

Stochastic Simulation

Generation of random variables

Continuous sample space

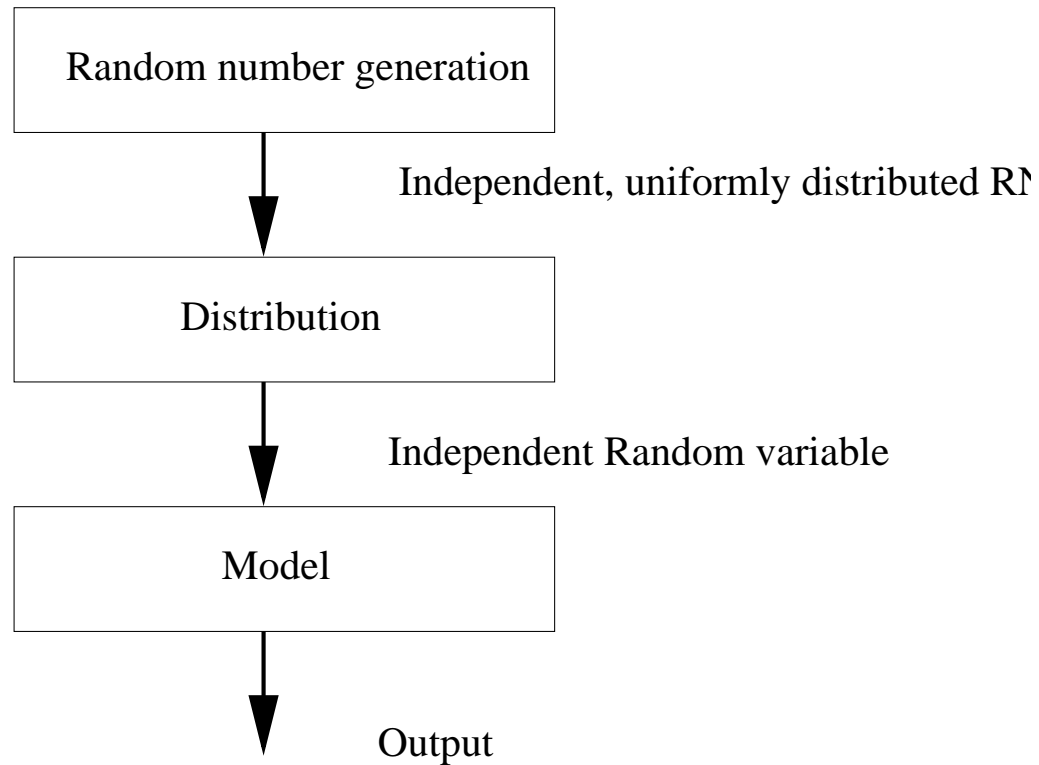
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Plan W1.1-2



Random variables



Aim

- The scope is the generation of **independent** random variables X_1, X_2, \dots, X_n with a **given distribution**, $F_x(x)$, (or probability density function [pdf]).
- We assume we have access to a supply (U_i) of random numbers, independent samples from the uniform distribution on $]0, 1[$.
- Our task is to transform U_i into X_i .

