# Supplementary Material for EMORF-II

This document contains the supplementary material for the manuscript titled "EMORF-II: ADAP-TIVE EM-BASED OUTLIER-ROBUST FILTERING WITH CORRELATED MEASUREMENT NOISE" submitted to the "International Workshop on Machine Learning and Signal Processing 2025". It provides notations used in the paper, details of derivations of  $\hat{\mathcal{I}}_k^i$  and  $\hat{b}_k$  in the paper and discussion of some more numerical evaluations.

#### Notation

As a general notation in this work,  $\mathbf{r}^{\top}$  is the transpose of the vector  $\mathbf{r}$ ,  $r^i$  denotes the *i*-th element of a vector  $\mathbf{r}$ ;  $\mathbf{r}^{-i}$  is the vector  $\mathbf{r}$  with its *i*-th element removed; the subscript k is used for the time index;  $\mathbf{r}_k$  is the vector  $\mathbf{r}$  at time instant k;  $\mathbf{r}_k^-$  is the sequence of vectors  $\mathbf{r}$  over all time steps except k;  $R^{i,j}$  is the element of the matrix  $\mathbf{R}$  in the *i*-th row and *j*-th column;  $\mathbf{R}^{-1}$  is the inverse of  $\mathbf{R}$ ;  $|\mathbf{R}|$  is the determinant of  $\mathbf{R}$ ;  $\delta(\cdot)$  represents the delta function;  $\langle \cdot \rangle_{q(\psi_k)}$  denotes expectation with respect to the distribution  $q(\psi_k)$ ;  $\mathrm{tr}(\cdot)$  is the trace operator;  $a \bmod b$  denotes the remainder of a/b; the superscripts - and + mark predicted and updated filtering parameters, respectively. Furthermore, for any matrix  $\mathbf{R}$  we obtain the sub-matrix  $\mathbf{R}^{-i,-i}$  by removing its i-th column and row.  $\mathbf{R}_k(\mathcal{I}_k^i, \hat{\mathcal{I}}_k^{i-})$  denotes  $\mathbf{R}_k(\mathcal{I}_k)$  evaluated at  $\mathcal{I}_k$  with its i-th element as  $\mathcal{I}_k^i$  and remaining entries  $\hat{\mathcal{I}}_k^{i-}$ .

## Details of estimating $\hat{\mathcal{I}}_k^i$

Invoking the EM algorithm, the estimates  $\hat{\mathcal{I}}_k^i$  can be obtained with the following maximization

$$\hat{\mathcal{I}}_{k}^{i} = \underset{\mathcal{I}^{i}}{\operatorname{argmax}} \left\langle \ln(p(\mathbf{x}_{k}, \mathcal{I}_{k}^{i}, \hat{\boldsymbol{\mathcal{I}}}_{k}^{i-}, \hat{b}_{k} | \mathbf{y}_{1:k}) \right\rangle_{q(\mathbf{x}_{k})}$$

This leads to Equation (21) in the manuscript given as

$$\hat{\mathcal{I}}_{k}^{i} = \underset{\mathcal{I}_{k}^{i}}{\operatorname{argmax}} \left\{ \underbrace{-\frac{1}{2} \operatorname{tr} \left( \mathbf{W}_{k} \mathbf{R}_{k}^{-1} (\mathcal{I}_{k}^{i}, \hat{\mathcal{I}}_{k}^{i-}) \right) - \frac{1}{2} \ln \left| \mathbf{R}_{k} (\mathcal{I}_{k}^{i}, \hat{\mathcal{I}}_{k}^{i-}) \right| + \ln \left( (1 - \theta_{k}) f(a_{k}, \hat{b}_{k}) (\mathcal{I}_{k}^{i})^{a_{k} - 1} e^{-\hat{b}_{k} \mathcal{I}_{k}^{i}} + \theta_{k} \delta(\mathcal{I}_{k}^{i} - 1) \right) + k_{1}^{i}} \right\} \\
= \underset{\ln q(\mathcal{I}_{k}^{i})}{\operatorname{tr}} \left\{ \underbrace{-\frac{1}{2} \operatorname{tr} \left( \mathbf{W}_{k} \mathbf{R}_{k}^{-1} (\mathcal{I}_{k}^{i}, \hat{\mathcal{I}}_{k}^{i-}) \right) - \frac{1}{2} \ln \left| \mathbf{R}_{k} (\mathcal{I}_{k}^{i}, \hat{\mathcal{I}}_{k}^{i-}) \right| + \ln \left( (1 - \theta_{k}) f(a_{k}, \hat{b}_{k}) (\mathcal{I}_{k}^{i})^{a_{k} - 1} e^{-\hat{b}_{k} \mathcal{I}_{k}^{i}} + \theta_{k} \delta(\mathcal{I}_{k}^{i} - 1) \right) + k_{1}^{i} \right\}$$
(21)

