

5. $E: 0 \neq \#E < +\infty$, $E^N = \{\omega = (\omega_i)_{i \in \mathbb{N}} : \omega_i \in E, \forall i \in \mathbb{N}\}$

$$[\omega_1, \dots, \omega_m] := \{\omega' \in E^N : \omega'_i = \omega_i, \forall i = 1, \dots, m\} \quad \mu([\omega_1, \dots, \omega_m]) = \prod_{i=1}^m p_{\omega_i}$$

$P = (p_e)_{e \in E}$ - PROBABILITY VECTOR $\Leftrightarrow \prod_{e \in E} p_e = 1$

Product Measure
Bernoulli Measure

$$E = \{H, T\} \quad (H = \text{HEAD}, T = \text{TAIL})$$

$$\Omega = E^N \quad \Omega' := \{\omega \in \Omega : \exists i, j, i \neq j, \omega_i = \omega_j = H\}$$

$$\mathcal{A} = \sigma(\bigcup_{m=0}^{+\infty} A_m) \quad \mathcal{A}' := \mathcal{A} \mid \Omega'$$

$$p_H = p_T = \frac{1}{2} \Rightarrow \mu_\Omega$$

$$\Omega_m := \{\omega \in \{H, T\}^m : \exists i, j, i \neq j, \omega_i = \omega_j = H\}, m \geq 2$$

$$\Omega := \bigcup_{m=2}^{+\infty} \Omega_m \quad P(\Omega_2) = P(\{(H, H)\}) = \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{2^2}$$

$$P(\Omega_3) = P(\{(T, H, H), (H, T, H)\}) = \frac{2}{2^3}$$

$$P(\Omega_4) = P(\{(H, T, T, H), (T, H, T, H), (T, T, H, H)\}) = \frac{3}{2^4}$$

:

$$P(\Omega_m) = \frac{m-1}{2^m}$$

5. $E = \{H, T\} \quad (H = \text{HEAD}, T = \text{TAIL})$

$$\Omega_m := E^m = \{(\omega_1, \dots, \omega_m) : \omega_i \in E, \forall i = 1, \dots, m\}, m \geq 1, E^0 := \{\emptyset\}$$

$$\Omega = \bigcup_{m=0}^{+\infty} \Omega_m, \quad P(\{\omega\}) = \frac{1}{2^m} \Leftrightarrow \omega \in \Omega_m, m \geq 1$$

SAMPLE SPACE S :

$$\begin{aligned} \Omega'_2 &:= \{(H, H)\} & \rightarrow P(\Omega'_2) &= 1 \cdot \frac{1}{2^2} \\ \Omega'_3 &:= \{(T, H, H), (H, T, H)\} & \rightarrow P(\Omega'_3) &= 2 \cdot \frac{1}{2^3} \\ \Omega'_4 &:= \{(T, T, H, H), (T, H, T, H), (H, T, T, H)\} & \rightarrow P(\Omega'_4) &= 3 \cdot \frac{1}{2^4} \\ &\vdots &&\vdots \\ \Omega'_m &:= \{(T, \dots, T, H, H), \dots, (H, T, \dots, T, H)\} & \rightarrow P(\Omega'_m) &= (m-1) \cdot \frac{1}{2^m} \end{aligned}$$

$$\text{SAMPLE SPACE } S := \bigcup_{m=2}^{+\infty} \Omega'_m$$

$$P(\Omega'_k) = \frac{k-1}{2^{k-1}}$$

$$6. \Omega = \{1, 2, \dots\}$$

SUPPOSE THERE IS A UNIFORM PROBABILITY MEASURE \underline{P} OVER $(\Omega, \mathcal{P}(\Omega))$ SUCH THAT IT SATISFY: $A, B \subseteq \Omega, |A|=|B| \Rightarrow \underline{P}(A)=\underline{P}(B)$.

