

Sean Z, Z' variables aleatorias tales que nos interesa estimar la cantidad

$$\mathbb{E}[Z]\mathbb{E}[Z']$$

mediante la simulación de $Z_1, \dots, Z_n \overset{iid}{\sim} \mathcal{L}(Z)$ y $Z'_1, \dots, Z'_n \overset{iid}{\sim} \mathcal{L}(Z')$, tales que $(Z_1, \dots, Z_n) \perp (Z'_1, \dots, Z'_n)$,
y el uso de estimadores

$$\left(\frac{1}{n} \sum_{i=1}^n Z_i\right) \left(\frac{1}{n} \sum_{i=1}^n Z'_i\right) \quad \text{y} \quad \frac{1}{n} \sum_{i=1}^n Z_i Z'_i$$

$\nearrow \mathbb{E}[Z] \quad \nearrow \mathbb{E}[Z'] \quad \nearrow \mathbb{E}[ZZ'] = \mathbb{E}[Z]\mathbb{E}[Z']$

Calculemos las varianzas de estos estimadores y compáremoslas (¿Cuál es más grande?)

Lema: Dados X, Y v.a. t. $X \perp Y$, tenemos que

$$\text{Var}(XY) = \text{Var}(X)\text{Var}(Y) + \text{Var}(X)\mathbb{E}^2[Y] + \text{Var}(Y)\mathbb{E}^2[X]$$

$$\text{Dem: } \text{Var}(XY) = \mathbb{E}[(XY)^2] - \mathbb{E}^2[XY]$$

$$= \mathbb{E}[X^2]\mathbb{E}[Y^2] - \mathbb{E}^2[X]\mathbb{E}^2[Y]$$

$$= (\text{Var}(X) + \mathbb{E}^2[X])(\text{Var}(Y) + \mathbb{E}^2[Y]) - \mathbb{E}^2[X]\mathbb{E}^2[Y]$$

$$\text{Var}(X) = \mathbb{E}[X^2] - \mathbb{E}^2[X]$$



$$\mathbb{E}[X^2] = \text{Var}(X) + \mathbb{E}^2[X]$$

$$(\text{Var}(x) + \mathbb{E}^2[x]) (\text{Var}(y) + \mathbb{E}^2[y]) - \mathbb{E}^2[x] \mathbb{E}^2[y]$$

$$= \text{Var}(x) \text{Var}(y) + \text{Var}(x) \mathbb{E}^2[y] + \mathbb{E}^2[x] \text{Var}(y) + \mathbb{E}^2[x] \mathbb{E}^2[y]$$

$$= \text{Var}(x) \text{Var}(y) + \text{Var}(x) \mathbb{E}^2[y] + \text{Var}(y) \mathbb{E}^2[x] - \mathbb{E}^2[x] \mathbb{E}^2[y]$$

$$\text{Var}\left(\left(\frac{1}{n}\sum_{i=1}^n Z_i\right)\left(\frac{1}{n}\sum_{i=1}^n Z_i'\right)\right) \sim \frac{1}{n}\sum_{i=1}^n Z_i Z_i'$$

$$\bar{Z} = \frac{1}{n}\sum_{i=1}^n Z_i, \quad \bar{Z}' = \frac{1}{n}\sum_{i=1}^n Z_i'$$

$$\begin{aligned}\text{Var}(\bar{Z}\bar{Z}') &= \text{Var}(\bar{Z})\text{Var}(\bar{Z}') + \text{Var}(\bar{Z})\mathbb{E}^2[\bar{Z}'] + \text{Var}(\bar{Z}')\mathbb{E}^2[\bar{Z}] \\ &= \frac{\text{Var}(Z)\text{Var}(Z')}{n} + \frac{\text{Var}(Z)\mathbb{E}^2[Z']}{n} + \frac{\text{Var}(Z')\mathbb{E}^2[Z]}{n} \\ &\quad \frac{\text{Var}(Z)\text{Var}(Z')}{n^2} + \frac{\text{Var}(Z)\mathbb{E}^2[Z']}{n} + \frac{\text{Var}(Z')\mathbb{E}^2[Z]}{n}\end{aligned}$$

$$\left(\frac{1}{n} \sum_{i=1}^n Z_i\right) \left(\frac{1}{n} \sum_{i=1}^n Z_i'\right) \quad y \quad \frac{1}{n} \sum_{i=1}^n Z_i Z_i'$$

