

## Lezione 3 -

Note on notation

$$I_A(x) = \begin{cases} 1 & \text{if } x \in A \\ 0 & \text{if } x \notin A \end{cases}$$

The expectation is the weighted center

The variance is the dispersion of the random variable.

Sometimes we are interested in some function which manipulates the random variable

$$y = g(x) \rightarrow y = x^2, y = \frac{1}{x}, y = \log(x) \text{ if } x > 0$$

↳ In this course we will be looking at cases where  $y$  remains a random variable.

We can calculate the expectation of  $y$ , knowing  $X$ , as

$$E[y] = \int_R g(x) \cdot f_x(x) dx$$

$$E[X^n] = \int_R x^n \cdot f_x(x) dx \rightarrow n^{\text{th}} \text{ moment of } x$$

$$\text{Var}(X) = E(X^2) - \underbrace{E(X)}_{\substack{\text{Second} \\ \text{moment}}}^2$$

Square of first moment

### Exercise

Let  $X$  be a absolutely continuous random variable with density

$$f_x = \begin{cases} cx(1-x) & \text{if } 0 < x < 1 \\ 0 & \text{otherwise} \end{cases} = cx(1-x) I_{(0,1)}(x)$$

1. Find  $c$  such that  $f_x$  is a density of a absolutely continuous random variable

continuity  $\hookrightarrow i) c x(1-x) \geq 0 \quad \forall x \in (0,1) \Rightarrow c \geq 0$

Integral  $\hookrightarrow ii) \int_{\mathbb{R}} f_x(x) dx = \int_0^1 c x(1-x) dx = c \int_0^1 x - x^2 dx$   
 $= c \left[ \frac{x^2}{2} - \frac{x^3}{3} \right]_0^1 = \frac{c}{6} \Rightarrow c = 6 \geq 0 \checkmark$

2.  $E X = \int_{\mathbb{R}} x \cdot f_x(x) dx = \int_0^1 x \cdot 6 \cdot x(1-x) dx = 6 \int_0^1 (x^2 - x^3) dx$   
 $= 6 \left[ \frac{x^3}{3} - \frac{x^4}{4} \right]_0^1 = \frac{1}{2}$

$$\text{Var}(X) = E(X^2) - E(X)^2 = \frac{3}{10} - \left(\frac{1}{2}\right)^2 = \frac{3}{10} - \frac{1}{4} = \frac{1}{20}$$

$$E(X^2) = \int_0^1 x^2 \cdot 6x(1-x) dx = 6 \int_0^1 x^3 - x^4 dx = 6 \left[ \frac{x^4}{4} - \frac{x^5}{5} \right]_0^1 = \frac{3}{10}$$

$$\text{sd}(X) = \sqrt{\text{Var}(X)} = \frac{1}{\sqrt{20}}$$

Definition : Median

d.f.  $F_x$  is  $\hookrightarrow$  the median  $m$  of an absolutely continuous random variable with

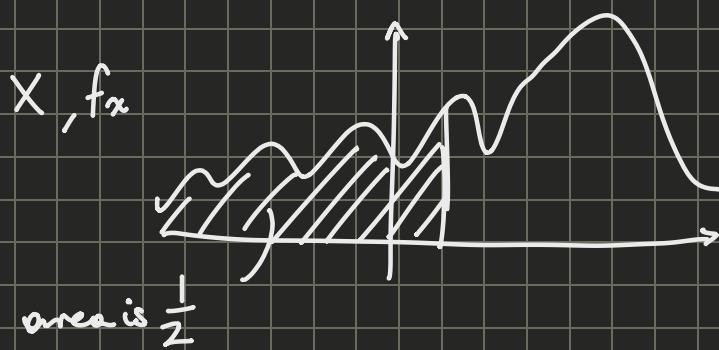
$$P(X < m) = P(X \leq m) = \frac{1}{2}$$

