Stochastic simulation exercises

ex123

June 2024

1 Exercises june 7th

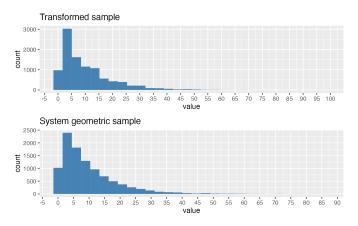
1.1 Exercise 1 (Sampling discrete random variables)

1.1.1 Simulate a geometric distribution

To simulate the geometric distribution we generated 10000 continuous uniform samples and then used the inverse distribution function technique to transform the uniform samples to samples following a geometric distribution. The transformation used was:

$$X = floor((\frac{\log(U)}{\log(1-p)}) + 1 \tag{1}$$

To compare the result we compared the resulting histograms of our transformed sample and a system available sampling from a geometric distribution.

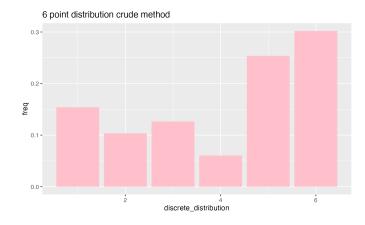


The results are not quite the same as hoped since there was a smaller tail on our transformed sample compared to the geometric sample. However the compared frequencies did not seem too far off.

A good way to test sample would be too perform a chi squared test up until the highest observed value that has over 5 observation. Time did not allow this unfortunately.

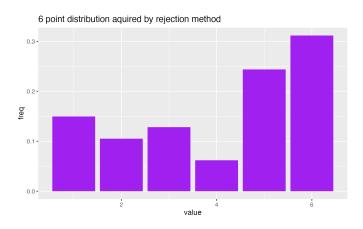
1.1.2 Simulate the 6 point distribution

• The direct crude method: This method revealed the frequencies show in the histogram (with frequencies on y) below:



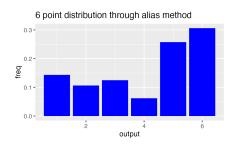
The resulting p value from a chi squared test revealed a p value of 0.997 which means our observation is extremely likely to come from the given 6 point distribution

• The rejection method: this method revealed the following histogram (with frequencies on the y axis)



A chi squared test gave the p value of 0.61 which is lower than the crude method but still concludes our sample to be likely from the given 6 point distribution

• Alias method: Using the Alias method the following frequencies were found using 10000 samples.



Verifying with a chi squared test revealed a p value of 0.32 "verifying" that our observation or something more extreme is likely

1.1.3 Compare the methods using adequate criteria and discuss results

To compare the methods we use running time to estimate which are best.

- Direct crude method: This method had a running time of 0.64 seconds which is nicely fast considering n = 10000
- The rejection method: The rejection method had a runtime of 0.26 seconds which is even faster than the crude method.
- The Alias method: The alias method revealed a runtime of 0.19 seconds. Thus the alias method was the overall fastest algorithm.

Using computational time, the alias method woulp appear to be the best method overall for constructing and abitrary discrete distribution. However, there could very well be a possibility that the code could be written more effectively (in terms of runtime) for each method.

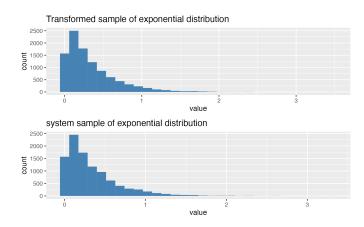
1.1.4 Give recommendations of how to choose the best suited method in different settings

- The alias method: This is great if we don't have too many classes since the setup process takes $\mathcal{O}(n)$ time, but the lookup process is of course $\mathcal{O}(1)$ so in general if the number of classes are not too high, the alias method is good for a high speed generation.
- The rejection method: This method can also be fast and it doesn't have a "long" setup time like the alias method. However, for the sample to converge quckly it need to have a high acceptance rate. If the acceptance rate is too low (as a consequence of having a high c value) then the algorithm will take longer to finish.
- The crude method: If there are many classes or some classes have probability mass very close to each
 other and maybe some class with a very high probability you might need a high sample size to get the real
 probabilities.

