast the IPCA of June 2021, for  $h = 1$ , we use the inflation expectations available on May 31, 2021. Similarly, in order to forecast the IPCA of June 2021, for  $h = 2$ , we use the Focus and BEI data available on April 30, 2021.

<sup>34</sup> Defined as the sum of the nominal exchange rate (R\$/US\$) monthly percentage variation and the U.S. inflation (assumed, for simplicity, to be 2.0% per year).

<sup>35</sup> The output gap is based on the seasonally adjusted IBC-BR index of economic activity. The Hodrick-Prescott (HP) filter is employed to generate the output gap in a recursive estimation scheme; that is, we reconstruct the entire output gap series for each new observation added to the estimation sample along the out-of-sample exercise (and then re-estimate the PC to build new forecasts).

proxied here by the *Focus* survey. We impose the usual coefficient restriction<sup>36</sup> to guarantee a vertical long-run PC. The forecasts for administered price inflation and headline inflation follow the same procedures described in the previous approach.

**Factor model 1 (direct forecast):** The idea that time variations in a large number of variables can be summarized by a small number of factors is empirically attractive, and it is employed in several studies in economics and finance (see [Forni et al. 2000](#), and [Stock and Watson 2002](#)). Let  $x_{i,t} \in \tilde{x}_t$  be the observed data for the  $i$ -th cross-section unit at time  $t$ , for  $i = 1, \dots, N$  and  $t = 1, \dots, T_2$ , and consider the following factor representation of the data:  $x_{i,t} = \lambda'_i F_t + e_{i,t}$ , where  $F_t$  is a vector of common factors,  $\lambda_i$  is a vector of factor loadings associated with  $F_t$ , and  $e_{i,t}$  is the idiosyncratic component of  $x_{i,t}$ .

Note that  $\lambda_i$ ,  $F_t$  and  $e_{i,t}$  are unknown because only  $x_{i,t}$  is observable. Here, we estimate the factors and respective loadings using principal components analysis (PCA). The number of components is determined by the [Bai and Ng \(2002\)](#) criterion. After the PCA estimation of the common factors  $F_t$ , we employ the *direct forecast* approach to model the inflation rate at time  $t + h$  as follows:  $y_{t+h} = \beta_h F_t + \epsilon_{t+h}$ .

Therefore, the inflation forecast from the (direct) factor model above, using a sample of  $t = 1, \dots, T_2$  observations, is given by  $f_{y_{T_2+h}}^{fm\text{-direct}} = \widehat{\beta}_h \widehat{F}_{T_2}$ , for  $h = 1, \dots, H$ .

**Factor model 2 (iterated forecast):** This approach is a variant of the previous one, but it uses an iterated forecast method instead of the direct forecast approach. The idea is, again, to employ common factors, but to model the inflation rate in a contemporaneous way with respect to the factors, i.e.,  $y_t = \gamma F_t + v_t$ .

Following the literature (e.g., [Bañbura et al. 2013](#)), we specify the factors as following a VAR process, i.e.,  $F_t = \Phi(L)F_t + u_t$ . The inflation forecast from this *iterated* factor model, using a sample of  $t = 1, \dots, T_2$  observations, is given by  $f_{y_{T_2+h}}^{fm\text{-iterated}} = \widehat{\gamma} \widehat{F}_{T_2+h|T_2}$ , for  $h = 1, \dots, H$ , where  $\widehat{F}_{T_2+h|T_2}$  are the  $h$ -step ahead (out-of-sample) forecasts of the common factors, using the VAR model estimated in a recursive scheme.

**Factor models 3 and 4 (with targeted predictors, direct or iterated):** These are the previous factor models, but now based on a *subset* of predictors that are selected by taking into account the fact that our variable of interest is the inflation rate. Here, we follow the idea of [Bai and Ng \(2008\)](#), who showed that the factor model's out-of-sample forecasting performance could be improved by selecting (or targeting) the predictors in advance.

The core idea is that irrelevant predictors employed to build a factor model only add noise into the analysis and thus produce factors with a poor predictive performance. In this sense, we use only pre-selected variables in the factor model, as follows:

- (i) in the direct forecast case, we first regress the inflation rate  $y_{t+h}$  (or  $y_t$  in the iterated case) on the intercept and the candidate variable  $x_{i,t} \in \tilde{x}_t$ , for all  $i = 1, \dots, N$ ;
- (ii) calculate the  $t$ -statistics for the coefficient associated to  $x_{i,t}$ ;
- (iii) include  $x_{i,t}$  in the set of predictors (used to extract the factors) only if it is statistically significant at a 5% level; and
- (iv) proceed as before, in the direct or iterated cases, to build the respective forecasts.

#### 2.2.8. Disaggregated forecasts

According to the [Brasil, 2021a](#), the three main groups of market prices (services, industrial goods, and food at home) show important differences in terms of average inflation level, dynamics and determinants.<sup>37</sup> Taking into account such differences in the data-generating process of the inflation subgroups can (potentially) lead to accuracy gains when forecasting the aggregated inflation.<sup>38</sup> This motivates a *bottom-up* approach, in which we separately model and forecast inflation for each one of the main three subgroups of market prices, besides administered prices. To do so, we employ three models: ARMA, adalasso, and RF. Then, we aggregate<sup>39</sup> the four individual forecasts (administered prices, services, industrial goods, and food at home) to form the forecast of headline inflation (IPCA).

#### 2.2.9. Forecast combination

Since the seminal work of [Bates and Granger \(1969\)](#), it has been observed that combining forecasts across multiple models often produces better forecasts compared to a single model. Nowadays, the accuracy gains of forecast combination over individual forecasts are well documented in the literature. According to [Elliott et al. \(2015\)](#), “*forecast combination offers one approach for dealing with the effects of estimation error, model uncertainty, and instability in the underlying data generating process. By diversifying across multiple models, combinations typically deliver more stable forecasts than those associated with individual models.*”

Here, we employ different forecast combination methods based on three main sets of individual forecasts. The first set of forecasts (set1) entails individual forecasts from models 1–27. The second and third sets of forecasts (set2 and set3) include only the superior models of the model confidence set (MCS) proposed by [Hansen et al. \(2011\)](#), considering models 1–27 or 1–29, respectively.

The MCS consists of a sequence of tests allowing the construction of a set of superior models, where the null hypothesis of equal predictive ability is not rejected at a given confidence level. The test statistic can be evaluated for several loss functions, such as mean

<sup>36</sup> The sum of coefficients on past inflation, expected inflation and imported inflation must be equal to one.

<sup>37</sup> The Inflation Report says that “*disaggregated approaches are useful for extending the scope of the analysis and broadening the understanding of inflation developments and its prospects.*” According to the report, the greatest inertial component is related to the services sector, which is not directly impacted by exchange rate or commodity prices. In fact, these factors are relevant in the sectors of industrial goods and food at home.

<sup>38</sup> Furthermore, in the equity-premium literature, [Ferreira and Santa-Clara \(2011\)](#) report that forecasting the three components of stock market returns (dividend yield, earnings growth, and price-earnings ratio growth) separately yields huge forecasting gain.

<sup>39</sup> Using real-time weights of each IPCA subgroup.

squared error (MSE) or mean absolute error (MAE). The MCS is a sequential testing procedure that eliminates the worst model at each step until the hypothesis of equal predictive ability is accepted for all the models belonging to a set of superior models. The MCS method is focused not on the selection of optimal weights but on the selection of superior models. In other words, it trims out the worst performing models based on a statistical significance test. We implement the MCS procedure considering the MSE loss function and the 95% confidence level (see [Shang and Haberman 2018](#), for further details).

Thus, based on either set1, set2, or set3 of forecasts, we use the following forecast combination approaches: mean, median, adalasso, RF, [Granger and Ramanathan \(1984\)](#), constrained least squares (CLS), and complete subset regression (CSR). The first two combination approaches are simply the mean and median forecasts computed across a given set of individual forecasts (set1, set2, or set3). The adalasso and RF approaches are used here as forecast combination devices (i.e., instead of using a set of *predictors*, they are now based on a given set of individual *forecasts*).

In turn, [Granger and Ramanathan \(1984\)](#) set out the foundations of optimal forecast combinations under symmetric and quadratic loss functions. The authors show that, under MSE loss, the optimal weights can be estimated through an ordinary least squares (OLS) regression of the target variable (inflation, in our case) on a given set of forecasts, plus an intercept to account for possible model bias. However, if the loss function differs from the MSE, then the computation of optimal weights may require methods other than a simple OLS.

Although the OLS combination of [Granger and Ramanathan \(1984\)](#) enables us to correct for bias through its intercept term, [Nowotarski et al. \(2014\)](#) point out that such unbiasedness comes at the expense of poorer performance for highly correlated regressors. The CLS approach advocated by [Nowotarski et al. \(2014\)](#) is a variant of the previous OLS combination with additional constraints: no intercept term is imposed and the coefficients have to be non-negative and be summed up to 1.

In a distinct approach, [Elliott et al. \(2013, 2015\)](#) propose a method for combining forecasts based on CSR. The method combines forecasts based on predictive regressions with  $k$  number of predictors (in our setup, a given set of individual forecasts).<sup>40</sup> Hence, assuming  $k = 1$  corresponds to an equal-weighted average of all possible forecasts from univariate prediction models, whereas  $k = 2$  corresponds to equal-weighted averages of all possible forecasts from bivariate prediction models. The CSR approach has the computational advantage that it can be applied even when the number of predictors exceeds the sample size.<sup>41</sup>

### 2.3. Fan chart

Providing confidence intervals is a relevant issue when building point forecasts. Our setup of competing forecasts can be easily adapted to provide a measure of uncertainty around future predictions of inflation. To summarize the idea, we first compute the forecast errors of a given method of interest for a set of horizons  $h = 1, \dots, H$ . Next, we estimate the forecast variance (for each  $h$ ) and smooth out these variances using a *spline* function to obtain a *smooth* term structure of variances along the horizons  $h$ . Finally, we generate the out-of-sample conditional quantiles for a grid of quantile levels  $\tau$  assuming a Gaussian distribution.<sup>42</sup>

Thus, in order to produce density forecasts of the inflation rate  $y_{t+h}$ , using the information set  $\mathcal{F}_t$  available at period  $t$ , we assume the conditional distribution of  $y_{t+h}$  is Gaussian, with conditional mean  $\mu_{t+h|t}$  and conditional variance  $\sigma_{t+h|t}^2$ , i.e.,  $(y_{t+h} | \mathcal{F}_t) \sim N(\mu_{t+h|t}, \sigma_{t+h|t}^2)$ . The conditional quantile of  $y_{t+h}$ , evaluated at quantile level  $\tau \in (0, 1)$ , is given as follows:

$$Q_\tau(y_{t+h} | \mathcal{F}_t) = \mu_{t+h|t} + \sigma_{t+h|t} \Phi^{-1}(\tau). \quad (12)$$

Now, let  $f_{t+h|t}^m$  be the model  $m$  estimate of the conditional mean of  $y_{t+h}$ . Thus,  $f_{t+h|t}^m = \widehat{\mu_{t+h|t}}$ , where  $\mu_{t+h|t} = \mathbb{E}(y_{t+h} | \mathcal{F}_t)$ . Also, let  $\widehat{\sigma_{t+h|t}^2}$  be the model  $m$  estimate of the conditional variance of  $y_{t+h}$ , i.e.,  $\sigma_{t+h|t}^2$ , computed using the [Newey and West's \(1987\)](#) HAC covariance matrix estimator from a regression of the forecast error of  $f_{t+h|t}^m$  on the intercept.<sup>43</sup>

Provided that  $[\widehat{\mu_{t+h|t}}, \widehat{\sigma_{t+h|t}^2}]$  are consistent estimates of  $[\mu_{t+h|t}, \sigma_{t+h|t}^2]$ , one can obtain consistent estimates of the conditional quantiles of  $y_{t+h}$ , along a grid of quantile levels  $\tau \in \Gamma$ , using equation (12). Therefore, the multi-step ahead density forecasts of  $y_{t+h}$  can be summarized by a *fan chart* graph, based on the estimated conditional quantiles, over the horizons  $h = 1, \dots, H$ , and the grid of quantile levels  $\tau \in \Gamma$ .<sup>44</sup>

<sup>40</sup> The optimal value of  $k$  can be determined from the covariance matrix of the potential regressors and, thus, can be selected recursively in time.

<sup>41</sup> According to [Elliott et al. \(2015\)](#), Monte Carlo simulations show that CSR offers a favorable bias-variance trade-off in the presence of many weak predictors. However, one drawback is that the number of regressions to be estimated increases very quickly for large datasets. In such cases, [Garcia et al. \(2017\)](#) adopt a pre-testing procedure similar to the *targeting predictors* approach of [Bai and Ng \(2008\)](#). In this paper, since the number of predictors is very large, we employ the CSR method with  $k = 2$  as a forecast combination device, based on a given set of individual forecasts, instead of considering subset regressions on all candidate variables.

<sup>42</sup> The objective here is not to produce a density forecast based on a more complex approach (e.g., allowing for asymmetry and fat tails), but only to attach a simple measure of uncertainty to the path of future inflation according to each model's past forecast errors.

<sup>43</sup> Keep in mind that the forecast error  $(f_{t+h|t}^m - y_{t+h})$  is computed along a (pseudo) out-of-sample forecasting exercise, i.e., considering  $t = t_{oos1}, \dots, t_{oosT}$  and a given  $h$ .

<sup>44</sup> See [Costa et al. \(2021\)](#) for a similar exercise of density forecasting of future oil prices.

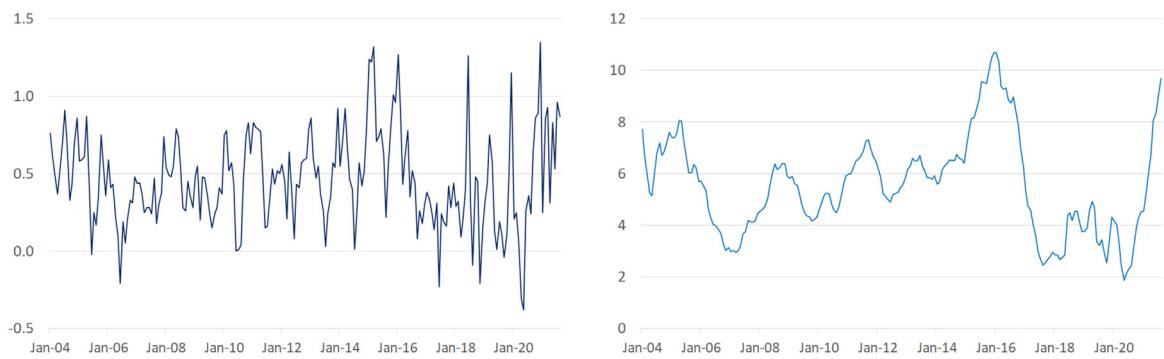


Fig. 4. Inflation rates (IPCA), % per month (left), % accumulated in 12 months (right). Source: Authors' own elaboration.

### 3. Empirical exercise

#### 3.1. Data

We focus the analysis on the IPCA, which is the consumer price index (CPI) measured by the Brazilian Institute of Geography and Statistics (IBGE), used to compute the official inflation measure and the target of monetary policy in Brazil. The dependent variable is either the monthly percentage change of the IPCA index or this measure accumulated over the last 12 months (12-month inflation). The forecast horizon ( $h$ ) varies from 1 to 18 months. The sample period spans over 17 years of data, from January 2004 to August 2021 ( $T = 212$  observations).<sup>45</sup> Fig. 4 shows the IPCA inflation in our sample period,<sup>46</sup> which starts a decade after the Brazilian monetary stabilization plan in mid-1994 and 5 years after the implementation of the inflation-targeting regime in mid-1999. Note that the inflation rate levels are comparable to the ones observed in many emerging market economies that adopt an inflation-targeting regime.

One of the key features driving the inflation dynamics in emerging economies is the degree of persistence (or inertia).<sup>47</sup> Besides past inflation, other predictors suggested in the literature for forecasting inflation often include economic slack measures (e.g., unemployment rate in a PC), variables related to production (Stock and Watson, 1999), financial variables (Forni et al., 2003), and surveys of expectations (Ang et al., 2007; Faust and Wright, 2013), among others.

In this paper, we use a diverse set of macroeconomic and financial variables drawn from a number of categories.<sup>48</sup> Our database consists of  $n = 167$  contemporaneous monthly variables, including, for instance, price indexes, interest rates, financial market variables, economic activity, labor market variables, government debt, imports and exports of goods and services, and international variables that are potentially related to the Brazilian economy.<sup>49</sup> The main data sources are Anbima, BCB, EPU (Baker et al., 2015),

<sup>45</sup> According to Machado and Portugal (2014), the limited sample problem is a well-known constraint for inference in Brazilian studies, particularly in inflation dynamics, where different policy regimes have been the case. In this sense, selecting the sample since 2004 helps us avoid large structural regime breaks. Indeed, the Bai and Perron (2003) structural break test using a HAC covariance matrix indicates no breakpoints in monthly inflation based on the sample since 2004.

