

Impact of food availability, larval concentration and light on *P. lividus* larval growth - data analysis

```
library('dplyr')
library('readr')
library('ggplot2')
library('ggdist')
library('knitr')
library('tidyr')
library('emmeans')
library('rstan')
rstan::rstan_options(auto_write = TRUE)
library('brms')
options(mc.cores = parallel::detectCores()) # run all cores
library('bayesplot')
library('marginaleffects')
library('ggdist')
```

```
nchain = 4
niter = 2500
moment_matching = TRUE
```

1. Data overview

Making a data set containing only the data on *P. lividus*.

```
Pl_df <- read_delim("larval_morphology.csv", delim = ",",
                  col_types = "fffnfifffiniif")
Pl_df = drop_na(Pl_df, length)
# make Pl_df$larva by concatenating Pl_df$species with Pl_df$larva
Pl_df$larva <- as.factor(paste0(Pl_df$species, Pl_df$larva))
Pl_df <- Pl_df[Pl_df$species == "P1",]
Pl_df <- Pl_df[Pl_df$length > 0,]
Pl_df <- Pl_df[! is.na(Pl_df$length),]
#ensure correct order for levels
Pl_df<-Pl_df %>% mutate(lit = factor(lit, levels = c("DD", "LD", "LL")))
Pl_df<-Pl_df %>% mutate(rod = factor(rod, levels = c("BR", "PO", "ALA")))

print(paste0('There are ', dim(Pl_df)[1], ' measures from ', length(unique(Pl_df$larva)), ' individual larvae.'))
```

```
## [1] "There are 1152 measures from 211 individual larvae."
```

```
head(Pl_df)
```

```
## # A tibble: 6 x 15
##   larva side rod   length ate Food_conc Food_species fed lit condition
##   <fct> <fct> <fct>   <dbl> <fct>    <int> <fct>      <fct> <fct> <fct>
## 1 P11   R     BR     199. YES      10 D_tertiolecta Fed  DD  FSW
## 2 P11   R     PO     160. YES      10 D_tertiolecta Fed  DD  FSW
```

```
## 3 P11 R ALA 189. YES 10 D_tertiolecta Fed DD FSW
## 4 P11 L BR 195. YES 10 D_tertiolecta Fed DD FSW
## 5 P11 L PO 143. YES 10 D_tertiolecta Fed DD FSW
## 6 P11 L ALA 195. YES 10 D_tertiolecta Fed DD FSW
## # i 5 more variables: larvae_per_well <int>, lar_ml <dbl>, hpf <int>,
## # dpf <int>, species <fct>
```

For the statistical analysis a few transformation are required. First both length and larvae concentration (lar_ml) are scaled (L and C respectively).

```
meanL <- mean(P1_df$length)
sdL <- sd(P1_df$length)
meanC <- mean(P1_df$lar_ml)
sdC <- sd(P1_df$lar_ml)
P1_df$L <- as.numeric(scale(P1_df$length))
P1_df$C <- as.numeric(scale(P1_df$lar_ml))
P1_df <- droplevels(P1_df) # drop factor levels which are absent
head(P1_df)
```

```
## # A tibble: 6 x 17
## larva side rod length ate Food_conc Food_species fed lit condition
## <fct> <fct> <fct> <dbl> <fct> <int> <fct> <fct> <fct> <fct>
## 1 P11 R BR 199. YES 10 D_tertiolecta Fed DD FSW
## 2 P11 R PO 160. YES 10 D_tertiolecta Fed DD FSW
## 3 P11 R ALA 189. YES 10 D_tertiolecta Fed DD FSW
## 4 P11 L BR 195. YES 10 D_tertiolecta Fed DD FSW
## 5 P11 L PO 143. YES 10 D_tertiolecta Fed DD FSW
## 6 P11 L ALA 195. YES 10 D_tertiolecta Fed DD FSW
## # i 7 more variables: larvae_per_well <int>, lar_ml <dbl>, hpf <int>,
## # dpf <int>, species <fct>, L <dbl>, C <dbl>
```

The chunk below produces a data summary for each condition. In column n we calculated also the number of observations.

```
P1_df %>% group_by(species, dpf, condition, lit, lar_ml, rod, fed) %>%
  summarise(mean = mean(length, na.rm = TRUE), stdev = sd(length, na.rm = TRUE),
            n = n())
```

```
## `summarise()` has grouped output by 'species', 'dpf', 'condition', 'lit',
## 'lar_ml', 'rod'. You can override using the `.groups` argument.
```

```
## # A tibble: 36 x 10
## # Groups:   species, dpf, condition, lit, lar_ml, rod [18]
## species dpf condition lit lar_ml rod fed mean stdev n
## <fct> <int> <fct> <fct> <dbl> <fct> <fct> <dbl> <dbl> <int>
## 1 P1 3 FSW DD 12.5 BR Fed 186. 18.8 10
## 2 P1 3 FSW DD 12.5 BR Starved 183. 9.27 12
## 3 P1 3 FSW DD 12.5 PO Fed 143. 12.5 9
## 4 P1 3 FSW DD 12.5 PO Starved 150. 26.5 10
## 5 P1 3 FSW DD 12.5 ALA Fed 178. 13.5 10
## 6 P1 3 FSW DD 12.5 ALA Starved 171. 17.1 11
## 7 P1 3 FSW DD 25 BR Fed 189. 16.8 42
## 8 P1 3 FSW DD 25 BR Starved 185. 20.0 35
## 9 P1 3 FSW DD 25 PO Fed 135. 22.6 44
## 10 P1 3 FSW DD 25 PO Starved 147. 23.0 32
## # i 26 more rows
```

