

GSERM - Ljubljana (2023)

Analyzing Panel Data

January 11, 2023

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1 contributor

9369 lines (9369 sloc) 354 KB

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	Country	CountryCode	Year	POLITY	PercentLiterate	UnivEnrollmentPerK	GDP	TotalTrade
1	United States	2	1945	10	NA	NA	NA	NA
2	United States	2	1946	10	97.1	11.8	1360	10982
3	United States	2	1947	10	97.3	14.8	1464	15665

Can also use (e.g.) `read_csv` (in `readr`):

```
> library(readr)
> Data<-read_csv("https://github.com/PrisonRodeo/GSERM-Ljubljana-APD-git/raw/main/Exercises/GSERM-APD-Exercise-January-2023.csv")
```

Generalized Least Squares Models

Start with a focus on residuals... For:

$$Y_{it} = \mathbf{X}_{it}\beta + u_{it}$$

i.i.d. OLS u_{it} s require:

$$\begin{aligned} \mathbf{u}\mathbf{u}' \equiv \mathbf{\Omega} &= \sigma^2 \mathbf{I} \\ &= \begin{pmatrix} \sigma^2 & 0 & \dots & 0 \\ 0 & \sigma^2 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \sigma^2 \end{pmatrix} \end{aligned}$$

This means that within units:

- $\text{Var}(u_{it}) = \text{Var}(u_{is}) \forall t \neq s$ (temporal homoscedasticity)
- $\text{Cov}(u_{it}, u_{is}) = 0 \forall t \neq s$ (no within-unit autocorrelation)

and between units:

- $\text{Var}(u_{it}) = \text{Var}(u_{jt}) \forall i \neq j$ (cross-unit homoscedasticity)
- $\text{Cov}(u_{it}, u_{jt}) = 0 \forall i \neq j$ (no between-unit / spatial correlation)

Estimator:

$$\hat{\beta}_{GLS} = (\mathbf{X}'\Omega^{-1}\mathbf{X})^{-1}\mathbf{X}'\Omega^{-1}\mathbf{Y}$$

with:

$$\widehat{V(\beta_{GLS})} = (\mathbf{X}'\Omega^{-1}\mathbf{X})^{-1}$$

Two approaches:

- Use OLS \hat{u}_{it} s to get $\hat{\Omega}$ (“feasible GLS” / “weighted least squares”)
- Use substantive knowledge about the data to structure Ω

The variance-covariance matrix is:

$$\begin{aligned}\text{Var}(\hat{\beta}_{WLS}) &= \sigma^2(\mathbf{X}'\mathbf{\Omega}^{-1}\mathbf{X})^{-1} \\ &\equiv (\mathbf{X}'\mathbf{W}^{-1}\mathbf{X})^{-1}\end{aligned}$$

A common case is:

$$\text{Var}(u_i) = \sigma^2 \frac{1}{N_i}$$

where N_i is the number of observations upon which (aggregate) observation i is based.

“Robust” Variance Estimators

Recall that, if $\sigma_i^2 \neq \sigma_j^2 \forall i \neq j$,

$$\begin{aligned}\text{Var}(\beta_{\text{Het.}}) &= (\mathbf{X}'\mathbf{X})^{-1}(\mathbf{X}'\mathbf{W}^{-1}\mathbf{X})(\mathbf{X}'\mathbf{X})^{-1} \\ &= (\mathbf{X}'\mathbf{X})^{-1} \mathbf{Q} (\mathbf{X}'\mathbf{X})^{-1}\end{aligned}$$

where $\mathbf{Q} = (\mathbf{X}'\mathbf{W}^{-1}\mathbf{X})$ and $\mathbf{W} = \sigma^2\mathbf{\Omega}$.

We can rewrite \mathbf{Q} as

$$\begin{aligned}\mathbf{Q} &= \sigma^2(\mathbf{X}'\mathbf{\Omega}^{-1}\mathbf{X}) \\ &= \sum_{i=1}^N \sigma_i^2 \mathbf{x}_i \mathbf{x}_i'\end{aligned}$$

Estimate $\hat{\mathbf{Q}}$ as:

$$\hat{\mathbf{Q}} = \sum_{i=1}^N \hat{u}_i^2 \mathbf{x}_i \mathbf{x}_i'$$

Yields:

$$\begin{aligned} \widehat{\text{Var}}(\hat{\boldsymbol{\beta}})_{\text{Robust}} &= (\mathbf{X}'\mathbf{X})^{-1}(\mathbf{X}'\hat{\mathbf{Q}}^{-1}\mathbf{X})(\mathbf{X}'\mathbf{X})^{-1} \\ &= (\mathbf{X}'\mathbf{X})^{-1} \left[\mathbf{X}' \left(\sum_{i=1}^N \hat{u}_i^2 \mathbf{x}_i \mathbf{x}_i' \right)^{-1} \mathbf{X} \right] (\mathbf{X}'\mathbf{X})^{-1} \end{aligned}$$

“Robust” VCV estimates:

- are heteroscedasticity-consistent, but
- are biased in small samples, and
- are less efficient than “naive” estimates when $\text{Var}(u) = \sigma^2 \mathbf{I}$.

“Clustering”

Huber / White

?????????

WLS / GLS

I know very little
about my error
variances...

I know a great
deal about my
error variances...