Generation of (pseudo)random variates

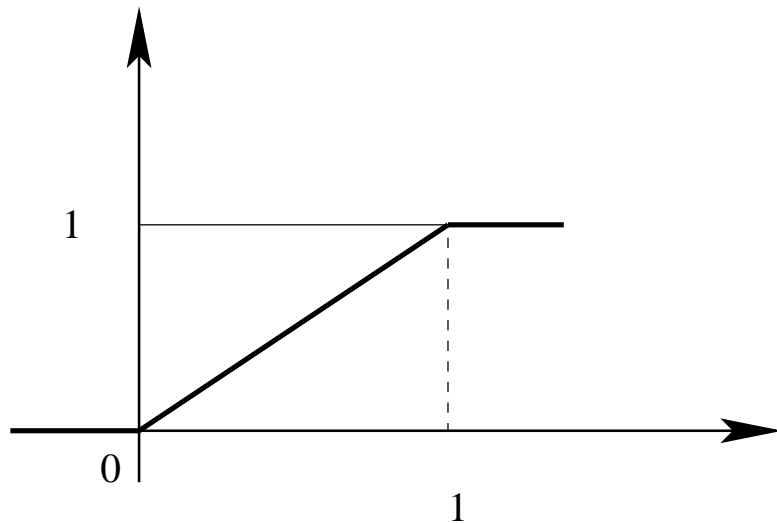


- Inverse transformation techniques
- Composition methods
- Acceptance/rejection methods
- Mathematical methods