2. Experimental set-up and aim

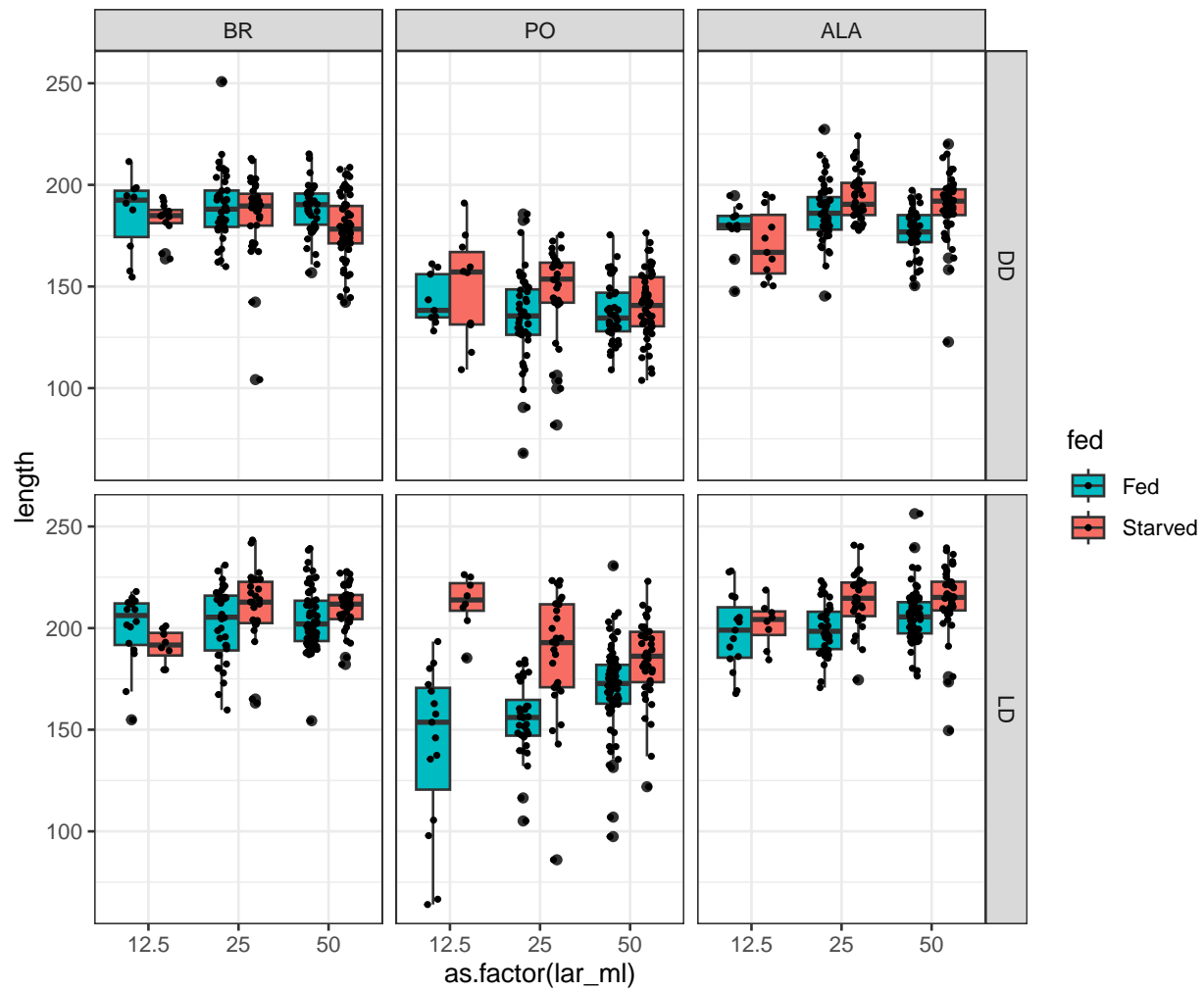
We are interested in investigating how the light dark cycle (lit) and the larval concentration (lar_ml) influence the phenotypic response to food availability (fed): larvae grow much shorter arms when food is abundant to save maternal storages; on the contrary, when food is scarce they develop much longer arms to maximize their capability to collect food. Three sets of spicules have been measured (rod): Body Rod (BR), Post Oral (PO), and Anterolateral (ALA) arms .

PLOT DATA OVERVIEW

```
plot1<- Pl_df %>%
  ggplot(aes(y=length,x=as.factor(lar_ml), fill=fed)) +
  facet_grid(lit~rod ) +
    geom_boxplot(position = position_dodge(width = 0.75)) +
    geom_jitter(position = position_jitterdodge(jitter.width = 0.25, dodge.width = 0.75), size=0.7)+
  scale_fill_manual(values=c("#00BBC1", "#F86D63"))+
  #geom_violin()+
  theme_bw()+
  theme(axis.text.x = element_text(angle = 0, hjust = 0.5)) +
  ggtitle("P. lividus, light and larval concentration impact on phenotypic response")

plot1
```

