

# Time Series Analysis

## Homework 2

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Due: Friday, February 5th

### 1 Problem 5

Let  $X_1, X_2, \dots, X_N$  be independent and identically distributed (iid) random vectors with  $EX_i = 0$  and  $EX_i^2 = 1$ . Compute

$$E\left(\sum_{i=1}^n \frac{1}{i} X_i\right)^2$$

and show

$$\left|\sum_{i=1}^n \frac{1}{i} X_i\right| = Op(1).$$

We know  $X_i \sim \text{iid}(0, 1)$ , let's let

$$Y_n = \sum_{i=1}^n \frac{1}{i} X_i = (1 \cdot X_1 + \frac{1}{2} \cdot X_2 + \frac{1}{3} \cdot X_3 + \dots + \frac{1}{n} \cdot X_n).$$

We need to solve for  $E[Y_n^2]$ . We compute by the following:

$$\text{Var}(Y_n) = E(Y_n^2) - E(Y_n)^2 \Rightarrow E(Y_n^2) = E(Y_n)^2 + \text{Var}(Y_n)$$

$$E(Y_n) = E\left[\sum_{i=1}^n \frac{1}{i} X_i\right] \stackrel{\text{iid}}{=} \sum_{i=1}^n \frac{1}{i} E(X_i) = 0$$

We know that  $E(X_i) = 0$ , therefore,  $\sum_{i=1}^n \frac{1}{i} E(X_i) = 0$ . This leads to the understanding that  $E(Y_n)^2 = 0$ . Now we solve for  $\text{Var}(Y_n)$  as follows:

$$E(Y_n^2) = 0 + \text{Var}(Y_n) = \text{Var}\left[\sum_{i=1}^n \frac{1}{i} X_i\right] \stackrel{\text{iid}}{=} \sum_{i=1}^n \frac{1}{i^2} \text{Var}(X_i) = \sum_{i=1}^n \frac{1}{i^2}$$

We know that  $\text{Var}(X_i) = 1$ , therefore,  $\sum_{i=1}^n \frac{1}{i^2}$  is a harmonic series.

Now we show that  $Y_n$  is bounded in probability using the Chebyshev's Inequality. Note: Used Wolfram Alpha calculation to find what the harmonic series converges to.

$$\begin{aligned} P\{|Y_n| > c\} &\leq \frac{1}{c^2} E(Y_n^2) \\ P\{|Y_n| > c\} &\leq \frac{1}{c^2} \left(\sum_{i=1}^n \frac{1}{i^2}\right) \end{aligned}$$

Now as  $n \rightarrow \infty$ , the series converges to  $\frac{1}{c^2} \left(\frac{\pi^2}{6}\right)$ , which means that  $Y_n$  is bounded by  $\frac{\pi^2}{6c^2}$ .

## 2 Problem 7

Let  $X_1, X_2, \dots, X_n$  be a sequence of random variables with  $EX_i = 0$ ,

$$EX_i X_j = \begin{cases} \sigma^2, & i = j, \\ \rho, & |i - j| = 1, \\ 0, & |i - j| > 1. \end{cases}$$

Compute

$$E\left(\sum_{i=1}^n X_i\right)^2$$

and show that

$$\frac{1}{n} \sum_{i=1}^n X_i \xrightarrow{P} 0.$$

If we let  $Y_n = \sum_{i=1}^n X_i = (X_1 + X_2 + \dots + X_n)$ , we want to solve for  $E(Y_n^2)$  again. Recall from above  $\text{Var}(Y_n) = E(Y_n^2) - E(Y_n)^2$  and  $E(Y_n^2) = E(Y_n)^2 + \text{Var}(Y_n)$ . Using the same process,

$$E(Y_n) = E\left[\sum_{i=1}^n X_i\right] \stackrel{\text{iid}}{=} \sum_{i=1}^n E(X_i) = 0$$

Therefore,  $E(Y_n)^2 = 0$ .

$$E(Y_n^2) = 0 + \text{Var}(Y_n) = E\left[\sum_{i=1}^n \sum_{j=1}^n X_i X_j\right] \stackrel{\text{iid}}{=} \sum_{i=1}^n \sum_{j=1}^n E(X_i X_j)$$

We arrive at the following:  $E(Y_n^2) = n\sigma^2 + 2(n-1)p$ . It is important to note that if  $Y_n$  is a random sample from a distribution with a finite mean and variance, then the sequence of sample mean converges in probability to  $\mu$ , as follows  $\bar{Y}_n \xrightarrow{P} \mu$ .

We need to show:  $\frac{1}{n} \sum_{i=1}^n X_i \xrightarrow{P} 0$ , hence we start with this:  $\frac{1}{n} \sum_{i=1}^n X_i = \bar{Y}_n$ .

Given finite mean 0 and variance ( $\sigma^2$ ), we are going to use (1) stochastic convergence to a constant (c) theorem and (2) Chebishev's inequality.

(1) By definition, if  $\lim_{n \rightarrow \infty} P[|Y_n - c| > \epsilon] = 0$  for all  $\epsilon > 0$  the sequence has convergence in probability to  $c$ .

(2) Likewise by definition,  $P(|\bar{Y}_n - \mu| \geq \epsilon) \leq \frac{\sigma^2}{\epsilon^2}$ . Thus,

$$P[|Y_n - \mu| \geq \epsilon] \geq 1 - \frac{\sigma^2}{\epsilon^2} \text{ and } P[|\bar{Y}_n - \mu| \geq \epsilon] \geq 1 - \frac{\sigma^2}{\epsilon^2 n}, \text{ and } \bar{Y}_n \xrightarrow{P} \mu \text{ as } n \rightarrow \infty.$$

As  $\mu = 0$  and  $\text{Var}(\bar{Y}_n) = (\frac{1}{n^2})(n\sigma^2 + 2(n-1)p)$ ,

$$P[|\bar{Y}_n| \geq \epsilon] \geq 1 - \frac{n\sigma^2 + 2(n-1)p}{\epsilon^2 n^2}.$$

As  $\frac{\text{Var}(\bar{Y}_n)}{n\epsilon^2} \rightarrow 0$  as  $n \rightarrow \infty$ ,  $\bar{Y}_n \xrightarrow{P} \mu$ .