or equivalently

$$m = F_x^{-1}\left(\frac{1}{2}\right) \Leftrightarrow F(m) = \frac{1}{2}$$

We can also define an  $\alpha$ -quantile, we need a real value  $x_\alpha$

$$P(X \leq x_\alpha) = \alpha$$

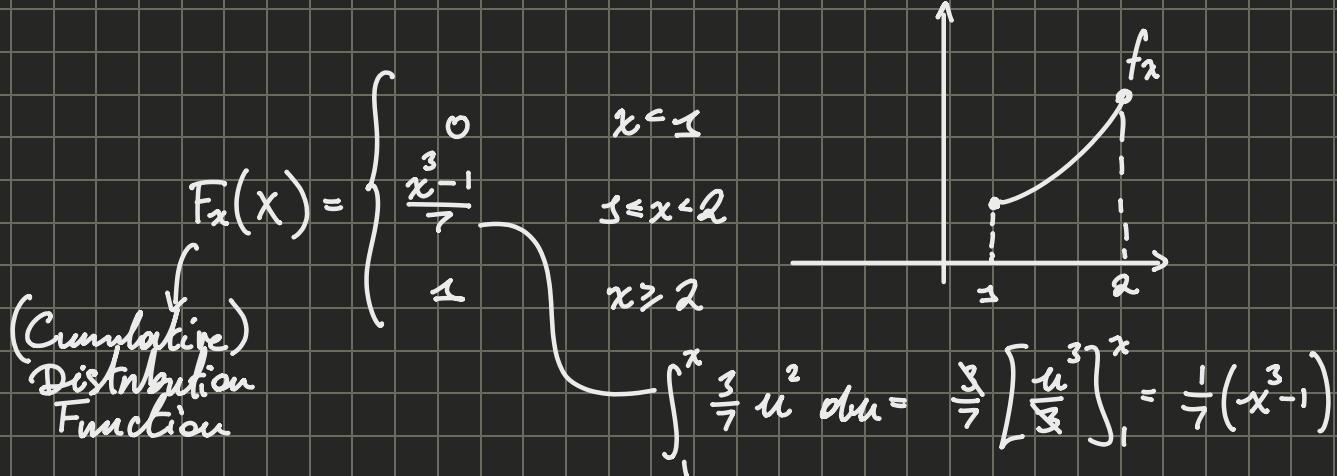


if  $\alpha = \frac{1}{2} \Rightarrow \alpha$ -quantile is the median



### Exercise

Let  $X$  be abs. continuous r.v., with density  $f_x(x) = \frac{3}{7}x^2 I_{[1,2]}(x)$



$$F_x(m) = \frac{1}{2} \Leftrightarrow \frac{m^3 - 1}{7} = \frac{1}{2} \Rightarrow m = \sqrt[3]{\frac{7}{2} + 1} = \sqrt[3]{\frac{9}{2}}$$

Homework : Compute the 25% quantile of  $X$ .

Density of  $Y = g(X)$  where  $X$  is an abs. continuous r.v and density  $f_X$

$\rightarrow$  Another abs. cont...

$Y$  is absolutely continuous, when  $g$  and  $g'$  are differentiable

Step 3: we compute the distribution function of  $y$ :

$$F_Y(y) = P(Y \leq y) = P(g(X) \leq y) = P(X \in g^{-1}((-\infty, y]))$$

$g(X) \in (-\infty, y]$

We need to find the intervals in which  $g(x) \leq y$

$\rightarrow$  Basically solving this inequality.