Uniform distribution I

Our basic distribution or building block, $U_i \sim U(0, 1)$

$$f(x) = 1 \quad F(x) = x \quad \text{for} \quad 0 \leq x \leq 1$$

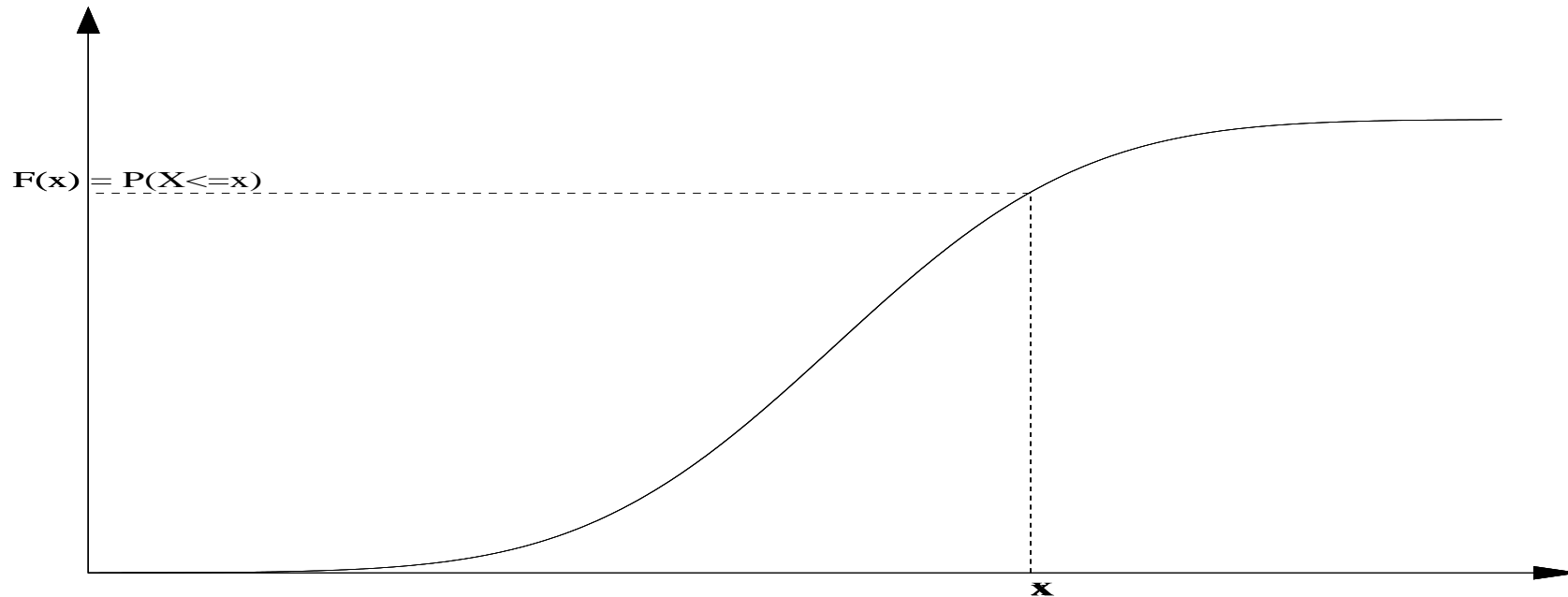


$$E(U_i) = \frac{1}{2} \quad \text{Var}(U_i) = \frac{1}{12}$$

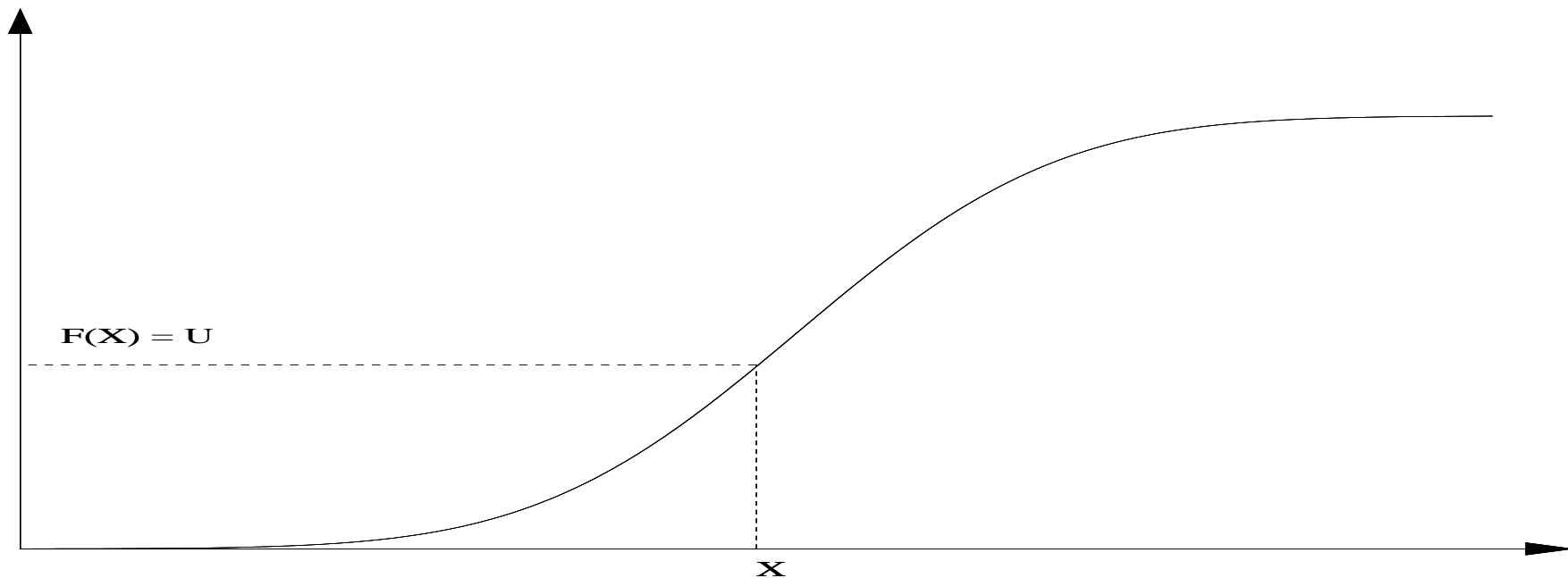
Inverse transformation

The cumulative distribution function (CDF)