$$\text{Var}\left(\frac{1}{n} \sum_{i=1}^n Z_i Z_i'\right) = \frac{1}{n^2} \text{Var}\left(\sum_{i=1}^n Z_i Z_i'\right)$$

$$= \frac{n \text{Var}(Z Z')}{n^2}$$

$$= \frac{\text{Var}(Z Z')}{n}$$

$$= \frac{\text{Var}(Z) \text{Var}(Z') + \text{Var}(Z) E^2(Z') + \text{Var}(Z') E^2(Z)}{n}$$

$$\text{Comb.} \quad \frac{\text{Var}(z)\text{Var}(z')}{n^2} \leq \frac{\text{Var}(z)\text{Var}(z')}{n}$$

$$\text{Int.} \quad \text{Var}(\bar{z}\bar{z}') \leq \text{Var}\left(\frac{1}{n} \sum_{i=1}^n z_i z_i'\right)$$

$$(2) \quad \frac{\hat{\sigma}_{\text{Anth}}}{R} = \frac{\hat{\sigma}_{\text{cmc}}}{R} (1 + \rho)$$

$$\Rightarrow \hat{\sigma}_{\text{Anth}} = \hat{\sigma}_{\text{cmc}} (1 + \rho)$$

$$R = 2M$$

$$(z_1, z_2) \dots (z_{R-1}, z_R)$$

En integración Monte Carlo, una selección estándar es considerar $z_1 = g(u)$,

$z_2 = g(1-u)$ para estimar $z = \int_0^1 g$

Si $g(x) = x$, ent. tenemos que

$$z_1 + z_2 = u + (1-u) = 1, \text{ por}$$

lo que el estimador antitético tiene
varianza cero (es una cte).

Si $g(x) = x^2$, ent. tenemos

$$\rho = \frac{\text{Cov}(U^2, (1-U)^2)}{\sqrt{\text{Var}(U^2) \text{Var}((1-U)^2)}}$$

$$= \frac{E[U^2(1-U)^2] - E[U^2]E[(1-U)^2]}{\sqrt{\text{Var}(U^2)\text{Var}(U^2)}}$$

$$= \frac{E[u^2(1-u)^2] - E[u^2]E[(1-u)^2]}{\sqrt{\text{Var}(u^2)\text{Var}(u^2)}}$$

$$= \frac{E[u^2(1-2u+u^2)] - E^2[u^2]}{\text{Var}(u^2)}$$

$$= \frac{E[u^2] - 2E[u^3] + E[u^4] - E^2[u^2]}{\text{Var}(u^2)}$$

$$\text{Var}(u^2) = E[u^4] - E^2[u^2]$$

$$= 1 + \frac{E[U^2] - 2(E[U])^2}{\text{Var}(U)}$$

$$= 1 + \frac{\frac{1}{3} - \frac{2}{4}}{\frac{1}{5} - \left(\frac{1}{3}\right)^2}$$

$$= 1 + \frac{\frac{4-9}{12}}{\frac{9-5}{45}}$$

So doja ol lector

$$\approx \frac{-7}{8}$$

$U(a,b)$ Si $a=0$,
 $a^0=1$

$$m_k = \frac{1}{k+1} \sum_{i=0}^k a_i b^{k-i}$$

$$\begin{aligned}
 \hat{T}_{Anth} &= \hat{T}_{cmc} (1 + \rho) \\
 &= \hat{T}_{cmc} \left(1 - \frac{7}{8}\right) \\
 &= \frac{1}{8} \hat{T}_{cmc}
 \end{aligned}$$

$$\textcircled{3} \quad \text{Var}(\mathbb{E}[X|Y]) \leq \text{Var}(X)$$

Sean X, τ v.a. con las sig.

distribuciones,

$$X|\tau \sim \text{Normal}(0, 1/\tau)$$

$$\tau \sim \text{Gamma}(\alpha, \beta)$$

¿Cuál es la densidad marginal de X ?