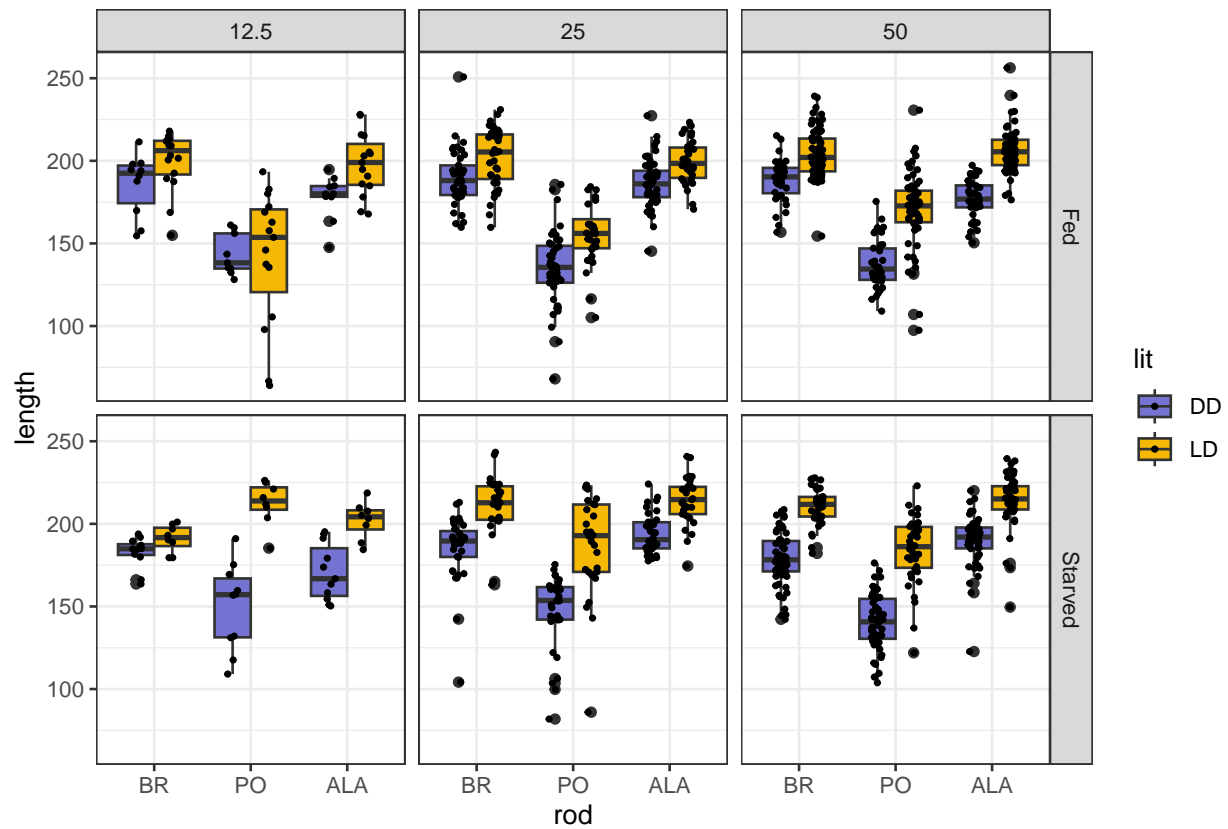
P. lividus, light and larval concentration impact on phenotypic response



```

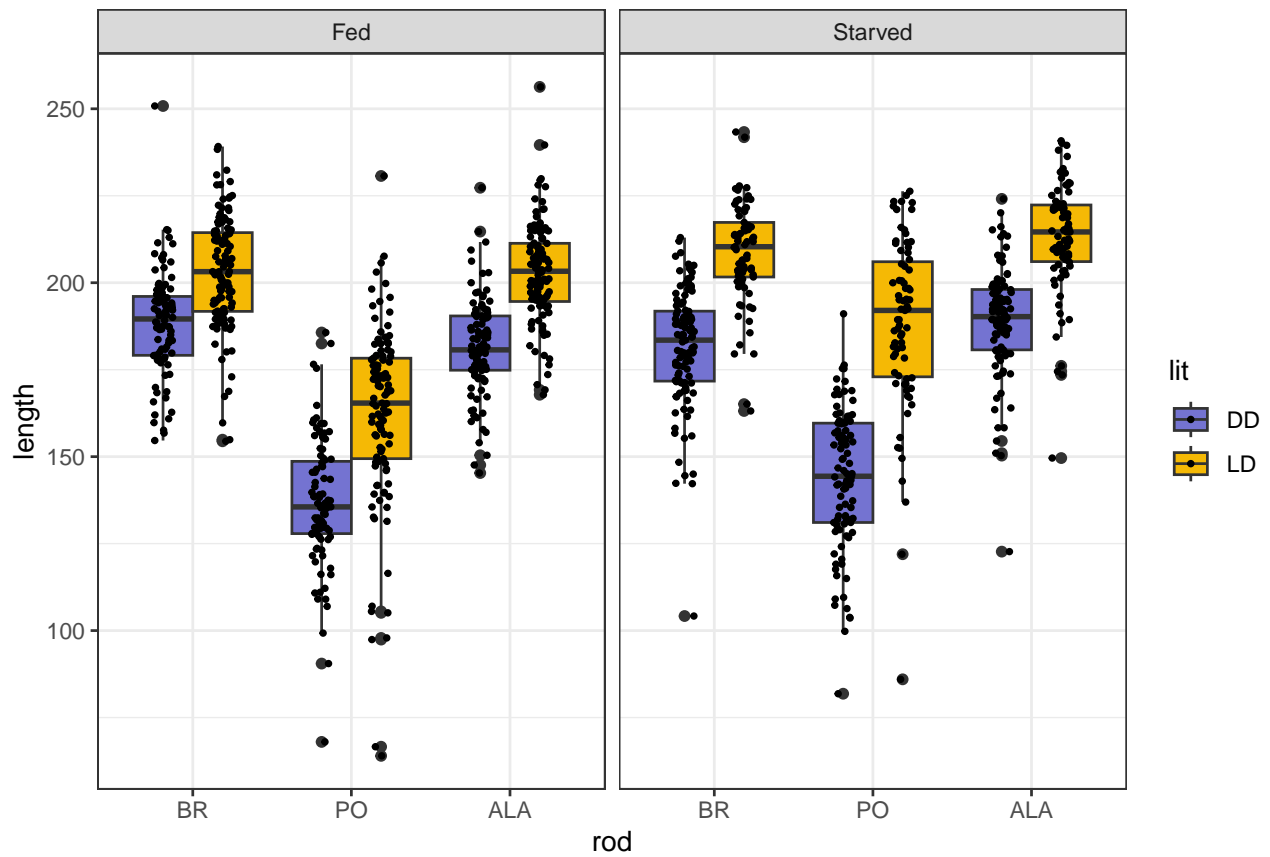
P1_df %>%
  ggplot(aes(y=length, x=rod, fill=lit, )) +
    facet_grid(fed~lar_ml) +
    geom_boxplot(position = position_dodge(width = 0.75)) +
    geom_jitter(position = position_jitterdodge(jitter.width = 0.25, dodge.width = 0.75), size=0.7) +
    scale_fill_manual(values=c("#7473d1", "#f5b905", "#d62222")) +
    #geom_violin() +
    theme_bw() +
    theme(axis.text.x = element_text(angle = 0, hjust = 0.5)) +
    ggtitle("P. lividus, impact of light on larval growth")
  
