

# Foundations of data science

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## 1. Exercise sheet Handin as announced on eCampus

**Exercise 1.1** (Computer scientist's random variables). fits to 01 8 points)

Consider the following algorithm:

1. Throw a coin  $C \stackrel{\text{f}}{\leftarrow} \{0, 1\}$ .
2. Choose  $X \stackrel{\text{f}}{\leftarrow} [0, 1]$  uniformly.
3. Roll a die  $D \stackrel{\text{f}}{\leftarrow} \{1, 2, 3, 4, 5, 6\}$ .
4. If  $C = 0$  then
5.     Let  $Z \leftarrow D + X$ .
6. Else
7.     Let  $Z \leftarrow D - X$ .
8. Return  $Z$

The output  $Z$  of this algorithm is a random variable.

- (i) Compute its expectation  $E(Z)$ .
- (ii) Compute its second moment  $E(Z^2)$ .
- (iii) Compute its variance  $\text{var}(Z)$ .

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**Solution.** Notice

- $E(X) = \int_0^1 x \, dx = \frac{1}{2}$ ,
- $E(X^2) = \int_0^1 x^2 \, dx = \frac{1}{3}$  and
- $E(D) = \sum_{d \in \{1, 2, 3, 4, 5, 6\}} d \, \text{prob}(D = d) = \frac{21}{6} = \frac{7}{2}$ ,
- $E(D^2) = \sum_{d \in \{1, 2, 3, 4, 5, 6\}} x^2 \, \text{prob}(D = d) = \frac{1+2^2+3^2+4^2+5^2+6^2}{6} = \frac{91}{6}$ .

Thus conditioned on  $C = 0$  we find

$$E(Z \mid C = 0) = E(D + X) = E(D) + E(X) = 4$$

and conditioned on  $C = 1$  we find

$$E(Z \mid C = 0) = E(D - X) = E(D) - E(X) = 3.$$

**Interludium total probability:** We can always split probabilities into pieces via conditional probabilities on the pieces of a partition of the universe: If  $\Omega = \bigsqcup_i \Omega_i$  then

$$\text{prob}(V \in \mathcal{A}) = \sum_i \text{prob}(V \in \mathcal{A} \mid \Omega_i) \text{prob}(\Omega_i)$$

This translates to expected values:

$$E(V) = \sum_i E(V \mid \Omega_i) \text{prob}(\Omega_i).$$

You can easily check this by using the definitions. This is *the* tool to deal with case distinctions in programs in general.

**(i)** In the case at hand the interludium means that we simply compute

$$E(Z) = \frac{1}{2} E(Z \mid C = 0) + \frac{1}{2} E(Z \mid C = 1).$$

and so

$$E(Z) = E(D) = \frac{7}{2}.$$

**(ii)** Similarly we get

$$\begin{aligned} E(Z^2) &= E((D + X)^2) \text{prob}(C = 0) + E((D - X)^2) \text{prob}(C = 1) \\ &= \frac{1}{2} (E(D^2) + 2E(D)E(X) + E(X^2)) \\ &\quad + \frac{1}{2} (E(D^2) - 2E(D)E(X) + E(X^2)) \\ &= E(D^2) + E(X^2) = \frac{91}{6} + \frac{1}{3} = \frac{31}{2}. \end{aligned}$$

**(iii)** Consequently,  $\text{var}(Z) = E(Z^2) - E(Z)^2 = \frac{31}{2} - \left(\frac{7}{2}\right)^2 = \frac{13}{4}$ . ○

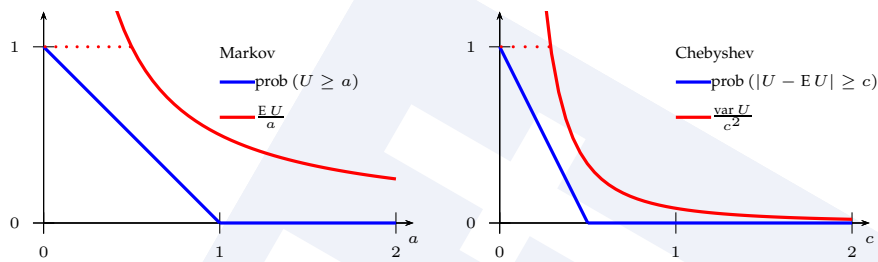
**Exercise 1.2** (Apply Markov's and Chebyshev's inequalities). fits to 02 **(12)** points

In this exercise you shall test the quality of Markov's and Chebyshev's inequalities.

- (i) Consider a uniform real number from the unit interval, ie.  $U \overset{\text{def}}{\sim} [0, 1]$  uniformly chosen.

- 1 (a) Plot  $\text{prob}(U \geq a)$  and  $\frac{E U}{a}$  as functions of  $a$ .
- 1 (b) Plot  $\text{prob}(|U - E U| \geq c)$  and  $\frac{\text{var } U}{c^2}$  as functions of  $c$ .
- 2 (c) Compare and interpret.

**Solution.**

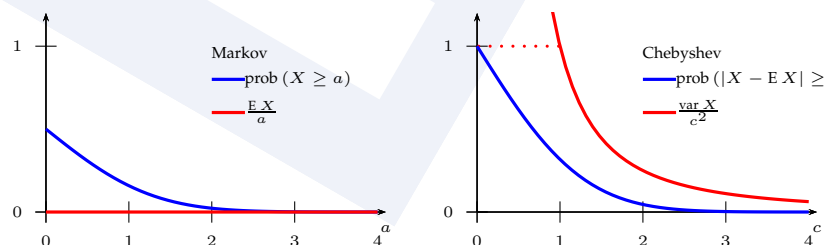


We can see that the Markov bound for the variable  $X$  is very loose and does not contain much information. For example, in the interval  $[0, \frac{1}{2}]$  the values of  $E(U)$  are larger than 1, even if we know that all probabilities are naturally bounded by 1. The tail bound is also extremely loose, as even if  $X > 1$  is impossible, we still see a significant value of the Markov bound there.

