

# Bayesian Hierarchical Modeling in JAGS

Marcel Niklaus

10. Juli 2015

# Not content of this talk

- Bayesian Hypothesis test
- Model comparison

# Why Bayes

- We are actually interested in the probability of a model given data
- p-values are based on probability of (unobserved) data given model (conceptually hard)

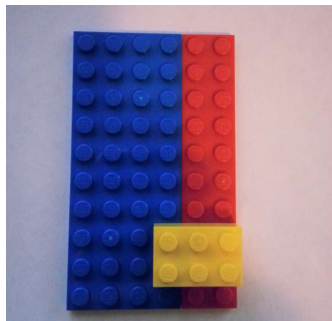
# Why Bayes

- We are actually interested in the probability of a model given data
- p-values are based on probability of (unobserved) data given model (conceptually hard)
- Surprise exercise; think of a situation you are interested in data given model!

# General Principles of Bayesian Analysis

- Uncertainty (of parameter estimate) is quantified by probability (distributions)
- Observed data is used to update the **prior** information to yield **posterior** information
- Bayesian workflow: set prior beliefs  $\Rightarrow$  get data  $\Rightarrow$  summarize posterior beliefs

# Formal Bayes Theorem



- Bayes rule is rooted in conditional probability and is uncontroversial.
- $P(\text{Red}|\text{Yellow}) = P(\text{Yellow}|\text{Red}) * P(\text{Red})/P(\text{Yellow})$
- $P(\text{Red}|\text{Yellow}) = (4/20) * (20/60)/(6/60)$
- $2/3 = 1/5 * 1/3 * 10$

# Formal Bayes Theorem (only slide with many formulas)

- Bayes Theorem tells us how to rationally revise prior beliefs in light of the data to yield posterior beliefs.
- $P(R|Y) = P(Y|R) * \frac{P(R)}{P(Y)}$
- Yellow = data, red = hypothesis  $\theta$
- $p(\theta|data) = \frac{p(data|\theta)*p(\theta)}{p(data)}$
- $posterior = \frac{likelihood * prior}{marginal\ likelihood}$
- $p(\theta|D) \propto p(D|\theta) * p(\theta)$
- Posterior is the likelihood weighted by the prior

# Let's play a game: Rock Paper Scissor

`http://87.106.45.173:  
3838/felix/BayesLessons/BayesianLesson1.Rmd`

- 6 Trials of Rock Paper Scissor
- Let's ignore the ties!



Pick your Prior: Who is going to win?

- $\theta$ : the rate with which the female wins
- $\theta = 1$ : female always wins
- $\theta = 0$ : male always wins
- $\theta = 0.5$ : equal odds
- Uncertainty (of parameter estimate) is quantified by probability (distributions)
- We assume our beliefs can be represented by a beta distribution
- A beta distribution is constraint to lie within 0-1 (perfect for proportion)

Pick your Prior: Who is going to win?

- $\theta \sim \text{Beta}(1, 1)$ : Uninformed prior: Do you think (before seeing the data) that it is equally likely that  $\theta = 0.01$  and  $\theta = 0.5$ ?
- $\theta \sim \text{Beta}(3, 3)$  : Slightly informed prior

We are now ready to play: Go!

# Likelihood

The likelihood is the workhorse of Bayesian inference. It represents the data part.

- What's the likelihood we observe the data ( $y$  wins given  $n$  trials) given the parameter  $\theta$
- Probability density of the data, considered as a function of  $\theta$
- The likelihood of a hypothesis ( $\theta$ ),  $L(H|data)$ , conditions on the data as if they are fixed while allowing the hypotheses (our parameter  $\theta$  to vary
- Binomial likelihood:  $L(p|n, y) = \binom{n}{y} p^y (1 - p)^{n-y}$
- [http://shiny.stat.calpoly.edu/MLE\\_Binomial/](http://shiny.stat.calpoly.edu/MLE_Binomial/)

After playing

- The posterior distribution summarizes our state of uncertainty about the true value of  $\theta$  after having observed the game.