```

P. lividus, impact of light on larval growth



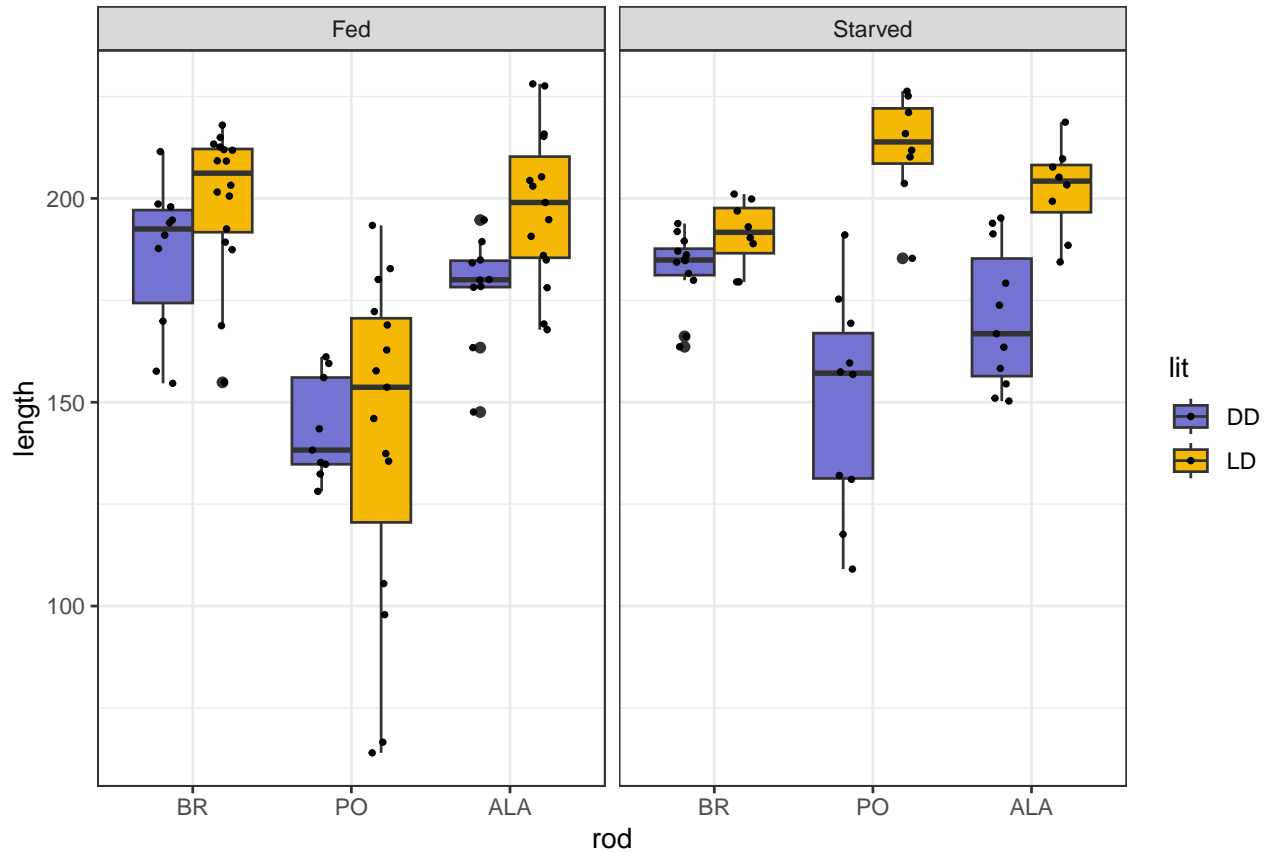
```
P1_df %>%
  ggplot(aes(y=length,x=rod, fill=lit, )) +
    facet_grid(~fed) +
    geom_boxplot(position = position_dodge(width = 0.75)) +
    geom_jitter(position = position_jitterdodge(jitter.width = 0.25, dodge.width = 0.75), size=0.7)+
    scale_fill_manual(values=c("#7473d1", "#f5b905", "#d62222" )) +
    theme_bw()+
    theme(axis.text.x = element_text(angle = 0, hjust = 0.5)) +
    ggtitle("P. lividus, impact of light on larval growth")
```

P. lividus, impact of light on larval growth



```
P1_df %>%
  filter(lar_ml==12.5) %>%
  ggplot(aes(y=length,x=rod, fill=lit, )) +
    facet_grid(~fed) +
    geom_boxplot(position = position_dodge(width = 0.75)) +
    geom_jitter(position = position_jitterdodge(jitter.width = 0.25, dodge.width = 0.75), size=0.7)+
    scale_fill_manual(values=c("#7473d1", "#f5b905", "#d62222" ))+
  #geom_violin()+
  theme_bw()+
  theme(axis.text.x = element_text(angle = 0, hjust = 0.5)) +
  ggtitle("P. lividus, impact of light on larval growth at 12.5 larvae/ml")
```

P. lividus, impact of light on larval growth at 12.5 larvae/ml



3. Statistics

Prior predictive test

Set priors on slope. Weakly informative priors are used to guide the model.

M0 Single intercept

Each treatment should be replicated sufficiently within each larva to capture the treatment effect accurately. Typically, having at least 3-5 measurements per treatment per larva can provide a reasonable balance between model complexity and data sufficiency.

Ideally, we would represent the data as a nested structure, with larvae nested within species. This would allow us to estimate the variance components for the species and larva levels. However, given the low number of measures per larva, it might be challenging to estimate these variance components reliably. Therefore, we will start with a simpler model that includes only the larva level as a random effect.

```
get_prior(bf(L ~ 1 + (1|larva), sigma ~ 1), data=P1_df)
```

##	prior	class	coef	group	resp	dpar	nlpar	lb	ub
##	student_t(3, 0.2, 2.5)	Intercept							
##	student_t(3, 0, 2.5)	sd						0	
##	student_t(3, 0, 2.5)	sd		larva				0	
##	student_t(3, 0, 2.5)	sd	Intercept	larva				0	
##	student_t(3, 0, 2.5)	Intercept				sigma			
##	source								

```
##      default
##      default
## (vectorized)
## (vectorized)
##      default

si_priors <- c(
  set_prior("normal(0, 0.5)", class = "Intercept"), # prior for intercept
  set_prior("student_t(3, 0, 0.5)", class = "sd"),
  set_prior("student_t(3, 0, 0.5)", class = "sd", group = "larva"), # Student's t prior for group-level
  set_prior("student_t(3, 0, 0.5)", dpar = "sigma", class = "Intercept") # Student's t prior for resid
)
```

The zeroth model is built to check the effect of the group level and see how the data is distributed overall. The nesting structure explicitly acknowledges the hierarchical nature of the data. This helps to avoid pseudoreplication and ensures that the estimates of variance components are not biased.

```
Pl_intercept_mod <- brm(
  bf(L ~ 1 + (1|larva), sigma ~ 1),
  family = gaussian,
  data=Pl_df, prior = si_priors,
  chains = nchain,
  iter = niter, warmup = niter/2,
  save_pars = save_pars(all = TRUE)
)
Pl_intercept_mod
```

```
## Family: gaussian
## Links: mu = identity; sigma = log
## Formula: L ~ 1 + (1 | larva)
##          sigma ~ 1
## Data: Pl_df (Number of observations: 1152)
## Draws: 4 chains, each with iter = 2500; warmup = 1250; thin = 1;
##          total post-warmup draws = 5000
##
## Multilevel Hyperparameters:
## ~larva (Number of levels: 211)
##          Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sd(Intercept)    0.47      0.04    0.40    0.55 1.00    2292    3149
##
## Regression Coefficients:
##          Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## Intercept      -0.00      0.04   -0.09    0.08 1.00    4211    3828
## sigma_Intercept -0.12      0.02   -0.16   -0.08 1.00    5969    3649
##
## Draws were sampled using sampling(NUTS). For each parameter, Bulk_ESS
## and Tail_ESS are effective sample size measures, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

