

01NAEX - Lecture 10

Introduction to Mixed Linear Models

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Introduction to Mixed Linear Models

Throughout the lecture we have assumed that the factors in an experiment were **fixed factors** and therefore the statistical inference made about these factors are confined to the specific levels studied.

In some situations, the factor levels are chosen at random from a larger population of possible levels, and we wish to draw a conclusions about the entire population of levels, not just those that were used in the experimental design. In this situation, the factor is said to be a **random factor**.

In this lesson we focus on random effects and on mixed effects model. Nested and split-plot designs, two situations where random factors are frequently encountered in practice, will be presented in following lectures.

Introduction in to Mixed Linear Models

What is mixed models?

- ▶ Mixed models model the covariance structure of data and they are generalization of linear models where observations are not independent.

What are the most common mixed models?

- ▶ Random effects models: certain effects arise from a distribution and add additional source of variation
- ▶ Random coefficients: used for repeated measures to model relationship with time, estimates are correlated.

Mixed Effects models

Mixed-effects models describe the relationship between a response variable and one or more covariates recorded with it.

- ▶ **Fixed-effects:** the factors in the experiment have a predetermined set of levels and the only inferences are for the levels of the factors actually used in the experiment.
- ▶ **Random effects:** the levels of factors used in the experiment are randomly selected from a population of possible levels. The inferences from the data in the experiment are for all levels of the factors in the population from which the levels were selected and not only the levels used in the experiment.
- ▶ **Mixed-effects:** the levels of some of the factors used in the experiment are randomly selected from a population of possible levels, whereas the levels of the other factors in the experiment are predetermined.

Introduction into Mixed Linear Models

Why use mixed models?

- ▶ To avoid mistakes.
- ▶ To get more appropriate fixed effect estimates.
- ▶ To broaden inference over wider population.
- ▶ To deal with missing data.
- ▶ To be able to handle correlation structure in data.
- ▶ To be able to handle heteroscedasticity between treatment groups.
- ▶ More flexible for repeated measure and longitudinal data (allows missing data and does not require measurements at the same time points).

Why is the terminology so various ...

Different names arose partially because these methods were re-invented in a variety of fields (psychology, education, agronomy, economics, ...), each with different slants and emphasis.

You can find ...

- ▶ Mixed-effects models
- ▶ Hierarchical linear models
- ▶ Multilevel models
- ▶ Variance component models
- ▶ Random-effects models
- ▶ Random-coefficients regression models
- ▶ ...

Fixed vs. random factors difference

Comparison of Fixed and Mixed model

	Fixed Model	Mixed Model
Levels	Given number of possibilities	Selected at random from a population
New experiment	Use same levels	Use different levels from same population
Goal	Estimates means of fixed levels	Estimate variance of population of means
Inference	$H_0 : \mu_1 = \mu_2 = \dots = \mu_k$ $H_1 : \exists i, j \in \{1, \dots, k\} \mu_i \neq \mu_j$	$H_0 : \sigma_\mu^2 = 0$ $H_1 : \sigma_\mu^2 > 0$

Packages for Linear Mixed Effects models in R

There are several packages in R that deal with linear mixed models, most widely used are:

- ▶ **nlme** - Non-Linear Mixed Effects, Part of R, `library(nlme)`
Fit only Gaussian outcomes, it is possible to specify the variance-covariance matrix for the random effects.
- ▶ **lme4** - Linear Mixed Effects, `library(lme4)`
Can be used to fit generalized mixed-effects regression models, it is not possible to specify the variance-covariance matrix for the random effects, but can handle with diagonal covariance structures or unstructured covariance matrices.
- ▶ **Other useful packages:** `merTools`, `glmmTMB`, `brms`, `modelr`

All of them has some advantages and some disadvantages. Notations, specifications, special functions and classes are different.

For power analysis for random effects in mixed models is possible to use package called **pamm**.

Simple example:

```
library(nlme)
library(lme4)
library(MASS)

data(oats)
names(oats) = c('block', 'variety', 'nitrogen', 'yield')
oats$mainplot = oats$variety
oats$subplot = oats$nitrogen
attach(oats)

# SIMPLE MIXED EFFECTS MODEL
modell1a=lme( yield~variety*nitrogen, random = ~ 1|block)
modell2a=lmer(yield~variety*nitrogen + (1|block))

# NESTED MIXED EFFECTS MODEL (will be covered next lecture)
modell1b=lme( yield~variety*nitrogen, random=~1|block/mainplot)
modell2b=lmer(yield~variety*nitrogen + (1|block/mainplot))
```

Simple example - comparison of classical and Mixed Models approach

Description of data set:

In a pharmaceutical company the use of NIR (Near Infrared Reflectance) spectroscopy was investigated as an alternative to the more cumbersome (and expensive) HPLC method to determine the content of active substance in tablets.

