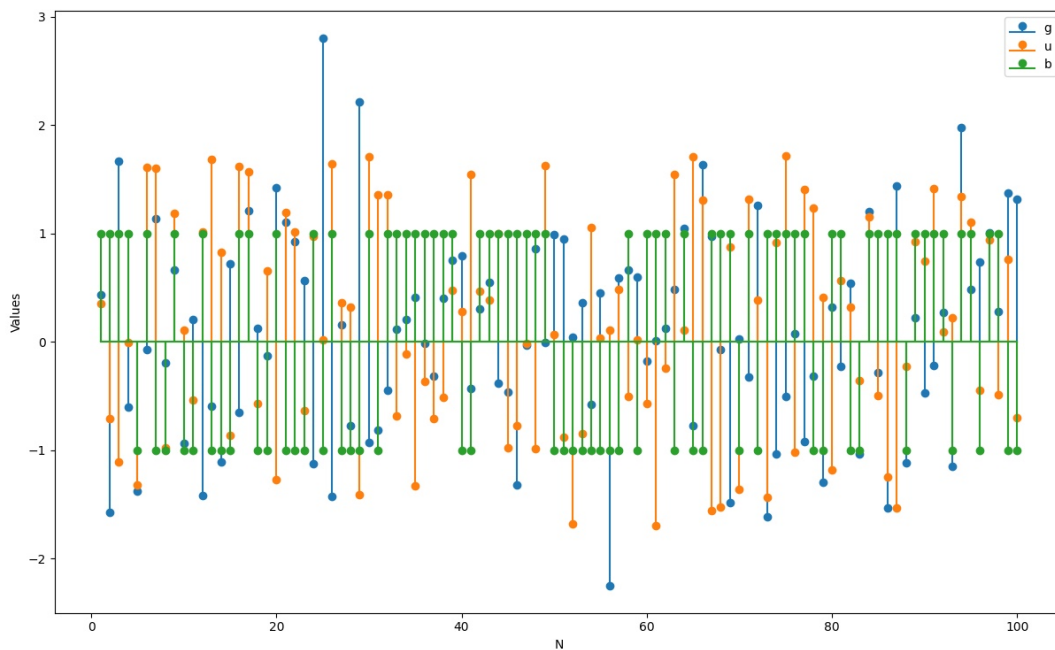


## problem 1)

g) Unit - Variance:



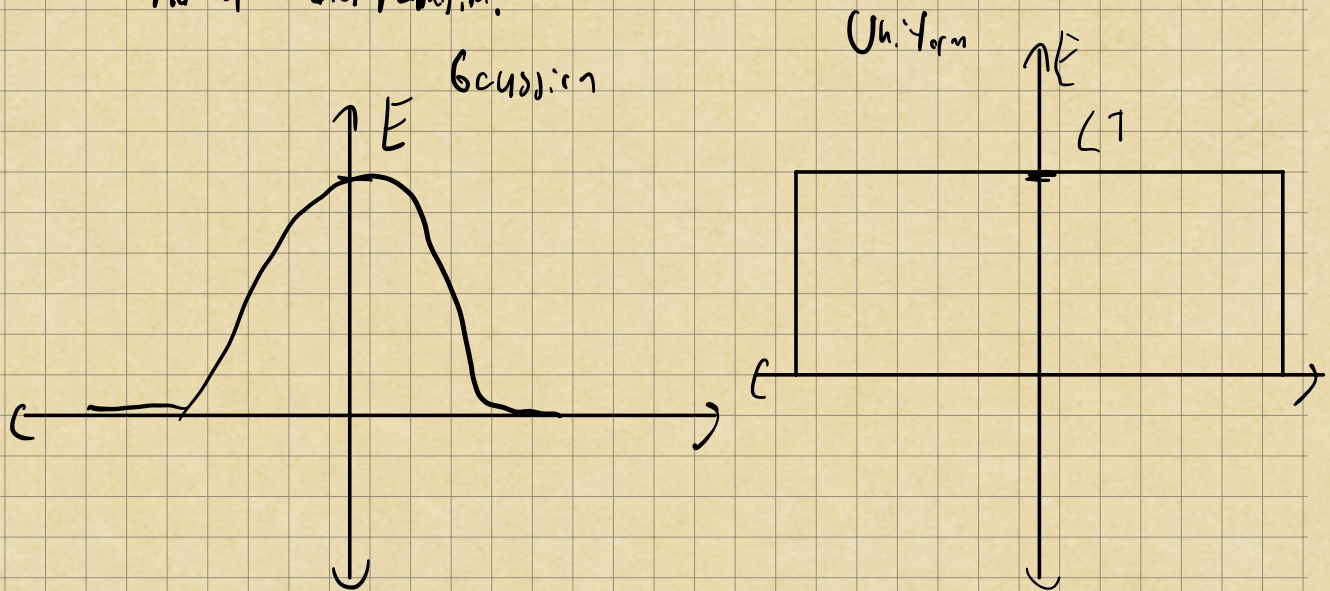
- All noisy samples are uncorrelated.

