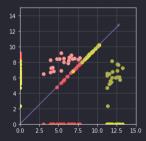
Linear Discriminant Analysis

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- LDA as a feature extractor
 - $f_E : \mathbb{R}^n \mapsto \mathbb{R}^m$ via linearly combining the original features
- LDA as a classification technique
 - Maximization of some "class discriminatory information"



-Measure of separability

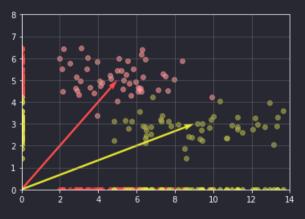
$$\bullet \ \omega_i = \{x_j = (x_{j1}, \ldots, x_{jn}) | j = 1, \ldots, p_i\}$$

- Consider a vector w
- $\tilde{\omega}_i = \{ y_i := \langle w, x_i \rangle | j = 1, \dots, p_i \}$
- $\omega_i = \{y_j := \langle w, x_j \rangle | j = 1, \dots, p_j \}$
- $\mu_i = E[x], \tilde{\mu_i} = E[y]$
 - $\blacksquare \ \widetilde{\mu}_i = E[w^t x] = w^t E[x] = \langle w, \mu_i \rangle$

-Measure of separability

- One possible way is to consider the separation between the means.
- $\mathcal{J}(w) = |\tilde{\mu_1} \tilde{\mu_2}|$, for i = 1, 2
- Possibly $\mathcal{J}(w) = \sum_{i,j:i \neq j} |\tilde{\mu_i} \tilde{\mu_j}|$, for a p class set-up
- Not a convenient measure, however.
- Well separated means does not necessarily imply well-separated class clusters (fig in the following slide).

-Measure of separability



Drawback of mean-based separation

-Measure of separability: Fisher's LDA

- To account for the *spread* of a class, consider the within-class scatter.
- ullet $ilde{s}_i^2 := \sum_{\mathbf{v} \in \omega_i} (y ilde{\mu}_i)^2$
- Within-class scatter $\tilde{s}_W := \sum_i \tilde{s}_i^2$
- Fisher linear discriminant analysis: find w that maximizes

$$\mathcal{J}(w) := rac{| ilde{\mu}_1 - ilde{\mu}_2|^2}{ ilde{s}_W}$$

Can be interpreted as maximally separated means with squeezed classes.

-Measure of separability: Fisher's LDA

- We try to express \mathcal{J} in terms of w.
- $\tilde{s}_i^2 = \sum_{y \in \omega_i} (y \tilde{\mu}_i)^2 = \sum_{x \in \omega_i} (w^t x w^t \mu_i)^2 = \sum_{x \in \omega_i} w^t (x \mu_i) (x \mu_i) w = w^t S_i w$
- $\sum_{i} \tilde{s}_{i}^{2} = \sum_{i} w^{t} S_{i} w = w^{t} S_{W} w$
 - $S_W := \sum_i (x \mu_i)(x \mu_i)^t$
 - Termed as the within-class scatter matrix
- $(\tilde{\mu}_1 \tilde{\mu}_2)^2 = (w^t \mu_1 w^t \mu_2)^2 = w^t (\mu_1 \mu_2)(\mu_1 \mu_2)^t w = w^t S_B w$
 - \blacksquare Where S_B is the between-class scatter

-Solving for Fisher's linear discriminant

•
$$\mathcal{J}(w) = \frac{w^t S_B w}{w^t S_W w}$$

 $\implies S_{RW} = \mathcal{J}(w)S_{WW}$

• In order to maximize \mathcal{J} wrt w, we equate $\frac{\partial \mathcal{J}}{\partial w}$ to 0.