A common case:

$$Y_{ij} = \mathbf{X}_{ij}\beta + u_{ij}$$

with

$$\sigma_{ij}^2 = \sigma_{ik}^2.$$

“Robust, clustered” estimator:

$$\widehat{\text{Var}}(\beta)_{\text{Clustered}} = (\mathbf{X}'\mathbf{X})^{-1} \left\{ \mathbf{X}' \left[\sum_{i=1}^N \left(\sum_{j=1}^{n_j} \hat{u}_{ij}^2 \mathbf{X}_{ij} \mathbf{X}_{ij}' \right) \right]^{-1} \mathbf{X} \right\} (\mathbf{X}'\mathbf{X})^{-1}$$

Robust / Clustered SEs: A Simulation

```
url_robust <- "https://raw.githubusercontent.com/IsidoreBeautrelet/economictheoryblog/master/robust_summary.R"
eval(parse(text = getURL(url_robust, ssl.verifypeer = FALSE)),
      envir=.GlobalEnv)
```

```
> set.seed(3844469)
> X <- rnorm(10)
> Y <- 1 + X + rnorm(10)
> df10 <- data.frame(ID=seq(1:10),X=X,Y=Y)
> fit10 <- lm(Y~X,data=df10)
> summary(fit10)
```

Call:

```
lm(formula = Y ~ X, data = df10)
```

Residuals:

Min	1Q	Median	3Q	Max
-1.318	-0.766	0.195	0.378	1.590

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	0.954	0.311	3.06	0.016 *
X	0.589	0.291	2.03	0.077 .

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.985 on 8 degrees of freedom

Multiple R-squared: 0.339, Adjusted R-squared: 0.257

F-statistic: 4.11 on 1 and 8 DF, p-value: 0.0772

```
> rob10 <- vcovHC(fit10,type="HC1")
> sqrt(diag(rob10))
(Intercept)      X
    0.315      0.285
```

Robust / Clustered SEs: A Simulation (continued)

```
> # "Clone" each observation 100 times:
>
> df1K <- df10[rep(seq_len(nrow(df10)),each=100),]
> df1K <- pdata.frame(df1K, index="ID")
> fit1K <- lm(Y~X,data=df1K)

> summary(fit1K)

Call:
lm(formula = Y ~ X, data = df1K)

Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept)   0.9536     0.0279   34.2   <2e-16 ***
X              0.5893     0.0260   22.6   <2e-16 ***
---
Signif. codes:  0 *** 0.001 ** 0.01 * 0.05 . 0.1 1

Residual standard error: 0.882 on 998 degrees of freedom
Multiple R-squared:  0.339, Adjusted R-squared:  0.339
F-statistic: 513 on 1 and 998 DF, p-value: <2e-16

> summary(fit1K, cluster="ID")

Call:
lm(formula = Y ~ X, data = df1K)

Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept)   0.954     0.297     3.21  0.0014 **
X              0.589     0.269     2.19  0.0286 *
---
Signif. codes:  0 *** 0.001 ** 0.01 * 0.05 . 0.1 1

Residual standard error: 0.882 on 998 degrees of freedom
Multiple R-squared:  0.339, Adjusted R-squared:  0.339
F-statistic: 4.8 on 1 and 9 DF, p-value: 0.0561
```

Example:

$$Y_t = \beta_0 + \beta_1 X_t + u_t$$

$$u_t = \rho u_{t-1} + e_t$$

with $e_t \sim i.i.d. N(0, \sigma_u^2)$ and $\rho \in [-1, 1]$ (typically).

→ “First-order autoregressive” (“AR(1)”) errors.

Serially Correlated Errors and OLS

Detection

- *Plot* of residuals vs. lagged residuals
- *Runs* test (Geary test)
- Durbin-Watson d
 - Calculated as:

$$d = \frac{\sum_{t=2}^N (\hat{u}_t - \hat{u}_{t-1})^2}{\sum_{t=1}^N \hat{u}_t^2}$$

- Non-standard distribution ($d \in [0, 4]$)
- Null: No autocorrelation
- Only detects first-order autocorrelation

Serially Correlated Errors and OLS

What to do about it?

- GLS, incorporating ρ / $\hat{\rho}$ into the equation
- First-difference equations (regressing changes of Y on changes of \mathbf{X})
- Cochrane-Orcutt / Prais-Winsten:
 1. Estimate the basic equation via OLS, and obtain residuals
 2. Use the residuals to consistently estimate $\hat{\rho}$ (i.e. the empirical correlation between u_t and u_{t-1})
 3. Use this estimate of $\hat{\rho}$ to estimate the *difference equation*:

$$(Y_t - \rho Y_{t-1}) = \beta_0(1 - \rho) + \beta_1(X_t - \rho X_{t-1}) + (u_t - \rho u_{t-1})$$

4. Save the residuals, and use them to estimate $\hat{\rho}$ again
5. Repeat this process until successive estimates of $\hat{\rho}$ differ by a very small amount

Running Example Redux

The World Development Indicators:

- Cross-national country-level time series data
- $N = 215$ countries, $T = 72$ years (1960-2021) + missingness
- Full descriptions are listed in the Github repo [here](#)

Regression model:

$$\text{WBLI}_{it} = \beta_0 + \beta_1 \text{Population Growth}_{it} + \beta_2 \text{Urban Population}_{it}^2 + \beta_3 \text{Fertility Rate}_{it} + \beta_4 \ln(\text{GDP Per Capita})_{it} + \beta_5 \text{Natural Resource Rents}_{it} + \beta_6 \text{Cold War}_t + u_{it}$$

Descriptive Statistics:

```
> describe(subset,fast=TRUE)
```

	vars	n	mean	sd	min	max	range	se
WomenBusLawIndex	1	7566	59.99	18.79	17.50	100.00	82.50	0.22
PopGrowth	2	7566	1.70	1.46	-6.77	17.51	24.28	0.02
UrbanPopulation	3	7566	51.19	23.90	2.85	100.00	97.16	0.27
FertilityRate	4	7566	3.67	1.91	0.90	8.61	7.70	0.02
NaturalResourceRents	5	7566	6.82	10.45	0.00	87.51	87.51	0.12
ColdWar	6	7566	0.32	0.47	0.00	1.00	1.00	0.01
lnGDPPerCap	7	7566	8.28	1.44	5.32	11.63	6.31	0.02

How Much Autocorrelation in **X**?