In other words, if you were given one sample there is no information about the value of the next one.

- Binary is limited to 2-values 1 and -1, where as the uniform and gaussian has a interval.



- Uniform is equally likely to take any value within a specified range but in contrast the gaussian follows a normal distribution.



b)

For binary

$$p_X = \begin{cases} 0.5, & x = \pm 1 \\ 0, & \text{else} \end{cases}$$

Gaussian:

$$p_X = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$$

Uniform:

$$p_X = \begin{cases} \frac{1}{2\sqrt{3}}, & x \in (-\sqrt{3}, \sqrt{3}) \\ 0, & \text{else} \end{cases}$$



Compute mean - Value:

$$\mu_x = E\{X[n]\} = \int_{-\infty}^{\infty} x f_X(x) dx = 0$$

Gaussian:

$$\mu_x = \int_{-\infty}^{\infty} \underbrace{x}_{\text{odd}} \cdot \underbrace{\frac{1}{\sqrt{2}} e^{-\frac{x^2}{2}}}_{\text{even}} dx$$

Since it is an odd function and symmetric interval

$$\Rightarrow \underline{\underline{\mu_x = 0}}$$

Uniform:

$$\begin{aligned} \mu_x &= \int_{-\infty}^{\infty} \frac{1}{2\sqrt{3}} \cdot x \, dx = \left[ \frac{x^2}{4\sqrt{3}} \right]_{-\sqrt{3}}^{\sqrt{3}} \\ &= \frac{3}{4\sqrt{3}} - \left( -\frac{3}{4\sqrt{3}} \right) \end{aligned}$$



$$= \underline{\underline{0}}$$

For binary

$$\mu_x = p(1) \cdot 1 + p(-1) \cdot (-1)$$

$$= 0.5 - 0.5 = \underline{\underline{0}}$$

Autocorrelation function:

$$R_X(k) = E[X(n)X(n-k)] = \sigma^2 \delta(k)$$

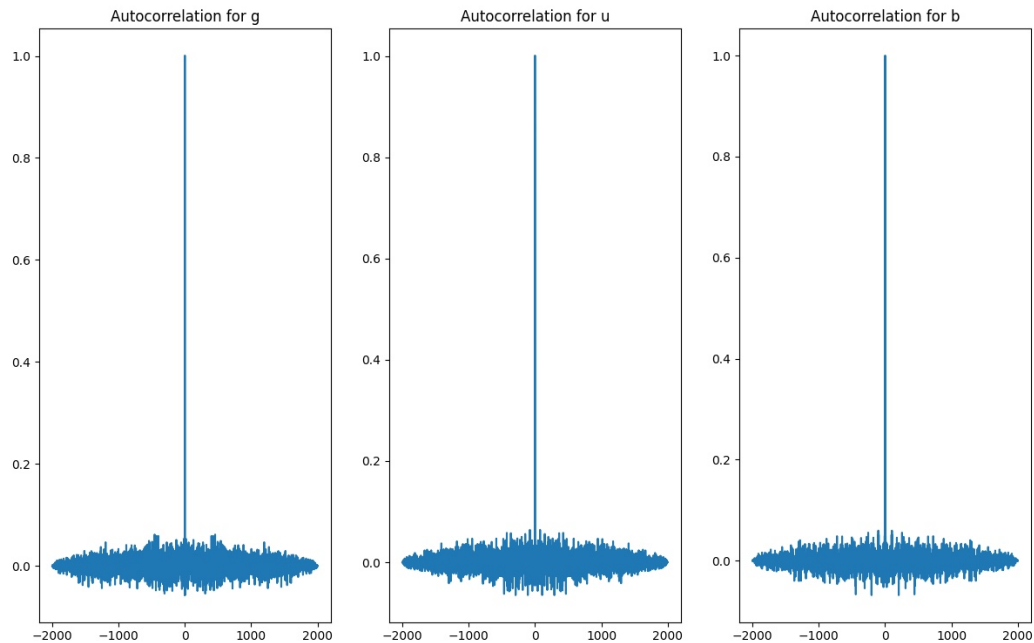
Power density correlation:

$$S_X = \text{DTFT} \{ R_X(k) \} = \underline{\underline{\sigma^2}}$$



c)

```
Eier@dhcp-10-24-21-119 DIGSIG 0ving / % /usr/bi
Sample Mean for g: -0.006095568348404223
Theoretical Mean for g: 0
Sample Mean for u: 0.00400903886484537
Theoretical Mean for u: 0
Sample Mean for b: 0.001
Theoretical Mean for b: 0
```