The Chebyshev inequality is also quite meaningless for small values of  $c$ . Even for  $c$  close to 0.5, which is the end of the range that  $c$  can take, we still have a lot of difference between the bound and the actual function. The tail bounds here are much better than in the Markov case and tend to 0 very quickly.  $\circ$

- (ii) Consider a normally distributed real number with mean 0 and variance 1, ie.  $X \stackrel{\text{def}}{\sim} \mathcal{N}(0, 1)$ .
  - 1 (a) Plot  $\text{prob}(X \geq a)$  and  $\frac{E X}{a}$  as functions of  $a$ .
  - 1 (b) Plot  $\text{prob}(|X - E X| \geq c)$  and  $\frac{\text{var } X}{c^2}$  as functions of  $c$ .
  - 2 (c) Compare and interpret.

**Solution.**



The most important thing to note is that Markov's inequality does not apply here. This is due to the fact, that  $X$  is a random variable that can take negative values. Markov is only applicable for non-negative random variables.

Chebychev's inequality yields a decent bound, except for very small  $c$ . ○

(iii) Consider a fair die, ie.  $D \stackrel{\text{def}}{\sim} \{1, 2, 3, 4, 5, 6\}$  with uniform distribution.

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(a) Plot  $\text{prob}(D \geq a)$  and  $\frac{E D}{a}$  as functions of  $a$ .

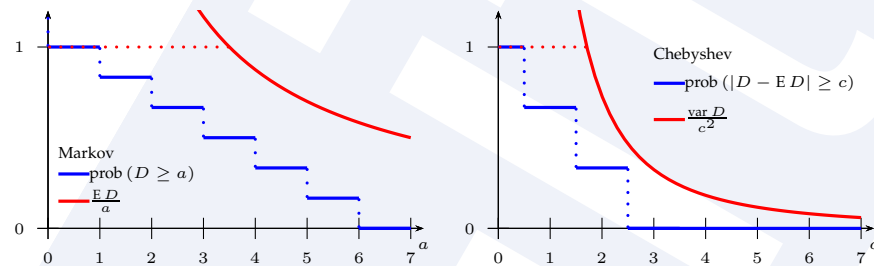
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(b) Plot  $\text{prob}(|D - E D| \geq c)$  and  $\frac{\text{var } D}{c^2}$  as functions of  $c$ .

2

(c) Compare and interpret.

**Solution.**



We see again that the Markov inequality is a quite rough bound. The tail bound of Chebychev is better, but it's still not very good. ○

**Exercise 1.3** (When is Markov sharp?).

(0+4 points)

+4

Show that for any  $a > 0$  there exists a probability distribution such that the Markov inequality is sharp, ie.  $\text{prob}(X \geq a) = \frac{E(X)}{a}$ . Use the  $\text{prob}(\dots)$ -notation to write down the distribution explicitly.

*Hint:* Recall the proof from the lecture. If it helps you, restrict to the discrete setting.

**Solution.** We use that  $\text{prob}(X \in I) = E(\mathbb{1}_{\{X \in I\}})$  since

$$\begin{aligned} E(\mathbb{1}_{\{X \in I\}}) &= \sum_{i=0}^1 i \cdot \text{prob}(\mathbb{1}_{\{X \in I\}} = i) \\ &= 1 \cdot \text{prob}(\mathbb{1}_{\{X \in I\}} = 1) \\ &= \text{prob}(X \in I). \end{aligned}$$

Applied to the equation, we find that the following are equivalent:

$$\begin{aligned}
 \text{prob}(X \geq a) &= \frac{E(X)}{a}. \\
 a \cdot \text{prob}(X \geq a) &= E(X). \\
 E(a \cdot \mathbb{1}_{\{X \geq a\}}) &= E(X \cdot \mathbb{1}_{\{X \geq 0\}}). \\
 E(X \cdot \mathbb{1}_{\{X \geq 0\}} - a \cdot \mathbb{1}_{\{X \geq a\}}) &= 0. \\
 E(X \cdot \mathbb{1}_{\{0 \leq X < a\}} + (X - a) \cdot \mathbb{1}_{\{X \geq a\}}) &= 0. \\
 E(X \cdot \mathbb{1}_{\{0 \leq X < a\}}) = 0 \quad \wedge \quad E((X - a) \cdot \mathbb{1}_{\{X \geq a\}}) &= 0. \\
 \text{prob}(0 < X < a) = 0 \quad \wedge \quad \text{prob}(0 < X - a) &= 0.
 \end{aligned}$$

For the last step we use: for any set  $M$  and any random variable  $Z$  with  $Z \geq 0$  on  $M$  we have that  $E(Z \cdot \mathbb{1}_M) = 0$  implies  $\text{prob}(Z \in M \setminus \{0\}) = 0$ . As a consequence Markov is sharp iff  $\text{prob}(X \in \{0, a\}) = 1$ .

In other words, each distribution that satisfies above requirement is given by some  $p \in [0, 1]$  and

$$\text{prob}(X = a) = p, \quad \text{prob}(X = 0) = 1 - p. \quad \bigcirc$$