Note that:

$$d = 2(1 - \rho)$$

which means that we can calculate:

$$\rho = 1 - \frac{d}{2}.$$

Autocorrelation in the Predictors

	Variable	Rho
1	Population Growth	0.942
2	Urban Population	0.973
3	Fertility Rate	0.968
4	GDP Per Capita	0.976
5	Natural Resource Rents	0.917
6	Cold War	0.913

Baseline Model: OLS (+ D-W Test)

```
> OLS<-plm(WomenBusLawIndex~PopGrowth+UrbanPopulation+FertilityRate+log(GDPPerCapita)+NaturalResourceRents+ColdWar,
+          data=WDI,model="pooling")

> summary(OLS)
.
.
.
Coefficients:
              Estimate Std. Error t-value    Pr(>|t|)
(Intercept)      54.8395      1.7261   31.77 < 2e-16 ***
PopGrowth        -3.1926      0.1437  -22.21 < 2e-16 ***
UrbanPopulation  -0.0584      0.0109   -5.37 0.000000083 ***
FertilityRate     -1.7928      0.1652  -10.85 < 2e-16 ***
log(GDPPerCapita)  3.1544      0.1993   15.83 < 2e-16 ***
NaturalResourceRents -0.3486      0.0162  -21.54 < 2e-16 ***
ColdWar          -11.3437      0.3716  -30.53 < 2e-16 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Total Sum of Squares:    2670000
Residual Sum of Squares: 1280000
R-Squared:                0.519
Adj. R-Squared: 0.519
F-statistic: 1361.01 on 6 and 7559 DF, p-value: <2e-16

> # Durbin-Watson test:
>
> pdwtest(OLS)

Durbin-Watson test for serial correlation in panel models

data:  WomenBusLawIndex ~ PopGrowth + UrbanPopulation + FertilityRate + ...
DW = 0.14, p-value <2e-16
alternative hypothesis: serial correlation in idiosyncratic errors
```

Example: Prais-Winsten

```
> PraisWinsten<-panelAR(WomenBusLawIndex~PopGrowth+UrbanPopulation+
+      FertilityRate+log(GDPPerCapita)+NaturalResourceRents+
+      ColdWar, data=WDI,panelVar="ISO3",timeVar="YearNumeric",
+      autoCorr="ar1",panelCorrMethod="none",
+      rho.na.rm=TRUE)
```

```
> summary(PraisWinsten)
```

Panel Regression with AR(1) Prais-Winsten correction and homoskedastic variance

Unbalanced Panel Design:

```
Total obs.:      7566 Avg obs. per panel 40.677
Number of panels: 186 Max obs. per panel 50
Number of times:  50  Min obs. per panel 1
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	71.81494	3.16208	22.71	< 2e-16 ***
PopGrowth	-0.03691	0.08543	-0.43	0.6657
UrbanPopulation	-0.03721	0.02468	-1.51	0.1317
FertilityRate	-5.64038	0.25230	-22.36	< 2e-16 ***
log(GDPPerCapita)	1.41254	0.36363	3.88	0.0001 ***
NaturalResourceRents	-0.01931	0.00868	-2.23	0.0261 *
ColdWar	-0.90520	0.22070	-4.10	0.000041 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

R-squared: 0.3621

Wald statistic: 1075.5886, Pr(>Chisq(6)): 0

