

# Formulas for the model built on catch declarations

## 1 Notations

We model observations at the level of catch declarations  $D_j$  as a sum of  $Y_{ij}$  lognormal observations each one realised at one fishing position  $x_{ij}$ .

We note :

- $\mathcal{P}_j$  : the vector of all fishing positions related to the  $j^{th}$  declaration
- $j \in \{1, \dots, n\}$  with  $n$  the number of declarations
- $i \in \{1, \dots, m_j\}$  with  $m_j$  the number of fishing positions belonging to the  $j^{th}$  declaration.

$$D_j = \sum_{i \in \mathcal{P}_j} Y_{ij}$$

## 2 Alternative observation models accounting for reallocation

### 2.1 First assumption about convolution of lognormal random variables

We assume a sum of lognormal random variables is still lognormal and that  $D_j \sim \mathcal{LN}(M(D_j), \Phi(D_j))$  with  $M(\cdot)$  being a function representing the central value of  $D_j$  and  $\Phi(\cdot)$  the variance function of  $D_j$ .

### 2.2 Reparameterization of the lognormal distribution

The standard lognormal distribution is written as :

$$D \sim \mathcal{LN}(\rho; \sigma^2) \text{ with } D = e^{\rho + \sigma N} \text{ and } N \sim \mathcal{N}(0, 1)$$

In this case,  $E(D) = e^{\rho + \frac{\sigma^2}{2}}$  and  $Var(D) = (e^{\sigma^2} - 1)e^{2\rho + \sigma^2}$ .

We choose to slightly reparameterize the lognormal distribution so that  $E(D)$  and  $Var(D)$  have more simple expressions.

Let's define  $\rho = \ln(\mu) - \frac{\sigma^2}{2}$ , then :

- $D = \mu e^{\sigma N - \frac{\sigma^2}{2}}$
- $E(D) = \mu$
- $Var(D) = \mu^2(e^{\sigma^2} - 1) \Leftrightarrow \sigma^2 = \ln(\frac{Var(D)}{E(D)^2} + 1)$

Then,  $D \sim \mathcal{LN}(M(D), \Phi(D))$  with  $M(D) = E(D)$  and  $\Phi(D) = \ln(\frac{Var(D)}{E(D)^2} + 1)$ .

### 2.3 $D_j$ probability distribution and moments

We have to express the probability distribution of  $D_j$  and its moments as a function of  $Y_{ij}$  with  $i \in \mathcal{P}_j$ .

Let's assume  $Y_{ij} = C_{ij} \cdot Z_{ij}$  is a zero-inflated lognormal distribution with  $C_{ij}$  a binary random variable and  $Z_{ij}$  a lognormal random variable.

$$C_{ij} \sim \mathcal{B}(1 - p_{ij})$$

with  $p_{ij} = \exp(-e^\xi \cdot S(x_{ij}))$  the probability to obtain a zero value

$$Z_{ij} \sim \mathcal{LN}\left(\frac{S(x_{ij})}{1 - p_{ij}}, \sigma^2\right)$$

Here,  $Y_{ij}$ ,  $C_{ij}$  and  $Z_{ij}$  are observations of a latent field  $S(x_{ij})$  at a sampled point  $x_{ij}$ .

#### 2.3.1 Probability of obtaining a zero declaration

$$\begin{aligned} P(D_j = 0 | S, X) &= \prod_{i \in \mathcal{P}_j} P(Y_{ij} = 0 | S, X), \\ &= \exp \left\{ - \sum_{i \in \mathcal{P}_j} e^\xi \cdot S(x_{ij}) \right\} = \pi_j. \end{aligned}$$

#### 2.3.2 Expectancy of a positive declaration

Following calculations are supposed to be conditionnal on  $S$  and  $X$ .

$$E(D_j | D_j > 0) = \sum_{i \in \mathcal{P}_j} E(C_{ij} Z_{ij} | \exists i \in \mathcal{P}_j, C_{ij} = 1)$$

As  $C_{ij}$  and  $Z_{ij}$  are assumed to be independant.

$$\begin{aligned} E(D_j | D_j > 0) &= E(D_j 1_{\{D_j > 0\}}) / P(D_j > 0), \\ &= E(D_j 1_{\{D_j > 0\}}) / (1 - \pi_j). \end{aligned}$$