$$\bullet \frac{\partial \mathcal{J}}{\partial w} = 0$$

$$\Rightarrow \frac{\partial (w^t S_B w)}{\partial w} (w^t S_W w) = (w^t S_B w) \frac{\partial (w^t S_W w)}{\partial w}$$

$$\Rightarrow (2S_B w) \cdot (w^t S_W w) = (w^t S_B w) \cdot (2S_W w)$$

$$\Rightarrow S_B w = \frac{w^t S_B w}{w^t S_W w} S_W w$$

-Solving for Fisher's linear discriminant

S_B is singular

- $S_B = uu^t$, where $u = (\mu_1 \mu_2)$
- Say, $u = (a_1, \ldots, a_n)^t$
- $r_2[uu^t] = a_2(a_1, \dots, a_n) = \frac{a_2}{a_1}a_1(a_1, \dots, a_n) = \frac{a_2}{a_1}r_1[uu^t]$
- Thus $\operatorname{rank}(S_B) \leq 1$ and $\operatorname{rank}(S_B) = 1$, when $\mu_1 \neq \mu_2$
- det $S_B = 0 \implies S_B^{-1}$ is undefined

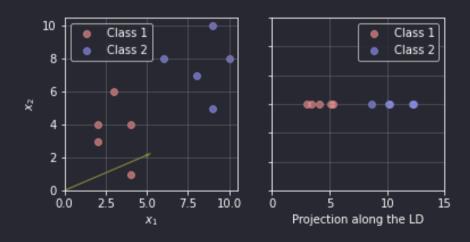
-Solving for Fisher's linear discriminant

- Continuing with $S_B w = \mathcal{J} S_W w ...$
- $\bullet \ S_W^{-1}S_Bw=\mathcal{J}(w)w$
 - \blacksquare S_B is not invertible
 - $\blacksquare \mathcal{J}(w)$ is a scalar
 - \blacksquare S_W is has to be nonsingular
- Note that $\operatorname{rank}(S_W^{-1}S_B) \leq \min \left[\operatorname{rank}(S_W^{-1}), \operatorname{rank}(S_B)\right]$
- Also, # distinct nonzero eigenvalues \le rank

-Solving for Fisher's linear discriminant

- w^* be the solution
- For a given distribution $(\mu_1 \mu_2)^t w^* = \lambda$
- $egin{array}{ccc} lacksquare & S_W^{-1} S_B w^* = S_W^{-1} (\mu_1 \mu_2) (\mu_1 \mu_2)^t w^* = \mathcal{J}(w^*) w^* \ \Longrightarrow & rac{\lambda}{\mathcal{J}(w)} S_W^{-1} (\mu_1 \mu_2) = w^* \end{array}$
- As only the direction is of concern, we find $w^* = S_W^{-1}(\mu_1 - \mu_2)$

-A 1D example



Multi-class LDA: Generalization

- As before, $\omega_i = \{x_j = (x_{j1}, \dots, x_{jn}) | j = 1, \dots, p_i\}, i = 1, \dots, c$
- Consider $y_j = (y_{j1}, \dots, y_{j(c-1)})$ where $y_{ji} = \langle w_i, x_j \rangle$
 - Thus, $y = W^t x$, $W = [w_1 | w_2 | \dots | w_{c-1}]$
- $S_W = \sum_{i=1}^c S_i = \sum_{i=1}^c \sum_{x \in \omega_i} (x \mu_i)(x \mu_i)^t$ • $\mu_i = \frac{1}{2} \sum_{k=1}^{p_i} x_k$
- $S_B := \sum_{i=1}^c p_i (\mu_i \mu) (\mu_i \mu)^t$
 - $\blacksquare \mu = E(p_i \mu_i), i = 1, \dots, c$
- ullet We proceed to find $ilde{S_W}, ilde{S_B}$

Multi-class LDA: Generalization

•
$$v = W^t x$$

•
$$\tilde{S_W} = \sum_{i=1}^c (y - \tilde{\mu_i})(y - \tilde{\mu_i})^t = \sum_{i=1}^c W^t S_i W = W^t S_W W$$

• $\tilde{\mu_i} = \tilde{V_{ou}} = W^t \tilde{X_{ou}} = W^t \mu_i$

$$\bullet \tilde{S_B} = \sum_{i=1}^c p_i (\tilde{\mu_i} - \tilde{\mu}) (\tilde{\mu_i} - \tilde{\mu})^t$$

$$= \sum_{i=1}^c p_i W^t (\mu_i - \mu) (\mu_i - \mu)^t W = W^t S_B W$$

•
$$\mathcal{J}(W) = \frac{|\tilde{S_B}|}{|\tilde{S_W}|} = \frac{|W^t S_B W|}{|W^t S_W W|}$$