```
> PraisWinsten$panelStructure$rho
[1] 0.9523
```

Better in a Table

	OLS	Prais-Winsten
Intercept	54.84* (1.73)	71.81* (3.16)
Population Growth	-3.19* (0.14)	-0.04 (0.08)
Urban Population	-0.06* (0.01)	-0.04 (0.02)
Fertility Rate	-1.79* (0.17)	-5.64* (0.25)
ln(GDP Per Capita)	3.15* (0.20)	1.41* (0.36)
Natural Resource Rents	-0.35* (0.02)	-0.02* (0.008)
Cold War	-11.34* (0.37)	-0.91* (0.22)
$\hat{\rho}$		0.95
R ²	0.52	0.36
Adj. R ²	0.52	
<i>NT</i>	7566	7566
<i>N</i> panels		186

* $p < 0.05$

Some Panel Data Challenges

Consider the error terms in the model:

$$Y_{it} = \mathbf{X}_{it}\beta + u_{it}$$

Issues:

<u>In Words:</u>	<u>In a Formula:</u>
<u>Variances:</u>	
Unit-Wise Heteroscedasticity	$\text{Var}(u_{it}) \neq \text{Var}(u_{jt})$
Temporal Heteroscedasticity	$\text{Var}(u_{it}) \neq \text{Var}(u_{is})$
<u>Covariances:</u>	
Contemporary Cross-Unit Correlation	$\text{Cov}(u_{it}, u_{jt}) \neq 0$
Within-Unit Serial Correlation	$\text{Cov}(u_{it}, u_{is}) \neq 0$
Non-Contemporaneous Cross-Unit Correlation	$\text{Cov}(u_{it}, u_{js}) \neq 0$

Parks' (1967) Approach

Assume:

- $\text{Var}(u_{it}, u_{jt}) = \sigma^2$ or σ_i^2 (Common or unit-specific error variances)
- $\text{Var}(u_{it}) = \text{Var}(u_{is}) \forall t \neq s$ (Temporal homoscedasticity)
- $\text{Cov}(u_{it}, u_{jt}) = \sigma_{ij} \forall i \neq j$ (Pairwise contemporaneous cross-unit correlation)
- $\text{Cov}(u_{it}, u_{is}) = \rho$ or ρ_i (Common or unit-specific temporal correlation)
- $\text{Cov}(u_{it}, u_{js}) = 0 \forall i \neq j, t \neq s$ (No non-contemporaneous cross-unit correlation)

(B&K: “panel error assumptions”).

Then:

1. Use OLS to generate $\hat{u}s \rightarrow \hat{\rho} (\rightarrow \hat{\Omega})$,
2. Use $\hat{\rho}$ for Prais-Winsten.

This method was widely used prior to Beck & Katz (1995)

$$\mathbf{\Omega} = \begin{pmatrix} \Sigma & 0 & \cdots & 0 \\ 0 & \Sigma & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \Sigma \end{pmatrix} = \Sigma \otimes \mathbf{I}_T$$

where

$$\Sigma_{N \times N} = \begin{pmatrix} \sigma_1^2 & \sigma_{12} & \cdots & \sigma_{1N} \\ \sigma_{12} & \sigma_2^2 & \cdots & \sigma_{2N} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{1N} & \sigma_{2N} & \cdots & \sigma_N^2 \end{pmatrix}$$

Means:

- $\frac{N(N-1)}{2}$ distinct contemporaneous covariances σ_{ij} ,
- NT observations,
- $\rightarrow 2T/(N+1)$ observations per $\hat{\sigma}$

From PROC PANEL in SAS:

Standard Corrections

For the PARKS option, the first-order autocorrelation coefficient must be estimated for each cross section. Let ρ be the $N \times 1$ vector of true parameters and $R = (r_1, \dots, r_N)'$ be the corresponding vector of estimates. Then, to ensure that only range-preserving estimates are used in PROC PANEL, the following modification for R is made:

$$r_i = \begin{cases} r_i & \text{if } |r_i| < 1 \\ \max(.95, r_{\max}) & \text{if } r_i \geq 1 \\ \min(-.95, r_{\min}) & \text{if } r_i \leq -1 \end{cases}$$

where

$$r_{\max} = \begin{cases} 0 & \text{if } r_i < 0 \text{ or } r_i \geq 1 \quad \forall i \\ \max_j[r_j : 0 \leq r_j < 1] & \text{otherwise} \end{cases}$$

and

$$r_{\min} = \begin{cases} 0 & \text{if } r_i > 0 \text{ or } r_i \leq -1 \quad \forall i \\ \max_j[r_j : -1 < r_j \leq 0] & \text{otherwise} \end{cases}$$

Whenever this correction is made, a warning message is printed.

Panel-Corrected Standard Errors

Key to PCSEs:

$$\hat{\sigma}_{ij} = \frac{\sum_{t=1}^T \hat{u}_{it} \hat{u}_{jt}}{T}$$

Define:

$$\mathbf{U}_{T \times N} = \begin{pmatrix} \hat{u}_{11} & \hat{u}_{21} & \cdots & \hat{u}_{N1} \\ \hat{u}_{12} & \hat{u}_{22} & \cdots & \hat{u}_{N2} \\ \vdots & \vdots & \ddots & \vdots \\ \hat{u}_{1T} & \hat{u}_{2T} & \cdots & \hat{u}_{NT} \end{pmatrix}$$

$$\hat{\Sigma} = \frac{(\mathbf{U}'\mathbf{U})}{T}$$

$$\hat{\Omega}_{PCSE} = \frac{(\mathbf{U}'\mathbf{U})}{T} \otimes \mathbf{I}_T$$

Panel-Corrected Standard Errors

Correct formula:

$$\text{Cov}(\hat{\beta}_{PCSE}) = (\mathbf{X}'\mathbf{X})^{-1}[\mathbf{X}'\mathbf{\Omega}\mathbf{X}](\mathbf{X}'\mathbf{X})^{-1}$$

General Issues:

- PCSEs do not fix unit-level heterogeneity (a la “fixed” / “random” effects)
- They also do not deal with dynamics
- They depend critically on the “panel data assumptions” of Park / Beck & Katz

Panel Assumptions and Numbers of Parameters

Panel Assumptions	No AR(1)	Common $\hat{\rho}$	Separate $\hat{\rho}_i$ s
$\sigma_i^2 = \sigma^2, \text{Cov}(\sigma_{it}, \sigma_{jt}) = 0$	$k + 1$	$k + 2$	$k + N + 1$
$\sigma_i^2 \neq \sigma^2, \text{Cov}(\sigma_{it}, \sigma_{jt}) = 0$	$k + N$	$k + N + 1$	$k + 2N$
$\sigma_i^2 \neq \sigma^2, \text{Cov}(\sigma_{it}, \sigma_{jt}) \neq 0$	$\frac{N(N-1)}{2} + k + N$	$\frac{N(N-1)}{2} + k + N + 1$	$\frac{N(N-1)}{2} + k + 2N$

Example: GLS with Homoscedastic AR(1) Errors

```
> GLS<-glS(WomenBusLawIndex~PopGrowth+UrbanPopulation+FertilityRate+
+          log(GDPPerCapita)+NaturalResourceRents+ColdWar,
+          data=WDI,correlation=corAR1(form=~1|ISO3),
+          na.action="na.omit")
```

```
> summary(GLS)
```

Generalized least squares fit by REML

Model: WomenBusLawIndex ~ PopGrowth + UrbanPopulation + FertilityRate + log(GDPPerCapita) + NaturalResourceRents + ColdWar

Data: WDI

AIC BIC logLik

35750 35813 -17866

Correlation Structure: AR(1)

Formula: ~1 | ISO3

Parameter estimate(s):

Phi

0.9888

Coefficients:

	Value	Std.Error	t-value	p-value
(Intercept)	53.93	4.195	12.855	0.0000
PopGrowth	0.04	0.079	0.469	0.6393
UrbanPopulation	0.18	0.041	4.318	0.0000
FertilityRate	-4.22	0.319	-13.222	0.0000
log(GDPPerCapita)	1.66	0.446	3.720	0.0002
NaturalResourceRents	0.01	0.008	1.080	0.2804
ColdWar	-0.40	0.207	-1.929	0.0538

Example: PCSEs