$$\text{IF } \underline{P}(\{\infty\})=0 \rightarrow 1 = \underline{P}(\Omega) = \sum_{m=1}^{+\infty} \underline{P}(\{m\}) = 0 \quad (\Leftrightarrow)$$

$$\text{IF } \underline{P}(\{\infty\})>\delta>0 \rightarrow 1 = \underline{P}(\Omega) = \sum_{m=1}^{+\infty} \underline{P}(\{m\}) > \lim_{m \rightarrow +\infty} \left[\sum_{i=1}^m \delta \right] = +\infty \quad (\Leftrightarrow)$$

$$7. A_1, A_2, \dots \in \mathcal{A} \text{ (EVENTS)}$$

$$B_m := A_m \setminus \bigcup_{k=1}^{m-1} A_k, m \geq 1 \quad \text{AND} \quad B_1 := A_1$$

By DEFINITION: IF $x \in B_m$, THEN $x \notin B_1, \dots, B_{m-1}$. Thus, IF $x \in B_m \cap B_m$ AND $m < m'$, $x \in B_m \rightarrow x \notin B_{m'} \rightarrow x \notin B_m$ (\Leftrightarrow). THEN $B_i \cap B_j = \emptyset \forall i, j, i \neq j$.

$$\star \bigcup_{i=1}^m A_i = \bigcup_{i=1}^m B_i \quad \text{OR} \quad \bigcup_{i=1}^{+\infty} A_i = \bigcup_{i=1}^{+\infty} B_i$$

$$(?) x \in \bigcup_{i=1}^m B_i \rightarrow \exists j: x \in B_j \rightarrow x \in A_j \rightarrow x \in \bigcup_{i=1}^m A_i$$

$$(\subseteq) x \in \bigcup_{i=1}^m A_i \rightarrow i_0 := \min \{i \in \{1, \dots, m\}: x \in A_i\} \rightarrow x \in A_{i_0} \text{ AND } x \notin A_j: j=1, \dots, i_0-1 \\ \rightarrow x \in B_{j_0} \rightarrow x \in \bigcup_{i=1}^m B_i$$

$$\begin{aligned} \underline{P}\left(\bigcup_{m=1}^{+\infty} A_i\right) &= \underline{P}\left(\bigcup_{m=1}^{+\infty} B_m\right) = \sum_{m=1}^{+\infty} \underline{P}(B_m) = \lim_{m \rightarrow +\infty} \left[\sum_{i=1}^m \underline{P}(B_i) \right] = \lim_{m \rightarrow +\infty} \left[\sum_{i=1}^m \left(\underline{P}(A_i) - \underline{P}\left(\bigcup_{k=1}^{i-1} A_k\right) \right) \right] \\ &= \lim_{m \rightarrow +\infty} \left[\sum_{i=1}^m \underline{P}(A_i) - \sum_{i=2}^m \underline{P}\left(\bigcup_{k=1}^{i-1} A_k\right) \right] \leq \lim_{m \rightarrow +\infty} \left[\sum_{i=1}^m \underline{P}(A_i) \right] = \sum_{i=1}^{+\infty} \underline{P}(A_i) \end{aligned}$$

$$8. \underline{P}(A_i) = 1, \forall i \in \mathbb{N} \rightarrow \underline{P}(A_i^c) = 0, \forall i \in \mathbb{N}$$

$$\underline{P}\left(\bigcup_{m=1}^{+\infty} A_i^c\right) \leq \sum_{i=1}^{+\infty} \underline{P}(A_i^c) = 0 \quad \Rightarrow \quad \underline{P}\left(\bigcup_{m=1}^{+\infty} A_i^c\right) = 0$$

↳ EXERCISE 7

$$\Rightarrow 1 = \underline{P}\left(\left(\bigcup_{m=1}^{+\infty} A_i^c\right)^c\right) = \underline{P}\left(\bigcap_{m=1}^{+\infty} A_i\right)$$