•
$$W^* = \operatorname{argmax} \mathcal{J}(W)$$

Multi-class LDA: Solution

- The optimal solution for $W^* = [w_1^* | \dots | w_{c-1}^*]$ is given by- $(S_B - \lambda_i S_W) w_i^* = 0$
 - How so?
- If S_W is non-singular, W^* would contain columns that are eigenvectors of $S_{M}^{-1}S_{B}$

Multi-class LDA: Solution

- As discussed earlier, rank $(uu^t) \leq 1$ where u is any arbitrary vector
- $\operatorname{rank}(S_B) = \operatorname{rank}(\sum_{i=1}^c p_i(\mu_i \mu)(\mu_i \mu)^t) \le c 1$
 - $\blacksquare a_i := (\mu_i \mu)$
 - $lacksquare \sum_{i=1}^{c} p_i a_i a_i^t = \sum_{i=1}^{c-1} p_i a_i a_i^t + p_c (\mu_c \mu) (\mu_c \mu)^t$
 - $\blacksquare \mu_c \mu = (c-1)\mu \sum_{i=1}^{c-1} \mu_i$
 - \blacksquare rank $(S_B) \leq c 1$
- $\operatorname{rank}(S_W) = \operatorname{rank}(\sum_{i=1}^c \sum_{x \in \omega_i} (x \mu_i)(x \mu_i)^t) \leq N c$
 - $\blacksquare N = \sum_i p_i$

Multi-class LDA: Solution and limitaions

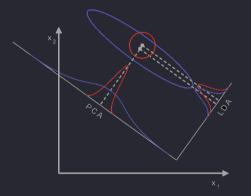
-For c-class set-up

- # distinct nonzero eigenvalues < rank < c-1
- LDA, thus, can project upon maximum c-1 features, considering c-1 eigenvectors, corresponding to the largest eigenvalues as w_i 's

Limitations

- The earlier discussion brings us to the first limitation: *feature-space* with restricted dimensions
- LDA might not be very effective in separating complex distributions
- It is bound to fail when the discriminatory information is not explained by the means

Multi-class LDA: Limitations

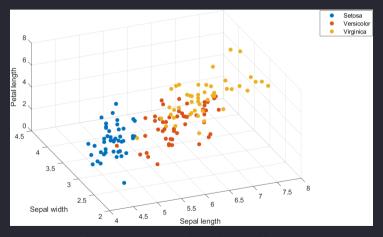


Discriminatory information in variance¹

¹Credit: LDA. CSCE 666 Pattern Analysis. Ricardo Gutierrez-Osuna

Multi-class LDA:

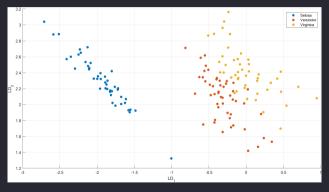
-A 3-class example on the Iris data-set



Param: sepal length, width, petal length

Multi-class LDA:

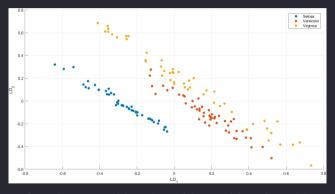
-A 3-class example on the Iris data-set



LDA considering three parameters mentioned earlier

Multi-class LDA:

-A 3-class example on the Iris data-set



Considering sepal length, width; petal length, width

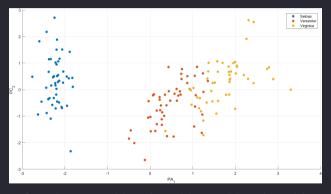
PCA against LDA:

- PCA is unsupervised; LDA needs class-labels
- PCA represents the maximum variance;
 LDA is about maximizing the discriminatory information
- "PCA does more of feature classification and LDA does data classification"

²Comparision of PCA and LDA for Face Recognition. Begum, Sajjan. 2013.

PCA against LDA

-On the Iris data-set



Considering sepal length, width; petal length, width