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~~After~~ Always good to come to a place where you are not known,
you make so many new friends.

o MY RESEARCH INTERESTS ARE IN THE FIELD OF MULTIPARAMETER ESTIMATION,
I'D LIKE TO PRESENT ONE OR TWO IDEAS FROM THAT AREA.

~~EST~~ WE KNOW A LOT ABOUT ESTIMATION OF POISSON MEANS. WHAT MORE IS LEFT TO BE STUDIED?

MUCH STATISTICAL THEORY IS BASED ON ASYMPTOTICS

BOOKS ON LOG-LINEAR MODELS

I AM GOING TO TALK TODAY WITHIN THE

ARE VERY COMPREHENSIVE.

END

ESTIMATION

MUCH STATISTICAL THEORY CONCERN'S ITSELF WITH OBTAINING EFFICIENT,

FIXED NUMBER

2. EFFICIENT METHODS FOR ESTIMATING PARAMETERS AS THE AMOUNT OF
DATA AVAILABLE INCREASES (ASYMPTOTICS) MLEST

1. ~~TALK~~ MY INTEREST TODAY CONCERN'S ANOTHER CASE

ESTIMATION METHODS FOR LARGE SPARSE DATA SETS DATA

ULTERO, IF ASYMPTOTICS ARE APPROPRIATE, $n, p \rightarrow \infty$. without information
building up about individual parameter

o QUITE DIFFERENT MET

1. ~~LARGE SPARSE DATA SETS~~ RATES UN

~~EXAMPLES~~

IN STUDIES COUNTING NUMBERS OF FAVOURABLE OUTCOMES
AND DETERMINING WHAT RISK FACTORS ARE RELATED
WE DIVIDE STUDY SUBJECTS INTO DIFFERENT SUBGROUPS

2. ~~As TECHNOLOGY IMPROV DEVELOPES, EQUIPMENT~~ IN IMAGE RECONSTR

RECORDING EQUIPMENT TYPICALLY COLLECTS MORE PRECISE DATA

~~128x128 DETECTORS~~ Now WHERE THERE WERE 64 LAST YEAR,

BUT THE IMAGE IS REQUIRED IN GREATER RESOLUTION ~~128x128~~

INTRODUCING MORE PARAMETERS. 128x128 grid

replacing 64x64 .

o QUITE DIFFERENT METHODS ARE REQUIRED,

SMOOTHING OR SHRINKAGE .

o FEW MORE WANT NOW TO DEVELOP A MORE SPECIFIC CONTEXT, TO
AND EXPERTLY FORMULATE MY PROBLEM

EXAMPLE

Numbers of study subjects

National servicemen

Army

Corps

Non-veteran

Veteran

Infantry

5400 (86)

8300 (140)

Engineer

2600 (42)

2800 (44)

Armour / Art'y

2700 (44)

2500 (40)

Minor field pres.

5900 (93)

2300 (37)

Non-field

9100 (140)

3400 (55)

Response data - number of deaths 10-12 yrs since

Cross classifying factors - Vet. status, Corps, Education...

Estimate - risk of death in each "cell" (p)

Number of deaths

Army corps

Non-veteran

Veteran

Infantry

80

122

Engineer

17

45

Armour / Art'y

35

34

Minor field pres.

38

23

Non-field

93

36

Elements

x_i , the cell count, subject to Poisson variation

$\mu = E[x]$, the expected count, to be estimated

(varies from one cell to the next)

μ_0 , another expected count,

either known constant } not to be
or model based estimate trusted!

The counts in different cells statistically independent

My objectives

Develop a methodology (MAP) for such problems

Mention other applications

Develop theory for MAP estimation

free of model assumptions (priors)

Two approaches

(a) Classical statistical theory - unrelated parameters

$$l(x|\mu) = \frac{e^{-\mu} \mu^x}{x!} \quad \hat{\mu} = x \quad \text{MVUE}.$$

(b) Prior distributions on μ

Ex $\mu^* = \mu_0^* + \varepsilon$ ^{known}
 where $\varepsilon \sim N(0, \sigma_0^2)$

STOCHASTIC
MODEL
or PRIOR

independent errors ε in different cells

MAP estimator will maximize, by choice of μ ,

$$p(\mu^* | \text{prior}, x) \propto l(x|\mu^*) p(\mu^*)$$

↑ ↓
LIKELIHOOD PRIOR DENSITY

Ex (cont.)

$$p(\mu^* | x) \propto e^{-\mu} \mu^x \exp \left\{ -\frac{1}{2\sigma_0^2} (\mu^* - \mu_0^*)^2 \right\}$$

Maximum when

$$\mu = x - \frac{1}{\sigma_0^2} (\mu^* - \mu_0^*)$$

↑ ↓
known Implicit

EXAMPLE

Contingency table cell estimates

- An insurance company wishes to set fair premiums for different classes of policy holders.
Data may be classified by age group, region,
... What is the risk for each subgroup.

