

Lectures 17

Order Statistics. We end the discussion on transformations of random variables with order statistics. Let X_1, X_2, \dots, X_n be n random variables, the order statistics denoted by $X^{(1)}, X^{(2)}, \dots, X^{(n)}$ of these random variables are defined as follows.

$$X^{(k)} = \min_{J \subseteq [n], \#J=k} \max_{j \in J} X_j,$$

where $[n]$ denote $\{1, 2, \dots, n\}$. Observe that

$$X^{(1)} = \min\{X_1, \dots, X_n\}, \quad X^{(n)} = \max\{X_1, \dots, X_n\}.$$

For the rest of the section, we assume that X_1, \dots, X_n are i.i.d. random variables with distribution function of X_1 denoted by F .

Distribution function of marginals: First we will compute the marginals of $X^{(k)}$. Let $F_{X^{(k)}}$ denote the distribution function of $X^{(k)}$. For $x \in \mathbb{R}$,

$$\begin{aligned} P(X^{(k)} \leq x) &= P(X_i \text{ are less than or equal to } x \text{ for atleast } k \text{ } i' \text{s}) \\ &= P\left(\bigcup_{l=k}^n \bigcup_{\{i_1, i_2, \dots, i_l\} \subseteq [n]} \{X_{i_1} \leq x, X_{i_2} \leq x, \dots, X_{i_l} \leq x, X_i > x, i \notin \{i_1, \dots, i_l\}\}\right) \\ &= \sum_{l=k}^n \sum_{\{i_1, i_2, \dots, i_l\} \subseteq [n]} P(X_{i_1} \leq x, \dots, X_{i_l} \leq x, X_i > x \forall i \notin \{i_1, \dots, i_l\}) \\ &= \sum_{l=k}^n \binom{n}{l} P(X_{i_1} \leq x, \dots, X_{i_l} \leq x, X_i > x \forall i \notin \{i_1, \dots, i_l\}) \\ &= \sum_{l=k}^n \binom{n}{l} F(x)^l (1 - F(x))^{n-l}. \end{aligned}$$

Hence

$$\boxed{F_{X^{(k)}}(x) = \sum_{l=k}^n \binom{n}{l} F(x)^l (1 - F(x))^{n-l}.} \quad (0.1)$$

For the rest of the discussion, we assume that the pdf of (X_1, \dots, X_n) exists and denote it by f . Also the distribution function F is assumed to be differentiable on $\{x : 0 < F(x) < 1\}$.

Differentiate (0.1) we get

$$\begin{aligned}
\frac{dF_{X^{(k)}}(x)}{dx} &= \sum_{l=k}^n \binom{n}{l} l F'(x) F(x)^{l-1} (1-F(x))^{n-l} \\
&\quad - \sum_{l=k}^{n-1} \binom{n}{l} (n-l) F'(x) F(x)^l (1-F(x))^{n-l-1} \\
&= \sum_{l=k}^n \frac{n!}{(l-1)!(n-l)!} f(x) F(x)^{l-1} (1-F(x))^{n-l} \\
&\quad - \sum_{l=k}^{n-1} \frac{n!}{l!(n-l-1)!} f(x) F(x)^l (1-F(x))^{n-l-1} \\
&= \sum_{l=k}^n \frac{n!}{(l-1)!(n-l)!} f(x) F(x)^{l-1} (1-F(x))^{n-l} \\
&\quad - \sum_{l=k+1}^n \frac{n!}{(l-1)!(n-l)!} f(x) F(x)^{l-1} (1-F(x))^{n-l} \\
&= \frac{n!}{(k-1)!(n-k)!} f(x) F(x)^{k-1} (1-F(x))^{n-k}.
\end{aligned}$$

Hence

$$f_{X^{(k)}}(x) = \frac{n!}{(k-1)!(n-k)!} f(x) F(x)^{k-1} (1-F(x))^{n-k}, x \in \mathbb{R}.$$

Distribution of $(X^{(1)}, \dots, X^{(n)})$:

Let $\mathbb{R}_{<}^n = \{(x_1, x_2, \dots, x_n) | x_1 < x_2 < \dots < x_n\}$. Then

$$P((X^{(1)}, \dots, X^{(n)}) \in \mathbb{R}_{<}^n) = 1.$$

Hence if $f_{12\dots n}$ denote the pdf of $(X^{(1)}, \dots, X^{(n)})$, then

$$f_{12\dots n}(x_1, x_2, \dots, x_n) = 0 \text{ outside } \mathbb{R}_{<}^n.$$

For $(x_1, x_2, \dots, x_n) \in \mathbb{R}_{<}^n$,

$$\begin{aligned}
&\{X^{(1)} \leq x_1, X^{(2)} \in (x_1, x_2], \dots, X^{(n)} \in (x_{n-1}, x_n]\} \\
&= \{X_{\pi(1)} \leq x_1, X_{\pi(2)} \in (x_1, x_2], \dots, X_{\pi(n)} \in (x_{n-1}, x_n] \text{ for some} \\
&\quad \text{permutation } \pi\} \\
&= \bigcup_{\pi \in S_n} \{X_{\pi(1)} \leq x_1, X_{\pi(2)} \in (x_1, x_2], \dots, X_{\pi(n)} \in (x_{n-1}, x_n]\},
\end{aligned}$$

where S_n denote the set of permutations of $(1, 2, \dots, n)$. Therefore

$$\begin{aligned} P(X^{(1)} \leq x_1, \dots, X^{(n)} \in (x_{n-1}, x_n]) &= \sum_{\pi \in S_n} P(X_{\pi(1)} \leq x_1, X_{\pi(2)} \in (x_1, x_2], \dots, X_{\pi(n)} \in (x_{n-1}, x_n]) \\ &= n! F(x_1)(F(x_2) - F(x_1)) \cdots (F(x_n) - F(x_{n-1})). \end{aligned}$$

Since (exercise¹)

$$\frac{\partial^n}{\partial x_1 \cdots \partial x_n} \left(P(X^{(1)} \leq x_1, \dots, X^{(n)} \in (x_{n-1}, x_n]) \right) = f_{12 \dots n}(x_1, \dots, x_n),$$

we get

$$f_{12 \dots n}(x_1, \dots, x_n) = \begin{cases} n! f(x_1) f(x_2) \cdots f(x_n) & \text{if } (x_1, \dots, x_n) \in \mathbb{R}_{<}^n, \\ 0 & \text{otherwise.} \end{cases}$$

Distribution of $(X^{(k)}, X^{(k+1)})$, $k = 1, 2, \dots, n-1$:

We can use $f_{12 \dots n}$ to compute the joint pdf f_{kk+1} of $(X^{(k)}, X^{(k+1)})$.

For $(x_2, \dots, x_n) \in \mathbb{R}_{<}^{n-1}$, i.e. $x_2 < x_3 < \dots < x_n$, the pdf $f_{23 \dots n}$ of $(X^{(2)}, \dots, X^{(n)})$ is given by

$$\begin{aligned} f_{23 \dots n}(x_2, \dots, x_n) &= \int_{-\infty}^{\infty} f_{12 \dots n}(y, x_2, \dots, x_n) dy \\ &= n! f(x_2) \cdots f(x_n) \int_{-\infty}^{x_2} f(y) dy \\ &= n! f(x_2) \cdots f(x_n) F(x_2). \end{aligned}$$

Therefore

$$f_{23 \dots n}(x_2, \dots, x_n) = \begin{cases} n! f(x_2) \cdots f(x_n) F(x_2) & \text{if } (x_2, \dots, x_n) \in \mathbb{R}_{<}^{n-1}, \\ 0 & \text{otherwise.} \end{cases}$$

For $(x_3, \dots, x_n) \in \mathbb{R}_{<}^{n-2}$, i.e., $x_3 < x_4 < \dots < x_n$, the pdf $f_{34 \dots n}$ of

¹Hint:

$$\begin{aligned} P(X^{(1)} \leq x_1, \dots, X^{(n)} \in (x_{n-1}, x_n]) &= P((X^{(1)}, \dots, X^{(n)}) \in (-\infty, x_1] \times \cdots \times (x_{n-1}, x_n]) \\ &= \int_{-\infty}^{x_1} \int_{x_1}^{x_2} \cdots \int_{x_{n-1}}^{x_n} f_{12 \dots n}(y_1, y_2, \dots, y_n) dy_n \cdots dy_1. \end{aligned}$$

Now take the n th order partial derivative of the above to conclude.