```
get_prior(bf(L ~ rod + (1|larva), sigma ~ rod), data=Pl_df)
```

```
##          prior      class      coef group resp  dpar nlpar lb ub
##          (flat)         b          rodALA
##          (flat)         b          rodPO
## student_t(3, 0.2, 2.5) Intercept
```



```
##      student_t(3, 0, 2.5)      sd      0
##      student_t(3, 0, 2.5)      sd      larva      0
##      student_t(3, 0, 2.5)      sd Intercept larva      0
##              (flat)      b      sigma
##              (flat)      b      rodALA      sigma
##              (flat)      b      rodPO      sigma
##      student_t(3, 0, 2.5) Intercept      sigma
##      source
##      default
## (vectorized)
## (vectorized)
##      default
##      default
## (vectorized)
## (vectorized)
##      default
## (vectorized)
## (vectorized)
##      default
## (vectorized)
## (vectorized)
##      default

priors <- c(
  set_prior("student_t(5, 0, 2)", class = "Intercept"), # prior for intercept
  set_prior("normal(0, 2)", class = "b"),

  set_prior("student_t(3, 0, 0.5)", class = "sd", group = "larva"),

  set_prior("student_t(5, 0, 2)", dpar = "sigma", class = "Intercept"), # prior for residual sd
  set_prior("normal(0, 1)", class = "b", dpar="sigma")
)
```

M1 Rod model

```
Pl_rod_mod <- brm(
  bf(L ~ rod + (1|larva), sigma ~ rod),
  family = gaussian,
  data=Pl_df, prior = priors,
  chains = nchain,
  iter = niter, warmup = niter/2,
  save_pars = save_pars(all = TRUE)
)
Pl_rod_mod

## Family: gaussian
## Links: mu = identity; sigma = log
## Formula: L ~ rod + (1 | larva)
##      sigma ~ rod
## Data: Pl_df (Number of observations: 1152)
## Draws: 4 chains, each with iter = 2500; warmup = 1250; thin = 1;
##      total post-warmup draws = 5000
##
## Multilevel Hyperparameters:
## ~larva (Number of levels: 211)
##      Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sd(Intercept)      0.56      0.03      0.50      0.63 1.00      917      1660
```

```
##
## Regression Coefficients:
##           Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## Intercept          0.41      0.04    0.33    0.49 1.01      879    1649
## sigma_Intercept    -1.12      0.09   -1.31   -0.95 1.01      677    1347
## rodPO              -1.33      0.04   -1.42   -1.24 1.00     7099    3863
## rodALA              0.00      0.03   -0.06    0.06 1.00     7849    3814
## sigma_rodPO         0.90      0.11    0.69    1.13 1.01      720    1605
## sigma_rodALA        0.43      0.14    0.17    0.70 1.01      693    1425
##
## Draws were sampled using sampling(NUTS). For each parameter, Bulk_ESS
## and Tail_ESS are effective sample size measures, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

M2 C Rod model

```
Pl_rod_C_mod <- brm(
  bf(L ~ rod*C + (1|larva), sigma ~ rod),
  family = gaussian,
  data=Pl_df, prior = priors,
  chains = nchain,
  iter = niter, warmup = niter/2,
  save_pars = save_pars(all = TRUE)
)
Pl_rod_C_mod

## Family: gaussian
## Links: mu = identity; sigma = log
## Formula: L ~ rod * C + (1 | larva)
##           sigma ~ rod
## Data: Pl_df (Number of observations: 1152)
## Draws: 4 chains, each with iter = 2500; warmup = 1250; thin = 1;
##           total post-warmup draws = 5000
##
## Multilevel Hyperparameters:
## ~larva (Number of levels: 211)
##           Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sd(Intercept)    0.56      0.03    0.50    0.63 1.01      775    1873
##
## Regression Coefficients:
##           Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## Intercept          0.41      0.04    0.33    0.50 1.00      592    1295
## sigma_Intercept    -1.11      0.09   -1.30   -0.93 1.01      545    1085
## rodPO              -1.33      0.05   -1.42   -1.24 1.00     6722    3445
## rodALA              0.00      0.03   -0.06    0.06 1.00     6494    3719
## C                   0.02      0.04   -0.07    0.10 1.00      670    1233
## rodPO:C             0.04      0.05   -0.05    0.13 1.00     7031    4041
## rodALA:C            0.05      0.03   -0.01    0.11 1.00     6711    3491
## sigma_rodPO        0.89      0.11    0.68    1.13 1.01      600    1172
## sigma_rodALA        0.41      0.14    0.15    0.68 1.01      550    1132
##
## Draws were sampled using sampling(NUTS). For each parameter, Bulk_ESS
## and Tail_ESS are effective sample size measures, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

M3 C Rod Fed model

```
Pl_rod_C_fed_mod <- brm(
  bf(L ~ rod*C*fed + (1|larva), sigma ~ rod*fed),
  family = gaussian,
  data=Pl_df, prior = priors,
  chains = nchain,
  iter = niter, warmup = niter/2,
  save_pars = save_pars(all = TRUE)
)
```

```
## Warning: Bulk Effective Samples Size (ESS) is too low, indicating posterior means and medians may be
## Running the chains for more iterations may help. See
## https://mc-stan.org/misc/warnings.html#bulk-ess
```

```
Pl_rod_C_fed_mod
```

```
## Family: gaussian
## Links: mu = identity; sigma = log
## Formula: L ~ rod * C * fed + (1 | larva)
##          sigma ~ rod * fed
## Data: Pl_df (Number of observations: 1152)
## Draws: 4 chains, each with iter = 2500; warmup = 1250; thin = 1;
##          total post-warmup draws = 5000
##
## Multilevel Hyperparameters:
## ~larva (Number of levels: 211)
##           Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sd(Intercept)      0.59      0.03    0.53    0.65 1.00      767    1517
##
## Regression Coefficients:
##           Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS
## Intercept           0.47      0.06    0.35    0.58 1.00      743
## sigma_Intercept     -1.05      0.09   -1.23   -0.86 1.00     1051
## rodPO              -1.57      0.06   -1.68   -1.46 1.00     4204
## rodALA             -0.14      0.04   -0.22   -0.07 1.00     4225
## C                   0.05      0.06   -0.07    0.16 1.01      507
## fedStarved         -0.11      0.09   -0.27    0.06 1.00      526
## rodPO:C              0.19      0.05    0.09    0.30 1.00     3982
## rodALA:C             0.02      0.04   -0.06    0.09 1.00     3941
## rodPO:fedStarved     0.54      0.09    0.36    0.71 1.00     3965
## rodALA:fedStarved    0.31      0.06    0.18    0.43 1.00     4384
## C:fedStarved        -0.06      0.09   -0.24    0.11 1.01      341
## rodPO:C:fedStarved  -0.36      0.09   -0.52   -0.19 1.00     4116
## rodALA:C:fedStarved  0.05      0.06   -0.07    0.17 1.00     3990
## sigma_rodPO          0.72      0.12    0.49    0.96 1.00     1162
## sigma_rodALA          0.16      0.14   -0.12    0.45 1.00     1031
## sigma_fedStarved     -0.46      0.15   -0.74   -0.14 1.00      787
## sigma_rodPO:fedStarved 0.60      0.19    0.23    0.96 1.00      862
## sigma_rodALA:fedStarved 0.87      0.21    0.44    1.27 1.00      816
##
## Tail_ESS
## Intercept           1417
## sigma_Intercept     1931
## rodPO               3879
## rodALA              3737
```