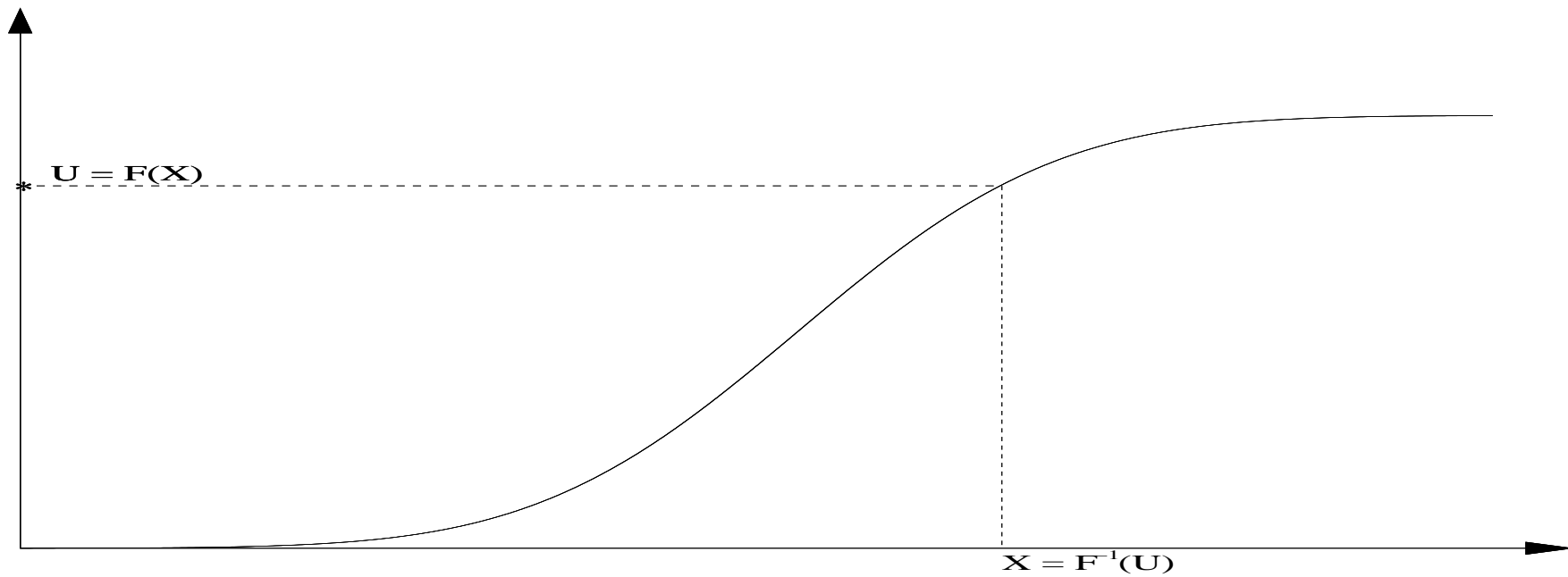
$$F(x) = P(X \leq x)$$



The random variable $F(X)$

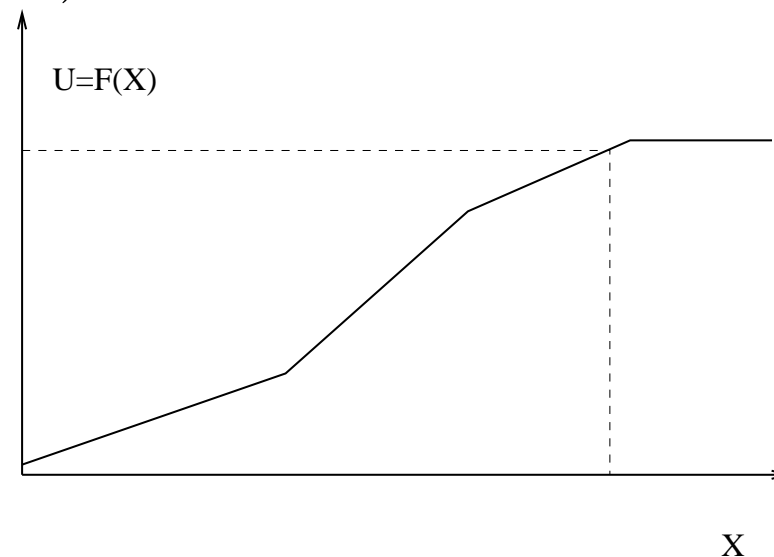


From U to X



Inversion method

The random variable $U = F(X)$



$$U = F(X) \quad F(x) = P(X \leq x)$$

$$P(U \leq u) = P(F(X) \leq u) = P(X \leq F^{-1}(u)) = F(F^{-1}(u)) = u$$

I.e. $F(X)$ is uniformly distributed.

