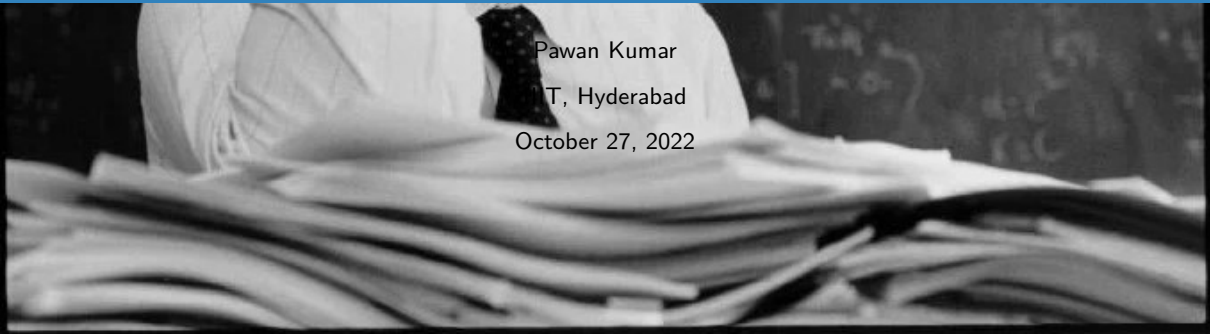




Probability and Statistics (Monsoon 2022)

Lecture-19



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October 21, 2022

Outline

① Statistical Inference

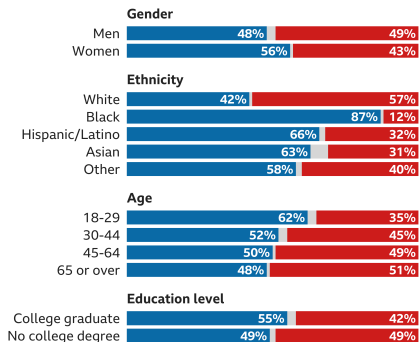
Maximum Likelihood Estimation

Motivation for Statistical Inference...

National exit poll

Support by gender, ethnicity, age group and education

■ Biden ■ Other ■ Trump



Sample size: 15,318 respondents

All figures have a margin of error which is wider for smaller sub-groups

Source: Edison Research/NEP via Reuters, 4 Nov, 17.00 EST (22.00 GMT)



- On the left, US exit poll results
- Poll on Trump Vs Biden
- Sample size of 15,318
- Error margin shown in grey
- Draw conclusions from the sample data
- Will inference fail? How much it can fail?
- How confident we are of this?

Motivation for Statistical Inference...

POLL OF ALL POLLS				
	NDA	MAHAGATHBANDHAN	LJP	OTHERS
JAN KI BAAT	104	128	6	5
C-VOTER	116	120	1	6
MY AXIS	80	150	4	9
TV9 BHARATVARSH	115	120	4	4
POLL OF POLLS	104	129	4	6

BIHAR ASSEMBLY ELECTIONS RESULTS 2020				
TOTAL SEATS 243				
NDA 125	MGB 110	OTH	8	
BJP 74	RJD 75	LJP 1		
JD(U) 43	CONG 19	AIMIM 5		
HAM 4	CPI-ML 11	BSP 1		
VIP 4	CPM 3	OTHERS 1		
	CPI 2			

- On the left, poll of polls showing clear majority for MAHAGATHBANDHAN
- After election, NDA has full majority
- How do we estimate such errors?

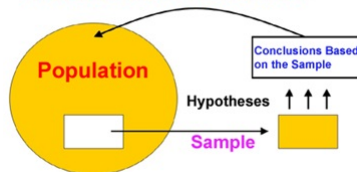
Definition of Statistical Inference...

Definition of Statistical Inference

Statistical inference is a collection of methods that deal with drawing conclusions from data that are prone to random variation.

- knowledge of probability is used
- we need to work with real data
- distribution of the data may not be known

Statistical Inference



- Two types: Frequentist and Bayesian

Statistical Inference Problem

To determine an unknown quantity, get some data, and then estimate the required quantity using this data.

Frequentist or Classical Inference...

Recall: A statistical inference problem is to estimate an unknown quantity

Frequentist Inference

Here the **unknown** quantity is assumed to be fixed quantity and not random. So, the unknown quantity θ is to be estimated by the observed data.

