# Portfolio 1

### Experimental Methods 3: Multilevel models and machine learning

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Parts of this assignment have been made in a group consisting of Freddy Wulf (FW), Ida Møller (IM), Maria Mujemula (MM), Sabrina Zaki (SZ) & Sara Kjær Kristensen (SK).

E.g. Introduction (MM, SK)

Maria wrote the main part and Sara adapted it.

The main parts of the assignment have been discussed in the mentioned group, and written by SK. The group discussions will be addressed as 'we' or 'the group' hereafter.

Character count with spaces w/o plots: 9.050

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#### Introduction

This portfolio will cover handed out data<sup>1</sup> investigating autistic and neurotypical children's language development over a series of 6 visits with several months apart. At each visit, different measures were taken including the binary diagnostic grouping, either autistic or neurotypical, and the child's mean length of utterance during that visit. To investigate the dataset R (version 4.2.0 & 4.2.1) with RStudio (version 2022.07.1) and statistical packages for Bayesian workflow BRMS (Bürkner, 2021).

#### Simulating the data at hand

To better understand the data and the models, that are going to describe it, a new dataset is simulated using the literature estimates as a guideline. The dataset is structured with the most important variables and leave the others for consideration; number of visits (6); number of participants in the two diagnostic conditions (50, sample size = 100); and mean length of utterance (MLU).

To estimate the priors, we used the values given in the assignment as a starting point for the MLU for each diagnostic group at visit 1, average individual variability in this initial MLU, the average change in MLU across visits for the diagnostic groups, and average individual variability for the diagnostic groups. To modify the priors, we make histograms to check if the values give a realistic impression of the data, by trial and error we estimate the priors to be informed priors as they assume a difference in the development of MLU. Generally, we use a lognormal distribution to exclude the possibility of the MLU going below zero.

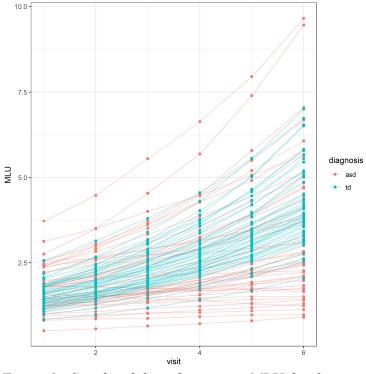


Figure 1 - Simulated data showcasing MLU development over time

<sup>&</sup>lt;sup>1</sup> From Fusaroli, et al (2019)

To assess the simulated data and what predictors are best to describe the change in MLU, we formulate 3 models:

Model	Equation	Description			
no.					
1	$MLU \sim 0 + diagnosis$	MLU modulated by diagnosis			
2	$MLU \sim 0 + diagnosis$	MLU modulated by diagnosis with a difference in			
	+ diagnosis: visit	change per visit pending on diagnosis			
3	$MLU \sim 0 + diagnosis$	MLU modulated by diagnosis with a difference in			
	+ diagnosis: visit	change per visit pending on diagnosis and individ-			
	+(1+visit ID)	ual differences per visit			

The priors are specified as needed for the different models and fitted to their model:

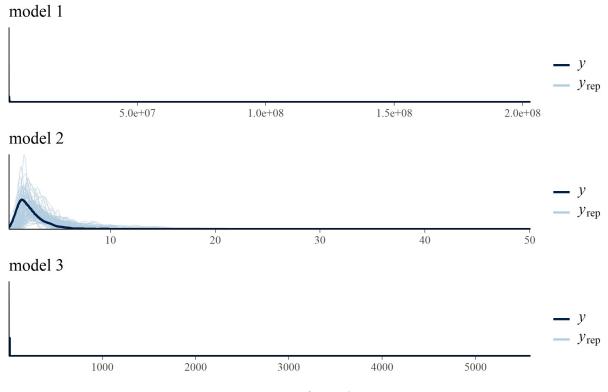


Figure 2 - Informed priors

Fitting the data to the posterior distributions gives these prior posterior update checks. This way the data's influence can be inspected.

