# A short primer on (G)LMMs

Marco D. Visser Sean McMahon Lisa Huelsmann

June 29, 2019

# Workshop schedule

- 1. Philosophy of multilevel modeling (Sean)
- 2. Computer lab: simple regression to multilevel models (Marco)
- 3. GLMM model diagnostics with DHARMa (Lisa)

#### Obligatory quote

- "The computational ease with which an abundance of parameters can be estimated should not be allowed to obscure the probable unwisdom of such estimation from limited data"
- Arthur P. Dempster in "Covariance selection", Biometrics 28 (1), 157-175 (March 1972)

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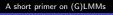
# This workshop about

#### This workshop is not about:

- 1. The math behinds mixed effect models
- 2. The technical details of model optimization

This workshop IS about how to fit mixed models in practice

Disclaimer: information given here is usually done in multiple semesters



This presentation, with all  ${\bf knitr}$  code examples is available at  ${\it github.com/MarcoDV}$  is ser/GLMMworkshop

# General regression modeling steps



# General regression modeling steps

- 1. Explore your data, make exploratory plots.
  - decide which model is appropriate to fit
  - distrust the model deeply
- 2. Evaluate the model fit
  - Check basic assumptions
  - Run model diagnostics
- 3. Is the model decent?
  - distrust the model slightly less
- 4. Does the model fail any test?
  - discard the model and start over

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#### **Tree Allometery**

Exercise 1: find an unbiased function for predicting crown area from diameter

# Tropical tree height and crown allometries for the Barro Colorado Nature Monument, Panama: a comparison of alternative hierarchical models incorporating interspecific variation in relation to life history traits

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#### Load the data

Step 1: load the data

```
allo <- read.csv("DiameterHeightCrownGLMMworkshop.csv")
traits <- read.csv("SpeciesTraits20190104.csv")</pre>
## look at the data
colnames(allo)
## [1] "X.2"
                               "X.1"
                                                     пХп
## [4] "SpeciesName"
                               "OriginalSource"
                                                     "Site"
## [7] "Date"
                               "Tag"
                                                     "HeightOfMeasurement"
## [10] "Diameter"
                               "Height"
                                                     "CrownArea"
## [13] "sp.code"
```

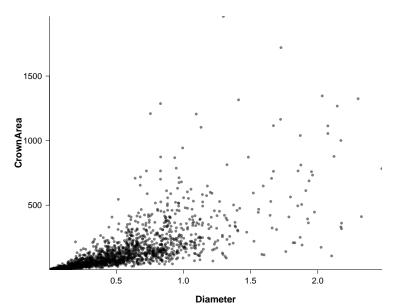
- Step 2: fit a model
- ► Step 3: evaluate the model

#### ▶ Plot the relationship

```
## Explore the data
par(cex.main = 1.5, mar = c(4.4, 5, 2, 1) + 0.1,
    mgp = c(3.5, 1, 0), cex.lab = 1.5,
    font.lab = 2, cex.axis = 1.3,
    bty = "1", las = 1,
    mfrow=c(1,1),xaxs="i",yaxs="i")

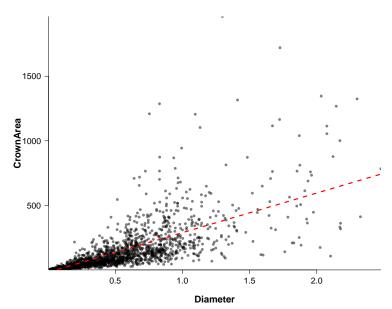
## Look at relationship
plot(CrownArea Diameter,allo,
    pch=16,col=rgb(0,0,0,alpha=0.5))
```

▶ Plot the relationship



#### ▶ Fit a model

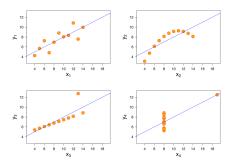
► Fit a model



#### Scrutinize the model

```
summary (mod)
##
## Call:
## lm(formula = CrownArea ~ Diameter, data = allo)
##
## Residuals:
     Min 1Q Median 3Q Max
## -519.93 -38.71 5.29 15.26 1582.22
##
## Coefficients:
            Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) -18.187 3.071 -5.922 3.64e-09 ***
## Diameter 306.802 5.483 55.952 < 2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 108.6 on 2423 degrees of freedom
## Multiple R-squared: 0.5637, Adjusted R-squared: 0.5635
## F-statistic: 3131 on 1 and 2423 DF, p-value: < 2.2e-16
```

# Anscombe's Quartet



## Anscombe's Quartet

Property	Value	Accuracy
Mean of x	9	exact
Sample variance of x	11	exact
Mean of y	7.50	to 2 decimal places
Sample variance of y	4.125	±0.003
Correlation between x and y	0.816	to 3 decimal places
Linear regression line	y = 3.00 + 0.500x	to 2 and 3 decimal places, respectively
Coefficient of determination of the linear regression	0.67	to 2 decimal places



What are regression assumptions?