As  $E(D_j 1_{\{D_j > 0\}}) = E(D_j)$ ,

$$\begin{aligned} E(D_j | D_j > 0) &= (1 - \pi_j)^{-1} E(D_j), \\ &= (1 - \pi_j)^{-1} \sum_{i \in \mathcal{P}_j} E(C_{ij} Z_{ij}), \\ &= (1 - \pi_j)^{-1} \sum_{i \in \mathcal{P}_j} (1 - p_{ij}) \frac{S(x_{ij})}{1 - p_{ij}}, \\ &= (1 - \pi_j)^{-1} \sum_{i \in \mathcal{P}_j} S(x_{ij}). \end{aligned}$$

### 2.3.3 Variance of a positive declaration

$$\text{Var}(D_j|D_j > 0) = E(D_j^2|D_j > 0) - E(D_j|D_j > 0)^2.$$

$$E(D_j^2|D_j > 0) = (1 - \pi_j)^{-1} E(D_j^2 1_{\{D_j > 0\}}) = (1 - \pi_j)^{-1} E(D_j^2)$$

$$E(D_j|D_j > 0)^2 = ((1 - \pi_j)^{-1} E(D_j 1_{\{D_j > 0\}}))^2 = (1 - \pi_j)^{-2} E(D_j)^2$$

And then,

$$\text{Var}(D_j|D_j > 0) = (1 - \pi_j)^{-1} E(D_j^2) - (1 - \pi_j)^{-2} E(D_j)^2 = (1 - \pi_j)^{-1} \text{Var}(D_j) - \frac{\pi_j}{(1 - \pi_j)^2} E(D_j)^2.$$

As the  $(Y_{ij})_{i \in \mathcal{P}_j}$  are independent,  $\text{Var}(D_j) = \sum_{i \in \mathcal{P}_j} \text{Var}(Y_{ij}) = \sum_{i \in \mathcal{P}_j} \text{Var}(C_{ij} \cdot Z_{ij})$ .

$$\begin{aligned} \text{Var}(C_{ij} Z_{ij}) &= E(C_{ij}^2 Z_{ij}^2) - E(C_{ij} Z_{ij})^2, \\ &= E(C_{ij}^2) E(Z_{ij}^2) - E(C_{ij})^2 E(Z_{ij})^2, \\ &= (1 - p_{ij}) E(Z_{ij}^2) - (1 - p_{ij})^2 E(Z_{ij})^2, \\ &= (1 - p_{ij}) (\text{Var}(Z_{ij}) + E(Z_{ij})^2) - (1 - p_{ij})^2 E(Z_{ij})^2, \\ &= \frac{S(x_{ij})^2}{1 - p_{ij}} (e^{\sigma^2} - 1) + \frac{S(x_{ij})^2}{1 - p_{ij}} - S(x_{ij})^2, \\ &= \frac{S(x_{ij})^2}{1 - p_{ij}} (e^{\sigma^2} - (1 - p_{ij})) \end{aligned}$$

### 2.3.4 Conclusion

$$P(D_j = 0|S, X) = \exp \left\{ - \sum_{i \in \mathcal{P}_j} e^{\xi} \cdot S(x_{ij}) \right\} = \pi_j$$

$$D_j|D_j > 0 \sim \mathcal{LN}(E(D_j|D_j > 0); \ln(\frac{\text{Var}(D_j|D_j > 0)}{E(D_j|D_j > 0)^2} + 1))$$

$$E(D_j|D_j > 0) = \frac{\sum_{i \in \mathcal{P}_j} S(x_{ij})}{1 - \pi_j}$$

$$\text{Var}(D_j|D_j > 0) = \frac{\sum_{i \in \mathcal{P}_j} \text{Var}(Y_{ij})}{1 - \pi_j} - \frac{\pi_j}{(1 - \pi_j)^2} E(D_j)^2$$

$$\text{Var}(Y_{ij}) = \frac{S(x_{ij})^2}{1 - p_{ij}} (e^{\sigma^2} - (1 - p_{ij}))$$

### 2.3.5 Simulations