<sup>46</sup> Inflation accumulated over 12 months fluctuates between 4.0% and 8.0% from 2004 until 2014, remaining within the inflation target tolerance interval during this period. In 2015, inflation rises due to administered price inflation and exchange rate depreciation, achieving its highest value of 10.7% by January 2016. In the 2016–2017 period, inflation gradually declines toward 2.5%, ending 2017 slightly below the lower limit of the tolerance interval, due to declines in prices of services and industrial goods and the behavior of food prices, dominated by record levels of agricultural production. After the impact of COVID-19 in Brazil in March 2020, inflation increases until the end of the sample due to the strong rise in prices of tradable goods in local currency, rise in electricity prices due to water scarcity, and demand-supply imbalances given the bottlenecks in global production chains.

<sup>47</sup> In Brazil, the relevance of past inflation has been vastly documented. For instance, Kohlschein (2012) suggests that models in which past inflation have greater weight in the expectations formation process are more accurate than others purely based on the assumption of rational expectations. In turn, Gaglianone et al. (2018) point out the relevance of considering a time-varying inertia when building more accurate inflation forecasting models.

<sup>48</sup> Besides the usual macro series, we included many financial variables that are shown in the literature (Forni et al., 2003) to be predictors that help forecast inflation. For example, financial market-based implied (breakeven) inflation, which provides a closer monitoring of inflation expectation (since it is updated on a continuously intra-day basis) and is competitive in terms of short-run predictive ability compared to survey expectations (Araujo and Vicente, 2017). We also included many non-standard variables in the database, for instance, to capture supply shocks arising from climate factors (e.g., amount of rainfall in Brazilian cities or even Pacific Ocean temperatures, to capture El Niño or La Niña effects).

<sup>49</sup> As a suggestion for future research, one could also include a Financial Conditions Index (FCI) in the database, besides the inflation target and the central bank inflation projections, as a way to provide exogenous, real-time relevant information from policymakers.

FGV, Funcex, IBGE, Inmet, IpeaData, and Reuters (Refinitiv Eikon Datastream). [Appendix A](#) presents the full list of variables used as potential predictors for the inflation rate in Brazil.

In order to ensure stationarity, we conduct an individual time series transformation, following the procedure adopted in the FRED-MD database of [McCracken and Ng \(2015\)](#). We consider six possibilities, as follows: (1) no transformation; (2)  $\Delta x_t$ ; (3)  $\Delta^2 x_t$ ; (4)  $\ln(x_t)$ ; (5)  $\Delta \ln(x_t)$ ; and (6)  $\Delta^2 \ln(x_t)$ . The transformation adopted for each series is presented in [Appendix A](#).

After transformations, the stationarity of each time series is checked using the Kwiatkowski-Phillips-Schmidt-Shin (KPSS) test, considering a 5% significance level. The  $n = 167$  contemporaneous variables are then lagged  $s = 3$  periods,<sup>50</sup> forming a final database containing 501 series. This way, besides an intercept, [Eq. \(1\)](#) entails  $\dim(\tilde{x}_t') = 501$  variables, used as potential predictors for inflation in Brazil.<sup>51</sup>

Recall that the first part of the sample ( $t = 1, \dots, T_1$ ) is used to estimate the models, whereas the second part of the sample ( $t = T_1 + 1, \dots, T_2$ ) is used for estimation of the forecast combination weights (where applicable). The remaining observations ( $t = T_2 + 1, \dots, T$ ) are reserved for genuine out-of-sample forecast comparison. We consider  $T_1 = 72$  months (6 years),  $T_2 = 120$  months (10 years), and  $P = T - T_2 = 212 - 120 = 92$  out-of-sample observations. Thus, the evaluation period for  $h = 1$  ranges from January 2014 to August 2021 (92 forecasts), whereas for  $h = 18$  it ranges from June 2015 to August 2021 (75 forecasts).<sup>52</sup>

All models are re-estimated every month in a *recursive* estimation scheme (i.e., expanding sample size) as we incorporate every new time-series observation, one at a time. In this context, each model is initially estimated using the first  $T_1$  observations (or  $T_2$  in the case of forecast combination), and the out-of-sample point forecasts are generated. We then add an additional observation at the end of the *training set*, re-estimate the models and generate out-of-sample forecasts again. This process is repeated for the remaining data (*test set*).<sup>53</sup>

We also conduct an extensive *robustness* analysis by considering two alternative training/test sets: (i)  $T_1 = 96$  months (8 years) and  $T_2 = 144$  months (12 years), and (ii)  $T_1 = 72$  months (6 years) and  $T_2 = 144$  months (12 years). The results are presented in [Appendix B](#).

The empirical exercise is implemented using R software (version 4.1.0, 64-bit).<sup>54</sup> The ridge regression, lasso and elastic net models are estimated using the R package *glmnet*, which fits a generalized linear model via penalized maximum likelihood. The adalasso model is implemented using the R package *HDeconometrics*. The same R package is used to compute the BIC information criterion for selection of hyperparameters. In order to implement the RF and the QRF methods,<sup>55</sup> we use the R package *ranger*, whereas the XGBoost approach is based on the R package *xgboost*. Finally, the R package *rnn* is used to implement the recurrent neural networks, and the R package *MCS* is used to implement the *model confidence set* procedure of [Hansen et al. \(2011\)](#).

### 3.2. Results

The observed inflation rate (% per month, or simply % p.m.) and the respective out-of-sample forecasts ( $h = 1$ ) of the 50 approaches covered in this paper are shown in [Fig. 5](#).<sup>56</sup>

For each horizon, the forecast errors are squared and averaged to form the out-of-sample MSE. In addition, we compute the *p*-value of the [Diebold and Mariano \(1995\)](#) test for non-nested models<sup>57</sup> using the forecasts from the ARMA model as a *benchmark*. Besides the MSE, we compare the best model with the benchmark, at each horizon, in terms of the  $R^2$  out-of-sample statistics ([Rapach et al., 2010](#)).<sup>58</sup> [Table 2](#) presents the results.

<sup>50</sup> Inflation models usually comprise a rich lag structure (particularly in emerging countries, more prone to inflation inertia). Such a structure should capture the dynamic relationship between inflation, past inflation and key macroeconomic variables. Here, we adopt 3 lags to avoid overfitting (besides, forecasting exercises with more lags generally produced higher MSEs).

<sup>51</sup> We standardize data (zero mean and unity variance) in the penalized-regression models (elastic net, lasso, adaptive lasso, and ridge), factor models, RNN and hybrid models. In turn, tree-based methods (RF, QRF, and XGBoost) do not require feature scaling (since common practice in such methods is not to standardize features, we follow this approach here).

<sup>52</sup> Our forecasting exercise is not implemented on a *strict* real-time basis to avoid unnecessary complications in the execution of the exercise. Practical limitations arise due to data revisions and/or various release delays of features in real time. The literature on inflation forecasting is generally focused on *pseudo* real-time exercises, such as the one conducted in this paper. Besides, the real-time issue loses relevance when considering longer horizons, such as 12 or 18 months.

<sup>53</sup> We adopt such an estimation scheme due to the greater efficiency, in general, of recursive regressions compared to rolling-window estimations. However, the latter approach could be justified under a setup with structural changes. See [Morales-Arias and Moura \(2013\)](#) for a good discussion on this issue.

<sup>54</sup> Unless indicated in the text, hyperparameter selection of ML methods is based on default software options. Detailed empirical results are available upon request.

<sup>55</sup> We used 1000 trees in both the RF and the QRF. In the latter method, we adopted the following grid of quantile levels:  $\tau \in (0.05, 0.10, \dots, 0.95)$ .

<sup>56</sup> Forecasts of the inflation rate accumulated in 12 months are not shown here to save space but are available upon request.

<sup>57</sup> The null hypothesis assumes the equal forecasting accuracy of two competing forecasts. The variances entering the test statistics use the [Newey and West \(1987\)](#) HAC covariance estimator.

<sup>58</sup> It is defined as follows:  $R^2_{oos} = 100 \times \left( 1 - \left( \frac{\sum_{t=T_2+1}^T (y_{t+h} - \hat{f}_{t+h|t})^2}{\sum_{t=T_2+1}^T (y_{t+h} - \hat{f}_{t+h|t}^{BMR})^2} \right) \right)$ , where  $\hat{f}_{t+h|t}$  is the forecast of  $y_{t+h}$  from model  $i$  using information up to period  $t$  and  $\hat{f}_{t+h|t}^{BMR}$  is the benchmark forecast. Positive (negative) values for  $R^2_{oos}$  means forecast  $\hat{f}_{t+h|t}$  beats (is beaten by)  $\hat{f}_{t+h|t}^{BMR}$ .

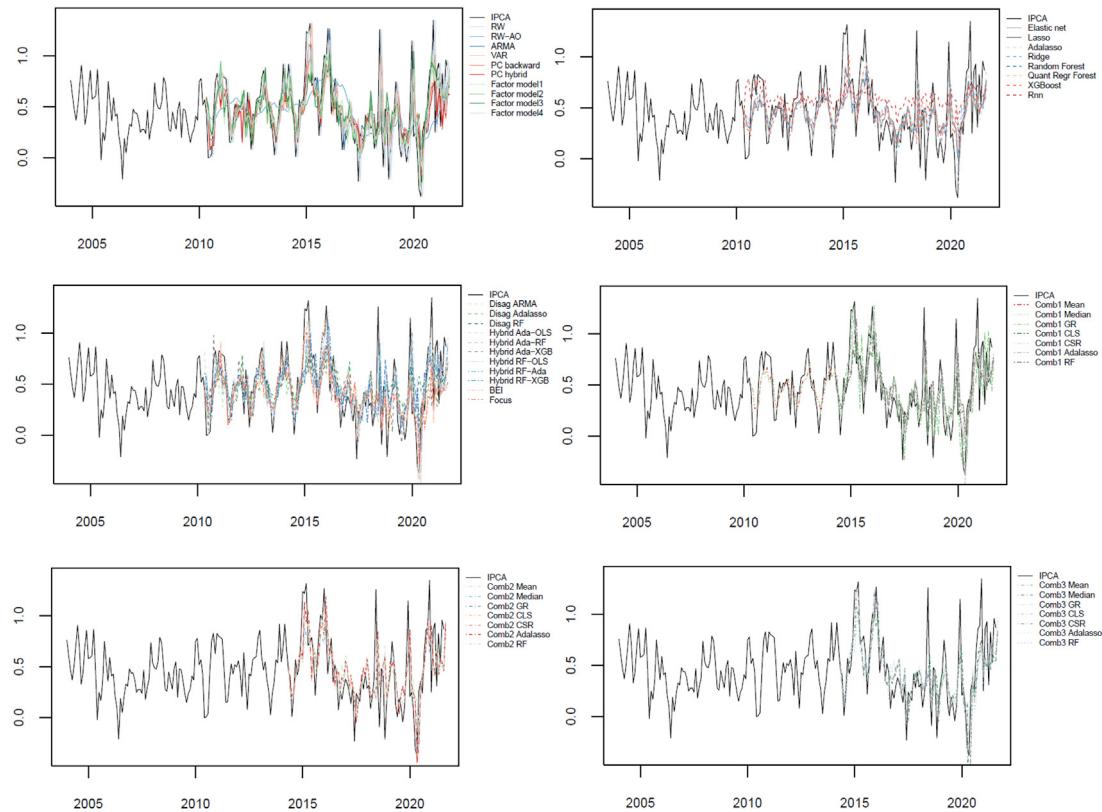


Fig. 5. Inflation rate (% p.m.) and forecasts,  $h = 1$ . Source: Authors' own elaboration.

Considering the monthly inflation rate (left panel in Table 2), at the shortest horizon ( $h = 1$ ), the best model is model 46 (comb3 GR), which provides an accuracy gain of 47%, in terms of the  $R^2_{oos}$  statistic, compared to the ARMA model. Such an expressive result is statistically significant (at the 1% level, using the Diebold-Mariano test), and it is achieved by using the MCS of Hansen et al. (2011) on the largest set of forecasts (i.e., including BEI and Focus) and the Granger and Ramanathan (1984) method to estimate the weights used to combine the MCS *superior* forecasts. The top forecasts at  $h = 1$  also include BEI and Focus, besides other forecast combinations using the third set of forecasts (comb3).<sup>59</sup>

For longer horizons, the accuracy gains of the best models over the benchmark decrease, ranging from 10% to 27%. Among the ML methods, it is worth highlighting the good performance of the regression tree-based methods (RF, QRF<sup>60</sup> and XGBoost), in particular, for longer horizons ( $h \geq 12$ ). The recurrent neural network (RNN) also shows a good result at longer horizons.<sup>61</sup> These results reflect the importance of nonlinear methods when modeling the inflation dynamics in Brazil.

On the other hand, traditional inflation forecasting (linear) models, such as ARMA and VAR, never enter the set of Top10 forecasts (yellow cells in Table 2) at any horizon. Results for the *disaggregated* forecasts are a bit disappointing,<sup>62</sup> since they beat the benchmark (ARMA) forecast in just a few cases. The results for the *hybrid* models seem to be a little more promising at some horizons.<sup>63</sup>

<sup>59</sup> Both Focus and BEI have the advantage of incorporating recent information about events that may affect future inflation, such as currency appreciation or devaluation, crop failures, relevant strikes, etc. In addition, these measures also take into account the future dynamics of the central bank's inflation target.

<sup>60</sup> For medium/long horizons, note that QRF shows a non-negligible accuracy improvement over RF, which is due to the role quantiles play at improving conditional mean forecasts.

<sup>61</sup> Deep learning models usually stand out when based on a large database, which is not necessarily the case in our monthly series setup.

<sup>62</sup> Here, the gains of separately forecasting inflation components seem not to offset model misspecification and parameter uncertainty, among other things, when estimating multiple individual models.

<sup>63</sup> In both inflation rates, Ada-OLS usually performs worse than adalasso (and RF-OLS worse than RF), suggesting nonlinearities are important in forecasting inflation. When comparing adalasso with RF, there is no clear indication of the best variable selection method (adalasso dominates RF-Ada in many horizons, whereas RF usually dominates Ada-RF for  $h \geq 6$ ). However, when including XGBoost, the results become clear for the 12-month inflation: XGBoost dominates Ada-XGB and RF-XGB in all horizons (similar results are obtained in many cases with monthly inflation), indicating XGBoost is a strong variable selection method.

Table 2

Mean Squared Error (MSE).