- Let θ be the percentage of people who will vote for a given candidate
 - $\hat{\theta} = \frac{Y}{n}$, Y is the number of people among randomly chosen ones who will vote for candidate
 - Although, θ is non random, we estimate it via $\hat{\theta}$, a random variable
 - Here $\hat{\theta}$ is random variable, because it depends on random sample

Bayesian Inference...

What is Bayesian Inference?

Here the unknown quantity Θ is assumed to be a random variable. Furthermore, we assume to have some **initial** guess about the distribution of Θ . After we **observe** the data, we can update the distribution of Θ using **Bayes rule**.

- Consider communication systems in which information is transmitted in the form of bits
- In each transmission, the transmitter sends a 1 with probability p , and sends a 0 with probability $1 - p$
- Hence, if Θ is the transmitted bit, then $\Theta \sim \text{Bernoulli}(p)$
- Let us assume that at receiver end we get the output X
- The **problem** now is to estimate Θ from the noisy output X
- We use the **prior** knowledge that $\Theta \sim \text{Bernoulli}(p)$

What is Random Sampling? Motivation with an example...



Simple Random Sampling



- 1 Choose a random sample of size $n : X_1, \dots, X_n$ **with replacement**
 - 2 We chose a person uniformly at random from the population and let X_1 be the height of that person. Here, every person in the population has the same chance of being chosen
 - 3 To determine the value of X_2 , again we choose a person uniformly (and independently from the first person) at random and let X_2 be the height of that person. Again, every person in the population has the same chance of being chosen
- In general, X_i is the height of the i th person that is chosen uniformly and independently from the population
 - **why do we do the sampling with replacement?**
 - if the population is large, then the probability of choosing one person twice is extremely low
 - big advantage of sampling **with replacement** is that X_i 's will be independent
 - that is, working with **independently and identically distributed** makes analysis simpler

Definition of Random Sample...

Definition of Random sample

The collection of random variables $X_1, X_2, X_3, \dots, X_n$ is said to be a **random sample** of size n if they are **independent** and **identically distributed (i.i.d.)**, i.e.,

- 1 $X_1, X_2, X_3, \dots, X_n$ are **independent** random variables, and
- 2 they have the **same** distribution, i.e.,

$$F_{X_1}(x) = F_{X_2}(x) = \dots = F_{X_n}(x), \quad \text{for all } x \in \mathbb{R}$$

Point Estimator and Sample Mean...

Definition of Sample Mean

Let X_1, X_2, \dots, X_n be **random sample**. That is, here X_1, X_2, \dots, X_n are **i.i.d.** That is, following holds true for **i.i.d.** random variables

- 1 The X_i 's are independent (since they are i.i.d.)
- 2 $F_{X_1}(x) = F_{X_2}(x) = \dots = F_{X_n}(x) = F_X(x)$ (the CDFs are same)
- 3 $E[X_i] = E[X] = \mu < \infty$
- 4 $0 < \text{Var}(X_i) = \text{Var}(X) = \sigma^2 < \infty$

Then the **sample mean** is defined as follows

$$\bar{X} = \frac{X_1 + X_2 + \dots + X_n}{n}$$

Recall: Properties of Sample Mean...

Properties of sample mean, \bar{X}

1 $E[\bar{X}] = \mu, \quad \text{Var}(\bar{X}) = \frac{\sigma^2}{n}$

2 Weak law of large numbers (WLLN)

$$\lim_{n \rightarrow \infty} P(|\bar{X} - \mu| \geq \epsilon) = 0$$

3 Central limit theorem: The random variable

$$Z_n = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \frac{X_1 + X_2 + \cdots + X_n - n\mu}{\sqrt{n}\sigma}$$

converges in distribution to the standard normal random variable

$$\lim_{n \rightarrow \infty} P(Z_n \leq x) = \Phi(x), \quad \text{for all } x \in \mathbb{R},$$

where $\Phi(x)$ is standard normal CDF.

Order Statistics and its PDF and CDF...