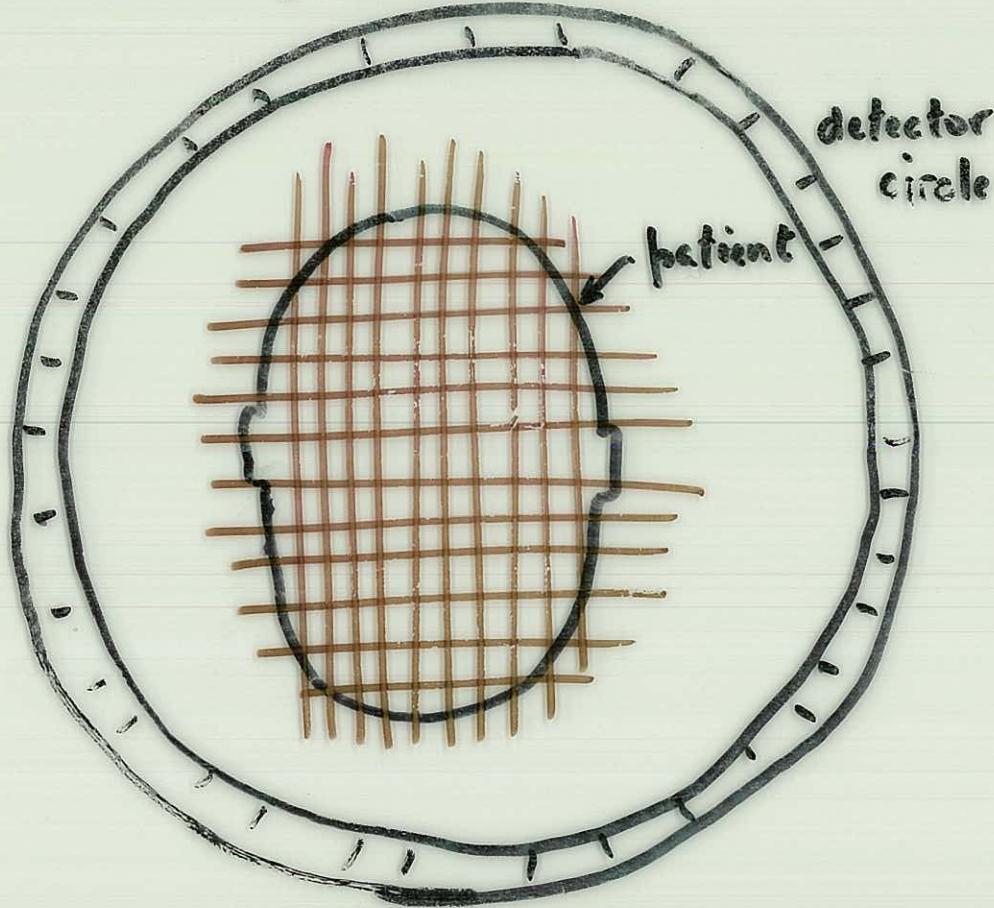
In each cell

x = number of events (risk)

M_0 = an expected number of events
(estimated by GLM?)

σ^2_0 measures expected variations in risk
between cells.

Ex Emission Tomography (ideal)



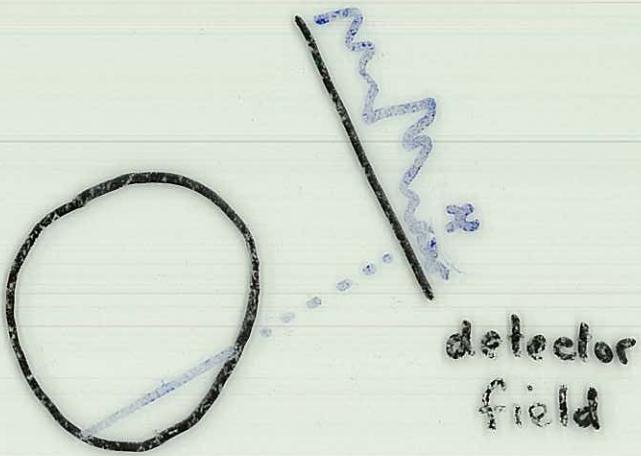
x_j = count of the photon emissions detected
which originated in pixel j

x_1, x_2, \dots are indept. Poisson r.v.'s with means
 μ_1, μ_2, \dots

$\{\mu_j\}$ is the image we wish to reconstruct

Model : independent variations of μ_j around
a known "expected" image (e.g. a
uniform background).