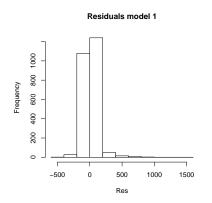
- ▶ What are regression assumptions?
- 1. Linearity & unbiasedness (no correlation in  $\epsilon$ )
- 2. Independence
- 3. Sample variation & little collinearity
- 4. Normality (of the residuals)
  - $Y \sim N(\mu = B_0 + B_1 X, \sigma = \epsilon)$
- 5. Homoskedasticity

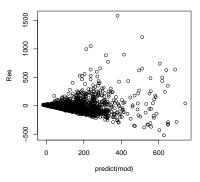
# Basic model diagnostics

► Normality & lack of fit

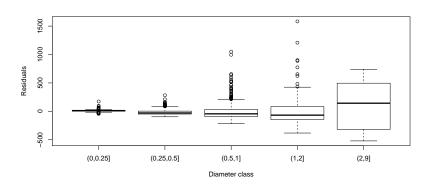
```
## Test model (Diagnostics)
Res <- residuals(mod)
par(mfrow=c(1,2))
hist(Res,main="Residuals model 1")
plot(Res^predict(mod))</pre>
```

# Basic model diagnostics





# Basic model diagnostics



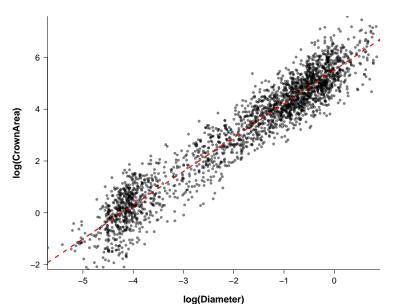
$$Y \sim N(\mu = B_0 + B_1 X, \sigma = B_2 \epsilon)$$

#### A better model

#### Transformation are non linear models

- $Y \sim N(B_0 X^{B_1}, \sigma = B_2 \epsilon)$

► Fit a slightly better model



#### A better model

```
summary (mod)
##
## Call:
## lm(formula = log(CrownArea) ~ log(Diameter), data = allo)
##
## Residuals:
##
      Min
              10 Median
                              30
                                     Max
## -2.1498 -0.3962 0.0053 0.4332 1.9921
##
## Coefficients:
                Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 5.505493 0.018918 291.0 <2e-16 ***
## log(Diameter) 1.302883 0.007868 165.6 <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.6178 on 2423 degrees of freedom
## Multiple R-squared: 0.9188, Adjusted R-squared: 0.9188
## F-statistic: 2.742e+04 on 1 and 2423 DF, p-value: < 2.2e-16
```

$$CA = \beta_0 DBH^{\beta_1} \tag{1}$$

# Evaluate fit at the species level

```
## Lets look at the fit per species
sp <- unique(allo$sp.code)</pre>
pdf("speciesLevelDiagnostics.pdf")
par(mfrow=c(2,2))
## get predictions and observations
Pred <- predict(mod)</pre>
Obser <- allo$CrownArea
for(i in sp){
    inc <- allo$sp.code%in%i
    plot(Pred[inc]~Obser[inc],
         ylab="Predicted",xlab="Observed",
         main=i)
    abline(0,1,lwd=2,ltv=2)
dev.off()
```