where the term to be maximized is supposed as  $\ln q(\mathcal{I}_k^i)$ . We perform the following manipulations to obtain Equations (23)-(25) in the manuscript. Manipulating  $\ln q(\mathcal{I}_k^i)$  we can get

$$\ln q\left(\mathcal{I}_{k}^{i}\right) = -\frac{1}{2}\ln\left|\mathbf{R}_{k}\left(\mathcal{I}_{k}^{i},\hat{\boldsymbol{\mathcal{I}}}_{k}^{-i}\right)\right| - \frac{1}{2}\operatorname{tr}\left(\mathbf{W}_{k}\,\mathbf{R}_{k}^{-1}\left(\mathcal{I}_{k}^{i},\hat{\boldsymbol{\mathcal{I}}}_{k}^{-i}\right)\right) + \ln\left(\left(1-\theta_{k}^{i}\right)f\left(a_{k},\hat{b}_{k}\right)\left(\mathcal{I}_{k}^{i}\right)^{a_{k}-1}e^{-\hat{b}_{k}\,\mathcal{I}_{k}^{i}} + \theta_{k}^{i}\,\delta\left(\mathcal{I}_{k}^{i}-1\right)\right) + k_{1}^{i}$$
which leads to

$$q(\mathcal{I}_{k}^{i}) = k_{2}^{i} \exp(-\frac{1}{2} \ln |\mathbf{R}_{k}(\mathcal{I}_{k}^{i}, \hat{\mathcal{I}}_{k}^{-i})| - \frac{1}{2} \operatorname{tr}(\mathbf{W}_{k} \mathbf{R}_{k}^{-1}(\mathcal{I}_{k}^{i}, \hat{\mathcal{I}}_{k}^{-i}))) \times (1 - \theta_{k}^{i}) f(a_{k}, \hat{b}_{k}) (\mathcal{I}_{k}^{i})^{a_{k} - 1} e^{-\hat{b}_{k} \mathcal{I}_{k}^{i}} + k_{2}^{i} \exp(-\frac{1}{2} \ln |\mathbf{R}_{k}(\mathcal{I}_{k}^{i} = 1, \hat{\mathcal{I}}_{k}^{-i})| - \frac{1}{2} \operatorname{tr}(\mathbf{W}_{k} \mathbf{R}_{k}^{-1}(\mathcal{I}_{k}^{i} = 1, \hat{\mathcal{I}}_{k}^{-i}))) \theta_{k}^{i} \delta(\mathcal{I}_{k}^{i} - 1)$$

where  $k_2^i$  is another proportionality constant. Further manipulation of  $q(\mathcal{I}_k^i)$  yields

$$q(\mathcal{I}_{k}^{i}) = k_{2}^{i}(R_{k}^{ii}/\mathcal{I}_{k}^{i})^{-\frac{1}{2}} \left| \hat{\mathbf{R}}_{k}^{-i,-i} \right|^{-\frac{1}{2}} \exp\left(-\frac{1}{2} W_{k}^{ii} \frac{\mathcal{I}_{k}^{i}}{R_{k}^{ii}} - \frac{1}{2} \operatorname{tr}\left(\mathbf{W}_{k}^{-i,-i} (\hat{\mathbf{R}}_{k}^{-i,-i})^{-1}\right)\right) \times (1 - \theta_{k}^{i}) f(a_{k}, \hat{b}_{k}) (\mathcal{I}_{k}^{i})^{a_{k}-1} e^{-\hat{b}_{k}\mathcal{I}_{k}^{i}} + k_{2}^{i} \exp\left(-\frac{1}{2} \ln\left|\mathbf{R}_{k}(\mathcal{I}_{k}^{i} = 1, \hat{\mathcal{I}}_{k}^{-i})\right| - \frac{1}{2} \operatorname{tr}\left(\mathbf{W}_{k} \mathbf{R}_{k}^{-1}(\mathcal{I}_{k}^{i} = 1, \hat{\mathcal{I}}_{k}^{-i})\right)\right) \theta_{k}^{i} \delta(\mathcal{I}_{k}^{i} - 1)$$