9. $B \in \mathcal{A}$, $\underline{P}(B) > 0 \Rightarrow \mu := \underline{P}(\cdot | B)$ IS A PROBABILITY MEASURE

$$\begin{aligned} \text{(I)} \quad \mu(A) &= \underline{P}(A|B) = \frac{\underline{P}(A \cap B)}{\underline{P}(B)} \geq 0 \quad \rightarrow \mu\left(\bigcup_{m=1}^{+\infty} A_m\right) = \underline{P}\left(\bigcup_{m=1}^{+\infty} A_m | B\right) = \\ \text{(II)} \quad \mu(\Omega) &= \underline{P}(\Omega | B) = \frac{\underline{P}(B)}{\underline{P}(B)} = 1 \quad = \sum_{i=1}^{+\infty} \frac{\underline{P}(A_m \cap B)}{\underline{P}(B)} = \sum_{i=1}^{+\infty} \frac{\underline{P}(A_m \cap B)}{\underline{P}(B)} \\ \text{(III)} \quad A_1, A_2, \dots &\text{ DISJOINT EVENTS} \quad = \sum_{i=1}^{+\infty} \mu(A_m) \end{aligned}$$

10. MONTY HALL PROBLEM (3 DOORS)

You ALWAYS PICK DOOR 1

$$\Omega = \{(\omega_1, \omega_2) : \omega_1, \omega_2 \in \{1, 2, 3\}\}, \quad \begin{array}{l} \omega_1 - \text{WHERE THE PRIZE IS} \\ \omega_2 - \text{IS THE DOOR MONTY OPENS} \end{array}$$

SAMPLE SPACE

$$\Omega' := \left\{ \begin{array}{l} (1, 2), (1, 3) \\ (2, 3) \\ (3, 2) \end{array} \right\} \quad \begin{array}{l} \underline{P}(\{(1, 2)\}) = \underline{P}(\{(1, 3)\}) = \frac{1}{3} \cdot \frac{1}{2} \\ \underline{P}(\{(2, 3)\}) = \frac{1}{3} \\ \underline{P}(\{(3, 2)\}) = \frac{1}{3} \end{array}$$

* You CONTINUE WITH DOOR 1

$$\underline{P}(\omega_1=1 | \omega_2=2) = \frac{\underline{P}(\omega_1=1, \omega_2=2)}{\underline{P}(\omega_2=2)} = \frac{\underline{P}(\{(1, 2)\})}{\underline{P}(\{(1, 2), (3, 2)\})} = \frac{\frac{1}{3} \cdot \frac{1}{2}}{\frac{1}{3} \cdot \frac{1}{2} + \frac{1}{3}} = \frac{1}{3}$$

$$\underline{P}(\omega_1=1 | \omega_2=3) = \frac{\underline{P}(\omega_1=1, \omega_2=3)}{\underline{P}(\omega_2=3)} = \frac{\underline{P}(\{(1, 3)\})}{\underline{P}(\{(1, 3), (2, 3)\})} = \frac{\frac{1}{3} \cdot \frac{1}{2}}{\frac{1}{3} \cdot \frac{1}{2} + \frac{1}{3}} = \frac{1}{3}$$

* You CHANGE THE DOOR

$$\underline{P}(\omega_1=2 | \omega_2=3) = \frac{\underline{P}(\omega_1=2, \omega_2=3)}{\underline{P}(\omega_2=3)} = \frac{\underline{P}(\{(2, 3)\})}{\underline{P}(\{(1, 3), (2, 3)\})} = \frac{\frac{1}{3}}{\frac{1}{3} \cdot \frac{1}{2} + \frac{1}{3}} = \frac{2}{3}$$

$$\underline{P}(\omega_1=3 | \omega_2=2) = \frac{\underline{P}(\omega_1=3, \omega_2=2)}{\underline{P}(\omega_2=2)} = \frac{\underline{P}(\{(3, 2)\})}{\underline{P}(\{(1, 2), (3, 2)\})} = \frac{\frac{1}{3}}{\frac{1}{3} \cdot \frac{1}{2} + \frac{1}{3}} = \frac{2}{3}$$

MONTY HALL PROBLEM (4 DOORS)