| dep. var. = IPCA % p.m.  | h = 1    | h = 2    | h = 3 | h = 6 | h = 9  | h = 12  | h = 15  | h = 18 | dep. var. = IPCA % 12 months | h = 1    | h = 2    | h = 3   | h = 6  | h = 9  | h = 12 | h = 15 | h = 18 |
|--------------------------|----------|----------|-------|-------|--------|---------|---------|--------|------------------------------|----------|----------|---------|--------|--------|--------|--------|--------|
| (1) RW                   | 0.129    | 0.187    | 0.200 | 0.278 | 0.204  | 0.255   | 0.305   | 0.279  | (1) RW                       | 0.259    | 0.795    | 1.448   | 3.854  | 6.551  | 10.059 | 11.966 | 13.254 |
| (2) RW-AO                | 0.124    | 0.137    | 0.146 | 0.164 | 0.185  | 0.194   | 0.197   | 0.200  | (2) RW-AO                    | 3.238    | 4.106    | 4.983   | 7.513  | 9.925  | 11.948 | 13.169 | 13.821 |
| (3) ARMA                 | 0.104    | 0.137    | 0.138 | 0.157 | 0.165  | 0.157   | 0.172   | 0.173  | (3) ARMA                     | 0.189    | 0.674    | 1.268   | 3.829  | 6.579  | 11.005 | 14.012 | 17.600 |
| (4) VAR                  | 0.107    | 0.138    | 0.141 | 0.157 | 0.156  | 0.145   | 0.147*  | 0.150  | (4) VAR                      | 0.268    | 0.844    | 1.574   | 4.460  | 7.755  | 11.917 | 14.411 | 15.237 |
| (5) PC backward          | 0.102    | 0.136    | 0.143 | 0.162 | 0.153  | 0.139*  | 0.149   | 0.172  | (5) PC backward              | 0.233    | 0.734    | 1.334   | 3.669  | 5.664  | 8.904  | 12.134 | 15.115 |
| (6) PC hybrid            | 0.082*** | 0.115**  | 0.123 | 0.142 | 0.146* | 0.146   | 0.146   | 0.159  | (6) PC hybrid                | 0.228    | 0.697    | 1.240   | 3.558  | 5.897  | 9.053  | 12.642 | 16.194 |
| (7) Factor model1        | 0.088**  | 0.113**  | 0.139 | 0.156 | 0.164  | 0.145   | 0.148*  | 0.150  | (7) Factor model1            | 0.768    | 1.306    | 1.922   | 4.081  | 6.890  | 9.305  | 8.832  | 8.476  |
| (8) Factor model2        | 0.085*** | 0.117*   | 0.133 | 0.154 | 0.154  | 0.139*  | 0.142*  | 0.146* | (8) Factor model2            | 0.515    | 0.997    | 1.574   | 3.814  | 6.305  | 9.131  | 10.262 | 10.384 |
| (9) Factor model3        | 0.088*   | 0.108*** | 0.133 | 0.155 | 0.148* | 0.150   | 0.141** | 0.158  | (9) Factor model3            | 0.439    | 0.993    | 1.747   | 4.720  | 8.157  | 8.707  | 7.887  | 7.066  |
| (10) Factor model4       | 0.094    | 0.120**  | 0.140 | 0.154 | 0.153  | 0.141   | 0.141*  | 0.146* | (10) Factor model4           | 0.408    | 0.899    | 1.456   | 3.470  | 6.205  | 9.960  | 11.569 | 11.382 |
| (11) Elastic net         | 0.092**  | 0.129    | 0.149 | 0.152 | 0.153  | 0.145   | 0.141*  | 0.150  | (11) Elastic net             | 0.253    | 0.854    | 1.662   | 4.474  | 9.741  | 11.590 | 11.326 | 7.789  |
| (12) Lasso               | 0.092*   | 0.128    | 0.151 | 0.152 | 0.153  | 0.146   | 0.142*  | 0.150  | (12) Lasso                   | 0.251    | 0.814    | 1.648   | 4.334  | 9.492  | 11.603 | 11.627 | 7.734  |
| (13) Adalasso            | 0.088*** | 0.119*   | 0.152 | 0.153 | 0.157  | 0.153   | 0.148   | 0.149  | (13) Adalasso                | 0.220    | 0.785    | 1.568   | 5.050  | 10.209 | 14.386 | 17.436 | 12.252 |
| (14) Ridge               | 0.100    | 0.122**  | 0.139 | 0.149 | 0.161  | 0.147   | 0.153   | 0.155  | (14) Ridge                   | 1.211    | 1.748    | 2.442   | 4.902  | 7.448  | 8.386  | 8.195  | 7.154  |
| (15) Random Forest       | 0.105    | 0.124    | 0.141 | 0.151 | 0.154  | 0.139*  | 0.139*  | 0.155  | (15) Random Forest           | 0.936    | 1.667    | 2.432   | 4.807  | 7.158  | 8.106  | 7.731  | 6.928  |
| (16) Quant. Regr. Forest | 0.106    | 0.126    | 0.143 | 0.150 | 0.153  | 0.138*  | 0.136** | 0.150  | (16) Quant. Regr. Forest     | 0.929    | 1.677    | 2.459   | 4.812  | 7.099  | 8.022  | 7.569  | 6.738  |
| (17) XGBoost             | 0.098    | 0.116**  | 0.143 | 0.161 | 0.159  | 0.146   | 0.133** | 0.172  | (17) XGBoost                 | 0.291    | 0.873    | 1.657   | 4.681  | 7.110  | 7.751  | 6.867  | 5.881  |
| (18) RNN                 | 0.134    | 0.134    | 0.144 | 0.152 | 0.143* | 0.136*  | 0.128*  | 0.146  | (18) RNN                     | 3.079    | 4.492    | 3.190   | 8.099  | 5.560  | 9.873  | 10.701 | 8.591  |
| (19) Disag. ARMA         | 0.110    | 0.144    | 0.153 | 0.168 | 0.174  | 0.168   | 0.169   | 0.162  | (19) Disag. ARMA             | 0.186    | 0.654    | 1.256   | 3.774  | 6.392  | 10.531 | 12.405 | 14.232 |
| (20) Disag. Adalasso     | 0.097    | 0.135    | 0.161 | 0.143 | 0.161  | 0.146   | 0.146*  | 0.141* | (20) Disag. Adalasso         | 0.234    | 0.833    | 1.970   | 5.245  | 10.187 | 10.668 | 12.970 | 12.201 |
| (21) Disag. RF           | 0.114    | 0.133    | 0.144 | 0.156 | 0.152  | 0.141** | 0.148   | 0.164  | (21) Disag. RF               | 0.982    | 1.781    | 2.623   | 5.132  | 7.140  | 7.967  | 8.092  | 7.879  |
| (22) Hybrid Ada-OLS      | 0.090*** | 0.120*   | 0.139 | 0.155 | 0.172  | 0.174   | 0.127** | 0.167  | (22) Hybrid Ada-OLS          | 0.209    | 0.729    | 1.509   | 4.513  | 11.412 | 14.469 | 14.719 | 9.173  |
| (23) Hybrid Ada-RF       | 0.100    | 0.112**  | 0.135 | 0.168 | 0.170  | 0.157   | 0.127** | 0.165  | (23) Hybrid Ada-RF           | 0.561    | 1.244    | 1.974   | 4.647  | 7.444  | 9.297  | 9.089  | 8.279  |
| (24) Hybrid Ada-XGB      | 0.107    | 0.120    | 0.154 | 0.219 | 0.201  | 0.172   | 0.145   | 0.185  | (24) Hybrid Ada-XGB          | 0.379    | 1.043    | 1.918   | 5.126  | 8.049  | 9.478  | 8.775  | 6.883  |
| (25) Hybrid RF-OLS       | 0.101    | 0.120*   | 0.213 | 0.203 | 0.196  | 0.208   | 0.208   | 0.172  | (25) Hybrid RF-OLS           | 0.182    | 0.746    | 1.438   | 5.200  | 16.010 | 17.836 | 20.383 | 10.608 |
| (26) Hybrid RF-Ada       | 0.092*   | 0.117**  | 0.187 | 0.173 | 0.160  | 0.157   | 0.190   | 0.157  | (26) Hybrid RF-Ada           | 0.188    | 0.790    | 1.658   | 4.788  | 16.140 | 18.435 | 18.502 | 10.201 |
| (27) Hybrid RF-XGB       | 0.093    | 0.112**  | 0.147 | 0.158 | 0.174  | 0.153   | 0.125** | 0.175  | (27) Hybrid RF-XGB           | 0.319    | 1.116    | 2.166   | 4.932  | 7.886  | 8.412  | 7.207  | 6.626  |
| (28) BEI                 | 0.067**  | 0.140    | 0.126 | 0.150 | 0.161  | 0.154   | 0.147   | 0.157  | (28) BEI                     | 0.072*** | 0.396**  | 0.636** | 2.412* | 4.848* | 7.239  | 7.120  | 7.290  |
| (29) Focus               | 0.073*** | 0.116    | 0.128 | 0.141 | 0.152  | 0.141   | 0.138   | 0.141  | (29) Focus                   | 0.080*** | 0.367*** | 0.748** | 2.381* | 4.349* | 6.306  | 6.305  | 6.023  |
| (30) Comb1 Mean          | 0.089*   | 0.114*** | 0.134 | 0.147 | 0.151  | 0.145   | 0.137** | 0.147  | (30) Comb1 Mean              | 0.331    | 0.880    | 1.520   | 3.939  | 6.851  | 8.172  | 7.689  | 6.004  |
| (31) Comb1 Median        | 0.090**  | 0.116**  | 0.136 | 0.149 | 0.153  | 0.141*  | 0.136** | 0.149  | (31) Comb1 Median            | 0.241    | 0.824    | 1.521   | 4.068  | 7.084  | 8.447  | 8.041  | 6.625  |
| (32) Comb1 GR            | 0.084**  | 0.144    | 0.157 | 0.239 | 0.238  | 0.293   | 0.285   | 0.350  | (32) Comb1 GR                | 0.244    | 0.824    | 1.896   | 5.346  | 10.638 | 21.204 | 20.188 | 15.492 |
| (33) Comb1 CLS           | 0.089*** | 0.114*** | 0.136 | 0.147 | 0.152  | 0.144   | 0.139** | 0.148  | (33) Comb1 CLS               | 0.316    | 0.893    | 1.610   | 4.513  | 6.774  | 8.360  | 7.687  | 6.026  |
| (34) Comb1 CSR           | 0.090*   | 0.118**  | 0.140 | 0.165 | 0.192  | 0.272   | 0.197   | 0.224  | (34) Comb1 CSR               | 0.343    | 0.852    | 1.645   | 5.372  | 10.008 | 12.918 | 14.097 | 14.011 |
| (35) Comb1 Adalasso      | 0.081*** | 0.118    | 0.133 | 0.174 | 0.198  | 0.319   | 0.291   | 0.284  | (35) Comb1 Adalasso          | 0.236    | 0.785    | 1.541   | 6.691  | 14.700 | 28.671 | 27.811 | 16.241 |
| (36) Comb1 RF            | 0.095    | 0.126    | 0.151 | 0.190 | 0.197  | 0.205   | 0.227   | 0.217  | (36) Comb1 RF                | 0.420    | 1.057    | 2.084   | 5.653  | 9.517  | 12.677 | 14.247 | 12.175 |
| (37) Comb2 Mean          | 0.092**  | 0.115*** | 0.144 | 0.146 | 0.154  | 0.147   | 0.138** | 0.160  | (37) Comb2 Mean              | 0.182    | 0.783    | 1.459   | 4.158  | 6.543  | 11.062 | 12.655 | 12.650 |
| (38) Comb2 Median        | 0.092**  | 0.115*** | 0.139 | 0.149 | 0.155  | 0.147   | 0.138** | 0.158  | (38) Comb2 Median            | 0.182    | 0.773    | 1.430   | 4.314  | 6.721  | 11.062 | 12.655 | 12.650 |
| (39) Comb2 GR            | 0.093**  | 0.122    | 0.150 | 0.171 | 0.223  | 0.184   | 0.235   | 0.227  | (39) Comb2 GR                | 0.206    | 0.748    | 1.552   | 4.096  | 12.533 | 19.296 | 22.484 | 24.721 |
| (40) Comb2 CLS           | 0.090*** | 0.117**  | 0.138 | 0.151 | 0.154  | 0.149   | 0.135** | 0.168  | (40) Comb2 CLS               | 0.186    | 0.715    | 1.441   | 4.523  | 6.952  | 13.451 | 15.088 | 12.206 |
| (41) Comb2 CSR           | 0.094**  | 0.122    | 0.146 | 0.165 | 0.181  | 0.184   | 0.204   | 0.227  | (41) Comb2 CSR               | 0.192    | 0.805    | 1.620   | 5.619  | 10.698 | 19.295 | 22.619 | 27.822 |
| (42) Comb2 Adalasso      | 0.092**  | 0.123    | 0.152 | 0.172 | 0.207  | 0.174   | 0.234   | 0.223  | (42) Comb2 Adalasso          | 0.193    | 0.754    | 1.435   | 5.168  | 13.552 | 17.963 | 20.351 | 28.134 |
| (43) Comb2 RF            | 0.101    | 0.118*   | 0.153 | 0.173 | 0.192  | 0.201   | 0.200   | 0.205  | (43) Comb2 RF                | 0.294    | 1.106    | 2.061   | 6.143  | 9.288  | 10.815 | 10.402 | 13.972 |
| (44) Comb3 Mean          | 0.067*** | 0.120    | 0.124 | 0.142 | 0.155  | 0.139   | 0.136** | 0.146  | (44) Comb3 Mean              | 0.070*** | 0.362*** | 0.684** | 3.606  | 6.252  | 9.025  | 9.517  | 9.962  |
| (45) Comb3 Median        | 0.067*** | 0.120    | 0.124 | 0.146 | 0.155  | 0.139   | 0.136** | 0.146  | (45) Comb3 Median            | 0.070*** | 0.362*** | 0.684** | 3.992  | 6.574  | 9.025  | 9.517  | 9.962  |
| (46) Comb3 GR            | 0.055*** | 0.121    | 0.128 | 0.180 | 0.161  | 0.166   | 0.215   | 0.162  | (46) Comb3 GR                | 0.084*** | 0.405**  | 0.794*  | 3.073  | 8.126  | 17.610 | 21.024 | 18.308 |
| (47) Comb3 CLS           | 0.068**  | 0.132    | 0.130 | 0.146 | 0.155  | 0.141   | 0.136*  | 0.146  | (47) Comb3 CLS               | 0.073*** | 0.409**  | 0.646** | 3.044  | 5.759  | 11.262 | 9.682  | 10.483 |
| (48) Comb3 CSR           | 0.071*   | 0.123    | 0.129 | 0.156 | 0.161  | 0.166   | 0.178   | 0.162  | (48) Comb3 CSR               | 0.085*** | 0.413**  | 0.698** | 4.287  | 9.300  | 18.846 | 27.321 | 25.655 |
| (49) Comb3 Adalasso      | 0.070*   | 0.121    | 0.129 | 0.160 | 0.163  | 0.166   | 0.185   | 0.162  | (49) Comb3 Adalasso          | 0.071*** | 0.424*   | 0.702** | 3.656  | 9.908  | 18.472 | 26.671 | 23.528 |
| (50) Comb3 RF            | 0.076*** | 0.124    | 0.131 | 0.158 | 0.190  | 0.182   | 0.175   | 0.174  | (50) Comb3 RF                | 0.180    | 0.629    | 1.241   | 5.042  | 8.396  | 11.425 | 13.284 | 15.828 |
| number of observations   | 88       | 87       | 86    | 83    | 80     | 77      | 74      | 71     | number of observations       | 88       | 87       | 86      | 83     | 80     | 77     | 74     | 71     |
| best model               | 46       | 9        | 6     | 29    | 18     | 18      | 27      | 29     | best model                   | 45       | 45       | 28      | 29     | 29     | 29     | 29     | 17     |
| R2 oos (%)               | 47       | 21       | 10    | 10    | 14     | 13      | 27      | 18     | R2 oos (%)                   | 62       | 46       | 49      | 37     | 33     | 42     | 55     | 66     |

Notes: Yellow cells denote Top10 models (lowest MSEs) in each horizon. \*\*\*, \*\*, and \* indicate rejection at 1%, 5%, and 10% levels, respectively, using the Diebold and Mariano (1995) test and considering model 3 as a benchmark. The  $R^2$  out-of-sample statistics (R2 oos) refer to the best model in each horizon. Source: Authors' own elaboration.

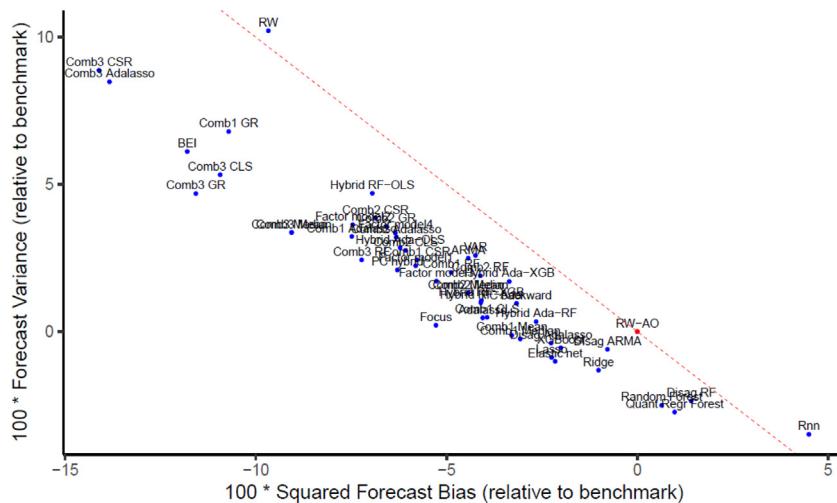
In turn, the PCs and the factor models (the latter especially for  $h \geq 9$ ) enter more often into the set of Top10 forecasts. Focus and BEI also belong to the set of best forecasts in several cases. Regarding combinations, the third set of forecasts (set3, including Focus and BEI in the pool of forecasts) clearly provides a superior<sup>64</sup> information set when compared to set1 and set2. In this sense, the quality of the information set embodied in the pool of forecasts seems to be more important here than the forecast combination method<sup>65</sup> used to weight the individual forecasts (i.e., there is no clear winner method in comb3 to be used in all horizons).

Now, considering the inflation rate in 12 months (right panel in Table 2), note the excellent performance of BEI and Focus, which belong to the set of Top10 forecasts in almost all horizons. Furthermore, both forecasts statistically beat the benchmark (at least) at the 10% level, using the Diebold-Mariano test, from  $h = 1$  to 9 months. The other forecasts that can also beat the benchmark (but only for horizons up to 3 months) are the forecast combinations (comb3, excepting model 50) based on the third set of forecasts. Again, comb3 dominates the other two sets of combinations at short/medium horizons, which can be attributed to the quality of the information set (e.g., BEI, Focus) coupled with the MCS method.

Again, the superior accuracy of the ML forecasts stands out, compared to most of the traditional forecasting approaches. For instance, for horizons from  $h = 12$  to 18 months, the tree-based methods dominate the other competing methods (with a few exceptions) in set1 of forecasts (models 1 to 27). In particular, for the longest horizon ( $h = 18$ ), the XGBoost method is the best among all the 50 candidates, providing an accuracy gain of 66%, in terms of the  $R^2_{oos}$  statistic, compared to the benchmark (ARMA) model.