Inversion method



Assume g continuous and increasing and let:

$$X = g(Y) \quad Y \sim F_Y(y) = \mathbf{P}(Y \leq y)$$

then

$$F_X(x) = \mathbf{P}(X \leq x) = \mathbf{P}(g(Y) \leq x) = \mathbf{P}(Y \leq g^{-1}(x)) = F_Y(g^{-1}(x))$$

If $Y = U$ then $F_U(u) = u$, and $F_X(x) = g^{-1}(x)$.

If

$$X = F^{-1}(U)$$

then X will have the cdf $F(x)$ ($(F^{-1}(x))^{-1} = F(x)$).

Uniform distribution II



Now, focus on $U(a, b)$.

$$f(x) = \frac{1}{b - a} \quad a \leq x \leq b$$

$$F(x) = \frac{x - a}{b - a} \quad F^{-1}(u) = a + (b - a)u$$

$$X = a + (b - a)U \quad \sim \quad U(a, b)$$

Exponential distribution

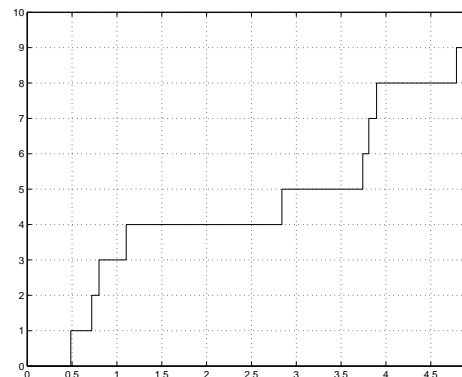
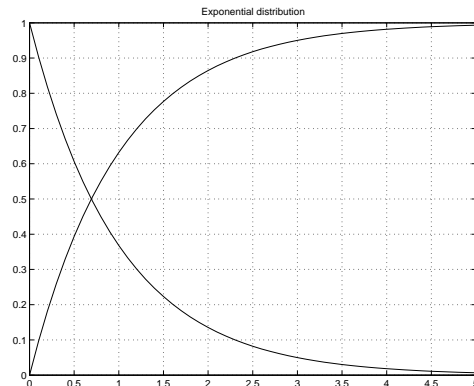
The time between events in a Poisson process is exponentially distributed. (Arrival time)



$$F(x) = 1 - e^{(-\lambda x)} \quad \mathbb{E}(X) = \frac{1}{\lambda} \quad F^{-1}(u) = -\frac{1}{\lambda} \log(1 - u)$$

So (both $1 - U$ and U are uniformly distributed)

$$X = -\frac{\log(U)}{\lambda} \sim \exp(\lambda)$$



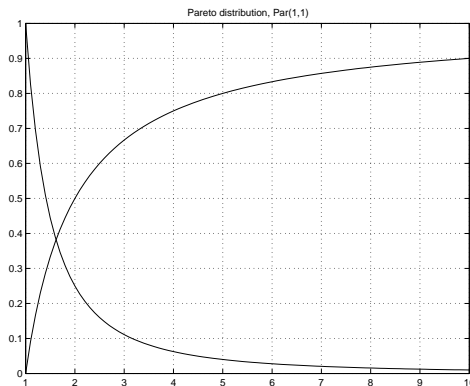
Pareto

Is often used in connection to description of income (over a certain level).



$$X \sim Pa(k, \beta) \quad F(x) = 1 - \left(\frac{\beta}{x}\right)^k \quad x \geq \beta \quad X = \beta \left(U^{-\frac{1}{k}}\right)$$

$$E(X) = \frac{k}{k-1}\beta \quad \text{Var}(X) = \frac{k}{(k-1)^2(k-2)}\beta^2 \quad k > 1, 2$$