EXAMPLE : EMISSION TOMOGRAPHY (existing capability)



z is the count of the number of emissions directed towards one detector on the detector field.

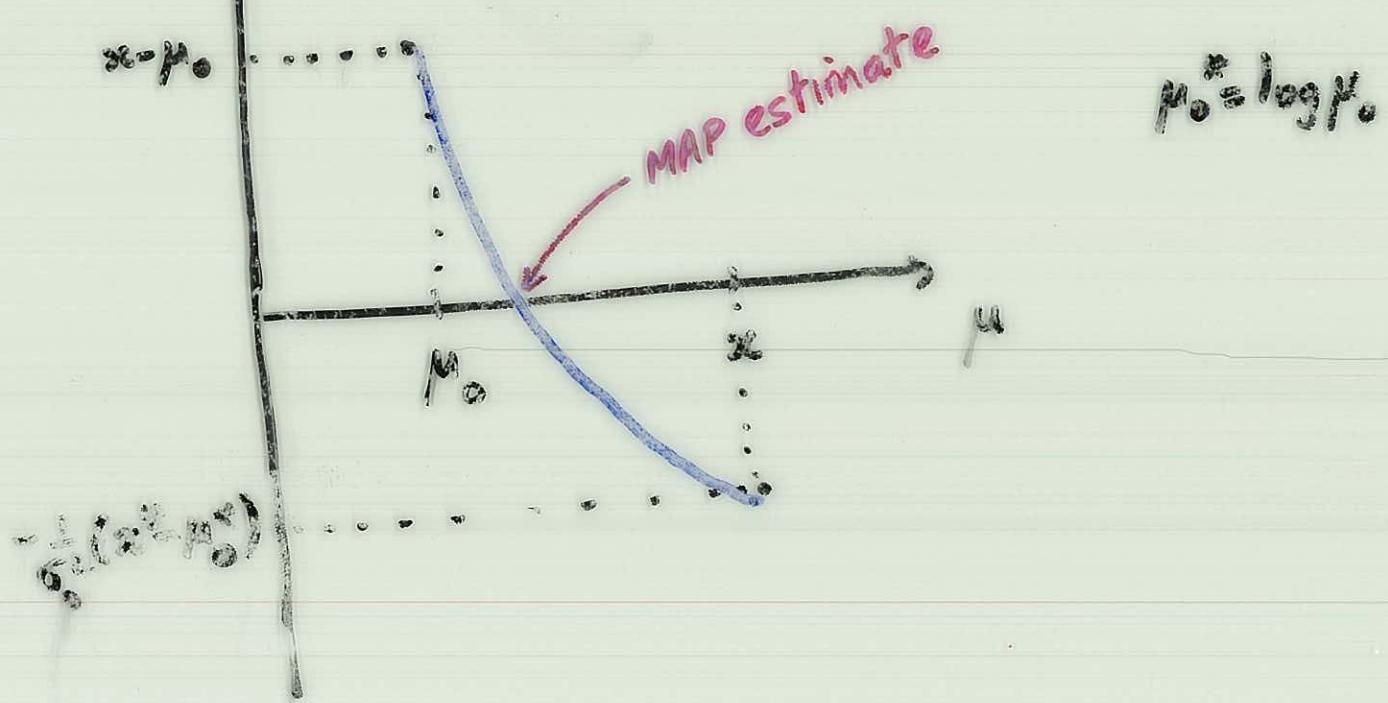
The ϕ row projection data includes Poisson variation, causing the (M.L.) image to be too rough.

Smoothing, by MAP, before reconstructing the image (Fourier techniques) may be appropriate. Set μ_0 to be an ~~over~~ expected count. σ_0^2 is a smoothing parameter.