$(X^{(3)}, \dots, X^{(n)})$ is given by

$$\begin{aligned} f_{3\dots n}(x_3, \dots, x_n) &= \int_{-\infty}^{\infty} f_{2\dots n}(y, x_3, \dots, x_n) dy \\ &= n! f(x_3) \cdots f(x_n) \int_{-\infty}^{x_3} f(y) F(y) dy \\ (\text{use the substitution } u = F(y)) &= n! f(x_3) \cdots f(x_n) \frac{1}{2} F(x_3)^2. \end{aligned}$$

Therefore

$$f_{3\dots n}(x_3, \dots, x_n) = \begin{cases} \frac{n!}{2!} f(x_3) \cdots f(x_n) F(x_3)^2 & \text{if } (x_3, \dots, x_n) \in \mathbb{R}_{<}^{n-2}, \\ 0 & \text{otherwise.} \end{cases}$$

Repeating the above procedure upto the pdf $f_{kk+1\dots n}$ of $(X^{(k)}, \dots, X^{(n)})$ we get

$$f_{k\dots n}(x_k, \dots, x_n) = \begin{cases} \frac{n!}{(k-1)!} f(x_k) \cdots f(x_n) F(x_k)^{k-1} & \text{if } (x_k, \dots, x_n) \in \mathbb{R}_{<}^{n-k+1}, \\ 0 & \text{otherwise.} \end{cases}$$

Now for $x_k < x_{k+1} < \dots < x_{n-1}$, the pdf $f_{k\dots(n-1)}$ is given by

$$\begin{aligned} f_{k\dots(n-1)}(x_k, \dots, x_{n-1}) &= \int_{-\infty}^{\infty} f_{k\dots n}(x_k, \dots, x_{n-1}, y) dy \\ &= \frac{n!}{(k-1)!} f(x_k) \cdots f(x_{n-1}) F(x_k)^{k-1} \int_{x_{n-1}}^{\infty} f(y) dy \\ &= \frac{n!}{(k-1)!} f(x_k) \cdots f(x_{n-1}) F(x_k)^{k-1} (1 - F(x_{n-1})). \end{aligned}$$

Hence

$$f_{k\dots(n-1)}(x_k, \dots, x_{n-1}) = \begin{cases} \frac{n!}{(k-1)!} f(x_k) \cdots f(x_{n-1}) F(x_k)^{k-1} (1 - F(x_{n-1})) & \text{if } (x_k, \dots, x_{n-1}) \in \mathbb{R}_{<}^{n-k}, \\ 0 & \text{otherwise.} \end{cases}$$

Again for $x_k < x_{k+1} < \dots < x_{n-2}$, the pdf $f_{k\dots(n-2)}$ is

$$\begin{aligned} f_{k\dots(n-2)}(x_k, \dots, x_{n-2}) &= \int_{-\infty}^{\infty} f_{k\dots(n-1)}(x_k, \dots, x_{n-2}, y) dy \\ &= \frac{n!}{(k-1)!} f(x_k) \cdots f(x_{n-2}) F(x_k)^{k-1} \int_{x_{n-1}}^{\infty} f(y) (1 - F(y)) dy \\ (u = 1 - F(y)) &= \frac{n!}{(k-1)!} f(x_k) \cdots f(x_{n-1}) F(x_k)^{k-1} \int_0^{1-F(x_{n-2})} u du \\ &= \frac{n!}{(k-1)! 2!} f(x_k) \cdots f(x_{n-2}) F(x_k)^{k-1} (1 - F(x_{n-2}))^2. \end{aligned}$$

Hence the pdf $f_{k \dots (n-2)}$ of $(X^{(k)}, \dots, X^{(n-2)})$ is given by

$$f_{k \dots (n-2)}(x_k, \dots, x_{n-2}) = \begin{cases} \frac{n!}{(k-1)!2!} f(x_k) \cdots f(x_{n-2}) F(x_k)^{k-1} (1 - F(x_{n-2}))^2 & \text{if } (x_k, \dots, x_{n-2}) \in \mathbb{R}_{<}^{n-k-1}, \\ 0 & \text{otherwise.} \end{cases}$$