# Markov chain Monte Carlo

- Posterior distribution = Prior distribution \* Likelihood density
- Analytical calculation of posterior only possible for simple models (high-dimensionality integration problem)
- Draw samples from posterior and summarize the distribution of those samples (it works!)
- Metropolis-Hastings algorithm and the Gibbs sampler
- JAGS WinBugs and STAN can do this for you
- Algorithm to approximate an unknown distribution

- $\leftarrow$  is equal to:  $y \leftarrow a + b$
- $\sim$  is distributed as..
- Binomial likelihood:  $y \sim \text{dbin}(\theta, nAttempts)$
- Normal likelihood:  $y \sim \text{dnorm}(mean, precision)$
- $precision = 1/Var$
- Loop: `for (i in 1:3){bla[i]=1+i}`
- `bla = 2,3,4, bla[2] = 3`

# Soccer Shootout

Let's estimate soccer players' ability to score a penalty in a world cup!

- Open Soccer.R in your R console
- Set your working directory (line 8)
- Jags model: ShootoutAbility.txt
- What is their ability  $\theta$ , [0-1]?
- $Y|\theta$  Prior: remember the game

Convergence: Has the MCMC Gibbs sampler converged on posterior (after starting from random values)

- Trace plot: "tight horizontal band"
- Autocorrelation plot: "Chains should forget previous visits with time": drop off quickly
- Gelman-Rubin-Brooks diagnostic: Between and within chain variance: 1 indicated convergence (F-value); less than 1.2. See `gelman.diag` [CODA]
- Geweke test of non-stationarity; Heidelberger-Welch test etc.



# Soccer Shootout: Afrika vs. Europa and America

- Jags model: ShootoutAbilityDifference.txt
- $\theta[\text{Cindex}[i]]$  can either be  $\theta[1]$ , or  $\theta[2]$

# Hierarchical Modeling

- With nested data (e.g. data for participants is organized on more than 1 level), a multilevel model is appropriate.
- The classic: students grouped in classes, which nest in school districts, which in turn nest in states.
- All Bayesian Models are hierarchical because every parameter has a prior.

# Easy IQ example: IQ measurements

- Scenario: We measure some IQs
- $y \sim N(\mu, \tau)$
- $\mu \sim N(100, 0.004)$
- $\tau \sim G(0.001, 0.001)$

)tau  $\sim G(0.001, 0.001)$

# Linear Mixed Models

- Linear Regression model
- Mixed because it includes coefficients that vary over group (random: participants, items) and some that don't (fixed, treatment group).
- Random Effects: zero mean restriction  $N(0, \omega)$
- Random intercepts model: random intercept, fixed slope
- Random intercepts and slope model: random intercept, random slope

In Bayesian Statistics, all parameters are random, i.e. drawn from an overarching distribution!

## Easy IQ example: IQ regressed on math grade

- $y \sim N(\mu, \tau)$
- $\mu \sim x_0 + x_1 * \text{math}$

## Easy IQ example: IQ regressed on math grade; Random Intercept

- $y \sim N(\mu[i], \tau)$
- $\mu[i] \sim -x0[G] + x1 * math$
- for all groups  $j$ :  $x0[j] \sim dnorm(100, 0.0044)$
- $x1 \sim dnorm(0, 0.001)$

## Easy IQ example: IQ regressed on math grade; Random Intercept and Slope

- $y \sim N(\mu, \tau)$
- $\mu = -x0[G] + x1[G] * math$
- for all groups  $j$ :  $x0[j] \sim dnorm(100, 0.0044)$
- for all groups  $j$ :  $x1[j] \sim dnorm(0, 0.01)$



# Simple linear regression

Let's estimate soccer player's ability to score a penalty in a world cup!

- Open Exam\_ple.R in your R console
- Set your working directory (line 7)
- Jags model: ExamSimple.txt

# Random intercept model

- $b0[school[i]]$
- for all schools:  $b0[j] \sim dnorm(school.b0, school.tau)$
- $school.b0 \sim dnorm(0, .0001)$

# Random intercept and slope model

- $b0[school[i]] + b1[school[i]] * lrt[i]$
- for all schools:  $b0[j] \sim dnorm(school.b0, school.tau)$
- for all schools:  $b1[j] \sim dnorm(school.b1, school.tau.b1)$
- $school.b0 \sim dnorm(0,.0001)$
- $school.b1 \sim dnorm(0,.0001)$

# Multi-Level Models

Covariates that capture the way groups vary are included

- Level 1: Regression for London Reading Test
  - Level 2: Group Level with covariates: Regression for entry score
- Ein