```
## C 1026
## fedStarved 942
## rodP0:C 3592
## rodALA:C 3640
## rodP0:fedStarved 3342
## rodALA:fedStarved 3360
## C:fedStarved 865
## rodP0:C:fedStarved 3697
## rodALA:C:fedStarved 3887
## sigma_rodP0 2203
## sigma_rodALA 1792
## sigma_fedStarved 1246
## sigma_rodP0:fedStarved 1471
## sigma_rodALA:fedStarved 1286
##
## Draws were sampled using sampling(NUTS). For each parameter, Bulk_ESS
## and Tail_ESS are effective sample size measures, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

M4 C Rod Lit model

```
Pl_rod_C_lit_mod <- brm(
  bf(L ~ rod*C*lit + (1|larva), sigma ~ rod*lit),
  family = gaussian,
  data=Pl_df, prior = priors,
  chains = nchain,
  iter = niter, warmup = niter/2,
  save_pars = save_pars(all = TRUE)
)
Pl_rod_C_lit_mod

## Family: gaussian
## Links: mu = identity; sigma = log
## Formula: L ~ rod * C * lit + (1 | larva)
##          sigma ~ rod * lit
## Data: Pl_df (Number of observations: 1152)
## Draws: 4 chains, each with iter = 2500; warmup = 1250; thin = 1;
##          total post-warmup draws = 5000
##
## Multilevel Hyperparameters:
## ~larva (Number of levels: 211)
##          Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sd(Intercept)    0.38      0.03    0.32    0.44 1.00      610     1370
##
## Regression Coefficients:
##          Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## Intercept         0.07      0.05   -0.03    0.16 1.00      2368     3301
## sigma_Intercept   -0.82      0.08   -0.97   -0.67 1.00      2482     3121
## rodP0             -1.51      0.05   -1.62   -1.41 1.00      5500     4217
## rodALA            -0.00      0.04   -0.09    0.09 1.00      4655     3799
## C                 -0.06      0.05   -0.15    0.04 1.00      1856     2828
## litLD              0.70      0.06    0.58    0.83 1.00      2166     2966
## rodP0:C             0.03      0.05   -0.08    0.13 1.00      4562     4029
## rodALA:C           0.06      0.04   -0.03    0.15 1.00      4148     3802
```

```
## rodPO:litLD      0.38      0.09      0.21      0.56 1.00      7314      4378
## rodALA:litLD     0.03      0.06     -0.09      0.16 1.00      4616      3697
## C:litLD          0.14      0.07      0.01      0.27 1.00      1764      2448
## rodPO:C:litLD    0.01      0.09     -0.16      0.18 1.00      5958      4284
## rodALA:C:litLD   -0.04      0.06     -0.16      0.09 1.00      4525      3989
## sigma_rodPO      0.30      0.10      0.10      0.51 1.00      2831      3738
## sigma_rodALA      0.00      0.12     -0.22      0.23 1.00      2663      3610
## sigma_litLD      -0.38      0.18     -0.75     -0.07 1.00        608      1172
## sigma_rodPO:litLD 0.82      0.21      0.43      1.26 1.00        674      1334
## sigma_rodALA:litLD 0.50      0.25      0.05      1.01 1.00        671      1404
##
## Draws were sampled using sampling(NUTS). For each parameter, Bulk_ESS
## and Tail_ESS are effective sample size measures, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

M5 C Rod Fed Lit model

```
Pl_rod_C_fed_lit_mod <- brm(
  bf(L ~ rod*C*fed*lit + (1|larva), sigma ~ rod*fed*lit),
  family = gaussian,
  data=Pl_df, prior = priors,
  chains = nchain,
  iter = niter, warmup = niter/2,
  save_pars = save_pars(all = TRUE)
)
Pl_rod_C_fed_lit_mod