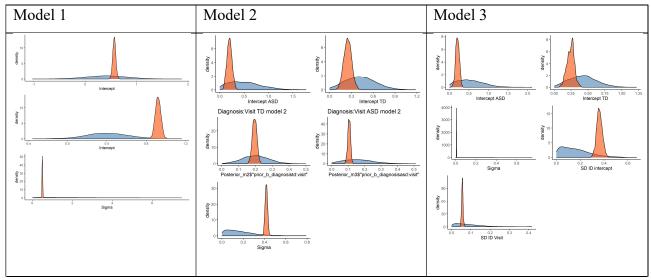


Table 1 - blue for prior and orange for posterior

Checking the quality of the models is assessed by looking for a ratio spread for the Markov Chain Monte Carlo (Rhat) value between 0.90 and 1 and as high values as possible for effective sample sizes for both the bulk (Bulk\_ESS) and tail (Tail\_ESS).

Model	Summary	Esti-	Est.Er-	1-95%	u-95%	Rhat	Bulk ESS	Tail ESS	
no.	output	mate	ror	CI	CI		_	_	
	Population-Level Effect								
	diagno-	0.56	0.03	0.51	0.62	1.00	2088	1534	
	sisasd								
1	diagno-	0.92	0.03	0.8	0.98	1.00	1803	1845	
	sistd								
	Family Spec	cific Parar	neters						
	Sigma	0.50	0.01	0.47	0.53	1.00	1839	1667	
2	Population-Level Effect								
	diagno-	0.20	0.05	0.10	0.30	1.00	1320	1262	
	sisasd								
	diagno-	0.25	0.05	0.14	0.35	1.00	1207	1205	
	sistd								
	diagno-	0.11	0.01	0.08	0.13	1.00	1288	1887	
	sisasd:visit								
	diagno-	0.20	0.01	0.17	0.22	1.00	1552	1374	
	sistd:visit								
	Family Specific Parameters								
	Sigma	0.42	0.01	0.40	0.44	1.00	2009	2251	
3	Population-Level Effects								

diagno- sisasd	0.20	0.05	0.11	0.29	1.01	103	239
diagno- sistd	0.24	0.05	0.14	0.35	1.05	44	78
diagno- sisasd:visit	0.11	0.01	0.09	0.12	1.01	117	132
diagno- sistd:visit	0.20	0.01	0.18	0.21	1.13	16	152
Group-Level Effects (~ID)							
sd(Inter- cept)	0.36	0.02	0.31	0.41	1.02	105	167
sd(visit)	0.06	0.00	0.05	0.06	1.02	144	295
cor(Inter- cept:visit)	0.03	0.09	-0.14	0.23	1.02	114	179
Family Specific Parameters							
Sigma	0.01	0.00	0.01	0.01	1.00	754	1206

Table 2 - Comparing model summary output

Model 1's parameters is overall okay, but the estimates are very different from our priors. Model 2's parameters are better both for estimates and mcmc specific parameters. Lastly, model 3's estimates are fine, but the mcmc specific parameters are concerning; the Rhat is above 1, Bulk\_ESS is in many case below a 100 samples same goes for the Tail\_ESS. Keeping these values in mind, divergency plots are consulted:

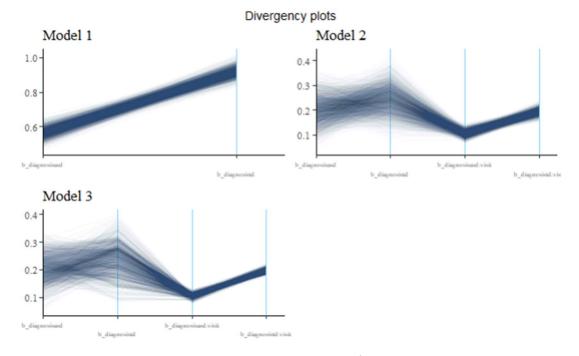


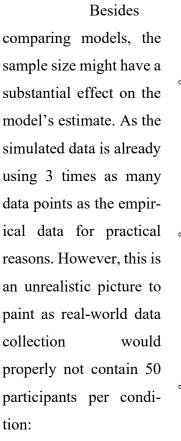
Figure 3 - Divergency plots

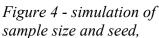
The divergency plots points towards model 2 instead of the other models because of the spread in the plots. However, consulting other seeds and my group gives another view on the models and by doing model comparison with the leave-one-out method, is a bit more in favour of model 3:

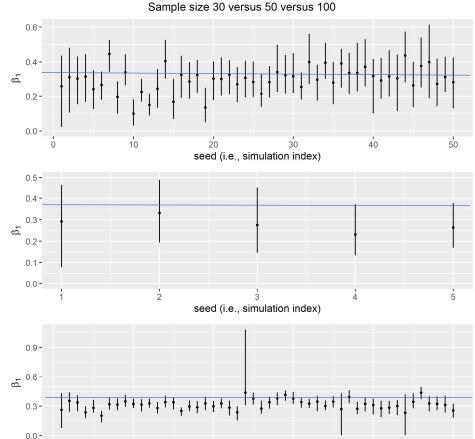
Model no.	elpd_diff	se_diff	weight
1	-2270.2	23.8	0.000
2	-2166.4	26.1	0.000
3	0.0	0.0	1.000

Table 3 - Model comparison with LOO

Expected log pointwise predictive density difference (elpd\_diff) compares models in ascending order, and the negative numbers suggest that the function favours the previous. Logically follows that model 3 is the better one as the leave-one-out log score for predictive density (weight) is maximized by using this model solely. Taking all of this into account and adding the theoretical layer, it makes sense to have children within a certain diagnostic group developing at their own pace. On top of that, the group's general view of model 3 as being the best describing model. Model 3 is chosen for further empirical analysis.







seed (i.e., simulation index)

the blue line resembles estimates for model 3's standard deviation of intercepts (0.36)

40

These simulation plots give a slight insight into how different seeds for random sampling for 30, 50 and 100 participants per group gives different estimates of the different ASD intercepts. Overall, more data points give more certainty, but 30 participants per group seem to be a sufficient place to start. Looking at the power estimates does not give further enlightenment:

Variable	Power estimate
Diagnosis ASD	1
Diagnosis TD	1
Diagnosis ASD : visit	1
Diagnosis TD : visit	1

Table 4 - Power estimates

Besides simulating across seeds, changing the priors to weakly informed prior to make the data work more to make a convincing case of change and difference between visits and diagnostic groups. On that note, it is time to dive into the empirical data.

# Empirical data

Like many other cases, the simulation does not capture all the noise in the data set. Comparing the simulated and empirical data the difference is clear as is the general tendency in language development.

Again, our prior belief is updated to a posterior distribution of the estimates, this time model 3's parameters are updated on the empirical data. Prior-posterior update plots are as follows:

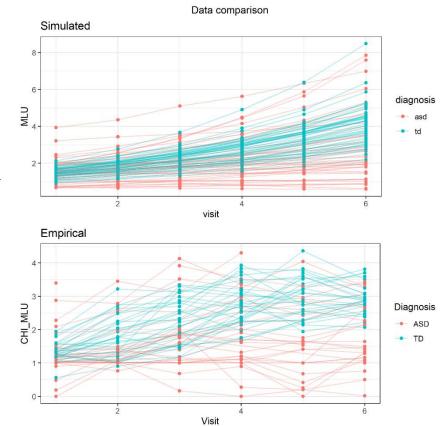


Figure 5 - comparing data's MLU development over time

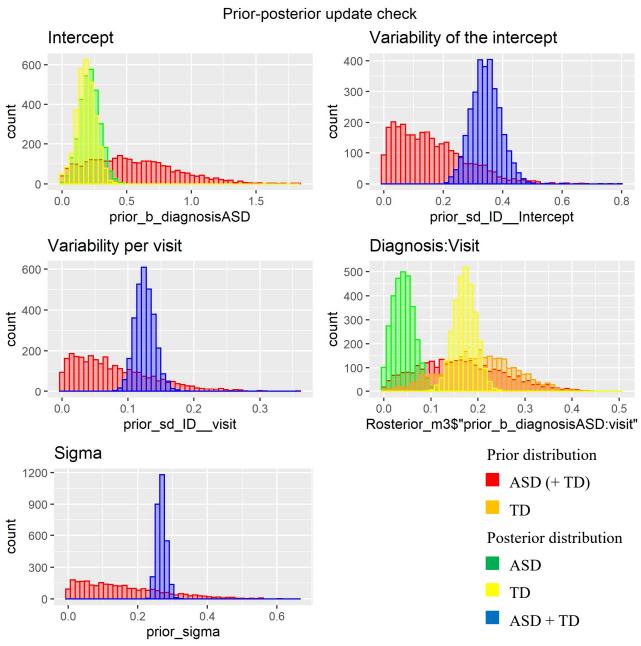


Figure 6 - Prior-Posterior update plots with empirical data

Before we dive into the exciting difference in change in MLU over time, as indicated in plot *Diagnosis:Visit*, we will take a look at the model's qualities. First of all, the model:

$$MLU \sim 0 + diagnosis + diagnosis: visit + (1 + visit|ID)$$

predicts the mean length of utterance (MLU), calculated as the mean of morphemes in the utterance divided by the number of words in the utterance during the visit (30 min), by the binary diagnostic grouping (autism spectrum disorder (ASD) or typically developed (TD)), interaction effect of diagnostic group and visit as participants in the different groups may vary differently in acquired language over time, and lastly with individual random effects for each participant over visits. This way the

participants' MLU can account for just bad days or overall smaller changes in MLU. The model's summary highlights the RHat, Bulk\_ESS and Tail\_ESS as mentioned before.

Group-Level Effects						
Parameters	Rhat	Bulk_ESS	Tail_ESS			
sd(Intercept)	1	1023	1574			
sd(visit)	1.01	344	823			
cor(Intercept,visit)	1.01	285	689			
Population-Level Ef	fects					
diagnosisASD	1.00	897	1074			
diagnosisTD	1.00	781	424			
diagnosisASD:visit	1.00	580	721			
diagnosisTD:visit 1.00 514 804						
Family Specific Parameters						
Sigma	1.00	1590	1836			

Table 5 - Model summary

The values do not give rise to much concern except the standard deviation for visit and correlation between intercept and visit in the group-level effects, as their Rhat-values are above 1. With this concern we consult sensitivity plots for the variability over time for the two diagnostic groups to inspect our prior's impact on the posterior estimates:

Sara Kjær Kristensen Study no. 202105320

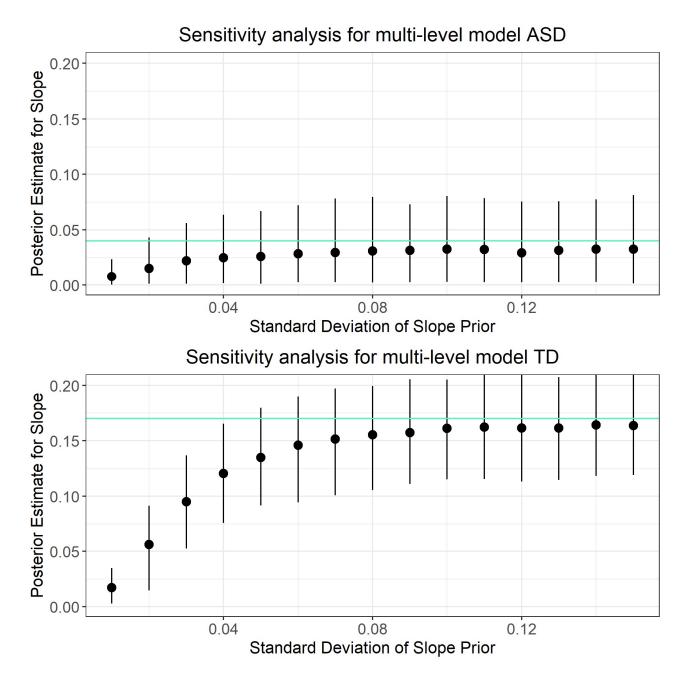


Figure 7 - Sensitivity check of prior-posterior estimates for variability over time for ASD and TD (MU). Lines reflects model posterior estimate diagnosisASD:visit (0.04) and diagnosisTD:visit (0.17) (SK).

The informed prior's standard deviation specified for our model and model summary above was set to 0.1 and 0.08 for ASD and TD respectively. In these plots, the posterior estimate for ASD (0.04) and TD (0.17) is better estimated with optimistic priors. For this analysis, the prior slope standard deviation was set to 0.1 for ASD and 0.08 for TD. Given the prior for slope standard deviation was

set to 0.1 for ASD the model gets a good estimate with a coherent credible interval. As do the prior standard deviation for TD (0.08) gives a decent interval of the posterior estimate of the slope.

