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# Guest lecture: MCMC with Discrete Parameters

Jakob Torgander



# Outline

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1. Discrete parameters - Introduction & discussion
2. Describe three methods for computing posteriors with discrete latent parameters
  - Marginalization
  - Gibbs sampling
  - Continuous approximation using Gumbel-Softmax-distribution
3. (Short) demonstration of methods.



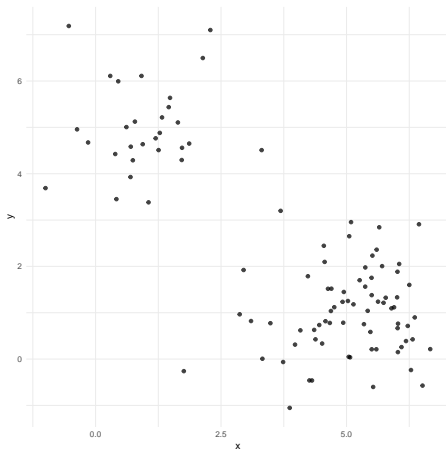
# Motivation

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- Discrete variables are everywhere
  - Count data: e.g. number of car accidents
  - Categorical data
  - Decision/classification problems: (eg. yes/no)
  - Factor analysis
- In many problems latent (hidden) discrete variables exists: conclusions changes if data is segmented into groups
- While current state-of-the-art method Hamiltonian Monte Carlo (HMC) works for discrete *data* **HMC does not directly work for discrete *parameters*.**



## Case study - Gaussian mixture model



- Latent class variable  $C$
- $p(y) = \sum_{k=1}^K \mathbb{1}(C = k) \mathcal{N}(y | \mu_k, \sigma_k)$ ,
- Task: identify cluster assignments, probabilities and centers



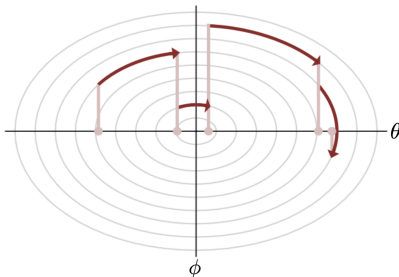
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## Section 2

### Computation



# Hamiltonian Monte Carlo (HMC) - Recap



(?, ?): Given a current parameter-momentum pair  $(\theta_i, \phi_i)$ , Hamiltonian  $H$  and mass matrix  $M$ :

1. Sample a new momentum variable  $\phi_{i+1} \sim \mathcal{N}(0, M)$
2. Lift  $\theta_i$  onto the joint phase space  $(\theta, \phi)$
3. Integrate the flow defined by  $H(\theta_i, \phi_{i+1}) = \text{constant}$  using Hamilton's equations
4. Project back to original parameter space to receive new parameter sample  $\theta_{i+1}$



## Recap: the leapfrog integrator

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Step 3 of HMC is based on the leapfrog algorithm

$$\psi \leftarrow \psi + \frac{1}{2}\epsilon \frac{d \log q(\theta|y)}{d\theta} \quad \text{1st momentum update}$$

$$\theta \leftarrow \theta + \epsilon M^{-1} \psi \quad \text{Parameter update}$$

$$\psi \leftarrow \psi + \frac{1}{2}\epsilon \frac{d \log q(\theta|y)}{d\theta} \quad \text{2nd momentum update,}$$

where  $q$  denotes the target posterior density.



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where  $q$  denotes the target posterior density. Q: Why does this fail when  $q$  is discrete?





## HMC does not work for discrete posteriors

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Main problem: Computation of the gradient  $\frac{d \log q(\theta|y)}{d\theta}$  requires the limits (partial derivatives)

$$\frac{\partial q}{\partial \theta_i} = \lim_{h \rightarrow 0} \frac{q(\theta + h\mathbf{e}_i) - q(\theta)}{h}$$

to exist. This only happens when  $q$  is continuous!



## Method 1: Marginalization

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Idea: Sum (marginalize) out the latent discrete parameters (?, ?).