<sup>64</sup> Recall that market professionals devote considerable resources to inflation forecasting and use a broad range of information.

<sup>65</sup> Regarding the use of ML methods as forecast combination devices, results are not so encouraging when compared to more traditional combination approaches.



**Fig. 6.** Scatterplot of relative forecast variance and squared forecast bias ( $h = 1$ ). Notes: The red line represents forecasts with the same MSE as the RW-AO; points to the right are forecasts outperformed by the RW-AO, and points to the left represent forecasts that outperform the RW-AO. Source: Authors' own elaboration. (For interpretation of the references to colour in this figure legend, the reader is referred to the web version of this article.)

In general, the MSEs from the robustness exercises, presented in [Appendix B](#), lead to similar conclusions. For instance, factor models show the best results among traditional models, ML methods perform better at medium/long horizons, and forecast combination including all individual forecasts and using the MCS approach delivers an improved accuracy when compared to the other sets of forecasts.

These results add to those previously reported by the inflation forecasting literature. For instance, regarding Brazilian inflation, [Medeiros et al. \(2016\)](#) report that ML (lasso-based) methods show the smallest errors at short horizons, whereas the AR is the best model for long horizons, followed by the factor model, in some cases. [Garcia et al. \(2017\)](#) forecast Brazilian inflation using some of the methods employed here, based on a sample from January 2003 to December 2015, with 59 macroeconomic variables and 34 variables linked to specialist forecasts. The authors document the superiority of high-dimensional econometric models, such as shrinkage and the CSR method, compared to alternative ML methods, and show that forecast combination based on model confidence sets can achieve superior predictive performances. In turn, [Medeiros et al. \(2021\)](#) present a similar investigation using U.S. data, based on a sample from January 1960 to December 2015, with 122 variables from the FRED-MD database. They point out that random forest is the ML method that deserves more attention, since it dominates all other models in several cases. Furthermore, they show that ML models with a large number of covariates are systematically more accurate than the benchmarks for several forecasting horizons, both in the 1990s and the 2000s.

Next, we deepen the forecast error analysis<sup>66</sup> by investigating the trade-off between bias and variance.<sup>67</sup> Following Lima and Meng (2017), we decompose the MSE into two parts: the forecast variance and the squared forecast bias. To do so, we calculate the

MSE of any forecast  $\widehat{f}_{y_{t+h}}$  as  $\frac{1}{T^*} \sum_t (y_{t+h} - \widehat{f}_{y_{t+h}})^2$  and the forecast variance as  $\frac{1}{T^*} \sum_t \left( \widehat{f}_{y_{t+h}} - \frac{1}{T^*} \sum_t \widehat{f}_{y_{t+h}} \right)^2$ , where  $T^*$  is the number of out-of-sample observations. The squared forecast bias is computed as the difference between MSE and the forecast variance.

Fig. 6 shows the relative forecast variance and squared forecast bias of all 50 forecasting methods considering the monthly inflation rate as the dependent variable. This analysis is particularly important in model selection and helps understand why some methods display a better forecast accuracy compared to others. The relative forecast variance (squared bias) is calculated as the difference between the forecast variance (squared bias) of the  $i$ -th model and the forecast variance (squared bias) of the moving-average approach RW-AO. Thus, the value of relative forecast variance (squared bias) for the RW-AO is necessarily equal to zero. Moreover, each point on the red line represents a forecast with the same MSE as the RW-AO. Points to the right of the line are forecasts outperformed by the RW-AO, and points to the left represent forecasts that outperform the RW-AO. Since the RW-AO is a simple moving average of inflation, it will have a low forecast variance but will likely be biased.

Note that for  $h = 1$ , all forecasts (excepting RW and RNN) outperform the RW-AO. Since forecast variances are generally greater than RW-AO's variance (note the blue dots are usually above zero in the vertical axis of Fig. 6), the higher accuracy compared to RW-AO relies almost exclusively on a method's ability to lower the bias relative to the RW-AO. Also, note that tree-based methods

<sup>66</sup> According to Ng (2015), time spent coming up with diagnostics for learning algorithms is time well spent.

<sup>67</sup> Underfitting usually occurs when a model is too simple (e.g., few predictors), showing low forecast variance but more bias toward wrong outcomes. In turn, models that are more complex are often able to reduce the bias, but at the cost of a higher forecast variance. This trade-off between models that are too simple (high bias) versus too complex (high variance) is a key issue in statistics and affects all supervised ML methods.

deliver a lower forecast variance, essentially, due to data bootstrapping in multiple learning, besides a random subset of variables used in each tree.

On the other hand, forecast combinations seem to increase the forecast variance when compared to individual forecasting methods. In particular, the best method for  $h = 1$  (Comb3 GR) delivered the lowest MSE by reducing the bias (e.g., compared to Focus) without increasing the variance too much. The main message is that the forecasting methods that yield a sizeable reduction in the forecast bias while keeping variance under control are able to improve forecasting accuracy over the lowest-variance approach (RW-AO).

The MSE decomposition for other cases are not shown here to save space, but are available upon request. For the 12-month inflation ( $h = 12$  or  $18$ ), the tree-based methods again show a lower variance, whereas forecast combinations in general exhibit a sizeable bias (one possible reason is multicollinearity at long horizons due to highly correlated forecasts).<sup>68, 69</sup>

The previous analysis enabled a discussion on relative (average) forecast accuracy. However, such measures alone do not convey any information on how the performance of the competing methods evolves *over time*. To tackle this issue, we compute the cumulative squared prediction error (CSPE) of each method, compared to the benchmark, along the pseudo out-of-sample exercise; see Rapach et al. (2010) and Lima and Meng (2017).

Fig. 7 shows the differences over time between the CSPEs of the benchmark (ARMA) and each competing method. When the blue curve increases, the considered method is outperformed by the benchmark, while the opposite holds when the curve decreases. Also, if the curve is higher at the end of the evaluation period, the method has a higher MSE compared to the benchmark, considering all out-of-sample observations. Note in Fig. 7 that the best forecast (model 46) consistently outperformed the benchmark (smooth decline of the blue line). On the other hand, in several cases (e.g., BEI, Comb3 CSR) there is a concentrated accuracy loss compared to the benchmark (sharp increase of the blue line) at the beginning of 2020 due to the COVID-19 pandemic. In other cases (RF and QRF), the blue line fluctuates slightly above the zero horizontal line, indicating a bit worse performance compared to the benchmark.

Another interesting analysis is the identification of the most important variables chosen by the ML methods to predict inflation. A first approach is to track the number of variables selected (or not) over time, along the pseudo out-of-sample forecasting exercise. Fig. 8 reveals, for illustrative purposes, among the 501 potential predictors for inflation, which ones were indeed selected (and when), according to the adalasso and elastic net methods.

Fig. 8 shows how the selection procedure works over time. The horizontal axis represents the end of the estimation sample, along the out-of-sample forecasting exercise, and the vertical axis denotes all of the 501 regressors. A blue dot indicates that variable  $i$  has a nonzero coefficient in the adalasso estimation (a red dot, in the elastic net) with a sample ending at period  $t$ , used to build forecasts for  $y_{t+h}$ . This allows us to discover how the models change in response to different economic conditions over time. In other words, Fig. 8 shows that the statistical significance of the coefficients varies considerably over time for some variables, while remaining relatively stable for several others. For instance, note for  $h = 1$  that adalasso quite regularly selects three variables along the forecasting exercise. As later revealed in Fig. 9, these variables are the IPCA headline and IPC-Fipe inflation rates, lagged 1 month, both representing the inertial component of inflation, besides the commercial consumption of electricity, also lagged 1 month, which is a variable directly related to economic activity.

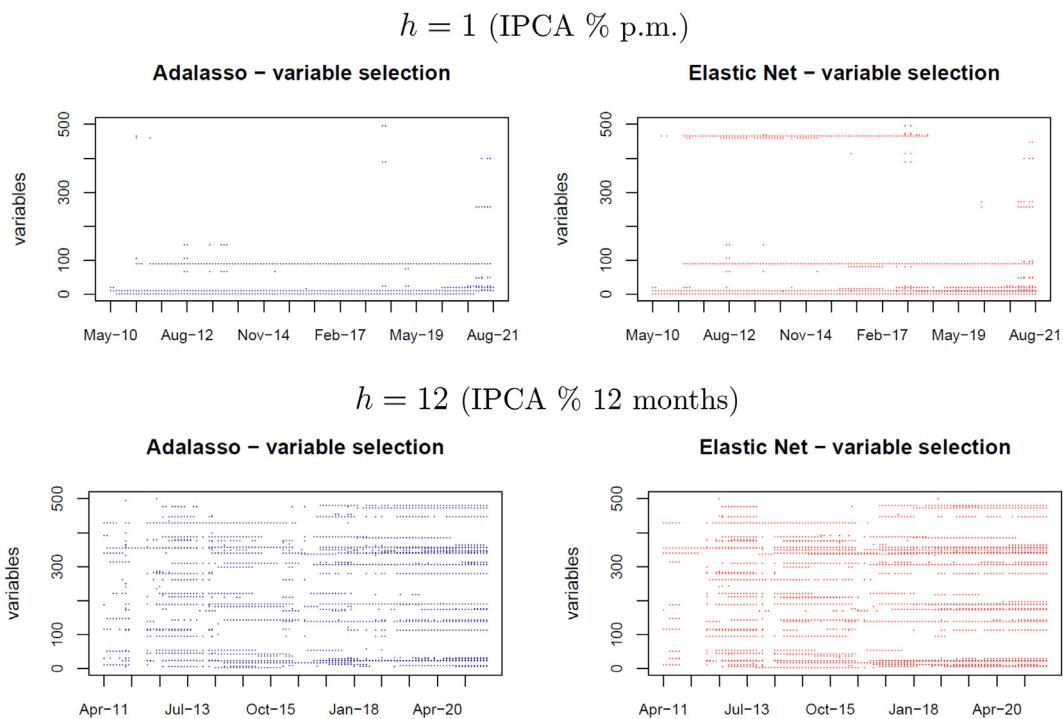
Now we identify which variables are more frequently chosen by some ML algorithms. Although we do not attempt to economically interpret the driving forces behind these forecasts, further inspecting such methods allows us to better understand how they are making forecasts, which may reveal new statistical relationships in the data previously overlooked by standard linear models. Fig. 9 shows *word cloud* graphs, which are images composed by the names of the variables<sup>70</sup> most frequently selected by some method, where the size of each word indicates the frequency a variable is selected (e.g., nonzero coefficient) along the out-of-sample forecasting exercise. Thus, most frequent variables have a larger font size, whereas variables with the same frequency of selection have the same size and color.

The variables identified by adalasso and elastic net (for  $h = 1$ ) include past inflation (inertial inflationary dynamics) and variables related to the real economy (e.g., commercial electricity consumption, agriculture exports). Regarding the 12-month inflation rate (for  $h = 12$ ), the set of most frequent variables now also includes interest rates, fiscal variables, external sector variables (current account, dollar index, CRB food, and oil price), and new variables (not traditionally used to forecast inflation), such as the temperature of the Pacific Ocean (oceanic nino index),<sup>71</sup> due to the role that the *El Niño* and *La Niña* might play in food inflation.<sup>72</sup>

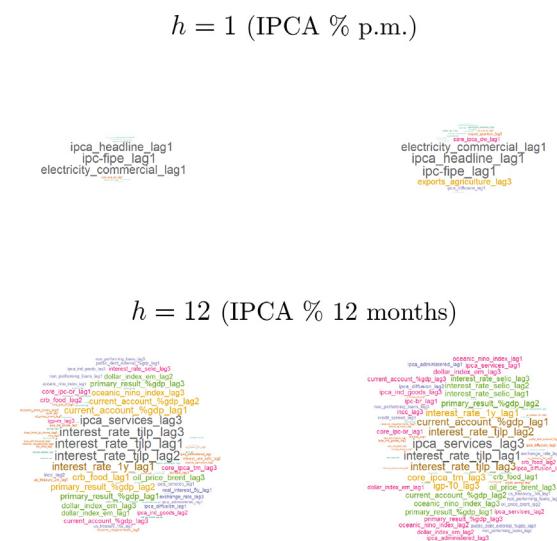
Another way to inspect the ML results is to build the so-called *variable importance* (or feature importance) graphs. The idea is to once again build a rank of variables but based on their usefulness in predicting inflation at a given horizon.<sup>73</sup> Word clouds can again be used to summarize the results. For the penalized-regression models (elastic net, lasso, adalasso, and ridge), we compute the rank of variable importance based on the absolute value of the estimated coefficients (adjusted for the original scale of each variable) multiplied by the standard deviation of the respective variable. For tree-based methods, the rank can be computed using either the



**Fig. 7.** Cumulative Square Prediction Error (CSPE) for  $h = 1$  (IPCA % p.m.) Notes: A positively sloped curve in each panel indicates the conditional model is outperformed by the benchmark, whereas the opposite holds true for a downward sloping curve. Moreover, if the curve is positive (negative) at the end of the period, then the competing method has a higher (lower) MSE than the benchmark over the evaluation period. Source: Authors' own elaboration.



**Fig. 8.** Variable selection over time in adalasso and elastic net. Source: Authors' own elaboration.



**Fig. 9.** Word cloud (*frequency*), adalasso (left), and elastic net (right). Note: The size of each word indicates the frequency a variable is selected along the out-of-sample forecasting exercise. Variables with the same frequency of selection have the same size and color. Source: Authors' own elaboration.

$h = 1$ , elastic net (left), random forest (right)



$h = 3$ , ridge regression (left), random forest (right)



**Fig. 10.** Word cloud (*importance*), selected models, IPCA % p.m. Note: The size of each word indicates the variable importance. Most relevant variables have larger font size, and variables with the same importance have the same size and color. Source: Authors' own elaboration.

methods of permutation or impurity (used here). Fig. 10 presents variable importance results for monthly inflation (selecting the best methods in Table 2, among models 11–17).

Figs. 11–12 also show variable importance results but only considering the Top 20 most important variables. Note that, in some cases (e.g., adalasso,  $h = 1$ ), there are just a few relevant predictors employed by the forecasting method. In Figs. 10 and 11, considering  $h = 1$  and the monthly inflation rate, the lagged inflation measured by the IPC-Fipe stands out as the most relevant predictor according to three methods (adalasso, elastic net, and RF), being second place in XGBoost. For  $h = 3$  (Fig. 10), the 1-year interest rate gains relevance, together with variables related to the real economy and the external sector, besides unusual predictors of Brazilian inflation, such as the amount of rainfall in some Brazilian cities. Also, note in Fig. 10 that penalized-regressions (elastic net and ridge) give more importance to fewer predictors, compared to RF. In turn, Fig. 12 shows the results for the 12-month inflation rate, and  $h = 12$ , confirming that the usual variables often employed to forecast inflation in Brazil (e.g., past inflation, lagged interest rates, exchange rate, fiscal variables) are indeed the most relevant in our empirical exercise.

Finally, we depart from the point forecast setup and build simple density forecasts associated with the point forecasts at multiple horizons. Fig. 13 shows such density forecasts (fan charts) for selected methods, which allows us to conduct *risk management* analysis.<sup>74</sup> For example, according to model 30 (comb1 mean), the probability of the monthly inflation rate to be greater than 0.8% p.m. (per month) in December 2022 is 20% (and to be above 1.0% p.m. is 9%). In turn, the chance of the 12-month rate, for instance, being above 5.0% p.y. (per year) in December 2022 is 36% according to BEI (and 29% according to Focus). Such analysis can be useful when evaluating the chance of future inflation to be above/below the target.<sup>75</sup>

## 4. Conclusions

Machine learning is constantly evolving with new methods being developed every day. Progress has been made in macroeconomics over the recent years on the usage of such methods with big data. However, model interpretability is usually lost in such approaches. According to Occam's razor, models should be simple and explainable. Nonetheless, ML methods are not easily interpreted, for instance, due to a highly nonlinear setup or a large set of inputs.

<sup>68</sup> Recall that the benefits of combination come from a set of forecasts containing low pairwise correlations

<sup>69</sup> Besides the bias-variance decomposition, other metrics could be used to investigate the performance of ML algorithms, such as overfit and stability measures (e.g., relative overfitting rate).

<sup>70</sup> The names displayed in the word clouds are the nicknames listed in Appendix A.

