

# 1 Reservoir Sampling.

```
function ReservoirSample (stream, k) {  
    sample = [] /* empty list for the output sample */  
    i = 0 /* index used to go through the stream */  
    while i < k :  
        sample[i] = stream[i]  
        i += 1  
    while stream[i] is not empty :  
        p = random [0, i)  
        if p < k :  
            sample[p] = stream[i]  
        i += 1  
    return sample  
}
```

Prove for Correctness by Induction:

At the beginning, my algorithm copies first  $k$  elements from stream to the sample, which creates the basis for proof by induction.

Base case: the algorithm trivially works for  $k=1$ .

Induction assumption: for a stream with  $k$  elements, all elements are chosen with the same final probability  $\frac{1}{k}$ .

Inductive step: to show for a stream with  $k+1$  elements, all elements have same probability of  $\frac{1}{k+1}$  to be sampled.

From my second loop, the probability to choose the next element is  $\frac{1}{k+1}$ , and all other elements can stay with prob  $\frac{k}{k+1}$  by assumption.

So, the current reservoir element stay with prob  $1 - \frac{1}{k+1} = \frac{k}{k+1}$ .

Therefore, all previous elements have final prob of  $\frac{1}{k} \cdot \frac{k}{k+1} = \frac{1}{k+1}$  to be the reservoir element after this iteration. Thus, all elements still have same prob of being selected as the reservoir element.

## 2 Median trick

Theorem: Let  $X$  be an unbiased estimator of a quantity  $Q$ . Let  $\{X_{ij}\}_{i \in [t], j \in [k]}$  be a collection of independent RVs with  $X_{ij}$  distributed identically to  $X$ , where

$$t = O\left(\log \frac{1}{\delta}\right), k = O\left(\frac{\text{Var}[X]}{\epsilon^2 E[X]^2}\right)$$

Let  $Z = \text{median}_{i \in [t]} \frac{1}{k} \sum_{j=1}^k X_{ij}$ . Then,  $\Pr(|Z - Q| \geq \epsilon Q) \leq \delta$ .

median of means for each row.

**Proof sketch:** Chebyshev and Chernoff. (Homework problem)

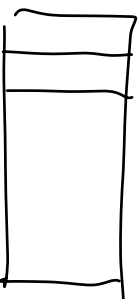
$$t = c_1 \log \frac{1}{\delta}, k = c_2 \frac{\text{Var}(X)}{\epsilon^2 E(X)^2}$$

$$\text{Chebyshev: } P(|Z - E(Z)| \geq k) \leq \frac{\text{Var}(Z)}{k^2}.$$

$$\text{Chernoff: } P(|Z - E(Z)| \geq \delta E(Z)) \leq 2e^{-\frac{\delta^2}{3} E(Z)}$$

$$\text{for each row, } P(|\bar{X} - E(X)| > kQ) \leq \frac{\text{Var}(X)}{k^2 E(X)^2}$$

$$\text{Since } k = c_2 \frac{\text{Var}(X)}{\epsilon^2 E(X)^2}, P(|\bar{X} - E(X)| > kQ) \leq \frac{1}{c_2}$$

$\begin{matrix} 1 \\ 2 \\ \vdots \\ t \end{matrix}$ 

 $\leq \frac{1}{c_2}$  Let  $S$  represent the estimated median for all rows  
 $\leq \frac{1}{c_2}$  Let  $R_i = \begin{cases} 1 & \text{if row } i \text{ has mean greater than median} \\ 0 & \text{o.w.} \end{cases}$   

$$S = \sum_{i=1}^t R_i$$

$$E(S) = np = t \cdot \frac{1}{c_2} = \frac{t}{c_2}$$

$$\begin{aligned} \Pr(|Z - Q| \geq \epsilon Q) &= P(S \geq \frac{t}{2}) \\ &= P(|S - E(S)| \geq \frac{t}{2} - E(S)) \\ &= P(|S - E(S)| \geq \frac{t}{2} - \frac{t}{c_2}) \end{aligned}$$

$$\frac{t}{2} - \frac{t}{c_2} = \frac{(c_2 - 2)t}{2c_2} = \left\lfloor \frac{c_2 - 2}{2} \right\rfloor \cdot \frac{t}{c_2}$$

Then, use Chernoff bound:  $P(|S - E(S)| \geq \delta E(S))$

$$0 < \delta = \frac{C_2 - 2}{2} < 1$$

$$0 < C_2 - 2 < 2$$

$$2 < C_2 < 4, \text{ which forces } \delta \text{ to be in bound } (0, 1)$$

$$\text{Using Chernoff bound, } P(|S - E(S)| \geq \delta E(S)) \leq \underline{2e^{-\frac{\delta^2}{3} E(S)}}$$

$$\text{where } \delta = \frac{C_2 - 2}{2}$$

$$\begin{aligned} \underline{2e^{-\frac{\delta^2}{3} E(S)}} &= 2e^{-\frac{\delta^2}{3} \cdot \frac{t}{C_2}} = 2e^{-\frac{(C_2 - 2)^2}{3 \cdot 4} \cdot \frac{C_1}{C_2} \log(\frac{1}{\delta})} \\ &= 2e^{-\frac{C_1 (C_2 - 2)^2}{12 C_2} \log \delta} \\ &= 2\delta^{\frac{C_1 (C_2 - 2)^2}{12 C_2}} \\ &= 2\delta^{\frac{C_1}{36}} \end{aligned}$$