- As we can see from the plots and values, the theoretical values are pretty accurate.



## Problem 2)

Random signal  $X[n]$  is generated by filtering white Gaussian noise  $W[n]$  with variance  $\sigma_w^2 = \frac{3}{4}$  by a

Causal filter with transfer function

$$H(z) = \frac{1}{1 + \frac{1}{2}z^{-1}}$$

a)

- The mean value of  $m[n] = 0$

- The auto correlation of  $r_{ww}[k] = \sigma_w^2 \delta(k) = \frac{3}{4} \delta(k)$

-  $r_{ww}(k) = \frac{3}{4}$  for  $k=0$

This gives values for  $X[n]$ :

$$m_x = E[X[n]] = m_w \sum_{k=-\infty}^{\infty} h(k) = 0$$

$$r_{xx}(m) = r_{ww}(m) \cdot r_{hh}(m)$$



Where

$$r_{hh}(m) = \sum_{-\infty}^{\infty} h(n) h(n+m)$$

$$\text{and } h(n) = \mathcal{Z}^{-1}\{H(z)\} = \left(-\frac{1}{2}\right)^n u(n)$$

This gives

$$r_{hh}(m) = \sum_{-\infty}^{\infty} \left(-\frac{1}{2}\right)^n u(n) \left(-\frac{1}{2}\right)^{n+m} u(n+m)$$

We first find for  $m \geq 0$

$$r_{hh}(m) = \sum_0^{\infty} \left(-\frac{1}{2}\right)^n u(n) \left(-\frac{1}{2}\right)^{n+m} u(n+m)$$

$$u(n)=1 \Rightarrow r_{hh}(m) = \sum_0^{\infty} \left(-\frac{1}{2}\right)^{2n+m}$$

$u(n+m)=1$

$$= \left(-\frac{1}{2}\right)^m \sum_0^{\infty} \left(-\frac{1}{4}\right)^n$$

$$= \underline{\underline{\frac{4}{3} \left(-\frac{1}{2}\right)^m}}$$



for  $m < 0$

we know:  $r_{hh}(m) = r_{hh}(-m)$

$$\Rightarrow r_{hh} = \frac{4}{3} \left(-\frac{1}{2}\right)^{|m|}, \quad \forall m$$

$$\begin{aligned}\Rightarrow r_{xx}(m) &= r_{ww}(m) * r_{hh}(m) = \frac{3}{4} \delta(m) * \left(\frac{4}{3}\right) \left(-\frac{1}{2}\right)^{|m|} \\ &= \left(-\frac{1}{2}\right)^{|m|}\end{aligned}$$

Power density spectrum:

$$P_{xx}(\omega) = P_{ww}(\omega) |H(\omega)|^2$$

$$\text{where } H(\omega) = H(z) \Big|_{z=e^{j\omega}}$$

$$\Rightarrow |H(\omega)|^2 = H(\omega) H(\omega)^* = \frac{4}{5 + 4 \cos \omega}$$



We know that:

$$\Gamma_{ww}(w) = \sigma_w^2(w) = \frac{3}{4}$$

$$\Rightarrow \Gamma_{xx}(w) = \frac{3}{4} \cdot \frac{4}{5 + 4 \cos w} = \underline{\underline{\frac{3}{5 + 4 \cos w}}}$$

power of the signal:

$$p_x = E[x^2(n)] = \Gamma_{xx}(0) = \underline{\underline{1}}$$

b)