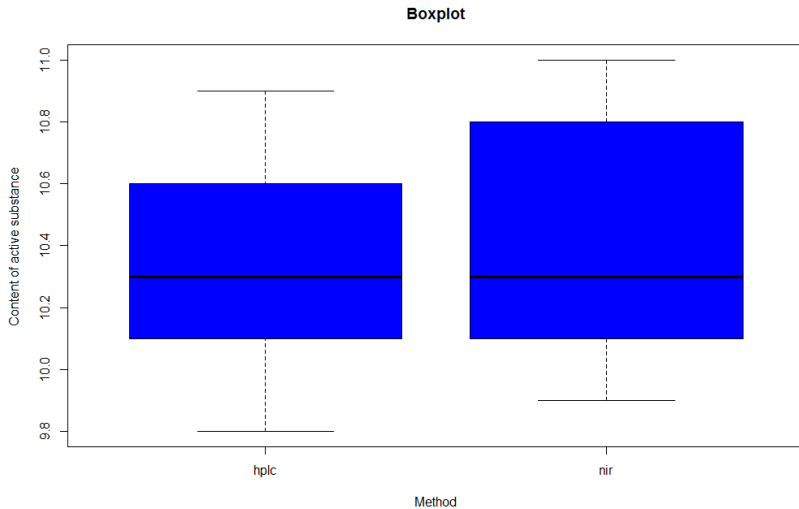
Source of data set:

Brockhoff and Thierry-Carstensen 2003, Test set validation using simple statistical methods.

	hplc	nir	difference
Tablet_1	10.4	10.1	0.3
Tablet_2	10.6	10.8	-0.2
Tablet_3	10.2	10.2	0.0
Tablet_4	10.1	9.9	0.2
Tablet_5	10.3	11.0	-0.7
Tablet_6	10.7	10.5	0.2
Tablet_7	10.3	10.2	0.1
Tablet_8	10.9	10.9	0.0
Tablet_9	10.1	10.4	-0.3
Tablet_10	9.8	9.9	-0.1

The aim is to study the method differences.

Simple example - comparison of classical and Mixed Models approach



Simple example - comparison of classical and Mixed Models approach

Simple analysis of the pharmaceutical data by the paired t-test

```
> mean(d)    -0.05
> var(d)      0.08722222
> sd(d)       0.2953341
t.test(d)
# alternatively:  t.test(hplc, nir, paired = TRUE)
Paired t-test
data:  hplc and nir
t = -0.5354, df = 9, p-value = 0.6054
alternative hypothesis: true difference in means is not equal to 0
95 percent confidence interval:  -0.2612693  0.1612693
sample estimates: mean of the differences -0.05
```

The standard error of the mean \bar{d} (ie. the uncertainty of the estimated difference) :

$$SE_{\bar{d}} = \frac{s_d}{\sqrt{n}} = \frac{0.295}{\sqrt{10}} = 0.0934$$

$$\text{t-statistics: } t = \frac{\bar{d}}{SE_{\bar{d}}} = \frac{-0.05}{0.0934} = -0.5354$$

Final regression model: $d_i = \mu + \varepsilon_i \quad \varepsilon_i \sim N(0, \sigma^2)$

Estimated model parameters: $\hat{\mu} = \bar{d}, \hat{\sigma} = s_d$

Simple example - comparison of classical and Mixed Models approach

Simple analysis of the pharmaceutical data by the ANOVA

We have randomized balanced design with two treatments - methods and 10 blocks - tablets.