## Family: gaussian
## Links: mu = identity; sigma = log
## Formula: L ~ rod * C * fed * lit + (1 | larva)
##          sigma ~ rod * fed * lit
## Data: Pl_df (Number of observations: 1152)
## Draws: 4 chains, each with iter = 2500; warmup = 1250; thin = 1;
##        total post-warmup draws = 5000
##
## Multilevel Hyperparameters:
## ~larva (Number of levels: 211)
##           Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
## sd(Intercept)    0.40      0.03    0.34    0.45 1.00      850      1198
##
## Regression Coefficients:
##           Estimate Est.Error 1-95% CI u-95% CI Rhat
## Intercept         0.21      0.07    0.06    0.35 1.00
## sigma_Intercept   -0.87      0.09   -1.05   -0.69 1.00
## rodPO             -1.76      0.07   -1.91   -1.62 1.00
## rodALA            -0.27      0.06   -0.39   -0.16 1.00
## C                 -0.01      0.07   -0.16    0.13 1.00
## fedStarved        -0.26      0.10   -0.45   -0.07 1.00
## litLD             0.47      0.09    0.29    0.66 1.00
## rodPO:C           0.02      0.08   -0.14    0.17 1.00
## rodALA:C          -0.10      0.06   -0.22    0.01 1.00
## rodPO:fedStarved  0.47      0.10    0.27    0.68 1.00
## rodALA:fedStarved 0.48      0.08    0.31    0.65 1.00
## C:fedStarved     -0.07      0.10   -0.25    0.13 1.00
```

## rodPO:litLD	0.34	0.11	0.12	0.55	1.00
## rodALA:litLD	0.24	0.08	0.08	0.39	1.00
## C:litLD	0.07	0.09	-0.11	0.25	1.00
## fedStarved:litLD	0.47	0.13	0.20	0.72	1.00
## rodPO:C:fedStarved	-0.03	0.11	-0.23	0.18	1.00
## rodALA:C:fedStarved	0.26	0.09	0.09	0.43	1.00
## rodPO:C:litLD	0.28	0.11	0.06	0.50	1.00
## rodALA:C:litLD	0.18	0.08	0.03	0.33	1.00
## rodPO:fedStarved:litLD	0.24	0.17	-0.09	0.58	1.00
## rodALA:fedStarved:litLD	-0.31	0.13	-0.55	-0.06	1.00
## C:fedStarved:litLD	0.15	0.13	-0.11	0.40	1.00
## rodPO:C:fedStarved:litLD	-0.61	0.17	-0.94	-0.28	1.00
## rodALA:C:fedStarved:litLD	-0.40	0.13	-0.66	-0.15	1.00
## sigma_rodPO	0.32	0.13	0.06	0.57	1.00
## sigma_rodALA	-0.19	0.14	-0.47	0.09	1.00
## sigma_fedStarved	-0.13	0.16	-0.45	0.18	1.00
## sigma_litLD	-0.24	0.16	-0.58	0.06	1.00
## sigma_rodPO:fedStarved	0.19	0.20	-0.19	0.61	1.00
## sigma_rodALA:fedStarved	0.51	0.23	0.07	0.98	1.00
## sigma_rodPO:litLD	0.56	0.20	0.17	0.98	1.00
## sigma_rodALA:litLD	0.45	0.24	-0.01	0.94	1.00
## sigma_fedStarved:litLD	-0.63	0.26	-1.11	-0.08	1.00
## sigma_rodPO:fedStarved:litLD	0.71	0.32	0.05	1.30	1.00
## sigma_rodALA:fedStarved:litLD	0.68	0.35	-0.04	1.33	1.00
##	Bulk_ESS	Tail_ESS			
## Intercept	1369	1902			
## sigma_Intercept	2481	3319			
## rodPO	2513	3828			
## rodALA	2283	3228			
## C	1290	1777			
## fedStarved	1420	1922			
## litLD	1381	2050			
## rodPO:C	2743	3610			
## rodALA:C	2494	3365			
## rodPO:fedStarved	3230	3823			
## rodALA:fedStarved	2566	2993			
## C:fedStarved	1324	2315			
## rodPO:litLD	3070	3676			
## rodALA:litLD	2467	3422			
## C:litLD	1343	2112			
## fedStarved:litLD	1222	1845			
## rodPO:C:fedStarved	2950	3803			
## rodALA:C:fedStarved	2850	3497			
## rodPO:C:litLD	2918	3683			
## rodALA:C:litLD	2715	3455			
## rodPO:fedStarved:litLD	4099	4102			
## rodALA:fedStarved:litLD	2969	3725			
## C:fedStarved:litLD	1261	1985			
## rodPO:C:fedStarved:litLD	3701	4022			
## rodALA:C:fedStarved:litLD	3396	3697			
## sigma_rodPO	2425	2951			
## sigma_rodALA	2731	3629			
## sigma_fedStarved	1234	1931			
## sigma_litLD	1114	2066			

```
## sigma_rodPO:fedStarved      1436      2269
## sigma_rodALA:fedStarved     1381      2304
## sigma_rodPO:litLD           1291      2148
## sigma_rodALA:litLD          1184      1982
## sigma_fedStarved:litLD       1015      1294
## sigma_rodPO:fedStarved:litLD 1182      1397
## sigma_rodALA:fedStarved:litLD 1211      1496
##
## Draws were sampled using sampling(NUTS). For each parameter, Bulk_ESS
## and Tail_ESS are effective sample size measures, and Rhat is the potential
## scale reduction factor on split chains (at convergence, Rhat = 1).
```

4. Models evaluation using LOO

```
Pl_intercept_mod = add_criterion(Pl_intercept_mod, criterion = "loo",
                                moment_match = TRUE, relloo = TRUE)
```

```
## No problematic observations found. Returning the original 'loo' object.
```

```
Pl_rod_mod = add_criterion(Pl_rod_mod, criterion = "loo",
                           moment_match = TRUE, relloo = TRUE)
```

```
## Warning: Some Pareto k diagnostic values are too high. See help('pareto-k-diagnostic') for details.
```

```
## 7 problematic observation(s) found.
```

```
## The model will be refit 7 times.
```

```
##
```

```
## Fitting model 1 out of 7 (leaving out observation 114)
```

```
## Start sampling
```

```
##
```

```
## Fitting model 2 out of 7 (leaving out observation 279)
```

```
## Start sampling
```

```
##
```

```
## Fitting model 3 out of 7 (leaving out observation 340)
```

```
## Start sampling
```

```
##
```

```
## Fitting model 4 out of 7 (leaving out observation 440)
```

```
## Start sampling
```

```
##
```

```
## Fitting model 5 out of 7 (leaving out observation 490)
```

```
## Start sampling
```

```
##
```

```
## Fitting model 6 out of 7 (leaving out observation 683)
```

```
## Start sampling
```

```
##
```

```
## Fitting model 7 out of 7 (leaving out observation 1021)
```

```
## Start sampling
```

```
Pl_rod_C_mod = add_criterion(Pl_rod_C_mod, criterion = "loo",  
                             moment_match = TRUE, relloo = TRUE)
```

```
## Warning: Some Pareto k diagnostic values are too high. See help('pareto-k-diagnostic') for details.
```

```
## 5 problematic observation(s) found.
```

```
## The model will be refit 5 times.
```

```
##
```

```
## Fitting model 1 out of 5 (leaving out observation 114)
```

```
## Start sampling
```

```
##
```

```
## Fitting model 2 out of 5 (leaving out observation 215)
```

```
## Start sampling
```

```
##
```

```
## Fitting model 3 out of 5 (leaving out observation 279)
```

```
## Start sampling
```

```
##
```

```
## Fitting model 4 out of 5 (leaving out observation 508)
```

```
## Start sampling
```

```
##
```

```
## Fitting model 5 out of 5 (leaving out observation 1021)
```

```
## Start sampling
```

```
Pl_rod_C_lit_mod = add_criterion(Pl_rod_C_lit_mod, criterion = "loo",  
                                 moment_match = TRUE, relloo = TRUE)
```

```
## Warning: Some Pareto k diagnostic values are too high. See help('pareto-k-diagnostic') for details.
```

```
## 6 problematic observation(s) found.
```

```
## The model will be refit 6 times.
```

```
##
```

```
## Fitting model 1 out of 6 (leaving out observation 114)
```

```
## Start sampling
```

```
##
```

```
## Fitting model 2 out of 6 (leaving out observation 279)
```

```
## Start sampling
```

```
##
```

```
## Fitting model 3 out of 6 (leaving out observation 683)
```

```
## Start sampling
```

```
##
```

```
## Fitting model 4 out of 6 (leaving out observation 733)
```

```
## Start sampling
```

```
##
```

```
## Fitting model 5 out of 6 (leaving out observation 736)
```

```
## Start sampling
```



```

##
## Fitting model 6 out of 6 (leaving out observation 1021)
## Start sampling
Pl_rod_C_fed_mod = add_criterion(Pl_rod_C_fed_mod, criterion = "loo",
                                moment_match = TRUE, relloo = TRUE)

## Warning: Some Pareto k diagnostic values are too high. See help('pareto-k-diagnostic') for details.
## 12 problematic observation(s) found.
## The model will be refit 12 times.