SAMPLE SPACE

$$\Omega' := \left\{ \begin{array}{l} (1,2), (1,3), (1,4) \\ (2,3), (2,4) \\ (3,2), (3,4) \\ (4,2), (4,3) \end{array} \right\} \rightsquigarrow \begin{aligned} P(\{\{1,2\}\}) &= P(\{\{1,3\}\}) = P(\{\{1,4\}\}) = \frac{1}{4} \cdot \frac{1}{3} \\ P(\{\{2,3\}\}) &= P(\{\{2,4\}\}) = \frac{1}{4} \cdot \frac{1}{2} \\ P(\{\{3,2\}\}) &= P(\{\{3,4\}\}) = \frac{1}{4} \cdot \frac{1}{2} \\ P(\{\{4,2\}\}) &= P(\{\{4,3\}\}) = \frac{1}{4} \cdot \frac{1}{2} \end{aligned}$$

* You CONTINUE WITH DOOR 1

$$P(w_1=1 | w_2=2) = \frac{P(w_1=1, w_2=2)}{P(w_2=2)} = \frac{P(\{\{1,2\}\})}{P(\{\{1,2\}, \{3,2\}, \{4,2\}\})} = \begin{pmatrix} \text{OTHER} \\ \text{CASES ARE} \\ \text{THE SAME} \end{pmatrix}$$

$$= \frac{\frac{1}{4} \cdot \frac{1}{3}}{\frac{1}{4} \cdot \frac{1}{3} + \frac{1}{4} \cdot \frac{1}{2} + \frac{1}{4} \cdot \frac{1}{2}} = \frac{\frac{1}{3}}{\frac{1}{3} + 1} = \frac{\frac{1}{3}}{\frac{4}{3}} = \frac{3}{4} \cdot \frac{1}{3} = \frac{1}{4} = 0.25$$

* You CHANGE THE DOOR

$$P(w_1=2 | w_2=3) = \frac{P(w_1=2, w_2=3)}{P(w_2=3)} = \frac{P(\{\{2,3\}\})}{P(\{\{1,3\}, \{2,3\}, \{4,3\}\})} =$$

$$= \frac{\frac{1}{4} \cdot \frac{1}{2}}{\frac{1}{4} \cdot \frac{1}{3} + \frac{1}{4} \cdot \frac{1}{2} + \frac{1}{4} \cdot \frac{1}{2}} = \frac{\frac{1}{2}}{\frac{1}{3} + 1} = \frac{\frac{1}{2}}{\frac{4}{3}} = \frac{3}{4} \cdot \frac{1}{2} = \frac{3}{8} = 0.375$$

MONTY HALL PROBLEM (M DOORS)

$$\Omega' = \left\{ \begin{array}{l} (1,2), (1,3), \dots, (1,m-1), (1,m) \\ (2,3), (2,4), \dots, (2,m) \\ \vdots \\ (m,2), (m,3), \dots, (m,m-1) \end{array} \right\} \begin{aligned} P(\{\{1,i\}\}) &= \frac{1}{m} \cdot \frac{1}{m-1}, \quad i=2,3,\dots,m \\ P(\{\{j,k\}\}) &= \frac{1}{m} \cdot \frac{1}{m-2}, \quad j \neq k, \quad j,k=2,3,\dots,m \end{aligned}$$

$$P(w_1=1 | w_2=2) = \frac{P(w_1=1, w_2=2)}{P(w_2=2)} = \frac{P(\{\{1,2\}\})}{P(\{\{1,2\}, \{3,2\}, \dots, \{m,2\}\})} = \frac{\frac{1}{m} \cdot \frac{1}{m-1}}{\frac{1}{m} \cdot \frac{1}{m-1} + (m-2) \cdot \frac{1}{m} \cdot \frac{1}{m-2}}$$

$$= \frac{1}{1 + (m-1)} = \frac{1}{m}$$

$$P(w_1=2 | w_2=3) = \frac{P(w_1=2, w_2=3)}{P(w_2=3)} = \frac{P(\{\{2,3\}\})}{P(\{\{1,3\}, \{2,3\}, \{4,3\}, \dots, \{m,3\}\})} = \frac{\frac{1}{m} \cdot \frac{1}{m-2}}{\frac{1}{m} \cdot \frac{1}{m-1} + (m-2) \cdot \frac{1}{m} \cdot \frac{1}{m-2}}$$

$$= \frac{\frac{1}{m-2}}{\frac{1}{m-1} + 1} = \frac{\frac{m-1}{m-2}}{1 + (m-1)} = \frac{m-1}{m(m-2)}$$