Used regression model: $y_{ij} = \mu + \alpha_i + \beta_j + \varepsilon_{ij}$ $\varepsilon_{ij} \sim N(0, \sigma^2)$

```
summary(aov(y ~ method+tablet))
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
method	1	0.0125	0.01250	0.287	0.6054
tablet	9	2.0005	0.22228	5.097	0.0118
Residuals	9	0.3925	0.04361		

The uncertainty of the average method difference:

$$SE_{\bar{y}_2 - \bar{y}_1} = \sqrt{MS_E \left(\frac{1}{n_1} + \frac{1}{n_2} \right)} = \sqrt{0.0436 \left(\frac{1}{10} + \frac{1}{10} \right)} = 0.0934$$

We have the same p-value as in the t-test approach and the uncertainty since the t statistic for testing $H_0 : \mu_1 = \mu_2$ is

$$t = \frac{\bar{y}_2 - \bar{y}_1}{\sqrt{MS_E \left(\frac{1}{n_1} + \frac{1}{n_j} \right)}}$$

Simple example - comparison of classical and Mixed Models approach

Misleading information from the classical ANOVA approach

The ANOVA approach compare treatments means and is based on knowledge obtained only from certain measurements:

In pharmaceutical data example, the **ANOVA approach** is valid only for statements about the 10 specific tablets in the experiment, not for tablets in general.

The uncertainty of the average NIR value is given by using only the 10 NIR measurements: $SE_{\bar{y}_1} = \frac{s_1}{\sqrt{10}} = 0.127$, $s_1 = sd(NIR) = 0.4012$.

On the other hand, the **Mixed Models approach** is valid for tablets (blocks) in general, since we consider the 10 tablets as a random sample.

Let us consider Mixed Model with tablet as a random effects:

$$y_{ij} = \mu + a_i + \beta_j + \varepsilon_{ij} \quad \varepsilon_{ij} \sim N(0, \sigma^2), \quad a_i \sim N(0, \sigma_T^2).$$

Simple example - comparison of classical and Mixed Models approach

Comparison of Fixed and Mixed model

The observations in Mixed Model are no longer independent and the tablet differences become a part of the variance structure.

1. The expected value of the observation y_{ij} .
2. The variance of the observation y_{ij} .
3. The relation between two observations.

Comparison table of fixed and mixed model

	Fixed Model	Mixed Model
1 $E[y_{ij}]$	$\mu + \alpha_i + \beta_j$	$\mu + \beta_j$
2 $var[y_{ij}]$	σ^2	$\sigma_T^2 + \sigma^2$
3 $cov(y_{i_1 j_1}, y_{i_2 j_2}) \text{ } i_1 \neq i_2$	0	σ_T^2 (if $i_1 = i_2$), 0 (if $i_1 \neq i_2$)

Simple example - comparison of classical and Mixed Models approach

Analysis of the pharmaceutical data by the Mixed Models

Used mixed model:

$$y_{ij} = \mu + a_i + \beta_j + \varepsilon_{ij} \quad \varepsilon_{ij} \sim N(0, \sigma^2), \quad a_i \sim N(0, \sigma_T^2).$$

```
# ways how to estimate mixed models
# function lme (need library nlme)
# function lmer (need library lme4)
> model<-lme( y~method, random = ~1|tablet, data=temp)
> model<-lme( y~method, random=list(tablet=~1), data=temp)
> model<-lmer(y~method+(1|tablet), data=temp)
> summary(model)
```


Simple example - comparison of classical and Mixed Models approach

Analysis of the pharmaceutical data by the Mixed Models

Used mixed model: $y_{ij} = \mu + a_i + \beta_j + \varepsilon_{ij}$ $\varepsilon_{ij} \sim N(0, \sigma^2)$, $a_i \sim N(0, \sigma_T^2)$.