Repeat this, we get

$$f_{k(k+1)}(x_k, x_{k+1}) = \begin{cases} \frac{n!}{(k-1)!(n-k-1)!} f(x_k) f(x_{k+1}) F(x_k)^{k-1} (1 - F(x_{k+1}))^{n-k-1} & \text{if } (x_k, x_{k+1}) \in \mathbb{R}_{<}^2, \\ 0 & \text{otherwise.} \end{cases}$$

Distribution of $(X^{(1)}, X^{(n)})$:

Since $P((X^{(1)}, X^{(n)}) \in \mathbb{R}_{<}^2) = 1$, we compute $P(X^{(1)} \leq x, X^{(n)} \leq y)$ for $(x, y) \in \mathbb{R}_{<}^2$. Observe that

$$P(X^{(1)} \leq x, X^{(n)} \leq y) = P(X^{(n)} \leq y) - P(X^{(1)} > x, X^{(n)} \leq y).$$

$$\begin{aligned} P(X^{(1)} > x, X^{(n)} \leq y) &= P(x < X_i \leq y, \text{ for all } i = 1, \dots, n) \\ &= (F(y) - F(x))^n. \end{aligned}$$

Therefore

$$P(X^{(1)} \leq x, X^{(n)} \leq y) = F_{X^{(n)}}(y) - (F(y) - F(x))^n = F(y)^n - (F(y) - F(x))^n.$$

Observing that

$$f_{1n}(x_1, x_n) = \frac{\partial^2}{\partial x_1 \partial x_n} P(X^{(1)} \leq x_1, X^{(n)} \leq x_n)$$

we get

$$f_{1n}(x_1, x_n) = \begin{cases} n(n-1)f(x_1)f(x_n)(F(x_n) - F(x_1))^{n-2} & \text{if } (x_1, x_n) \in \mathbb{R}_{<}^2, \\ 0 & \text{otherwise.} \end{cases}$$

Example 0.1 Let X_1, \dots, X_n be i.i.d. Uniform $(0, 1)$ random variables. Compute the marginal distributions of the 2nd order statistics. Also find the joint distribution of $\min\{X_1, \dots, X_n\}$ and $\max\{X_1, \dots, X_n\}$. Hence find $P(\min\{X_1, \dots, X_n\} + \frac{1}{2} > \max\{X_1, \dots, X_n\})$

Note the pdf f_k of the k th order statistics $X^{(k)}$ is give by

$$\begin{aligned} f_k(x) &= \frac{n!}{(k-1)!(n-k)!} f(x) F(x)^{k-1} (1-F(x))^{n-k}, x \in \mathbb{R} \\ &= \frac{n!}{(k-1)!(n-k)!} x^{k-1} (1-x)^{n-k}, 0 < x < 1, = 0, \text{ otherwise.} \end{aligned}$$

i.e. $X^{(k)}$ has Beta $(k, n-k+1)$ distribution² The joint distribution of $X := \min\{X_1, \dots, X_n\}$, $Y := \max\{X_1, \dots, X_n\}$ is given by

$$\begin{aligned} f_{1n}(x, y) &= \frac{n!}{(n-2)!} f(x) f(y) (F(y) - F(x))^{n-2}, x < y, \\ &= 0, \text{ otherwise} \\ &= n(n-1)(y-x)^{n-2}, 0 < x < y < 1, = 0, \text{ otherwise.} \end{aligned}$$

Hence the required probability is given by

$$\begin{aligned} P(Y < X + \frac{1}{2}) &= n(n-1) \int_0^{\frac{1}{2}} \int_{x_1}^{x+\frac{1}{2}} (y-x)^{n-2} dy dx \\ &\quad + n(n-1) \int_{\frac{1}{2}}^1 \int_x^1 (y-x)^{n-2} dy dx \\ &= n(n-1) \left[\frac{1}{2^n(n-1)} + \frac{1}{2^n n(n-1)} \right] = \frac{n+1}{2^n}. \end{aligned}$$

²Beta (n, m) random variable is defined through the pdf

$$f(x) = \begin{cases} \frac{(n+m-1)!}{(m-1)!(n-1)!} x^{m-1} (1-x)^{n-1}, & 0 < x < 1 \\ 0 & \text{otherwise} \end{cases}$$