Pareto with $X \geq 0$

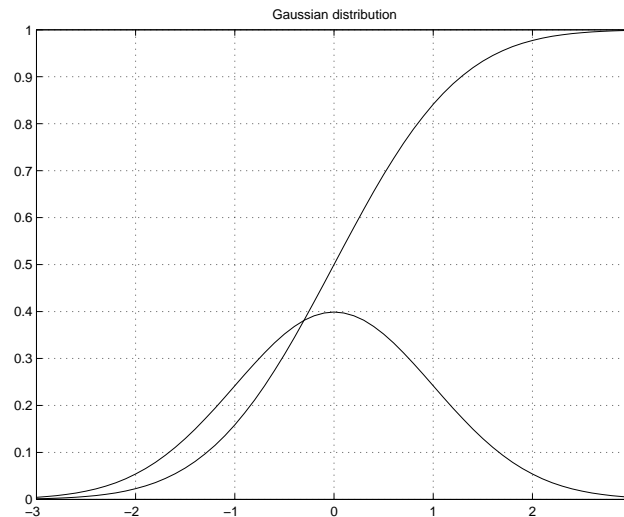
$$F(x) = 1 - \left(1 + \frac{x}{\beta}\right)^{-k} \quad X = \beta \left(U^{-\frac{1}{k}} - 1\right)$$

Gaussian

X a result of many (∞) independent sources (Central limit theorem)

$$X \sim N(\mu, \sigma^2)$$

$$Z \sim N(0, 1) \quad X = \mu + \sigma Z \quad Z = \Phi^{-1}(U)$$



Rayleigh Fordeling

- For $X_i \sim N(0, 1)$, define $R = \sqrt{X_1^2 + X_2^2}$. Now

$$F_R(r) = P(R \leq r) = 1 - e^{-\frac{r^2}{2}}$$

With $S = R^2 = X_1^2 + X_2^2$

$$F_S(s) = P(S \leq s) = 1 - e^{-\frac{s}{2}}$$

i.e. $S \sim \exp\left(\frac{1}{2}\right)$.

- $X_1 = R \cos(2\pi\theta)$ and $X_2 = R \sin(2\pi\Theta)$ for $\Theta \sim \text{Unif}(0; 1)$ independent of R .

Mathematical Method

- By means of transformation and other techniques we can obtain a random variable with a certain distribution.

The Box-Muller method A transformation from polar $(\theta = 2\pi U_2, r = \sqrt{-2 \log(U_1)})$ into Cartesian coordinates $(X = Z_1 \text{ and } Y = Z_2)$.

$$\begin{bmatrix} Z_1 \\ Z_2 \end{bmatrix} = \sqrt{-2 \log(U_1)} \begin{bmatrix} \cos(2\pi U_2) \\ \sin(2\pi U_2) \end{bmatrix} \quad Z_1, Z_2 \sim N(0, 1)$$

Central limit theorem

$$X = \sum_{i=1}^n U_i - \frac{n}{2} \quad \text{eg. } n = 6$$

Generation of cos and sin



Sine and cosine can be calculated by the following acceptance/rejection algorithm m:

1. Generate $V_1, V_2 \sim U(-1, 1)$
2. Generate $R^2 = V_1^2 + V_2^2$
3. If $R^2 > 1$ goto 1.
4. $\cos(2\pi U_2) = \frac{V_1}{R}, \sin(2\pi U_2) = \frac{V_2}{R}$

$\text{LN}(\alpha, \beta^2)$



Logarithmic Gaussian, $\text{LN}(\alpha, \beta^2)$

$$Y \sim \text{LN}(\alpha, \beta^2) \quad \log(Y) \sim \text{N}(\alpha, \beta^2)$$

$$Y = e^X \quad X = \alpha + \beta Z \quad Z \sim \text{N}(0, 1)$$

General and multivariate normal distribution

- Generate n independent values from an $N(0, 1)$ distribution, $Z_i \sim N(0, 1)$.
- $X_i = \mu_i + \sum_{j=1}^i c_{ij} Z_j$
- Where c_{ij} are the elements in the Cholesky factorisation of Σ , $\Sigma = CC'$

Composition methods - hyperexponential distribution



$$F(x) = 1 - \sum_{i=1}^m p_i e^{-\lambda_i x} = \sum_{i=1}^m p_i (1 - e^{-\lambda_i x})$$

