

# Introduction

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- Predict with uncertainty
- Optimize any black box function
- Efficiently create a training set
- Generative modelling

# Predict with Uncertainty: Classification

# Predict with Uncertainty: Regression

# Questions

- We used squared error loss function for linear regression. Why?
- We used cross entropy loss function for logistic regression. Why?
- How does `np.random.randn` work?
- `np.std(x)` and `pd.std(x)` give different results. Why?

## How: Bayes Rule

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

= Rewriting it using the ML notation:

$$P(\theta|D) = \frac{P(D|\theta) \cdot P(\theta)}{P(D)}$$

- $P(\theta|D)$  is called the posterior
- $P(D|\theta)$  is called the likelihood
- $P(\theta)$  is called the prior
- $P(D)$  is called the evidence

# Univariate Normal Distribution

The probability density function of a univariate normal distribution is given by:

$$f(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right) \quad (1)$$

Let us assume we have a dataset  $D = \{x_1, x_2, \dots, x_n\}$ , where each  $x_i$  is an independent sample from the above distribution. We want to estimate the parameters  $\theta = \{\mu, \sigma\}$  from the data.

Our likelihood function is given by:

$$P(D|\theta) = \mathcal{L}(\mu, \sigma^2) = \prod_{i=1}^n f(x_i|\mu, \sigma^2) \quad (2)$$

# Log Likelihood Function

Log-likelihood function:

$$\log \mathcal{L}(\mu, \sigma^2) = \sum_{i=1}^n \log f(x_i | \mu, \sigma^2) \quad (3)$$

Simplifying the above equation, we get:

$$\begin{aligned} \log \mathcal{L}(\mu, \sigma^2) &= \sum_{i=1}^n \log f(x_i | \mu, \sigma^2) \\ &= \sum_{i=1}^n \log \left( \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left( -\frac{(x_i - \mu)^2}{2\sigma^2} \right) \right) \\ &= \sum_{i=1}^n \left( \log \left( \frac{1}{\sqrt{2\pi\sigma^2}} \right) + \log \left( \exp \left( -\frac{(x_i - \mu)^2}{2\sigma^2} \right) \right) \right) \end{aligned}$$



$$\begin{aligned}
 \log \mathcal{L}(\mu, \sigma^2) &= \sum_{i=1}^n \left( \log \left( \frac{1}{\sqrt{2\pi\sigma^2}} \right) - \frac{(x_i - \mu)^2}{2\sigma^2} \right) \\
 &= \sum_{i=1}^n \left( -\frac{1}{2} \log(2\pi\sigma^2) - \frac{(x_i - \mu)^2}{2\sigma^2} \right) \\
 &= -\frac{n}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2
 \end{aligned}$$

### Log Likelihood Function for Univariate Normal Distribution

Log-likelihood function for normally distributed data is:

$$\log \mathcal{L}(\mu, \sigma^2) = -\frac{n}{2} \log(2\pi) - n \log(\sigma) - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2$$

## Maximum Likelihood Estimate for $\mu$

To find the MLE for  $\mu$ , we differentiate the log-likelihood function with respect to  $\mu$  and set it to zero:

$$\frac{\partial \log \mathcal{L}(\mu, \sigma^2)}{\partial \mu} = \frac{\partial}{\partial \mu} \left( -\frac{n}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2 \right) = 0$$
$$\frac{\partial}{\partial \mu} \left( \sum_{i=1}^n (x_i - \mu)^2 \right) = 0$$

### Maximum Likelihood Estimate for $\mu$

MLE of  $\mu$ , denoted as  $\hat{\mu}_{\text{MLE}}$ , is given by:

$$\hat{\mu}_{\text{MLE}} = \frac{1}{n} \sum_{i=1}^n x_i$$

## MLE for $\sigma$ for normally distributed data

Recall that the log-likelihood function is given by:

$$\log \mathcal{L}(\mu, \sigma^2) = \sum_{i=1}^n \log f(x_i | \mu, \sigma^2) \quad (4)$$

Let us find the maximum likelihood estimate of  $\sigma^2$  now. We can do this by taking the derivative of the log-likelihood function with respect to  $\sigma^2$  and equating it to zero.

$$\frac{\partial \log \mathcal{L}(\mu, \sigma^2)}{\partial \sigma^2} = \sum_{i=1}^n \frac{\partial \log f(x_i | \mu, \sigma^2)}{\partial \sigma^2} = 0 \quad (5)$$

## MLE for $\sigma$ for normally distributed data

### Log Likelihood Function for Univariate Normal Distribution

Log-likelihood function for normally distributed data is:

$$\log \mathcal{L}(\mu, \sigma^2) = -\frac{n}{2} \log(2\pi) - n \log(\sigma) - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2$$

Now, we can differentiate the log-likelihood function with respect to  $\sigma$  and equate it to zero.