#### Look at the PDF

### Exercise 2: Improve species level fits

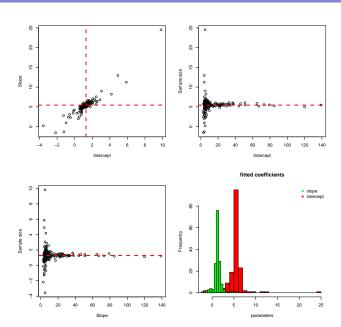
#### Start from the code below

```
## Lets look at the fit per species
sp <- unique(allo$sp.code)
pdf("speciesLevelDiagnostics.pdf")
par(mfrow=c(2,2))
## get predictions and observations
Pred <- predict(mod)</pre>
Obser <- allo$CrownArea
for(i in sp){
   inc <- allo$sp.code%in%i
    plot(Pred[inc]~Obser[inc],
         vlab="Predicted",xlab="Observed",
         main=i)
    abline(0,1,lwd=2,lty=2)
dev.off()
```

```
## Fit for each species
sp <- unique(allo$sp.code)</pre>
pdf("speciesLevelFits.pdf")
par(mfrow=c(2,2))
## get predictions and observations
Dbh <- allo$Diameter
Obser <- allo$CrownArea
SpFits <- array(dim=c(length(sp),3))
N <- as.numeric(table(allo$sp.code))
colnames(SpFits) <- c("Intercept", "Slope", "N")</pre>
for(i in sp){
    inc <- allo$sp.code%in%i
    plot(log(Obser[inc])~log(Dbh[inc]),
         vlab="Predicted",xlab="Observed",
         main=i)
    mod <- lm(log(Obser[inc])~log(Dbh[inc]))</pre>
    abline(mod, lwd=2, lty=2)
    SpFits[which(sp==i),] <- c(coef(mod),N[which(sp==i)])</pre>
dev.off()
## pdf
## 2
```

#### Look at the PDF

```
## look at estimated coefficients compared to sample size
par(mfrow=c(2,2))
MeanStats <- colMeans(SpFits) # mean values
## slope and intercept
plot(SpFits[.1]~SpFits[.2].xlab="Intercept".vlab="Slope")
abline (h=MeanStats [1], v=MeanStats [2], col="red", lty=2, lwd=3)
## intercept and sample size
plot(SpFits[,1]~SpFits[,3],xlab="Intercept",ylab="Sample size")
abline(h=MeanStats[1],col="red",lty=2,lwd=3)
## intercept and sample size
plot(SpFits[,2]~SpFits[,3],xlab="Slope",ylab="Sample size")
abline(h=MeanStats[2],col="red",lty=2,lwd=3)
## distributions
hist(SpFits[,1],freq=TRUE,col="red",breaks=20,xlab="parameters",
     main="fitted coefficients")
hist(SpFits[,2],freq=TRUE,col="green",breaks=20,add=TRUE)
legend("topright",legend=c("slope","intercept"),
       bty="n",pch=16,col=c("green", "red"))
```

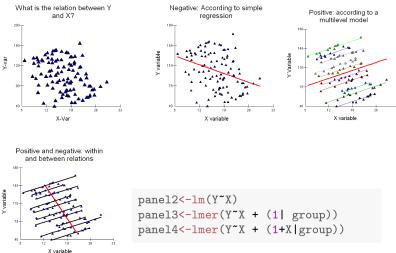


# Multi-level modeling

$$Y_{ij} \sim N(\mu = B_{0j} + B_1 X_{ij}, \sigma = \sigma_{\epsilon})$$

$$B_{0j} \sim N(\mu = \gamma, \sigma = \sigma_{\gamma})$$

#### In a hierarchical world some model can be misleading!



How to find out which model to use?

- ▶ What are the regression assumptions of multi-level modeling?
  - $Y_{ij} \sim N(\mu = B_{0i} + B_1 X_{ii}, \sigma = \sigma_{\epsilon})$
  - $\triangleright$   $B_{0j} \sim N(\mu = \gamma, \sigma = \sigma_{\gamma})$
  - 1. Linearity & unbiasedness (no correlation in  $\epsilon$  in i or j)
  - 2. Independence ("at the test level")
  - 3. Sample variation & little collinearity
  - 4. Normality (of the residuals)