```
> summary(model)
```

Linear mixed model fit by REML

Formula: $y \sim \text{method} + (1 \mid \text{tablet})$

Random effects:

Groups	Name	Variance	Std.Dev.
tablet	(Intercept)	0.089333	0.29889
Residual		0.043611	0.20883

Number of obs: 20, groups: tablet, 10

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	10.39000	0.11530	90.11
methodhplc	-0.05000	0.09339	-0.54

Correlation of Fixed Effects:

	(Intr)
methodhplc	-0.405

anova(model)	numDF	denDF	F-value	p-value
(Intercept)	1	9	9666.573	<.0001
method	1	9	0.287	0.6054

Simple example - comparison of classical and Mixed Models approach

Analysis of the pharmaceutical data by the Mixed Models

Used mixed model:

$$y_{ij} = \mu + a_i + \beta_j + \varepsilon_{ij} \quad \varepsilon_{ij} \sim N(0, \sigma^2), \quad a_i \sim N(0, \sigma_T^2),$$

with estimated Random effects components:

Groups	Name	Variance	Std.Dev.
tablet	(Intercept)	0.089333	0.29889
Residual		0.043611	0.20883

$$\hat{\sigma}^2 = 0.043611, \quad \hat{\sigma}_T^2 = 0.089333,$$

and Fixed effects components:

	Estimate	Std. Error	t value
(Intercept)	10.39000	0.11530	90.11
methodhplc	-0.05000	0.09339	-0.54

$$\hat{\mu} = 10.39000, \quad \hat{\beta}_1 = 0, \quad \hat{\beta}_2 = -0.05000$$

Example with missing values

We have the same pharmaceutical data set, but we add 5 non-paired measurement from each method.

n	tablet	method	y	n	tablet	method	y	n	tablet	method	y
1	1	hplc	10.4	14	7	nir	10.2	27	14	hplc	NA
2	1	nir	10.1	15	8	hplc	10.9	28	14	nir	10.3
3	2	hplc	10.6	16	8	nir	10.9	29	15	hplc	NA
4	2	nir	10.8	17	9	hplc	10.1	30	15	nir	9.7
5	3	hplc	10.2	18	9	nir	10.4	31	16	hplc	10.3
6	3	nir	10.2	19	10	hplc	9.8	32	16	nir	NA
7	4	hplc	10.1	20	10	nir	9.9	33	17	hplc	9.6
8	4	nir	9.9	21	11	hplc	NA	34	17	nir	NA
9	5	hplc	10.3	22	11	nir	10.8	35	18	hplc	10.0
10	5	nir	11.0	23	12	hplc	NA	36	18	nir	NA
11	6	hplc	10.7	24	12	nir	9.8	37	19	hplc	10.2
12	6	nir	10.5	25	13	hplc	NA	38	19	nir	NA
13	7	hplc	10.3	26	13	nir	10.5	39	20	hplc	9.9
								40	20	nir	NA

The mixed model in a direct way may give information about the key issues in a data set, that a straightforward fixed ANOVA does not.

Example with missing values

Analysis by fixed effects ANOVA

The fixed effect analysis only uses the information in the first 10 tablets.

```
> model1<-lm(y~tablet+method)
```

```
> anova(model1)
```

Analysis of Variance Table

Response: y

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
tablet	19	3.7230	0.195947	4.4931	0.01288 *
method	1	0.0125	0.012500	0.2866	0.60537
Residuals	9	0.3925	0.043611		

Note that only the Tablets row of the table has changed compared to the previous analysis.

Example with missing values

Analysis by mixed model

```
> model2<-lme(y~method,random=~1| tablet,data=temp2t)
> anova(model2)
```

	numDF	denDF	F-value	p-value
(Intercept)	1	19	15715.981	<.0001
method	1	9	0.687	0.4285