Estimates:

$$\underline{\hat{m}_x = \frac{1}{N} \sum_0^{N-1} x(n)}$$

power:

$$\underline{\underline{\hat{p}_x = \frac{1}{N} \sum_{n=0}^{N-1} x^2(n)}}$$



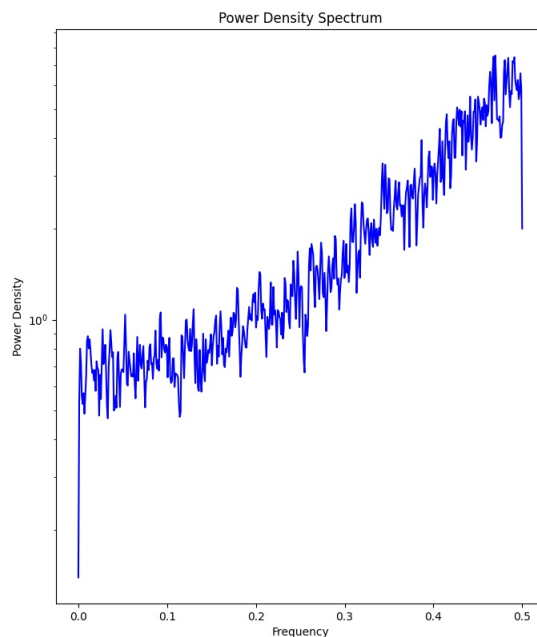
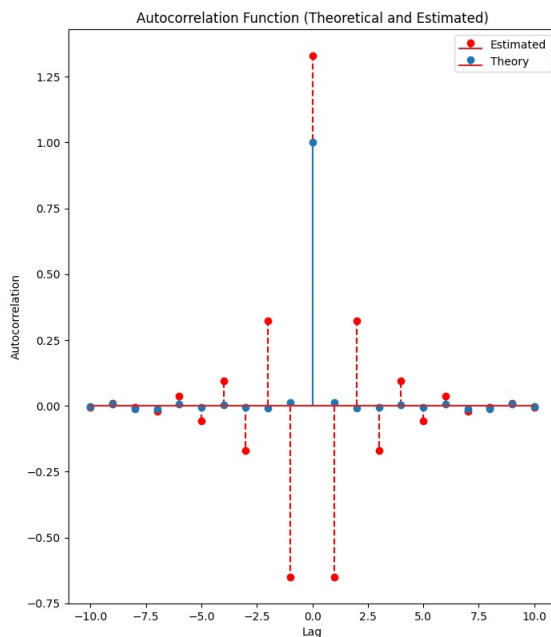
Autocorrelation:

$$\hat{\gamma}_{xx}(l) = \begin{cases} \frac{1}{N} \sum_0^{N-l} x(n)x(n+l), & 0 \leq l \leq N-l \\ 0, & N \leq l \end{cases}$$

Power spectral density:

$$\hat{f}_{xx} = \text{DFT}(\hat{\gamma}_{xx}(l))$$

c)





d) Couldn't figure this one out.

e) How done this

### problem 3)

g)

```
#Task3a)
K = 20

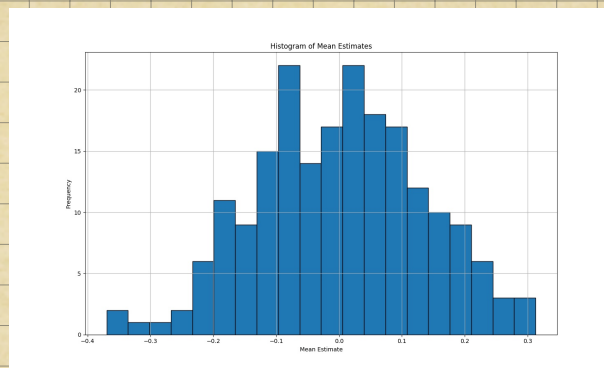
# Initialize an array to store the mean estimates
mean_estimates = []

# Repeat the mean calculation for 200 segments
for i in range(200):
    # Extract a segment of length K from the signal
    segment = x[i * K: (i + 1) * K]

    # Calculate the mean of the segment
    segment_mean = np.mean(segment)

    # Append the mean estimate to the list
    mean_estimates.append(segment_mean)
```

b)





c)

Result from Code:

```
Estimated Mean: 0.0017699369465278897  
Estimated Variance: 0.02033342792956164
```

d)

```
For K = 20:  
Estimated Mean: 0.009194283868357246  
Estimated Variance: 0.015414014133601595  
  
For K = 40:  
Estimated Mean: 0.00949630515671533  
Estimated Variance: 0.014223619659367224  
  
For K = 100:  
Estimated Mean: 0.025640978212525256  
Estimated Variance: 0.01276697397870169
```

e)

Increasing the segment length  $k$  aligns with theoretical expectation, by reducing the bias and variance, this results in a more accurate and less variable mean estimates for  $\hat{\mu}_x$ .