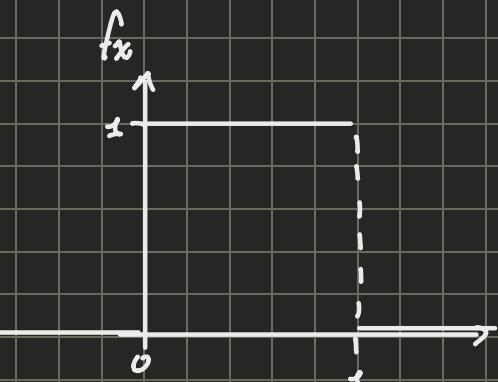
Step 2:  $F_Y'(y)$  in  $y$  where the derivative of  $F(y)$  exists

$$F_Y'(y) = f_Y(y)$$

Example 1

$$X, f_X(x) = I_{(0,1)}(x)$$

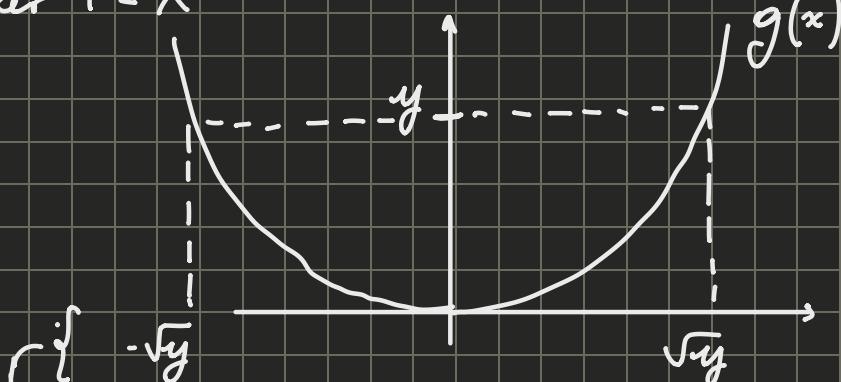
$$F_X(x) = \begin{cases} 0 & \text{if } x < 0 \\ x & \text{if } 0 \leq x < 1 \\ 1 & \text{if } x \geq 1 \end{cases}$$



It's a density because it's always positive and the area is 1.

Let us consider  $Y = X^2$

$$g(x) = x^2$$



Step 1  $\rightarrow$  for  $y < 0$   $F_Y(y) = 0$

$$F_Y(y) = P(Y \leq y) = P(X^2 \leq y) =$$

$y > 0$

$$X^2 \leq y \Leftrightarrow -\sqrt{y} \leq X \leq \sqrt{y}$$

when  $y > 0$

$$= P(-\sqrt{y} \leq X \leq \sqrt{y}) \xrightarrow{\text{for here we defined } f_X} = P(0 \leq X \leq \sqrt{y}) = F_X(\sqrt{y})$$

$$F_Y(y) = \begin{cases} 0 & \text{if } y = 0 \\ \sqrt{y} & \text{if } 0 \leq \sqrt{y} \leq 1 \\ 1 & \text{if } \sqrt{y} \geq 1 \end{cases} = \begin{cases} 0 & \text{if } y < 0 \\ \sqrt{y} & \text{if } 0 \leq y < 1 \\ 1 & \text{if } y \geq 1 \end{cases}$$

Step 2

$$F_Y(y) = \begin{cases} 0 & \text{if } y < 0, y \geq 1 \\ \frac{1}{2\sqrt{y}} & \text{if } 0 \leq y < 1 \end{cases} \Rightarrow f_Y(y) = \begin{cases} \frac{1}{2\sqrt{y}} & 0 < y < 1 \\ 0 & \text{if } y \leq 0, y \geq 1 \end{cases}$$

Example 2 (Linear Transformation)