$\frac{m-1}{m-2} > 1 \quad \text{and} \quad \frac{1}{m} < \frac{1}{m} \cdot \frac{m-1}{m-2}$

11. (Ω, A, \mathbb{P}) $A, B \in \mathcal{A}$ ARE INDEPENDENTS

LET'S SHOW THAT A^c, B ARE INDEPENDENTS

$$\mathbb{P}(B) = \mathbb{P}(B \cap (A \cup A^c)) = \mathbb{P}(B \cap A) + \mathbb{P}(B \cap A^c) = \mathbb{P}(B) \cdot \mathbb{P}(A) + \mathbb{P}(B \cap A^c)$$

$$\begin{aligned} \hookrightarrow \mathbb{P}(B) - \mathbb{P}(B) \cdot \mathbb{P}(A) &= \mathbb{P}(B \cap A^c) \rightarrow \mathbb{P}(B) \cdot (1 - \mathbb{P}(A)) = \mathbb{P}(B \cap A^c) \\ \rightarrow \mathbb{P}(B) \cdot \mathbb{P}(A^c) &= \mathbb{P}(B \cap A^c) \end{aligned}$$

SAME ARGUMENTS WE CONCLUDE THAT A^c, B^c ARE INDEPENDENTS.

12.

FRONT BACK

$$\mathbb{P}(\{(C_1, G, G)\}) = \frac{1}{3}$$

C_1

G	6
---	---

$$\mathbb{P}(\{(C_2, R, R)\}) = \frac{1}{3}$$

C_2

R	R
---	---

$$\mathbb{P}(\{(C_3, G, R)\}) = \mathbb{P}(\{(C_3, R, G)\}) = \frac{1}{3} \cdot \frac{1}{2}$$

C_3

G	R
---	---

R	G
---	---

$$\begin{aligned} \mathbb{P}(\text{BACK} = G \mid \text{FRONT} = G) &= \frac{\mathbb{P}(\text{FRONT} = G, \text{BACK} = G)}{\mathbb{P}(\text{BACK} = G)} = \frac{\mathbb{P}(\{(C_1, G, G)\})}{\mathbb{P}(\{(C_1, G, G)\}) + \mathbb{P}(\{(C_3, G, R)\})} = \\ &= \frac{\frac{1}{3}}{\frac{1}{3} + \frac{1}{3} \cdot \frac{1}{2}} = \frac{1}{1 + \frac{1}{2}} = \frac{2}{3} \end{aligned}$$

13.

(a) SAMPLE SPACE

$$\Omega'_2 := \{(H, T), (T, H)\}$$

$$\Omega'_3 := \{(H, H, T), (T, T, H)\}$$

$$\Omega'_4 := \{(H, H, H, T), (T, T, T, H)\}$$

:

$$\Omega'_m := \{(H, \dots, H, T), (T, \dots, T, H)\}$$

$$\rightsquigarrow \Omega' := \bigcup_{m=2}^{+\infty} \Omega'_m \quad \text{SAMPLE SPACE}$$

(b) PROBABILITY

$$\rightsquigarrow \mathbb{P}(\Omega'_2) = 2 \cdot \frac{1}{2^2}$$

$$\rightsquigarrow \mathbb{P}(\Omega'_3) = 2 \cdot \frac{1}{2^3}$$

$$\rightsquigarrow \mathbb{P}(\Omega'_4) = 2 \cdot \frac{1}{2^4}$$

$$\rightsquigarrow \mathbb{P}(\Omega'_m) = 2 \cdot \frac{1}{2^m} = \frac{1}{2^{m-1}}$$

$$\mathbb{P}(\Omega'_3) = \frac{1}{2^2}$$

14. $P(A) \in \{0, 1\} \rightarrow A$ IS INDEPENDENT OF EVERY OTHER EVENT

$E \in \mathcal{A}$ - EVEN SPACE

IF $P(A) = 0$, THEN $P(A \cap E) \leq P(A) = 0 \rightarrow P(A \cap E) = 0 = 0 \cdot P(E) = P(A) \cdot P(E)$ (i)