By the law of total probability:

$$p(y) = \sum_{k=1}^K p(y|c_k)p(c_k).$$

Then,  $p(y)$  is continuous if  $p(y|c_k), p(c_k)$  are.



# Marginalization

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**Example:** for the GM-model:

$$p_Y(y, |\pi, \mu, \sigma) = \sum_{k=1}^K \underbrace{\pi_k}_{p(c_k)} \underbrace{\mathcal{N}(y|\mu_k, \sigma_k)}_{p(y|c_k)},$$

where  $\pi_k$  are (continuous) parameters.



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where  $\pi_k$  are (continuous) parameters. **Remark:** Compare with the original model formulation

$$p(y|\pi, \mu, \sigma) = \sum_{k=1}^K \mathbb{1}(C = k) \mathcal{N}(y|\mu_k, \sigma_k)$$



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## Method 2: Gibbs sampling

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**Recall:** Gibbs sampling: conditional (or block) sampling of  $\theta$

$$\theta_j \sim p(\theta_j | \theta_{-j}, y)$$



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For the GM-model, with  $\sigma$  known:

1. For each observation  $y$ , sample classes  $c_i$  with probability

$$p(c_i | \mu, \sigma, y) = \frac{p(y | c_i) p(c_i)}{\sum_{j=1}^K p(y | c_j) p(c_j)} = \frac{p(c_i) \mathcal{N}(y | \mu_i, \sigma)}{\sum_{j=1}^K p(c_j) \mathcal{N}(y | \mu_j, \sigma)}$$



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2. Sample means  $\mu_i$  using the conditional distributions  $p(\mu_i | y, c_i)$  (normal if likelihood and prior for  $\mu$  is )



## Method 3: Continuous approximation Gumbel-Softmax

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### Ideas:

- Approximate a discrete (categorical) distribution with a continuous distribution.
- The approximated distribution can then be used with HMC
- Use the "Gumbel trick" ( $\epsilon$ ,  $\epsilon$ ) from the field of deep learning





## "The Gumbel trick"

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**Proposition:** Let  $Z$  be a categorical r.w with probability distribution  $\pi = (\pi_1, \dots, \pi_K)$  and let  $G_i$  be Gumbel(0,1)-distributed with density

$$f_{G_i} = e^{-x-e^{-x}}.$$

Then the random variable

$$U = \arg \max_i G_i + \log \pi_i,$$

follows the same distribution as  $Z$ .



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Then the random variable

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follows the same distribution as  $Z$ . **Q: How to use the argmax-function in a density?**



## Gumbel-Softmax distribution

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Idea by ? (?): Approximate argmax with the softmax function.

$$Y_i = \frac{\exp((\log(\pi_i) + G_i)/\tau)}{\sum_{j=1}^k \exp((\log(\pi_j) + G_j)/\tau)},$$

where  $G_i$  are Gumbel(0,1)-distributed and  $\tau$  is a "temperature" parameter.



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As  $\tau$  approaches 0,  $Y = (Y_1, \dots, Y_K)$  then tends to a "one-hot" vector on the form

$$[0, \dots, 0, 1, 0, \dots, 0],$$

where a "1" in position  $m$  indicates the  $m$ -th class.



## Gumbel-Softmax distribution

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This yields the Gumbel-Softmax (GS) density function:

$$p_{\pi, \tau}(y_1, \dots, y_K) = (K-1)! \cdot \tau^{K-1} \left( \sum_{i=1}^K \pi_i / y_i^{\tau} \right)^{-k} \prod (\pi_i / y_i^{\tau+1}).$$

Continuous! Can hence be used with HMC and Stan.