```
> PCSE<-panelAR(WomenBusLawIndex~PopGrowth+UrbanPopulation+FertilityRate+
+               log(GDPPerCapita)+NaturalResourceRents+ColdWar,
+               data=WDI,panelVar="ISO3",timeVar="YearNumeric",
+               autoCorr="ar1",panelCorrMethod="pcse",
+               rho.na.rm=TRUE)
```

```
> summary(PCSE)
```

Panel Regression with AR(1) Prais-Winsten correction and panel-corrected standard errors

Unbalanced Panel Design:

```
Total obs.:      7566 Avg obs. per panel 40.677
Number of panels: 186 Max obs. per panel 50
Number of times:  50  Min obs. per panel 1
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	71.8149	4.6453	15.46	<2e-16 ***
PopGrowth	-0.0369	0.0978	-0.38	0.706
UrbanPopulation	-0.0372	0.0250	-1.49	0.136
FertilityRate	-5.6404	0.3983	-14.16	<2e-16 ***
log(GDPPerCapita)	1.4125	0.4559	3.10	0.002 **
NaturalResourceRents	-0.0193	0.0129	-1.49	0.135
ColdWar	-0.9052	0.5804	-1.56	0.119

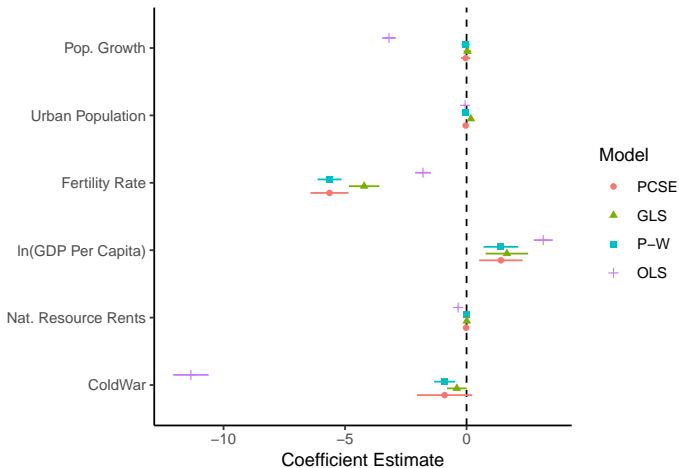
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

R-squared: 0.3621

Wald statistic: 394.4807, Pr(>Chisq(6)): 0

```
> PCSE$panelStructure$rho
[1] 0.9523
```

Model Comparisons



Dynamics!

Time Series: Stationarity

Stationarity: A constant d.g.p. over time.¹

Mean stationarity:

$$E(Y_t) = \mu \quad \forall t$$

Variance stationarity:

$$\text{Var}(Y_t) = E[(Y_t - \mu)^2] \equiv \sigma_Y^2 \quad \forall t$$

Covariance stationarity:

$$\text{Cov}(Y_t, Y_{t-s}) = E[(Y_t - \mu)(Y_{t-s} - \mu)] = \gamma_s \quad \forall s$$

¹A stricter form of stationarity requires that the joint probability distribution (in other words, *all* the moments) of series of observations $\{Y_1, Y_2, \dots, Y_t\}$ is the same as that for $\{Y_{1+s}, Y_{2+s}, \dots, Y_{t+s}\}$ for all t and s .

The “ARIMA” Approach

“ARIMA” = *Autoregressive Integrated Moving Average*...

A (first-order) integrated series (“random walk”) is:

$$Y_t = Y_{t-1} + u_t, \quad u_t \sim i.i.d.(0, \sigma_u^2)$$

...a/k/a a “random walk”:

$$\begin{aligned} Y_t &= Y_{t-2} + u_{t-1} + u_t \\ &= Y_{t-3} + u_{t-2} + u_{t-1} + u_t \\ &= \sum_{t=0}^T u_t \end{aligned}$$

I(1) series are not stationary.

Variance:

$$\text{Var}(Y_t) \equiv E(Y_t)^2 = t\sigma^2$$

Autocovariance:

$$\text{Cov}(Y_t, Y_{t-s}) = |t - s|\sigma^2.$$

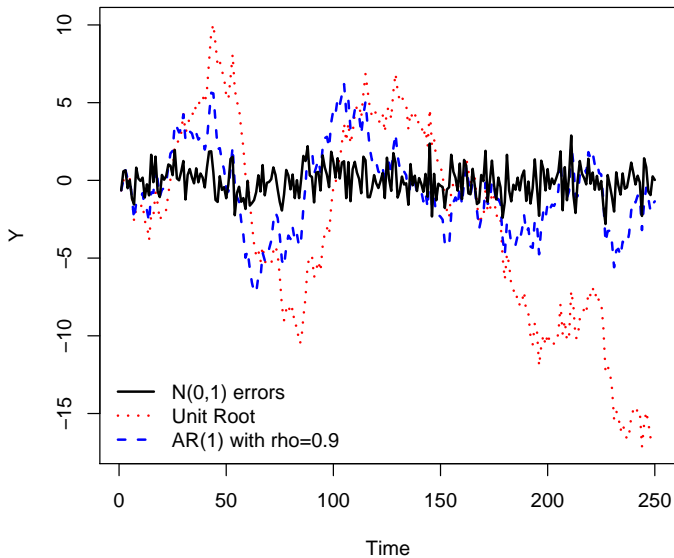
Both depend on t ...