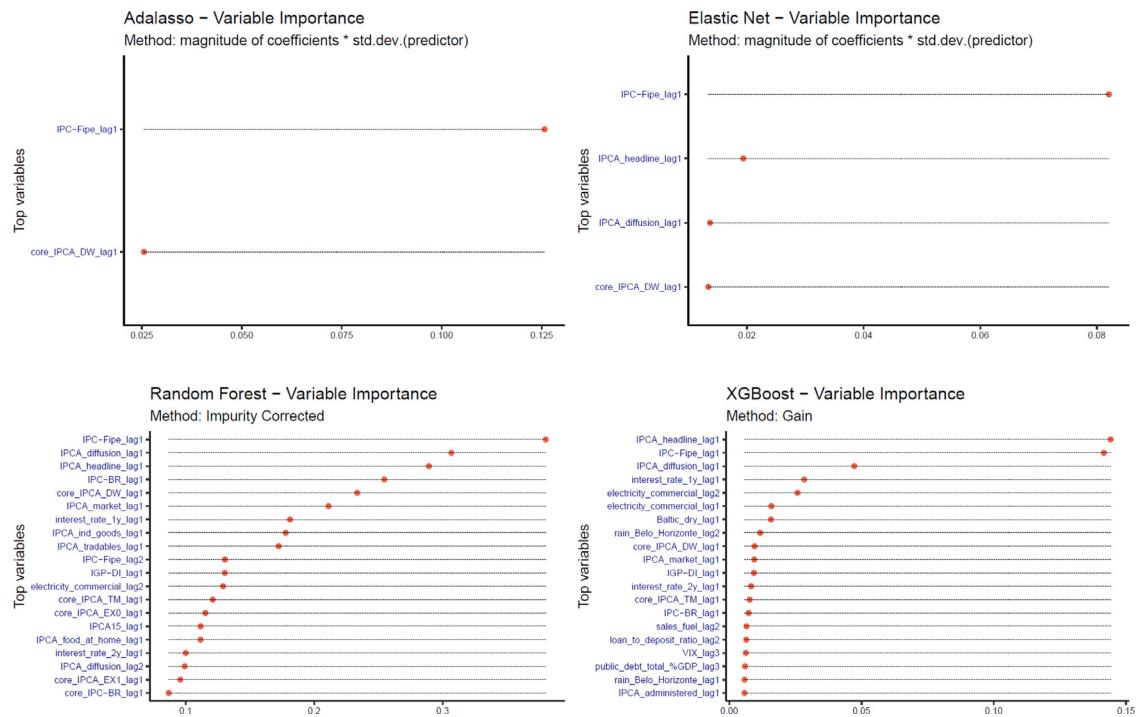
<sup>71</sup> This series is the Oceanic Niño Index (ONI), provided by the Climate Prediction Center, linked to the National Oceanic and Atmospheric Administration (NOAA, USA).

<sup>72</sup> The semi-structural models of BCB include this variable (see Brasil, 2021b).

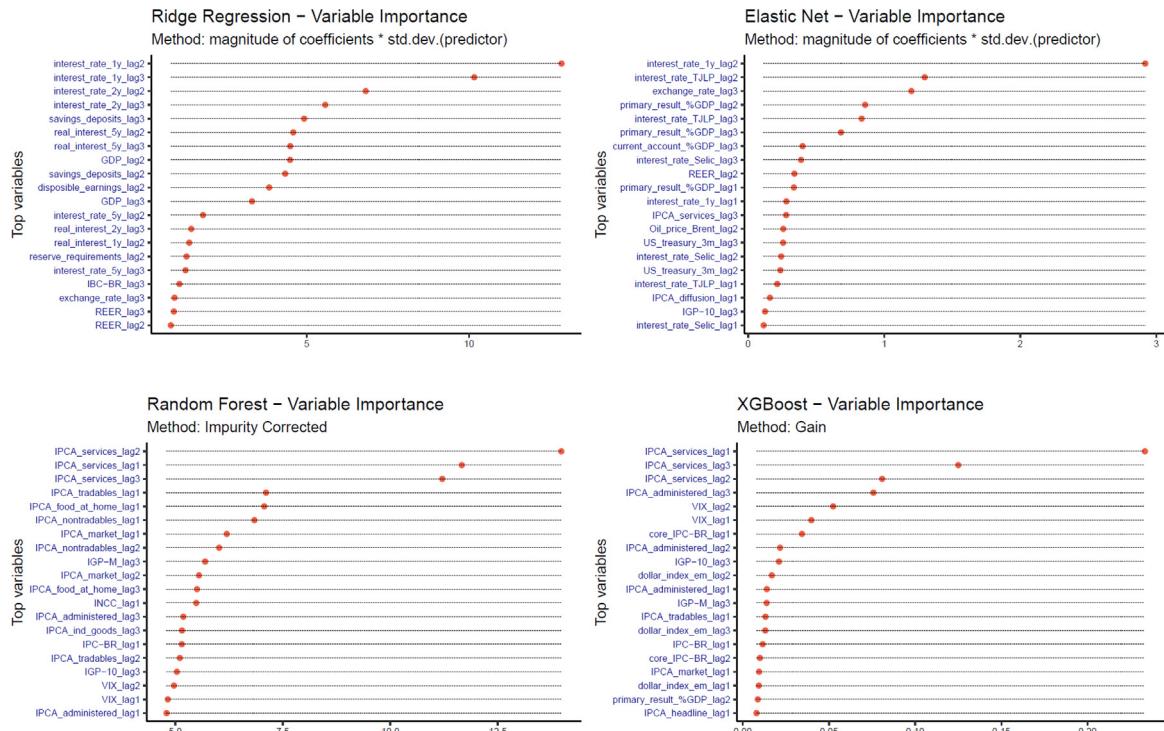
<sup>73</sup> Variable (or feature) importance refers here to techniques that assign scores to variables based on how useful they are at improving the forecast accuracy of a given model.

<sup>74</sup> Recall that, for simplicity, we assumed a Gaussian distribution, where the conditional mean is the point forecast of a given method and the conditional variance comes from past forecast errors of such method.

<sup>75</sup> Note in both charts that the long-run median forecast is close to 0.5% p.m., which is in line with the unconditional forecast, proxied by the sample average of inflation, equal to 0.47% p.m. (or 5.57% in 12 months).



**Fig. 11.** Variable importance, top variables,  $h = 1$ , IPCA % p.m. Note: Higher figures represent more relevant features to predict inflation. Source: Authors' own elaboration.



**Fig. 12.** Variable importance, top variables,  $h = 12$ , IPCA % 12 months. Note: Higher figures represent more relevant features to predict inflation. Source: Authors' own elaboration.

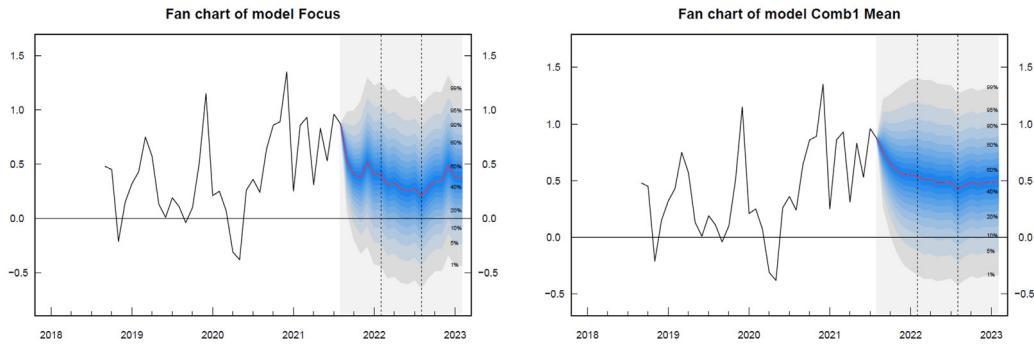


Fig. 13. Fan charts (IPCA % p.m.). Source: Authors' own elaboration.

In this paper, we tackle this issue and take a step toward transparency (turning the *black box* into a *gray box*) by providing complementary tools (e.g., word clouds, variable importance graphs, fan charts) to better understand the ML outcomes. The tools we provide for identifying the most important variables to predict inflation also allow us to shift the discussion from *big data* to *good data*, in the sense that finding high-quality data is more important than the quantity of data (Ng, 2015; 2021).

In this context, we study the inflation forecasting accuracy of 50 competing methods, including some new ML techniques proposed here (hybrid models, and a quantile-combination method based on quantile regression forest), traditional econometric models (e.g., VAR), reduced-form structural models (Phillips curves), factor models, survey-based forecasts, regularization procedures (e.g., elastic net), and forecast combinations, among others.

The variable of interest is the Brazilian inflation as measured by the IPCA. To evaluate the predictive power of each method, we conduct a pseudo out-of-sample empirical exercise (*horse-race*) based on 501 time series, coming from 167 macroeconomic and financial variables, where each method produces point forecasts for horizons  $h = 1, \dots, 18$  months ahead.<sup>76</sup>

According to Wolpert (1996), there is no universal best model. In other words, the set of assumptions that works well in one domain may work poorly in another setup. The empirical results documented in this paper go in this direction, suggesting that some ML algorithms are able to consistently outperform traditional econometric models in terms of MSE. However, there is no supreme model for all cases as the performance depends on the forecast horizon and whether inflation is measured by its monthly rate or accumulated in 12 months.

The main takeaways are the following: (i) ML methods, often designed to work under low *noise-to-signal* ratio setups (e.g., image classification, voice recognition) and large datasets, can do a pretty good job under a medium/high noise-to-signal ratio and a dataset not so large in the time dimension (e.g., applied macroeconomics); (ii) ML methods consistently beat the benchmark (ARMA) model and, in many cases, exhibit two-digit accuracy gains in terms of the  $R^2$  out-of-sample statistic; (iii) nonlinearities captured by ML methods (Varian, 2014), such as recurrent neural networks or random forest, are important to forecast inflation in Brazil; (iv) at shorter horizons, forecast combinations are useful, especially when using big data to build individual forecasts and the model confidence set of Hansen et al. (2011) to select the superior forecasts to be combined; (v) at longer horizons, tree-based methods, such as random forest and XGBoost, perform quite well and dominate other models in several cases; (vi) Focus and BEI also belong to the set of top forecasts in many horizons (for monthly inflation, they improve the quality of the information set used as input in combinations, whereas for the 12-month inflation rate, they belong to the set of top forecasts in almost all horizons); (vii) XGBoost is a competing variable selection method; and (viii) fan charts can easily be constructed from point forecasts, which allows us to estimate the probability of future inflation to be above/below the target.

To sum it up, we analyze ML methods that forecasters should have in their toolkit when predicting inflation in Brazil. Moreover, we reveal the top features that an econometrician should have in mind when building inflation forecast models. These findings represent a valuable contribution to academics, practitioners, and policymakers interested in macroeconomic forecasting using machine learning, specifically focused on Brazilian inflation.

#### Declaration of Competing Interest

The authors declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

<sup>76</sup> As suggestion for future research, a Monte Carlo experiment could be employed to investigate the stability and accuracy of the competing models under different circumstances. In addition, one could investigate if the empirical results presented in the paper apply to other countries or regions. Also, the set of forecasting methods could include MIDAS models using both monthly and daily data. See Andreou et al. (2013) and Monteforte and Moretti (2010) for further details. Finally, one could check the impact of post-COVID-19 observations in the empirical results.

## Appendix A. Database

**Table A1**

List of macroeconomic and financial variables.

|    | Category       | Name   | Source       | Original unit | Nickname            | tcode |
|----|----------------|--|--------------|---------------|---------------------|-------|
| 1  | Inflation      | IPCA (consumer price index)  | IBGE         | % p.m.        | IPCA_headline       | 1     |
| 2  | Inflation      | IPCA (consumer price index, market prices)                           | IBGE         | % p.m.        | IPCA_market         | 1     |
| 3  | Inflation      | IPCA (consumer price index, administered prices)                     | IBGE         | % p.m.        | IPCA_administered   | 1     |
| 4  | Inflation      | IPCA (consumer price index, tradables)                               | BCB          | % p.m.        | IPCA_tradables      | 1     |
| 5  | Inflation      | IPCA (consumer price index, nontradables)                            | BCB          | % p.m.        | IPCA_nontradables   | 1     |
| 6  | Inflation      | IPCA (consumer price index, services)                                | BCB          | % p.m.        | IPCA_services       | 1     |
| 7  | Inflation      | IPCA (consumer price index, industrial goods)                        | BCB          | % p.m.        | IPCA_ind_goods      | 1     |
| 8  | Inflation      | IPCA (consumer price index, food at home)                            | BCB          | % p.m.        | IPCA_food_at_home   | 1     |
| 9  | Inflation      | IPCA diffusion index   | BCB          | %             | IPCA_diffusion      | 1     |
| 10 | Inflation      | IPCA-15 (consumer price index-extended 15)                           | IBGE         | % p.m.        | IPCA15              | 1     |
| 11 | Inflation      | IPC-Fipe (consumer price index)                                      | Fipe         | % p.m.        | IPC-Fipe            | 1     |
| 12 | Inflation      | IPC-Br (consumer price index)  | FGV          | % p.m.        | IPC-BR              | 1     |
| 13 | Inflation      | IGP-DI (general price index)   | FGV          | % p.m.        | IGP-DI              | 1     |
| 14 | Inflation      | IGP-M (general price index)  | FGV          | % p.m.        | IGP-M               | 1     |
| 15 | Inflation      | IGP-10 (general price index)   | FGV          | % p.m.        | IGP-10              | 1     |
| 16 | Inflation      | INCC (national index of building costs)                              | FGV          | % p.m.        | INCC                | 1     |
| 17 | Inflation      | Core IPC-Br (core inflation)   | FGV          | % p.m.        | core_IPC-BR         | 1     |
| 18 | Inflation      | Core IPCA - Exclusion EX0 (core inflation)                           | BCB          | % p.m.        | core_IPCA_EX0       | 1     |
| 19 | Inflation      | Core IPCA - Exclusion EX1 (core inflation)                           | BCB          | % p.m.        | core_IPCA_EX1       | 1     |
| 20 | Inflation      | Core IPCA - Double Weight (core inflation)                           | BCB          | % p.m.        | core_IPCA_DW        | 1     |
| 21 | Inflation      | Core IPCA - Trimmed Means Smoothed (core inflation)                  | BCB          | % p.m.        | core_IPCA_TM        | 1     |
| 22 | Interest rates | Nominal policy interest rate (Selic)                                 | BCB          | % p.a.        | interest_rate_Selic | 2     |
| 23 | Interest rates | Nominal policy interest rate (long-term interest rate, TJLP)         | BCB          | % p.a.        | interest_rate_TJLP  | 2     |
| 24 | Interest rates | Nominal market interest rate (prefixed, 1 year)                      | Anbima       | % p.a.        | interest_rate_1y    | 2     |
| 25 | Interest rates | Nominal market interest rate (prefixed, 2 years)                     | Anbima       | % p.a.        | interest_rate_2y    | 2     |
| 26 | Interest rates | Nominal market interest rate (prefixed, 5 years)                     | Anbima       | % p.a.        | interest_rate_5y    | 2     |
| 27 | Interest rates | Real market interest rate (indexed IPCA, 1 year)                     | Anbima       | % p.a.        | real_interest_1y    | 2     |
| 28 | Interest rates | Real market interest rate (indexed IPCA, 2 years)                    | Anbima       | % p.a.        | real_interest_2y    | 2     |
| 29 | Interest rates | Real market interest rate (indexed IPCA, 5 years)                    | Anbima       | % p.a.        | real_interest_5y    | 2     |
| 30 | Interest rates | U.S. Treasury 3 months nominal yield                                 | Reuters      | % p.a.        | US_treasury_3m      | 2     |
| 31 | Interest rates | U.S. Treasury 2 years nominal yield                                  | Reuters      | % p.a.        | US_treasury_2y      | 2     |
| 32 | Interest rates | U.S. Treasury 10 years nominal yield                                 | Reuters      | % p.a.        | US_treasury_10y     | 2     |
| 33 | Interest rates | U.S. Treasury 5 years TIPS (Treasury Inflation-Protected Securities) | Reuters      | % p.a.        | US_treasury_5y_tips | 2     |
| 34 | Money          | Monetary base  | BCB          | R\$ thousand  | monetary_base       | 5     |
| 35 | Money          | Money supply (currency outside banks)                                | BCB          | R\$ thousand  | money_supply        | 5     |
| 36 | Money          | Demand deposits  | BCB          | R\$ thousand  | demand_deposits     | 5     |
| 37 | Money          | Savings deposits   | BCB          | R\$ thousand  | savings_deposits    | 5     |
| 38 | Money          | M1   | BCB          | R\$ thousand  | M1                  | 5     |
| 39 | Money          | M2   | BCB          | R\$ thousand  | M2                  | 6     |
| 40 | Money          | M3   | BCB          | R\$ thousand  | M3                  | 6     |
| 41 | Money          | M4   | BCB          | R\$ thousand  | M4                  | 6     |
| 42 | Banking        | Credit spread (nonearmarked credit rate - Selic rate)                | BCB, authors | basis points  | credit_spread       | 2     |

(continued on next page)