## MLE for $\sigma$ for normally distributed data

$$\frac{\partial}{\partial \sigma} \log \mathcal{L}(\mu, \sigma^2) = -\frac{n}{\sigma} + \frac{1}{\sigma^3} \sum_{i=1}^n (x_i - \mu)^2 = 0$$

Multiplying through by  $\sigma^3$ , we have:

$$-n\sigma^2 + \sum_{i=1}^n (x_i - \mu)^2 = 0$$

Maximum Likelihood Estimate for  $\sigma^2$

MLE of  $\sigma^2$ , denoted as  $\hat{\sigma}_{\text{MLE}}^2$ , is given by:

$$\sigma^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \mu)^2$$

## Bias of an Estimator

The bias of an estimator  $\hat{\theta}$  of a parameter  $\theta$  is defined as:

$$\text{Bias}(\hat{\theta}) = \mathbb{E}(\hat{\theta}) - \theta$$

where  $\mathbb{E}(\hat{\theta})$  is the expected value of the estimator  $\hat{\theta}$ .

- An estimator is said to be unbiased if  $\text{Bias}(\hat{\theta}) = 0$ .
- An estimator is said to be biased if  $\text{Bias}(\hat{\theta}) \neq 0$ .

## Bias of an Estimator: $\hat{\mu}_{MLE}$

Question: What is the expectation of  $\hat{\mu}_{MLE}$  calculated over? What is the source of randomness?

Let us assume that the true underlying distribution is  $\mathcal{N}(\mu, \sigma^2)$ .

Let  $\mathcal{D}^1 = \{x_1^1, x_2^1, \dots, x_n^1\}$  be a dataset obtained from this distribution.

The MLE of  $\mu$  based on  $\mathcal{D}^1$  is given by:

$$\hat{\mu}_{MLE}^1 = \frac{1}{n} \sum_{i=1}^n x_i^1$$

If we obtained another dataset  $\mathcal{D}^2 = \{x_1^2, x_2^2, \dots, x_n^2\}$  from the same distribution, the MLE of  $\mu$  based on  $\mathcal{D}^2$  would be:

$$\hat{\mu}_{MLE}^2 = \frac{1}{n} \sum_{i=1}^n x_i^2$$

## Bias of an Estimator: $\hat{\mu}_{MLE}$

If we repeat this process and obtain datasets  $\mathcal{D}^1, \mathcal{D}^2, \dots, \mathcal{D}^k$ , we would have  $k$  different estimates of  $\mu$ .

Taking the expectation of these  $k$  estimates gives us the expected value of  $\hat{\mu}_{MLE}$ :

$$\mathbb{E}(\hat{\mu}_{MLE}) = \frac{1}{k} \sum_{i=1}^k \hat{\mu}_{MLE}^i$$

Simplifying further, we have:

$$\mathbb{E}(\hat{\mu}_{MLE}) = \frac{1}{kn} \sum_{i=1}^k \sum_{j=1}^n x_j^i$$

This expectation is calculated over multiple datasets  $\mathcal{D}^1, \mathcal{D}^2, \dots, \mathcal{D}^k$ , where each dataset represents a different realization of the random variables from the underlying distribution.



## Bias of an Estimator: $\hat{\mu}_{MLE}$

To show that the estimator  $\hat{\mu}_{MLE}$  is unbiased, we need to demonstrate that  $\mathbb{E}(\hat{\mu}_{MLE}) = \mu$ .

Recall that each  $x_j^i$  is a random variable following  $\mathcal{N}(\mu, \sigma^2)$ . Therefore, the sum  $\sum_{i=1}^k x_j^i$  follows  $\mathcal{N}(k\mu, k\sigma^2)$ .

Thus, we can write:

$$\begin{aligned}\mathbb{E}(\hat{\mu}_{MLE}) &= \frac{1}{kn} \sum_{i=1}^k \sum_{j=1}^n x_j^i = \frac{1}{kn} \sum_{j=1}^n \left( \sum_{i=1}^k x_j^i \right) \\ &= \frac{1}{kn} \sum_{j=1}^n (k\mu) = \frac{1}{kn} (kn\mu) = \mu\end{aligned}$$

Estimator  $\hat{\mu}_{MLE}$  is unbiased

$$\mathbb{E}(\hat{\mu}_{MLE}) = \mu$$

## Bias of $\sigma_{MLE}^2$

The MLE of  $\sigma^2$  is given by

$\hat{\sigma}_{MLE}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \mu)^2$  where  $\mu$  is the MLE of the mean.

$$\begin{aligned}\mathbb{E}(\hat{\sigma}_{MLE}^2) &= \mathbb{E} \left[ \frac{1}{n} \sum_{i=1}^n (x_i - \mu)^2 \right] = \frac{1}{n} \sum_{i=1}^n \mathbb{E}[(x_i - \mu)^2] \\ &= \frac{1}{n} \sum_{i=1}^n \mathbb{E}[x_i^2] - 2\mu \mathbb{E}[x_i] + \mu^2 = \frac{1}{n} \sum_{i=1}^n \sigma^2 + \mu^2 - 2\mu\mu \\ &= \frac{n-1}{n} \sigma^2 + \mu^2 - \mu^2 = \frac{n-1}{n} \sigma^2\end{aligned}$$

Estimator  $\hat{\sigma}_{MLE}^2$  is biased

$$\mathbb{E}(\hat{\sigma}_{MLE}^2) = \frac{n-1}{n} \sigma^2$$