Properties of MAP estimator

MAP is a compromise between the data and the prior value.

$$f(\mu) = x - \mu - \frac{1}{\sigma_0^2} (\mu^* - \mu_0^*)$$



As $\sigma_0^2 \uparrow \infty$, MAP approaches MLE.

As $\sigma_0^2 \downarrow 0$, MAP approaches prior value μ_0 .

Limitations of MAP

- Model dependent in choice of prior
(intensities a random sample from a log normal distn).
How do we know the process generating the intensities?
- Hyper parameters ~~are~~ (μ_0, σ^2) are assumed known. How are they to be chosen? What properties will MAP have if they are data determined (e.g. fitted values).
- MAP estimates require the solution, cell by cell, of an implicit equation.

Unbiased estimator of error mean square

• Fundamental identity for Poisson

$$\mu E^M f(x) = E^M x f(x-1)$$

$E/f(x) < \infty$

$$\text{i.e. } E^M (x-\mu)f(x) = E^M x(f(x)-f(x-1))$$

• Application 1-dimension

$$E^M (x-\mu)^2 - E^M (g(x)-\mu)^2$$

$$= E^M \{ -2(x-\mu)f(x) - f^2(x) \}$$

$$g(x) = x + f(x)$$

$$= E^M \{ \underbrace{-2x[f(x)-f(x-1)]}_{\Psi(x)} - f^2(x) \}$$

$$\Psi(x)$$

• More than 1 mean

$$X = (x_1, \dots, x_p); \quad x_i \sim P(\mu_i) \text{ indept}; \quad \mu = (\mu_1, \dots, \mu_p)$$

$$E^M \|X-\mu\|^2 - E^M \|\underbrace{g(x)-\mu}\|^2$$

$$= E^M \{ \underbrace{-2 X \cdot \nabla f - \|f(x)\|^2}_{\Psi(X)} \}$$

$$g_i(x) = x_i + f_i(x)$$

We can assess the performance of any estimator g on a given set of data X , by evaluating $\Psi(X)$.

Application to MAP

- consider a simplified approximation to MAP.

$$\hat{\mu} = \mu - \tau (x^* - \mu_0^*)$$

explicit calculation

τ is a constant controlling smoothing

$$x^* = \log\left(\frac{x + \gamma}{\gamma}\right)$$

↳ Euler's constant $\approx 0.58...$

- how should we choose τ ?

my choice $\tau = \frac{p-2}{\sum (x^* - \mu_0^*)^2} = \frac{p-2}{S}$ where

the sum extends over all cells, and p represents the number of cells with non-zero counts.

- this estimator has smaller error mean square than MLE:

for any configuration of means (almost)

for any expected values μ_0

because for any number of cells, provided 3 or more.

$$E^P \|x - \mu\|^2 - E^P \|\hat{\mu} - \mu\|^2 \geq E^P \left\{ \frac{(p-2)^2}{S} \right\}$$

Choosing expected values (μ_{ij}) which are data determined

- Choose a simple log-linear model

$$\mu_{oi}^* = \underline{a}_i' \underline{\beta} \quad \underline{\beta} = (\beta_1, \dots, \beta_q)$$

\underline{a}_i known

- Estimate β by ordinary least squares to minimize

$$S = \sum (x_i^* - \underline{a}_i' \underline{\beta})^2$$

i.e. $\hat{\beta} = (\mathbf{A}'\mathbf{A})^{-1}\mathbf{A}'\underline{x}^*$ for given design \mathbf{A} .

- Then the estimator

$$\hat{\mu} = x - \frac{(b-q-2)}{S} \cdot (x^* - \underline{a}_i' \hat{\beta})$$

has error mean square reduction exceeding

$$E^P \left\{ \frac{(b-q-2)^2}{S} \right\}.$$

Conclusion

- o In problems with large numbers of parameters, standard statistical theory is unhelpful. Stein's contribution.
- o Smoothing procedures must be employed.
- o The unbiased estimator of error mean square can provide estimates of smoothing parameters.
- o This choice leads to estimation of cell means that are
 - (a) efficient in that they incorporate prior information
 - (b) safe the extent of use of prior information depends on how consistent it is with data.
- o "Log-linear estimators" have explicit, small sample error mean square properties.

$$c = \sum_{i=1}^D c_{ij} \quad \text{pixel } j.$$

$$= \sum_{i=1}^D 1_{\{j \text{ projects onto } i\}} = N_{ANG}.$$

Uniform grey image - bad news for patient -

apparent extra variability in projections (random fluctuations).