Order Statistics and its PDF and CDF

Let X_1, X_2, \dots, X_n be random sample from a continuous distribution with CDF $F_X(x)$. If we order the random variables from smallest to largest i.e., $X_{(1)} < X_{(2)} < \dots < X_{(n)}$ with

$$X_{(1)} = \min(X_1, X_2, \dots, X_n) \quad \text{and} \quad X_{(n)} = \max(X_1, X_2, \dots, X_n),$$

then $X_{(i)}$'s is called **order statistics**. The **CDF** and **PDF** of $X_{(i)}$ are given by

$$f_{X_{(i)}} = \frac{n!}{(i-1)!(n-i)!} f_X(x) [F_X(x)]^{i-1} [1 - F_X(x)]^{n-i}$$
$$F_{X_{(i)}} = \sum_{k=i}^n \binom{n}{k} [F_X(x)]^k [1 - F_X(x)]^{n-k}$$

Also, the **joint PDF** of $X_{(1)}, X_{(2)}, \dots, X_{(n)}$ is given by

$$f_{X_{(1)}, \dots, X_{(n)}}(x_1, \dots, x_n) = \begin{cases} n! f_X(x_1) f_X(x_2) \dots f_X(x_n) & \text{for } x_1 \leq x_2 \leq \dots \leq x_n \\ 0 & \text{otherwise} \end{cases}$$

Example of Order Statistics...

Example (Order Statistics)

Let X_1, X_2, \dots, X_4 be a random variable from the Uniform(0,1) distribution, and let $X_{(1)}, X_{(2)}, X_{(3)}, X_{(4)}$ be the **order statistics** of X_1, X_2, \dots, X_4 .

Find the PDFs of $X_{(1)}, X_{(2)}$, and $X_{(4)}$.

Answer to previous problem...

Point Estimator, Biased and Unbiased Estimators...

Definitions: point estimator, bias and unbiased estimators

- 1 Let θ be an **unknown** parameter to be estimated. For example, $\theta = E[X]$
- 2 Let X_1, X_2, \dots, X_n be a random sample using which we want to estimate θ . Here X_i 's have same distribution
- 3 To estimate θ we define **point estimator** $\hat{\Theta}$ as follow

$$\hat{\Theta} = h(X_1, X_2, \dots, X_n)$$

- 4 There can be many possible point estimators, which one to choose?

- For example if $\theta = E[X]$, then $\hat{\Theta} = h(X_1, \dots, X_n) = \frac{X_1 + \dots + X_n}{n}$

- 5 **Bias:** The **bias** of a point estimator $\hat{\Theta}$ is defined as

$$B(\hat{\Theta}) = E[\hat{\Theta}] - \theta$$

- If bias is close to 0, then $\hat{\Theta}$ is closer to θ
- We say that $\hat{\Theta}$ is an **unbiased estimator** for a parameter θ if

$$B(\hat{\Theta}) = 0, \quad \text{for all possible values of } \theta$$

Example

Example

Example Let X_1, \dots, X_n be a random sample. Show that the sample mean

$$\hat{\theta} = \bar{X} = \frac{X_1 + X_2 + \dots + X_n}{n} \quad (1)$$

is an unbiased estimator of $\theta = EX_i$.

Unbiased Estimator is not Necessarily a Good Estimator...

Fact

Show that unbiased estimator is **not** necessarily a good estimator.

Solution:

In the above example, if we choose $\hat{\theta}_1 = X_1$, then $\hat{\theta}_1$ is also an unbiased estimator of θ .

- However, we suspect that $\hat{\theta}_1$ is probably not as good as the sample mean \bar{X} .
- We need other measures to ensure that an estimator is a good estimator.

Mean Squared Error...

Mean squared error

The **mean squared error (MSE)** of a point estimator $\hat{\theta}$ denoted by $\text{MSE}(\hat{\theta})$ is defined as

$$\text{MSE}(\hat{\theta}) = E[(\hat{\theta} - \theta)^2]$$

Example (Application of MSE)

Let X_1, X_2, \dots, X_n be a random sample from a distribution with mean $E[X_i] = \theta$, and variance $\text{Var}(X_i) = \sigma^2$. For the following two estimators for θ

1 $\hat{\theta}_1 = X_1$

2 $\hat{\theta}_2 = \bar{X} = \frac{X_1 + X_2 + \dots + X_n}{n}$

Find $\text{MSE}(\hat{\theta}_1)$ and $\text{MSE}(\hat{\theta}_2)$ and show that for $n > 1$

$$\text{MSE}(\hat{\theta}_1) > \text{MSE}(\hat{\theta}_2).$$

Answer to previous problem...

Answer to previous problem...

Relationship of MSE, Variance, and Bias...