I(1) series (continued)

More generally:

- $|\rho| > 1$
 - Series is nonstationary / *explosive*
 - Past shocks have a greater impact than current ones
 - Uncommon
- $|\rho| < 1$
 - *Stationary* series
 - Effects of shocks die out exponentially according to ρ
 - Is mean-reverting
- $|\rho| = 1$
 - Nonstationary series
 - Shocks persist at full force
 - Not mean-reverting; variance increases with t

Time Series Types, Illustrated



I(1) Series: Differencing

For an I(1) series:

$$Y_t - Y_{t-1} = u_t$$

which we often write in terms of the difference operator Δ (or sometimes ∇):

$$\Delta Y_t = u_t$$

The differenced series is just the (stationary, ergodic) white-noise process u_t .

Unit Root Tests Review: Dickey-Fuller

Two steps:

- Estimate $Y_t = \rho Y_{t-1} + u_t$,
- test the hypothesis that $\hat{\rho} = 1$, *but*
- this requires that the us are uncorrelated.

But suppose:

$$\Delta Y_t = \sum_{i=1}^p d_i \Delta Y_{t-i} + u_t$$

which yields

$$Y_t = Y_{t-1} + \sum_{i=1}^p d_i \Delta Y_{t-i} + u_t.$$

D.F. tests will be incorrect.

Augmented Dickey-Fuller Tests:

- Estimate

$$\Delta Y_t = \rho Y_{t-1} + \sum_{i=1}^p d_i \Delta Y_{t-i} + u_t$$

- Test $\hat{\rho} = 1$

Phillips-Perron Tests:

- Estimate:

$$\Delta Y_t = \alpha + \rho Y_{t-1} + u_t$$

- Calculate modified test statistics (Z_ρ and Z_t)
- Test $\hat{\rho} = 0$

Issues with Unit Roots in Panel Data

- Short series + Asymptotic tests → “borrow strength”
- Requires uniform unit roots across cross-sectional units
- Many tests require balanced panels...
- Various alternatives:
 - Maddala and Wu (1999)
 - Hadri (2000)
 - Levin, Lin and Chu (2002)
 - Im, Pesaran, and Shin (2003)
- What to do?
 - Difference the data...
 - Error-correction models

Panel Unit Root Tests: R

```
[data wrangling...]

> purtest(WBLI.W,exo="trend",test="levinlin",pmax=2)

Levin-Lin-Chu Unit-Root Test (ex. var.: Individual Intercepts and
Trend)

data: WBLI.W
z = -2.7, p-value = 0.003
alternative hypothesis: stationarity

> purtest(WBLI.W,exo="trend",test="hadri",pmax=2)

Hadri Test (ex. var.: Individual Intercepts and Trend)
(Heterosked. Consistent)

data: WBLI.W
z = 189, p-value <2e-16
alternative hypothesis: at least one series has a unit root

> purtest(WBLI.W,exo="trend",test="madwu",pmax=2)

Maddala-Wu Unit-Root Test (ex. var.: Individual Intercepts and
Trend)

data: WBLI.W
chisq = 332, df = 376, p-value = 0.9
alternative hypothesis: stationarity

> purtest(WBLI.W,exo="trend",test="ips",pmax=2)

Im-Pesaran-Shin Unit-Root Test (ex. var.: Individual Intercepts
and Trend)

data: WBLI.W
Wtbar = 3.6, p-value = 1
alternative hypothesis: stationarity
```

Table: Panel Unit Root Tests: WBRI

Test	Alternative	Statistic	Estimate	P-Value
Levin-Lin-Chu Test	stationarity	z	-2.286	0.0111
Hadri Test (Heterosked. Consistent)	\geq one series has a unit root	z	192.036	< 0.001
Maddala-Wu Test	stationarity	χ^2	782.604	< 0.001
Im-Pesaran-Shin Test	stationarity	W_t	3.342	0.9996

Note: All assume individual intercepts and trends.

$$Y_{it} = \phi Y_{it-1} + \mathbf{X}_{it}\beta_{LDV} + \epsilon_{it}$$

If ϵ_{it} is perfect...

- $\hat{\beta}_{LDV}$ is biased (but consistent),
- $O(\text{bias}) = \frac{-1+3\beta_{LDV}}{T}$

If ϵ_{it} is autocorrelated...

- $\hat{\beta}_{LDV}$ is biased and inconsistent
- IV is one (bad) option...

Lagged Y s and GLS-ARMA

Can rewrite:

$$\begin{aligned}Y_{it} &= \mathbf{X}_{it}\boldsymbol{\beta}_{AR} + u_{it} \\u_{it} &= \phi u_{it-1} + \eta_{it}\end{aligned}$$

as

$$\begin{aligned}Y_{it} &= \mathbf{X}_{it}\boldsymbol{\beta}_{AR} + \phi u_{it-1} + \eta_{it} \\&= \mathbf{X}_{it}\boldsymbol{\beta}_{AR} + \phi(Y_{it-1} - \mathbf{X}_{it-1}\boldsymbol{\beta}_{AR}) + \eta_{it} \\&= \phi Y_{it-1} + \mathbf{X}_{it}\boldsymbol{\beta}_{AR} + \mathbf{X}_{it-1}\psi + \eta_{it}\end{aligned}$$

where $\psi = \phi\boldsymbol{\beta}_{AR}$ and $\psi = 0$ (by assumption).