IF $P(A) = 1$, THEN $P(A^c) = 0$ AND

$$\begin{aligned} P(E) \cdot P(A) &= P(E) \cdot 1 = P(E) = P(E \cap (A \cup A^c)) = P(E \cap A) + P(E \cap A^c) \\ &\stackrel{(i)}{=} P(E \cap A) + 0 = P(E \cap A) \end{aligned}$$

15. BE = "BLUE EYES" NBE = "NOT BLUE EYES" / $P(BE) = \frac{1}{4}$ $P(NBE) = \frac{3}{4}$

(a) $\Omega = \{(x, y, z) : x, y, z \in \{BE, NBE\}\}$ SAMPLE SPACE

$\mathcal{A} = \mathcal{P}(\Omega) = \{X : X \subseteq \Omega\}$ EVENT SPACE

X = NUMBER OF CHILDREN WITH BLUE EYES

$$P(X \geq 2 | X \geq 1) = ?$$

$$P(X \geq 1) = 1 - P(X=0) = 1 - \left(\frac{3}{4}\right)^3 = \frac{64 - 27}{64} = \frac{37}{64}$$

$$\begin{aligned} P(X \geq 2) &= P(X=2) + P(X=3) = 3 \cdot \left(\frac{1}{4}\right)^2 \cdot \left(\frac{3}{4}\right)^2 + \left(\frac{1}{4}\right)^3 = \left(\frac{1}{4}\right)^3 (3 \cdot 1 \cdot 3 + 1) \\ &= \left(\frac{1}{4}\right)^3 (9 + 1) = \frac{10}{64} \end{aligned}$$

$$P(X \geq 2 | X \geq 1) = \frac{P(X \geq 2, X \geq 1)}{P(X \geq 1)} = \frac{P(X \geq 2)}{P(X \geq 1)} = \frac{\frac{10}{64}}{\frac{37}{64}} = \frac{10}{37}$$

(b) $Z = \{(x, y, BE) : x, y \in \{BE, NBE\}\}$

$$\begin{aligned} P(X \geq 2 | Z) &= \frac{P(X \geq 2, Z)}{P(Z)} = \frac{P(\{(BE, NBE, BE), (NBE, BE, BE), (BE, BE, BE)\})}{P(Z)} \\ &= \frac{\left(\frac{1}{4}\right)^3 (2 \cdot 1^2 \cdot 3^2 + 1)}{\frac{1}{4}} = \frac{\frac{7}{4^2}}{\frac{1}{4}} = \frac{7}{16} \end{aligned}$$

16.

$\{A, B\}$ ARE INDEPENDENT EVENTS $\Leftrightarrow \underline{\underline{P}}(A \cap B) = \underline{\underline{P}}(A) \cap \underline{\underline{P}}(B)$

$\{A, B\}$ ARE INDEPENDENT EVENTS $\Rightarrow \underline{\underline{P}}(A|B) = \frac{\underline{\underline{P}}(A \cap B)}{\underline{\underline{P}}(B)} = \frac{\underline{\underline{P}}(A) \cdot \underline{\underline{P}}(B)}{\underline{\underline{P}}(B)} = \underline{\underline{P}}(A)$ ($\text{if } \underline{\underline{P}}(B) > 0$)