Property

If $\hat{\Theta}$ is a point estimator for θ ,

$$\text{MSE}(\hat{\Theta}) = \text{Var}(\hat{\Theta}) + B(\hat{\Theta})^2$$

Solution:

$$\text{MSE}(\hat{\Theta}) = E[(\hat{\Theta} - \theta)^2] \tag{2}$$

$$= \text{Var}(\hat{\Theta} - \theta) + (E[\hat{\Theta} - \theta])^2 \tag{3}$$

$$= \text{Var}(\hat{\Theta}) + B(\hat{\Theta})^2. \tag{4}$$

Consistent Estimator...

Definition of Consistent Estimator

Let $\hat{\Theta}_1, \hat{\Theta}_2, \dots, \hat{\Theta}_n, \dots$, be a sequence of point estimators of θ . We say that $\hat{\Theta}_n$ is a **consistent estimator** of θ , if

$$\lim_{n \rightarrow \infty} P(|\hat{\Theta}_n - \theta| \geq \epsilon) = 0, \quad \text{for all } \epsilon > 0$$

Theorem

Let $\hat{\Theta}_1, \hat{\Theta}_2, \dots$, be a sequence of point estimators of θ . If

$$\lim_{n \rightarrow \infty} \text{MSE}(\hat{\Theta}_n) = 0$$

then $\hat{\Theta}_n$ is a **consistent estimator** of θ

Answer to previous problem...

Example

Example

Let X_1, X_2, \dots, X_n be a random sample with mean $E[X_i] = \theta$ and variance $\text{Var}(X_i) = \sigma^2$. Show that $\hat{\Theta}_n = \bar{X}$ is a consistent estimator of θ .

Definition of Sample Variance and Sample Standard Deviation...

Sample Variance and Sample Standard Deviation

Let X_1, X_2, \dots, X_n be a random variable with mean $E[X_i] = \mu < \infty$, and variance $0 < \text{Var}(X_i) < \sigma^2 < \infty$. The **sample variance** of this random sample is defined as

$$S^2 = \frac{1}{n-1} \sum_{k=1}^n (X_k - \bar{X})^2 = \frac{1}{n-1} \left(\sum_{k=1}^n X_k^2 - n\bar{X}^2 \right)$$

We can check that sample variance is an unbiased estimator of σ^2 . The **sample standard deviation** is defined as

$$S = \sqrt{S^2}$$

and it is usually used as an estimator for σ . Also, S is an unbiased estimator of σ .

Example

Example

Example Let X_1, \dots, X_n be a random variable with mean $E[X_i] = \mu$ and variance $\text{Var}(X_i) = \sigma^2$. Suppose that we use

$$\bar{S}^2 = \frac{1}{n} \sum_{k=1}^n (X_k - \bar{X})^2 = \frac{1}{n} \left(\sum_{k=1}^n X_k^2 - n\bar{X}^2 \right) \quad (5)$$

to estimate σ^2 . Find the bias of this estimator

$$B(\bar{S}^2) = E[\bar{S}^2] - \sigma^2. \quad (6)$$

Example (Sample Mean, Sample Variance, Sample Standard Deviation)

Let T be the time that is needed for a specific task in a factory to be completed. In order to estimate the mean and variance of T , we observe a random sample T_1, T_2, \dots, T_6 . Thus, T_i 's are i.i.d. and have the same distribution as T . We obtain the following values (in minutes):

18, 21, 17, 16, 24, 20.

Find the values of the sample mean, the sample variance, and the sample standard deviation for the observed sample.

Answer to previous problem...

Example (Prediction Using Data)

I have a bag that contains 3 balls. Each ball is either red or blue, but I have no information in addition to this. Thus, the number of blue balls, call it θ , might be 0, 1, 2, or 3. I am allowed to choose 4 balls at random from the bag with replacement. We define the random variables X_1, X_2, X_3 , and X_4 as follows

$$X_i = \begin{cases} 1 & \text{if the } i\text{th chosen ball is blue} \\ 0 & \text{if the } i\text{th chosen ball is red} \end{cases}$$

We observe here that X_i 's are i.i.d. and $X_i \sim \text{Bernoulli}\left(\frac{\theta}{3}\right)$. After the experiment, we observe the values for X_i 's

$$x_1 = 1, x_2 = 0, x_3 = 1, x_4 = 1.$$

From above, we have 3 blue balls and 1 red ball. Answer the following

- 1 Find the probability of the observed sample $(x_1, x_2, x_3, x_4) = (1, 0, 1, 1)$ for each possible θ
- 2 Find the value of θ that maximizes the probability of the observed sample

Answer to previous problem...

Answer to previous problem...