Lagged Y s and World Domination

In:

$$Y_{it} = \phi Y_{it-1} + \mathbf{X}_{it}\beta_{LDV} + \epsilon_{it}$$

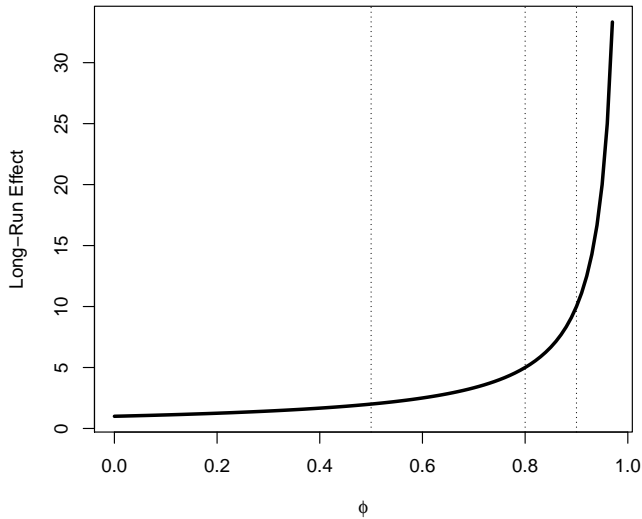
Achen: Bias “deflates” $\hat{\beta}_{LDV}$ relative to $\hat{\phi}$, “suppress” the effects of \mathbf{X} ...

Keele & Kelly (2006):

- Contingent on ϵ s having autocorrelation
- Key: In LDV, *long-run impact of a unit change in X is:*

$$\hat{\beta}_{LR} = \frac{\hat{\beta}_{LDV}}{1 - \hat{\phi}}$$

Long-Run Impact for $\hat{\beta} = 1$



Lagged Y s and Unit Effects

Consider:

$$Y_{it} = \phi Y_{it-1} + \mathbf{X}_{it}\boldsymbol{\beta} + \alpha_i + u_{it}.$$

If we omit the unit effects, we have:

$$Y_{it} = \phi Y_{it-1} + \mathbf{X}_{it}\boldsymbol{\beta} + u_{it}^*$$

with

$$u_{it}^* = \alpha_i + u_{it}$$

Lagging yields:

$$Y_{it-1} = \phi Y_{it-2} + \mathbf{X}_{it-1}\boldsymbol{\beta} + \alpha_i + u_{it-1}$$

which means

$$\text{Cov}(Y_{it-1}, u_{it}^*) \neq 0. \rightarrow \text{bias in } \hat{\phi}, \hat{\boldsymbol{\beta}}$$

Bias in $\hat{\phi}$ is

- toward zero when $\phi > 0$,
- increasing in ϕ .

Including unit effects still yields bias in $\hat{\phi}$ of $O(\frac{1}{T})$, and bias in $\hat{\beta}$.

Solutions:

- Difference/GMM estimation
- Bias correction approaches

First Difference Estimation

$$\begin{aligned}Y_{it} - Y_{it-1} &= \phi(Y_{it-1} - Y_{it-2}) + (\mathbf{X}_{it} - \mathbf{X}_{it-1})\beta + (\alpha_i - \alpha_i) + (u_{it} - u_{it-1}) \\ \Delta Y_{it} &= \phi \Delta Y_{it-1} + \Delta \mathbf{X}_{it}\beta + \Delta u_{it}\end{aligned}$$

Anderson/Hsiao: If \nexists autocorrelation, then use ΔY_{it-2} or Y_{it-2} as instruments for ΔY_{it-1} ...

- Consistent in theory,
- in practice, the former is preferred, and both have issues if ϕ is high;
- both are inefficient.

Arellano & Bond (also Wawro): Use *all* lags of Y_{it} and \mathbf{X}_{it} from $t - 2$ and before.

- “Good” estimates, better as $T \rightarrow \infty$,
- Easy to handle higher-order lags of Y ,
- Easy software (p1m in R , xtabond in Stata).
- Model *is* fixed effects...
- \mathbf{Z}_i has $T - p - 1$ rows, $\sum_{i=p}^{T-2} i$ columns \rightarrow difficulty of estimation declines in p , grows in T .

Kiviet (1995, 1999; Bun and Kiviet 2003; Bruno 2005a,b): Derive the bias in $\hat{\phi}$ and $\hat{\beta}$, then correct it...

- More accurate than the instrumental-variables/GMM estimators of A&H/A&B...
- ...especially when T is small; but not as T gets reasonably large ($T \approx 20$)

Some Dynamic Models

	Lagged Y	First Difference	FE	Lagged Y + FE
Intercept	2.260* (0.335)	0.641* (0.040)		
Lagged WBLI	0.986* (0.002)			0.948* (0.004)
Population Growth	-0.051 (0.027)	0.035 (0.077)	0.073 (0.119)	0.011 (0.037)
Urban Population	0.002 (0.002)	-0.040 (0.062)	0.248* (0.021)	0.009 (0.007)
Fertility Rate	-0.085* (0.030)	-1.023* (0.373)	-2.066* (0.166)	-0.292* (0.052)
ln(GDP Per Capita)	-0.036 (0.037)	0.780 (0.476)	9.161* (0.310)	0.276* (0.102)
Natural Resource Rents	-0.010* (0.003)	0.020* (0.008)	0.035 (0.018)	-0.003 (0.006)
Cold War	-0.298* (0.072)	-0.021 (0.204)	-7.192* (0.295)	-0.445* (0.094)
R ²	0.984	0.003	0.535	0.956
Adj. R ²	0.984	0.002	0.523	0.954
Num. obs.	7463	7380	7566	7463

* $p < 0.05$

Anderson-Hsiao, Arellano-Bond, etc.

In R:

- Anderson-Hsiao can be fit using `lm` or (more easily) `p1m` in the `p1m` package
- Arellano-Bond is most easily fit using `pgmm` (“panel gmm”) in the `p1m` package
- See Cribari and Millo (2018, Chapter 7) for statistics + code details
- [This post](#) is also useful...

Stata:

- `xtabond` / `xtdpdsys` `xtdpd` fit both A-H and A-B / Blundell-Bond models (among others)
- [This](#) is also a good (slightly dated) reference

What if Y is *trending* over time?