```
> summary(model2)
```

Linear mixed-effects model fit by REML

Random effects:

Formula: ~1 | tablet

	(Intercept)	Residual
StdDev:	0.3192429	0.2085067

Fixed effects: y ~ method

	Value	Std.Error	DF	t-value	p-value
(Intercept)	10.283857	0.09259174	19	111.06668	0.0000
methodhplc	-0.072111	0.08697180	9	-0.82913	0.4285

Correlation:

	(Intr)
methodhplc	-0.47

Example with missing values

Compariosn of Fixed and Mixed model approach

We use the R function `estimable`.

FIXED:

	Estimate	Std. Error	t value	DF	Pr(> t)	Lower.CI	Upper.CI
LS HPLC	10.2125	0.06177356	165.321550	9	0.0000000	10.0727585	10.3522415
LS NIR	10.2625	0.06177356	166.130957	9	0.0000000	10.1227585	10.4022415
LS DIF	-0.0500	0.09339284	-0.535373	9	0.6053664	-0.2612693	0.1612693

MIXED:

	Estimate	Std. Error	t value	DF	Pr(> t)	Lower.CI	Upper.CI
LS HPLC	10.21174590	0.09259174	110.2878719	9	0.0000000	10.0022888	10.4212030
LS NIR	10.28385689	0.09259174	111.0666778	19	0.0000000	10.0900602	10.4776536
LS DIF	-0.07211099	0.08697180	-0.8291307	9	0.4284719	-0.2688549	0.1246329

Note that apart from giving a slightly different value, the Mixed model estimation is also more precise, than the one only based on tablets 1-10. This is the kind of analysis that the mixed model for this situation leads to.

Simple Introduction to Theory of Mixed Models

Recall simple linear regression model of **fixed effects** approach:

$$y_{ij} = \mu + \alpha_i + \beta_j + \varepsilon_{ij} \quad \varepsilon_{ij} \sim N(0, \sigma^2), \quad i \in \{1, 2\}, \quad j \in \{1, 2, 3\},$$

$$\underbrace{\begin{pmatrix} y_{11} \\ y_{21} \\ y_{12} \\ y_{22} \\ y_{13} \\ y_{23} \end{pmatrix}}_{\mathbf{y}} = \underbrace{\begin{pmatrix} 1 & 1 & 0 & 1 & 0 & 0 \\ 1 & 0 & 1 & 1 & 0 & 0 \\ 1 & 1 & 0 & 0 & 1 & 0 \\ 1 & 0 & 1 & 0 & 1 & 0 \\ 1 & 1 & 0 & 0 & 0 & 1 \\ 1 & 0 & 1 & 0 & 0 & 1 \end{pmatrix}}_{\mathbf{X}} \cdot \underbrace{\begin{pmatrix} \mu \\ \alpha_1 \\ \alpha_2 \\ \beta_1 \\ \beta_2 \\ \beta_3 \end{pmatrix}}_{\boldsymbol{\beta}} + \underbrace{\begin{pmatrix} e_{11} \\ e_{21} \\ e_{12} \\ e_{22} \\ e_{13} \\ e_{23} \end{pmatrix}}_{\mathbf{e}}.$$

equivalently,

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{e},$$

where \mathbf{y} is a vector of all observations - response variables, \mathbf{X} is a known matrix of predictors (usually called design matrix), $\boldsymbol{\beta}$ is a vector of unknown coefficients - fixed effects parameters and \mathbf{e} is a vector of unknown independent measurement errors $\mathbf{e} \sim N(\mathbf{0}, \sigma^2 \mathbf{I})$.

Simple Introduction to Theory of Mixed Models

Regression model of **Mixed linear model** approach:

$$y_{ij} = \mu + \alpha_i + b_j + \varepsilon_{ij} \quad b_j \sim N(0, \sigma_B^2), \varepsilon_{ij} \sim N(0, \sigma^2) \quad i \in \{1, 2\}, j \in \{1, 2, 3\},$$

$$\underbrace{\begin{pmatrix} y_{11} \\ y_{21} \\ y_{12} \\ y_{22} \\ y_{13} \\ y_{23} \end{pmatrix}}_{\mathbf{y}} = \underbrace{\begin{pmatrix} 1 & 1 & 0 \\ 1 & 0 & 1 \\ 1 & 1 & 0 \\ 1 & 0 & 1 \\ 1 & 1 & 0 \\ 1 & 0 & 1 \end{pmatrix}}_{\mathbf{X}} \cdot \underbrace{\begin{pmatrix} \mu \\ \alpha_1 \\ \alpha_2 \end{pmatrix}}_{\boldsymbol{\beta}} + \underbrace{\begin{pmatrix} 1 & 0 & 0 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \end{pmatrix}}_{\mathbf{Z}} \cdot \underbrace{\begin{pmatrix} b_1 \\ b_2 \\ b_3 \end{pmatrix}}_{\mathbf{u}} + \underbrace{\begin{pmatrix} e_{11} \\ e_{21} \\ e_{12} \\ e_{22} \\ e_{13} \\ e_{23} \end{pmatrix}}_{\mathbf{e}},$$

equivalently,

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\mathbf{u} + \mathbf{e},$$

where \mathbf{y} is a vector of all observations - response variables, \mathbf{X} is a known matrix of predictors (usually called design matrix), $\boldsymbol{\beta}^0$ is a vector of unknown coefficients - fixed effects parameters, \mathbf{Z} is the design matrix for random effects, \mathbf{u} is the vector of random effects $\mathbf{u} \sim N(\mathbf{0}, \mathbf{G})$, $\text{cov}(u_i, u_j) = G_{ij}$ and \mathbf{e} is a vector of unknown independent measurement errors $\mathbf{e} \sim N(\mathbf{0}, \mathbf{R})$, $\text{cov}(e_i, e_j) = R_{ij}$, typically \mathbf{R} is diagonal.

Simple Introduction to Theory of Mixed Models

Distribution of response variable \mathbf{y} in Mixed Models:

Let us consider linear mixed model:

$$\mathbf{Y} = \mathbf{X}\beta + \mathbf{Z}u + \mathbf{e}, \quad \text{where } \mathbf{u} \sim N(\mathbf{0}, \mathbf{G}), \quad \mathbf{e} \sim N(\mathbf{0}, \mathbf{R}).$$

Distribution of \mathbf{y} is multivariate normal

$$\mathbf{y} \sim N(\mu, \mathbf{V})$$

with

$$\mu = E[\mathbf{X}\beta + \mathbf{Z}u + \mathbf{e}] = \mathbf{X}\beta$$

and