##
## Fitting model 1 out of 12 (leaving out observation 114)
## Start sampling

##
## Fitting model 2 out of 12 (leaving out observation 215)
## Start sampling

##
## Fitting model 3 out of 12 (leaving out observation 218)
## Start sampling

##
## Fitting model 4 out of 12 (leaving out observation 279)
## Start sampling

##
## Fitting model 5 out of 12 (leaving out observation 432)
## Start sampling

##
## Fitting model 6 out of 12 (leaving out observation 440)
## Start sampling

##
## Fitting model 7 out of 12 (leaving out observation 490)
## Start sampling

##
## Fitting model 8 out of 12 (leaving out observation 505)
## Start sampling

##
## Fitting model 9 out of 12 (leaving out observation 508)
## Start sampling

##
## Fitting model 10 out of 12 (leaving out observation 645)
## Start sampling

##
## Fitting model 11 out of 12 (leaving out observation 683)
## Start sampling

```

```

##
## Fitting model 12 out of 12 (leaving out observation 1021)
## Start sampling
Pl_rod_C_fed_lit_mod = add_criterion(Pl_rod_C_fed_lit_mod, criterion = "loo",
                                     moment_match = TRUE, relloo = TRUE)

## Warning: Some Pareto k diagnostic values are too high. See help('pareto-k-diagnostic') for details.
## 20 problematic observation(s) found.
## The model will be refit 20 times.

##
## Fitting model 1 out of 20 (leaving out observation 114)
## Start sampling

##
## Fitting model 2 out of 20 (leaving out observation 279)
## Start sampling

##
## Fitting model 3 out of 20 (leaving out observation 495)
## Start sampling

##
## Fitting model 4 out of 20 (leaving out observation 508)
## Start sampling

##
## Fitting model 5 out of 20 (leaving out observation 642)
## Start sampling

##
## Fitting model 6 out of 20 (leaving out observation 645)
## Start sampling

##
## Fitting model 7 out of 20 (leaving out observation 683)
## Start sampling

##
## Fitting model 8 out of 20 (leaving out observation 716)
## Start sampling

##
## Fitting model 9 out of 20 (leaving out observation 728)
## Start sampling

##
## Fitting model 10 out of 20 (leaving out observation 733)
## Start sampling

##
## Fitting model 11 out of 20 (leaving out observation 736)
## Start sampling

```

```
##
## Fitting model 12 out of 20 (leaving out observation 782)
## Start sampling
##
## Fitting model 13 out of 20 (leaving out observation 785)
## Start sampling
##
## Fitting model 14 out of 20 (leaving out observation 800)
## Start sampling
##
## Fitting model 15 out of 20 (leaving out observation 803)
## Start sampling
##
## Fitting model 16 out of 20 (leaving out observation 807)
## Start sampling
##
## Fitting model 17 out of 20 (leaving out observation 810)
## Start sampling
##
## Fitting model 18 out of 20 (leaving out observation 813)
## Start sampling
##
## Fitting model 19 out of 20 (leaving out observation 991)
## Start sampling
##
## Fitting model 20 out of 20 (leaving out observation 1021)
## Start sampling
```

```
# Perform LOO comparison
loo_results <- loo_compare(Pl_intercept_mod,
                           Pl_rod_mod,
                           Pl_rod_C_mod,
                           Pl_rod_C_fed_mod,
                           Pl_rod_C_lit_mod,
                           Pl_rod_C_fed_lit_mod
                           )

loo_results
```

```
##               elpd_diff se_diff
## Pl_rod_C_fed_lit_mod      0.0      0.0
## Pl_rod_C_fed_mod        -24.7     15.6
## Pl_rod_C_lit_mod         -64.8     16.2
## Pl_rod_mod              -81.2     19.5
## Pl_rod_C_mod            -87.8     19.2
## Pl_intercept_mod       -672.2    33.5
```

```

best_model_name <- rownames(loo_results)[1]
best_model <- get(best_model_name)

# Save the best model to an RDS file
saveRDS(best_model, file = paste0(
  "./model_objects/", best_model_name, ".rds"))

# Print the name of the best model
print(paste("The best model is:", best_model_name))

## [1] "The best model is: Pl_rod_C_fed_lit_mod"

```

Model Equation

The model assumes a normal distribution for the response variable:

$$Y_i \sim \mathcal{N}(\mu_i, \sigma_i)$$

where:

Linear Predictor for the Mean (μ_i):

$$\mu_i = \beta_0 + X_i\beta + u_{J_1[i]}Z_{1,i}$$

- β_0 (Intercept): The population-level intercept.
- $X_i\beta$: Fixed effects (population-level predictors) with centered design matrix.
- $u_{J_1[i]}$: Random effect for group-level predictor, where $J_1[i]$ is the grouping index.
- $Z_{1,i}$: Group-level predictor values.
- $u_{J_1[i]}$ follows a normal distribution:

$$u_{J_1[i]} \sim \mathcal{N}(0, \sigma_u)$$

where σ_u is the standard deviation of the group-level effect.

Linear Predictor for the Standard Deviation (σ_i):

$$\log(\sigma_i) = \alpha_0 + X_{\sigma,i}\alpha$$

- α_0 (Intercept_sigma): Population-level intercept for the variance structure.
- $X_{\sigma,i}\alpha$: Fixed effects for the variance model.

Prior Distributions:

$$\beta \sim \mathcal{N}(0, 2), \quad \beta_0 \sim t_5(0, 2)$$

$$\alpha \sim \mathcal{N}(0, 1), \quad \alpha_0 \sim t_5(0, 2)$$

$$\sigma_u \sim t_3(0, 0.5)$$

This structure allows the model to estimate both the mean and the variance of growth (Y) while accounting for hierarchical effects from group-level predictors.

```

# Save the brms model to an RDS file
#saveRDS(best_model, file = paste0("./model_objects/", best_model_name, ".rds"))

```