Let  $X$  be an abs. conti. r.v. with  $f_X$  density, and  $a, b \in \mathbb{R}$  with  $a \neq 0$

let us take  $Y = aX + b$

e.g.  $X = \text{temp. } C^\circ \Rightarrow Y = \text{temp. } F^\circ$

$$Y = \frac{9}{5}X + 32 \quad a = \frac{9}{5} \quad b = 32$$

Step 1

$$F_Y(y) = P(X \leq y) = P(aX + b \leq y) = P(aX \leq y - b)$$

$$P(X \leq \frac{y-b}{a}) \quad \text{if } a > 0$$

$$P(X \geq \frac{y-b}{a}) \quad \text{if } a < 0$$

$$F_X\left(\frac{y-b}{a}\right)$$

$$P(X > \frac{y-b}{a}) \\ \text{if } a < 0$$

$$1 - F_X\left(\frac{y-b}{a}\right)$$

Step 2

$$F_Y'(y) = \begin{cases} \text{if } a > 0 & f_X\left(\frac{y-b}{a}\right) \cdot \frac{1}{a} \\ \text{if } a < 0 & -f_X\left(\frac{y-b}{a}\right) \cdot \frac{1}{a} = f_X\left(\frac{y-b}{a}\right) \cdot \left(-\frac{1}{a}\right) \end{cases}$$

If we want to write one since expression we can write the following:

$$f_Y(y) = f_X\left(\frac{y-b}{a}\right) \cdot \frac{1}{|a|}$$

$$Y = aX + b$$

### Notable Absolutely Continuous Distributions

Distribution

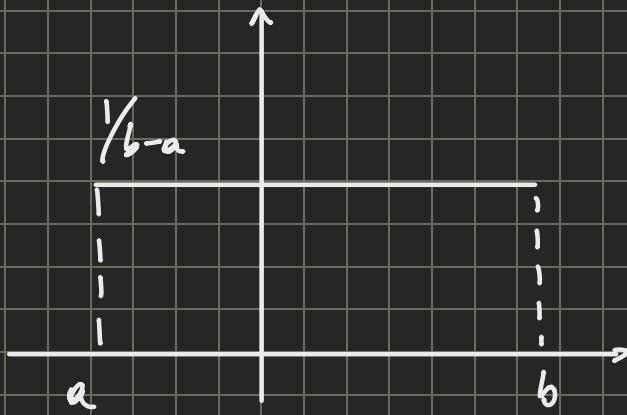
Uniform on interval  $(a, b)$ :

if  $X \sim U((a, b))$  is r.v with density

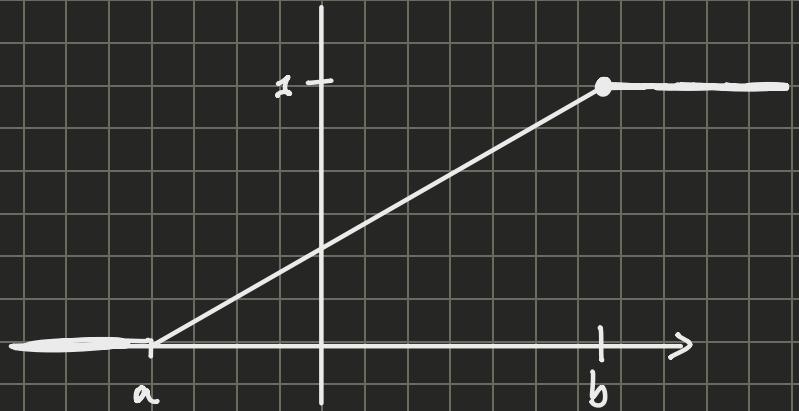
$$f_X(x) = \frac{1}{b-a} I_{(a, b)}(x)$$

Uniform on  $(a, b)$

$$f_x = \begin{cases} \frac{1}{b-a} & \text{if } a < x < b \\ 0 & \text{otherwise} \end{cases}$$



$$F_x(x) = \begin{cases} 0 & \text{if } x < a \\ \frac{x-a}{b-a} & \text{if } a \leq x < b \\ 1 & \text{if } x \geq b \end{cases}$$



$$EX = \frac{a+b}{2}$$

$$\text{Var}(X) = \frac{(b-a)^2}{12}$$



The probability of it being in an interval around an  $x$  is independent of  $x$ .

if  $a=0$  &  $b=1$ ,

the distribution is the same as before.