Formally we can express

$Z = X_I$ where $I \sim \{1, 2, \dots, m\}$ with $P(I = i) = p_i$ and $X_I \sim \exp(\lambda_I)$

1. Choose $I \sim \{1, 2, \dots, m\}$ with probabilities p_i 's
2. $Z = -\frac{1}{\lambda_I} \log(U)$

Composition methods - Erlang distribution

- The Erlang distribution is a special case of the Gamma distribution with integer valued shape parameter
- An Erlang distributed random variable can be interpreted as a sum of independent exponential variables
- We can generate an Erlang- n distributed random variate by adding n exponential random variates.

$$Y \sim \text{Erl}_n(\lambda) \quad \mathbb{E}(Y) = \frac{n}{\lambda} \quad \text{Var}(Y) = \frac{n}{\lambda^2}$$

with $\lambda_i = \lambda$

$$Y = \sum_{i=1}^n X_i = \sum_{i=1}^n -\frac{1}{\lambda} \log(U_i) = -\frac{1}{\lambda} \log(\prod_{i=1}^n U_i)$$

Composition methods II



Generalization:

$$f(x) = \int f(x|y)f(y)dy$$

$$X \text{ given } Y : f(x|y) \quad Y : f(y)$$

Y is typically a parameter (eg. the conditional distribution of X given $Y = \mu$ is $N(\mu, \sigma^2)$)

Generate:

- Generate Y from $f(y)$.
- Generate X from $f(x|y)$ where Y is used.

Composition methods example of generalisation



$$f(y) = \mu e^{-\mu y}, \quad f_X(x|Y = y) = ye^{-yx}$$

$$f(x) = \int_0^{\infty} ye^{-yx} \mu e^{-\mu y} dy = \mu \int_0^{\infty} ye^{-(\mu+x)y} dy = \frac{\mu}{(\mu+x)^2}$$

$$= \frac{\frac{1}{\mu}}{\left(1 + \frac{x}{\mu}\right)^2}$$

$$F(x) = 1 - \left(1 + \frac{x}{\mu}\right)^{-1}$$

a Pareto distribution. The example can easily be generalised with gamma distributions, the algebra is slightly more involved.

Acceptance/rejection



Problem: we would like to generate X from pdf f , but it is much faster to generate Y

with pdf g . NB. X and Y have the same sample space. If

$$\frac{f(y)}{g(y)} \leq c \quad \text{for all } y \text{ and some } c$$

- Step 1. Generate Y having density g .
- Step 2. Generate a random number U
- If $U \leq \frac{f(Y)}{cg(Y)}$ set $X = Y$. Otherwise return to step 1.


$$g(y)dy \frac{f(y)}{cg(y)} = \frac{f(y)dy}{c}$$

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Random Number Generators.

- betarnd - Beta random numbers.
- binornd - Binomial random numbers.
- chi2rnd - Chi square random numbers.
- exprnd - Exponential random numbers.
- frnd - F random numbers.
- gamrnd - Gamma random numbers.
- geornd - Geometric random numbers.
- hygernd - Hypergeometric random numbers.
- iwishrnd - Inverse Wishart random matrix.
- lognrnd - Lognormal random numbers

Exercise 3

1. Generate simulated values from the following distributions 
 - (a) Exponential distribution
 - (b) Normal distribution (at least with standard Box-Mueller)
 - (c) Pareto distribution, with $\beta = 1$ and experiment with different values of k values: $k = 2.05$, $k = 2.5$, $k = 3$ and $k = 4$.

Verify the results by comparing histograms with analytical results and perform tests for distribution type.

2. For the Pareto distribution with support on $[\beta, \infty[$ compare mean value and variance, with analytical results, which can be calculated as $E(X) = \beta \frac{k}{k-1}$ (for $k > 1$) and $\text{Var}(X) = \beta^2 \frac{k}{(k-1)^2(k-2)}$ (for $k > 2$). Explain problems if any.
3. For the normal distribution generate 100 95% confidence intervals for the mean and variance, each based on 10 observations. Discuss the results.
4. Simulate from the Pareto distribution using composition.