- First Question: Why?
 - Organic growth (e.g., populations)
 - Temporary / short-term factors
 - Covariates...
- Second question: Should we care?
(A: Yes, usually... → “spurious regressions”)
- Third question: What to do?
 - Ignore it...
 - Include a counter / trend term...

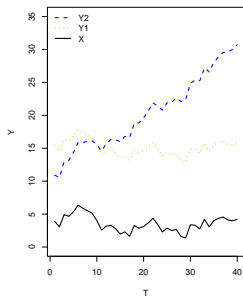
In general, adding a trend term will *decrease* the magnitudes of $\hat{\beta}$...

Trends Matter, Illustrated

Data generating processes:

$$Y_{1t} = 10 + (1 \times X_t) + u_t$$

$$Y_{2t} = 5 + (1 \times X_t) + (0.5 \times T) + u_t$$



	Y ₁	Y ₂	
		No Trend	Trend
X	0.921*** (0.245)	-0.382 (0.786)	0.874*** (0.255)
T			0.482*** (0.026)
Constant	10.300*** (0.917)	20.200*** (2.950)	5.860*** (1.200)
Observations	40	40	40
R ²	0.272	0.006	0.905
Adjusted R ²	0.253	-0.020	0.900
Residual Std. Error	1.800 (df = 38)	5.790 (df = 38)	1.810 (df = 37)

Note:

* p<0.1; ** p<0.05; *** p<0.01

Trends Matter, Part II

Table: FE Models of WBLI

	FE	FE.trend	FE.intx
Population Growth	0.073 (0.119)	-0.287*** (0.100)	-0.242** (0.100)
Urban Population	0.248*** (0.021)	-0.024 (0.018)	-0.003 (0.018)
Fertility Rate	-2.066*** (0.166)	1.080*** (0.150)	1.018*** (0.149)
ln(GDP Per Capita)	9.161*** (0.310)	2.867*** (0.283)	2.585*** (0.283)
Natural Resource Rents	0.035* (0.018)	0.009 (0.015)	0.008 (0.015)
Cold War	-7.192*** (0.295)	1.660*** (0.293)	9.300*** (0.944)
Trend (1950=0)		0.749*** (0.013)	0.783*** (0.014)
Cold War x Trend			-0.220*** (0.026)
Observations	7,566	7,566	7,566
R ²	0.535	0.674	0.678
Adjusted R ²	0.523	0.666	0.669
F Statistic	1,414.000*** (df = 6; 7374)	2,182.000*** (df = 7; 7373)	1,937.000*** (df = 8; 7372)

*p<0.1; **p<0.05; ***p<0.01

Another Approach: Orthogonalization

Note: We're rarely substantively interested in the fixed effects $\hat{\alpha}$...

- \rightarrow reparameterize the α s so that they are *information-orthogonal* to the other parameters in the model (including the β s and ϕ)
- Key idea: Transform the α s so that (for example):

$$E \left(\frac{\partial^2 L_i}{\partial \alpha \partial \beta} \right) = 0$$

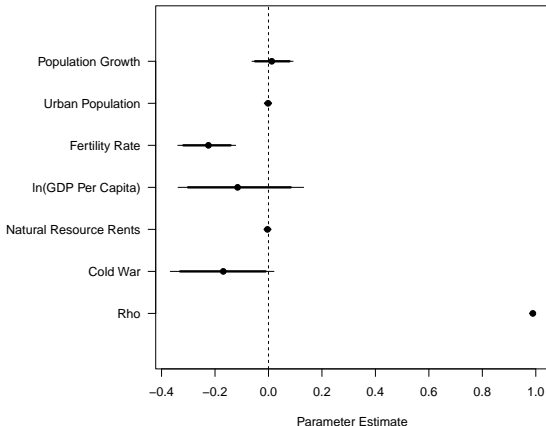
- Can do this via imposition of priors, in a Bayesian framework...
- **In general**, this approach is less assumption-laden and more efficient than the IV/GMM-based approaches discussed above.
- Provides consistent-in- N estimates for T as low as 2...

References:

- Lancaster, T. 2002. "Orthogonal Parameters and Panel Data." *Review of Economic Studies* 69:647-666.
- Pickup et al. (2017) [the "orthogonalized panel model" ("OPM")]

FE + Dynamics Using Orthogonalization

```
> library(OrthoPanels)
> set.seed(7222009)
> OPM.fit <- opm(WomenBusLawIndex~PopGrowth+UrbanPopulation+FertilityRate+
  lnGDPPerCap+NaturalResourceRents+ColdWar,data=WDI,
  index=c("ISO3","Year"),n.samp=1000)
```



OPM Results: Short- and Long-Run Effects

For $\hat{\phi} \approx 0.98$:

Parameter	Short-Run	Long-Run
Population Growth	0.0122	0.9148
Urban Population	-0.0016	-0.1420
Fertility Rate	-0.2247	-19.0090
ln(GDP Per Capita)	-0.1155	-9.9996
Natural Resource Rents	-0.0037	-0.3086
Cold War	-0.1691	-14.3630

R :

- the `plm` package (`purtest` for unit roots; `plm` for first-difference models; `pgmm` for Arellano-Bond)
- the `panelAR` package (GLS-ARMA models)
- the `glS` package (GLS)
- the `dynpanel` package (A&H, A&B; minimal...)

Stata :

- `xtglS` (GLS)
- `xtpcse` (PCSEs)
- `xtabond` / `xtdpd` (A&H A&B dynamic models)

Final Thoughts: Dynamic Panel Models

Things to consider:

- N vs. T ...
- Are dynamics nuisance or substance?
- What problem(s) do you *really* care about?