1. (a) For any  $\epsilon > 0$  and  $n \in \mathbb{Z}$ , we have

$$P(|Z_n| > \epsilon^{-1/2}) \le \operatorname{Var}(Z_n)\epsilon = \epsilon,$$

by Chebyshev inequality. By definition,  $Z_n = O_p(1)$ .

Notice that  $X_n = m + \sigma_n Z_n = m + \sigma_n O_p(1)$  and  $\sigma_n^2 \to 0$ , then  $X_n = m + o_p(1)$ .

(b) By Taylor expansion, we have

$$f(X_n) = f(m) + f'(m)(X_n - m) + o_p(|X_n - m|).$$

Then,  $Y_n - Z_n = (f'(m)\sigma_n)^{-1}(o_p(|X_n - m|)) = o_p(1)$  by the fact  $X_n = m + o_p(1)$  and  $\sigma_n^2 \to 0$ .

- (c) Suppose there exists a random variable Z such that  $Z_n \stackrel{P}{\to} Z$ . By the definition,  $o_p(1) \stackrel{P}{\to} 0$ . Then  $Y_n = Z_n + o_p(1) \stackrel{p}{\to} Z$ . Suppose there exists a random variable Z such that  $Z_n \stackrel{d}{\to} Z$ .  $Y_n = Z_n + o_p(1) \stackrel{d}{\to} Z$  by the Slutsky Theorem.
- (d) By the Central Limit Theorem , we have

$$\sqrt{n}(\frac{S_n}{n}-p) \stackrel{d}{\to} N(0,p(1-p)).$$

By (c),

$$f(\frac{S_n}{n}) \stackrel{d}{\to} N\left(f(p), \frac{p(1-p)(f'(p))^2}{n}\right).$$

- 2. By the Slutsky Theorem,  $X_n \stackrel{d}{\to} \mu$ . Since  $\mu$  is a constant,  $X_n \stackrel{P}{\to} \mu$ .
- 3. (a) Sufficiency: By the condition, we have

$$\frac{X_n - \mu_n}{\sigma_n} = \frac{X_n - \tilde{\mu}_n}{\tilde{\sigma}_n} \frac{\tilde{\sigma}_n}{\sigma_n} + \frac{\tilde{\mu}_n - \mu_n}{\sigma_n}.$$

It is easy to proof that  $N(0,1) \stackrel{d}{=} N(\frac{\tilde{\mu}_n - \mu_n}{\sigma_n}, \left(\frac{\tilde{\sigma}_n}{\sigma_n}\right)^2)$ . Thus,  $\frac{\tilde{\sigma}_n}{\sigma_n} \to 1$  and  $\frac{\tilde{\mu}_n - \mu_n}{\sigma_n} \to 0$ .

Necessity: The result follows the Slutsky Theorem and the equation

$$\frac{X_n - \mu_n}{\sigma_n} = \frac{X_n - \tilde{\mu}_n}{\tilde{\sigma}_n} \frac{\tilde{\sigma}_n}{\sigma_n} + \frac{\tilde{\mu}_n - \mu_n}{\sigma_n}.$$

- (b) The proof is similar to (a).
- (c) The former result follows (b). It is easy to proof that  $n^{-1/2}X_n$  is  $\operatorname{An}(n^{1/2},2)$ .
- 4. Wrong (Deduct 1 point): If  $n \to +\infty$ , (a)  $\bar{X}_n^2 \stackrel{d}{\to} \mu^2$ ; (b)  $\bar{X}_n^2 \stackrel{d}{\to} 0$ . Denote  $\bar{X}_n = n^{-1} \sum_{j=1}^n X_j$ , we have  $\bar{X}_n \sim N(\mu, \sigma^2/n) \stackrel{d}{\to} \mu$ . By the Continuous Mapping Theorem, (a)  $\bar{X}_n^2 \sim \frac{\sigma^2}{n} \chi_1^2(\lambda)$ , where  $\lambda = \mu^2$  is the location parameter; (b)  $\bar{X}_n^2 \sim \frac{\sigma^2}{n} \chi_1^2$ .

Or using delta method to obtain the convergent form of the normal distribution. (a)  $\bar{X}_n^2 \sim N(\mu^2, 4\mu\sigma^2/n)$ ; (b)  $\bar{X}_n^2 \sim \frac{\sigma^2}{n}\chi_1^2$ . Moment method or Characteristic function method also works.