Table A1 (continued)

|    | Category        | Name   | Source       | Original unit | Nickname                      | tcode |
|----|-----------------|--|--------------|---------------|-------------------------------|-------|
| 43 | Banking         | Non-Performing Loans (NPL) of total credit   | BCB, authors | %             | non_performing_loans          | 2     |
| 44 | Banking         | Loan-to-Deposit ratio (LTD)  | BCB, authors | Units         | loan_to_deposit_ratio         | 3     |
| 45 | Banking         | Reserve requirements ratio (financial inst. reserve requirements / total deposits)   | BCB, authors | Units         | reserve_requirements          | 2     |
| 46 | Banking         | Nonearmarked credit operations outstanding   | BCB, authors | R\$ million   | credit_outstanding            | 6     |
| 47 | Capital markets | Ibovespa (Brazil)  | Reuters      | Index         | Ibovespa                      | 5     |
| 48 | Capital markets | Euro Stoxx 50 price index  | Reuters      | Index         | euro_stoxx50                  | 5     |
| 49 | Capital markets | MSCI emerging countries (EM, US\$)   | Reuters      | Index         | MSCI_emerging                 | 5     |
| 50 | Capital markets | MSCI developed countries (World, US\$)   | Reuters      | Index         | MSCI Developed                | 5     |
| 51 | FX and risk     | FX-rate (nominal exchange rate, R\$/US\$)  | Reuters      | Units         | exchange_rate                 | 5     |
| 52 | FX and risk     | REER (Real effective exchange rate, IPA-13 currencies)                               | Reuters      | Index         | REER                          | 5     |
| 53 | FX and risk     | U.S. dollar index (DXY, geometric average of 6 currencies in respect to US\$)        | Reuters      | Index         | dollar_index                  | 5     |
| 54 | FX and risk     | U.S. dollar emerging market index (Federal Reserve, 19 countries)                    | Reuters      | Index         | dollar_index_em               | 5     |
| 55 | FX and risk     | Embi+Br (Emerging Markets Bond Index Plus Brazil, spread)                            | Reuters      | basis points  | embi+br                       | 5     |
| 56 | FX and risk     | CDS (Credit Default Swap, Brazil 5 years)  | Reuters      | basis points  | CDS_5y                        | 5     |
| 57 | FX and risk     | U.S. corporate bonds Moody's seasoned BAA  | Reuters      | % p.a.        | US_corp_bonds                 | 2     |
| 58 | FX and risk     | VIX CBOE volatility index (30-day expected volatility of the SP500)                  | Reuters      | Index         | VIX                           | 1     |
| 59 | Labor           | Unemployment rate (open)   | IBGE         | %             | unemployment_rate             | 3     |
| 60 | Labor           | Formal employment created - South  | MTb          | Units         | employment_south              | 2     |
| 61 | Labor           | Formal employment created - Southeast  | MTb          | Units         | employment_southeast          | 2     |
| 62 | Labor           | Formal employment created - North  | MTb          | Units         | employment_north              | 2     |
| 63 | Labor           | Formal employment created - Northeast  | MTb          | Units         | employment_northeast          | 2     |
| 64 | Labor           | Formal employment created - Central-West   | MTb          | Units         | employment_central_west       | 2     |
| 65 | Labor           | Minimum wage   | MTb          | R\$           | minimum_wage                  | 5     |
| 66 | Labor           | Hours worked in production (Rio Grande do Sul)                                       | Fiergs       | Index         | hours_worked_prod_RS          | 5     |
| 67 | Labor           | Disposable overall earnings (accumulated in 12 months)                               | BCB          | R\$ million   | disposable_earnings           | 6     |
| 68 | Industry        | Industrial production (mineral extraction)   | IBGE         | Index         | ind_prod_mineral_extract      | 5     |
| 69 | Industry        | Industrial production (manufacturing industry)                                       | IBGE         | Index         | ind_prod_manufacturing        | 5     |
| 70 | Industry        | Industrial production (capital goods)  | IBGE         | Index         | ind_prod_capital_goods        | 5     |
| 71 | Industry        | Industrial production (intermediate goods)   | IBGE         | Index         | ind_prod_interm_goods         | 5     |
| 72 | Industry        | Industrial production (consumer goods)   | IBGE         | Index         | ind_prod_consumer_goods       | 5     |
| 73 | Industry        | Industrial production (durable goods)  | IBGE         | Index         | ind_prod_durable_goods        | 5     |
| 74 | Industry        | Industrial production (semidurable and nondurable goods)                             | IBGE         | Index         | ind_prod_non-durable          | 5     |
| 75 | Industry        | Installed capacity utilization (Rio Grande do Sul)                                   | Fiergs       | %             | capacity_utilization_RS       | 2     |
| 76 | Industry        | Capacity utilization (manufacturing industry, FGV)                                   | FGV          | %             | capacity_utilization_industry | 2     |
| 77 | Industry        | Steel production   | BCB          | Index         | steel_production              | 5     |
| 78 | Industry        | Vehicles production (total)  | Anfavea      | Units         | vehicles_production           | 5     |
| 79 | Industry        | Truck production   | Anfavea      | Units         | truck_production              | 5     |
| 80 | Industry        | Bus production   | Anfavea      | Units         | bus_production                | 5     |
| 81 | Industry        | Production of agricultural machinery (total)   | Anfavea      | Units         | agricultural_machinery        | 5     |
| 82 | Sales           | Sales volume index in the retail sector (total)                                      | IBGE         | Index         | sales_total                   | 5     |
| 83 | Sales           | Sales volume index in the retail sector (fuel and lubricants)                        | IBGE         | Index         | sales_fuel                    | 5     |
| 84 | Sales           | Sales volume index in the retail sector (hyperm., superm., food, bever. and tobacco) | IBGE         | Index         | sales_hypermarket             | 5     |
| 85 | Sales           | Sales volume index in the retail sector (textiles, clothing and footwear)            | IBGE         | Index         | sales_textiles                | 5     |
| 86 | Sales           | Sales volume index in the retail sector (furniture and white goods)                  | IBGE         | Index         | sales_furniture               | 5     |

(continued on next page)

Table A1 (continued)

|     | Category          | Name  | Source     | Original unit | Nickname                  | tcode |
|-----|-------------------|---|------------|---------------|---------------------------|-------|
| 87  | Sales             | Sales volume index in the retail sector (vehicles and motorcycles, spare parts) | IBGE       | Index         | sales_vehicles1           | 5     |
| 88  | Sales             | Vehicle sales (total)   | Anfavea    | Units         | sales_vehicles2           | 5     |
| 89  | Sales             | Domestic vehicle sales  | Anfavea    | Units         | domestic_sales_vehicles   | 5     |
| 90  | Energy            | Electric energy consumption (commercial)  | Eletrobras | GWh           | electricity_commercial    | 5     |
| 91  | Energy            | Electric energy consumption (residential)                                       | Eletrobras | GWh           | electricity_residential   | 5     |
| 92  | Energy            | Electric energy consumption (industrial)  | Eletrobras | GWh           | electricity_industrial    | 5     |
| 93  | Energy            | Electric energy consumption (other)   | Eletrobras | GWh           | electricity_other         | 5     |
| 94  | Energy            | Electric energy consumption (total)   | Eletrobras | GWh           | electricity_total         | 5     |
| 95  | Climate           | El Niño-Southern Oscillation, as measured by the Oceanic Niño Index (ONI)       | NOAA       | Index         | oceanic_nino_index        | 2     |
| 96  | Climate           | Total monthly precipitation (mm) in Belém                                       | INMET      | mm            | rain_Belem                | 2     |
| 97  | Climate           | Total monthly precipitation (mm) in Belo Horizonte                              | INMET      | mm            | rain_Belo_Horizonte       | 2     |
| 98  | Climate           | Total monthly precipitation (mm) in Curitiba                                    | INMET      | mm            | rain_Curitiba             | 2     |
| 99  | Climate           | Total monthly precipitation (mm) in Florianópolis                               | INMET      | mm            | rain_Florianopolis        | 2     |
| 100 | Climate           | Total monthly precipitation (mm) in Goiânia                                     | INMET      | mm            | rain_Goiania              | 2     |
| 101 | Climate           | Total monthly precipitation (mm) in Manaus                                      | INMET      | mm            | rain_Manaus               | 2     |
| 102 | Climate           | Total monthly precipitation (mm) in Palmas                                      | INMET      | mm            | rain_Palmas               | 2     |
| 103 | Climate           | Total monthly precipitation (mm) in Porto Alegre                                | INMET      | mm            | rain_Porto_Alegre         | 2     |
| 104 | Climate           | Total monthly precipitation (mm) in Recife                                      | INMET      | mm            | rain_Recife               | 2     |
| 105 | Climate           | Total monthly precipitation (mm) in Rio Branco                                  | INMET      | mm            | rain_Rio_Branco           | 2     |
| 106 | Climate           | Total monthly precipitation (mm) in Rio de Janeiro                              | INMET      | mm            | rain_Rio_de_Janeiro       | 2     |
| 107 | Climate           | Total monthly precipitation (mm) in Salvador                                    | INMET      | mm            | rain_Salvador             | 2     |
| 108 | Climate           | Total monthly precipitation (mm) in São Luís                                    | INMET      | mm            | rain_Sao_Luis             | 2     |
| 109 | Climate           | Total monthly precipitation (mm) in São Paulo                                   | INMET      | mm            | rain_Sao_Paulo            | 2     |
| 110 | Climate           | Total monthly precipitation (mm) in Vitória                                     | INMET      | mm            | rain_Vitoria              | 2     |
| 111 | Public sector     | Primary result of consolidated public sector (current monthly flow)             | BCB        | R\$ million   | primary_result            | 2     |
| 112 | Public sector     | Primary result of consolidated public sector (flow accum. in 12 months)         | BCB        | R\$ million   | primary_result_12m        | 3     |
| 113 | Public sector     | Primary result of consolidated public sector (flow accum. in 12 months, % GDP)  | BCB        | %             | primary_result_%GDP       | 2     |
| 114 | Public sector     | Net public debt (total, federal government and central bank, % GDP)             | BCB        | %             | public_debt_total_%GDP    | 3     |
| 115 | Public sector     | Net public debt (internal, federal government and central bank, % GDP)          | BCB        | %             | public_debt_internal_%GDP | 3     |
| 116 | Public sector     | Net public debt (external, federal government and central bank, % GDP)          | BCB        | %             | public_debt_external_%GDP | 2     |
| 117 | Public sector     | Net public debt (total, consolidated public sector, balances in reais)          | BCB        | R\$ million   | public_debt_total         | 6     |
| 118 | Public sector     | Net public debt (internal, consolidated public sector, balances in reais)       | BCB        | R\$ million   | public_debt_internal      | 6     |
| 119 | Public sector     | Net public debt (external, consolidated public sector, balances in reais)       | BCB        | R\$ million   | public_debt_external      | 2     |
| 120 | Economic activity | IBC-BR (central bank economic activity index)                                   | BCB        | Index         | IBC-BR                    | 5     |
| 121 | Economic activity | GDP (accumulated in the last 12 months, current prices)                         | BCB        | R\$ million   | GDP                       | 6     |
| 122 | Economic activity | Consumer confidence index   | Fecomercio | Index         | consum_confidence         | 2     |
| 123 | Exterior          | Import price index  | Funex      | Index         | import_price              | 6     |
| 124 | Exterior          | Import quantum index  | Funex      | Index         | import_quantum            | 5     |
| 125 | Exterior          | Export price index  | Funex      | Index         | export_price              | 6     |
| 126 | Exterior          | Export quantum index  | Funex      | Index         | export_quantum            | 5     |

(continued on next page)

Table A1 (continued)

|     | Category           | Name   | Source       | Original unit | Nickname               | tcode |
|-----|--------------------|--|--------------|---------------|------------------------|-------|
| 127 | Exterior           | Imports (agriculture, forestry and fishing)                    | MDIC/Secex   | US\$ FOB      | imports_agriculture    | 5     |
| 128 | Exterior           | Imports (mining and quarrying)                                 | MDIC/Secex   | US\$ FOB      | imports_mining         | 5     |
| 129 | Exterior           | Imports (manufacturing)  | MDIC/Secex   | US\$ FOB      | imports_manufacturing  | 5     |
| 130 | Exterior           | Imports (other products)                                       | MDIC/Secex   | US\$ FOB      | imports_others         | 5     |
| 131 | Exterior           | Imports (total)  | MDIC/Secex   | US\$ FOB      | imports_total          | 5     |
| 132 | Exterior           | Exports (agriculture, forestry and fishing)                    | MDIC/Secex   | US\$ FOB      | exports_agriculture    | 5     |
| 133 | Exterior           | Exports (mining and quarrying)                                 | MDIC/Secex   | US\$ FOB      | exports_mining         | 5     |
| 134 | Exterior           | Exports (manufacturing)  | MDIC/Secex   | US\$ FOB      | exports_manufacturing  | 5     |
| 135 | Exterior           | Exports (other products)                                       | MDIC/Secex   | US\$ FOB      | exports_others         | 5     |
| 136 | Exterior           | Exports (total)  | MDIC/Secex   | US\$ FOB      | exports_total          | 5     |
| 137 | Exterior           | International reserves (total)                                 | BCB          | US\$ million  | international_reserves | 6     |
| 138 | Exterior           | Current account (monthly, net)                                 | BCB          | US\$ million  | current_account        | 2     |
| 139 | Exterior           | Current account (accumulated in 12 months, in relation to GDP) | BCB          | %             | current_account_%GDP   | 2     |
| 140 | Exterior           | FDI (Foreign Direct Investment, accumulated in 12 months)      | BCB, authors | US\$ million  | FDI                    | 2     |
| 141 | Exterior           | FPI (Foreign Portfolio Investment, accumulated in 12 months)   | BCB, authors | US\$ million  | FPI                    | 2     |
| 142 | Commodities        | CRB all commodities index                                      | Reuters      | Index         | CRB                    | 5     |
| 143 | Commodities        | CRB foodstuffs index   | Reuters      | Index         | CRB_food               | 5     |
| 144 | Commodities        | CRB metals index   | Reuters      | Index         | CRB_metals             | 5     |
| 145 | Commodities        | Baltic exchange dry index                                      | Reuters      | Index         | Baltic_dry             | 5     |
| 146 | Commodities        | Oil price (Brent, Europe)                                      | Reuters      | US\$/barrel   | Oil_price_Brent        | 5     |
| 147 | Commodities        | Oil price (WTI, Oklahoma-USA)                                  | Reuters      | US\$/barrel   | Oil_price_WTI          | 5     |
| 148 | Global uncertainty | Economic Policy Uncertainty index for Australia                | EPU          | Index         | EPU_Australia          | 2     |
| 149 | Global uncertainty | Economic Policy Uncertainty index for Brazil                   | EPU          | Index         | EPU_Brazil             | 2     |
| 150 | Global uncertainty | Economic Policy Uncertainty index for Canada                   | EPU          | Index         | EPU_Canada             | 2     |
| 151 | Global uncertainty | Economic Policy Uncertainty index for Chile                    | EPU          | Index         | EPU_Chile              | 2     |
| 152 | Global uncertainty | Economic Policy Uncertainty index for China                    | EPU          | Index         | EPU_China              | 2     |
| 153 | Global uncertainty | Economic Policy Uncertainty index for Colombia                 | EPU          | Index         | EPU_Colombia           | 2     |
| 154 | Global uncertainty | Economic Policy Uncertainty index for France                   | EPU          | Index         | EPU_France             | 2     |
| 155 | Global uncertainty | Economic Policy Uncertainty index for Germany                  | EPU          | Index         | EPU_Germany            | 2     |
| 156 | Global uncertainty | Economic Policy Uncertainty index for Greece                   | EPU          | Index         | EPU_Greece             | 2     |
| 157 | Global uncertainty | Economic Policy Uncertainty index for India                    | EPU          | Index         | EPU_India              | 2     |
| 158 | Global uncertainty | Economic Policy Uncertainty index for Ireland                  | EPU          | Index         | EPU_Ireland            | 2     |
| 159 | Global uncertainty | Economic Policy Uncertainty index for Italy                    | EPU          | Index         | EPU_Italy              | 2     |
| 160 | Global uncertainty | Economic Policy Uncertainty index for Japan                    | EPU          | Index         | EPU_Japan              | 2     |
| 161 | Global uncertainty | Economic Policy Uncertainty index for Korea                    | EPU          | Index         | EPU_Korea              | 2     |
| 162 | Global uncertainty | Economic Policy Uncertainty index for Netherlands              | EPU          | Index         | EPU_Netherlands        | 2     |
| 163 | Global uncertainty | Economic Policy Uncertainty index for Russia                   | EPU          | Index         | EPU_Russia             | 2     |
| 164 | Global uncertainty | Economic Policy Uncertainty index for Spain                    | EPU          | Index         | EPU_Spain              | 2     |
| 165 | Global uncertainty | Economic Policy Uncertainty index for Singapore                | EPU          | Index         | EPU_Singapore          | 2     |
| 166 | Global uncertainty | Economic Policy Uncertainty index for UK                       | EPU          | Index         | EPU_UK                 | 2     |
| 167 | Global uncertainty | Economic Policy Uncertainty index for USA                      | EPU          | Index         | EPU_USA                | 2     |

Note: The column "tcode" denotes the following data transformations: (1) no transformation; (2)  $\Delta x_t$ ; (3)  $\Delta^2 x_t$ ; (4)  $\ln(x_t)$ ; (5)  $\Delta \ln(x_t)$ ; (6)  $\Delta^2 \ln(x_t)$ .