Moving the relevant terms in the exponents we get

$$q(\mathcal{I}_{k}^{i}) = k_{2}^{i}(R_{k}^{ii})^{-\frac{1}{2}} \left| \hat{\mathbf{R}}_{k}^{-i,-i} \right|^{-\frac{1}{2}} \exp\left(-\frac{1}{2} \operatorname{tr}\left(\mathbf{W}_{k}^{-i,-i}(\hat{\mathbf{R}}_{k}^{-i,-i})^{-1}\right)\right) \times (1 - \theta_{k}^{i}) f(a_{k}, \hat{b}_{k}) (\mathcal{I}_{k}^{i}) \underbrace{(a_{k} + \frac{1}{2})}_{-1} e^{-\underbrace{(\hat{b}_{k} + \frac{1}{2} \frac{W_{k}^{ii}}{R_{k}^{2i}})}_{+ k_{2}^{i} \exp\left(-\frac{1}{2} \ln\left|\mathbf{R}_{k}(\mathcal{I}_{k}^{i} = 1, \hat{\mathcal{I}}_{k}^{-i})\right| - \frac{1}{2} \operatorname{tr}\left(\mathbf{W}_{k} \mathbf{R}_{k}^{-1}(\mathcal{I}_{k}^{i} = 1, \hat{\mathcal{I}}_{k}^{-i})\right)\right) \theta_{k}^{i} \delta(\mathcal{I}_{k}^{i} - 1)}$$

 $\alpha_k$  and  $\beta_k^i$  can easily be identified as follows given as Equations (26)-(27) in the manuscript

$$\alpha_k = a_k + 0.5 \tag{26}$$

$$\beta_k^i = \hat{b}_k + 0.5 \frac{W_k^{ii}}{R_k^{ii}} \tag{27}$$

Through further manipulation we can write  $q(\mathcal{I}_k^i)$  as

$$q(\mathcal{I}_{k}^{i}) = k_{2}^{i} \underbrace{\left(R_{k}^{ii}\right)^{-\frac{1}{2}} \left|\hat{\mathbf{R}}_{k}^{-i,-i}\right|^{-\frac{1}{2}} \exp\left(-\frac{1}{2}\operatorname{tr}\left(\mathbf{W}_{k}^{-i,-i}(\hat{\mathbf{R}}_{k}^{-i,-i})^{-1}\right)\right) (1-\theta_{k}^{i}) \frac{\Gamma(\alpha_{k})\hat{b}_{k}^{i}^{\alpha_{k}}}{\Gamma(a_{k})\beta_{k}^{i}^{\alpha_{k}}}}_{G_{k}^{i}} \underbrace{f(\alpha_{k},\beta_{k}^{i}) \left(\mathcal{I}_{k}^{i}\right)^{\alpha_{k}-1} e^{-\beta_{k}^{i}\mathcal{I}_{k}^{i}}}_{G_{k}^{i}} + k_{2}^{i} \underbrace{\exp\left(-\frac{1}{2}\ln\left|\mathbf{R}_{k}(\mathcal{I}_{k}^{i}=1,\hat{\mathcal{I}}_{k}^{-i})\right| - \frac{1}{2}\operatorname{tr}\left(\mathbf{W}_{k}\mathbf{R}_{k}^{-1}(\mathcal{I}_{k}^{i}=1,\hat{\mathcal{I}}_{k}^{-i})\right)\right) \theta_{k}^{i}}_{H^{i}} \delta(\mathcal{I}_{k}^{i}-1)$$