$$\text{Let } C_2 = 3,$$

$$\text{Let } C_1 = 36,$$

$$= \underline{2\delta} > \delta$$

Therefore, by Chernoff bound and Chebyshev bounds used before

$$Pr(|Z - \mu| \geq \epsilon Q) = P(|S - E(S)| \geq \delta E(S)) \leq 2e^{-\frac{\delta^2}{3} E(S)} = \delta$$

$$\Rightarrow Pr(|Z - \mu| \geq \epsilon Q) \leq \delta$$

### 3. Variance of Morris Counter

#### Properties of Morris algorithm

Prove  $\text{Var}(Z) = \frac{m(m-1)}{2}$

1. The expectation of the variable  $Z=2^{X_m}$  satisfies the following:

$$E[Z] = m+1$$

**Corollary:** Morris algorithm outputs an unbiased estimator of  $m$ .

1. The variance of  $Z$  is equal to  $\text{Var}[Z] = m(m-1)/2$

**Observation:** No improvement in terms of concentration as  $m$  grows since  $\text{Var}(Z)/E(Z)^2$  is constant.

$$\begin{aligned}\text{Var}(Z) &= E(Z^2) - (E(Z))^2 \\ &= E(2^{2X_m}) - (m+1)^2\end{aligned}$$

To find  $E(2^{2X_m})$ , use definition of Expectation:

$$E(2^{2X_m}) = \sum_{i=1}^{\infty} 2^{2i} P(X_m = i)$$

By Morris algorithm,

$$P(X_m = i) = \frac{1}{2^{i-1}} P(X_{m-1} = i-1) + (1 - \frac{1}{2^{i-1}}) P(X_{m-1} = i)$$

$$\begin{aligned}\therefore E(2^{2X_m}) &= \sum_{i=1}^{\infty} 2^{2i} P(X_{m-1} = i-1) + \sum_{i=1}^{\infty} 2^{2i} P(X_{m-1} = i) - \sum_{i=1}^{\infty} 2^i P(X_{m-1} = i) \\ &= 2^2 E(2^{X_{m-1}}) + E(2^{2X_{m-1}}) - E(2^{X_{m-1}}) \\ &= 3 E(2^{X_{m-1}}) + E(2^{2X_{m-1}}) \\ &= 3(m-1+1) + E(2^{2X_{m-1}}) \\ &= 3m + E(2^{2X_{m-1}})\end{aligned}$$

$$\Rightarrow E(2^{2X_m}) = 3m + E(2^{2X_{m-1}})$$

$$\text{for } m=0, E(2^{2X_0}) = 1$$

$$m=1, E(2^{2X_1}) = 3 + E(2^{2X_0})$$

$$m=2, E(2^{2X_2}) = 6 + 3 + E(2^{2X_0})$$

$\vdots$

$$\Rightarrow E(2^{2X_m}) = 1 + \sum_{i=1}^m 3i = 1 + \frac{3}{2} m(m+1)$$

Therefore, plug in the value,

$$\begin{aligned}\text{Var}(Z) &= E(2^{2X_m}) - (m+1)^2 \\ &= 1 + \frac{3}{2} m(m+1) - (m+1)^2\end{aligned}$$

$$\begin{aligned}
&= x + \frac{3}{2}m^2 + \frac{3}{2}m - m^2 - 2m - x \\
&= \frac{1}{2}m^2 - \frac{1}{2}m \\
&= \frac{m(m-1)}{2}
\end{aligned}$$

4. Uniform RVs.

Let  $V_k$  be the  $k$ -th smallest hashed value

$$\begin{aligned}
P(V_k \leq x) &= \Pr[\text{at least } k \text{ samples in } [0, x]] \\
&= \sum_{l=k}^n \binom{n}{l} \cdot x^l (1-x)^{n-l} = \Delta
\end{aligned}$$

$$\begin{aligned}
(a) \quad \frac{d\Delta}{dx} &= \sum_{l=k}^n \binom{n}{l} \cdot (l \cdot x^{l-1} \cdot (1-x)^{n-l} - (n-l)(1-x)^{n-l-1} \cdot x^l) \\
&= \sum_{l=k}^n \binom{n}{l} \cdot l \cdot x^{l-1} (1-x)^{n-l} - \sum_{l=k}^{n-1} \binom{n}{l} (n-l)(1-x)^{n-l-1} x^l \\
&= \sum_{l=k}^n n \cdot \binom{n-1}{l-1} x^{l-1} (1-x)^{n-l} - \sum_{l=k}^{n-1} n \cdot \binom{n-1}{l} (1-x)^{n-l-1} x^l \\
&= \underbrace{n \cdot \binom{n-1}{k-1} \cdot x^{k-1} \cdot (1-x)^{(n-1)-(k-1)}}_{\text{pdf of beta distribution.}} \\
&= \frac{n!}{(k-1)!(n-k)!} \cdot x^{k-1} \cdot (1-x)^{(n-1)-(k-1)}
\end{aligned}$$

(b) Therefore,  $E_{V_k \sim \text{Beta}(\alpha, \beta)}[V_k] = \frac{\alpha}{\alpha + \beta}$ , where  $\alpha = k$ ,  $\beta = n - k + 1$

$$E(V_k) = \frac{k}{k + n - k + 1} = \frac{k}{n+1}$$