$$\begin{aligned} \underline{\underline{P}}(A|B) &= \frac{\underline{\underline{P}}(A \cap B)}{\underline{\underline{P}}(B)} \rightarrow \underline{\underline{P}}(A \cap B) = \underline{\underline{P}}(A|B) \cdot \underline{\underline{P}}(B) \\ \underline{\underline{P}}(B|A) &= \frac{\underline{\underline{P}}(B \cap A)}{\underline{\underline{P}}(A)} \rightarrow \underline{\underline{P}}(B \cap A) = \underline{\underline{P}}(B|A) \cdot \underline{\underline{P}}(A) \end{aligned} \quad \left. \begin{array}{l} \underline{\underline{P}}(A \cap B) = \underline{\underline{P}}(B \cap A) = \\ = \underline{\underline{P}}(A|B) \cdot \underline{\underline{P}}(B) \\ = \underline{\underline{P}}(B|A) \cdot \underline{\underline{P}}(A) \end{array} \right\}$$

17.

$$\underline{\underline{P}}(A \cap B \cap C) = \underline{\underline{P}}(A \cap (B \cap C)) \stackrel{16.}{=} \underline{\underline{P}}(A|B \cap C) \cdot \underline{\underline{P}}(B \cap C) \stackrel{16.}{=} \underline{\underline{P}}(A|B \cap C) \cdot \underline{\underline{P}}(B|C) \cdot \underline{\underline{P}}(C)$$

18.

LET'S PROVE BY CONTRADICTION. SUPPOSE $\underline{\underline{P}}(A_i|B) < \underline{\underline{P}}(A_i), \forall i = 1, \dots, k$, WE ALSO HAVE $\underline{\underline{P}}(A_1|B) < \underline{\underline{P}}(A_1)$ BY HYPOTHESIS.

$$\underline{\underline{P}}(A_i|B) < \underline{\underline{P}}(A_i), \forall i = 1, \dots, k \quad (*)$$

$$\begin{aligned} \rightsquigarrow \underline{\underline{P}}(B) &= \underline{\underline{P}}(B \cap (\bigcup_{i=1}^k A_i)) = \sum_{i=1}^k \underline{\underline{P}}(B \cap A_i) = \sum_{i=1}^k \underline{\underline{P}}(A_i|B) \cdot \underline{\underline{P}}(B) \\ &= \left(\sum_{i=1}^k \underline{\underline{P}}(A_i|B) \right) \cdot \underline{\underline{P}}(B) < \left(\sum_{i=1}^k \underline{\underline{P}}(A_i) \right) \cdot \underline{\underline{P}}(B) = \underline{\underline{P}}\left(\bigcup_{i=1}^k A_i\right) \cdot \underline{\underline{P}}(B) = 1 \cdot \underline{\underline{P}}(B) \end{aligned}$$

$$\rightsquigarrow \underline{\underline{P}}(B) < \underline{\underline{P}}(B) \quad (\rightarrow \leftarrow)$$

19.

MAC	WIN	LIN
30%	50%	20%

65% MAC GOT VIRUS	82% WIN GOT VIRUS	50% LIN GOT VIRUS
-------------------	-------------------	-------------------

$$\underline{\underline{P}}(\underline{\underline{P}}_{\text{C}} = \text{WIN} \mid \text{VIRUS} = \text{TRUE}) = ?$$

$$\begin{aligned} \underline{\underline{P}}(\underline{\underline{P}}_{\text{C}} = \text{WIN} \mid \text{VIRUS} = \text{TRUE}) &= \frac{\underline{\underline{P}}(\text{VIRUS} = \text{TRUE} \mid \underline{\underline{P}}_{\text{C}} = \text{WIN}) \cdot \underline{\underline{P}}(\underline{\underline{P}}_{\text{C}} = \text{WIN})}{\sum_{\text{PC}=\text{x}} \underline{\underline{P}}(\text{VIRUS} = \text{TRUE} \mid \underline{\underline{P}}_{\text{C}} = \text{x}) \cdot \underline{\underline{P}}(\underline{\underline{P}}_{\text{C}} = \text{x})} \\ &= \frac{(0.82) \cdot (0.5)}{(0.65) \cdot (0.3) + (0.82) \cdot (0.5) + (0.5) \cdot (0.2)} \simeq 0.5815 = 58.15\% \end{aligned}$$