Now returning to plot *Diagnosis:Visit* in the prior posterior update checks, the model summary is investigated.

Prior			Posterior					
Estimate	Error	Variable	Estimate	Error	1-95% CI	u-95% CI		
Populatio	n-level					·		
0.41	0.41	diagnosisASD	0.21	0.08	0.07	0.36		
0.41	0.22	diagnosisTD	0.18	0.08	0.03	0.32		
0.15	0.1	diagnosisASD:visit	0.04	0.02	0.00	0.09		
0.2	0.08	diagnosisTD:visit	0.17	0.02	0.13	0.22		
Group lev	Group level							
0	0.2	sd(Intercept)	0.34	0.05	0.25	0.44		
0	0.1	sd(visit)	0.13	0.01	0.10	0.16		
		cor(Intercept,visit)	-0.46	0.14	-0.68	-0.15		

Table 6 - Prior and posterior estimates for population and group level effects

To test the first hypothesis about whether the development across diagnostic groups, the hypothesis() function is used for "diagnosisTD:visit is greater than diagnosisASD:visit". The output gives an estimate of 0.13 with a 95% credible interval of 0.08-0.18, and posterior probability of a 100%. The posterior probability reflects sampling 1 random participant from each group and finding that the TD has a higher MLU is 100 % likely. This result is reflected in the prior-posterior update plots for diagnosis(ASD/TD):visit above and in the conditional effect plot below.

To test the second hypothesis for the starting value across diagnostic groups, same function for "diagnosisASD < diagnosisTD". The output of 0.03 with a 95% credible interval of -0.13-0.20, and posterior probability of 40%. Formulating last the hypothesis as the individual development for ASD is smaller than the development for TD. The output gives an estimate of 0.13 with a 95% credible interval of 0.08-0.18, and posterior probability of a 100% across all children. Which funny enough is the same as for the population-level. However, all in all indicating that the diagnostic slopes are statistically different at the population-level:

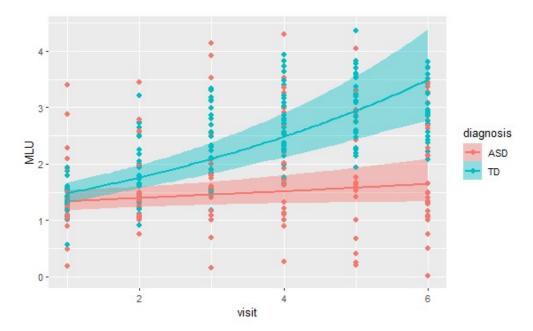


Figure 8 - conditional effects of the development of MLU over time depending on the diagnostic group

Some additional factors would be interesting to add to the model, as the child's own effect on her language can hardly account for her linguistic development. E.g., the mother's MLU would be interesting to consider as well as it we know that children learn from their environment and if the environment or environment controller (in this case the mother) decides what the child is exposed to and therefore decides the syllabus on the upbringings linguistic course. If the child is inhibited from exposure to more diverse linguistic environments, the child could quickly reach a plateau in the linguistic acquirement or make up her own language. The latter part is probably more likely if a 3-year-old could make up her own language and syntax. On the other hand, if the child's environment is stimulating and has a lot to offer, then the development in MLU would reflect more on the child's ability to acquire linguistic concepts rather than it being a lack of resources.

# Reference

Fusaroli, R., Weed, E., Fein, D., & Naigles, L. (2019). Hearing me hearing you: Reciprocal effects between child and parent language in autism and typical development. *Cognition*, *183*, 1–18. <a href="https://doi.org/10.1016/j.cognition.2018.10.022">https://doi.org/10.1016/j.cognition.2018.10.022</a>

#### Materials

RStudio Team (2022). RStudio: Integrated Development for R. RStudio, PBC, Boston, MA URL <a href="http://www.rstudio.com/">http://www.rstudio.com/</a>.

Bürkner P (2021). "Bayesian Item Response Modeling in R with brms and Stan." *Journal of Statistical Software*, **100**(5), 1–54. doi:10.18637/jss.v100.i05.