## Appendix B. Robustness Analysis

**Table B1**

 Mean Squared Error (MSE),  $T_1 = 96$  months,  $T_2 = 144$  months .

| <i>dep. var. = IPCA % p.m.</i> | <i>h = 1</i> | <i>h = 2</i> | <i>h = 3</i> | <i>h = 6</i> | <i>h = 9</i> | <i>h = 12</i> | <i>h = 15</i> | <i>h = 18</i> | <i>dep. var. = IPCA % 12 months</i> | <i>h = 1</i> | <i>h = 2</i> | <i>h = 3</i> | <i>h = 6</i> | <i>h = 9</i> | <i>h = 12</i> | <i>h = 15</i> | <i>h = 18</i> |
|--------------------------------|--------------|--------------|--------------|--------------|--------------|---------------|---------------|---------------|-------------------------------------|--------------|--------------|--------------|--------------|--------------|---------------|---------------|---------------|
| (1) RW                         | 0.147        | 0.194        | 0.193        | 0.260        | 0.173        | 0.276         | 0.291         | 0.197         | (1) RW                              | 0.301        | 0.932        | 1.708        | 4.552        | 7.829        | 10.303        | 10.471        | 9.399         |
| (2) RW-AO                      | 0.128        | 0.144        | 0.153        | 0.171        | 0.180        | 0.209         | 0.187         | 0.177         | (2) RW-AO                           | 3.806        | 4.836        | 5.870        | 8.845        | 11.415       | 12.964        | 13.108        | 12.423        |
| (3) ARMA                       | 0.113        | 0.130        | 0.129        | 0.144        | 0.139        | 0.155         | 0.155         | 0.152         | (3) ARMA                            | 0.216        | 0.770        | 1.447        | 4.382        | 6.675        | 9.485         | 8.673         | 7.074         |
| (4) VAR                        | 0.117        | 0.138        | 0.137        | 0.150        | 0.152        | 0.167         | 0.162         | 0.164         | (4) VAR                             | 0.282        | 0.862        | 1.564        | 4.160        | 6.704        | 9.845         | 10.358        | 9.060         |
| (5) PC backward                | 0.110        | 0.131        | 0.134        | 0.159        | 0.136        | 0.153         | 0.162         | 0.204         | (5) PC backward                     | 0.278        | 0.873        | 1.587        | 4.327        | 6.438        | 8.826         | 10.590        | 11.585        |
| (6) PC hybrid                  | 0.092***     | 0.124        | 0.136        | 0.157        | 0.143        | 0.163         | 0.170         | 0.192         | (6) PC hybrid                       | 0.271        | 0.830        | 1.481        | 4.270        | 6.941        | 9.553         | 11.995        | 13.619        |
| (7) Factor model1              | 0.091**      | 0.106***     | 0.137        | 0.145        | 0.150        | 0.155         | 0.159         | 0.176         | (7) Factor model1                   | 0.642        | 1.102        | 1.698        | 3.512*       | 5.599        | 7.091         | 6.824         | 6.582         |
| (8) Factor model2              | 0.091***     | 0.117        | 0.131        | 0.147        | 0.144        | 0.155         | 0.151*        | 0.154         | (8) Factor model2                   | 0.365        | 0.780        | 1.316        | 3.479*       | 5.746        | 8.081*        | 8.035         | 6.817         |
| (9) Factor model3              | 0.095*       | 0.106**      | 0.134        | 0.145        | 0.139        | 0.165         | 0.151         | 0.187         | (9) Factor model3                   | 0.352        | 0.865        | 1.561        | 4.620        | 5.538        | 5.619*        | 4.353         | 5.215         |
| (10) Factor model4             | 0.103        | 0.125        | 0.141        | 0.145        | 0.142        | 0.158         | 0.150**       | 0.154         | (10) Factor model4                  | 0.363        | 0.791        | 1.245        | 2.756        | 4.680        | 7.584         | 7.972         | 6.951         |
| (11) Elastic net               | 0.096**      | 0.121        | 0.144        | 0.142        | 0.147        | 0.156         | 0.158         | 0.161         | (11) Elastic net                    | 0.260        | 0.845        | 1.641        | 4.284        | 7.626        | 8.500         | 6.842         | 5.388         |
| (12) Lasso                     | 0.098**      | 0.120        | 0.145        | 0.142        | 0.146        | 0.159         | 0.159         | 0.162         | (12) Lasso                          | 0.255        | 0.825        | 1.663        | 4.164        | 7.764        | 8.708         | 7.094         | 5.367         |
| (13) Adalasso                  | 0.093***     | 0.113*       | 0.141        | 0.143        | 0.150        | 0.169         | 0.163         | 0.160         | (13) Adalasso                       | 0.230        | 0.764        | 1.536        | 4.546        | 8.137        | 8.959         | 8.594         | 7.620         |
| (14) Ridge                     | 0.101        | 0.122        | 0.138        | 0.142        | 0.161        | 0.169         | 0.181         | 0.187         | (14) Ridge                          | 0.939        | 1.389        | 1.997        | 4.017        | 5.784        | 5.950**       | 5.660         | 5.398         |
| (15) Random Forest             | 0.105        | 0.123        | 0.142        | 0.150        | 0.149        | 0.155         | 0.153         | 0.172         | (15) Random Forest                  | 0.771        | 1.449        | 2.087        | 4.008        | 5.816        | 6.533         | 5.903         | 5.257         |
| (16) Quant Regr. Forest        | 0.105        | 0.122        | 0.141        | 0.146        | 0.148        | 0.154         | 0.149         | 0.167         | (16) Quant Regr. Forest             | 0.767        | 1.452        | 2.094        | 3.967        | 5.671        | 6.379         | 5.657         | 5.010         |
| (17) XGBoost                   | 0.102        | 0.118        | 0.148        | 0.168        | 0.160        | 0.159         | 0.150         | 0.185         | (17) XGBoost                        | 0.331        | 0.960        | 1.799        | 4.592        | 6.804        | 6.548         | 5.088         | 4.230         |
| (18) RNN                       | 0.142        | 0.143        | 0.169        | 0.234        | 0.146        | 0.133*        | 0.163         | 0.186         | (18) RNN                            | 3.087        | 8.169        | 5.318        | 26.928       | 1.916        | 10.601        | 12.438        | 2.992         |
| (19) Disag. ARMA               | 0.118        | 0.148        | 0.155        | 0.169        | 0.175        | 0.199         | 0.181         | 0.169         | (19) Disag. ARMA                    | 0.214        | 0.762        | 1.476        | 4.522        | 7.471        | 10.794        | 9.828         | 7.874         |
| (20) Disag. Adalasso           | 0.099*       | 0.127        | 0.147        | 0.137        | 0.156        | 0.150         | 0.152         | 0.153         | (20) Disag. Adalasso                | 0.268        | 0.790        | 1.616        | 5.400        | 8.028        | 7.460*        | 7.151         | 6.533         |
| (21) Disag. RF                 | 0.120        | 0.139        | 0.148        | 0.160        | 0.150        | 0.156         | 0.163         | 0.181         | (21) Disag. RF                      | 0.972        | 1.743        | 2.510        | 4.734        | 6.123        | 6.739         | 7.071         | 6.814         |
| (22) Hybrid Ada-OLS            | 0.098**      | 0.124        | 0.137        | 0.144        | 0.157        | 0.185         | 0.136**       | 0.201         | (22) Hybrid Ada-OLS                 | 0.218        | 0.774        | 1.669        | 5.472        | 11.954       | 9.838         | 7.371         | 6.019         |
| (23) Hybrid Ada-RF             | 0.105        | 0.114        | 0.131        | 0.159        | 0.147        | 0.165         | 0.144*        | 0.195         | (23) Hybrid Ada-RF                  | 0.391        | 1.129        | 1.974        | 4.297        | 5.955        | 7.007         | 6.951         | 6.367         |
| (24) Hybrid Ada-XGB            | 0.113        | 0.122        | 0.159        | 0.228        | 0.183        | 0.180         | 0.161         | 0.204         | (24) Hybrid Ada-XGB                 | 0.340        | 1.044        | 2.111        | 5.271        | 6.772        | 7.111         | 6.182         | 5.215         |
| (25) Hybrid RF-OLS             | 0.110        | 0.115        | 0.160        | 0.156        | 0.211        | 0.250         | 0.205         | 0.202         | (25) Hybrid RF-OLS                  | 0.193        | 0.821        | 1.544        | 6.175        | 13.578       | 15.006        | 12.185        | 6.459         |
| (26) Hybrid RF-Ada             | 0.100*       | 0.112*       | 0.145        | 0.136*       | 0.153        | 0.169         | 0.173         | 0.178         | (26) Hybrid RF-Ada                  | 0.199        | 0.861        | 1.784        | 5.656        | 14.539       | 14.741        | 11.851        | 6.218         |
| (27) Hybrid RF-XGB             | 0.096        | 0.116        | 0.157        | 0.159        | 0.163        | 0.165         | 0.146*        | 0.208         | (27) Hybrid RF-XGB                  | 0.372        | 1.283        | 2.469        | 5.110        | 7.086        | 6.846         | 4.780         | 5.121         |
| (28) BEI                       | 0.082*       | 0.168        | 0.143        | 0.158        | 0.162        | 0.161         | 0.162         | 0.173         | (28) BEI                            | 0.088***     | 0.464*       | 0.666**      | 2.323**      | 4.641***     | 6.126*        | 6.192         | 6.404         |
| (29) Focus                     | 0.083**      | 0.125        | 0.138        | 0.150        | 0.157        | 0.161         | 0.159         | 0.167         | (29) Focus                          | 0.090***     | 0.373***     | 0.742**      | 2.193**      | 3.706**      | 4.687*        | 4.564*        | 4.269         |
| (30) Comb1 Mean                | 0.095***     | 0.114**      | 0.138        | 0.145        | 0.145        | 0.160         | 0.152         | 0.164         | (30) Comb1 Mean                     | 0.329        | 0.869        | 1.568        | 4.473        | 6.465        | 7.452         | 6.690         | 5.067         |
| (31) Comb1 Median              | 0.095**      | 0.113**      | 0.134        | 0.145        | 0.146        | 0.157         | 0.151         | 0.167         | (31) Comb1 Median                   | 0.247        | 0.808        | 1.514        | 4.117        | 6.177        | 7.156*        | 6.437         | 5.520         |
| (32) Comb1 GR                  | 0.099        | 0.147        | 0.175        | 0.273        | 0.295        | 0.265         | 0.354         | 0.714         | (32) Comb1 GR                       | 0.292        | 1.034        | 1.809        | 5.305        | 16.059       | 31.148        | 25.465        | 28.669        |
| (33) Comb1 CLS                 | 0.094***     | 0.114**      | 0.137        | 0.152        | 0.149        | 0.175         | 0.158         | 0.170         | (33) Comb1 CLS                      | 0.348        | 0.797        | 1.553        | 5.392        | 6.601        | 8.251         | 6.979         | 5.235         |
| (34) Comb1 CSR                 | 0.099**      | 0.125        | 0.155        | 0.193        | 0.188        | 0.202         | 0.216         | 0.250         | (34) Comb1 CSR                      | 0.338        | 0.848        | 1.609        | 6.025        | 11.365       | 18.254        | 20.333        | 21.751        |
| (35) Comb1 Adalasso            | 0.091**      | 0.122        | 0.155        | 0.195        | 0.243        | 0.189         | 6.423         | 0.198         | (35) Comb1 Adalasso                 | 0.301        | 0.927        | 1.563        | 9.059        | 29.811       | 36.481        | 31.575        | 27.743        |
| (36) Comb1 RF                  | 0.099*       | 0.135        | 0.178        | 0.228        | 0.195        | 0.198         | 0.222         | 0.216         | (36) Comb1 RF                       | 0.477        | 1.186        | 2.172        | 6.225        | 9.113        | 12.270        | 16.555        | 13.888        |
| (37) Comb2 Mean                | 0.092***     | 0.116        | 0.139        | 0.173        | 0.160        | 0.201         | 0.156         | 0.175         | (37) Comb2 Mean                     | 0.210        | 0.776        | 1.453        | 4.297        | 6.872        | 8.582         | 9.470         | 5.994         |
| (38) Comb2 Median              | 0.092***     | 0.116        | 0.139        | 0.173        | 0.156        | 0.201         | 0.156         | 0.177         | (38) Comb2 Median                   | 0.210        | 0.776        | 1.453        | 4.297        | 6.872        | 8.582         | 9.470         | 7.254         |
| (39) Comb2 GR                  | 0.096***     | 0.129        | 0.163        | 0.196        | 0.221        | 0.177         | 0.200         | 0.214         | (39) Comb2 GR                       | 0.229        | 0.834        | 1.608        | 5.697        | 12.588       | 10.144        | 24.401        | 24.850        |
| (40) Comb2 CLS                 | 0.093***     | 0.120        | 0.136        | 0.166        | 0.154        | 0.167         | 0.162         | 0.178         | (40) Comb2 CLS                      | 0.214        | 0.788        | 1.513        | 4.387        | 7.085        | 8.805         | 10.766        | 8.046         |
| (41) Comb2 CSR                 | 0.096***     | 0.129        | 0.163        | 0.196        | 0.192        | 0.177         | 0.200         | 0.234         | (41) Comb2 CSR                      | 0.219        | 0.818        | 1.611        | 5.719        | 12.678       | 10.142        | 33.575        | 24.581        |
| (42) Comb2 Adalasso            | 0.099**      | 0.127        | 0.158        | 0.190        | 0.181        | 0.172         | 0.204         | 0.214         | (42) Comb2 Adalasso                 | 0.224        | 0.810        | 1.626        | 5.353        | 12.348       | 10.611        | 32.409        | 23.788        |
| (43) Comb2 RF                  | 0.105        | 0.132        | 0.179        | 0.218        | 0.179        | 0.192         | 0.202         | 0.237         | (43) Comb2 RF                       | 0.331        | 1.106        | 2.130        | 5.908        | 10.036       | 11.546        | 18.029        | 12.989        |
| (44) Comb3 Mean                | 0.075**      | 0.135        | 0.135        | 0.151        | 0.159        | 0.160         | 0.163         | 0.169         | (44) Comb3 Mean                     | 0.082***     | 0.395***     | 0.951**      | 4.297        | 6.872        | 8.582         | 9.470         | 8.579         |
| (45) Comb3 Median              | 0.075**      | 0.135        | 0.135        | 0.151        | 0.159        | 0.160         | 0.163         | 0.169         | (45) Comb3 Median                   | 0.082***     | 0.395***     | 0.951**      | 4.297        | 6.872        | 8.582         | 9.470         | 8.579         |
| (46) Comb3 GR                  | 0.066***     | 0.135        | 0.146        | 0.188        | 0.182        | 0.172         | 0.186         | 0.191         | (46) Comb3 GR                       | 0.088***     | 0.444**      | 0.829*       | 5.697        | 12.588       | 10.144        | 24.401        | 31.376        |
| (47) Comb3 CLS                 | 0.083*       | 0.152        | 0.138        | 0.157        | 0.161        | 0.167         | 0.170         | 0.169         | (47) Comb3 CLS                      | 0.089***     | 0.476*       | 0.727**      | 4.387        | 7.085        | 8.805         | 10.766        | 10.846        |
| (48) Comb3 CSR                 | 0.096        | 0.165        | 0.150        | 0.188        | 0.182        | 0.172         | 0.186         | 0.191         | (48) Comb3 CSR                      | 0.105**      | 0.471*       | 0.731**      | 5.719        | 12.678       | 10.142        | 33.575        | 33.768        |
| (49) Comb3 Adalasso            | 0.089        | 0.163        | 0.152        | 0.186        | 0.178        | 0.171         | 0.188         | 0.188         | (49) Comb3 Adalasso                 | 0.087***     | 0.506        | 0.710**      | 5.353        | 12.348       | 10.611        | 32.409        | 51.737        |
| (50) Comb3 RF                  | 0.075**      | 0.141        | 0.161        | 0.198        | 0.194        | 0.187         | 0.198         | 0.204         | (50) Comb3 RF                       | 0.165        | 0.645        | 1.388        | 5.908        | 10.036       | 11.546        | 18.029        | 19.453        |
| number of observations         | 64           | 63           | 62           | 59           | 56           | 53            | 50            | 47            | number of observations              | 64           | 63           | 62           | 59           | 56           | 53            | 50            | 47            |
| best model                     | 46           | 7            | 3            | 26           | 5            | 18            | 22            | 3             | best model                          | 45           | 29           | 28           | 29           | 18           | 29            | 9             | 18            |
| R2 oos (%)                     | 41           | 18           | 0            | 5            | 2            | 13            | 12            | 0             | R2 oos (%)                          | 62           | 51           | 53           | 49           | 71           | 50            | 49            | 57            |

Notes: Yellow cells denote Top10 models (lowest MSEs) in each horizon. \*\*\*, \*\*, and \* indicate rejection at 1%, 5%, and 10% levels, respectively, using the [Diebold and Mariano \(1995\)](#) test, considering model3 (ARMA) as a benchmark. The  $R^2$  out-of-sample statistics (R2 oos) refer to the best model in each horizon. Source: Authors' own elaboration.