$$Y \sim N(\mu = B_0 + B_1 X, \sigma = \epsilon)$$

Homoskedasticity

A short primer on (G)LMMs

#### **Tree Allometery**

Exercise 3: find the best multilevel model to predict crown area from diameter

### Basic model diagnostics

```
## fit multi-level model
require(lme4)

## Loading required package: lme4
## Loading required package: Matrix

LMmod <- lm(log(CrownArea)^log(Diameter),allo)
LMMmod <- lmer(log(CrownArea)^log(Diameter)+(1+log(Diameter)|sp.code),allo)</pre>
```

```
## Summarize model
summary (LMMmod)
## Linear mixed model fit by REML ['lmerMod']
## Formula: log(CrownArea) ~ log(Diameter) + (1 + log(Diameter) | sp.code)
##
     Data: allo
##
## REML criterion at convergence: 4062.1
##
## Scaled residuals:
      Min 1Q Median 3Q
                                    Max
## -3.2354 -0.6382 -0.0163 0.6593 3.4589
##
## Random effects:
                     Variance Std.Dev. Corr
## Groups Name
## sp.code (Intercept) 0.06302 0.2510
           log(Diameter) 0.01715 0.1310 0.36
##
## Residual
                         0.27748 0.5268
## Number of obs: 2425, groups: sp.code, 162
##
## Fixed effects:
##
               Estimate Std. Error t value
## (Intercept) 5.54102 0.03275 169.19
## log(Diameter) 1.31477 0.01847 71.17
##
## Correlation of Fixed Effects:
              (Intr)
##
## log(Diamtr) 0.648
```

## Multi-level modeling

$$Y_{ij} \sim N(\mu = B_{0j} + B_1 X_{ij}, \sigma = \sigma_{\epsilon})$$

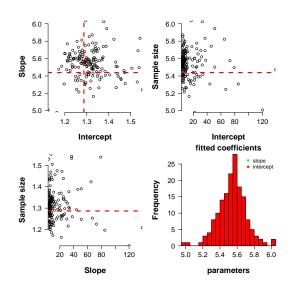
$$B_{0j} \sim N(\mu = \gamma, \sigma = \sigma_{\gamma})$$

$$B_{1j} \sim N(\mu = \alpha, \sigma = \sigma_{\alpha})$$

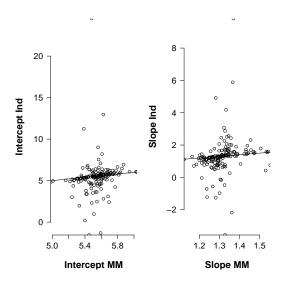
$$B_j \sim N(\mu = [\gamma, \alpha], \sigma = \Sigma)$$

```
## extract coefficients
LMMcoef <- as.data.frame(coef(LMMmod)$sp.code)
LMMcoef$n <- N</pre>
```

```
## Look at coefficients
par(mfrow=c(2,2))
MeanStats <- colMeans(LMMcoef)
print(MeanStats)
print(fixef(LMMmod))
plot(LMMcoef[,1]~LMMcoef[,2],xlab="Intercept",ylab="Slope")
abline (h=MeanStats [1], v=MeanStats [2], col="red", lty=2, lwd=3)
plot(LMMcoef[.1]~LMMcoef[.3].xlab="Intercept",vlab="Sample size")
abline(h=MeanStats[1],col="red",ltv=2,lwd=3)
plot(LMMcoef[,2]~LMMcoef[,3],xlab="Slope",ylab="Sample size")
abline(h=MeanStats[2],col="red",ltv=2,lwd=3)
hist(LMMcoef[,1],freq=TRUE,col="red",breaks=20,xlab="parameters",
     main="fitted coefficients")
hist(LMMcoef[,2],freg=TRUE,col="green",breaks=20,add=TRUE)
legend("topright",legend=c("slope","intercept"),
       bty="n",pch=16,col=c("green","red"))
```



```
## Compared earlier model fits mixed model coefficients
par(mfrow=c(1,2))
plot(SpFits[,1]~LMMcoef[,1],xlab="Intercept MM",ylab="Intercept Ind")
abline(0,1)
plot(SpFits[,2]~LMMcoef[,2],xlab="Slope MM",ylab="Slope Ind")
abline(0,1)
```



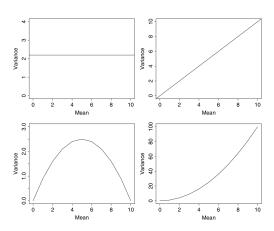
#### LMM towards GLMM

- What are the regression assumptions of multi-level modeling?
  - $ightharpoonup Y_{ii} \sim N(\mu = B_{0i} + B_1 X_{ii}, \sigma = \sigma_{\epsilon})$
  - $\triangleright$   $B_{0j} \sim N(\mu = \gamma, \sigma = \sigma_{\gamma})$
  - 1. Linearity & unbiasedness (no correlation in  $\epsilon$  in i or j)
  - 2. Independence ("at the test level")
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$$Y \sim N(\mu = B_0 + B_1 X, \sigma = \epsilon)$$

5. Homoskedasticity

# **GLMs**



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