Since maximizing logarithm of a positive function is equivalent to maximizing the function, the resulting decision criterion (from Equation 21) depends on which part -  $G_k^i$  or  $H_k^i$  - of the multi-modal function  $q(\mathcal{I}_k^i)$  has higher weight. If  $H_k^i/G_k^i \geq 1$  we declare  $\hat{\mathcal{I}}_k^i = 1$ . Otherwise, if  $H_k^i/G_k^i < 1$ , we assign  $\hat{\mathcal{I}}_k^i$  as the mode of  $Gamma(\alpha_k, \beta_k^i)$ . The criterion is summarized in Equation (23) in the manuscript reproduced as

$$\mathcal{I}_{k}^{i} = \begin{cases} 1 & \text{if } H_{k}^{i}/G_{k}^{i} \ge 1\\ (\alpha_{k} - 1)/\beta_{k}^{i} & \text{if } H_{k}^{i}/G_{k}^{i} < 1 \end{cases}$$
(23)

## Details of estimating $\hat{b}_k$

Employing the EM algorithm the estimate for  $b_k$  can be calculated through the following maximization

$$\hat{b}_k = \operatorname*{argmax}_{b_k^i} \left\langle \ln(p(\mathbf{x}_k, \hat{\mathcal{I}}_k, b_k | \mathbf{y}_{1:k})) \right\rangle_{q(\mathbf{x}_k)}$$

which leads to

$$\hat{b}_k = \underset{b_k}{\operatorname{argmax}} \sum_{i=1}^m \ln \left[ (1 - \theta_k^i) f(a_k, b_k) (\hat{\mathcal{I}}_k^i)^{a_k - 1} e^{-b_k \hat{\mathcal{I}}_k^i} \right. \\ \left. + \theta_k^i \delta(\hat{\mathcal{I}}_k^i - 1) \right] + \ln(f(A_k, B_k) b_k^{A_k - 1} e^{-B_k b_k})$$

Since  $\hat{\mathcal{I}}_k^i = 1$  do not affect the estimates  $\hat{b}_k$ , we can further write

$$\hat{b}_k = \underset{b_k}{\operatorname{argmax}} \sum_{\{i: \hat{\mathcal{I}}_k^i \neq 1\}} \ln \left[ (1 - \theta_k^i) f(a_k, b_k) (\hat{\mathcal{I}}_k^i)^{a_k - 1} e^{-b_k \hat{\mathcal{I}}_k^i} \right] + \ln(f(A_k, B_k) b_k^{A_k - 1} e^{-B_k b_k})$$

which can be simplified as

$$\hat{b}_k = \operatorname*{argmax}_{b_k} \sum_{\{i: \hat{\mathcal{I}}_i^i \neq 1\}} (a_k \ln b_k - b_k \mathcal{I}_k^i) - B_k b_k + (A_k - 1) \ln b_k$$

Maximizing the above equation through differentiation leads to the estimate  $\hat{b}_k$  given as

$$\hat{b}_k = \frac{M_k a_k + A_k - 1}{B_k + \sum_{\{i: \hat{\mathcal{I}}_k^i \neq 1\}} \hat{\mathcal{I}}_k^i}$$

where  $M_k = \#\{i : \hat{\mathcal{I}}_k^i \neq 1\}$  i.e. the count of  $\mathcal{I}_k$  elements not equal to one. Further defining  $\overline{A}_k = M_k a_k + A_k$  and  $\overline{B}_k = B_k + \sum_{\{i : \mathcal{I}_k^i \neq 1\}} \hat{\mathcal{I}}_k^i$  results in Equation (28) of the manuscript

$$\hat{b}_k = \frac{\overline{A}_k - 1}{\overline{B}_k} \tag{28}$$

### Other numerical experiments

We further extend our numerical experiments by introducing outliers generated from Uniform and Laplace distributions, replacing the previously considered Gaussian-based outliers. Across a range of parameter variations, the comparative performance trends remained similar to those presented in the main manuscript. Further evaluation of EMORF-II was conducted by testing alternative prior parameters  $\theta_k^i$ , including values sampled from a Uniform distribution over the range [0.05, 0.95]. The results show that EMORF-II maintains robust performance across variations in  $\theta_k^i$ . Nevertheless, setting  $\theta_k^i = 0.5$  is recommended as a default choice in the absence of specific prior knowledge regarding outlier rates.