**Table B2**Mean Squared Error (MSE),  $T_1 = 72$  months,  $T_2 = 144$  months.

| dep. var. = IPCA % p.m. | h = 1    | h = 2    | h = 3 | h = 6  | h = 9 | h = 12   | h = 15  | h = 18 | dep. var. = IPCA % 12 months | h = 1    | h = 2    | h = 3   | h = 6    | h = 9    | h = 12  | h = 15 | h = 18 |
|-------------------------|----------|----------|-------|--------|-------|----------|---------|--------|------------------------------|----------|----------|---------|----------|----------|---------|--------|--------|
| (1) RW                  | 0.147    | 0.194    | 0.193 | 0.260  | 0.173 | 0.276    | 0.291   | 0.197  | (1) RW                       | 0.301    | 0.952    | 1.708   | 4.552    | 7.329    | 10.303  | 10.471 | 9.399  |
| (2) RW-AO               | 0.128    | 0.144    | 0.153 | 0.171  | 0.180 | 0.209    | 0.187   | 0.177  | (2) RW-AO                    | 3.806    | 4.836    | 5.870   | 8.845    | 11.415   | 12.964  | 13.108 | 12.423 |
| (3) ARMA                | 0.113    | 0.130    | 0.129 | 0.144  | 0.139 | 0.155    | 0.155   | 0.152  | (3) ARMA                     | 0.216    | 0.770    | 1.447   | 4.382    | 6.675    | 9.485   | 8.673  | 7.074  |
| (4) VAR                 | 0.117    | 0.138    | 0.137 | 0.150  | 0.152 | 0.167    | 0.162   | 0.164  | (4) VAR                      | 0.282    | 0.862    | 1.564   | 4.160    | 6.704    | 9.845   | 10.358 | 9.060  |
| (5) PC backward         | 0.110    | 0.131    | 0.134 | 0.159  | 0.136 | 0.158    | 0.162   | 0.204  | (5) PC backward              | 0.278    | 0.873    | 1.587   | 4.327    | 6.438    | 8.826   | 10.590 | 11.585 |
| (6) PC hybrid           | 0.092*** | 0.124    | 0.136 | 0.157  | 0.143 | 0.163    | 0.170   | 0.192  | (6) PC hybrid                | 0.271    | 0.830    | 1.481   | 4.270    | 6.941    | 9.553   | 11.995 | 13.619 |
| (7) Factor model1       | 0.091**  | 0.106*** | 0.137 | 0.145  | 0.150 | 0.155    | 0.159   | 0.176  | (7) Factor model1            | 0.642    | 1.102    | 1.698   | 3.512*   | 5.599    | 7.091   | 6.824  | 6.582  |
| (8) Factor model2       | 0.091*** | 0.117    | 0.131 | 0.147  | 0.144 | 0.155    | 0.151*  | 0.154  | (8) Factor model2            | 0.365    | 0.780    | 1.316   | 3.479*   | 5.746    | 8.081*  | 8.035  | 6.817  |
| (9) Factor model3       | 0.095*   | 0.106**  | 0.134 | 0.145  | 0.139 | 0.165    | 0.151   | 0.187  | (9) Factor model3            | 0.352    | 0.865    | 1.561   | 4.620    | 5.588    | 5.619*  | 4.353  | 5.215  |
| (10) Factor model4      | 0.103    | 0.125    | 0.141 | 0.145  | 0.142 | 0.156    | 0.150** | 0.154  | (10) Factor model4           | 0.363    | 0.791    | 1.245   | 2.756    | 4.680    | 7.584   | 7.972  | 6.951  |
| (11) Elastic net        | 0.096**  | 0.121    | 0.144 | 0.142  | 0.147 | 0.156    | 0.158   | 0.161  | (11) Elastic net             | 0.260    | 0.845    | 1.641   | 4.284    | 7.626    | 8.500   | 6.942  | 5.388  |
| (12) Lasso              | 0.098**  | 0.120    | 0.145 | 0.142  | 0.146 | 0.159    | 0.159   | 0.162  | (12) Lasso                   | 0.255    | 0.825    | 1.663   | 4.164    | 7.764    | 8.708   | 7.094  | 5.367  |
| (13) Adalasso           | 0.093*** | 0.113*   | 0.141 | 0.143  | 0.150 | 0.169    | 0.165   | 0.160  | (13) Adalasso                | 0.230    | 0.764    | 1.536   | 4.546    | 8.137    | 8.959   | 8.594  | 7.620  |
| (14) Ridge              | 0.101    | 0.122    | 0.138 | 0.142  | 0.161 | 0.169    | 0.181   | 0.187  | (14) Ridge                   | 0.939    | 1.389    | 1.997   | 4.017    | 5.784    | 5.950** | 5.660  | 5.398  |
| (15) Random Forest      | 0.105    | 0.123    | 0.142 | 0.150  | 0.149 | 0.155    | 0.153   | 0.172  | (15) Random Forest           | 0.771    | 1.449    | 2.087   | 4.005    | 5.816    | 6.533   | 5.903  | 5.257  |
| (16) Quant Regr. Forest | 0.105    | 0.122    | 0.141 | 0.146  | 0.148 | 0.154    | 0.149   | 0.167  | (16) Quant Regr. Forest      | 0.767    | 1.452    | 2.094   | 3.967    | 5.671    | 6.379   | 5.657  | 5.010  |
| (17) XGBoost            | 0.102    | 0.118    | 0.148 | 0.168  | 0.160 | 0.159    | 0.150   | 0.185  | (17) XGBoost                 | 0.331    | 0.960    | 1.799   | 4.592    | 6.804    | 6.548   | 5.088  | 4.230  |
| (18) RNN                | 0.146    | 0.138    | 0.148 | 0.149  | 0.134 | 0.135*** | 0.136*  | 0.166  | (18) RNN                     | 2.420    | 2.349    | 1.940   | 4.051    | 2.635    | 3.716   | 6.300  | 3.050  |
| (19) Disag. ARMA        | 0.118    | 0.148    | 0.155 | 0.159  | 0.175 | 0.199    | 0.183   | 0.169  | (19) Disag. ARMA             | 0.214    | 0.762    | 1.476   | 4.522    | 7.471    | 10.794  | 9.828  | 7.874  |
| (20) Disag. Adalasso    | 0.099*   | 0.127    | 0.147 | 0.137  | 0.156 | 0.150    | 0.152   | 0.155  | (20) Disag. Adalasso         | 0.268    | 0.790    | 1.616   | 5.400    | 8.028    | 7.460*  | 7.151  | 6.553  |
| (21) Disag. RF          | 0.120    | 0.139    | 0.148 | 0.160  | 0.150 | 0.156    | 0.163   | 0.181  | (21) Disag. RF               | 0.972    | 1.743    | 2.510   | 4.734    | 6.123    | 6.739   | 7.071  | 6.814  |
| (22) Hybrid Ada-OLS     | 0.098**  | 0.124    | 0.137 | 0.144  | 0.157 | 0.185    | 0.136** | 0.201  | (22) Hybrid Ada-OLS          | 0.218    | 0.774    | 1.669   | 5.472    | 11.954   | 9.838   | 7.371  | 6.019  |
| (23) Hybrid Ada-RF      | 0.105    | 0.114    | 0.131 | 0.159  | 0.147 | 0.168    | 0.144   | 0.195  | (23) Hybrid Ada-RF           | 0.391    | 1.129    | 1.974   | 4.297    | 5.955    | 7.007   | 6.951  | 6.367  |
| (24) Hybrid Ada-XGB     | 0.113    | 0.122    | 0.159 | 0.228  | 0.183 | 0.180    | 0.161   | 0.204  | (24) Hybrid Ada-XGB          | 0.340    | 1.044    | 2.111   | 5.271    | 6.772    | 7.111   | 6.182  | 5.215  |
| (25) Hybrid RF-OLS      | 0.110    | 0.115    | 0.160 | 0.156  | 0.211 | 0.250    | 0.205   | 0.202  | (25) Hybrid RF-OLS           | 0.193    | 0.821    | 1.544   | 6.175    | 13.578   | 15.006  | 12.185 | 6.459  |
| (26) Hybrid RF-Ada      | 0.100*   | 0.112*   | 0.145 | 0.136* | 0.153 | 0.169    | 0.173   | 0.178  | (26) Hybrid RF-Ada           | 0.199    | 0.861    | 1.784   | 5.658    | 14.539   | 14.741  | 11.851 | 6.218  |
| (27) Hybrid RF-XGB      | 0.096    | 0.116    | 0.157 | 0.159  | 0.163 | 0.165    | 0.146   | 0.208  | (27) Hybrid RF-XGB           | 0.372    | 1.283    | 2.469   | 5.110    | 7.086    | 6.846   | 4.780  | 5.121  |
| (28) BEI                | 0.082*   | 0.168    | 0.143 | 0.158  | 0.162 | 0.161    | 0.162   | 0.173  | (28) BEI                     | 0.088*** | 0.464*   | 0.666** | 2.322**  | 4.641*** | 6.126*  | 6.192  | 6.404  |
| (29) Focus              | 0.083**  | 0.125    | 0.138 | 0.150  | 0.157 | 0.161    | 0.159   | 0.167  | (29) Focus                   | 0.090*** | 0.373*** | 0.742** | 2.193*** | 3.706**  | 4.687*  | 4.564* | 4.269  |
| (30) Combi Mean         | 0.095**  | 0.114**  | 0.133 | 0.144  | 0.144 | 0.160    | 0.152   | 0.164  | (30) Combi Mean              | 0.311    | 0.864    | 1.546   | 4.095    | 6.488    | 7.324*  | 6.541  | 5.020  |
| (31) Combi Median       | 0.095**  | 0.113**  | 0.134 | 0.144  | 0.146 | 0.156    | 0.150   | 0.166  | (31) Combi Median            | 0.247    | 0.807    | 1.497   | 4.078    | 6.192    | 7.060*  | 6.432  | 5.524  |
| (32) Combi GR           | 0.094*   | 0.138    | 0.154 | 0.251  | 0.225 | 0.326    | 0.374   | 0.429  | (32) Combi GR                | 0.265    | 0.919    | 2.022   | 5.911    | 9.757    | 25.777  | 24.184 | 15.396 |
| (33) Combi CLS          | 0.094*** | 0.114**  | 0.135 | 0.143  | 0.144 | 0.160    | 0.155   | 0.166  | (33) Combi CLS               | 0.318    | 0.882    | 1.650   | 4.952    | 6.428    | 7.748*  | 6.671  | 5.103  |
| (34) Combi CSR          | 0.098**  | 0.123    | 0.149 | 0.182  | 0.177 | 0.196    | 0.205   | 0.243  | (34) Combi CSR               | 0.341    | 0.851    | 1.644   | 5.626    | 10.313   | 14.179  | 16.628 | 17.264 |
| (35) Combi Adalasso     | 0.089*** | 0.123    | 0.143 | 0.196  | 0.171 | 0.217    | 0.241   | 0.323  | (35) Combi Adalasso          | 0.274    | 0.792    | 1.472   | 7.815    | 13.681   | 31.862  | 35.690 | 20.999 |
| (36) Combi RF           | 0.100*   | 0.140    | 0.170 | 0.197  | 0.201 | 0.200    | 0.180   | 0.208  | (36) Combi RF                | 0.436    | 1.144    | 2.216   | 5.985    | 9.291    | 13.587  | 16.528 | 13.533 |
| (37) Comb2 Mean         | 0.092*** | 0.120    | 0.142 | 0.159  | 0.159 | 0.172    | 0.155   | 0.175  | (37) Comb2 Mean              | 0.219    | 0.753    | 1.435   | 4.501    | 6.872    | 8.556   | 10.809 | 10.900 |
| (38) Comb2 Median       | 0.092*** | 0.120    | 0.142 | 0.159  | 0.155 | 0.165    | 0.155   | 0.179  | (38) Comb2 Median            | 0.219    | 0.753    | 1.435   | 4.501    | 6.872    | 8.556   | 10.809 | 10.900 |
| (39) Comb2 GR           | 0.097*** | 0.131    | 0.163 | 0.199  | 0.180 | 0.208    | 0.191   | 0.341  | (39) Comb2 GR                | 0.230    | 0.801    | 1.530   | 4.896    | 8.957    | 12.357  | 23.214 | 13.804 |
| (40) Comb2 CLS          | 0.093*** | 0.122    | 0.140 | 0.157  | 0.162 | 0.178    | 0.154   | 0.187  | (40) Comb2 CLS               | 0.216    | 0.779    | 1.518   | 4.723    | 6.982    | 9.342   | 11.881 | 10.148 |
| (41) Comb2 CSR          | 0.096*** | 0.131    | 0.163 | 0.199  | 0.178 | 0.186    | 0.191   | 0.263  | (41) Comb2 CSR               | 0.220    | 0.784    | 1.522   | 4.898    | 8.957    | 12.357  | 24.944 | 13.804 |
| (42) Comb2 Adalasso     | 0.096*** | 0.132    | 0.161 | 0.196  | 0.178 | 0.184    | 0.187   | 0.359  | (42) Comb2 Adalasso          | 0.217    | 0.795    | 1.585   | 5.026    | 7.780    | 12.339  | 25.242 | 13.455 |
| (43) Comb2 RF           | 0.102**  | 0.129    | 0.171 | 0.201  | 0.180 | 0.180    | 0.208   | 0.258  | (43) Comb2 RF                | 0.313    | 0.950    | 1.803   | 5.725    | 7.752    | 13.480  | 15.294 | 15.459 |
| (44) Comb3 Mean         | 0.075**  | 0.137    | 0.187 | 0.151  | 0.159 | 0.160    | 0.163   | 0.169  | (44) Comb3 Mean              | 0.082*** | 0.395*** | 0.698** | 3.795    | 6.199    | 8.556   | 8.378  | 10.900 |
| (45) Comb3 Median       | 0.075**  | 0.137    | 0.187 | 0.151  | 0.159 | 0.160    | 0.163   | 0.169  | (45) Comb3 Median            | 0.082*** | 0.395*** | 0.698** | 4.024    | 6.188    | 8.556   | 8.378  | 10.900 |
| (46) Comb3 GR           | 0.066*** | 0.143    | 0.147 | 0.171  | 0.178 | 0.198    | 0.185   | 0.187  | (46) Comb3 GR                | 0.086*** | 0.436**  | 0.798** | 4.161    | 11.209   | 12.357  | 16.751 | 13.804 |
| (47) Comb3 CLS          | 0.083*   | 0.154    | 0.145 | 0.154  | 0.160 | 0.161    | 0.169   | 0.169  | (47) Comb3 CLS               | 0.089*** | 0.476*   | 0.666** | 3.055**  | 5.808    | 9.342   | 7.330  | 10.148 |
| (48) Comb3 CSR          | 0.096    | 0.146    | 0.148 | 0.171  | 0.178 | 0.198    | 0.185   | 0.187  | (48) Comb3 CSR               | 0.105**  | 0.487*   | 0.707** | 3.991    | 7.982    | 12.357  | 16.751 | 13.804 |
| (49) Comb3 Adalasso     | 0.089    | 0.144    | 0.148 | 0.166  | 0.177 | 0.197    | 0.188   | 0.185  | (49) Comb3 Adalasso          | 0.086*** | 0.502    | 0.718** | 4.063    | 11.143   | 12.339  | 13.966 | 13.455 |
| (50) Comb3 RF           | 0.075**  | 0.144    | 0.147 | 0.184  | 0.189 | 0.178    | 0.188   | 0.186  | (50) Comb3 RF                | 0.150    | 0.635    | 1.194   | 5.202    | 8.642    | 13.480  | 15.786 | 15.459 |
| number of observations  | 64       | 63       | 62    | 59     | 56    | 53       | 50      | 47     | number of observations       | 64       | 63       | 62      | 59       | 56       | 53      | 50     | 47     |
| best model              | 46       | 7        | 3     | 26     | 18    | 18       | 18      | 3      | best model                   | 45       | 29       | 28      | 29       | 18       | 18      | 9      | 18     |
| R2 oos (%)              | 41       | 18       | 0     | 5      | 3     | 12       | 12      | 0      | R2 oos (%)                   | 62       | 51       | 53      | 49       | 60       | 60      | 49     | 56     |

Notes: Yellow cells denote Top10 models (lowest MSEs) in each horizon. \*\*\*, \*\*, and \* indicate rejection at 1%, 5%, and 10% levels, respectively, using the [Diebold and Mariano \(1995\)](#) test, considering model3 (ARMA) as a benchmark. The  $R^2$  out-of-sample statistics (R2 oos) refer to the best model in each horizon. Source: Authors' own elaboration.

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