$$\mathbf{V} = \text{var}[\mathbf{X}\beta + \mathbf{Z}u + \mathbf{e}] = \text{var}[\mathbf{Z}u] + \text{var}[\mathbf{e}] = \mathbf{ZGZ}^T + \mathbf{R}.$$

$$\mathbf{y} \sim N(\mathbf{X}\beta, \mathbf{ZGZ}^T + \mathbf{R})$$

Notice that if \mathbf{R} is diagonal and we have random block effect model then \mathbf{V} is block diagonal matrix.

Simple Introduction to Theory of Mixed Models

Example of a \mathbf{y} distribution for simple linear Mixed Effects Models:

Let us consider linear mixed effect model:

$$y_{ij} = \mu + \alpha_i + b_j + \varepsilon_{ij} \quad b_j \sim N(0, \sigma_B^2), \quad \varepsilon_{ij} \sim N(0, \sigma^2) \quad i \in \{1, 2\}, \quad j \in \{1, 2, 3\},$$

$$\boldsymbol{\mu} = \begin{pmatrix} \mu + \alpha_1 \\ \mu + \alpha_2 \\ \mu + \alpha_1 \\ \mu + \alpha_2 \\ \mu + \alpha_1 \\ \mu + \alpha_2 \end{pmatrix} \quad \mathbf{V} = \begin{pmatrix} \sigma^2 + \sigma_B^2 & \sigma_B^2 & 0 & 0 & 0 & 0 \\ \sigma_B^2 & \sigma^2 + \sigma_B^2 & 0 & 0 & 0 & 0 \\ 0 & 0 & \sigma^2 + \sigma_B^2 & \sigma_B^2 & 0 & 0 \\ 0 & 0 & \sigma_B^2 & \sigma^2 + \sigma_B^2 & 0 & 0 \\ 0 & 0 & 0 & 0 & \sigma^2 + \sigma_B^2 & \sigma_B^2 \\ 0 & 0 & 0 & 0 & \sigma_B^2 & \sigma^2 + \sigma_B^2 \end{pmatrix}$$

Notice that two observations from the same block are correlated !!

Theory of Mixed Models

The likelihood function L for mixed effect models:

For given parameter values, the likelihood function L returns a measure of the probability of observing response variables \mathbf{y} .

The negative log likelihood function for a mixed model is given by:

$$l(\mathbf{y}, \beta, \gamma) = \frac{1}{2} \left(n \ln(2\pi) + \ln |\mathbf{V}(\gamma)| + (\mathbf{y} - \mathbf{X}\beta)^T \mathbf{V}(\gamma)^{-1} (\mathbf{y} - \mathbf{X}\beta) \right)$$

In the simple one way ANOVA with random block effect model is

$$\gamma = (\sigma^2, \sigma_b^2)^T \text{ and } \beta = (\alpha_1, \alpha_2)^T$$

The maximum likelihood estimation is given by

$$(\hat{\beta}^{(ML)}, \hat{\gamma}^{(ML)}) = \operatorname{argmin}_{(\beta, \gamma)} l(\mathbf{y}, \beta, \gamma)$$

Assume γ is known then ML estimation of $\hat{\beta}(\gamma)$ is given by weighted least squares estimation

$$\hat{\beta}^{ML}(\gamma) = \operatorname{argmin}_{(\beta)} (\mathbf{y} - \mathbf{X}\beta)^T \mathbf{V}(\gamma)^{-1} (\mathbf{y} - \mathbf{X}\beta),$$

by differentiate and equal to zero we obtain:

$$\hat{\beta}^{ML}(\gamma) = (\mathbf{X}^T \mathbf{V}(\gamma)^{-1} \mathbf{X})^{-1} \mathbf{X}^T \mathbf{V}(\gamma)^{-1} \mathbf{y}$$

Theory of Mixed Models

The restricted likelihood method for mixed effect models:

Since the maximum likelihood estimation is biased, we want to modify it to obtain unbiased estimator.

Idea of Restricted (residual) maximum likelihood (REML) is in linear transform of data which eliminates mean.

REML method is given by modification of classical ML function by

$$l_{REML}(\mathbf{y}, \beta, \gamma) = \frac{1}{2} \left(n \ln(2\pi) + \ln |\mathbf{V}(\gamma)| + (\mathbf{y} - \mathbf{X}\beta)^T \mathbf{V}^{-1}(\gamma) (\mathbf{y} - \mathbf{X}\beta) + \ln |\mathbf{X}^T \mathbf{V}^{-1}(\gamma) \mathbf{X}| \right).$$

This REML method gives (at least in balanced case) the unbiased estimates and is generally preferred in mixed models.

Theory of Mixed Models

Restricted maximum likelihood (REML)

- ▶ The default parameter estimation criterion for linear mixed models in `lme` and `lmer` functions is REML.
- ▶ Maximum likelihood (ML) estimates (sometimes called full maximum likelihood) can be requested by specifying `REML=FALSE`.
- ▶ Generally REML estimates of variance components are preferred. ML estimates are known to be biased. Although REML estimates are not guaranteed to be unbiased, they are usually less biased than ML estimates.

For more info and REML proof see:

- ▶ **lmer vignettes**
- ▶ **proof by Kevin Liu**

Example of Mixed effect model analysis

Drying of beech wood planks

To investigate the effect of drying of beech wood on the humidity percentage, the following experiment was conducted. Each of 20 planks was dried in a certain period of time. Then the humidity percentage was measured in 5 depths (1,3,5,7,9) and 3 widths (1,2,3) for each plank.

Variables:

- ▶ plank - Numbered 1-20
- ▶ width - Numbered 1,2,3
- ▶ depth - Numbered 1,3,5,7,9
- ▶ humidity - Humidity percentage

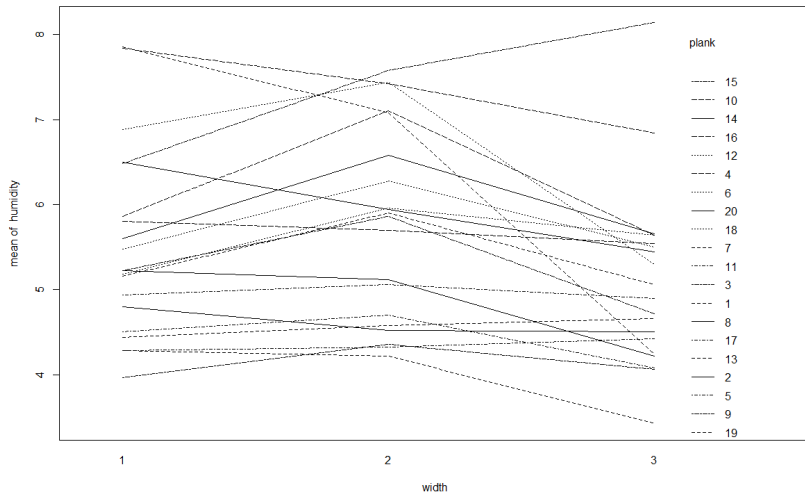
Source: The Royal Veterinary and Agricultural University, Denmark.

Number of observations: 300 (20 planks)

depth 1: close to the top
depth 5: in the center
depth 9: close to the bottom
depth 3: between 1 and 5
depth 7: between 5 and 9
width 1: close to the side
width 3: in the center
width 2: between 1 and 3

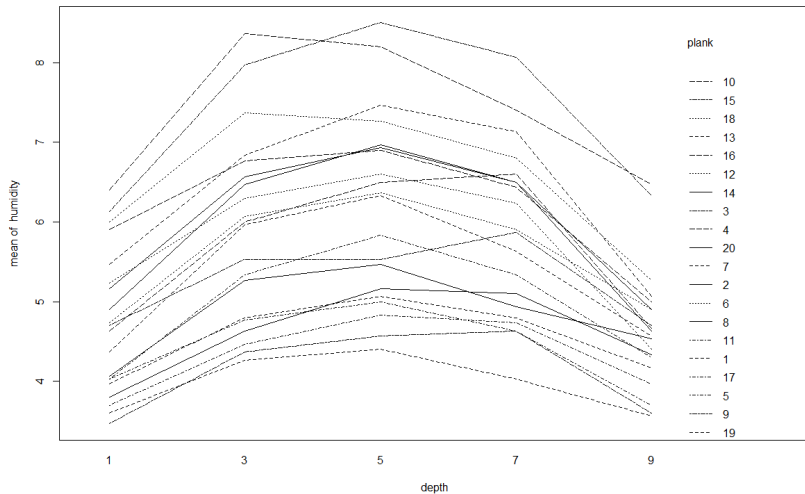
Exercise - Drying of beech wood planks