1.2 Exercise 2 (Sampling continuous distributions)

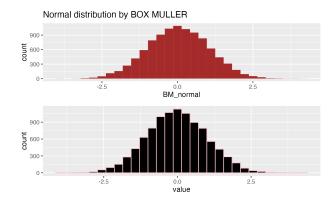
1.2.1 Generate samples from different continuous distributions

• Exponential distribution: Here we use the inverse transformation technique. Setting $\lambda = 3$ we compared again the transformed results with the system available generator.



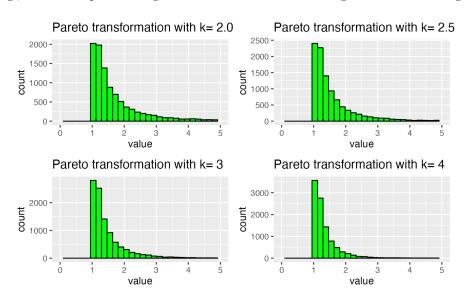
To test these results we used kolmogorov smirnov test. we arrived at a p value of 0.14 which is on the lower side, but still we cannot reject that the transformed observations are from the actual exponential distribution

• Normal distribution: Using the box muller method to transform our uniform samples to normal revealed the following results again compared to the system available normal distribution. We only compared one of the resulting transformed parameters.



A kolmogorov smirnov test revealed a p value of 0.71 meaning our transformed result was very likely to come from an actual standard normal distribution

• Pareto: Setting $\beta = 1$ and experimenting with the different values for k gave us the following histograms



We again tested these with the kolmogorov smirnov test. for k = 2.05 we got p value 0.45.

for k = 2.5 we got a p value of 0.07

for k = 3 we got p value of 0.18

for k = 4 we got p value of 0.17

A thing to not here is that the variance of the p values were very high! Each time we tested we frequently got results in either the very low range or the very high range.

1.2.2 Compare sample means with theoretical means for the pareto distribution

The means for every value of k in the first exercise were precise op to 3 digits. One could maybe expect a lower precision here since the tails are so large on the pareto distribution.

The variance however was very distant from the theoretical. This must be because of the have tails as a high value very far from the mean would would explode the variance as it is the average squared distance from the mean. The results can be seen on the table below. From the observed one could theorise that if we bootstrapped the variance parameter the variance of this would indeed also have a high variance and maybe also that the variance would increase for smaller values of k.

value <dbl></dbl>	sample_var <dbl></dbl>	theory_var <dbl></dbl>
2.05	5.483863	39.0476190
2.50	1.509980	3.3333333
3.00	0.613147	1.5000000
4.00	0.200078	0.6666667

1.2.3 Generate 100 95% confidence intervals for the mean based on 10 observations

Generating 100 confidence intervals for the mean and variance revealed the results below. Green line contain the actual parameter value and red lines don't.



To calculate the confidence intervals for the mean, we've used the fact that the sample variance given iid conditions will be normally distributed and since we have a low sample size, the t distribution is used to accommodate in terms of the quantiles used in the interval. For the confidence intervals for the variance, we use the fact that the $\frac{(n-1)s^2}{\sigma^2} \sim \chi_{n-1}^2$ and thus can extract a confidence interval from this. So in the end we've used the intervals:

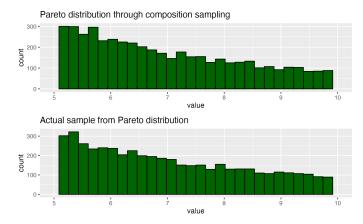
$$\bar{X} \pm t_{\alpha/2}(9) \frac{s}{\sqrt{n}} \tag{2}$$

$$\frac{(10-1)s^2}{\chi_{1-(\alpha/2)}^2(9)} \le \sigma^2 \le \frac{(10-1)s^2}{\chi_{(\alpha/2)}^2(9)} \tag{3}$$

To conclude on the results we see that small samples sizes impact the precision of the CI for the mean. However, the intervals are also much very long.

1.2.4 Simulate from the pareto distribution through composition

From slides we know that if we choose some β and let $Y \sim exp(\beta)$. Then the density $f(X|Y=y) = ye^{-yx}$ (I.e $X_i \sim exp(y_i)$ will describe a Pareto distribution. Generating this with n=10000 and $\beta=5$ and comparing with a system available pareto distribution, gave the following histograms.



Testing the distributions against each other with a kolmogorov smirnov test gave us a p-value of 0.75 which makes our composition sample very likely for a pare to distribution with $\beta = 5$.