$$g_{x|\tau}(x|\tau) = \sqrt{\frac{\tau}{2\pi}} \exp\left(-\frac{\tau x^2}{2}\right)$$

$$g_{\tau}(\tau) = \frac{\beta^{\alpha}}{\Gamma(\alpha)} \tau^{\alpha-1} \exp(-\beta\tau)$$

$$f_X(x) = \int g_{x|\tau}(x|\tau) g_{\tau}(\tau) d\tau$$

$$g_{X|T}(x|\tau) = \sqrt{\frac{\tau}{2\pi}} \exp\left(-\frac{\tau x^2}{2}\right)$$

$$g_T(\tau) = \frac{\beta^\alpha}{\Gamma(\alpha)} \tau^{\alpha-1} \exp(-\beta\tau)$$

$$= \int \frac{\tau^{1/2}}{\sqrt{2\pi}} \exp\left(-\frac{\tau x^2}{2}\right) \frac{\beta^\alpha}{\Gamma(\alpha)} \tau^{\alpha-1} \exp(-\beta\tau)$$

$$= \frac{\beta^\alpha}{\sqrt{2\pi} \Gamma(\alpha)} \int \tau^{\overbrace{(\alpha-1/2)-1}^{\alpha'}} \exp\left(-\tau \underbrace{\left(\frac{x^2}{2} + \beta\right)}_{\beta'}\right)$$

$$= \frac{\beta^\alpha}{\sqrt{2\pi} \Gamma(\alpha)} \frac{\Gamma(\alpha')}{\beta'^{\alpha'}} \int \frac{\beta'^{\alpha'}}{\Gamma(\alpha')} \tau^{\alpha'-1} \exp(-\tau \beta')$$

$$= \frac{\beta^\alpha}{\sqrt{2\pi} \Gamma(\alpha)} \Gamma\left(\alpha + \frac{1}{2}\right) \left(\frac{x^2}{2} + \beta\right)^{-(\alpha + \frac{1}{2})}$$

$$= \frac{\Gamma\left(\frac{2\alpha+1}{2}\right)}{\sqrt{2\pi} \Gamma\left(\frac{2\alpha}{2}\right)} \cancel{\beta^\alpha} \cancel{\beta^{-\alpha-\frac{1}{2}}} \left(\frac{x^2}{2\beta} + 1\right)^{-(\alpha + \frac{1}{2})}$$

$$= \frac{\Gamma\left(\frac{2\alpha+1}{2}\right)}{\sqrt{2\pi} \sqrt{\beta} \Gamma\left(\frac{2\alpha}{2}\right)} \left(\frac{x^2}{2\beta} + 1\right)^{-(\frac{2\alpha+1}{2})}$$

$$v = 2\alpha$$

$$= \frac{\Gamma\left(\frac{v+1}{2}\right)}{\sqrt{2\pi} \sqrt{\beta} \Gamma\left(\frac{v}{2}\right)} \left(\frac{x^2}{\frac{v}{\alpha}\beta} + 1\right)^{-\left(\frac{v+1}{2}\right)}$$

$$\hat{\sigma} = \sqrt{\frac{\beta}{\alpha}} = \frac{\Gamma\left(\frac{v+1}{2}\right)}{\sqrt{\frac{2\alpha}{\beta}} \sqrt{\beta} \Gamma\left(\frac{v}{2}\right)} \left(\frac{\frac{1}{v} \frac{x^2}{\frac{\beta}{\alpha}} + 1\right)^{-\left(\frac{v+1}{2}\right)}$$

$$= \frac{\Gamma\left(\frac{\nu+1}{2}\right)}{\sqrt{\pi\alpha} \Gamma\left(\frac{\nu}{2}\right)} \left(\frac{1}{\nu} \left(\frac{x}{\sigma} \right)^2 + 1 \right)^{-\left(\frac{\nu+1}{2}\right)}$$

$$X \sim t(\nu=2\alpha, \mu=0, \sigma=\sqrt{\frac{\beta}{\alpha}})$$

$$\text{S; } \alpha=\beta=1$$

$$X \sim t(\nu=2, \mu=0, \sigma=1)$$

$$\sim t(2)$$

$$\therefore E[X] = 0$$

Ahora, si nos fijamos en $E[X|\tau]$,

$$E[X|\tau] = 0$$

Dado $\tau = \tau, X|\tau = \tau \sim N(0, 1/\tau)$

$$X|\tau \sim N(0, 1/\tau)$$

$$E[X|\tau] = 0$$