```
with(planks, interaction.plot(width,plank,humidity,legend=T))
```



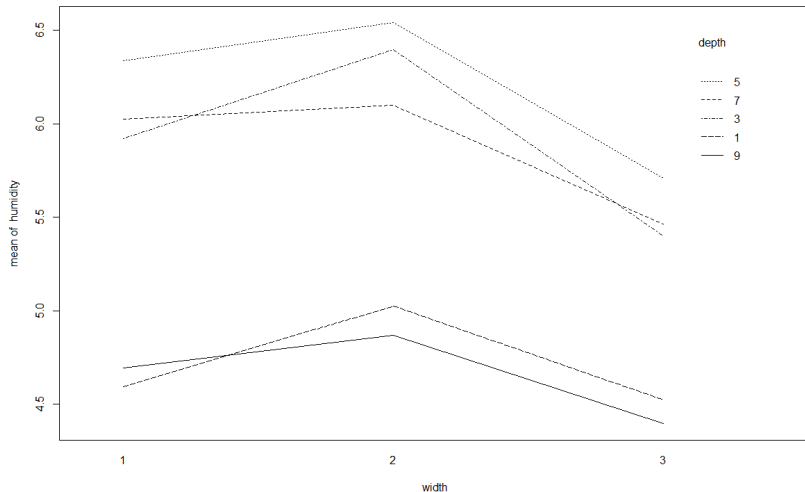
Exercise - Drying of beech wood planks