## Methods - Summary

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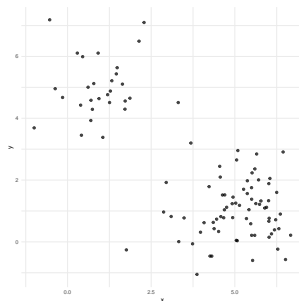
Method	Pros	Cons
Marginalization	Works efficiently with HMC	Does not return the discrete parameter
Gibbs	Returns classes, reliable for "simple" distributions	Difficult (& sometimes less efficient) for non-conjugate distributions
Gumbel-Softmax.	Returns classes, works with HMC	High dependency on temperature $\tau$ , leapfrog (very) unstable for low temperatures



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## Section 3

### Demonstration

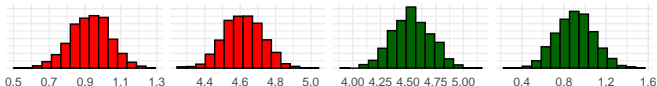


- Data: simulated gaussian mixtures with means  $\mu_1 = (1, 5), \mu_2 = (5, 1)$  and  $\sigma_1 = \sigma_2 = \mathbf{I}$
- Weakly informative  $\mathcal{N}(0, 10)$ -prior used for all  $\mu$
- Dirichlet(1,1)-prior (see e.g. ( ?, ?, p. 69)) used for HMC methods (1 and 3)
- 2000 samples generated for each method

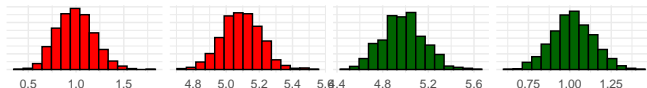




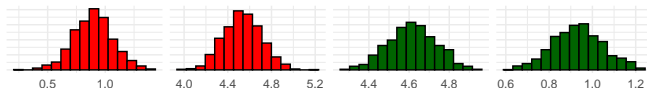
Method: Integration



Method: Gibbs



Method: Cont. approx.

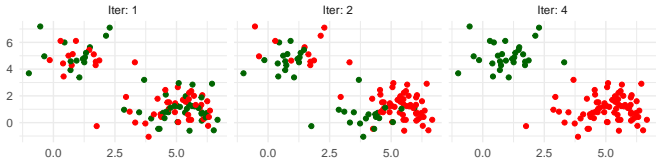


- All three methods correctly identify the centers (red =  $\mu_1$ , green =  $\mu_2$ )
- Gibbs sampler closer to "ground truth" in this case
- Difference possible due to weakly informative Dirichlet prior of Method 1,3. Needs further investigation..

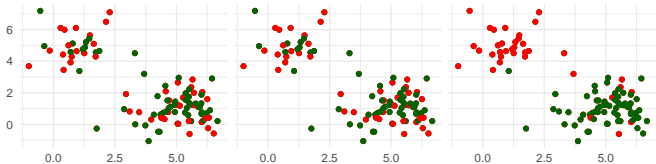


# Class assignments - convergence

Method: Gibbs



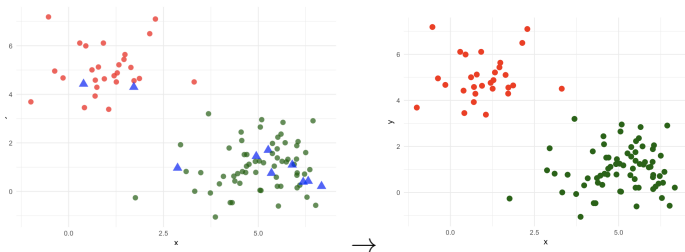
Method: Cont. approx.



- Both Gibbs and Continuous approx. converges quickly to the correct classes
- Gibbs sampler one iteration quicker



## Use case: imputing missing values



- Gibbs, and Gumbel-Softmax method can be used to impute missing values (classes)
- Idea: Generate class parameter if non-present in the data and use the actual class otherwise
- For general tips about handling missing values, see (?, ?)



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## Future research

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- How do the methods scale with data size and dimension?
- How can  $\tau$  in the GS-approximation be selected and tuned?
- Performance and convergence of methods on more complicated, high-dimensional posteriors?



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Thank you!



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# References

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