```
with(planks, interaction.plot(depth,plank,humidity,legend=T))
```



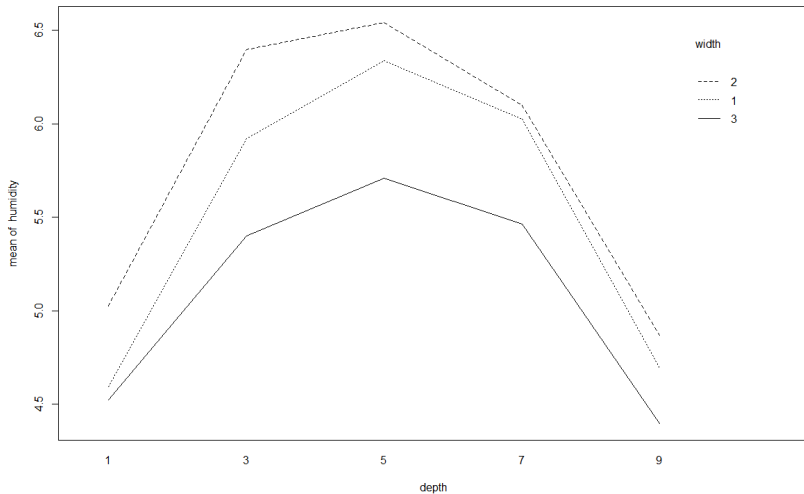
Exercise - Drying of beech wood planks

```
with(planks, interaction.plot(width,depth,humidity,legend=T))
```



Exercise - Drying of beech wood planks

```
with(planks, interaction.plot(depth,width,humidity,legend=T))
```



Exercise

Analyze data from the **Drying of beech wood planks** experiment.

- ▶ Plot four average humidity profiles:
2 interaction plots for width and 2 for depth (done).
- ▶ Carrying out the fixed effects model analysis.
- ▶ Carry out the mixed model analysis.
- ▶ Run the post hoc analysis
- ▶ Compare the fixed parameters and use the p-value correction (TukeyHSD).
Hint: Use function `lsmeans` from the package `lsmeans` with `adjust="tukey"`.
- ▶ Summarize results.